

Lecture 6:

Variational Bayes, Expectation Maximization



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Bayesian Data Analysis and
Machine Learning for Physical
Sciences



Course Map

Module 1	Maximum Entropy and Information, Bayes Theorem
Module 2	Naive Bayes, Bayesian Parameter Estimation, MAP
Module 3	MLE, Lin Regression, Model selection: Comparing Distributions
Module 4	Model Selection: Bayesian Signal Detection
Module 5	Variational Bayes, Expectation Maximization
Module 6	Stochastic Processes
Module 7	Monte Carlo Methods
Module 8	Markov Models, Graphs
Module 9	Machine Learning Overview, Supervised Methods
Module 10	Unsupervised Methods
Module 11	ANN: Perceptron, Backpropagation
Module 12	ANN: Basic Architecture, Regression vs Classification, Backpropagation again
Module 13	Convolution and Image Classification and Segmentation
Module 14	TBD (GNNs)
Module 15	TBD (RNNs and LSTMs)
Module 16	TBD (Transformer and LLMs)



Outline

The Problem

K-means

Actual EM

Variational Bayes



Outline

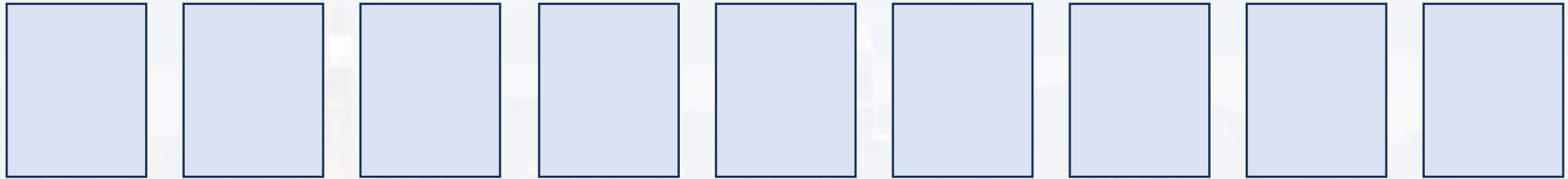
The Problem

K-means

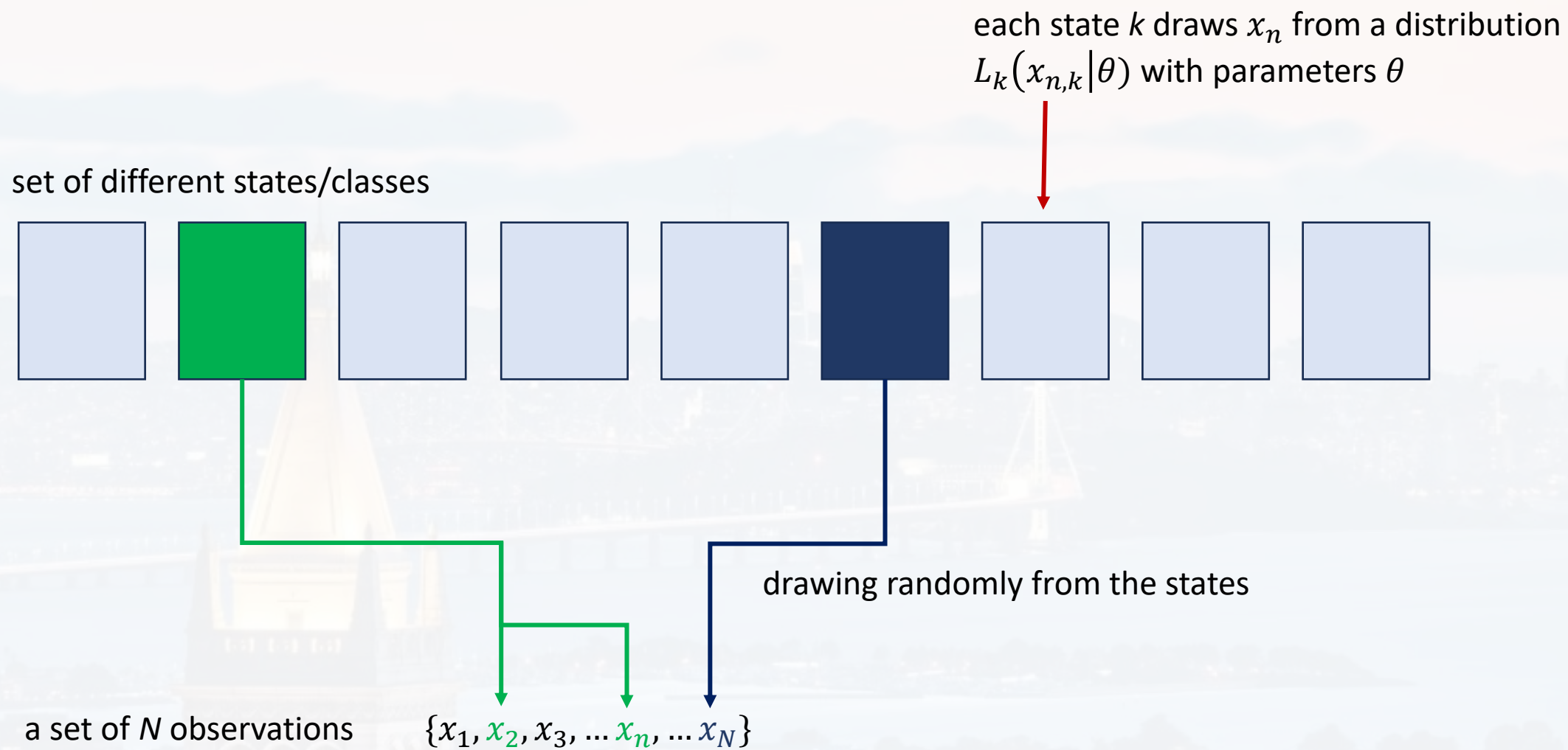
Actual EM

Variational Bayes

set of different states/classes

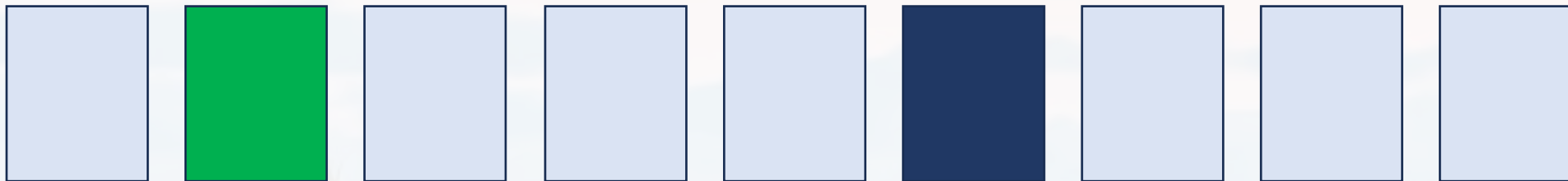


a set of N observations $\{x_1, x_2, x_3, \dots x_n, \dots x_N\}$





set of different states/classes

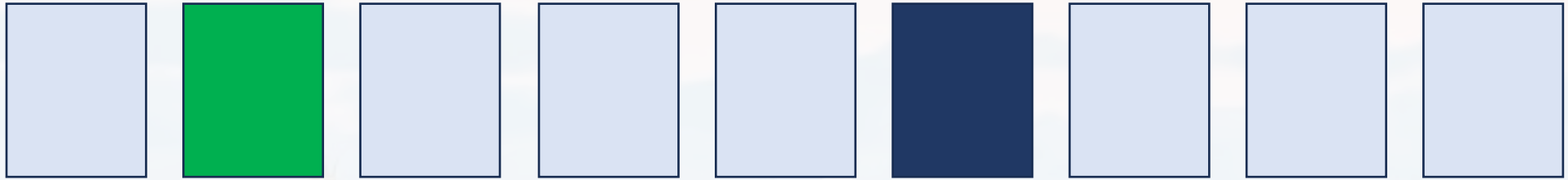


a set of N observations $\{x_1, x_2, x_3, \dots, x_n, \dots, x_N\}$

problem: - we have a model of $L_k(x_{n,k} | \theta)$, but
- we don't know θ and
- we don't know from which class/state k x_n has been generated

goal: - find an estimator for θ and find the class/state k of each x_n

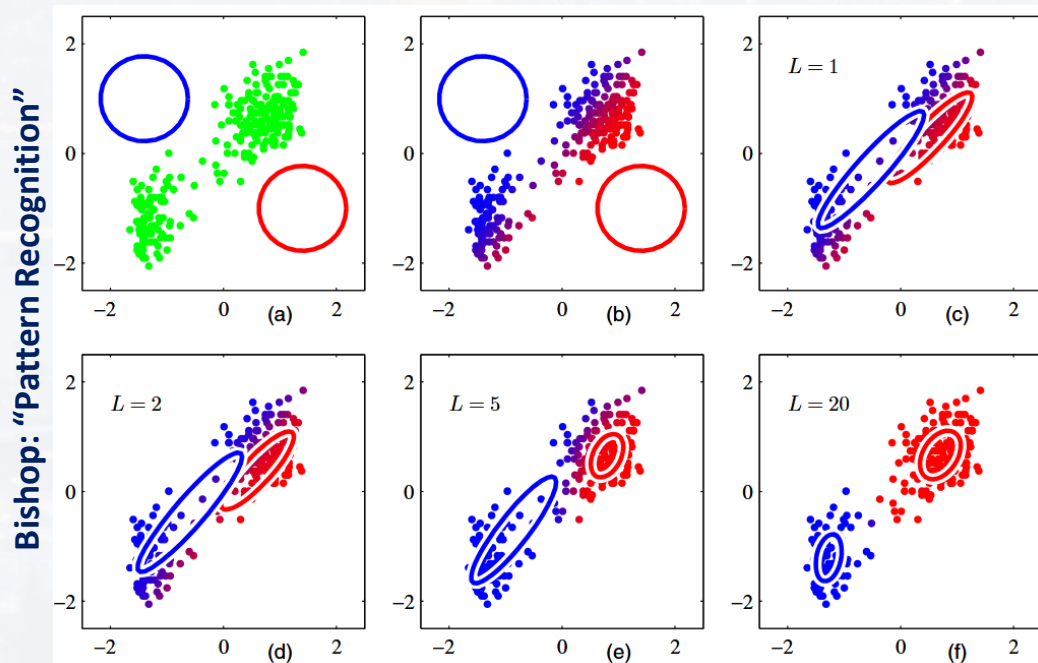
set of different states/classes



$$L_k(x_{n,k}|\theta)$$

a set of N observations

$$\{x_1, x_2, x_3, \dots, x_n, \dots, x_N\}$$



- Gaussian Mixture Models (GMM)
- K-means (clustering, image segmentation)
- HMM

→ unsupervised



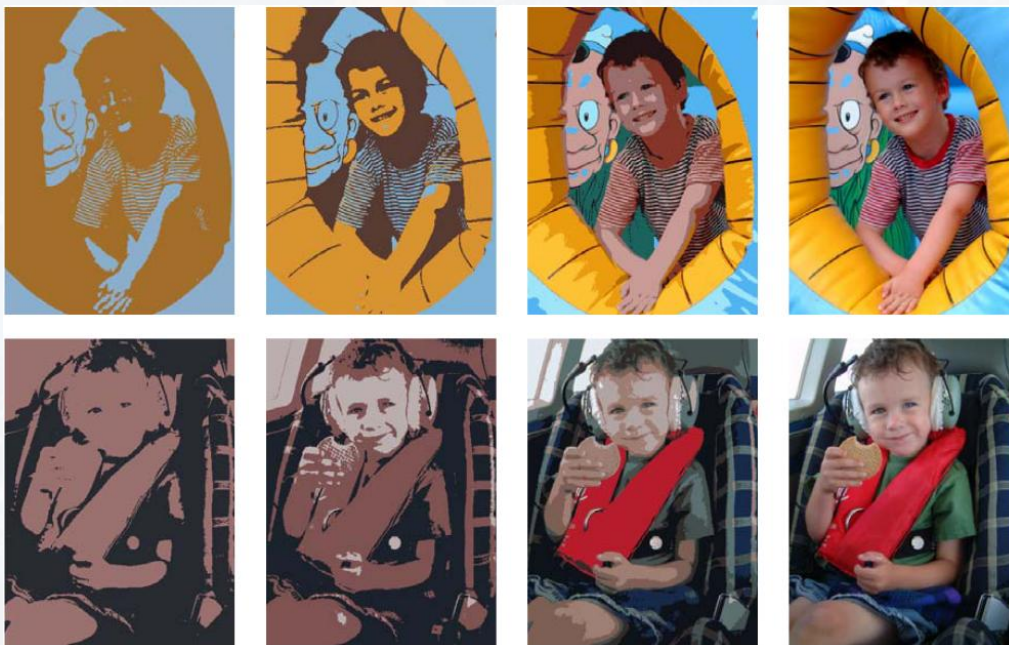
set of different states/classes



a set of N observations

$\{x_1, x_2, x_3, \dots, x_n, \dots, x_N\}$

Bishop: "Pattern Recognition"

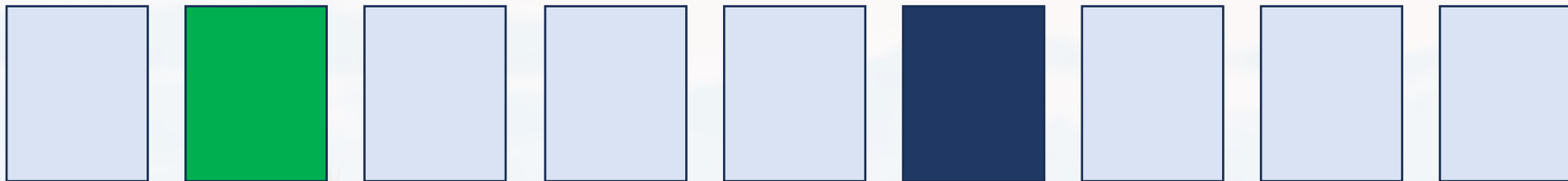


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- K-means (clustering, image segmentation)
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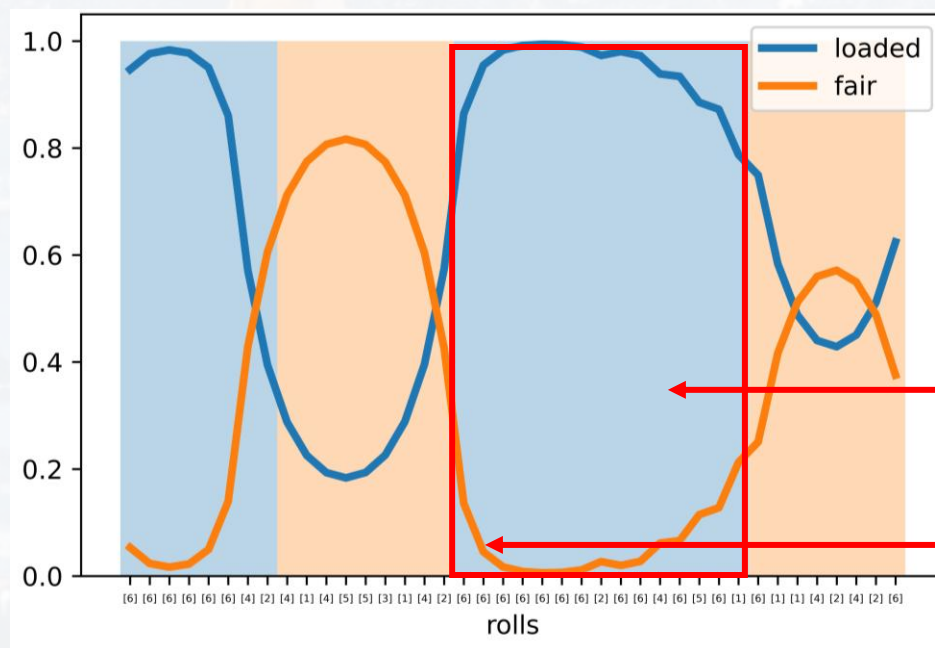


set of different states/classes



a set of N observations

$\{x_1, x_2, x_3, \dots, x_n, \dots, x_N\}$



- Gaussian Mixture Models (GMM)
- K-means (clustering, image segmentation)
- HMM

→ unsupervised

actual (unknown,
aka hidden state)

predicted state,
given observation



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indicator function $r_{n,k} \in \{0, 1\}$

goal: minimizing

$$J = \sum_{n=1}^N \sum_{k=1}^K r_{n,k} \|x_n - \mu_k\|^2$$

μ_k : barycenter of cluster k

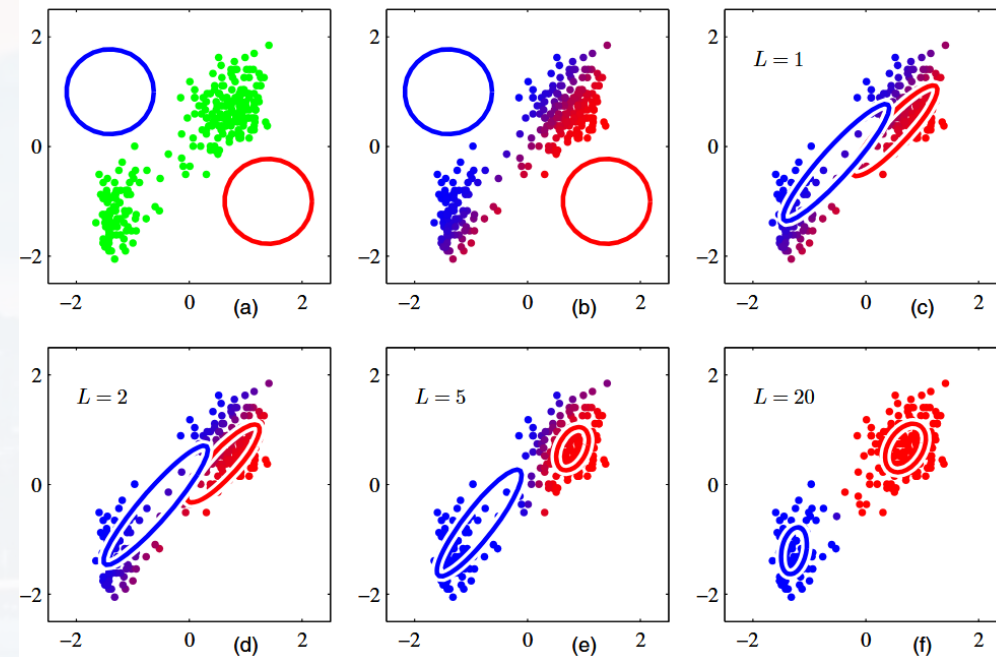
assigning x_n to its closed mean

$$r_{n,k} = \begin{cases} 1, & \text{if } k = \operatorname{argmin}_j \|x_n - \mu_j\|^2 \\ 0, & \text{else} \end{cases}$$

$$\frac{\partial J}{\partial \mu_k} = 0 \quad \text{MLE} \quad \longrightarrow \quad \mu_k = \frac{\sum_{n=1}^N r_{n,k} x_n}{\sum_{n=1}^N r_{n,k}}$$

K – means is an iterative process

K : number of cluster
 N : number of observations

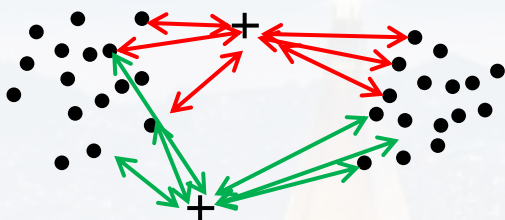




K – means is an iterative process



a) assign k means randomly



b) calculate *distance* from each point to each mean



c) assign each point to its closest mean



d) update the means accordingly

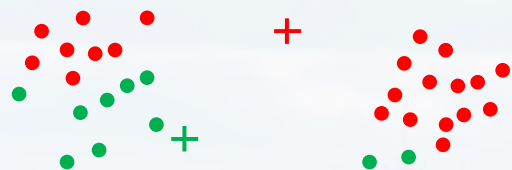
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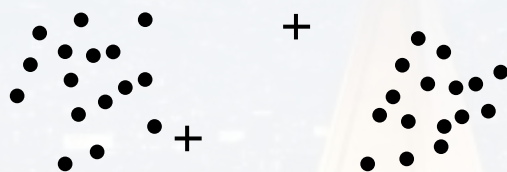
$$J = \sum_{n=1}^N \sum_{k=1}^K r_{n,k} \|x_n - \mu_k\|^2$$



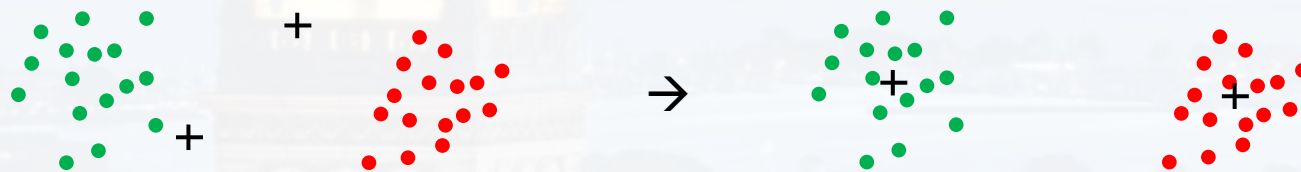
K – means is an iterative process



d) update the means accordingly



e) go back to b)



K : number of cluster
 N : number of observations
 μ_k : barycenter of the cluster

$$r_{n,k} = \begin{cases} 1, & \text{if } k = \operatorname{argmin}_j \|x_n - \mu_j\|^2 \\ 0, & \text{else} \end{cases}$$

$$J = \sum_{n=1}^N \sum_{k=1}^K r_{n,k} \|x_n - \mu_k\|^2$$



problem: $K = \text{number of cluster}$, is a hyperparameter. How do I know the correct value for K ?

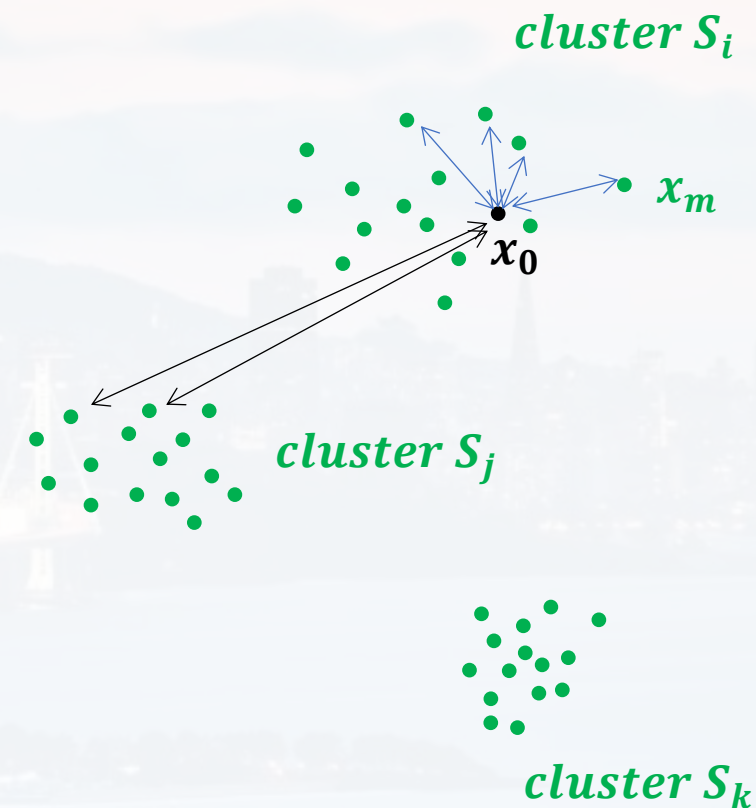
→ silhouette Ψ

- distance d_1 of a data point x_0 to *its assigned cluster* S_i
vs distance d_2 to *closest cluster (here S_j)*

$$\Psi(x_0) = \begin{cases} 0 & \text{if } d_1 = 0 \\ \frac{d_2 - d_1}{\max[d_1; d_2]} \end{cases}$$

- average over all points → ψ_{tot}

if	$\psi_{tot} = 0.75 \dots 1.00$	→ well clustered
	$\psi_{tot} = 0.50 \dots 0.75$	→ medium clustered
	$\psi_{tot} = 0.25 \dots 0.50$	→ poorly clustered
	$\psi_{tot} < 0.25$	→ data has no structure





problem: $K = \text{number of cluster}$, is a hyperparameter. How do I know the correct value for K ?

→ silhouette Ψ

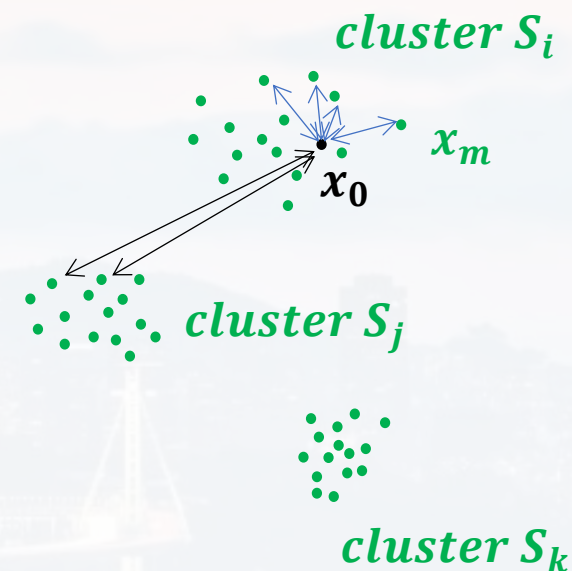
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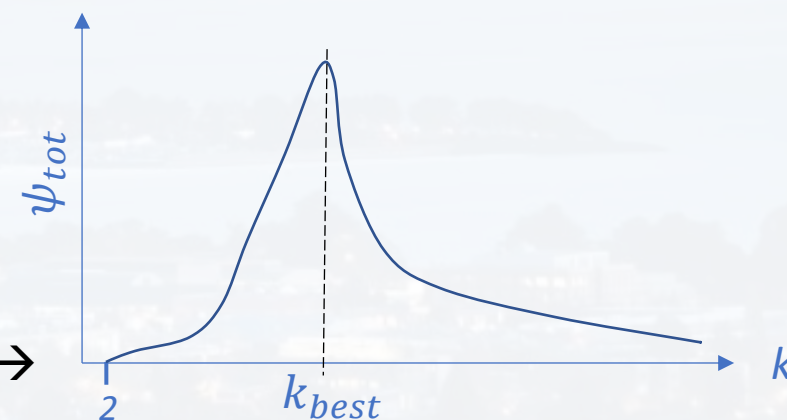
- average over all points → ψ_{tot}

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	$\psi_{tot} = 0.25 \dots 0.50$	→ poorly clustered
	$\psi_{tot} < 0.25$	→ data has no structure

see `Walk_Through_Kmeans.ipynb`



ideal world →





the actual problem:

observation x_n has been drawn from any of the cluster C_k

$$P(x_n) = \sum_{k=1}^K P(x_n|C_k) P(C_k) \qquad \sum_{k=1}^K P(C_k) = 1$$

$P(x_n|C_k)$ likelihood function

$P(C_k)$ mixing coefficient

general:

$$P(x_n) = \sum_z P(x_n|z) P(z)$$

z : latent variable (i. e. not observable, but can be inferred from $\{x_n\}$)

K : number of cluster
 N : number of observations
 μ_k : barycenter of the cluster
 $r_{n,k} = \begin{cases} 1, & \text{if } k = \operatorname{argmin}_j \|x_n - \mu_j\|^2 \\ 0, & \text{else} \end{cases}$

$$J = \sum_{n=1}^N \sum_{k=1}^K r_{n,k} \|x_n - \mu_k\|^2$$



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Actual EM

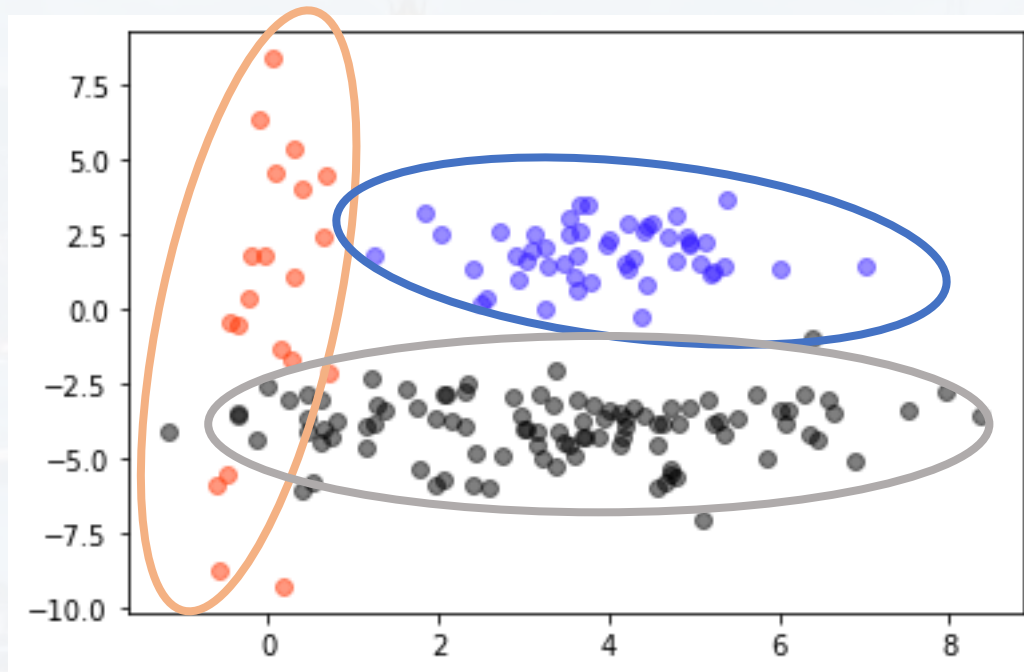
Variational Bayes

example: **G**aussian **M**ixture **M**odels

K : number of cluster
 π_k : mixing coefficient

f features

$$\mathcal{N}_k(x|\mu_k, \Sigma_k) = \frac{1}{(2\pi)^{f/2} \det(\Sigma_k)^{1/2}} \exp \left[-\frac{1}{2} (x - \mu_k)^T \Sigma_k^{-1} (x - \mu_k) \right]$$



two features, $K = 3$ components

$$\begin{aligned}
 P(x) &= \sum_z P(x|z) P(z) \\
 &= \sum_{k=1}^K \mathcal{N}_k(x|\mu_k, \Sigma_k) \pi_k
 \end{aligned}$$

example: **G**aussian **M**ixture **M**odels

$$P(x) = \sum_z P(x|z) P(z) = \sum_{k=1}^K \mathcal{N}_k(x|\mu_k, \Sigma_k) \pi_k$$

N	: number of observations
K	: number of cluster
π_k	: mixing coefficient

indicator variable $z_k \in \{0, 1\}$

goal:
$$P(z_k = 1|x) = \frac{P(z_k = 1) P(x|z_k = 1)}{P(x)} = \frac{\pi_k \mathcal{N}_k(x|\mu_k, \Sigma_k)}{\sum_{j=1}^K \mathcal{N}_j(x|\mu_j, \Sigma_j) \pi_j}$$

via maximizing likelihood by finding best θ
$$L = \ln[P(x|\pi, \mu, \Sigma)] = \sum_{n=1}^N \ln \left\{ \sum_{k=1}^K \pi_k \mathcal{N}_k(x_n|\mu_k, \Sigma_k) \right\}$$

model parameter $\theta = \{\pi, \mu, \Sigma\}$

example: **G**aussian **M**ixture **M**odels

N	: number of observations
K	: number of cluster
π_k	: mixing coefficient

1) initialize $\theta = \{\pi, \mu, \Sigma\}$ and $P(z_k = 1|x_n)$

2) **E**xpectation step t :

$$P(z_k = 1|x_n) = \frac{\pi_k \mathcal{N}_k(x_n|\mu_k, \Sigma_k)}{\sum_{j=1}^K \mathcal{N}_j(x_n|\mu_j, \Sigma_j) \pi_j} \quad \text{i.e. evaluate } P(z|x, \theta)$$

3) **M**aximization step t (for example MLE):

$$\mu_k^t = \frac{1}{N_k} \sum_{n=1}^N P(z_k = 1|x_n) x_n \quad \Sigma_k^t = \frac{1}{N_k} \sum_{n=1}^N P(z_k = 1|x_n) (x_n - \mu_k^t) (x_n - \mu_k^t)^T$$

$$N_k^t = \sum_{n=1}^N P(z_k = 1|x_n) \quad \pi_k^t = \frac{N_k}{N} \quad \theta^t = \underset{\theta}{\operatorname{argmax}} \{L(\theta^{t-1})\}$$

example: **G**aussian **M**ixture **M**odels

N	: number of observations
K	: number of cluster
π_k	: mixing coefficient

1) initialize $\theta = \{\pi, \mu, \Sigma\}$ and $P(z_k = 1|x_n)$

2) **E**xpectation step t :

$$P(z_k = 1|x_n) = \frac{\pi_k \mathcal{N}_k(x_n|\mu_k, \Sigma_k)}{\sum_{j=1}^K \mathcal{N}_j(x_n|\mu_j, \Sigma_j) \pi_j} \quad \text{i.e. evaluate } P(z|x, \theta)$$

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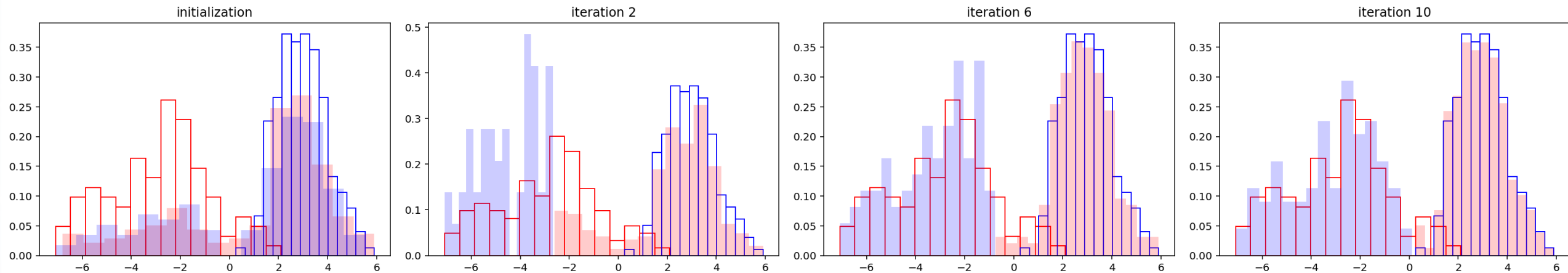
$$\mu_k^t = \frac{1}{N_k} \sum_{n=1}^N P(z_k = 1|x_n) x_n \quad \Sigma_k^t = \frac{1}{N_k} \sum_{n=1}^N P(z_k = 1|x_n) (x_n - \mu_k^t) (x_n - \mu_k^t)^T \quad N_k = \sum_{n=1}^N P(z_k = 1|x_n) \quad \pi_k^t = \frac{N_k}{N}$$

$$\theta^t = \underset{\theta}{\operatorname{argmax}} \left\{ \sum_z P(z|x, \theta^{t-1}) \ln[P(x, z|\theta^{t-1})] \right\}$$

4) evaluate log likelihood

$$\ln[P(x|\pi, \mu, \Sigma)] = \sum_{n=1}^N \ln \left\{ \sum_{k=1}^K \pi_k \mathcal{N}_k(x_n|\mu_k, \Sigma_k) \right\}$$

- note:**
- one can show that EM indeed converges and maximizes the likelihood function (Bishop, Sec 9.4)
 - no guarantee to find a **global** minimum
 - applications: HMM, unsupervised clustering, image segmentation (before CNNs)
 - iterative process: $\theta_{t+1} = \theta_t + \Delta\theta \rightarrow$ connection to gradient descent (see later)
 - see also EM__Example.py



blue and red swapped
 \rightarrow unsupervised learning



Outline

The Problem

K-means

Actual EM

Variational Bayes



- | | |
|------------------------------|--|
| EM: | <ul style="list-style-type: none">- likelihood function given- parameter via MLE (point estimate) |
| problem: | <ul style="list-style-type: none">- integrals can be complicated, calculation takes too long etc (“intractable”)- need pdf of desired parameter (see BPE) without ad-hoc constrain |
| only two assumptions: | <ul style="list-style-type: none">- 1) maximum entropy- 2) pdf factorizes wrt its parameters (\rightarrow mean field approximation) |

D : data set

$Z = \{Z_1, \dots, Z_n\}$: set of (latent) parameter

goal:

- find an approximation $Q(Z)$ for the posterior $P(Z|D)$ via **max ent**
- the more data, $Q(Z) \rightarrow P(Z|D)$ (“learning”)

idea:

- $Q(Z) = \prod_{i=1}^n q_i(Z_i|D)$ mean field approximation



find an approximation $Q(Z)$ for the posterior $P(Z|D)$

D	: data set
Z	: set of n (latent) parameter

KL divergence: tells us how much our information is “off” if we work with the approximation $Q(Z)$

$$KL(P||Q) = - \int P(x) \log \left[\frac{Q(x)}{P(x)} \right] dx \quad (\text{module 1})$$

goal: find the $Q(x)$ that **minimizes $KL(P||Q)$** \rightarrow variational calculus (Euler-Lagrange)

same principle: **maximizing evidence** $P(D)$ \rightarrow variational calculus (Euler-Lagrange)

we run the 2nd idea \rightarrow find an equation we know from stat TD
see also reading Paisley, Blei & Jordan

$$\ln[P(D)] = \ln \left[\int P(D, Z) dZ \right] = \ln \left[\int P(D, Z) \frac{Q(Z)}{Q(Z)} dZ \right] \geq \int Q(Z) \ln \left[\frac{P(D, Z)}{Q(Z)} \right] dZ$$

“Jensen inequality” \rightarrow evidence lower bound (“ELBO”)



find an approximation $Q(Z)$ for the posterior $P(Z|D)$

maximizing evidence $P(D)$

D	: data set
Z	: set of n (latent) parameter

$$\ln[P(D)] = \ln \left[\int P(D, Z) dZ \right] = \ln \left[\int P(D, Z) \frac{Q(Z)}{Q(Z)} dZ \right] \geq \int Q(Z) \ln \left[\frac{P(D, Z)}{Q(Z)} \right] dZ$$

“Jensen inequality”

$$\ln[P(D)] \geq \int Q(Z) \ln[P(D, Z)] dZ - \int Q(Z) \ln[Q(Z)] dZ \quad \text{entropy of } Q(Z)$$

$$\ln[P(D)] \geq \int \prod_{i=1}^n q_i(Z_i|D) \ln[P(D, Z)] dZ - \underbrace{\int \prod_{i=1}^n q_i(Z_i|D) \ln \left[\prod_{i=1}^n q_i(Z_i|D) \right] dZ}_{S_q} \quad \text{assumption 2}$$



find an approximation $Q(Z)$ for the posterior $P(Z|D)$

maximizing evidence $P(D)$

D	: data set
Z	: set of n (latent) parameter

$$S_q = \int \prod_{i=1}^n q_i(Z_i|D) \ln \left[\prod_{i=1}^n q_i(Z_i|D) \right] dZ$$

one n -dimensional problem (hard)

$$= \int \{q_1(Z_1|D) q_2(Z_2|D) \dots q_N(Z_N|D)\} \ln[q_1(Z_1|D)] dZ_1 dZ_2 \dots dZ_n +$$

$$\int \{q_1(Z_1|D) q_2(Z_2|D) \dots q_N(Z_N|D)\} \ln[q_2(Z_2|D)] dZ_1 dZ_2 \dots dZ_n +$$

... +

$$\int \{q_1(Z_1|D) q_2(Z_2|D) \dots q_N(Z_N|D)\} \ln[q_n(Z_n|D)] dZ_1 dZ_2 \dots dZ_n$$



find an approximation $Q(Z)$ for the posterior $P(Z|D)$

maximizing evidence $P(D)$

D
 Z

: data set

: set of n (latent) parameter

$$\begin{aligned} S_q &= \int \{q_1(Z_1|D) q_2(Z_2|D) \dots q_N(Z_N|D)\} \ln[q_1(Z_1|D)] dZ_1 dZ_2 \dots dZ_n + \\ &\quad \int \{q_1(Z_1|D) q_2(Z_2|D) \dots q_N(Z_N|D)\} \ln[q_2(Z_2|D)] dZ_1 dZ_2 \dots dZ_n + \dots + \\ &\quad \int \{q_1(Z_1|D) q_2(Z_2|D) \dots q_N(Z_N|D)\} \ln[q_n(Z_n|D)] dZ_1 dZ_2 \dots dZ_n \\ &= \int q_1(Z_1|D) \ln[q_1(Z_1|D)] dZ_1 \int q_2(Z_2|D) dZ_2 \dots \int q_n(Z_n|D) dZ_n + \\ &\quad \int q_2(Z_2|D) \ln[q_2(Z_2|D)] dZ_2 \int q_1(Z_1|D) dZ_1 \dots \int q_n(Z_n|D) dZ_n + \dots + \end{aligned}$$



find an approximation $Q(Z)$ for the posterior $P(Z|D)$

maximizing evidence $P(D)$

 D
 Z

: data set

: set of n (latent) parameter

$$\begin{aligned} S_q &= \int q_1(Z_1|D) \ln[q_1(Z_1|D)] dZ_1 \int q_2(Z_2|D) dZ_2 \dots \int q_n(Z_n|D) dZ_n + \\ &\quad \int q_2(Z_2|D) \ln[q_2(Z_2|D)] dZ_2 \int q_1(Z_1|D) dZ_1 \dots \int q_n(Z_n|D) dZ_n + \dots + \\ &= \sum_{i=1}^n \int q_i(Z_i|D) \ln[q_i(Z_i|D)] dZ_i \left\{ \prod_{j \neq i}^{n-1} \int q_j(Z_j|D) dZ_j \right\} \\ &= \sum_{i=1}^n \langle \ln[q_i(Z_i|D)] \rangle \left\{ \prod_{j \neq i}^{n-1} \int q_j(Z_j|D) dZ_j \right\} = \mathbf{1} \text{ (the } q_j \text{ are all a pdf of } Z_j\text{)} \end{aligned}$$

$$S_q = \sum_{i=1}^n \langle \ln[q_i(Z_i|D)] \rangle = \sum_{i=1}^n \int q_i(Z_i|D) \ln[q_i(Z_i|D)] dZ_i \quad \text{in one-dimensional problems (not so hard)}$$



find an approximation $Q(Z)$ for the posterior $P(Z|D)$

maximizing evidence $P(D)$

D	: data set
Z	: set of n (latent) parameter

$$\ln[P(D)] \geq \int \prod_{i=1}^n q_i(Z_i|D) \ln[P(D|Z)] dZ - \underbrace{\int \prod_{i=1}^n q_i(Z_i|D) \ln \left[\prod_{i=1}^n q_i(Z_i|D) \right] dZ}_{S_q}$$

$$\ln[P(D)] \geq \int \prod_{i=1}^n q_i(Z_i|D) \ln[P(D, Z)] dZ - \sum_{i=1}^n \int q_i(Z_i|D) \ln[q_i(Z_i|D)] dZ_i$$

$$E(D, Z) := \ln[P(D, Z)]$$

$$\ln[P(D)] \geq \int \prod_{i=1}^n q_i(Z_i|D) E(D, Z) dZ - \sum_{i=1}^n \int q_i(Z_i|D) \ln[q_i(Z_i|D)] dZ_i$$



find an approximation $Q(Z)$ for the posterior $P(Z|D)$

maximizing evidence $P(D)$

D
 Z

: data set

: set of n (latent) parameter

$$\ln[P(D)] \geq \int \prod_{i=1}^n q_i(Z_i|D) E(D, Z) dZ - \sum_{i=1}^n \int q_i(Z_i|D) \ln[q_i(Z_i|D)] dZ_i$$

has the structure of $F = U - TS$, but with $-U + TS$ term (aka **negative variational free energy**)

solution:

$$q_i(Z_i|D) = \frac{1}{Z} \exp \left(\langle E(Z_i, \{Z_{j \neq i}\}, D) \rangle_{\{j \neq i\}} \right)$$
$$Z = \int \exp \left(\langle E(Z_i, \{Z_{j \neq i}\}, D) \rangle_{\{j \neq i\}} \right) dZ_{\{j \neq i\}}$$



example: we can measure μ and $\frac{1}{\sigma^2} = \tau$ from D

goal: find $q_\mu(\mu|D)$ and $q_\tau(\tau|D)$

D	: data set
Z	: set of n (latent) parameter
σ^2	: variance
μ	: mean
$\frac{1}{\sigma^2} = \tau$: precision

$$q_i(Z_i|D) = \frac{1}{Z} \exp \left(\langle E(Z_i, \{Z_{j \neq i}\}, D) \rangle_{\{j \neq i\}} \right)$$

$$\ln[q_\mu(\mu|D)] = \langle E(\mu, \tau, D) \rangle_\tau - \ln(Z)$$

$$= \langle \ln[P(D|\mu, \tau)P(\mu|\tau)P(\tau)] \rangle_\tau - \ln(Z)$$

$$= \langle \ln[P(D|\mu, \tau)] \rangle_\tau + \langle \ln[P(\mu|\tau)] \rangle_\tau + \langle \ln[P(\tau)] \rangle_\tau - \ln(Z)$$



example: we can measure μ and $\frac{1}{\sigma^2} = \tau$ from D

goal: find $q_\mu(\mu|D)$ and $q_\tau(\tau|D)$

D	: data set of size K
Z	: set of n (latent) parameter
σ^2	: variance
μ	: mean
$\frac{1}{\sigma^2} = \tau$: precision

$$\ln[q_\mu(\mu|D)] = \langle \ln[P(D|\mu, \tau)] \rangle_\tau + \langle \ln[P(\mu|\tau)] \rangle_\tau + \langle \ln[P(\tau)] \rangle_\tau - \ln(Z)$$

if no constrain: gaussian

$$P(D|\mu, \tau) = \prod_{k=1}^K \mathcal{N}(x_k|\mu, \tau)$$

support is $[0, +\infty) \rightarrow$ max ent

$$P(\tau) = \Gamma(\tau|a, b) = \frac{b^a \tau^{a-1} e^{-b\tau}}{\int_0^\infty t^{a-1} e^{-t} dt}$$

with yet unknown parameter a and b

if μ has support $(-\infty, +\infty)$ it is drawn from a gaussian of yet unknown μ_0 and precision τ_0 (= small pos number, max ent)

$$P(\mu|\tau) = \mathcal{N}(\mu|\mu_0, \tau_0)$$



example: we can measure μ and $\frac{1}{\sigma^2} = \tau$ from D

goal: find $q_\mu(\mu|D)$ and $q_\tau(\tau|D)$

D	: data set of size K
Z	: set of n (latent) parameter
σ^2	: variance
μ	: mean
$\frac{1}{\sigma^2} = \tau$: precision

$$\ln[q_\mu(\mu|D)] = \langle \ln[P(D|\mu, \tau)] \rangle_\tau + \langle \ln[P(\mu|\tau)] \rangle_\tau + \langle \ln[P(\tau)] \rangle_\tau - \ln(Z)$$

$$P(D|\mu, \tau) = \prod_{k=1}^K \mathcal{N}(x_k|\mu, \tau) \quad P(\mu|\tau) = \mathcal{N}(\mu|\mu_0, \tau_0) \quad P(\tau) = \Gamma(\tau|a, b) = \frac{b^a \tau^{a-1} e^{-b\tau}}{\int_0^\infty t^{a-1} e^{-t} dt}$$

after some ([lengthy algebra](#)):

$$\ln[q_\mu(\mu|D)] = -\frac{\langle \tau \rangle_\tau}{2} \left\{ \sum_k (x_k - \mu)^2 + \tau_0 (\mu - \mu_0)^2 \right\} + \text{constant terms}$$

$$q_\mu(\mu|D) \sim \mathcal{N}(\mu|\mu_K, \lambda_K^{-1})$$

where

$$\mu_K = \frac{\tau_0 \mu_0 + K \bar{x}}{\tau_0 + K}$$
$$\lambda_K = (\tau_0 + K) \langle \tau \rangle_\tau$$
$$\bar{x} = \frac{1}{K} \sum_{k=1}^K x_k$$



example: we can measure μ and $\frac{1}{\sigma^2} = \tau$ from D

goal: find $q_\mu(\mu|D)$ and $q_\tau(\tau|D)$

D	: data set of size K
Z	: set of n (latent) parameter
σ^2	: variance
μ	: mean
$\frac{1}{\sigma^2} = \tau$: precision

$$\ln[q_\mu(\mu|D)] = \langle \ln[P(D|\mu, \tau)] \rangle_\tau + \langle \ln[P(\mu|\tau)] \rangle_\tau + \langle \ln[P(\tau)] \rangle_\tau - \ln(Z)$$

$$P(D|\mu, \tau) = \prod_{k=1}^K \mathcal{N}(x_k|\mu, \tau) \quad P(\mu|\tau) = \mathcal{N}(\mu|\mu_0, \tau_0) \quad P(\tau) = \Gamma(\tau|a, b) = \frac{b^a \tau^{a-1} e^{-b\tau}}{\int_0^\infty t^{a-1} e^{-t} dt}$$

same for $q_\tau(\tau|D)$ ([lengthy algebra](#))

$$q_\tau(\tau|D) \sim \Gamma(\tau|a_K, b_K)$$

where

$$a_K = a + \frac{K+1}{2}$$
$$b_K = b + \frac{1}{2} \langle \sum_k (x_k - \mu)^2 + \tau_0 (\mu - \mu_0)^2 \rangle_\mu$$



$$q_{\mu}(\mu|D) \sim \mathcal{N}(\mu|\mu_K, \lambda_K^{-1})$$

where

$$\mu_K = \frac{\tau_0 \mu_0 + K \bar{x}}{\tau_0 + K}$$

$$\lambda_K = (\tau_0 + K) \langle \tau \rangle_{\tau}$$

$$\bar{x} = \frac{1}{K} \sum_{k=1}^K x_k$$

D	: data set of size K
Z	: set of n (latent) parameter
σ^2	: variance
μ	: mean
$\frac{1}{\sigma^2} = \tau$: precision

$$q_{\tau}(\tau|D) \sim \Gamma(\tau|a_K, b_K)$$

where

$$a_K = a + \frac{K+1}{2}$$

$$b_K = b + \frac{1}{2} \langle \sum_k (x_k - \mu)^2 + \tau_0 (\mu - \mu_0)^2 \rangle_{\mu}$$

$$\langle \tau \rangle_{\tau} = \int \tau q_{\tau}(\tau|D) d\tau = \frac{a_K}{b_K}$$

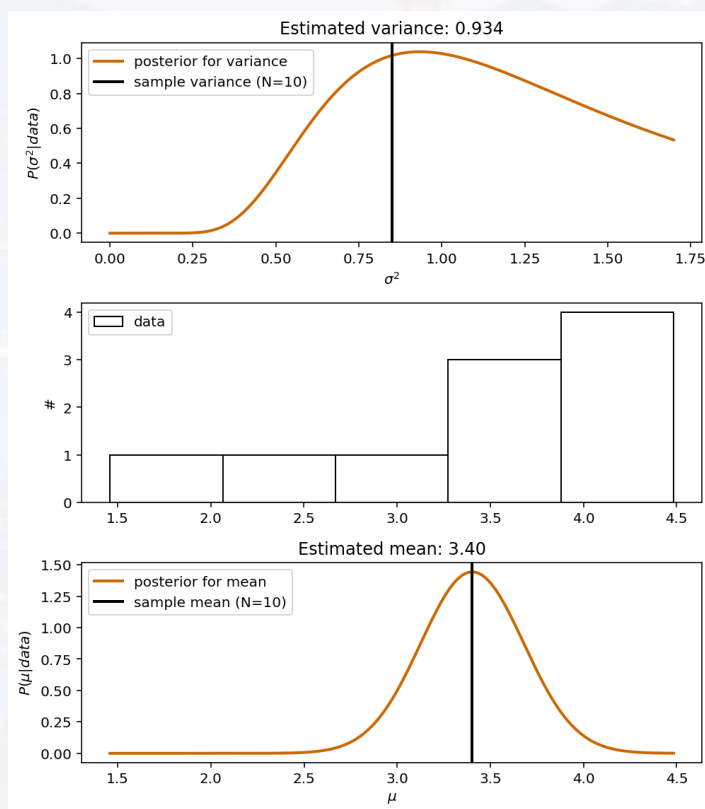
We start with setting τ_0 , μ_0 , a and b to **small positive values** (largest ignorance, broad peaks) and use the circular dependencies (like actual EM)

- from $\langle \tau \rangle_{\tau}$ (here integral over the gamma distribution) we can calculate $q_{\mu}(\mu|D)$
- from that we can calculate b_K and $\langle \mu \rangle_{\mu} = \mu_K$ and $\langle \mu^2 \rangle_{\mu} = \frac{1}{\lambda_K} + \mu_K^2$ etc

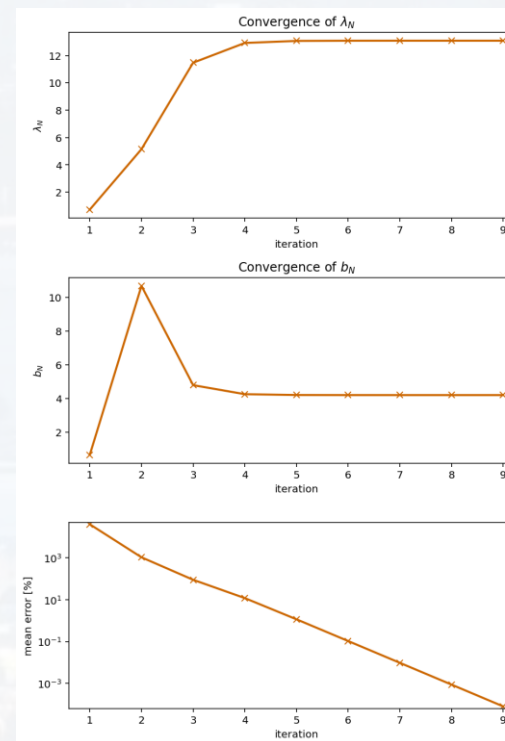


- note:**
- calculations for $q_i(Z_i|D)$ will lead to an **iterative procedure** like for actual EM
 - instead of point estimates for Z_i (MLE) as before, **we get the actual posterior $q_i(Z_i|D)$**
 - these distributions get more accurate the larger $D \rightarrow$ **learning** (see BPE)
 - we only use maximum entropy

see Var_Bayes_Example.py



```
data = np.random.normal(3, 1, (10,))  
Var_Bayes_Example(data)
```





Thank you very much for your attention!

