

Kalman Filter

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1. Basic Idea and Terminology

Here's the basic procedure associated with the Kalman Filter:

1. Start with a prior for some variable of interest in the current period, $p(x)$.
2. Observe the current measurement y_t .
3. "Filter" out the noise and compute the filtering distribution: $p_t(x|y)$.
4. Compute the predictive distribution $p_{t+1}(x)$ from the filtering distribution and your model.
5. Increment t by one, and go back to step 1, taking the predictive distribution as your prior.

2. Normal Example

Suppose we want to measure some variable x . We will assume a *prior* that is multivariate normal such that

$$x \sim N(\hat{x}, \Sigma)$$

Next, we "measure" x by matching it to an observable in a *measurement equation*:

$$y = Gx + v \quad v \sim N(0, R)$$

where R is positive definite, while G and R are both 2×2 . This forms the *likelihood*.

We then "filter" out the noise, updating our view of x in light of the data in the filtering step using Bayes' Rule:

$$\begin{aligned} p(x | y) &= \frac{p(y | x) \cdot p(x)}{p(y)} \propto p(y | x) \cdot p(x) \\ &\propto \exp \left\{ -\frac{1}{2} (y - Gx)' R^{-1} (y - Gx) \right\} \exp \left\{ -\frac{1}{2} (x - \hat{x})' \Sigma^{-1} (x - \hat{x}) \right\} \end{aligned} \tag{1}$$

Now let's expand the term the lefthand exponential:

$$\begin{aligned} A &= (y - Gx)' R^{-1} (y - Gx) = (y' - x'G') R^{-1} (y - Gx) \\ &= (y'R^{-1} - x'G'R^{-1}) (y - Gx) \\ &= (y'R^{-1}y - y'R^{-1}Gx - x'G'R^{-1}y + x'G'R^{-1}Gx) \end{aligned}$$

And now the same for the righthand exponential:

$$\begin{aligned} B &= (x - \hat{x})' \Sigma^{-1} (x - \hat{x}) = (x' - \hat{x}') \Sigma^{-1} (x - \hat{x}) \\ &= (x' \Sigma^{-1} - \hat{x}' \Sigma^{-1}) (x - \hat{x}) \\ &= x' \Sigma^{-1} x - x' \Sigma^{-1} \hat{x} - \hat{x}' \Sigma^{-1} x + \hat{x}' \Sigma^{-1} \hat{x} \end{aligned} \quad (2)$$

Adding the two exponentials, we get:

$$\begin{aligned} C &= A + B = x' (\Sigma^{-1} + G'R^{-1}G) x - x' (\Sigma^{-1} \hat{x} + G'R^{-1}y) - (\hat{x}' \Sigma^{-1} + y'R^{-1}G) x \\ &\quad + \hat{x}' \Sigma^{-1} \hat{x} + y'R^{-1}y \end{aligned}$$

Now notice that Expression 1 is the probability distribution of x *conditional* on y and pretty much anything else that isn't x . And because of the wonderful properties of the exponential function and the black-hole nature of the proportionality constant, we'll be able to simplify things nicely (and we'll worry that the distribution $p(x | y)$ integrates to one later on).

Specifically, in the expression for C , the two terms in the second row *don't* depend upon x . Therefore, letting $C(x)$ be the portion of C that depends upon x , and letting $C(\neg x)$ be the additive terms which don't depend upon x , we can simplify

$$\begin{aligned} p(x | y) &\propto \exp \left\{ -\frac{1}{2} C \right\} = \exp \left\{ -\frac{1}{2} [C(x) + C(\neg x)] \right\} \\ &\propto \exp \left\{ -\frac{1}{2} C(x) \right\} + \exp \left\{ -\frac{1}{2} C(\neg x) \right\} \\ &\propto \exp \left\{ -\frac{1}{2} C(x) \right\} \end{aligned}$$

We just absorb the portion not relevant to $p(x | y)$ into the proportionality constant. This means our the work we did above to get C simplifies our target expression to

$$p(x | y) \propto \exp \left\{ -\frac{1}{2} [x' (\Sigma^{-1} + G'R^{-1}G) x - x' (\Sigma^{-1} \hat{x} + G'R^{-1}y) - (\hat{x}' \Sigma^{-1} + y'R^{-1}G) x] \right\} \quad (3)$$

Now this doesn't look too helpful, but with a little bit of work, we can turn this into the probability distribution for a multivariate normal random variable. So let's do it.

First, the variance of the normal distribution corresponding to $p(x \mid y)$ can be derived by examining Equation 3 and likening it to Equation 2 (which has the contents of the exponential in the prior distribution of x). Namely, the inverse of the new variance, which we'll denote as Σ_F will be sandwiched in between x' and x in Equation 3, just as it was sandwiched between x' and x in Equation 2. We use this fact, along with the the Woodbury matrix identity, stated in the appendix, to derive:

$$\begin{aligned} \Sigma_F &= (\Sigma^{-1} + G'R^{-1}G)^{-1} \\ \text{Woodbury Identity} \Rightarrow &= \Sigma - \Sigma G'(R + G\Sigma G')^{-1}G\Sigma \end{aligned}$$

Next, we want to get the mean of the distribution of $p(x|y)$. Again, once we take a second and compare Expression 3 to Expression 2, it's becomes clear from inspection that we must have

$$(\Sigma^{-1}\hat{x} + G'R^{-1}y) = (\Sigma^{-1} + G'R^{-1}G) Z \quad (4)$$

To see this, liken the lefthand side of Equation 4 to the result of the matrix multiplication $\Sigma^{-1}\hat{x}$ in Equation 2. To get the righthand side, use the fact that we *know* the Equation 4 analogue to Equation 2's Σ^{-1} , which we just derived.

So all that's left to do is solve for Z in Equation 4. And so we solve Equation 2 by using the Woodbury matrix identity representation from above:

$$\begin{aligned} (\Sigma^{-1}\hat{x} + G'R^{-1}y) &= (\Sigma^{-1} + G'R^{-1}G) Z \\ \Rightarrow Z &= (\Sigma^{-1} + G'R^{-1}G)^{-1} (\Sigma^{-1}\hat{x} + G'R^{-1}y) \\ Z &= (\Sigma - \Sigma G'(R + G\Sigma G')^{-1}G\Sigma) (\Sigma^{-1}\hat{x} + G'R^{-1}y) \end{aligned}$$

Now let's simplify this representation a bit, particularly using the trick $(AB)^{-1} =$

$B^{-1}A^{-1}$ and especially in reverse:

$$\begin{aligned}
Z &= (\Sigma - \Sigma G'(R + G\Sigma G')^{-1}G\Sigma) (\Sigma^{-1}\hat{x} + G'R^{-1}y) \\
&= \Sigma\Sigma^{-1}\hat{x} + \Sigma G'R^{-1}y - \Sigma G'(R + G\Sigma G')^{-1}G\Sigma\Sigma^{-1}\hat{x} \\
&\quad - \Sigma G'(R + G\Sigma G')^{-1}G\Sigma G'R^{-1}y \\
&= \hat{x} + (\Sigma G'R^{-1} - \Sigma G'(R + G\Sigma G')^{-1}G\Sigma G'R^{-1}) y \\
&\quad - \Sigma G'(R + G\Sigma G')^{-1}G\hat{x} \\
&= \hat{x} + (\Sigma G'R^{-1} - (G'^{-1}\Sigma^{-1})^{-1}(R + G\Sigma G')^{-1}G\Sigma G'R^{-1}) y \\
&\quad - \Sigma G'(R + G\Sigma G')^{-1}G\hat{x} \\
&= \hat{x} + \left(\Sigma G'R^{-1} - \{ (R + G\Sigma G')(G'^{-1}\Sigma^{-1}) \}^{-1} G\Sigma G'R^{-1} \right) y \\
&\quad - \Sigma G'(R + G\Sigma G')^{-1}G\hat{x} \\
&= \hat{x} + \left(\Sigma G'R^{-1} - \{ RG'^{-1}\Sigma^{-1} + G\Sigma G'G'^{-1}\Sigma^{-1} \}^{-1} G\Sigma G'R^{-1} \right) y \\
&\quad - \Sigma G'(R + G\Sigma G')^{-1}G\hat{x} \\
&= \hat{x} + \left(\Sigma G'R^{-1} - \{ RG'^{-1}\Sigma^{-1} + G\Sigma^{-1} \}^{-1} G\Sigma G'R^{-1} \right) y \\
&\quad - \Sigma G'(R + G\Sigma G')^{-1}G\hat{x} \\
&= \hat{x} + \left(\Sigma G'R^{-1} - \{ G^{-1}(RG'^{-1}\Sigma^{-1} + G\Sigma^{-1}) \}^{-1} \Sigma G'R^{-1} \right) y \\
&\quad - \Sigma G'(R + G\Sigma G')^{-1}G\hat{x} \\
&= \hat{x} + \left(\Sigma G'R^{-1} - \{ G^{-1}RG'^{-1}\Sigma^{-1} + \Sigma^{-1} \}^{-1} \Sigma G'R^{-1} \right) y \\
&\quad - \Sigma G'(R + G\Sigma G')^{-1}G\hat{x}
\end{aligned}$$

A. Woodbury Matrix Identity

For matrices A , U , C , and V :

$$(A + UCV)^{-1} = A^{-1} - A^{-1}U(C^{-1} + VA^{-1}U)^{-1}VA^{-1} \quad (5)$$

Now consider the special case we have above with the Kalman filter:

$$(A + V'CV)^{-1} = A^{-1} - A^{-1}V'(C^{-1} + VA^{-1}V')^{-1}VA^{-1} \quad (6)$$

We'll use the trick $(AB)^{-1} = B^{-1}A^{-1}$ often, and typically in reverse, going from the right side of the equality to the left to simplify Equation 6.

$$\begin{aligned} (A + V'CV)^{-1} &= A^{-1} - A^{-1}V'(C^{-1} + VA^{-1}V')^{-1}VA^{-1} \\ &= A^{-1} - (V'^{-1}A)^{-1}(C^{-1} + VA^{-1}V')^{-1}VA^{-1} \\ &= A^{-1} - \{(C^{-1} + VA^{-1}V')(V'^{-1}A)\}^{-1}VA^{-1} \\ &= A^{-1} - \{C^{-1}(V'^{-1}A) + VA^{-1}V'(V'^{-1}A)\}^{-1}VA^{-1} \\ &= A^{-1} - \{C^{-1}(V'^{-1}A) + V\}^{-1}(AV^{-1})^{-1} \\ &= A^{-1} - \{(AV^{-1})[C^{-1}(V'^{-1}A) + V]\}^{-1} \\ &= A^{-1} - \{AV^{-1}C^{-1}(V'^{-1}A) + AV^{-1}V\}^{-1} \\ &= A^{-1} - \{AV^{-1}C^{-1}V'^{-1}A + A\}^{-1} \\ &= A^{-1} - \{AV^{-1}C^{-1}V'^{-1}A + A\}^{-1} \end{aligned}$$