# ECON 4360: Empirical Finance

Regression-Based Tests of Linear Factor Models

Sherry Forbes

University of Virginia

Empirics Lecture #05

## What are we doing today?

- Regression-Based Tests of Linear Factor Models
  - Time-Series Regressions
  - Cross-Section Regressions

### What's a Linear Factor Model?

For our central asset pricing equation

$$p = E(mx)$$
,

we have a *linear* factor model if we can express the SDF as

$$m = b'f$$
.

- We're not going to work with the model in this specification just yet (we will soon though, never fear!)...
- But recall: A linear model for the discount factor is equivalent to an expected-return beta representation, so we're going to start here...

## Recall: Expected Return-Beta Representation

- The model:
  - Expected Return-Beta Representation of linear factor pricing models

$$E\left(R^{i}\right) = \gamma + \beta_{i,a}\lambda_{a} + \beta_{i,b}\lambda_{b} + \cdots, i = 1 \dots N$$

- The general idea:
  - Explain the variation in average returns across assets due to each assets exposure - as measured by the betas - to various risks, as priced by the lambdas.

### Recall: One of Our Derivations

 Recall how we can use our central asset pricing equation to get a basic pricing equation for returns

$$1=E\left[ mR^{i}\right]$$

And recall how we can manipulate this expression to re-write it as

$$1 = E[mR^{i}]$$

$$1 = E(m)E(R^{i}) + cov(m, R^{i})$$

$$E(R^{i}) = R^{f} - R^{f}cov(m, R^{i})$$

$$E(R^{i}) = R^{f} + \left(\frac{cov(R^{i}, m)}{var(m)}\right)\left(-\frac{var(m)}{E(m)}\right)$$

• Or, equivalently

$$E\left(R^{i}\right) = R^{f} + \beta_{i,m}\lambda_{m}$$

### Recall: What We Learned From This...

 This derivation gave us an expression for an expected return-beta model...

$$E\left(R^{i}\right) = R^{f} + \beta_{i,m}\lambda_{m}$$

 In the consumption-based model, we figured out that assets whose returns covary positively with consumption growth must have higher expected returns as compensation for risk

### Recall: What We Learned From This...

 This derivation gave us an expression for an expected return-beta model...

$$E\left(R^{i}\right) = R^{f} + \beta_{i,m}\lambda_{m}$$

- In the consumption-based model, we figured out that assets whose returns covary positively with consumption growth must have higher expected returns as compensation for risk
  - Think about/interpret  $\lambda_m$  as consumption growth

### Recall: What We Learned From This...

 This derivation gave us an expression for an expected return-beta model...

$$E\left(R^{i}\right) = R^{f} + \beta_{i,m}\lambda_{m}$$

- In the consumption-based model, we figured out that assets whose returns covary positively with consumption growth must have higher expected returns as compensation for risk
  - ullet Think about/interpret  $\lambda_m$  as consumption growth
  - Think about the  $\beta_{i,m}$ 's as regression coefficients telling us whether returns for a particular asset are typically high in good times or high in bad times.

### So what can we do now?

- ullet The idea is that we're trying to find a proxy or proxies for  $\lambda_m$ 
  - It doesn't necessarily have to be consumption growth!
- What we can see now is that if we can find "proxies" for consumption growth, good times/bad times, etc., we will have a "linear factor model" in expected return-beta form!

Generally speaking, an asset pricing model of the form

$$E\left(R^{i}\right) = R^{f} + \beta_{i,m}\lambda_{m}$$

- Says an expected return should be proportional to its beta i.e., the regression coefficient  $\beta_{i,m}$  of an asset's return  $R^i$  on risk factors  $\lambda_m$
- What is  $\lambda_m$ ?

• Generally speaking, an asset pricing model of the form

$$E\left(R^{i}\right) = R^{f} + \beta_{i,m}\lambda_{m}$$

- Says an expected return should be proportional to its beta i.e., the regression coefficient  $\beta_{i,m}$  of an asset's return  $R^i$  on risk factors  $\lambda_m$
- What is  $\lambda_m$ ?
  - ullet Interpreted as the price of risk, the same for all i assets

Generally speaking, an asset pricing model of the form

$$E\left(R^{i}\right) = R^{f} + \beta_{i,m}\lambda_{m}$$

- Says an expected return should be proportional to its beta i.e., the regression coefficient  $\beta_{i,m}$  of an asset's return  $R^i$  on risk factors  $\lambda_m$
- What is  $\lambda_m$ ?
  - Interpreted as the price of risk, the same for all i assets
- What is  $\beta_{i,m}$ ?

Generally speaking, an asset pricing model of the form

$$E\left(R^{i}\right) = R^{f} + \beta_{i,m}\lambda_{m}$$

- Says an expected return should be proportional to its beta i.e., the regression coefficient  $\beta_{i,m}$  of an asset's return  $R^i$  on risk factors  $\lambda_m$
- What is  $\lambda_m$ ?
  - Interpreted as the price of risk, the same for all i assets
- What is  $\beta_{i,m}$ ?
  - Interpreted as the quantity of risk in each asset

### Introduction to Evaluating Linear Factor Models

- Estimating and evaluating these types of models are what we will look at next...
- Linear factor models are the most common in empirical work.
  - How should we estimate / evaluate them?
- Next-Up:
  - Times-series and Cross-section Regressions (Yay!)

## Lots of Econometric Techniques

- Same questions:
  - How do we estimate parameters?
    - How do we calculate their standard errors?
  - How can we calculate standard errors of the pricing errors?
  - How can we test the model?
- Recall, we've already addressed these issues in a GMM Framework.
  - Now, we're going to look at them through times-series and cross-section regressions (Yay!)

- Let's start with an example of the simplest model a linear (1) factor model
- We can evaluate the model

$$E\left(R^{ei}\right) = \beta_i E\left(f\right)$$

by running OLS time-series regressions

$$R_t^{ei} = \alpha_i + \beta_i f_t + \varepsilon_t^i, \quad t = 1, 2, ..., T.$$

- If the factor is also a return like the excess return on the market portfolio,  $f_t = R_t^{em} R_t^f$  we have the familiar CAPM.
- The theory says that

$$E\left(R_{t}^{ei}\right)=\beta_{i}E\left(f_{t}\right)$$

• Since the theory says

$$E\left(R_{t}^{ei}\right)=\beta_{i}E\left(f_{t}\right)$$

the implications for

$$R_t^{ei} = \alpha_i + \beta_i f_t + \varepsilon_t^i, \qquad t = 1, 2, ..., T.$$

are that all the regression intercepts  $\alpha_i$  should be zero!

Since the theory says

$$E\left(R_{t}^{ei}\right)=\beta_{i}E\left(f_{t}\right)$$

the implications for

$$R_t^{ei} = \alpha_i + \beta_i f_t + \varepsilon_t^i, \quad t = 1, 2, ..., T.$$

are that all the regression intercepts  $\alpha_i$  should be zero!

• The regression intercepts here are equivalent to the "pricing errors"

Since the theory says

$$E\left(R_{t}^{ei}\right)=\beta_{i}E\left(f_{t}\right)$$

the implications for

$$R_t^{ei} = \alpha_i + \beta_i f_t + \varepsilon_t^i, \quad t = 1, 2, ..., T.$$

are that all the regression intercepts  $\alpha_i$  should be zero!

- The regression intercepts here are equivalent to the "pricing errors"
- If the theory is correct, then  $\alpha_i$  should = 0. If that condition is "true", then our model correctly prices that asset.

- For our simple model, we have a factor pricing model with a single factor
  - The factor is an excess return, e.g.,  $R^{em} = R^m R^f$
  - And all the test assets are excess returns
- Recall, first: If we are using excess returns

$$E(R^{i}) = \gamma + \beta_{i,a}\lambda_{a} + \beta_{i,b}\lambda_{b} + \cdots, i = 1 \dots N$$

but

$$E\left(R^{ei}\right) = \beta_{i,a}\lambda_a + \beta_{i,b}\lambda_b + \cdots$$

I.e., the intercept drops out.

• Recall, second: If the factors themselves are returns, like  $f = R^{em} = R^m_t - R^f_t$  for the CAPM, the model should apply to the factor as well.

- Recall, second: If the factors themselves are returns, like  $f = R^{em} = R^m_t R^f_t$  for the CAPM, the model should apply to the factor as well.
- Then the time-series regression for the factor would be

$$R_t^{em} = \alpha + \beta_{i,m} R_t^{em} + \varepsilon_t^i$$

- Recall, second: If the factors themselves are returns, like  $f = R^{em} = R^m_t R^f_t$  for the CAPM, the model should apply to the factor as well.
- Then the time-series regression for the factor would be

$$R_t^{em} = \alpha + \beta_{i,m} R_t^{em} + \varepsilon_t^i$$

ullet So  $eta_{i,m}=1$  - i.e., the factor has a beta of one on itself

- Recall, second: If the factors themselves are returns, like  $f = R^{em} = R^m_t R^f_t$  for the CAPM, the model should apply to the factor as well.
- Then the time-series regression for the factor would be

$$R_t^{em} = \alpha + \beta_{i,m} R_t^{em} + \varepsilon_t^i$$

- ullet So  $eta_{i.m}=1$  i.e., the factor has a beta of one on itself
- And we can write an estimate of the factor risk premium as

$$\widehat{\lambda} = E_T(f) = E_T(R^{em})$$

Really Easy

- Really Easy
  - 1. Estimate the factor risk premium by the sample mean of the factor, e.g.,

$$\widehat{\lambda} = E_T(f) = E_T(R^{em})$$

- Really Easy
  - 1. Estimate the factor risk premium by the sample mean of the factor, e.g.,

$$\widehat{\lambda} = E_T(f) = E_T(R^{em})$$

• 2. Run OLS time-series regressions

$$R_t^{ei} = \alpha_i + \beta_i f_t + \varepsilon_t^i, \qquad t = 1, 2, ..., T$$

for each test asset.

- Really Easy
  - 1. Estimate the factor risk premium by the sample mean of the factor, e.g.,

$$\widehat{\lambda} = E_T(f) = E_T(R^{em})$$

• 2. Run OLS time-series regressions

$$R_t^{ei} = \alpha_i + \beta_i f_t + \varepsilon_t^i, \quad t = 1, 2, ..., T$$

for each test asset.

• If you have 10 test assets, how many regressions do you run? How many  $\beta$ 's do you have?

# Testing A Single Pricing Error

• What can you use to test whether a pricing error is zero?

- If the regression errors are uncorrelated and homoskedastic, you can use the standard distributions.
  - What if they are not? Do you know how to handle that?

# Testing A Single Pricing Error

- What can you use to test whether a pricing error is zero?
  - A t-test
- If the regression errors are uncorrelated and homoskedastic, you can use the standard distributions.
  - What if they are not? Do you know how to handle that?

## Testing Many Pricing Errors

• Can you anticipate how to test that all pricing errors are jointly zero?

## Testing Many Pricing Errors

- Can you anticipate how to test that all pricing errors are jointly zero?
  - A chi-square test

# Testing Many Pricing Errors

• To test that all pricing errors are jointly zero, the form is

$$T\left[1+\left(\frac{E_{T}\left(f\right)}{\widehat{\sigma}\left(f\right)}\right)^{2}\right]^{-1}\widehat{\alpha}'\widehat{\Sigma}^{-1}\widehat{\alpha}\sim\chi_{N}^{2}$$

- Very intuitive vis-a-vis what we've seen already
  - The meat of the test  $\widehat{\alpha}'\widehat{\Sigma}^{-1}\widehat{\alpha}$  is just a quadratic form of the pricing errors
  - (Details are in the book, if you're interested in the derivation.)

### Another Kind of Test

- We've already talked about the fact that the CAPM won't work well if we don't have a good proxy for  $R^{em}$ .
- Generally speaking, a single-beta representation (a one factor model!) exists iff the reference return is on the M-V frontier.
- So a test of this model can also be found through a statistic interpreted as a test of whether f is actually ex ante mean-variance efficient.

### Another Kind of Test

- We've already talked about the fact that the CAPM won't work well
  if we don't have a good proxy for R<sup>em</sup>.
- Generally speaking, a single-beta representation (a one factor model!) exists iff the reference return is on the M-V frontier.
- So a test of this model can also be found through a statistic interpreted as a test of whether f is actually ex ante mean-variance efficient.
  - Why not a test of f on the ex post m-v frontier?

### Another Kind of Test

- We've already talked about the fact that the CAPM won't work well
  if we don't have a good proxy for R<sup>em</sup>.
- Generally speaking, a single-beta representation (a one factor model!) exists iff the reference return is on the M-V frontier.
- So a test of this model can also be found through a statistic interpreted as a test of whether f is actually ex ante mean-variance efficient.
  - Why not a test of f on the ex post m-v frontier?
  - Even if f is on the m-v frontier using population moments, it may not be on the m-v frontier using sample moments. It may be outperformed by others due to luck, but it shouldn't be "too far" inside the sample / ex post m-v frontier.

## Why a Multi-Factor Model?

- One factor may be insufficient. Why?
- If we have k factors (still excess returns), we can just write

$$E\left(R_{t}^{ei}\right)=\beta_{i}^{\prime}E\left(f_{t}\right)$$

and use regression equations

$$R_t^{ei} = \alpha_i + \beta_i' f_t + \epsilon_t^i, \qquad t = 1, 2, ..., T.$$

• The basic difference is that we now have f and  $\beta$  as  $k \times 1$  vectors and we get a little bit more algebra.

### Again: The Expected Return-Beta Representation

 Let's again think about an Expected Return-Beta Representation of linear factor pricing models

$$E(R^{i}) = \gamma + \beta_{i,a}\lambda_{a} + \beta_{i,b}\lambda_{b} + \cdots, i = 1...N$$

- Since we want to explain the variation in average returns across
  assets due to each assets exposure as measured by the betas to
  various risks, as priced by the lambdas...
  - We're now going to look at cross-sectional regressions...

 Think about a scatterplot of average returns verses betas in a single-factor model like the CAPM.

- Think about a scatterplot of average returns verses betas in a single-factor model like the CAPM.
  - Average returns should be proportional to betas

- Think about a scatterplot of average returns verses betas in a single-factor model like the CAPM.
  - Average returns should be proportional to betas
- If we fit a line through the datapoints, the slope of that line would be our estimate of  $\lambda$

- Think about a scatterplot of average returns verses betas in a single-factor model like the CAPM.
  - Average returns should be proportional to betas
- ullet If we fit a line through the datapoints, the slope of that line would be our estimate of  $\lambda$ 
  - Returns won't lie exactly on that line, of course, so we will have a scatter plot

- Think about a scatterplot of average returns verses betas in a single-factor model like the CAPM.
  - Average returns should be proportional to betas
- ullet If we fit a line through the datapoints, the slope of that line would be our estimate of  $\lambda$ 
  - Returns won't lie exactly on that line, of course, so we will have a scatter plot
  - Deviations from that line are the pricing errors (i.e., the residuals in a cross-sectional regression).

• A cross-sectional regression just fits a line through such a scatterplot.

- A cross-sectional regression just fits a line through such a scatterplot.
- First step: We find estimates of the betas for each of our *N* assets from the time-series regressions

$$R_t^{ei} = \alpha_i + \beta_i f_t + \varepsilon_t^i, \quad t = 1, 2, ..., T.$$

- A cross-sectional regression just fits a line through such a scatterplot.
- First step: We find estimates of the betas for each of our *N* assets from the time-series regressions

$$R_t^{ei} = \alpha_i + \beta_i f_t + \varepsilon_t^i, \quad t = 1, 2, ..., T.$$

ullet Second step: We then estimate the factor risk premia  $\lambda$  from a regression across assets of average returns on the *betas we just estimated* 

$$E_T(R^{ei}) = \beta'_i \lambda + \alpha_i, \quad i = 1, 2, ..., N.$$

- A cross-sectional regression just fits a line through such a scatterplot.
- First step: We find estimates of the betas for each of our *N* assets from the time-series regressions

$$R_t^{ei} = \alpha_i + \beta_i f_t + \varepsilon_t^i, \quad t = 1, 2, ..., T.$$

ullet Second step: We then estimate the factor risk premia  $\lambda$  from a regression across assets of average returns on the *betas we just estimated* 

$$E_T(R^{ei}) = \beta'_i \lambda + \alpha_i, \quad i = 1, 2, ..., N.$$

• In the second step, what are we trying to estimate? Whare are the RHS variables? What are the pricing errors? Can / should you run this regression without a constant?

• In the second step regression:

$$E_T(R^{ei}) = \beta'_i \lambda + \alpha_i, \quad i = 1, 2, ..., N.$$

- The betas are the RHS variables, the alphas are the pricing errors (the residuals), and the lambas are the regression coefficients we are trying to estimate.
- You can run the regression with or without a constant, but the theory says the constant (zero-beta excess return!) should be zero.

## More on Cross-Sectional Regressions

- Estimates:
  - For

$$E_T\left(R^{ei}\right)=eta_i'\lambda+lpha_i, ~~i=1,2,...,N.$$

OLS cross-section estimates are

$$\widehat{\lambda} = (\beta'\beta)^{-1} \beta' E_t (R^e)$$

$$\widehat{\alpha} = E_T (R^e) - \widehat{\lambda} \beta.$$

- Distribution Theory:
  - Note: Normally, when we use OLS in a cross-sectional regression, the assumption is that the RHS variables are fixed.
    - The betas (our RHS variables here) are not fixed we estimated them from the time-series regressions, and this matters for the asymptotic distributions.
  - Your book has various derivations, but the correct asymptotic standard errors are given in Equations (12.19) and (12.20).
    - The correction is due to Shanken (1992), and basically incorporate the variance-covariance matrix of the factors into the correction.

#### Test of the Model?

- Again, a test of the model is a test of whether or not the pricing errors are close to zero.
  - Here, this can again be done with a chi-square test on the pricing errors here, those pricing errors are the residuals.
- If that sounds strange, it probably should.
  - How can you test residuals in OLS regressions? What does it mean for the residuals to be zero?
  - Normally, we wouldn't have any information about the residuals, other than the residuals themselves.
  - However, now, the first-stage time-series regressions give us independent information about the size of  $cov(\alpha\alpha')$  that we cannot get from looking at the cross-section residuals by themselves.

- How are the approaches different?
  - You can run cross-sectional regressions when the factors are not returns.
  - The time-series regressions require returns so you can estimate factor risk premia by  $\widehat{\lambda}=E_{T}\left(f\right)$
- Our asset-pricing model predicts restrictions on the intercepts in the time-series regressions (that we can test, of course).
  - If we impose the restriction  $E\left(R^{ei}\right)=\beta_i'\lambda$  we can write the time-series regression as

$$R_{t}^{ei}=eta_{i}^{\prime}\lambda+eta_{i}^{\prime}\left(f_{t}-E\left(f
ight)
ight)+arepsilon_{t}^{i},\hspace{0.5cm}t=1,2,...,T,$$

so the intercept restriction is

$$a_i = \beta_i' (\lambda - E(f))$$

- This gives us our zero intercept condition (as expected).
  - Interpretation: Mean returns should be proportional to betas; the intercept controls the mean return.

• Even if factors are returns, the two approaches are still not necessarily the same.

- Even if factors are returns, the two approaches are still not necessarily the same.
- The time-series regression estimates the factor risk premium as the sample mean of the factor.

- Even if factors are returns, the two approaches are still not necessarily the same.
- The time-series regression estimates the factor risk premium as the sample mean of the factor.
  - The factor receives a zero pricing error in each sample.

- Even if factors are returns, the two approaches are still not necessarily the same.
- The time-series regression estimates the factor risk premium as the sample mean of the factor.
  - The factor receives a zero pricing error in each sample.
- The OLS cross-section regression picks the slope (estimate of the factor risk premium) and intercept to best fit all the data points.

- Even if factors are returns, the two approaches are still not necessarily the same.
- The time-series regression estimates the factor risk premium as the sample mean of the factor.
  - The factor receives a zero pricing error in each sample.
- The OLS cross-section regression picks the slope (estimate of the factor risk premium) and intercept to best fit all the data points.
- What about GLS?

- Even if factors are returns, the two approaches are still not necessarily the same.
- The time-series regression estimates the factor risk premium as the sample mean of the factor.
  - The factor receives a zero pricing error in each sample.
- The OLS cross-section regression picks the slope (estimate of the factor risk premium) and intercept to best fit all the data points.
- What about GLS?
- If the factor is a return, GLS cross-section is equivalent to the time-series regression since GLS puts all its weight on the asset with the lowest residual variance.

- Even if factors are returns, the two approaches are still not necessarily the same.
- The time-series regression estimates the factor risk premium as the sample mean of the factor.
  - The factor receives a zero pricing error in each sample.
- The OLS cross-section regression picks the slope (estimate of the factor risk premium) and intercept to best fit all the data points.
- What about GLS?
- If the factor is a return, GLS cross-section is equivalent to the time-series regression since GLS puts all its weight on the asset with the lowest residual variance.
  - If the factor is included as a test asset, it has zero residual variance.

- Even if factors are returns, the two approaches are still not necessarily the same.
- The time-series regression estimates the factor risk premium as the sample mean of the factor.
  - The factor receives a zero pricing error in each sample.
- The OLS cross-section regression picks the slope (estimate of the factor risk premium) and intercept to best fit all the data points.
- What about GLS?
- If the factor is a return, GLS cross-section is equivalent to the time-series regression since GLS puts all its weight on the asset with the lowest residual variance.
  - If the factor is included as a test asset, it has zero residual variance.
  - Interpretation: The "efficient" cross-section regression (GLS) ignores
    all information in other asset returns and uses only information in the
    factor returns to estimate the factor risk premium.

• Fama-MacBeth (1973) - historically important, computationally simple, and widely used.

- Fama-MacBeth (1973) historically important, computationally simple, and widely used.
- Procedure:

- Fama-MacBeth (1973) historically important, computationally simple, and widely used.
- Procedure:
  - First, find estimates of the betas with time-series regressions.

- Fama-MacBeth (1973) historically important, computationally simple, and widely used.
- Procedure:
  - First, find estimates of the betas with time-series regressions.
  - Second, run a cross-sectional regression at each time period.

$$R_t^{ei} = \beta_i' \lambda_t + \alpha_{it}, \qquad i = 1, 2, ..., N$$

- Fama-MacBeth (1973) historically important, computationally simple, and widely used.
- Procedure:
  - First, find estimates of the betas with time-series regressions.
  - Second, run a cross-sectional regression at each time period.

$$R_t^{ei} = \beta_i' \lambda_t + \alpha_{it}, \qquad i = 1, 2, ..., N$$

 (This is done instead of estimating a single cross-sectional regression with sample averages.)

- Fama-MacBeth (1973) historically important, computationally simple, and widely used.
- Procedure:
  - First, find estimates of the betas with time-series regressions.
  - Second, run a cross-sectional regression at each time period.

$$R_t^{ei} = \beta_i' \lambda_t + \alpha_{it}, \quad i = 1, 2, ..., N$$

- (This is done instead of estimating a *single* cross-sectional regression with sample averages.)
- Third, estimate  $\lambda$  and  $a_i$  as averages of the cross-sectional regression estimates

$$\widehat{\lambda} = \frac{1}{T} \sum_{t=1}^{T} \widehat{\lambda}_{t}$$

$$\widehat{\alpha}_i = \frac{1}{T} \sum_{t=1}^T \widehat{\alpha}_{it}$$

## Fama-MacBeth: Sampling Errors

• Fama-MacBeth (1973) sampling errors are then

$$\sigma^{2}\left(\widehat{\lambda}\right) = \frac{1}{T^{2}} \sum_{t=1}^{T} \left(\widehat{\lambda}_{t} - \widehat{\lambda}\right)^{2}$$
$$\sigma^{2}\left(\widehat{\alpha}_{i}\right) = \frac{1}{T^{2}} \sum_{t=1}^{T} \left(\widehat{\alpha}_{it} - \widehat{\alpha}_{i}\right)^{2}$$

- Intuitively appealing, since sampling error is about how a statistic might vary from one sample to the next if observations were repeated.
- Fama-MacBeth uses variation in the statistic  $\widehat{\lambda}_t$  over time to deduce its variation across samples.

 Note that this technique does not correct for a time-series that is autocorrelated, but your book gives standard formulas for how to do this.

- Note that this technique does not correct for a time-series that is autocorrelated, but your book gives standard formulas for how to do this.
  - Asset return data are usually not highly correlated

- Note that this technique does not correct for a time-series that is autocorrelated, but your book gives standard formulas for how to do this.
  - Asset return data are usually not highly correlated
  - Corporate finance data or other regressions where the cross-sectional estimates are highly correlated over time would require this correction.

- Note that this technique does not correct for a time-series that is autocorrelated, but your book gives standard formulas for how to do this.
  - Asset return data are usually not highly correlated
  - Corporate finance data or other regressions where the cross-sectional estimates are highly correlated over time would require this correction.
- Fama-MacBeth standard errors also do not correct for the fact that the betas are generated regressors. Your book also details the Shanken correction for these standard errors are well.

- Note that this technique does not correct for a time-series that is autocorrelated, but your book gives standard formulas for how to do this.
  - Asset return data are usually not highly correlated
  - Corporate finance data or other regressions where the cross-sectional estimates are highly correlated over time would require this correction.
- Fama-MacBeth standard errors also do not correct for the fact that the betas are generated regressors. Your book also details the Shanken correction for these standard errors are well.
- Historical import: The FM procedure allows for changing betas, which a single unconditional cross-sectional regression or a time-series regression test cannot easily handle.

### Fama and French 1992 and 1993

- The Cross-Section of Expected Stock Returns
  - Journal of Finance (1992)
- Common Risk Factors in the Returns on Stocks and Bonds
  - Journal of Financial Economics (1993)
- Read the articles for next time! They are on Collab.

### End of Today's Lecture.

- That's all for today. Today's material corresponds roughly to Chapter 12 in Cochrane (2005).
- Please read both of the Fama-French articles posted on Collab before next class!