

TOLES 1999 The Buffalo News. Reprinted with permission of UNIVERSAL PRESS SYNDICATE. All rights reserved.

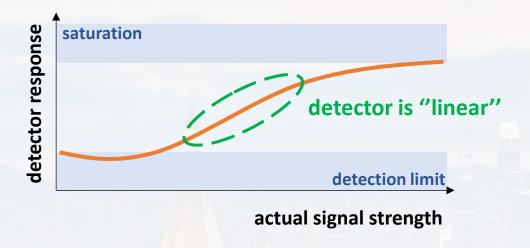
Outline:

Error Estimation





- **systematic errors:** calibration, non-linearity of the detector



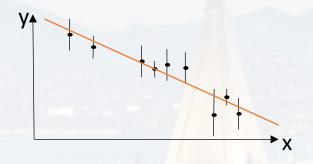
statistical errors: limited precision, natural variance of the data
 → spread of the data around an average value



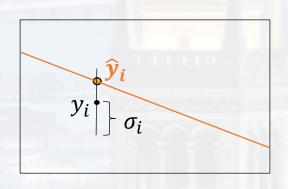


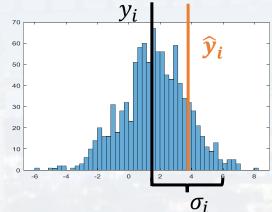
statistical errors: limited precision, natural variance of the data→ spread of the data around an average value

assumption: far from the detection limit and the saturation → the spread follows a normal distribution



$$P(x|\mu,\sigma) = \frac{1}{\sqrt{2\pi \sigma^2}} exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right]$$



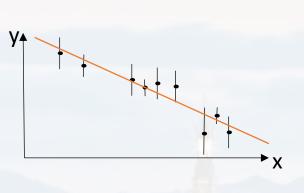


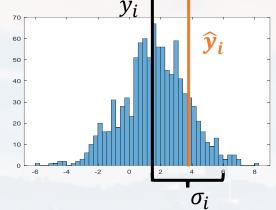
$$p_i(y_i|\hat{\mathbf{y}_i}, \sigma_i) = \frac{1}{\sqrt{2\pi}\sigma_i} exp \left[-\frac{1}{2} \frac{(y_i - \hat{\mathbf{y}_i})^2}{{\sigma_i}^2} \right]$$





<u>assumption</u>: far from the detection limit and the saturation → the spread follows a **normal distribution**



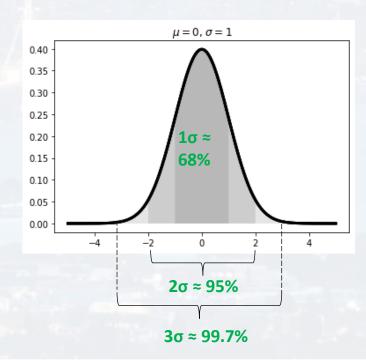


$$p_i(y_i|\hat{\mathbf{y}_i}, \sigma_i) = \frac{1}{\sqrt{2\pi}\sigma_i} exp \left[-\frac{1}{2} \frac{(y_i - \hat{\mathbf{y}_i})^2}{{\sigma_i}^2} \right]$$

for large (> 50...100) N (number of data points):

≈ 2/3 of the data points should be consistent with the model within their 1σ error bars

≈ 95% of the data points should be consistent with the model within their 2σ error bars







for large (> 50...100) N (number of data points):

≈ 2/3 of the data points should be consistent with the model within their 1σ error bars

 \approx 95% of the data points should be consistent with the model within their 2 σ error bars

$$\chi_{red}^2 = \frac{1}{df} \sum_{i=1}^{N} \left(\frac{y_i - \hat{y}_i}{\sigma_i} \right)^2 \qquad df = N - p - 1$$

rule of thumb:

reduced $\chi^2 \approx 1.0$ excellent fit

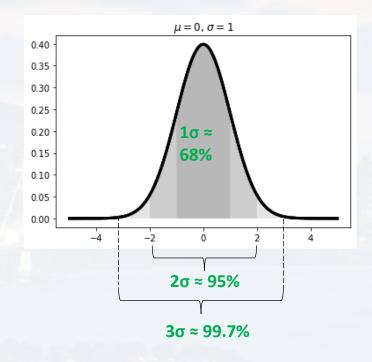
1.0...1.5 acceptable fit

1.5...1.7 bad fit

>2.0 not acceptable

<<1.0 suspicious, errors are overestimated!

$$p_{i}(y_{i}|\widehat{y}_{i},\sigma_{i}) = \frac{1}{\sqrt{2\pi}\sigma_{i}}exp\left[-\frac{1}{2}\frac{(y_{i}-\widehat{y}_{i})^{2}}{\sigma_{i}^{2}}\right]$$



 y_i : measured value of data point

 σ_i : statistical error of y_i (often aka ey_i)

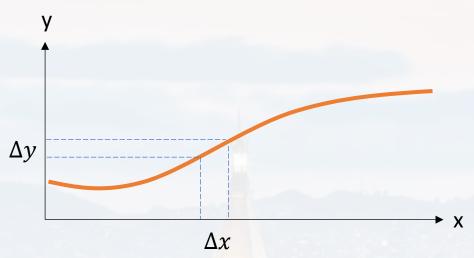
 \hat{y}_i : prediction by the model *after the fit*

N : number of data points

p: number of fit parameter







$$\frac{dy}{dx} \approx \frac{\Delta y}{\Delta x}$$

for $\Delta x \ll x$

$$\Delta x \left| \frac{dy}{dx} \right| \approx \Delta y$$

example:

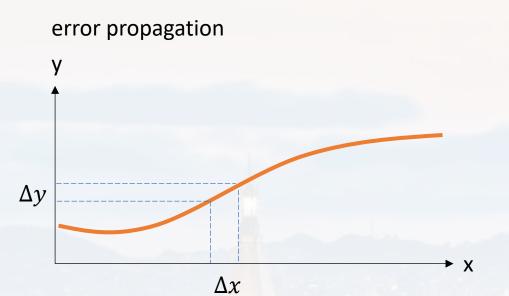
$$V = \frac{4}{3} \pi r^3$$
 $\Delta V = ?$ given Δr

$$\Delta V = \frac{dV}{dr} \Delta r = 4 \pi r^2 \Delta r$$

$$\frac{\Delta V}{V} = 3 \frac{\Delta r}{r}$$
 $\Delta r, \Delta V \approx 1\sigma$







general:

$$\Delta f(max) = \sum_{i=1}^{I} \left| \frac{\partial f}{\partial x_i} \right| \Delta x_i$$
 maximum error estimation

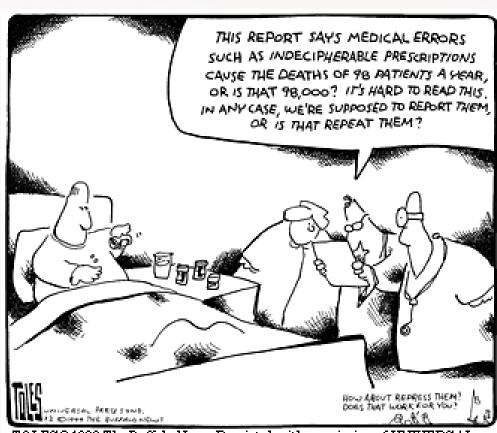
if x_i do not correlate, i. e. are mutually independent:

$$\Delta f^2 = \sum_{i=1}^{I} \left| \frac{\partial f}{\partial x_i} \right|^2 (\Delta x_i)^2$$

Note:
$$\Delta f(max)^2 > \Delta f^2$$
 because of the mixed terms $\left| \frac{\partial f}{\partial x_i} \right| \left| \frac{\partial f}{\partial x_j} \right| \Delta x_i \Delta x_j$ in $\Delta f(max)^2$







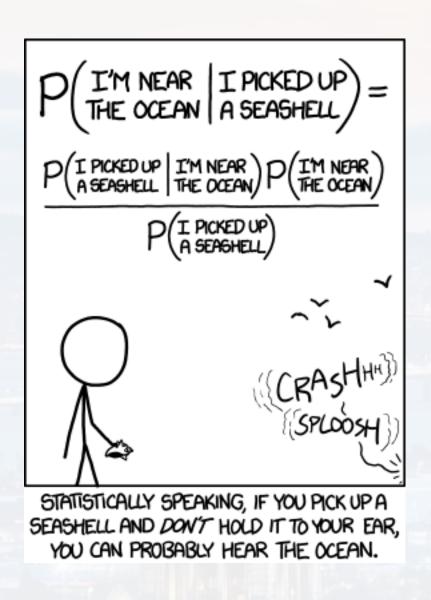
TOLES 1999 The Buffalo News. Reprinted with permission of UNIVERSAL PRESS SYNDICATE. All rights reserved.

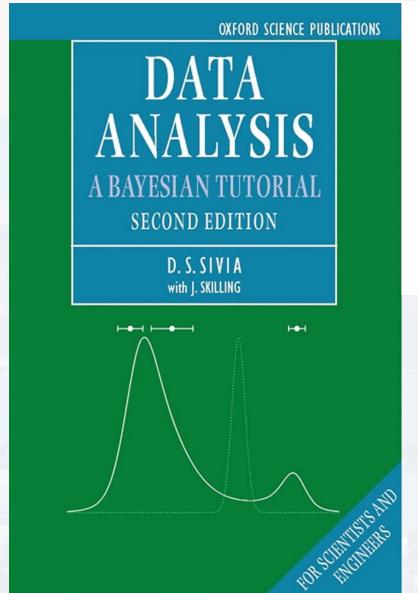
Outline:

Bayesian Statistics













 $P(A \cap B)$ probability **P** that the events **A** and **B** occur

so far: A and B were independent $P(A \cap B) = P(A)P(B) = P(B)P(A)$

now: conditional probabilities | "given" or "under the condition"



Thomas Bayes (1701 - 1761)

$$P(A \cap B) = P(A|B)P(B)$$

$$= P(B|A)P(A)$$

$$= P(B|A)P(A)$$

$$= P(B|A)P(B)$$

Bayes Theorem

posterior
$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$
 prior





$$P(A|B)P(B) = P(B|A)P(A)$$

Bayes Theorem

posterior
$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$
 prior



Thomas Bayes (1701 - 1761)

$$X_1$$
 X_2
 \vdots
 X_N

$$P(B) = \sum_{n=1}^{N} P(B|X_n)P(X_n)$$

$$P(B) = \int P(B|X)P(X) dX$$

marginalization

example:

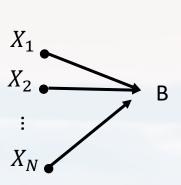
model: N

data:

 $P(D|M) = \int P(D|all \ model \ param, M) \ P(all \ model \ param|M) \ d \ all \ model \ param$







$$P(B) = \sum_{n=1}^{N} P(B|X_n)P(X_n)$$

$$P(B) = \int P(B|X)P(X) dX$$

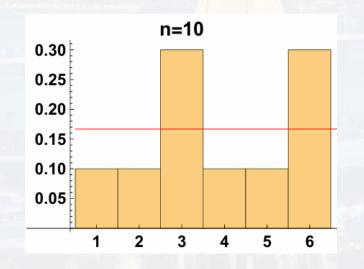
marginalization



Thomas Bayes (1701 - 1761)

for a normal distribution $\mathcal{N}(\mu, \sigma)$

$$P(D|\mathcal{N}(\mu,\sigma)) = \int P(D|\mathcal{N}(\mu,\sigma)) P(\mu,\sigma|\mathcal{N}(\mu,\sigma)) d \Omega_{\mu,\sigma}$$



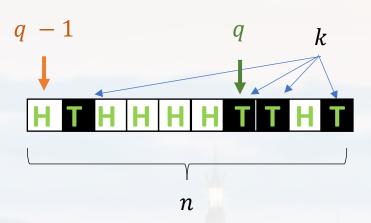
$$\sigma = 2, \mu = 5.0$$

$$\sigma = 1.5$$
, $\mu = 3.5$

$$\sigma = 7.0$$
, $\mu = 1.0$







probability of having a sequence of k tails and n-k heads

$$p_{tot} = \prod_{i} q_i^{n_i} = q^k (1 - q)^{n - k}$$

probability of having any sequence of k tails and n-k heads

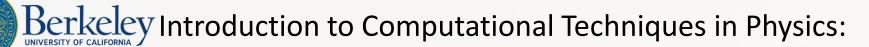
$$P(k|q,n) = \binom{n}{k} q^k (1-q)^{n-k}$$

binomial distribution

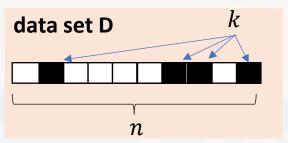


fair coin? q = 0.5 ???

$$\frac{n!}{k!(n-k)!} =: {n \choose k}$$
 "in choose k"



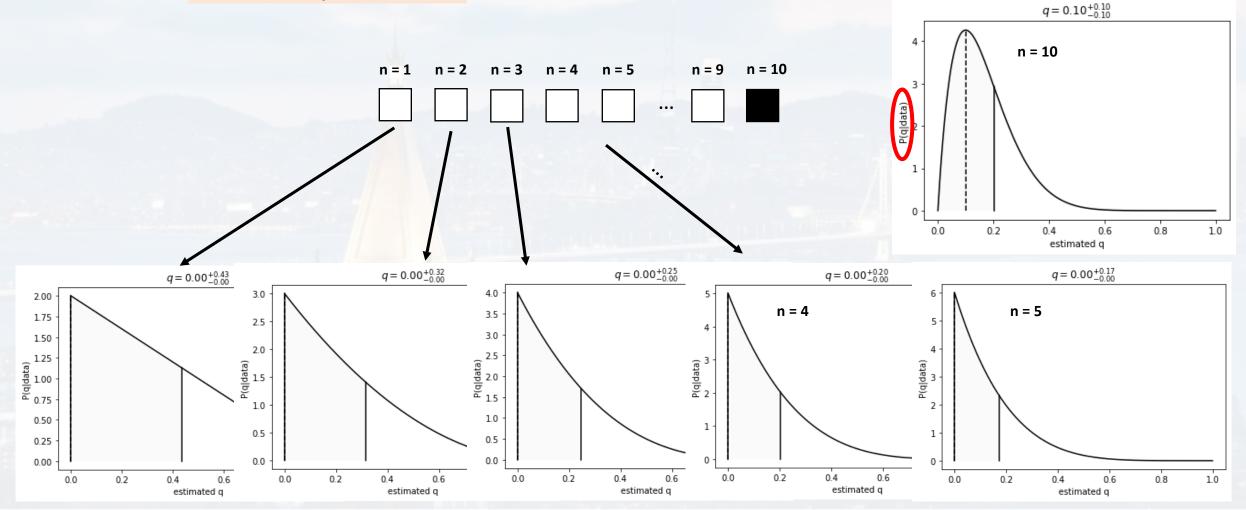




q = ?

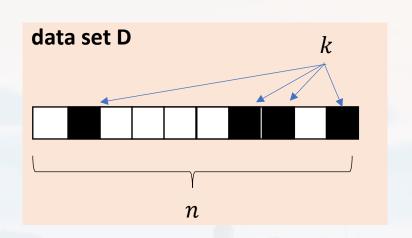
goal:

- P(q|D) Parameter Estimation
- the larger **D**, the more certain **q** → learning









q = ?

goal:

- P(q|D)

the larger *D*, the more certain *q* → learning

$$P(k|n,q) = \binom{n}{k} q^k (1-q)^{n-k}$$

Bayes' theorem:

likelihood function (here: binomial)

$$P(q|data set) = \frac{P(data set|q)P(q) \text{ prior}}{P(data set) \text{ evidence (const wrt q)}}$$

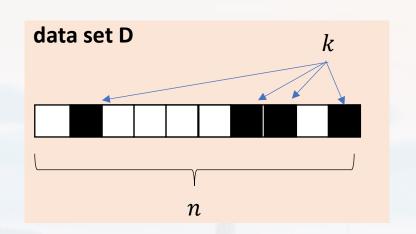
$$=\frac{\binom{n}{k}q^k(1-q)^{n-k}}{P(D)}P(q)$$

$$\sim q^k (1-q)^{n-k} P(q)$$

P(D) and $\binom{n}{k}$ are no functions of q







$$q = ?$$

goal: - P(q|D)

the larger *D*, the more certain *q* → learning

$$P(k|n,q) = \binom{n}{k} q^k (1-q)^{n-k}$$

$$P(q|data set) = \frac{P(data set|q)P(q)}{P(data set)}$$

$$= \frac{\binom{n}{k}q^{k}(1-q)^{n-k}}{P(D)}P(q)$$

$$\sim q^k(1-q)^{n-k}P(q)$$

max. entropy: P(q) = const if no prior information about q

$$\sim q^k (1-q)^{n-k}$$

$$P(q|data set) = \frac{q^{k}(1-q)^{n-k}}{\int_{0}^{1} q^{k}(1-q)^{n-k} dq}$$





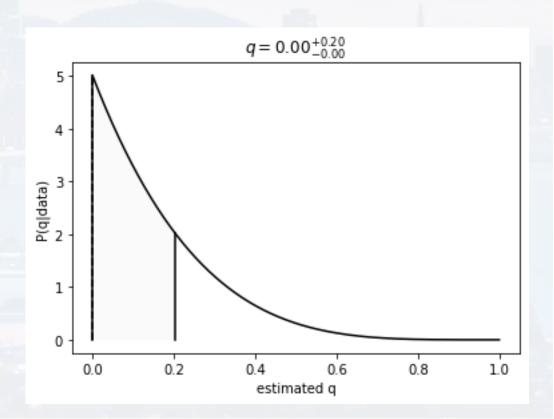
check out bayesian_bino.py

$$n1 = 4$$

k1 = np.random.binomial(n1, 0.25)

creating artificial data set note: in reality **q** is unknown!

$$P(q|data set) = \frac{q^{k}(1-q)^{n-k}}{\int_{0}^{1} q^{k}(1-q)^{n-k} dq}$$



 $q = 0.2^{+0.05}_{-0.06}$

0.5

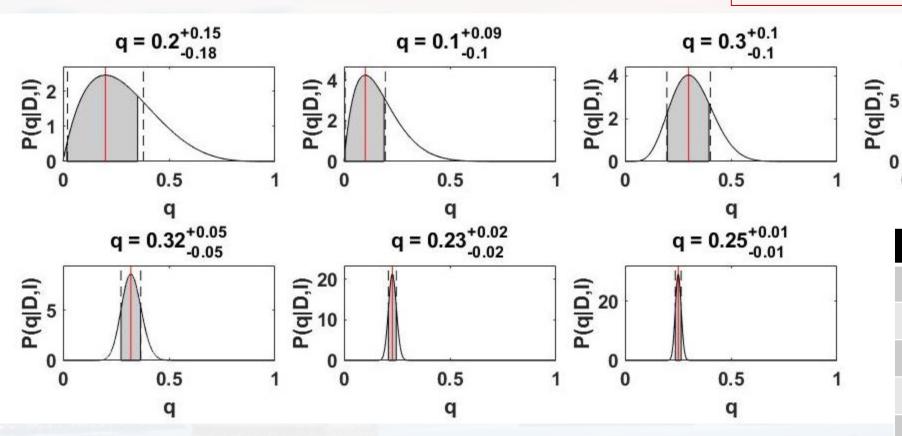


Berkeley Introduction to Computational Techniques in Physics:



check out bayesian_bino.py

$$P(q|data set) = \frac{q^{k}(1-q)^{n-k}}{\int_{0}^{1} q^{k}(1-q)^{n-k} dq}$$



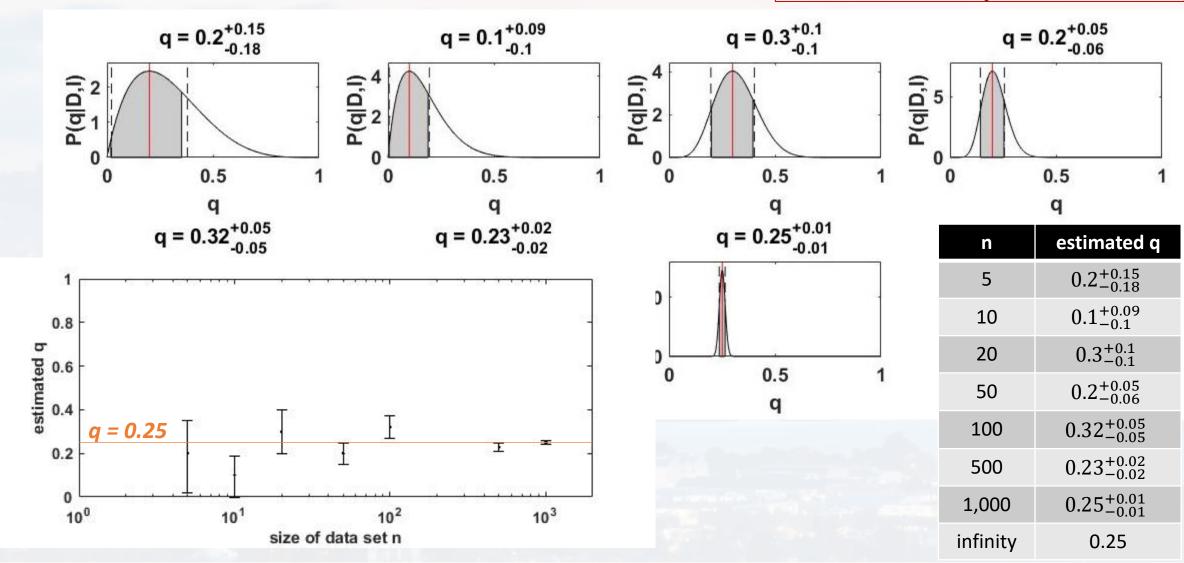
q	
n	estimated q
5	$0.2^{+0.15}_{-0.18}$
10	$0.1^{+0.09}_{-0.1}$
20	$0.3^{+0.1}_{-0.1}$
50	$0.2^{+0.05}_{-0.06}$
100	$0.32^{+0.05}_{-0.05}$
500	$0.23^{+0.02}_{-0.02}$
1,000	$0.25^{+0.01}_{-0.01}$
infinity	0.25





check out bayesian_bino.py

$$P(q|data set) = \frac{q^{k}(1-q)^{n-k}}{\int_{0}^{1} q^{k}(1-q)^{n-k} dq}$$







Of course, Bayesian Parameter Estimation works with **any other pdf**

goal: - P(q|D)

the larger *D*, the more certain *q* → learning

likelihood function

$$P(q|data set) = \frac{P(data set|q)P(q) \text{ prior}}{P(data set) \text{ evidence (const wrt q)}}$$

What is the average number of WhatsUp messages I get every day?

Mon: 5 event - has no duration

- is rare

data = np.random.poisson(lam = 0.4, 15)
poissfit(data)

Wed: 1

Tue:

Thu: 3

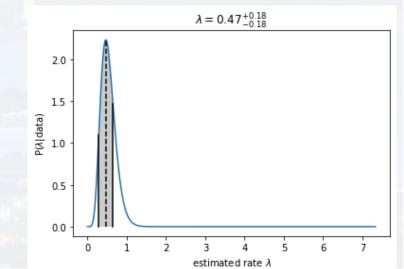
Fri: 9

Sat: 2

Sun: 5

→ Poissonian

$$P(k|\lambda) = \frac{\lambda^k}{k!} e^{-\lambda}$$







Of course, Bayesian Parameter Estimation works with **any other pdf**

goal: - P(q|D)

- the larger D, the more certain q \rightarrow learning

What is the average number of WhatsUp messages I get every day?

Mon: 5 event - has no duration

Tue: 7 - is rare

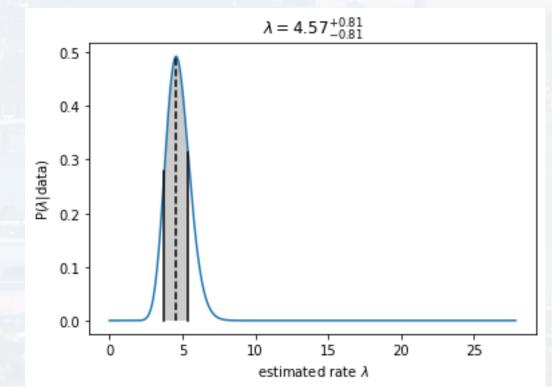
Wed: 1

Thu: 3 \rightarrow Poissonian

Fri: 9

Sat: 2 Sun: 5 $P(k|\lambda) = \frac{\lambda^k}{k!} e^{-\lambda}$

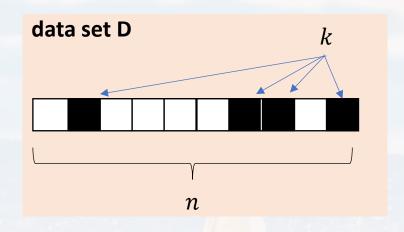
poissfit([5, 7, 1, 3, 9, 2, 5])



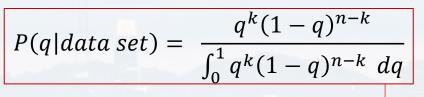


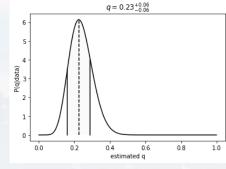


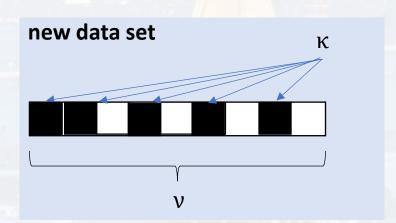
What if there is new data?











if there **is** prior information **I** about **q**:

$$P(q|new\ data\ set,I) = \frac{P(new\ data\ set|q,I)\ P(q,I)}{P(new\ data\ set)}$$





What if there is new data?

$$P(q|new\ data\ set,I) = \frac{P(new\ data\ set|q,I)}{P(new\ data\ set)}$$

$$P(q|data set) = \frac{q^{k}(1-q)^{n-k}}{\int_{0}^{1} q^{k}(1-q)^{n-k} dq}$$

$$= \frac{q^{\kappa} (1-q)^{\nu-\kappa}}{\int_0^1 q^{\kappa} (1-q)^{\nu-\kappa}} \frac{q^k (1-q)^{n-k}}{q^k (1-q)^{n-k}} dq$$

$$= \frac{q^{k+\kappa}(1-q)^{\nu-\kappa+n-k}}{\int_0^1 q^{k+\kappa}(1-q)^{\nu-\kappa+n-k} dq}$$

$$= \frac{q^{k+\alpha-1}(1-q)^{n-k+\beta-1}}{\int_0^1 q^{k+\alpha-1}(1-q)^{n-k+\beta-1} dq}$$

often: $\kappa = \alpha - 1$ $\beta = \nu - \kappa - 1$

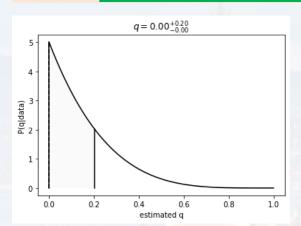
Beta function



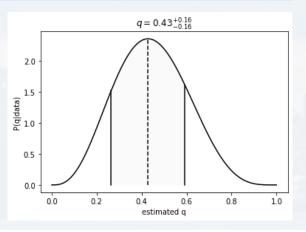


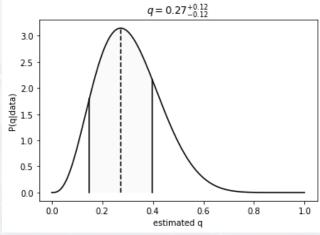
What if there is new data?

$$P(q|new\ data\ set, I) = \frac{q^{\kappa}(1-q)^{\nu-\kappa}}{\int_{0}^{1} q^{\kappa}(1-q)^{\nu-\kappa}} \frac{q^{k}(1-q)^{n-k}}{q^{k}(1-q)^{n-k}} \frac{q^{k}(1-q)^{n-k}}{q^{k}(1-q)^{n$$



$$P(q,I) = \frac{q^{k}(1-q)^{n-k}}{\int_{0}^{1} q^{k}(1-q)^{n-k} dq}$$





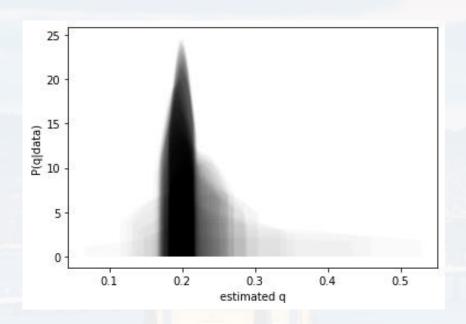




What if there is new data?

$$P(q|new\ data\ set, I) = \frac{q^{\kappa}(1-q)^{\nu-\kappa}}{\int_{0}^{1} q^{\kappa}(1-q)^{\nu-\kappa}} \frac{q^{k}(1-q)^{n-k}}{q^{k}(1-q)^{n-k}} dq$$

The posterior from the previous experiment is the prior of the next experiment



- → we become more certain about the model parameters
- → learning!
- → see e.g. Variational Auto Encoders

2D images → 3D objects



credit: StableAI

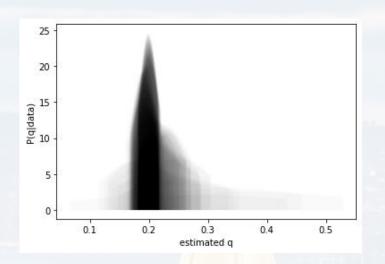




What if there is new data?

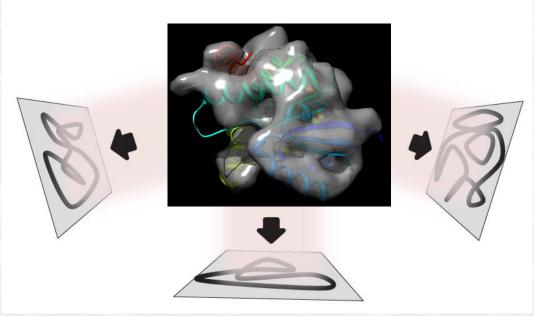
$$P(q|new\ data\ set,I) = \frac{q^{\kappa}(1-q)^{\nu-\kappa}}{\int_{0}^{1} q^{\kappa}(1-q)^{\nu-\kappa}} \frac{q^{k}(1-q)^{n-k}}{q^{k}(1-q)^{n-k}} dq$$

The posterior from the previous experiment is the prior of the next experiment



- → we become more certain about the model parameters
- → learning!
- → see e.g. Variational Auto Encoders 2D images → 3D objects

Cryo – EM: 3D structure from 2D images/projections



(image courtesy: Thomas Becker, GC LMU Munich)



Thank you very much for your attention!



