The file XOM5YrsDaily.txt contains daily returns in percent for Exxon Mobil over the five year (1258 trading day) period 1/3/2012 - 12/30/2016. In this project you will estimate historical XOM price volatility over this period using a hidden Markov chain approach. The underlying (imperfectly observed) Markov chain is XOM's time-dependent daily volatility in percent, σ_t . We assume it is undergoing an exponentiated random walk: $\sigma_t = \sigma \exp(\alpha W_t)$, where σ and α are model parameters. Here $0 \le t \le 1258$ is elapsed time in days and $(W_t : 0 \le t \le 1258)$ is a simple symmetric random walk on $\mathbb Z$ starting at $W_0 = 0$. The volatility process is not observed but we do observe a realization of XOM's daily percent returns $R_t \sim \sigma_t N_t$ where $(N_t : 0 \le t < 1258)$ are iid standard normals that are independent of the W process. Call this realization $(r_t : 0 \le t < 1258)$; r_t is the realized return from time t to time t+1. Here t is in (trading) days.

Preliminaries. Let us label the states of the MC according to the value of the random walk, so the state space is $S = \{k : -1258 \le k \le 1258\}$. At time t the only states with positive probability are

$$S_t = \{k : -t \le k \le t \text{ in steps of } 2\},\$$

e.g., $S_0 = \{0\}$, $S_1 = \{-1, 1\}$, $S_2 = \{-2, 0, 2\}$, and so forth. For $k \in S$, let $s_k = \sigma \exp(\alpha k)$, the daily volatility in percent when the MC is in state k. To estimate σ_t at time t — the volatility of R_t — the only data we may legitimately use is r_0, \ldots, r_{t-1} . So put

$$D_t = \{R_0 = r_0, R_1 = r_1, \dots, R_{t-1} = r_{t-1}\}, \text{ so also } D_{t+1} = D_t \cap \{R_t = r_t\}.$$

For any event E, let $P_0[E] = P[E]$ (no data yet, so conditional probability = unconditional probability) and, for $t \ge 1$, let $P_t[E] = P[E \mid D_t]$. We note that $P_{t+1}[E] = P_t[E \mid R_t = r_t]$ (updating for $D_{t+1} = P_t[E \mid R_t = r_t]$).

Estimating Volatility. Suppose that, with our Bayesian prowess, we have calculated the numbers $p_k = P_t[W_t = k]$ for $k \in S_t$. We have this for t = 0: $p_0 = P_0[W_0 = 0] = P[W_0 = 0] = 1$. Using MAP estimation, put $k^* = \operatorname{argmax}_k(P_t[W_t = k] : k \in S_t)$ and take $\hat{\sigma}_t = s_{k^*}$, so, e.g., $\hat{\sigma}_0 = s_0 = \sigma$.

Bayesian Updating. Next we seek to update the p_k s, which reflect data through r_{t-1} , to reflect the new data $\{R_t = r_t\}$. To this end, for $k \in S_{t+1}$, put

$$\widetilde{p}_k \ = \ P_{t+1}[W_{t+1} = k] \ = \ P_t[W_{t+1} = k \, | \, R_t = r_t] \ = \ \frac{P_t[R_t = r_t, W_{t+1} = k]}{P_t[R_t = r_t]}.$$

Since the walk steps up or down by one each day, if $W_{t+1} = k$ we must also have either $W_t = k + 1$ or $W_t = k - 1$. We partition the numerator into those two cases. Given $\{W_t = j\}$, $R_t \sim \text{Normal}(0, s_i^2)$, so

$$P_{t}[R_{t} = r_{t}, W_{t+1} = k, W_{t} = k+1] = P_{t}[R_{t} = r_{t}, W_{t+1} = k \mid W_{t} = k+1] P_{t}[W_{t} = k+1]$$

$$= f(r_{t}, s_{k+1}) \cdot \frac{1}{2} \cdot p_{k+1},$$

where f(r,s) is the density function of a Normal $(0,s^2)$ evaluated at r: $f(r,s) = \frac{1}{\sqrt{2\pi}s}e^{-r^2/2s^2}$. Similarly,

$$P_t[R_t = r_t, W_{t+1} = k, W_t = k-1] = f(r_t, s_{k-1}) \cdot \frac{1}{2} \cdot p_{k-1}.$$

It follows that

$$P_t[R_t = r_t, W_{t+1} = k] = \frac{1}{2} f(r_t, s_{k-1}) p_{k-1} + \frac{1}{2} f(r_t, s_{k+1}) p_{k+1}.$$
 (1)

To complete the calculation of \widetilde{p}_k for $k \in S_{t+1}$, we compute the denominator:

$$P_t[R_t = r_t] = \sum_{j \in S_t} P_t[R_t = r_t \mid W_t = j] P_t[W_t = j] = \sum_{j \in S_t} f(r_t, s_j) p_j,$$
 (2)

yielding

$$\widetilde{p}_{k} = \frac{f(r_{t}, s_{k-1})p_{k-1} + f(r_{t}, s_{k+1})p_{k+1}}{2\sum_{j \in S_{t}} f(r_{t}, s_{j})p_{j}}.$$
(3)

In (1), (2), and (3), f(r,s) is a density rather than a probability, but it gives the right answer; why?

The numbers \widetilde{p}_k then become the new p_k when we advance to updating for the next new data, namely $\{R_{t+1} = r_{t+1}\}$. In this manner we may compute the numbers $(p_k : k \in S_t)$ for t = 1, then 2, then 3, and so forth to 1258 — and compute the corresponding $\widehat{\sigma}_t$.

Your mission is to implement this model. The model has two free parameters, σ and α , which must be estimated using the data. I have done that for you! Take $\sigma = 1.11$ and $\alpha = 0.038$. The code HiddenMC.cpp will get you started. As it stands, this code reads in the time series $(r_t: 0 \le t < 1258)$ from the data file XOM5YrsDaily.txt. It also generates some output files as discussed below.

- Implement this model and generate the sequence ($\hat{\sigma}_t : 0 \le t \le 1258$). Do not report it, but ...
- Graphically present the annualized time series $(\sqrt{252}\hat{\sigma}_t: 51 \le t \le 1258)$ and compare this to the same data that EWMA produced in class (this data is in the file EWMA.txt). You may wish to use the TeX file HistoricalVols.tex to do this, in which case you must put the hidden MC data in a file called HMC.txt in the same format as EWMA.txt. (The Report function in the starter code does this for you!)
- Generate a scatter plot of the points $((x_t, y_t): 51 \le t \le 1258)$ where x_t is the EWMA annualized daily volatility in percent and $y_t = \sqrt{252}\widehat{\sigma}_t$ is the HMC annualized daily volatility in percent. You may wish to use the TeX file ScatterPlot.tex to do this, in which case you must put the data in a file called ScatterData.txt. (The Report function in the starter code does this for you!)
- Generate a (normal) histogram of the standardized returns $(r_t/\widehat{\sigma}_t : 51 \le t < 1258)$. You may wish to use the NormalHistogram function in Functions.h to do this. Use 20 "buckets". (The Report function in the starter code does this for you!)
- Graphically present the time series of standardized returns. You may wish to use the TeX file ReturnTS.tex to do this. (The required data file StandardizedXOM.txt is generated in the Report function!)