

Interpreting UniFrac with Absolute Abundance: A Conceptual and Practical Guide

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Data Availability: All data and code used to produce the manuscript are available at https://github.com/MarschmiLab/Pendleton_2025_Absolute_Unifrac_Paper, in addition to a reproducible **renv** environment. All packages used for analysis are listed in Table S1.

Abstract

Across all systems, microbial ecologists use β -diversity metrics to compare communities, but metrics vary in which ecological dimensions they measure and their ability to separate communities. UniFrac distances (both unweighted and weighted) are highly popular within microbial ecology as they incorporate phylogenetic similarity in their calculation, but they assume normalization to proportions (relative abundance) by default. This overlooks a particularly important dimension of variation across microbial communities: microbial load. As methods for estimating absolute abundance rise in popularity, there will be a need to incorporate absolute abundances into β -diversity for more ecologically-relevant analyses. Here, we present *Absolute Unifrac*, wherein absolute abundances are used rather than proportions, and use simulations and a case study with real freshwater communities to demonstrate conceptual differences between Absolute Unifrac and other popular β -diversity metrics. Using simple scenarios, we identify how Absolute Unifrac can discern both greater and lower distance between communities than other metrics, dependent on differences in both microbial load alongside phylogenetic similarity. Using real data, we demonstrate that Absolute Unifrac can improve statistical power in separating microbial communities. However, we identify an important caveat to Absolute Unifrac: it can be highly sensitive to differences in microbial load. As such, we recommend modulating this impact using an absolute derivation of Generalized Unifrac with a tunable α parameter.

37 Main Text

38 Microbial ecologists routinely compare communities using β -diversity metrics derived
 39 from relative abundances. Yet this approach overlooks a critical ecological dimension:
 40 microbial load. High-throughput sequencing produces compositional data, in which each
 41 taxon’s abundance is constrained by all others [1]. However, quantitative profiling studies
 42 show that cell abundance, not only composition, can drive major community differences
 43 [2]. In low-biomass samples, relying on relative abundance can allow contaminants to
 44 appear biologically meaningful despite absolute counts too low for concern [3].

45 To overcome compositional constraints, researchers increasingly use flow cytometry,
 46 qPCR, and genomic spike-ins to quantify microbial load [4, 5]. These tools improve detec-
 47 tion of functionally relevant taxa and mitigate the compositional constraints imposed by
 48 sequencing [1, 2]. Most studies using absolute data have used Bray-Curtis dissimilarity,
 49 which does not necessarily expect normalization to proportions [e.g. 4, 6]. But in the
 50 field of microbial ecology, UniFrac distances remain popular when working with relative
 51 abundance data. Here, we present *Absolute UniFrac*, a direct extension of Weighted
 52 UniFrac that incorporates total abundance, and evaluate its impact across simulated and
 53 real-world datasets.

54 The UniFrac distance was first introduced by Lozupone & Knight (2005), and has
 55 since become enormously popular as a measure of β -diversity within the field of microbial
 56 ecology [7]. A benefit of the UniFrac distance is that it considers phylogenetic informa-
 57 tion when estimating the distance between two communities. After first generating a
 58 phylogenetic tree representing species (or amplicon sequence variants, “ASVs”) from all
 59 samples, the UniFrac distance computes the fraction of branch-lengths which is *shared*
 60 between communities, relative to the total branch length represented in the tree. UniFrac
 61 can be both unweighted, in which only the incidence of species is considered, or weighted,
 62 wherein a branch’s contribution is weighted by the proportional abundance of taxa on
 63 that branch [8]. The weighted UniFrac is derived:

$$U^R = \frac{\sum_{i=1}^n b_i |p_i^a - p_i^b|}{\sum_{i=1}^n b_i (p_i^a + p_i^b)}$$

64 where we weight the length of each branch, b_i , by the difference in the relative abun-
 65 dance of all species (p_i) descended from that branch in sample a or sample b . Here,
 66 we denote this distance as U^R , for “Relative Unifrac”. Popular packages which calculate
 67 weighted Unifrac-including **diversity-lib** QIIME plug-in and the R packages **phyloseq**
 68 and **GUniFrac-run** this normalization by default.

69 Because U^R is most sensitive to changes in abundant lineages, it can sometimes
 70 obscure compositional differences driven by rare to moderately-abundant taxa [9]. To
 71 address this weakness, Chen et al. (2012) introduced the generalized UniFrac distance
 72 (GU^R), in which the impact of abundant lineages can be mitigated by decreasing the
 73 parameter α :

$$GU^R = \frac{\sum_{i=1}^n b_i (p_i^a + p_i^b)^\alpha \left| \frac{p_i^a - p_i^b}{p_i^a + p_i^b} \right|}{\sum_{i=1}^n b_i (p_i^a + p_i^b)^\alpha}$$

where α ranges from 0 (close to unweighted UniFrac) up to 1 (identical to U^R , above). However, if one wishes to use absolute abundances, both U^R and GU^R can be derived without normalizing to proportions:

$$U^A = \frac{\sum_{i=1}^n b_i |c_i^a - c_i^b|}{\sum_{i=1}^n b_i (c_i^a + c_i^b)} \quad GU^A = \frac{\sum_{i=1}^n b_i (c_i^a + c_i^b)^\alpha \left| \frac{c_i^a - c_i^b}{c_i^a + c_i^b} \right|}{\sum_{i=1}^n b_i (c_i^a + c_i^b)^\alpha}$$

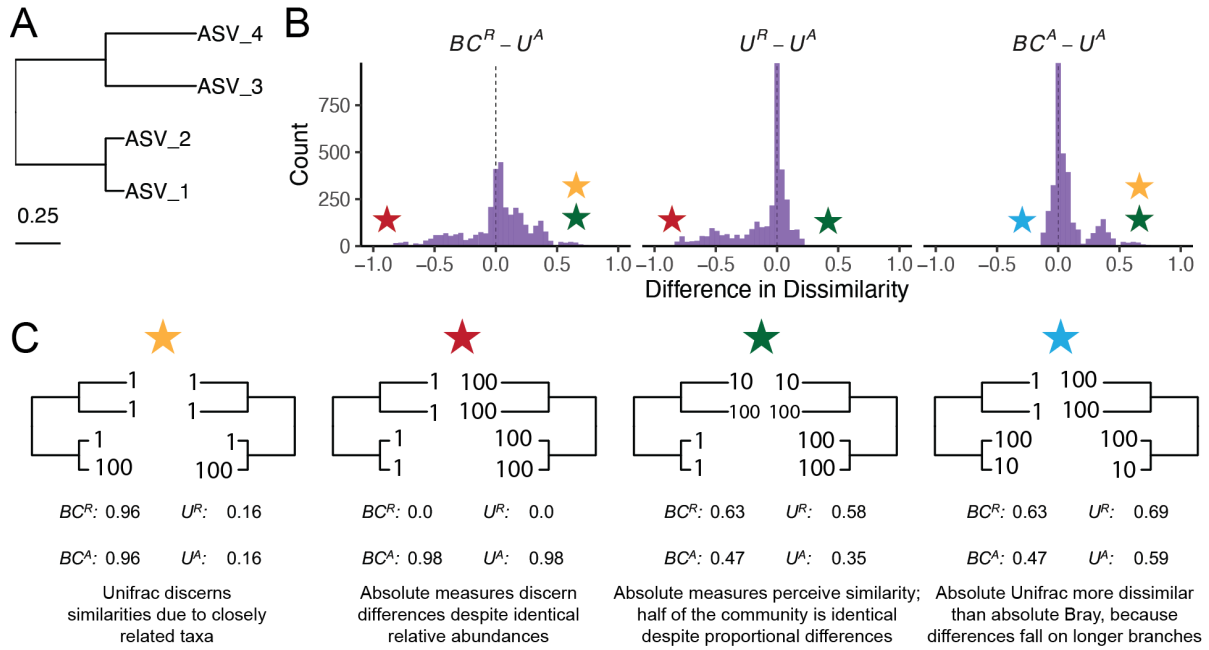
Where c_i^a and c_i^b stands for the absolute counts of species descended from branch b_i in community a and b , respectively. We refer to these distances as “Absolute Unifrac” and “Generalized Absolute Unifrac” (U^A and GU^A).

To illustrate how U^A behaves, we constructed a simulated community of four ASVs arranged in a simple phylogeny (Fig. 1A). By varying the absolute abundance of each ASV (1, 10, or 100), we generated 81 samples and 3,240 pairwise comparisons. For each pair, we computed four dissimilarity metrics: Bray-Curtis with relative abundance (BC^R), Bray-Curtis with absolute abundance (BC^A), Weighted UniFrac with relative abundance (U^R), and Weighted UniFrac with absolute abundance (U^A).

U^A does not consistently yield higher or lower distances but instead varies depending on how abundance and phylogeny intersect (Fig. 1B). In the improbable scenario that all branch lengths are equal, U^A is always less than or equal to BC^A (Fig. S1). These comparisons emphasize that incorporating phylogeny and absolute abundance reshapes distance estimates in nontrivial ways.

To better understand how these metrics diverge, we examined individual sample pairs (Fig. 1C). Scenario 1 (gold star) illustrates the classic advantage of UniFrac: ASV_1 and ASV_2 are phylogenetically close, so U^R and U^A discern greater similarity between samples than BC^R and BC^A , which ignore phylogenetic structure. Scenario 2 (red star) highlights a limitation of relative metrics: two samples with identical relative composition but a 100-fold difference in biomass appear identical to BC^R and U^R , but not to their absolute counterparts. In Scenario 3 (green star), incorporating absolute abundance decreases dissimilarity. BC^A and U^A are lower than their relative counterparts because half the community is identical in absolute terms, despite proportional differences. In contrast, Scenario 4 (blue star) shows that U^A can increase dissimilarity relative to BC^A when abundance differences occur on long branches, amplifying phylogenetic dissimilarity.

Across all 3,240 pairwise comparisons, U^A is usually smaller than and strongly correlated with BC^A (Pearson's $r = 0.82$, $p < 0.0001$) than with BC^R ($r = 0.41$) and U^R ($r = 0.55$), reflecting the effect illustrated in Scenario 1. However, exceptions like Scenario 4 show that U^A can also yield larger distances than BC^A when abundance differences occur on long branches. These scenarios demonstrate that U^A integrates ecological realism by capturing differences in both lineage identity and total biomass.



108

109 *Figure 1. Simulated communities reveal how absolute abundance affects phylogenetic and non-phylogenetic*
 110 *-diversity measures.* (A) We constructed a simple four-ASV community with a known phylogeny and
 111 generated all permutations of each ASV having an absolute abundance of 1, 10, or 100, resulting in 81
 112 unique communities and 3,240 pairwise comparisons. (B) Distributions of pairwise differences between
 113 weighted UniFrac using absolute abundance (U^A) and three other metrics: Bray-Curtis using relative
 114 abundance (BC^R), weighted UniFrac using relative abundance (U^R), and Bray-Curtis using absolute
 115 abundance (BC^A). (C) Example comparisons illustrating specific scenarios where U^A yields greater
 116 or smaller dissimilarity than other metrics. Colored stars indicate where each scenario falls within the
 117 distributions shown in panel B. Actual values for each metric are displayed beneath each scenario.

118 We next evaluated how U^A influences group separation in a real-world dataset. Using a
 119 previously published 16S rRNA gene dataset from Lake Ontario, we analyzed 66 samples
 120 and >7,000 ASVs. Samples clustered into three groups defined by depth and month,
 121 reflecting shifts in both taxonomic composition and microbial load (Fig. S2, [10]). Our
 122 goal was to determine whether weighting phylogenetic distances by absolute abundance
 123 enhances interpretability and statistical power to distinguish sample groups.

124 We calculated generalized absolute UniFrac (GU^A) across three levels of α : 0.0 (approx-
 125 imating unweighted UniFrac), 0.5, and 1.0 (equivalent to U^A). As α increased, PCoA
 126 ordinations revealed stronger similarity between Shallow May and Shallow September
 127 samples, reflecting their higher cell counts compared to the Deep samples (Fig. 2A). No-
 128 tably, the proportion of variation explained by the first PCoA axis increased substantially
 129 with α , going from 18.3% at $\alpha = 0$ up to 76.7% $\alpha = 1$. This trend was also true for U^R
 130 across multiple α , but to a much weaker degree (Fig. S3).

131 To quantify the impact on group differentiation, we performed PERMANOVA across
 132 depth-month groupings using GU^R , GU^A , BC^R , and BC^A at varying α (Fig. 2B–C).
 133 Across all metrics, incorporating absolute abundance increased both the proportion of
 134 explained variance (R^2) and the *pseudo* – F -statistic. GU^A achieved a maximum R^2 of
 135 75.8% and a *pseudo* – F -statistic 1.56 \times greater than GU^R , highlighting the ability of
 136 GU^A to detect group differences driven by microbial load.

137 However, a major caveat emerged: at high α values, GU^A became strongly correlated
 138 with total cell count alone (Fig. 2D). Mantel tests showed that absolute distance metrics

are much more sensitive to differences in cell counts than their relative counterparts. This is intuitive, and to a degree, intentional: these metrics aim to detect changes in biomass even when composition remains constant. Yet at $\alpha = 1$, U^A is nearly a proxy for sample biomass. In ordination space, this manifests as Axis 1 being almost perfectly correlated with absolute abundance (Spearman's $\rho = -1.0$), whereas at $\alpha = 0.0$, the correlation is moderate ($\rho = 0.58$). The structure observed in Fig. 2A at $\alpha = 1$ may largely reflect a horseshoe effect [11], where a strong gradient (here, cell counts) becomes curved in ordination space, potentially distorting ecological interpretation.

We urge careful calibration of α based on research goals, thereby mitigating this effect using GU^A rather than U^A . Researchers should consider how much emphasis they want their dissimilarity metric to place on microbial load. In this dataset, we recommend an intermediate α of 0.5, consistent with prior guidance [9], but especially important when using absolute abundance data.

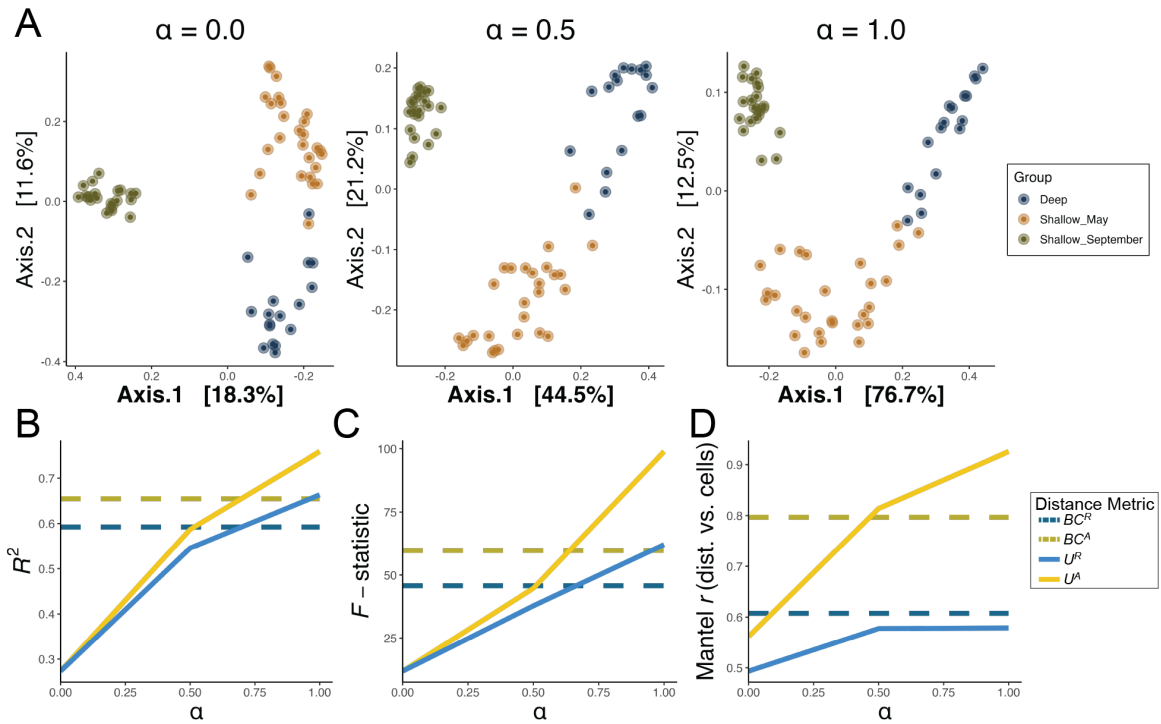


Figure 2. Absolute abundance sharpens ecological signal in freshwater microbial communities but increases sensitivity to biomass. (A) Principal Coordinates Analysis (PCoA) of microbial communities from Lake Ontario, sampled in May and September at two depths (Shallow and Deep) [10]. Ordinations are based on generalized UniFrac with absolute abundance (GU^A) at α values of 0.0, 0.5, and 1.0. Variance explained by each axis is shown in brackets. The x-axis is reversed in the first panel to provide visual symmetry across ordinations. (B-C) Results of PERMANOVA analyses quantifying (B) variance explained (R^2) and (C) statistical power (*pseudo* – F -statistic) across depth-month groups for four distance metrics: GU^R , GU^A , BC^R , and BC^A , each evaluated at multiple α values. (D) Mantel correlations between each distance matrix and differences in cell abundances.

The incorporation of absolute abundance allows microbial ecologists to assess more realistic, ecologically-relevant differences in microbial communities, especially in contexts where microbial load matters. For example, the temporal development of the infant microbiome involves both a rise in absolute abundance and compositional changes [6]; bacteriophage predation in wastewater bioreactors can be understood only when microbial load is considered [12]; and antibiotic-driven declines in specific swine gut taxa were

missed using relative abundance approaches [13]. As β -diversity metrics (and UniFrac specifically) remain central to microbial ecology, we encourage researchers to adopt GU^A when absolute abundance data are available. While demonstrated here with 16S rRNA data, the approach is generalizable to other marker genes or (meta)genomic features, provided absolute abundance estimates are available. In doing so, GU^A offers not only a more grounded picture of lineage differences but also sensitivity to both biomass variation and phylogenetic depth, enabling detection of subtle yet ecologically meaningful shifts.

That said, interpretation of GU^A requires care. When biomass differences dominate, ordinations may largely reflect microbial load rather than lineage turnover, particularly at $\alpha = 1$ and with long phylogenetic branches. In such cases, higher statistical power may come at the cost of biological nuance. We also do not address related concerns, such as how sequencing depth influences richness estimates or whether rarefaction should be applied before calculating GU^A [14]. As with any β -diversity study, researchers should interpret results critically, explore sensitivity across metrics, and justify their choice of α [15]. Our results suggest that an intermediate α value offers a practical compromise that balances sensitivity to biomass with robustness to overdominance by total load, especially when lineage turnover is also of interest. We anticipate that GU^A will become an essential tool for microbiome researchers seeking to incorporate absolute abundance information into ecologically grounded beta diversity comparisons.

References

1. Gloor GB et al. Microbiome Datasets Are Compositional: And This Is Not Optional. *Front Microbiol* 2017;**8**:2224. <https://doi.org/10.3389/fmicb.2017.02224>
2. Vandeputte D et al. Quantitative microbiome profiling links gut community variation to microbial load. *Nature* 2017;**551**:507–511. <https://doi.org/10.1038/nature24460>
3. Barlow JT, Bogatyrev SR, Ismagilov RF. A quantitative sequencing framework for absolute abundance measurements of mucosal and lumenal microbial communities. *Nat Commun* 2020;**11**:2590. <https://doi.org/10.1038/s41467-020-16224-6>
4. Props R et al. Absolute quantification of microbial taxon abundances. *ISME J* 2017;**11**:584–587. <https://doi.org/10.1038/ismej.2016.117>
5. Wang X et al. Current Applications of Absolute Bacterial Quantification in Microbiome Studies and Decision-Making Regarding Different Biological Questions. *Microorganisms* 2021;**9**:1797. <https://doi.org/10.3390/microorganisms9091797>
6. Rao C et al. Multi-kingdom ecological drivers of microbiota assembly in preterm infants. *Nature* 2021;**591**:633–638. <https://doi.org/10.1038/s41586-021-03241-8>
7. Lozupone C, Knight R. UniFrac: a New Phylogenetic Method for Comparing Microbial Communities. *Applied and Environmental Microbiology* 2005;**71**:8228–8235. <https://doi.org/10.1128/AEM.71.12.8228-8235.2005>
8. Lozupone CA et al. Quantitative and Qualitative Diversity Measures Lead to Different Insights into Factors That Structure Microbial Communities. *Applied and Environmental Microbiology* 2007;**73**:1576–1585. <https://doi.org/10.1128/AEM.01996-06>
9. Chen J et al. Associating microbiome composition with environmental covariates using generalized UniFrac distances. *Bioinformatics* 2012;**28**:2106–2113. <https://doi.org/10.1093/bioinformatics/bts342>

- 197 10. Pendleton A, Wells M, Schmidt ML. Upwelling periodically disturbs the ecological
assembly of microbial communities in the Laurentian Great Lakes.
- 198 11. Morton JT et al. Uncovering the Horseshoe Effect in Microbial Analy-
ses. *mSystems* 2017;**2**:10.1128/msystems.00166-16. [https://doi.org/10.1128/
msystems.00166-16](https://doi.org/10.1128/msystems.00166-16)
- 199 12. Shapiro OH, Kushmaro A, Brenner A. Bacteriophage predation regulates micro-
bial abundance and diversity in a full-scale bioreactor treating industrial wastew-
ater. *The ISME Journal* 2010;**4**:327-336. <https://doi.org/10.1038/ismej.2009.118>
- 200 13. Wagner S et al. Absolute abundance calculation enhances the significance of
microbiome data in antibiotic treatment studies. *Front Microbiol* 2025;**16**. [https:
//doi.org/10.3389/fmicb.2025.1481197](https://doi.org/10.3389/fmicb.2025.1481197)
- 201 14. Schloss PD. Waste not, want not: revisiting the analysis that called into question
the practice of rarefaction. *mSphere* 2023;**9**:e00355-23. [https://doi.org/10.1128/
msphere.00355-23](https://doi.org/10.1128/msphere.00355-23)
- 202 15. Kers JG, Saccenti E. The Power of Microbiome Studies: Some Considerations
on Which Alpha and Beta Metrics to Use and How to Report Results. *Front
Microbiol* 2022;**12**. <https://doi.org/10.3389/fmicb.2021.796025>