

# AI HW #6

Maximilian Schulten (mls384 )

2024-10-12

## Exercise 1

a)

We check this by verifying the axioms of probability. Namely, every element  $0 \leq P(x = i, y = j) \leq 1$ ,  $\forall i, j \in \{1, 2, 3, 4\}$ : This clearly holds by inspection. Moreover that  $\sum_{i,j} P(x = i, y = j) = 1$ . Clearly demonstrated below:

```
0.06+0.08+0.04+0.02+0.12+0.16+0.08+0.04+0.09+0.12+0.06+0.03+0.03+0.04+0.02+0.01
```

```
## [1] 1
```

b)

Define  $P_x(x)$  as the marginal distribution of  $x$ .

$$P_x(x = i) := \sum_j p(x = i, y = j), \forall i \in \{1, 2, 3, 4\}$$

So we fix  $x$  and let  $y$  vary. Summing the rows we find:

```
joint_prob <- matrix(c(  
  0.06, 0.08, 0.04, 0.02,  
  0.12, 0.16, 0.08, 0.04,  
  0.09, 0.12, 0.06, 0.03,  
  0.03, 0.04, 0.02, 0.01  
) , nrow = 4, byrow = TRUE)
```

```
px <- rowSums(joint_prob)  
py <- colSums(joint_prob)
```

```
px
```

```
## [1] 0.2 0.4 0.3 0.1
```

```
py
```

```
## [1] 0.3 0.4 0.2 0.1
```

So we conclude:

$$P_x(x) = \begin{cases} 0.2, & x = 1 \\ 0.4, & x = 2 \\ 0.3, & x = 3 \\ 0.1, & x = 4 \end{cases}$$

c)

$X$  and  $Y$  are independent iff their joint probabilities can be factored as a product of their marginal probabilities. Namely we need to check that  $P(x = i, y = j) = P_x(x = i) \cdot P_y(y = j)$  for every element of the table. We can easily see that this is the case. Hence  $X$  and  $Y$  are independent.

## Exercise 2

a)

*Proposition:*  $P(a|b, c) = P(b|a, c) \implies P(a|c) = P(b|c)$ .

$P(a|b, c) = P(b|a, c) \implies \frac{P(a, b, c)}{P(b, c)} = \frac{P(a, b, c)}{P(a, c)} \implies P(a, c) = P(b, c) \implies P(a|c)P(c) = P(b|c)P(c) \implies P(a|c) = P(b|c)$ ; **QED**.

b)

*Proposition:*  $P(a|b, c) = P(a) \implies P(b, c) = P(b)$ .

$P(a|b, c) = P(a) \iff \frac{P(a, b, c)}{P(b, c)} = P(a) \iff P(b, c) = \frac{P(a, b, c)}{P(a)} = P(b, c|a) \iff P(c|b)P(b) = P(b, c|a) = \frac{P(a|b, c)P(b, c)}{P(a)} \iff P(c|b)P(b) = P(b, c)$

$\therefore P(a|b, c) = P(a) \implies P(b, c) = P(b) \iff P(c|b) = 1$  i.e.  $c$  must be guaranteed to occur when  $b$  occurs. Hence the proposition is false.

c)

*Proposition:*  $P(a|b) = P(a) \implies P(a|b, c) = P(a|c)$ .

$P(a|b, c) = \frac{P(a, b, c)}{P(b, c)} = \frac{P(a, c)P(b|c)}{P(b|c)P(c)} = P(a|c)$ ; **QED**.

The statement  $P(a, b, c) = P(a, c)P(b|c)$  follows from:

$P(a, b, c) = P(a, c)P(b|a, c)$ . Since  $P(a|b) = P(a)$ ,  $a$  and  $b$  are independent. Hence  $P(b|a, c) = P(b|c)$ . So,  $P(a, b, c) = P(a, c)P(b|c)$ .

## Exercise 3

As a proxy for how indicative the tests are I will say that the test with the higher probability of a patient having the disease given that the test returns positive is the more indicative of the two. To that end I will use Bayesian inference, and define  $D$  as the event that a patient is sick.

We are given that, For  $A$ :

$$P(T = \text{true}|D = \text{true}) = 0.95, P(T = \text{true}|D = \text{false}) = 0.1, P(D = \text{true}) = 0.01$$

Using this information we find:

$$P(D = \text{true}|T = \text{true}) = \frac{P(T = \text{true}|D = \text{true})P(D = \text{true})}{P(T = \text{true})} = \frac{0.95 \cdot 0.01}{\sum_D P(T = \text{true}, D)}$$

Notice that:

$$\sum_D P(T = \text{true}, D) = \sum_D P(T = \text{true}|D)P(D) = 0.95 \cdot 0.01 + 0.1 \cdot 0.99 = 0.1085.$$

Hence:

$$P(D = \text{true}|T = \text{true}) = \frac{0.95 \cdot 0.01}{0.1085} = 0.08756 = 8.756\%$$

Following a very similar procedure for B, we are given that:

$$P(T = \text{true}|D = \text{true}) = 0.9, P(T = \text{true}|D = \text{false}) = 0.05, P(D = \text{true}) = 0.01$$

Using this information we find:

$$P(D = \text{true}|T = \text{true}) = \frac{P(T = \text{true}|D = \text{true})P(D = \text{true})}{P(T = \text{true})} = \frac{0.9 \cdot 0.01}{\sum_D P(T = \text{true}, D)}$$

Notice that:

$$\sum_D P(T = \text{true}, D) = \sum_D P(T = \text{true}|D)P(D) = 0.9 \cdot 0.01 + 0.05 \cdot 0.99 = 0.0585.$$

Hence:

$$P(D = \text{true}|T = \text{true}) = \frac{0.9 \cdot 0.01}{0.0585} = 0.15385 = 15.385\%$$

We conclude that test  $B$  is more indicative of the patient having the virus. This is due to the probabilistic reasoning above yielding almost twice the likelihood of a correct test result given that the disease is being carried when using  $B$  versus  $A$ .

## Exercise 4

a)

Let  $T$  represent what color was testified under oath, and let  $C$  be the true color of the taxi. We look to find a way to compute the most likely color. We can do this by taking  $\max\{P(C = \text{green}|T = \text{green}), 1 - P(C = \text{green}|T = \text{green})\}$ . Using Bayes' theorem we can say:

$$P(C = \text{green}|T = \text{green}) = \frac{P(T = \text{green}|C = \text{green})P(C = \text{green})}{P(T = \text{green})}$$

However, we have no information on the distribution of taxis in Athens; namely  $P(C)$ . So, we cannot find this probability!

b)

Now we have  $P(C = \text{green}) = 0.9$  and implicitly  $P(C = \text{blue}) = 0.1$ . Returning back to the above we find:

$$P(C = \text{green}|T = \text{green}) = \frac{P(T = \text{green}|C = \text{green})P(C = \text{green})}{P(T = \text{green})} = \frac{0.75 \cdot 0.9}{\sum_C P(T = \text{green}, C)}$$

We know:

$$P(T = \text{green}) = \sum_D P(T = \text{green}, D) = \sum_D P(T = \text{green}|C)P(C) = 0.75 \cdot 0.9 + 0.25 \cdot 0.1 = 0.7$$

So:

$$P(C = \text{green}|T = \text{green}) = \frac{0.75 \cdot 0.9}{0.7} = 0.9643 = 96.43\%$$

It goes without saying then that  $\max\{P(C = \text{green}|T = \text{green}), 1 - P(C = \text{green}|T = \text{green})\} = P(C = \text{green}|T = \text{green}) = 0.9643$ . Therefore, given that we testify the taxi is green, there is a 96.43% chance the taxi is in fact green.

## Exercise 5

We have:

$p(L, G, V)$	$L$	$G$	$V$
0.87318	$F$	$F$	$F$
0.0072	$F$	$F$	$T$
0.049	$F$	$T$	$F$
0.0004	$F$	$T$	$T$
0.00882	$T$	$F$	$F$
0.0108	$T$	$F$	$T$
0.049	$T$	$T$	$F$
0.0016	$T$	$T$	$T$

a)

$$P(L) = \sum_{G,V} P(L, G, V) = 0.00882 + 0.0108 + 0.049 + 0.0016 = 0.07022$$

$P(L)$	$L$
0.0696	$T$
0.9304	$F$

b)

$$P(G) = \sum_{L,V} P(L, G, V) = 0.049 + 0.0004 + 0.049 + 0.0016 = 0.1$$

$P(G)$	$G$
0.1	$T$
0.9	$F$

c)

$$P(L = T | G = T) = \frac{P(L=T, G=T)}{P(G=T)} = \frac{0.049+0.0016}{0.1} = 0.5$$

$$P(L = F | G = T) = \frac{P(L=F, G=T)}{P(G=T)} = \frac{0.049+0.0004}{0.1} = 0.5$$

$$P(L = T | G = F) = \frac{P(L=T, G=F)}{P(G=F)} = \frac{0.0082+0.0108}{0.9} = 0.0215556$$

$$P(L = F | G = F) = \frac{P(L=F, G=F)}{P(G=F)} = \frac{0.87318+0.0072}{0.9} = 0.9784444$$

$P(L G)$	$L$	$G$
0.9784	$F$	$F$
0.494	$F$	$T$
0.0215556	$T$	$F$
0.506	$T$	$T$

d)

$$P(G = T | L = T \vee V = T) = \frac{P(G=T, L=T \vee V=T)}{P(L=T \vee V=T)} = \frac{0.0004+0.049+0.0016}{0.0072+0.0004+0.00882+0.0108+0.049+0.0016} = 0.6554$$

e)

Let us translate the problem statement. We do not have any information on  $V$ , hence we treat it as we would if it hadn't been mentioned.  $P(L = F | G = T) = 0.0215556$

## Exercise 6

a)

Every hypothesis makes a statement about the probability of drawing a certain flavor. In a sense we can see this as a series of Bernoulli trials modeled via a binomial distribution. As we gather more and more data (i.e. draw more candy from the bag) the probability of success, whichever way we may define it, resembles the actual sample more and more for one of the hypotheses. Hence we become more confident in our guesses. From the text we have the formula:

$$P(\mathbf{d}|h_i) = \prod_j P(d_j|h_i)$$

The reason why certain posterior probabilities begin at 0.2, others at 0.1, and one at 0.4 is defined in the problem. When we haven't observed any candy we have no ability to infer beyond the probabilities of each bag type, which is what is modeled at  $n = 0$ .

b)