

# Lévy processes and approximation with compound Poisson process: theory, implementation and code.

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## Abstract

In this article, referring to the book of Cont and Tankov [2], we detailed show how to approximate an infinite activity Lévy process with a compound Poisson process. We also provide an application to the gamma process and we implement the algorithm in Python.

## 1 Introduction

In this article we provide a simple example of the approximation of an infinite activity Lévy process by a compound Poisson process. We mainly refer to Cont and Tankov [2, Chapter 6.3]. We try also to make this work self-contained, providing all the necessary elements we need to understand this topic. We will show an application fo this approximation to the gamma process and we also provide the Python code which implements from scratch the simulation algorithm.

## 2 Preliminaries

In this section we introduce all the theory we need to better understand the example we provide in Section 3. The interested reader can refer to Cont and Tankov [2] and Schoutens [6] to better understand applications of Lévy processes to finance or to Sato [5] and Applebaum [1] for a more theoretical discussion of the topic.

### 2.1 Poisson processes and its counting measure

**Definition 2.1.** *Cont and Tankov [2, Definition 2.2] Let  $E \subset \mathbb{R}^d$ . A Radon measure on  $(E, \mathcal{B}(E))$  is a measure  $\mu$  such that for every compact<sup>1</sup> measurable set  $B \in \mathcal{B}(\mathbb{R}^d)$ ,  $\mu(B) < \infty$ .*

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<sup>1</sup>A compact subset of  $\mathbb{R}^d$  is simply a bounded closed subset.

Consider now a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ . Starting from a given positive Radon measure  $\mu$  on a general measurable space  $(E, \mathcal{E})$  we can construct a Poisson random measure which is an integer valued measure on  $\Omega \times \mathcal{E} \rightarrow \mathbb{N}$  “counting the number of events occurring in a given element of the  $\sigma$ -algebra  $\mathcal{E}$ ”. The measure is Random since it depends on  $\omega \in \Omega$ .

In order to better understand this concept, we first consider a simple example, by referring to the Poisson process (see Cont and Tankov [2, Chapter 2.5]).

**Definition 2.2.** Let  $(\tau_i)_{i \geq 1}$  be a sequence of independent exponential random variables with parameter  $\lambda > 0$  and  $T_n = \sum_{i=1}^n \tau_i$ . The process  $(N_t)_{t \geq 0}$  defined by:

$$N_t = \sum_{n \geq 1} \mathbb{1}_{T_n \leq t} \quad (1)$$

is called *Poisson process with intensity  $\lambda$* .

The Poisson process is a counting process: it counts the number of random times  $T_n$  which occur in the time interval  $[0, t]$ . Hence:

$$N_t = \# \{i \geq 1, T_i \in [0, t]\}.$$

The jumps times  $T_1, T_2, \dots$  form a random configuration of points in  $[0, \infty)$  and the process  $N_t$  counts the number of such points in the interval  $[0, t]$ . This counting procedure can be used to define a measure  $M$  on  $\mathbb{R}^+$ : for any measurable set  $A \subset \mathbb{R}^+$  let:

$$M(\omega, A) = \# \{i \geq 1, T_i \in A\}. \quad (2)$$

With fixed  $\omega \in \Omega$ , namely fixing the random variables  $T_i(\omega)$ ,  $M(\omega, \cdot)$  is a positive, integer valued measure and  $M(\cdot, A)$  is finite with probability 1 for every bounded set  $A$  (this follow from the first point of Cont and Tankov [2, Proposition 2.12]). Note that  $M(\omega, \cdot)$  depends on  $\omega$ : it is thus a *random measure*.

Observe also that, from Cont and Tankov [2, Proposition 2.12] we have that  $N_t$  follows a Poisson distribution with parameter  $\lambda t$ . Hence

$$\mathbb{E}[M(\cdot, [s, t])] = \lambda(t - s),$$

or for a more general measurable set  $A \subset \mathbb{R}^+$

$$\mathbb{E}[M(\cdot, A)] = \lambda|A|,$$

where  $|A|$  denotes the Lebesgue measure of  $A$ .  $M$  is called the random jump measure<sup>2</sup> associated to the Poisson process  $\{N_t\}_{t \geq 0}$ .

The Poisson process can be expressed in terms of the random measure  $M$  in the following way:

$$N_t(\omega) = M(\omega, [0, t]) = \int_{[0, t]} M(\omega, ds).$$

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<sup>2</sup>Actually, it is a Poisson random measure and it is just a special case of the construction we derive in Section 2.3

*Remark 1.* The notation  $\int_{[0,t]} M(\omega, ds)$  means that you are counting the events occurring on the interval  $[0, t]$  according to the random measure  $M$ . Observe that this quantity is a random variable (it depends on  $\omega$ ). Intuitively, starting from a general random measure  $M$  for each  $t > 0$  it defines a random variable and hence it can be used to define a stochastic process.

*Remark 2.* The properties of the Poisson process translate into the following properties for  $M$ . For disjoint intervals  $[t_1, t'_1], \dots, [t_n, t'_n]$ :

- $M([t_k, t'_k])$  is the number of jumps of the Poisson process in  $[t_k, t'_k]$ : it is a Poisson random variable with parameter  $\lambda(t'_k - t_k)$ .
- If  $j \neq i$ :  $M([t_k, t'_k])$  and  $M([t_i, t'_i])$  are independent random variables.
- More generally for any measurable set  $A$ ,  $M(\cdot, A)$  follows a Poisson distribution with parameter  $\lambda|A|$  where  $|A|$  is the Lebesgue measure of  $A$ .

Hence we have a random configuration of points in  $\mathbb{R}^+$  and for any  $A \subset \mathbb{R}^+$  measurable the measure  $M$  counts the number of these random points which belongs to  $A$ .

**We have defined the Poisson process, which induces a counting (random) measure  $M$  on  $\mathbb{R}^+$  which, for every measurable  $A \subset \mathbb{R}^+$ , counts the number of random points belonging to  $A$ .**

## 2.2 Integration with respect a counting measure

In this section we give an intuition of what does it means integrate with respect an integer measure. Consider  $(E, \mathcal{E}, \nu)$ ,  $E \subset \mathbb{R}$ , with  $\nu : E \rightarrow \mathbb{N}$  be a counting measure, namely for each  $A$  measurable subset of  $E$   $\nu(A)$  represents the cardinality of  $A$ . Consider  $f : E \rightarrow \mathbb{R}$ . We want to define  $\int_E f d\nu$ . Assume  $f$  is a simple function of the form:

$$f = \sum_{i=1}^n c_i \mathbb{1}_{A_i}.$$

Then we define the integral in the usual way:

$$\int_E f d\nu = \sum_{i=1}^n c_i \nu(A_i).$$

We have that this is equal to:

$$\int_E f d\nu = \sum_{k \geq 1} f(x_k).$$

where  $E = \bigcup_{k \geq 1} \{x_k\}$ . Hence by proceeding in the usual way we can define the integral for  $f$  positive and hence for general functions  $f : E \rightarrow \mathbb{R}$ .

### 2.3 Poisson random measure

Is it possible, given a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ , a measurable space  $(E, \mathcal{E}, \mu)$ , with  $E \subset \mathbb{R}^d$  and  $\mu$  a (positive) Radon measure, to construct an integer valued random measure:

$$\begin{aligned} M : \Omega \times \mathcal{E} &\rightarrow \mathbb{N}, \\ (\omega, A) &\mapsto M(\omega, A). \end{aligned}$$

such that its properties generalize the properties we obtained when we consider the random measure associated to a Poisson process? The answer is yes, and it turns out that the random measure we obtained in the previous chapter is just a special case of this construction when we take  $E = \mathbb{R}^+$  and  $\mu$  as the Lebesgue measure on  $\mathbb{R}^+$ .

First we define what a Poisson random measure is.

**Definition 2.3.** *Cont and Tankov [2, Definition 2.18] Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space,  $E \subset \mathbb{R}^d$  and  $\mu$  a given (positive) Radon measure on  $(E, \mathcal{E})$ . A Poisson random measure on  $E$  with intensity measure  $\mu$  is an integer valued random measure:*

$$\begin{aligned} M : \Omega \times \mathcal{E} &\rightarrow \mathbb{N}, \\ (\omega, A) &\mapsto M(\omega, A). \end{aligned}$$

such that:

1. For (almost all)  $\omega \in \Omega$ ,  $M(\omega, \cdot)$  is an integer-valued Radon measure on  $E$ : for any bounded measurable  $A \subset E$ ,  $M(\cdot, A) < \infty$  is an integer-valued random variable.
2. For each measurable set  $A \subset E$ ,  $M(\cdot, A) = M(A)$  is a Poisson random variable with parameter  $\mu(A)$ :

$$\forall k \in \mathbb{N} : \quad \mathbb{P}(M(A) = k) = e^{-\mu(A)} \frac{(\mu(A))^k}{k!}.$$

3. For disjoint measurable sets  $A_1, \dots, A_n \in \mathcal{E}$ , the variables  $M(A_1), \dots, M(A_n)$  are independent.

The following proposition shows how to construct a Poisson random measure starting from a general Radon measure  $\mu$ .

**Proposition 2.1.** *Cont and Tankov [2, Proposition 2.18] For any Radon measure  $\mu$  on  $E \subset \mathbb{R}^d$ , there exists a Poisson random measure  $M$  on  $E$  with intensity  $\mu$ .*

*Proof.* We consider only the case when  $\mu(E) < \infty$ . For the other case see Cont and Tankov [2, Proposition 2.18].

1. Consider a sequence of iid random variables  $\{X_i\}_{i \geq 1}$  so that  $\mathbb{P}(X_i \in A) = \frac{\mu(A)}{\mu(E)}$ .
2. Take  $M(\cdot, E)$  be a Poisson random variable on  $(\Omega, \mathcal{F}, \mathbb{P})$  with mean  $\mu(E)$  and independent of the  $X_i$ .

3. Define:

$$M(\cdot, A) = \sum_{i=1}^{M(\cdot, E)} \mathbb{1}_A(X_i).$$

$M$  is a Poisson random measure with intensity measure  $\mu$ . For simplicity denote  $M(A) = M(\cdot, A)$  and  $M(E) = M(\cdot, E)$ .

1. Fix  $\omega \in \Omega$ .  $M(\omega, A) = \sum_{i=1}^{M(\omega, E)} \mathbb{1}_A(X_i(\omega)) < \infty$  since the sum is finite.  $M(\omega, A) \in \mathbb{N}$  by definition.
2. Fix  $k \in \mathbb{N}$ .

$$\begin{aligned} \mathbb{P}(M(A) = k) &= \mathbb{P}\left(\sum_{i=1}^{M(E)} \mathbb{1}_A(X_i)\right) \\ &= \sum_{n=0}^{\infty} \mathbb{P}\left(\sum_{i=1}^n \mathbb{1}_A(X_i) = k \mid M(E) = n\right) \mathbb{P}(M(\cdot, E) = n) \\ &= \sum_{n=k}^{\infty} \binom{n}{k} \left(\frac{\mu(A)}{\mu(E)}\right)^k \left(1 - \frac{\mu(A)}{\mu(E)}\right)^{n-k} \frac{\mu(E)^n}{n!} e^{-\mu(E)} \\ &= \frac{\mu(A)^k}{k!} e^{-\mu(A)}. \end{aligned}$$

3. Independence follows from the fact that the joint probability factorizes, namely if  $A_1 \cap A_2 = \emptyset$  then:

$$\begin{aligned} &\mathbb{P}(M(A_1) = k, M(A_2) = m) \\ &= \sum_{n \geq k+m} \mathbb{P}\left(\sum_{i=1}^n \mathbb{1}_{A_1}(X_i) = k, \sum_{i=1}^n \mathbb{1}_{A_2}(X_i) = m\right) \mathbb{P}(M(E) = n) \\ &= \sum_{n \geq k+m} \left(\frac{\mu(A_1)}{\mu(E)}\right)^k \left(\frac{\mu(A_2)}{\mu(E)}\right)^m \binom{n}{k} \binom{n-k}{m} \left(1 - \frac{\mu(A_1)}{\mu(E)} - \frac{\mu(A_2)}{\mu(E)}\right)^{n-k-m} \frac{e^{-\mu(E)} \mu(E)^n}{n!} \\ &= \frac{\mu(A_1)^k}{k!} \frac{\mu(A_2)^m}{m!} e^{-\mu(A_1)} e^{-\mu(A_2)} \\ &= \mathbb{P}(M(A_1) = k) \cdot \mathbb{P}(M(A_2) = m). \end{aligned}$$

■

The above proof shows that any Poisson random measure on  $E$  can be represented as counting measure associate to a random sequence of points in  $E$ : there exists  $\{X_n(\omega)\}_{n \geq 1}$  such that:

$$\forall A \in \mathcal{E}, \quad M(\omega, A) = \sum_{n \geq 1} \mathbb{1}_A(X_n(\omega)) = \sum_{n \geq 1} \delta_{X_n(\omega)}(A).$$

$M$  is thus the sum of Dirac masses located at random points  $(X_n)_{n \geq 1}$ .

*Remark 3.* For the Poisson process defined in Equation (1),  $E = [0, T] \times \{1\}$  and the sequence is given by  $(T_n, 1)_{n \geq 1}$ . It represents a sequence of points of size 1 located at time  $T_n$ . For an intuitive introduction to Poisson point process see Lanchier [4].

## 2.4 Building jump processes from Poisson random measure: Marked point process

Consider  $(\Omega, \mathcal{F}, \mathbb{P})$ ,  $E = [0, T] \times \mathbb{R}^d \setminus \{0\}$  and a given random measure  $M$  on  $E$ . As in the proof of Proposition 2.1, the random measure can be described as the counting measure associated to a configuration of points  $(T_n, Y_n) \in E$ :

$$M(\omega, A) = \sum_{n \geq 1} \delta_{(T_n(\omega), Y_n(\omega))}(A), \quad A \in \mathcal{E}.$$

The random sequence  $(T_n, Y_n)$  is called marked point process.

**Definition 2.4.** A marked point process on  $(\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0}, \mathbb{P})$  is a sequence  $(T_n, Y_n)_{n \geq 1}$  where:

- $(T_n)_{n \geq 1}$  is an increasing sequence of non anticipating random times with  $T_n \rightarrow \infty$  almost surely when  $n \rightarrow \infty$ .
- $(Y_n)_{n \geq 1}$  is a sequence of random variables taking values in  $\mathbb{R}^d \setminus \{0\}$ .
- The value of  $Y_n$  is revealed at  $T_n$ :  $Y_n$  is  $\mathcal{F}_{T_n}$  measurable.

The first condition guarantees that the number of events in an interval of the form  $[0, T]$  is finite.

Fixed  $\omega \in \Omega$ ,  $M(\omega, \cdot)$  defines a measure on  $E$ . Hence we can construct integrals with respect to this measure. Consider  $f : E \rightarrow \mathbb{R}$ .

- $f : E \rightarrow \mathbb{R}^+$  where  $f = \sum_{i=1}^n c_i \mathbb{1}_{A_i}$ , with  $c_i \geq 0$  and  $A_i \subset E$  disjoint measurable subset of  $E$ . Define the integral as:

$$M(f) = \sum_{i=1}^n c_i M(A_i).$$

**Observe that  $M(f)$  is a random variable** with expectation given by<sup>3</sup>:

$$\mathbb{E}[M(f)] = \sum_{i=1}^n c_i \mu(A_i).$$

For each  $\omega \in \Omega$  you have a different configuration of points  $(T_n(\omega), Y_n(\omega))_{n \geq 1}$  in  $E$ . For each measurable  $A_i \subset E$  you simply count the number of points belonging to  $A_i$ , by using  $M$ . If you change  $\omega$  the configuration of points changes and also  $M(\omega, A_i)$  changes and hence the value of the integral  $M(f)$  consequently.

If  $f_n = \sum_{j=1}^{m(n)} c_j \mathbb{1}_{A_j}$  is simple we can define the integral with respect to  $\mu$  as:

$$\mu(f_n) = \sum_{j=1}^{m(n)} c_j \mu(A_j),$$

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<sup>3</sup>Remember that a Poisson random measure has intensity measure  $\mu$  defined on  $(E, \mathcal{E})$ .

and hence we can define  $\mu(f) = \sup \{\mu(\phi), \phi \text{ simple function}, \phi \leq f\}$  for a general positive function  $f : E \rightarrow [0, \infty)$ .

It follows that:

$$\mathbb{E}[M(f_n)] = \sum_{i=1}^{m(n)} c_i \mu(A_i) = \mu(f_n).$$

- For  $f : E \rightarrow [0, \infty)$  you can take  $(f_n)_{n \geq 1}$  simple increasing to  $f$  and define:

$$M(f) = \lim_{n \rightarrow \infty} M(f_n).$$

$M(f)$  is a random variable since it is the limit of random variables. It follows that:

$$\mathbb{E}[M(f)] = \mathbb{E}\left[\lim_{n \rightarrow \infty} M(f_n)\right] = \lim_{n \rightarrow \infty} \mathbb{E}[M(f_n)] = \lim_{n \rightarrow \infty} \mu(f_n) = \mu(f),$$

where the limit is to be considered as the one we use to define the integral of  $f$ . Expectation and limit can be exchanged because of the monotone convergence theorem.

- If  $f : E \rightarrow R$  it can be decompose as:  $f = f_+ - f_-$ . We also request that:

$$\mu(|f|) = \int_{[0, T] \times \mathbb{R}^d \setminus \{0\}} |f(s, y)| \mu(ds \times dy) < \infty.$$

Thanks to this condition the random variables  $M(f_+)$  and  $M(f_-)$  have finite expectation and hence are almost surely finite. We can define  $M(f) = M(f_+) - M(f_-)$ .  $M(f)$  is a random variable and its expectation is given by:

$$\mathbb{E}[M(f)] = \mu(f) = \int_{[0, T] \times \mathbb{R}^d \setminus \{0\}} f(s, y) \mu(ds \times dy).$$

Integrating  $f$  with respect to  $M$  up to time  $t$  yields to a nonanticipating stochastic process:

$$X_t = \int_{[0, t] \times \mathbb{R}^d \setminus \{0\}} f(s, y) M(ds \times dy) = \sum_{\{n, T_n \in [0, t]\}} f(T_n, Y_n).$$

Hence, starting from a Poisson random measure we have constructed a jump process.  $(X_t(f))_{t \in [0, T]}$  is a jump process whose jumps happen at random time  $T_n$  and have amplitude given by  $f(T_n, Y_n)$ .

## 2.5 Jump measure of a cadlag process

So far we have seen that, given a point process we can construct a Poisson random measure (see proof of proposition 2.1).

Given a cadlag process  $(X_t)_{t \in [0, T]}$  with values in  $\mathbb{R}^d$  one can associate a random measure  $J_X$  on  $[0, T] \times \mathbb{R}^d$  called the jump measure in the following way. Since  $X$  is a cadlag process it has a countable number of jumps:  $\{t \in [0, T], \Delta X_t = X_t - X_{t-} \neq 0\}$  is countable and its elements can be arranged in a sequence  $\{T_n\}_{n \in \mathbb{N}}$  which are the (random) jump times of  $X$ . At time  $T_n$ , the process  $X$  has a discontinuity  $Y_n = X_{T_n} - X_{T_n-} \in \mathbb{R}^d \setminus \{0\}$ .

Hence  $(T_n, Y_n)_{n \geq 1}$  defines a marked point process on  $[0, T] \times \mathbb{R} \setminus \{0\}$  which contains all the information about the jumps of the process  $X$ : jumps time  $T_n$  and jump sizes  $Y_n$ . The associated random measure, which is denoted by  $J_X$ , is called **jump measure** of the process  $X$ :

$$J_X(\omega, \cdot) = \sum_{n \geq 1} \delta_{(T_n(\omega), Y_n(\omega))} = \sum_{t \in [0, T], \Delta X_t \neq 0} \delta_{t, \Delta X_t}.$$

Intuitively, for any measurable set  $A \subset \mathbb{R}^d$ :

$J_X([0, T] \times A)$  = number of jumps of  $X$  occurring between 0 and  $t$  whose amplitude belongs to  $A$ .

The random measure  $J_X$  does not contains information about the continuous part of  $X$ , but it describes only its jumps, in particular when they occur and how big they are.

All quantities involving the jumps of  $X$  can be computed by integrating various functions against  $J_X$ . For example if  $f(t, y) = y^2$  we have:

$$\int_{[0, T] \times \mathbb{R}} y^2 J_X(dt \times dy) = \sum_{t \in [0, T]} (\Delta X_t)^2.$$

These expression may involve infinite sums and hence its convergence must be defined in some sense.

**Example 2.1.** Consider a Poisson process. Its jumps measure is given by:

$$J_N = \sum_{n \geq 1} \delta_{(T_n, 1)},$$

namely:

$$J_N([0, t] \times A) = \begin{cases} \# \{i \geq 1, T_i \in [0, t]\} & \text{if } 1 \in A \\ 0 & \text{if } 1 \notin A. \end{cases} \quad (3)$$

## 2.6 Lévy processes: compound Poisson process

In this section we briefly outline some important properties of Lévy processes. The interested reader can refer to Cont and Tankov [2, Chapter 3].

There is an important connection between Lévy processes and infinitely divisible laws.

**Proposition 2.2.** Let  $(X_t)_{t \geq 0}$  be a Lévy process. Then for every  $t$ ,  $X_t$  has an infinitely divisible distribution. Conversely if  $F$  is an infinitely divisible distribution, then there exists a Lévy process  $(X_t)_{t \geq 0}$  such that the distribution of  $X_1$  is given by  $F$ .

Moreover, the characteristic function of a Lévy process has specific form.

**Proposition 2.3.** Let  $(X_t)_{t \geq 0}$  be a Lévy process on  $\mathbb{R}^d$ . There exists a continuous function  $\psi : \mathbb{R}^d \rightarrow \mathbb{R}$  called the characteristic exponent of  $X$ , such that:

$$\mathbb{E}[e^{iu \cdot X_t}] = \exp\{t\psi(u)\}, \quad u \in \mathbb{R}^d.$$

A particular Lévy process is called the Compound Poisson process.



**Definition 2.5.** A compound Poisson process with intensity  $\lambda > 0$  and jump size distribution  $f$  is a stochastic process  $(X_t)_{t \geq 0}$  defined as:

$$X_t = \sum_{i=1}^{N_t} Y_i, \quad (4)$$

where jumps sizes  $Y_i$  are iid with distribution  $f$  and  $\{N(t), t \geq 0\}$  is a Poisson process with intensity  $\lambda$ , independent from  $(Y_i)_{i \geq 1}$ .

The characteristic function of a compound Poisson process is given by (see Cont and Tankov [2, Proposition 3.4]):

$$\mathbb{E} [e^{iu \cdot X_t}] = \exp \left\{ t \lambda \int_{\mathbb{R}^d} (e^{iu \cdot x} - 1) f(dx) \right\}, \quad u \in \mathbb{R}^d. \quad (5)$$

Introducing a new measure  $\nu(A) = \lambda f(A)$  Formula (6) can be rewritten as:

$$\mathbb{E} [e^{iu \cdot X_t}] = \exp \left\{ \lambda \int_{\mathbb{R}^d} (e^{iu \cdot x} - 1) \nu(dx) \right\}, \quad u \in \mathbb{R}^d. \quad (6)$$

$\nu$  is called Lévy measure of the process  $(X_t)_{t \geq 0}$ .

Since a compound Poisson process  $(X_t)_{t \geq 0}$  is a cadlag continuous time jump process we can associate to it its jump measure, as observed in the previous section. This measure  $J_X$  fully describes the structure of jumps. It is a random measure on  $[0, \infty) \times \mathbb{R}^d$  and for any measurable set  $B \subset \mathbb{R}^d \times [0, \infty)$  we have:

$$J_X(B) = \# \{(t, X_t - X_{t-}) \in B\}.$$

In particular, for any measurable set  $A \subset \mathbb{R}^d$ ,  $J_X([t_1, t_2] \times A)$  counts the number of jumps of  $X$  in the interval  $[t_1, t_2]$  which size belongs to  $A$ .

**Proposition 2.4.** Let  $(X_t)_{t \geq 0}$  be a compound Poisson process with intensity  $\lambda$  and jump size distribution  $f$ . Its jump measure  $J_X$  is a Poisson random measure on  $\mathbb{R}^d \times [0, \infty)$  with intensity measure  $\mu(dx \times dt) = \nu(dx)dt = \lambda f(dx)dt$ .

This proposition suggest to interpret the Lévy measure as the average number of jumps per unit of time.

**Definition 2.6.** Let  $(X_t)_{t \geq 0}$  be a Lévy process on  $\mathbb{R}^d$ . The measure  $\nu$  on  $\mathbb{R}^d$  defined by:

$$\nu(A) = \mathbb{E} [\# \{t \in [0, 1] : \Delta X_t \neq 0, \Delta X_t \in A\}], \quad A \in \mathcal{B}(\mathbb{R}^d).$$

is called Lévy measure of  $X$ .  $\nu(A)$  is the expected number, per unit time, of jumps whose size belongs to  $A$ .

Proposition 2.4 implies that every compound Poisson process can be represented in the following form:

$$X_t = \sum_{s \in [0, t]} \Delta X_s = \int_{[0, t] \times \mathbb{R}^d} x J_x(ds \times dx), \quad (7)$$

where  $J_X$  is a Poisson random measure with intensity measure  $\nu(dx)dt$ . This is a special case of the **Lévy-Itô decomposition**: we have only rewritten the process  $X$  as the sum of its jumps. Since a compound Poisson process has almost surely a finite number of jumps in interval  $[0, t]$  the stochastic integral is just a finite sum and hence there are no convergence problems.

## 2.7 Lévy-Itô decomposition

Consider  $\{X_t^0\}_{t \geq 0}$  represented as is Equation (7). Consider its Lévy measure defined by:

$$\nu(A) = \mathbb{E} [\# \{t \in [0, 1] : \Delta X_t^0 \neq 0, \Delta X_t^0 \in A\}], \quad A \in \mathcal{B}(\mathbb{R}^d).$$

Take a Brownian motion with drift  $\gamma t + W_t$ , independent from  $X_0$ .  $X_t = X_t^0 + \gamma t + W_t$  defines another Lévy process which can be represented as:

$$X_t = \gamma t + W_t + \sum_{s \in [0, t]} \Delta X_s = \gamma t + W_t + \int_{[0, t] \times \mathbb{R}^d} x J_x(ds \times dx), \quad (8)$$

where  $J_X$  is a Poisson random measure on  $[0, \infty) \times \mathbb{R}^d$  with intensity measure  $\nu(dx)dt$ .

Given a general Lévy process  $X_t$ , we can still define its Lévy measure  $\nu$  as above.  $\nu(A)$  is still finite for any compact set  $A$  not containing 0: if this were not true the process would have an infinite number of jumps of finite size on  $[0, T]$  which contradicts the cadlag property. So  $\nu$  defines a Radon measure on  $\mathbb{R} \setminus \{0\}$ . But  $\nu$  is not necessarily a finite measure: the above restriction still allows it to blow up at zero and  $X$  may have an infinite number of small jumps on  $[0, T]$ . In this case the sum of the jumps becomes an infinite series and its convergence imposes some conditions on the measure  $\nu$ , under which we obtain a decomposition of  $X$  similar to the one above.

**Proposition 2.5.** *Lévy-Itô decomposition* Let  $(X_t)_{t \geq 0}$  be a Lévy process on  $\mathbb{R}^d$  and  $\nu$  its Lévy measure.

- $\nu$  is a Radon measure on  $\mathbb{R}^d \setminus \{0\}$  and verifies:

$$\int_{|x| \leq 1} |x|^2 \nu(dx) < \infty, \quad \int_{|x| > 1} \nu(dx) < \infty.$$

- The jump measure of  $X$ , denoted by  $J_X$ , is a Poisson random measure on  $[0, \infty) \times \mathbb{R}^d$  with intensity measure  $\nu(dx)dt$ .
- There exists a vector  $\gamma$  and a  $d$ -dimensional Brownian motion  $(B_t)_{t \geq 0}$  with covariance matrix  $A$  such that:

$$X_t = \gamma t + B_t + X_t^l + \lim_{\epsilon \rightarrow 0} \tilde{X}_t^\epsilon.$$

where:

$$\begin{aligned} X_t^l &= \int_{|x| \geq 1, s \in [0, t]} x J_X(ds \times dx), \\ \tilde{X}_t^\epsilon &= \int_{\epsilon \leq |x| \leq 1, s \in [0, t]} x \{J_X(ds \times dx) - \nu(dx)ds\}. \end{aligned}$$

All these terms are independent and the convergence of the last term is almost sure and uniform in  $t$  in  $[0, T]$ .

$(\nu, A, \gamma)$  is unique and it is called the Lévy triplet of the process  $X$ .

The so called Lévy-Khinchin representation follows from this result.

**Proposition 2.6.** *Let  $(X_t)_{t \geq 0}$  be a Lévy process on  $\mathbb{R}^d$  with characteristic triplet  $(A, \nu, \gamma)$ . Then:*

$$\mathbb{E} [e^{iz \cdot X_t}] = e^{t\psi(z)}, \quad z \in \mathbb{R}^d, \quad (9)$$

$$\psi(z) = -\frac{1}{2}z \cdot Az + i\gamma \cdot z + \int_{\mathbb{R}^d} (e^{iz \cdot x} - 1 - iz \cdot x \mathbb{1}_{|x| \leq 1}) \nu(dx) \quad (10)$$

$\gamma$  is not the drift and it depends on the choice of the truncation function of the large jumps, which in this case is given by  $g(x) = \mathbb{1}_{|x| \leq 1}$ .

When  $\nu(\mathbb{R}) = \infty$  (infinite activity case), the set of jump times on every trajectory of the Lévy process is countably finite and dense in  $[0, \infty)$ . Intuitively, an infinite activity Lévy process has a countably infinite number of small jumps in each time interval  $[t_1, t_2]$ .

## 2.8 Path properties of Lévy processes

Properties of sample paths of a Lévy process  $X$  can be deduced by analytical properties of its characteristic triplet  $(A, \gamma, \nu)$ .

**Proposition 2.7.** *Lévy process with piece-wise constant trajectories*

*A Lévy process has piecewise constant trajectories if and only if its characteristic triplet satisfies the following conditions:  $A = 0$ ,  $\int_{\mathbb{R}^d} \nu(dx) < \infty$  and  $\gamma = \int_{|x| \leq 1} x \nu(dx)$  or, equivalently, if its characteristic exponent is of the form:*

$$\psi(z) = \int_{-\infty}^{\infty} (e^{iuz} - 1) \nu(du), \quad \text{with } \nu(\mathbb{R}) < \infty.$$

*Remark 4.* Recall that almost all trajectories of a Lévy process are piecewise constant if and only if it is of compound Poisson type (see Cont and Tankov [2, Proposition 3.3]).

A Lévy process is said to be of finite variation if its trajectories are functions of finite variation with probability 1.

**Proposition 2.8.** *Lévy process of finite variation A Lévy process is of finite variation if and only if its characteristic triplet  $(A, \gamma, \nu)$  satisfies:*

$$A = 0, \quad \int_{|x| \leq 1} |x| \nu(dx) < \infty.$$

In this case we have also that if  $\{X_t\}_{t \geq 0}$  is a Lévy process with characteristic function  $(\nu, 0, \gamma)$ :

$$X_t = bt + \int_{[0, t] \times \mathbb{R}^d} x J_X(ds \times dx) = bt + \sum_{s \in [0, t], \Delta X_s \neq 0} \Delta X_s.$$

and its characteristic function can be expressed as:

$$\mathbb{E} [e^{iz \cdot X_t}] = \exp \left\{ ib \cdot z + \int_{\mathbb{R}^d} (e^{iz \cdot x} - 1) \nu(dx) \right\},$$

where  $b = \gamma - \int_{|x| \leq 1} x \nu(dx)$ . Note that the characteristic triplet is not given by  $(b, 0, \nu)$  but by  $(\gamma, 0, \nu)$ . In fact,  $\gamma$  is not an intrinsic quantity and depends on the truncation function used in the Lévy-Khinchin representation while  $bt$  as an intrinsic interpretation as the continuous part of  $X$ .

**Proposition 2.9. Subordinators** Let  $(X_t)_{t \geq 0}$  be a Lévy process on  $\mathbb{R}$ . The following conditions are equivalent:

- $X_t \geq 0$  a.s. for some  $t \geq 0$ .
- $X_t \geq 0$  a.s. for every  $t \geq 0$ .
- Sample paths of  $X$  are almost surely non decreasing:  $s \leq t \implies X_s \leq X_t$ .
- The characteristic triplet of  $(X_t)_{t \geq 0}$  satisfies the following conditions:  $A = 0$ ,  $\nu((-\infty, 0]) = 0$ ,  $\int_0^\infty (x \wedge 1) \nu(dx) < \infty$  and  $b \geq 0$ . Hence  $X$  has no diffusion component, only positive jumps of finite variation and positive drift.

### 3 Approximation of an infinite activity Lévy process by a compound Poisson

Let  $(X_t)_{t \geq 0}$  be an infinite activity Lévy process with characteristic triplet  $(0, \nu, \gamma)$ . From the Lévy-Itô decomposition we know that  $X$  can be represented as a sum of compound Poisson process and an almost sure limit of compensated compound processes:

$$X_t = \gamma t + \sum_{s \leq t} \Delta X_s \mathbb{1}_{|\Delta X_s| \geq 1} + \lim_{\epsilon \rightarrow 0} N_t^\epsilon,$$

where

$$N_t^\epsilon = \sum_{s \leq t} \Delta X_s \mathbb{1}_{\epsilon \leq |\Delta X_s| \leq 1} - t \int_{\epsilon \leq |x| \leq 1} x \nu(dx).$$

Therefor a natural idea is to approximate  $X$  with:

$$X_t^\epsilon = \gamma t + \sum_{s \leq t} \Delta X_s \mathbb{1}_{|\Delta X_s| \geq 1} + N_t^\epsilon.$$

For a more detailed exposition see Cont and Tankov [2, Chapter 6.3].

#### 3.1 The finite variation case

Assume that  $X$  is a Lévy process of finite variation. Hence we have:

$$A = 0 \int_{|x| \leq 1} |x| \nu(dx) < \infty \quad \gamma,$$

$$\mathbb{E} [e^{izX_t}] = \exp \left\{ ibz + \int_{\mathbb{R}^d} (e^{izx} - 1) \nu(dx) \right\},$$

with  $b = \gamma - \int_{|x| \leq 1} x \nu(dx)$ .

Hence we have that:

$$\begin{aligned} X_t^\epsilon &= \gamma t + \sum_{s \leq t} \Delta X_s \mathbb{1}_{|\Delta X_s| \geq 1} + N_t^\epsilon \\ &= \gamma t + \sum_{s \leq t} \Delta X_s \mathbb{1}_{|\Delta X_s| \geq 1} + \sum_{s \leq t} \Delta X_s \mathbb{1}_{\epsilon \leq |\Delta X_s| < 1} - t \int_{\epsilon \leq |x| < 1} \\ &\quad + t \int_{0 < |x| < \epsilon} x \nu(dx) - t \int_{0 < |x| < \epsilon} x \nu(dx) \\ &= bt + \sum_{s \leq t} \Delta X_s \mathbb{1}_{|\Delta X_s| \geq \epsilon} + \int_0^t \int_{|x| < \epsilon} x \nu(dx) ds. \end{aligned}$$

Recall that we can write:

$$\sum_{s \leq t} \Delta X_s \mathbb{1}_{|\Delta X_s| < \epsilon} = \int_{[0, t] \times \mathbb{R}^d} x \mathbb{1}_{|x| < \epsilon} J_X(ds \times dx),$$

and that  $\mathbb{E} [J(\cdot, A)] = \mu(A)$ . Hence we have that:

$$\mathbb{E} \left[ \sum_{s \leq t} \Delta X_s \mathbb{1}_{|\Delta X_s| < \epsilon} \right] = \int_{[0, t] \times \mathbb{R}^d} x \mathbb{1}_{|x| < \epsilon} \nu(dx) ds$$

and finally:

$$X_t^\epsilon = bt + \sum_{s \leq t} \Delta X_s \mathbb{1}_{|\Delta X_s| \geq \epsilon} + \mathbb{E} \left[ \sum_{s \leq t} \Delta X_s \mathbb{1}_{|\Delta X_s| < \epsilon} \right].$$

In the finite variation case the approximation is constructed by replacing small jumps with their expectation. In the next section we show the complete procedure to approximate a gamma process with a compound Poisson process. Other examples can be found in Cont and Tankov [2, Chapter 6.3].

### 3.2 An example: the gamma process

Consider the Gamma process, which characteristic triplet is given by:

$$\gamma = \frac{\alpha t}{\beta} (1 - e^{-\beta}), \quad \sigma = 0, \quad \nu(x) = \alpha x^{-1} e^{-\beta x} \mathbb{1}_{x \geq 0}.$$

Its characteristic exponent is given by:

$$\psi(u) = \int_0^\infty (e^{iux} - 1) \alpha x^{-1} e^{-\beta x} dx = \log \left( 1 - \frac{i u}{\beta} \right)^{-\alpha}.$$

with  $b = \gamma - \int_{|x| \leq 1} x \nu(dx) = 0$ . The process is of finite variation, hence we use the approximation introduced in the previous section:

$$X_t^\epsilon = bt + \sum_{s \leq t} \Delta X_s \mathbb{1}_{|\Delta X_s| \geq \epsilon} + \mathbb{E} \left[ \sum_{s \leq t} \Delta X_s \mathbb{1}_{|\Delta X_s| < \epsilon} \right]. \quad (11)$$

We have to investigate the two terms. The first one is a compound Poisson process (since we have truncated the original Lévy measure around the origin). The second term must be computed, possibly explicitly.

First we truncate the Lévy measure:

$$\nu^\epsilon(x) = \alpha \frac{e^{-\beta x}}{x} \mathbb{1}_{x \geq \epsilon}(x).$$

Since for a compound Poisson process the intensity measure of the jump measure  $J_X$  is given by:

$$\mu(dx \times dt) = \nu(dx)dt = \lambda f(dx)dt,$$

where  $f(dx)$  is the distribution (*cdf*)<sup>4</sup> of jumps, we can obtain the intensity  $\lambda(\epsilon)$  by integrating the Lévy measure over  $\mathbb{R}$ .

$$\lambda(\epsilon) = \int_{\mathbb{R}} \nu^\epsilon(dx) = \alpha E_1(\beta\epsilon),$$

where  $E_1(z)$  is the Exponential integral defined by:

$$E_1(z) = \int_z^\infty \frac{e^{-t}}{t} dt, \quad |\arg z| \leq \pi.$$

On the other hand, the *pdf*  $p^\epsilon(x)$  of jumps can be obtained by:

$$p^\epsilon(dx) = \frac{\nu^\epsilon(dx)}{\lambda(\epsilon)}.$$

In order to simulate from jumps following a distribution with *pdf* given by  $p^\epsilon$ , one can observe that:

$$p^\epsilon(x) = \frac{\alpha e^{-\beta x}}{x} \mathbb{1}_{x \geq \epsilon}(x) \frac{1}{U(\epsilon)} \leq \beta e^{-\beta x} \frac{\alpha}{\epsilon \beta U(\epsilon)} = g(x)M.$$

Hence the *pdf*  $p^\epsilon$  is bounded by the *pdf* of and exponential random variable with parameter  $\beta$ ,  $\mathcal{E}(\beta)$ , multiplied by a constant term  $M = \frac{\alpha}{\epsilon \beta U(\epsilon)}$ . For this reason, in order to sample jumps, a simple acceptance rejection method can be implemented (see Devroye [3]).

Now, what we have to do is to compute the second term in Equation (11). This is very easy and we get:

$$\mathbb{E} \left[ \sum_{s \leq t} \Delta X_s \mathbb{1}_{|\Delta X_s| < \epsilon} \right] = \int_0^t \int_0^\epsilon \frac{x \alpha e^{-\beta x}}{x} dx ds = t b_\epsilon,$$

---

<sup>4</sup>June 25, 2025: Probably here you have to use the distribution of jumps and not the *pdf* but it appears it is the same. Check it.

---

**Algorithm 1** Acceptance-rejection method

---

```
Sample  $y \sim \mathcal{E}(\beta)$ .  
Sample  $u \sim U[0, 1]$   
if  $u \geq \frac{p^\epsilon(y)}{Mg(y)}$  then  
    Return  $u$  as a valid sample from  $p^\epsilon(x)$ .  
else  
    Repeat.  
end if
```

---

where  $b_\epsilon = \frac{\alpha}{\beta} [1 - e^{-\beta\epsilon}]$ .

Finally, the Compound Poisson process with intensity  $\lambda(\epsilon)$ , jumps distribution  $p^\epsilon(x)$  and drift  $b_\epsilon$  can be easily simulated by using Cont and Tankov [2, Algorithm 6.2] which is reported in Algorithm 2.

---

**Algorithm 2** Algorithm for compound Poisson process (with drift  $b$ )

---

- 1: Simulate a random variable  $N$  from Poisson distribution with parameter  $\lambda(\epsilon)T$ .  $N$  gives the total number of jumps in the interval  $[0, T]$ .
- 2: Simulate  $N$  independent *rv*  $U_i$ , uniformly distributed on the interval  $[0, T]$ . These variables correspond to the jump times.
- 3: Simulate jump sizes:  $N$  independent *rv*  $Y_i$  with law  $\frac{\nu(dx)}{\lambda}$ . The trajectory is given by:

$$X_t = bt + \sum_{i=1}^N \mathbb{1}_{U_i < t} Y_i.$$

---

In the next section, we report the python code which implements this approximation and we compare it with the simulation of the gamma process obtained by simulating the increments.

## 4 Python Code Example

The Python code<sup>5</sup> is the following

```
1 import numpy as np
2 import matplotlib.pyplot as plt
3 import scipy
4 from scipy.special import exp1
5 from scipy.special import gamma
6
7
8 def exponential_density(x, mu):
9     """
10     Density of an exponential random variable
11     :param x: point at which the pdf is evaluated
12     :param mu: rate > 0
13     :return:
14     """
15     if x > 0:
16         return mu*np.exp(-x*mu)
17     else:
18         return 0
19
20
21 def gamma_density(x, alpha, beta):
22     """
23
24     :param x: point at which the pdf is evaluated
25     :param alpha: shape parameter
26     :param beta: rate parameter
27     :return:
28     """
29     if x > 0:
30         return (beta**alpha)/gamma(alpha)*x**(alpha-1)*np.exp(-beta
31 *x)
32     else:
33         return 0
34
35 class GammaJumpsDistribution:
36     """
37     This class implements the distribution of jumps
38     """
39
40     def __init__(self, alpha, beta, epsilon):
41         self.alpha = alpha
42         self.beta = beta
43         self.epsilon = epsilon
44
```

---

<sup>5</sup>Please note that this code aims not to be a perfect exercise on object oriented programming nor aims at showing how to implement the algorithm following Python best practices. It is just a didactic example, which try to help to reader to move from theory to the numerical implementation.



```

45     def _density(self, x):
46         """
47         Density of jumps.
48         :param x: point at which the pdf is evaluated
49         :return:
50         """
51
52         # This is what in the article is called lambda(epsilon)
53         u_epsilon = self.alpha * scipy.special.exp1(self.beta*self.
epsilon)
54
55         if x <= self.epsilon:
56             return 0
57         else:
58             return self.alpha/x*np.exp(-self.beta*x)/u_epsilon
59
60     def density(self, x):
61         """
62         Vectorized the density function
63         :param x: points at which the pdf is evaluated
64         :return:
65         """
66         g = np.vectorize(self._density, otypes=[np.ndarray])
67         return g(x)
68
69     def sample(self, n_sim=1):
70         """
71         Sample from the distribution using an acceptance rejection
method
72         :param n_sim: number of sample
73         :return:
74         """
75
76         u_epsilon = self.alpha * scipy.special.exp1(self.beta *
self.epsilon)
77         M = self.alpha/(self.epsilon*self.beta*u_epsilon)
78         x = np.zeros(n_sim)
79
80         for i in range(0, n_sim):
81             sampled = False
82             while not sampled:
83                 y = np.random.exponential(scale=1/self.beta)
84                 u = np.random.uniform(low=0, high=1, size=1)
85                 if exponential_density(y,self.beta) > 0:
86                     if u <= self._density(y)/(M*exponential_density
(y,self.beta)):
87                         x[i] = y
88                         sampled = True
89
90         return x
91
92     def plot_density(self, x, x_sample=None):

```

```

93         """
94         Given empirical data x_sample the function plots and
95         histogram of empirical data and compares it with the
96         theoretical probability density function.
97         :param x: point at which the density has to be plotted
98         :param x_sample: empirical data
99         :return:
100         """
101         f = self.density(x)
102         plt.plot(x, f)
103
104         if not (x_sample is None):
105             plt.hist(x_sample, bins=500, density=True, edgecolor='
106             black', alpha=0.7)
107
108             plt.xlabel('x')
109             plt.ylabel('f(x)')
110             plt.title('pdf')
111             plt.grid(True)
112             plt.show()
113
114     class GammaProcess:
115         """
116         This class implements the gamma process
117         """
118
119         def __init__(self, alpha, beta, T):
120             self.alpha = alpha
121             self.beta = beta
122             self.T = T
123             self.sim = None
124             self.time_grid = None
125             self.number_jumps_smaller_than_epsilon = None
126             self.epsilon = 0
127
128         def set_epsilon(self, epsilon):
129             """
130             Set the truncation parameter
131             :param epsilon:
132             :return:
133             """
134             self.epsilon = epsilon
135
136
137         def simulate(self, n_sim, n_steps):
138             """
139             Simulate from a gamma process
140             :param n_sim: number of paths
141             :param n_steps: number of time steps
142             :return:

```

```

143         """
144
145         self.time_grid = np.linspace(0, self.T, n_steps)
146         X = np.zeros((n_sim, n_steps))
147
148         self.number_jumps_smaller_than_epsilon = np.zeros(n_sim)
149
150         dt = self.T/n_steps
151
152         for i in range(n_sim):
153             for j in range(1,n_steps):
154                 gamma_increments = np.random.gamma(self.alpha*dt,
155 1.0 / self.beta)
156                 if gamma_increments < self.epsilon:
157                     self.number_jumps_smaller_than_epsilon[i] +=1
158
159                 X[i, j] = X[i, j-1] + gamma_increments
160
161         self.sim = X.transpose()
162
163 class CompoundApproximator():
164     """
165     This class implements the compound Poisson process which
166     approximate the gamma process.
167     """
168
169     def __init__(self, alpha, beta, epsilon, T):
170         self.alpha = alpha
171         self.beta = beta
172         self.epsilon = epsilon
173         self.T = T
174         self.sim = None
175         self.time_grid = None
176
177     def simulate(self, n_sim, n_steps):
178         """
179         Simulate from the compound Poisson process
180         :param n_sim: number of trajectories
181         :param n_steps: number of time steps
182         :return:
183         """
184         intensity = self.alpha * scipy.special.exp1(self.beta *
185 self.epsilon)
186         Y = GammaJumpsDistribution(self.alpha, self.beta, self.
187 epsilon)
188
189         b_epsilon = self.alpha/self.beta*(1-np.exp(-self.beta*self.
190 epsilon))
191
192         self.time_grid = np.linspace(0,self.T, n_steps)

```

```

190         X = np.zeros((n_sim, n_steps))
191
192         for i in range(n_sim):
193             # Generate event times
194             num_events = np.random.poisson(intensity * self.T)
195             event_times = np.sort(np.random.uniform(0, self.T,
num_events))
196
197             # Generate random values
198             random_values = Y.sample(n_sim=num_events)
199
200             # Compute compound process values
201             X[i, :] = b_epsilon*self.time_grid
202
203             for index, jump_time in enumerate(event_times):
204                 idx_time = self.time_grid >= jump_time
205                 X[i,idx_time] = X[i,idx_time] + random_values[index
]
206
207         self.sim = X.transpose()
208
209
210 if __name__ == "__main__":
211     alpha = 5.0
212     beta = 2.0
213     epsilon = 0.1 # truncation parameter
214     T = 2.0 # time horizon of the simulation
215
216     # number of paths and number of time steps
217     n_sim = 5000
218     n_steps = 500
219
220     # Simulate using the approximation with a compound Poisson
221     X_eps = CompoundApproximator(alpha=alpha, beta=beta, epsilon=
epsilon, T=T)
222     print("Approximator sim")
223     X_eps.simulate(n_sim=n_sim, n_steps=n_steps)
224
225     # Exact simulation of the gamma process
226     print("Exact process sim")
227     X = GammaProcess(alpha=alpha, beta=beta, T=T)
228     X.set_epsilon(epsilon)
229     X.simulate(n_sim=n_sim, n_steps=n_steps)
230
231     # Plot the distribution at time t = T/2 of the compound Poisson
232     # process
233     data = X.sim[250, :]
234     data_eps = X_eps.sim[250, :]
235
236     t = T/n_steps*250
237     vectorized_func = np.vectorize(gamma_density)

```

```

238
239     # Plot histograms
240     plt.figure()
241     plt.hist(data_eps, bins=50, density=True, alpha=0.9, label='
compound Poisson process')
242     plt.hist(data, bins=50, density=True, alpha=0.9, label='gamma
process')
243     x = np.linspace(np.min(data), np.max(data), 1000)
244     y = vectorized_func(x, alpha*t, beta)
245     plt.plot(x, y, label="density", color='black', linewidth=2)
246
247     # Add labels, title, and legend
248     plt.xlabel('x')
249     plt.ylabel('Density')
250     plt.title('Distribution of the processes at time t=1')
251     plt.legend()
252
253     # Save the figure in EPS format
254     plt.savefig('distribution.eps', format='eps')
255
256     # Show the plot
257     plt.show()

```

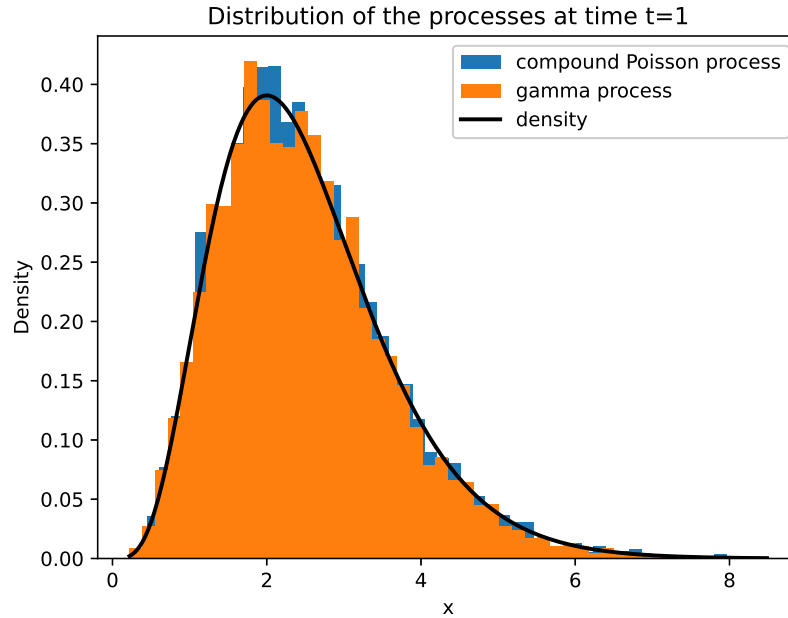


Figure 1: Distribution of the gamma process and of its approximating compound Poisson process at time  $t = 1$ . The processes have been simulated over the interval  $[0, T]$  with  $T = 2$ . The parameters are:  $\alpha = 5$ ,  $\beta = 2$  and truncation level  $\epsilon = 0.1$ .

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