

Cattle Supply Chains and Deforestation in Brazil

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Abstract

Tropical forests are crucial to the global environment, serving as biodiversity hubs, storing and sequestering carbon, yet they face severe threats from agricultural expansion. Recognizing their role in deforestation, countries like the UK, US, and EU have implemented trade policies to curb deforestation by requiring proof that the production of a set of agricultural commodities is deforestation-free. A key question is whether these policies will encourage farmers to adopt sustainable practices or instead push them toward less regulated buyers.

In this paper, we adopt a structural approach to evaluate the effectiveness and potential leakages of sustainable supply-chain policies in the context of Brazil's cattle sector. First, we leverage a unique dataset on animal transport, matched with property boundaries, to document key empirical patterns of deforestation and the supply network of cattle at the farm level in the state of Pará, Brazil. Second, we build a structural model of farm-level land use decisions with an endogenous supply network. Third, guided by model-derived gravity equations, we estimate the role of existing domestic supply-chain zero-deforestation policies in shaping the trade of cattle. Our results indicate a strong deforestation trade penalty, whereby a farm that deforests is 16% less likely to supply to a zero-deforestation committed slaughterhouse, and each 10% increase in deforestation is associated with an 8.3% decrease in trade. In the next steps, these results will serve as foundation to evaluate counterfactual scenarios on the potential effectiveness of supply-chain policies such as the EU's upcoming Regulation on Deforestation-free products (EUDR).

1 Introduction

Tropical deforestation is advancing at a high rate, threatening our capacity to reduce greenhouse gas emissions and to preserve ecosystems. Agriculture is a major driver of this deforestation, representing up to 90% of deforestation ([Pendrill et al., 2022](#)). Acknowledging the limitations of public forest conservation policies at curbing deforestation, there has been a recent surge in strategies to reduce deforestation through supply chain policies. Supply chain policies place responsibility on large sourcing firms

to purchase deforestation-free commodities, potentially excluding suppliers who fail to meet environmental standards, such as avoiding deforestation. Some of these have taken the form of private voluntary standards (e.g. the G4 in Brazil or RSPO in Indonesia) or government-led certification schemes. Such supply chain policies will soon become mandatory for exporters to the European Union (EU) with the new deforestation regulation (EUDR) coming into effect at the end of 2025. The EUDR will prohibit the importation of commodities like coffee, cattle, cocoa, palm oil, and soy unless companies can prove their production is not associated with deforestation. In this paper, we analyze the potential effectiveness of supply-chain policies in the cattle sector in the state of Pará in Brazil.

Despite their increasing popularity, the efficacy of supply chain policies has remained an open question in the literature. Several studies point out to the many possibilities of evading such policies, be it by shifting production to non-targeted regions, by shifting sales towards buyers without restrictions, or by shifting deforestation upstream in the supply chain (Alix-Garcia and Gibbs, 2017; Levy et al., 2023; Gollnow et al., 2022). Cattle laundering, whereby illegally raised cattle enter deforestation-free supply chain, has been a pervasive concern due to the difficulty for companies to monitor producers beyond their first direct suppliers. This situation is likely to change rapidly with the EUDR, which mandates companies to monitor their entire supply chain. Whether this is going to effectively incentivize suppliers to reduce their deforestation or instead shift their production toward less restrictive importers remains a challenging yet important question to address. Most current land use models are not well-suited to understand these phenomena, as they tend to abstract away from the supply-chain organization of agriculture. Two recent papers are the exception, Domínguez-Iino (2024) and Barrozo (2024). They highlight the importance of slaughterhouse market power in mediating the impact of existing anti-deforestation policies. We contribute to this growing literature by looking at the multi-tiered nature of the cattle market. We consider farm-to-farm as well as farm-to-slaughterhouse transaction and model how farms endogenously decide whether to be direct or indirect suppliers.

We leverage a granular farm-level dataset of the cattle supply chain in the state of Pará in Brazil, to analyze the effectiveness of supply chain policies. This dataset contains information on farm-to-farm and farm-to-slaughterhouse cattle transport over five years and is matched with farm geolocation, enabling us to track the deforestation history of each farm in our sample over twenty years. We develop a structural model that jointly analyzes land use and supply network participation decisions. Specifically, the model seeks to analyze: (i) the direct effects of supply chain policies within the current supply network, but also (ii) how policies such as the EUDR would change the supply network, and (iii) the aggregate effects on deforestation, cattle production, and carbon emissions.

We contribute to three broad branches of literature. First, we contribute to a large literature analyzing the impacts and scope of zero-deforestation commitments in supply chains such as cattle in Brazil (Alix-Garcia and Gibbs, 2017; Levy et al., 2023; Skidmore et al., 2021; Gibbs et al., 2016; Villoria et al., 2022), soy in Brazil (Gollnow et al., 2022;

Heilmayr et al., 2020b), timber in Chile (Heilmayr and Lambin, 2016), and palm oil in Indonesia (Carlson et al., 2018; Heilmayr et al., 2020a; Lee et al., 2020). While this literature has provided valuable insights into the effectiveness of supply chain commitments, a key limitation is its reliance on reduced-form estimates that assume the supply chain is fixed and exogenous. Our main contribution is the development of a structural framework that jointly models land use and supply chain decisions, enabling counterfactual analyses of supply chain policies while accounting for strategic reallocation within the supply chain. Second, we contribute to a growing literature employing structural models to analyze deforestation (Hsiao, 2024; Domiguez-Iino, 2024; Farrokhi et al., 2024). These models have made substantial contributions to the understanding of deforestation as a dynamic process, and to the importance of the equilibrium consequences of policies via leakage and changes in comparative advantage in agriculture. Our main departure from these models is to consider farms, instead of pixels, as the unit of analysis. In our framework, decisions about how much to deforest, how to produce, and whom to sell to, are made at the level of the property, rather than at the level of the pixel. This enables us to better explain the observed heterogeneity in land use and sales. Third, we contribute to a growing literature on the role of production networks (Antras and Chor, 2022; Bernard and Moxnes, 2018; Caliendo and Parro, 2014; Eaton et al., 2022). The closest papers to our setting are Bernard et al. (2019), Arkolakis et al. (2023), and Panigrahi (2022) which consider spatial frictions in the context of firm-to-firm trade and allow for firm heterogeneity in productivity. The majority of models in this literature assume that trade flows can be represented by a dense trade matrix —an assumption that contradicts the cattle flows observed in our dataset, where a typical farm sells cattle to only a handful of other farms. To address this, we follow closely the setup in Panigrahi (2022) as it deals with the sparsity of the production network as discussed in Bernard and Zi (2022), and delivers predictions for the probability of trade with firm-level granularity.

2 Data and Stylized facts

Our granular dataset of farm-level deforestation and supply chain decisions comes from three distinct sources, 1) the animal transport dataset (GTA) for the state of Para, 2) the property boundary dataset (CAR) and 3) deforestation maps from MapBiomass. This results into a dataset of 20 years of land use and deforestation decisions (2001-2020), five years of animal transactions (2016-2020) and property size information. Analysis of this dataset reveals key facts on the nature of the cattle supply network.

Fact 1: Direct suppliers of ZDC slaughterhouses have significantly reduced their levels of deforestation. The farms that supply cattle to ZDC slaughterhouses during the period that we observe, between 2015 and 2019, have significantly reduced their rate of deforestation relative to other farms. In the figure below, we plot the trends in deforestation rates, defined as the percentage of forest left that is deforested in any given year, across four types of farms defined by their transaction in the period 2015-2019. First,

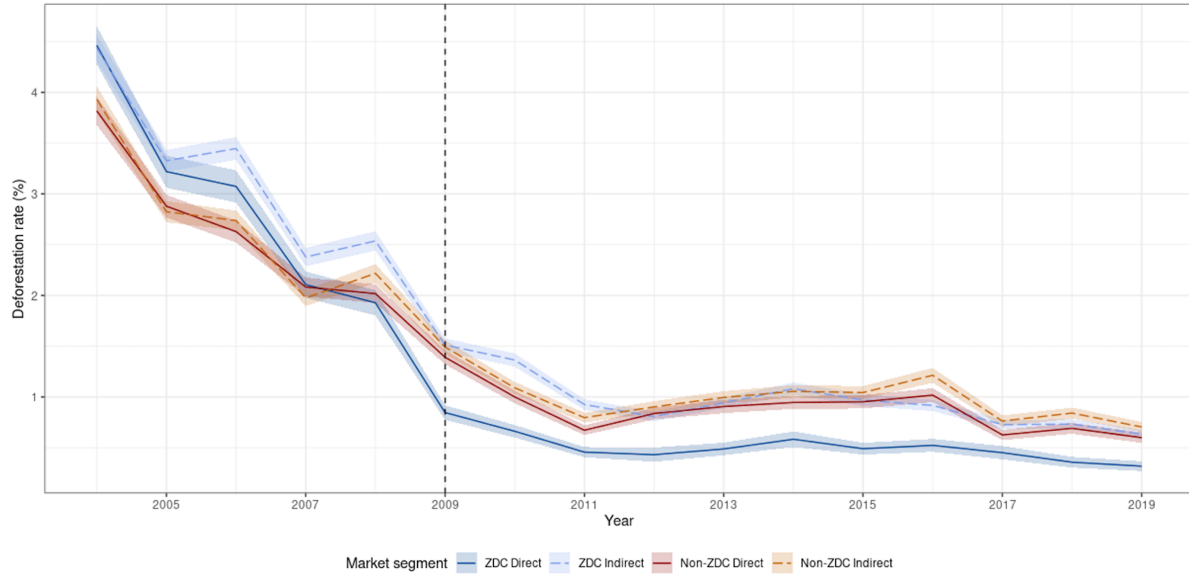


Figure 1: Deforestation Trends according to sales to ZDC slaughterhouse

“ZDC Direct” are those farms that sell at least cow to a ZDC slaughterhouse, and all others do not sell any to ZDC slaughterhouses. They may also sell to other non certified slaughterhouses or to farms. “ZDC Indirect” only sell cows to other farms, and they sell cows to farms that sell to a ZDC slaughterhouse. “Non-ZDC Direct” sell cows to slaughterhouses but not to any that hasn’t signed a zero-deforestation commitment. “Non-ZDC Indirect” are all the others. That is, they sell cows to farms that sell to other farms and to non certified slaughterhouses only. The pattern shown by the graph is striking. Despite ZDC Direct and ZDC Indirect having the highest levels of deforestation in 2004, and indistinguishable levels from other farms over the period 2004-2008, they have a more steeply decreasing rate of deforestation and from 2009 onwards they have persistently lower deforestation than the other farms, which have double their average deforestation rates.

From this fact, it is tempting to conclude that the policy has had the desired effect. However, since we do not have data on the supply chain before the policy was enacted in 2009, we cannot establish causality from this fact. Since the supply network is endogenous and highly variable, as we will document in stylized fact 4, it may be that farms first decide how much to deforest according to factors orthogonal to the policy and then ZDC slaughterhouses are more likely to buy only from those with lower rates of deforestation. With this endogeneity of a supply chain in mind, a difference-in-differences exercise based on the trends above will overestimate the effect of the policy. This fact is consistent with a policy that only affects sorting but not deforestation rates.

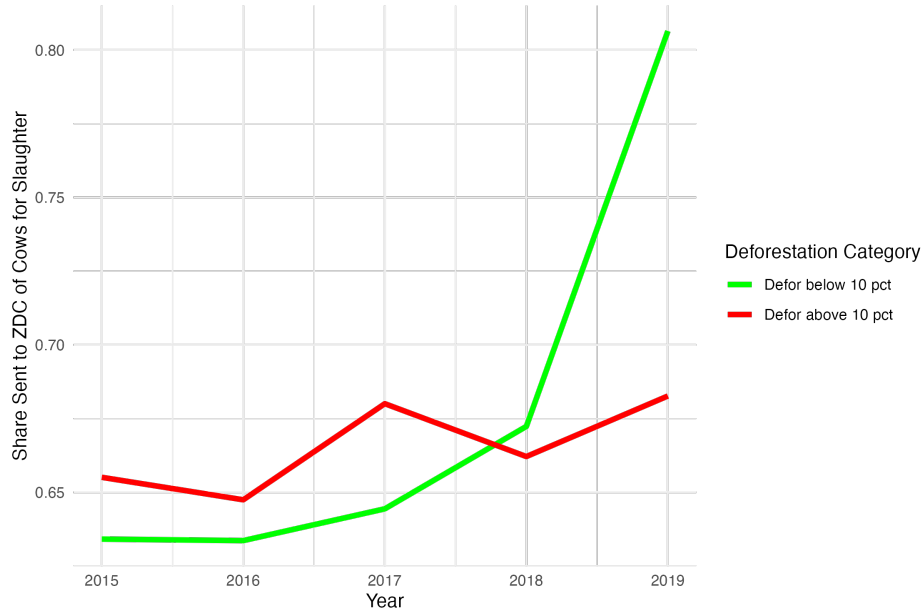


Figure 2: Share sent to ZDC slaughterhouse by low /high deforestation

Fact 2: Probability of being direct supplier of ZDC slaughterhouses is affected by deforestation. Second, we provide evidence for the endogeneity of the supply chain. The figure below split farms in two groups. “Low deforestation” farms are those that in the period 2009-2019 deforested less than 10% of their initial (in 2009) forest areas. Those that deforested more than 10% of their 2009 forest area are referred to as “high deforestation”. We then plot, out of the total number of cows sold for slaughter by each group, what share is sold to ZDC slaughterhouses. Over the period 2015-2019, we notice that the low deforestation farms greatly increase their share of sales to ZDC slaughterhouses until they greatly high deforestation farms in 2019. This is only descriptive evidence, as there may be many other underlying differences between low and high deforestation farms in terms of size, location, and initial area in forest. However, it is indisputable of the fact that supply chains change over time. Farms that were selling less than 65% of their cows to ZDC slaughterhouse end up selling 80%. And this change is different for farms with different deforestation histories. High deforestation farms barely increase their sales to ZDC slaughterhouses by 5 percentage points over the same period. In section 4.1 we do a more rigorous estimation of the way in which deforestation influences farms’ supply chain decisions.

If this is the case, we still expect that farmers making a rational decision of how much to deforest will take into account the effects on the supply chain, and hence reduce their level of deforestation as a result of the policy. This effect is, however, likely to be much smaller than that implied by a diff-in-diff as in Fact 1. Moreover, it is likely to be highly heterogeneous according to the potential buyers of each farms as given by the farm’s location.



Figure 3: Distribution of the share of cows sent to a slaughterhouse

Fact 3: Farm-to-farm trade is very important and few farms are fully direct. In 2017, 77% of all recorded transactions were farm-to-farm, and they accounted for 63% of cows sold. Even farms that sold some cows to slaughterhouses often choose to sell a large fraction of their cows to other farms too. This suggests that there does not seem to be a stark division between upstream and downstream farms or between. Figure 3 is a histogram of the share of a farms' cattle sent to a slaughterhouse, aggregated over the five years of data. It shows that a considerable fraction of the farms, 41.7%, sell all of their cows to other farms. It also shows that out of the farms that sell to slaughterhouses, very few sell exclusively to slaughterhouses, only 3.7%. Instead, most farms do a mix of both, and selling large fractions of one's cows for slaughter is rare.

The lack of a stark division between and indirect has two important implications for our analysis. First, that deforesting farms that sell for slaughter and want to avoid ZDC regulations may redirect their sales to farms and not just to non-ZDC slaughterhouses. Second, it suggests that there are purely indirect farms. We interpret this, in line with conversations with people in the field, as evidence of the fact that some farms specialize in breeding calves and selling to farms that will rear and fatten for slaughter. This will play a role when deciding whether to let a farm's production costs be the same for cows for slaughter and for cows for rearing. In our preferred specification, we allow for these two costs to be different (but the same for producing cows to a ZDC slaughterhouse or a non-ZDC slaughterhouse) and potentially infinite for the farms that do not sell any cows for slaughter.

Fact 4: Churn rates are high. There does not seem to be a high level of persistence among trade partners. Out of 146,888 active buyer-supplier links in 2015, 83% are not active in 2016, and 75% are not active ever again in our sample. This is similar when we look at other years. The numbers are not much lower if we restrict attention only to those links in which the supplier sells appears in all five years of our data –79% of links disappear in the following year and 68% disappear for the following four years. This indicates that not only do sellers break buyer-seller links and form new ones easily. In 2019, our last year of transactions data, 70% of the buyer-seller pairs were new, that is, they had not happened in earlier years. The number goes down to 64% when restricting attention to suppliers that appear in our data every year since 2015.

3 Model

We develop a spatial structural model of the cattle market with land use change and endogenous supply network formation. Our goal is to estimate how the regulations at the slaughterhouse level¹ affect the decision of farmers of how much to deforest, and who to sell their cows to. Such a structural approach allows us to see ZDC regulation as affecting the entire network of farms and slaughterhouses, and not just the current suppliers of regulated slaughterhouses. As a result, we can use the model to quantify the extent to which ZDC regulation led to deforestation reductions as well as changes in the supply chain. Our stylized facts suggest that both phenomena are likely to be at play. On one hand, we know that farms that currently sell to ZDC slaughterhouses have reduced their levels of deforestation since the onset of the policy in 2009. On the other hand, we see that farms change who they sell to after deforesting, and in particular they sell less to ZDC slaughterhouses. Since there is no pre-ZDC regulation supply-network data, a model is needed to provide evidence on the relative strength of supply-network reshuffling versus actual deforestation reductions. Equipped with this model, we will be able to simulate different counterfactual scenarios. For example, one in which there are no zero-deforestation commitments, or one in which more slaughterhouses become ZDC. We could also use the model to get a sense of what would make for an efficient spatial targeting of such a regulation.

The model features three types of agents: landowners, cattle farmers, and slaughterhouses. Each landowner owns a farm, chooses how much to deforest within it, and rents it out to cattle farmers.² When choosing how much to deforest, the landowner internal-

¹We hold fixed location and farm size as these move more slowly than land use change and cattle transaction decisions and this model is about the short-run (5 years). Given the high churn rates, we also abstract away from stickiness in farmer-farmer and farmer-slaughterhouse relationships since we see frequent changes in trading partners year-to-year.

²The device of distinguishing the landowner from the farmer follows common practice in the deforestation literature. It is not needed that farmers and landowners are different people. It is employed as a way of separating the two decisions and simplifying the model by considering deforestation and trade sequentially.

izes how deforestation will reduce the farmers' ability to sell to a ZDC slaughterhouse, thus reducing their willingness to pay for that land. This is especially true if the main demand in their region comes from this type of slaughterhouse. Farmers produce cows for sale by mixing the pasture land that they rent and cows that are bought from other farms. The total value of these bought cows is increased when they are fed pasture. In practice this is reflected by higher weight of the cows or by an increased number of cows, since they can be reproduced in the farm. The fattened or reproduced cows are then sold to other farms or slaughterhouses, which can be ZDC or non ZDC. Cows are thus an intermediate output. For simplicity, we let each slaughterhouse's expenditure in cattle be exogenous.

3.1 Deforestation

The set of all farms is \mathcal{S} with size S . Each landowners owns a farm $s \in \mathcal{S}$ with total amount of land $\bar{l}(s)$, and an amount of already deforested land, that is the area in pastures, $l^0(s)$. The amount they choose to deforest, $l^{\text{defor}}(s)$, bounded above by $\bar{l}(s) - l(s)$ dictates the amount of pasture land that they will have available to rent³,

$$l(s) = l^0(s) + l^{\text{defor}}(s).$$

Deforestation, even in the absence of slaughterhouse regulations, is a costly activity. These costs are partly due to the physical effort and resources required to fell down trees and partly due to the potential risk of being directly caught by environmental authorities and fined, as the vast majority of deforestation in the Amazon is illegal. We assume that the cost of deforestation, $\phi(l^{\text{defor}}(s))$, is increasing and convex in the amount of land deforested.

Optimal deforestation decision. The landowner with farm s of size $\bar{l}(s)$, of which $\underline{l}(s) \leq \bar{l}(s)$ is pasture in the initial period (2009), is choosing how much of the remaining forested land to clear, $l^{\text{defor}}(s) \leq \bar{l}(s) - \underline{l}(s)$ to have additional pasture area.⁴ The resulting pasture area, $l(s) = \underline{l}(s) + l^{\text{defor}}(s)$ gets rented out to cattle farmers at an endogenous rental rate $r(s)$. In equilibrium, the rental rate will depend on the total area in pasture of farm s and on how much of that pasture has been deforested since 2009. Considering this static framework in which the decision to deforest gets made once and it determines the rental rate $r(s, l(s), l^{\text{defor}}(s))$ for the foreseeable future, the landowner solves the following maximization problem

$$\max_{l^{\text{defor}}(s)} \left\{ \frac{1}{1 - \rho} r(s, l(s), l^{\text{defor}}(s)) \times l(s) - \phi(s, l^{\text{defor}}(s)) \right\} \quad (1)$$

In this framework, slaughterhouses' ZDC regulations influence the landowners' deforestation decision by making the rental rate of the land more sensitive to deforestation.

³We abstract away from forest regeneration, pasture degradation, and the planting of other crops for the time being.

⁴The choice to reforest or abandon pasture is currently rules out of the model.

As the area of deforested land in farm s goes up, the demand for cattle raised in s goes down, and thus cattle farmers have less demand for that landowner's land, lowering the equilibrium rental rate $r(s)$. In mathematical terms, more stringent or more widespread ZDC regulations further decrease $\partial r / \partial l^{\text{defor}(s)} < 0$. This is substantially different from the effect on deforestation of command-and-control policies that affects landowners' direct costs of deforestation, as highlighted in [Alix-Garcia and Gibbs \(2017\)](#). Those types of regulations can be seen as decreasing the marginal costs of deforestation, $\partial \phi / \partial l^{\text{defor}} > 0$.

A key feature of ZDC regulations as an anti-deforestation policy is that the precise way in which it affects a farm's deforestation decisions depends on how much it depends on the demand from ZDC slaughterhouses. This is difficult to understand without a model, as it depends on many factors that interact. First, the level of competition from other farms that do not deforest. Second, paradoxically, having a lot of non-deforesting farms nearby can also enable a farm's deforestation because it means there are farms with high demand for cows as input. Third, on how much demand from non-ZDC slaughterhouses there is. In general, all of this factors are heavily mediated by geography since cattle trade is particularly sensitive to distance, as shown in stylized fact 3.

3.2 Cattle farming technology

Given a farm's amount of pasture land and its intrinsic productivity, the farm chooses how to participate in the cattle supply network, that is, how many cows to buy and sell, who to buy them from, and who to sell them to. Slaughterhouses buy mature cattle and produce beef they sell to domestic and international consumers. Farms and slaughterhouses differ in their location and productivity.

Cows are produced with a Cobb-Douglas production function that combines a bundle of cows from other farms $m(s)$ and land $l(s)$ with a constant returns to scale (CRS) technology and Hicks-neutral productivity $z(s)$,

$$y(s) = z(s) \left(\frac{l(s)}{\beta} \right)^{\beta} \left(\frac{m(s)}{1 - \beta} \right)^{1 - \beta}$$

for a share of land in production $0 < \beta < 1$ and a current amount of land in pastures $l(s)$. The bundle of cows bought as intermediate inputs is an aggregate of a finite number of purchases K from other farms⁵

$$m(s) = \prod_{k=1}^K m(s, k)^{1/K}.$$

Farmers choose the amount of land they rent and the amount of cows they buy and whom to buy cows from in order to minimize their unit costs of production. The rental rate of

⁵The assumption of complementarities between purchases is made to ensure tractability. In practice, it can be thought of as coming from the fact that cows reproduce and in that sense augment each others' productivity. Moreover, genetic diversity is desirable.

land and the price at which suppliers sell them cows will dictate the costs of production, which in turn dictates the price at which they sell to their buyers, leading to a recursive formula for production costs.

3.3 Slaughterhouse demand

Slaughterhouses buy cows for slaughter and produce beef that is sold to consumers. We assume that a slaughterhouse b of type $q \in \{\text{non ZDC s.h.}, \text{ZDC s.h.}\}$ in location d has a fixed expenditure in cows $X_d^q(b)$. This can be rationalized by (s) consumers with an exogenously fixed income and Cobb-Douglas preferences, (ii) no substitution between beef from different slaughterhouses, and (iii) Cobb-Douglas production of beef. The second assumption is harder to justify as an accurate depiction of the real world, but it is helpful in order to quantify the channels of main interest to us. Our primary goal is to understand how the impact of a slaughterhouse having stricter environmental regulation can be mitigated by laundering and avoidance, even if that slaughterhouse continues to spend the same amount on cattle. It is straightforward to extend the model to allow for substitution between slaughterhouses but for now we avoid doing so in order to isolate the role of the upstream margin of the supply chain of beef.

3.4 Production network formation in space

The choice of who to supply from is done by choosing the least cost supplier for each of the finite number, K , of transactions that a buyer makes. First, a few words on notation. We index buyers (farms or slaughterhouses) by b and sellers (farms) by s . The *type* of the buyer, which can be either *farm*, *non ZDC slaughterhouse*, or *ZDC slaughterhouse*, is indexed by q . The location of the seller, also referred to as the origin, is indexed by o , and the location of the buyer, or destination, is indexed by d . Although we have the precise location of all farms and slaughterhouses, we coarsen geography to a grid of 10 km wide squares⁶. This helps computationally by greatly reducing the dimensionality of the trade cost matrix to be estimated.

Sourcing decisions. For each purchase k that a buyer b of type q at destination d makes, it sources from the supplier farm s from origin o that offers the lowest price out of those that are visible to b . Out of all the farms in the economy, buyer b sees only a subset, $\mathcal{S}(b) \subseteq \mathcal{S}$ such that each supplier belongs to that subset with probability $\frac{\lambda}{5}$ via independent Bernoulli trials. Thus the chosen supplier for purchase k by buyer b equals,

$$s^*(b, k) = \arg \min_{s \in \mathcal{S}(b)} p_{od}(s, b, k).$$

where the price $p_{od}(s, b, k)$ offered by suppliers in turn depends on geography, supply chain anti-deforestation policy, production costs, and other idiosyncratic shocks. For-

⁶As a robustness exercise we try bigger grid-cells.

mally,

$$p_{od}(s, b, k) = \frac{\bar{m}_{od}(s, b, k) c_o(s) \kappa^q(s) \tau_{od}^q}{a_{od}(s, b, k)}$$

where $\mathcal{S}(b)$ is the set of suppliers visible to b , each visible with iid probability λ , $c_o(s)$ is the marginal cost of production of supplier farm s in origin o , $\kappa^q(s)$ is the deforestation penalty, τ_{od}^q is the iceberg trade cost between o and d , $a_{od}(s, b, k)$ is a random transaction-pair productivity shock, and $\bar{m}_{od}(s, b, k)$ is a the mark-up charged. Since only a finite number of potential suppliers are visible, the lowest cost supplier can sell at the price charged by the second cheapest, leading to markups. We assume that $a_{od}(a, b, k)$ are i.i.d. and follow a Pareto-distribution with shape parameter ζ and scale parameter a_0 ,

$$a_{od}(s, b, k) \text{ i.i.d. } \sim F_a(a) = 1 - (a/a_0)^{-\zeta}.$$

The Deforestation Penalty. Our main departure from production network models in the literature is to add a seller-buyer type-specific trade cost, the deforestation penalty $\kappa^q(l^{\text{defor}}(s))$. The deforestation penalty operates as an iceberg trade cost between sellers and buyers.

We assume that $\kappa^q(0) = 1$ and will show that it is increasing in l^{defor} and that

$$\kappa^{\text{farm}}(x) < \kappa^{\text{non ZDC s.h.}}(x) < \kappa^{\text{ZDC s.h.}}(x).$$

We estimate how deforesting will persistently reduce a farm's ability to trade with different types of buyers. In particular, we can see how deforesting more, but keeping all else equal, lowers the likelihood of selling to ZDC slaughterhouses. After closing the model, we will estimate how the deforestation penalty will reduce landowners' incentives to deforest in equilibrium.

Recursive formulation of costs. This functional form for suppliers' prices leads to a recursive formula for farms' production costs where the production cost of each farm b , $c_d(b)$, depends on the rental rate of land r_d and the prices charges by its suppliers for inputs, which in turn depend on their costs,

$$c_d(b) = \frac{r_d^\beta}{z_d(b)} \times \prod_{k=1}^K \min_{s \in \mathcal{S}_d(b)} \left\{ \frac{\bar{m}_{od}(s, b, k) \tau_{od}^{\text{farm}}}{a_{od}(s, b, k)} \times c_o(s) \kappa_o^{\text{farm}}(s) \right\}^{\frac{1-\beta}{K}}.$$

From [Panigrahi \(2022\)](#)'s Proposition 1, it follows that, before the idiosyncratic match-specific productivities $a_{od}(s, b, k)$ are realized, b supplies any task k from s with probability

$$\mathbb{E}[\pi_{od}^q(s, b)] = \frac{(c_o(s) \kappa_o^q(s) \tau_{od}^q)^{-\zeta}}{\sum_{s'} (c_{o'}(s') \kappa_{o'}^q(s') \tau_{o'd}^q)^{-\zeta}}. \quad (2)$$

This equation summarizes how trade probabilities between a buyer and a supplier depend on distance, on deforestation histories, and on farm-level productivities. The multiplicative nature of the relationship allows for its decomposition into components that vary at different levels of aggregation of the transaction data. As will see in the next section, this is the main equation that we will rely on to estimate the (i) deforestation penalty $\kappa_o^q(s)$, (ii) trade frictions τ_{od}^q and how they depend on distance, and, eventually, (iii) back out the the productivity of each farm.

3.5 Closing the model

In equilibrium, the land rent payments of farmers s towards their landlord amount to a fraction β , the cost share of land, of their expected total costs of cattle farming, which are equal to the expected revenues divided by the markup which is, in expectation, equal to $\frac{\zeta+1}{\zeta} > 1$. The expected revenues received by a farmer, or seller, s , in location o , amount to the sum of all payments that s received from other farms and slaughterhouses. The payments from a buyer of type q in a location d can be calculated as their expenditures in cows, $X_d^q(b)$, times their expected share sourced from s , $\mathbb{E}(\pi_{od}^q(s, -))$. For slaughterhouses, this is taken to be exogenous data. For farms, this is equal to $\frac{1-\beta}{\beta}$ times the rental rate payments that those buying farms make, $r_d(b)l_d(b)$. Therefore,

$$\begin{aligned}
r_o(s)l_o(s) &= \frac{\zeta}{1+\zeta}\beta\left(\mathbb{E}Y^{\text{farm}}(s) + \mathbb{E}Y^{\text{non ZDC s.h.}}(s) + \mathbb{E}Y^{\text{ZDC s.h.}}(s)\right) \\
&= \frac{\beta\zeta}{\zeta+1}\sum_d \mathbb{E}(\pi_{od}^{\text{farm}}(s, -)) \underbrace{\sum_{b \in \mathcal{B}_d^{\text{farm}}} \frac{1-\beta}{\beta} r_d(b)l_d(b)}_{X_d^{\text{farm}}} \\
&\quad + \frac{\beta\zeta}{\zeta+1}\sum_d \mathbb{E}(\pi_{od}^{\text{non ZDC SH}}(s, -)) X_d^{\text{non ZDC s.h.}} \\
&\quad + \frac{\beta\zeta}{\zeta+1}\sum_d \mathbb{E}(\pi_{od}^{\text{ZDC SH}}(s, -)) X_d^{\text{ZDC s.h.}}
\end{aligned} \tag{3}$$

The market clearing equation above serves three main purposes. First, it can be used to invert the model and solve for the rental rates of each farm in the baseline equilibrium, in which ZDC policies are in place. Notice that it is a system of S equations with S unknowns, where S is the number of farms and $\{r_o(s)\}_{s \in \mathcal{S}}$ are the unknowns. This requires data on: (i) the land in pastures of each farm, $l_o(s)$, which we observe, (ii) the expected trade shares of each type of buyer in each destination for each seller, $\mathbb{E}(\pi_{od}^q(s, -))$, which we can estimate using the structure of the model and data on observed trade shares $\pi_{od}^q(s, b)$, and (iii) the aggregate expenditures of slaughterhouses in each destination on cows, $X_d^{\text{non ZDC s.h.}}$ and $X_d^{\text{ZDC s.h.}}$. This is done in the next section, where we take the model to data.

Second, equation (3) can be used to model how we expect the land rents from farm s to vary with deforestation, and how that depends on ZDC regulation in partial equilibrium. We can rewrite the landowners' optimal deforestation decision laid out in (1) using (3) to understand the dependence of the land rents from a farm, $r_o(s)l_o(s)$, on its level of deforestation, $l_o^{\text{defor}}(s)$. There are two main competing channels: (i) deforestation increases revenues by increasing the amount of pasture land available to rent out to cattle farmers, $l_o(s) = \underline{l}_o(s) + l_o^{\text{defor}}(s)$, albeit with decreasing marginal returns since $r_o(s)$ is decreasing in $l_o(s)$, and (ii) deforestation reduces the rental rates of land $r_o(s)$ by lowering the ability of farmers to sell cows to ZDC slaughterhouses. The second channel is less prevalent for farms that expect to sell relatively less to ZDC slaughterhouses, either because they specialize in calving, and hence to no direct sales to slaughterhouses, because they are far from ZDC slaughterhouses relative to other types of buyers, or because they have already deforested a large area since 2009.

Third, having the rental rates in the baseline equilibrium and a model for how landowners make optimal deforestation decisions while taking ZDC regulation into account, we can simulate how deforestation decisions would have been different in an alternative market equilibrium without such policies.

3.6 Deforestation in equilibrium

In equilibrium, the landowner is deciding how much to deforest, $l_o^{\text{defor}}(s)$, in order to maximize their profits as outlined in maximisation problem (1), while accounting for the following three channels. First, that more land in pastures means receiving more rents, $\frac{\partial r_o(s)l_o(s)}{\partial l_o(s)} = r_o(s)$. Second, that more land in pastures lowers rental rates per hectare, $\frac{\partial r_o(s)}{\partial l_o(s)}l_o(s) = -\frac{1}{1+\beta\zeta}r_o(s)$. And third, that deforestation lowers rental rates per hectare due to ZDC regulation. If only ZDC slaughterhouses penalize deforesting farms, i.e. $\kappa^{\text{farm}}(s) = \kappa^{\text{non ZDC sh}}(s) = 1$, then this channel can be quantified as follows,

$$\frac{\partial r_o(s)}{\partial l_o^{\text{defor}}(s)}l_o(s) = \frac{\beta\zeta}{1+\beta\zeta} \frac{1}{\kappa_o^{\text{ZDC}}(s)} \frac{\partial \kappa_o^{\text{ZDC}}(s)}{\partial l_o^{\text{defor}}(s)} \mathbb{E}Y_o^{\text{ZDC}}(s),$$

where $\mathbb{E}Y_o^{\text{ZDC}}(s)$ are the sales of farm s to ZDC slaughterhouses in expectation.

Assuming that $\kappa_o^{\text{ZDC}}(s) = \exp(\alpha \cdot \text{lhs}(l_o^{\text{defor}}(s)))$, then, as shown in appendix section C,

$$\begin{aligned} \frac{d\text{Landowner Rent}_o(s)}{dl_o^{\text{defor}}(s)} &= \frac{\partial r_o(s)l_o(s)}{\partial l_o(s)} + \frac{\partial r_o(s)}{\partial l_o(s)}l_o(s) + \frac{\partial r_o(s)}{\partial l_o^{\text{defor}}(s)}l_o(s) \\ &\approx \frac{\beta\zeta}{1+\beta\zeta} \left(r_o(s) - \frac{\alpha\zeta}{1+\zeta} \frac{1}{\sqrt{l_o^{\text{defor}}(s)^2 + 1}} \mathbb{E}Y_o^{\text{ZDC}}(s) \right). \end{aligned} \quad (4)$$

4 Taking the model to data

We rely on five years of trade flows between farms and slaughterhouses, their geographical locations, and the land use history of each farm, to take the model to data as follows. In an initial step, we estimate key parameters of the supply-network formation model. The main two are: (i) the distance-elasticity of trade costs τ_{od} , estimated through a gravity equation, and (ii) the deforestation penalty $\kappa_o^q(s)$. Second, we invert the model to back out farm-level rental rates at baseline. In a third step, these rental rates are used to back out the implied marginal costs of deforestation. This is a function of deforestation that rationalizes observed deforestation rates given the estimated profitability of cattle production coming from the supply-network model.

Using structural equation (2), we can use data on observed trade flows to parse out the role of characteristics of the seller, the origin location, the buyer, the destination location, and the origin-destination pair. We re-write (2) to show this more explicitly,

$$\mathbb{E}[\pi_{od}^q(s, b)] = \underbrace{(c_o(s)\kappa_o^q(s))^{-\zeta}}_{\text{Supplier-Buyer Type}} \times \underbrace{(\tau_{od}^q)^{-\zeta}}_{\text{Origin-Destination-Buyer Type}} \times \underbrace{\frac{1}{\sum_{s'} (c_{o'}(s')\kappa_o^q(s')\tau_{o'd}^q)^{-\zeta}}}_{\text{Destination-Buyer Type}}. \quad (5)$$

The structural determinants of expected trade probabilities are of three types. First, the characteristics of the supplier that vary by buyer type, thus including the costs of production of cows and the deforestation penalty. This carries the implicit assumption that all buyers of the same type are on average equally strict on their rules towards buyers that deforest. Second, the role of geography, varying at the origin and destination location pair, which we will assume follow a specific increasing function of distance. And third, the role of demand from a certain type of buyer at each destination, which varies at the destination and buyer type level.

4.1 Estimating the Deforestation Penalty

The deforestation penalty $\kappa_o^q(s)$ governs the degree to which a farm's deforestation history lowers its attractiveness to each of the three types q of buyers: farms, non-ZDC slaughterhouses, and ZDC slaughterhouses differentially. As motivation, we showed that the share of cows that a farm sells to each type of seller varies as a result of the amount of deforestation it does. This is true even when we control for farm fixed effects using the panel of transactions for years 2015-2019. In this section, we take a structural approach to estimate a parametric form for the deforestation penalty as a function of the area deforested since 2009. We assume the following functional form for the deforestation penalty,

$$\kappa_o^q(s) = \kappa_o(l^{\text{defor}}(s)) = \exp\left(\alpha^q \cdot \text{lhs}(l^{\text{defor}}(s))\right),$$

where α^q captures how much farms are penalized for each additional increase in deforestation in selling to buyer of type q . It approximates an elasticity, so that a 1% increase

in deforestation should reflect an $\alpha^q\%$ increase in (iceberg) costs of trading with buyers of type q . We also make the normalization assumption that $\kappa_o^{\text{farm}}(s) = 1$, that is $\alpha^{\text{farm}} = 0$.

For this estimation we focus on the variation in transaction volumes that comes from the supplier-buyer type combination. First, we decompose the farm unit production costs into location and farm components

$$c_o(s) = \underbrace{\tilde{c}_o(s)}_{\text{Supplier level}} \times \underbrace{c_o}_{\text{Origin level}},$$

where the term $\tilde{c}_o(s)$ is normalized such that $\sum_{s \in o} (\tilde{c}_o(s))^{-\zeta} = 1$ ⁷. Second, we do a similar decomposition of the deforestation penalty,

$$\kappa^q(s) = \underbrace{\tilde{\kappa}^q(s)}_{\text{Supplier-Buyer type level}} \times \underbrace{\kappa_o^q}_{\text{Origin-Buyer type level}},$$

where the term $\tilde{\kappa}^q(s)$ is normalized such that $\sum_{s \in o} (\tilde{c}_o(s) \tilde{\kappa}_o^q(s))^{-\zeta} = 1$ ⁸. Then, relying on the same logic as Proposition 2 from [Panigrahi \(2022\)](#), we can show that the closed-form formula for the seller-buyer type-year fixed effect estimator is

$$(\tilde{c}_o(s)^{-\zeta} \tilde{\kappa}_o^q(s)^{-\zeta})^* = \frac{\sum_b \pi_{od}^q(s, b)}{\sum_{s' \in o} \sum_{b'} \pi_{od}^q(s', b')} \equiv \frac{\text{intensity of use}^q(s)}{\sum_{s' \in o} \text{intensity of use}^q(s')}. \quad (6)$$

In order to estimate α^q , we rely on the estimator of the farm-specific fixed effects $(\tilde{c}_o(s)^{-\zeta} \tilde{\kappa}_o^q(s)^{-\zeta})^*$ given above. This estimator captures the characteristics of a farm that make it a more attractive supplier to buyers of type q relative to other farms in the same origin o , since it is normalized. This includes both the farm's (relative) productivity as well as the deforestation penalty. In order to estimate α^q we would like to know how these fixed effects vary with the level of deforestation of the farm, all else equal. Our preferred identification strategy uses the 2015-2019 panel data on transactions to document how changes in cumulative deforestation from 2009 leads to changes in the farm-specific fixed effects for each type of buyer. Specifically, we run the following PPML regressions of the fixed effect estimator on the inverse hyperbolic sine of deforestation with different sets of controls and fixed effects,

$$(\tilde{c}_{ot}(s)^{-\zeta} \tilde{\kappa}_{ot}^q(s)^{-\zeta})^* = \exp \left((-\zeta) \underbrace{(\alpha^q \cdot \text{ih}(\text{defor}_t(s)))}_{\log(\kappa_{ot}^q(s))} + \beta^q Z_t(s) + \eta_{ot}^q + \epsilon_{ot}^q(s) \right). \quad (7)$$

The results from different specifications of that equation are shown in table 1.

⁷Define $c_o \equiv \left(\sum_{s \in o} (c_o(s))^{-\zeta} \right)^{-1/\zeta}$ and let $\tilde{c}_o(s) \equiv c_o(s)/c_o$.

⁸Define $\kappa_o^q \equiv \left(\sum_{s \in o} (\tilde{c}_o(s) \kappa_o^q(s))^{-\zeta} \right)^{-1/\zeta}$ and let $\tilde{\kappa}_o^q(s) \equiv \kappa_o^q(s)/\kappa_o^q$.

	Dep. Var: $\tilde{c}^q(s) - \zeta \kappa^q(s) - \zeta$				
	(1)	(2)	(3)	(4)	(5)
ihs(Defor. since 2009) x To Farm	0.32*** (0.04)	-0.10** (0.04)	-0.13*** (0.04)	0.13 (0.13)	0.08 (0.14)
ihs(Defor. since 2009) x To non ZDC sh	0.21** (0.07)	-0.54*** (0.08)	-0.56*** (0.08)	0.08 (0.41)	0.24 (0.42)
ihs(Defor. since 2009) x To ZDC sh	0.33*** (0.10)	-0.86*** (0.13)	-0.86*** (0.13)	-1.48** (0.48)	-1.50** (0.49)
ihs(Defor. 2003-2009) x To Farm		0.24*** (0.01)	0.21*** (0.03)		
ihs(Defor. 2003-2009) x To non ZDC sh		0.43*** (0.03)	0.42*** (0.05)		
ihs(Defor. 2003-2009) x To ZDC sh		0.75*** (0.02)	0.50*** (0.09)		
log(Prop. Area) x To Farm		0.14*** (0.03)	0.07** (0.03)		
log(Prop. Area) x To non ZDC sh		0.16** (0.06)	0.01 (0.06)		
log(Prop. Area) x To ZDC sh		0.01 (0.06)	-0.09 (0.06)		
Origin cell x Buyer Type x Year FE	X	X	X	X	X
Land Use Controls x Buyer Type FE			X		X
Seller x Buyer Type FE				X	X
Num. Obs	513610	513610	423651	291631	245158
Num. Origin cell x Buyer Type x Year	12524	12524	12209	12524	12209
Num. Seller x Buyer Type				62035	52801

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 1: Fixed Effects Poisson Regression, D=20 km

All columns include origin-year-buyer type fixed effects in order to control for the characteristics of the farm that vary by location and year such as climate and weather, access to markets, and economic shocks that may be differently relevant for different types of buyer. For instance, there may be years in which ZDC slaughterhouses have more demand in certain locations. It is also well documented that certain years have much more demand for cow slaughter than others, in which most transactions happen between farms, this is often referred to as the cattle cycle. Columns (1)-(3) rely on cross-sectional comparisons of farms within the same 10km wide grid-cell, whereas columns (4) and (5) have farm-level fixed effects. Column (1) has no other controls, and shows that farms that deforest more area sell more cows, but less so for ZDC sales. Column (2) controls for the size of the farm, which significantly changes the results. Since larger farms sell more cows to all farms, and in particular to ZDC slaughterhouses, and also deforest more, when we control for farm size it becomes clear that additional deforestation is associated with significant reduction in sales to slaughterhouses, and especially to ZDC slaughterhouses, with an elasticity of 1. The results don't change dramatically in column (3), where we control for the area in each land is (pasture, forest, crops) interacted with the buyer type. The main difference is that deforestation is more associated with reduction in sales when we control for the fact that deforestation also increases pasture area. Columns (4) and (5) are our preferred specifications since they exploit within-farm variation. That is, we look at how farms that deforested in the period 2015-2019 changed in how attractive they were to different types of buyers. Column (5) controls for changes in land use, which is important to avoid confounding increases in area deforested since 2009 with increases in area in pasture⁹. Column (5) is the most credible specification and the coefficient that would be used to estimate $\alpha^{\text{ZDC s.h.}}$ is large and statistically significant at the 10% level. It implies that a 10% increase in deforested area is associated with a 14% increase in the implied iceberg trade costs when selling to ZDC slaughterhouses. The frictions when selling to farms and to ZDC slaughterhouses are small or insignificant. However, columns (3)-(5) tell a consistent story.

Identifying assumptions and validity. The main assumption underlying the estimation of α^j in column (5) is that the timing of deforestation is uncorrelated with other time-varying factors that may be driving a farm's relative attractiveness to different types of buyers. We are controlling for characteristics that

⁹These two are not collinear because of the choice of functional forms. Since we are taking logarithms, a 1% increase in deforestation is not necessarily a 1% increase in pasture area, it would depend on the baseline pasture area.

4.2 Structural Gravity Equation

In order to estimate trade costs, we rely on the closed-form expression for the estimator of the (o, d, t) specific term, which equals¹⁰

$$\left(\frac{(c_o \kappa_o^q \tau_{od}^q)^{-\zeta}}{\sum_{o'} (c_{o'} \kappa_{o'}^q \tau_{o'd}^q)^{-\zeta}} \right)^* = \frac{1}{B_d^q} \sum_{b \in d} \sum_{s \in o} \pi_{od}^q(s, b) = \text{Average Sourcing Intensity}_{od}^q.$$

This allows running a structurally-derived gravity equation with the average sourcing intensity as the dependent variable on origin-year fixed effects ν_{ot} and destination-year fixed effects η_{dt} , so that the remaining variation is due to origin-destination-year variation, namely trade costs. Trade costs will depend on distance but their exact value will also depend on the quality of roads, traffic, and the availability of vehicles to transport the cattle. We run a simple gravity equation on log of distance¹¹. We run this equation pooling all types of buyers q and also separately for different buyer types, as the trade costs for different types of buyers may be different,

$$\text{Average Sourcing Intensity}_{odt}^q = \exp \left(\nu_{ot}^q + \eta_{dt}^q + \beta^q \cdot \log(\text{Distance}_{od}) \right) + u_{odt}$$

The gravity equation results, shown in table 2, show that cattle trade steeply decreases with distance. This elasticity is high relative to what is usually found in the trade literature for all traded goods. This makes sense since cattle needs to be fed and hydrated during transport, making each additional increase in travel distance more salient.

4.3 Baseline Rental Rate Estimation

The next step in this quantification exercise is to back out the expected rental rates of all farms at baseline. The estimation procedure is described in detail in appendix section D. It consists of four steps. First, we calculate the expected sourcing intensities, $\pi_{od}^{q*}(\bullet, -) = \sum_{s \in S_o} \pi_{od}^{q*}(s, b)$, or expected aggregate trade shares. These represent the expected probability that a buyer in a destination d chooses to source from a buyer in an origin o . In order to approximate these, we use the fitted values from cross-sectional gravity equations that pool together transactions from all years of data. Second, we aggregate the market clearing equation at the level of the grid-cell location so that the only unknowns are the aggregate rental payments to landowners in each location, $R_o \equiv \sum_{s \in o} r_o(s) l_o(s)$. Then, using as data (i) the expected trade shares, and (ii) the aggregate expenditure of slaughterhouses in cows, we can solve the linear system of N unknown where N is the number of locations or grid-cells, and estimate R_o for all $o = 1, 2, \dots, N$. In the third step, we use

¹⁰This is also shown in Proposition 2 in Panigrahi (2022).

¹¹The distance for two establishments within the same square is taken to be the average distance between two random points in a square of width D , which equals approximately $4.276 \times D$, and for all other point is equals the distance between their centroids plus approximately $5.79 \times D$. Therefore it is always greater than zero and so we can take logarithms.

	Dep. Var: Sourcing Intensity (oqt)	
	(1)	(2)
log(Distance)	−2.17*** (0.01)	
log(Distance) x To Farm		−2.17*** (0.01)
log(Distance) x To non ZDC sh		−2.03*** (0.10)
log(Distance) x To ZDC sh		−2.22*** (0.29)
Num. Obs	8381294	8381294
Num. Origin cell x Buyer Type x Year	12524	12524
Num. Destination cell x Buyer Type x Year	6468	6468

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 2: Fixed Effects Poisson Regression, D=20 km

transactions data to estimate the seller-buyer type fixed effects, $FE_o^q(s) = (\tilde{c}_o(s)\kappa_o^q(s))^{-\zeta}$. In the fourth and final step, we return to the disaggregated market clearing equation, and, relying on the estimated farm-level expected trade shares and the aggregate rental payments, we can calculate farm-level rents $r_o(s)l_o(s)$.

Estimating farm-level fixed effects

The most significant step is the estimation of the farm-level fixed effects, $FE_o^q(s)$. They capture the attractiveness of a farm s as a supplier relative to neighboring farms. The fixed effect captures both how productive the farm is and how suitable it is for each type of buyer. One way of estimating them, is taking farms' relative intensity of use as in equation (6). This is described in step 3A of appendix section D. The weakness of this strategy is that if a farm did not sell to a ZDC slaughterhouse in the data, it will not sell to a ZDC in any counterfactual simulation either, as this would only be rationalized by an infinite penalty. Alternatively, we can impose the following functional form on the deforestation penalty, let it be $\kappa^q(s) = \exp(\alpha^q \cdot \text{lhs}(l^{\text{defor}}(s)))$. Using this functional form and that unit costs of production are the same for cows sold to slaughterhouses as for cows sold to farms, the fixed effects can be re-estimated as shown in Step 3B in the appendix section D. This estimation strategy has the opposite weakness of the first one, though. It implies that in expectation all farms would sell some cows to farms and some cows to slaughterhouses, whereas in reality we see specialization. In our preferred specification, described in Step 3C of section D, we rely on the empirical observation that farms tend to partially specialize in stages of the cows' life-cycle. Farms seem to be either purely indirect, that is, they sell only to other farms, or a mix of direct and indirect, in which case they sell to

slaughterhouses and farms. With this in mind, we let farms have two productivities, a productivity to produce cows to sell to other farms $z^{\text{farm}}(s)$ and a productivity to produce cows to sell to slaughterhouses $z^{\text{sh}}(s)$. As a result, the unit costs for each type of production differ, $c^{\text{farm}}(s)$ and $c^{\text{sh}}(s)$. Thus, we can estimate the fixed effects separately for farms and for slaughterhouses as

$$FE_o^{\text{farm}*}(s) \equiv (\tilde{c}_o^{\text{farm}}(s)^{-\zeta} \tilde{\kappa}^{\text{farm}}(s)^{-\zeta})^* = \frac{\text{Intensity of use}^{\text{farm}}(s)}{\sum_{s' \in o} \text{Intensity of use}^{\text{farm}}(s')}$$

and

$$FE_o^{q*}(s) \equiv (\tilde{c}_o^{\text{sh}}(s)^{-\zeta})^* (\tilde{\kappa}_o^q(s)^{-\zeta})^* = \frac{(\tilde{c}_o^{\text{sh}}(s)^{-\zeta})^* \cdot \exp(-\zeta \alpha^{q*} \cdot \text{ih}(\text{defor}(s)))}{\sum_{s' \in o} (\tilde{c}_o^{\text{sh}}(s')^{-\zeta})^* \cdot \exp(-\zeta \alpha^{q*} \cdot \text{ih}(\text{defor}(s')))},$$

where the estimated relative costs of production for cows for slaughter is given by pooling together the intensity of use of cows for slaughter net of deforestation cost,

$$(\tilde{c}_o^{\text{sh}}(s)^{-\zeta})^* = \sum_{q \text{ is sh}} \frac{\text{Intensity of use}^q(s)}{\exp(-\zeta \alpha^{q*} \cdot \text{ih}(\text{defor}(s)))} \bigg/ \sum_{q \text{ is sh}} \sum_{s' \in o} \frac{\text{Intensity of use}^q(s')}{\exp(-\zeta \alpha^{q*} \cdot \text{ih}(\text{defor}(s')))}.$$

for $q \in \{\text{non ZDC s.h.}, \text{ZDC s.h.}\}$.

5 Counterfactuals

The first counterfactual scenario that we are interested in, is one in which there is no deforestation penalty, $\alpha' = 0$. We denote counterfactual prices and quantities with a prime, so that the counterfactual value of x is x' , and the changes with hats, so that $\hat{x} \equiv x'/x$. In order to estimate how deforestation changes in equilibrium, we assume that deforestation satisfies the first order conditions of maximization problem (1), that is that the marginal benefits of deforestation equals the marginal costs of deforestation,

$$\frac{d\phi(l_o^{\text{defor}}(s))}{dl_o^{\text{defor}}(s)} = \frac{1}{1-\rho} \frac{d \text{Landowner Rent}_o(s)}{dl_o^{\text{defor}}(s)},$$

where $\text{Landowner Rent}_o(s) = r_o(s)l_o(s)$. Taking the total derivative of landowner rents with respect to deforestation from equation (4), we can estimate the marginal deforestation costs in equilibrium. For the counterfactual, we assume locally linear deforestation costs,

Assumption: Locally linear deforestation costs. Assume that between baseline deforestation $l^{\text{defor}}(s)$ and counterfactual deforestation $l^{\text{defor}}(s)'$ the difference in marginal deforestation rates is negligible, that is, that the deforestation cost function is well approximated by a line in between $l^{\text{defor}}(s)$ and $l^{\text{defor}}(s)'$. Formally,

$$\frac{d\phi(l_o^{\text{defor}}(s))}{dl_o^{\text{defor}}(s)} = \frac{d\phi(l_o^{\text{defor}}(s)')}{dl_o^{\text{defor}}(s)}.$$

Then, if both the old and new equilibria have interior solutions for the optimal amount of deforestation,

$$\begin{aligned} \frac{d\phi(l_o^{\text{defor}}(s))}{dl_o^{\text{defor}}(s)} &= \frac{\beta\zeta}{(1-\rho)(1+\beta\zeta)} \left(r_o(s) - \frac{\alpha\zeta}{1+\zeta} \frac{1}{\sqrt{l_o^{\text{defor}}(s)^2 + 1}} \mathbb{E}Y_o^{\text{ZDC}}(s) \right) \\ &= \frac{d\phi(l_o^{\text{defor}}(s)')}{dl_o^{\text{defor}}(s)} = \frac{\beta\zeta}{(1-\rho)(1+\beta\zeta)} \left(r_o(s)' \right). \end{aligned} \quad (8)$$

Thus,

$$\hat{r}_o(s) \approx 1 - \frac{1}{r_o(s)} \frac{\alpha\zeta}{1+\zeta} \frac{1}{\sqrt{l_o^{\text{defor}}(s)^2 + 1}} \mathbb{E}Y_o^{\text{ZDC}}(s).$$

It may seem counterintuitive that with less deforestation restrictions the rental rates decrease, when we said that ZDC regulation makes deforested land less valuable by reducing the demand for it. This is because ZDC regulation does not directly reduce rental rates, which are an equilibrium outcome. What ZDC regulation reduces is *the sensitivity of rental rates to deforestation*, that is, $dr_o(s)/dl_o^{\text{defor}}(s) < 0$ further decreases in α . In equilibrium, the level of deforestation will change to maintain the optimality of the landowners' choices. Without ZDC regulation affecting the decision of how much to deforest, the only thing making deforestation undesirable is the direct costs for the landowner, $\phi(\cdot)$, which we assume to stay constant on the margin. Under the assumption above, the rental rates need to be such that the marginal benefits of deforestation don't change between baseline and counterfactual. Deforestation has a mechanical effect on landowner rents $r_o(s)l_o(s)$ by increasing $l_o(s)$ and an effect that is given by the supply chain policies, $l_o(s) \times dr_o(s)/dl_o^{\text{defor}}(s) < 0$. The latter one reduces the marginal benefits of deforestation, so if we want them to stay constant, we need that the mechanical benefits of deforestation,

$$\frac{dr_o(s)l_o(s)}{dl_o(s)} = \frac{\beta\zeta}{1+\beta\zeta} r_o(s),$$

be smaller in the counterfactual scenario, to compensate for the lack of deforestation penalty. That is, deforestation will go up until the point where the counterfactual rental rates are low enough to compensate for the marginal losses from the deforestation penalty observed at baseline.

From $\hat{r}_o(s)$, we can back out the counterfactual level of deforestation, although it requires a few more steps. This is because the link between rental rates and deforestation is mediated by the solution of the new trade probabilities in the new equilibrium. See appendix section E to see the next steps that need to be followed in order to estimate the changes in trade probabilities and the changes in the counterfactual scenario and thus the counterfactual levels of deforestation.

Partial Counterfactual A “partial” counterfactual exercise of interest could be to use the counterfactual rental rates with baseline old trade probabilities. If we use the baseline trade probabilities and the expenditure of slaughterhouses does not change then the revenues of farms, $r_o(s)l_o(s)$ also do not change, hence

$$l_o(s)' = l_o(s) \frac{r_o(s)}{r_o(s)'}$$

Having the counterfactual area in pasture, we have the counterfactual deforestation rate, which I define as the fraction of the area in pasture in 2019 that was deforested after 2009. As for the aggregate level of deforestation, it goes down by 23% with the policy, from at 19,500 km² in the no policy counterfactual to 15000 km² in the observed baseline data. In the full counterfactual, we expect this number to change. The places with more deforestation may also have lower revenues as a result, which may reduce the counterfactual $l_o(s)$. Therefore, we think that our interim results are likely to be an upper bound on the full effect of the policy. The full counterfactual estimation procedure is outline in the appendix section E and it is not likely to be too computationally demanding so we will to have it ready for the conference in January, hopefully alongside more counterfactual scenarios.

6 Conclusion

The cattle sector plays a central role in the process of deforestation in the Amazon. Given the high market concentration of the market at the level of slaughterhouses, there is a growing number of public and private initiatives that seek to regulate slaughterhouses in order to reduce deforestation. This means devolving the authority to penalize upstream deforestation through supply chain exclusion and directing public monitoring and enforcement efforts at slaughterhouses instead of individual farms.

These policies introduce frictions in a farm’s ability to sell to a segment of the market, regulated slaughterhouses, thus lowering the potential profits from deforestation. The main shortcoming of supply-chain policies is, however, that their effectiveness is thought to be easily undermined by leakage. First, because not all slaughterhouses are regulated, and thus farmers sell to the unregulated ones. Second, because, in supply-chain policies as they are currently implemented, slaughterhouses are only required to guarantee that their direct suppliers are not deforesting. However, a large fraction of farms sell exclusively to other farms, and thus they are not directly affected by this policy. These two forms of leakage pose serious concerns about the effectiveness of supply chain policies.

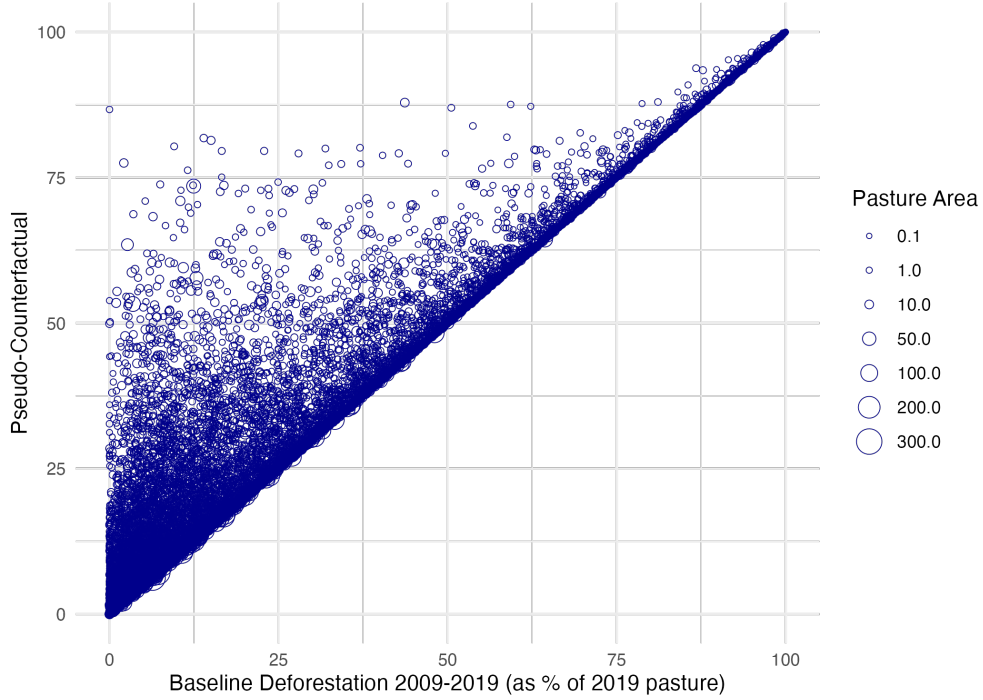


Figure 4: “Partial” counterfactual share of deforestation done after 2009

We develop a framework to study the supply chain of beef that takes cattle farms as unit of analysis and considers both farm-to-farm as well as farm-to-slaughterhouse transactions. We adapt a state-of-the-art model of firm-to-firm trade with endogenous network formation and trade frictions to our setting. The key contribution to models in the production network literature is to add deforestation as a farm-level decision that increases trade frictions for a set of buyers. In this framework, there are two stages in every period. First, deforestation decisions are made. Then, with all land use and deforestation histories taken as given, production and trade happen between farms and slaughterhouses.

Leveraging a unique dataset of farm-to-farm and farm-to-slaughterhouses sales matched with geo-localized farm deforestation, our results show that trade is heavily shaped by farm’s deforestation histories. In particular, farms in the same 10km wide square, with the same property size, and the same area in pasture and forest, are significantly less likely to trade with slaughterhouses, and especially ZDC slaughterhouses, if they have deforested. Moreover, the amount of deforestation matters: more deforestation leads to less trade with ZDC slaughterhouses. We model this as stemming from an iceberg trade cost that depends on farms’ deforestation since the policy’s starting period (2009), and refer to it as the deforestation penalty. Our results indicate that ZDC regulation changes the consequences of deforestation for farms. These effects are robust to different specifications, statistically significant, and large. We estimate that a farm that deforests is 16% less likely to be chosen as a ZDC’s slaughterhouse supplier, and each additional 10% increase in deforestation is associated with an 8.3% decrease in trade probability with ZDC

slaughterhouses.

These encouraging results will serve as foundation for the next step in our analysis: modeling how the deforestation trade penalty influences deforestation incentives and thus deforestation decisions. Incorporating this component and closing the model, either analytically or using hat algebra, will enable us to address the paper’s ultimate goal: estimating the counterfactual impact of supply chain policy changes, such as the EUDR, on deforestation while accounting for endogenous responses within the supply chain network.

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A Continuum Approximation for Large Network Economies

The following definition formalizes the notion of the limiting economy in the context of this paper. RE-DO with my notation

Definition. Consider a sequence of finite economies $\{\mathcal{E}_t : t \in \mathbb{N}\}$ where $\mathcal{E}_t \equiv \{\mathcal{S}_t, \mathcal{L}_t, \mathcal{J}_t\}$ is such that the t^{th} economy has the form $\mathcal{M}_t = \{m_1, \dots, m_{M_t}\} \subset [0, 1]$, $\mathcal{L}_t = \{\ell_1, \dots, \ell_{L_t}\} \subset [0, 1]$ and $\mathcal{J}_t = \mathcal{J}$. The uniform distribution on \mathcal{M}_t is given by $\mathcal{U}_t^M(\mathcal{M}_t^0) = \frac{M_t^0}{M_t}$ for all $\mathcal{M}_t^0 \subset \mathcal{M}_t$. Similarly, the uniform distribution on \mathcal{L}_t is given by $\mathcal{U}_t^L(\mathcal{L}_t^0) = \frac{L_t^0}{L_t}$ for all $\mathcal{L}_t^0 \subset \mathcal{L}_t$. Then, $\{\mathcal{E}_t : t \in \mathbb{N}\}$ is a discretizing sequence of economies if it satisfies:

- (1) $\mathcal{M}_t \subset \mathcal{M}_{t+1}$ and $\mathcal{L}_t \subset \mathcal{L}_{t+1}$ for all t ,
- (2) $\lim_{t \rightarrow \infty} \mathcal{U}_t^M(\mathcal{M}_t \cap [a_l, a_h]) = \mathcal{U}([a_l, a_h])$,
- (3) $\lim_{t \rightarrow \infty} \mathcal{U}_t^L(\mathcal{L}_t \cap [a_l, a_h]) = \mathcal{U}([a_l, a_h])$,

where $\mathcal{U}(\bullet)$ denotes the uniform distribution with support over $[0, 1]$ and $[a_l, a_h] \subset [0, 1]$.

Assumption 2. The discretizing sequence of economies $\{\mathcal{E}_t : t \in \mathbb{N}\}$ satisfies the following conditions: ¹²

- (1) $\{\lambda_t, a_{0,t} : t \in \mathbb{N}\}$ is such that $\lambda_t = o(M_t)$ and $\lambda_t a_{0,t}^\zeta = \Theta(1)$
- (2) $\{M_{d,t}, L_{d,t} : d \in \mathcal{J}, t \in \mathbb{N}\}$ is such that $M_{d,t} = \Theta(M_t)$ and $L_{d,t} = \Theta(L_t)$ for all $d \in \mathcal{J}$

B Deriving Market Clearing Equation

Joint Distribution of the Lowest and the Second Lowest Effective Costs. We begin by characterizing the joint distribution of the lowest and second lowest effective cost available to buyer b of type q located at d ,

$$\tilde{F}_{p_d^q}(p^{(1)}, p^{(2)}) = \mathbb{P}(p_d^{q*}(b, k) \leq p^{(1)}, p_d^q(b, k) \geq p^{(2)}).$$

To do so, we evaluate the probability with which b receives exactly one offer with an effective cost no greater than $p^{(1)}$ and no other offers less than $p^{(2)}$ ($> p^{(1)}$). The lowest cost offer $p^{(1)}$ can be from any one of the locations. We evaluate the probability with which this offer is from any given location o and sum it across all locations. The

¹²For any two functions $f(n)$ and $g(n)$, $f(n) = o(g(n)) \implies \lim_{n \rightarrow \infty} \frac{f(n)}{g(n)} = 0$ and $f(n) = \Theta(g(n)) \implies \limsup_{n \rightarrow \infty} \frac{|f(n)|}{|g(n)|} < \infty$ and $\limsup_{n \rightarrow \infty} \left| \frac{f(n)}{g(n)} \right| > 0$.

probability with which b receives one offer with an effective cost no greater than $p^{(1)}$ from o and no other offers less than $p^{(2)}$ across all locations is given by:

$$\left\{ \begin{array}{ll} \left(\binom{S_o}{1} \frac{\lambda}{S} \mathbb{P} \left(\frac{c_o^q(s) \kappa_o^q(s) \tau_{od}^q}{a_{od}(s, b, k)} \leq p^{(1)} \right) \right. & \text{if } o \neq d \\ \times \left(1 - \frac{\lambda}{S} \mathbb{P} \left(\frac{c_o^q(s) \kappa_o^q(s) \tau_{od}^q}{a_{od}(s, b, k)} \leq p^{(2)} \right) \right)^{S_o-1} & \\ \times \left(1 - \frac{\lambda}{S} \mathbb{P} \left(\frac{c_d(s) \kappa_o^q(s) \tau_{dd}^q}{a_{dd}(s, b, k)} \leq p^{(2)} \right) \right)^{S_d-1} & \\ \times \prod_{o' \notin \{o, d\}} \left(1 - \frac{\lambda}{M} \mathbb{P} \left(\frac{c_{o'}^q(s) \tau_{o'd}^q}{a_{o'd}(s, b, k)} \leq p^{(2)} \right) \right)^{M_{o'}} & \\ \left(\binom{S_o-1}{1} \frac{\lambda}{S} \mathbb{P} \left(\frac{c_o^q(s) \kappa_o^q(s) \tau_{od}^q}{a_{od}(s, b, k)} \leq p^{(1)} \right) \right. & \text{if } o = d \\ \times \left(1 - \frac{\lambda}{S} \mathbb{P} \left(\frac{c_o^q(s) \kappa_o^q(s) \tau_{od}^q}{a_{od}(s, b, k)} \leq p^{(2)} \right) \right)^{S_o-2} & \\ \times \prod_{o' \neq o} \left(1 - \frac{\lambda}{S} \mathbb{P} \left(\frac{c_{o'}^q(s) \kappa_o^q(s) \tau_{o'd}^q}{a_{o'd}(s, b, k)} \leq p^{(2)} \right) \right)^{S_{o'}} & \end{array} \right.$$

Under Assumption 2, the probability with which b encounters exactly one supplier who can deliver at a cost no greater than $p^{(1)}$ and encounters no other suppliers with offers less than $p^{(2)}$ across all locations is given by:

$$\tilde{F}_{p_d^q}(p^{(1)}, p^{(2)}) = \sum_o \lambda \mu_o \mathbb{P} \left(\frac{c_o^q(s) \kappa_o^q(s) \tau_{od}^q}{a_{od}(s, b, k)} \leq p^{(1)} \right) \exp \left(- \sum_{o'} \lambda \mu_{o'} \mathbb{P} \left(\frac{c_{o'}^q(s) \kappa_o^q(s) \tau_{o'd}^q}{a_{o'd}(s, b, k)} \leq p^{(2)} \right) \right)$$

Using the limit $\lim_{t \rightarrow \infty} \lambda_t a_{0,t}^\zeta \rightarrow 1$, this can be further simplified as $A_d^q(p^{(1)})^\zeta \exp \left(-A_d^q(p^{(2)})^\zeta \right)$ where $A_d^q = \sum_o \mu_o (\tau_{od}^q)^{-\zeta} \mathbb{E} [(c_o^q(\cdot) \kappa_o^q(\cdot))^{-\zeta}]$ is obtained as follows:

$$\begin{aligned} A_d^q p^\zeta &= \sum_o \lambda \mu_o \mathbb{P} \left(\frac{c_o^q(s) \kappa_o^q(s) \tau_{od}^q}{a_{od}(s, b, k)} \leq p \right) \\ &= \sum_o \lambda \mu_o \mathbb{E}_{\{c_o^q\}} \left[1 - F_a \left(\frac{c_o^q(s) \kappa_o^q(s) \tau_{od}^q}{p} \right) \right] \\ &= \left(\sum_o \mu_o (\tau_{od}^q)^{-\zeta} \mathbb{E} [(c_o^q(\cdot) \kappa_o^q(\cdot))^{-\zeta}] \right) p^\zeta \\ \implies A_d^q &= \sum_o \mu_o (\tau_{od}^q)^{-\zeta} \mathbb{E} [(c_o^q(\cdot) \kappa_o^q(\cdot))^{-\zeta}] \end{aligned}$$

The density function is then obtained by the negative cross-derivative of $\tilde{F}_{p_d^q}(p^{(1)}, p^{(2)})$ as follows:

$$\begin{aligned}\tilde{F}'_{p_d^q}(p^{(1)}, p^{(2)}) &= -\frac{\partial^2 F_{p_d^q}(p^{(1)}, p^{(2)})}{\partial p^{(1)} \partial p^{(2)}} \\ &= -\frac{\partial \left(A_d^q (p^{(1)})^\zeta \right)}{\partial p^{(1)}} \frac{\partial \left(\exp \left(-A_d^q (p^{(2)})^\zeta \right) \right)}{\partial p^{(2)}} \\ &= \zeta^2 (A_d^q)^2 (p^{(1)} p^{(2)})^{\zeta-1} e^{-A_d^q (p^{(2)})^\zeta}\end{aligned}$$

Distribution of Effective Prices. We derive an expression for $F_{p_d^q}(p)$, that is, the probability with which any firm b of type q located in d faces an effective price no greater than p for one of its tasks k . Firm b faces an effective price no greater than p if the second-lowest cost available to it is no less than p . This is obtained as:

$$\begin{aligned}F_{p_d^q}(p) &= \int_0^p \left(\int_0^{p^{(2)}} F'_{p_d^q}(p^{(1)}, p^{(2)}) dp^{(1)} \right) dp^{(2)} \\ &= 1 - A_d^q p^\zeta \exp(-A_d^q p^\zeta) - \exp(-A_d^q p^\zeta)\end{aligned}$$

Derivation of Market Access. Given supplier's rental rate $r_o(s)$, the cost of production is given by

$$\begin{aligned}c_o^q(s) &= \frac{1}{z_o^q(s)} r_o(s)^\beta \left(\prod_{k=1}^K p_o^{\text{farm}}(s, k)^{1/K} \right)^{1-\beta} \\ \Rightarrow \mathbb{E} \left[c_o^q(\cdot)^{-\zeta} \kappa_o^q(\cdot)^{-\zeta} \right] &= \mathbb{E} \left[\left(\frac{r_o(\cdot)^\beta \left(\prod_{k=1}^K p_o^{\text{farm}}(\cdot, k)^{1/K} \right)^{1-\beta}}{z_o^q(\cdot)} \right)^{-\zeta} \kappa_o^q(\cdot)^{-\zeta} \right] \\ &= \mathbb{E} \left[\prod_{k=1}^K p_o^{\text{farm}}(\cdot, k)^{-(1-\beta)\zeta/K} \right] \mathbb{E} \left[r_o(\cdot)^{-\zeta\beta} z_o^q(\cdot)^\zeta \kappa_o^q(\cdot)^{-\zeta} \right] \\ &= \prod_{k=1}^K \mathbb{E} \left[p_o^{\text{farm}}(\cdot, k)^{-(1-\beta)\zeta/K} \right] \mathbb{E} \left[r_o(\cdot)^{-\zeta\beta} z_o^q(\cdot)^\zeta \kappa_o^q(\cdot)^{-\zeta} \right] \\ &= \overline{r_o^{-\zeta\beta} (z_o^q)^\zeta (\kappa_o^q)^{-\zeta}} \Gamma \left(2 - \frac{1-\beta}{K} \right)^K (A_o^{\text{farm}})^{1-\beta}\end{aligned}$$

This is because, using the distribution of effective prices derived above, the probability density function equals

$$f_{p_d^q}(p) = (A_d^q)^2 p^{2\zeta-1} \exp\left(-A_d^q p^\zeta\right),$$

so that

$$\begin{aligned} \mathbb{E}\left[p_o^q(\cdot, k)^{-(1-\beta)\zeta/K}\right] &= \int_0^\infty p^{\zeta(2-(1-\beta)/K)-1} (A_d^q)^2 \exp\left(-A_d^q p^\zeta\right) dp \\ &= (A_d^q)^{\frac{1-\beta}{K}} \Gamma\left(2 - \frac{1-\beta}{K}\right) \end{aligned}$$

This implies that the market access of each location for farms $\{A_d^{\text{farm}}\}_d$ solves the following fixed point problem:

$$A_d^{\text{farm}} = \sum_o \mu_o r_o^{-\zeta\beta} (z_o^{\text{farm}})^\zeta (\kappa_o^{\text{farm}})^{-\zeta} \Gamma\left(2 - \frac{1-\beta}{K}\right)^K (\tau_{od}^{\text{farm}})^{-\zeta} (A_o^{\text{farm}})^{1-\beta}$$

and the market access terms for slaughterhouses equal

$$A_d^q = \sum_o \mu_o r_o^{-\zeta\beta} (z_o^q)^\zeta (\kappa_o^q)^{-\zeta} \Gamma\left(2 - \frac{1-\beta}{K}\right)^K (\tau_{od}^q)^{-\zeta} (A_o^{\text{farm}})^{1-\beta}$$

Sourcing probabilities (origin-destination). For any realization of farm-level productivities, deforestation levels, and trade costs, the probability with which any firm at d of type q sources from firms at o for any of its tasks is given by

$$\begin{aligned} \pi_{od}^{q0}(\bullet, -) &= \left(\lim_{t \rightarrow \infty} \frac{S_o}{S}\right) \left(\lim_{t \rightarrow \infty} \frac{1}{S_o} \sum_{s \in \mathcal{S}_o} \pi_{od}^{q0}(s, -)\right) \\ &= \left(\lim_{t \rightarrow \infty} \frac{S_o}{S}\right) \left(\lim_{t \rightarrow \infty} \frac{1}{S_o} \sum_{s \in \mathcal{S}_o} \frac{c_o(s)^{-\zeta} \kappa_o^q(s)^{-\zeta} (\tau_{od}^q)^{-\zeta}}{A_d^q}\right) \\ &= \frac{\mu_o \mathbb{E}\left[c_o(\cdot)^{-\zeta} \kappa_o^q(\cdot)^{-\zeta}\right] (\tau_{od}^q)^{-\zeta}}{A_d^q} \\ &= \frac{\mu_o r_o^{-\zeta\beta} (z_o^q)^\zeta (\kappa_o^q)^{-\zeta} \Gamma\left(2 - \frac{1-\beta}{K}\right)^K (\tau_{od}^q)^{-\zeta} (A_o^{\text{farm}})^{1-\beta}}{A_d^q} \end{aligned}$$

The law of large numbers implies that in the limiting economy

$$\frac{1}{B_d^q} \sum_{b \in \mathcal{B}_d^q} \pi^q(\bullet, b) \xrightarrow{t \rightarrow \infty} \pi^{q0}(\bullet, -).$$

Thus we can close the model to characterize the rental rates of land in equilibrium in the limiting economy.

Sourcing probabilities (supplier-destination). For any realization of farm-level productivities, deforestation levels, and trade costs, the probability with which any firm at d of type q sources from firms at o for any of its tasks is given by

$$\begin{aligned}\pi_{od}^{q0}(s, -) &= \tilde{c}_o^q(s)^{-\zeta} \tilde{\kappa}_o^q(s)^{-\zeta} \pi_{od}^{q0}(\bullet, -) \\ &= \tilde{c}_o^q(s)^{-\zeta} \tilde{\kappa}_o^q(s)^{-\zeta} \frac{\overline{\mu_o r_o^{-\zeta \beta} z_o^\zeta (\kappa_o^q)^{-\zeta} \Gamma \left(2 - \frac{1-\beta}{K} \right)^K (\tau_{od}^q)^{-\zeta} (A_o^{\text{farm}})^{1-\beta}}{A_d^q}\end{aligned}$$

Closing the model. For any realization of σ , land demand by farm b at d can be expressed as:

$$r_d(b)l_d(b) = \beta c_d^{\text{farm}}(b)y_d(b)$$

Where $l_d(b)$ is data. Goods market clearing condition for firm s located at o can be simplified as:

$$\begin{aligned}y_o(s) &= \sum_d \sum_q \sum_{b \in \mathcal{B}_d^q} \sum_{k=1}^K \frac{\tau_{od}^q(s) \kappa^q(s) m_{od}^q(s, b, k)}{a_{od}(s, b, k)} \\ \implies c_o^{\text{farm}}(s)y_o(s) &= \sum_d (1 - \beta) \sum_{b \in \mathcal{S}_d} \left(\frac{1}{K} \sum_{k=1}^K \frac{\mathbf{1}\{s = s_d^*(b, k)\}}{\bar{m}_d(b, k)} \right) c_d(b)y_d(b) \\ &\quad + \sum_d \sum_{q \in \{SH, ZDC\}} \sum_{b \in \mathcal{B}_d^q} \left(\frac{1}{K} \sum_{k=1}^K \frac{\mathbf{1}\{s = s_d^*(b, k)\}}{\bar{m}_d(b, k)} \right) X^q(b)\end{aligned}$$

We can simplify the LHS by making use of the land market clearing condition as:

$$\text{Supply}(s) = c_o^{\text{farm}}(s)y_o(s) = \frac{r_o(s)l_o(s)}{\beta}$$

We can the demand for cattle from farms as follows

Farm cattle demand (s)

$$\begin{aligned}
&= \sum_d (1 - \beta) \sum_{b \in \mathcal{B}_d^{\text{farm}}} \left(\frac{1}{K} \sum_{k \in \mathcal{K}} \frac{\mathbf{1}\{s_d^*(b, k) = s\}}{\bar{m}_d(b, k)} \right) c_d(b) y_d(b) \\
&\quad \underbrace{\left(\frac{1}{B_d^{\text{farm}}} \sum_{b \in \mathcal{B}_d^{\text{farm}}} \left(\frac{1}{K} \sum_{k=1}^K \frac{\mathbf{1}\{s_d^*(b, k) = s\}}{\bar{m}_d(b, k)} \right) c_d(b) y_d(b) \right)}_{(A)} \\
&\quad \frac{\sum_d (1 - \beta)}{\underbrace{\frac{1}{B_d^{\text{farm}}} \sum_{b \in \mathcal{B}_d^{\text{farm}}} c_d(b) y_d(b)}_{(B)}} \times \sum_{b \in \mathcal{B}_d^{\text{farm}}} c_d(b) y_d(b)
\end{aligned}$$

Term (A) can be simplified as follows:

$$\begin{aligned}
(A) &= \frac{1}{B_d^{\text{farm}}} \sum_{b \in \mathcal{B}_d^{\text{farm}}} \left(\frac{1}{K} \sum_{k \in \mathcal{K}} \frac{\mathbf{1}\{s_d^*(b, k) = s\}}{\bar{m}_d(b, k)} \right) c_d(b) y_d(b) \\
&\xrightarrow{t \rightarrow \infty} \mathbb{E} \left[\left(\frac{1}{K} \sum_{k \in \mathcal{K}} \frac{\mathbf{1}\{s_d^*(\cdot, k) = s\}}{\bar{m}_d(\cdot, k)} \right) c_d(\cdot) y_d(\cdot) \right] \\
&= \mathbb{E} \left[\left(\frac{1}{K} \sum_{k \in \mathcal{K}} \frac{\mathbf{1}\{s_d^*(\cdot, k) = s\}}{\bar{m}_d(\cdot, k)} \right) \right] \mathbb{E} [c_d(\cdot) y_d(\cdot)] \\
&= \mathbb{E} \left[\left(\frac{1}{K} \sum_{k \in \mathcal{K}} \frac{\mathbf{1}\{s_d^*(\cdot, k) = s\}}{\bar{m}_d(\cdot, k)} \right) \right] \mathbb{E} [c_d(\cdot) y_d(\cdot)] \\
&= \frac{1}{K} \sum_{k \in \mathcal{K}} \mathbb{E} \left[\frac{\mathbf{1}\{s_d^*(\cdot, k) = s\}}{\bar{m}_d(\cdot, k)} \right] \mathbb{E} [c_d(\cdot) y_d(\cdot)] \\
&= \mathbb{E} \left[\frac{\mathbf{1}\{s_d^*(\cdot, \cdot) = s\}}{\bar{m}_d(\cdot, \cdot)} \right] \mathbb{E} [c_d(\cdot) y_d(\cdot)] \\
&= \mathbb{E} \left[\frac{1}{\bar{m}_d(\cdot, \cdot)} \right] \mathbb{E} [\mathbf{1}\{s_d^*(\cdot, \cdot) = s\}] \mathbb{E} [c_d(\cdot) y_d(\cdot)] \\
&= \frac{\zeta}{\zeta + 1} \mathbb{E}(\pi_{od}^{\text{farm}}(s, -)) \mathbb{E} [c_d(\cdot) y_d(\cdot)]
\end{aligned}$$

Term (B) can be simplified as follows:

$$\begin{aligned}
(B) &= \frac{1}{B_d^{\text{farm}}} \sum_{b \in \mathcal{B}_d^{\text{farm}}} c_d(b) y_d(b) \\
&\xrightarrow{t \rightarrow \infty} \mathbb{E} [c_d(\cdot) y_d(\cdot)]
\end{aligned}$$

Substituting (A) and (B) back in the Farm Input Demand, we obtain:

$$\text{Farm Input Demand (s)} = \sum_d \frac{1-\beta}{\beta} \frac{\zeta}{\zeta+1} \mathbb{E}(\pi_{od}^{\text{farm}}(s, -)) \sum_{b \in \mathcal{B}_d^{\text{farm}}} r_d(b) l_d(b)$$

We can simplify term Slaughterhouses' demand as

$$\begin{aligned} & \text{Slaughterhouse of Type q Demand (s)} \\ &= \sum_d \sum_{b \in \mathcal{B}_d^q} \left(\frac{1}{K} \sum_{k=1}^K \frac{\mathbf{1}\{s_d^*(b, k) = s\}}{\bar{m}_d(b, k_1)} \right) X_d^q(b) \\ &\xrightarrow{t \rightarrow \infty} \sum_d \mathbb{E} \left[\frac{\mathbf{1}\{s_d^*(\cdot, k) = s\}}{\bar{m}_d(\cdot, k)} \right] \sum_{b \in \mathcal{B}_d^q} X_d^q(b) \\ &= \sum_d \mathbb{E} \left[\frac{1}{\bar{m}_d(\cdot, \cdot)} \right] \mathbb{E} [\mathbf{1}\{s_d^*(\cdot, \cdot) = s\}] X_d^q \\ &= \sum_d \frac{\zeta}{\zeta+1} \mathbb{E}(\pi_{od}^q(s, -)) X_d^q \end{aligned}$$

Market clearing equation. Putting these together we can further simplify the cattle market clearing condition. If supplier s charges an average markup of $\frac{\zeta+1}{\zeta} > 1$, then, after multiplying both sides of the market clearing equation by β and substituting in the Farm and Slaughterhouse input demand formulas,

$$\begin{aligned} r_o(s) l_o(s) &= \frac{\zeta}{\zeta+1} (1-\beta) \sum_d \mathbb{E}(\pi_{od}^{\text{farm}}(s, -)) \sum_{b \in \mathcal{B}_d^{\text{farm}}} r_d(b) l_d(b) \\ &+ \frac{\beta\zeta}{\zeta+1} \sum_d \mathbb{E}(\pi_{od}^{\text{non ZDC SH}}(s, -)) X_d^{\text{non ZDC s.h.}} \\ &+ \frac{\beta\zeta}{\zeta+1} \sum_d \mathbb{E}(\pi_{od}^{\text{ZDC SH}}(s, -)) X_d^{\text{ZDC s.h.}} \end{aligned}$$

The first term in the RHS equal the total payments from farms who buy the cows from s as intermediate inputs, divided by the markup and multiplied by the share that goes towards input costs. This is because the payments farm b makes towards cow inputs equal $\frac{1-\beta}{\beta}$ times the payments it makes towards land rental rates. Out of those, in expectation, a fraction $\pi_{od}^{\text{farm}}(s, -)$ goes towards s . As the number of buyers in d increases, by the law of large numbers, the income from sales to farms in d approaches its expected value. The second and third terms are the total amount of payments the non ZDC and ZDC slaughterhouses make, collectively, towards cows from s .

This can also be written in a simplified form that makes the dependence on $r_o(s)$ more explicit,

$$r_o(s)^{1+\beta\zeta} l_o(s) = G_o(s)$$

where

$$\begin{aligned} G_o(s) \equiv & \frac{\zeta}{\zeta+1} (p_o)^{-\zeta(1-\beta)} \left((1-\beta) \sum_d \frac{\tau_{od}^{-\zeta} z_o^{\text{farm}}(s)^\zeta}{\sum_{o'} (c_{o'}^{\text{farm}} \tau_{o'd})^{-\zeta}} \sum_{b \in \mathcal{B}_d^{\text{farm}}} r_d(b) l_d(b) \right. \\ & + \beta \sum_d \frac{\tau_{od}^{-\zeta} z_o^{\text{sh}}(s)^\zeta}{\sum_{o'} (c_{o'}^{\text{sh}} \tau_{o'd})^{-\zeta}} X_d^{\text{non ZDC s.h.}} \\ & \left. + \beta \sum_d \frac{\tau_{od}^{-\zeta} z_o^{\text{sh}}(s)^\zeta \kappa_o^{\text{ZDC}}(s)^{-\zeta}}{\sum_{o'} (c_{o'}^{\text{sh}} \kappa_{o'}^{\text{ZDC}} \tau_{o'd})^{-\zeta}} X_d^{\text{ZDC s.h.}} \right) \end{aligned}$$

C Optimal level of deforestation

The level of deforestation in equilibrium is assumed to be optimal, that is to maximize the long term profits of the landowner, who receives land rents and pays the direct costs of cutting down the trees. Thus, we need to (i) understand how land owner rents vary with the amount of pasture land in a farm, and (ii) the deforestation cost function.

First let us calculate how the land rents of a landowner change with changing levels of area in pasture, $l_o(s)$. There will be a mechanical increase as there is more land to produce in, but also a decrease in the rental rate of land as more land of that quality and in that farm is being made available. In what follows we describe how the equilibrium land rents change with $l_o(s)$.

$$\text{Landowner Rent}_o(s) \equiv r_o(s) l_o(s),$$

so that the profits of a landowner equal

$$\text{Landowner Profits}_o(s) \equiv \frac{1}{1-\rho} r_o(s) l_o(s) - \phi(l_o(s) - \underline{l}_o(s))$$

where $\phi(\cdot)$ is the deforestation cost function, $\underline{l}_o(s)$ is the amount of already deforested land so that $l_o^{\text{defor}} = l_o(s) - \underline{l}_o(s)$, and ρ is a time discount factor to account for the fact that rent increases are perceived over a lifetime.

First order conditions of Landowner. If there is an interior solution (the landowner deforests a positive amount but not the entire farm, then the marginal

$$\frac{1}{1-\rho} \frac{d \text{Landowner Rent}(s)}{d l_o^{\text{defor}}(s)} = \phi'(l_o^{\text{defor}}(s))$$

We then employ the structure of our model to estimate how landowner rents depend on the amount of land deforested.

Case 1. No ZDC policy. Without deforestation penalties for any buyer, in equilibrium,

$$\frac{d\text{Landowner Rent}(s)}{dl_o^{\text{defor}}(s)} = \frac{\partial r_o(s)}{\partial l_o(s)} l_o(s) + r_o(s) \approx r_o(s) \left(-\frac{1}{1 + \beta\zeta} + 1 \right) = \frac{\beta\zeta}{1 + \beta\zeta} r_o(s).$$

Elasticity interpretation. A 1% increase in pasture land leads to a $\frac{\beta\zeta}{1+\beta\zeta}$ % increase in land rents received by the owner.

Case 2. ZDC policy. If we also consider deforestation penalties for ZDC buyers, then

$$\begin{aligned} \frac{d\text{Landowner Rent}(s)}{dl_o^{\text{defor}}(s)} &\approx \frac{\beta\zeta}{1 + \beta\zeta} r_o(s) + \frac{\partial r_o(s) l_o(s)}{\partial l_o^{\text{defor}}(s)} \\ &\approx \frac{\beta\zeta}{1 + \beta\zeta} \left(r_o(s) - \frac{1}{\kappa_o^{\text{ZDC}}(s)} \underbrace{\frac{\partial \kappa_o^{\text{ZDC}}(s)}{\partial l_o^{\text{defor}}(s)}}_{>0} Y^{\text{ZDC}*}(s) \right). \end{aligned}$$

where $Y^{\text{ZDC}}(s)$ is the total amount of sales of farm s to ZDC buyers (endogenous)

Assuming that $\kappa_o^{\text{ZDC}}(s)$ equals $\kappa_o^{\text{ZDC}}(s) = \exp(\alpha \cdot \ln(l_o^{\text{defor}}(s)))$, which asymptotically approximates $\kappa l_o^{\text{defor}}(s)^\alpha$, we get that

$$\frac{1}{\kappa_o^{\text{ZDC}}(s)} \frac{\partial \kappa_o^{\text{ZDC}}(s)}{\partial l_o^{\text{defor}}(s)} = \alpha \frac{1}{\sqrt{l_o^{\text{defor}}(s)^2 + 1}},$$

and hence,

$$\frac{d\text{Landowner Rent}(s)}{dl_o^{\text{defor}}(s)} \approx \frac{\beta\zeta}{1 + \beta\zeta} \left(r_o(s) - \frac{\alpha\zeta}{1 + \zeta} \frac{1}{\sqrt{l_o^{\text{defor}}(s)^2 + 1}} Y^{\text{ZDC}*}(s) \right).$$

Then the first order condition, in the general case which allows for a deforestation penalty,

$$\frac{d\phi(l_o^{\text{defor}}(s))}{dl_o^{\text{defor}}(s)} = \frac{\beta\zeta}{(1 - \rho)(1 + \beta\zeta)} \left(r_o(s) - \frac{\alpha\zeta}{1 + \zeta} \frac{1}{\sqrt{l_o^{\text{defor}}(s)^2 + 1}} Y^{\text{ZDC}*}(s) \right)$$

Approximation needed to estimate $\frac{\partial r_o(s)}{\partial l_o(s)} l_o(s)$. In order to estimate $\frac{\partial r_o(s)}{\partial l_o(s)} l_o(s)$ we rely on an approximation. Let's write the market clearing equation as

$$r_o(s)^{1+\beta\zeta} l_o(s) = G_o(s).$$

Notice that $G_o(s)$ comes from writing explicitly the dependence of the trade shares on the rental rate and decomposing it.

$$\mathbb{E}[\pi_{od}^q(s, b)] = \frac{(c_o(s) \tau_{od} \kappa_o^q(s))^{-\zeta}}{\sum_{s'} (c_{o'}(s') \tau_{o'd} \kappa_o^q(s'))^{-\zeta}} = r_o(s)^{-\beta\zeta} \frac{\left(\frac{1}{z_o(s)} (p_o^*)^{1-\beta} \tau_{od} \kappa_o^q(s) \right)^{-\zeta}}{\sum_{s'} (c_{o'}(s') \tau_{o'd} \kappa_o^q(s'))^{-\zeta}}$$

By doing this, we are taking into account how the numerator changes with $r_o(s)$, via $c_o(s)$, but not how the denominator does, since s is on of the terms added in the denominator. More over, $r_o(s)$ could also cause changes in p^* in equilibrium. As such, the full formula for how rental rates change with land in pasture, calculated using the simplified market clearing equation, solving for $l_o(s)$ and using the formula for the derivative of an inverse function, and multiplying by $l_o(s)$, equals

$$\frac{\partial r_o(s)}{\partial l_o(s)} l_o(s) = \frac{r_o(s)}{-(1 + \beta\zeta) + \frac{dG_o(s)}{dr_o(s)} \frac{r_o(s)}{G_o(s)}} \approx -\frac{r_o(s)}{1 + \beta\zeta}$$

The assumption required for this approximation is that

$$\left| \frac{\partial G_o(s)}{\partial r_o(s)} \frac{r_o(s)}{G_o(s)} \right| \ll 1 + \beta\zeta.$$

That is, that the elasticity of $G_o(s)$ with respect to $r_o(s)$ is sufficiently small. Intuitively, it means that farms are small players and do not perceive for their effects on equilibrium prices.

D Model inversion

Step 1: Compute the expected aggregate trade shares $\pi_{od}^{q*}(\bullet, -)$.

Rather than using the observed aggregate trade shares, which are often zero, we use the fitted shares from the gravity regressions

$$\pi_{od}^q(\bullet, -) = \exp \left(v_{ot}^q + \eta_{dt}^q + \beta^q \cdot \log (\text{Distance}_{od}) \right) + u_{odt}$$

where the left hand side is defined as

$$\pi_{od}^q(\bullet, -) = \text{Average Sourcing Intensity}_{od}^q = \frac{1}{N_d^q} \sum_{b \in d} \sum_{s \in o} \pi_{od}^q(s, b).$$

Step 2: Compute aggregate rental rate payments in initial state.

In the initial state, the vector of rental rates $(r_o(s)l_o(s) : s \in \mathcal{S})$ is the solution to the following system of equations:

$$\begin{aligned} \frac{r_o(s)l_o(s)}{\beta} &= \frac{\zeta}{\zeta + 1} \frac{1 - \beta}{\beta} \sum_d \pi_{od}^{\text{farm}*}(s, -) \sum_{b \in \mathcal{B}_d^{\text{farm}}} r_d(b)l_d(b) \\ &+ \frac{\zeta}{\zeta + 1} \sum_d \pi_{od}^{\text{nonZDCs.h.*}}(s, -) X_d^{\text{nonZDCs.h.}} \\ &+ \frac{\zeta}{\zeta + 1} \sum_d \pi_{od}^{\text{ZDCs.h.*}}(s, -) X_d^{\text{ZDCs.h.}} \end{aligned}$$

Solving this linear system of size 60,000 would be computationally challenging. Fortunately, this is not necessary, as the RHS only has the region-level aggregate land rents. Define region-level land rents,

$$R_d \equiv \sum_{b \in \mathcal{B}_d} r_d(b)l_d(b).$$

Then, adding up over all farms s in location o who are buyers, and denoting the total expected expenditure of slaughterhouses on cows from origin o by $Y_o^{\text{s.h.*}}$,¹³ we get the following aggregate market clearing equation,

$$R_o = \frac{\zeta}{\zeta + 1} (1 - \beta) \sum_d \pi_{od}^{\text{farm}*}(\bullet, -) R_d + \frac{\zeta}{\zeta + 1} \beta Y_o^{\text{s.h.*}}$$

since, by construction, the farm fixed effects within a location add up to 1.

Let $\Pi^{\text{farm}*}$ be the matrix with entries $\pi_{od}^{\text{farm}*}(\cdot, -)$, the fitted values from the gravity equations, then aggregate land rents are the solution to the matrix equation

¹³This differs from actual expenditures because we use expected rather than observed trade shares.

$$\begin{aligned}\vec{R} &= \frac{\zeta}{1+\zeta}(1-\beta)\mathbf{\Pi}^{\text{farm}*}\vec{R} + \frac{\zeta}{1+\zeta}\beta\vec{Y}^{\text{s.h.*}} \\ \Rightarrow \vec{R} &= \left(\frac{1+\zeta}{\zeta}\mathbf{I} - (1-\beta)\mathbf{\Pi}^{\text{farm}*}\right)^{-1}\beta\vec{Y}^{\text{s.h.*}}\end{aligned}$$

So if there are N regions, we need to invert an $N \times N$ matrix.

Step 3A: Estimate farm-buyer type level fixed effects with buyer-specific unit costs

Using this strategy, farm level unit costs $\tilde{c}_o^q(s)$ vary by buyer type and are estimated so that they exactly match the farm fixed effect by buyer type. Thus, whatever the level of deforestation does not explain, is a fixed characteristic of the farm that differentially affects its ability to sell to a type of buyer,

$$\text{FE}_o^{q*}(s) \equiv (\tilde{c}_o^q(s)^{-\zeta} \tilde{\kappa}_o^q(s)^{-\zeta})^* = \frac{\text{Intensity of use}^q(s)}{\sum_{s' \in o} \text{Intensity of use}^q(s')}.$$

Therefore, using the deforestation penalty as with the assumed functional form and the estimated parameters α^{q*} ,

$$(\kappa_o^q(s)^{-\zeta})^* = \exp(-\zeta\alpha^{q*} \cdot \text{lhs}(l^{\text{defor}}(s))).$$

The normalized penalty equals, by definition,

$$(\tilde{\kappa}_o^q(s)^{-\zeta})^* = \frac{(\kappa_o^q(s)^{-\zeta})^*}{((\kappa_o^q)^{-\zeta})^*} = \frac{\exp(-\zeta\alpha^{q*} \cdot \text{lhs}(l^{\text{defor}}(s)))}{\sum_{s' \in \mathcal{S}_o} (\tilde{c}_o^q(s')^{-\zeta} \tilde{\kappa}_o^q(s')^{-\zeta})^*}.$$

And the normalized unit costs equal

$$(\tilde{c}_o^q(s)^{-\zeta})^* = \frac{(\tilde{c}_o^q(s)^{-\zeta} \tilde{\kappa}_o^q(s)^{-\zeta})^*}{(\tilde{\kappa}_o^q(s)^{-\zeta})^*}.$$

In this case, costs of producing for a type q will be estimated to be infinite, or $(\tilde{c}_o(s)^{-\zeta})^* = 0$, if s does not sell to type q buyers. This has very important implications for counterfactual scenarios, as no extensive-margin switching of buyers will be allowed.

Step 3B: Estimate farm-level fixed effects with equal unit costs across buyers

An alternative strategy is to attribute all differences in a farm's ability to sell to different types of buyers to its deforestation penalty. In this case, we force unit costs to be the same across all buyer types. For that, we use a "net-of-deforestation-penalty" intensity of use measure to get the unit costs,

$$\text{Net-of-defor Intensity of use}^q(s) = \frac{\text{Intensity of use}^q(s)}{(\tilde{\kappa}_o^q(s)^{-\zeta})^*}.$$

And using this, we can aggregate to get a total (across all types of buyers) intensity of use of farm s net of deforestation. This can be thought of as the intensity of use that s would get if it did not deforest and everything else stayed constant

$$\text{Net-of-defor Intensity of use}(s) = \sum_q \frac{\text{Intensity of use}^q(s)}{(\tilde{\kappa}_o^q(s) - \zeta)^*}.$$

Then we use the net-of-deforestation-penalty intensity of use measures to estimate farm-level unit costs,

$$(\tilde{c}_o(s) - \zeta)^* = \frac{\text{Net-of-defor Intensity of use}(s)}{\sum_{s' \in \mathcal{S}_o} \text{Net-of-defor Intensity of use}(s')}$$

If the estimated deforestation penalty is assumed to be $(\kappa_o^q(s) - \zeta)^* = \exp(-\zeta \alpha^{q*} \cdot \text{lhs}(l^{\text{defor}}(s)))$, then the normalized deforestation penalty, $(\tilde{\kappa}_o^q(s) - \zeta)^* = (\kappa_o^q(s) - \zeta)^* \cdot \exp(-\zeta \alpha^{q*} \cdot \text{lhs}(l^{\text{defor}}(s)))$. Thus,

$$(\tilde{c}_o(s) - \zeta)^* = \sum_q \frac{\text{Intensity of use}^q(s)}{\exp(-\zeta \alpha^{q*} \cdot \text{lhs}(l^{\text{defor}}(s)))} \bigg/ \sum_q \sum_{s' \in \mathcal{O}} \frac{\text{Intensity of use}^q(s')}{\exp(-\zeta \alpha^{q*} \cdot \text{lhs}(l^{\text{defor}}(s')))}.$$

To calculate the normalized deforestation penalty, notice

$$(\tilde{\kappa}_o^q(s) - \zeta)^* = \frac{(\kappa_o^q(s) - \zeta)^*}{((\kappa_o^q(s) - \zeta)^*)} = \frac{(\kappa_o^q(s) - \zeta)^*}{\sum_{s' \in \mathcal{O}} (\tilde{c}_o(s') - \zeta)^* (\kappa_o^q(s') - \zeta)^*}.$$

Thus the farm-buyer type fixed effect needed to get the farm-buyer type level expected trade shares equal

$$FE_o^{q*}(s) \equiv (\tilde{c}_o(s) - \zeta)^* (\tilde{\kappa}_o^q(s) - \zeta)^* = \frac{(\tilde{c}_o(s) - \zeta)^* \cdot \exp(-\zeta \alpha^{q*} \cdot \text{lhs}(l^{\text{defor}}(s)))}{\sum_{s' \in \mathcal{O}} (\tilde{c}_o(s') - \zeta)^* \cdot \exp(-\zeta \alpha^{q*} \cdot \text{lhs}(l^{\text{defor}}(s')))}.$$

In this case, all farms are able to sell to different types of buyers, and all that influences their relative ability to sell to one versus the other (holding geographical distances fixed) is their level of deforestation.

Step 3C: Estimate farm-buyer type level fixed effects with productivity for sale to slaughterhouses and farms

In this case, there are two marginal costs: the marginal cost of producing cows for slaughter, $c^{\text{sh}}(s)$, and the marginal cost of producing cows for rearing/fattening, $c^{\text{farm}}(s)$, i.e. to sell to other farms. Then we estimate the fixed effect relating to sales to farms as in the no-switching case,

$$FE_o^{\text{farm}*}(s) \equiv (\tilde{c}_o^{\text{farm}}(s) - \zeta \tilde{\kappa}^{\text{farm}}(s) - \zeta)^* = \frac{\text{Intensity of use}^{\text{farm}}(s)}{\sum_{s' \in \mathcal{O}} \text{Intensity of use}^{\text{farm}}(s')}$$

And for the sales to slaughterhouses, the marginal cost is multiplied by the deforestation penalty, $(\kappa_o^q(s)^{-\zeta})^* = \Delta_o^q \cdot \exp(-\zeta \alpha^{q*} \cdot \text{ihf}(l^{\text{defor}}(s)))$, estimated up to an unknown origin-buyer type level constant Δ_o^q , so that the normalized deforestation penalty equals $(\tilde{\kappa}_o^q(s)^{-\zeta})^* = \tilde{\Delta}_o^q \exp(-\zeta \alpha^{q*} \cdot \text{ihf}(l^{\text{defor}}(s)))$. Thus the normalised costs of production of cows for slaughter equals

$$(\tilde{c}_o^{\text{sh}}(s)^{-\zeta})^* = \sum_{q \text{ is sh}} \frac{\text{Intensity of use}^q(s)}{\exp(-\zeta \alpha^{q*} \cdot \text{ihf}(l^{\text{defor}}(s)))} \Bigg/ \sum_{q \text{ is sh}} \sum_{s' \in o} \frac{\text{Intensity of use}^q(s')}{\exp(-\zeta \alpha^{q*} \cdot \text{ihf}(l^{\text{defor}}(s')))}.$$

To calculate the normalized deforestation penalty, notice

$$(\tilde{\kappa}_o^q(s)^{-\zeta})^* = \frac{(\kappa_o^q(s)^{-\zeta})^*}{((\kappa_o^q)^{-\zeta})^*} = \frac{(\kappa_o^q(s)^{-\zeta})^*}{\sum_{s' \in o} (\tilde{c}_o^{\text{sh}}(s')^{-\zeta})^* (\kappa_o^q(s')^{-\zeta})^*}.$$

Thus the farm-slaughterhouse-type fixed effect needed to get the farm-slaughterhouse-type level expected trade shares equal, for $q \in \{\text{non ZDC s.h.}, \text{ZDC s.h.}\}$,

$$\text{FE}_o^{q*}(s) \equiv (\tilde{c}_o^{\text{sh}}(s)^{-\zeta})^* (\tilde{\kappa}_o^q(s)^{-\zeta})^* = \frac{(\tilde{c}_o^{\text{sh}}(s)^{-\zeta})^* \cdot \exp(-\zeta \alpha^{q*} \cdot \text{ihf}(l^{\text{defor}}(s)))}{\sum_{s' \in o} (\tilde{c}_o^{\text{sh}}(s')^{-\zeta})^* \cdot \exp(-\zeta \alpha^{q*} \cdot \text{ihf}(l^{\text{defor}}(s')))}.$$

Step 4: Compute farm-level values equilibrium rental rates in the initial state.

Having the baseline values of aggregate land rents, we can plug them into the RHS (demand side) of the disaggregated equation, so that we solve for $r_o(s)$,

$$\begin{aligned} r_o(s) = \frac{\zeta}{\zeta + 1} \frac{1}{l_o(s)} & \left((1 - \beta) \sum_d \pi_{od}^{\text{farm}*}(s, -) R_d \right. \\ & + \beta \sum_d \pi_{od}^{\text{nonZDCs.h.*}}(s, -) X_d^{\text{nonZDCs.h.}} \\ & \left. + \beta \sum_d \pi_{od}^{\text{ZDCs.h.*}}(s, -) X_d^{\text{ZDCs.h.}} \right) \end{aligned}$$

which can be re-written as

$$\begin{aligned} r_o(s) = \frac{\zeta}{\zeta + 1} \frac{1}{l_o(s)} & \left((1 - \beta) \cdot \text{FE}_o^{\text{farm}*}(s) \cdot \sum_d \pi_{od}^{\text{farm}*}(\bullet, -) R_d \right. \\ & + \beta \cdot \text{FE}_o^{\text{nonZDCs.h.*}}(s) \cdot Y_o^{\text{nonZDCs.h.*}} \\ & \left. + \beta \cdot \text{FE}_o^{\text{ZDCs.h.*}}(s) \cdot Y_o^{\text{ZDCs.h.*}} \right) \end{aligned}$$

where the expected purchases of cows from in origin o by ZDC slaughterhouses, Y_o^{q*} , equal

$$Y_o^{q*} \equiv \sum_d \pi_{od}^{q*}(\bullet, -) X_d^{\text{ZDCs.h.}}$$

Step 5: Estimate marginal deforestation costs.

If the optimal level of deforestation is a local maximum, we can rely on the first order conditions of the deforestation problem to notice that, if the deforestation decision problem had an interior solution, it satisfies:

$$\frac{d\phi(l_o^{\text{defor}}(s))}{dl_o^{\text{defor}}(s)} = \frac{\beta\zeta}{(1-\rho)(1+\beta\zeta)} \left(r_o(s) - \frac{\alpha\zeta}{1+\zeta} \frac{1}{\sqrt{l_o^{\text{defor}}(s)^2 + 1}} Y^{\text{ZDC}*}(s) \right)$$

Since everything in the left hand side we either already have from the calibration, is data, or we have assumed outside of the model, we know the marginal costs of deforestation $\frac{d\phi(l_o^{\text{defor}}(s))}{dl_o^{\text{defor}}(s)}$ for each farm at the baseline equilibrium.

E Counterfactual estimation

E.1 No ZDC policy counterfactuals, i.e. $\alpha = 0$

Step 1: Derive equation from optimization problem

Assuming locally linear deforestation costs, and denoting the new equilibrium prices and quantities with tildes, and the changes with hats, if both the old and new equilibria have interior solutions for the optimal amount of deforestation,

$$\begin{aligned} \frac{d\phi(l_o^{\text{defor}}(s))}{dl_o^{\text{defor}}(s)} &= \frac{\beta\zeta}{(1-\rho)(1+\beta\zeta)} \left(r_o(s) - \frac{\alpha\zeta}{1+\zeta} \frac{1}{\sqrt{l_o^{\text{defor}}(s)^2 + 1}} \mathbb{E}Y_o^{\text{ZDC}}(s) \right) \\ &\approx \frac{d\phi(l_o^{\text{defor}}(s)')}{dl_o^{\text{defor}}(s)} = \frac{\beta\zeta}{(1-\rho)(1+\beta\zeta)} \left(r_o(s)' \right) \end{aligned} \quad (9)$$

Therefore, the change in rental rates $\hat{r}_o(s)$ is given by:

$$\hat{r}_o(s) \approx 1 - \frac{1}{r_o(s)} \frac{\alpha\zeta}{1+\zeta} \frac{1}{\sqrt{l_o^{\text{defor}}(s)^2 + 1}} \mathbb{E}Y_o^{\text{ZDC}}(s)$$

Step 2: Get policy shifters

First define policy shifters, both farm specific and aggregated at the region level, as follows:

$$\widehat{\delta}_o^q(s) \equiv \widehat{r}_o(s)^{-\beta\zeta} \widehat{\kappa}_o^q(s)^{-\zeta}, \quad \text{and}$$

$$\widehat{\delta}_o^q = \sum_{o \in \mathcal{S}_o} \text{FE}_o^{q*}(s) \widehat{\delta}_o^q(s) = \sum_{o \in \mathcal{S}_o} (\widehat{c}_o^q(s)^{-\zeta})^* (\widehat{\kappa}_o^q(s)^{-\zeta})^* \widehat{r}_o(s)^{-\beta\zeta} \widehat{\kappa}_o^q(s)^{-\zeta}.$$

The change in rental rates, $\widehat{r}_o(s)$, comes from the previous step, and the change in deforestation penalty $\widehat{\kappa}_o^q(s)$ can be directly computed without needing to know the counterfactual deforestation rate. In this case, the deforestation penalty for farms and non ZDC slaughterhouses stays null, so there is no change, i.e. $\widehat{\kappa}_o^q(s) = 1$ if $q \in \{\text{farm, non ZDC s.h.}\}$. For ZDC slaughterhouses, the penalty goes away, so all we need is to know the current rate of deforestation and not the counterfactual one to compute the change,

$$\widehat{\kappa}_o^{\text{ZDC}}(s) = (\kappa_o^{\text{ZDC}}(s))^{-1} = \exp(-\alpha \cdot \text{lhs}(l_o^{\text{defor}}(s))).$$

Step 3: Get aggregate sourcing probabilities

Notice that, since

$$\mathbb{E}\pi_{od}^q(\bullet, -)' = \mathbb{E}\pi_{od}^q(\bullet, -) \frac{\widehat{\delta}_o^q(\widehat{A}_o^{\text{farm}})^{1-\beta}}{\widehat{A}_d^q}$$

And therefore the changes in market access for cows as intermediate inputs solve the system of equations

$$\widehat{A}_d^{\text{farm}} = \sum_o \mathbb{E}\pi_{od}^{\text{farm}}(\bullet, -) \widehat{\delta}_o^{\text{farm}}(\widehat{A}_o^{\text{farm}})^{1-\beta}$$

where $\mathbb{E}\pi_{od}^q(\bullet, -)$ are the baseline sourcing intensities as predicted by the gravity equations. Then the market access changes for slaughterhouses of type q can be computed as

$$\widehat{A}_d^q = \sum_o \mathbb{E}\pi_{od}^q(\bullet, -) \widehat{\delta}_o^q(\widehat{A}_o^{\text{farm}})^{1-\beta}.$$

Having solved for changes in market access, we have everything we need to back out the counterfactual sourcing probabilities.

Step 4: Back out new landlord rents $r_o(s)l_o(s)$

Using the farm-level shifters, we can compute the counterfactual farm-level fixed effects,

$$\text{FE}_o^q(s)' \equiv (\widehat{c}_o^q(s)' \widehat{\kappa}_o^q(s)')^{-\zeta} = \frac{\text{FE}_o^q(s) \widehat{\delta}_o^q(s)}{\sum_{s' \in \mathcal{S}_o} \text{FE}_o^q(s') \widehat{\delta}_o^q(s')}.$$

	Dep. Var: Share sold to each buyer type				
	(1)	(2)	(3)	(4)	(5)
ihs(Defor. since 2009) x To Farm	0.64*** (0.04)	-0.16*** (0.04)	-0.20*** (0.04)	0.32* (0.13)	0.27* (0.14)
ihs(Defor. since 2009) x To non ZDC sh	0.21** (0.07)	-0.59*** (0.08)	-0.61*** (0.08)	0.22 (0.38)	0.31 (0.38)
ihs(Defor. since 2009) x To ZDC sh	0.31** (0.10)	-0.90*** (0.14)	-0.90*** (0.13)	-1.57** (0.50)	-1.57** (0.51)
ihs(Defor. 2003-2009) x To Farm		0.49*** (0.02)	0.46*** (0.03)		
ihs(Defor. 2003-2009) x To non ZDC sh		0.48*** (0.02)	0.43*** (0.05)		
ihs(Defor. 2003-2009) x To ZDC sh		0.77*** (0.02)	0.51*** (0.09)		
log(Prop. Area) x To Farm		0.20*** (0.03)	0.09** (0.03)		
log(Prop. Area) x To non ZDC sh		0.12* (0.06)	-0.03 (0.06)		
log(Prop. Area) x To ZDC sh		-0.00 (0.06)	-0.11 (0.06)		
Origin cell x Buyer Type x Year FE	X	X	X	X	X
Land Use Controls x Buyer Type FE			X		X
Seller x Buyer Type FE				X	X
Num. Obs	513610	513610	423651	291631	245158
Num. Origin cell x Buyer Type x Year	12524	12524	12209	12524	12209
Num. Seller x Buyer Type				62035	52801

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 3: Fixed Effects Poisson Regression, D=20 km

Making use of these and the counterfactual aggregate sourcing probabilities, $\mathbb{E}\pi_{od}^q(\bullet, -)'$, we can repeat the procedure used in the model inversion to back out $r_o(s)'l_o(s)'$, and since we had the counterfactual rental rates $r_o(s)'$, from step 1, this means we can also back out the amount of land $l_o(s)'$, and counterfactual level of deforestation!

F Additional Results

	Dep. Var: Any sales to Buyer Type				
	(1)	(2)	(3)	(4)	(5)
ihs(Defor. since 2009) x To Farm	0.14*** (0.01)	-0.06*** (0.01)	-0.07*** (0.01)	0.19*** (0.06)	0.14* (0.06)
ihs(Defor. since 2009) x To non ZDC sh	0.17*** (0.05)	-0.48*** (0.06)	-0.49*** (0.06)	0.11 (0.25)	0.15 (0.25)
ihs(Defor. since 2009) x To ZDC sh	0.22*** (0.06)	-0.78*** (0.08)	-0.75*** (0.08)	-0.83** (0.26)	-0.79** (0.27)
ihs(Defor. 2003-2009) x To Farm		0.08*** (0.00)	-0.02 (0.01)		
ihs(Defor. 2003-2009) x To non ZDC sh		0.32*** (0.01)	0.22*** (0.03)		
ihs(Defor. 2003-2009) x To ZDC sh		0.56*** (0.02)	0.33*** (0.06)		
log(Prop. Area) x To Farm		0.08*** (0.01)	0.04*** (0.01)		
log(Prop. Area) x To non ZDC sh		0.14*** (0.04)	0.04 (0.04)		
log(Prop. Area) x To ZDC sh		-0.03 (0.04)	-0.09* (0.04)		
Origin cell x Buyer Type x Year FE	X	X	X	X	X
Land Use Controls x Buyer Type FE			X		X
Seller x Buyer Type FE				X	X
Num. Obs	513610	513610	423651	291631	245158
Num. Origin cell x Buyer Type x Year	12524	12524	12209	12524	12209
Num. Seller x Buyer Type				62035	52801

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 4: Fixed Effects Poisson Regression, D=20 km