

Problem (1)

Consider i.i.d. real random variables Y_1, Y_2, \dots with common density f given by

$$f(y) = \begin{cases} \frac{1}{2}e^{-y}, & y \geq 0 \\ e^{2y} & y < 0 \end{cases}$$

Let $S_0 = 0$ and define $(S_n)_{n \geq 1}$ as

$$S_n = Y_1 + \dots + Y_n, \quad \text{for } n \in \mathbb{N}$$

Also define

$$\mathcal{F}_n = \begin{cases} \{\emptyset, \Omega\} & \text{for } n = 0 \\ \sigma(Y_1, \dots, Y_n) & \text{for } n \in \mathbb{N} \end{cases}$$

- (a) For $n \in \mathbb{N}_0$ determine $\mathbb{E}[e^{\alpha S_{n+1}} | \mathcal{F}_n]$ for all $\alpha \in \mathbb{R}$.
 (b) Determine $\mathbb{P}(S_n < 0 < S_{n+1} | \mathcal{F}_n)$ for $n \in \mathbb{N}$.

Solution

- (a) Let $n \in \mathbb{N}_0$,

$$\begin{aligned} \mathbb{E}[e^{\alpha S_{n+1}} | \mathcal{F}_n] &= \mathbb{E}[e^{\alpha(Y_1 + \dots + Y_n + Y_{n+1})} | \mathcal{F}_n] \\ &= \mathbb{E}[e^{\alpha(Y_1 + \dots + Y_n)} e^{\alpha Y_{n+1}} | \mathcal{F}_n] \\ &\stackrel{D.10(6)}{=} e^{\alpha(Y_1 + \dots + Y_n)} \mathbb{E}[e^{\alpha Y_{n+1}} | \mathcal{F}_n] \\ &\stackrel{D.10(8)}{=} e^{\alpha(Y_1 + \dots + Y_n)} \mathbb{E}[e^{\alpha Y_{n+1}}] \\ &\stackrel{LOTUS}{=} e^{\alpha(Y_1 + \dots + Y_n)} \lim_{n \rightarrow \infty} \int_{-n}^n e^{\alpha y} f(y) \, dy \\ &= e^{\alpha(Y_1 + \dots + Y_n)} \left(\lim_{n \rightarrow \infty} \int_{-n}^0 e^{\alpha y} e^{2y} \, dy + \lim_{n \rightarrow \infty} \int_0^n e^{\alpha y} \frac{1}{2} e^{-y} \, dy \right) \\ &= e^{\alpha S_n} \left(\lim_{n \rightarrow \infty} \int_{-n}^0 e^{\alpha y} e^{2y} \, dy + \lim_{n \rightarrow \infty} \int_0^n e^{\alpha y} \frac{1}{2} e^{-y} \, dy \right) \end{aligned}$$

For $\lim_{n \rightarrow \infty} \int_{-n}^0 e^{\alpha y} e^{2y} \, dy$ we have:

$$\begin{aligned} \lim_{n \rightarrow \infty} \int_{-n}^0 e^{\alpha y} e^{2y} \, dy &= \int_{-\infty}^0 e^{(\alpha+2)y} \, dy \\ &= \lim_{n \rightarrow \infty} \left[\frac{e^{(\alpha+2)y}}{\alpha+2} \right]_{-n}^0 \\ &= \begin{cases} \frac{1}{\alpha+2}, & \text{if } \alpha > -2 \\ \infty, & \text{otherwise} \end{cases} \end{aligned}$$

For $\lim_{n \rightarrow \infty} \int_0^n e^{\alpha y} \frac{1}{2} e^{-y} dy$ we have:

$$\begin{aligned} \lim_{n \rightarrow \infty} \int_0^n e^{\alpha y} \frac{1}{2} e^{-y} dy &= 2 \int_0^\infty e^{(\alpha-1)y} dy \\ &= \frac{1}{2} \lim_{n \rightarrow \infty} \left[\frac{e^{(\alpha-1)y}}{\alpha-1} \right]_0^n \\ &= \begin{cases} \frac{1}{2\alpha-2}, & \text{if } \alpha < 1 \\ \infty, & \text{otherwise} \end{cases} \end{aligned}$$

Combining these we get:

$$\mathbb{E} [e^{\alpha S_{n+1}} | \mathcal{F}_n] = \begin{cases} e^{\alpha S_n} \left(\frac{1}{\alpha+2} + \frac{1}{2\alpha-2} \right), & \text{if } -2 < \alpha < 1 \\ \infty, & \text{otherwise} \end{cases}$$

(b) Let $n \in \mathbb{N}$:

$$\begin{aligned} \mathbb{P}(S_n < 0 < S_{n+1} | \mathcal{F}_n) &\stackrel{D.1(4)}{=} \mathbb{E}[\mathbb{1}_{\{S_n < 0 < S_{n+1}\}} | \mathcal{F}_n] \\ &= \mathbb{E}[\mathbb{1}_{\{S_n < 0\}} \mathbb{1}_{\{0 < S_{n+1}\}} | \mathcal{F}_n] \\ &= \mathbb{E}[\mathbb{1}_{\{S_n < 0\}} \mathbb{1}_{\{-S_n < Y_{n+1}\}} | \mathcal{F}_n] \\ &\stackrel{\perp}{=} \mathbb{1}_{\{S_n < 0\}} \mathbb{E}[\mathbb{1}_{\{-S_n < Y_{n+1}\}} | \mathcal{F}_n] \\ &\stackrel{D.7}{=} \mathbb{1}_{\{S_n < 0\}} \mathbb{E}[\mathbb{1}_{\{-s < Y_{n+1}\}} | \mathcal{F}_n]_{s=S_n} \\ &= \mathbb{1}_{\{S_n < 0\}} \lim_{n \rightarrow \infty} \int_{-s}^n \frac{1}{2} e^{-y} dy \\ &= \mathbb{1}_{\{S_n < 0\}} \frac{e^s}{2} \\ &= \begin{cases} \frac{e^{S_n}}{2}, & \text{if } S_n < 0 \\ 0, & \text{otherwise} \end{cases} \end{aligned}$$

Problem (2)

Let the state space be $S = \mathbb{N}_0 = \{0, 1, \dots\}$ and for $i \in S$ let $p_i \in [0, 1]$. Consider the Markov chain with state space S specified as follows: A particle which in state $i \in S$ moves to state $i + 1$ in one step with probability p_i and to state 0 with probability $1 - p_i$

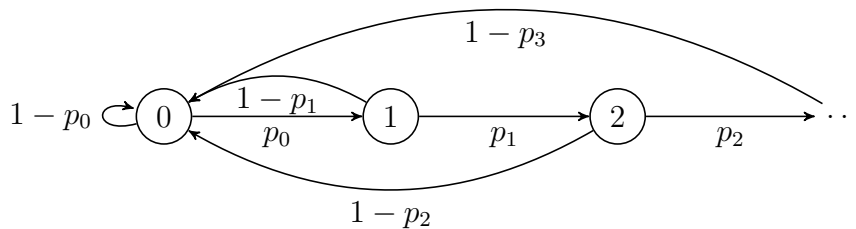
- Argue that the chain is irreducible if $p_i \in (0, 1)$ for all $i \in S$. Also, indicate the transition diagram in this case.
- Determine p_{02}^4 . For $n \in \mathbb{N}$ determine $\mathbb{P}_0(T_0 \geq n)$ and $\mathbb{P}_0(T_0 = n)$.
- Show that state 0 is recurrent if and only if $\prod_{i=0}^{\infty} p_i = 0$.
- Give a concrete example of probabilities $(p_i)_{i=0}^{\infty}$ such that $p_i \in (0, 1)$ for all $i \in S$ and the chain is transient.
- Assume that $p_i \in (0, 1)$ for all $i \in S$. Show that there exists a unique stationary distribution if and only if $\sum_{n=0}^{\infty} (p_0 p_1 \cdots p_n) < \infty$. Determine the stationary distribution when it exists.

Solution

- To argue that the chain is irreducible, we need to argue that any state can be reached from any other state with positive probability.

If $p_i \in (0, 1) \forall i \in S$, we have probability $p_i > 0$ for moving to state $i + 1$ and probability $1 - p_i > 0$ for moving to state 0.

For states i, j , we have that if $j > i$, then to go from i to j we can move from i to $i + 1, i + 2, \dots, j$, with positive probability. If $j < i$ we can move to 0 from any state (with positive probability) and then move to state $1, 2, \dots, i$ with positive probability. As this covers all cases, the chain is irreducible. Here is a sketch for the transition diagram:



- We want to determine the probability to go from state 0 to state 2 in exactly 4 steps. There are only 2 combinations where this is possible as represented by the following paths:

$$0 \rightarrow 1 \rightarrow 0 \rightarrow 1 \rightarrow 2$$

with probability $p_0^2 p_1 (1 - p_1)$, or

$$0 \rightarrow 0 \rightarrow 0 \rightarrow 1 \rightarrow 2$$

with probability $(1 - p_0)^2 p_0 p_1$,

So we have that

$$p_{02}^4 = p_0^2 p_1 (1 - p_1) + (1 - p_0)^2 p_0 p_1$$

We now calculate $\mathbb{P}_0(T_0 \geq n)$ and $\mathbb{P}_0(T_0 = n)$. Let $n \in \mathbb{N}$. T_0 is the return time to state 0, meaning that if $T_0 = n$ then we move to state 0 at time $n - 1$ from a state i :

For $\mathbb{P}_0(T_0 \geq n)$, we want to calculate the probability of the first return time being at n or later, if we start in state 0, i.e. not returning to 0 for the first $n - 2$ steps and then no restrictions from time $n - 1$ and onwards. Intuitively this is

$$\mathbb{P}_0(T_0 \leq n) = \prod_{i=0}^{n-2} p_i$$

For $\mathbb{P}_0(T_0 = n)$ we want the probability of the first return time being exactly at time n . This is the same as not returning for the first $n - 2$ steps and then exactly returning at the $n - 1$ 'th step

$$\mathbb{P}_0(T_0 = n) = \left(\prod_{i=0}^{n-2} p_i \right) (1 - p_{n-1})$$

(c) Assume state 0 is recurrent:

Then by definition,

$$\mathbb{P}_0(T_0 < \infty) = 1$$

This implies that

$$\mathbb{P}_0(T_0 \geq \infty) = \prod_{i=0}^{\infty} p_i = 0$$

On the other hand, assume that $\prod_{i=0}^{\infty} p_i = 0$, then there is probability 0 of the first time being after $n = \infty$. I.e. that

$$\mathbb{P}_0(T_0 \geq \infty) = \prod_{i=0}^{\infty} p_i = 0 \implies \mathbb{P}_0(T_0 < \infty) = 1$$

Which is the definition of state 0 being recurrent.

(d) To ensure that the chain is transient we must have that

$$\mathbb{P}_0(T_0 = \infty) = \prod_{i=0}^{\infty} p_i > 0 \implies \mathbb{P}_0(T_0 < \infty) < 1$$

We therefore choose $p_i = e^{-1/2^i}$, which is always in $(0, 1)$ and

$$\prod_{i=0}^{\infty} p_i > 0$$

meaning that the chain is transient.

- (e) Assume that there exists a unique stationary distribution, π . By Theorem 6.22 we have that $\pi_j = 1/\mathbb{E}_j[T_j]$ and that it is positive recurrent. We then have:

$$\mathbb{E}_0[T_0] < \infty \Rightarrow \sum_{n=0}^{\infty} \prod_{i=0}^n p_i < \infty$$

Now, assume that $\sum_{n=0}^{\infty} \prod_{i=0}^n p_i < \infty$:

$$\sum_{n=0}^{\infty} \prod_{i=0}^n p_i < \infty \Rightarrow \mathbb{E}_0 \left[\sum_{n=0}^{\infty} \prod_{i=0}^n p_i \right] < \infty$$

By remark 6.20, P is positive recurrent, and by 6.22(4) and 6.22(1), we then have that there exists a unique stationary distribution.

To determine the stationary distribution we know that $\pi_k = \pi_{k-1} \prod_{i=0}^{k-1} p_i$, and by extension, $\pi_k = \pi_0 \prod_{i=0}^{k-1} p_i$. So we can conclude that

$$\pi_0 = \frac{1}{\sum_{k=0}^{\infty} \prod_{i=0}^{k-1} p_i} \quad \text{for } i = 0$$

and

$$\pi_k = \pi_0 \prod_{i=0}^{k-1} p_i = \frac{\prod_{i=0}^{k-1} p_i}{\sum_{k=0}^{\infty} \prod_{i=0}^{k-1} p_i}, \quad \text{for } i \in \mathbb{N}$$

Problem (3)

Consider Proposition 6.3 in the lecture notes.

(a) Prove (6.4).

Let $i \in S$ and μ denote a probability measure on S . Assume $\mathbb{P}_\mu(T_i < \infty) = 1$ and $\mathbb{P}_i(T_i < \infty) = 1$.

(b) Show that T_i^n is finite \mathbb{P}_μ -a.s. for every $n \in \mathbb{N}$.

(c) Show that T_i^1 and $T_i^2 - T_i^1$ are independent under \mathbb{P}_μ and that the distribution of $T_i^2 - T_i^1$ under \mathbb{P}_μ is the same as the distribution of T_i^1 under \mathbb{P}_i .

Essentially, the above proves Proposition 6.3. The only remaining step is to extend (c) above by induction to get Proposition 6.3(2)(b) – (c).

Solution

(a) We need to prove that

$$\mathbb{E}_\mu = \mathbb{P}_\mu(T_i < \infty)(1 + \mathbb{E}_i[N_i])$$

Let $i \in S$ and μ a probability measure on S :

$$\begin{aligned} \mathbb{E}_\mu[N_i] &\stackrel{6.2}{=} \mathbb{E}_\mu \left[((1 + N_i) \circ \theta^{T_i}) \mathbb{1}_{\{T_i < \infty\}} \right] \\ &\stackrel{\text{Tower}}{=} \mathbb{E}_\mu \left[\mathbb{E}_\mu \left[((1 + N_i) \circ \theta^{T_i}) \mathbb{1}_{\{T_i < \infty\}} \right] \middle| \mathcal{F}_{T_i}^X \right] \\ &= \mathbb{E}_\mu \left[\mathbb{1}_{\{T_i < \infty\}} \mathbb{E}_\mu \left[((1 + N_i) \circ \theta^{T_i}) \middle| \mathcal{F}_{T_i}^X \right] \right] \\ &\stackrel{4.25}{=} \mathbb{E}_\mu \left[\mathbb{1}_{\{T_i < \infty\}} \mathbb{E}_{X_{T_i}} [(1 + N_i)] \right] \\ &= \mathbb{E}_\mu \left[\mathbb{1}_{\{T_i < \infty\}} \mathbb{E}_i [(1 + N_i)] \right] \\ &= \mathbb{E}_\mu \left[\mathbb{1}_{\{T_i < \infty\}} \right] \mathbb{E}_i [(1 + N_i)] \\ &= \mathbb{P}_\mu(T_i < \infty) \mathbb{E}_i [(1 + N_i)] \\ &= \mathbb{P}_\mu(T_i < \infty) (1 + \mathbb{E}_i[N_i]) \end{aligned}$$

Which is the desired result.

(b) We assume that $\mathbb{P}_\mu(T_i < \infty) = 1$ and $\mathbb{P}_i(T_i < \infty) = 1$. From (6.3) we know that

$$\mathbb{P}_\mu(T_i^{n+1} < \infty) = \mathbb{P}_\mu(T_i < \infty) [\mathbb{P}_i(T_i < \infty)]^n$$

Which is equivalent with

$$\mathbb{P}_\mu(T_i^n < \infty) = \mathbb{P}_\mu(T_i < \infty) [\mathbb{P}_i(T_i < \infty)]^{n-1}$$

So we have:

$$\mathbb{P}_\mu(T_i^n < \infty) = 1 \cdot 1^{n-1} = 1$$

Which means that $T_i^n < \infty$ \mathbb{P}_μ -a.s.

(c) We first show that they are equally distributed. We need to show that

$$\mathbb{P}_\mu(T_i^2 - T_i^1 \in B) = \mathbb{P}_i(T_i^1 \in B) \quad \forall B \in \mathcal{B}(S).$$

Let $B \in \mathcal{B}(S)$:

$$\begin{aligned} \mathbb{P}_\mu(T_i^2 - T_i^1 \in B) &= \mathbb{E}_\mu \left[\mathbb{1}_{\{T_i^2 - T_i^1 \in B\}} \right] \\ &\stackrel{6.1}{=} \mathbb{E}_\mu \left[\mathbb{1}_{\{T_i \circ \theta^{T_i^1} \in B\}} \right] \\ &\stackrel{\text{Assumption}}{=} \mathbb{E}_\mu \left[\mathbb{1}_{\{T_i^1 < \infty\}} \mathbb{1}_{\{T_i \circ \theta^{T_i^1} \in B\}} \right] \\ &\stackrel{\text{Tower}}{=} \mathbb{E}_\mu \left[\mathbb{1}_{\{T_i^1 < \infty\}} \mathbb{E}_\mu \left[\mathbb{1}_{\{T_i \circ \theta^{T_i^1} \in B\}} \middle| \mathcal{F}_{T_i^1}^X \right] \right] \\ &\stackrel{4.25}{=} \mathbb{E}_\mu \left[\mathbb{1}_{\{T_i^1 < \infty\}} \mathbb{E}_{X_{T_i^1}} \left[\mathbb{1}_{\{T_i^1 \in B\}} \right] \right] \\ &= \mathbb{E}_\mu \left[\mathbb{1}_{\{T_i^1 < \infty\}} \mathbb{P}_{X_{T_i^1}}(T_i^1 \in B) \right] \\ &\stackrel{X_{T_i^1=i}}{=} \mathbb{P}_\mu(T_i^1 < \infty) \mathbb{P}_i(T_i^1 \in B) \\ &\stackrel{\text{Assumption}}{=} \mathbb{P}_i(T_i^1 \in B) \end{aligned}$$

We now show that they are independent. Let $B_1, B_2 \in \mathcal{B}(S)$:

$$\begin{aligned} \mathbb{P}_\mu(T_1 \in B_1, T_2 - T_1 \in B_2) &= \mathbb{E}_\mu \left[\mathbb{1}_{\{T_1 \in B_1\}} \mathbb{1}_{\{T_2 - T_1 \in B_2\}} \right] \\ &\stackrel{\text{Tower}}{=} \mathbb{E}_\mu \left[\mathbb{E}_\mu \left[\mathbb{1}_{\{T_1 \in B_1\}} \mathbb{1}_{\{T_2 - T_1 \in B_2\}} \middle| \mathcal{F}_{T_1}^X \right] \right] \\ &\stackrel{D.10(6)}{=} \mathbb{E}_\mu \left[\mathbb{1}_{\{T_1 \in B_1\}} \mathbb{E}_\mu \left[\mathbb{1}_{\{T_2 - T_1 \in B_2\}} \middle| \mathcal{F}_{T_1}^X \right] \right] \\ &= \mathbb{E}_\mu \left[\mathbb{1}_{\{T_1^1 \in B_1\}} \mathbb{E}_\mu \left[\mathbb{1}_{\{T_i^1 \circ \theta^{T_i^1} \in B_2\}} \mathbb{1}_{\{T_i^1 < \infty\}} \middle| \mathcal{F}_{T_1}^X \right] \right] \\ &\stackrel{4.25}{=} \mathbb{E}_\mu \left[\mathbb{1}_{\{T_1^1 \in B_1\}} \mathbb{1}_{\{T_i^1 < \infty\}} \mathbb{E}_{X_{T_i^1}} \left[\mathbb{1}_{\{T_i^1 \in B_2\}} \middle| \mathcal{F}_{T_1}^X \right] \right] \\ &= \mathbb{P}_\mu(T_1^1 \in B_1) \mathbb{P}_\mu(T_i^1 < \infty) \mathbb{P}_i(T_i^1 \in B_2) \\ &= \mathbb{P}_\mu(T_i^1 \in B_1) \mathbb{P}_i(T_i^1 \in B_2) \end{aligned}$$

Which are the desired results.

Problem (4)

Let $S = \{1, 2, 3, 4, 5, 6\}$ and

$$P = \begin{pmatrix} 0.1 & 0.4 & 0.1 & 0 & 0 & 0.4 \\ 0 & 0.2 & 0.3 & 0.5 & 0 & 0 \\ 0 & 0 & 0.2 & 0.3 & 0.5 & 0 \\ 0 & 0 & 0.5 & 0.2 & 0.2 & 0.1 \\ 0 & 0 & 0.8 & 0 & 0.2 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 \end{pmatrix}$$

Argue that there exists a unique stationary distribution, determine it, and determine $\mathbb{E}_i[T_i]$ for all $i \in S$

Solution

We have a transition matrix, P . We see that the states 1,2 are transient states, since you cannot return to them if you leave them. By exercise 60 this means that $\pi_1 = \pi_2 = 0$ with $\mathbb{E}[T_1] = \mathbb{E}[T_2] = \infty$. We now look at $S^* = \{3, 4, 5, 6\}$. We see that this chain is irreducible, and that all states are recurrent.

We have the following transition matrix:

$$P^* = \begin{pmatrix} 0.2 & 0.3 & 0.5 & 0 \\ 0.5 & 0.2 & 0.2 & 0.1 \\ 0.8 & 0 & 0.2 & 0 \\ 0 & 0 & 1 & 0 \end{pmatrix}$$

We determine the stationary distribution π^* for $S^* = \{3, 4, 5, 6\}$ by using that it satisfies the following:

$$\pi^* = \pi^* P^*$$

We solve this by using CAS to obtain:

$$\pi^* = (0.4591 \quad 0.1722 \quad 0.3515 \quad 0.0172)$$

To determine $\mathbb{E}_i[T_i]$ we know that

$$\pi_i = \frac{1}{\mathbb{E}_i[T_i]} \quad \forall i \in S$$

We thus take each element in π to the power of -1 to obtain the following vector of expected values:

$$E^* = (2.1782 \quad 5.8072 \quad 2.8450 \quad 58.1395)$$

Since we have that P^* is irreducible and that $\mathbb{E}[T_i] < \infty \quad \forall i \in S^*$, P^* is positive recurrent, and by 6.22 has a unique stationary distribution.