

A comparative study on PCR, PLS, Envelope and BayesPLS models

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- Background
- Estimation methods under comparison
- Data Simulation
- Analysis, Results and Discussions

- **PLS Population Model** [Helland, 1990] which further discussed by [Naes and Helland, 1993, Helland, 2001]

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- Cook et al. [2013] said that PLS is fundamentally an envelope in the population model

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- Using `simrel` [Sæbø et al., 2015] R-package, data with diverse nature are simulated.
- `simrel` allows to have control over latent structure (relevant component) of the data, fine analysis of strength and weakness of a models is possible

The common ground of all the methods is to best describe (fit) the multivariate linear model below,

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon} \quad (1)$$

where,

\mathbf{y}	:	Response
\mathbf{X}	:	Matrix of p predictor variable
$\boldsymbol{\beta}$:	Regression Coefficients
$\boldsymbol{\epsilon}$:	Error $\boldsymbol{\epsilon} \sim \text{NID}(0, \sigma^2)$

Here, both \mathbf{y} and \mathbf{X} are considered to be centered.

All the models under this study consider a **subspace of predictor variables that is relevant for response**. They differ in the ways of finding the subspace and corresponding model estimates. The true estimates can also be written as,

$$\beta = \Sigma_{XX}^{-1} \sigma_{Xy} = \sum_{j=1}^p \frac{1}{\alpha_j} \mathbf{e}_j \mathbf{e}_j^t \sigma_{Xy} = \sum_{j=1}^p \gamma_j \mathbf{e}_j$$

where,

$$\gamma_j : \frac{\mathbf{e}_j^t \sigma_{Xy}}{\lambda_j}$$

\mathbf{e}_j : Eigenvector of Σ_{xx}

λ_j : Eigenvalue of Σ_{xx}

σ_{Xy} : Covariance between y and X

So, True regression estimates are the space spanned by the eigenvectors of population covariance matrix Σ_{xx} .

PCR

- * Regression of response on latent space of predictor
 - * No strict assumption
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Envelope (MLE)

- * Estimation using Maximum Likelihood
 - * Can not be used when predictor is larger than observations
-

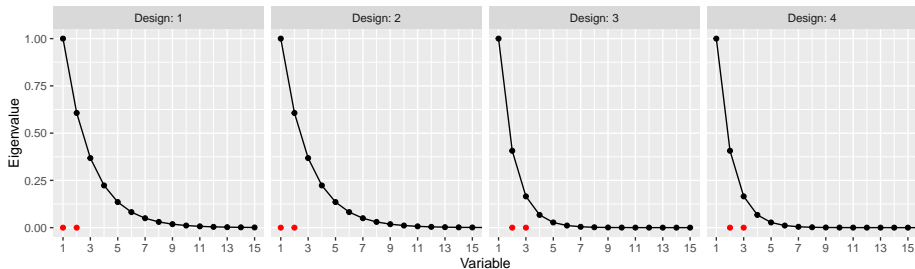
Bayes

- * Estimation through MCMC approach with rotation of relevant space
 - * Heavy Computation when p is large
-

Models are analysed under diverse nature of data. Data are simulated using `simrel` package (R). In this study, I have included following four design;

n	p	R2	relpos	gamma
50	15	0.5	1, 2	0.5
50	40	0.5	1, 2	0.5
50	15	0.9	2, 3	0.9
50	40	0.9	2, 3	0.9

n	:	Number of observations
p	:	Number of variables
R2	:	Variation explained by the model
relpos	:	Position of relevant components
gamma	:	Reduction factor of eigenvalue of X

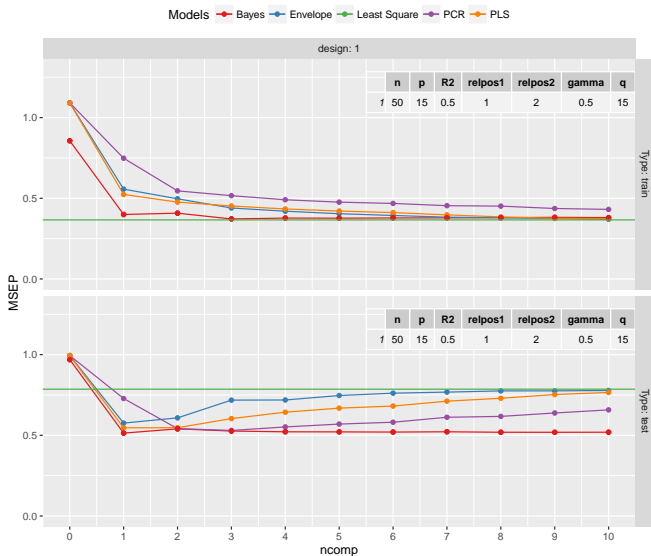


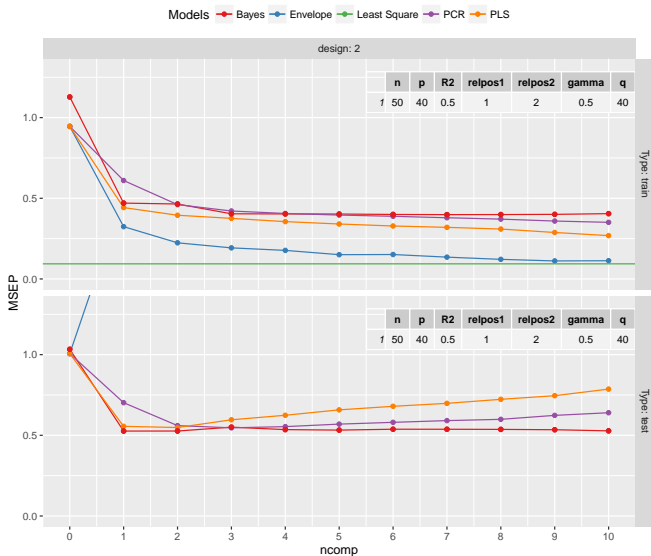
- When Relevant components are at the position of high eigenvalues, the situation is easier to model
- When Relevant components are at the position of low eigenvalues, for example 5, 10, then the most variation present in X are not relevant for Y and this will become a very difficult situation.

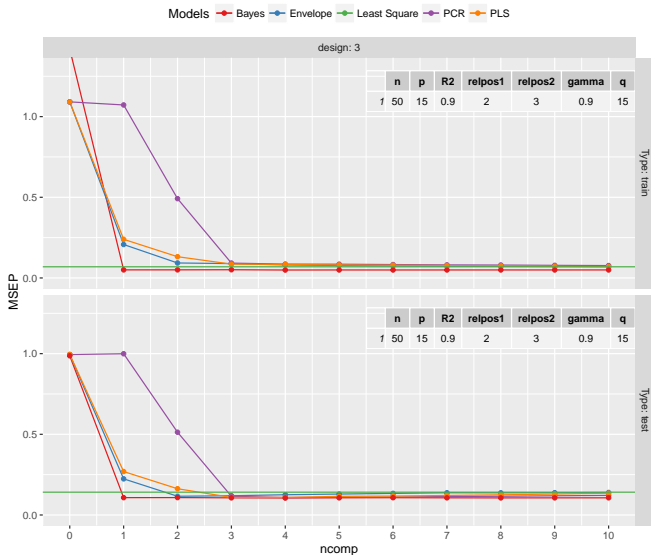
Models are compared on the basis of their prediction ability by measuring *test* and *training* **Mean Square Error of Prediction (MSEP)**. Mean prediction error is calculated as,

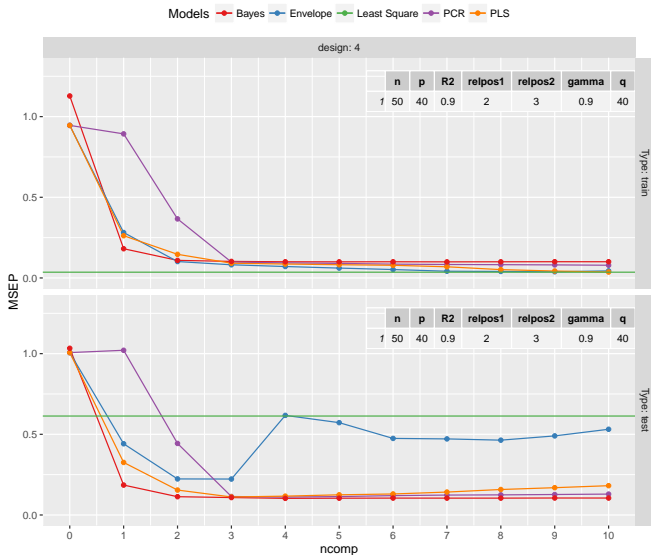
$$(\text{Prediction Error})_{\text{training}} = \frac{1}{n} \sum_{i=1}^n (y_i - \hat{y}_i)^2 = \frac{1}{n} \sum_{i=1}^n \left(y_i - \left(\hat{\beta}_0 + \hat{\beta} X_i \right) \right)^2$$

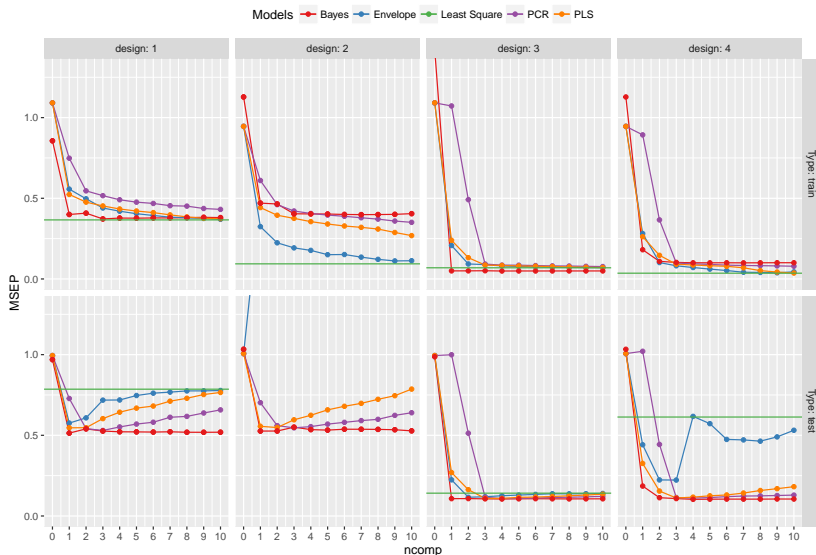
$$\begin{aligned} (\text{Prediction Error})_{\text{test}} &= \frac{1}{n} \sum_{i=1}^{\text{ntest}} \left(y_{i(\text{test})} - \hat{y}_{i(\text{test})} \right)^2 \\ &= \frac{1}{n} \sum_{i=1}^{\text{ntest}} \left(y_{i(\text{test})} - \left(\hat{\beta}_0 + \hat{\beta} X_{i(\text{test})} \right) \right)^2 \end{aligned}$$











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- All the models are performing better than the least square solution

[illegible]

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