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Question 1

(a) ① To proof non-negative:

$$p(\theta) = 30\theta^2 \cdot (1-\theta)^2$$

$$\because \theta^2 > 0, (1-\theta)^2 > 0$$

$$\therefore p(\theta) > 0$$

\therefore non-negative

② To proof normalised:

$$\because \theta \in [0, 1]$$

$$\therefore \int_0^1 p(\theta) d\theta = \int_0^1 30\theta^2(1-\theta)^2 d\theta$$

$$= \int_0^1 30\theta^2 - 60\theta^3 + 30\theta^4 d\theta$$

$$= 10\theta^3 \Big|_0^1 - 15\theta^4 \Big|_0^1 + 6\theta^5 \Big|_0^1$$

$$= 1$$

\therefore normalised

\therefore Proof $p(\theta)$ is valid

$$(b) \text{ ① } p(\theta | x_{1:2} = 01) = \frac{p(x_{1:2} = 01 | \theta) \cdot p(\theta)}{p(x_{1:2} = 01)} \quad (\text{Bayes' Theorem})$$

\because flips are independent

$$\therefore p(x_{1:2} = 01 | \theta) = p(x_1 = 0 | \theta) \cdot p(x_2 = 1 | \theta) = (1-\theta) \cdot \theta$$

$$\therefore p(x_{1:2} = 01) = \int_0^1 p(x_{1:2} = 01 | \theta) \cdot p(\theta) d\theta$$

$$\Rightarrow P(\theta | x_{1:2} = 01) = \frac{(1-\theta) \cdot \theta \cdot 30\theta^2 (1-\theta)^2}{\int_0^1 (1-\theta)^3 \cdot 30\theta^3 d\theta} = 30 \times \frac{14}{3} \theta^3 \cdot (1-\theta)^3 = 140 \theta^3 (1-\theta)^3$$

$$\begin{aligned} \textcircled{2} P(\theta | x_{1:4} = 0101) &= \frac{P(x_{1:4} = 0101 | \theta) \cdot P(\theta)}{P(x_{1:4} = 0101)} && \text{Similarly as } \textcircled{1} \\ &= \frac{(1-\theta)^2 \theta^2 \cdot 30\theta^2 (1-\theta)^2}{\left(\int_0^1 (1-\theta)^4 \cdot 30\theta^4 d\theta \right)} = 21 \times 30 \cdot \theta^4 (1-\theta)^4 \\ &= 630 \theta^4 (1-\theta)^4 \end{aligned}$$

$$(c) \quad \mu = \int_0^1 p(\theta) \cdot \theta \cdot d\theta$$

$$\begin{aligned} \textcircled{1} \mu_0 &= \int_0^1 30\theta^2 (1-\theta)^2 \cdot \theta d\theta \\ &= \int_0^1 30\theta^5 \cdot 600\theta^4 + 30\theta^3 d\theta \\ &= \frac{1}{2} = 0.5 \end{aligned}$$

$$\begin{aligned} \textcircled{2} \mu_{012} &= \int_0^1 140\theta^3 (1-\theta)^3 \cdot \theta d\theta \quad (\text{using CAS to assist with integration}) \\ &= \frac{1}{2} = 0.5 \end{aligned}$$

$$\begin{aligned} \textcircled{3} \mu_{01234} &= \int_0^1 630\theta^4 (1-\theta)^4 \cdot \theta d\theta \\ &= \frac{1}{2} = 0.5 \end{aligned}$$

(d) For $\sigma^2 = \mu(\theta^2) - [\mu(\theta)]^2$ as $\mu(\theta)$ calculated in (c)

$$\begin{aligned}\textcircled{1} \sigma_{\theta}^2 &= \int_0^1 \theta^2 \cdot p(\theta) d\theta - \frac{1}{4} \\ &= \int_0^1 \theta^2 \cdot 30\theta^2(1-\theta)^2 d\theta - \frac{1}{4} \\ &= \frac{2}{7} - \frac{1}{4} = \frac{1}{28}\end{aligned}$$

$$\begin{aligned}\textcircled{2} \sigma_{12}^2 &= \int_0^1 \theta^2 \cdot p(\theta | x_{1,2} = 01) \cdot d\theta - \left(\frac{3}{7}\right)^2 \\ &= \int_0^1 \theta^2 \cdot 140\theta^3(1-\theta)^3 d\theta - \frac{1}{4} \\ &= \frac{5}{18} - \frac{1}{4} \\ &= \frac{1}{36}\end{aligned}$$

$$\begin{aligned}\textcircled{3} \sigma_{1234}^2 &= \int_0^1 \theta^2 \cdot p(\theta | x_{1234} = 0101) \cdot d\theta - \left(\frac{8}{21}\right)^2 \\ &= \int_0^1 630\theta^6(1-\theta)^4 d\theta - \frac{1}{4} \\ &= \frac{2}{11} - \frac{1}{4} \\ &= \frac{1}{44}\end{aligned}$$

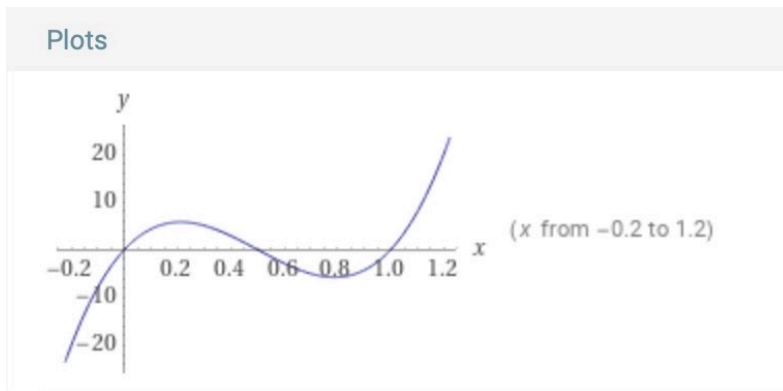
$$(e) \theta_{\text{map}} = \arg \max_{\theta} p(\theta|x) = \arg \max_{\theta} \frac{p(x|\theta) \cdot p(\theta)}{p(x)} \propto \arg \max_{\theta} p(x|\theta) \cdot p(\theta)$$

$$\textcircled{1} \theta_{\text{map}, \theta} = \arg \max_{\theta} 30\theta^2(1-\theta)^2 \quad (\text{where } \theta \in [0,1])$$

$$\Rightarrow 60\theta \cdot (2\theta^2 - 3\theta + 1) = 0$$

(Using CAS to assist with differentiation)

the plots is shown below:



$$\therefore \text{When } \theta < 0.5, (30\theta^2(1-\theta)^2)' > 0.$$

$$\theta > 0.5, (30\theta^2(1-\theta)^2)' < 0$$

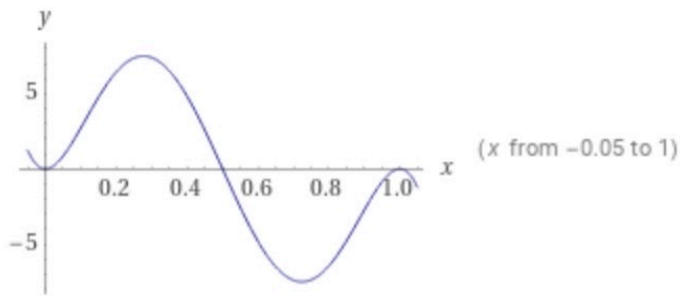
$$\therefore \text{we can get } \theta_{\text{map}, \theta} = 0.5 \text{ by } 2\theta^2 - 3\theta + 1 = 0$$

$$\textcircled{2} \theta_{\text{map}, 1,2} = \arg \max_{\theta} 140\theta^3(1-\theta)^3 \quad (\theta \in [0,1])$$

$$\Rightarrow -420(\theta-1)^2\theta^2(2\theta-1)$$

the plots is shown below

Plots



Similarly as ①

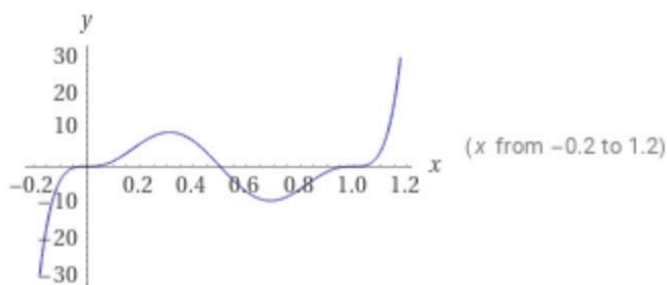
we can get $\Theta_{map} = 0.5$

$$\textcircled{3} \quad \Theta_{map, 1234} = \arg \max_{\Theta} 650 \Theta^4 (1-\Theta)^4$$

$$\Rightarrow 2500 \Theta^3 (1-\Theta)^3 (1-\Theta) = 0$$

the plots is shown below:

Plots

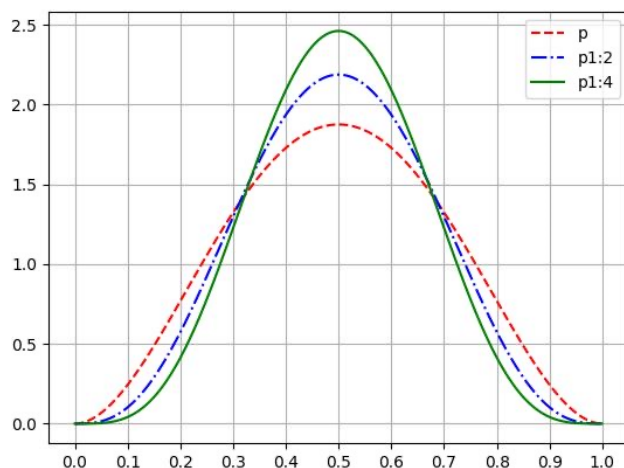


Similarly as ① ②:

we can get $\Theta_{map, 1234} = 0.5$

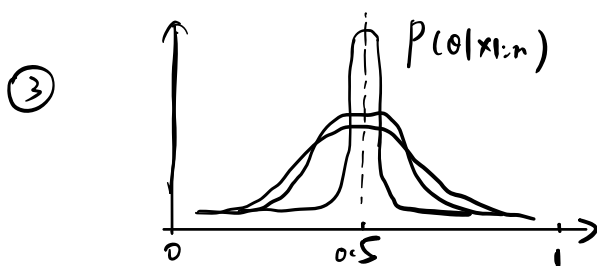
Posterior	PDF	μ	σ^2	θ_{MAP}
$P(\theta)$	$30\theta^2(1-\theta)^2$	$\frac{1}{2}$	$\frac{1}{28}$	$\frac{1}{2}$
$P(\theta X_{1:2}=01)$	$140\theta^3(1-\theta)^3$	$\frac{1}{2}$	$\frac{1}{36}$	$\frac{1}{2}$
$P(\theta X_{1:4}=0101)$	$630\theta^4(1-\theta)^4$	$\frac{1}{2}$	$\frac{1}{44}$	$\frac{1}{2}$

(f)



(g) ① When $X_{1:n} = 0101\dots$, $P(\theta|X_{1:n})$ will have more chance on 0.5, which means more and more close to $P(X=1) = P(X=0) = \frac{1}{2}$

② For large n , μ still $\frac{1}{2}$. σ^2 getting smaller and smaller. θ_{MAP} still $\frac{1}{2}$.



Question 2

(a) if $\phi = 1$, then $P(\hat{X}=1 | X=1) = 1$

$$P(\hat{X}=0 | X=0) = 1$$

always return correct answer
without errors

if $\phi = 0.5$, then each return's confidence is unreliable
as it has same chance report right or wrong
answers.

if $\phi = 0$ then it always report the opposite answer.

(b) Known $P(X=0|\theta) = 1-\theta$, $P(X=1|\theta) = \theta$

$\because X=0$ and $X=1$ i.i.d

$$\begin{aligned}\therefore P(\hat{X}=0|\theta) &= P(\hat{X}=0|\theta, X=0) + P(\hat{X}=0|\theta, X=1) \\ &= (1-\theta) \cdot \phi + \theta \cdot (1-\phi)\end{aligned}$$

$$\begin{aligned}P(\hat{X}=1|\theta) &= P(\hat{X}=1|\theta, X=0) + P(\hat{X}=1|\theta, X=1) \\ &= (1-\theta) \cdot (1-\phi) + \theta \cdot \phi\end{aligned}$$

(c) Bayes' Theorem

$$P(\theta | \hat{X}=0) = \frac{P(\hat{X}=0|\theta) \cdot P(\theta)}{P(\hat{X}=0)}$$

$$= \frac{[(1-\theta)\phi + \theta \cdot (1-\phi)] \cdot p(\theta)}{p(\hat{x}=0)}$$

where $p(\hat{x}=0) = \int_0^1 p(\hat{x}=0|\theta) \cdot p(\theta) \cdot d\theta \quad (\theta \in [0,1])$

$$= \int_0^1 [(1-\theta)\phi + \theta(1-\phi)] \cdot p(\theta) d\theta$$

$$\therefore p(\theta|\hat{x}=0) = \frac{[(1-\theta)\phi + \theta \cdot (1-\phi)] \cdot p(\theta)}{\int_0^1 [(1-\theta)\phi + \theta(1-\phi)] \cdot p(\theta) d\theta}$$

① when $\phi=1$

$$p(\theta|\hat{x}=0) = \frac{(1-\theta) \cdot p(\theta)}{\int_0^1 (1-\theta) \cdot p(\theta) d\theta}$$

② when $\phi=0.5$

$$p(\theta|\hat{x}=0) = \frac{\frac{1}{2} \cdot p(\theta)}{\int_0^1 \frac{1}{2} p(\theta) d\theta} = \frac{p(\theta)}{\int_0^1 p(\theta) d\theta} = p(\theta)$$

($\because \int_0^1 p(\theta) d\theta = 1$)

③ when $\phi=0$

$$p(\theta|\hat{x}=0) = \frac{\theta p(\theta)}{\int_0^1 \theta p(\theta) d\theta}$$

Observation:

From $\phi=1$ to $\phi=0.5$, $\phi=0$. Correct \rightarrow Wrong report.

$\phi=1$ rely on appropriate corresponding $\theta \rightarrow X=0$.

$\phi=0.5$ totally rely on prior $p(\theta)$

$\phi=0$ rely on wrong corresponding $\theta \rightarrow X=1$

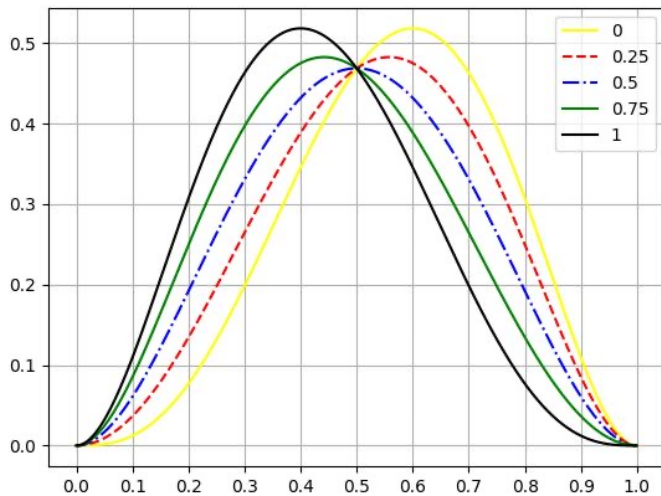
(d) As calculated in (c)

$$\begin{aligned} p(\theta | \hat{x}=0) &= \frac{[(1-\theta)\phi + \theta(1-\phi)] \cdot 30\theta^2(1-\theta)^2}{\int_0^1 [(1-\theta)\phi + \theta(1-\phi)] \cdot 30\theta^2(1-\theta)^2 d\theta} \\ &= \frac{[(1-\theta)\phi + \theta(1-\phi)] \cdot 30\theta^2(1-\theta)^2}{\int_0^1 [30\theta^2(1-\theta)^3\phi + 30\theta^3(1-\theta)^2(1-\phi)] d\theta} \end{aligned}$$

Using CAS to assist with integration, denominator can be calculated as $-\frac{1}{2}$

$$\begin{aligned} \therefore p(\theta | \hat{x}=0) &= 2 \left[[(1-\theta)\phi + \theta(1-\phi)] \cdot 30\theta^2(1-\theta)^2 \right] \\ &= 60 \left[\theta^2(1-\theta)^3\phi + \theta^3(1-\theta)^2(1-\phi) \right] \\ &= 60\theta^2(1-\theta)^2 \cdot [(1-\theta)\phi + \theta(1-\phi)] \end{aligned}$$

(2)



①

The shape of distribution will shift to the left, which means $P(\theta | \hat{x})$ is changing on $P(\hat{x} | x)$ distribution.

②

ϕ more close to 1 means the observation \hat{x} has more probability to be real x . so $P(\theta | \hat{x})$ is more close to real $P(\theta | x)$.

Question 3

$$F(x) = P(X \leq x) = \int x^2 + \frac{2}{3}x + \frac{1}{3} dx$$

$$= \frac{1}{3}x^3 + \frac{1}{3}x^2 + \frac{1}{3}x$$

$$\because Y = X^2 + 2, \quad x \in [0, 1]$$

$$\therefore x = \sqrt{Y-2}$$

$$\text{For } F(y) = P(X \leq \sqrt{y-2}) = F(\sqrt{y-2}) = \frac{1}{3}(y-2)^{\frac{3}{2}} + \frac{1}{3}(y-2) + \frac{1}{3}(y-2)^{\frac{1}{2}}$$

$$\therefore P(Y) = \frac{1}{2}\sqrt{y-2} + \frac{1}{3} + \frac{1}{6}\frac{1}{\sqrt{y-2}}$$