CSDS 391 HW8

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Exercise 1

a) By Bayes Theorem we establish the following:

$$P(\theta|y,n) = \frac{P(y|\theta,n)P(\theta)}{P(y|n)} \tag{1}$$

Now, we assume a uniform prior so $P(\theta) = 1$. So from equation (1) we find that:

$$P(\theta|y,n) = \frac{P(y|\theta,n)}{\int_0^1 P(y|\theta,n)d\theta} = P(y|\theta,n)(n+1)$$
 (2)

Finally, given the known probability density function of a binomial distribution, inserting into (2) we conclude:

$$P(\theta|y,n) = \binom{n}{y} \theta^y (1-\theta)^{n-y} (n+1)$$
(3)

b)

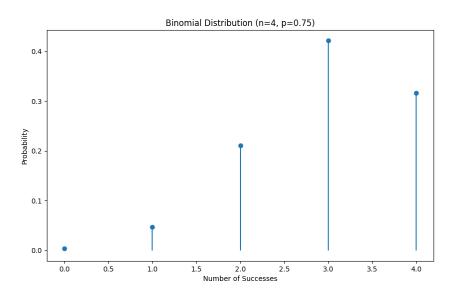


Figure 1: Likelihood of Binomial Distribution

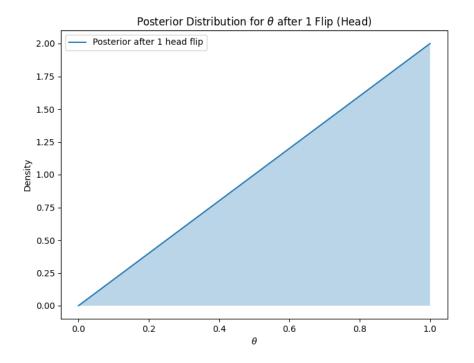


Figure 2: Posterior Distribution After 1 Heads

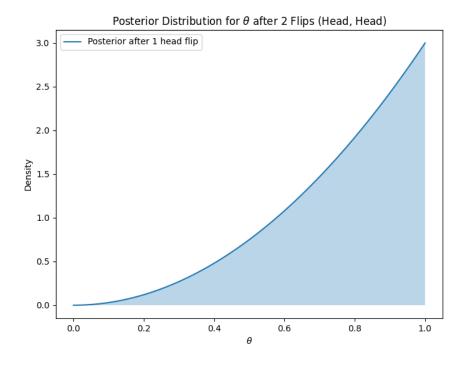


Figure 3: Posterior Distribution After 2 Heads

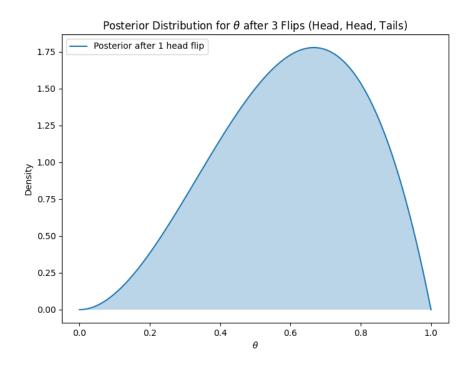


Figure 4: Posterior Distribution After 2 Heads one Tails

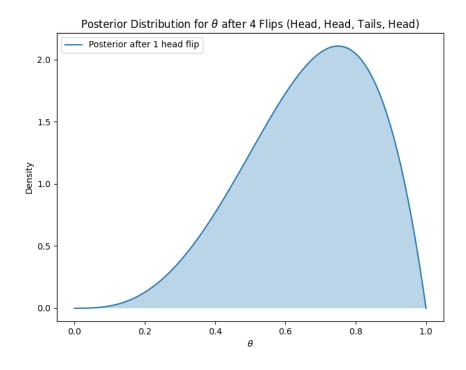


Figure 5: Posterior Distribution After 3 Heads one Tails

Exercise 2

Suppose we have x_1, \ldots, x_n independent, identically distributed observations. We establish the probability density function of the Gaussian distribution as:

$$f(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi\sigma_3^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$
 (4)

Given our assumptions, we apply the likelihood function for μ and θ :

$$L(\mu, \sigma^2) = \prod_{i=1}^n f(x_i | \mu, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x_i - \mu)^2}{2\sigma^2}\right)$$
 (5)

From here we can optimize the likely hood function with respect to θ and μ to find their respective maximum likelihood estimates. Let us begin by simplifying our expression as much possible.

$$L(\mu, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x_i - \mu)^2}{2\sigma^2}\right)$$

$$\ln L(\mu, \sigma^2) = \ln \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x_i - \mu)^2}{2\sigma^2}\right) = \sum_{i=1}^n \left(\ln \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x_i - \mu)^2}{2\sigma^2}\right)\right)$$

$$\ln L(\mu, \sigma^2) = \sum_{i=1}^n \left(\ln \left(\frac{1}{\sqrt{2\pi\sigma^2}}\right) - \frac{(x_i - \mu)^2}{2\sigma^2}\right) = \sum_{i=1}^n \left(-\frac{1}{2}\ln \left(2\pi\sigma^2\right) - \frac{(x_i - \mu)^2}{2\sigma^2}\right)$$

$$\ln L(\mu, \sigma^2) = -\frac{n}{2}\ln \left(2\pi\sigma^2\right) - \frac{1}{2\sigma^2}\sum_{i=1}^n (x_i - \mu)^2$$

Now we optimize with respect to μ .

$$\frac{\partial}{\partial \mu} \left(\ln L(\mu, \sigma^2) = -\frac{n}{2} \ln \left(2\pi \sigma^2 \right) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2 \right)$$

$$\frac{\partial}{\partial \mu} \ln L(\mu, \sigma^2) = \frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \mu)$$

$$\frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \mu) = 0$$

$$\sum_{i=1}^n (x_i - \mu) = 0$$

$$(x_1 - \mu) + (x_2 - \mu) + \dots + (x_n - \mu) = 0$$

$$\sum_{i=1}^n x_i = n\mu$$

$$\mu = \frac{1}{n} \sum_{i=1}^n x_i$$

We conclude that $\hat{\mu}_{\text{MLE}} = \frac{1}{n} \sum_{i=1}^{n} x_i$. Now for σ .

$$\frac{\partial}{\partial \sigma} \left(\ln L(\mu, \sigma^2) = -\frac{n}{2} \ln \left(2\pi \sigma^2 \right) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2 \right)$$

$$\frac{\partial}{\partial \sigma} \ln L(\mu, \sigma^2) = -\frac{n}{2} \frac{4\pi\sigma}{2\pi\sigma^2} + \frac{1}{\sigma^3} \sum_{i=1}^n (x_i - \mu)^2 = 0$$

$$-\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^n (x_i - \mu)^2 = 0$$

$$\frac{1}{\sigma^3} \sum_{i=1}^n (x_i - \mu)^2 = \frac{n}{\sigma}$$

$$\sum_{i=1}^n (x_i - \mu)^2 = n\sigma^2$$

$$\sigma^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2$$

We conclude that $\hat{\sigma}_{\text{MLE}}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2$.

Exercise 3

a) We establish that $C_1 \sim \mathcal{N}(\mu_1, \sigma_1^2)$ and $C_2 \sim \mathcal{N}(\mu_2, \sigma_2^2)$. Moreover, we are given that, from the priors: $P(C_1) = 2P(C_2)$. By the axioms of probability:

$$P(C_1) + P(C_2) = 1$$

$$2P(C_2) + P(C_2) = 1$$

$$P(C_2) = \frac{1}{3}$$

$$P(C_1) = \frac{2}{3}$$

We can also derive the probability of P(x) as:

$$P(x) = P(x|C_1)P(C_1) + P(x|C_2)P(C_2)$$

$$P(x) = P(x|C_1)\frac{2}{3} + P(x|C_2)\frac{1}{3}$$

$$P(x) = \frac{2}{3}\frac{1}{\sqrt{2\pi (\sigma_1)^2}} \exp\left(-\frac{(x-\mu_1)^2}{2(\sigma_1)^2}\right) + \frac{1}{3}\frac{1}{\sqrt{2\pi (\sigma_2)^2}} \exp\left(-\frac{(x-\mu_2)^2}{2(\sigma_2)^2}\right)$$

b) Denote the probability of error as $P(\mathbf{E}) := P(\text{Classify } x \in C_1 | x \in C_2) P(C_2) + P(\text{Classify } x \in C_2 | x \in C_1) P(C_1)$ in general terms. We are given $\mu_1 < \mu_2$. Interpreting this visually, we expect the probability distribution of $P(x|C_1)$ to lie "below", if you will, that of $P(x|C_2)$ by virtue of them following a Gaussian distribution. As defined in the problem, let $x = \theta$ be the decision boundary, and let us say that $x \in C_1 \iff x < \theta$ and $x \in C_2 \iff x \ge \theta$. So we can derive the probability of misclassifying an $x \in C_1$ as a C_2 element as $P(x \ge \theta | x \in C_1)$ and $P(x < \theta | x \in C_2)$ for the converse. Here we can apply the Cumulative Distribution Function for the Gaussian distribution. Namely:

$$P(x \ge \theta | C_1) = 1 - \int_{-\infty}^{\theta} P(x | C_1) dx = P(\text{Classify } x \in C_2 | x \in C_1)$$
 (6)

$$P(x < \theta | C_2) = \int_{-\infty}^{\theta} P(x | C_2) dx = P(\text{Classify } x \in C_1 | x \in C_2)$$
 (7)

Finally, we can establish from (6) and (7) that:

$$P(\mathbf{E}) = \frac{2}{3} \left(1 - \int_{-\infty}^{\theta} P(x|C_1) \, dx \right) + \frac{1}{3} \left(\int_{-\infty}^{\theta} P(x|C_2) \, dx \right)$$

$$P(\mathbf{E}) = \frac{2}{3} \left(1 - \int_{-\infty}^{\theta} \frac{1}{\sqrt{2\pi (\sigma_1)^2}} \exp\left(-\frac{(x - \mu_1)^2}{2 (\sigma_1)^2} \right) \, dx \right) + \frac{1}{3} \left(\int_{-\infty}^{\theta} \frac{1}{\sqrt{2\pi (\sigma_2)^2}} \exp\left(-\frac{(x - \mu_2)^2}{2 (\sigma_2)^2} \right) \, dx \right)$$

Exercise 4

a) We are given the objective function:

$$D = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{n,k} \|\mathbf{x_n} - \mu_{\mathbf{k}}\|^2$$

Since we are taking the 2-norm of column vectors, we can consider the function represented by D a scalar valued function. Hence we take the partial derivative with respect to $\mu_{k,i}$ (the *i*th value of μ_k vector) and

optimize. Moreover, by linearity of the derivative operation, we may (essentially) bypass the sums.

$$D = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{n,k} \|\mathbf{x}_{n} - \mu_{i}\|^{2} = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{n,k} \sum_{j=1}^{I} (x_{n,j} - \mu_{k,j})^{2}$$
$$\frac{\partial D}{\partial \mu_{k,i}} = \frac{\partial}{\partial \mu_{k,i}} \left(-2 \sum_{n=1}^{N} \sum_{k=1}^{K} r_{n,k} \sum_{j=1}^{I} (x_{n,j} - \mu_{k,j}) \right)$$

Now, since we are taking the partial derivative with respect to the *i*th scalar value of the *k*th mean vector we can ignore all terms where: a) $j \neq i$ and b) terms where we are not in the *k*the cluster. Hence we can drop those 2 sums and replace all instances of j with i as we are considering the case where j = i only.

$$\begin{split} \frac{\partial D}{\partial \mu_{k,i}} &= -2 \sum_{n=1}^{N} r_{n,k} \left(x_{n,i} - \mu_{k,i} \right) = 0 \\ &\sum_{n=1}^{N} r_{n,k} \left(x_{n,i} - \mu_{k,i} \right) = 0 \\ &\sum_{n=1}^{N} r_{n,k} \left(x_{n,i} - \mu_{k,i} \right) = r_{1,k} x_{1,i} - r_{1,k} \mu_{k,i} + \dots + r_{N,k} x_{N,i} - r_{N,k} \mu_{k,i} = 0 \\ &r_{1,k} x_{1,i} + \dots + r_{N,k} x_{N,i} = r_{1,k} \mu_{k,i} + \dots + r_{N,k} \mu_{k,i} \\ &\sum_{n=1}^{N} r_{n,k} x_{n,i} = \mu_{k,i} \sum_{n=1}^{N} r_{n,k} \\ &\mu_{i,k} = \frac{\sum_{n=1}^{N} r_{n,k} x_{n,i}}{\sum_{n=1}^{N} r_{n,k}} \end{split}$$

b) We can generalize this rule for all entries of μ_k as:

$$\mu_{\mathbf{k}} = \frac{\sum_{n=1}^{N} r_{n,k} \mathbf{x_n}}{\sum_{n=1}^{N} r_{n,k}}$$

Exercise 5

a) Code Snippet

```
import pandas as pd
import numpy as np

# Take euclidean distance of two points
def distance(p1, p2):
    p1, p2 = np.array(p1), np.array(p2)
    dist = np.sqrt(np.sum(np.square(p1 - p2)))
    return dist
```

Assign each point to a cluster

```
def assign_clusters(X, centroids):
    clusters = []
    for x in X.values:
        min_dist = float('inf')
        closest_centroid = -1
        for idx, mu in enumerate(centroids):
            dist = distance(x, mu)
            # Update if this centroid is closer
            if dist < min_dist:</pre>
                min_dist = dist
                closest_centroid = idx
        # Append the index of the closest centroid for this data point
        clusters.append(closest_centroid)
    return clusters
def update_centroids(X, clusters, k):
    # Initialize a list to hold the new centroids
   new_centroids = []
    # Iterate over each cluster index
    for cluster_idx in range(k):
        # Points in cluster
       pts = []
        for idx, pt in enumerate(clusters):
            if pt == cluster_idx:
                pts.append(X.iloc[idx])
        # If there are points in the cluster, calculate the mean
        if pts:
            # Convert pts to a DataFrame to calculate the mean of each feature
            pts_df = pd.DataFrame(pts)
            new_centroid = pts_df.mean().values
            new_centroids.append(new_centroid)
        else:
            new_centroids.append(np.zeros(X.shape[1]))
    return np.array(new_centroids)
# Main procedure
def main(k, data, max_iters=999, err=1e-4):
    # Read csv
    data = pd.read_csv(data)
    # Set seed for reproducibility
    np.random.seed(123)
    # Find number of features
    # Assume the last column is the target column
    features = data.shape[1]-1
    # Get feature columns
    X = data.iloc[:, :features]
```

```
# Select K random data points as the initial centroids
centroids = np.array(X.loc[np.random.choice(X.shape[0], k, replace=False)])

for _ in range(max_iters):
    clusters = assign_clusters(X, centroids)

    new_centroids = update_centroids(X, clusters, k)

    diff = np.abs(new_centroids - centroids)
    if np.all(diff <= err):
        print('Convergence Reached!')
        return centroids, clusters
    centroids = new_centroids

print("No convergence...")
return centroids, clusters</pre>
```

Implementation Explanation

Most of this stuff is by the book, exactly as described in the slides. Upon testing, this seed converged after 5 iterations.

b) Code Snippet

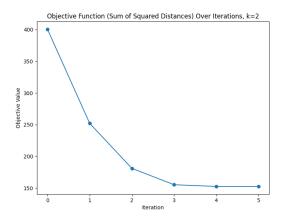
```
# Find objective function for plotting
def calculate_objective(X, clusters, centroids):
    total_distance = 0
    for idx, point in enumerate(X.values):
        centroid = centroids[clusters[idx]]
        total_distance += distance(point, centroid) ** 2
    return total_distance

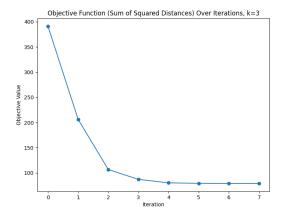
plt.figure(figsize=(8, 6))
plt.plot(objective_values, marker='o')
plt.title(f"Objective Function (Sum of Squared Distances) Over Iterations, k={k}")
plt.xlabel("Iteration")
plt.ylabel("Objective Value")
plt.show()
```

Plots

c) Code Snippet

```
# Plotting function for data points and centroids
def plot_clusters(X, centroids, clusters, title):
    plt.figure(figsize=(8, 6))
    for cluster_idx in range(len(centroids)):
        cluster_points = X.values[np.array(clusters) == cluster_idx]
        plt.scatter(cluster_points[:, 0], cluster_points[:, 1], label=f'Cluster {cluster_idx + 1}')
    centroids = np.array(centroids)
    plt.scatter(centroids[:, 0], centroids[:, 1], s=50, c='black', marker='x', label='Centroids')
    plt.xlabel('Feature 1')
    plt.ylabel('Feature 2')
    plt.title(title)
    plt.legend()
    plt.show()
```

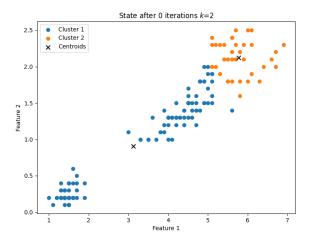




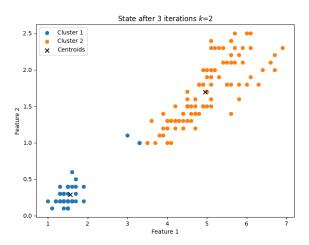
- (a) ${\cal D}$ as a function of the no. of iterations, k=2
- (b) ${\cal D}$ as a function of the no. of iterations, k=3

Figure 6: Comparison of Objective Functions for Different k Values

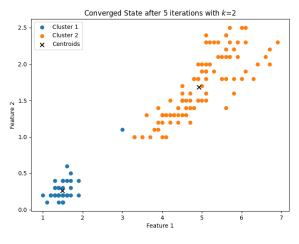
 ${\bf Plots}$



(a) Clusters and Centroids at initialization

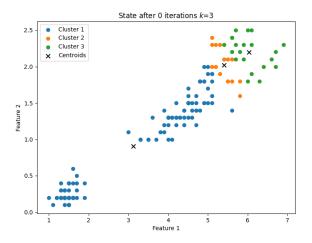


(b) Clusters and Centroids after 3 iterations

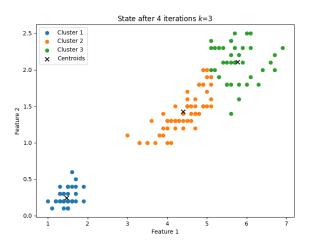


(c) Clusters and Centroids after convergence

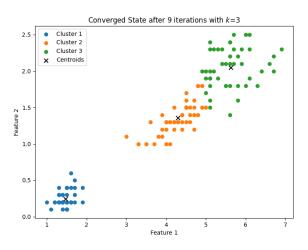
Figure 7: Centroids, Clusters, and Data plotted at different states of k-means clustering where k=2



(a) Clusters and Centroids at initialization



(b) Clusters and Centroids after 3 iterations



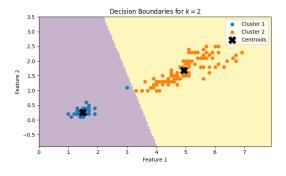
(c) Clusters and Centroids after convergence

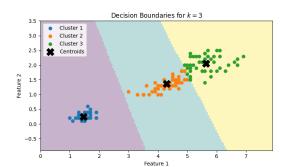
Figure 8: Centroids, Clusters, and Data plotted at different states of k-means clustering where k=3

d) The intuition I followed here is that: If we identify the points on the plot where said point is equidistant to one or more centroids then it is part of a decision boundary. This is what I implemented as:

```
def plot_decision_boundaries_approx(X, centroids, clusters, k, title):
   # Define the range of the grid based on the first two features
   x_{\min}, x_{\max} = X.iloc[:, 0].min() - 1, X.iloc[:, 0].max() + 1
   y_min, y_max = X.iloc[:, 1].min() - 1, X.iloc[:, 1].max() + 1
   xx, yy = np.meshgrid(np.arange(x_min, x_max, 0.05),
                         np.arange(y_min, y_max, 0.05))
   # Prepare the grid points to assign clusters
   grid_points = np.c_[xx.ravel(), yy.ravel()]
   grid_df = pd.DataFrame(grid_points, columns=[X.columns[0], X.columns[1]])
   # Assign each grid point to the nearest centroid
   grid_clusters = assign_clusters(grid_df, centroids)
   grid_clusters = np.array(grid_clusters).reshape(xx.shape)
   # Plot decision boundaries by coloring each region
   plt.figure(figsize=(8, 6))
   plt.imshow(grid_clusters, extent=(x_min, x_max, y_min, y_max),
               origin='lower', cmap='viridis', alpha=0.3, interpolation='nearest')
   # Plot the actual data points and centroids
   for cluster_idx in range(k):
        cluster_points = X.values[np.array(clusters) == cluster_idx]
       plt.scatter(cluster_points[:, 0], cluster_points[:, 1], label=f'Cluster {cluster_idx + 1}')
   centroids = np.array(centroids)
   plt.scatter(centroids[:, 0], centroids[:, 1], s=200, c='black', marker='X', label='Centroids')
   plt.xlabel('Feature 1')
   plt.ylabel('Feature 2')
   plt.title(title)
   plt.legend()
   plt.show()
```

Plots





(a) Decision Boundaries k=2

(b) Decision Boundaries k=3

Figure 9: Decision Boundaries for different k values