

Adaptive State Estimation

Robert Stengel

Optimal Control and Estimation MAE 546

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- Nonlinearity of adaptation
- Parameter-adaptive filtering
- Test for whiteness of the residual
- Bias estimation and noise-adaptive filtering
- Multiple model estimation



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<http://www.princeton.edu/~stengel/MAE546.html>

<http://www.princeton.edu/~stengel/OptConEst.html>

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Nonlinearity of Adaptation

- Adaptation required if
 - System parameters are unknown
 - System structure is unknown
 - Disturbance/measurement statistics are uncertain
- Adaptive estimators are **fundamentally nonlinear**, even if the system is linear
 - Parameters to be estimated multiply the state
 - Statistics are derived from measurement residuals
 - Estimator gain depends on parameter estimates

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“Process Noise”

White noise disturbance input (“process noise”) is similar to random parameter variation

$$\begin{aligned}\dot{\mathbf{x}}(t) &= \mathbf{F}(\mathbf{p})\mathbf{x}(t) + \mathbf{G}\mathbf{u}(t) + \mathbf{L}\mathbf{w}(t) \\ &\simeq \mathbf{F}(\mathbf{p}_o)\mathbf{x}(t) + \left\{ \frac{\partial \mathbf{F}(\mathbf{p}_o)}{\partial \mathbf{p}} \Delta \mathbf{p}(t) \right\} \mathbf{x}(t) + \mathbf{G}\mathbf{u}(t) + \mathbf{L}\mathbf{w}(t) \\ &\triangleq \mathbf{F}(\mathbf{p}_o)\mathbf{x}(t) + \mathbf{G}\mathbf{u}(t) + \begin{bmatrix} \mathbf{L}_w & \mathbf{L}_{\Delta p}[\mathbf{x}(t)] \end{bmatrix} \begin{bmatrix} \mathbf{w}(t) \\ \Delta \mathbf{p}(t) \end{bmatrix}\end{aligned}$$

$$\left\{ \frac{\partial \mathbf{F}(\mathbf{p}_o)}{\partial \mathbf{p}} \Delta \mathbf{p}(t) \right\} \text{ is an } (n \times n) \text{ matrix}$$

$$\mathbf{L}_{\Delta p}[\mathbf{x}(t)] \text{ is } (n \times s) \text{ matrix}$$

$$\mathbf{w}'(t) \triangleq \begin{bmatrix} \mathbf{w}(t) \\ \Delta \mathbf{p}(t) \end{bmatrix}; \quad E[\mathbf{w}'(t)] = \mathbf{0}; \quad E\{\mathbf{w}'(t)\mathbf{w}'^T(\tau)\} = \mathbf{Q}'\delta(t - \tau)$$

... however, model is approximate and nonlinear

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Parameter-Adaptive Estimation (Parameter Identification via Extended Kalman Filter)

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Parameter-Dependent Linear System

Linear system with parameter-dependent sensitivity matrices

$$\begin{aligned}\dot{\mathbf{x}}(t) &= \mathbf{F}(\mathbf{p})\mathbf{x}(t) + \mathbf{G}(\mathbf{p})\mathbf{u}(t) + \mathbf{L}(\mathbf{p})\mathbf{w}(t) \\ \mathbf{z}(t) &= \mathbf{H}\mathbf{x}(t) + \mathbf{n}(t) \\ E \begin{bmatrix} \mathbf{w}(t) \\ \mathbf{n}(t) \end{bmatrix} &= \mathbf{0}; \quad E \left\{ \begin{bmatrix} \mathbf{w}(t) \\ \mathbf{n}(t) \end{bmatrix} \begin{bmatrix} \mathbf{w}^T(\tau) & \mathbf{n}^T(\tau) \end{bmatrix} \right\} = \begin{bmatrix} \mathbf{Q} & \mathbf{0} \\ \mathbf{0} & \mathbf{R} \end{bmatrix} \delta(t - \tau)\end{aligned}$$

Parameter vector, $\mathbf{p}(t)$, could be

Known: a prescribed function of time

Unknown: the output of a random dynamic process

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Dynamic Model for Parameter Estimation

- Augment vector to include original state and parameter vector

$$\mathbf{x}_A(t) \triangleq \begin{bmatrix} \mathbf{x}(t) \\ \mathbf{p}(t) \end{bmatrix}$$

- Augment system model for parameter identification
- System is nonlinear because parameter is contained in the augmented state

$$\begin{aligned}\dot{\mathbf{x}}_A(t) &= \begin{bmatrix} \mathbf{F}[\mathbf{p}(t)]\mathbf{x}(t) + \mathbf{G}[\mathbf{p}(t)]\mathbf{u}(t) + \mathbf{L}[\mathbf{p}(t)]\mathbf{w}_x(t) \\ \mathbf{f}_p[\mathbf{p}(t), \mathbf{w}_p(t)] \end{bmatrix} \triangleq \mathbf{f}_A[\mathbf{x}(t), \mathbf{p}(t), \mathbf{u}(t), \mathbf{w}(t)] \\ \mathbf{z}(t) &= \mathbf{H}_A[\mathbf{p}(t)] \begin{bmatrix} \mathbf{x}(t) \\ \mathbf{p}(t) \end{bmatrix} + \mathbf{n}(t) = \begin{bmatrix} \mathbf{H}[\mathbf{p}(t)] & \mathbf{0} \end{bmatrix} \begin{bmatrix} \mathbf{x}(t) \\ \mathbf{p}(t) \end{bmatrix} + \mathbf{n}(t) + \mathbf{b}[\mathbf{p}(t)]\end{aligned}$$

$\mathbf{H}[\mathbf{p}(t)]$: Unknown scale factors and coupling terms (TBD)

$\mathbf{b}[\mathbf{p}(t)]$: Unknown bias error (TBD)

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Parameter Vector Must Have a Dynamic Model

Unknown constant: $\mathbf{p}(t) = \text{constant}$

$$\dot{\mathbf{p}}(t) = \mathbf{0}; \quad \mathbf{p}(0) = \mathbf{p}_o; \quad \mathbf{P}_p(0) = \mathbf{P}_{p_o}$$

Random $\mathbf{p}(t)$ (integrated white noise)

$$\begin{aligned} \dot{\mathbf{p}}(t) &= \mathbf{w}_p(t); \quad \mathbf{p}(0) = \mathbf{0}; \quad \mathbf{P}_p(0) = \mathbf{P}_{p_o} \\ E[\mathbf{w}_p(t)] &= \mathbf{0}; \quad E[\mathbf{w}_p(t)\mathbf{w}_p^T(\tau)] = \mathbf{Q}_p\delta(t-\tau) \end{aligned}$$

Linear dynamic system (Markov process)

$$\dot{\mathbf{p}}(t) = \mathbf{A}\mathbf{p}(t) + \mathbf{B}\mathbf{w}_p(t) \triangleq \mathbf{f}_p[\mathbf{p}(t), \mathbf{w}_p(t)]; \quad \mathbf{w}_p(t) \sim N(\mathbf{0}, \mathbf{Q}_p)$$

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Dynamic Models for the Parameter Vector

Doubly integrated white noise

$$\dot{\mathbf{p}}_M(t) = \begin{bmatrix} \dot{\mathbf{p}}(t) \\ \dot{\mathbf{p}}_D(t) \end{bmatrix} = \begin{bmatrix} \mathbf{0} & \mathbf{I} \\ \mathbf{0} & \mathbf{0} \end{bmatrix} \begin{bmatrix} \mathbf{p}(t) \\ \mathbf{p}_D(t) \end{bmatrix} + \begin{bmatrix} \mathbf{0} \\ \mathbf{w}_p(t) \end{bmatrix}$$

$$\begin{bmatrix} \mathbf{p}(t) \\ \mathbf{p}_D(t) \end{bmatrix} = \begin{bmatrix} \text{Parameter vector} \\ \text{Parameter rate of change} \end{bmatrix}$$

Triply integrated white noise

$$\dot{\mathbf{p}}_M(t) = \begin{bmatrix} \dot{\mathbf{p}}(t) \\ \dot{\mathbf{p}}_D(t) \\ \dot{\mathbf{p}}_A(t) \end{bmatrix} = \begin{bmatrix} \mathbf{0} & \mathbf{I} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{I} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} \end{bmatrix} \begin{bmatrix} \mathbf{p}(t) \\ \mathbf{p}_D(t) \\ \mathbf{p}_A(t) \end{bmatrix} + \begin{bmatrix} \mathbf{0} \\ \mathbf{0} \\ \mathbf{w}_p(t) \end{bmatrix}$$

$$\begin{bmatrix} \mathbf{p}(t) \\ \mathbf{p}_D(t) \\ \mathbf{p}_A(t) \end{bmatrix} = \begin{bmatrix} \text{Parameter vector} \\ \text{Parameter rate of change} \\ \text{Parameter acceleration} \end{bmatrix}$$

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Weathervane Example (4.7-1)

2nd - order system

Constant parameter, $\omega_n^2 = a \equiv 4$,

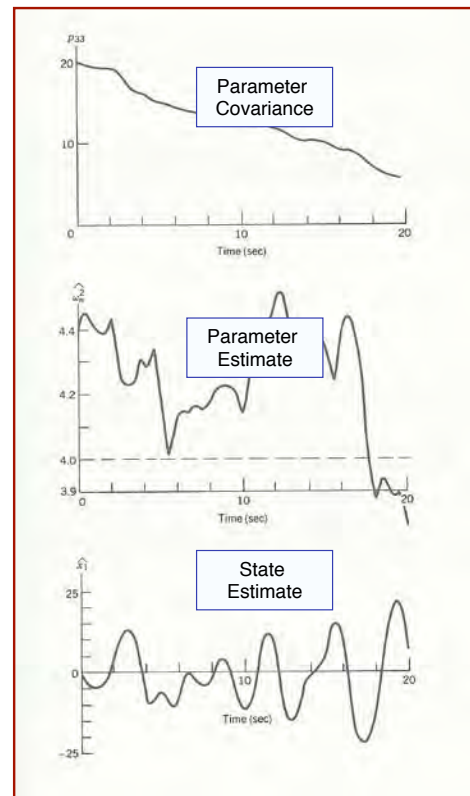
Assumed to be 4.4

$b = 0.4$ (known)

$$\begin{bmatrix} \dot{x}_1 \\ \dot{x}_2 \\ \dot{a} \end{bmatrix} = \begin{bmatrix} 0 & 1 & 0 \\ a & b & 0 \\ 0 & 0 & 0 \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \\ a \end{bmatrix} + \begin{bmatrix} 0 \\ -a \\ 0 \end{bmatrix} w$$

$$Q = 1000; \quad \mathbf{R} = \begin{bmatrix} 10 & 0 \\ 0 & 10 \end{bmatrix}; \quad \mathbf{P}_p(0) = 20$$

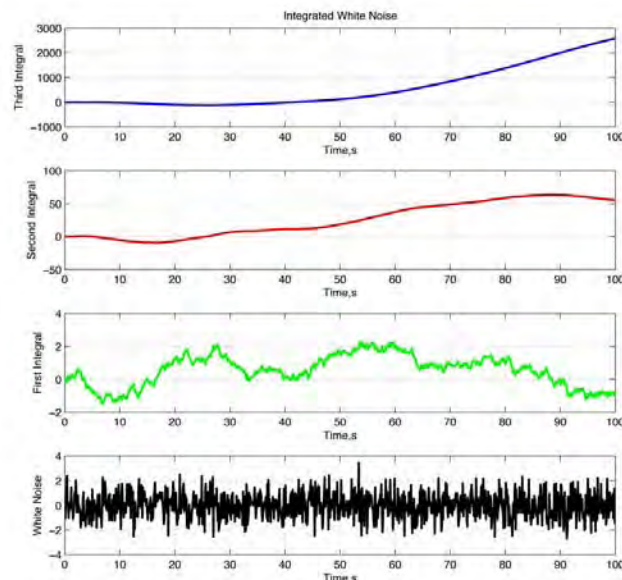
Additional details in text



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Integrated White Noise Models of a Parameter

- **Third integral** models slowly varying, smooth parameter
- **Second integral** is smoother but still has fast changes
- **First integral** of white noise has abrupt jumps, valleys, and peaks
- **White noise**



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Hybrid Filter for Parameter Estimation

Extrapolation of Augmented State

$$\hat{\mathbf{x}}_A[t_k(-)] = \mathbf{x}_A[t_{k-1}(+)] + \int_{t_{k-1}}^{t_k} \mathbf{f}_A[\hat{\mathbf{x}}_A(\tau), \mathbf{u}(\tau)] d\tau$$

Covariance Extrapolation

$$\mathbf{P}_A[t_k(-)] = \mathbf{P}_A[t_{k-1}(+)] + \int_{t_{k-1}}^{t_k} [\mathbf{F}_A(\tau)\mathbf{P}_A(\tau) + \mathbf{P}_A(\tau)\mathbf{F}_A^T(\tau) + \mathbf{L}_A(\tau)\mathbf{Q}'_C(\tau)\mathbf{L}_A^T(\tau)] d\tau$$

Filter Gain Calculation

$$\mathbf{K}(t_k) = \mathbf{P}_A[t_k(-)]\mathbf{H}_A^T(t_k) [\mathbf{H}_A(t_k)\mathbf{P}_A[t_k(-)]\mathbf{H}_A^T(t_k) + \mathbf{R}(t_k)]^{-1}$$

$(\mathbf{F}_A, \mathbf{H}_A)$ must be locally observable except at isolated points

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Hybrid Filter for Parameter Estimation

State Update

$$\hat{\mathbf{x}}_A[t_k(+)] = \hat{\mathbf{x}}_A[t_k(-)] + \mathbf{K}(t_k) \langle \mathbf{z}(t_k) - \mathbf{h}\{\hat{\mathbf{x}}[t_k(-)]\} \rangle$$

Covariance Update

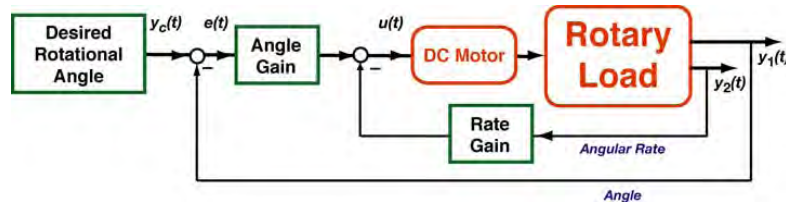
$$\mathbf{P}_A[t_k(+)] = [\mathbf{I}_n - \mathbf{K}(t_k)\mathbf{H}_A(t_k)]\mathbf{P}_A[t_k(-)]$$

$(\mathbf{F}_A, \mathbf{H}_A)$ must be locally observable except at isolated points

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Example: Estimation of Variable Rotary Load for a Robot Arm Elbow



Closed-loop dynamic equation

$$\begin{bmatrix} \dot{x}_1(t) \\ \dot{x}_2(t) \end{bmatrix} = \begin{bmatrix} 0 & 1 \\ -c_1 / J(t) & -c_2 / J(t) \end{bmatrix} \begin{bmatrix} x_1(t) \\ x_2(t) \end{bmatrix} + \begin{bmatrix} 0 \\ c_1 / J(t) \end{bmatrix} y_c$$

Parameter, $p(t) \triangleq$ Rotary load inertia, $J(t)$; c_1, c_2 : Control gains (given)

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Augmented State and Measurement for Unknown Inertia Modeled as Doubly Integrated Parameter

$$\mathbf{x}_A(t) = \begin{bmatrix} x_1(t) \\ x_2(t) \\ J(t) \\ \dot{J}(t) \end{bmatrix} \triangleq \begin{bmatrix} x_1(t) \\ x_2(t) \\ p(t) \\ p_D(t) \end{bmatrix} \begin{matrix} \text{Angle} \\ \text{Angular Rate} \\ \text{Rotary Load Inertia} \\ \text{Inertia Rate} \end{matrix}$$

$$\mathbf{z}(t) = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \end{bmatrix} \begin{bmatrix} x_1(t) \\ x_2(t) \\ p(t) \\ p_D(t) \end{bmatrix} + \begin{bmatrix} n_1(t) \\ n_2(t) \end{bmatrix}$$

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Nonlinear Dynamic Equation with Unknown Inertia Modeled as Doubly Integrated White Noise



$$\dot{\mathbf{x}}_A(t) = \begin{bmatrix} \dot{x}_1(t) \\ \dot{x}_2(t) \\ \dot{p}(t) \\ \dot{p}_D(t) \end{bmatrix} = \begin{bmatrix} 0 & 1 & 1 & 0 \\ -c_1 / p(t) & -c_2 / p(t) & 0 & 0 \\ 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 0 \end{bmatrix} \begin{bmatrix} x_1(t) \\ x_2(t) \\ p(t) \\ p_D(t) \end{bmatrix} + \begin{bmatrix} 0 \\ c_1 / p(t) \\ 0 \\ 0 \end{bmatrix} y_c + \begin{bmatrix} 0 \\ w_2(t) \\ 0 \\ w_{p_D}(t) \end{bmatrix} \triangleq \mathbf{f}_A[\bullet]$$

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Stability and Measurement Matrices

$$\mathbf{F}_A = \begin{bmatrix} 0 & 1 & 1 & 0 \\ -c_1 / \hat{p}(t) & -c_2 / \hat{p}(t) & \left[\frac{c_1}{2\hat{p}^2(t)} \hat{x}_1(t) + \frac{c_2}{2\hat{p}^2(t)} \hat{x}_2(t) \right] & 0 \\ 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 0 \end{bmatrix}$$

$$\mathbf{H}_A = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \end{bmatrix}$$

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Means and Covariances

$$\hat{\mathbf{x}}_A(0) = E \begin{bmatrix} x_1(0) \\ 0 \\ p(0) \\ 0 \end{bmatrix} \triangleq \begin{bmatrix} \hat{x}_1(0) \\ 0 \\ \hat{p}(0) \\ 0 \end{bmatrix}$$

$$\mathbf{P}_A(0) = E \left\{ [\mathbf{x}_A(0) - \hat{\mathbf{x}}_A(0)] [\mathbf{x}_A(0) - \hat{\mathbf{x}}_A(0)]^T \right\}$$

$$\begin{aligned} \hat{\mathbf{w}}(t) &= \mathbf{0}; \quad \mathbf{Q}(t) = E[\mathbf{w}(t)\mathbf{w}^T(\tau)] \\ \hat{\mathbf{n}}(t) &= \mathbf{0}; \quad \mathbf{R}(t) = E[\mathbf{n}(t)\mathbf{n}^T(\tau)] \end{aligned}$$

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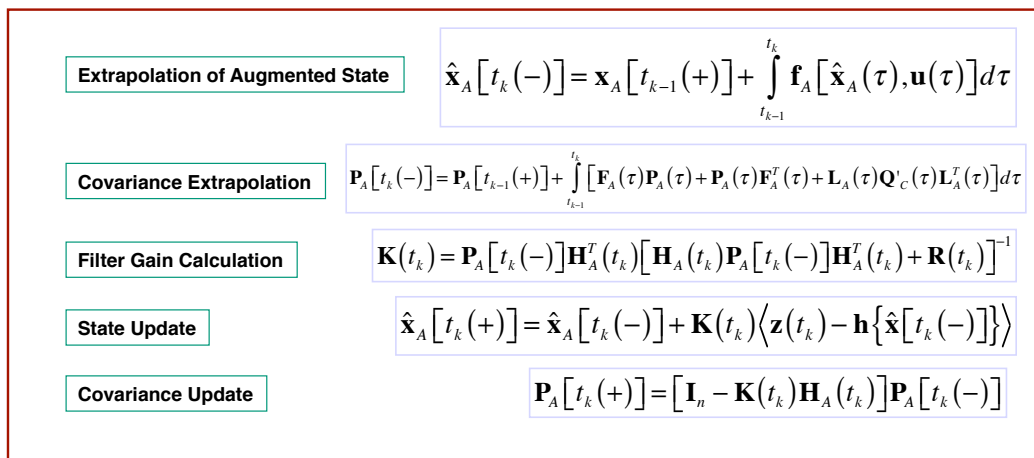
Hybrid Filter for Robot Elbow State and Load Estimation

Convergence is problem-dependent

Qualitative observability of parameter

Actual and assumed uncertainty covariances

Accuracy of dynamic model



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Bias Estimation and Noise-Adaptive Filtering

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Residuals and Optimal Filtering

Linear-optimal filtering has the innovations property

Optimal estimation extracts **all the available information**
from the measurements

Measurement residual should be **zero-mean white noise**

State estimate should be **orthogonal to the error**

Residual and its statistics

$$\mathbf{r}_k(-) = \mathbf{z}_k - \mathbf{H}\hat{\mathbf{x}}_k(-) = \mathbf{H}[\mathbf{x}_k(-) - \hat{\mathbf{x}}_k(-)] + \mathbf{n}_k$$

$$\begin{aligned} E[\mathbf{r}_k(-)] &= \hat{\mathbf{r}}_k(-) = \mathbf{0} \\ E[\mathbf{r}_k(-)\mathbf{r}_k^T(-)] &\triangleq \mathbf{S}_k(-) = \mathbf{H}\mathbf{P}_k(-)\mathbf{H}^T + \mathbf{R}_k \\ E(\hat{\mathbf{x}}_k\mathbf{n}_k^T) &= \mathbf{0} \end{aligned}$$

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Residual Should Be White Noise if State Estimate is Optimal

Test for whiteness using normalized autocovariance function

Sampled (batch process) estimate of the autocovariance function matrix, $\mathbf{C}(k)$

$$\mathbf{C}(k) = \left(\frac{1}{N} \right) \sum_{n=k}^N \mathbf{r}_n \mathbf{r}_{n+k}^T, \quad k \ll N$$

$$\dim[\mathbf{C}(k)] = r \times r$$

Normalize diagonal elements of $\mathbf{C}(k)$ by their zero-lag ($k = 0$) values

$$\rho_{ij}(k) = \frac{c_{ij}(k)}{c_{ii}(0)}$$

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Test for Residual Whiteness

If \mathbf{r}_k is white

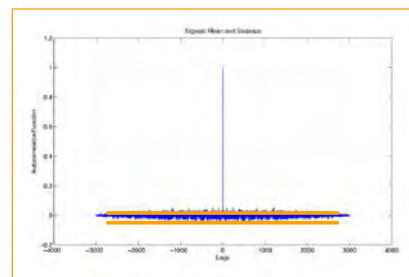
$$\rho_{ij}(k) = \begin{cases} 1, & i = j \text{ and } k = 0 \\ 0, & i \neq j \text{ or } k \neq 0 \end{cases}, \quad N \rightarrow \infty$$

Off-diagonal terms should be negligible

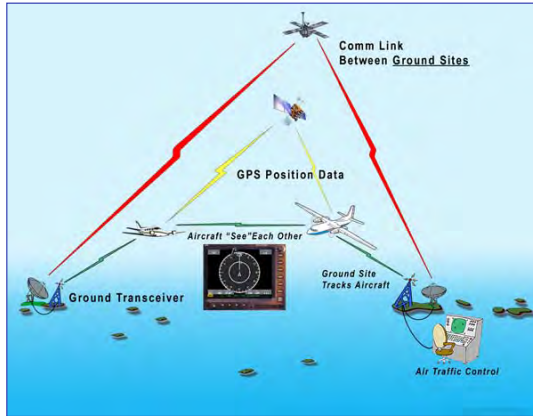
For finite sample, 95% confidence limit based on diagonal elements

$$|\rho_{ii}(k)| \leq \frac{1.96}{\sqrt{N}}, \quad k \neq 0$$

Test is passed if 19 out of 20 non-zero-lag samples are within the limit



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Bias Example: Advanced Dependent Surveillance (ADS-B) System for Air Traffic Control

- Surveillance and tracking radars on ground
- GPS/Inertial/Air data measurements in aircraft
- Satellite/Line-of-sight communications links
- Air traffic control centers (ATCC)
 - Prevent collisions
 - Maintain efficient flow of air traffic

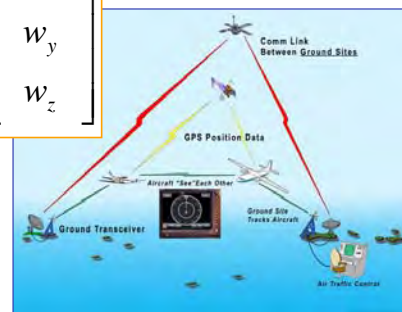
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Dynamics of Individual Aircraft

Surveillance equations of motion

Neglect fast dynamics of aircraft

$$\begin{bmatrix} \dot{x} \\ \dot{y} \\ \dot{z} \\ \dot{v}_x \\ \dot{v}_y \\ \dot{v}_z \end{bmatrix} = \begin{bmatrix} 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 \end{bmatrix} \begin{bmatrix} x \\ y \\ z \\ v_x \\ v_y \\ v_z \end{bmatrix} + \begin{bmatrix} 0 \\ 0 \\ 0 \\ w_x \\ w_y \\ w_z \end{bmatrix}$$



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ATCC Measurements and Communicated Locations of Individual Aircraft

$$\mathbf{z} = \begin{bmatrix} \text{Ground-based radar measurements} \\ \text{Aircraft} \begin{cases} \text{GPS measurements} \\ \text{Inertial navigation measurements} \\ \text{Air data measurements} \end{cases} \end{bmatrix} = (\mathbf{H}\mathbf{x} + \mathbf{n}) + \mathbf{b}$$

$$\mathbf{b} = \text{Biases} = \begin{bmatrix} \text{Radar Range Bias} \\ \text{Radar Azimuth Bias} \\ \text{Aircraft Altimeter Bias} \\ \text{ADS-B Longitude Bias} \\ \text{ADS-B Latitude Bias} \\ \text{ADS-B Altitude Bias} \\ \text{ADS-B North Velocity Bias} \\ \text{ADS-B East Velocity Bias} \end{bmatrix}$$

Bias is a quasi-constant error that is not related to random noise
Least-squares estimator does not reduce bias error

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Measurement Bias and Covariance Estimation

Batch processing estimates

$$\bar{\mathbf{r}}(-) \triangleq \hat{\mathbf{b}}(-) = \frac{1}{N} \sum_{i=1}^N \mathbf{r}_i(-) = \frac{1}{N} \sum_{i=1}^N [\mathbf{z}_i - \mathbf{H}\hat{\mathbf{x}}_i(-)] \quad \text{Bias Estimate}$$

$$\hat{\mathbf{S}} = \frac{1}{N-1} \sum_{i=1}^N (\mathbf{r}_i - \bar{\mathbf{r}})(\mathbf{r}_i - \bar{\mathbf{r}})^T \quad \text{Sample Covariance Matrix}$$

$$\hat{\mathbf{R}} = \hat{\mathbf{S}} - \frac{N-1}{N} \mathbf{H}\mathbf{P}_k(-)\mathbf{H}^T \quad \text{Measurement Noise Estimate}$$

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Running Estimate of Measurement Bias and Error Covariance

$$\begin{aligned}\hat{\mathbf{b}}_k &= \hat{\mathbf{b}}_{k-1} + k_{bias} \left\{ [\mathbf{z}_k - \mathbf{H}\hat{\mathbf{x}}_k(-)] - \hat{\mathbf{b}}_{k-1} \right\} \\ &= \hat{\mathbf{b}}_{k-1} + k_{bias} [\mathbf{r}_k - \hat{\mathbf{b}}_{k-1}] \\ &= [1 - k_{bias}] \hat{\mathbf{b}}_{k-1} + k_{bias} \mathbf{r}_k\end{aligned}$$

$$k_{bias}, k_{noise} < 1$$

$$\hat{\mathbf{R}}_k = \hat{\mathbf{R}}_{k-1} + k_{noise} \left\{ \left[(\mathbf{r}_k - \hat{\mathbf{b}}_k)(\mathbf{r}_k - \hat{\mathbf{b}}_k)^T - \mathbf{H}\mathbf{P}_k(-)\mathbf{H}^T \right] - \hat{\mathbf{R}}_{k-1} \right\}$$

Options

- Choose add hoc recursive gain
- Use weighted least-squares estimator
- Incorporate in an integrated parameter-adaptive filter

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Disturbance Bias Estimation

System equation

$$\mathbf{x}_{k+1} = \Phi \mathbf{x}_k + \mathbf{w}_k$$

Disturbance residual

$$\mathbf{w}_k = \Phi \mathbf{x}_k - \mathbf{x}_{k+1} \approx \Phi \hat{\mathbf{x}}_k(+) - \hat{\mathbf{x}}_{k+1}(+)$$

Sample mean

$$\bar{\mathbf{w}} \triangleq \hat{\mathbf{w}}(+) = \frac{1}{N} \sum_{i=1}^N [\Phi \hat{\mathbf{x}}_i(+) - \hat{\mathbf{x}}_{i+1}(+)] = \frac{1}{N} \sum_{i=1}^N \mathbf{w}_i$$

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Disturbance Covariance Estimation

Sample covariance

$$\hat{\mathbf{W}} = \frac{1}{N-1} \sum_{i=1}^N (\mathbf{w}_i - \bar{\mathbf{w}})(\mathbf{w}_i - \bar{\mathbf{w}})^T$$

Disturbance covariance estimate

$$\hat{\mathbf{Q}} = \hat{\mathbf{W}} - \frac{N-1}{N} \Phi \mathbf{P}_k (+) \Phi^T$$

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Running Estimate of Disturbance Bias and Covariance

$$\hat{\mathbf{w}}_k (+) = \hat{\mathbf{w}}_{k-1} (+) + k_{bias} \left\{ \mathbf{w}_k (+) - \hat{\mathbf{w}}_{k-1} (+) \right\}$$

$$\hat{\mathbf{Q}}_k = \hat{\mathbf{Q}}_{k-1} + k_{noise} \left\{ \left[(\mathbf{w}_k - \hat{\mathbf{w}}_k)(\mathbf{w}_k - \hat{\mathbf{w}}_k)^T - \frac{N-1}{N} \Phi \mathbf{P}_k (+) \Phi^T \right] - \hat{\mathbf{Q}}_{k-1} \right\}$$

$$k_{bias}, k_{noise} < 1$$

- Options as before
 - Choose add hoc recursive gain
 - Use weighted least-squares estimator
 - Incorporate in an integrated parameter-adaptive filter

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Noise-and-Bias Adaptive Filter

Use separately estimated means and covariances in
Kalman filter

$$\hat{\mathbf{x}}_k(-) = \Phi_{k-1} \hat{\mathbf{x}}_{k-1}(+) + \Gamma_{k-1} \mathbf{u}_{k-1} + \hat{\mathbf{w}}_{k-1}(+)$$

$$\mathbf{P}_k(-) = \Phi_{k-1} \mathbf{P}_{k-1}(+) \Phi_{k-1}^T + \hat{\mathbf{Q}}_{k-1}$$

$$\mathbf{K}_k = \mathbf{P}_k(-) \mathbf{H}_k^T \left[\mathbf{H}_k \mathbf{P}_k(-) \mathbf{H}_k^T + \hat{\mathbf{R}}_k \right]^{-1}$$

$$\hat{\mathbf{x}}_k(+) = \hat{\mathbf{x}}_k(-) + \mathbf{K}_k \left[\mathbf{z}_k - \mathbf{H}_k \hat{\mathbf{x}}_k(-) + \hat{\mathbf{b}}_k \right]$$

$$\mathbf{P}_k(+) = \left[\mathbf{P}_k^{-1}(-) + \mathbf{H}_k^T \hat{\mathbf{R}}_k^{-1} \mathbf{H}_k \right]^{-1}$$

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Multiple Model Estimation

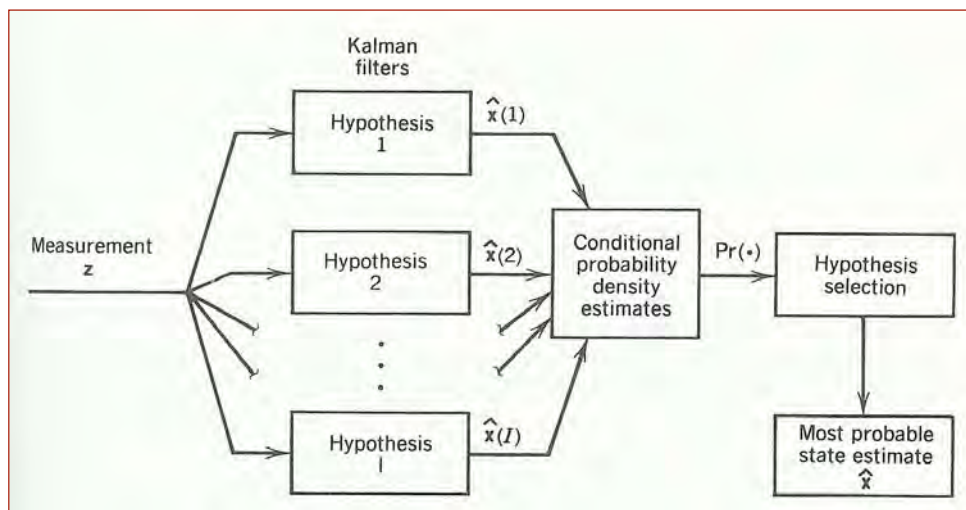
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Multiple Model Estimation

- Bank of Kalman filters, each “tuned” to a different hypothesis
 - Different model parameters or structures
 - Different uncertainty models
 - Different initial conditions
- Best performance determined by a hypothesis test, e.g., Maximum Likelihood
- State estimate chosen accordingly

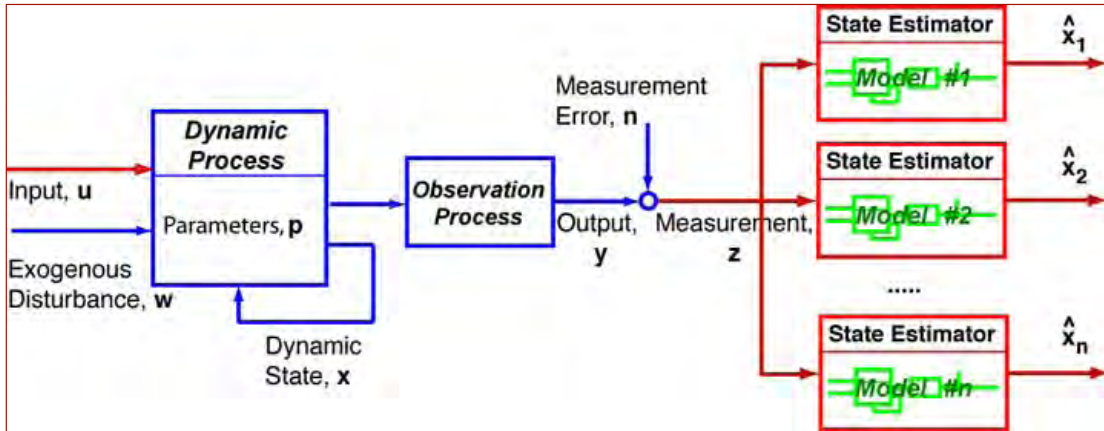
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Hypothesis Testing



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Multiple Model Estimation



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Multiple Model Estimation

- Consider **J systems** distinguished by **J** parameter vectors
- Conditional probability mass function for the **jth** parameter set

$$\Pr(\mathbf{p}_j | \mathbf{z}_k) = \frac{\Pr(\mathbf{z}_k | \mathbf{p}_j) \Pr(\mathbf{p}_j)_{k-1}}{\sum_{i=1}^J [\Pr(\mathbf{z}_k | \mathbf{p}_i) \Pr(\mathbf{p}_i)_{k-1}]}$$

- Probability that the measurement at **k - 1** was obtained is one; therefore,

$$\Pr(\mathbf{p}_j)_{k-1} = \Pr(\mathbf{p}_j | \mathbf{z}_{k-1})$$

- ... and the equation forms the basis for a recursion

$$\Pr(\mathbf{p}_j | \mathbf{z}_k) = \frac{\Pr(\mathbf{z}_k | \mathbf{p}_j)}{\sum_{i=1}^J [\Pr(\mathbf{z}_k | \mathbf{p}_i) \Pr(\mathbf{p}_i)_{k-1}]} \Pr(\mathbf{p}_j | \mathbf{z}_{k-1})$$

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Multiple Model Estimation

Conditional probability density function \mathbf{z}_k must be found
With the true parameter set

$$\begin{aligned}\mathbf{z}_k &= \mathbf{H}\mathbf{x}_k + \mathbf{n}_k \\ \mathbf{x}_k &= \Phi\mathbf{x}_{k-1} + \Gamma\mathbf{u}_{k-1} + \mathbf{w}_{k-1}, \quad \mathbf{x}_0 = \mathbf{x}(0)\end{aligned}$$

If the true state were known

$$\begin{aligned}\text{pr}(\mathbf{z}_k | \mathbf{p}) &= \text{pr}[\mathbf{z}_k | \mathbf{x}_k(\mathbf{p})] \\ &= \frac{1}{(2\pi)^{n/2} |\mathbf{R}_k|^{1/2}} e^{-\frac{1}{2}(\mathbf{z}_k - \mathbf{H}\mathbf{x}_k)^T \mathbf{R}_k^{-1} (\mathbf{z}_k - \mathbf{H}\mathbf{x}_k)} \\ &= \frac{1}{(2\pi)^{n/2} |\mathbf{R}_k|^{1/2}} e^{-\frac{1}{2}\mathbf{n}_k^T \mathbf{R}_k^{-1} \mathbf{n}_k}\end{aligned}$$

However, only an estimate of \mathbf{x}_k is available

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Multiple Model Estimation

- The state is estimated by a Kalman filter for the “true” parameter

$$\text{pr}[\mathbf{z}_k | \hat{\mathbf{x}}_k(\mathbf{p})] = \frac{1}{(2\pi)^{n/2} |\mathbf{S}_k|^{1/2}} e^{-\frac{1}{2}\mathbf{r}_k^T \mathbf{S}_k^{-1} \mathbf{r}_k}$$

with

$$\begin{aligned}\mathbf{r}_k(-) &= \mathbf{z}_k - \mathbf{H}\hat{\mathbf{x}}_k(-) \\ \mathbf{S}_k &= \mathbf{H}\mathbf{P}_k(-)\mathbf{H}^T + \mathbf{R}_k\end{aligned}$$

- The bank of J Kalman filters is formed with each filter assuming that different parameters are the true parameter

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Conditional Probabilities and the Adaptive State Estimate

Conditional probabilities for each hypothesis

$$\Pr(\mathbf{p}_j | \mathbf{z}_k) = \frac{\Pr[\mathbf{z}_k | \hat{\mathbf{x}}_k(\mathbf{p}_j)] \Pr(\mathbf{p}_j | \mathbf{z}_{k-1})}{\sum_{i=1}^J \left\{ \Pr[\mathbf{z}_k | \hat{\mathbf{x}}_k(\mathbf{p}_i)] \Pr(\mathbf{p}_i | \mathbf{z}_{k-1}) \right\}}, \quad j = 1, J$$

State estimate is chosen to be the one for which the conditional probability is highest, or a weighted sum of the state estimates

$$\hat{\mathbf{x}}_k(+) = \sum_{i=1}^J \left\{ \Pr(\mathbf{p}_i | \mathbf{z}_k) \hat{\mathbf{x}}_k[\mathbf{p}_i, (+)] \right\}$$

Parameter vector is chosen accordingly

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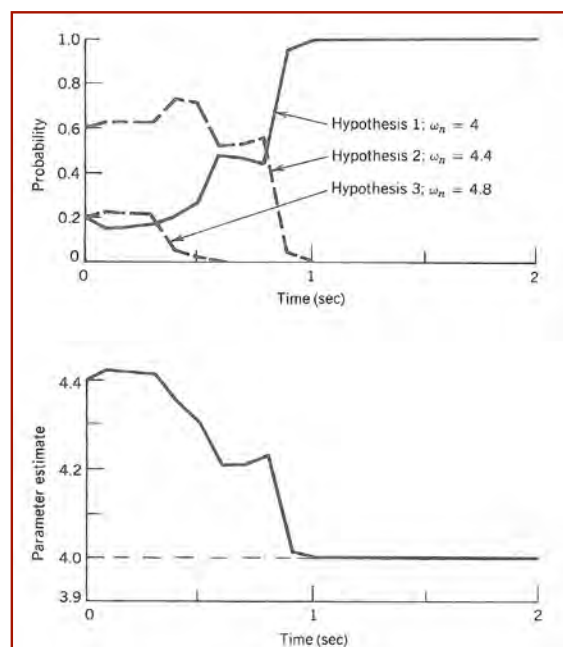
Weathervane Example (4.7-2)

- 2nd-order system with three hypothesized natural frequencies

$$\omega_n^2 = \begin{cases} 4 & \text{[Correct]} \\ 4.4 & \text{[Expected]} \\ 4.8 \end{cases}$$

$$\zeta = 0.1$$

- Algorithm searches, then homes in on correct solution



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***Next Time:
Stochastic Optimal
Control***

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Supplemental Material

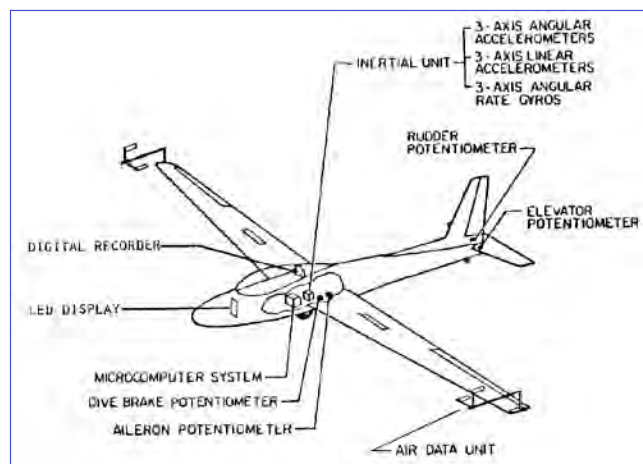
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Estimating Parameters of Nonlinear Systems

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Aerodynamic Coefficients of a Sailplane from Flight Data*



- Princeton University Flight Research Laboratory

* Sri-Jayantha and Stengel, 1988

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Dynamic Equations of the Sailplane

Body-axis velocity and angular rate equations using quaternions

$$\begin{bmatrix} \dot{x}_1 \\ \dot{x}_2 \\ \dot{x}_3 \\ \dot{x}_4 \\ \dot{x}_5 \\ \dot{x}_6 \end{bmatrix} = \begin{bmatrix} \dot{u} \\ \dot{v} \\ \dot{w} \\ \dot{p} \\ \dot{q} \\ \dot{r} \end{bmatrix} = \begin{bmatrix} rv - qw + 2g(e_2e_4 - e_1e_3) + X \\ pw - ru + 2g(e_2e_3 + e_1e_4) + Y \\ qu - pv + g(e_1^2 + e_2^2 - e_3^2 - e_4^2) + Z \\ pqC_1 + qrC_2 + qC_3 + L + NC_4 \\ prC_5 + (r^2 - p^2)C_6 - rC_7 + M \\ pqC_8 + qrC_9 + qC_{10} + LC_{11} + N \end{bmatrix}$$

Propagation of quaternions from angular rates

$$\begin{bmatrix} \dot{x}_7 \\ \dot{x}_8 \\ \dot{x}_9 \\ \dot{x}_{10} \end{bmatrix} = \begin{bmatrix} \dot{e}_1 \\ \dot{e}_2 \\ \dot{e}_3 \\ \dot{e}_4 \end{bmatrix} = (1/2) \begin{bmatrix} 0 & -r & -q & -p \\ r & 0 & -p & q \\ q & p & 0 & -r \\ p & -q & r & 0 \end{bmatrix} \begin{bmatrix} e_1 \\ e_2 \\ e_3 \\ e_4 \end{bmatrix}$$



Schweizer 2-32

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Definitions of Terms

Definitions of terms in dynamic equations

$$\begin{aligned} X &= qSC_X/m, Y = qSC_Y/m, Z = qSC_Z/m \\ L &= qSb(C_l/I_{XX})\{I_{XX}I_{ZZ}/(I_{XX}I_{ZZ} - I_{XZ}^2)\} \\ M &= qSc(C_m/I_{YY}) + \{(Xm\dot{e}_1 - Zm\dot{e}_4)/I_{YY}\} \\ N &= qSb(C_n/I_{ZZ})\{I_{XX}I_{ZZ}/(I_{XX}I_{ZZ} - I_{XZ}^2)\} \\ C_1 &= \{I_{XZ}(I_{ZZ} + I_{XX} - I_{YY})\}/I^2 \\ C_2 &= I_{ZZ}(I_{YY} - I_{ZZ}) - I_{XZ}^2/I^2 \\ C_3 &= 0 \text{ (case with no rotating engine components)} \\ C_4 &= I_{XZ}/I_{XX} \\ C_5 &= (I_{ZZ} - I_{XX})/I_{YY} \\ C_6 &= I_{XZ}/I_{YY} \\ C_7 &= 0 \\ C_8 &= \{I_{XX}(I_{XX} - I_{YY}) + I_{XZ}^2\}/I^2 \\ C_9 &= I_{XZ}(I_{YY} - I_{ZZ} - I_{XX})/I^2 \\ C_{10} &= 0 \\ C_{11} &= I_{XZ}/I_{ZZ} \\ q &= (1/2)\rho V^2, I^2 = \{I_{XX}I_{ZZ} - I_{XZ}^2\} \end{aligned}$$



Schweizer 2-32

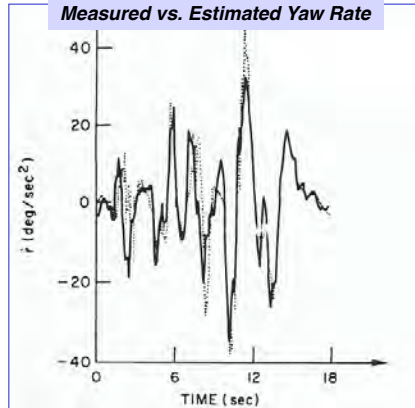
Specific force and moment definitions for parameter identification

$$\begin{aligned} x_{12} = b_{10} = X &= \text{Axial specific force, ft/s}^2 \\ x_{15} = b_{20} = Y &= \text{Side specific force, ft/s}^2 \\ x_{18} = b_{30} = Z &= \text{Normal specific force, ft/s}^2 \\ x_{21} = b_{40} = L &= \text{Roll specific moment, rad/s}^2 \\ x_{24} = b_{50} = M &= \text{Pitch specific moment, rad/s}^2 \\ x_{27} = b_{60} = N &= \text{Yaw specific moment, rad/s}^2 \end{aligned}$$

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Aerodynamic Coefficients of a Sailplane from Flight Data

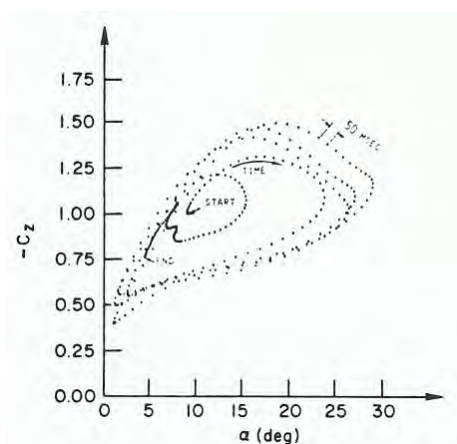
- Estimation-before-modeling technique
 - Estimate the state
 - Extended Kalman Filter (forward pass)
 - Modified Bryson-Frazier Smoother (backward pass)
 - Use multivariate regression to model aerodynamic coefficients



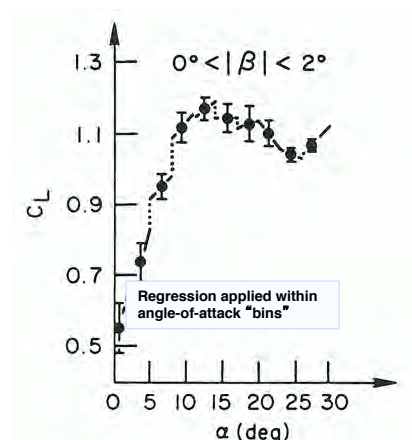
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Aerodynamic Coefficients of a Sailplane from Flight Data

Smoothed Estimate of Normal-Force Coefficient (C_z) History



Estimate of Lift Coefficient (C_L) vs. Angle of Attack



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