

# The SEND Model

Mastodon C

December 20, 2018

## Contents

<b>1</b>	<b>SEND Population Model</b>	<b>1</b>
1.1	Model Entities, Terminology, and Prediction . . . . .	1
1.2	Visualising and Initialising System State . . . . .	2
<b>2</b>	<b>SEND Model's Population Prediction Algorithm</b>	<b>4</b>
2.1	Movers, Remainers, Leavers and Joiners . . . . .	4
2.2	Leavers . . . . .	5
2.3	Joiners . . . . .	6
2.4	Movers and Remainers . . . . .	8
2.5	The Full Algorithm . . . . .	9
<b>3</b>	<b>A Case Study</b>	<b>10</b>
<b>4</b>	<b>Other Things To Do</b>	<b>12</b>
<b>A</b>	<b>To MC or not to MC</b>	<b>12</b>
A.1	Markov Chain . . . . .	13
A.2	Markov Transition Matrix . . . . .	13
A.3	SEND is not MC . . . . .	14

## 1 SEND Population Model

### 1.1 Model Entities, Terminology, and Prediction

From Simulation, Modeling and Analysis [1, ch. 1,p. 3]:

“The state of a system is defined to be the collection of variables necessary to describe a system at a particular point in time, relative to the objective of a study.”

and

“A system is defined to be the collection of entities that act and interact towards the accomplishment of some goal.”

In the SEND Model the collection of entities that make up the system are drawn from  $S \times N \times AY$  where

$$S = \{MMSIB, ISS, ISSR, \dots, MU\}$$

$$N = \{CL, \dots, NONSEND\}$$

$$AY = \{4, \dots, 24\}$$

For a given local authority,  $S$  represents the set of valid types of school setting,  $N$  represents the set of valid need types, and  $AY$  is the set of allowable academic years in SEND. In the case of our model, an ‘entity’ is a unique combination of setting/need/academic year, and we are interested in knowing the ‘population’ (number of pupils) in that entity.

Academic Year is a somewhat overloaded term: however in this document, for historic reasons, it is used to represent the National Curriculum Year.

Each entity has a single state variable  $t$  which represents its total population - it will always be a non-negative integer.

The points in time for which we can examine an entity’s state is per calendar year.

The SEND simulation predicts the population of each of the system entities from one calendar year to the next using the following:

- The current entities population.
- Predicted future pupil populations at an Academic Year level (not at an entity level i.e., ONS data).
- Probability distributions generated from historical data.

One further note on the terminology; ‘state’ is a contextual word whose scope changes to encompass our current level of thinking e.g., the ‘system state’ is the collection of all entities with their particular state variables at a point in time. Whereas an entity’s ‘state’ is the collection of its state variables only - or in our model’s case, as there is only one state variable, it’s simply the population of the entity. Further, ‘state’ has been used in the code-base to represent a need-setting pair. In this document we will endeavour to always qualify ‘state’ for clarity.

## 1.2 Visualising and Initialising System State

By visualising the mathematical structures that govern system state we can gain an intuitive understanding of how the model works. In Fig.1,  $t_{s_x, n_y, ay_z}$  represents the total population of some setting  $s_x$ , need  $n_y$  and academic year

$ay_z$ . Where  $x, y, z$  are any integer index into the respective ordered sets:  $S, N, AY$ .

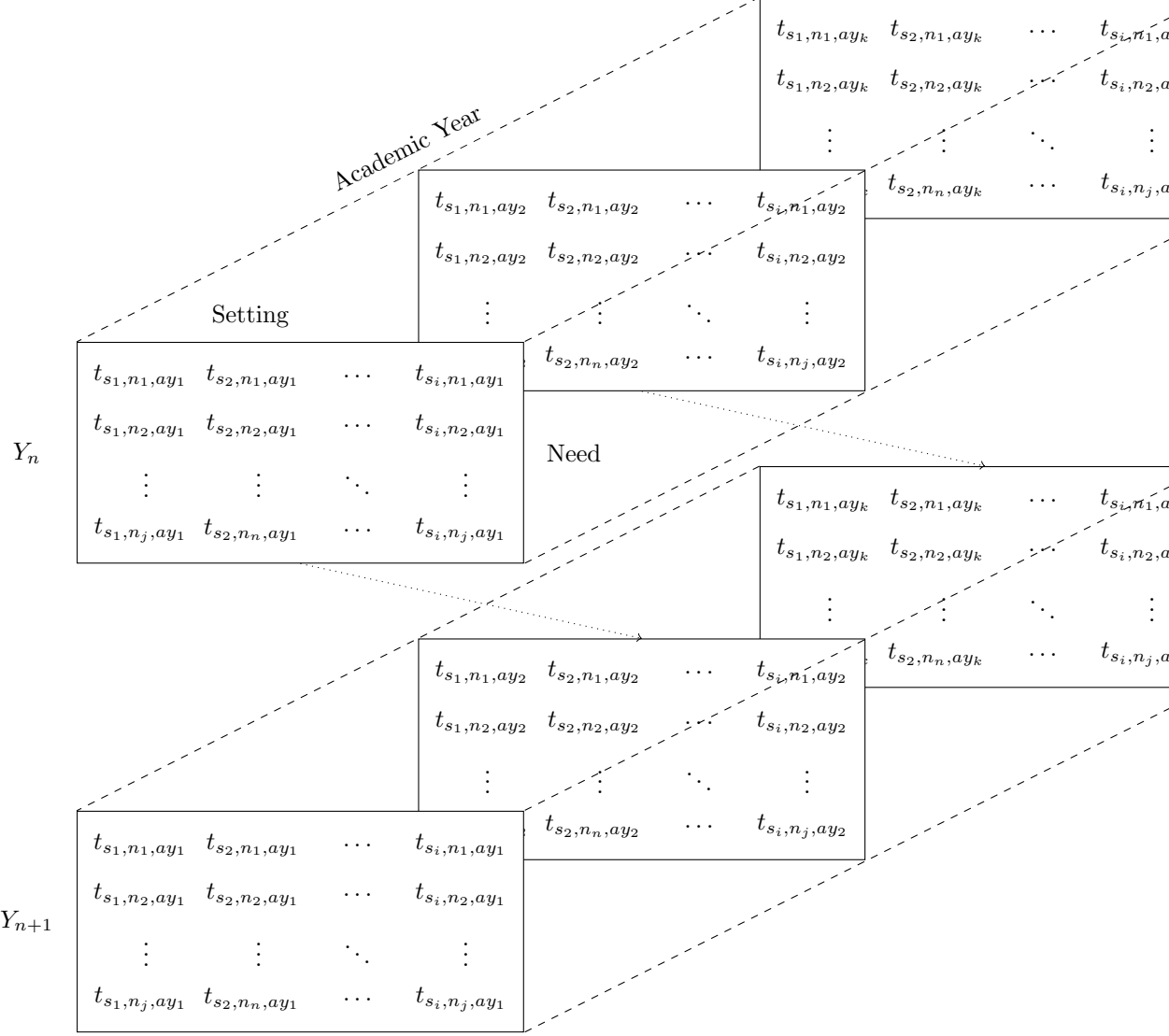


Figure 1: System state from one calendar year to the next

The initial system state, is constructed from the historic data provided by the client (we simply count the number of pupils that, for example, match  $(s_1, n_1, ay_1)$ , in a calendar year).

In Fig.1 you can think of  $Y_n$  as a 3-dimensional matrix of population values, representing the pupil populations of the set of unique entities in year  $n$ .  $Y_{n+1}$

is the next system state for the following calendar year.

## 2 SEND Model's Population Prediction Algorithm

The crux of the SEND Model's prediction algorithm is to take us from  $Y_n$  to  $Y_{n+1}$  in Fig.1. From 1.1 we know in addition to our starting state we also use external pupil population predictions and historic data to generate probability distributions. Thus:

$$Y_{n+1} = T(Y_n, P_{ext}, H)$$

$Y_n$  represents calendar year  $n$ 's system state.

$P_{ext}$  represents the external pupil population predictions.

$H$  represents the historic data.

$T$  is the function that calculates the next years system state.

To simplify our scope when trying to understand  $T$  let us consider first a single entity only. How will that entity's population  $t_{sx,ny,ayz}$  be used to predict the following years? How are  $P_{ext}$  and  $H$  used to support this?

In all the following sections except where noted it is assumed that the underlying processes governing leavers, joiners, and movers are independent and identically distributed meaning that we can combine historic calendar years data,  $H$ , as one.

### 2.1 Movers, Remainers, Leavers and Joiners

First we must understand the model's concepts of movers( $m$ ), remainers( $r$ ), leavers( $l$ ), and joiners( $j$ ).

Let  $m_{sx,ny,ayz}$  be defined as all the pupils that move **into** a particular entity from another entity that has a different need or setting. We deliberately exclude academic year from the definition as if only this changes then we consider the move to be part of the remainers.

Let  $r_{sx,ny,ayz}$  be defined as all the pupils that move out of a particular entity and into another entity with the same need and setting but different academic year.

Let  $l_{sx,ny,ayz}$  be defined as all the pupils that **leave** a particular entity and out of the SEND programme completely.

Let  $j_{sx,ny,ayz}$  be defined as all the pupils that join **into** a particular entity from the outside populace.

So  $T(Y_n, P_{ext}, H)_{s_x, n_y, ay_z} = m_{s_x, n_y, ay_z} + r_{s_x, n_y, ay_z} - l_{s_x, n_y, ay_z} + j_{s_x, n_y, ay_z}$

Reading this equation aloud makes obvious intuitive sense. The question now becomes how do we calculate each of the components of this function?

## 2.2 Leavers

From our historic data we can work out the percentages over the calendar years of how many pupils leave a particular entity,  $(s_1, n_1, ay_1)$ . Averaging this sets our expectation e.g. for  $(s_1, n_1, ay_1)$  we may expect 40% of the population to leave.

For the purposes of our simulation we do not want every run to have exactly 40% of the population leave but rather have some variance in the size. How these variances in the population propagate through the simulation lets us see confidence bounds when we come to aggregate the results of many runs of the model.

A first pass at introducing variance is to use the binomial distribution to alter the population that leaves per run.

Continuing our example given  $t_{s_1, n_1, ay_1} = 100$  and an expectation of 40% then a sample of ten values looks like so:

---

```
(sample 10 (binomial { :n 100 :p 0.4 } ))
(46 42 31 45 43 44 48 41 41 40)
```

---

We can further refine the introduction of variance by noting that the expectation of the percentage of leavers is stronger or weaker (more or less confident in) depending on the amount of data in  $H$  we have. Intuitively if we have lots of observations around the number of leavers then we can be much more confident in our expected value. If we have very few observations then we are much less confident its true underlying value.

The beta distribution can be used to express the variance in our expected leaver percentage (40% or 0.4 in the above example). The parameters for the distribution  $\alpha, \beta$  are simply the observed leavers and non-leavers respectively. Drawing from the beta distribution is equivalent to working out the percentage expectation (obviously slightly changed as it's purpose is to introduce variance in this value). For low values of observations there will be a wide range in the values of the expected leavers and for high values of observations there will be a much narrower range of expected values.

If we use the results of the Beta distribution as the probability parameter of the Binomial distribution (composition) we form what's called a hierarchical model. This chaining of distributions increases the size of variance in our model runs.

Thus we can write:

$$l_{s_x, n_y, ay_z} = \text{Binomial}(t_{s_x, n_y, ay_z}, \text{Beta}(\alpha_{s_x, n_y, ay_z}^l, \beta_{s_x, n_y, ay_z}^l))$$

where  $\alpha_{s_x, n_y, ay_z}^l$  represents the count of leavers over all historic calendar years for the entity  $(s_x, n_y, ay_z)$ . Similarly  $\beta_{s_x, n_y, ay_z}^l$  represents the count of all non leavers for the same entity over all historic calendar years.  $(t_{s_x, n_y, ay_z}$  represents the population of the entity.

To recap, the beta distribution using the historical leavers and non-leavers is drawn from once per calculation to find the probability someone leaves. Then using this probability and the population of the entity we use the binomial distribution to draw a value for pupils that leave. This is a common pattern throughout the model and is important to understand well.

## 2.3 Joiners

The calculation for working out an entity's joiners is split into two parts: first the expected rate of joiners is calculated per academic year, then this additional population is assigned a need and setting.

The sum of the joiners for a single academic year over all historic calendar years is first calculated. As for leavers we wish calculate the probability of joining SEND - we know that if we have the non-leavers we can use the beta distribution to draw a probability value of joining.

We work out the number of non-joiners by taking the joiners away from the predicted number of pupils for an AY across the entire school system: both send and non-send. This external data source,  $P_{ext}$ , might, for example, be the ONS data. Typically  $P_{ext}$  is only available at an AY level. Our desire to use it means that we can't work at the normal  $s_x, n_y, ay_z$  level - and explains why are considering joiners over an entire academic year rather than an entity.

We once more use the binomial distribution to provide variance in our population values once we have an expected probability from the beta distribution. Thus:

$$j_{ay_z} = \text{Binomial}(p_{ay_z}, \text{Beta}(\frac{\alpha_{ay_z}^j}{cy_{obs}}, \frac{p_{ay_z}^H - \alpha_{ay_z}^j}{cy_{obs}}))$$

where  $\alpha_{ay_z}^j$  is the total joiners for  $ay_z$  (from the historic calendar years),  $cy_{obs}$  is the number of observed calendar years,  $p_{ay_z}$  is the projected population from our additional population source for  $ay_z$ , and  $p_{ay_z}^H$  is the total population count over the historic calendar years.

You may notice an oddity in this equation, in our code-base we divide through the sum of the joiners and non-joiners by observed calendar years. This has the effect of increasing variance. It is in direct contrast to the other probability

distribution calculations that assume independent and identical distributions. Why this is done is in under-investigation.

The second part of the calculation is to distribute the new joiners to appropriate needs and settings. We count for each academic year exactly how many joiners there are from non-SEND to each of the need-settings. The proportions of which tell us how to distribute the joiners. Note: this time we don't divide through by the academic years. Armed with these proportions we can calculate the expectation as a percentage for each e.g 10% to  $(s_x, n_y, ay_z)$  and 90% to  $(s_j, n_k, ay_l)$ .

Again, rather than a direct calculation of the population to be assigned to each entity we introduce variance by using the multinomial distribution. It performs exactly the same job as the binomial except over multiple possible outcomes.

In a now familiar pattern, we also wish to introduce variance in the actual percentages used to calculate the assignment. Like before, we want high confidence in values we have lots of observed data for and low confidence for few observations. The Dirichlet distribution does this for multivariates in exactly the same way beta distribution does for two. We combine the two to produce another hierarchical model.

$$\text{Multinomial}(j_{ay_z}, \text{Dirichlet}(\alpha_{s_x, n_y, ay_z}^j, \dots, \alpha_{s_{x'''}, n_{y''}, ay_{z''}}^{j''}))$$

Here the  $\alpha^j$ 's represent the total number of transitions from *nonSEND* to  $s_x, n_y, ay_z$ .

The results of the multinomial function is however a vector of the counts of the joiners for each of the possible entities. Thus for any particular entity of interest we need to take the appropriate index.

$$j_{s_x, n_y, ay_z} = \left[ \text{Multinomial}(j_{ay_z}, \text{Dirichlet}(\alpha_{s_x, n_y, ay_z}^j, \alpha_{s_{x'}, n_{y'}, ay_{z'}}^{j'}, \dots, \alpha_{s_{x''}, n_{y''}, ay_{z''}}^{j''})) \right]_{s_x, n_y, ay_z}$$

We can write this out in full for single entity by expanding  $j_{ay_z}$ .

$$j_{s_x, n_y, ay_z} = \left[ \text{Multinomial}(\text{Binomial}(p_{ay_z}, \text{Beta}(\frac{\alpha_{ay_z}^j}{cy_{obs}}, \frac{p_{ay}^H - \alpha_{ay_z}^j}{cy_{obs}}), \text{Dirichlet}(\alpha_{s_x, n_y, ay_z}^j, \dots, \alpha_{s_{x''}, n_{y''}, ay_{z''}}^{j''}))) \right]_{s_x, n_y, ay_z}$$

Interestingly, we can immediately see that the number of joiners to an entity is not directly related to that entity's population state, rather its relation is to the projected population for an academic year and the non-SEND to entity transitions in the observed data.

## 2.4 Movers and Remainers

The core form of the calculation for movers should be by now familiar.

$$m_{s_x, n_y, ay_z}^{out} = \text{Multinomial}(\text{Binomial}(t_{s_x, n_y, ay_z} - l_{s_x, n_y, ay_z}, \text{Beta}(\alpha_{s_x, n_y, ay_z}^m, \beta_{s_x, n_y, ay_z}^m)), \\ \text{Dirichlet}(\alpha_{s'_x, n'_y, ay'_z}^{m'}, \dots, \alpha_{s''_x, n''_y, ay''_z}^{m''}))$$

The superscript *out* represents the fact this equation actually provides a vector of the pupils that leave  $s_x, n_y, ay_z$  not move into it. How we deal with that we'll come to later.

Like leavers the relationship is based on the current entity's population count (not the academic year's population count). However it is slightly modified in that we take the previously calculated leavers away for this entity (as we know they move to non-SEND).

The Beta params  $\alpha^m$  and  $\beta^m$  are the counts of the transitions out-of- $s_x, n_y, ay_z$ -and-into-some-other-setting and the transitions to-itself respectively. Unlike joiners they are not divided through by observed calendar years.

The Dirichlet params  $\alpha^{m'}$ 's are the counts from the observed data of our entity  $s_x, n_y, ay_z$  to the other entities it can move to  $s'_x, n'_y, ay'_z$  (represented by ever increasing dashes). As such the vector that is the result of the multinomial has no count of any movers to our entity  $s_x, n_y, ay_z$ . Instead we need to calculate  $m^{out}$  for every other entity and take from each generated vector the index that represents the population that moves to  $m_{s_x, n_y, ay_z}$ . The sum of these is the total movers into our entity.

$$m_{s_x, n_y, ay_z} = \sum_{j \neq x, k \neq y, l \neq z} \left[ \text{Multinomial}(\text{Binomial}(t_{s_j, n_k, ay_l} - l_{s_j, n_k, ay_l}, \text{Beta}(\alpha_{s_j, n_k, ay_l}^m, \beta_{s_j, n_k, ay_l}^m)), \right. \\ \left. \text{Dirichlet}(\alpha_{s'_j, n'_k, ay'_l}^{m'}, \dots, \alpha_{s''_j, n''_k, ay''_l}^{m''})) \right]_{s_x, n_y, ay_z}$$

where  $j, k, l$  belong to  $S, N, AY$  respectively.

As we know that remainers for an entity move to an entity with the same need setting but in the next academic year we just use the beta binomial to calculate these.



$$r_{s_x, n_y, ay_z} = t_{s_x, n_y, ay_z} - l_{s_x, n_y, ay_z} - \text{Binomial}(t_{s_x, n_y, ay_z} - l_{s_x, n_y, ay_z}, \text{Beta}(\alpha_{s_x, n_y, ay_z}^m, \beta_{s_x, n_y, ay_z}^m))$$

For consistency of the model that Binomial component of the above equation should be the same result as when it was used to calculate the movers.

## 2.5 The Full Algorithm

Calculate next years  $t$  for an entity given the current years state,  $Y_n$ , predicted pupil population,  $P_{ext}$ , and historic data  $H$ .

$$Y_{n+1, s_x, n_y, ay_z} = T(Y_n, P_{ext}, H)_{s_x, n_y, ay_z} = m_{s_x, n_y, ay_z} + r_{s_x, n_y, ay_z} - l_{s_x, n_y, ay_z} + j_{s_x, n_y, ay_z}$$

Where

$$m_{s_x, n_y, ay_z} = \sum_{j \neq x, k \neq y, l \neq z} \left[ \text{Multinomial}(\text{Binomial}(t_{s_j, n_k, ay_l} - l_{s_j, n_k, ay_l}, \text{Beta}(\alpha_{s_j, n_k, ay_l}^m, \beta_{s_j, n_k, ay_l}^m)), \text{Dirichlet}(\alpha_{s'_j, n'_k, ay'_l}^{m'}, \dots, \alpha_{s''_j, n''_k, ay''_l}^{m''})) \right]_{s_x, n_y, ay_z}$$

$$r_{s_x, n_y, ay_z} = t_{s_x, n_y, ay_z} - l_{s_x, n_y, ay_z} - \text{Binomial}(t_{s_x, n_y, ay_z} - l_{s_x, n_y, ay_z}, \text{Beta}(\alpha_{s_x, n_y, ay_z}^m, \beta_{s_x, n_y, ay_z}^m))$$

$$l_{s_x, n_y, ay_z} = \text{Binomial}(t_{s_x, n_y, ay_z}, \text{Beta}(\alpha_{s_x, n_y, ay_z}^l, \beta_{s_x, n_y, ay_z}^l))$$

$$j_{s_x, n_y, ay_z} = \left[ \text{Multinomial}(\text{Binomial}(p_{ay_z}, \text{Beta}(\frac{\alpha_{ay_z}^j}{cy_{obs}}, \frac{p_{ay}^H - \alpha_{ay_z}^j}{cy_{obs}})), \text{Dirichlet}(\alpha_{s_x, n_y, ay_z}^j, \dots, \alpha_{s_x'', n_y'', ay_z''}^{j''})) \right]_{s_x, n_y, ay_z}$$

The  $\alpha$  and  $\beta$  parameters to the Beta and Dirichlet distributions are calculated in advance from the historic data  $H$ .

The joiner calculation requires  $P_{ext}$  to provide the data for the general population.

The entire current state  $Y_n$  is required to work out a single entity's future population: this is because we use the entire AY population in the joiner calculation and we need to sum all movers into our entity from all the other entities.

### 3 A Case Study

We go through the calculation of an entity's next state (via the leaver and joiner calculations) to show how the algorithm works and to embed the meaning of the myriad of parameters.

Let us consider need  $a$ , setting  $b$ , in academic year 1, the calendar year of 2018 - this entity at this time has a population of 100. The algorithm will give us this entity's population for the next calendar year 2019.

$$\begin{aligned} Y_{2019,s_a,n_b,ay_1} &= T(Y_{2018}, P_{ext}, H)_{s_a,n_b,ay_1} \\ &= m_{s_a,n_b,ay_1} + r_{s_a,n_b,ay_1} - l_{s_a,n_b,ay_1} + j_{s_a,n_b,ay_1} \end{aligned}$$

and

$$t_{s_a,n_b,ay_1} = 100$$

First-off the leavers as it's the simplest:

$$l_{s_a,n_b,ay_1} = \text{Binomial}(t_{s_a,n_b,ay_1}, \text{Beta}(\alpha_{s_a,n_b,ay_1}^l, \beta_{s_a,n_b,ay_1}^l))$$

Let's say we already know the expected rate of leavers is 40% then

$$t_{s_a,n_b,ay_1} \times 40\% = 100 \times 40 = 40 \text{ pupils}$$

We want variance in that so we use the binomial distribution to draw some samples

$$\text{Binomial}(t_{s_a,n_b,ay_1}, 0.4) = \text{Binomial}(100, 0.4) = 38$$

Now we also want variance in the 0.4 value, so we use the Beta distribution. Assume that we have historic data,  $H$ , for the last 2 years that has 60 and 20 leavers, and 110 and 90 non-leavers respectively.

$$\text{Beta}(\alpha_{s_a,n_b,ay_1}^l, \beta_{s_a,n_b,ay_1}^l) = \text{Beta}(60 + 20, 110 + 90) = \text{Beta}(80, 200) = 0.38$$

Combining this all together

$$l_{s_a,n_b,ay_1} = \text{Binomial}(100, \text{Beta}(80, 200)) = 43$$

Next let's look at joiners.

$$j_{s_a, n_b, ay_1} = \left[ \text{Multinomial}(\text{Binomial}(p_{ay_1}, \text{Beta}(\frac{\alpha_{ay_1}^j}{cy_{obs}}, \frac{p_{ay}^H - \alpha_{ay_1}^j}{cy_{obs}}), \right. \\ \left. \text{Dirichlet}(\alpha_{s_x, n_y, ay_z}^j, \dots, \alpha_{s_{x''}, n_{y''}, ay_{z''}}^{j''})) \right]_{s_a, n_b, ay_1}$$

Let's assume our external population  $p_{ay_1}$  is predicted to be a 1000 for  $ay_1$  in 2019.

For our two years of historic data let's assume from  $P_{ext}$  the predicted populations are 900 and 1100.

Let's also say that the number of joiners in our historic data is 30 and 20 over all needs and settings for  $ay_1$ .

A first pass at expectation based on this data is:

$$\frac{30+20}{900+1100} = 0.025$$

As before we need would like some variance in this value, so we use the beta distribution again.

$$\text{Beta}(\frac{\alpha_{ay_1}^j}{cy_{obs}}, \frac{p_{ay_1} - \alpha_{ay_1}^j}{cy_{obs}}) = \text{Beta}(\frac{20+30}{2}, \frac{900+1100-20-30}{2}) = \text{Beta}(25, 975) = 0.02501$$

note: ignore the division through by calendar years for now (it doesn't effect expectation only the variance)

We once more use the binomial to get a varied value from the population and the expectation.

$$\text{Binomial}(p_{ay_1}, 0.02501) = \text{Binomial}(1000, 0.02501) = 25$$

Armed with the fact we have 25 joiners for academic year 1 in 2019 our equation reduces to:

$$j_{s_a, n_b, ay_1} = \left[ \text{Multinomial}(25, \text{Dirichlet}(\alpha_{s_a, n_b, ay_1}^j, \dots, \alpha_{s_{x''}, n_{y''}, ay_1}^{j''})) \right]_{s_a, n_b, ay_1}$$

This part deals with assigning those 25 pupils to specific entities. Assuming our historic data for academic year one has one other valid setting need c and setting d - what are the proportions of the population between the two entities?

Let's say ab1 has populations of 20 and 40 over the two years and cd1 has populations of 60 and 80. Then we can say 30% of the population is in ab1 and 70% of the population is in cd1.

We can use these percentages to distribute the 25 joiners appropriately. As before we'd like to vary this and the Dirichlet performs the same job as the Beta in this regard. So,

$$\begin{aligned}\text{Dirichlet}(\alpha_{s_a, n_b, ay_1}^j, \dots, \alpha_{s_{x''}, n_{y''}, ay_1}^{j''}) &= \text{Dirichlet}(20 + 40_{s_a, n_b, ay_1}, 60 + 80_{s_c, n_d, ay_1}) \\ &= \text{Dirichlet}(60_{s_a, n_b, ay_1}, 140_{s_c, n_d, ay_1}) \\ &= [0.29_{s_a, n_b, ay_1}, 0.71_{s_c, n_d, ay_1}]\end{aligned}$$

Again rather than simply calculate the distribution based on these percentages we introduce variance by using the multinomial to vary the results (this plays the same role as the binomial from before). Thus

$$\text{Multinomial}(25, [0.29_{s_a, n_b, ay_1}, 0.71_{s_c, n_d, ay_1}]) = [7_{s_a, n_b, ay_1}, 17_{s_c, n_d, ay_1}]$$

The result is a vector and as we are calculating for only  $s_a, n_b, ay_1$  the last thing we do is pull its index out (the square brackets in the algorithm and its index represent that). Thus we can now answer the question of how many joiners will there be in the next calendar year to  $s_a, n_b, ay_1$ .

$$j_{s_a, n_b, ay_1} = 7$$

These two walk-throughs should provide enough to understand the mechanism for the movers and remainers.

## 4 Other Things To Do

- Priors section for leavers, joiners, movers
- Scenarios
- Valid state implications
- The real simulation tick mechanism

## A To MC or not to MC

This appendix deals with why the SEND model is Markov Chain like but not actually Markov Chain based. Rather it is closer to event simulation.

## A.1 Markov Chain

This section and the next explain how the SEND model would look if it were implemented using just Markov Chains. If we understand this we understand why we can say the SEND Model is “Markov like”.

The initial state of our Markov Chain,  $Y_0$ , is constructed from the historic data provided by the client (we simply count the number of pupils that, for example, match  $(s_1, n_1, ay_1)$ ).

You can think of  $Y_n$  as a 3-dimensional matrix of population values, representing the pupil populations of the set of unique entities in year  $n$ .

In Fig.1  $Y_{n+1}$  is the next state of our Markov Chain.

How we transition from  $Y_n$  to  $Y_{n+1}$  is the subject of the next section. However, it’s worth noting now that the dotted arrows on the diagram represent how the  $t$ ’s of one academic year transition into the next academic year, in the following calendar year.

The complete state of the model is represented as  $S = \{Y_0, \dots, Y_n\}$  where  $n$  is the number of years we run the simulation for.

## A.2 Markov Transition Matrix

In mathematics a Markov Transition Matrix is formally a matrix containing the probability of transitioning between Markov states. Conceptually, for the SEND Model, it can be thought of as telling us what percentages of  $t_{s_x, n_y, ay_z}$  will move to each of the allowed entities  $t_{s_x, n_y, ay_z}$  can move to. Sometimes this is referred to as *rates*.

By looking at the historic data we could work out the expectation of each of our transitions and build this matrix. We visualise what this looks like in Fig.2.

In the Fig.2 every  $P_{s_i, n_j, ak_1}$  is a probability matrix that represents the probabilities of transitioning to any other state in the next academic year (see the diagram’s cut-out for an example.)

Armed with  $P$ , this rather complicated Markov Transition Matrix, we can apply it to the initial state

$$Y_1 = Y_0 \cdot P$$

In fact at this point as  $P$  is time independent we can work the populations in year  $n$  by

$$Y_n = Y_0 \cdot P^n$$

Of note is that this equation requires no tracking of  $t$ ’s we can simply use the above formula to work out  $Y_n$ . BUT we don’t do this! Why?

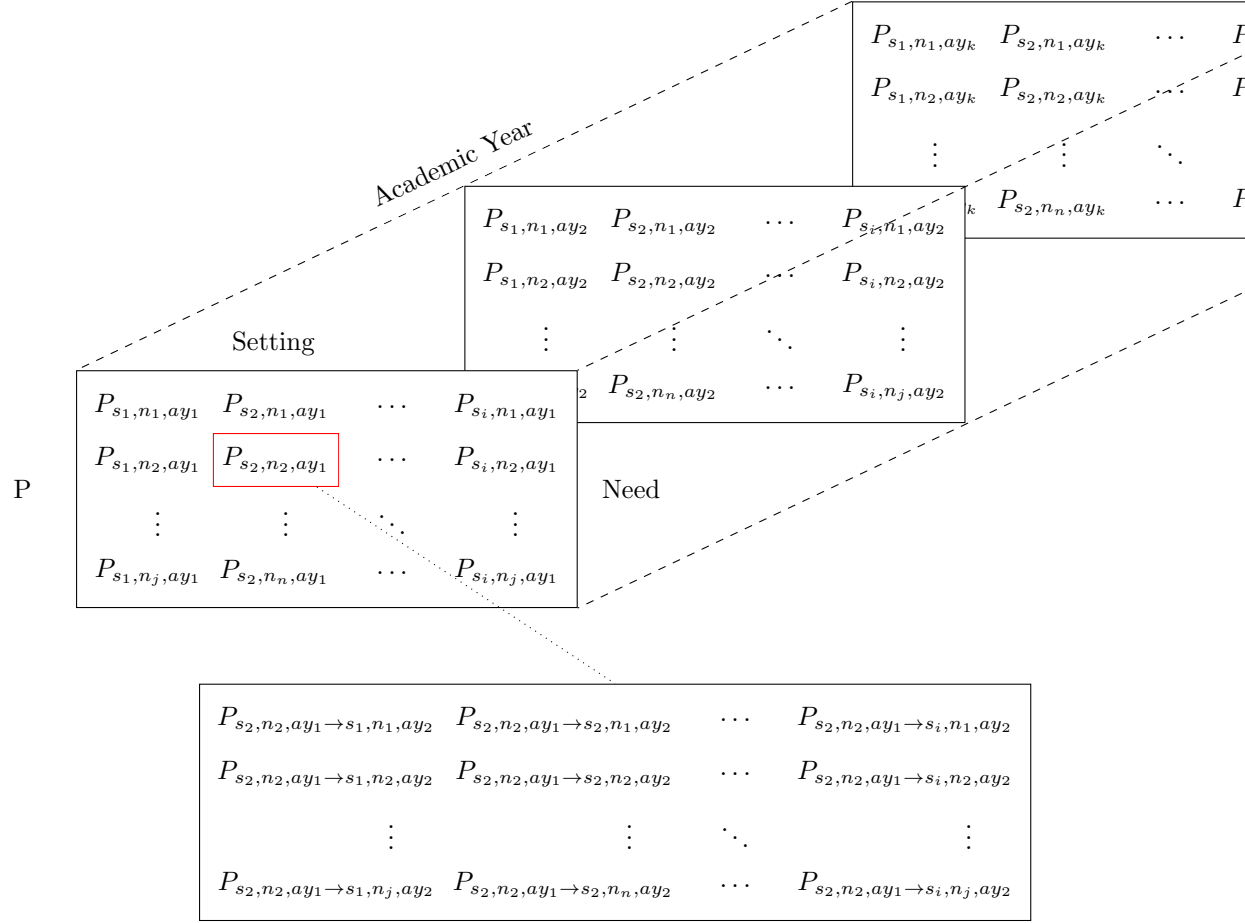


Figure 2: Markov Transition Matrix (many nested probability matrices)

### A.3 SEND is not MC

The SEND model actually sides step the creation of a Markov Transition Matrix and takes an algorithmic approach to calculating  $Y_{n+1}$ 's population.

This is done as some of the data is perhaps too sparse, but also the algorithmic approach allows us to introduce priors and scenarios rather more easily than modifying  $P$  directly.

Further, we also use more than the just the initial state, in our algorithmic predictions. The model incorporates external population data for each predicted calendar year.

As we will come to see the algorithmic approach to working out the  $t$ 's in  $Y_n$

introduces stochastic processes. Thus we introduce the need to use Monte Carlo methods to ensure confidence in our calculations.

The model can be thought of as a simulation with a simulation clock that ticks once a calendar year. On clock tick all the entities have their new population calculated: based on their existing state, pre-calculated historic probability distributions and projected population.

## References

- [1] Averill M. Law, *Simulation Modelling and Analysis*, Addison Wesley, Massachusetts, 4th edition, 2007.