

1. NALOGA

a)

Povprečni dohodek

Populacija je velikosti 43886. Vzamemo enostavni slučajni vzorec 400 enot.

Sledi: $N = 43886, n = 400$.

Naj bo X_k skupni dohodek v k -ti družini.

Torej je povprečni dohodek v Kindergradu:

$$\bar{X} = \frac{X_1 + \dots + X_n}{n}.$$

Standardna napaka

Vemo: $se(\bar{X}) = \sqrt{var(\bar{X})}$. Ker imamo enostavni slučajni vzorec, vemo tudi, da je

$$var(\bar{X}) = \frac{1}{n} \frac{N-n}{N-1} \sigma^2$$

kjer je σ^2 populacijska varjanca. Nepristranska cenilka za σ^2 je

$$\hat{\sigma}^2 = \frac{N-1}{N(n-1)} \sum_{k=1}^n (X_k - \bar{X})^2.$$

Sledi:

$$\widehat{se(\bar{X})} = \sqrt{\frac{1}{n} \frac{N-n}{N(n-1)} \sum_{k=1}^n (X_k - \bar{X})^2}.$$

Interval zaupanja

Iz navodil sledi, da je interval zaupanja enak $\bar{X} \pm 1,96 \cdot se(\bar{X})$.

Končne vrednosti

b)

Če stratificiramo, mora veljati

$$\frac{n_k}{n} = \frac{N_k}{N}, \sum_{k=1}^K n_k = n.$$

V našem primeru stratificiramo po četrtih, torej $k = 4$. Vemo:

$N_1 = 10149$ (*severna četrt*),

$N_2 = 10390$ (*vzgodna četrt*),

$N_3 = 13457$ (*južna četrt*),
 $N_4 = 9890$ (*zahodna četrt*). Če malo obrnemo zgornjo enakost, dobimo

$$n_k = \frac{N_k}{N}n.$$

Izračunamo za $k = 1, 2, 3, 4$ in upoštevamo vrednosti N_1, N_2, N_3, N_4 ter $N = 43886$.

Dobimo:

$$n_1 = \frac{10149}{43886} \cdot 400 = 92,5033 \rightarrow n_1 = 92,$$

$$n_2 = \frac{10390}{43886} \cdot 400 = 94,699 \rightarrow n_2 = 95,$$

$$n_3 = \frac{13457}{43886} \cdot 400 = 122,654 \rightarrow n_3 = 123,$$

$$n_4 = \frac{9890}{43886} \cdot 400 = 90,142 \rightarrow n_4 = 90.$$

Preverimo:

$$\sum_{k=1}^4 n_k = 92 + 96 + 123 + 90 = 400.$$

Naj bo sedaj X_{kj} povprečni dohodek j -te družine v k -tem stratumu. Povprečni dohodek družine se sedaj izraža kot:

$$\bar{X} = \frac{1}{n} \sum_{k=1}^{\#stratumov} \sum_{j=1}^{n_k} X_{kj}.$$

Standarna napaka $se(\bar{X}) = \sqrt{var(\bar{X})}$:

$$var(\bar{X}) = \sum_k w_k^2 var(\bar{X}_k) = \sum_k w_k^2 \cdot \frac{\hat{\sigma}_k^2}{n_k} \cdot \frac{N_k - n_k}{N_k - 1},$$

kjer je $w_k = \frac{N_k}{N}$ delež, σ_k^2 pa populacijska varjanca v k -tem stratumu. Torej:

$$\hat{\sigma}_k^2 = \frac{N_k - 1}{N_k(n_k - 1)} \sum_{j=1}^{n_k} (X_{kj} - \bar{X}_k)^2,$$

kjer je \bar{X}_k povprečje k -tega stratuma.

Interval zaupanja: $\bar{X} \pm 1,96 \cdot se(\bar{X})$.

Vstavimo podatke in dobimo:

c)

2. NALOGA

3. NALOGA

a)

$$X \sim f(x, \alpha) = \begin{cases} \frac{\Gamma(3\alpha)}{\Gamma(\alpha)\Gamma(2\alpha)} x^{\alpha-1} (1-x)^{2\alpha-1}, & 0 < x < 1. \\ 0, & \text{šicer.} \end{cases}$$

Vemo:

$$E(X) = \frac{1}{3}, \text{var}(X) = \frac{2}{9(3\alpha+1)}.$$

Ker je:

$$\text{var}(X) = E(X^2) - E(X)^2,$$

sledi:

$$E(X^2) = \text{var}(X) + E(X)^2 = \frac{2}{9(3\alpha+1)} + \frac{1}{9} = \frac{\alpha+1}{3(3\alpha+1)}$$

$E(X^2)$ je drugi moment slučajne spremenljivke X , torej:

$$E(X^2) = \frac{1}{n} \sum_{i=1}^n X_i^2.$$

Zgornji enačbi izenačimo in poračunamo α :

$$\frac{1}{n} \sum_{i=1}^n X_i^2 = \frac{\alpha+1}{3(3\alpha+1)}$$

$$\frac{9\alpha+3}{n} \sum_{i=1}^n X_i^2 = \alpha+1$$

$$\frac{9\alpha}{n} \sum_{i=1}^n X_i^2 - \alpha = 1 - \frac{3}{n} \sum_{i=1}^n X_i^2$$

$$\alpha = \frac{1 - \frac{3}{n} \sum_{i=1}^n X_i^2}{\frac{9}{n} \sum_{i=1}^n X_i^2 - 1}$$

b)

$$f_X(x, \alpha) = \begin{cases} \frac{\Gamma(3\alpha)}{\Gamma(\alpha)\Gamma(2\alpha)} x^{\alpha-1} (1-x)^{2\alpha-1}, & 0 < x < 1. \\ 0, & \text{šicer.} \end{cases}$$

$$L_1(\alpha|x) = \frac{\Gamma(3\alpha)}{\Gamma(\alpha)\Gamma(2\alpha)} x^{\alpha-1} (1-x)^{2\alpha-1}$$

$$L(\alpha|x_1, \dots, x_n) = L_1(\alpha|x_1) \cdot \dots \cdot L_1(\alpha|x_n) =$$

$$\left(\frac{\Gamma(3\alpha)}{\Gamma(\alpha)\Gamma(2\alpha)} \right)^n x_1^{\alpha-1} \cdot \dots \cdot x_n^{\alpha-1} (1-x_1)^{2\alpha-1} \cdot \dots \cdot (1-x_n)^{2\alpha-1}$$

$$l(\alpha|x_1, \dots, x_n) = \ln(L(\alpha|x_1, \dots, x_n)) = l_1(\alpha|x_1) + \dots + l_1(\alpha|x_n)$$

$$l_1(\alpha|x) = \ln(L_1(\alpha|x)) =$$

$$\ln(\Gamma(3\alpha)) - \ln(\Gamma(\alpha)) - \ln(\Gamma(2\alpha)) + (\alpha-1)\ln(x) + (2\alpha-1)\ln(1-x)$$

$$l(\alpha|x) =$$

$$\sum_{i=1}^n (\ln(\Gamma(3\alpha)) - \ln(\Gamma(\alpha)) - \ln(\Gamma(2\alpha)) + (\alpha-1)\ln(x_i) + (2\alpha-1)\ln(1-x_i))$$

$$\frac{\partial l}{\partial \alpha} = \sum_{i=1}^n \frac{1}{\Gamma(3\alpha)} \Gamma'(3\alpha) 3 - \frac{1}{\Gamma(\alpha)} \Gamma'(\alpha) - \frac{1}{\Gamma(2\alpha)} \Gamma'(2\alpha) 2 + \ln(x_i) + 2\ln(1-x_i)$$

$$\sum_{i=1}^n \frac{1}{\Gamma(3\alpha)} \Gamma'(3\alpha) 3 - \frac{1}{\Gamma(\alpha)} \Gamma'(\alpha) - \frac{1}{\Gamma(2\alpha)} \Gamma'(2\alpha) 2 = \ln\left(\frac{1}{x_i(1-x_i)}\right)$$

Cenilka obstaja, ko ima zgornja enačba rešitev.

Uporabimo funkcijo digamma in rešimo do konca? no clue.

c)

$$var(\hat{\alpha}) = \frac{1}{nI_1(\hat{\alpha})}.$$

$$I_1(\hat{\alpha}) = -E \left[\frac{\partial^2 l_1(\alpha|x)}{\partial \alpha^2} \right]$$

$$\frac{\partial^2 l_1(\alpha|x)}{\partial \alpha^2} =$$

$$3 \frac{\Gamma''(3\alpha)\Gamma(3\alpha) - \Gamma'(3\alpha)^2}{\Gamma(3\alpha)^2} - \frac{\Gamma''(\alpha)\Gamma(\alpha) - \Gamma'(\alpha)^2}{\Gamma(\alpha)^2} - 2 \frac{\Gamma''(2\alpha)\Gamma(2\alpha) - \Gamma'(2\alpha)^2}{\Gamma(2\alpha)^2}$$

$$var(\hat{\alpha}) = \frac{1}{n} \frac{1}{3 \frac{\Gamma''(3\alpha)\Gamma(3\alpha) - \Gamma'(3\alpha)^2}{\Gamma(3\alpha)^2} - \frac{\Gamma''(\alpha)\Gamma(\alpha) - \Gamma'(\alpha)^2}{\Gamma(\alpha)^2} - 2 \frac{\Gamma''(2\alpha)\Gamma(2\alpha) - \Gamma'(2\alpha)^2}{\Gamma(2\alpha)^2}}$$

4. NALOGA

5. NALOGA

Ocen po metodi največjega verjetja

Da bomo vedeli kako so porazdeljeni Y_i -ji, poračunamo njihovo pričakovano vrednost in varjanco.

$$E(Y_i) = E(\beta_0 + \beta_1 x_i + \epsilon_i) = \beta_0 + \beta_1 x_i$$

$$\text{var}(Y_i) = \sigma^2$$

Torej so Y_i porazdeljeni $N(\beta_0 + \beta_1 x_i, \sigma^2)$. Dobimo:

$$L_1 = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y_i - \beta_0 - \beta_1 x_i)^2}{2\sigma^2}}$$

Verjetnostna funkcija je sestavljena iz produktov porazdelitvenih funkcij. Torej dobimo:

$$\log(L) = -n\log(\sigma) - \frac{n}{2}\log(2\pi) - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)^2$$

. Najprej poračunamo za prvi parameter β_0 . Enačbo odvajamo po β_0 in enačimo z 0.

$$\begin{aligned} \frac{\partial}{\partial \beta_0} \log(L) &= \frac{\partial}{\partial \beta_0} (-n\log(\sigma) - \frac{n}{2}\log(2\pi) - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)^2) = \\ &= -\frac{1}{\sigma^2} \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i) \end{aligned}$$

Dobimo:

$$\sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i) = 0$$

Sledi:

$$\sum_{i=1}^n y_i - n\beta_0 - \beta_1 \sum_{i=1}^n x_i = 0$$

Torej je končni rezultat enak:

$$\beta_0 = \frac{1}{n} \sum_{i=1}^n y_i - \beta_1 \sum_{i=1}^n x_i.$$

Postopek ponovimo za drugi parameter β_1 :

$$\begin{aligned}\frac{\partial}{\partial \beta_1} \log(L) &= \frac{\partial}{\partial \beta_1} \left(-n \log(\sigma) - \frac{n}{2} \log(2\pi) - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \beta_o - \beta_1 x_i)^2 \right) = \\ &= -\frac{1}{\sigma^2} \sum_{i=1}^n x_i (y_i - \beta_o - \beta_1 x_i)\end{aligned}$$

Dobimo:

$$\sum_{i=1}^n x_i (y_i - \beta_o - \beta_1 x_i) = 0.$$

Sledi:

$$\begin{aligned}\sum_{i=1}^n x_i y_i - n \beta_o - \beta_1 \sum_{i=1}^n x_i &= 0. \\ \sum_{i=1}^n x_i y_i - n \beta_o - \beta_1 \sum_{i=1}^n x_i &= 0.\end{aligned}$$

Vstavimo vrednost za β_o :

$$\begin{aligned}\sum_{i=1}^n x_i y_i - n \left(\frac{1}{n} \sum_{i=1}^n y_i - \beta_1 \sum_{i=1}^n x_i \right) - \beta_1 \sum_{i=1}^n x_i &= 0. \\ \sum_{i=1}^n x_i y_i - \sum_{i=1}^n y_i - \beta_1 \left(n \sum_{i=1}^n x_i - \sum_{i=1}^n x_i \right) &= 0.\end{aligned}$$

Torej je končni rezultat enak:

$$\beta_1 = \frac{\sum_{i=1}^n x_i y_i - \sum_{i=1}^n y_i}{n \sum_{i=1}^n x_i - \sum_{i=1}^n x_i}.$$

Vstavimo še za β_o :

$$\begin{aligned}\beta_o &= \frac{1}{n} \sum_{i=1}^n y_i - \frac{\sum_{i=1}^n x_i y_i - \sum_{i=1}^n y_i}{n \sum_{i=1}^n x_i - \sum_{i=1}^n x_i} \sum_{i=1}^n x_i. \\ \beta_o &= \frac{1}{n} \sum_{i=1}^n y_i - \frac{\sum_{i=1}^n x_i \sum_{i=1}^n x_i y_i - \sum_{i=1}^n x_i \sum_{i=1}^n y_i}{n \sum_{i=1}^n x_i - \sum_{i=1}^n x_i}. \\ \beta_o &= \frac{\sum_{i=1}^n x_i \sum_{i=1}^n y_i - \frac{1}{n} \sum_{i=1}^n x_i \sum_{i=1}^n y_i - \sum_{i=1}^n x_i \sum_{i=1}^n x_i y_i + \sum_{i=1}^n x_i \sum_{i=1}^n y_i}{n \sum_{i=1}^n x_i - \sum_{i=1}^n x_i}.\end{aligned}$$

Ocena po metodi najmanjših kvadratov

Ker so šumi ϵ_i neodvisni za vsak $i = 1, 2, \dots, n$ in $\epsilon_i \sim N(0, \sigma^2)$ lahko uporabimo izrek Gauss-Markova za ocene parametrov β_0 in β_1 . Imamo:

$$\begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix} = \begin{bmatrix} 1 & x_1 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \end{bmatrix} + \begin{bmatrix} \epsilon_1 \\ \vdots \\ \epsilon_n \end{bmatrix}$$

Po izreku Gauss-Markova vemo da je najboljša cenilka za $\beta = (X^T X)^{-1} X^T Y$. Računamo:

$$\begin{aligned} & \left(\begin{bmatrix} 1 & \cdots & 1 \\ x_1 & \cdots & x_n \end{bmatrix} \begin{bmatrix} 1 & x_1 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix} \right)^{-1} \begin{bmatrix} 1 & \cdots & 1 \\ x_1 & \cdots & x_n \end{bmatrix} \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix} = \\ & = \left(\begin{bmatrix} n & \sum_{i=1}^n x_i \\ \sum_{i=1}^n x_i & \sum_{i=1}^n x_i^2 \end{bmatrix} \right)^{-1} \begin{bmatrix} \sum_{i=1}^n y_i \\ \sum_{i=1}^n x_i y_i \end{bmatrix} = \\ & = \frac{1}{n \sum_{i=1}^n x_i^2 - \sum_{i=1}^n x_i \sum_{i=1}^n x_i} \begin{bmatrix} \sum_{i=1}^n x_i^2 & -\sum_{i=1}^n x_i \\ -\sum_{i=1}^n x_i & n \end{bmatrix} \begin{bmatrix} \sum_{i=1}^n y_i \\ \sum_{i=1}^n x_i y_i \end{bmatrix} = \\ & = \frac{1}{n \sum_{i=1}^n x_i^2 - \sum_{i=1}^n x_i \sum_{i=1}^n x_i} \begin{bmatrix} \sum_{i=1}^n x_i^2 \sum_{i=1}^n y_i - \sum_{i=1}^n x_i \sum_{i=1}^n x_i y_i \\ -\sum_{i=1}^n x_i \sum_{i=1}^n y_i + n \sum_{i=1}^n x_i y_i \end{bmatrix} \end{aligned}$$

Z malo računanja in spretnosti vidimo da sta cenilki enaki.