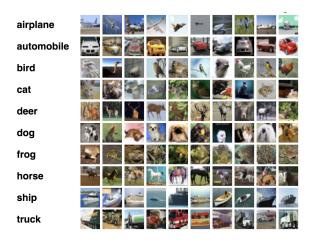
# Mathematics of Machine Learning - Summer School

Lecture 9
High-Dimensional Statistics. Gaussian Complexity

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## Recall. Offline Statistical Learning: Prediction



## Offline learning: prediction

Given a batch of observations (images & labels) interested in predicting the label of a new image

/12

## Recall. Offline Statistical Learning: Prediction

- 1. Observe training data  $Z_1, \ldots, Z_n$  i.i.d. from <u>unknown</u> distribution
- 2. Choose action  $A \in \mathcal{A} \subseteq \mathcal{B}$
- 3. Suffer an expected/population loss/risk r(A), where

$$a \in \mathcal{B} \longrightarrow r(a) := \mathbf{E}\,\ell(a,Z)$$

with  $\ell$  is an **prediction** loss function and Z is a new test data point

**Goal:** Minimize the estimation error defined by the following decomposition

$$\underbrace{r(A) - \inf_{a \in \mathcal{B}} r(a)}_{\text{excess risk}} = \underbrace{r(A) - \inf_{a \in \mathcal{A}} r(a) + \inf_{a \in \mathcal{A}} r(a) - \inf_{a \in \mathcal{B}} r(a)}_{\text{estimation error}} + \underbrace{\inf_{a \in \mathcal{A}} r(a) - \inf_{a \in \mathcal{B}} r(a)}_{\text{approximation error}}$$

as a function of n and notions of "complexity" of the set  $\mathcal A$  of the function  $\ell$ 

Note: Estimation/Approximation trade-off, a.k.a. complexity/bias

# Offline Statistical Learning: Estimation



## Offline learning: estimation

User 1

User 2

User 3

Given a batch of observations (users & ratings) interested in  $\underbrace{\text{estimating}}$  the missing ratings in a recommendation system

3/12

## Offline Statistical Learning: Estimation

- 1. Observe training data  $Z_1,\ldots,Z_n$  i.i.d. from distr. parametrized by  $a^\star\in\mathcal{A}$
- 2. Choose a parameter  $A \in \mathcal{A}$
- 3. Suffer a loss  $\ell(A, a^*)$  where  $\ell$  is an **estimation** loss function

**Goal:** Minimize the estimation loss  $\ell(A, a^*)$  as a function of n and notions of "complexity" of the set  $\mathcal A$  of the function  $\ell$ 

### Main differences:

- No test data (i.e., no population risk r). Only training data
- Underlying distribution is not completely unknown We consider a parametric model

Remark: We could also consider prediction losses with a new test data...

# Supervised Learning. High-Dimensional Estimation

1. Observe training data  $Z_1=(x_1,Y_1),\ldots,Z_n=(x_n,Y_n)\in\mathbb{R}^d\times\mathbb{R}$  i.i.d. from distr. parametrized by  $w^*\in\mathbb{R}^d$ :

$$Y_i = \langle x_i, \boldsymbol{w}^* \rangle + \sigma \xi_i$$
  $i \in [n]$   
 $Y = \mathbf{x} \boldsymbol{w}^* + \sigma \xi$  (data in matrix form:  $Y \in \mathbb{R}^n$  and  $\mathbf{x} \in \mathbb{R}^{n \times d}$ )

- 2. Choose a parameter  $W \in \mathcal{W}$
- 3. **Goal:** Minimize loss  $\ell(W, \mathbf{w}^*) = \|W \mathbf{w}^*\|_2$

**High-dimensional setting:** n < d (dimension greater than no. of data)

Assumptions (otherwise problem is ill-posed):

- ▶ Sparsity:  $\|w^*\|_0 := \sum_{i=1}^d 1_{|w_i^*|>0} \le k$
- **Low-rank:** Rank $(w^*) \le k$ , when  $w^*$  can be thought of as a matrix

## Non-Convex Estimator. Restricted Eigenvalue Condition

Assume that we know k, the upper bound on the sparsity  $(\|\mathbf{w}^{\star}\|_{0} \leq k)$ 

Algorithm:

$$W^0 := \underset{w:||w||_0 \le k}{\operatorname{argmin}} \frac{1}{2n} ||\mathbf{x}w - Y||_2^2$$

## Restricted eigenvalues (Assumption 12.2)

There exists  $\alpha>0$  such that for any vector  $w\in\mathbb{R}^d$  with  $\|w\|_0\leq 2k$  we have

$$\frac{1}{2n} \|\mathbf{x}w\|_2^2 \ge \alpha \|w\|_2^2$$

## Statistical Guarantees $\ell_0$ Recovery (Theorem 12.5)

If the restricted eigenvalue assumption holds, then

$$\|W^0 - \boldsymbol{w}^*\|_2 \le \sqrt{2} \frac{\sigma \sqrt{k}}{\alpha} \frac{\|\mathbf{x}^\top \boldsymbol{\xi}\|_{\infty}}{n}$$

## Proof of Theorem 12.5

▶ Let  $\Delta = W^0 - w^*$ . By the definition of  $W^0$ , we have

$$\|\mathbf{x}\Delta - \sigma\xi\|_2^2 = \|\mathbf{x}W^0 - Y\|_2^2 \le \|\mathbf{x}w^\star - Y\|_2^2 = \|\sigma\xi\|_2^2$$

so that, expanding the square, we find the basic inequality:

$$\|\mathbf{x}\Delta\|_2^2 \le 2\sigma \langle \mathbf{x}\Delta, \xi \rangle$$

▶ The restricted eigenvalue assumption yields, noticing that  $\|\Delta\|_0 \le 2k$ :

$$\alpha \|\Delta\|_2^2 \le \frac{1}{2n} \|\mathbf{x}\Delta\|_2^2 \le \frac{\sigma}{n} \langle \mathbf{x}\Delta, \xi \rangle = \frac{\sigma}{n} \langle \Delta, \mathbf{x}^\top \xi \rangle \le \frac{\sigma}{n} \|\Delta\|_1 \|\mathbf{x}^\top \xi\|_{\infty}$$

where the last inequality follows from Hölder's inequality.

▶ The proof follows by applying the Cauchy-Swartz's inequality:

$$\|\Delta\|_1 = \langle \operatorname{sign}(\Delta), \Delta \rangle \le \|\operatorname{sign}(\Delta)\|_2 \|\Delta\|_2 \le \sqrt{2k} \|\Delta\|_2$$

# Bounds in Expectation. Gaussian Complexity

Recall: 
$$\|W^0 - \mathbf{w}^{\star}\|_2 \leq \sqrt{2} \frac{\sigma \sqrt{k}}{\alpha} \frac{\|\mathbf{x}^{\top} \boldsymbol{\xi}\|_{\infty}}{n}$$

## Gaussian complexity (Definition 12.6)

The Gaussian complexity of a set  $\mathcal{T} \subseteq \mathbb{R}^n$  is defined as

$$\mathsf{Gauss}(\mathcal{T}) := \mathbf{E} \sup_{t \in \mathcal{T}} \frac{1}{n} \sum_{i=1}^{n} \xi_i t_i$$

where  $\xi_1, \dots, \xi_n$  are i.i.d. standard Gaussian random variables

$$\mathbf{E}\frac{\|\mathbf{x}^{\top}\boldsymbol{\xi}\|_{\infty}}{n} = \mathtt{Gauss}(\mathcal{A}_1 \circ \{x_1, \dots, x_n\})$$



## Proof of Corollary 12.7

▶ The  $\ell_{\infty}$  norm is the dual of the  $\ell_1$  norm:  $\|\mathbf{x}^{\top}\xi\|_{\infty} = \sup_{u \in \mathbb{R}^d: \|u\|_1 \le 1} \langle \mathbf{x}u, \xi \rangle$ 

Hölder's inequality yields  $\langle \mathbf{x}u, \xi \rangle = \langle u, \mathbf{x}^{\top} \xi \rangle \leq \|u\|_1 \|\mathbf{x}^{\top} \xi\|_{\infty}$  for any u, so

$$\|\mathbf{x}^{\top}\xi\|_{\infty} \ge \sup_{u \in \mathbb{R}^d: \|u\|_1 \le 1} \langle \mathbf{x}u, \xi \rangle$$

On the other hand, note that the choice  $u = e_j$ ,  $j \in [d]$ , satisfies  $||u||_1 = 1$  and yields  $\langle \mathbf{x}e_j, \xi \rangle = \langle e_j, \mathbf{x}^\top \xi \rangle = (\mathbf{x}^\top \xi)_j$ , so that the inequality is achieved by at least one of the vectors  $e_j$ ,  $j \in [d]$ .

We have

$$\langle \mathbf{x}u, \xi \rangle = \sum_{i=1}^{n} (\mathbf{x}u)_{i} \xi_{i} = \sum_{i=1}^{n} \langle u, x_{i} \rangle \xi_{i}$$

SO

$$\frac{1}{n}\mathbf{E}\|\mathbf{x}^{\top}\xi\|_{\infty} = \mathbf{E}\sup_{u \in \mathbb{R}^{d}: \|u\|_{1} \leq 1} \frac{1}{n} \sum_{i=1}^{n} \xi_{i} \langle u, x_{i} \rangle = \operatorname{Gauss}(\mathcal{A}_{1} \circ \{x_{1}, \dots, x_{n}\})$$



# Bounds in Probability. Gaussian Concentration

Recall: 
$$\|W^0 - \mathbf{w}^{\star}\|_2 \leq \sqrt{2} \frac{\sigma \sqrt{k}}{\alpha} \frac{\|\mathbf{x}^{\top} \boldsymbol{\xi}\|_{\infty}}{n}$$

## Column normalization (Assumption 12.8)

$$\mathbf{c}_{jj} = \left(\frac{\mathbf{x}^{\top} \mathbf{x}}{n}\right)_{jj} = \frac{1}{n} \sum_{i=1}^{n} x_{ij}^{2} \le 1$$

If the column normalization assumption holds, then

$$\mathbf{P}\bigg(\frac{\|\mathbf{x}^{\top}\boldsymbol{\xi}\|_{\infty}}{n} < \sqrt{\frac{\tau \log d}{n}}\bigg) \ge 1 - \frac{2}{d^{\tau/2 - 1}}.$$

# Proof of Corollary 12.9 (Part I)

Let  $V = \frac{\mathbf{x}^{\top} \boldsymbol{\xi}}{\sqrt{n}} \in \mathbb{R}^d$ . As each coordinate  $V_i$  is a linear combination of Gaussian random variables, V is a Gaussian random vector with mean

$$\mathbf{E}V = \frac{1}{\sqrt{n}}\mathbf{x}^{\top}\mathbf{E}\xi = 0$$

and covariance matrix given by

$$\mathbf{E}[VV^{\top}] = \frac{1}{n}\mathbf{E}[\mathbf{x}^{\top}\xi\xi^{\top}\mathbf{x}] = \frac{1}{n}\mathbf{x}^{\top}\mathbf{E}[\xi\xi^{\top}]\mathbf{x} = \frac{\mathbf{x}^{\top}\mathbf{x}}{n} = \mathbf{c}$$

as  $\xi$  is made of independent standard Gaussian components, so  $\mathbf{E}[\xi\xi^{\top}]=I$ 

That is,  $V \sim \mathcal{N}(0, \mathbf{c})$  and, in particular, the *i*-th component has distribution  $V_i \sim \mathcal{N}(0, \mathbf{c}_{ii})$ . By the union bound

$$\mathbf{P}\left(\frac{\|\mathbf{x}^{\top}\boldsymbol{\xi}\|_{\infty}}{\sqrt{n}} \geq \varepsilon\right) = \mathbf{P}(\|V\|_{\infty} \geq \varepsilon) = \mathbf{P}\left(\max_{i \in [n]} |V_{i}| \geq \varepsilon\right)$$
$$= \mathbf{P}\left(\bigcup_{i=1}^{d} \{|V_{i}| \geq \varepsilon\}\right) \leq \sum_{i=1}^{d} \mathbf{P}(|V_{i}| \geq \varepsilon) \leq d \max_{i \in [d]} \mathbf{P}(|V_{i}| \geq \varepsilon)$$

# Proof of Corollary 12.9 (Part II)

▶ By concentration for sub-Gaussian random variables (Proposition 6.6) and Assumption 12.8 we have

$$\mathbf{P}(|V_i| \ge \varepsilon) \le 2e^{-\frac{\varepsilon^2}{2\mathbf{c}_{ii}}} \le 2e^{-\frac{\varepsilon^2}{2}}$$

▶ Putting everything together we obtain

$$\mathbf{P}\left(\frac{\|\mathbf{x}^{\top}\boldsymbol{\xi}\|_{\infty}}{\sqrt{n}} \ge \varepsilon\right) \le 2de^{-\frac{\varepsilon^2}{2}}$$

By setting  $\varepsilon=\sqrt{\tau\log d}$  for  $\tau>2$ , we have  $2de^{-\frac{\varepsilon^2}{2}}=\frac{2}{d^{\tau/2-1}}$  so that

$$\mathbf{P}\left(\frac{\|\mathbf{x}^{\top}\boldsymbol{\xi}\|_{\infty}}{n} < \sqrt{\frac{\tau \log d}{n}}\right) \ge 1 - \frac{2}{d^{\tau/2 - 1}}$$