

**STA 2002, Summer 2024**  
**Probability and Statistics II**  
**Order Statistics and QQ plot**

# Why study order statistics

- **Issue 1:** Is there a method for us to check whether our data is statistically similar to a normal distribution?
  - (Yes, we can use QQ plots.)
- **Issue 2:** If the data is not statistically similar to normal distribution, is there a way to do hypothesis testing?
  - (Yes, **non-parametric / distribution-free** confidence interval and hypothesis testing.)

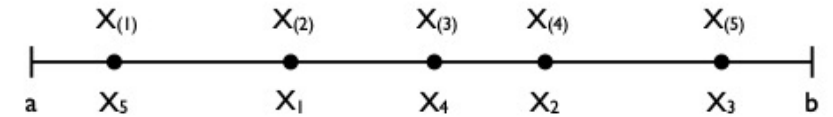
**Order Statistics** serves as an essential tool

# Order Statistics

Let  $X_1, X_2, \dots, X_n$  be i.i.d. random samples from a common distribution  $f$ .

$$X_1, X_2, \dots, X_n \sim f$$

## Definition



Denote the *ordered* sample values  $X_{(1)} < X_{(2)} < \dots < X_{(n)}$  as the **order statistic**

For each  $k$ , the  $k$ -th order statistic is

$$X_{(k)} = k\text{-th smallest of } X_1, X_2, \dots, X_n$$

Example

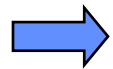
$$\begin{aligned} X_{(1)} &= \min\{X_1, X_2, \dots, X_n\}, \\ X_{(2)} &= \text{second smallest of } X_1, \dots, X_n, \\ X_{(n)} &= \max\{X_1, X_2, \dots, X_n\}. \end{aligned}$$

# Density of Order Statistics

Let  $X_1, X_2, \dots, X_n$  be i.i.d. random samples from a common distribution with CDF  $F(x)$  and PDF  $f(x)$ . What is the CDF and PDF of any order statistic  $X_{(k)}$ ?

A special case:  $X_{(n)} = \max\{X_1, X_2, \dots, X_n\}$ .

$$\begin{aligned} P(X_{(n)} \leq x) &= P(\max\{X_1, \dots, X_n\} \leq x) = P(X_1 \leq x, X_2 \leq x, \dots, X_n \leq x) \\ &= P(X_1 \leq x) \times P(X_2 \leq x) \times \dots \times P(X_n \leq x) \\ &= F(x)^n \end{aligned}$$



$$\text{CDF: } F_{X_{(n)}}(x) = F(x)^n$$

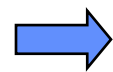
$$\text{PDF: } f_{X_{(n)}}(x) = nF(x)^{n-1}f(x)$$

# Density of Order Statistics

Let  $X_1, X_2, \dots, X_n$  be i.i.d. random samples from a common distribution with CDF  $F(x)$  and PDF  $f(x)$ . What is the CDF and PDF of any order statistic  $X_{(k)}$ ?

A special case:  $X_{(1)} = \min\{X_1, X_2, \dots, X_n\}$ ,

$$\begin{aligned} P(X_{(1)} > x) &= P(\min\{X_1, \dots, X_n\} > x) = P(X_1 > x, X_2 > x, \dots, X_n > x) \\ &= P(X_1 > x) \times P(X_2 > x) \times \dots \times P(X_n > x) \\ &= (1 - F(x))^n \end{aligned}$$



$$\text{CDF: } F_{X_{(1)}}(x) = 1 - (1 - F(x))^n$$

$$\text{PDF: } f_{X_{(1)}}(x) = n(1 - F(x))^{n-1} f(x)$$

# Density of Order Statistics

Let  $X_1, X_2, \dots, X_n$  be i.i.d. random samples from a common distribution with CDF  $F(x)$  and PDF  $f(x)$ . What is the CDF and PDF of any order statistic  $X_{(k)}$ ?

What is the CDF and PDF of  $X_{(k)}$  for any  $k$ ?

We define  $W \sim \text{Bin}(n, F(x))$  as the number of  $X_i$  that is smaller than  $x$ .

$$\begin{aligned} F_{X_{(k)}}(x) &= P(X_{(k)} \leq x) \\ &= P(W \geq k) \\ &= \sum_{l=k}^n P(W = l) \\ &= \sum_{l=k}^n \binom{n}{l} F(x)^l (1 - F(x))^{n-l} \\ &= \sum_{l=k}^{n-1} \binom{n}{l} F(x)^l (1 - F(x))^{n-l} + F(x)^n. \end{aligned}$$

# Density of Order Statistics

We define  $W \sim \text{Bin}(n, F(x))$  as the number of  $X_i$  that is smaller than  $x$ .

$$F_{X_{(k)}}(x) = \sum_{l=k}^n \binom{n}{l} F(x)^l (1 - F(x))^{n-l} = \sum_{l=k}^{n-1} \binom{n}{l} F(x)^l (1 - F(x))^{n-l} + F(x)^n.$$

$$\begin{aligned} f_{X_{(k)}}(x) &= \sum_{l=k}^{n-1} \binom{n}{l} l [F(x)]^{l-1} f(x) [1 - F(x)]^{n-l} + \sum_{l=k}^{n-1} \binom{n}{l} (n-l) [F(x)]^l (-f(x)) [1 - F(x)]^{n-l-1} \\ &\quad + n F(x)^{n-1} f(x) \\ &= \sum_{l=k}^n \binom{n}{l} l [F(x)]^{l-1} f(x) [1 - F(x)]^{n-l} + \sum_{l=k}^{n-1} \binom{n}{l} (n-l) [F(x)]^l (-f(x)) [1 - F(x)]^{n-l-1} \\ &= \sum_{l=k}^n \binom{n}{l} l [F(x)]^{l-1} f(x) [1 - F(x)]^{n-l} + \sum_{l'=k+1}^n \binom{n}{l'-1} (n-l'+1) [F(x)]^{l'-1} (-f(x)) [1 - F(x)]^{n-l'} \\ &= \sum_{l=k}^n \frac{n!}{(l-1)!(n-l)!} [F(x)]^{l-1} f(x) [1 - F(x)]^{n-l} + \sum_{l'=k+1}^n \frac{n!}{(l'-1)!(n-l')!} [F(x)]^{l'-1} (-f(x)) [1 - F(x)]^{n-l'} \\ &= \frac{n!}{(k-1)!1!(n-k)!} [F(x)]^{k-1} f(x) [1 - F(x)]^{n-k}. \end{aligned}$$

# Density of Order Statistics

Let  $X_1, X_2, \dots, X_n$  be i.i.d. random samples from a common distribution with CDF  $F(x)$  and PDF  $f(x)$ . What is the CDF and PDF of any order statistic  $X_{(k)}$ ?

**Theorem 1** Suppose that  $X_1, \dots, X_n$  are i.i.d. continuous random variables with common pdf  $f(x)$  and cdf  $F(x)$ . For  $k = 1, \dots, n$ , denote the cdf and pdf of the  $k$ th order statistic  $X_{(k)}$  to be respectively  $F_{X_{(k)}}$  and  $f_{X_{(k)}}$ . They can be written as

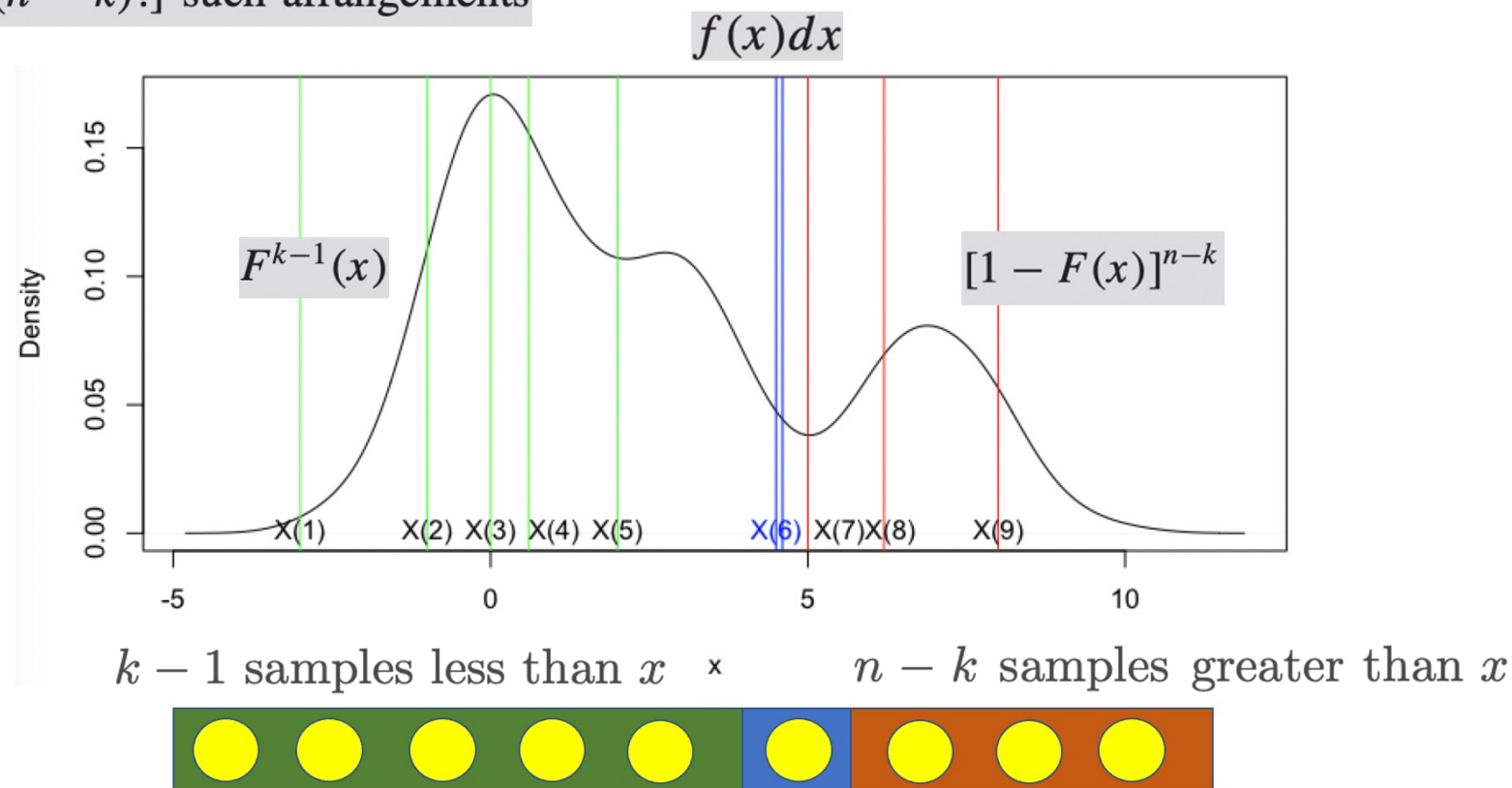
$$F_{X_{(k)}}(x) = \sum_{l=k}^n \binom{n}{l} F(x)^l (1 - F(x))^{n-l},$$
$$f_{X_{(k)}}(x) = \frac{n!}{(k-1)!1!(n-k)!} [F(x)]^{k-1} f(x) [1 - F(x)]^{n-k}.$$



# Density of Order Statistics

$$f_{X_{(k)}}(x) = \frac{n!}{(k-1)!1!(n-k)!} [F(x)]^{k-1} f(x) [1 - F(x)]^{n-k}.$$

$n! / [(k-1)!1!(n-k)!]$  such arrangements



# Example – Uniform distribution

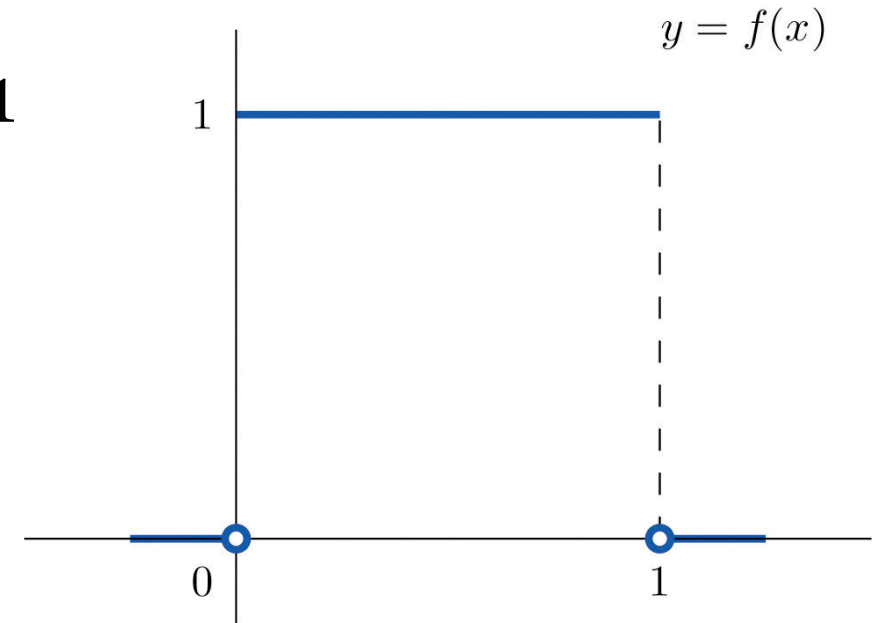
Let  $X_1, X_2, \dots, X_n$  be i.i.d. samples from the uniform distribution on  $[0,1]$ .

$$X_1, X_2, \dots, X_n \sim \text{Uniform}[0,1]$$

$$F(x) = x, \quad f(x) = 1, \forall 0 \leq x \leq 1$$

$$f_{X_{(k)}}(x) = \frac{n!}{(k-1)!1!(n-k)!} [F(x)]^{k-1} f(x) [1 - F(x)]^{n-k}.$$

$$f_{X_{(k)}}(x) = \frac{n!}{(k-1)!1!(n-k)!} x^{k-1} (1-x)^{n-k}$$



# Example – Uniform distribution

$$f_{X_{(k)}}(x) = \frac{n!}{(k-1)! 1! (n-k)!} x^{k-1} (1-x)^{n-k}$$

## Beta distribution

$$X \sim \text{Beta}(\alpha, \beta) \quad f(x) = \frac{x^{\alpha-1} (1-x)^{\beta-1}}{B(\alpha, \beta)}, 0 \leq x \leq 1$$

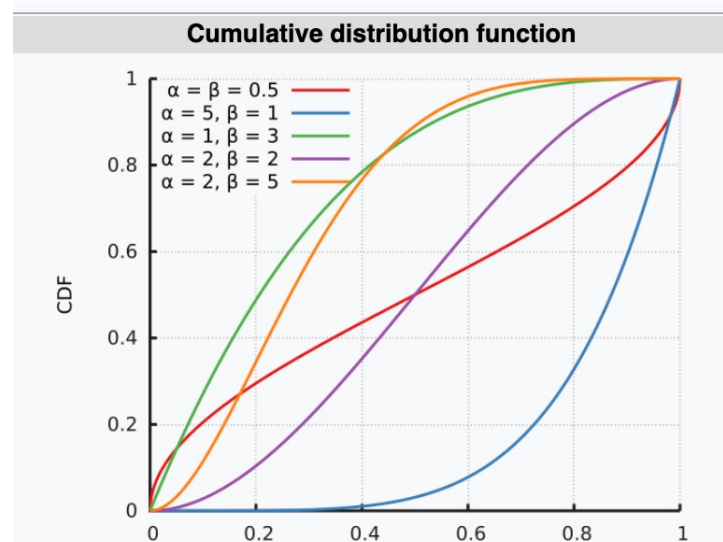
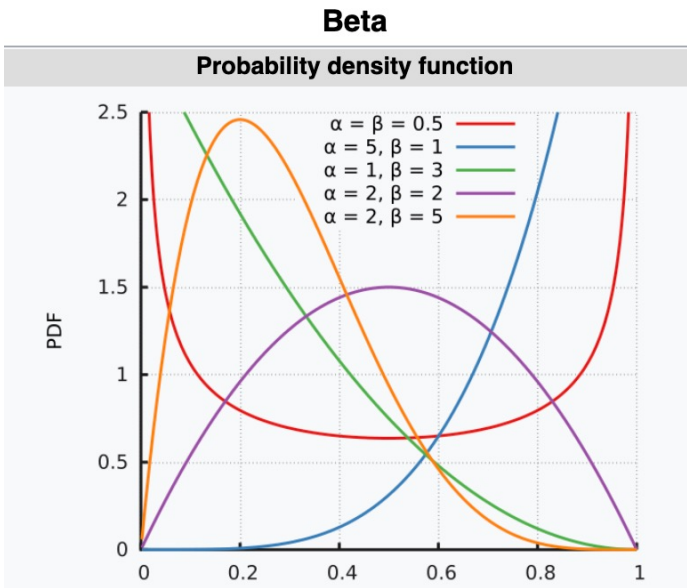
$$B(\alpha, \beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha + \beta)}$$

Gamma function:  $\Gamma(t) = \int_0^\infty y^{t-1} e^{-y} dy,$

$$\Gamma(n) = (n-1)!$$



$$X_{(k)} \sim \text{Beta}(k, n - k + 1)$$



Credit: Wiki

# Example – Uniform distribution

$$f_{X_{(k)}}(x) = \frac{n!}{(k-1)! 1! (n-k)!} x^{k-1} (1-x)^{n-k} \quad X_{(k)} \sim \text{Beta}(k, n-k+1)$$

**The mean and variance of  $X_{(k)}$**

$$E(X_{(k)}) = \int_0^1 x \frac{n!}{(k-1)! 1! (n-k)!} x^{k-1} (1-x)^{n-k} dx = \int_0^1 \frac{n!}{(k-1)! 1! (n-k)!} x^{k+1-1} (1-x)^{n+1-(k+1)} dx$$

$$= \frac{n!}{(k-1)! 1! (n-k)!} \times \frac{k! 1! (n+1-(k+1))!}{(n+1)!} = \boxed{\frac{k}{n+1}}$$

$$E(X_{(k)}^2) = \int_0^1 x^2 \frac{n!}{(k-1)! 1! (n-k)!} x^{k-1} (1-x)^{n-k} dx = \int_0^1 \frac{n!}{(k-1)! 1! (n-k)!} x^{k+2-1} (1-x)^{(n+2)-(k+2)} dx$$

$$= \frac{n!}{(k-1)! 1! (n-k)!} \times \frac{(k+1)! 1! (n+2-(k+2))!}{(n+2)!} = \frac{k(k+1)}{(n+1)(n+2)}$$

➡ 
$$\text{Var}(X_{(k)}^2) = \frac{k(k+1)}{(n+1)(n+2)} - \left(\frac{k}{n+1}\right)^2 = \boxed{\frac{k(n+1-k)}{(n+1)^2(n+2)}}$$

# Example – Uniform distribution

$$f_{X_{(k)}}(x) = \frac{n!}{(k-1)! 1! (n-k)!} x^{k-1} (1-x)^{n-k} \quad X_{(k)} \sim \text{Beta}(k, n-k+1)$$

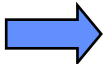
**The mean and variance of  $X_{(k)}$**

Beta distribution

$$X \sim \text{Beta}(\alpha, \beta) \quad f(x) = \frac{x^{\alpha-1} (1-x)^{\beta-1}}{B(\alpha, \beta)}, 0 \leq x \leq 1 \quad B(\alpha, \beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}$$

$$E(X) = \frac{\alpha}{\alpha+\beta} \quad \text{Var}(X) = \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$$

$$\alpha = k, \beta = n - k + 1$$


$$E(X_{(k)}) = \frac{\alpha}{\alpha+\beta} = \frac{k}{n+1}$$

$$\text{Var}(X_{(k)}) = \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)} = \frac{k(n+1-k)}{(n+1)^2(n+2)}$$

# Expectation of Order Statistics

- For **uniform** distribution

$$X_1, X_2, \dots, X_n \sim \text{Uniform}[0,1]$$

- The expectation

$$E(X_{(k)}) = \frac{\alpha}{\alpha + \beta} = \frac{k}{n + 1}$$

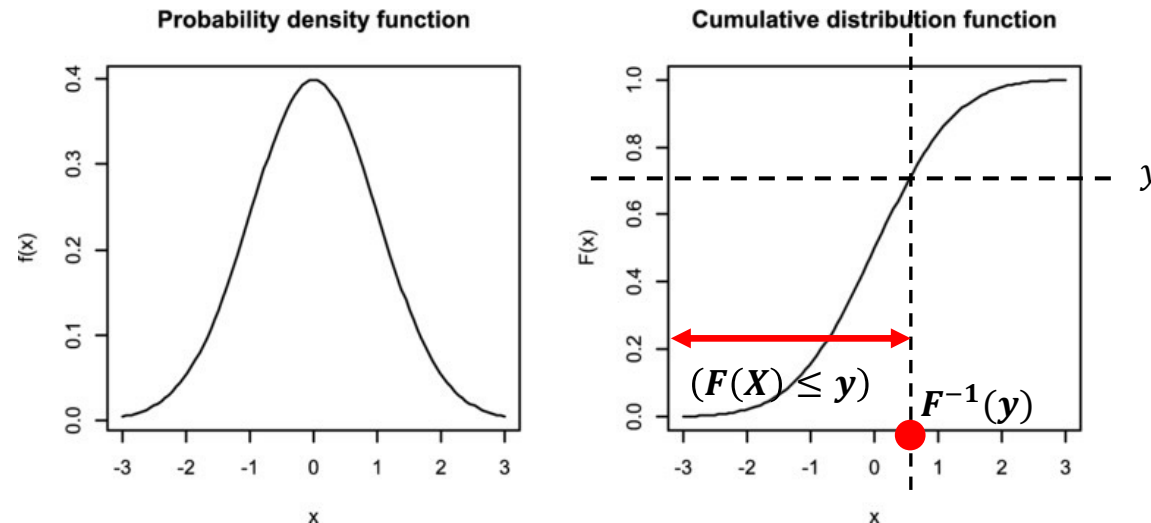
- How about other distributions (non-uniform)? Is this true for any other distributions? **No.**

# Expectation of $F(X_{(k)})$

- Let  $X$  be a random variable with PDF  $f(x)$  and strictly increasing CDF  $F(x)$ .
- Let  $Y = F(X)$  be a new random variable.  $0 \leq Y \leq 1$ .
- The CDF of  $Y$ :

$$P(Y \leq y) = P(F(X) \leq y) = P(X \leq F^{-1}(y)) = F(F^{-1}(y)) = y$$

- Therefore,  $Y \sim \text{Uniform}[0,1]$



# Expectation of $F(X_{(k)})$

Let  $X_1, X_2, \dots, X_n$  be i.i.d. random samples from a common distribution with PDF  $f(x)$  and strictly increasing CDF  $F(x)$ .

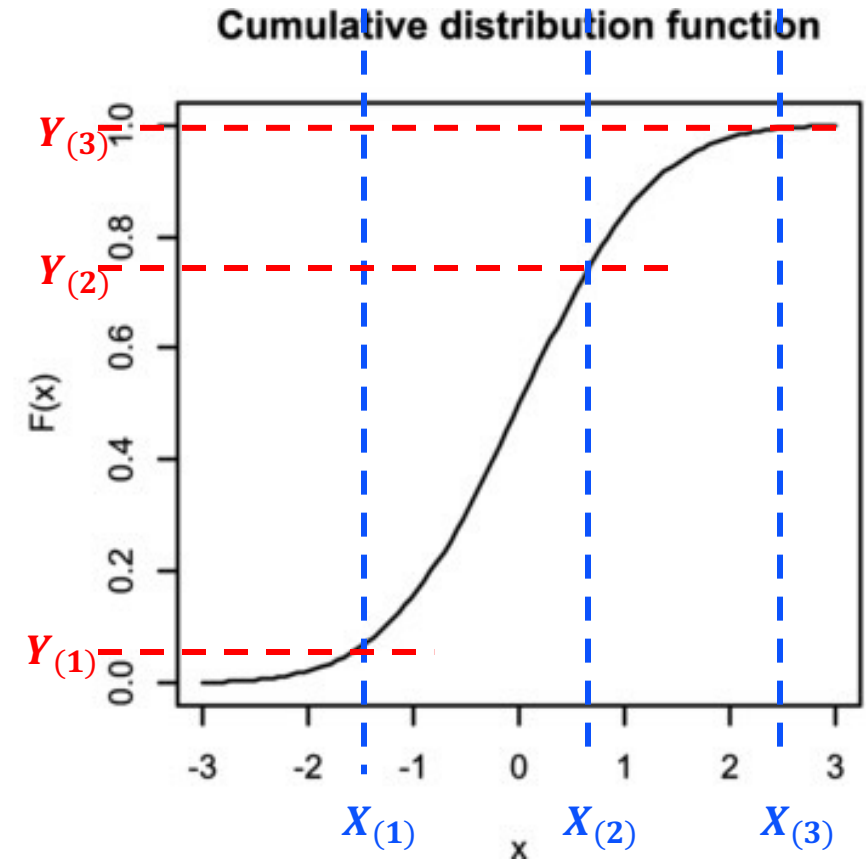
$$X_1, X_2, \dots, X_n \sim f$$

Let  $Y_1 = F(X_1), Y_2 = F(X_2), \dots, Y_n = F(X_n)$

$$Y_1, Y_2, \dots, Y_n \sim \text{Uniform}[0,1]$$

$Y_{(k)} = F(X_{(k)})$  is the k-th order statistic for uniform distributions:

$$E(Y_{(k)}) = E(F(X_{(k)})) = \frac{k}{n+1}$$





# Q-Q plot

# Theoretical vs Sample Quantiles

- The  $p$ -th **theoretical quantile** of the distribution  $F$  is  $\pi_p$ :

$$F(\pi_p) = p$$

$$\pi_p = F^{-1}(p)$$

- For example, when  $p = 0.5$ ,  $\pi_{0.5}$  is called the median of the distribution  $F$ .
- For standard normal distribution, we have  $\pi_p = Z_{1-p}$

How to estimate  $\pi_p$  from samples  $X_1, X_2, \dots, X_n \sim F$ ?

$$E\left(F(X_{(k)})\right) = \frac{k}{n+1}$$

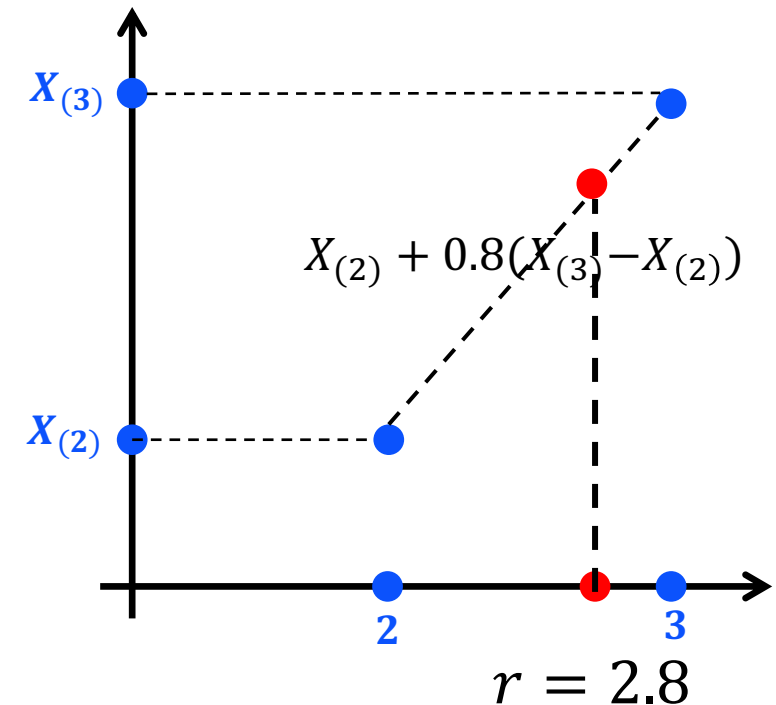
- For  $p = \frac{k}{n+1}$ ,  $X_{(k)}$  is an estimate for  $\pi_p$ .
- $X_{(k)}$ ,  $k = (n+1)p$  is the  $p$ -th **sample quantile**  $\hat{\pi}_p$ .

# Theoretical vs Sample Quantiles

## Calculation of the p-th sample quantile

- Order the samples  $X_{(1)} < X_{(2)} < \dots < X_{(n)}$
- $r = (n + 1)p$
- If  $r$  is an integer, the sample quantile is  $\hat{\pi}_p = X_{(r)}$
- If  $r$  is not an integer,  $r = \lfloor r \rfloor + (r - \lfloor r \rfloor)$ , the sample quantile is

$$\hat{\pi}_p = X_{(\lfloor r \rfloor)} + (r - \lfloor r \rfloor)(X_{(\lfloor r \rfloor + 1)} - X_{(\lfloor r \rfloor)})$$

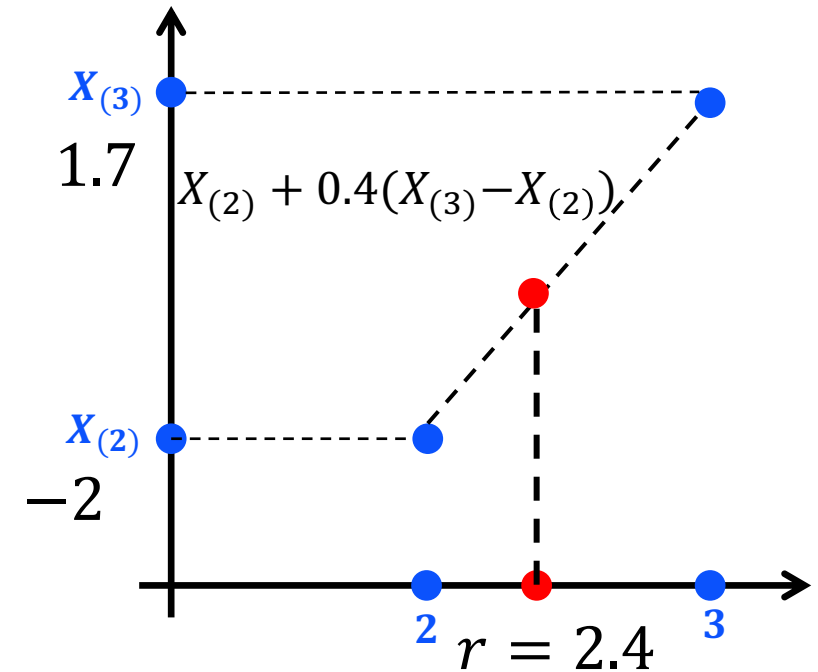


# Example

$i$	$x_{(i)}$	$i/(n+1)$
1	-3.9	1/6
2	-2.0	2/6
3	1.7	3/6
4	7.3	4/6
5	11.7	5/6

- $p = \frac{1}{2}$ , the sample median is  $X_{(3)} = 1.7$
- $p = 0.4 \Rightarrow r = (n+1)p = 2.4$

$$\hat{\pi}_{0.4} = X_{(2)} + 0.4(X_{(3)} - X_{(2)}) = -0.52$$



# Quantile-Quantile (Q-Q) plot

Given data  $x_1, x_2, \dots, x_n$ , we suspect that they are coming from **a common distribution  $F$** , say normal distribution or exponential distribution or gamma distribution. We can then compute the **theoretical quantiles**

$\pi_{i/(n+1)} = F^{-1}\left(\frac{i}{n+1}\right)$  for  $i = 1, 2, \dots, n$ . If the data is indeed statistically similar to the distribution, then we expect that

$$\underbrace{\hat{\pi}_{i/(n+1)} = x_{(i)}}_{\text{sample quantiles}} \approx \underbrace{\pi_{i/(n+1)}}_{\text{theoretical quantiles}}$$

If we plot a scatterplot of the pairs  **$(\pi_{i/(n+1)}, x_{(i)})$** , it should be close to the line  $y = x$ .

This is what we call **quantiles-quantiles plot or QQ plots**, as we plot the theoretical quantiles against the sample quantiles.

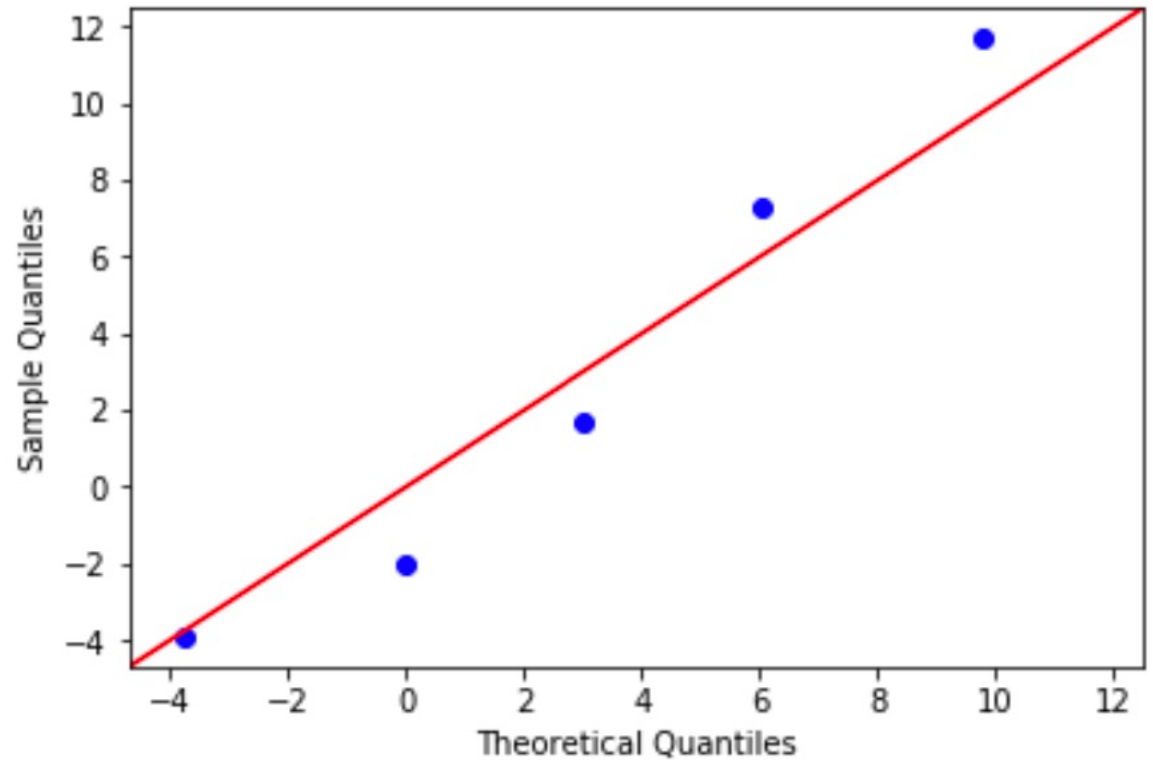
# Q-Q plot Examples

$i$	$x_{(i)}$	$i/(n+1)$
1	-3.9	1/6
2	-2.0	2/6
3	1.7	3/6
4	7.3	4/6
5	11.7	5/6

We suspect that the data is coming from  $N(3, 7^2)$ , and we would like to draw a QQ plot.

# Q-Q plot Examples

Sample quantiles			Theoretical quantiles	
$i$	$x_{(i)}$	$i/(n+1)$	$\pi_{i/(n+1)}$	
1	-3.9	1/6	-3.77	
2	-2.0	2/6	-0.02	
3	1.7	3/6	3	
4	7.3	4/6	6.02	
5	11.7	5/6	9.77	



Theoretical quantiles of  $N(3, 7^2)$

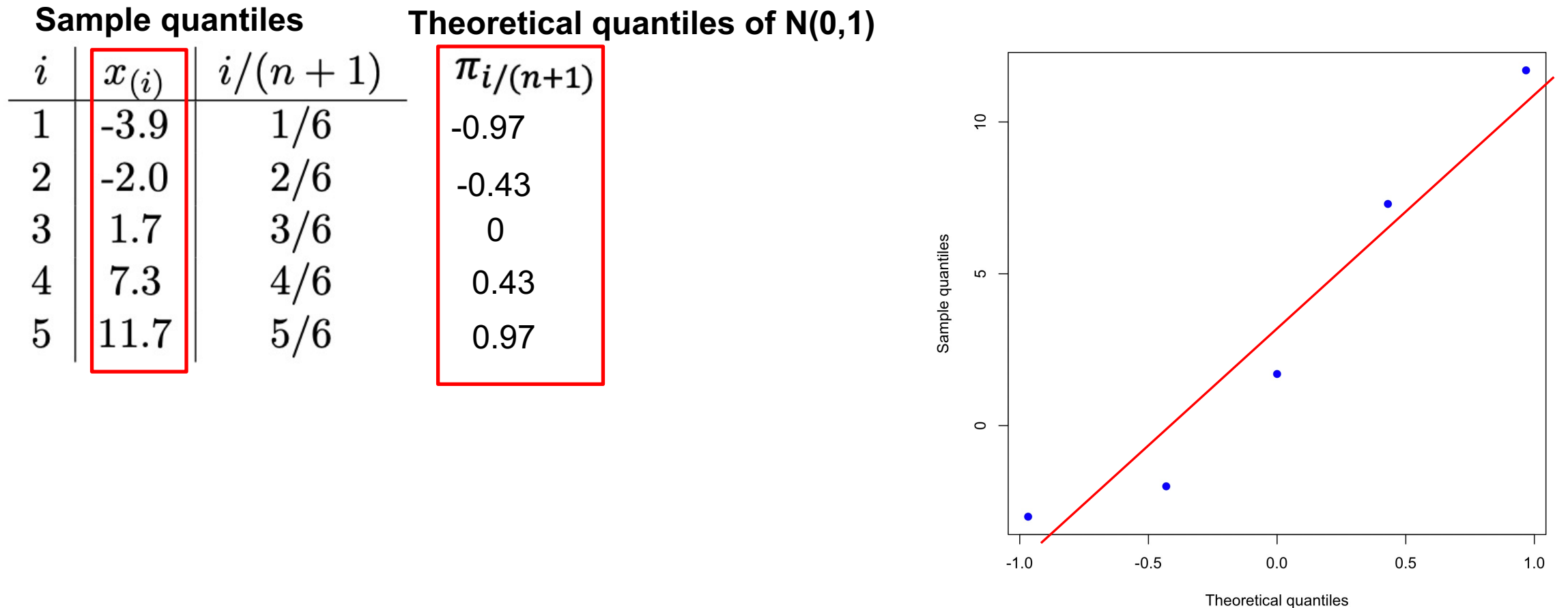
$$X \sim N(3, 7^2)$$

$$P(X < \pi_p) = P\left(\frac{X - 3}{7} \leq \frac{\pi_p - 3}{7}\right) = p$$

$$\Rightarrow \frac{\pi_p - 3}{7} = Z_{1-p} \quad \Rightarrow \quad \pi_p = 3 + 7Z_{1-p}$$

$$\pi_{i/(n+1)} = 3 + 7 \times Z_{1-i/(n+1)}$$

# Q-Q plot Examples



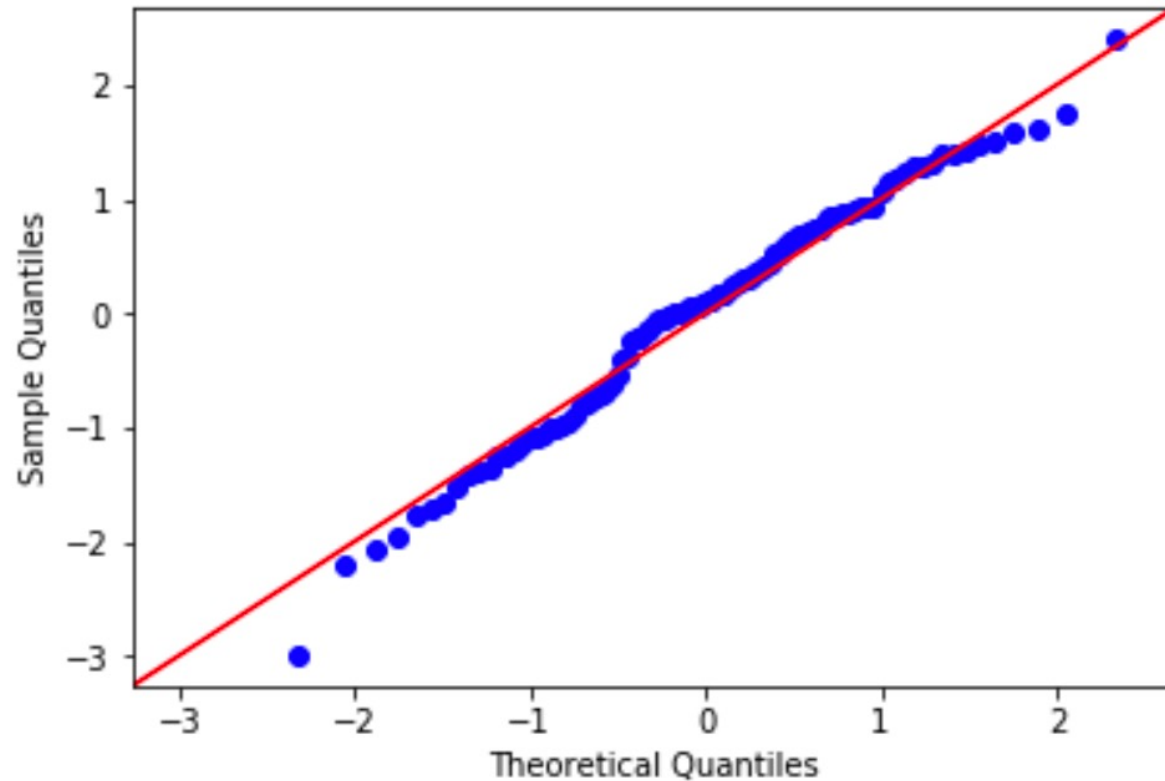
The distributions can be *shifted* and *stretched*.

As long as the shapes match, the Q-Q plot is close to a straight line.



# Example: A good fit

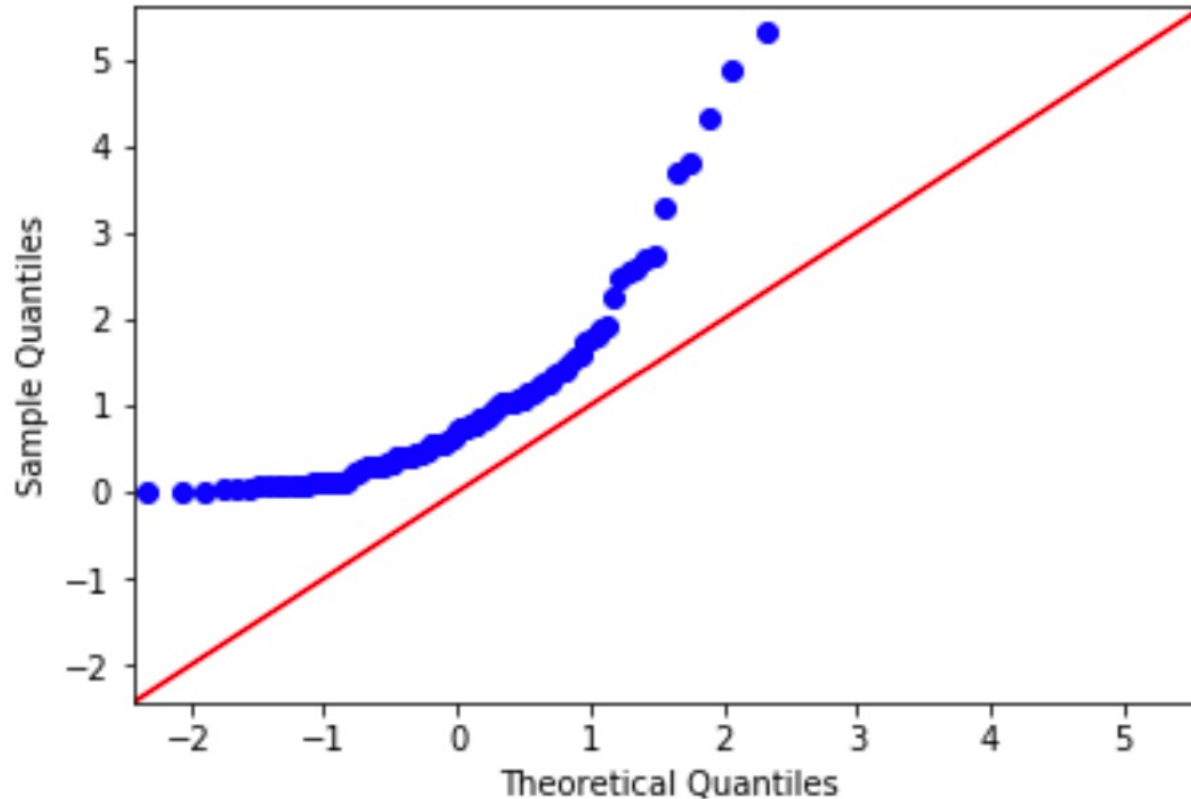
Hypothesis: the data is coming from standard normal.



This dataset looks like a good fit.

# Example: A poor fit

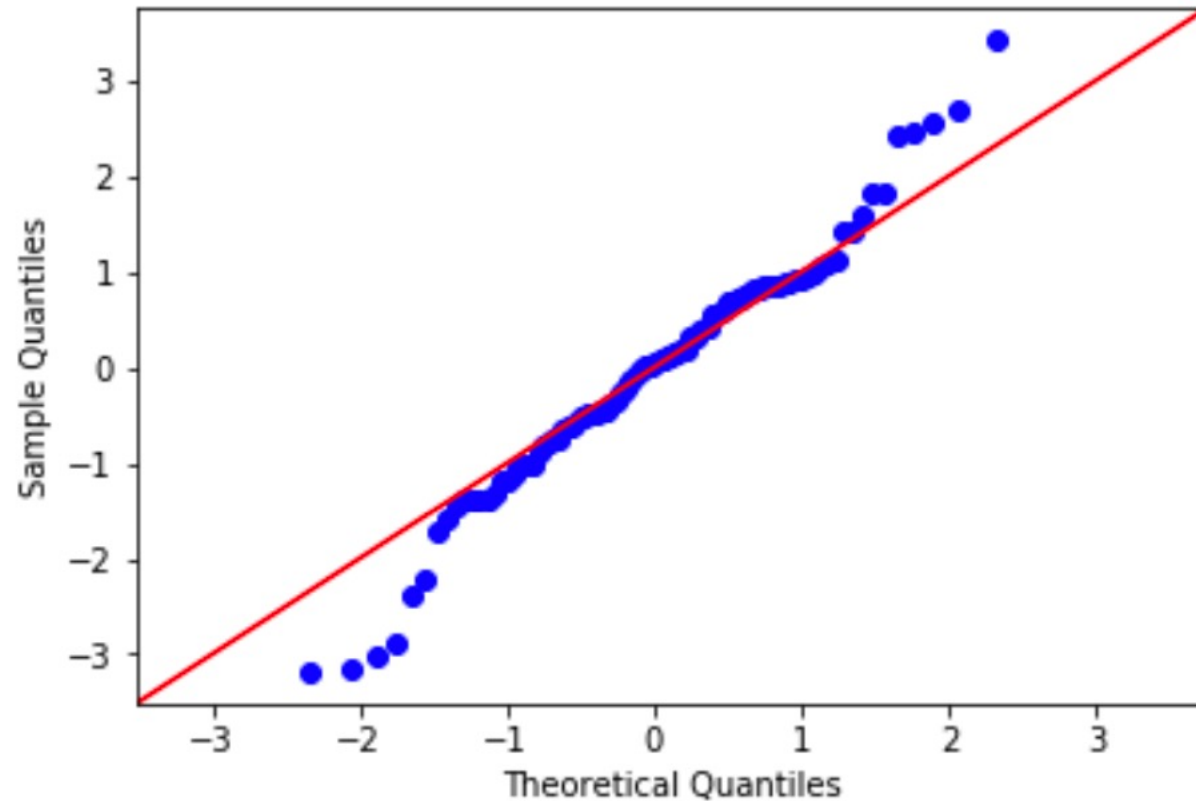
Hypothesis: the data is coming from standard normal.



This is a poor fit, since the data is clearly non-negative, and is more likely to take on larger values than normal distribution (larger “right tail”).

# Example: A poor fit

Hypothesis: the data is coming from standard normal.



This is a poor fit, since the data is more likely to take on extremely large or extremely small values than standard normal (a heavy-tailed distribution)