

Lecture 11:

Data Fitting and Regression



Markus Hohle
University California, Berkeley

Numerical Methods for Computational Science



Numerical Methods for Computational Science

Course Map

Week 1: Introduction to Scientific Computing and Python Libraries

Week 2: Linear Algebra Fundamentals

Week 3: Vector Calculus

Week 4: Numerical Differentiation and Integration

Week 5: Solving Nonlinear Equations

Week 6: Probability Theory Basics

Week 7: Random Variables and Distributions

Week 8: Statistics for Data Science

Week 9: Eigenvalues and Eigenvectors

Week 10: Simulation and Monte Carlo Method

Week 11: Data Fitting and Regression

Week 12: Optimization Techniques

Week 13: Machine Learning Fundamentals

Berkeley Numerical Methods for Computational Science:



<u>Outline</u>

Linear Regression

- The Math
- What is Linear?
- Stats
- a Python Example

Logistic Regression

Curve Fitting

Berkeley Numerical Methods for Computational Science:



<u>Outline</u>

Linear Regression

- The Math
- What is Linear?
- Stats
- a Python Example

Logistic Regression

Curve Fitting

problem: a parameter y depends on many **regressors** x_n

goal: finding a model that tells us how y depends on each x_n

deriving a model to **predict** y from new data points

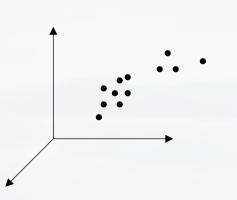
The Math

What is Linear? Stats

a Python Example

		N .					
		x_1	x_2	x_3	x_4	x_5	y_k
	Index	molecular_weight	electronegativity	bond_lengths	num_hydrogen_bonds	logP	toxicity_score
int	0	341.704	2.65585	3.09407	2	9.11147	80.9281
a point	1	335.951	3.22262	2.89039	7	8.92848	83.4911
h data	2	235.203	2.44115	2.48203	1	6.49731	61.8406
kth	3	246.505	2.76656	2.71547	7	7.45089	57.0538
	4	437.939	3.4801	3.59569	3	10.9156	131.326





idea: data point y_k in N dimensional space

$$\rightarrow y_k = f(x_1, \dots x_n, \dots x_N) + \epsilon$$

for each data point *k*

The Math

What is Linear?
Stats
a Python Example

ansatz:

$$y_k = \beta_0 + \sum_{n=1}^N \beta_n x_n + \epsilon$$

linear combination

y: response

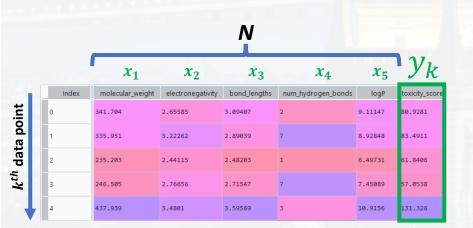
X: regressors (assumed to be independent)

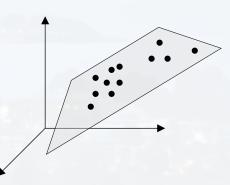
β: factors (how a regressor contributes to the response)

 β_0 : intercept

ε: error (stochasticity of the data, assumed to be

normally dist.)







linear ≠ not curved

$$y_k = \beta_0 + \sum_{n=1}^N \beta_n \, x_n^n + \epsilon$$

...is still linear just define:
$$\bar{x}_n \coloneqq x_n^n$$

The Math

What is Linear?

Stats

a Python Example

$$y_k = \beta_1 \, x_n^{\beta_2}$$

...is still linear

just use log:
$$\bar{y}_k = \log(y_k) = \log(\beta_1) + \beta_2 \log(x_n)$$

= $\bar{\beta}_1 + \beta_2 \bar{x}_n$

y: response x: regressors β: factors β_0 : intercept ε: error

As long as we can recover the linear structure by any transformation → it is linear

in part. log scaling is quite common examples:

- log fold change (DESeq/RNASeq)
- log odds ratio (comparing models, HMM)
- sound → dB is a log unit
- log incidence rates (medical studies)
- percentiles (medical studies)

-....

...what is **not** linear?

$$y_k = \beta_0 + \beta_1 x_n^{\beta_2}$$

log trick does not work here

general: linear refers to the factors

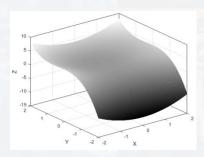
$$y_k = \beta_0 + \beta_1 x_1 + \beta_2 x_2$$

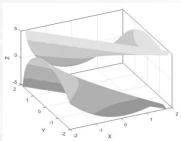
2D plane in 3D space

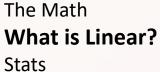
$$y_k = \beta_0 + \beta_1 x_1^2 + \beta_2 x_2^2$$

2D parabolic

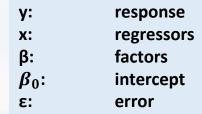
$$y_k = \beta_0 + \beta_1 x_1^2 - \beta_2 x_2^2$$





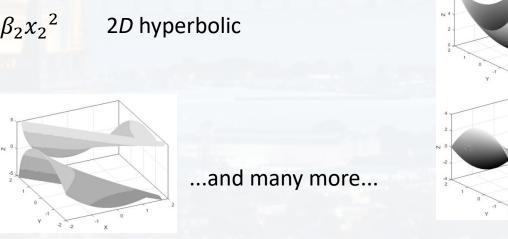


a Python Example



all linear

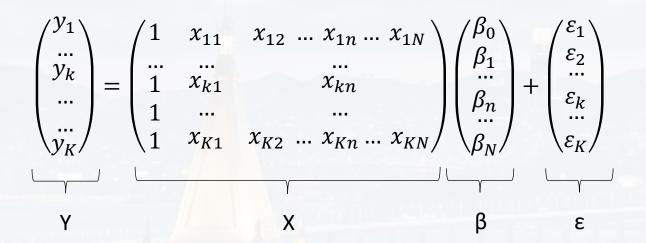
$$y_k = \beta_0 + \sum_{n=1}^N \beta_n x_n + \epsilon$$

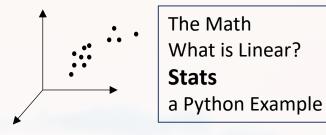




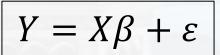
for K data points in N dimensional space

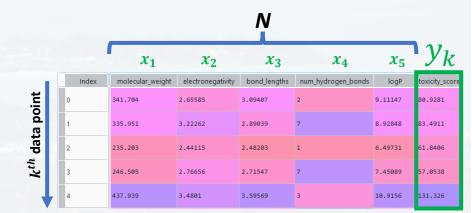
$$y_k = \beta_0 + \sum_{n=1}^N \beta_n x_n + \epsilon$$





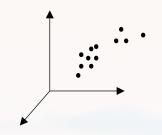
y: response x: regressors β : factors β_0 : intercept ϵ : error





for *K* data points in *N* dimensional space

$$y_k = \beta_0 + \sum_{n=1}^N \beta_n x_n + \epsilon$$



The Math What is Linear? **Stats** a Python Example

$$\begin{pmatrix} y_1 \\ \vdots \\ y_k \\ \vdots \\ y_K \end{pmatrix} = \begin{pmatrix} 1 & x_{11} & x_{12} \dots x_{1n} \dots x_{1N} \\ \vdots & x_{k1} & x_{kn} \\ 1 & x_{k1} & x_{k2} \dots x_{kn} \dots x_{kN} \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_n \\ \vdots \\ \beta_N \end{pmatrix} + \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_k \\ \vdots \\ \varepsilon_K \end{pmatrix}$$

$$Y = X\beta + \varepsilon$$

response regressors factors intercept error

fitting: finding the best β in terms of minimizing the errors

$$(Y - X\beta)^{T}(Y - X\beta) = \sum_{k} \varepsilon_{k}^{2}$$

the model

$$\frac{\partial}{\partial \beta} \sum_{k} \varepsilon_{k}^{2} = 0 \quad \beta_{best} = \hat{\beta} = (X^{T}X)^{-1}X^{T}Y \quad \widehat{Y} = X\hat{\beta} = X(X^{T}X)^{-1}X^{T}Y$$

X and Y are all observables

$$Y = X\beta + \varepsilon$$

finding the best β in terms of minimizing the errors

$$(Y - XB)^T (Y - XB) = \sum \varepsilon_{k}^2$$

$$(Y - X\beta)^T (Y - X\beta) = \sum_k \varepsilon_k^2$$
 the model $\widehat{Y} = X\widehat{\beta} = X(X^TX)^{-1}X^TY$

hat matrix **H**

The Math What is Linear? Stats

a Python Example

response regressors factors β_0 : intercept error

some properties of the hat matrix:

-
$$H = H^T$$
 (symmetry)

-
$$HH = H \rightarrow H^n = H$$
 (idempotency)

evaluating the fit:

$$\hat{\varepsilon} = Y - X\hat{\beta} = Y - \hat{Y} = (I - H)Y$$

$$\hat{\varepsilon}^T \hat{\varepsilon} = [(I - H)Y]^T (I - H)Y = Y^T (I - H)^T (I - H)Y = Y^T (I - H)Y$$

sum of squared errors (SSE)

$$\hat{\sigma}^2 = \frac{\hat{\varepsilon}^T \hat{\varepsilon}}{K - N}$$
 degrees of freedom

variance or

mean of squared errors (MSE)

summary:

the model:

the fit:
$$\hat{Y} = X\hat{\beta} = X(X^TX)^{-1}X^TY$$

 $\hat{\varepsilon}^T \hat{\varepsilon} = Y^T (I - H) Y$ SSE:

variance or MSE:
$$\hat{\sigma}^2 = \frac{\hat{\varepsilon}^T \hat{\varepsilon}}{K-N}$$
 (after the fit)

 $Y = X\beta + \varepsilon$

often fit quality is judged by

Ten fit quality is judged by
$$R^2 \coloneqq 1 - \frac{\sum_k (\hat{y}_k - y_k)^2}{\sum_k (y_k - \langle y \rangle)^2}$$

or adjusted
$$R^2$$
 $\bar{R}^2 := R^2 - (1 - R^2) \frac{N}{K - N - 1}$

and it is said that the fit is good if R^2 is close to one....

...but that is not true...

The Math What is Linear? **Stats**

a Python Example

y: response x: regressors β: factors β_0 : intercept

error

ε:

(after the fit)

$$R^{2} = 1 - \frac{\sum_{i=1}^{N} (y_{i} - \hat{y}_{i})^{2}}{\sum_{i=1}^{N} (y_{i} - \bar{y})^{2}}$$

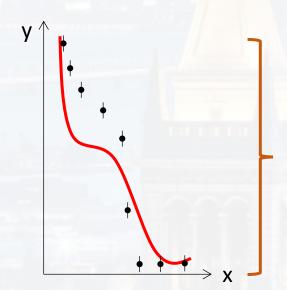
variance data vs model (aka residual sum of squares)

variance of the data (aka total sum of squares)

p:

The Math
What is Linear?
Stats
a Python Example

Note: do not confuse R^2 with Pearsons coefficient: $\rho = \frac{cov(x,y)}{\sqrt{var(x)var(y)}}$



data variance can be huge (i. e. exponential functions) $\rightarrow R^2$ could be around 1.0 even if fit is completely off!

 \bar{y} : mean of the data point values

 \overline{y} : mean of the data point values y_i : measured value of data point σ_i : statistical error of y_i (often aka ey_i) \widehat{y}_i : prediction by the model after the fit number of data points

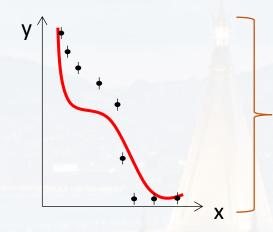
number of fit parameter



$$R^{2} = 1 - \frac{\sum_{i=1}^{N} (y_{i} - \hat{y}_{i})^{2}}{\sum_{i=1}^{N} (y_{i} - \bar{y})^{2}}$$

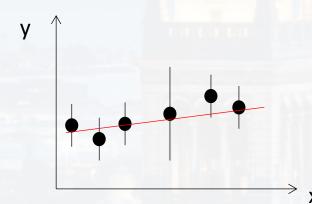
variance data vs model (aka residual sum of squares)

variance of the data (aka total sum of squares) The Math What is Linear? Stats a Python Example



 \bar{y} : mean of the data point values

data variance can be huge (i. e. exponential functions) $\rightarrow R^2$ could be around 1.0 even if fit is completely off!



variance data vs model (aka residual sum of squares)

variance of the data (aka total sum of squares)

although the fit is good!

$$\approx 1 \rightarrow R^2 = 0$$

 \overline{y} : y_i :

mean of the data point values

 σ_i :

measured value of data point statistical error of y_i (often aka ey_i)

 \hat{y}_i : N:

prediction by the model after the fit number of data points

p:

number of fit parameter

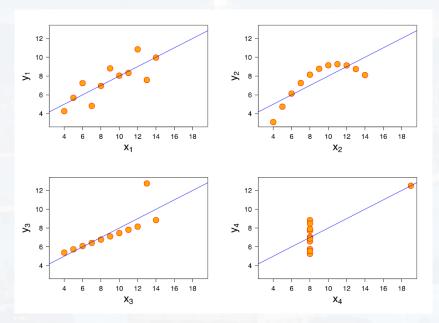


$$R^{2} = 1 - \frac{\sum_{i=1}^{N} (y_{i} - \hat{y}_{i})^{2}}{\sum_{i=1}^{N} (y_{i} - \bar{y})^{2}}$$

variance data vs model (aka residual sum of squares)

variance of the data (aka total sum of squares)

The Math
What is Linear?
Stats
a Python Example



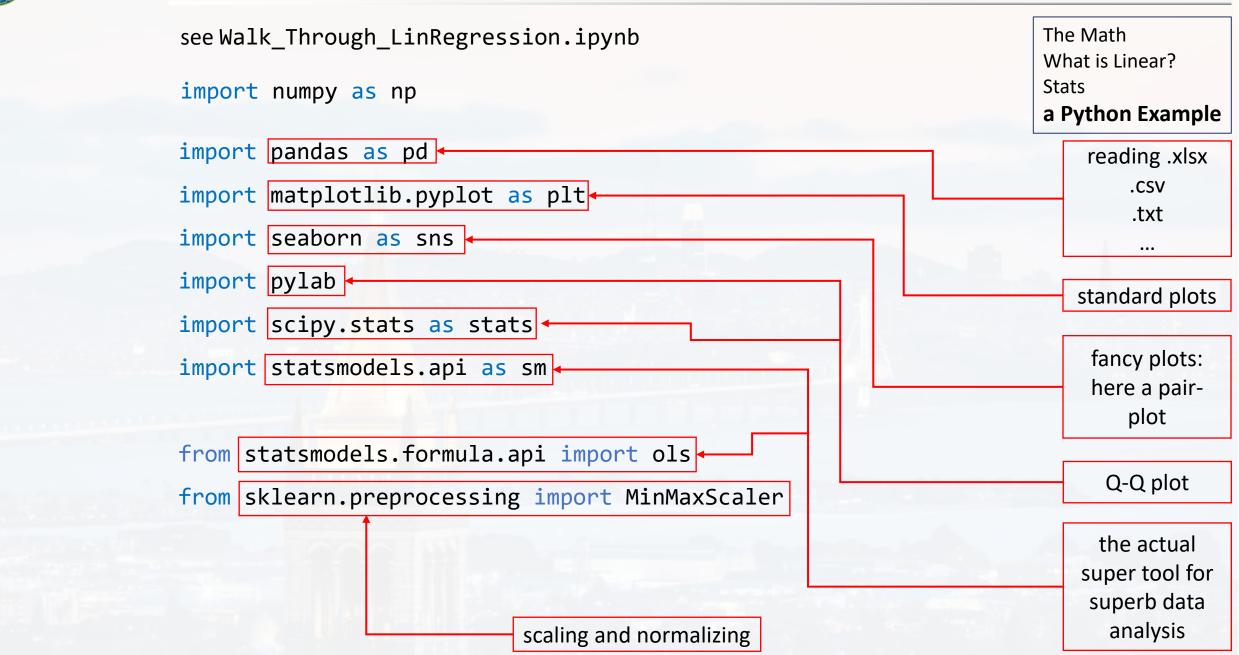
all plots: same R^2

Also, Wiki is full of examples...
...and warnings (see "caveats" therein)

conclusion:

 R^2 is **not** a measure of the fit quality (but χ^2 is)

Given a good fit, R^2 tells how strong the dependent variable responds to the independent variable



	x_1	$\boldsymbol{x_2}$	x_3	x_4	x_5	\mathcal{Y}_k
Index	molecular_weight	electronegativity	bond_lengths	num_hydrogen_bonds	logP	toxicity_score
0	341.704	2.65585	3.09407	2	9.11147	80.9281
1	335.951	3.22262	2.89039	7	8.92848	83.4911
2	235.203	2.44115	2.48203	1	6.49731	61.8406
3	246.505	2.76656	2.71547	7	7.45089	57.0538
4	437.939	3.4801	3.59569	3	10.9156	131.326

$$y_k = \beta_0 + \sum_{n=1}^N \beta_n x_n + \epsilon$$

y: toxicity_score

molecular_weight, electronegativity,
bond_lengths, num_hydrogen_bonds, logP

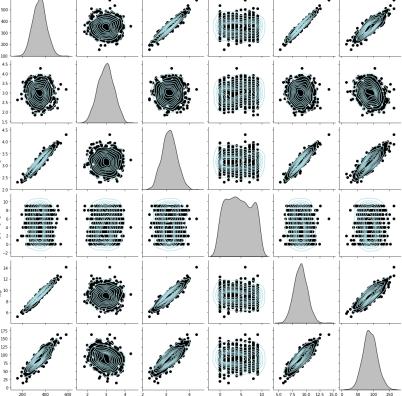
- 2) plotting data
- 3) scaling data
- 4) fitting model
- 5) evaluating model

```
Train = pd.read_csv("molecular_train_gbc.csv")
Test = pd.read_csv("molecular_test_gbc.csv")
```

```
1) loading data
```

- 2) plotting data
- 3) scaling data
- 4) fitting model
- 5) evaluating model





ev Data Fitting and Regression:

TrainS = scaler.fit_transform(Train)

TestS = scaler.transform(Test)

```
TrainS = pd.DataFrame(TrainS, columns = Train.columns)
TestS = pd.DataFrame(TestS, columns = Train.columns)
```

```
1) loading data
```

- 2) plotting data
- 3) scaling data
- 4) fitting model

the scaler returns an np.array

→ convert back to data frame

5) evaluating model



```
TrainS = pd.DataFrame(TrainS, columns = Train.columns)
TestS = pd.DataFrame(TestS, columns = Train.columns)
```

$$y_k = \beta_0 + \sum_{n=1}^N \beta_n x_n + \epsilon$$
 toxicity_score ~ molecular_weight + electronegativity + bond_lengths + num_hydrogen_bonds + logP

OLS (ordinary least squares)



my_model.summary()

number of data points is much larger than the number of regressors D

→ degree of freedom approx. no of obs

	OLS Regress	sion Results n	ot the fit quality!	
Dep. Variable:	toxicity_score	R-squared:	0.790	
Model:	OLS	Adj. R-squared:	0.789	
Method:	Least Squares	F-statistic:	597.5	
Date:	Fri, 13 Sep 2024	Prob (F-statistic):	3.34e-266	p-value for constant model
Time:	20:57:10	Log-Likelihood:	1013.0	
No. Observations:	800	AIC:	-2014.	
Df Residuals:	794	BIC:	-1986.	
Df Model:	5	p-valu	ues for	
Covariance Type:	nonrobust	factor	° S	

12.533

8.982

P>|t|

0.000

0.000

[0.025

0.126

0.622

0.975]

0.173

0.970

1)	loading	data
----	---------	------

- plotting data
- scaling data
- fitting model
- evaluating model

$$y_k = \beta_0 + \sum_{n=1}^{N} \beta_n x_n + \epsilon$$

 2σ conf range of factors

electronegativity	-0.1682	0.01	L5	-11.591	0.000	-0.197	-0.140
bond_lengths	0.0204	0.04	19	0.417	0.677	-0.076	0.116
num_hydrogen_bonds	0.0035	0.00	8	0.458	0.647	-0.011	0.018
logP	0.1246	0.07	72	1.723	0.085	-0.017	0.267
						 	-
Omnibus:		2.249	Dur	bin-Watson:		1.984	1
Prob(Omnibus):		0.325	Jar	que-Bera (JB):	2.246	9
Skew:		-0.129	Pro	b(JB):		0.32	5

std err

0.012

0.089

coef

0.1494

0.7961

Kurtosis: 65.6

Notes:

Intercept

molecular weight

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

more accurate: determining the p-values for the factors using ANOVA for the corresponding residuals

```
= sm.stats.anova_lm(my_model, typ = 1)
table
print(table)
```

- loading data
- plotting data
- scaling data
- fitting model
- evaluating model

$$y_k = \beta_0 + \sum_{n=1}^{N} \beta_n x_n + \epsilon$$

	df	sum_sq	mean_sq	F	PR(>F)
molecular_weight	1.0	13.346285	13.346285	2847.525516	8.024085e-265
electronegativity	1.0	0.640388	0.640388	136.631363	3.085962e-29
bond_lengths	1.0	0.000684	0.000684	0.145954	7.025342e-01
num_hydrogen_bonds	1.0	0.000703	0.000703	0.150055	6.985866e-01
logP	1.0	0.013917	0.013917	2.969353	8.524510e-02
Residual	794.0	3.721459	0.004687	NaN	NaN

vs from t-test

0.0000

0.0000

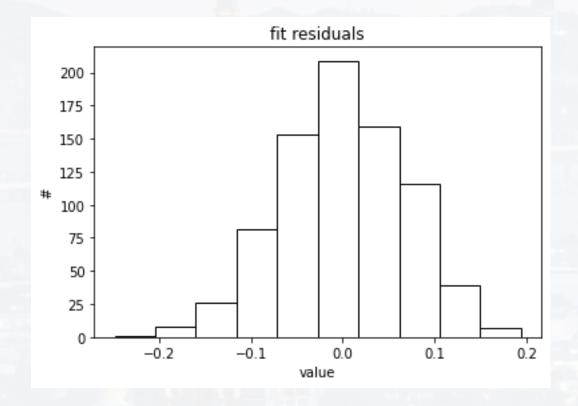
0.6766

0.6473

0.0852

```
residuals = my_model.resid

plt.hist(residuals, color = 'w', edgecolor = 'black')
plt.title('fit residuals')
plt.ylabel('#')
plt.xlabel('value')
plt.show()
```



1) loading data

2) plotting data

3) scaling data

4) fitting model

5) evaluating model

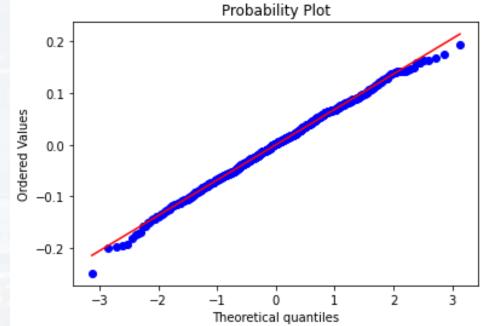
$$y_k = \beta_0 + \sum_{n=1}^{N} \beta_n x_n + \epsilon$$

residuals approx. normally distributed around $\mu = 0$

- 1) loading data
- 2) plotting data
- 3) scaling data
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- 5) evaluating model

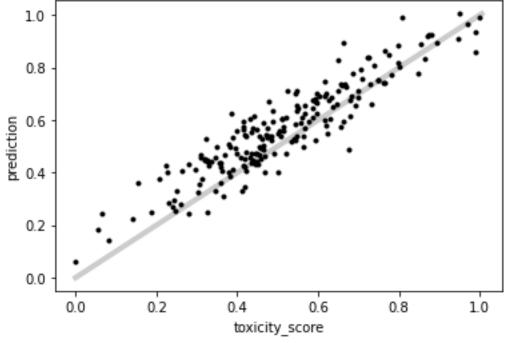
$$y_k = \beta_0 + \sum_{n=1}^{N} \beta_n x_n + \epsilon$$

stats.probplot(residuals, dist = "norm", plot = pylab)
pylab.show()



- 1) loading data
- 2) plotting data
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- 5) evaluating model

$$y_k = \beta_0 + \sum_{n=1}^{N} \beta_n x_n + \epsilon$$





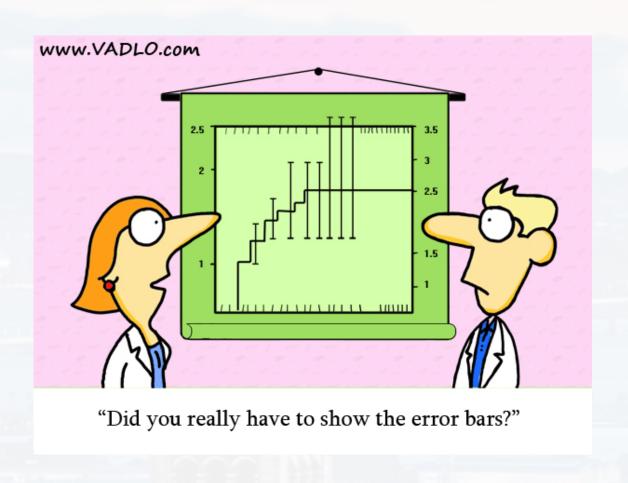
```
Ypred = my model.predict(TestS)
higher = np.max([Ypred, TestS.toxicity score])
lower = np.min([Ypred, TestS.toxicity score])
plt.plot([lower, higher], [lower, higher], c = [0, 0, 0, 0.2], linewidth = 4)
plt.scatter(TestS.toxicity_score, Ypred, marker = '.', c = 'k')
plt.ylabel('prediction')
plt.xlabel('toxicity score')
                                     1.0
plt.show()
                                     0.8
                                   0.6
0.4
                                     0.2
                                     0.0
                                                0.2
                                                       0.4
                                                              0.6
                                                                     0.8
                                                                           1.0
                                         0.0
                                                       toxicity_score
```

5) evaluating model

$$y_k = \beta_0 + \sum_{n=1}^N \beta_n x_n + \epsilon$$

```
mean_dev = np.sum( abs(TestS.toxicity_score - Ypred) )/len(Ypred)
print(mean_dev)
```

Berkeley Numerical Methods for Computational Science:



<u>Outline</u>

Linear Regression

- The Math
- What is Linear?
- Stats
- a Python Example

Logistic Regression

Curve Fitting

<u>linear model:</u> regressors are continuous or categorical,

response is continuous

<u>logistic model:</u> response is **categorical**

Index	molecular_weight	electronegativity	bond_lengths	num_hydrogen_bonds	logP	label
0	341.704	2.65585	3.09407	2	9.11147	Toxic
1	335.951	3.22262	2.89039	7	8.92848	Toxic
2	235.203	2.44115	2.48203	1	6.49731	Non-Toxic
3	246.505	2.76656	2.71547	7	7.45089	Non-Toxic
4	437.939	3.4801	3.59569	3	10.9156	Non-Toxic

<u>linear model:</u> regressors are continuous or categorical,

response is continuous

<u>logistic model:</u> response is **categorical**

y: response

x: regressors (assumed to be independent)

 β : factors β_0 : intercept

ε: error (stochasticity of the data, assumed to be

normally dist.)

dichotomic model:

probability to be in state A) $\rightarrow p$

probability to be in state B) $\rightarrow 1 - p$

Toxic
Toxic
Non-Toxic
Non-Toxic
Non-Toxic

ansatz:

$$\log\left(\frac{p}{1-p}\right) = \beta_0 + \sum_{n=1}^{N} \beta_n x_n + \epsilon$$

log odds ratio: linear model



dichotomic model:

probability to be in state A) $\rightarrow p$

response

regressors (assumed to be independent)

β: factors

intercept error (stochasticity of the data, assumed to be normally dist.)

probability to be in state B) $\rightarrow 1 - p$

Toxic Toxic Non-Toxic Non-Toxic Non-Toxic

ansatz:

$$\log\left(\frac{p}{1-p}\right) = \beta_0 + \sum_{n=1}^{N} \beta_n x_n + \epsilon$$

log odds ratio: linear model

→ probability for being in a certain state

$$p = \frac{e^{\beta_0 + \beta_1 x_1 + \dots}}{1 + e^{\beta_0 + \beta_1 x_1 + \dots}}$$

often:

$$logit(p) = log\left(\frac{p}{1-p}\right)$$

examples:

- probability that a gene has been mutated
- probability of being diseased (cancer, alzheimer etc) as function of age, environmental influence etc ...
- Verhulst equation: $N(t) = N_0 \frac{e^{rt}}{C + e^{rt}}$
- activation functions in ANNs

$$\log\left(\frac{p}{1-p}\right) = \beta_0 + \sum_{n=1}^{N} \beta_n x_n + \epsilon$$

y: response
x: regressors (assumed to be independent) β : factors β_0 : intercept ϵ : error (stochasticity of the data, assumed to be normally dist.)

Note: one can derive the logit function from max. entropy too!

$$p = \frac{e^{\beta_0 + \beta_1 x_1 + \dots}}{1 + e^{\beta_0 + \beta_1 x_1 + \dots}} = \frac{1}{1 + e^{-\beta_0 - \beta_1 x_1 - \dots}}$$

onset of Alzheimer's disease (AD) is a function of *age* and years spent in *education* (and other risk factors we ignore here for the sake of simplicity)

education: $d = x_1$ [yrs] age: $a = x_2$ [yrs]

model: $p_{AD} = \frac{1}{1 + e^{-\beta_0 - \beta_1 d - \beta_2 a}}$

+ data set + fit \Rightarrow $\beta_0 = +0.1$ $\beta_1 = -1.5$ $\beta_2 = +0.12$

- positive value → increasing p

- negative value → decreasing p

intercept: "background" prevalence, not
 related to environmental/internal conditions

model:
$$p_{AD} = \frac{1}{1 + e^{-\beta_0 - \beta_1 d - \beta_2 a}}$$

education:
$$d = x_1$$
 [yrs] $\beta_0 = +0.1$ age: $\beta_1 = -1.5$

$$\beta_2 = +0.12$$

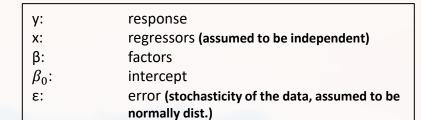
example: **65yrs** old person, **8yrs** spent in education $\Rightarrow p_{AD} = 1.6\%$

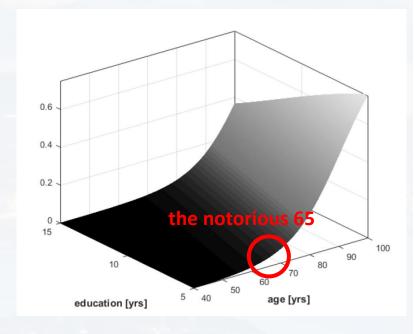
65yrs old person, **13yrs** spent in education $\rightarrow p_{AD} = 0.001\%$

How does education compensate aging?

$$p_{AD}(d+\bar{d},a+\bar{a})=p_{AD}(d,a)$$

$$\rightarrow \bar{a} = 12.5 \ \bar{d}$$





hence, one more year prolonged education compensates 12.5 years of aging

(warning: don't confuse correlation with causation here!)

model:
$$p_{AD} = \frac{1}{1 + e^{-\beta_0 - \beta_1 d - \beta_2 a}}$$

education:
$$d = x_1$$
 [yrs]

age:
$$a = x_2$$
 [yrs]

$$\beta_0 = +0.1$$

$$\beta_1 = -1.5$$

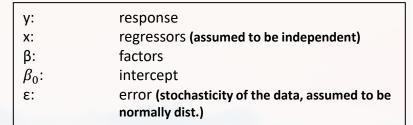
$$3_1 = -1.5$$

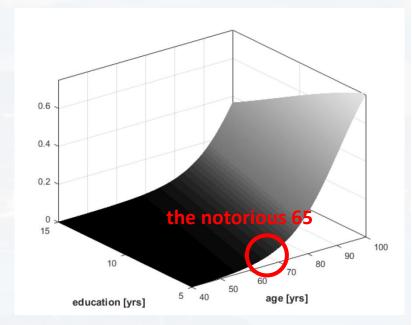
$$\beta_2 = +0.12$$

How does the risk of onset changes per year?

relative change:

$$\frac{p_{AD}(a+1) - p_{AD}(a)}{p_{AD}(a)} \approx e^{\beta_2} - 1 \approx 12.7\%$$





 $p_{AD}\ll 1$ (hence, for small Δa and "young" ages, i. e. below pprox 80yrs)

the risk of getting AD increases by 12.7% every year

(warning: does not mean that it increases by 127% in ten yrs - we made an approximation!)

model:
$$p_{AD} = \frac{1}{1 + e^{-\beta_0 - \beta_1 d - \beta_2 a}}$$

education:
$$d=x_1$$
 [yrs] $\beta_0=+0.1$ age: $\alpha=x_2$ [yrs] $\beta_1=-1.5$

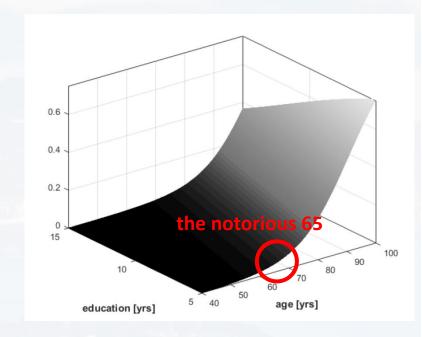
$$\beta_2 = +0.12$$

How does the risk of onset changes per year?

more precise: relative change of the odds ratio

$$\frac{\frac{\partial}{\partial x_i} \left(\frac{p_{AD}}{1 - p_{AD}} \right)}{\frac{p_{AD}}{1 - p_{AD}}} = \beta_i$$
 x_i is the desired regressor, for example, age again (x_2)

y:	response
x:	regressors (assumed to be independent)
β:	factors
β_0 :	intercept
β: β ₀ : ε:	error (stochasticity of the data, assumed to be normally dist.)



the factors β_i indicate how strong (and in which direction) p changes wrt a regressor x_i

let us return to the molecule data set:

Index	molecular_weight	electronegativity	bond_lengths	num_hydrogen_bonds	logP	label
0	341.704	2.65585	3.09407	2	9.11147	Toxic
1	335.951	3.22262	2.89039	7	8.92848	Toxic
2	235.203	2.44115	2.48203	1	6.49731	Non-Toxic
3	246.505	2.76656	2.71547	7	7.45089	Non-Toxic
4	437.939	3.4801	3.59569	3	10.9156	Non-Toxic

import statsmodels.api as sm



```
it is the same data set \rightarrow plotting and scaling is as before
                                                             adding the
                                                              intercept
X = sm.add_constant(TrainS)
                                                           Python needs
                                                            True/False
Y = pd get_dummies(Train['label'])
                                                            as categorical
In [48]: print(Y)
    Non-Toxic Toxic
         False
                 True
                                                            we have two
         False True
                                                            states: toxic /
         True False
         True False
                                                             non-toxic
         True False
my_model = sm.GLM(Y, X, family = sm.families.Binomial()).fit()
                                                           GLM: general
my_model.summary()
                                                            linear model
```

const

logP

molecular weight

electronegativity

num hydrogen bonds

bond lengths

it is the same data set \rightarrow plotting and scaling is as before

6.1641

3.2874

0.6736

-0.3082

-7.6090

-10.4920

```
X = sm.add_constant(TrainS)
Y = pd.get_dummies(Train['label'])

my_model = sm.GLM(Y, X, family = sm.families.Binomial()).fit()
my_model.summary()
```

Generalized Linear Model Regression Results ['Non-Toxic', 'Toxic'] Dep. Variable: No. Observations: 800 Df Residuals: Model: GLM 794 Model Family: Binomial Df Model: Link Function: Scale: Logit 1.0000 Log-Likelihood: Method: IRLS -332.82 Date: Sat, 14 Sep 2024 Deviance: 665.64 Time: 20:59:18 Pearson chi2: 1.14e+03 No. Iterations: Pseudo R-squ. (CS): p-values for 0.4243 Covariance Type: nonrobust factors [0.025 0.975] coef std err P> | z | Z

10.536

-2.893

0.352

-1.018

-2.555

5.492

0.000

0.004

0.000

0.725

0.309

0.011

5.017

2.114

-3.075

-0.902

-13.447

-17.599

7.311

-3.385

4.461

4.422

0.285

-1.771

0.585

3.626

0.599

1.913

0.303

2.978

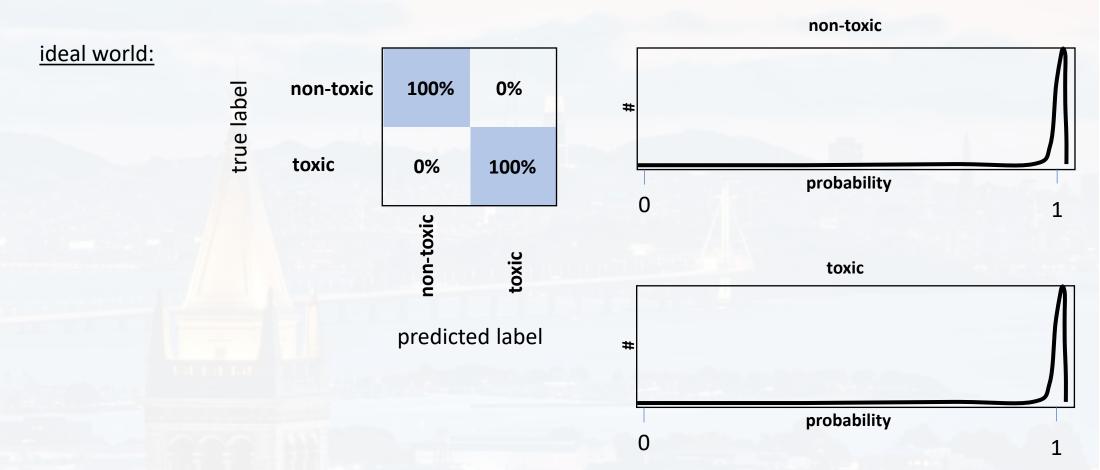
n –	$e^{\beta_0+\beta_1x_1+\cdots}$
p =	$1 + e^{\beta_0 + \beta_1 x_1 + \cdots}$

p-value for constant model

 2σ conf range of factors

accuracy: How *often* did the model make the correct prediction.

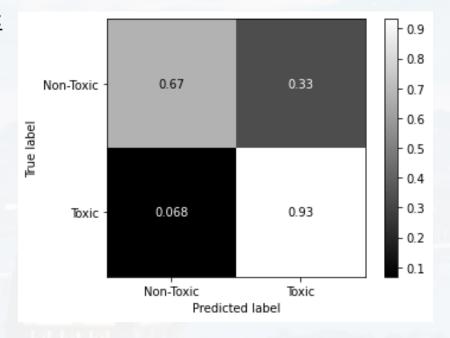
cross-entropy: How *certain* was the model when making the prediction.

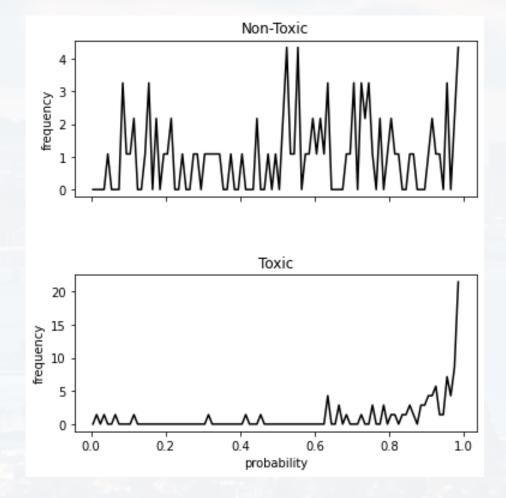


accuracy: How *often* did the model make the correct prediction.

cross-entropy: How *certain* was the model when making the prediction.

real world:





Berkeley Numerical Methods for Computational Science:



<u>Outline</u>

Linear Regression

- The Math
- What is Linear?
- Stats
- a Python Example

Logistic Regression

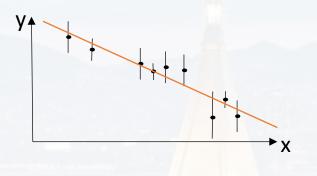
Curve Fitting

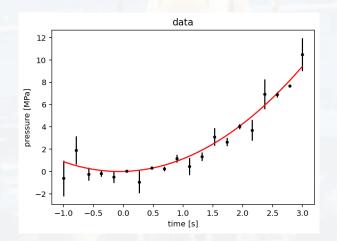
problem: set of data points $y_i = f(x_i)$, ideally each data point y_i has an error σ_i

we know $f(x_i)$

goal: find parameters of $f(x_i)$

once we know the parameters → prediction





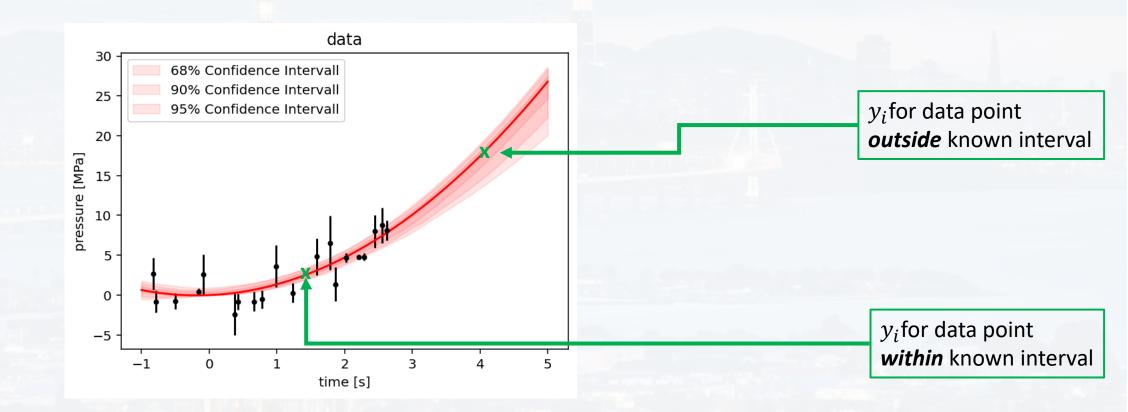
model: $f(x_i) = m x_i + n$ parameter: m, n

model: $f(x_i) = a x_i^2 + b x_i + c$ parameter: a, b, c **problem:** set of data points $y_i = f(x_i)$, ideally each data point y_i has an error σ_i

we know $f(x_i)$

goal: find parameters of $f(x_i)$

once we know the parameters → prediction

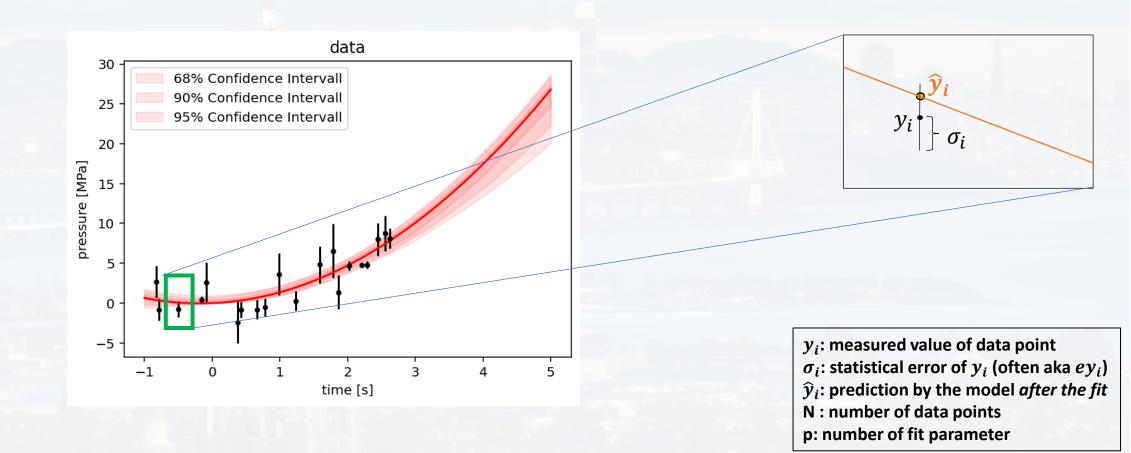


problem: set of data points $y_i = f(x_i)$, ideally each data point y_i has an error σ_i

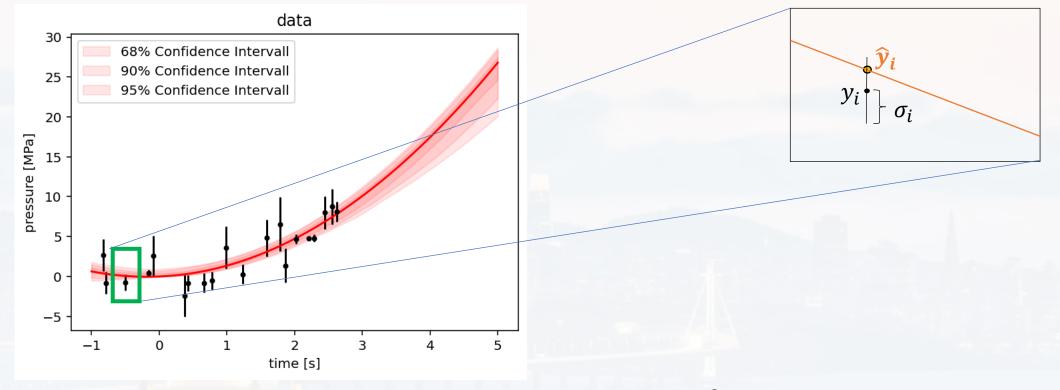
we know $f(x_i)$

goal: find parameters of $f(x_i)$

once we know the parameters → prediction







$$\chi_{red}^2 = \frac{1}{df} \sum_{i=1}^{N} \left(\frac{y_i - \hat{y}_i}{\sigma_i} \right)^2 \qquad df = N - p - 1$$
 see module 8

 y_i : measured value of data point σ_i : statistical error of y_i (often aka ey_i)

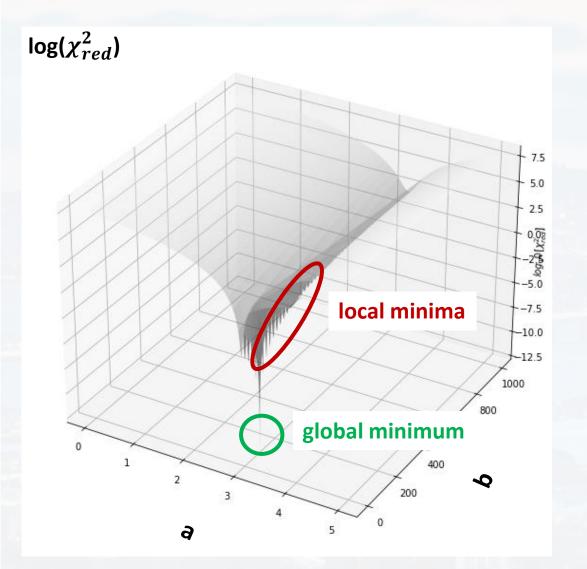
 \hat{y}_i : prediction by the model after the fit

N: number of data points p: number of fit parameter

or
$$MSE = \frac{1}{df} \sum_{i=1}^{N} (y_i - \hat{y}_i)^2$$

if no errors given

finding best parameters by minimizing



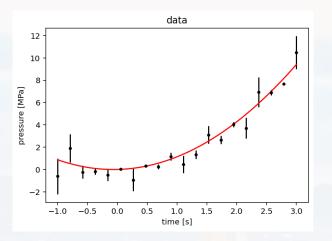
$$\chi_{red}^2 = \frac{1}{df} \sum_{i=1}^{N} \left(\frac{y_i - \hat{y}_i}{\sigma_i} \right)^2 \qquad df = N - p - 1$$
 see module 8

or
$$MSE = \frac{1}{df} \sum_{i=1}^{N} (y_i - \hat{y}_i)^2$$
 if no errors given

finding the global minimum of a higher dimensional non-analytical function

→ see also module 3

from module 8



$$\chi_{red}^2 = \frac{1}{df} \sum_{i=1}^{N} \left(\frac{y_i - \hat{y}_i}{\sigma_i} \right)^2$$

$$df = N - p - 1$$

N : number of data points

p: number of fit parameter (model)

given a fitted model: χ^2_{red} is a measure of the fit quality

for large (> 50...100) N (number of data points):

≈ 2/3 of the data points should be consistent with the model within their 1σ error bars

 \approx 95% of the data points should be consistent with the model within their 2 σ error bars

 \approx 99.7% of the data points should be consistent with the model within their 3σ error bars

$$\chi^2_{red} \approx$$

1.0 excellent fit

1.0...1.5 acceptable fit

1.5...1.7 bad fit

>2.0 not acceptable

<<1.0 suspicious, errors are overestimated!

```
from scipy.optimize import curve_fit
def fun_to_fit(x, a, b, c):
           return a*x**2 + b*x + c
ValsBest, Cov = curve_fit(fun_to_fit, x, y)
ValsBest
array([ 1.37347967, -0.31706189, -0.20746552])
                                b
                                                       C
In [16]: Cov
                                                                   \begin{pmatrix} \sigma_a^2 & cov(a,b) & cov(a,c) \\ cov(b,a) & \sigma_b^2 & cov(b,c) \\ cov(c,a) & cov(c,b) & \sigma_c^2 \end{pmatrix}
Out[16]:
array([[ 0.15053344, -0.28309719, -0.05890134],
       [-0.28309719, 0.6630454, -0.02965514],
       [-0.05890134, -0.02965514, 0.33825647]])
```

```
from scipy.optimize import curve_fit
                                                                       scipy has the optimization
                                                                            tool cuve fit
def fun_to_fit(x, a, b, c):
                                                                         defining the model as a
         return a*x**2 + b*x + c
                                                                            python function
ValsBest, Cov = curve_fit(fun_to_fit, x, y)
                                                                            run the curve fit
ValsBest
                                                                      best values by minimizing MSE
array([ 1.37347967, -0.31706189, -0.20746552])
                                                                               or related
                          b
                                            C
In [16]: Cov
                                                                  cov(a,b)
                                                                             cov(a,c)
Out[16]:
array([[ 0.15053344, -0.28309719, -0.05890134],
                                                                              cov(b,c)
      [-0.28309719, 0.6630454, -0.02965514],
      [-0.05890134, -0.02965514, 0.33825647]])
                               covariance matrix for the model parameters
                                  non diagonal values should be ≈ 0 (mutually independent)
                                  diagonal: squared errors (see module 8)
```

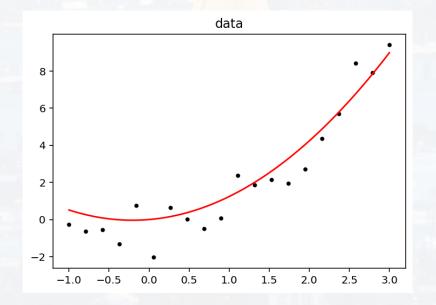
```
In [16]: Cov Out[16]: array([[ 0.15053344, -0.28309719, -0.05890134], [-0.28309719, 0.6630454], -0.02965514], [-0.05890134, -0.02965514, 0.33825647]])  \begin{pmatrix} \sigma_a^2 & cov(a,b) & cov(a,c) \\ cov(b,a) & \sigma_b^2 & cov(b,c) \\ cov(c,a) & cov(c,b) & \sigma_c^2 \end{pmatrix}
```

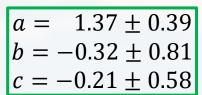
```
error_1sigma = np.sqrt(np.diagonal(Cov))
```

from scipy.optimize import curve_fit

ValsBest, Cov = curve_fit(fun_to_fit, x, y)

$$y = 1.37 x^2 - 0.32 x - 0.21$$



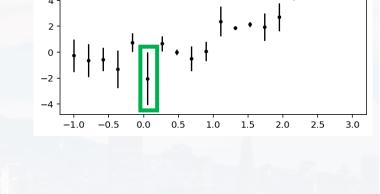


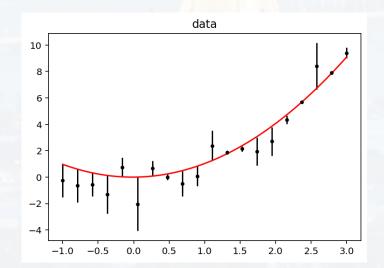
from scipy.optimize import curve_fit

ValsBest, Cov = curve_fit(fun_to_fit, x, y)

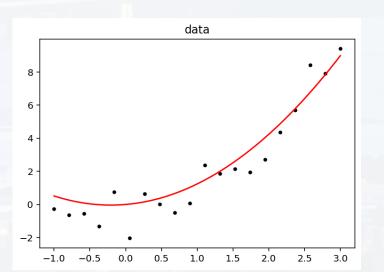
if error bars know → error weighted fit!

ValsBest, Cov = curve_fit(fun_to_fit, x, y, sigma = err, absolute_sigma = True)





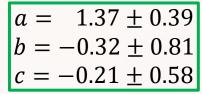


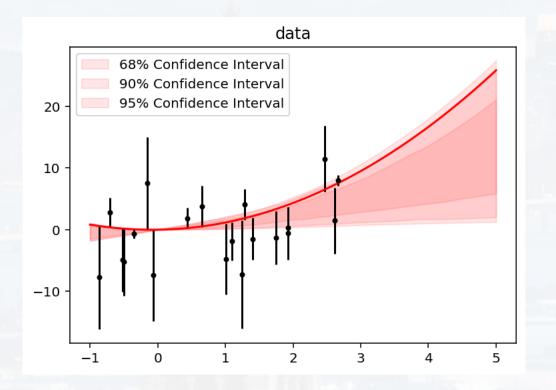


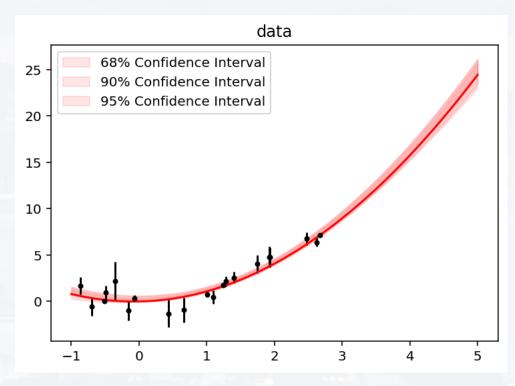


data points have errors and/or fit parameter have uncertainties too

- → implies uncertainties of the model
- → confidence band/interval









data points have errors and/or fit parameter have uncertainties too

- → implies uncertainties of the model
- → confidence band/interval
- → idea: **bootstrapping!**

$$a = 1.37 \pm 0.39$$

 $b = -0.32 \pm 0.81$
 $c = -0.21 \pm 0.58$

data points have errors and/or fit parameter have uncertainties too

 $a = 1.37 \pm 0.39$ $b = -0.32 \pm 0.81$ $c = -0.21 \pm 0.58$

- → implies uncertainties of the model
- → confidence band/interval
- → idea: **bootstrapping!**

if no errors of y_i known

- \rightarrow assuming that fitted parameters follow a **normal distribution**, i.e. $a=1.37\pm0.39$ where $\mu_a=1.37$ and $\sigma_a=0.39$ and so on...
- \rightarrow varying the parameters within their errors using np.random.normal(μ_a , σ_a , N) N times
- → for each N, generating a curve fit
- → from set of N curve fits → calculating percentiles for confidence band/ interval



data points have errors and/or fit parameter have uncertainties too

- → implies uncertainties of the model
- → confidence band/interval
- → idea: **bootstrapping!**

$$a = 1.37 \pm 0.39$$

 $b = -0.32 \pm 0.81$
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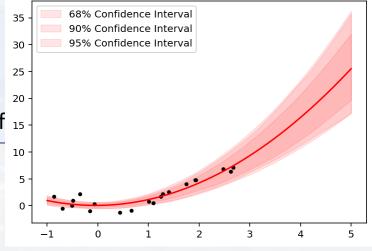
if no errors of y_i known

 \Rightarrow assuming that fitted parameters follow a **normal distribution**, i.e. $a=1.37\pm0.39$ where $\mu_a=1.37$ and $\sigma_a=0.39$ and so on...

→ varying the parameters within their errors using np.random normal(" a N)

N times

- → for each N, generating a curve fit
- → from set of N curve fits → calculating percentiles for conf 15



data points have errors and/or fit parameter have uncertainties too

 $a = 1.37 \pm 0.39$ $b = -0.32 \pm 0.81$ $c = -0.21 \pm 0.58$

- → implies uncertainties of the model
- → confidence band/interval
- → idea: **bootstrapping!**

if errors of y_i known

- ightarrow assuming that errors of y_i follow a normal distribution, i.e. $y_i(boot) = y_i \pm \sigma_i$
- \rightarrow varying all y_i within their errors using np.random.normal (y_i, σ_i, N) N times
- → for each N, generating a curve fit
- → from set of N curve fits → calculating percentiles for confidence band/ interval



data points have errors and/or fit parameter have uncertainties too

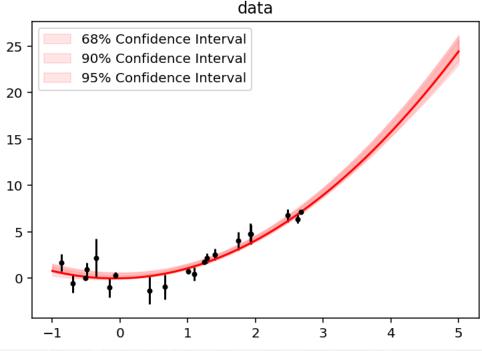
- → implies uncertainties of the model
- → confidence band/interval
- → idea: **bootstrapping!**

$$a = 1.37 \pm 0.39$$

 $b = -0.32 \pm 0.81$
 $c = -0.21 \pm 0.58$

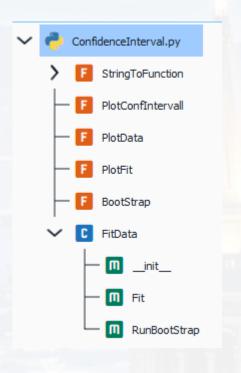
if errors of y_i known

- \rightarrow assuming that errors of y_i follow a normal distribution i.e. $v_i(hoot) = v_i + \sigma_i$
- \rightarrow varying all y_i within their errors using np. ra 25 N times
- → for each N, generating a curve fit
- \rightarrow from set of N curve fits \rightarrow calculating percen ¹⁰



bootstrapping

explore the .py file
ConfidenceInterval.py



```
USAGE:
····generating·a·test·sample:
\cdots \times \cdots = -np.linspace(-1,3,20)
\cdotserr\cdots=·np.random.normal(0,·1,·(len(x),))#1sigma·errorbars
····y····=·x**2·+·err
····errorbars·=·abs(err)
····1) · plotting · data
\cdots F1 = FitData(x, y)
\cdots F2 = FitData(x, y, errorbars)
····F3·=·FitData(x,·y,·errorbars,·time·=·'[s]',·pressure·=·'[MPa]')
\cdots 2) · fitting · data · (returns · best · values · of · fitted · params, · 1 sigma · confidence · and
    ·····reduced·chi2·if·errorbars·given,·MSE·else)
\cdotsres1 \cdots = F1.Fit()
····res2··=·F2.Fit()
····res2··=·F3.Fit()
\cdotsres12·=·F1.Fit("a*x**2",·[1],·(-0.5,·10))
····3)·Bootstrapping·(either·varying·within·errorbars·or·within·conf·of·fitted
        ····params)
····F1.RunBootStrap()
····F2.RunBootStrap()
····F3.RunBootStrap()
\cdotsF1.RunBootStrap(100, \cdot[90, \cdot95], \cdotnp.linspace(-1,5,200))
\cdotsF3. RunBootStrap(100, \cdot[90, \cdot95], \cdotnp. linspace(-1, 5, 200))
```

Berkeley Numerical Methods for Computational Science:

Thank you very much for your attention!

