

# Introduction to Statistical Inference

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# Outline

Warming up

Estimation

Confidence intervals

Statistical Inference and confidence intervals

# Warming up

# Population and Sample

The aim is to **understand** a phenomena

Research questions:

- will the new treatment for hypertension work better than the standard one?
- which are risk factors for IC admission due to COVID-19?
- are genetic factors related to the onset of breast cancer?
- ...

# Population and Sample (cont'd)

The **target population** is the precise definition of the total group of individuals for whom we want to draw conclusions.

- This is achieved by formulating the **inclusion criteria** for the study

Ideally, we collect data from the whole population (i.e., from all subjects), and proceed to analyze them:

- data are actually the realizations from the random variables of interest,
- e.g., blood pressure measurements

## Population and Sample (cont'd)

When all subjects from the population have the same chance to be included in the sample we obtain a **random sample**.

- Such a sample is guaranteed to provide us with valid statements about the target population.

## Population and Sample (cont'd)

To be able to make generalizations from our sample, we want it to be sufficiently **representative** of the target population.

A **representative sample** is a group of subjects from the target population that adequately replicates the population according to whatever characteristic or quality is under study.

A representative sample parallels key variables and characteristics of the larger population.

**Statistical inference** refers to the use of statistics to draw conclusions about an unknown aspect of a population based on a random sample.

# Estimation and Sampling Variability

**Sampling Error:** There will be a difference between the characteristic we measure in the sample and the same characteristic in the population.

**Sampling variability** is the variability in the analysis results caused by the fact that we work with the sample and not the whole population.



## Estimation and Sampling Variability (cont'd)

The **estimand** is the parameter of the target population we wish to estimate from a sample.

Example: the mean of the blood pressure  $\mu$ .

# Estimation and Sampling Variability (cont'd)

Example:

- let's assume that we will obtain a representative sample from this population of size  $n$  (i.e., the number of patients in our sample)
- each subject in the sample has a random variable  $X_i$  describing his/her blood pressure levels
- we could then estimate  $\mu$  using the sample mean

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

The **estimator** is a rule for estimating the parameter in the population using the data we will collect in a sample.

# Estimation and Sampling Variability (cont'd)

Example:

- when we have available specific values  $x_i$  from a **realized sample**, we calculate the realized value of the sample mean

$$\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$$

where  $x_i$  denotes the blood pressure measurements for patient  $i$

The estimate of a particular population characteristic we obtain from a specific sample using an estimator is called the **point estimate**.

## Estimation and Sampling Variability (cont'd)

But what would be the variability of this point estimate in all different samples of size  $n$  from this target population?

**Standard error of the sample mean**

$$se(\bar{X}) = \frac{\sigma}{\sqrt{n}}$$

where  $\sigma$  is the standard deviation of blood pressure values in our population.

# Estimation

# Estimation, Inference, and Sampling Distributions

Key concepts:

- **Population and Sample:** We aim to learn about a population by analysing random samples drawn from it.
- **Randomness:** The values in the sample are random variables due to chance, arising from sampling and biological variation.
- **Probability:** Statements about the population will involve probabilities due to the inherent randomness.

In essence, statistical inference uses sample data and probability to make conclusions about a population.

# Estimation

- Estimation involves assigning likely values to unknown population parameters based on sample data.
- Examples:
  - Sample mean ( $\bar{Y}$ ) as an estimator for population mean ( $\mu$ ).
  - Sample variance ( $s^2$ ) as an estimator for population variance ( $\sigma^2$ ).

## Estimation (cont'd)

For a given parameter, there can be **multiple possible estimators**.

- Example: Estimating population variance ( $\sigma^2$ ).
  - Standard sample variance:

$$s^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2$$

- Alternative estimator:

$$s_*^2 = \frac{1}{n} \sum_{i=1}^n (Y_i - \bar{Y})^2$$

- The alternative uses  $n$  instead of  $n-1$  as the divisor.



## Estimation (cont'd)

If multiple estimators exist for a parameter, the obvious question is: how do we choose the best one?

- Example: Choosing between different estimators for the population variance.

## Estimation (cont'd)

- Estimators like  $\bar{Y}$  and  $s^2$  have their own underlying populations with probability distributions.
- Consider the population of all possible  $\bar{Y}$  values from all possible samples of size  $n$ .

## Estimation (cont'd)

- The mean of the population of  $\bar{Y}$  values is equal to the population mean  $\mu$ :

$$\text{Mean of population of } \bar{Y} = \mu_{\bar{Y}} = \mu$$

in other words,  $E(\bar{Y}) = \mu$ .

- This means  $\bar{Y}$  is an **unbiased** estimator of  $\mu$ .
- Unbiasedness is a desirable property: The estimator's average value matches the true parameter.

## Estimation (cont'd)

- If the mean of the estimator's probability distribution equals the parameter being estimated, the estimator is **unbiased**.
- Example:  $\bar{Y}$  (sample mean) is an unbiased estimator of  $\mu$  (population mean).
- Unbiasedness is a key criterion for choosing among competing estimators.

## Estimation (cont'd)

Example:

- We have two estimators for population variance  $\sigma^2$ :

- $s^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2$

- $s_*^2 = \frac{1}{n} \sum_{i=1}^n (Y_i - \bar{Y})^2$

## Estimation (cont'd)

- $s_*^2$  (dividing by  $n$ ) has a mean of  $\left(\frac{n-1}{n}\right) \sigma^2$ , not  $\sigma^2$ .
- $s_*^2$  is **biased downward**.
- It underestimates  $\sigma^2$ , especially for small  $n$ .

## Estimation (cont'd)

- $s^2$  (dividing by  $n - 1$ ) is designed to be an **unbiased** estimator.
- Dividing by  $n - 1$  corrects the bias.
- Therefore,  $s^2$  is the accepted estimator for  $\sigma^2$  due to its unbiasedness.

## Estimation (cont'd)

- If two estimators are both unbiased, how do we choose the better one?
- Unbiasedness is desirable, but not the only criterion.



## Estimation (cont'd)

- We want estimators to be “close” to the true value.
- This means we want the estimator’s probability distribution to have **small variance**.
- Small variance means the estimator’s values vary little across different samples.

## Estimation (cont'd)

- If two estimators are unbiased, choose the one with the **smaller variance**.
- This ensures the estimator is more consistent and precise.

# Estimators for population mean and variance

It turns out that for normally distributed data  $Y$ , the estimators  $\bar{Y}$  (for  $\mu$ ) and  $s$  (for  $\sigma^2$ ) have this desirable property.

We use  $\bar{Y}$  and  $s$  as estimators for population mean and variance because they have the desirable properties of **unbiasedness and minimum variance**.

# Confidence intervals

## Estimators

- Estimators like  $\bar{Y}$  and  $s^2$  provide “likely” values for population parameters  $\mu$  and  $\sigma^2$ .
- Due to chance, they are rarely exactly equal to the true parameters.
- However, “good” estimators tend to be “close” to the true values.

# Confidence interval (cont'd)

## Interval Estimates

- Instead of reporting a single value, we report an **interval** based on the estimator.
- This interval is likely to contain the true parameter value.
- “Likely” implies that **probability** is involved.

Here, we discuss the notion of such an interval, known as a **confidence interval**.

## Confidence interval for $\mu$

Example: we wish to estimate  $\mu$  by the sample mean  $\bar{Y}$ .

- Assume  $Y \sim N(\mu, \sigma^2)$ , where  $\mu$  and  $\sigma^2$  are unknown.
- We have a random sample  $Y_1, \dots, Y_n$ .
- Our estimator for  $\mu$  is the sample mean,  $\bar{Y}$ .

## Confidence interval for $\mu$ (cont'd)

- To make probability statements about  $\bar{Y}$ , we need to account for the unknown  $\sigma^2$ .
- We must estimate  $\sigma^2$  even if it's not our primary interest.
- We use the statistic:

$$\frac{\bar{Y} - \mu}{s_{\bar{Y}}}$$

where  $s_{\bar{Y}}$  is the estimated standard error of  $\bar{Y}$ .



## Confidence interval for $\mu$ (cont'd)

- We want to quantify the uncertainty of estimating the fixed value  $\mu$  using the random variable  $\bar{Y}$ .
- The randomness arises from the random sample.
- Probability statements reflect the uncertainty of trying to get an understanding of  $\mu$  using  $\bar{Y}$ .

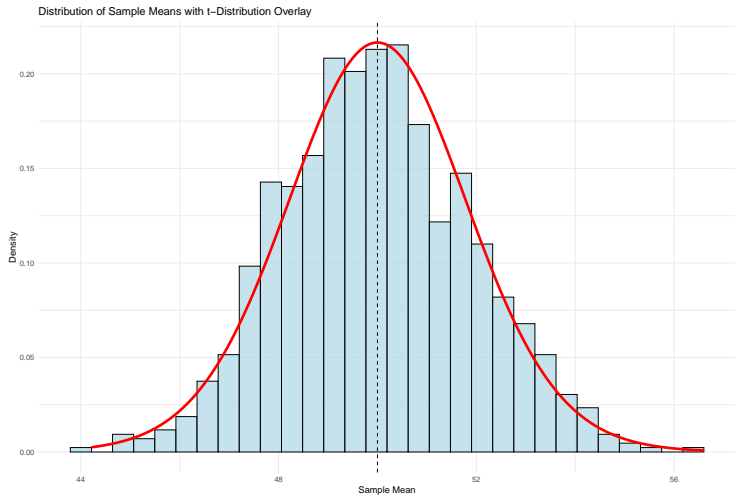
## Confidence interval for $\mu$ (cont'd)

Wait a minute, here's one more detail about probability distributions!

- Let  $t_{n-1, \alpha/2}$  be the point such that the shaded region under the Student's  $t$  density with  $(n - 1)$  degrees of freedom has area  $\alpha/2$ .
- Note that then unshaded region has area  $(1 - \alpha)/2$ .
- By symmetry,  $t_{n-1, \alpha/2}$  is such that each shaded region has area  $\alpha/2$ , with  $1 - \alpha$  in the middle.

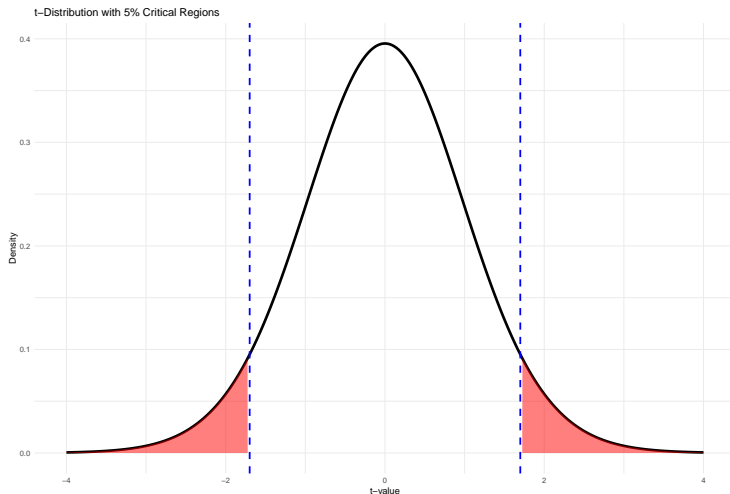
## Confidence interval for $\mu$ (cont'd)

Wait a minute, here's one more detail about probability distributions!



## Confidence interval for $\mu$ (cont'd)

Wait a minute, here's one more detail about probability distributions!



## Confidence interval for $\mu$ (cont'd)

- We aim to create an interval that likely contains the true population mean  $\mu$ .
- The formula is derived from the t-distribution:

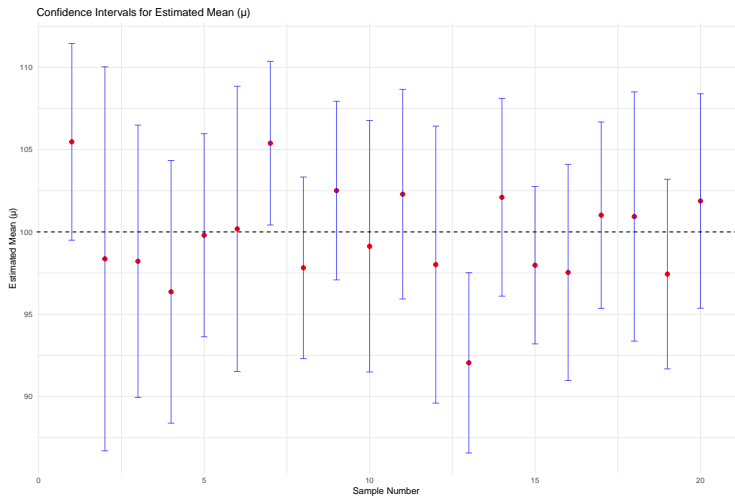
$$P(\bar{Y} - t_{n-1, \alpha/2} \cdot s_{\bar{Y}} \leq \mu \leq \bar{Y} + t_{n-1, \alpha/2} \cdot s_{\bar{Y}}) = 1 - \alpha$$

- $t_{n-1, \alpha/2}$  is the critical t-value,  $s_{\bar{Y}}$  is the standard error of the mean.

## Confidence interval for $\mu$ (cont'd)

- The probability  $(1 - \alpha)$  is **not** about  $\mu$  being in the interval.
- $\mu$  is a **fixed constant**.
- The probability refers to the **random interval**  $(\bar{Y} \pm t_{n-1, \alpha/2} \cdot s_{\bar{Y}})$ .
- It's the probability that this interval will contain the true  $\mu$ .

# Confidence interval for $\mu$ (cont'd)



## Confidence interval for $\mu$ (cont'd)

The interval

$$(\bar{Y} - t_{n-1, \alpha/2} \cdot s_{\bar{Y}}, \bar{Y} + t_{n-1, \alpha/2} \cdot s_{\bar{Y}})$$

is called a  $100 \cdot (1 - \alpha)\%$  **confidence interval** for  $\mu$ .

- For example, if  $\alpha = 0.05$ , then the interval would be called a 95% confidence interval.
- In general the value  $1 - \alpha$  is called the **confidence coefficient**.



## Confidence interval for $\mu$ (cont'd)

- The probability associated with a confidence interval is about the **interval's endpoints**, not about the population mean  $\mu$ .
- It's wrong to say “the probability that  $\mu$  falls in the interval is 95%”.
- $\mu$  is a **fixed but unknown** value.

## Confidence interval for $\mu$ (cont'd)

- The interval is calculated from the sample mean ( $\bar{Y}$ ) and its standard error ( $s_{\bar{Y}}$ ).
- It is a **random interval** because it varies with different samples.
- Different samples yield different  $\bar{Y}$  and  $s_{\bar{Y}}$ , and thus different intervals.

## Confidence interval for $\mu$ (cont'd)

- The probability (e.g., 95%) refers to the likelihood that the **sample** produces an interval that contains the true  $\mu$ .
- It's about the process of repeatedly taking samples and constructing intervals.
- In the long run, 95% of the intervals will contain  $\mu$ .

## Confidence interval for $\mu$ (cont'd)

The confidence coefficient  $(1 - \alpha)$ ...

- represents the probability a sample will produce an interval that covers the true population mean  $\mu$ .
- measures our “confidence” in the sampling procedure and interval construction.

## Confidence interval for $\mu$ (cont'd)

- The width of the confidence interval depends on the confidence coefficient and the size of the sample.
  - For a 90% interval would narrower than a 95% interval. For the 90% interval, we may be a little more stringent with the width as we aren't requiring such a high a level confidence.
  - If we wish greater confidence of 95%, we must widen the interval in order to be “more confident” that it will cover the fixed value  $\mu$ .

## Exercise: Dissolved Oxygen Concentration

- Data: 6 measurements of dissolved oxygen (mg/L).
- Values: 2.62, 2.65, 2.79, 2.83, 2.91, 3.57
- Assume data follows a normal distribution  $N(\mu, \sigma^2)$ .
- Goal: Find a confidence interval for the population mean  $\mu$ .

## Exercise: Dissolved Oxygen Concentration (cont'd)

- Sample size:  $n = 6$
- Sample mean:

$$\bar{Y} = \frac{2.62 + 2.65 + 2.79 + 2.83 + 2.91 + 3.57}{6} = 2.895 \text{ mg/L}$$

- Sample variance:

$$s^2 = \frac{0.075625 + 0.059025 + 0.011025 + 0.004225 + 0.000225 + 0.455625}{5} = 0.12115$$

- Sample standard deviation:  $s = 0.348 \text{ mg/L}$
- Standard error of the mean:  $s_{\bar{Y}} = \frac{0.348}{\sqrt{6}} = 0.142 \text{ mg/L}$

## Exercise: Dissolved Oxygen Concentration

### 95% Confidence Interval

- Using  $t_{5,0.025} = 2.571$  (t-distribution with 5 degrees of freedom).
- Confidence interval:

$$(2.895 \pm 2.571 \times 0.142) = (2.529, 3.261) \text{ mg/L}$$



## Exercise: Dissolved Oxygen Concentration

### 90% Confidence Interval

- Using  $t_{5,0.05} = 2.015$  (t-distribution with 5 degrees of freedom).
- Confidence interval:

$$(2.895 \pm 2.015 \times 0.142) = (2.609, 3.181) \text{ mg/L}$$

## Example in R

How can we calculate it using R?

```
# Data  
oxygen_values <- c(2.62, 2.65, 2.79, 2.83, 2.91, 3.57)
```

## Example in R

How can we calculate it using R?

```
# Sample statistics
n <- length(oxygen_values) # Sample size
mean_Y <- mean(oxygen_values) # Sample mean
s <- sd(oxygen_values) # Sample standard deviation
se_Y <- s / sqrt(n) # Standard error

df <- n - 1 # Degrees of freedom
```

## Example in R

How can we calculate it using R?

```
# 95% Confidence Interval
t_95 <- qt(0.975, df) # t-value for 95% confidence
i95_lower <- mean_Y - t_95 * se_Y
i95_upper <- mean_Y + t_95 * se_Y
c(i95_lower,i95_upper)
```

# Statistical Inference and confidence intervals

# Statistical Inference and confidence intervals

- The confidence interval procedure is a form of **statistical inference**.
- We are now defining this term more formally.

# Statistical Inference and confidence intervals (cont'd)

- Conclusions about the population are based on **sample evidence**.
- This introduces an element of **uncertainty**.
- We aim to assess and summarize this uncertainty.
- We assess the **quality of our sample evidence**.

## Statistical Inference and confidence intervals (cont'd)

- We temper inferences about the population with an indication of our assessment of sample evidence.
- This indication is expressed in terms of **probability**.
- Probability reflects how likely things are in the sample.



# Statistical Inference and confidence intervals (cont'd)

We can use confidence intervals as assessment

- A confidence interval is one way to assess sample evidence.
- It is a statement about the **quality of sample evidence** regarding the population mean.

# Statistical Inference and Sampling Distributions (cont'd)

- When the population is approximately normal, and we want to make inferences about the population mean ( $\mu$ ) using the sample mean ( $\bar{Y}$ ), we use:

$$\frac{\bar{Y} - \mu}{s_{\bar{Y}}} \sim t_{n-1}$$

- This statistic follows a  $t$ -distribution with  $n - 1$  degrees of freedom.
- We use this to make probability statements (confidence intervals) about the sample evidence.

# Statistical Inference and Sampling Distributions (cont'd)

- The probability distribution of a statistic is central to statistical methodology.
- The distribution of all possible values of a statistic across all possible samples is called a **sampling distribution**.
- Examples:
  - Student's t-distribution
  - Chi-squared ( $\chi^2$ ) distribution

# Statistical Inference and Sampling Distributions (cont'd)

- Formal statistical inference is based on comparing statistics to appropriate **sampling distributions**.
- This enables us to make formal probability statements.
- These probability statements characterize the quality of the sample evidence in answering our research question.

# Confidence interval for a difference of population means

- We often want to compare **multiple populations**, not just study one.
- Experiments are designed to compare different treatments or conditions.

# Confidence interval for a difference of population means (cont'd)

## Example: Fertilizer and Yield

- An experiment might compare crop yields with different fertilizer application rates.
- We want to determine if different treatments (fertilizer rates) lead to **truly different responses** (yields).

# Experimental Procedure

- Take two random samples of experimental units.
- Apply treatment 1 to the first sample and treatment 2 to the second sample.
- Analyze the difference in responses between the two treatments.

## Example: Toxic Agent on Rats

- Compare two concentrations of a toxic agent on rat weight loss.
- Randomly assign rats to receive either concentration 1 or 2.
- Variable of interest:  $Y$  = weight loss.



# Two Populations

- Before treatments, rats are from a common population:  $Y \sim N(\mu, \sigma^2)$ .
- After treatments, consider two populations:
  - Population 1: Y's under treatment 1
  - Population 2: Y's under treatment 2
- Samples are from these respective populations.

# Normal Distribution Assumption

- Assume two random variables,  $Y_1$  and  $Y_2$ , for each population.
- Assume they are normally distributed:
  - Population 1:  $Y_1 \sim N(\mu_1, \sigma_1^2)$
  - Population 2:  $Y_2 \sim N(\mu_2, \sigma_2^2)$

# Notation for Two Populations

- We introduce notation for observations from two populations.
- $Y_{ij}$  denotes the  $j$ th observation from the  $i$ th treatment.
- Data Representation:
  - Population 1:  $Y_{11}, Y_{12}, \dots, Y_{1n_1}$  ( $n_1$  observations)
  - Population 2:  $Y_{21}, Y_{22}, \dots, Y_{2n_2}$  ( $n_2$  observations)

# Comparing population means

- Question: Is there a difference in response between the two treatments?
- Formally: Is  $\mu_1$  different from  $\mu_2$ ?
- Consider the difference:  $(\mu_1 - \mu_2)$ 
  - $(\mu_1 - \mu_2) = 0$ : No difference
  - $(\mu_1 - \mu_2) \neq 0$ : Real difference

## Comparing population means (cont'd)

- Estimate the difference  $(\mu_1 - \mu_2)$  using the data from the two samples.

## Comparing population means (cont'd)

What's the distribution of the difference?

- If  $Y_1$  and  $Y_2$  are normally distributed, then:

$$(Y_1 - Y_2) \sim N(\mu_1 - \mu_2, \sigma_D^2)$$

where  $\sigma_D^2 = \sigma_1^2 + \sigma_2^2$ .

- The difference  $(Y_1 - Y_2)$  is also normally distributed.

## Comparing population means (cont'd)

- Define  $\hat{D} = \bar{Y}_1 - \bar{Y}_2$ , where:
  - $\bar{Y}_1 = \frac{1}{n_1} \sum_{j=1}^{n_1} Y_{1j}$
  - $\bar{Y}_2 = \frac{1}{n_2} \sum_{j=1}^{n_2} Y_{2j}$
- $\hat{D}$  is the difference in sample means.

## Comparing population means (cont'd)

- $\hat{D}$  is normally distributed:

$$\hat{D} \sim \mathcal{N}(\mu_1 - \mu_2, \sigma_D^2)$$

where  $\sigma_D^2 = \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}$ .

- The mean of the distribution of  $\hat{D}$  is  $\mu_1 - \mu_2$ .
- The statistic  $\frac{\hat{D} - (\mu_1 - \mu_2)}{\sigma_{\hat{D}}}$  follows a standard normal distribution.



## Comparing population means (cont'd)

- Use  $\hat{D}$  as an estimator for  $\mu_1 - \mu_2$ .
- We want to report a confidence interval to assess the quality of the sample evidence.

# Comparing population means (cont'd)

## Assumptions

- In practice,  $\sigma_1^2$  and  $\sigma_2^2$  are usually unknown.
- We consider the simplest case:
  - $n_1 = n_2 = n$  (equal sample sizes).
  - $\sigma_1^2 = \sigma_2^2 = \sigma^2$  (equal variances).

## Comparing population means (cont'd)

- We assume equal variances:  $\sigma_1^2 = \sigma_2^2 = \sigma^2$ .
- Equal sample sizes:  $n_1 = n_2 = n$ .
- Sample means:
  - $\bar{Y}_1 = \frac{1}{n} \sum_{j=1}^n Y_{1j}$
  - $\bar{Y}_2 = \frac{1}{n} \sum_{j=1}^n Y_{2j}$
- Variance of the difference:  $\sigma_D^2 = 2\left(\frac{\sigma^2}{n}\right)$ .

## Comparing population means (cont'd)

- Since  $\sigma_1^2 = \sigma_2^2$ , we use a pooled estimate:

$$s^2 = \frac{(n-1)s_1^2 + (n-1)s_2^2}{2(n-1)} = \frac{s_1^2 + s_2^2}{2}$$

- $s^2$  is the average of the two sample variances.

## Comparing population means (cont'd)

- Estimator for  $\sigma_{\bar{D}}^2$ :

$$s_{\bar{D}}^2 = 2\left(\frac{s^2}{n}\right)$$

- Test statistic:

$$\frac{\bar{D} - (\mu_1 - \mu_2)}{s_{\bar{D}}} \sim t_{2(n-1)}$$

where  $s_{\bar{D}} = \sqrt{\frac{2}{n}}s$ .

## Confidence interval for $\mu_1 - \mu_2$

- Confidence interval with confidence coefficient  $(1 - \alpha)$ :

$$\bar{D} \pm t_{2(n-1), \alpha/2} s_{\bar{D}}$$

- Probability statement:

$$P(\bar{D} - t_{2(n-1), \alpha/2} s_{\bar{D}} \leq \mu_1 - \mu_2 \leq \bar{D} + t_{2(n-1), \alpha/2} s_{\bar{D}}) = 1 - \alpha$$

## Confidence Interval for $\mu_1 - \mu_2$ (cont'd)

- The confidence interval for the difference of population means ( $\mu_1 - \mu_2$ ) is:

$$(\bar{D} - t_{2(n-1),\alpha/2} s_{\bar{D}}, \bar{D} + t_{2(n-1),\alpha/2} s_{\bar{D}})$$

- Where:
  - $\bar{D} = \bar{Y}_1 - \bar{Y}_2$  (difference of sample means)
  - $t_{2(n-1),\alpha/2}$  is the critical t-value with  $2(n-1)$  degrees of freedom
  - $s_{\bar{D}}$  is the standard error of the difference

## Confidence Interval for $\mu_1 - \mu_2$ (cont'd)

- The interpretation is the same as for a single sample and single population.
- The confidence interval is a statement about the quality of the evidence in the samples from the two populations.
- It assesses the evidence about the fixed population parameter  $(\mu_1 - \mu_2)$ .



## Example: Pig Rations

- Two rations, A and B, are tested on pigs for weight gain (lbs).
- 12 pigs per ration ( $n = 12$ ).
- Data:
  - Ration A: 31, 34, 29, 26, 32, 35, 38, 34, 30, 29, 32, 31
  - Ration B: 26, 24, 28, 29, 30, 29, 32, 26, 31, 29, 32, 28

## Example: Pig Rations (cont'd)

- $\bar{Y}_1 = 31.75$  (mean for Ration A)
- $\bar{Y}_2 = 28.6667$  (mean for Ration B)
- $(n-1)s_1^2 = 112.25$
- $(n-1)s_2^2 = 66.64$
- Pooled variance:  $s^2 = \frac{112.25+66.64}{2 \times 11} = 8.1314$
- $\bar{D} = 3.0833$  (difference of means)
- $s_{\bar{D}} = \sqrt{\frac{2(8.1314)}{12}} = 1.1641$

## Example: Pig Rations (cont'd)

Based on this, we compute the 95% confidence interval:

- $t_{22,0.025} = 2.074$  (for  $\alpha = 0.05$ )
- Confidence interval:

$$(3.0833 \pm 2.074 \times 1.1641) = (0.6689, 5.4978)$$

## Example: Pig Rations (cont'd)

- The 95% confidence interval (0.6689, 5.4978) does not contain 0.
- This suggests strong evidence that the mean weight gains for the two rations are different.
- The sample evidence supports a real difference in weight gain between the rations.

# References and other materials

1. D.R. Cox, E. J. Snell. Applied Statistics - Principles and Examples. Chapman & Hall, 1981.
2. Spiegelhalter, David. The art of statistics: Learning from data. Penguin UK, 2019.
3. James, G., Witten, D., Hastie, T., Tibshirani, R. An introduction to statistical learning. Springer, 2013.
4. Teaching courses from <https://www4.stat.ncsu.edu/~davidian/>
5. Teaching courses from <https://www.drizopoulos.com/>

That's all folks!



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