## Big Data Computing

Master's Degree in Computer Science 2020-2021

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  - Due to the curse of dimensionality

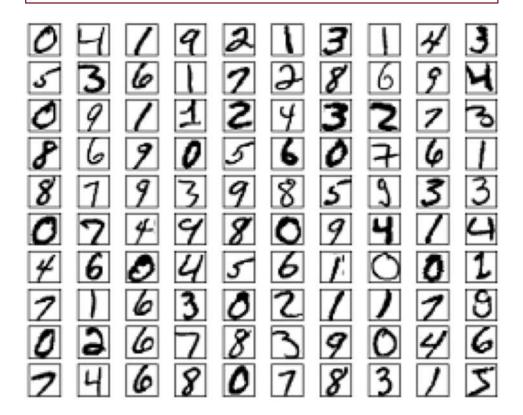
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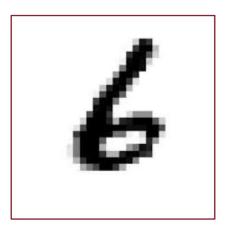
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- Clustering high-dimensional data may be problematic
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- Many other data sources (e.g., images) share the same issue
- Good news! High-dimensionality is often not real!
  - Due to the way in which we observe/collect data

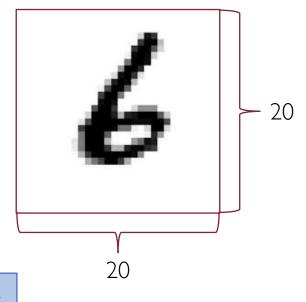
# DIMENSIONALITY REDUCTION

#### **Example**

Handwritten digit recognition



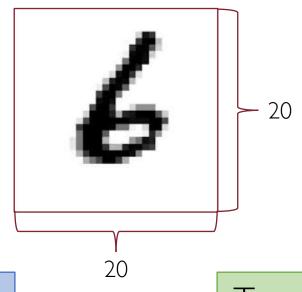




Modeled dimensionality

Each digit represented by 20x20 bitmap

400-dimensional binary vector



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Each digit represented by 20x20 bitmap

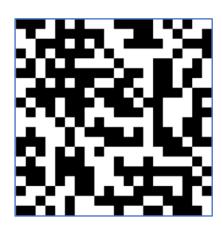
400-dimensional binary vector

True dimensionality

Actual digits just cover a tiny fraction of all this huge space

Small variations of the pen-stroke

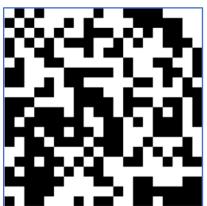
Random samples from 400-d space





Random samples from 400-d space





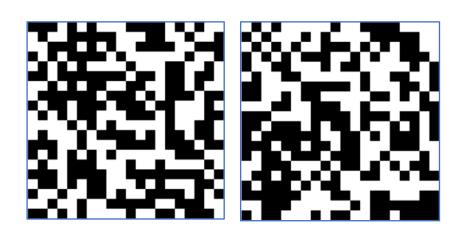
True digits living in a 400-d space





Random samples from 400-d space









We model data (i.e., digits) as very high dimensional...

... In fact, they are not so

#### The Curse of Dimensionality

As dimensionality grows fewer examples in each region of the feature space (assuming # examples is constant)

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Put it another way:

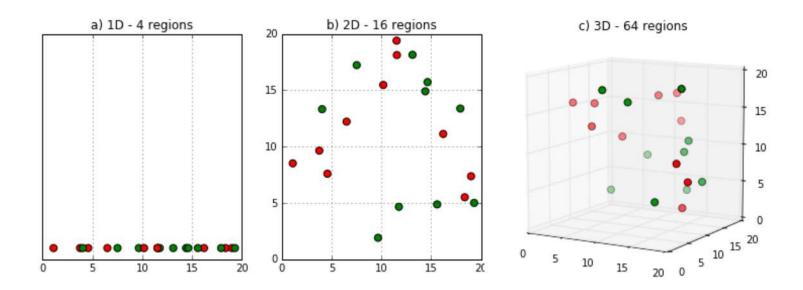
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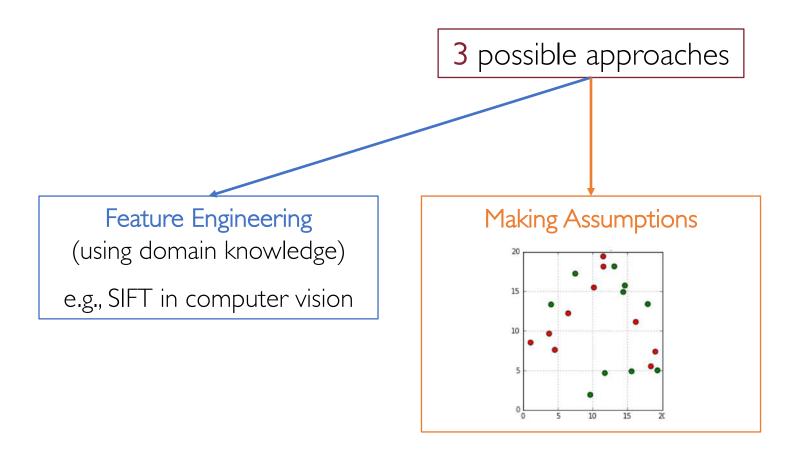
## Dealing with High Dimensionality

3 possible approaches

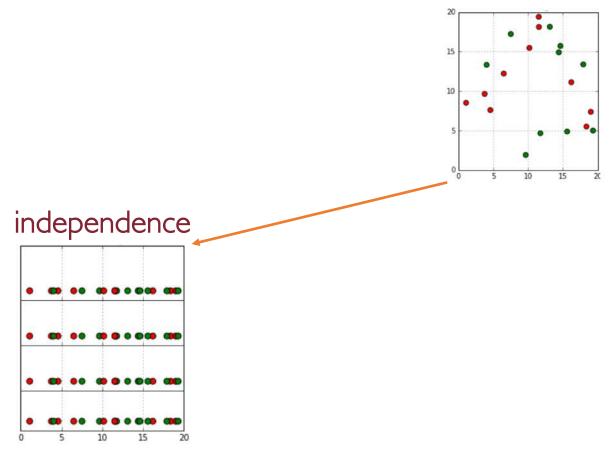
Feature Engineering (using domain knowledge)

e.g., SIFT in computer vision

## Dealing with High Dimensionality

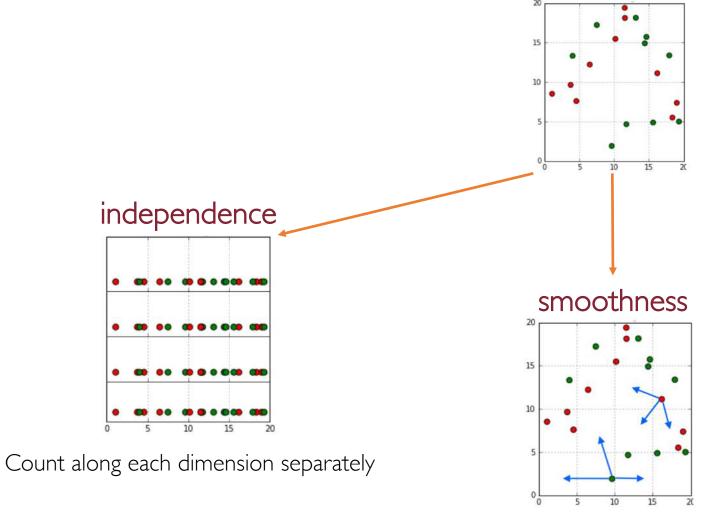


## Dealing with High Dimensionality: Assumptions

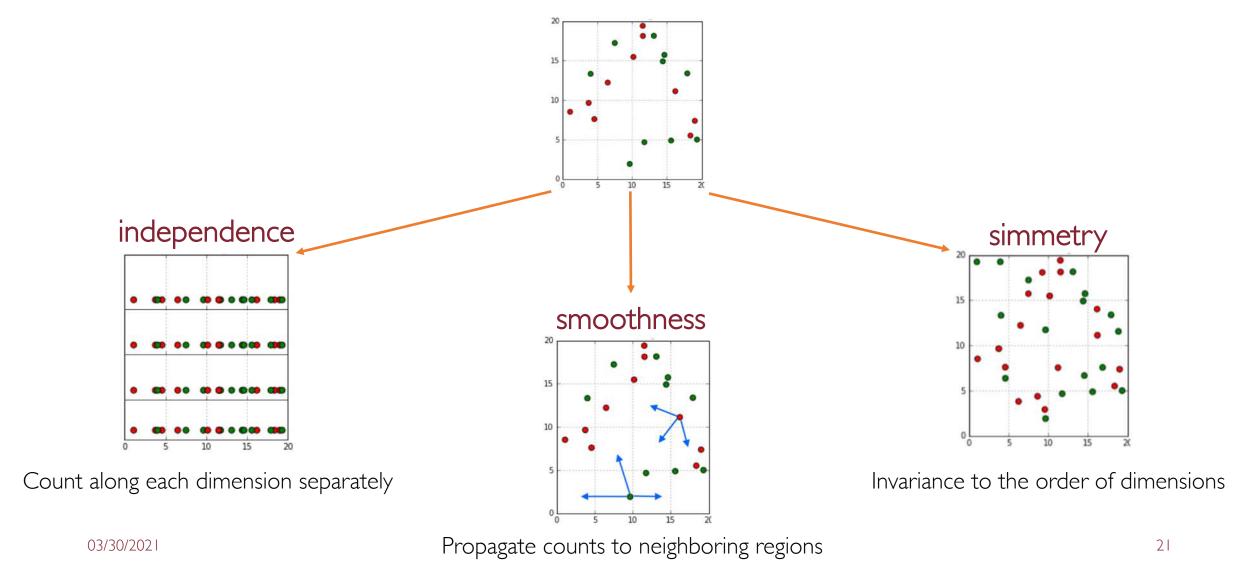


Count along each dimension separately

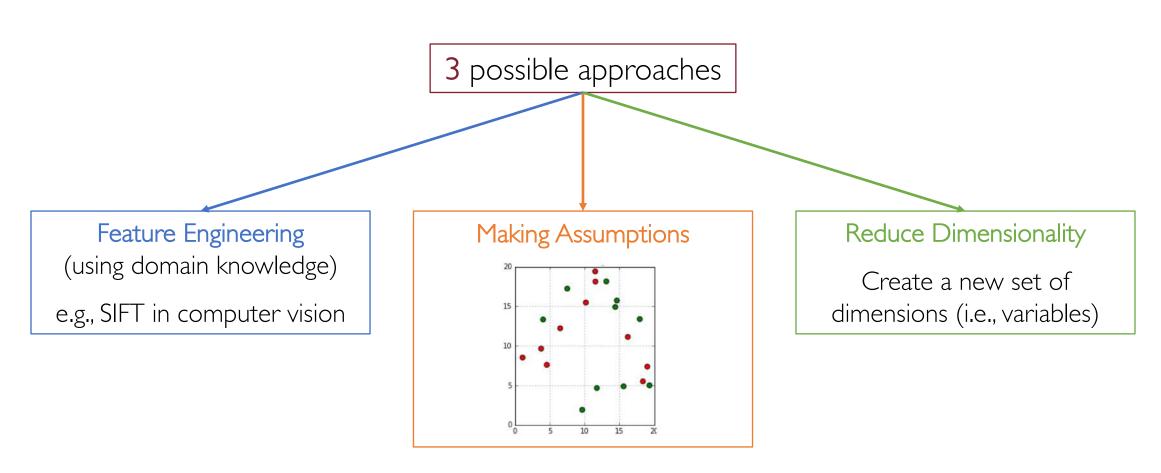
## Dealing with High Dimensionality: Assumptions



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#### Dealing with High Dimensionality



- A technique to unveil the actual (i.e., meaningful) dimensions of data
- A pre-processing step for representing data with fewer features
- Preserve as much "structure" of the data as possible
- Retained structure must be discriminative affecting data separability

"structure" here means variance

2 main approaches

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#### Feature Selection

Pick a subset of the original dimensions that are good predictors (e.g., using information gain)

$$X_1, X_2, ..., X_{j-1}, X_j, X_{j+1}, ..., X_{d-1}, X_d$$

2 main approaches

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#### Feature Extraction

Build a new set of k < d dimensions as a (linear) combination of the originals

$$e_1, e_2, \ldots, e_k$$

$$\mathbf{e}_{i} = f(\mathbf{x}_{1}, \mathbf{x}_{2}, \ldots, \mathbf{x}_{d})$$

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Dimensionality reduction technique based on feature extraction

High-dimensional data is in fact embedded into some lower dimensional space

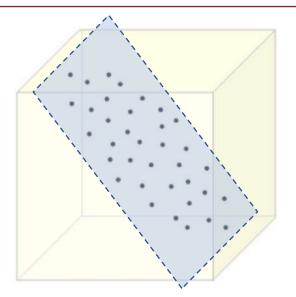
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Dimensionality reduction technique based on feature extraction

High-dimensional data is in fact embedded into some lower dimensional space

#### **Example**

A 3-d set of points embedded into a 2-d hyperplane



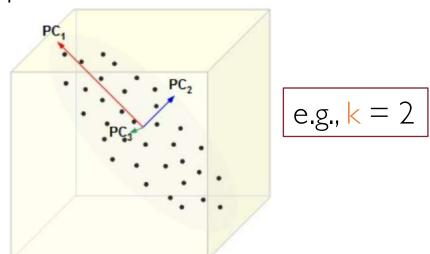
PCA defines a set of principal components as follows:

- Ist: direction of the greatest variance of data
- 2nd: perpendicular to 1st and greatest variance of what's left
- ... and so on until d

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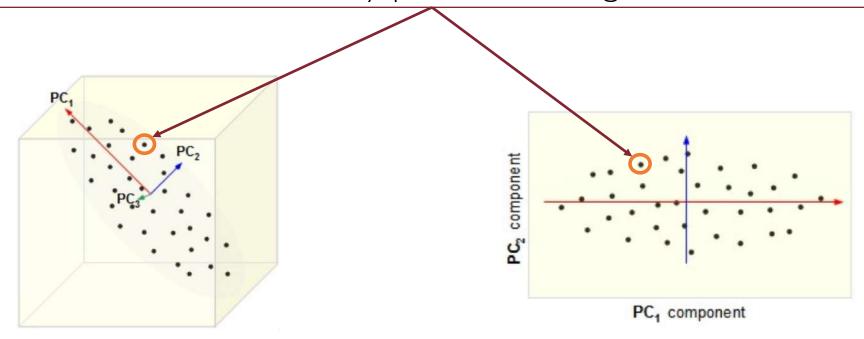
The top k < d components become the new dimensions



 $PC_1$  and  $PC_2$  are the top-2 principal components

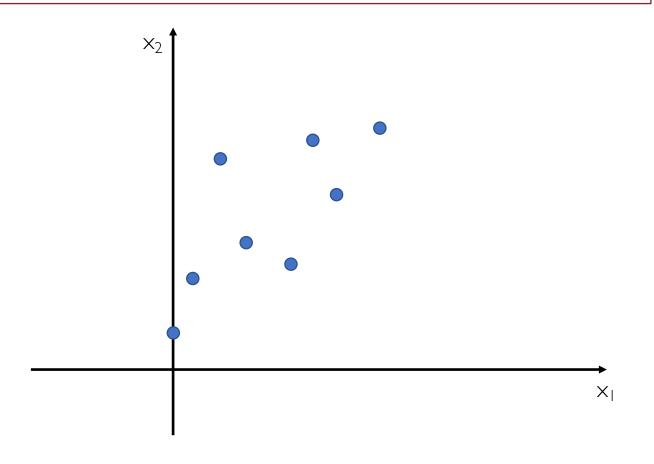
PC<sub>1</sub> and PC<sub>2</sub> are the top-2 principal components

Change the coordinates of every point according to the new dimensions



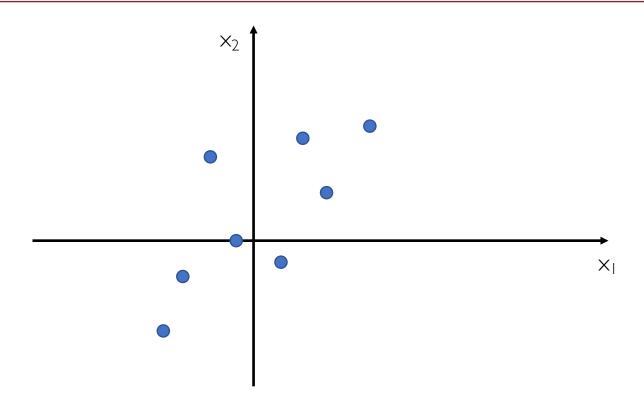
#### Why Do We Look for Greatest Variance?

Example: Reduce 2-dimensional data to 1-d



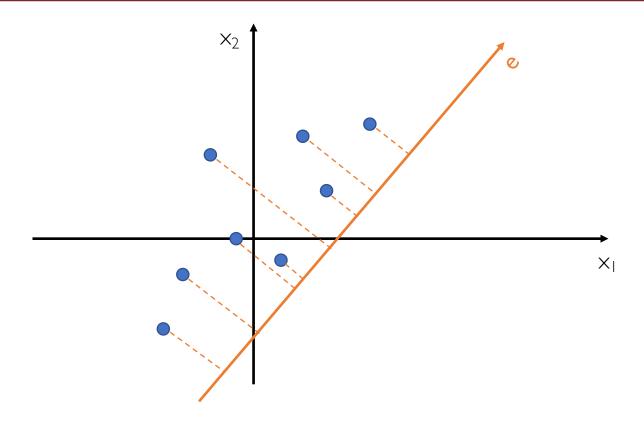
## Why Do We Look for Greatest Variance?

First of all, let's center the points around the mean along  $x_1$  and  $x_2$ 

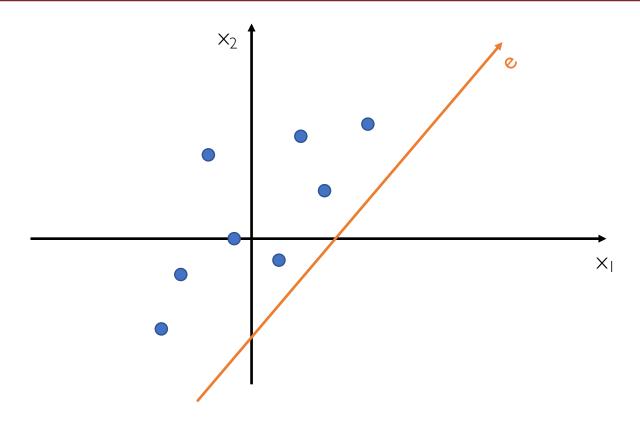


#### Why Do We Look for Greatest Variance?

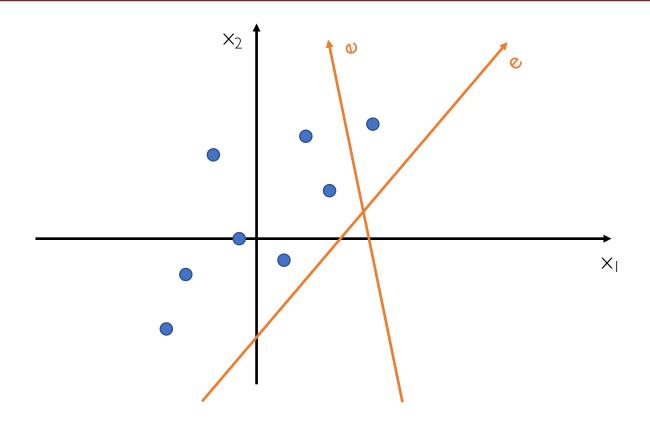
Map, i.e., project  $(x_1, x_2)$  to a new single dimension axis e



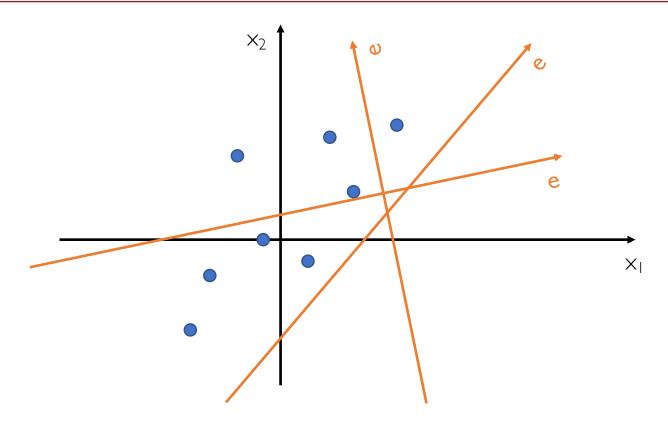
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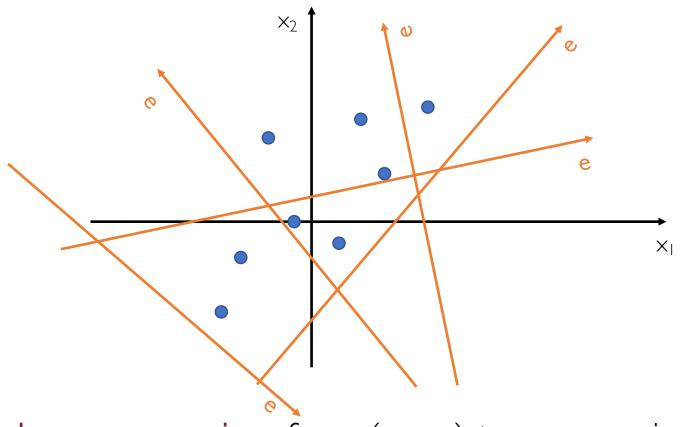
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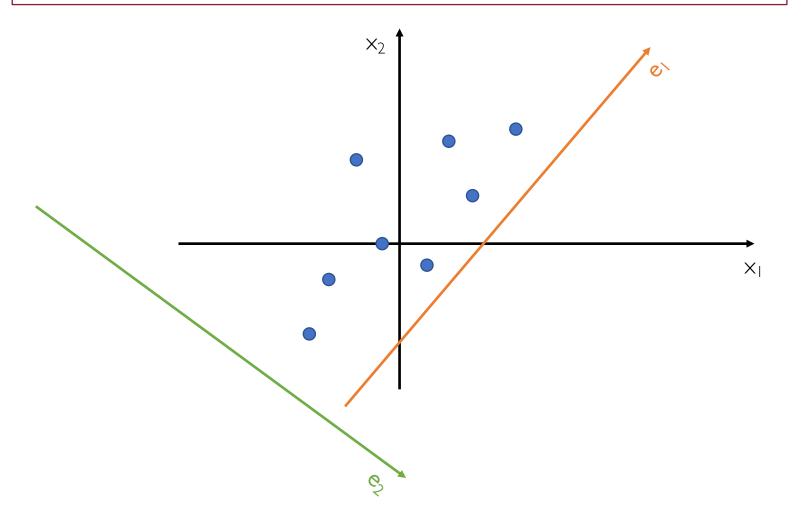


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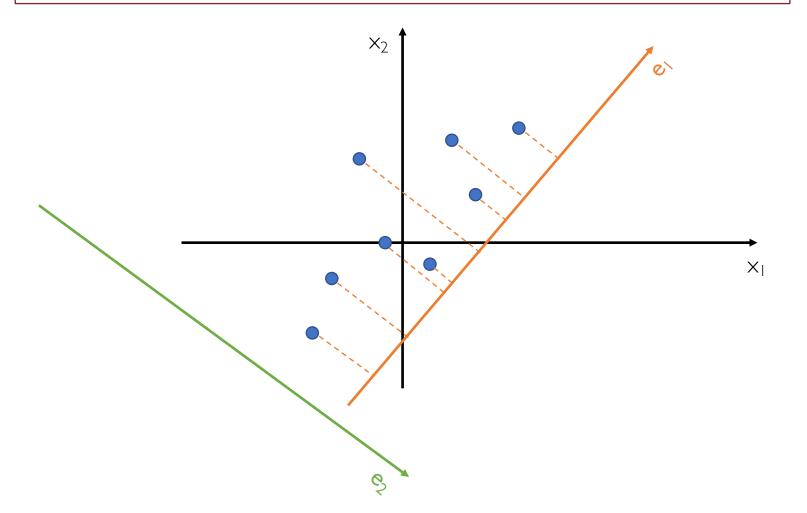


infinitely many mappings from  $(x_1, x_2)$  to a new axis e

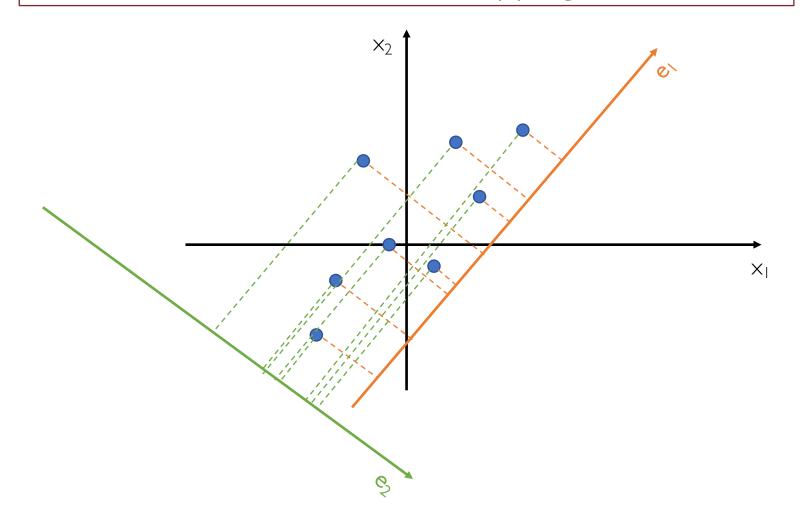
Let's consider 2 different mappings e<sub>1</sub> and e<sub>2</sub>



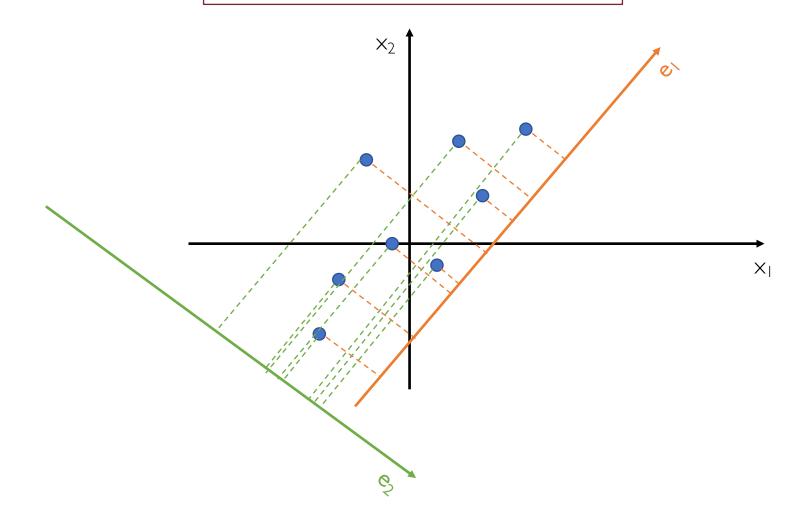
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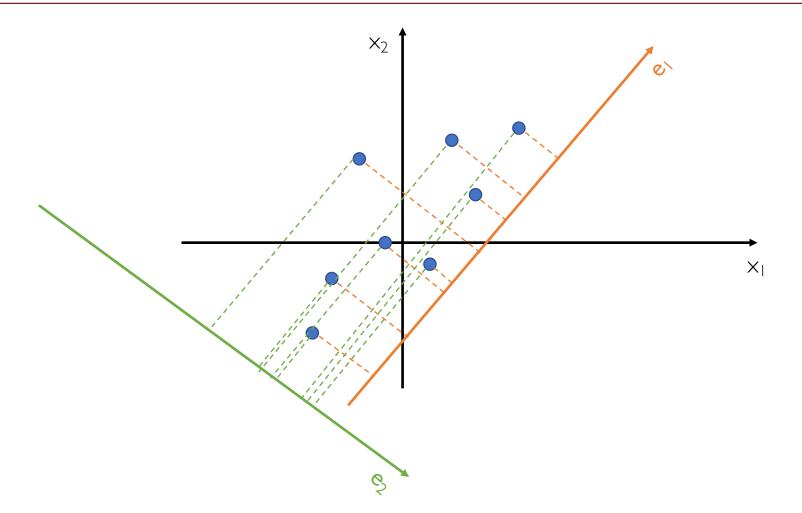
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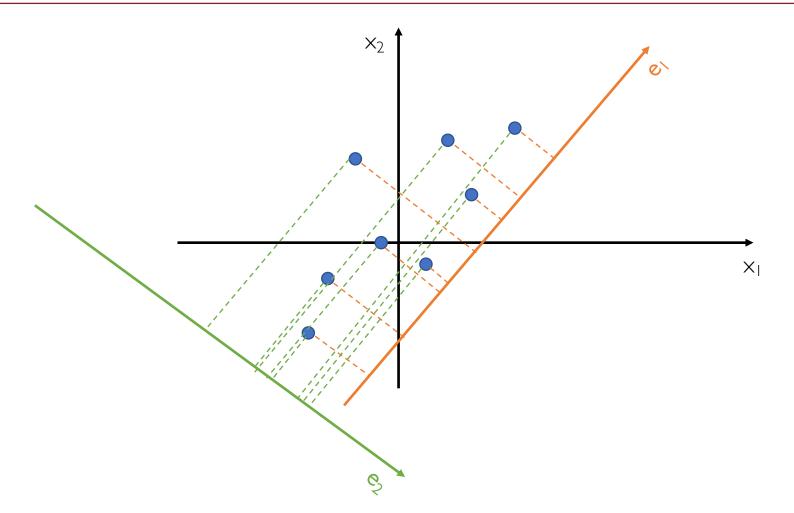
Which one is better?



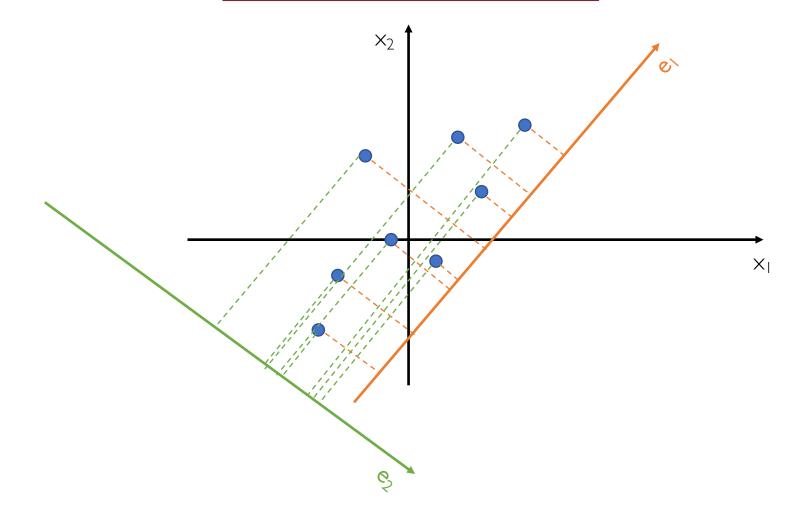
Points projected onto e look more spread-out than onto e2



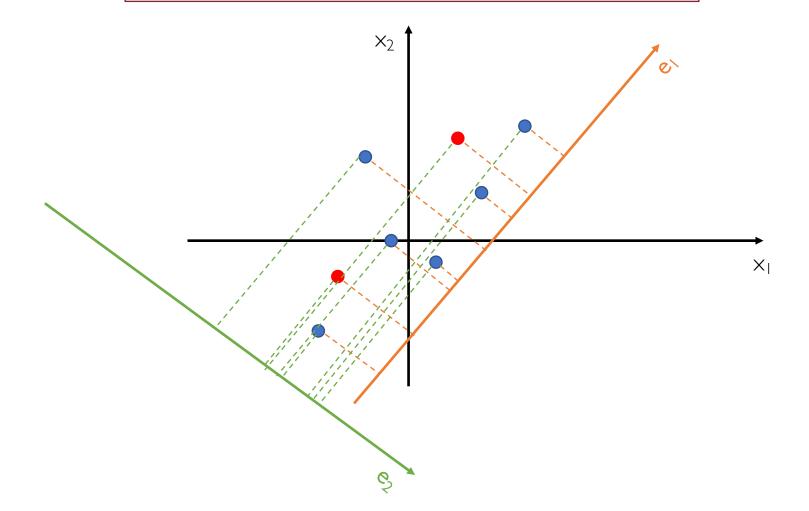
The variance along e<sub>1</sub> is larger than along e<sub>2</sub>



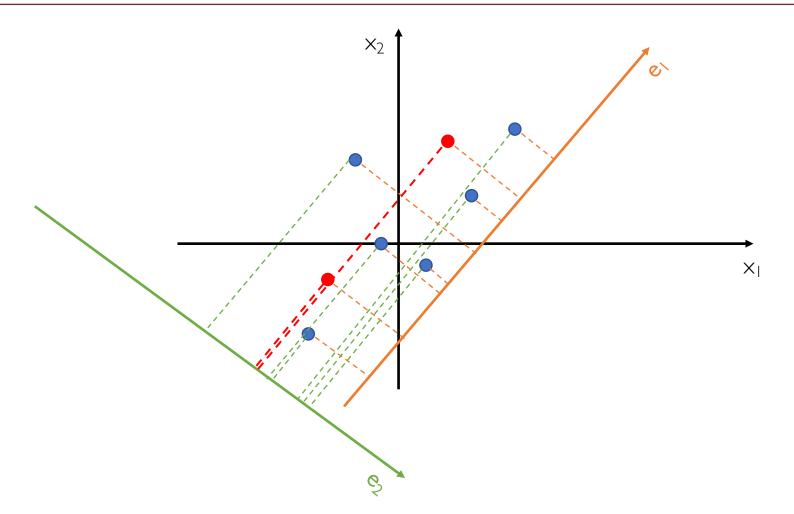
Why is that good?



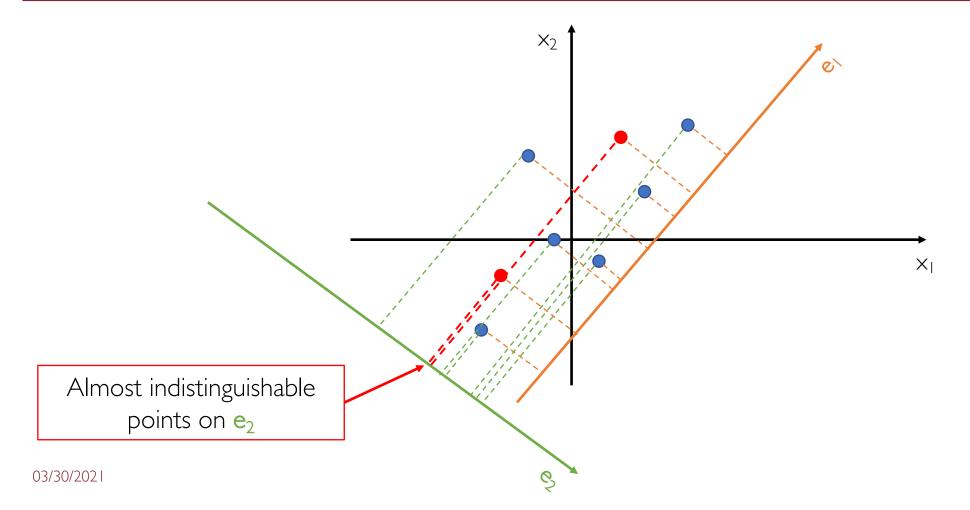
Consider the 2 red points below



On  $(x_1, x_2)$  far away from each other, end up close if projected onto  $e_2$ 

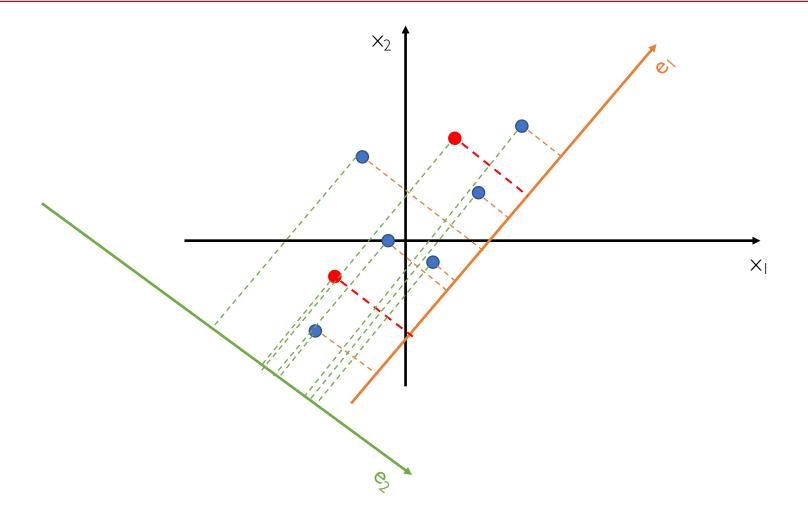


On  $(x_1, x_2)$  far away from each other, end up close if projected onto  $e_2$ 



50

If projected onto e<sub>1</sub> they better preserve their distance



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- Minimize distances between points as measured on  $(x_1, x_2)$  space and those measured on e

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#### **Solution**

Pick e so as to maximize variance of projected data

#### Variance of a Random Variable

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- Formally, it is the expected value of the squared deviation from its mean

$$Var(X) = E[(X - \mu)^2]$$

where 
$$\mu = E[X]$$

#### Covariance of Two Random Variables

- A measure of the joint variability of two random variables X and Y
  - Do X and Y increase/decrease together, or when one increases/decreases the other decreases/increases?

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- Formally, it is the expected value of the product of their deviations from their individual means

$$Cov(X, Y) = E[(X - \mu_X)(Y - \mu_Y)]$$

$$Cov(X, X) = Var(X)$$

where 
$$\mu_X = E[X]$$
 and  $\mu_Y = E[Y]$ 

### Covariance Matrix

• Given a random vector  $\mathbf{X} = (X_1, ..., X_d)$  its covariance matrix K is a dxd square matrix with the covariance between each pair of elements

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- Given a random vector  $\mathbf{X} = (X_1, ..., X_d)$  its covariance matrix K is a dxd square matrix with the covariance between each pair of elements
- In the matrix diagonal there are variances, i.e., the covariance of each element with itself

$$K[i, j] = Cov(X_i, X_j)$$

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- In our example, d = 2 and  $X = (X_1, X_2)$
- The covariance matrix K is a 2-by-2 matrix
- To ease the covariance computation, we center each data point at zero
  - Subtracting the mean of each attribute/dimension
  - The mean of each dimension becomes then 0

Let n be the total number of data points:  $\mathbf{x}_1, \dots, \mathbf{x}_n$ Each data point is represented by a  $(x_1, x_2)$  pair  $\mathbf{x}_i = (x_{i,1}, x_{i,2})$ 

We associate 2 random variables  $X_1, X_2$  to each dimension, and we compute:

$$\mu_1 = E[X_1] = \frac{1}{n} \sum_{i=1}^n x_{i,1}$$

$$\mu_2 = E[X_2] = \frac{1}{n} \sum_{i=1}^n x_{i,2}$$

$$\mathbf{x}_i = (x_{i,1} - \mu_1, x_{i,2} - \mu_2)$$

Let us rewrite each data point  $\mathbf{x}_i$  as follows:

$$\mathbf{x}_{i} = (x'_{i,1}, x'_{i,2})$$
 where:

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 $x'_{i,1} = x_{i,1} - \mu_1; x'_{i,2} = x_{i,2} - \mu_2$ 

$$\mu_1^{\text{new}} = E[X_1] = \frac{1}{n} \sum_{i=1}^n x'_{i,1} = \frac{1}{n} \sum_{i=1}^n (x_{i,1} - \mu_1)$$

$$\mu_2^{\text{new}} = E[X_2] = \frac{1}{n} \sum_{i=1}^n x'_{i,2} = \frac{1}{n} \sum_{i=1}^n (x_{i,2} - \mu_2)$$

$$\mu_1^{\text{new}} = \frac{1}{n} \sum_{i=1}^n (x_{i,1} - \mu_1) = \frac{1}{n} \left( \sum_{i=1}^n x_{i,1} - \sum_{i=1}^n \mu_1 \right) = 0$$

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0-mean

Scaling data so as to have 0-mean on all dimensions allow computing covariance much easily

$$Cov(X_1, X_2) = E[(X_1 - \underbrace{\mu_1^{\text{new}}}_{=0})(X_2 - \underbrace{\mu_2^{\text{new}}}_{=0})] = E[X_1 X_2]$$

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As a consequence, the covariance matrix is also easier to compute!

Let's assume the following is our 2-by-2 covariance matrix

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$$\begin{array}{c|c} \times_{1} & \times_{2} \\ \times_{1} & 2 & 4/5 \\ \times_{2} & 4/5 & 3/5 \end{array}$$
  $\begin{array}{c|c} \operatorname{Cov}(X_{1}, X_{2}) = \frac{1}{n} \sum_{i=1}^{n} x'_{i,1} * x'_{i,2} \\ \times_{2} & 4/5 & 3/5 \end{array}$ 

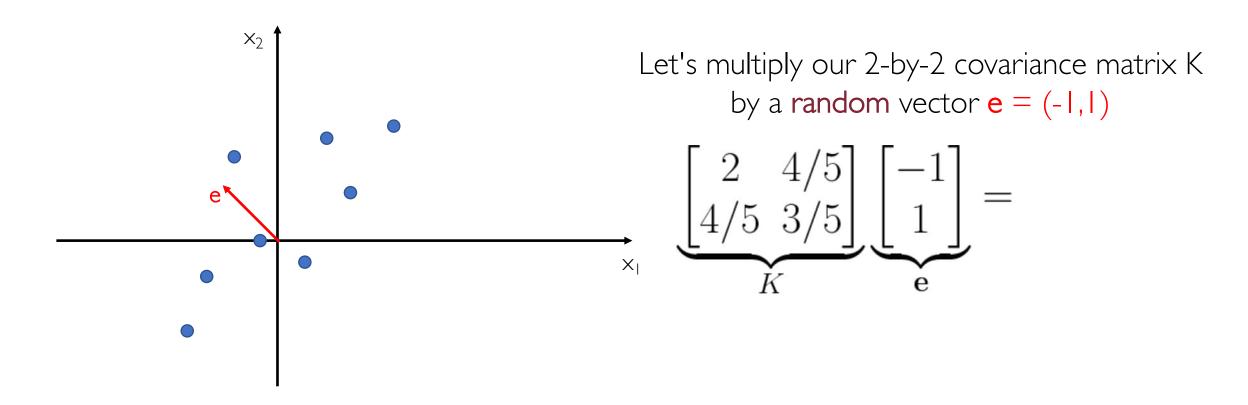
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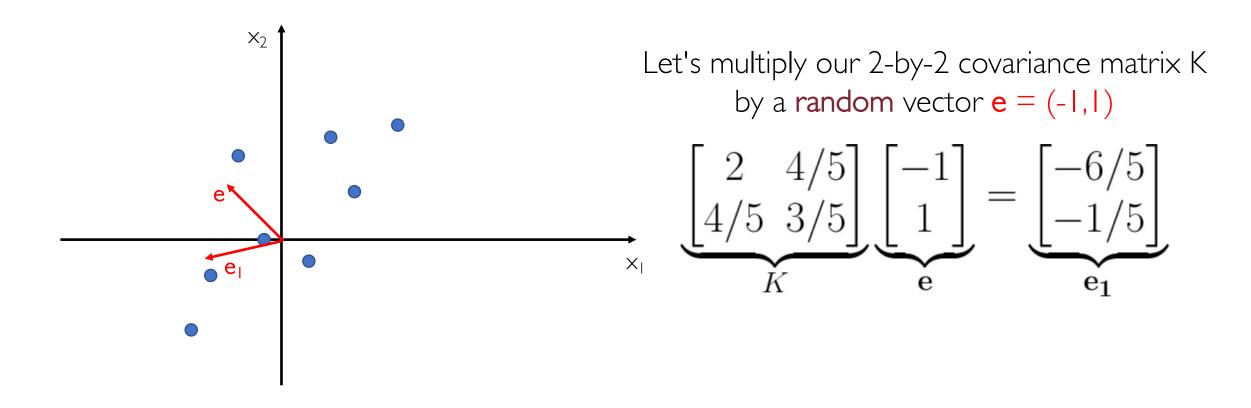
$$\begin{array}{c}
\times_{1} \\
\times_{2} \\
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\end{array}$$

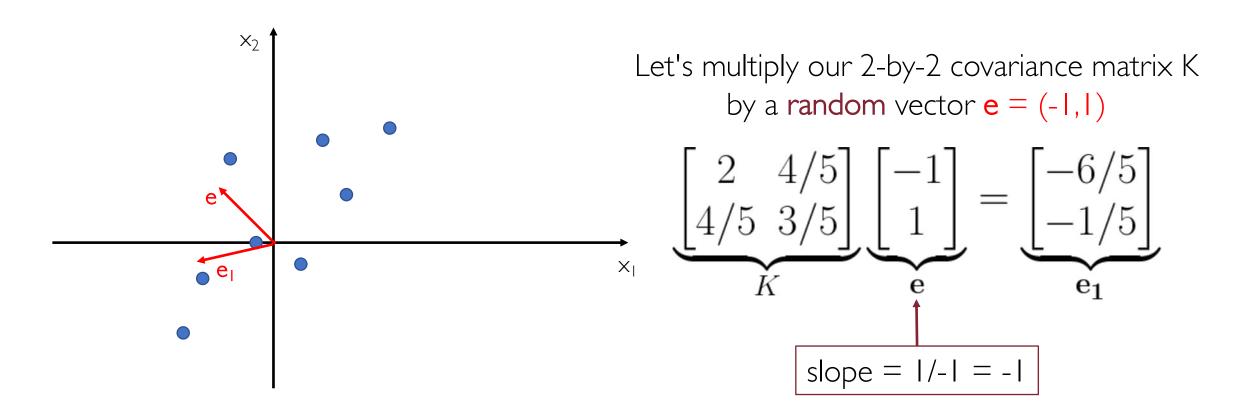
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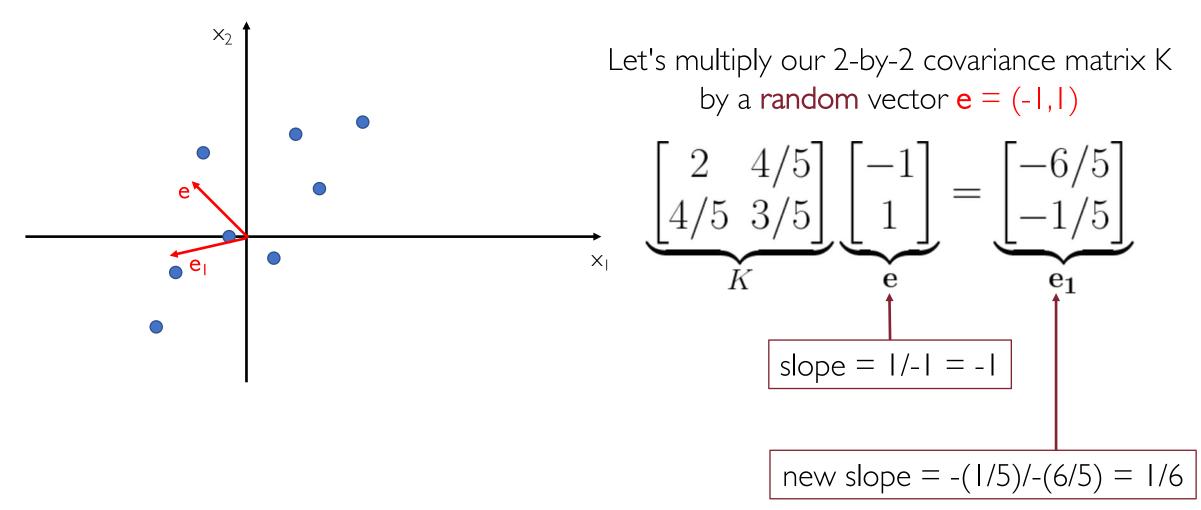
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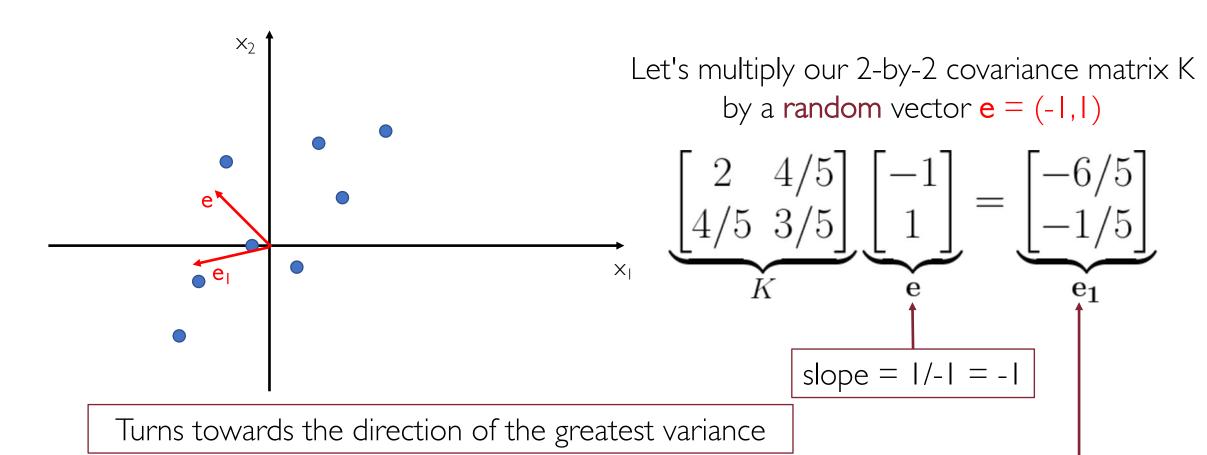
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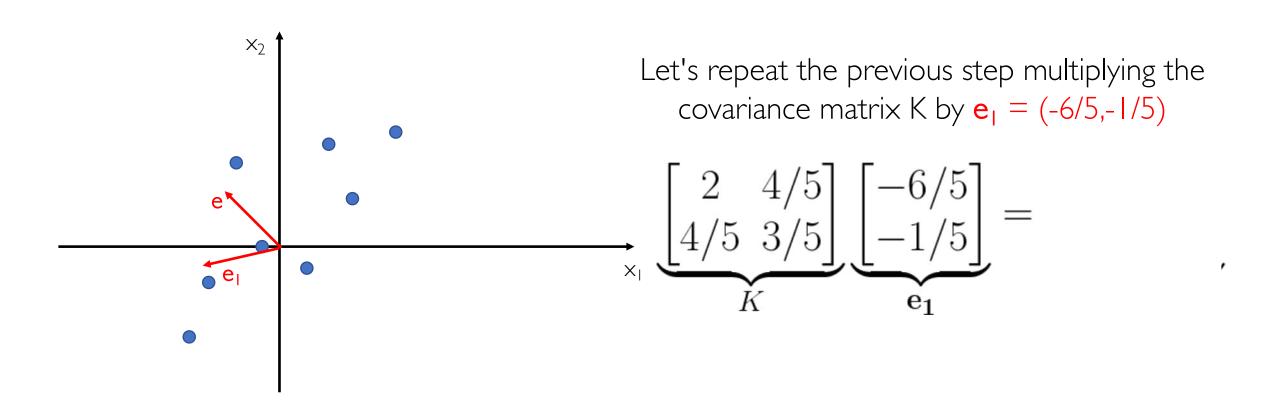


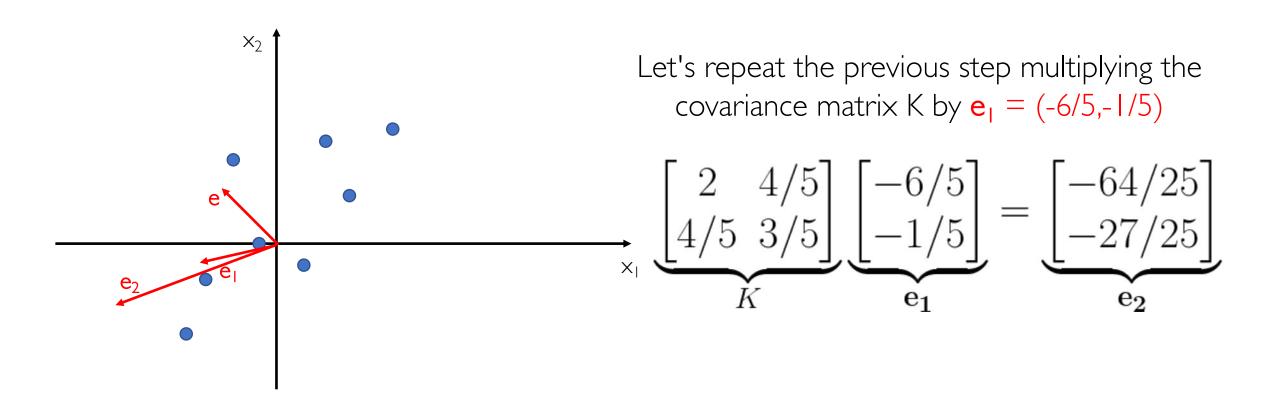


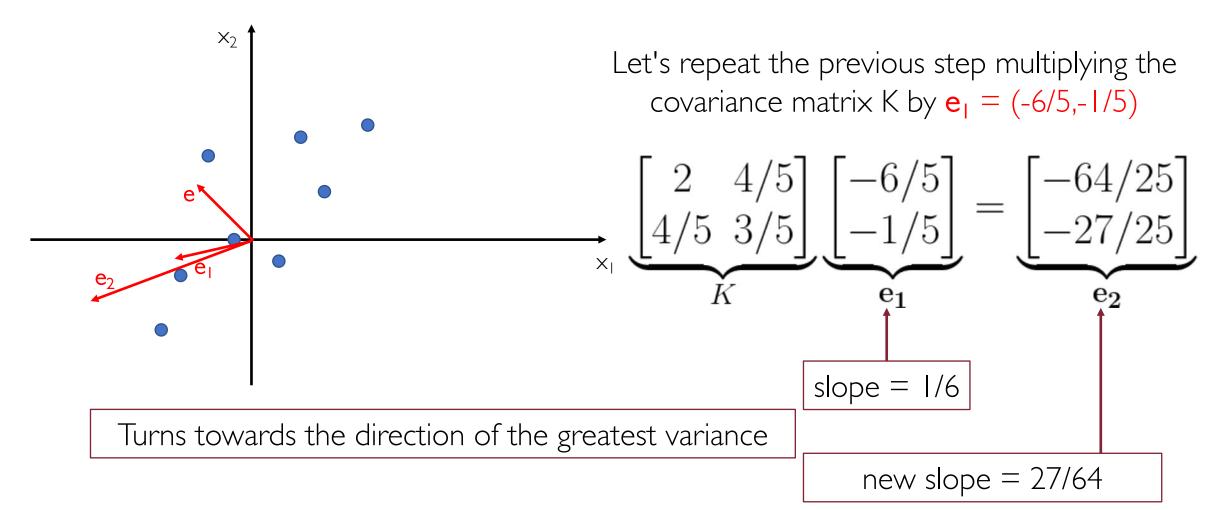




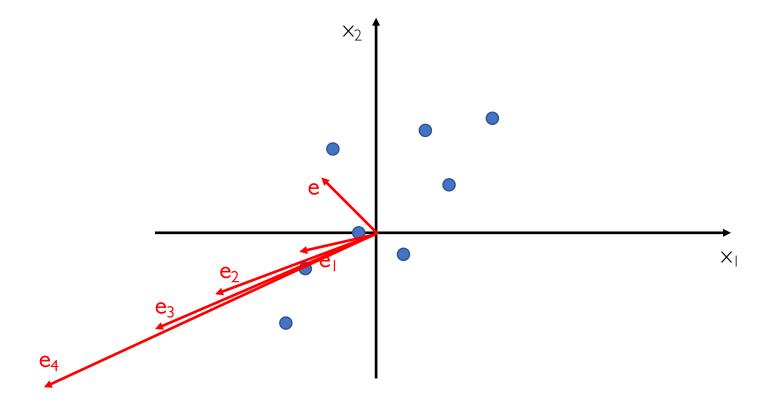
new slope = -(1/5)/-(6/5) = 1/6



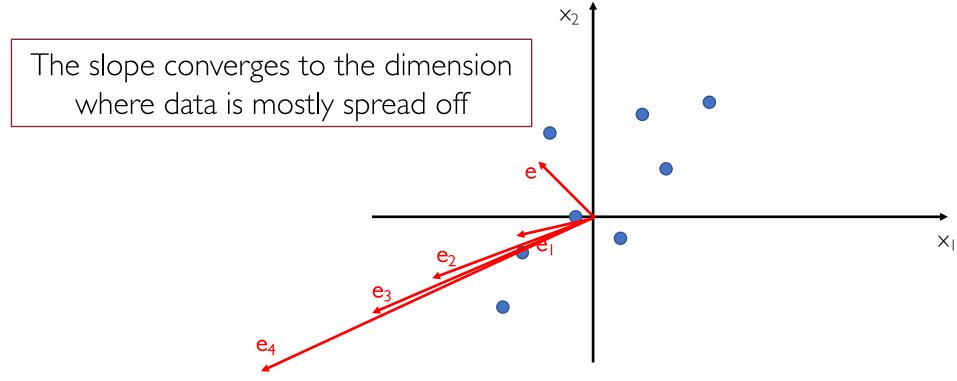




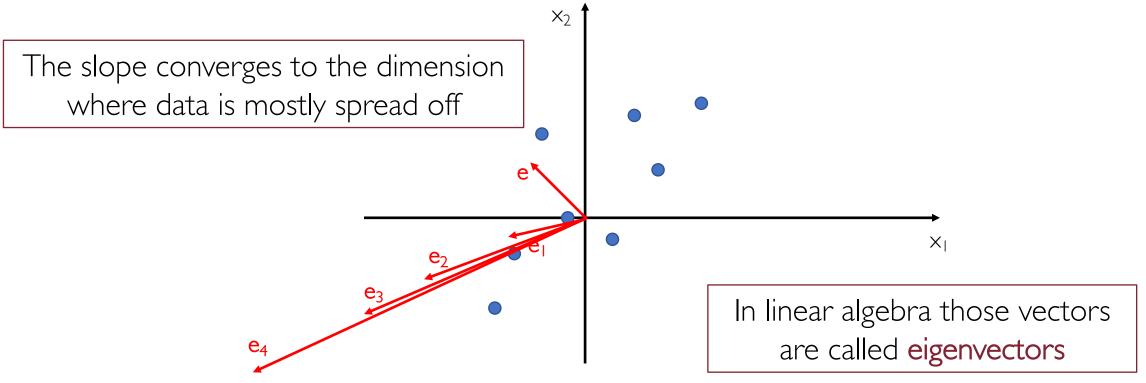
If we keep doing this the resulting vector is getting longer and turns towards the direction of the largest variance



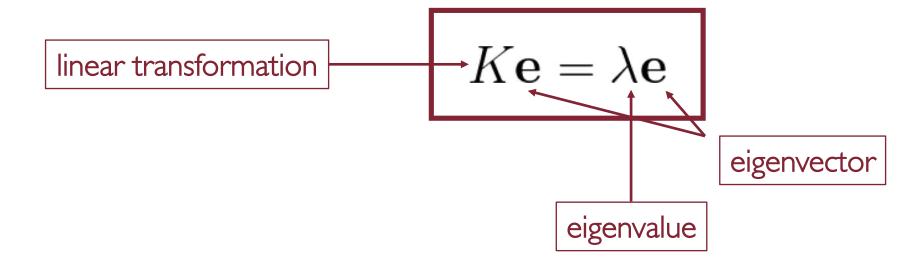
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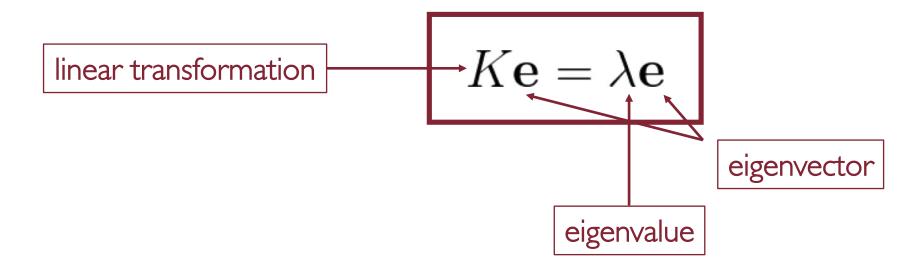


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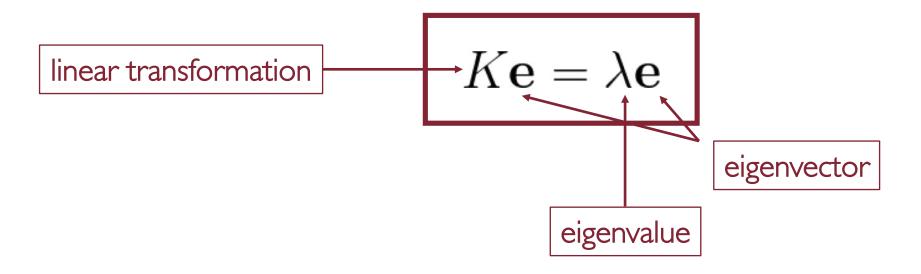


$$K\mathbf{e} = \lambda \mathbf{e}$$



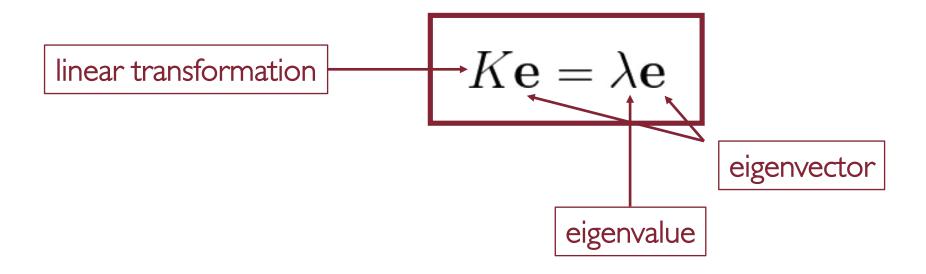


When you multiply a matrix by an eigenvector e the resulting vector does not change its direction, but it is only scaled by a factor  $\lambda$  (eigenvalue)



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In other words, eigenvectors encapsulate all the relevant information to describe a linear transformation (in our case, represented by the covariance matrix K)



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#### **Principal Components**

eigenvectors of the covariance matrix with the largest eigenvalues

Remember that we want to solve for **e** the following:

$$K\mathbf{e} = \lambda \mathbf{e}$$

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We can rewrite the system of equations above as:

$$K\mathbf{e} - \lambda \mathbf{e} = 0 \Rightarrow (K - \lambda I)\mathbf{e} = 0$$

I is the identity matrix

We therefore resort to solve the following homogeneous system:

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The only way for the homogeneous system above to have a **non-trivial** solution is for its matrix  $(K - \lambda I)$  to be **non-invertible**, otherwise:

$$(\underline{K} - \lambda I)(\underline{K} - \lambda I)^{-1} \mathbf{e} = 0(\underline{K} - \lambda I)^{-1}$$

A square matrix is invertible iff its determinant is not equal to 0

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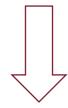
If the determinant of the matrix  $(K - \lambda I)$  is equal to 0, it is non-invertible

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A square matrix is invertible iff its determinant is not equal to 0



If the determinant of the matrix  $(K - \lambda I)$  is equal to 0, it is non-invertible



The corresponding homogeneous system will have a non-trivial solution

I. Find the eigenvalues by solving for  $\lambda$ : det(K –  $\lambda$ I) = 0

$$\det\left(\underbrace{\begin{bmatrix}2-\lambda & 4/5\\4/5 & 3/5-\lambda\end{bmatrix}}_{K-\lambda I}\right) = 0$$

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characteristic equation of K

$$\lambda_1 = \frac{13 + \sqrt{113}}{10} \approx 2.36; \quad \lambda_2 = \frac{13 - \sqrt{113}}{10} \approx 0.24$$

2. Plug each eigenvalue in to find the corresponding eigenvector

$$\underbrace{\begin{bmatrix} 2 & 4/5 \\ 4/5 & 3/5 \end{bmatrix}}_{K} \underbrace{\begin{bmatrix} e_{1,1} \\ e_{1,2} \end{bmatrix}}_{\mathbf{e}_{1}} = \lambda_{1} \underbrace{\begin{bmatrix} e_{1,1} \\ e_{1,2} \end{bmatrix}}_{\mathbf{e}_{1}}$$

$$\underbrace{\begin{bmatrix} 2 & 4/5 \\ 4/5 & 3/5 \end{bmatrix}}_{K} \underbrace{\begin{bmatrix} e_{2,1} \\ e_{2,2} \end{bmatrix}}_{\mathbf{e}_{2}} = \lambda_{2} \underbrace{\begin{bmatrix} e_{2,1} \\ e_{2,2} \end{bmatrix}}_{\mathbf{e}_{2}}$$

Let's see what happens for  $\lambda_1$ 

$$\begin{cases} 2e_{1,1} + 4/5e_{1,2} = \frac{13+\sqrt{113}}{10}e_{1,1} \\ 4/5e_{1,1} + 3/5e_{1,2} = \frac{13+\sqrt{113}}{10}e_{1,2} \end{cases}$$

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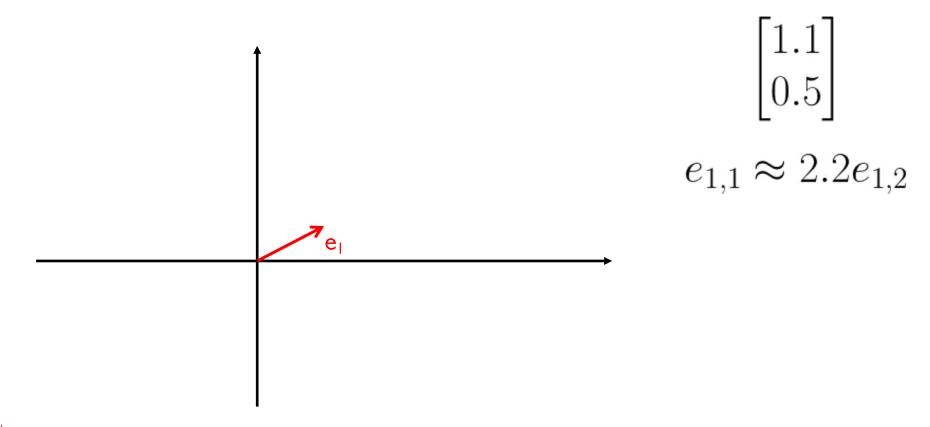
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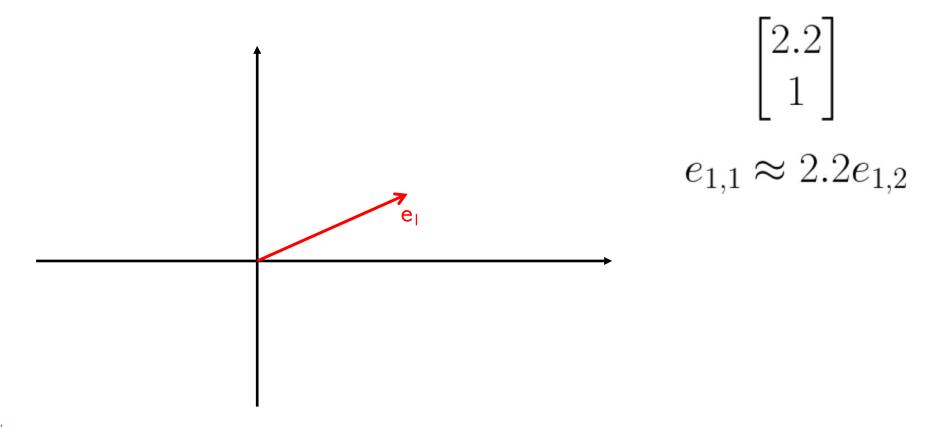
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The system has infintely many solutions

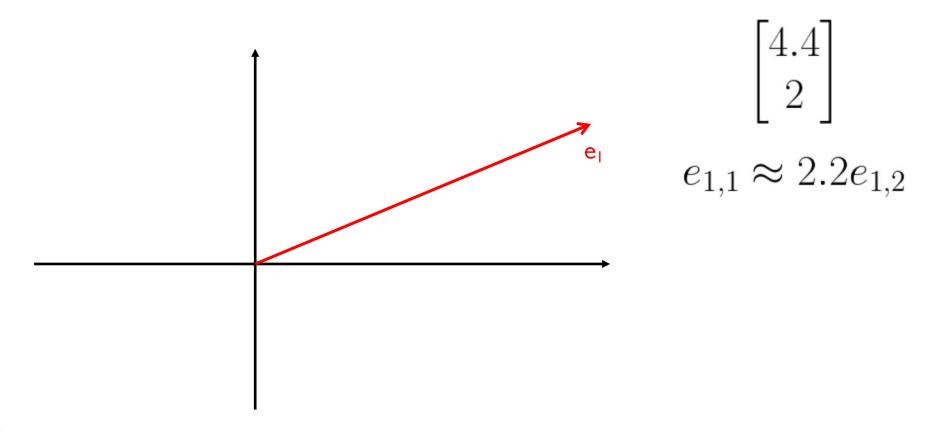
Any vector which satisfies the relationship above works!



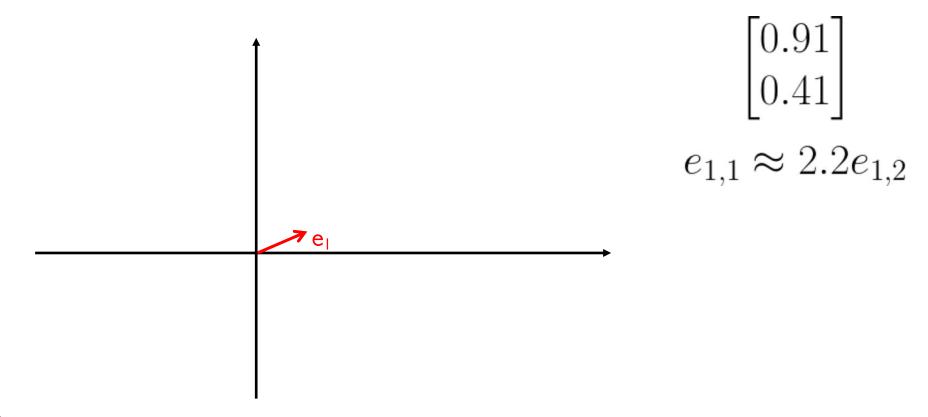
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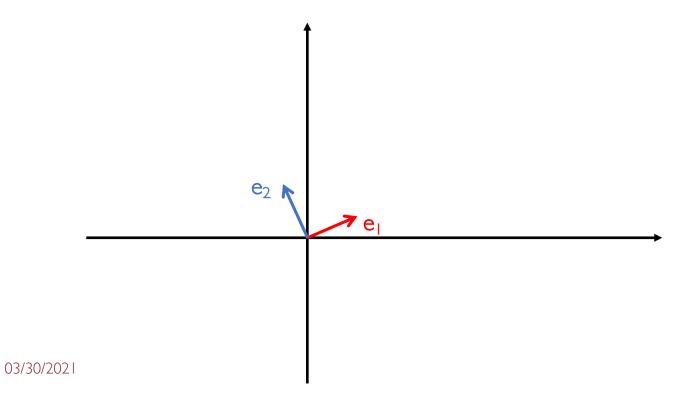
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By convention, we restrict to  $\|\mathbf{e}_{\mathbf{I}}\| = \mathbf{I}$ 



The second eigenvector  $e_2$  can be found by plugging in the smaller eigenvalue  $\lambda_2$ 

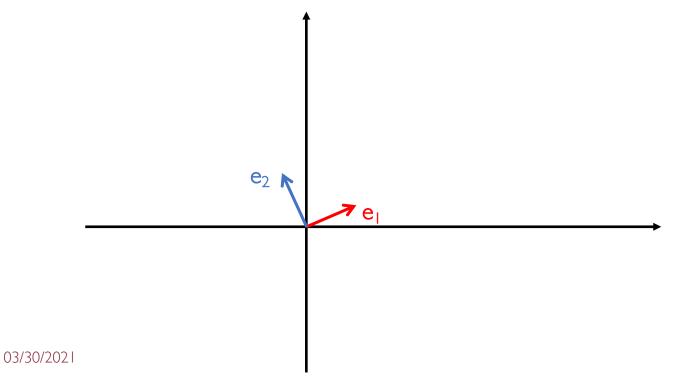


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# How Do We Compute Eigenvectors?

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This is just orthogonal to the previously found e

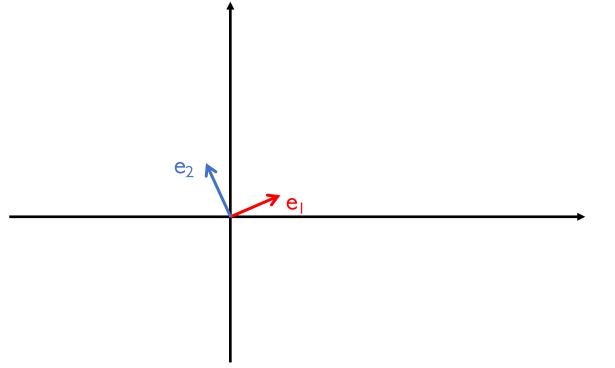


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# How Do We Compute Eigenvectors?

The second eigenvector  $e_2$  can be found by plugging in the smaller eigenvalue  $\lambda_2$ 

This is just orthogonal to the previously found e<sub>1</sub>



 $e_1$  and  $e_2$  are the new coordinate system replacing the original  $x_1$  and  $x_2$ 

$$\mathbf{e_1} = \begin{bmatrix} 0.91 \\ 0.41 \end{bmatrix} \mathbf{e_2} = \begin{bmatrix} -0.41 \\ 0.91 \end{bmatrix}$$

# Principal Components

$$\mathbf{e_1} = \begin{bmatrix} 0.91 \\ 0.41 \end{bmatrix} \mathbf{e_2} = \begin{bmatrix} -0.41 \\ 0.91 \end{bmatrix}$$

e<sub>I</sub> is the 1st principal component as it is the eigenvector corresponding to the largest eigenvalue

e<sub>2</sub> is the 2nd principal component as it is the eigenvector corresponding to the smallest eigenvalue

• e<sub>1</sub> and e<sub>2</sub> identify our new coordinate system (principal components)

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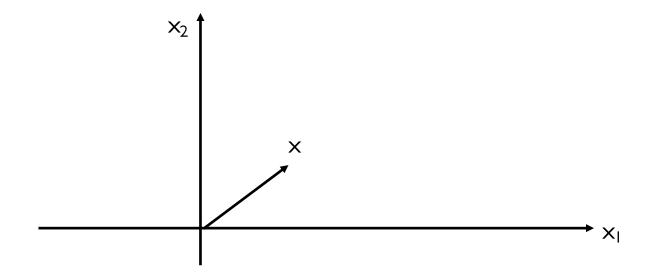


#### Goal

We want to represent x in the new  $(e_1, e_2)$ -coordinate system

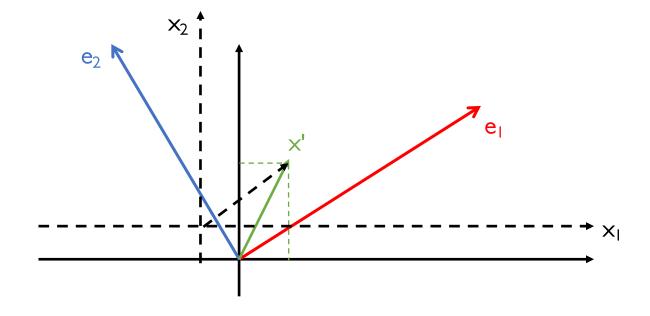
I. Center x around the mean of each dimension

$$\mathbf{x}' = \mathbf{x} - \boldsymbol{\mu} = (x_1 - \mu_1, x_2 - \mu_2)$$



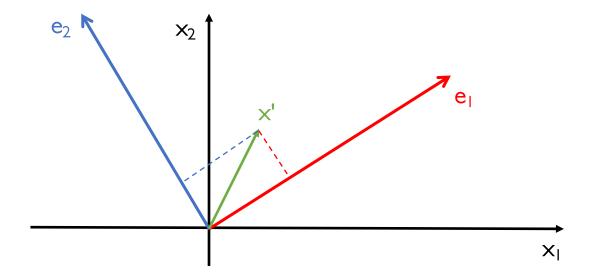
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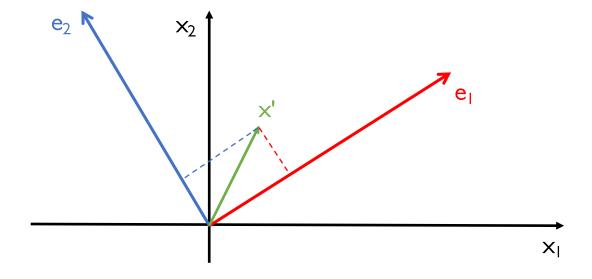


2. Project x' on each dimension  $e_1$  and  $e_2$ 

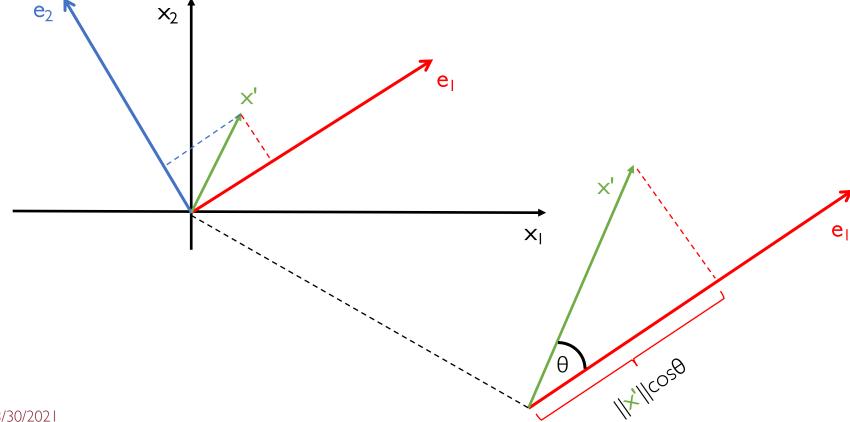
$$\mathbf{x}' = \underbrace{(\mathbf{x}'_1, \mathbf{x}'_2)}_{\text{coordinates of } \mathbf{x}' \text{ in the } (\mathbf{e}_1, \mathbf{e}_2) \text{-space}} = (\mathbf{x}'^T \mathbf{e}_1, \mathbf{x}'^T \mathbf{e}_2)$$



Why the dot product?

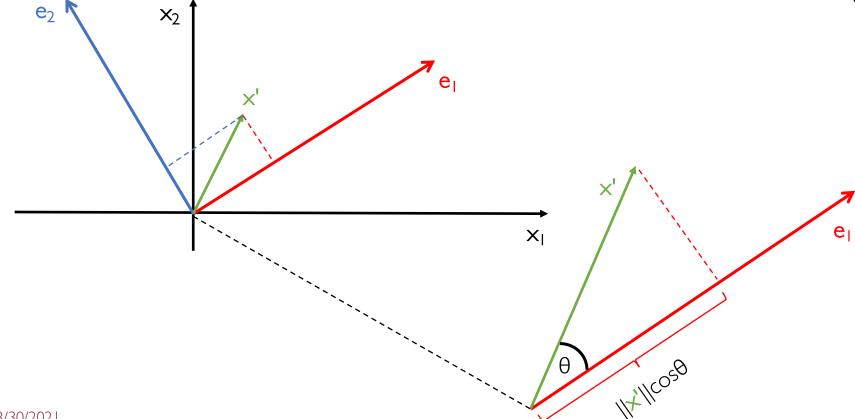


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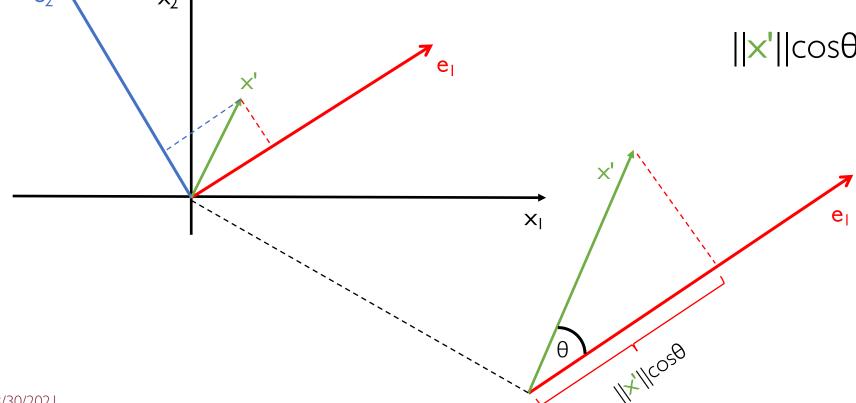
$$cos\theta = (x'e_I)/||x'||||e_I||$$

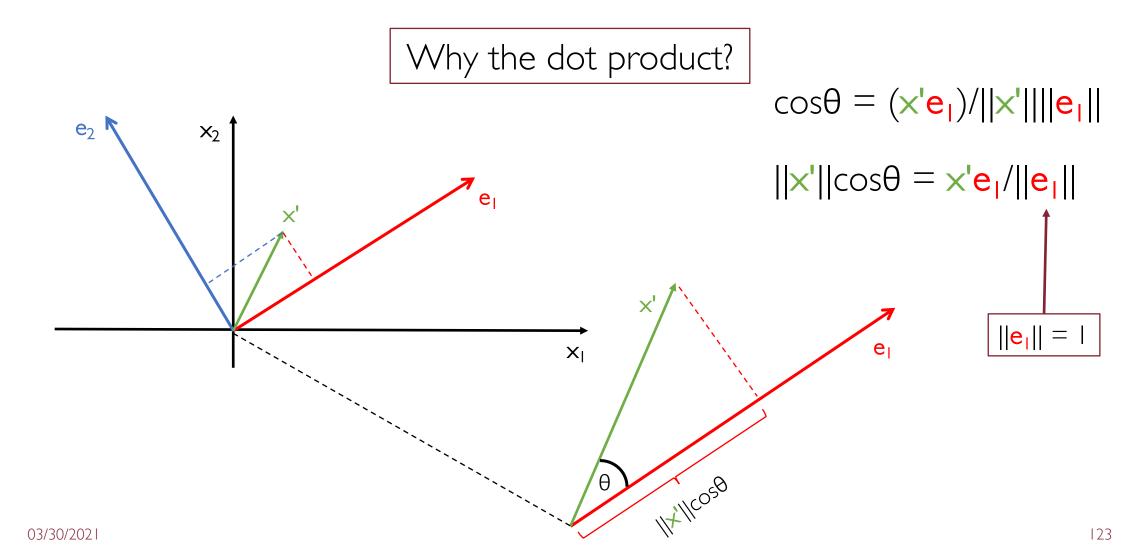


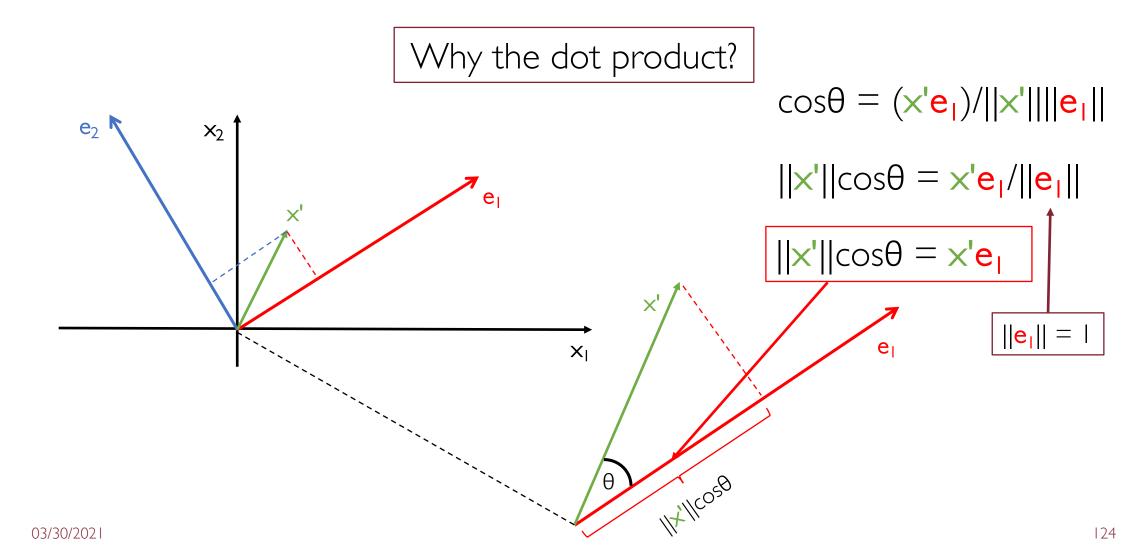
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$$||x'||\cos\theta = x'e_1/||e_1||$$

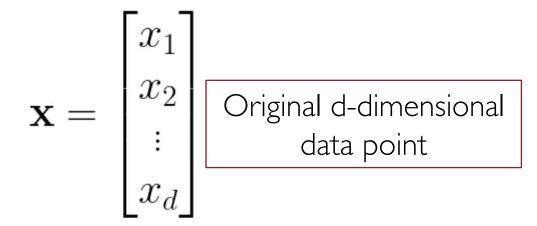






The new coordinates of the original data point x according to the eigenvectors  $e_1$  and  $e_2$  are as follows:

$$\mathbf{x}' = \begin{bmatrix} x_1' \\ x_2' \end{bmatrix} = \begin{bmatrix} \mathbf{x}'^T \mathbf{e}_1 \\ \mathbf{x}'^T \mathbf{e}_2 \end{bmatrix} = \begin{bmatrix} (x_1 - \mu_1)e_{1,1} + (x_2 - \mu_2)e_{1,2} \\ (x_1 - \mu_1)e_{2,1} + (x_2 - \mu_2)e_{2,2} \end{bmatrix}$$



$$\mathbf{x} = \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_d \end{bmatrix} \begin{array}{|c|c|c|c|} \hline \text{Original d-dimensional} & \mathbf{e}_1, \mathbf{e}_2, \dots, \mathbf{e}_k \\ k \ll d, \ \mathbf{e}_i \in \mathbb{R}^d & \text{principal components} \\ \hline \end{array}$$

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I. Mean centering

$$\mathbf{x}' = \mathbf{x} - \boldsymbol{\mu} = \begin{bmatrix} x_1 - \mu_1 \\ x_2 - \mu_2 \\ \vdots \\ x_d - \mu_d \end{bmatrix}$$

$$\mathbf{x} = egin{bmatrix} x_1 \ x_2 \ dots \ x_d \end{bmatrix}$$
 Original d-dimensional data point  $k \ll d, \ \mathbf{e}_i \in \mathbb{R}^d$  prince

$$\mathbf{e}_1, \mathbf{e}_2, \dots, \mathbf{e}_k$$
 $k \ll d, \ \mathbf{e}_i \in \mathbb{R}^d$ 

 $k \ll d$ principal components

2. Projection to principal components

$$\mathbf{x}' = \begin{bmatrix} x_1' \\ x_2' \\ \vdots \\ x_k' \end{bmatrix} = \begin{bmatrix} (\mathbf{x} - \boldsymbol{\mu})^T \mathbf{e}_1 \\ (\mathbf{x} - \boldsymbol{\mu})^T \mathbf{e}_2 \\ \vdots \\ (\mathbf{x} - \boldsymbol{\mu})^T \mathbf{e}_k \end{bmatrix} = \begin{bmatrix} (\mathbf{x} - \boldsymbol{\mu})^T \mathbf{e}_1 \\ (\mathbf{x} - \boldsymbol{\mu})^T \mathbf{e}_2 \\ \vdots \\ (\mathbf{x} - \boldsymbol{\mu})^T \mathbf{e}_k \end{bmatrix} = \begin{bmatrix} (x_1 - \mu_1)e_{1,1} + (x_2 - \mu_2)e_{1,2} + \dots + (x_d - \mu_d)e_{1,d} \\ (x_1 - \mu_1)e_{2,1} + (x_2 - \mu_2)e_{2,2} + \dots + (x_d - \mu_d)e_{2,d} \\ \vdots \\ (x_1 - \mu_1)e_{k,1} + (x_2 - \mu_2)e_{k,2} + \dots + (x_d - \mu_d)e_{k,d} \end{bmatrix}$$

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More details available here:

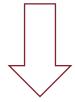
https://github.com/gtolomei/big-data-computing/raw/master/extra/Notes on Principal Component Analysis.pdf

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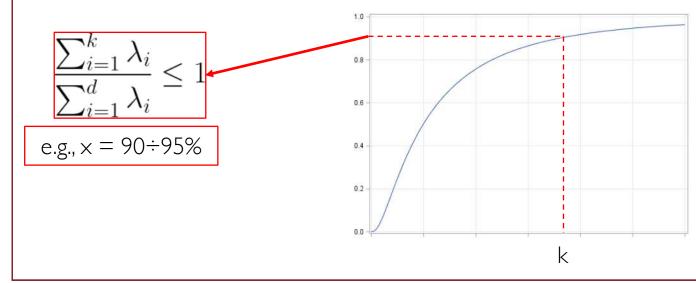
Pick the subset of k eigenvectors that "explain" the most variance

I. Sort eigenvectors by eigenvalues such that  $\lambda_1 \geqslant \lambda_2 \geqslant ... \geqslant \lambda_d$ 

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2. Pick the first k eigenvectors that explain x% of the total variance



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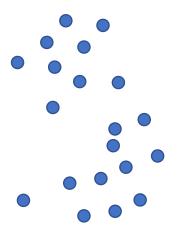
#### Solution

Normalize each dimension to 0-mean and 1-std-deviation

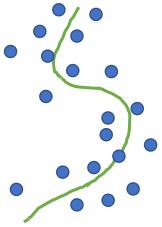
$$z = \frac{x - \mu}{\sigma}$$

- PCA assumes the projection subspace is linear, i.e., an hyperplane:
  - I-d → straight line, 2-d → flat surface, ...

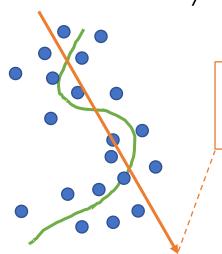
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PCA will find a straight line and will not mimic non-linearity

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- If data do not live on a linear subspace PCA may not work well