

# PHYS465: Statistical Data Analysis in Physics

*Week 2: Hypothesis Testing, Common  
distributions, Model testing*

Dr. Mathew Smith

[mat.smith@lancaster.ac.uk](mailto:mat.smith@lancaster.ac.uk)

Physics Building; C46

# Module Structure

*'model testing'*: does our data agree with our model?

**Week 11:** Fitting a model to data; estimating parameters

**Week 12:** Hypothesis testing; the likelihood; estimating uncertainties

**Week 13:** Posterior sampling; Bayesian statistics

*'data driven'*: what does our data tell us?

**Week 14:** Clustering and Classification algorithms

**Week 15:** Machine learning techniques

# Week 12: Previous session

*Last week we learnt:*

- Our measurements drawn from an underlying continuous distribution
  - The shape of this distribution is dependent on our experiment
  - The probability of our data can be described through a PDF
- Given a model we can find the best-fit parameters using least-squares fitting
- This is generalised to  $\chi^2$ -fitting in the presence of gaussian uncertainties
- A good model will have  $\chi^2_{red} \sim 1$ 
  - Where  $\chi^2_{red} = \chi^2/n_{dof}$
  - $\chi^2_{red} > 1$  suggests the need for more complicated models;  $\chi^2_{red} < 1$  is the reverse
- Our knowledge of model parameters is expressed through confidence intervals
- Question: how do we measure a confidence interval or parameter uncertainty?

# *Week 12: Learning aims*

*Today we will introduce*

- Probabilities and how they relate to measurements
- The Null Hypothesis
  - And how to test it
  - How our experiment will change how the data is distributed
  - How to test which distribution is correct
  - How do we estimate confidence intervals?
  - How to select which model is most likely

Practical examples on Friday!

# Revision: Probabilities

Given an experiment:

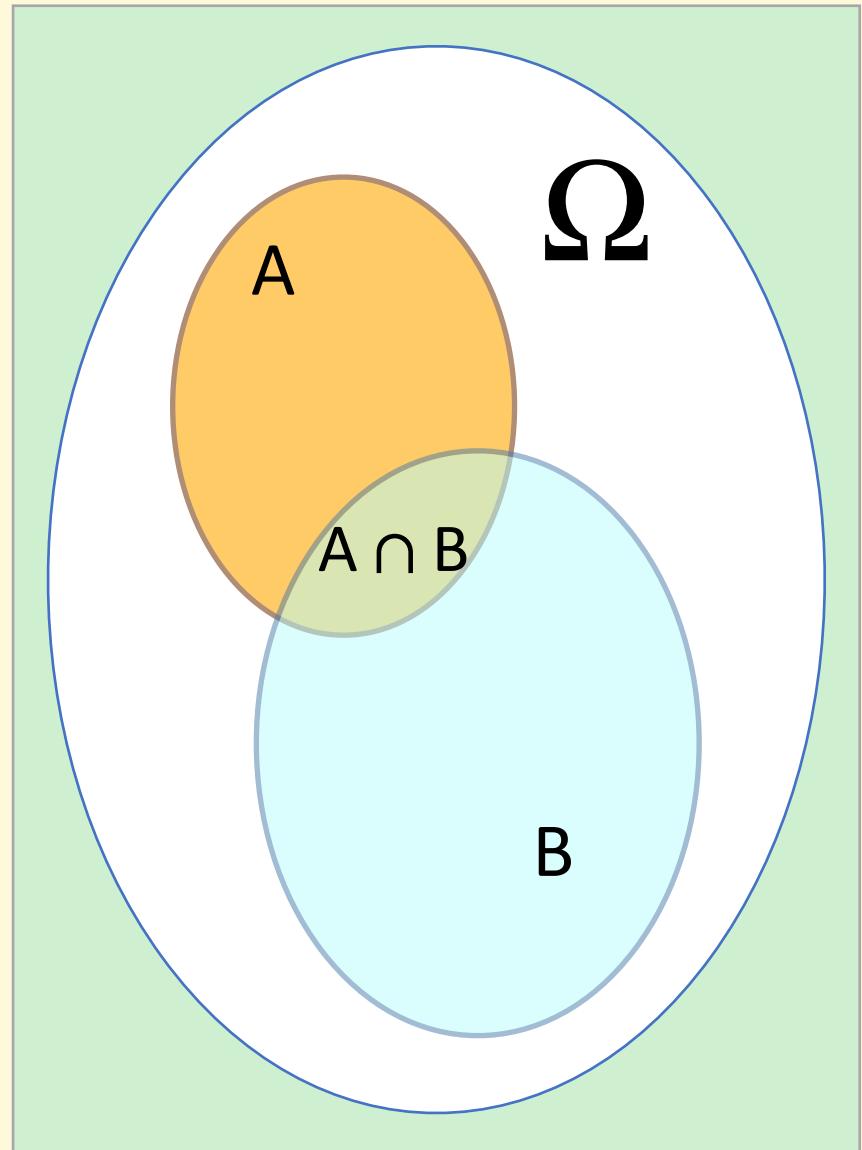
- $\Omega$  is the list of all possible outcomes

Axioms:

- $0 \leq P(A) \leq 1$
- $P(\Omega) = 1$
- $P(A_1 \cup A_2 \cup A_3 \dots) = P(A_1) + P(A_2) + P(A_3) + \dots$

Consequences:

- $P(A) + P(A^c) = 1$
- $P(A \cup B) = P(A) + P(B) - P(A \cap B)$
- $P(A \cap B) = P(A) \times P(B)$  if  $A, B$  are independent



# Conditional Probabilities

$$P(A | B) = \frac{P(A \cap B)}{P(B)}$$

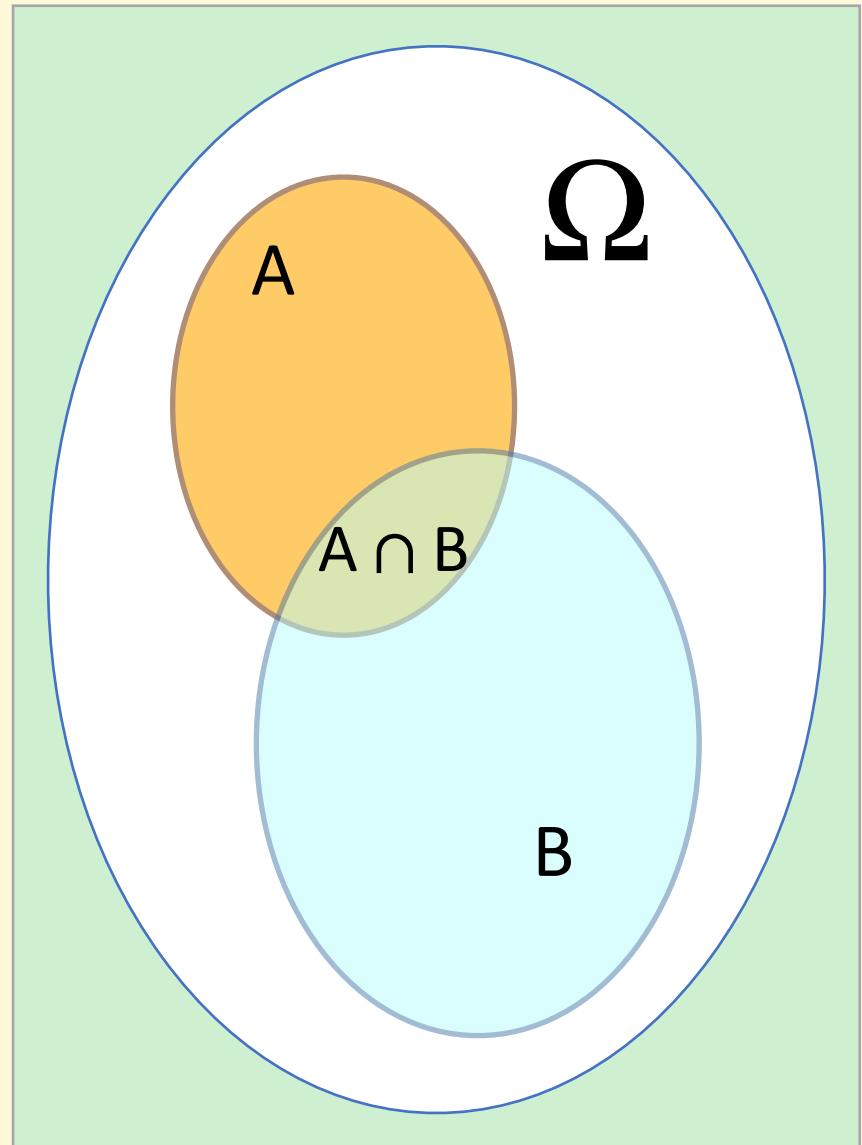
*The probability of A given B (an event or condition)*

when  $\theta$  is a model, or parameter

$P(A | \theta)$  is called the **likelihood**:  $\mathcal{L}(\theta | A)$

“the **likelihood** of obtaining the data ‘A’ given the model”

*Plenty more on this with Bayesian statistics next week*



# Hypothesis testing

How good is the model?

**Fundamental question:** Is the model that we are testing a good fit to our measured data?  
Is there a more likely alternative?

**Null Hypothesis:** Our starting assumption  
Can be either parameter values or choice of model

**In practice:** Compares our assumption (or prediction) with experimental measurements  
e.g. does the luminosity distribution of galaxies match a Schechter function?

**Outcome:** Make a statement on the probability of obtaining our result  
(see confidence level slides).

# Null Hypothesis testing

e.g. Given a model  $f(\theta)$ , with parameter  $\theta$ , is the value of  $\theta$  in a set of possible values  $\Theta_0$ ?

$H_0$  : The Null Hypothesis:  $\theta \in \Theta_0$

$H_1$  : An Alternative Hypothesis:

- $\theta \in \Theta_1$  where  $\Theta_0 \cap \Theta_1 = \emptyset$

$\alpha$  : threshold of rejection (e.g. 0.05)

With new observations  $X$ , such that  $p(X, \theta)$

$p(X) < \alpha$       reject  $H_0$  and accept  $H_1$

(N.B. This does not mean accept another model)

$p(X) > \alpha$       no evidence to reject  $H_0$

(N.B. This is not the same as accepting  $H_0$ )

		Outcome	
		$H_0$ not rejected	$H_0$ rejected
Truth	$H_0$ is true	Probability of this: $1 - \alpha$	Type I error Will happen $\alpha$ % of the time
	$H_1$ is true	Type II error Happens $1 - \beta$ % of the time	$P_{H_1}(H_0 \text{ is rejected}) = \beta$

A good experiment is defined around these requirements 8

# Confidence Intervals

*Given a probability threshold, what is the allowed range of the parameter:*

$$\hat{p} - \delta p \leq P_0 \leq \hat{p} + \delta p$$

**Interpretation:** if we repeat our experiment, how often will  $P_0$  be within our quoted range

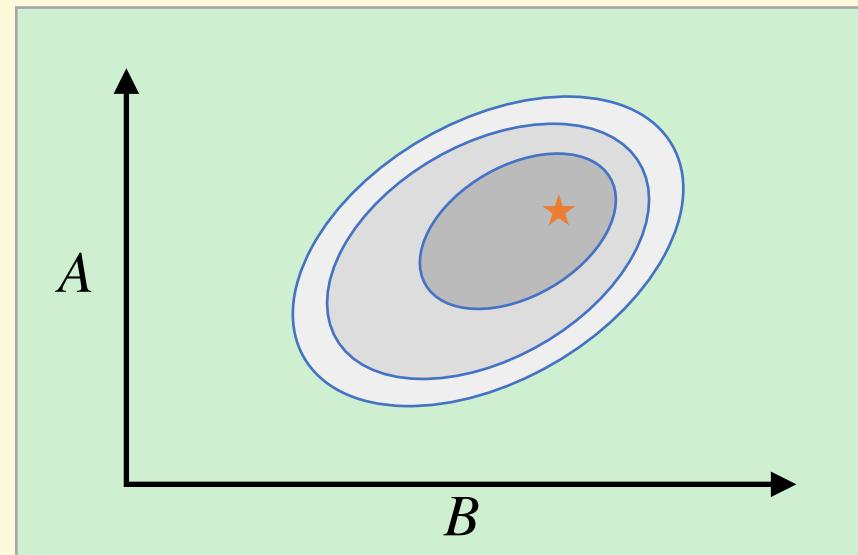
---

**For multiple variables ( $A, B$ ):**

An error ellipse containing  $\delta p$

of the probability

Typical values are 68%, 90%, 95%.



# Key Assumption:

The probability,  $p$ , needed for **Hypothesis testing** and **confidence intervals** depend on how the data and model are distributed

Are they both drawn from a Gaussian distribution?

# Key Distributions: Gaussian

Aka a normal

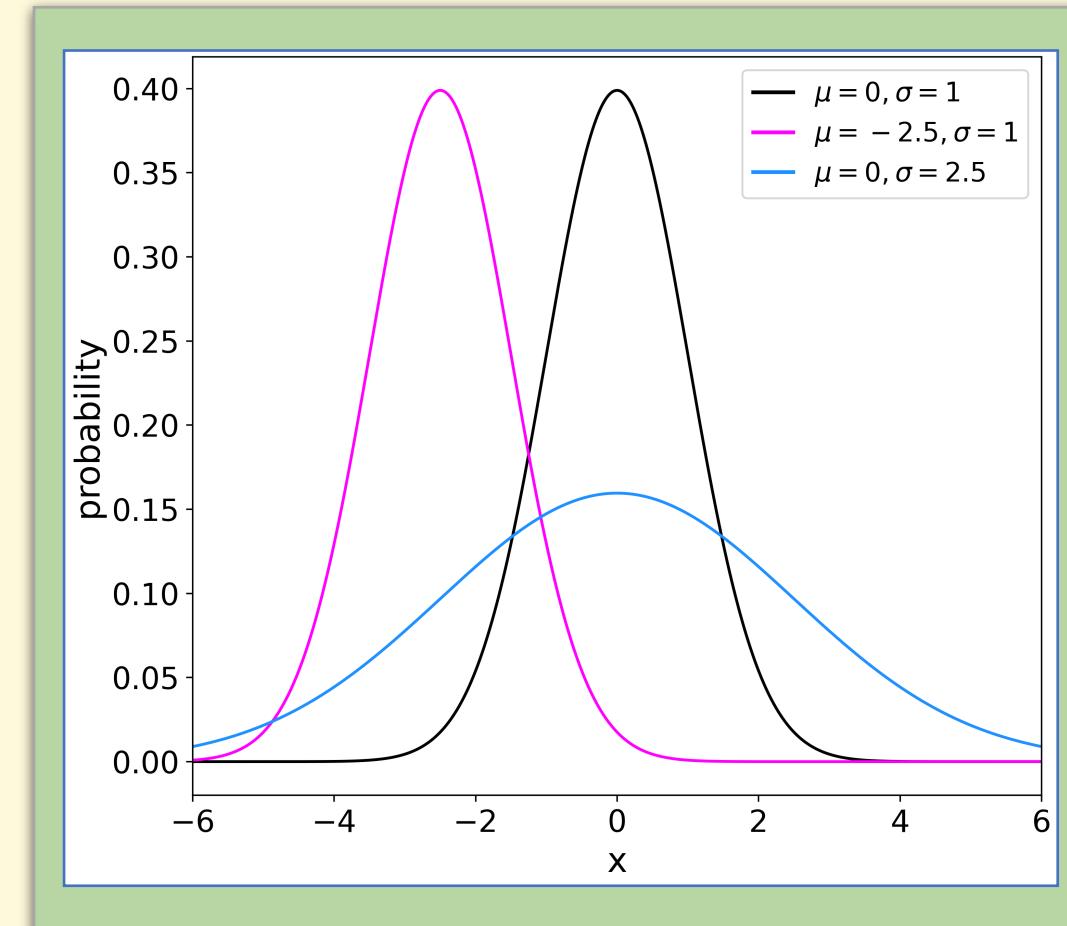
$$P(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\{- (x - \mu)^2 / 2\sigma^2\}$$

The limiting case ( $\rightarrow \infty$ ) of most other distributions

**Expectation value:** mean =  $\mu$

**Variance:**  $\sigma^2 = \sigma^2$

$$P(-t\sigma < x_0 < t\sigma) = \frac{1}{\sqrt{2}} \int_{-t}^t e^{-z^2/2} dz$$



Common assumption across physics

# Key Distributions: Binomial

Fixed number of outcomes

$$P(X = k) = \binom{n}{k} p^k (1 - p)^{n-k}$$

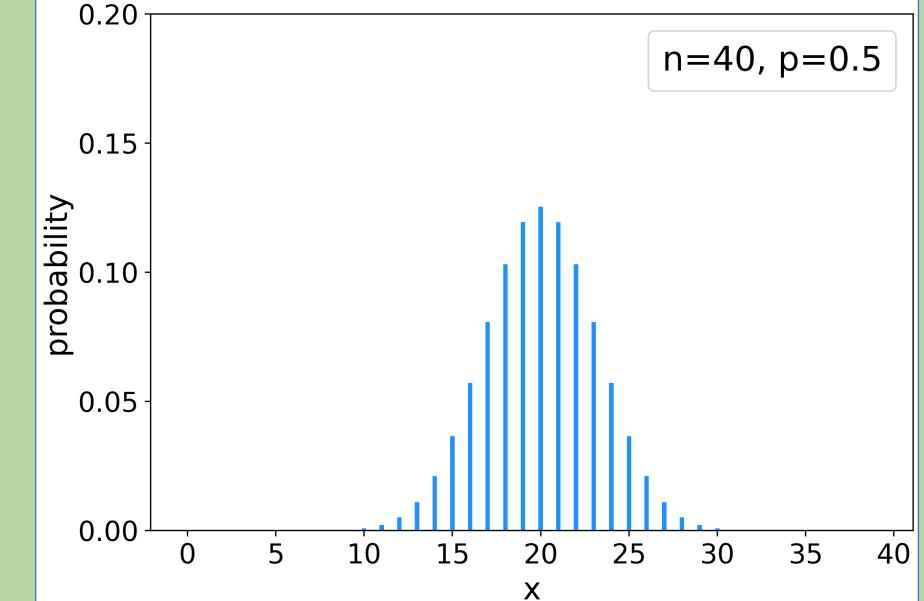
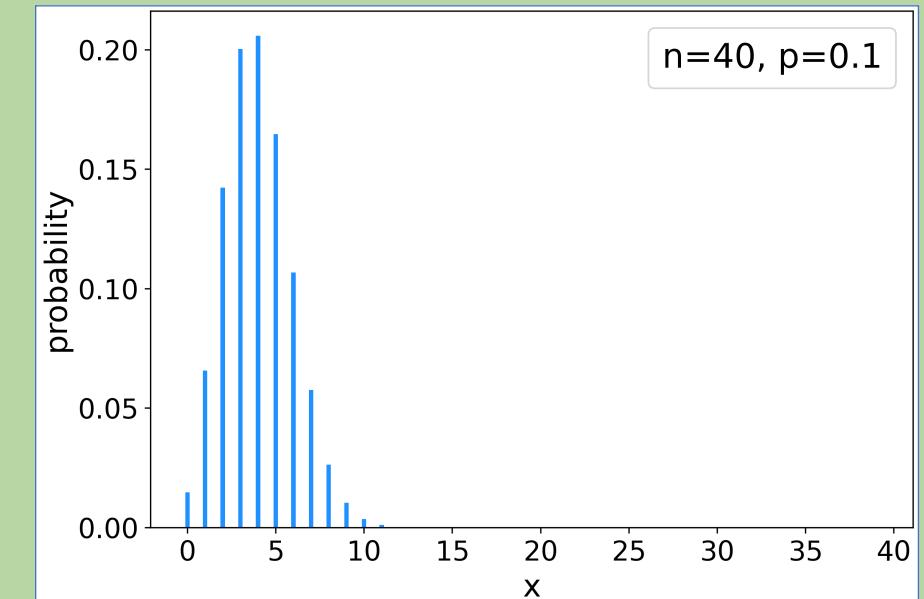
draw  $n$  events from  $k$  possibilities each with probability  $p$

**Expectation value:** mean =  $np$

**Variance:**  $np(1 - p)$

=> **Gaussian distribution as  $N \rightarrow \infty$**

**Example: coin tossing**



# Key Distributions: Poisson

## The counting distribution

$$P(X = k) = \frac{\lambda^k e^{-\lambda}}{k!}$$

observe  $k$  events drawn from  $\lambda$  possibilities

**Expectation value:**

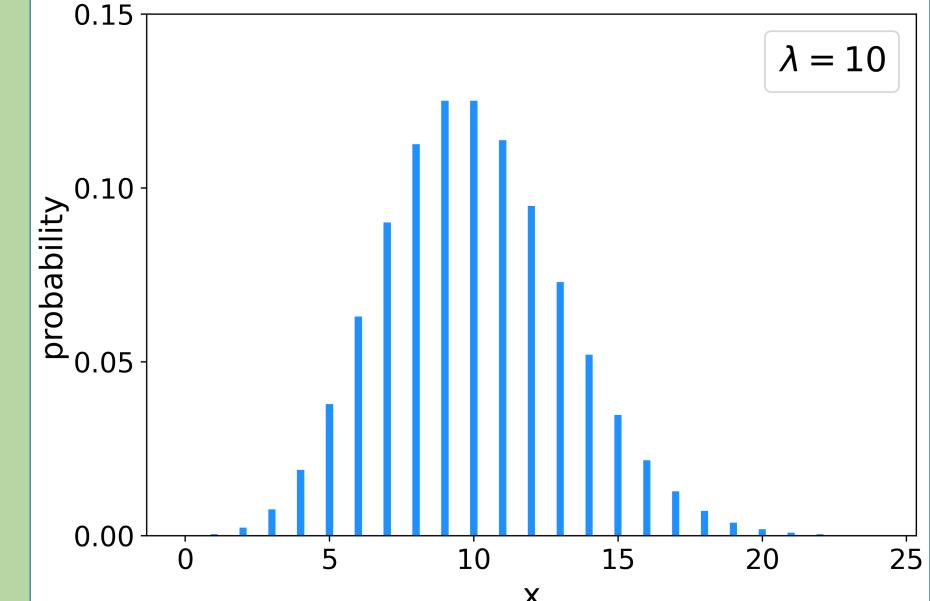
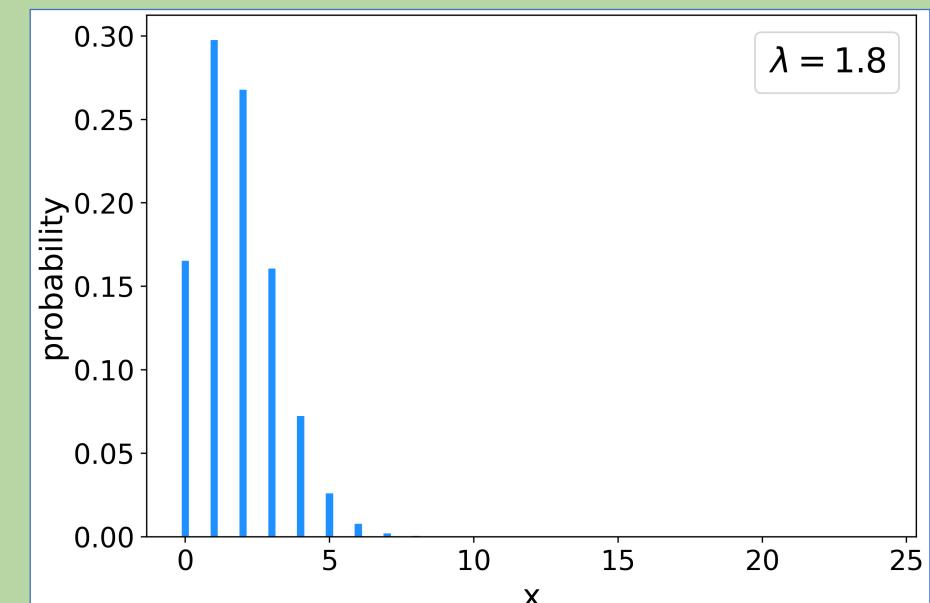
mean =  $\lambda$

**Variance:**

$\sigma^2 = \lambda$

=> Gaussian distribution as  $k \rightarrow \infty$

**Example: histogram (or photon) counting**



# Key Distributions: $\chi^2$

## Related to the Gamma distribution

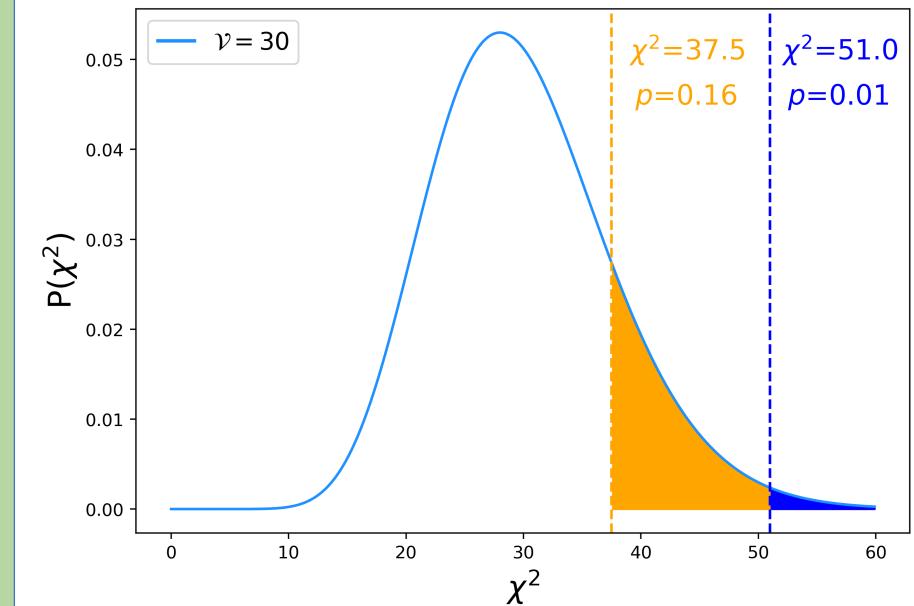
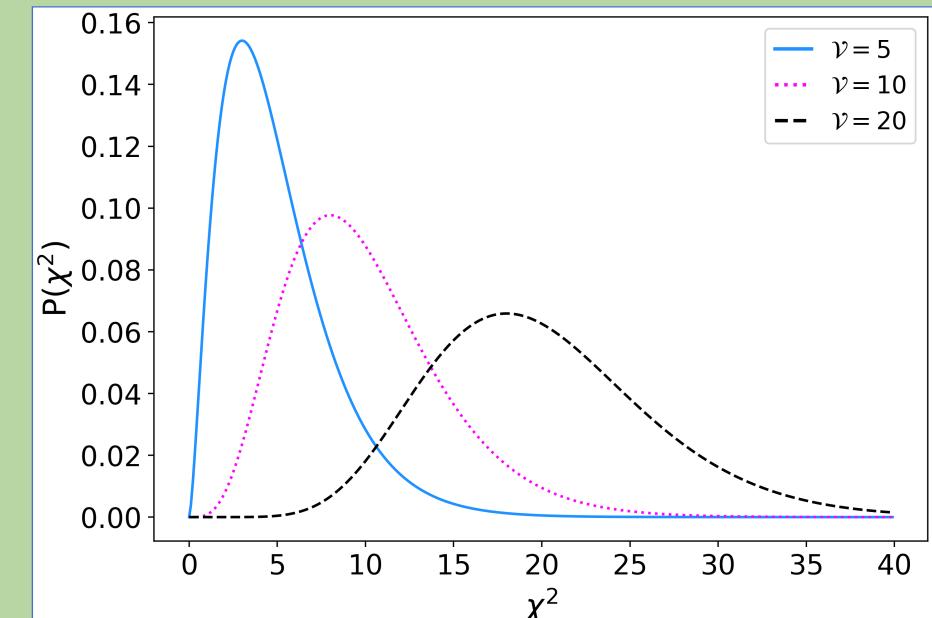
$$P \propto (\chi^2)^{\frac{\nu-2}{2}} \exp(-\chi^2/2)$$

$\nu$  is the number of degrees of freedom

**Expectation value:** mean =  $\nu = N - p$

**Variance:**  $\sigma^2 = 2\nu$

i.e. if the model is correct then we expect  $\chi^2 \sim \nu \pm \sqrt{2\nu}$   
 $\chi^2/\nu \sim 1$



# Gaussian Distribution: probabilities

Recall

$$P(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\{- (x - \mu)^2 / 2\sigma^2\}$$

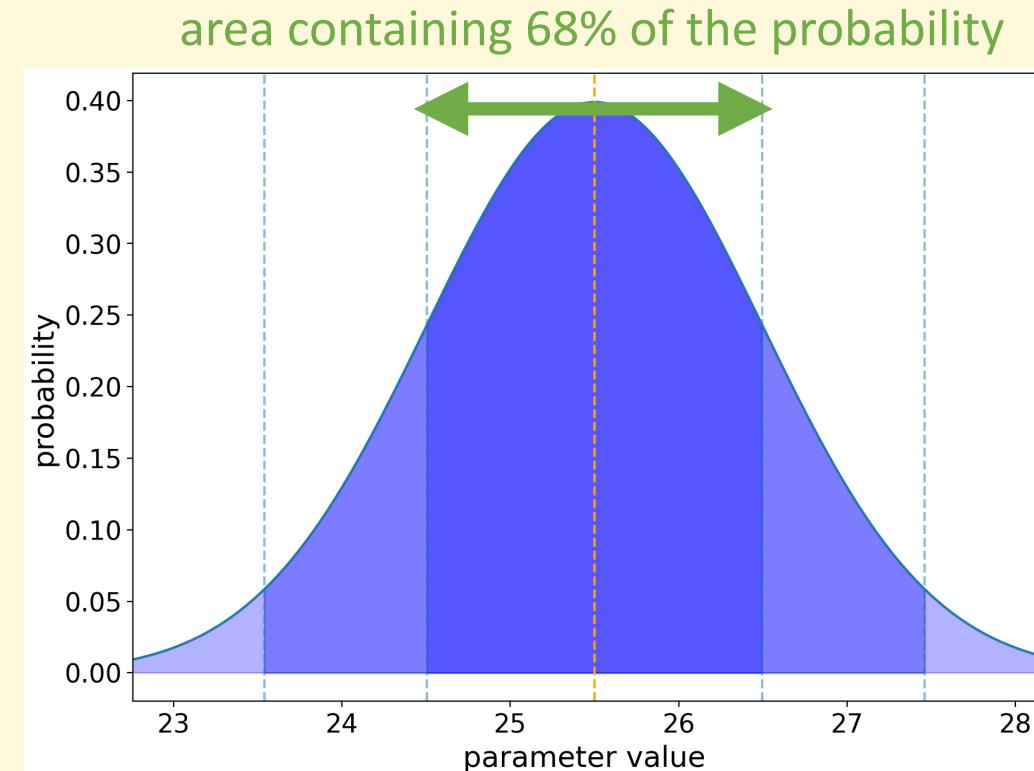
$$P(-t < x_0 < t) = \frac{1}{\sqrt{2}} \int_{-t}^t e^{-z^2/2} dz$$

$1\sigma$  : contains **68.3%** of the probability

$2\sigma$  : contains **95.5%** of the probability

$3\sigma$  : contains **99.7%** of the probability

$5\sigma$  : contains **99.99994%** of the probability



**WARNING:** This process knows nothing about physics: beware of unphysical regions

# Multiple parameters: Marginalisation

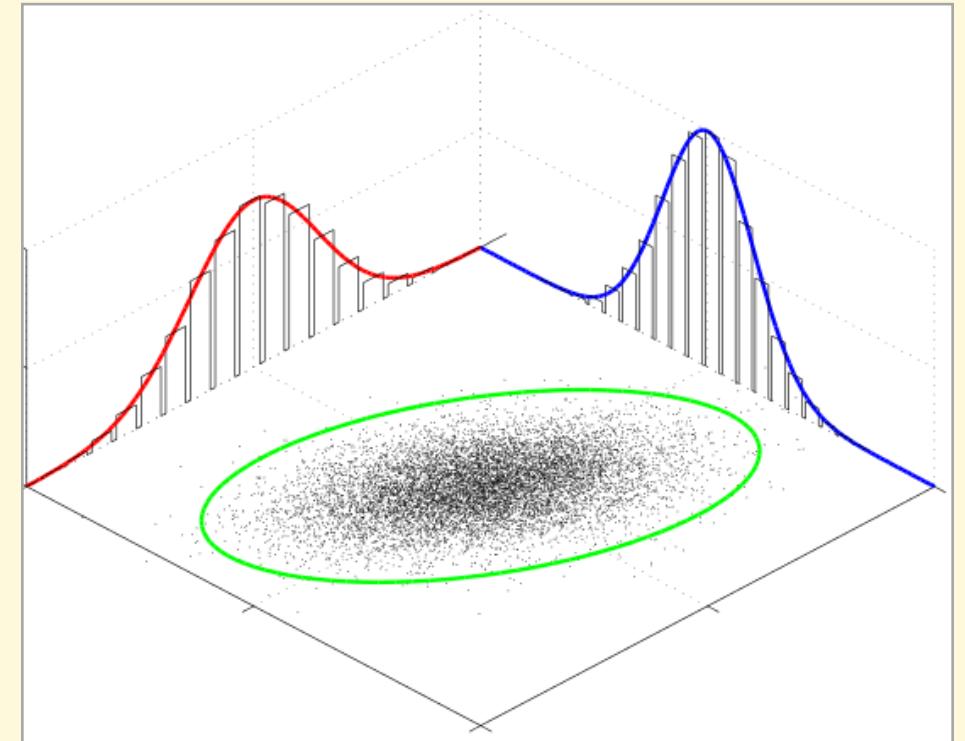
Distributions are often NOT the function of 1 parameter

$$p(x, y)$$

**NB: Coursework 1: is 'happiness' a function solely of GDP?**

The probability of  $x$  is calculated by integrating over  $y$

$$p(x) = \int p(x, y) dy$$



**NB: We wanted to measure  $p(\text{happiness} | \text{GDP})$ , but we measured  $p(\text{happiness, country} | \text{GDP})$**

$$p(\text{happiness} | \text{GDP}) = p(\text{happiness, UK} | \text{GDP}) + p(\text{happiness, France} | \text{GDP}) + p(\text{happiness, Spain} | \text{GDP}) + \dots$$

# An aside: Covariance

When two measurements are not independent, we call them covariant

For two connected variables ( $z = x + y$ ):

$$\sigma_z = \sqrt{\sigma_x^2 + \sigma_y^2}$$

if  $x$  is related to  $y$

$$\sigma_z = \sqrt{\left(\frac{\partial z}{\partial x} \sigma_x\right)^2 + \left(\frac{\partial z}{\partial y} \sigma_y\right)^2 + 2 \frac{\partial z}{\partial x} \frac{\partial z}{\partial y} cov(x, y)}$$

$$cov(x, y) = \frac{1}{N-1} \sum (x - \hat{x})(y - \hat{y})$$

**NB: Parameter transformations (e.g.  $y = \log(x)$ ) change the functional form of the distribution**

## An aside: Covariance

Often expressed in matrix form:

$$\begin{pmatrix} \sigma_x^2 & \sigma_{xy} \\ \sigma_{xy} & \sigma_y^2 \end{pmatrix}$$

Variance in the diagonal terms; covariance in the off-diagonals

# Finding the Distribution

## Kolmogorov-Smirnov (KS) testing

Null hypothesis: the two distributions are the same

The maximal distance between

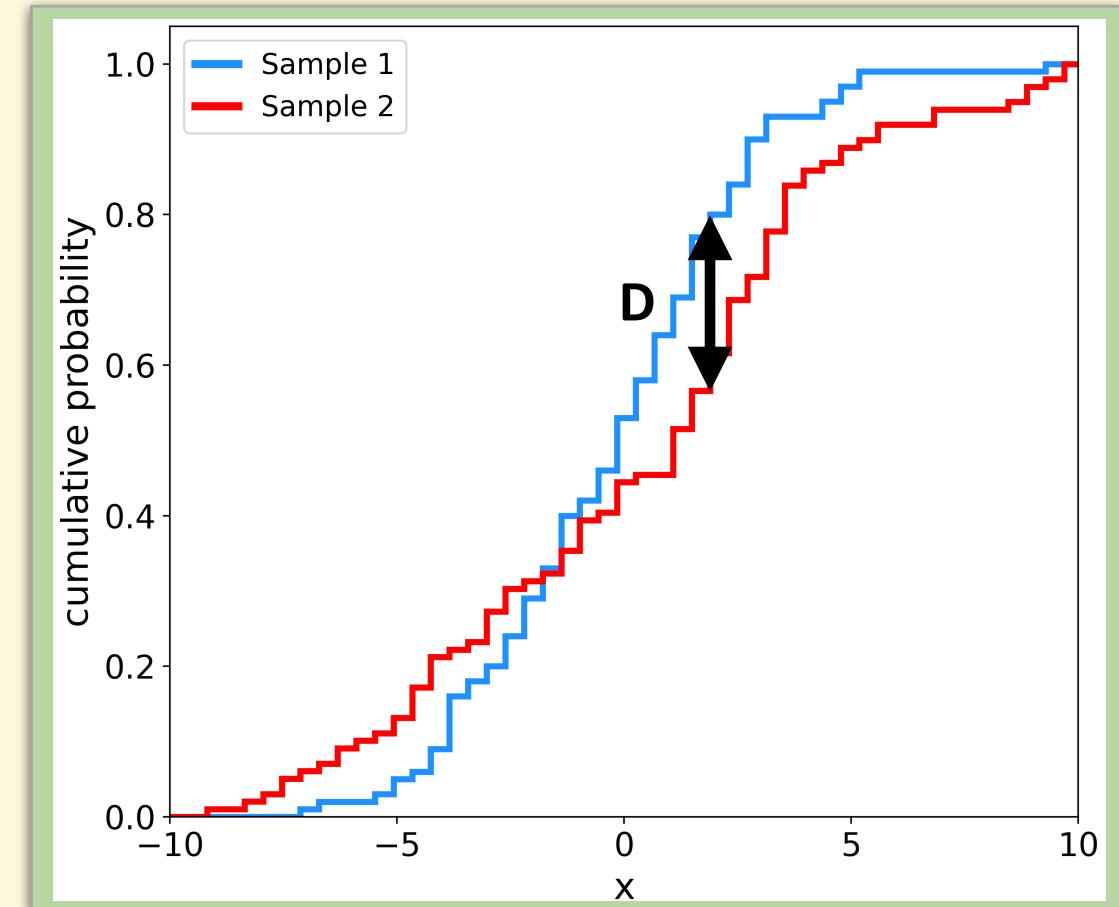
Cumulative Density Functions (CDF)

$$D = \max_{-\infty < x < \infty} |S_N(x) - F(x)|$$

Recall: a CDF is a rank-ordered PDF.

No assumptions about the underlying distribution

**NB:** See also Anderson-Darling test



# Testing for Correlations

## Spearman's Rank (SR)

Testing whether **two variables are related**

Compares the ordered rankings (R) of two datasets

$$p_s = 1 - \frac{\sum_{i=1}^N [R(x_i) - R(y_i)]^2}{N(N^2 - 1)}$$

Output will be a value between -1 and 1.

Null Hypothesis: **the two variables are uncorrelated**

*NB: A non-parametric version of the Pearson test*

# Likelihoods

Given a set of measurements  $x = (x_1, x_2, \dots)$

the *likelihood*:

$$\mathcal{L}(\theta | x) = \prod_{i=0}^N P(x_i | \theta)$$

Best-fit value:

The value that maximises  $\mathcal{L}$

Plausible values:

Determined by the width of  $\mathcal{L}$

Uncertainty:

For a Gaussian distribution  $1\sigma$

# Log Likelihoods

Given a set of measurements  $x = (x_1, x_2, \dots)$

the *log-likelihood*:

$$\ell(\theta) = \ln \mathcal{L}(\theta) = \sum_{i=0}^N P(x_i | \theta)$$

**Much easier to maximise**

Best-fit value:

The value that maximises  $\ell$

Plausible values:

Determined by the width of  $\ell$

Uncertainty:

$$\ell = \ell_{\max} - 0.5$$

**NB: Assumes Gaussianity**

# Log Likelihoods: the Gaussian case

Given a set of measurements  $x = (x_1, x_2, \dots)$

$$p(x_i) = \frac{1}{\sqrt{2\pi}\sigma_i} \exp\left(-\frac{(x_i - \mu_i(\theta))^2}{2\sigma_i^2}\right)$$

$$\mathcal{L}(\theta | x) = p(x_1 | \theta) * p(x_2 | \theta) * \dots$$

$$\mathcal{L}(\theta | x) = e^{-\chi^2/2} * \prod_{i=1}^N \frac{1}{\sigma_i} \times (2\pi)^{(-N/2)}$$

$$-2 \ln \mathcal{L}(\theta | x) = \chi^2 + 2 \sum_{i=1}^N \ln \sigma_i + N \ln(2\pi)$$

**Maximising  $\mathcal{L}(\theta | x)$  is equivalent to minimising  $\chi^2$**

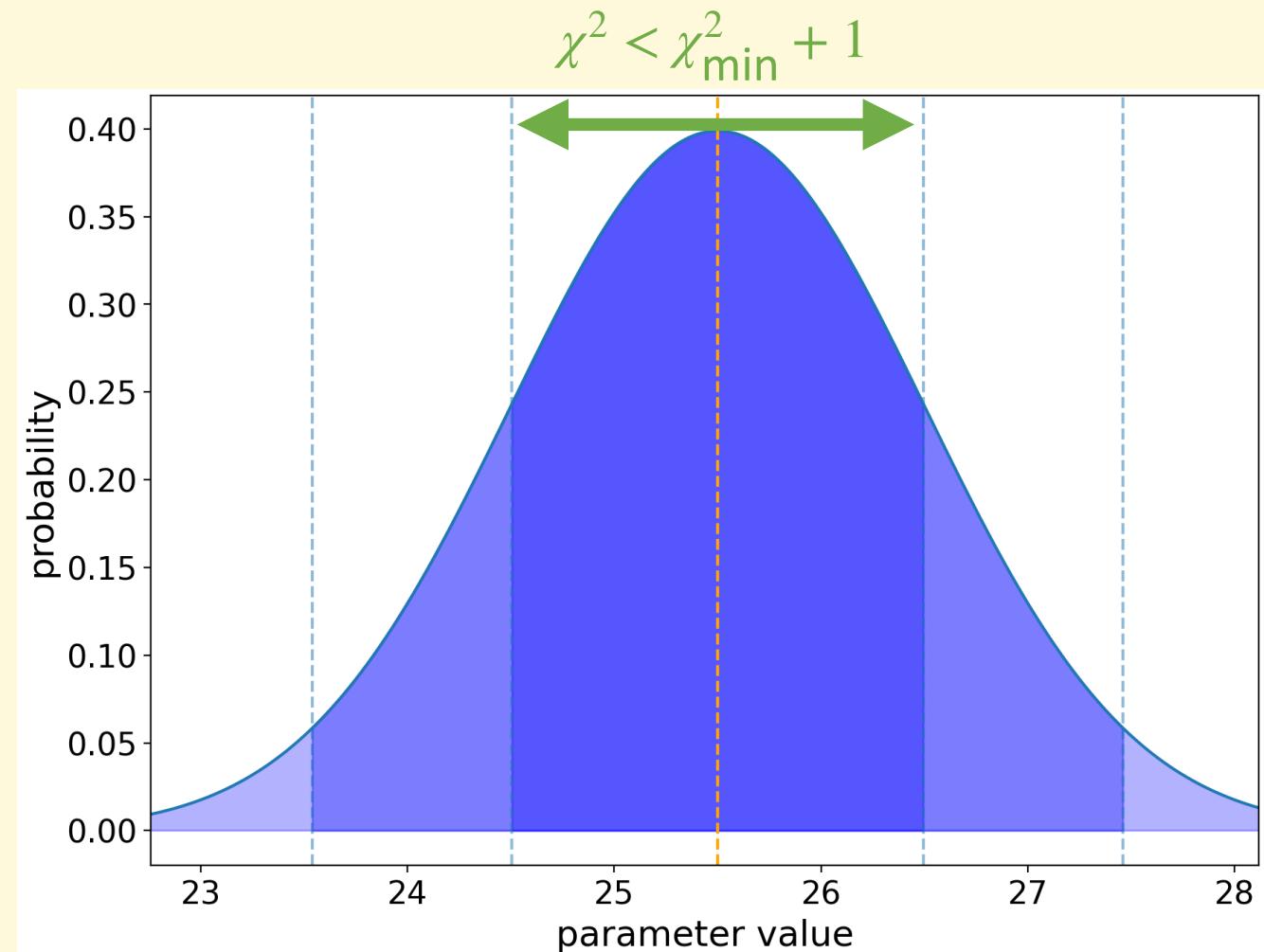
# Confidence levels & $\chi^2$

If our best-fit value has  $\chi^2 = \chi^2_{\min}$ :

Contours of equal probability are defined by

$$\chi^2 = \chi^2_{\min} + \Delta$$

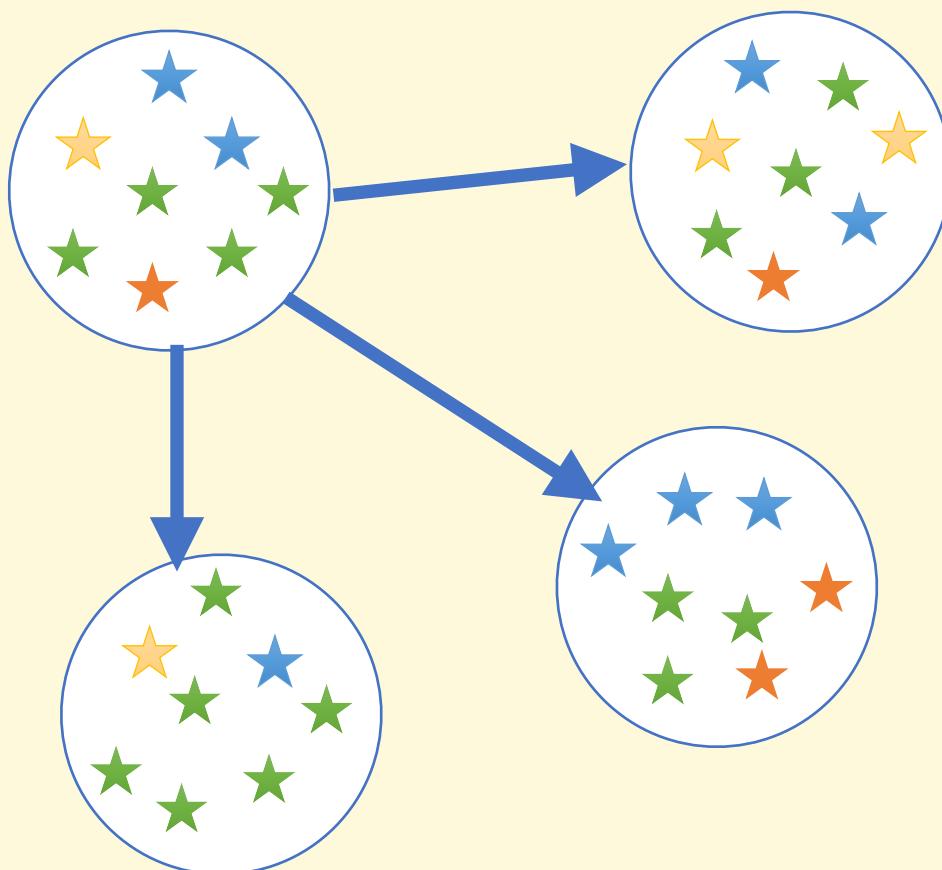
		<i>N params</i>		
		1	2	3
C	68%	1 $\sigma$	2.30	3.53
	95.4%	2 $\sigma$	6.17	8.02
	99.7%	3 $\sigma$	11.8	14.2



**WARNING:** This process knows nothing about physics: beware of unphysical regions

# $\chi^2$ : estimating uncertainties : bootstrapping

Resampling the data **with** replacement



(For a dataset with  $N$  measurements)

For each of  $k$  samples:

- > Randomly select  $N$  values
- > Calculate the  $\chi^2$

Calculate the mean, and variance on  $\chi^2$

*Does not assume a functional form for the distribution*

*NB: Can also sample from uncertainties : 'MonteCarlo'*

# Model Selection

Given two models  $M_1, M_2$  how do we know which one to choose?

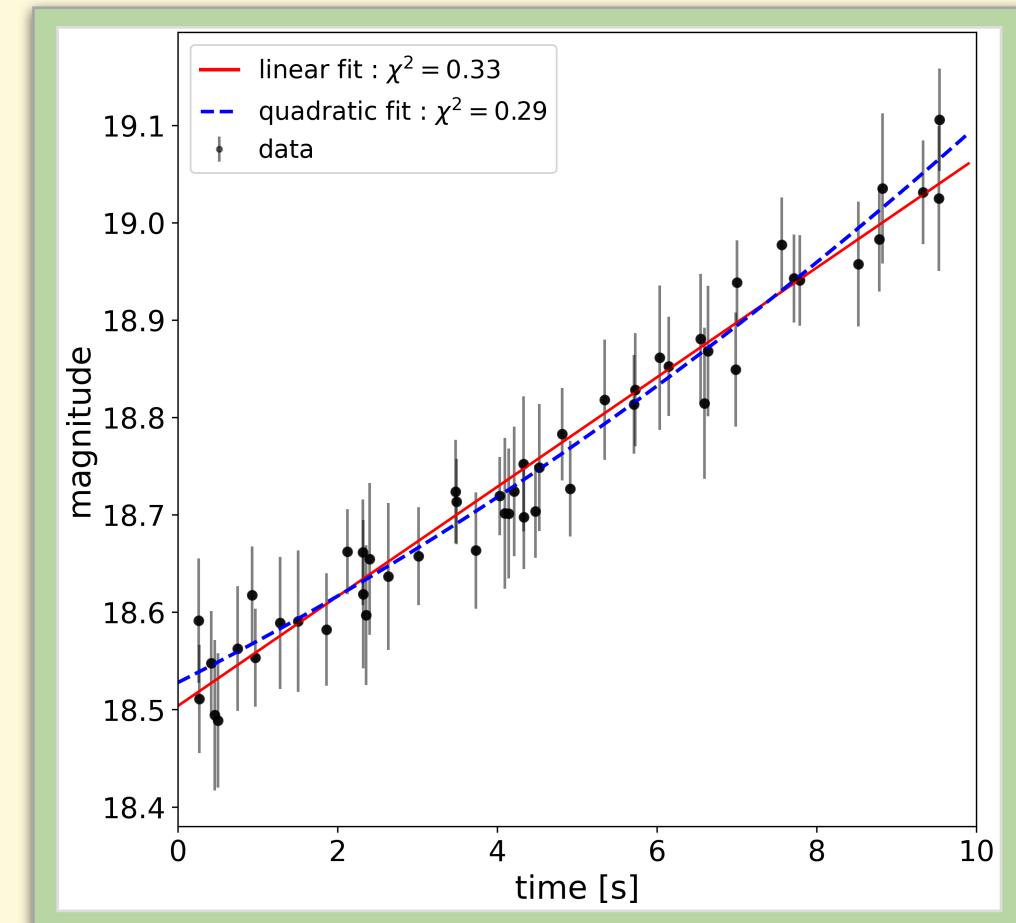
**Recall that for a good fit, we expect:  $\min(\chi^2_{\text{red}}) \sim 1$**

Akaike Information Criteria (AIC)

$$\text{AIC} = -2 \ln(\mathcal{L}(\theta)) + 2p$$

Bayes Information Criteria (BIC)

$$\text{BIC} = -2 \ln(\mathcal{L}(\theta)) + p \log N$$



$$(p = N_{\text{params}}; n = N_{\text{data}})$$

# *Week 12: Learning outcomes*

*Today you have learnt*

- How to relate probabilities to measurements: the likelihood
- How to define a null hypothesis
  - How to use probabilities to test our hypothesis
- Three key distributions in Physics
- Key tests of whether our hypothesis (assumptions) are correct
  - Distributions: KS test
  - Correlations: Spearman's Rank
- Goodness-of-fit metrics: calculating likelihoods
  - Selecting a model given a likelihood

Practical examples on Friday!