Interpreting UniFrac with Absolute Abundance: A Conceptual and Practical Guide

Augustus Pendleton[1]* & Marian L. Schmidt[1]*

[1]Department of Microbiology, Cornell University, 123 Wing Dr, Ithaca, NY 14850, USA

Corresponding Authors: Augustus Pendleton: arp277@cornell.edu; Marian L. Schmidt: marschmi@cornell.edu

Author Contribution Statement: Both authors contributed equally to the manuscript.

Preprint servers: This article was submitted to *bioRxiv* (doi:) under a CC-BY-NC-ND 4.0 International license.

Keywords: Microbial Ecology - Beta Diversity - Absolute Abundance - Bioinformatics - UniFrac

Data Availability: All data and code used to produce the manuscript are available at https://github.com/MarschmiLab/Pendleton_2025_Absolute_Unifrac_Paper, in addition to a reproducible `renv` environment. All packages used for analysis are listed in Table S1.

Microbial ecologists routinely compare communities using -diversity metrics derived from relative abundances. Yet this approach overlooks a critical ecological dimension: microbial load. Across ecosystems, ranging from gut to soil to water, total microbial abundance can vary by orders of magnitude, even among healthy systems. Relative data can misrepresent true shifts in biomass, making taxa appear to change when they have not. High-throughput sequencing produces compositional data, in which each taxon's abundance is constrained by all others [1]. In low-biomass samples, this distortion is especially pronounced, allowing contaminants to appear biologically meaningful when absolute counts are unknown.

Recognition of these limitations has reshaped how we interpret microbial variation. Quantitative profiling studies show that cell abundance, not only composition, can drive major community differences and alter co-occurrence networks [2]. To overcome compositional constraints, researchers increasingly use flow cytometry, qPCR, and genomic spike-ins to quantify microbial load [3, 4]. These tools improve detection of functionally relevant taxa and mitigate the compositional constraints imposed by sequencing [1, 2]. Most studies using absolute data rely on Bray-Curtis dissimilarity, which does not require normalization to proportions [e.g. 3, 5]. Weighted UniFrac remains the dominant phylogenetic -diversity metric, yet its behavior under absolute abundance frameworks remains unclear. Here, we present *Absolute UniFrac*, a direct extension of Weighted UniFrac that incorporates total abundance, and evaluate its impact across simulated and real-world datasets.

The UniFrac distance was first introduced by Lozupone & Knight (2005), and has since become enormously popular as a measure of β-diversity within the field of microbial ecology [6]. A benefit of the UniFrac distance is that is considers phylogenetic information when estimating the distance between two communities. After first generating a phylogenetic tree representing species (or amplicon sequence variants, "ASVs") from all samples, the UniFrac distance computes the fraction of branch-lengths which is *shared* between communities, relative to the total branch length represented in the tree. UniFrac can be both unweighted, in which only the incidence of species is considered, or weighted, wherein a branch's contribution is weighted by the proportional abundance of taxa on that branch [7]. The weighted UniFrac is derived:

$$U^R = \frac{\sum_{i=1}^{n} b_i |p_i^a - p_i^b|}{\sum_{i=1}^{n} b_i (p_i^a + p_i^b)}$$

where we weight the length of each branch, b_i, by the difference in the relative abundance of all species (p_i) descended from that branch in sample a or sample b. Here, we denote this distance as U^R, for "Relative Unifrac". Popular packages which calculate weighted Unifrac-including `diversity-lib` QIIME plug-in and the R packages `phyloseq` and `GUniFrac`-run this normalization by default.

Because U^R is most sensitive to changes in abundant lineages, it can sometimes obscure compositional differences driven by rare to moderately-abundant taxa [8]. To address this weakness, Chen et al. (2012) introduced the generalized UniFrac distance (GU^R), in which the impact of abundant lineages can be mitigated by decreasing the

parameter α:

$$GU^R = \frac{\sum_{i=1}^{n} b_i (p_i^a + p_i^b)^\alpha \left| \frac{p_i^a - p_i^b}{p_i^a + p_i^b} \right|}{\sum_{i=1}^{n} b_i (p_i^a + p_i^b)^\alpha}$$

where α ranges from 0 (close to unweighted UniFrac) up to 1 (identical to U^R, above). However, if one wishes to use absolute abundances, both U^R and GU^R can be derived without normalizing to proportions:

$$U^A = \frac{\sum_{i=1}^{n} b_i |c_i^a - c_i^b|}{\sum_{i=1}^{n} b_i (c_i^a + c_i^b)} \qquad GU^A = \frac{\sum_{i=1}^{n} b_i (c_i^a + c_i^b)^\alpha \left| \frac{c_i^a - c_i^b}{c_i^a + c_i^b} \right|}{\sum_{i=1}^{n} b_i (c_i^a + c_i^b)^\alpha}$$

Where c_i^a and c_i^b stands for the absolute counts of species descended from branch b_i in community a and b, respectively. We refer to these distances as "Absolute Unifrac" and "Generalized Absolute Unifrac" (U^A and GU^A) .

To illustrate how U^A behaves, we constructed a simulated community of four ASVs arranged in a simple, balanced phylogeny (Fig. 1A). By varying the absolute abundance of each ASV (1, 10, or 100), we generated 81 samples and 3,240 pairwise comparisons. For each pair, we computed four dissimilarity metrics: Bray-Curtis with relative abundance (BC^R), Bray-Curtis with absolute abundance (BC^A), Weighted UniFrac with relative abundance (U^R), and Weighted UniFrac with absolute abundance (U^A).

U^A does not consistently yield higher or lower distances but instead varies depending on how abundance and phylogeny intersect (Fig. 1B). In the improbable scenario that all branch lengths are equal, U^A is always less than or equal to BC^A (Fig. S1), but real-world trees rarely behave this way. These comparisons emphasize that incorporating phylogeny and absolute abundance reshapes distance estimates in nontrivial ways.

To better understand how these metrics diverge, we examined individual sample pairs (Fig. 1C). Scenario 1 (gold star) illustrates the classic advantage of UniFrac: ASV_1 and ASV_2 are phylogenetically close, so U^R and U^A discern greater similarity between samples than BC^R and BC^A, which ignore phylogenetic structure. Scenario 2 (red star) highlights a limitation of relative metrics: two samples with identical relative composition but a 100-fold difference in biomass appear identical to BC^R and U^R, but not to their absolute counterparts. In Scenario 3 (green star), incorporating absolute abundance decreases dissimilarity. BC^A and U^A are lower than their relative counterparts because half the community is identical in absolute terms, despite proportional differences. In contrast, Scenario 4 (blue star) shows that U^A can increase dissimilarity relative to BC^A when abundance differences occur on long branches, amplifying phylogenetic dissimilarity.

Across all 3,240 pairwise comparisons, U^A is usually smaller than and strongly correlated with BC^A (Pearsons $r = 0.82$, $p < 0.0001$) than with BC^R ($r = 0.41$) and U^R ($r = 0.55$), reflecting the effect illustrated in Scenario 1. However, exceptions like Scenario 4 show that U^A can also yield larger distances than BC^A when abundance differences occur on long branches. These scenarios demonstrate that U^A integrates ecological realism by

capturing differences in both lineage identity and total biomass, offering a nuanced view of community structure grounded in biological context, while remaining sensitive to long branches.

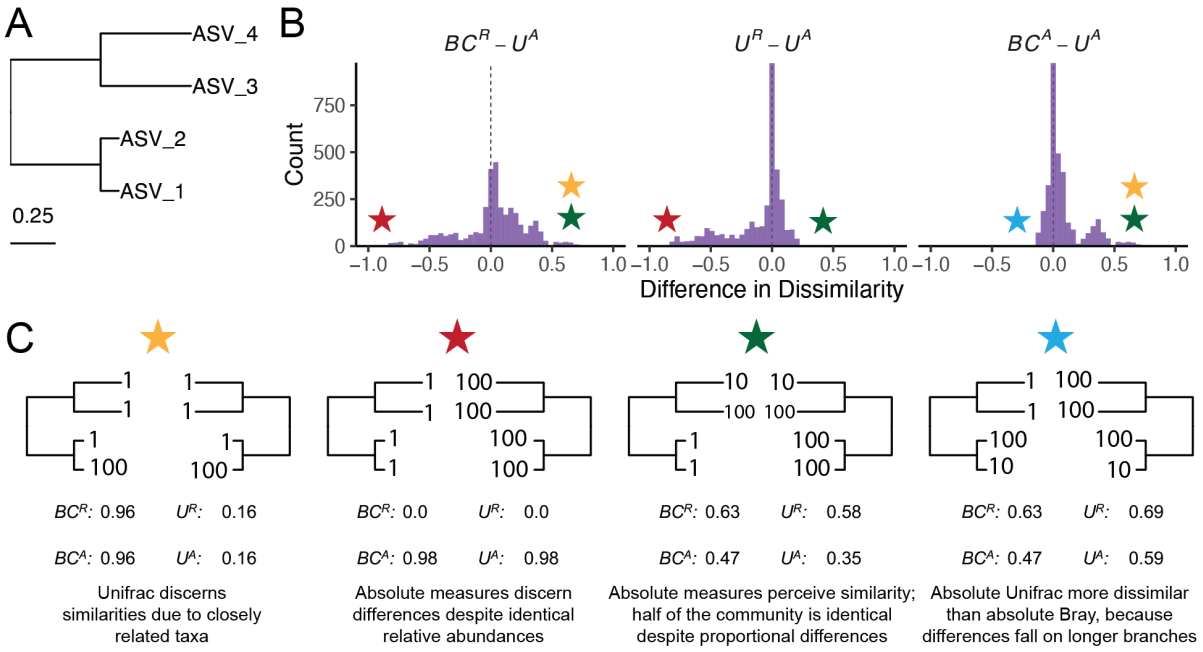

Figure 1. Simulated communities reveal how absolute abundance affects phylogenetic and non-phylogenetic -diversity measures. (A) We constructed a simple four-ASV community with a known phylogeny and generated all permutations of each ASV having an absolute abundance of 1, 10, or 100, resulting in 81 unique communities and 3,240 pairwise comparisons. (B) Distributions of pairwise differences between weighted UniFrac using absolute abundance (U^A) and three other metrics: Bray-Curtis using relative abundance (BC^R), weighted UniFrac using relative abundance (U^R), and Bray-Curtis using absolute abundance (BC^A). (C) Example comparisons illustrating specific scenarios where U^A yields greater or smaller dissimilarity than other metrics. Colored stars indicate where each scenario falls within the distributions shown in panel B. Actual values for each metric are displayed beneath each scenario.

We next evaluated how U^A influences group separation in a real-world dataset. Using a previously published 16S rRNA gene dataset from Lake Ontario, we analyzed 66 samples and >7,000 ASVs. Samples clustered into three groups defined by depth and month, reflecting shifts in both taxonomic composition and microbial load (Fig. S2, [9]). Our goal was to determine whether weighting phylogenetic distances by absolute abundance enhances interpretability and statistical power to distinguish sample groups.

We calculated generalized absolute UniFrac (GU^A) across three levels of α: 0.0 (approximating unweighted UniFrac), 0.5, and 1.0 (equivalent to U^A). As α increased, PCoA ordinations revealed stronger similarity between Shallow May and Shallow September samples, reflecting their higher cell counts compared to the Deep samples (Fig. 2A). Notably, the proportion of variation explained by the first PCoA axis increased substantially with α, going from 18.3% at $\alpha = 0$ up to 76.7% $\alpha = 1$. This trend was also true for U^R across multiple α, but to a much weaker degree (Fig. S3).

To quantify the impact on group differentiation, we performed PERMANOVA across depth-month groupings using GU^R, GU^A, BC^R, and BC^A at varying α (Fig. 2B–C). Across all metrics, incorporating absolute abundance increased both the proportion of explained variance (R^2) and the *pseudoF*-statistic. GU^A achieved a maximum R^2 of 75.8% and a *pseudoF*-statistic 1.56× greater than GU^R, highlighting the ability of GU^A

to detect group differences driven by microbial load.

However, a major caveat emerged: at high α values, GU^A became strongly correlated with total cell count alone (Fig. 2D). Mantel tests showed that absolute distance metrics are much more sensitive to differences in cell counts than their relative counterparts. This is intuitive, and to a degree, intentional; since these metrics aim to detect changes in biomass even when composition remains constant. Yet at $\alpha = 1$, U^A is nearly a proxy for sample biomass. In ordination space, this manifests as Axis 1 being almost perfectly correlated with absolute abundance (Spearman's ρ = -1.0), whereas at $\alpha = 0.0$, the correlation is moderate ($\rho = 0.58$). The structure observed in Fig. 2A at $\alpha = 1$ may largely reflect a horseshoe effect [10].

We urge careful calibration of α based on research goals, modulate this effect using GU^A rather than U^A. Therefore, researchers should consider how much emphasis they want their dissimilarity metric to place on microbial load. In this dataset, we recommend an intermediate α of 0.5, consistent with prior guidance [8], but especially important when using absolute abundance data.

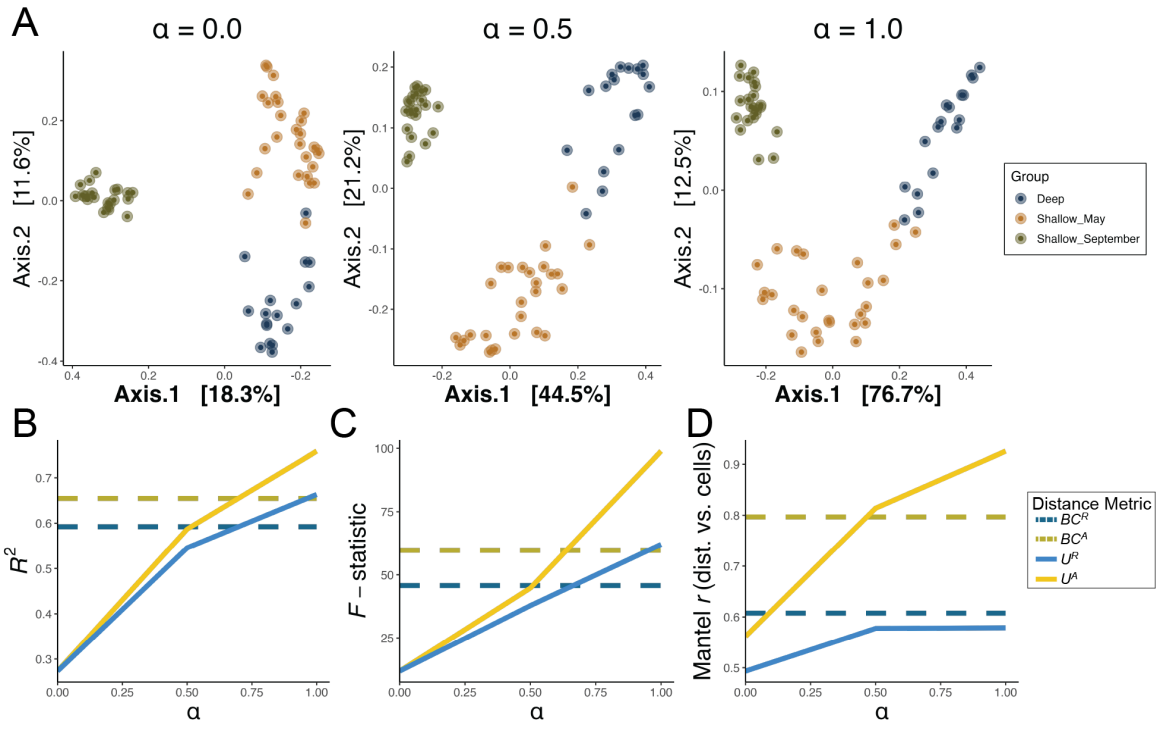

Figure 2. Impact of absolute abundance on separating freshwater microbial communities (A) Samples represent microbial communities from Lake Ontario taken at different times (May and September) and at different depths (Shallow and Deep), as originally described in [9]. We used Principal Coordinates Analysis (PCoA) to ordinate samples based on the GU^A at α values of 0.0, 0.5, and 1.0. The proportion of variance assigned to each axis is provided in brackets. Note that the x-axis is reversed in the first panel, to provide visual symmetry across ordinations. We then used PERMANOVAs to quantify the variance (R^2) (B) and statistical power (F-statistic) (C) our three depth-month groups can explain across different distance metrics (GU^R, GU^A, BC^R and BC^A) and different alpha values. (D) Correlations between different distance metrics and differences in over-all cell abundances as assessed by a Mantel test.

The incorporation of absolute abundance allows microbial ecologists to assess more realistic, ecologically-relevant differences in microbial communities in situations for where microbial load matters. For example, the temporal development of the infant microbiome is captured in both a rise in absolute abundance in addition to compositional changes

[5]; bacteriophage predation in wastewater bioreactors can be understood only when incorporating microbial load [11]; and the negative impact of antibiotics on the abundance of specific taxa in the swine gut was missed using relative abundance approaches [12]. As β-diversity metrics are an essential tool for every microbial ecologist-and UniFrac distances are highly popular within microbiome research specifically-we encourage other researchers to adopt the usage of GU^A in comparing their samples.

Our data show, however, that the interpretation of these metrics can be complex, and researchers should consider (and justify) the relative importance of absolute abundance in their estimates of β-diversity. We also do not address other concerns relevant to estimating β-diversity, including the impacts of sampling effort on richness and whether rarefaction should be applied when calculating GU^A [13]. As ever, interpreting the results of these metrics critically and exploring the sensitivity of your conclusions across multiple metrics is encouraged [14]. Through this work, we hope to facilitate the adoption of GU^A and its interpretation in an approachable manner.

References

1. Gloor GB et al. Microbiome Datasets Are Compositional: And This Is Not Optional. *Front Microbiol* 2017;**8**:2224. https://doi.org/10.3389/fmicb.2017.02224
2. Vandeputte D et al. Quantitative microbiome profiling links gut community variation to microbial load. *Nature* 2017;**551**:507–511. https://doi.org/10.1038/nature24460
3. Props R et al. Absolute quantification of microbial taxon abundances. *ISME J* 2017;**11**:584–587. https://doi.org/10.1038/ismej.2016.117
4. Wang X et al. Current Applications of Absolute Bacterial Quantification in Microbiome Studies and Decision-Making Regarding Different Biological Questions. *Microorganisms* 2021;**9**:1797. https://doi.org/10.3390/microorganisms9091797
5. Rao C et al. Multi-kingdom ecological drivers of microbiota assembly in preterm infants. *Nature* 2021;**591**:633–638. https://doi.org/10.1038/s41586-021-03241-8
6. Lozupone C, Knight R. UniFrac: a New Phylogenetic Method for Comparing Microbial Communities. *Applied and Environmental Microbiology* 2005;**71**:8228–8235. https://doi.org/10.1128/AEM.71.12.8228-8235.2005
7. Lozupone CA et al. Quantitative and Qualitative Diversity Measures Lead to Different Insights into Factors That Structure Microbial Communities. *Applied and Environmental Microbiology* 2007;**73**:1576–1585. https://doi.org/10.1128/AEM.01996-06
8. Chen J et al. Associating microbiome composition with environmental covariates using generalized UniFrac distances. *Bioinformatics* 2012;**28**:2106–2113. https://doi.org/10.1093/bioinformatics/bts342
9. Pendleton A, Wells M, Schmidt ML. Upwelling periodically disturbs the ecological assembly of microbial communities in the Laurentian Great Lakes.
10. Morton JT et al. Uncovering the Horseshoe Effect in Microbial Analyses. *mSystems* 2017;**2**:10.1128/msystems.00166–16. https://doi.org/10.1128/msystems.00166-16
11. Shapiro OH, Kushmaro A, Brenner A. Bacteriophage predation regulates microbial abundance and diversity in a full-scale bioreactor treating industrial wastewater. *The ISME Journal* 2010;**4**:327–336. https://doi.org/10.1038/ismej.2009.118

177 12. Wagner S et al. Absolute abundance calculation enhances the significance of microbiome data in antibiotic treatment studies. *Front Microbiol* 2025;**16**. https://doi.org/10.3389/fmicb.2025.1481197

178 13. Schloss PD. Waste not, want not: revisiting the analysis that called into question the practice of rarefaction. *mSphere* 2023;**9**:e00355–23. https://doi.org/10.1128/msphere.00355-23

179 14. Kers JG, Saccenti E. The Power of Microbiome Studies: Some Considerations on Which Alpha and Beta Metrics to Use and How to Report Results. *Front Microbiol* 2022;**12**. https://doi.org/10.3389/fmicb.2021.796025