

Section 3.8 Problems

1. 3.8.3 part (a) Let X and Y be two independent random variables. Given the marginal pdfs shown below, find the pdf of $X + Y$. In each case, check to see if $X + Y$ belongs to the same family of pdfs as do X and Y .

$$p_X(k) = e^{-\lambda} \frac{\lambda^k}{k!} \text{ and } p_Y(k) = e^{-\mu} \frac{\mu^k}{k!}, k = 0, 1, 2, \dots$$

Answer: We directly apply Theorem 3.8.3, where $W = X + Y$.

$$\begin{aligned} p_W(w) &= \sum_{\text{all } x} p_X(x) \cdot p_Y(w-x) \\ &= \sum_{\text{all } x} e^{-\lambda} \frac{\lambda^x}{x!} \cdot e^{-\mu} \frac{\mu^{w-k}}{(w-k)!} \\ &= e^{-\lambda-\mu} \sum_{\text{all } x} \frac{\lambda^x \mu^{w-k}}{x!(w-k)!} \\ &= \frac{e^{-\lambda-\mu}}{w!} (\lambda + \mu)^w = \frac{e^{-\lambda-\mu} (\lambda + \mu)^w}{w!} \end{aligned}$$

We note that $W = X + Y$ belongs to the same family of pdfs as X and Y .

2. 3.8.7 Let Y be a continuous nonnegative random variable. Show that $W = Y^2$ has pdf $f_W(w) = \frac{1}{2\sqrt{w}} f_Y(\sqrt{w})$. (Hint: First find $F_W(w)$.)

Answer: Observe that we can express $F_W(w)$ differently: $F_W(w) = P(W \leq w) = P(Y^2 \leq w) = P(Y \leq \sqrt{w}) = F_Y(\sqrt{w})$, so then $F_W(w) = F_Y(\sqrt{w})$.

Now we recall that $f_W(w) = \frac{d}{dw} F_W(w) = \frac{d}{dw} F_Y(\sqrt{w}) = \frac{1}{2\sqrt{w}} f_Y(\sqrt{w})$. So then $W = Y^2$ has pdf $f_W(w) = \frac{1}{2\sqrt{w}} f_Y(\sqrt{w})$

3. 3.8.8 Let Y be a uniform random variable over the interval $[0, 1]$. Find the pdf of $W = Y^2$.

Answer: We discovered in problem 2 that the pdf of a continuous nonnegative random variable $W = Y^2$ is $f_W(w) = \frac{1}{2\sqrt{w}} f_Y(\sqrt{w})$. So then $f_W(w) = \frac{1}{2\sqrt{w}}(0) + \frac{1}{2\sqrt{w}}(1) = \frac{1}{2\sqrt{w}}, 0 \leq w \leq 1$.

4. 3.8.13 Suppose that X and Y are two independent random variables, where $f_X(x) = xe^{-x}, x \geq 0$, and $f_Y(y) = e^{-y}, y \geq 0$. Find the pdf of Y/X .

Answer: We directly apply Theorem 3.8.4 to find the pdf, where $W = Y/X$:

$$\begin{aligned} f_W(w) &= \int_{-\infty}^{\infty} |x| f_X(x) f_Y(wx) dx \\ &= \int_0^{\infty} |x| x e^{-x} \cdot e^{-wx} dx = \int_0^{\infty} x^2 e^{-x(1+w)} dx \end{aligned}$$

Now we integrate by parts with $u = x^2, du = 2x, v = -\frac{e^{-x(1+w)}}{1+w}, dv = e^{-x(1+w)}$.

$$\begin{aligned} f_W(w) &= \int_0^{\infty} x^2 e^{-x(1+w)} dx = \frac{-x^2 e^{-x(1+w)}}{1+w} - \int_0^{\infty} \frac{-2x e^{-x(1+w)}}{1+w} dx \\ &= \frac{-x^2 e^{-x(1+w)}}{1+w} \Big|_0^{\infty} - \frac{2}{1+w} \int_0^{\infty} x e^{-x(1+w)} dx \\ &= \frac{2}{1+w} \int_0^{\infty} x e^{-x(1+w)} dx \end{aligned}$$

Again, we integrate by parts with $u = x, du = 1, v = -\frac{e^{-x(1+w)}}{1+w}, dv = e^{-x(1+w)}$.

$$\begin{aligned} f_W(w) &= \frac{2}{1+w} \left(-\frac{x e^{-x(1+w)}}{1+w} \Big|_0^{\infty} - \int_0^{\infty} -\frac{e^{-x(1+w)}}{1+w} dx \right) \\ &= \frac{2}{(1+w)^2} \int_0^{\infty} e^{-x(1+w)} dx = -\frac{2}{(1+w)^2} \left(\frac{e^{-x(1+w)}}{1+w} \right) \Big|_0^{\infty} = \frac{2}{(1+w)^3}, \text{ for all } w \geq 0. \end{aligned}$$

Section 3.9 Problems

5. 3.9.5 Suppose that X_i is a random variable for which $E(X_i) = \mu \neq 0, i = 1, 2, \dots, n$. Under what conditions will the following be true?

$$E\left(\sum_{i=1}^n a_i X_i\right) = \mu$$

Answer: We first rewrite the expression on the left-hand side of the above equation using the properties of expected value:

$$E\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i E[X_i] = \sum_{i=1}^n a_i \mu = \mu \sum_{i=1}^n a_i$$

We now make the observation that in order for the initial equation to be true, $\sum_{i=1}^n a_i = 1$.

6. (461 only) 3.9.2 Suppose that $f_{X,Y}(x, y) = \lambda^2 e^{-\lambda(x+y)}, 0 \leq x, 0 \leq y$. Find $E(X + Y)$.

Answer: First, we find the marginal PDFs of X and Y :

$$f_X(x) = \int_0^\infty \lambda^2 e^{-\lambda(x+y)} dy = \lambda^2 e^{-\lambda x} \int_0^\infty e^{-\lambda y} dy = \lambda^2 e^{-\lambda x} \left(-\frac{e^{-\lambda y}}{\lambda} \right) \Big|_0^\infty = \lambda^2 e^{-\lambda x} \left(0 + \frac{1}{\lambda} \right) = \lambda e^{-\lambda x}, \forall x \geq 0$$

$$f_Y(y) = \int_0^\infty \lambda^2 e^{-\lambda(x+y)} dx = \lambda^2 e^{-\lambda y} \int_0^\infty e^{-\lambda x} dx = \dots = \lambda e^{-\lambda y}, \forall y \geq 0$$

We now compute $E[X]$ and $E[Y]$:

$$E[X] = \int_0^\infty x \lambda e^{-\lambda x} dx = -\lambda^2 x e^{-\lambda x} \Big|_0^\infty - \int_0^\infty -\lambda^2 e^{-\lambda x} dx = 0 - \left(\frac{1}{\lambda e^{\lambda x}} \right) \Big|_0^\infty = \frac{1}{\lambda}$$

$$E[Y] = \int_0^\infty y \lambda e^{-\lambda y} dy = E[X] = \frac{1}{\lambda}$$

$$\therefore E[X + Y] = 2 \frac{1}{\lambda} = \frac{2}{\lambda}.$$

7. 3.9.14 Show that $\text{Cov}(aX + b, cY + d) = ac\text{Cov}(X, Y)$ for any constants a, b, c , and d .

Answer:

$$\begin{aligned} \text{Cov}(aX + b, cY + d) &= E[(aX + b)(cY + d)] - E[aX + b] \cdot E[cY + d] \\ &= E[acXY + adX + bcY + bd] - E[aX + b] \cdot E[cY + d] \\ &= E[acXY] + E[adX] + E[bcY] + E[bd] - E[aX + b] \cdot E[cY + d] \\ &= acE[XY] + adE[X] + bcE[Y] + bd - (aE[X] + b)(cE[Y] + d) \\ &= acE[XY] + adE[X] + bcE[Y] + bd - acE[X]E[Y] - adE[X] - bcE[Y] - bd \\ &= acE[XY] - acE[X]E[Y] \\ &= ac(E[XY] - E[X]E[Y]) \\ &= ac\text{Cov}(X, Y). \end{aligned}$$

8. 3.9.18 Suppose that $f_{X,Y}(x, y) = \frac{2}{3}(x + 2y), 0 \leq x \leq 1, 0 \leq y \leq 1$. Find $\text{Var}(X + Y)$.

Answer: Okay. This one's kinda lengthy. $\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y)$. So we will find $\text{Var}(X)$, $\text{Var}(Y)$, and $2\text{Cov}(X, Y)$ separately:

$$f_X(x) = \frac{2}{3} \int_0^1 x + 2y dy = \frac{2}{3} (xy + y^2) \Big|_0^1 = \frac{2}{3} (x + 1), \forall x \in [0, 1] \quad f_Y(y) = \frac{2}{3} \int_0^1 x + 2y dx = \frac{2}{3} \left(\frac{x^2}{2} + 2xy \right) \Big|_0^1 = \frac{2}{3} \left(\frac{1}{2} + 2y \right), \forall y \in [0, 1]$$

$$E[X] = \frac{2}{3} \int_0^1 x^2 + x dx = \frac{2}{3} \left(\frac{x^3}{3} + \frac{x^2}{2} \right) \Big|_0^1 = \frac{2}{3} \left(\frac{1}{3} + \frac{1}{2} \right) = \frac{5}{9} \quad E[Y] = \frac{2}{3} \int_0^1 \frac{y}{2} + 2y^2 dy = \frac{2}{3} \left(\frac{y^2}{4} + \frac{2y^3}{3} \right) \Big|_0^1 = \frac{11}{18}$$

$$E[X^2] = \frac{2}{3} \int_0^1 x^3 + x^2 dx = \frac{2}{3} \left(\frac{x^4}{4} + \frac{x^3}{3} \right) \Big|_0^1 = \frac{2}{3} \left(\frac{1}{4} + \frac{1}{3} \right) = \frac{10}{9} \quad E[Y^2] = \frac{2}{3} \int_0^1 \frac{y^2}{2} + 2y^3 dy = \frac{2}{3} \left(\frac{y^3}{6} + \frac{2y^4}{5} \right) \Big|_0^1 = \frac{4}{9}$$

$$\text{Var}(X) = \frac{7}{18} - \frac{25}{81} = \frac{13}{162} \quad \text{Var}(Y) = \frac{4}{9} - \frac{121}{324} = \frac{23}{324}$$

$$E[XY] = \int_0^1 \int_0^1 \frac{2x^2y + 4xy^2}{3} dy dx = \int_0^1 \left(\frac{x^2y^2}{3} + \frac{4xy^3}{9} \right) \Big|_0^1 dx = \int_0^1 \frac{x^2}{3} + \frac{4x}{9} dx = \frac{x^3 + 2x^2}{9} \Big|_0^1 = \frac{1}{3}$$

$$2\text{Cov}(X, Y) = 2 \left[\frac{1}{3} - \left(\frac{5}{9} \cdot \frac{11}{18} \right) \right] = -\frac{1}{81}$$

$$\text{Var}(X + Y) = \frac{13}{162} + \frac{23}{324} - \frac{1}{81} = \frac{5}{36}. \quad \text{Wow.}$$

9. 3.9.21 A *Poisson random variable* has pdf $p_X(k) = e^{-\lambda} \frac{\lambda^k}{k!}$, $k = 0, 1, 2, \dots$ and $\lambda > 0$. Also, $E(X) = \lambda$. Suppose the Poisson random variable U is the number of calls for technical assistance received by a computer company during the firm's nine normal workday hours, with the average number of calls per hour equal 7.0. Also suppose each call costs the company \$50. Let V be a Poisson random variable representing the number of calls for technical assistance received during a day's remaining fifteen hours. Assume the average number of calls per hour is four for that time period and that each such call costs the company \$60. Find the expected cost and the variance of the cost associated with the calls received during a twenty-four-hour day.

Answer: It is given that $E[U_i] = 7.0$, where $i = 1, 2, \dots, 9$ is the i -th hour in the workday. It is also given that $E[V_j] = 4$, where $j = 1, 2, \dots, 15$ is the j -th hour in the remaining 15 hours. If we set W to be the sum of $50 \cdot U_i$ for each i and $60 \cdot V_j$ for each j . So then we must find $E[W]$ and $\text{Var}(W)$...

$$E[W] = E \left[50 \sum_{i=1}^9 U_i + 60 \sum_{j=1}^{15} V_j \right] = 50 \sum_{i=1}^9 E[U_i] + 60 \sum_{j=1}^{15} E[V_j]$$

$$= 50 \sum_{i=1}^9 7 + 60 \sum_{j=1}^{15} 4 = 50(9)(7) + 60(15)(4) = \$6750$$

$$\text{Var}(W) = 50^2(9)(7) + 60^2(15)(4) = \$373,500$$

Optional R Basics

10. Set up R and start working through basic exercises (see R setup and R exercises on Canvas)

Answer: