

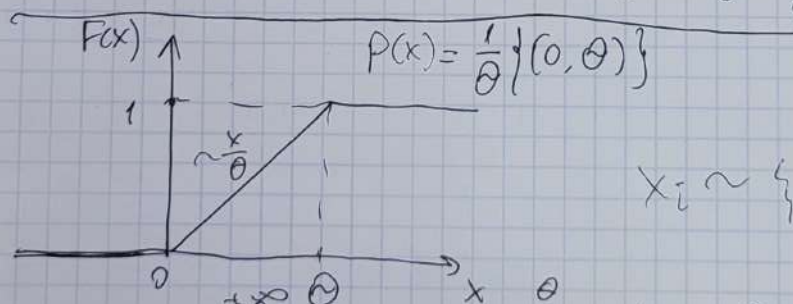
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Т1 Сл. вел. распределены равномерно на  $[0, \theta]$   
По выборке  $n$  найденные оценки параметра  $\theta$ :

$$\tilde{\theta}_1 = 2\bar{x}, \quad \tilde{\theta}_2 = x_{\min}, \quad \tilde{\theta}_3 = x_{\max},$$

$$\tilde{\theta}_5 = \left( x_1 + \frac{\sum_{k=2}^n x_k}{(n-1)} \right)$$

- а) Проверить оценки на несмещенность / сопоставленность  
б) Найти наиболее эффективную оценку



$$а) M[\xi] = \int_{-\infty}^{\infty} x p(x) dx = \int_0^{\theta} \frac{x}{\theta} dx = \frac{\theta}{2}$$

$$M[\xi^2] = \int_{-\infty}^{\infty} x^2 p(x) dx = \int_0^{\theta} \frac{x^2}{\theta} dx = \frac{\theta^2}{3}$$

$$D[\xi] = M[\xi^2] - M^2[\xi] = \frac{\theta^2}{3} - \frac{\theta^2}{4} = \frac{\theta^2}{12}$$

$$= \theta \cdot \left[1 - \frac{1}{n+1}\right] = \frac{\theta}{n+1} \text{ — сходимое}$$

$$\tilde{\theta}_2' = (n+1) X_{\min} \quad M[\tilde{\theta}_2'] = \theta$$

$$\begin{aligned} M[\tilde{\theta}_2^2] &= M[\tilde{\theta}_2^2(\bar{X}_n)] = \int_0^\theta n \left(1 - \frac{y}{\theta}\right)^{n-1} \frac{1}{\theta} y^2 dy = \\ &= \left\{ t = 1 - \frac{y}{\theta} \right\} = - \int_1^0 n \theta^2 (t^{n-1} - 2t^n + t^{n+1}) dt = \\ &= n \theta^2 \left( \frac{1}{n} - 2 \frac{1}{n+1} + \frac{1}{n+2} \right) = \frac{2\theta^2}{(n+1)(n+2)} \end{aligned}$$

$$D[\tilde{\theta}_2] = \frac{2\theta^2}{(n+1)(n+2)} - \frac{\theta^2}{(n+1)^2} = \frac{\theta^2 n}{(n+1)^2(n+2)}$$

$$D[\tilde{\theta}_2'] = (n+1)^2 D[\tilde{\theta}_2] = \frac{\theta^2 n}{n+2} \xrightarrow{n \rightarrow \infty} 0$$

$\tilde{\theta}_2'$  — no overestimate:

доп. укл  
не требуется

$$\forall \theta > 0 \quad \forall \varepsilon > 0 \quad P(|\tilde{\theta}_2' - \theta| \geq \varepsilon) \xrightarrow{n \rightarrow \infty} 0$$

$$P(|\tilde{\theta}_2' - \theta| \geq \varepsilon) \geq$$



$$\geq P(\tilde{\theta}_2' \geq \theta + \varepsilon) = P_{(n+1)X_{\min} \geq \theta + \varepsilon} =$$

$$= P(X_{\min} \geq \frac{\theta + \varepsilon}{n+1}) = 1 - P(X_{\min} < \frac{\theta + \varepsilon}{n+1}) =$$

$$= 1 - (1 - (1 - F(\frac{\theta + \varepsilon}{n+1}))^n) = (1 - F(\frac{\theta + \varepsilon}{n+1}))^n =$$

$$\geq \zeta \quad \varepsilon: \frac{\theta + \varepsilon}{n+1} < \theta \quad \left\{ = \left(1 - \frac{\theta + \varepsilon}{\theta} \cdot \frac{1}{n+1}\right)^n \xrightarrow{n \rightarrow \infty} e^{-\frac{\theta + \varepsilon}{\theta}} \right.$$



— не абсолютная состоятельность

$$\begin{aligned}\tilde{\theta}_2: P(|\tilde{\theta}_2 - \theta| \geq \varepsilon) &= P(\tilde{\theta}_2 < \theta - \varepsilon + \\ &+ P(\tilde{\theta}_2 > \theta + \varepsilon)) = P(X_{\min} < \theta - \varepsilon) = \\ &= \Phi(\theta - \varepsilon) = 1 - (1 - \frac{\theta - \varepsilon}{\theta})^n = 1 - (\frac{\varepsilon}{\theta})^n \xrightarrow{n \rightarrow \infty} 1\end{aligned}$$

→ не состоятельность.

$$\begin{aligned}3) \tilde{\theta}_3^* &= X_{\max}, \quad \tilde{\varphi}(y) = (F(y))^n \\ \tilde{\varphi}(y) &= n(F(y))^{n-1} p(y) = n(\frac{y}{\theta})^{n-1} \theta^{-1} = \frac{n y^{n-1}}{\theta^n} \\ M[\tilde{\theta}_3(X_n)] &= \int_{-\infty}^{\infty} y \cdot \tilde{\varphi}(y) dy = \int_0^{\theta} n \cdot \frac{y^n}{\theta^n} dy = \frac{n}{n+1} \theta\end{aligned}$$

— состоятельность.

$$\begin{aligned}\tilde{\theta}_3^1 &= \frac{n+1}{n} X_{\max} \\ M[\tilde{\theta}_3^2(X_n)] &= \frac{n y^{n+2}}{n+2 \theta^n} \Big|_0^{\theta} = \frac{n}{n+2} \theta^2 \\ D[\tilde{\theta}_3] &= \frac{n}{n+2} \theta^2 - \frac{n^2}{(n+1)^2} \theta^2 = \frac{n \theta^2}{(n+1)(n+1)^2} \\ D[\tilde{\theta}_3^1] &= \frac{(n+1)^2}{n^2} \cdot D[\tilde{\theta}_3] = \frac{\theta^2}{n(n+2)} \xrightarrow{n \rightarrow \infty} 0\end{aligned}$$

Поопы:  $\tilde{\theta}_3$ :

$$\forall \theta > 0 \quad \forall \varepsilon > 0 \quad P(|\tilde{\theta}_3 - \theta| \geq \varepsilon) = 0$$

→ не состоятельность:

$\tilde{\theta}_3^1$  — состоятельность

$$\begin{aligned}
 &= P_{\max}(X_{\max} \leq \theta - \varepsilon) + P(X_{\max} \geq \theta + \varepsilon) = \\
 &= (F(\theta - \varepsilon))^n; \quad 0 < \varepsilon < \theta: \quad P(|\tilde{\theta}_3 - \theta| \geq \varepsilon) = \left(\frac{\theta - \varepsilon}{\theta}\right)^n = \\
 &= \left(1 - \frac{\varepsilon}{\theta}\right)^n \xrightarrow{n \rightarrow \infty} 0
 \end{aligned}$$

$$\begin{aligned}
 \varepsilon \geq 0: \quad P(|\tilde{\theta}_3 - \theta| \geq \varepsilon) &= 0^n = 0 \xrightarrow{n \rightarrow \infty} 0 \\
 \Rightarrow \tilde{\theta}_3 - \text{по сур.} &\text{состоятельна}
 \end{aligned}$$

$\tilde{\theta}_3'$  - по сурегену:  $(\Pi \Theta \Lambda \rho \sigma \kappa)$

$$\begin{aligned}
 P(|\tilde{\theta}_3' - \theta| \geq \varepsilon) &= P\left(\left(\frac{n+1}{n}\right) X_{\max} \leq \theta + \varepsilon\right) + \\
 &+ P\left(X_{\max} \left(\frac{n+1}{n}\right) \geq \theta + \varepsilon\right) = \\
 &= 1 - P\left(X_{\max} < \frac{\theta + \varepsilon}{\frac{n+1}{n}}\right) + P\left(X_{\max} \leq \frac{\theta - \varepsilon}{\frac{n+1}{n}} \cdot n\right) = \\
 &= 1 - \underbrace{\left(F\left((\theta + \varepsilon) \cdot \frac{n}{n+1}\right)\right)^n}_I + \underbrace{\left(F\left((\theta - \varepsilon) \cdot \frac{n}{n+1}\right)\right)^n}_{II}
 \end{aligned}$$

$$\boxed{I}: \text{• при } \varepsilon \geq \frac{\theta}{n}: 1 - \left(F\left((\theta - \varepsilon) \cdot \frac{n}{n+1}\right)\right)^n \stackrel{II}{=} 0 \Rightarrow$$

$$\boxed{II}: \left(F\left((\theta - \varepsilon) \cdot \frac{n}{n+1}\right)\right)^n = 0, \varepsilon \geq 0 \quad \Rightarrow \lim_{n \rightarrow \infty} \frac{1}{n} = 0$$

$$0 < \varepsilon < \theta: \quad \left(\frac{\theta - \varepsilon}{\theta} \cdot \frac{n}{n+1}\right)^n = \left(1 - \frac{1}{n+1}\right)^n \cdot \left(1 - \frac{\varepsilon}{\theta}\right)^n \xrightarrow{n \rightarrow \infty} 0$$



$$\begin{aligned} & \varepsilon > 0 \Rightarrow \lim_{n \rightarrow \infty} \mathbb{P}(\varepsilon > 0) \\ & \Rightarrow P(|\hat{\theta}_3' - \theta_3| \geq \varepsilon) \xrightarrow{n \rightarrow \infty} 0 \quad \forall \varepsilon > 0 \\ & \Rightarrow \theta_3' - \text{состояние по определению} \end{aligned}$$

$$\begin{aligned} 4) \quad \tilde{\theta}_5 &= X_1 + \frac{1}{n-1} \sum_{k=2}^n X_k \quad \begin{matrix} M[\xi] \\ // \end{matrix} \\ M[\tilde{\theta}_5] &= M\left[X_1 + \frac{1}{n-1} \sum_{k=2}^n X_k\right] = M[X_1] + \frac{1}{n-1} \sum_{k=2}^n M[X_k] \\ &= \frac{\theta}{2} + \frac{1}{n-1} \cdot (n-1) \cdot \frac{\theta}{2} = \theta - \text{расчет} \\ D[\tilde{\theta}_5] &= D\left[X_1 + \frac{1}{n-1} \sum_{k=2}^n X_k\right] = D[\xi] + \frac{1}{n-1} D[\xi] \\ &= \frac{\theta^2}{12} \cdot \frac{n}{n-1} \xrightarrow{n \rightarrow \infty} 0 \end{aligned}$$

$$\left\{ \begin{aligned} \xi_n &\xrightarrow{P} \xi, \quad \eta_n \xrightarrow{P} \eta \Rightarrow \xi_n + \eta_n \xrightarrow{P} \xi + \eta \\ X_1 &\xrightarrow{P} \xi = X_1, \quad \frac{1}{n-1} \sum_{k=2}^n X_k \rightarrow M[\xi] = \frac{\theta}{2} \\ \text{По ЗБЧ Хинчина } \{X_k\}_{k=2}^n &\text{ - н.з. сгруппировать} \end{aligned} \right\}$$

$$\Rightarrow \tilde{\theta}_4 \longrightarrow X_1 + \frac{\theta}{2} \Rightarrow \text{не явл. состоянием}$$

$\uparrow$   
 с.л. вел

$$\delta) \quad \tilde{\Theta}_1 = 2\bar{x} \rightarrow D[\tilde{\Theta}_1] = \frac{\Theta^2}{3n}$$

$$\tilde{\Theta}_3' = \frac{n+1}{n} x_{\max} \rightarrow D[\tilde{\Theta}_3'] = \frac{\Theta^2}{n(n+2)}$$

$$\frac{\Theta^2}{n(n+2)} < \frac{\Theta^2}{3n}, \quad n > 1 - \underline{\text{выполнено}}$$

$\Rightarrow \tilde{\Theta}_3' - \text{превосходит } \Theta_1$

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