Time Series Analysis & Forecasting Using R



Outline

- 1 ARIMA models
- 2 Lab Session 16
- 3 Seasonal ARIMA models
- 4 Lab Session 17
- 5 Forecast ensembles

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AR: autoregressive (lagged observations as inputs)

I: integrated (differencing to make series stationary)

MA: moving average (lagged errors as inputs)

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An ARIMA model is rarely interpretable in terms of visible data structures like trend and seasonality. But it can capture a huge range of time series patterns.

Stationarity

Definition

If $\{y_t\}$ is a stationary time series, then for all s, the distribution of (y_t, \ldots, y_{t+s}) does not depend on t.

Stationarity

Definition

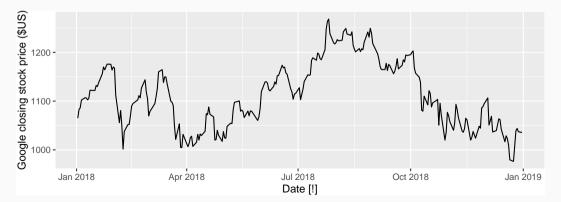
If $\{y_t\}$ is a stationary time series, then for all s, the distribution of (y_t, \ldots, y_{t+s}) does not depend on t.

A **stationary series** is:

- roughly horizontal
- constant variance
- no patterns predictable in the long-term

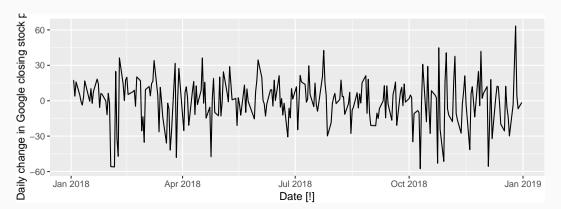
Stationary?

```
gafa_stock ▷
  filter(Symbol = "GOOG", year(Date) = 2018) ▷
  autoplot(Close) +
  labs(y = "Google closing stock price ($US)")
```



Stationary?

```
gafa_stock ▷
  filter(Symbol = "G00G", year(Date) = 2018) ▷
  autoplot(difference(Close)) +
  labs(y = "Daily change in Google closing stock price")
```



Differencing

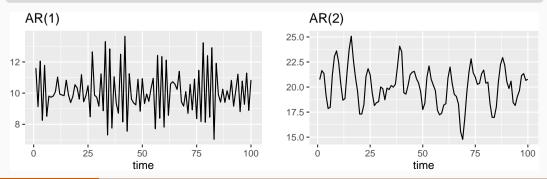
- Differencing helps to **stabilize the mean**.
- The differenced series is the *change* between each observation in the original series.
- Occasionally the differenced data will not appear stationary and it may be necessary to difference the data a second time.
- In practice, it is almost never necessary to go beyond second-order differences.

Autoregressive models

Autoregressive (AR) models:

$$y_t = c + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \varepsilon_t,$$

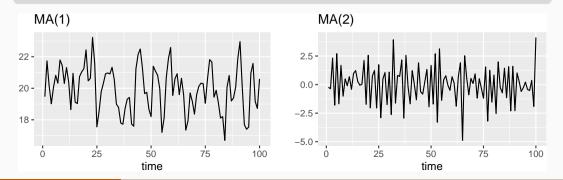
where ε_t is white noise. A multiple regression with **lagged values** of y_t as predictors.



Moving Average (MA) models

Moving Average (MA) models:

 $y_t = c + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2} + \cdots + \theta_q \varepsilon_{t-q},$ where ε_t is white noise. A multiple regression with **lagged** errors as predictors. Don't confuse with moving average smoothing!



Autoregressive Moving Average models:

$$y_{t} = c + \phi_{1}y_{t-1} + \dots + \phi_{p}y_{t-p} + \theta_{1}\varepsilon_{t-1} + \dots + \theta_{q}\varepsilon_{t-q} + \varepsilon_{t}.$$

Autoregressive Moving Average models:

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Predictors include both lagged values of y_t and lagged errors.

Autoregressive Moving Average models:

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Predictors include both lagged values of y_t and lagged errors.

Autoregressive Integrated Moving Average models

- Combine ARMA model with **differencing**.
- *d*-differenced series follows an ARMA model.
- Need to choose p, d, q and whether or not to include c.

ARIMA(p, d, q) model

- AR: p =order of the autoregressive part
 - I: d =degree of first differencing involved
- MA: q =order of the moving average part.
 - White noise model: ARIMA(0,0,0)
 - Random walk: ARIMA(0,1,0) with no constant
 - Random walk with drift: ARIMA(0,1,0) with const.
 - \blacksquare AR(p): ARIMA(p,0,0)
 - \blacksquare MA(q): ARIMA(0,0,q)

```
fit <- global economy ▷
 model(arima = ARIMA(Population))
fit
## # A mable: 263 x 2
## # Key: Country [263]
##
      Country
                                                   arima
##
      <fct>
                                                <model>
    1 Afghanistan
                                        <ARIMA(4,2,1)>
###
    2 Albania
                                        <ARIMA(0,2,2)>
###
    3 Algeria
                                        \langle ARIMA(2,2,2) \rangle
##
    4 American Samoa
                                        \langle ARIMA(2,2,2) \rangle
###
                              <ARIMA(2,1,2) w/ drift>
##
    5 Andorra
    6 Angola
                                        \langle ARIMA(4,2,1) \rangle
##
    7 Antigua and Barbuda <ARIMA(2,1,2) w/ drift>
```

44 0 Amah Wamid

```
fit ▷
 filter(Country = "Australia") ▷
 report()
## Series: Population
## Model: ARIMA(0,2,1)
###
## Coefficients:
###
           ma1
       -0.661
###
## s.e. 0.107
###
## sigma^2 estimated as 4.063e+09: log likelihood=-699
## AIC=1401 AICc=1402 BIC=1405
```

```
fit ▷
  filter(Country = "Australia") ▷
  report()
## Series: Population
## Model: ARIMA(0,2,1)
###
## Coefficients:
                                    y_t = 2y_{t-1} - y_{t-2} - 0.7\varepsilon_{t-1} + \varepsilon_t
                                                    \varepsilon_t \sim \mathsf{NID}(0,4 \times 10^9)
###
              ma1
## -0.661
## s.e. 0.107
##
## sigma^2 estimated as 4.063e+09: log likelihood=-699
## ATC=1401
                ATCc=1402 BTC=1405
```

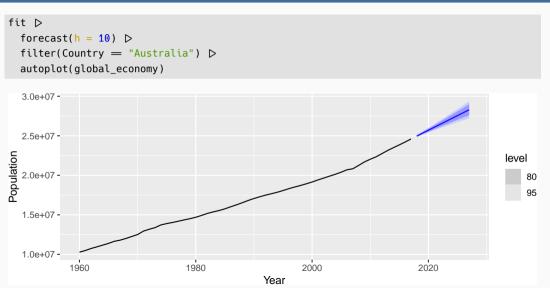
Understanding ARIMA models

- If c = 0 and d = 0, the long-term forecasts will go to zero.
- If c = 0 and d = 1, the long-term forecasts will go to a non-zero constant.
- If c = 0 and d = 2, the long-term forecasts will follow a straight line.
- If $c \neq 0$ and d = 0, the long-term forecasts will go to the mean of the data.
- If $c \neq 0$ and d = 1, the long-term forecasts will follow a straight line.
- If $c \neq 0$ and d = 2, the long-term forecasts will follow a quadratic trend.

Understanding ARIMA models

Forecast variance and *d*

- The higher the value of *d*, the more rapidly the prediction intervals increase in size.
- For d = 0, the long-term forecast standard deviation will go to the standard deviation of the historical data.



Hyndman and Khandakar (JSS, 2008) algorithm:

- Select no. differences *d* via KPSS test.
- Select *p*, *q* and inclusion of *c* by minimising AICc.
- Use stepwise search to traverse model space.

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AICc =
$$-2\log(L) + 2(p+q+k+1)\left[1 + \frac{(p+q+k+2)}{T-p-q-k-2}\right]$$
.

where *L* is the maximised likelihood fitted to the *differenced* data, k = 1 if $c \neq 0$ and k = 0 otherwise.

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.

where *L* is the maximised likelihood fitted to the *differenced* data, k = 1 if $c \neq 0$ and k = 0 otherwise.

Note: Can't compare AICc for different values of d.

Step1: Select current model (with smallest AICc) from:

ARIMA(2, d, 2)

ARIMA(0, d, 0)

ARIMA(1, d, 0)

ARIMA(0, d, 1)

- **Step 2:** Consider variations of current model:
 - vary one of p, q, from current model by ± 1 ;
 - p, q both vary from current model by ± 1 ;
 - Include/exclude *c* from current model.

Model with lowest AICc becomes current model.

```
Step1: Select current model (with smallest AICc) from:
```

ARIMA(2, d, 2)

ARIMA(0, d, 0)

ARIMA(1, d, 0)

ARIMA(0, d, 1)

Step 2: Consider variations of current model:

- vary one of p, q, from current model by ± 1 ;
- p, q both vary from current model by ± 1 ;
- Include/exclude *c* from current model.

Model with lowest AICc becomes current model.

Repeat Step 2 until no lower AICc can be found.

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Lab Session 16

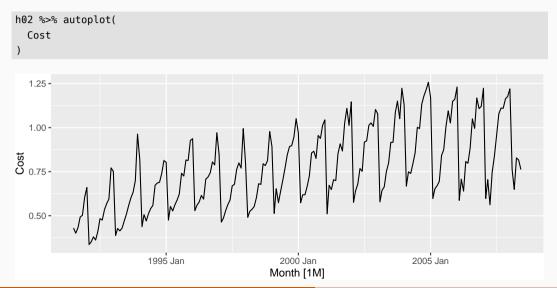
For the United States GDP data (from global_economy):

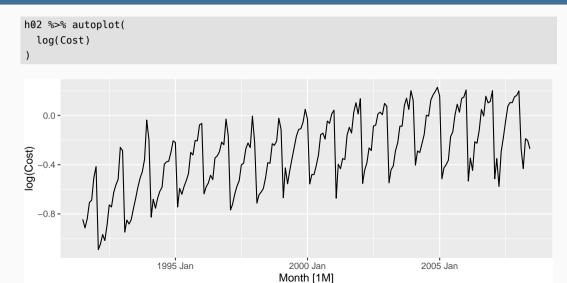
- Fit a suitable ARIMA model for the logged data.
- Produce forecasts of your fitted model. Do the forecasts look reasonable?

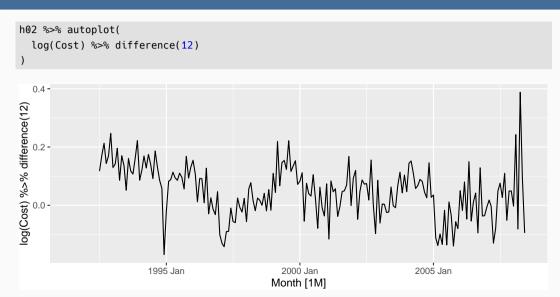
Outline

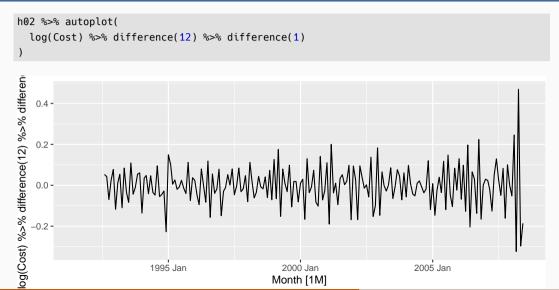
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```
h02 <- PBS %>%
filter(ATC2 = "H02") %>%
summarise(Cost = sum(Cost) / 1e6)
```







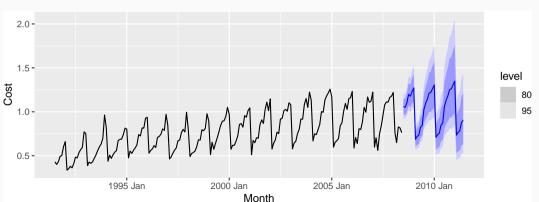


Example: US electricity production

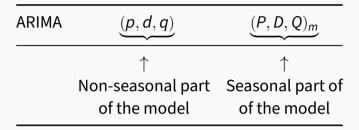
```
h02 ⊳
 model(arima = ARIMA(log(Cost))) >
 report()
## Series: Cost
## Model: ARIMA(2,1,0)(0,1,1)[12]
## Transformation: log(Cost)
##
## Coefficients:
                  ar2
###
             ar1
                              sma1
###
        -0.8491 -0.4207 -0.6401
## s.e. 0.0712 0.0714 0.0694
###
## sigma^2 estimated as 0.004387: log likelihood=245
## ATC=-483 ATCc=-483 BTC=-470
```

Example: US electricity production





Seasonal ARIMA models



- \blacksquare m = number of observations per year.
- *d* first differences, *D* seasonal differences
- p AR lags, q MA lags
- P seasonal AR lags, Q seasonal MA lags

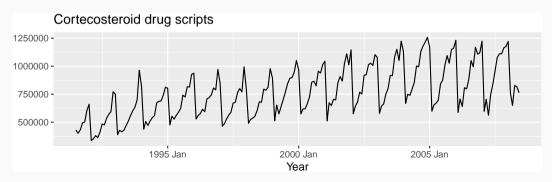
Seasonal and non-seasonal terms combine multiplicatively

Common ARIMA models

The US Census Bureau uses the following models most often:

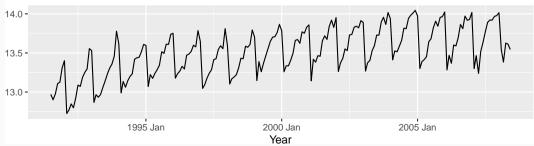
$ARIMA(0,1,1)(0,1,1)_m$	with log transformation
$ARIMA(0,1,2)(0,1,1)_m$	with log transformation
$ARIMA(2,1,0)(0,1,1)_m$	with log transformation
$ARIMA(0,2,2)(0,1,1)_m$	with log transformation
$ARIMA(2,1,2)(0,1,1)_m$	with no transformation

```
h02 <- PBS ▷
  filter(ATC2 = "H02") ▷
  summarise(Cost = sum(Cost))
h02 ▷ autoplot(Cost) +
  labs(x = "Year", y = "", title = "Cortecosteroid drug scripts")</pre>
```



```
h02 <- PBS ▷
  filter(ATC2 = "H02") ▷
  summarise(Cost = sum(Cost))
h02 ▷ autoplot(log(Cost)) +
  labs(x = "Year", y = "", title = "Log Cortecosteroid drug scripts")</pre>
```

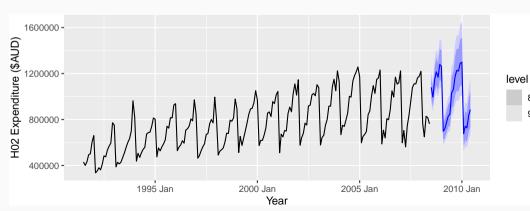
Log Cortecosteroid drug scripts



```
fit <- h02 >
  model(auto = ARIMA(log(Cost)))
report(fit)
## Series: Cost
## Model: ARIMA(2,1,0)(0,1,1)[12]
## Transformation: log(Cost)
###
## Coefficients:
                     ar2
                             sma1
###
            ar1
       -0.8491 -0.4207 -0.6401
###
## s.e. 0.0712 0.0714 0.0694
##
## sigma^2 estimated as 0.004399: log likelihood=245
## ATC=-483
             ATCc = -483
                         BIC = -470
```

```
fit <- h02 >
 model(best = ARIMA(log(Cost),
    stepwise = FALSE,
    approximation = FALSE,
    order constraint = p + q + P + 0 \leq 9
  ))
report(fit)
## Series: Cost
## Model: ARIMA(4,1,1)(2,1,2)[12]
## Transformation: log(Cost)
##
  Coefficients:
                                        ma1 sar1
###
            ar1
                  ar2
                         ar3
                                ar4
                                                     sar2
                                                             sma1
                                                                   sma2
        -0.0426 0.210 0.202
                             -0.227 -0.742 0.621
                                                   -0.383
##
                                                           -1.202
                                                                  0.496
## s.e. 0.2167 0.181 0.114
                             0.081
                                      0.207 0.242
                                                    0.118
                                                            0.249
                                                                  0.214
###
## sigma^2 estimated as 0.004061: log likelihood=254
## AIC=-489 AICc=-487
                        BIC=-456
```

```
fit D
forecast() D
autoplot(h02) +
labs(y = "H02 Expenditure ($AUD)", x = "Year")
```



80 95

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Lab Session 17

For the Australian tourism data (from tourism):

- Fit a suitable ARIMA model for all data.
- Produce forecasts of your fitted models.
- Check the forecasts for the "Snowy Mountains" and "Melbourne" regions. Do they look reasonable?

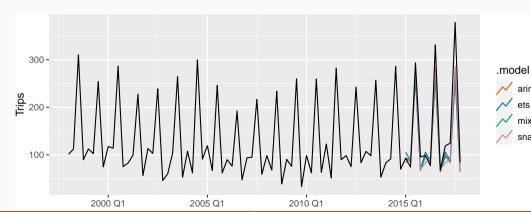
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```
train <- tourism >
  filter(year(Quarter) \leq 2014)
fit <- train >
  model(
  ets = ETS(Trips),
   arima = ARIMA(Trips),
   snaive = SNAIVE(Trips)
) >
  mutate(mixed = (ets + arima + snaive) / 3)
```

- Ensemble forecast mixed is a simple average of the three fitted models.
- forecast() will produce distributional forecasts taking into account the correlations between the forecast errors of the component models.

```
fc <- fit ▷ forecast(h = "3 years")</pre>
fc ▷
  filter(Region = "Snowy Mountains", Purpose = "Holiday") ▷
  autoplot(tourism, level = NULL)
```



arima ets mixed snaive

```
accuracy(fc, tourism)  
group_by(.model)  
summarise(
    RMSE = mean(RMSE),
    MAE = mean(MAE),
    MASE = mean(MASE)
)  
arrange(RMSE)
```

```
## # A tibble: 4 x 4
## .model RMSE MAE MASE
## <chr> <dbl> <dbl> <dbl> <dbl> <dbl> 
## 1 mixed 19.8 16.0 0.997
## 2 ets 20.2 16.4 1.00
## 3 snaive 21.5 17.3 1.17
## 4 arima 21.9 17.8 1.06
```

Can we do better than equal weights?

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- Hard to find weights that improve forecast accuracy.
- Known as the "forecast combination puzzle".
- Solution: FFORMA

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FFORMA (Feature-based FORecast Model Averaging)

- Vector of time series features used to predict best weights.
- A modification of xgboost is used.
- Method came 2nd in the 2018 M4 international forecasting competition.
- Main author: Pablo Montero-Manso (Monash U)
- Not (yet) available for fable.