

Parallel Tempering

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Bayesian Inference in Bioinformatics

- Suppose we can measure some quantity y. Assume, that parameter α describes y's distribution
- \rightarrow let both be random and their joint density $g(y, \alpha)$ factorise so that $g(y, \alpha) = h(y|\alpha)f(\alpha)$, where f is a priori distribution on the parameter
- \rightarrow f might result from an underlying physical theory
- Real sample points $\mathfrak{y} = [y_1, \dots, y_M]$ are observed
- The *a posteriori* distribution of α given the sample \mathfrak{y} , $f(\alpha|\mathfrak{y})$, describes how our knowledge about the studied quantity x is influenced by empirical evidence collected in \mathfrak{y}
- → Obtain it via the Bayes Formula

$$f(\alpha|\mathfrak{y}) = \frac{h(y_1|\alpha)\dots h(y_M|\alpha)f(\alpha)}{\int h(y_1|\beta)\dots h(y_M|\beta)f(\beta)d\beta}$$

Applications

- Hierarchical modelling for identification of co-expression patterns in microarray data by cluster analysis (Medvedovic et al., 2004; Stingo and M., 2010)
- Assessing the importance of explanatory variables (Stingo and M., 2010)
- Model Selection

MCMC

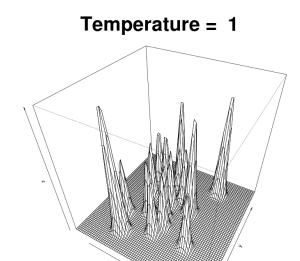
- MCMC algorithms are used to simulate samples out of analytically untractable posterior distributions
- → Most popular algorithm: Green-Metropolis-Hastings (Geyer, 2012)
- \rightarrow Generates a sequence of points that are thought of as being an instantiation of a Markov Chain, $X \equiv \{X^{[k]}\}_{k=0}^{\infty}$
- \rightarrow Each point $X^{[k]}$ is generated by accepting or rejecting at random a step proposal given the chains last position $X^{[k-1]}$
- → Approximates, thanks to Ergodic Theory, integrals

$$\mathcal{E}g(X) = \int_{\Omega} g(x)\pi(x) dx \approx \frac{1}{N} \sum_{i=1}^{N} g(X^{[i]}),$$

where π is the density of a posteriori distribution. In particular: approximates probabilities of any measurable set, $\mathcal{P}(A)$

- GMH estimates may suffer from poor mixing
- \rightarrow Chain X restricted to user-provided number of iterations N could get stuck in a probability cluster
- → Multimodial priors result in multimodial posteriors
- → Multimodial priors are selected when we suspect that the phenomenon under study is not concentrated around a particular point

Example

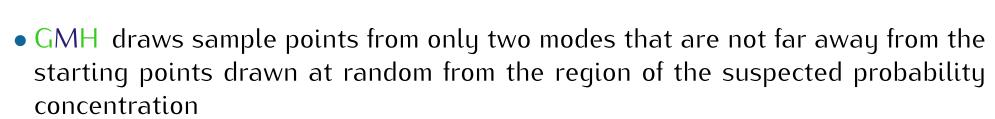


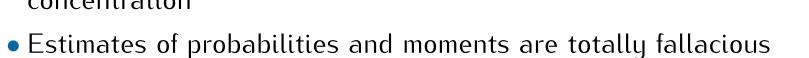
ullet Let π be a mixture of normal distributions

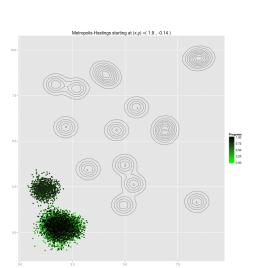
$$\pi(x) = \sum_{i=1}^{20} \frac{\omega_i}{\sigma_i \sqrt{2\pi}} \exp\left(-\frac{(x-\mu_i)^{t}(x-\mu_i)}{2\sigma_i^2}\right)$$

where σ_i are standard deviations, ω_i are weights, and μ_i are means (Baragatti et al., 2013)

• Some of the peaks mingle together to form bigger ones







¿Question?

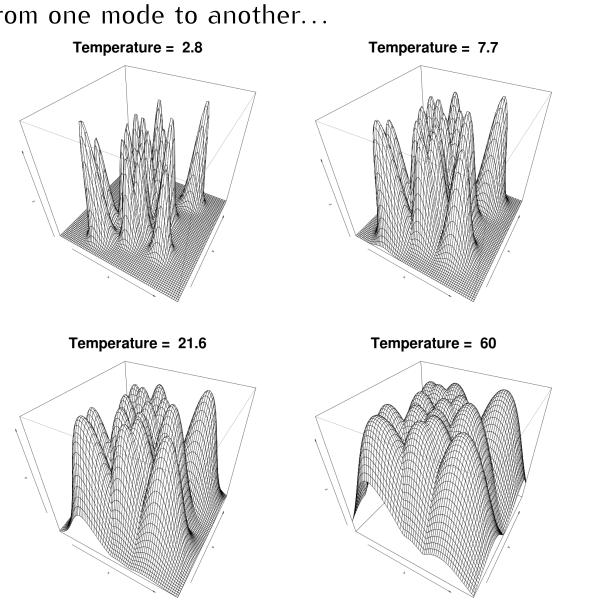
How can we enhance mixing so that the State Space is better searched for probability clusters?

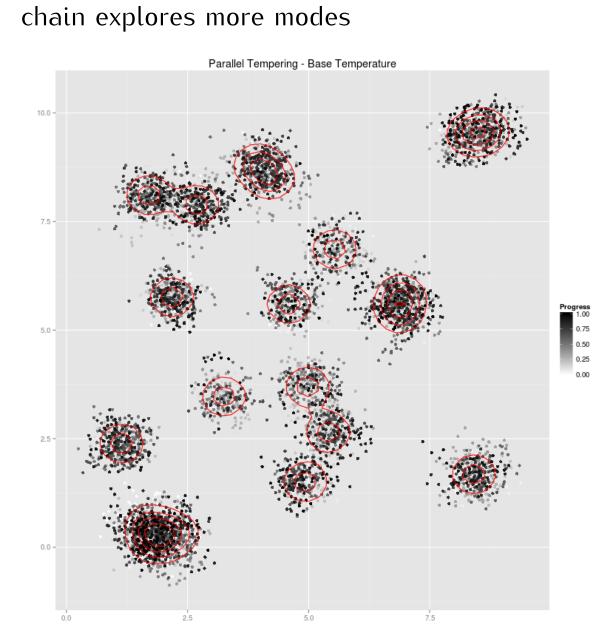
Parallel Tempering a.k.a. Replica Monte Carlo

- Foundations of PT laid by Swendsen and Wang (1986)
- Generate several chains $X = [X_1, \dots, X_L]$ and consists of two phases
- Ph I Drawing a point \tilde{X}_l from π^{β_l} , where $1 = \beta_1 > \cdots > \beta_l > 0$ are called inverse temperatures (note that first coordinate corresponds to our initial problem)
- Ph II Swaping some of \tilde{X} coordinates at random: the Swap Strategy
- Ph I mitigates the impact of multimodiality enlarging the probability of accepting steps from regions which the GMH would judge unlikely to draw from
- Ph II allows the passing of information from different chains: otherwise they would operate independently and first chain would gain nothing from other coordinates

Parallel Tempering at work

• ...and Ph II assures that the base temperature More tempered chains are at ease when passing chain explores more modes from one mode to another...





Different Swap Strategies

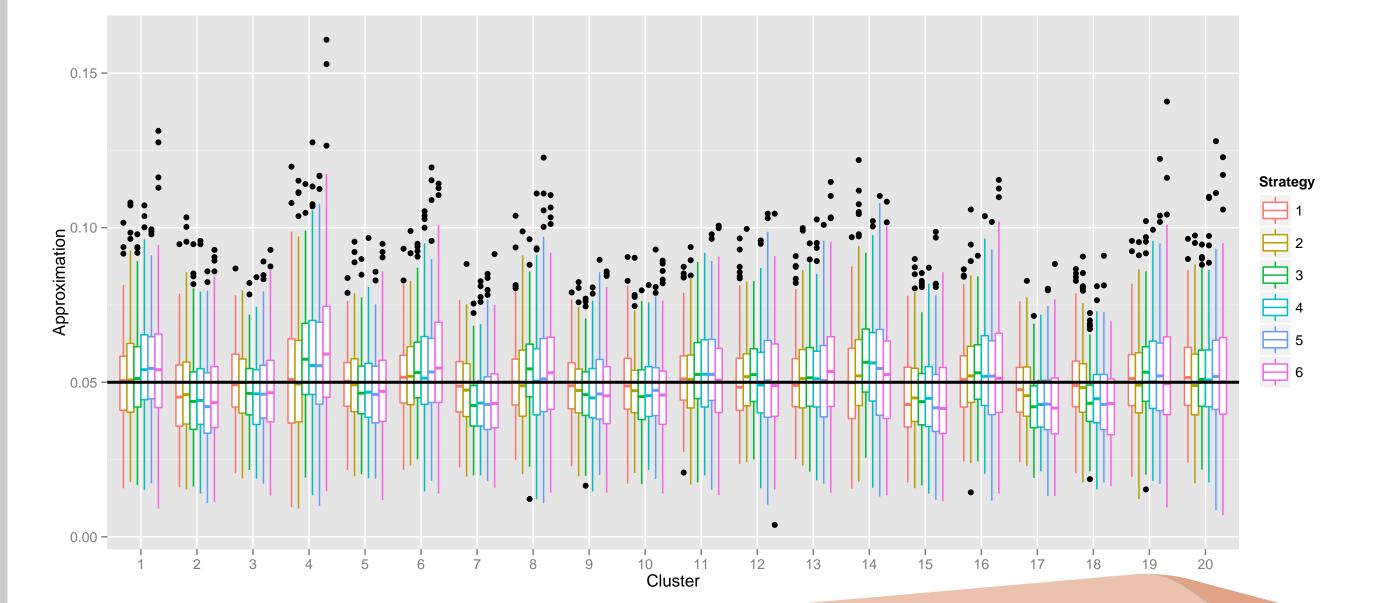
- In Ph II one can implement a multitude of Swap Strategies
- Suppose that $\tilde{X}^{[k]} = x$. Then the distribution on the indices can be described as p(i, j|x). We explored the following strategies (\propto denotes proportionality and \wedge - the minimum)

 $p(i,j|x) \propto \frac{\pi(x_j)}{\pi(x_i)} \wedge \frac{\pi(x_i)}{\pi(x_j)}$ Strategy 1 promotes swaps between coordinates or relatively the same level, i.e. $\pi(x_j) \approx \pi(x_i)$ Strategy 2 breaks the symmetry of the previous one, giving more

 $p(i, j|x) \propto \frac{\pi(x_j)}{\pi(x_i)} \wedge 1$ attention to swaps into regions of higher probability Strategy 3 softens the requirement that $\pi(x_i) \approx \pi(x_i)$ for $p(i,j|x) \propto \left(\frac{\pi(x_j)}{\pi(x_i)} \wedge \frac{\pi(x_i)}{\pi(x_i)}\right)^{\beta_i - \beta_j}$ similarly tempered coordinates, i.e. where $\beta_i - \beta_j \approx 0$:

swaps between adjacent chains get more probable Strategy 4 generalises the last one favouring more distant $p(i,j|x) \propto \left(\frac{\pi(x_j)}{\pi(x_i)} \wedge \frac{\pi(x_i)}{\pi(x_i)}\right)^{\frac{\beta_i - \beta_j}{1 + \rho(x_i, x_j)}}$ choices: ρ might any metric (e.g. euclidean)

- All the above strategies explicitly refer to values of π in points drawn in Ph I making the draws computationally cheap
- ullet Strategies 5 and 6 are independent of evaluations of π giving equal probability to all possible and all neighbouring swaps respectively
- Beneath we represent results obtained by all the strategies when trying to approximate the values of different modes: they should be equal roughly to 0.05, π being a mixture of 20 equally probable normal distributions. Result were obtained after 240 runs of PT with 2500 steps of burn-in and 7500 steps of simulations



- iGood Software Available Soon! • An R package, under the working name of StochasticSimulations, will be soon available for widespread use for free
- Among its features
- \rightarrow Division of the simulations into modules: Algorithm, State Space, Target Mea-SURE will provide a logic for the implementation of different models ightharpoonup Implementation of the most common choices for State Space: \mathbb{R}^S and Discrete
- \rightarrow Implementation of the GMH and PT algorithms with the above-mentioned
- STRATEGIES → GGPLOT 2 based visualisations
- ... and much, much more: stay tuned!

References

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