Simulation

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Setup

```
#install.packages('DMwR')
#install.packages(c('xts', 'quantmod', 'abind', 'ROCR'))
#install.packages( "DMwR_0.4.1.tar.gz", repos=NULL, type="source" )
library(DMwR)
```

Generate simulation data

```
# set parameters
set.seed(98)
reps = 200
n = 1000
beta = c(\log(0.05 / 0.95) - 4.15, 1, 1) # Was -2, but changed to -4.15 to get 0.05
X_{mean} = 1
X1_{var} = X_{mean} ^2
X2\_cond\_var = 1
X2_var = X2_cond_var + X1_var
# list of dataframes
sims <- vector("list", reps)</pre>
for(i in 1:reps){
  # generate covariates. X_1 is an exponential with rate 1 and
  \# X_2 is a normal with mean X_1 and SD = 1
  X1 <- rexp(n, X_mean)</pre>
  X2 <- rnorm(n, X1, X2_cond_var)</pre>
  # generate response variable. Y_i is distributed bernoulli
  \# s.t. \ logit(P(Y_i = 1)) = beta_0 + beta_1 * X_1 + beta_2 * X_2
  p = \exp(beta[1] + beta[2] * X1 + beta[3] * X2) /
   (1 + \exp(beta[1] + beta[2] * X1 + beta[3] * X2))
  Y = rbinom(n, 1, p)
  #Y = rep(NA, n)
  #for(j in 1:n){
  \# Y[j] = rbinom(1, 1, p[j])
  #}
  # can we do rbinom(n, 1, p) instead of the loop?
  df = data.frame("X1" = X1, "X2" = X2, "Y" = Y)
  sims[[i]] = df
}
```

Helper functions:

```
corr = function(a, b){
  # Computes the correlation between two vectors a and b
 num = mean(a * b) - mean(a) * mean(b)
  denom = sd(a) * sd(b)
 return(num / denom)
}
get_empty_df = function(sz = reps){
  # creates and returns an empty dataframe to
  # be populated by simulation results.
 df = data.frame(
   mean_X1 = rep(NA, sz),
   mean_X2 = rep(NA, sz),
   var_X1 = rep(NA, sz),
   var_X2 = rep(NA, sz),
   corr_X1_X2 = rep(NA, sz),
   beta0_star = rep(NA, sz),
   beta1_star = rep(NA, sz),
   beta2_star = rep(NA, sz),
   F1 = rep(NA, sz)
 )
 return(df)
get_perc_over = function(Y, desired_ratio){
  # Given a response variable Y and the desired proportion of 1s in Y,
  # this function calculates the required percentage of oversampling that
  # is required in order to achieve the desired proportion of 1s in Y.
  # This is designed so the output can be directly passed to the SMOTE function.
 num1 = sum(Y == 1)
 num0 = sum(Y == 0)
  # solved mathematically
 perc = 100 * desired_ratio * num0 / (num1 *(1 - desired_ratio))
 return(perc)
compute_metrics = function(df, df_test){
 # Given a data frame (Either the original generated dataset or the
  # augmented dataset) and a test data frame, this function calculates all
  # the metrics we are interested in and returns them.
  # This is designed to return a row that can be added to a
  # dataframe of results.
  # distribution metrics:
 mean X1 = mean(df$X1)
 mean X2 = mean(df$X2)
  var_X1 = var(df$X1)
  var_X2 = var(df$X2)
  # Correlation metric:
 corr_X1_X2 = corr(df$X1, df$X2)
```

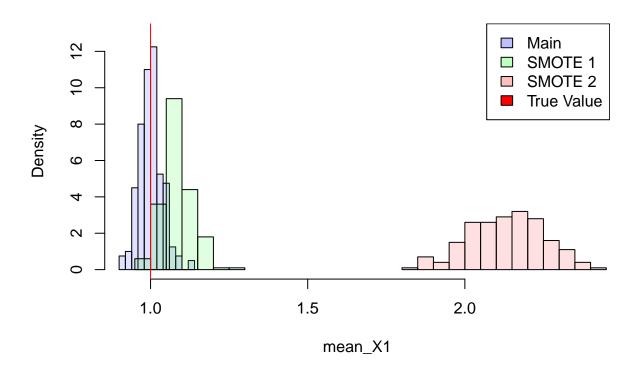
```
# Coefficient metrics:
model = glm(Y ~ X1 + X2, data = df, family = "binomial")
beta_star = model$coefficients
# F1 score on test data
Y_star = (predict(model, newdata = df_test, type = "response") > 0.5) * 1
TP = sum(Y_star == df_test$Y & Y_star == 1)
FP = sum(Y_star != df_test$Y & Y_star == 1)
FN = sum(Y_star != df_test$Y & Y_star == 0)
precision = TP / (TP + FP)
recall = TP / (TP + FN)
F1 = 2 * (precision * recall) / (precision + recall)
return(list(mean_X1,
            mean_X2,
            var_X1,
            var_X2,
            corr_X1_X2,
            beta_star[1],
            beta_star[2],
            beta_star[3],
            F1))
```

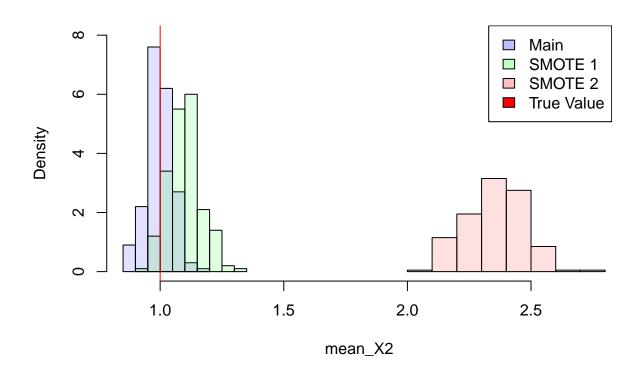
Simulation

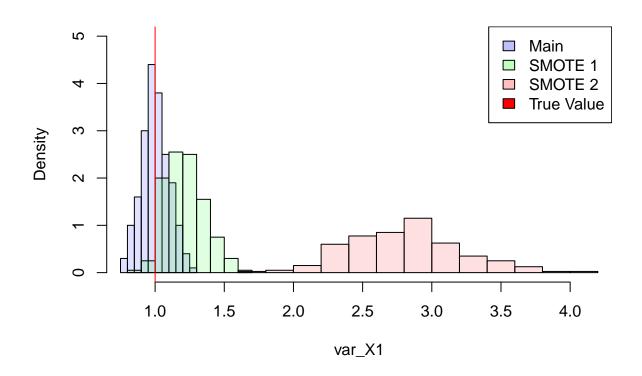
```
# The three dataframes of results
res_base = get_empty_df()
res_smote1 = get_empty_df()
res_smote2 = get_empty_df()
\#means = rep(NA, reps)
for (i in 1:reps){
  \# Get the i-th data frame and split it into train and test
 df = sims[[i]]
 df_{train} = df[-1:-(0.2 * n),]
 df_{test} = df[1:(0.2 * n),]
  # Calculate the metrics for the original dataset
 res_base[i, ] = compute_metrics(df_train, df_test)
  # Convert Y into factor as required for the SMOTE function
  df_train$Y = factor(df_train$Y)
  # Augment the data according to two oversampling rates
  df_smote1 = SMOTE(Y ~ .,
                    df train,
                    perc.over = get_perc_over(df_train$Y, 0.1),
                    perc.under = 1800,
                    k = 5)
  df_smote2 = SMOTE(Y ~ .,
                    df_train,
                    perc.over = get_perc_over(df_train$Y, 0.5),
                    perc.under = 106,
                    k = 5)
  # Convert from factor back to numbers
  df_{motel} = (df_{motel} = 1)*1
  df_{mote2}Y = (df_{mote2}Y == 1)*1
  \# Calculate the metrics for the two SMOTE datasets
  res_smote1[i, ] = compute_metrics(df_smote1, df_test)
 res_smote2[i, ] = compute_metrics(df_smote2, df_test)
```

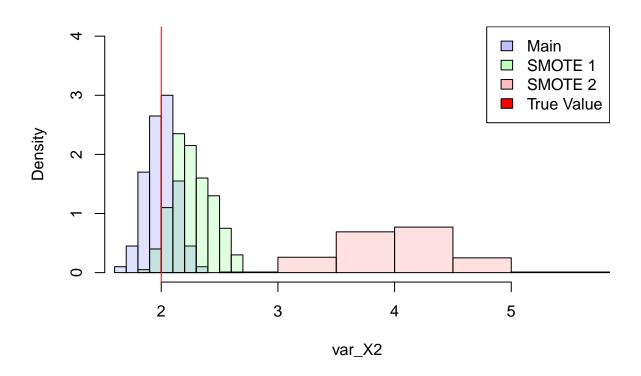
Results: Histograms

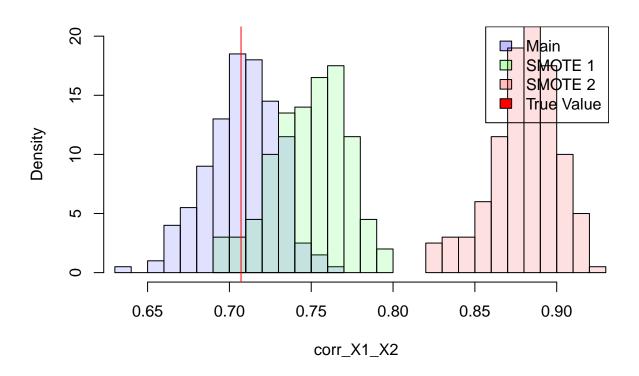
```
# Plot histograms of all variables according with their real values (if
# those exist.)
theoritical_values = c(1, 1, 1, 2, 1 / sqrt(2), beta[1], beta[2], beta[3], NA)
ylims = c(13, 8, 5, 4, 20, 0.6, 2, 1.4, 5)
for (i in 1:9){
  col_name = names(res_base)[i]
  # uncomment this next line to save the figure as png
  #pnq(file=paste(c("hist_", col_name, ".png"), collapse = ""),width=600, height=350)
 mn1 = min(res base[, col name], na.rm = T)
 mn2 = min(res_smote1[, col_name], na.rm = T)
  mn3 = min(res_smote2[, col_name], na.rm = T)
 mn = min(mn1, mn2, mn3)
 mx1 = max(res_base[, col_name], na.rm = T)
  mx2 = max(res_smote1[, col_name], na.rm = T)
 mx3 = max(res_smote2[, col_name], na.rm = T)
 mx = max(mx1, mx2, mx3)
 hist(res_base[, col_name],
       main = " ",
       xlab = col_name,
       col = rgb(0,0,1,1/8),
       xlim = c(mn, mx),
      ylim = c(0, ylims[i]),
      freq = F)
 hist(res_smote1[, col_name],
      xlab = col_name,
       add = T,
       col = rgb(0,1,0,1/8),
       freq = F)
  hist(res_smote2[, col_name],
       xlab = col_name,
       add = T,
       col = rgb(1,0,0,1/8),
       freq = F)
  if (!is.na(theoritical values[i])){
   abline(v = theoritical_values[i], col = "red")
   legend("topright",
         legend=c("Main", "SMOTE 1", "SMOTE 2", "True Value"),
         fill = c(rgb(0, 0, 1, 1/4),
                  rgb(0, 1, 0, 1/4),
                  rgb(1, 0, 0, 1/4),
                  "red"))
 }
```

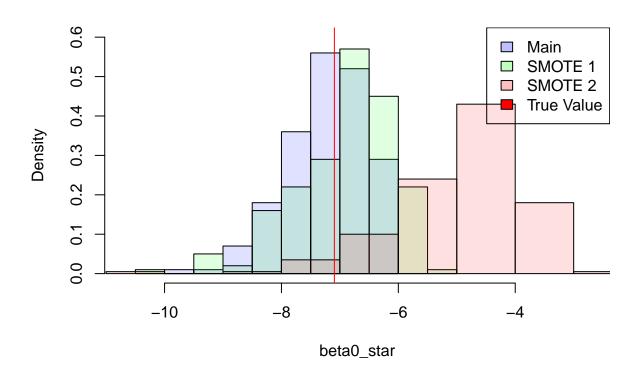


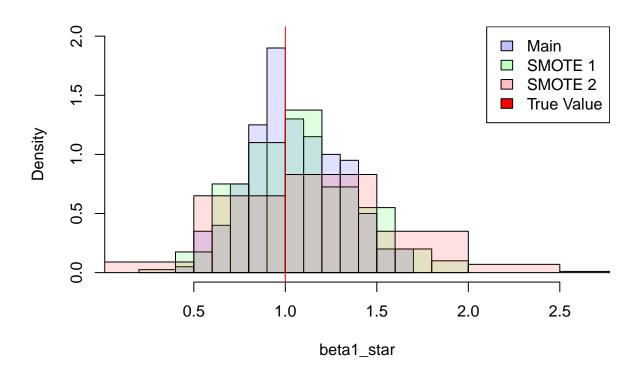


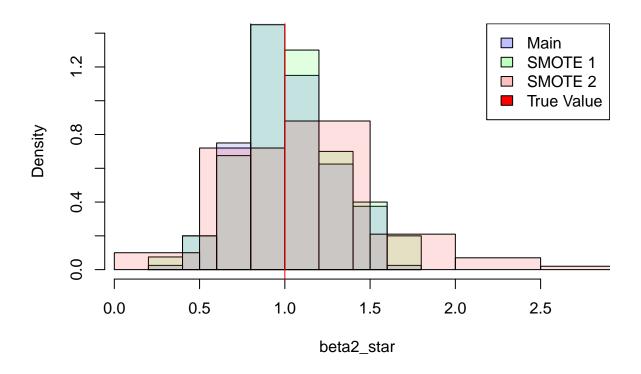


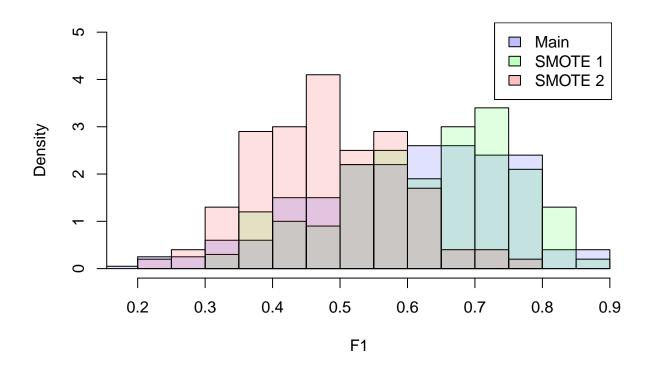












Results: Table

```
# Print a table of the results
get_mean = function(df){
 return(list(
   mean(df$mean X1),
   mean(df$mean_X2),
   mean(df$var_X1),
   mean(df$var_X2),
   mean(df$corr_X1_X2),
   mean(df$beta0_star),
   mean(df$beta1_star),
   mean(df$beta2_star),
   mean(df$F1)))
}
res_comparision = get_empty_df(3)
res_comparision[1, ] = get_mean(res_base)
res_comparision[2, ] = get_mean(res_smote1)
res_comparision[3, ] = get_mean(res_smote2)
res_comparision
##
      mean_X1
                         var_X1 var_X2 corr_X1_X2 beta0_star beta1_star
               mean_X2
## 1 1.000229 0.9958496 1.001767 1.999820 0.7077176 -7.238039 1.049425
## 2 1.085639 1.0986097 1.218497 2.268150 0.7492140 -6.914132 1.091067
## 3 2.129992 2.3523018 2.806397 4.010298 0.8800853 -4.904668 1.173547
   beta2_star
                     F1
## 1 0.9966389 0.6042718
## 2 1.0407546 0.6334307
## 3 1.1293139 0.4849455
```

Results: Hypothesis testing for difference in Beta Estimates and F1 scores

Now, we can conduct two ANOVA tests, to see if the difference in smote level as seen in the histograms is significant at $\alpha = 0.05$.

```
# conduct an anova test on the beta estimates
anova1 <- aov(beta1 ~ smote_level, data = df_params)
summary(anova1)</pre>
```

```
## Df Sum Sq Mean Sq F value Pr(>F)
## smote_level 2 1.60 0.7981 6.85 0.00114 **
## Residuals 597 69.56 0.1165
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

Testing for β_1 , we get a test statistic of $F_{2,597} = 6.85$ which corresponds to a p-value of 0.001. Thus, altering the oversampling rate does have a significant effect on β_1 estimates.

```
anova2 <- aov(beta2 ~ smote_level, data = df_params)
summary(anova2)</pre>
```

```
## Df Sum Sq Mean Sq F value Pr(>F)
## smote_level 2 1.83 0.9131 8.401 0.000252 ***
## Residuals 597 64.89 0.1087
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

Testing for β_2 , we get a test statistic of $F_{2,597} = 8.401$ which corresponds to a p-value of 0.00025. Thus, altering the oversampling rate does have a significant effect on β_2 estimates.

Focusing in on both smote options explored, with final ratios of 0.1 and 0.5, we can conduct a t-test on the F1 scores of each resulting model using Fisher's strict null hypothesis.

```
t.test(df_params$F1[df_params$smote_level == 0.1], df_params$F1[df_params$smote_level == 0.5])
```

```
##
## Welch Two Sample t-test
##
## data: df_params$F1[df_params$smote_level == 0.1] and df_params$F1[df_params$smote_level == 0.5]
## t = 12.588, df = 380.72, p-value < 2.2e-16
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## 0.1252915 0.1716788
## sample estimates:
## mean of x mean of y
## 0.6334307 0.4849455</pre>
```

Here, we get a test statistic of $t_{380} = 12.588$ which corresponds to a p-value of less than 2.2e-16, a highly significant result. Therefore, the F1 score of the less aggressive smote model with a 0.1 positive rate is significantly higher than that of the 0.5 positive oversampling smote.

We can also perform this test to compare between the original data and the smote model with a level of 0.1.

```
{\tt t.test(df\_params\$F1[df\_params\$smote\_level == 0.1], df\_params\$F1[df\_params\$smote\_level == 0])}
```

```
##
## Welch Two Sample t-test
##
## data: df_params$F1[df_params$smote_level == 0.1] and df_params$F1[df_params$smote_level == 0]
## t = 2.1585, df = 395.77, p-value = 0.03149
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## 0.002600593 0.055717275
## sample estimates:
## mean of x mean of y
## 0.6334307 0.6042718
```

While the F1 score on these two datasets is close, we can see that using smote with level 0.1 did significantly improve the F1 score.

Results: Boxplots

