

# STATS5620\_Q2\_Analysis

## STAT 5620 Project Outline

### Linking Diet and Reproductive Success in Northwest Atlantic Grey Seals

Max Henry-Adams (B00952295) and Peter Braithwaite (B00504978)

Data Analysis STAT 4620/5620 Winter 24-25"

Link to GitHub Repository: <https://github.com/MaxHA17/STAT5620.Project>

#### Abstract:

**Keywords:** Capital breeding strategy, Generalized Linear Models, Pinnipeds, Quantitative Fatty Acid Signature Analysis, Reproductive Ecology

## Introduction

For marine predators, prey distribution may vary unpredictably in time and space, undergoing seasonal, inter annual, and longer-term changes in environmental conditions. Understanding how top predators use different foraging strategies to navigate and adapt to patchy and unpredictable prey availability has important implications for individual fitness and population dynamics, yet studies exploring the consequences of such decisions on fitness are relatively rare (Nathan et al., 2008).

Vertebrates fuel the costs of reproduction along a continuum of an income-capital breeding strategy, where animals that use an income breeding strategy rely on concurrent energy accumulation during the breeding period as opposed to a capital breeding strategy, where energy acquired and stored prior to the breeding period is used solely to finance major reproductive costs (Stephens et al., 2009; Stephens et al., 2014). A capital breeding strategy is thought to be advantageous for large animals capable of carrying large energy stores and offers energetic benefits in areas of patchy or unreliable food availability and allows parents to direct more resources (e.g. time and energy) towards offspring, as opposed to foraging, during the breeding and lactation period (Stephens et al., 2009; Stephens et al., 2014). Given the importance of pre-breeding and lactation period foraging success and energy acquisition in capital breeders, these species offer interesting model systems to study the relationships between diet and reproductive success.

Phocid seals are wide-ranging, large-bodied, and long-lived marine predators, with individuals of many species foraging across a wide range of habitats over large temporal scales (Bowen and Jonsen, 2022). Grey seals (*Halichoerus grypus*) are a long-lived, sexually dimorphic, phocid species and are capital breeders. In the Sable Island population, pregnant females haul out in late December or early January, give birth to a single pup and nurse that pup for 16-18 days, relying solely on energy accumulated during the foraging period prior to parturition to support the costs of lactation (Iverson et al., 1993; Stephens et al., 2009). At or near the abrupt weaning of that pup, females are mated and return to sea to replenish spent body reserves and, after a several month period of delayed implantation, to support

gestation and most importantly preparation for the subsequent December or early January lactation period. As in all mammals, the costs of lactation, and its contribution to reproductive success, far exceed the costs of gestation. Like most other phocids, grey seals are generalist predators (Bowen and Jonsen, 2022), but females tend to feed on a narrower range of energy rich species (predominantly sand lance, redfish, and other pelagic fishes) following the lactation period and expand their prey diversity in the several months leading up to the December-January breeding and lactation period (Beck et al., 2007). Female grey seals reach sexual maturity at age 4-5 years and can continue to reproduce as late as into their early 40s (Bowen et al., 2006). During the brief 16-18-day lactation period, females utilize approximately 25-38% of parturition body mass and daily maternal mass loss is a significant predictor of pup growth rates during lactation, with heavier mothers at parturition weaning heavier pups (Iverson et al., 1993; Mellish et al., 1999). While pup survival is positively related to pup weaning mass up to the mean population weaning mass of 51.5 kg, body length at weaning is the best predictor of grey seal pup survival, as predicted by a bigger-is-better hypothesis for size selective mortality (Bowen et al., 2015). The phenomena of heavier females weaning larger offspring is evident even from primiparity where it is believed that heavier females have larger resource stores that they can mobilize during lactation to produce larger offspring at weaning (Iverson et al., 1993; Mellish et al., 1999; Bowen et al., 2006). Following the lactation period, parental care abruptly ends and the pup is left to fast and survive for several weeks on the energy acquired during lactation before undertaking its first foraging trip (Noren et al., 2008). Thus, answering questions relating to the diet of females prior to parturition will improve our understanding of the importance of habitat use and energy acquisition as they relate to life history characteristics, such as reproductive success, in these capital breeders.

Northwest Atlantic grey seals breeding on Sable Island, Nova Scotia, tend to be central place foragers and have high fidelity to Sable Island for breeding and as a haul-out site, making them excellent candidates for long-term studies (Bowen et al., 2015). Using diet estimations methods, such as quantitative fatty acid (FA) signature analysis (QFASA; Iverson et al., 2004; Beck et al., 2007), it becomes possible to explore the relationship between diet variation and reproductive success at the individual level. QFASA is used to estimate predator diet by comparing the FA signatures of predator adipose tissue to those of candidate prey species (Iverson, 2009; Karnovsky et al., 2012). Predator FA signatures alone have historically been used to provide a qualitative assessment of spatial and temporal patterns in diet diversity. However, since many FAs that are biosynthesized within unique prey species accumulate in predator adipose tissue over time, evaluating the amount of prey-specific FAs relative to the predator's overall FA stores can provide a quantitative estimate of diet, known as QFASA (Iverson et al., 2004). For accurate diet estimation, species-specific calibration coefficients must be experimentally derived to account for the altered incorporation of prey FAs into the predator adipose tissue due to lipid metabolic processes within the predator (Iverson et al., 2004; Karnovsky et al., 2012). When calibration coefficients are derived and the FA signature of many or all potential prey species in an ecosystem is known, as in the case of the Northwest Atlantic grey seal range (Iverson et al., 2004; Beck et al., 2007), QFASA can provide a statistical approach to accurately depicting diet and prey diversity over ecologically relevant time scales (e.g., weeks to several months).

This study will test the hypothesis that differences in diet will affect an individual's ability to store energy prior to parturition, and thus influence their reproductive success. Specifically, we will explore two research questions:

1. Does variation in diet influence maternal mass change over the foraging period leading up to parturition in Northwest Atlantic grey seals?
2. Does female diet predict pup weaning mass in Northwest Atlantic grey seals?

## Data description

---

*Response variables (organized by research question):*

1. Maternal mass change during the pre-breeding season foraging period: continuous variable ranging from -45 – 88 kg. Calculated by subtracting the maternal recovery mass from her deployment mass.
2. Pup weaning mass: continuous variable ranging from 34.5 – 63 kg. Denotes the mass of the pup after the mother has terminated the lactation period and departed the pup and/or Sable Island.

## Predictor or supporting variables

---

3. Year: discrete variable with values ranging from 1996-2015. Denotes the year in which each observation was collected.
4. Maternal ID: discrete variable with a unique identifier for each individual female.
5. Pup sex: categorical; male or female.
6. Maternal age: discrete variable with values ranging from 7-40 years old.
7. Maternal deployment mass: continuous variable with values ranging from 116 – 250 kg. Describes the female mass when she was located pregnant on Sable Island 3-6 months prior to parturition and prior to undergoing her pre-parturition foraging period.
8. Maternal mass at parturition: continuous variable ranging from ~ 137 – 250 kg. This is the mass of the female after returning to Sable Island and giving birth to her pup.
9. Maternal dominant prey species: discrete variable with 6 unique values. Represents the species making up the greatest proportion of the female diet.
10. Maternal dietary energy density: continuous variable ranging from 4.88 – 6.51 kJ/g. Energy density (kJ/g) of the female diet is calculated from the average energy content per gram of prey body tissue.
11. Maternal diet diversity: continuous variable 0.18 – 0.58, no specific units. Diversity is calculated using a standardized Shannon-Weaver diversity index to quantify individual diet diversity.

## Method

The initial step was data cleaning to ensure all readily apparent data errors were rectified prior to data analysis and model fitting. Most of the data errors discovered resulted from the absence of critical data and thus these observations were removed from the dataset. The first question dataset resulted in an initial dataset of 56 observations ( $n=56$ ) and the second question resulted in 76 observations ( $n=76$ ). It should also be noted that additional observations were filtered from the dataset throughout the analysis process as outliers with high leverage were also removed. Thus, the final dataset used for question one had 72 observations and the final dataset for question two had 50 observations.

Once cleaned and uploaded to RStudio, initial data plots were made using the 'flexplot' and 'plot\_explore,' the R function we have included in our Git Hub. These visualizations facilitated the exploration of variable distributions, specifically the response variables. From this it was determined that the response variables for both questions, Maternal Mass Change and Pup Weaning Mass, were Gaussian distributed. This normality allowed the exploration of the data through General Linear Models as well as more complex Generalized Linear Models and Mixed Models. Although our original analytical approach proposed Generalized Linear Mixed Models, it felt prudent to explore the efficiency, interpretability and computational simplicity of Linear Models if the assumption were met and the data suggested linearity.

As a result, both questions were analysed first as Linear Models (LM), then as Generalized Linear Models (GLM) with Gaussian distributions, and finally as Generalized Linear Mixed Models (GLMM) with Random Effects and Fixed Effects on some and all the categorical data. With each model type we began with all the predictor variables prior to using the backward and forward "Step" function to reduce the number of variables through elimination based on the p-values for Linear Models and AIC for GLM and GLMM.

Once we had established the best fit model for both questions using the AIC these models were tested using Cross Validation, question one with 7 folds and question two with 11 folds.

## Analysis

### Research Question 1

---

Here we are exploring whether variation in diet influence maternal mass change over the foraging period leading up to parturition in Northwest Atlantic grey seals?

The response variable of interest is maternal mass change. We will first look at the distribution of this variable:

```
#Load the data
library(readr)
Data_Q1<- read.csv("/Users/peterbraithwaite/Desktop/IDPhD Classes/Stat 5620_Updated/Final

#Ensure the data types are correctly assigned
Data_Q1$MomID <- as.factor(Data_Q1$MomID)
Data_Q1$Year <- factor(Data_Q1$Year, levels = sort(unique(Data_Q1$Year)), ordered = TRUE)
```

```
Data_Q1$`Dominant.prey.species` <- as.character(Data_Q1$`Dominant.prey.species`)
Data_Q1$`Diet.diversity` <- as.numeric(Data_Q1$`Diet.diversity`)
Data_Q1$`Dietary.energy.density` <- as.numeric(Data_Q1$`Dietary.energy.density`)

summary(Data_Q1)
```

```
      MomID      Year Dietary.energy.density Diet.diversity
4269   : 3   2013   :13      Min.   :4.881           Min.   :0.1788
24     : 2   2011   :11    1st Qu.:5.623           1st Qu.:0.3063
146    : 2   2010   : 9    Median :5.753           Median :0.3572
829    : 2   2012   : 9     Mean  :5.736           Mean   :0.3625
3271   : 2   2009   : 6    3rd Qu.:5.893           3rd Qu.:0.4181
3616   : 2   2015   : 5     Max.   :6.517           Max.   :0.5782
(Other):63  (Other):23
Dominant.prey.species  Mass.change
Length:76              Min.   : -40.50
Class :character       1st Qu.:   6.25
Mode  :character       Median : 22.75
                        Mean    : 26.02
                        3rd Qu.: 45.00
                        Max.    : 88.00
```

```
#rename columns
library(dplyr)
```

Attaching package: 'dplyr'

The following objects are masked from 'package:stats':

filter, lag

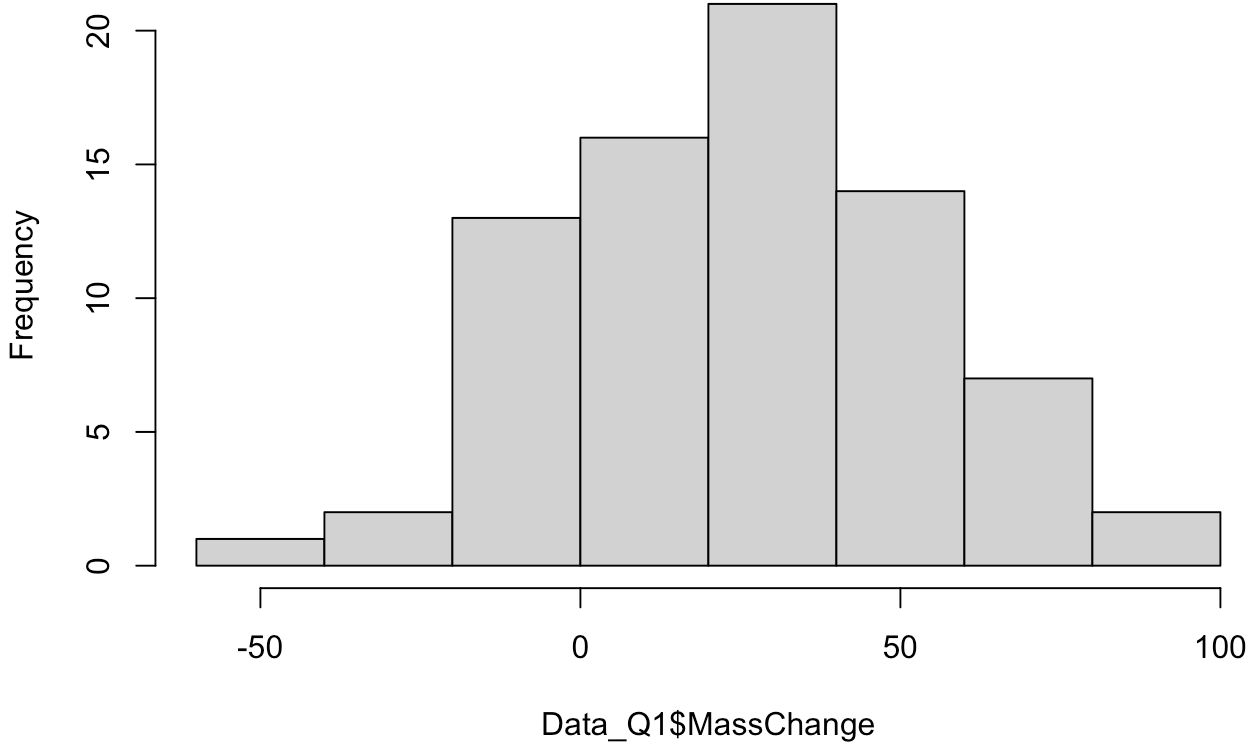
The following objects are masked from 'package:base':

intersect, setdiff, setequal, union

```
Data_Q1 <- Data_Q1 %>%
  rename(
    DietEngDen = `Dietary.energy.density`,
    DietDiv = `Diet.diversity`,
    DomSpp = `Dominant.prey.species`,
    MassChange = `Mass.change`)

hist(Data_Q1$MassChange)
```

## Histogram of Data\_Q1\$MassChange



There are a total of  $n=76$  observations and mass change appears to follow a normal distribution a Gaussian family distribution is a good starting point for an initial model.

One of the covariates, dominant prey species, is a categorical variable. We will first check the number of observations in each category:

```
table(Data_Q1$DomSpp)
```

AtlanticCod	Capelin	NorthernSandlance	Pollock
6	3	29	5
Redfish	WhiteHake		
32	1		

With only one observation of white hake and 3 observations of capelin, there is likely not enough data to estimate the effect of these species on the response variable and we are unable to draw meaningful conclusions. These points also cause extreme residual values. Capelin and white hake observations will be removed from the data set:

```
Data_Q1 <- Data_Q1 %>%  
  filter(DomSpp != "WhiteHake")
```

```
Data_Q1 <- Data_Q1 %>%  
  filter(DomSpp != "Capelin")
```

The initial model seeks to model maternal mass change as a function of dominant prey species, diet diversity, dietary energy density, and deployment year. We will now explore the apparent relationships between each continuous and categorical covariate using the `plot_explore` function built and contained in the `STAT5620.Project R` package:

Initially, it appears that diet diversity is negatively related to mass change, with energy density possibly having a slight negative relationship. It also appears that animals with cod, redfish or sand lance as the dominant prey species gain more mass during foraging than females who primarily prey on Pollock. There does not appear to be any clear trends in mass change over time.

We will begin with an initial model with all covariates of interest. Based on the distribution of mass change, we will begin with a simple model with a gaussian distribution and an identity link function:

```
library(lme4)
```

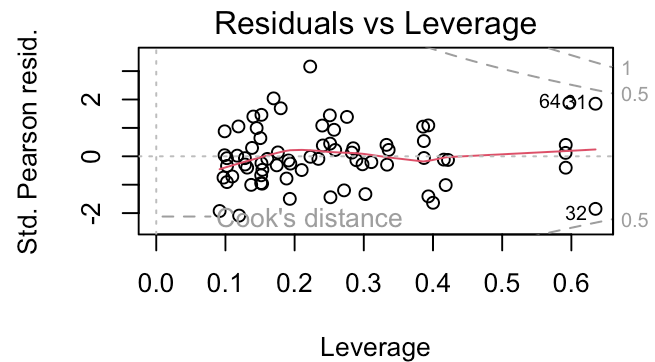
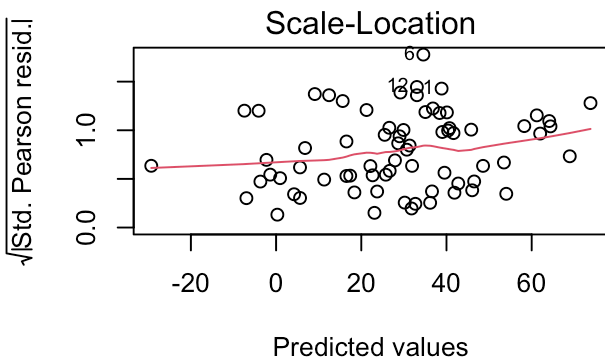
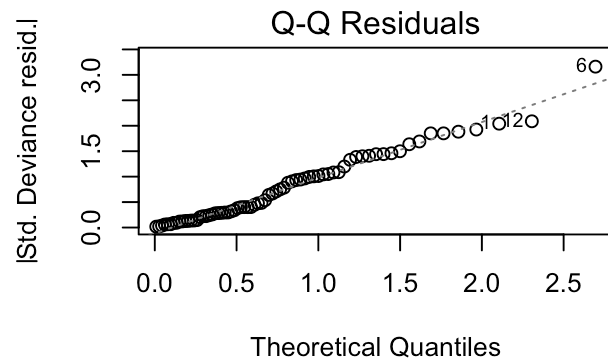
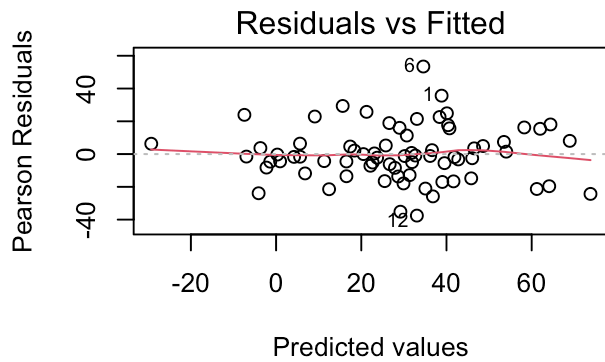
Loading required package: Matrix

```
mom_mod <- glm(data=Data_Q1, family = gaussian(link="identity"), formula = MassChange ~ D  
par(mfrow = c(2, 2))  
plot(mom_mod)
```

Warning: not plotting observations with leverage one:

45





The residual plots from this initial model show good homoscedasticity and normality. There is one observation (observation 6) that appears to be an extreme value. This observation comes from a female that experiences an 88 kg increase in mass over the pre-breeding period foraging period, which is well above the mean mass change value of 28.5. However, none of the covariate values associated with this observation are abnormal and this point does not have high leverage. As a results, the decision is to retain this data point since it is not highly influential and knowledge of the system suggests it is a valid data point.

We can now use step selection to preform repeated hypothesis tests for variable selection to determine which covariates best explain maternal mass change while minimizing model complexity and colinearity between covariates.

```
step(mom_mod, direction = "both")
```

Start: AIC=647.55

MassChange ~ DietDiv + DietEngDen + Year + DomSpp

	Df	Deviance	AIC
- DietEngDen	1	19508	645.67
<none>		19478	647.55
- DietDiv	1	20858	650.48
- DomSpp	3	23417	654.82
- Year	13	38977	671.50



Step: AIC=645.67

MassChange ~ DietDiv + Year + DomSpp

	Df	Deviance	AIC
<none>		19508	645.67
+ DietEngDen	1	19478	647.55
- DietDiv	1	20891	648.60
- DomSpp	3	23602	653.38
- Year	13	39769	670.95

```
Call: glm(formula = MassChange ~ DietDiv + Year + DomSpp, family = gaussian(link =  
"identity"),  
data = Data_Q1)
```

Coefficients:

(Intercept)	DietDiv	Year.L
70.599	-63.415	28.367
Year.Q	Year.C	Year^4
12.387	18.307	15.184
Year^5	Year^6	Year^7
-15.549	3.305	23.361
Year^8	Year^9	Year^10
2.392	-23.277	-11.186
Year^11	Year^12	Year^13
13.261	21.328	-39.502
DomSppNorthernSandlance	DomSppPollock	DomSppRedfish
-18.212	-47.479	-20.303

Degrees of Freedom: 71 Total (i.e. Null); 54 Residual

Null Deviance: 48570

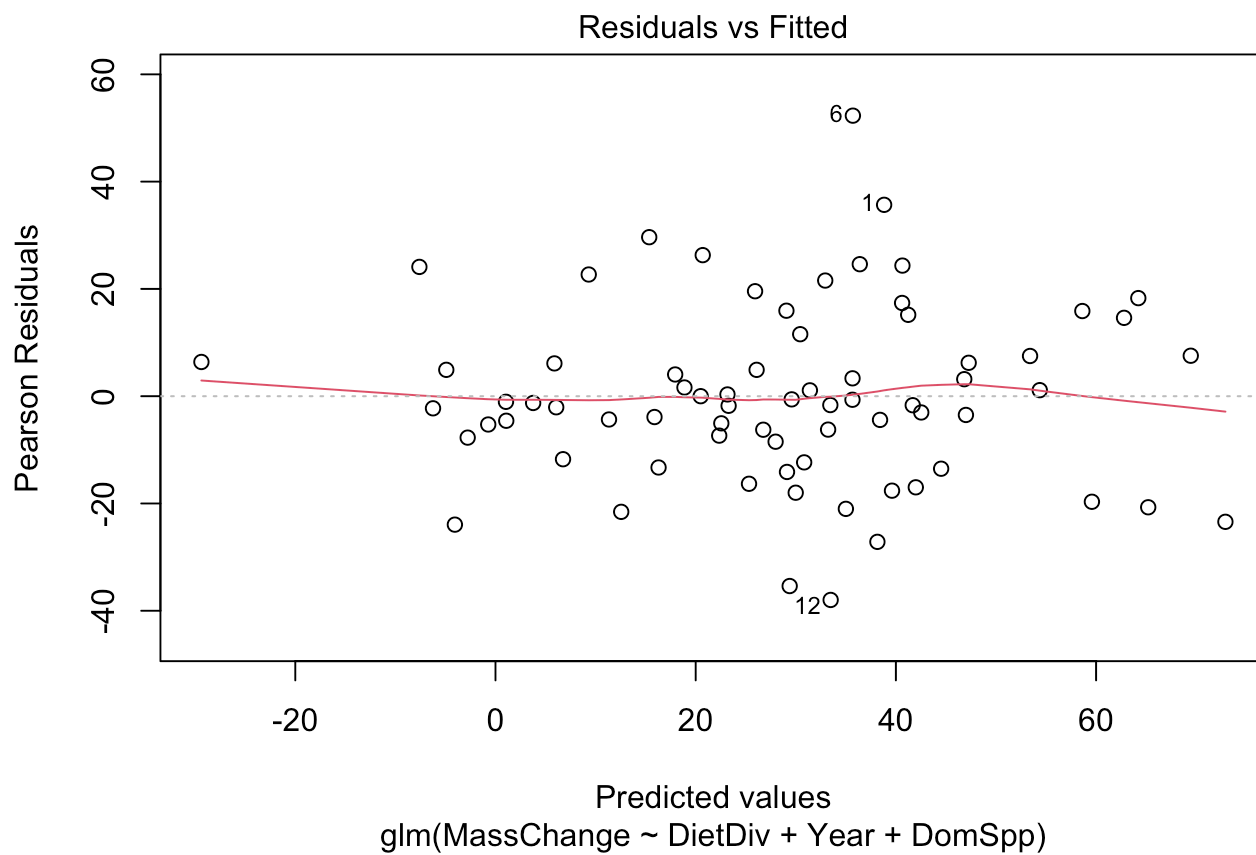
Residual Deviance: 19510 AIC: 645.7

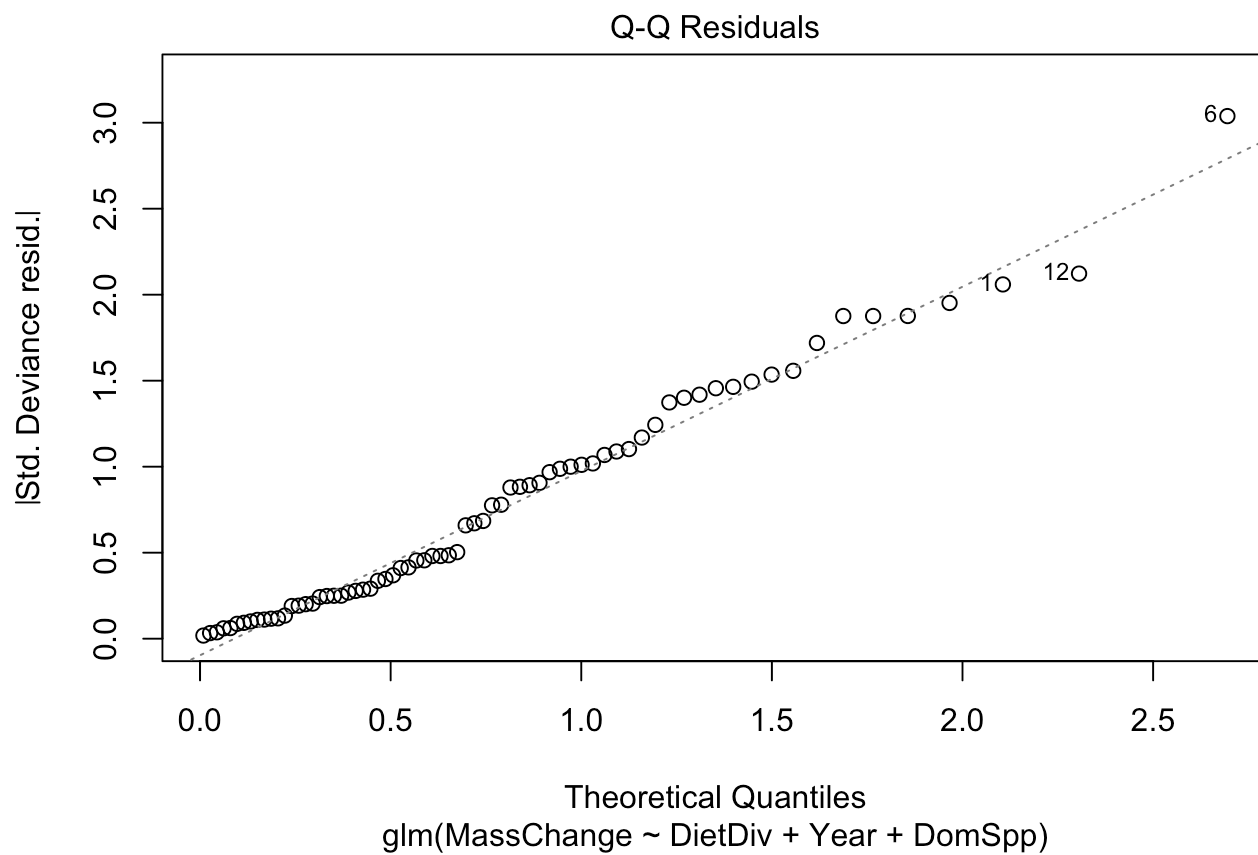
The lowest AIC model drops the dietary energy density covariate. We will update the initial model to drop the energy density covariate:

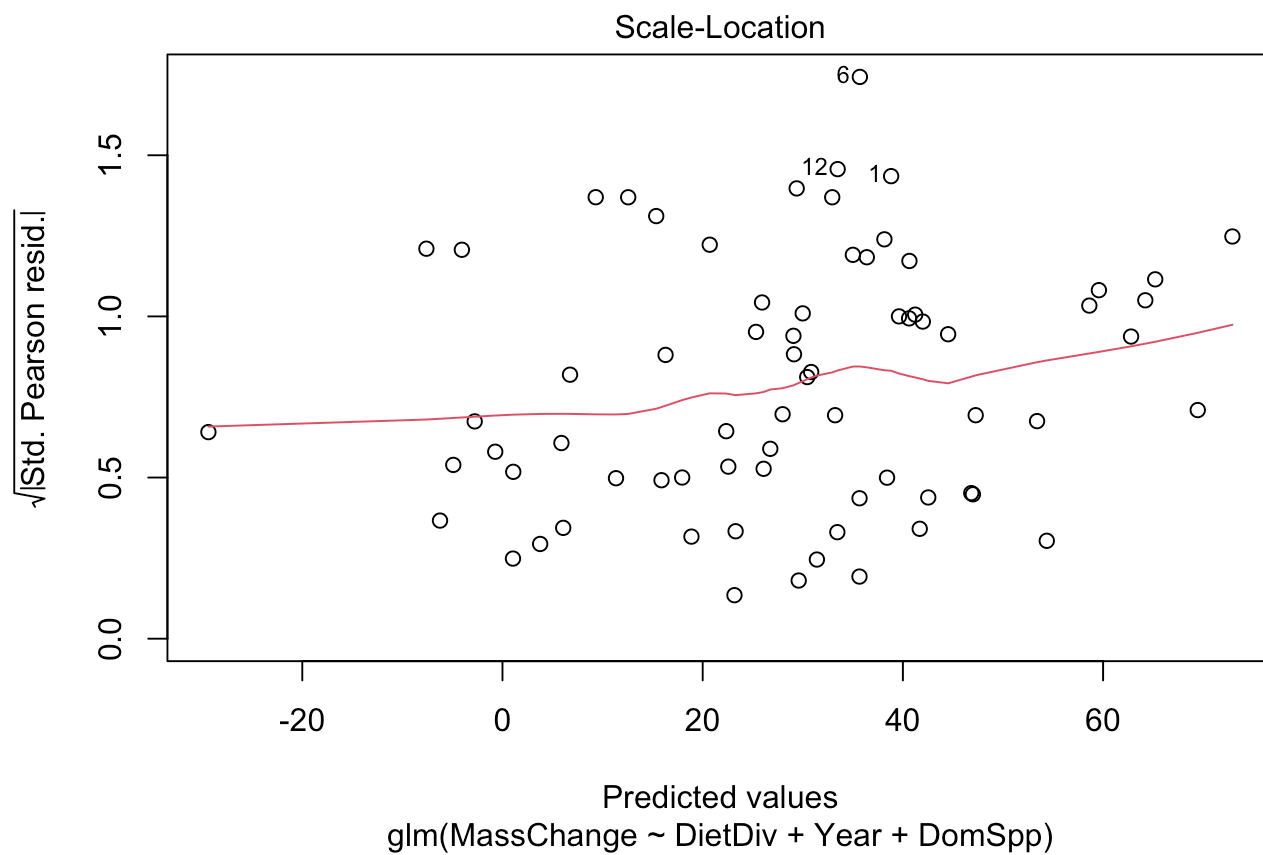
```
mom_mod2 <- glm(formula = MassChange ~ DietDiv + Year + DomSpp, family = gaussian(link =  
data = Data_Q1)  
plot(mom_mod2)
```

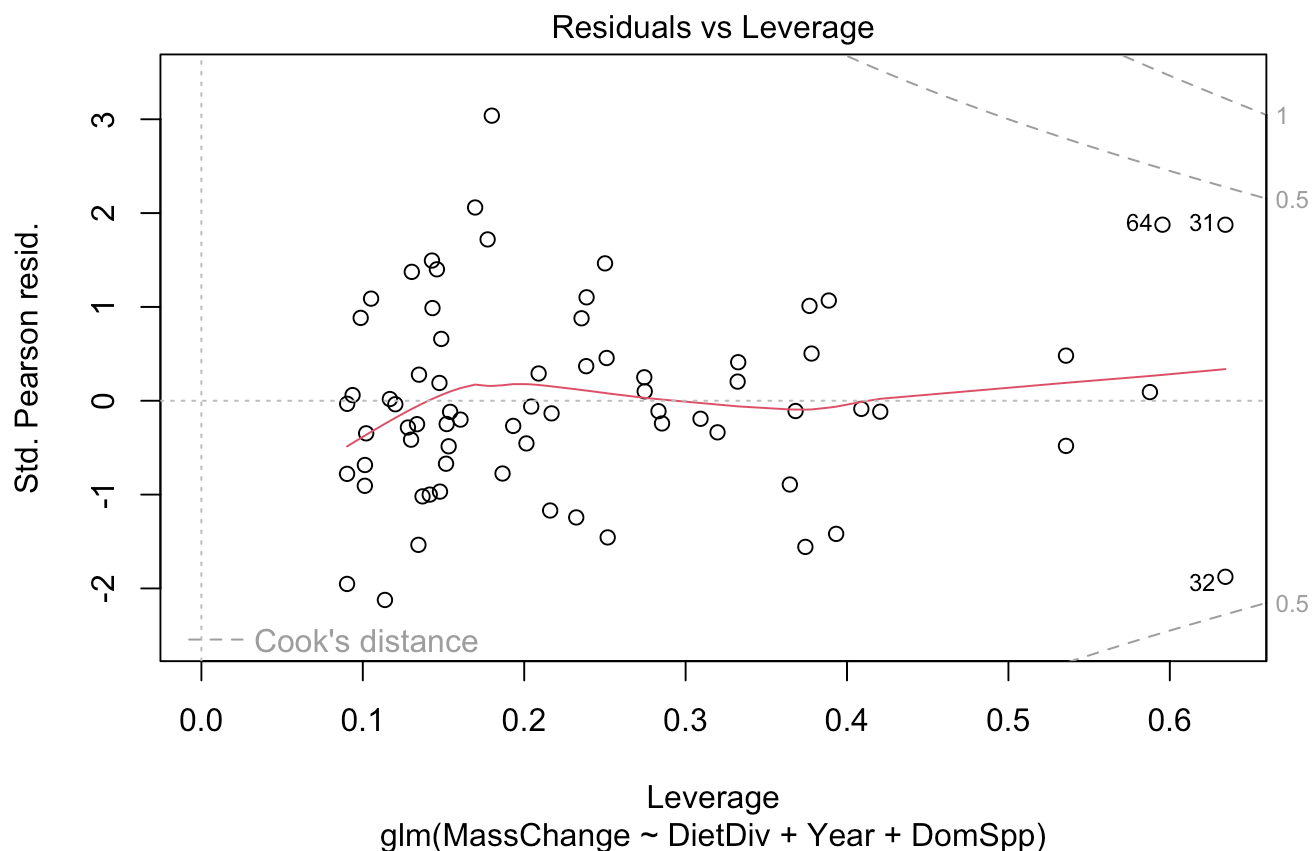
Warning: not plotting observations with leverage one:

45









All model residuals still look good. We can proceed to interpreting model output:

```
summary(mom_mod2)
```

Call:

```
glm(formula = MassChange ~ DietDiv + Year + DomSpp, family = gaussian(link = "identity"),
     data = Data_Q1)
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	70.599	16.241	4.347	6.15e-05 ***
DietDiv	-63.415	32.417	-1.956	0.055621 .
Year.L	28.367	15.657	1.812	0.075584 .
Year.Q	12.387	12.620	0.982	0.330717
Year.C	18.307	13.274	1.379	0.173517
Year^4	15.184	14.787	1.027	0.309065
Year^5	-15.549	12.334	-1.261	0.212857
Year^6	3.305	10.924	0.303	0.763373
Year^7	23.361	10.525	2.219	0.030674 *
Year^8	2.392	10.067	0.238	0.813123
Year^9	-23.277	10.057	-2.314	0.024475 *
Year^10	-11.186	10.451	-1.070	0.289261
Year^11	13.261	9.842	1.347	0.183483

Year <sup>12</sup>	21.328	8.874	2.403	0.019708 *
Year <sup>13</sup>	-39.502	9.864	-4.005	0.000191 ***
DomSppNorthernSandlance	-18.212	14.318	-1.272	0.208824
DomSppPollock	-47.479	16.434	-2.889	0.005550 **
DomSppRedfish	-20.303	13.938	-1.457	0.150992

----

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for gaussian family taken to be 361.2663)

Null deviance: 48567 on 71 degrees of freedom  
 Residual deviance: 19508 on 54 degrees of freedom  
 AIC: 645.67

Number of Fisher Scoring iterations: 2

While several years have mass change values that significantly differ from the expected mass change value, the effect of time is not of direct interest to our research question and is only present to account for changes in prey dynamics over time. As a result, the interpretation of this output is not relevant. However, we see that females who feed primarily on pollock experience lower mass change than the reference level, in this case being females who primarily feed on cod. Being an identity link function, the model is essentially a simple linear model and there is no transformations applied to the response variable. As such, coefficient estimates can be interpreted directly. In this case, individuals who feed on pollock experience 47.5% lower mass change than individuals who primarily feed on cod. This effect is presented in figure 1.

```
require(ggplot2)
```

Loading required package: ggplot2

```
ggplot(Data_Q1, aes(x = DomSpp, y = MassChange, fill = DomSpp)) +
  geom_boxplot() +
  stat_summary(
    fun = "mean", geom = "errorbar",
    aes(ymax = ..y.., ymin = ..y..),
    width = 0.75, color = "red", size = 1.2
  ) + geom_text(x = "Pollock", y = 20, label = "*",
    aes(x = x, y = y, label = label),
    color = "black", size = 6, fontface = "bold"
  ) +
  labs(
    title = "Figure 1. Median (black line) and mean (red line) mass change by dominant prey species",
    x = "Prey species",
    y = "Mass change (kg)", fill = "Dominant prey species"
  ) +
  theme_minimal() + theme(
    axis.line = element_line(color = "black", size = 1), # Add axis lines
    axis.text.x = element_text(size = 12, angle = 45, hjust = 1)
  ) +
```

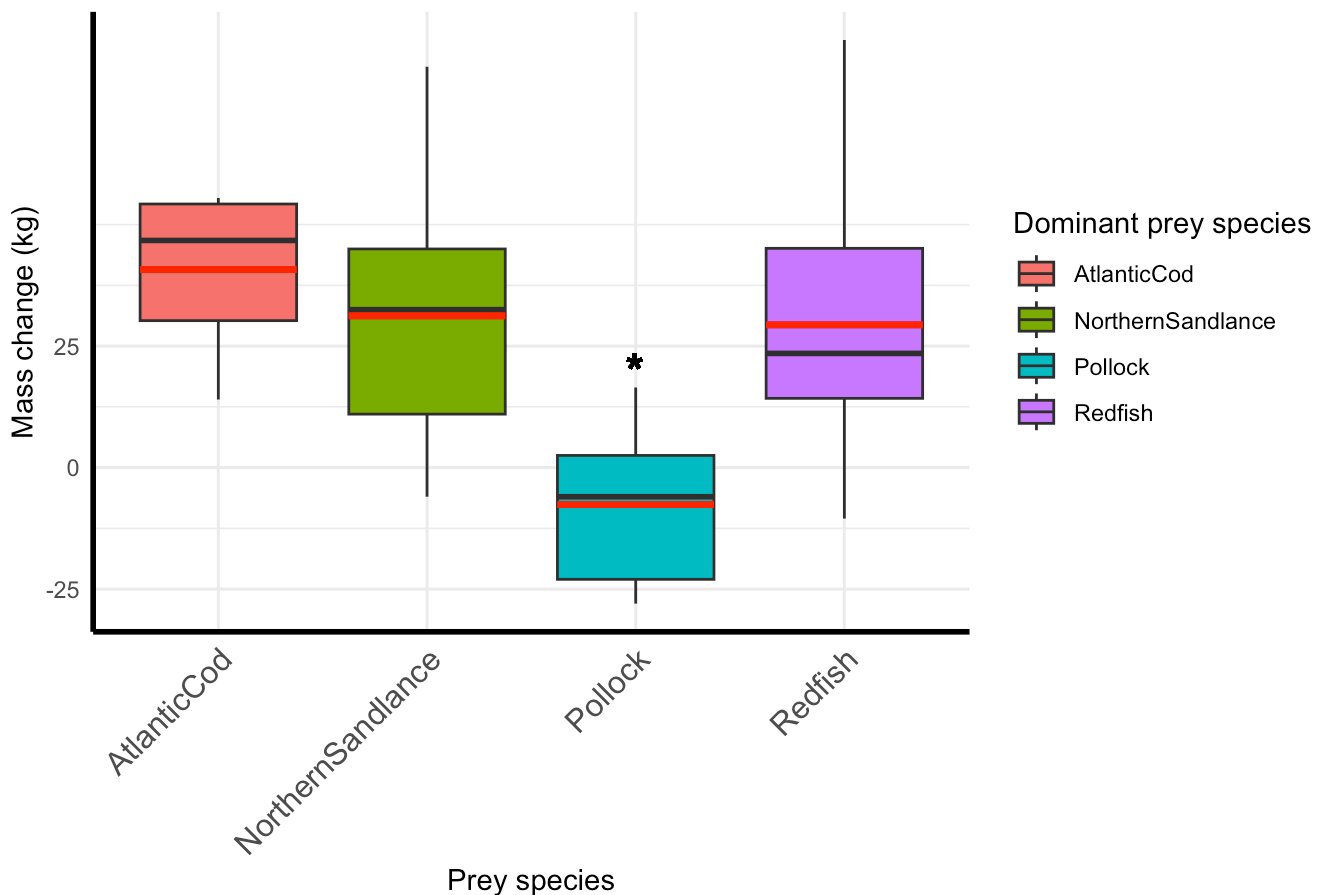
```
scale_y_continuous(
  breaks = seq(0, max(Data_Q1$MassChange), by = 25) - 50
)
```

Warning: Using `size` aesthetic for lines was deprecated in ggplot2 3.4.0.  
i Please use `linewidth` instead.

Warning: The `size` argument of `element\_line()` is deprecated as of ggplot2 3.4.0.  
i Please use the `linewidth` argument instead.

Warning: The dot-dot notation (`..y..`) was deprecated in ggplot2 3.4.0.  
i Please use `after\_stat(y)` instead.

Figure 1. Median (black line) and mean (red line) mass change by dominant prey



It may also be worth noting that diet diversity is almost significant at  $p = 0.0556$ . While this p-value is not significant, the negative relationship between diet diversity and mass change may support the observed effects of dominant prey species, where individuals who primarily prey on cod, sand lance, or redfish and have a lower diet diversity due to this focus on a few profitable species are more successful in gaining mass than individuals who feed on a more diverse array of prey species which include less profitable species.

We can perform cross-validation to test this model's predictions. In order to have each fold represent about 10% of the data, we will perform 7 fold cross validation:



```

set.seed(123)

#Number of folds
k <- 7

# Split the data into folds
DatCV <- Data_Q1 %>% mutate(fold = sample(1:k, length.out = n()))

# Create a vector to store RMSE for each fold
rmse_values <- numeric(k)

#Perform cross validation
for (i in 1:k) {
  train_data <- DatCV %>% filter(fold != i)
  test_data <- DatCV %>% filter(fold == i)

  # Fit the model on training data
  modCV <- glm(MassChange ~ DietDiv + Year + DomSpp, family = gaussian(link = "identity"),

  # Predict on the test data
  predictions <- predict(modCV, newdata = test_data)

  # Compute RMSE for this fold
  actuals <- test_data$MassChange
  rmse_values[i] <- sqrt(mean((predictions - actuals)^2))
}

# Summarize
rmse_values

```

```
[1] 22.44760 11.88442 18.05650 11.05129 18.98355 15.02906 14.37965
```

```

mean_rmse <- mean(rmse_values)
cat("Average RMSE across", k, "folds:", round(mean_rmse, 2), "\n")

```

Average RMSE across 7 folds: 15.98

The average root mean squared error is 15.98, indicating the model generally predicts values within 16 kg of the true value. Check the standard deviation to compare the average RMSE to:

```
sd(Data_Q1$MassChange)
```

```
[1] 26.15433
```

Compared to the standard deviation of mass change, the RMSE value is about 40% lower than the standard deviation, meaning the model predictions are much better than using the population mean to predict mass change. However, compared to the overall mean mass change within our dataset (28.5

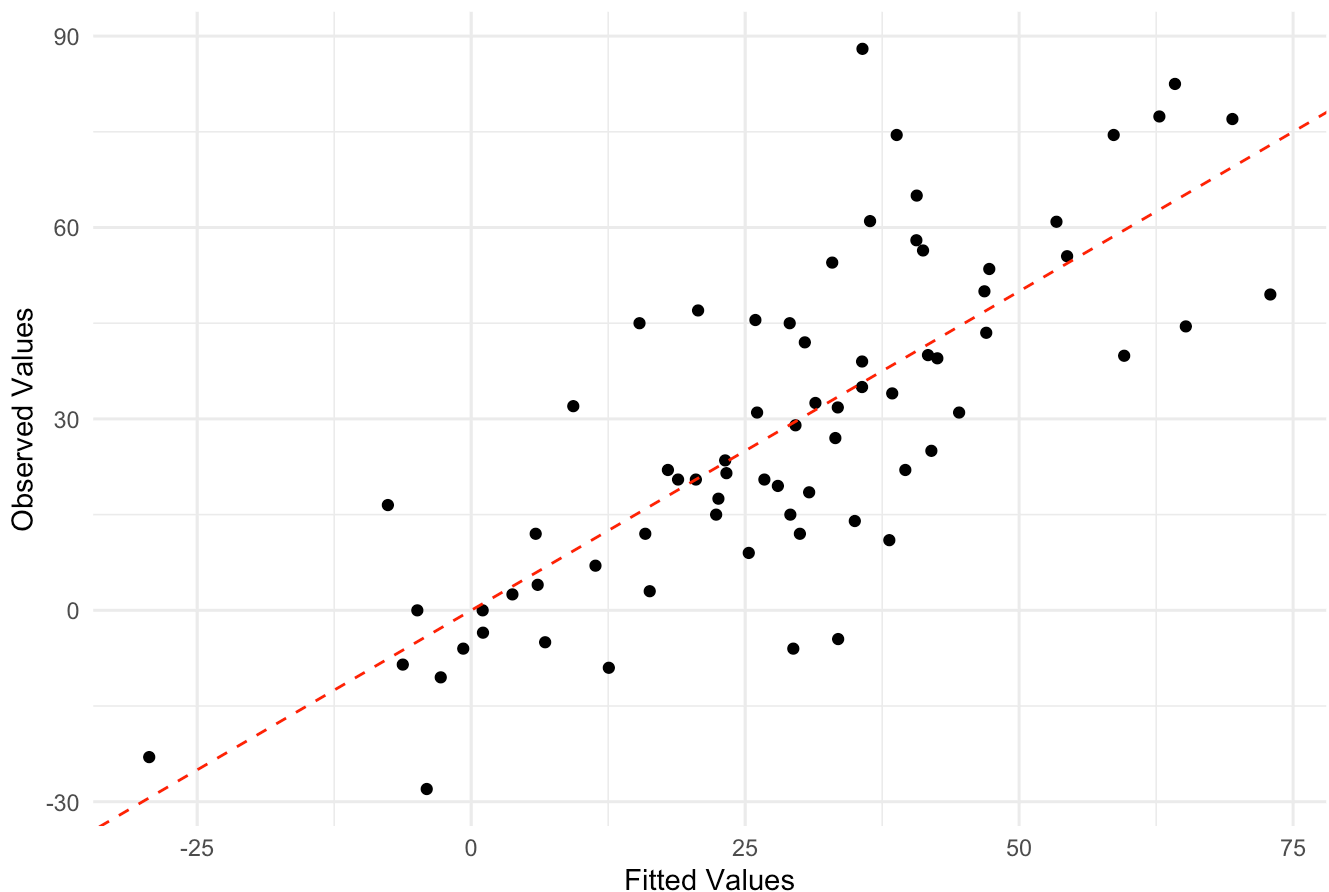
kg), there is relatively large prediction error. We can also visualize prediction accuracy with a fitted vs observed value plot (figure 2):

```
fit_mom <- fitted(mom_mod2)

# Get the observed response variable (actual values)
obs_mom <- Data_Q1$MassChange

# Plot the fitted vs. observed values
ggplot(data = Data_Q1, aes(x = fit_mom, y = obs_mom)) +
  geom_point(color = "black") +
  geom_abline(slope = 1, intercept = 0, color = "red", linetype = "dashed") + # Add 1:1
  labs(
    title = "Figure 2. Fitted vs. Observed Values with 1:1 line",
    x = "Fitted Values",
    y = "Observed Values"
  ) +
  theme_minimal()
```

Figure 2. Fitted vs. Observed Values with 1:1 line



While figure 2 shows that the model predictions have consistent accuracy across the range of fitted values, as suggested by the average RMSE value obtained from cross-validation, there is reasonably large prediction error. Given that the model residuals suggest good model fit, the relatively large prediction error suggest the set of covariates used in this model are not sufficient in predicting maternal

mass change. From this we can conclude that females who feed on redfish, sand lance, or cod experience greater mass change than individuals who feed on Pollock, but other factors outside of maternal diet, such as individual physiology or characteristics associated with individual foraging behavior, are responsible for explaining a large proportion of variation in maternal mass change.

## Research Question 2

### Does variation in diet of maternal female predict pup weaning mass in Northwest Atlantic grey seals?

#### Load Data

`require(tinytex)` `quarto install tinytex`

Cleaned data file was uploaded. `seal_data = read.csv("Data_Q2.csv", header = T)`

```
seal_data <- read.csv("/Users/peterbraithwaite/Desktop/IDPhD Classes/Stat 5620_Updated/Fi
summary(seal_data)
```

MomID	Year	Dietary.energy.density	Diet.diversity
Min. : 19	Min. :1996	Min. :5.207	Min. :0.1788
1st Qu.: 467	1st Qu.:2002	1st Qu.:5.676	1st Qu.:0.3108
Median : 4688	Median :2011	Median :5.784	Median :0.3564
Mean : 4999	Mean :2008	Mean :5.772	Mean :0.3622
3rd Qu.: 9412	3rd Qu.:2013	3rd Qu.:5.900	3rd Qu.:0.4181
Max. :10690	Max. :2015	Max. :6.517	Max. :0.5290
Dominant.prey.species	Mom.Age	Pup.sex	Pup.Weane.Mass
Length:56	Min. : 9.00	Min. :1.000	Min. :34.50
Class :character	1st Qu.:23.00	1st Qu.:1.000	1st Qu.:45.50
Mode :character	Median :25.00	Median :1.000	Median :50.75
	Mean :23.46	Mean :1.464	Mean :50.31
	3rd Qu.:26.25	3rd Qu.:2.000	3rd Qu.:55.12
	Max. :31.00	Max. :2.000	Max. :63.00

#### Data Exploration

The package "flexplot" was used to explore the predictor and response variables.

```
require(flexplot)
```

Loading required package: flexplot

Attaching package: 'flexplot'

The following object is masked from 'package:ggplot2':

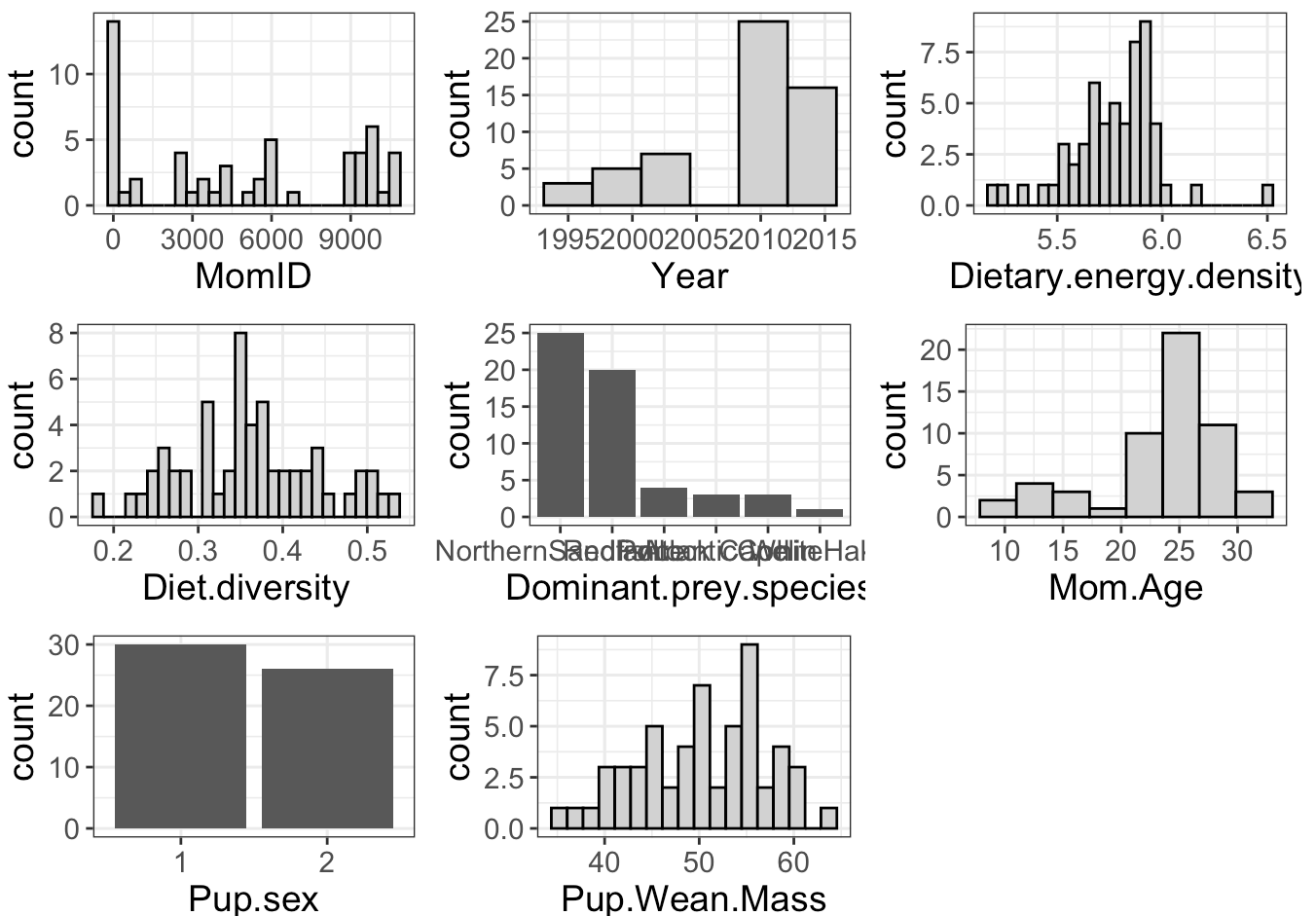
flip\_data

```
a = flexplot (MomID~1, data= seal_data)
b = flexplot (Year~1, data= seal_data)
c = flexplot (Dietary.energy.density~1, data= seal_data)
d = flexplot (Diet.diversity~1, data= seal_data)
e = flexplot (Dominant.prey.species~1, data= seal_data)
f = flexplot (Mom.Age~1, data= seal_data)
g = flexplot (Pup.sex~1, data= seal_data)
h = flexplot (Pup.Weann.Mass~1, data= seal_data)

require (cowplot)
```

Loading required package: cowplot

```
plot_grid(a,b,c,d,e,f,g,h)
```



## Figure 1 - Exploring Variable Distributions

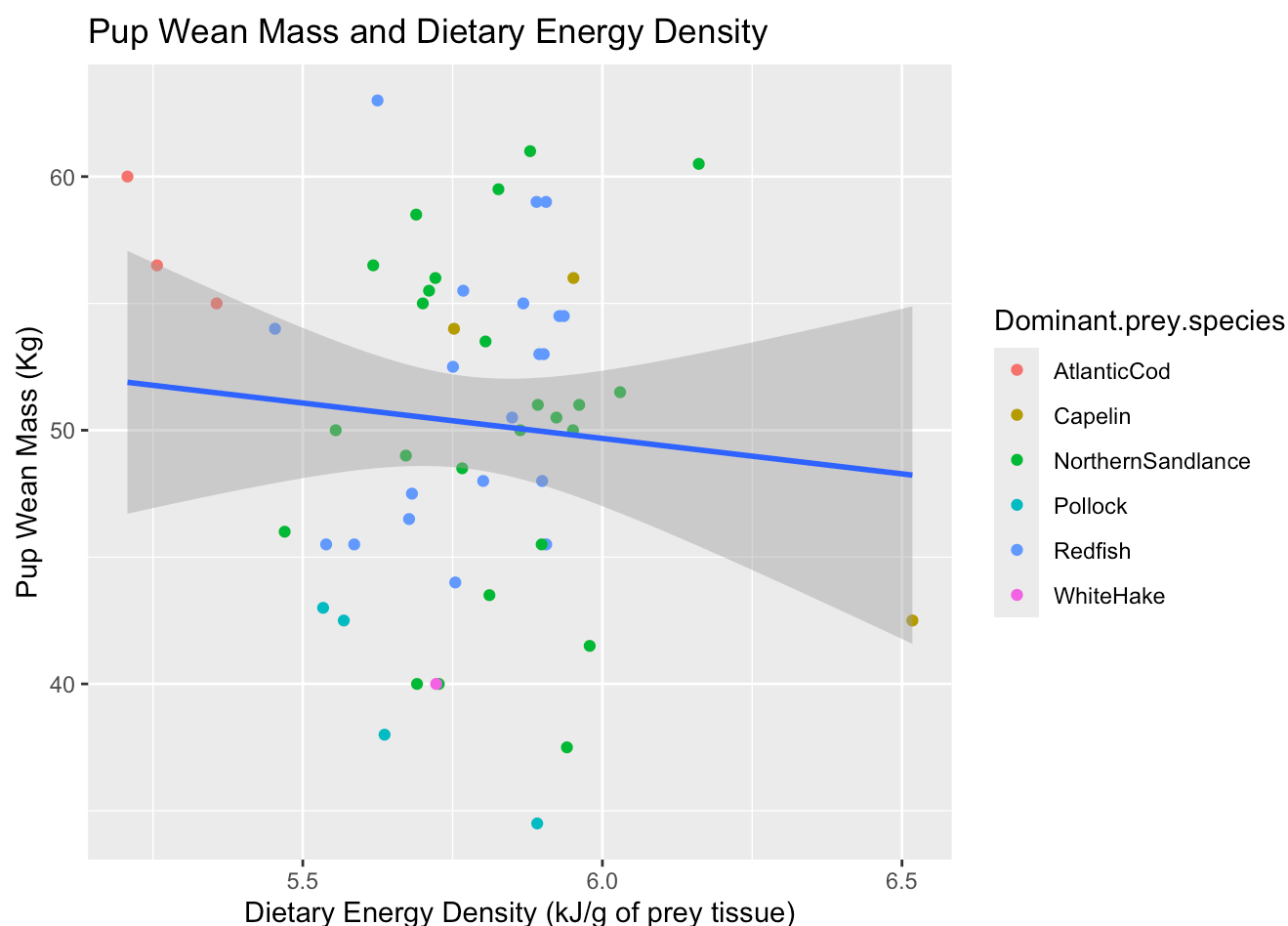
From the plots above it seems that the response variable (Pup Wean Mass) has a Gaussian distribution.

To further explore the relationship between variable the package "ggplot2" was used plot individual variables in relation to the response variables.

## Pup Wean Mass and Dietary Engery Density

```
require(ggplot2)
# Pup Wean Mass + Dietary.energy.density
ggplot(seal_data) + geom_point(aes(Dietary.energy.density, Pup.Wean.Mass, color = Dominan
```

```
`geom_smooth()` using formula = 'y ~ x'
```



## Figure 2 - Pup Wean Mass and Dietary Energy Density

From this plot we can see a negative or inverse relationship between Pup Wean Mass and Dietary Energy Density. This indicates that maternal mothers that feed on prey species resulting in greater dietary energy density tend to have pups with a lower wean mass. This is somewhat at odds with what might have been expected. We might have expected the greater the dietary energy density of the materanl mother would result in heavier and ultimately more fit pups.

We can also see that the Dominant Species Capelin has only three observation and one of these appears to be an outlier. It should also be noted that White Hake has only one observation and Atlantic Cod has only three. This lack of observations may be problematic while fitting a model and we may need to remove these Dominant Prey Species.

## Pup Wean Mass and Dietary Engery Density

```
# Pup Wean Mass + Diet Diveristy
ggplot(seal_data) + geom_point(aes(Diet.diversity, Pup.Wean.Mass, color = Dominant.prey.s
  labs(title = "Pup Wean Mass and Diet Diversity", x = "Diet Diveristy (no specific units
  geom_smooth(aes(Diet.diversity, Pup.Wean.Mass), method="lm", se=T)
```

`geom\_smooth()` using formula = 'y ~ x'

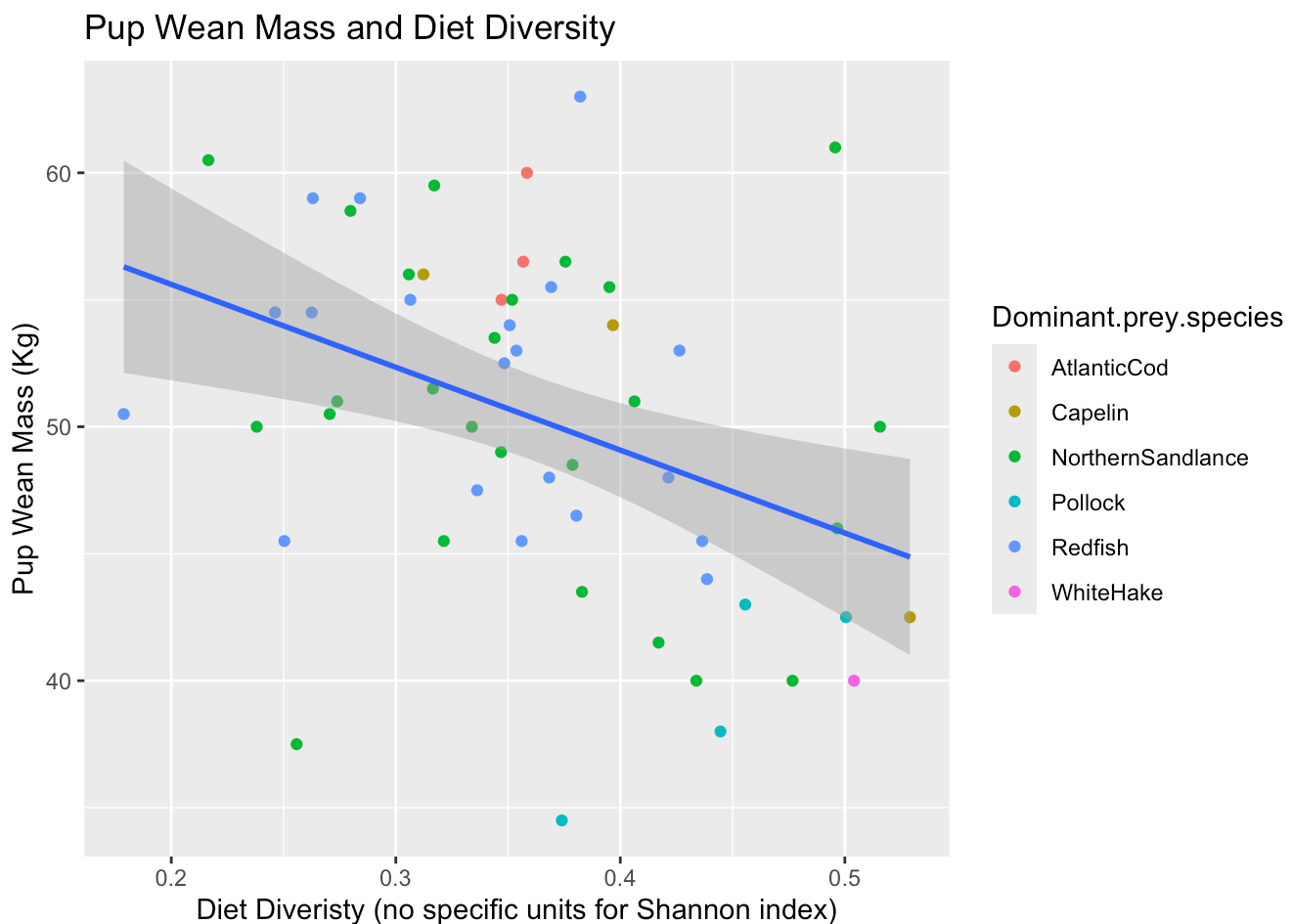


Figure 3 - Pup Wean Mass and Diet Diversity

From this plot we can see a negative or inverse relationship between Pup Wean Mass and Diet Diversity. This indicates that as the maternal mothers feed on more diverse prey species their pup wean mass decreases. This may indicates that mothers that reduce the diversity of their diet have healthier and more fit pups.

## Pup Wean Mass + Year

Exploring if specific years showed abnormally high or low Pup Wean Mass and if these were related to specific dominant prey species.

```
ggplot(seal_data) + geom_point(aes(Year, Pup.Wean.Mass, color = Dominant.prey.species)) +  
  labs(title = "Pup Wean Mass and Year", x = "Year", y = "Pup Wean Mass (Kg)") +  
  geom_smooth(aes(Year, Pup.Wean.Mass), method="lm", se=T)
```

`geom\_smooth()` using formula = 'y ~ x'

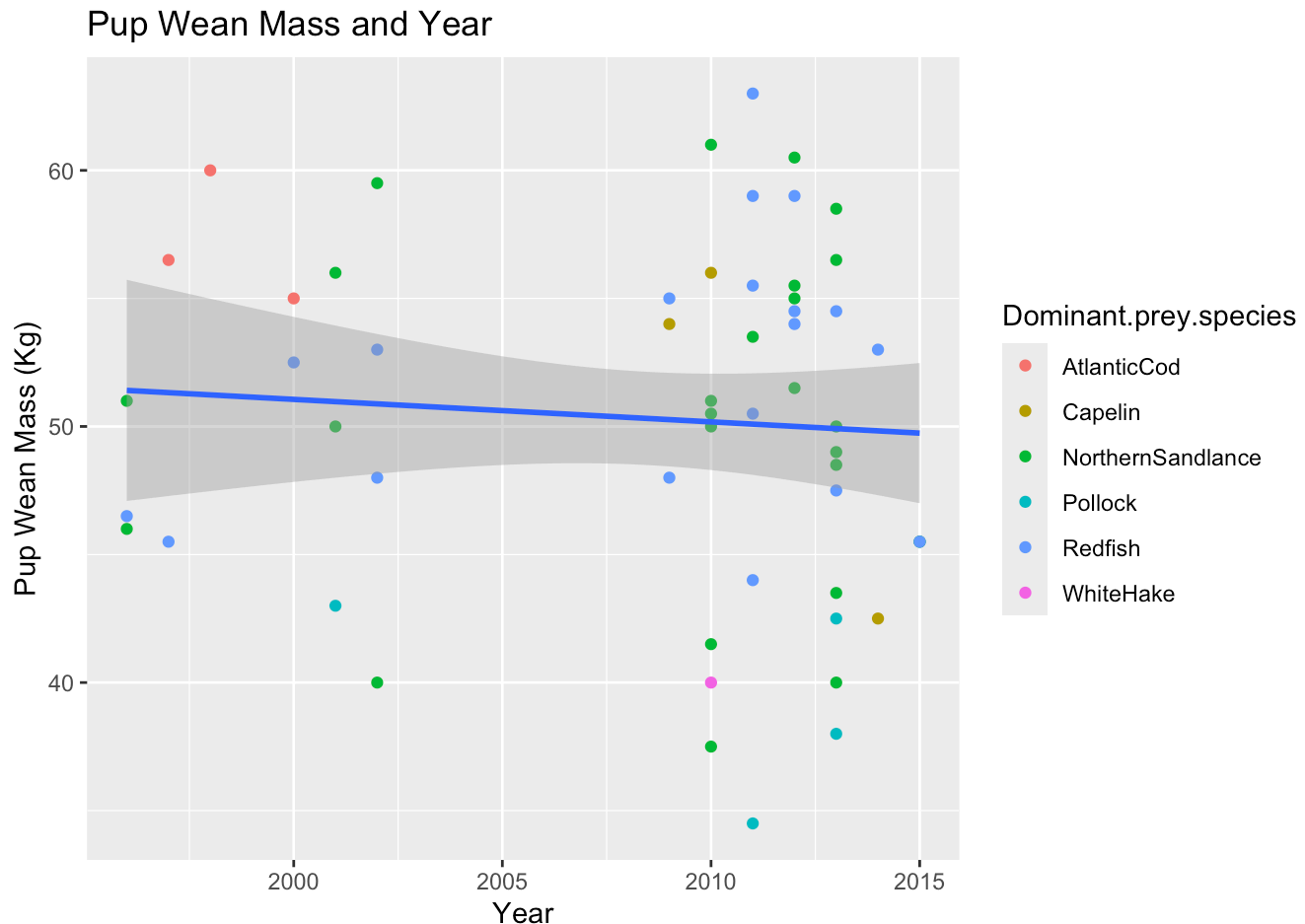


Figure 4 - Pup Wean Mass and Year

From this plot it seems Pup Wean Mass seems to be fairly evenly distributed over the observed years and it does not seem that a specific Dominant Prey Species had a disproportionate effect on the Pup Wean Mass.

## Pup Wean Mass + MomID

Exploring if Pup Wean Mass shows significant differences based on the maternal mother ID.



```
ggplot(seal_data) + geom_point(aes(MomID, Pup.Wean.Mass, color = Year )) +
  labs(title = "Pup Wean Mass and Mom ID", x = "Mom ID", y = "Pup Wean Mass (Kg)") +
  geom_smooth(aes(MomID, Pup.Wean.Mass), method="lm", se=T)
```

`geom\_smooth()` using formula = 'y ~ x'

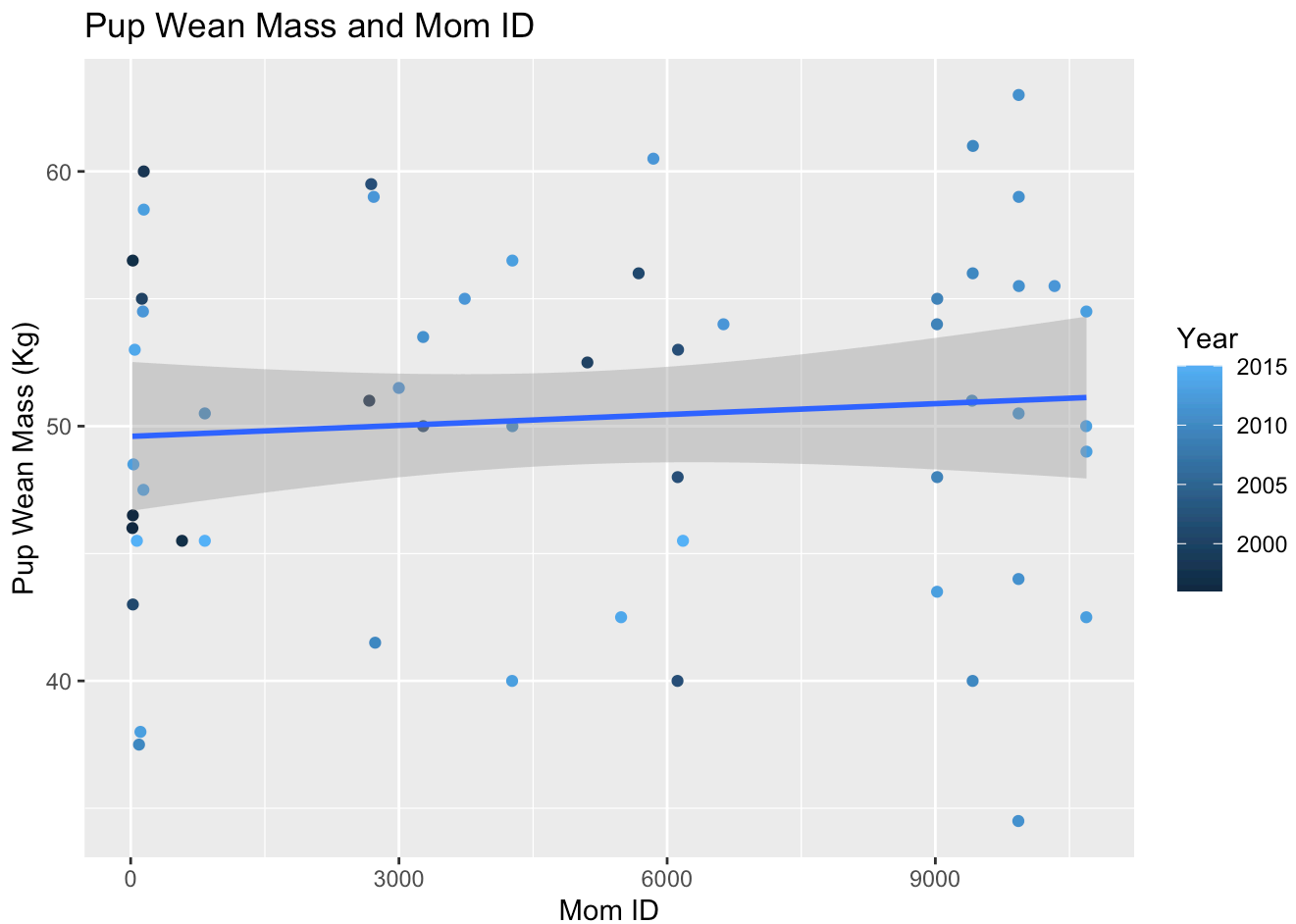
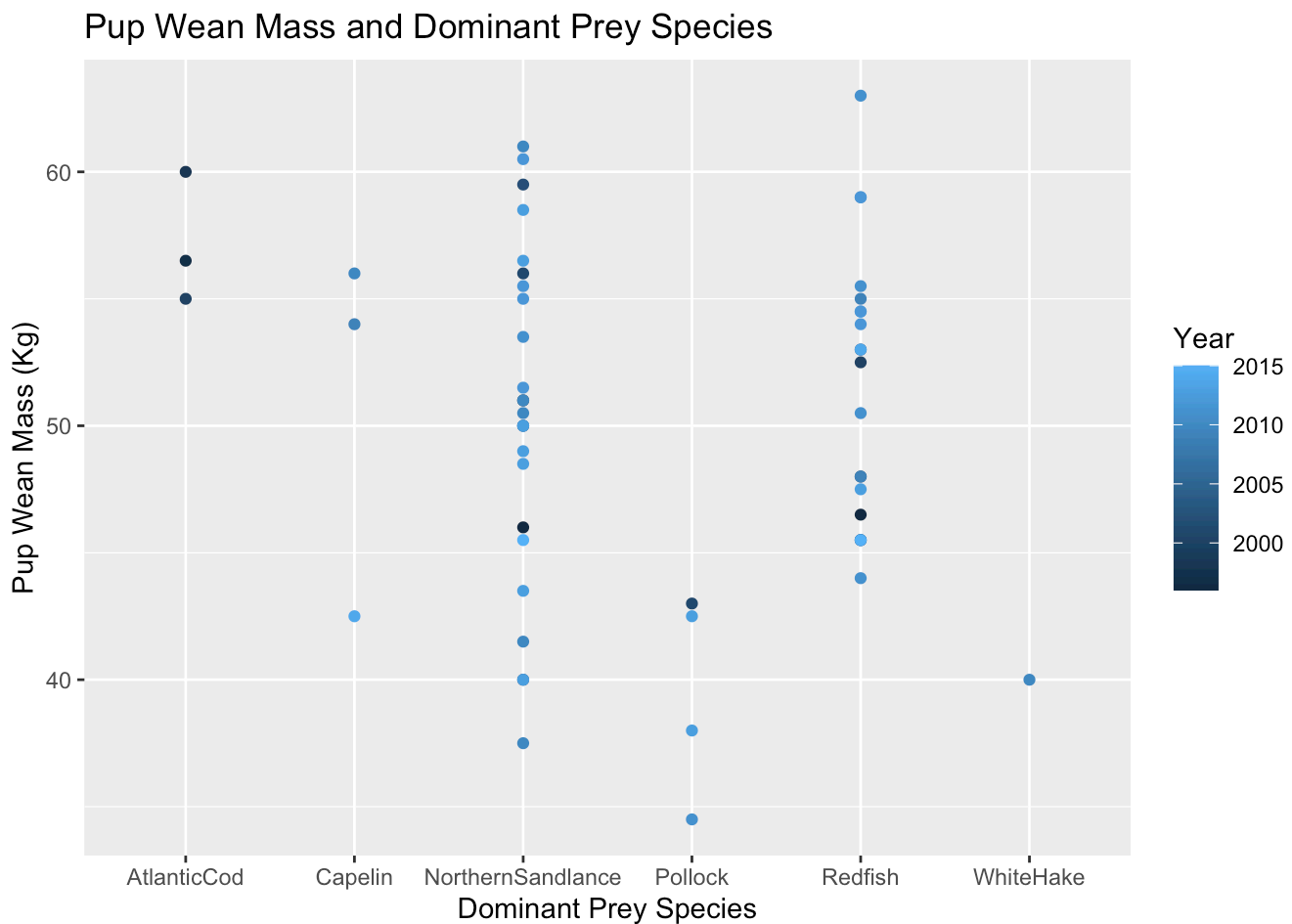


Figure 5 - Pup Wean Mass and Material Mother ID

With the vertical rows of the this plot representing unique maternal mothers, it does not seem any individual is creating disproportionately large or small pup weights. Rather, it seems most maternal mother are creating a fairly evenly distribution of pup weights dependent on the year.

## Pup Wean Mass + Dominant Prey Species

```
ggplot(seal_data) + geom_point(aes(Dominant.prey.species, Pup.Wean.Mass, color = Year ))
```



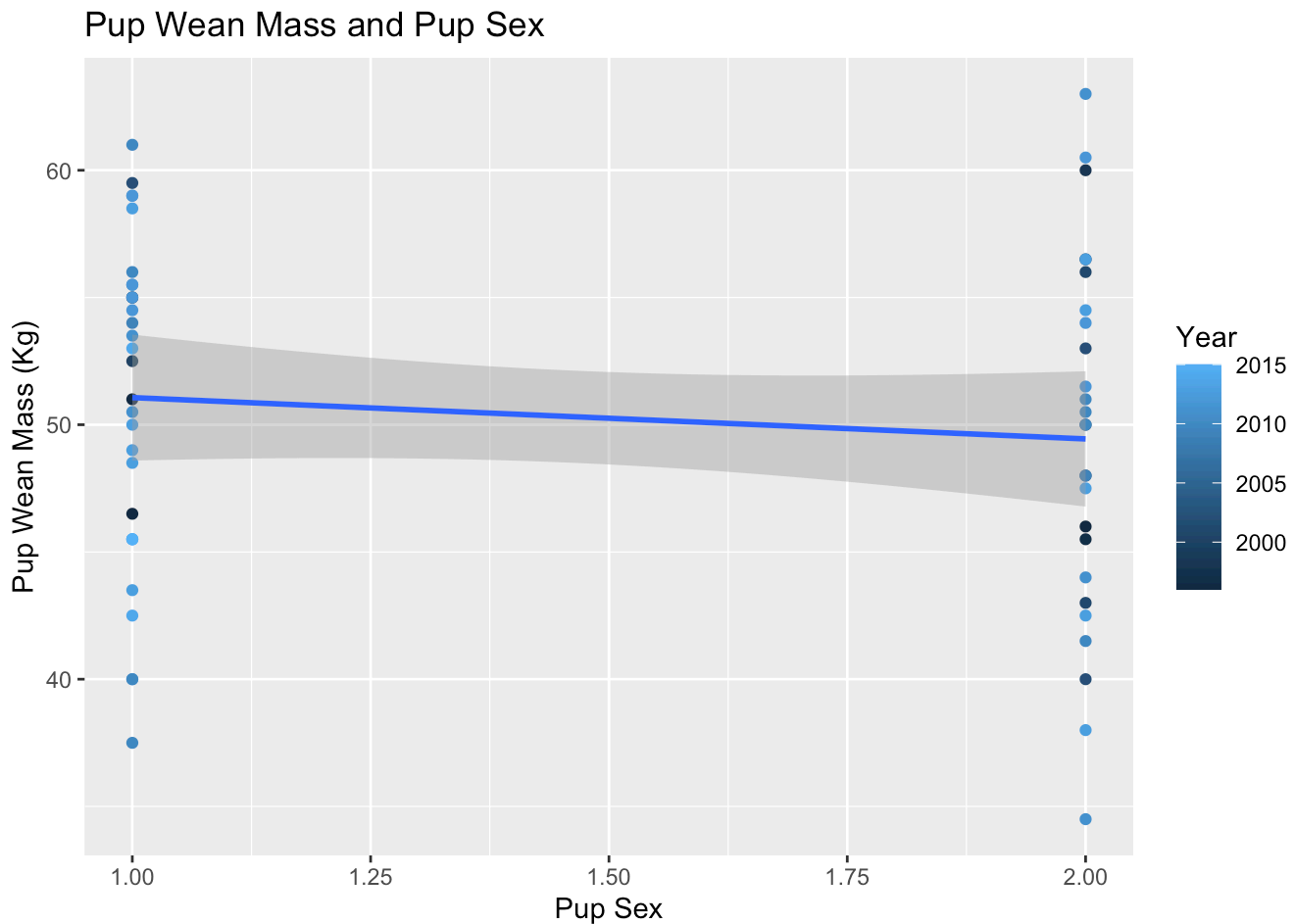
**Figure 6 - Pup Wean Mass and Dominant Prey Species**

A considerable amount of insight can be gained from this plot including: The Dominant Prey Species with the largest number of observations (Northern Sandlance and Redfish) show a variety of Pup Wean Masses and these occur over a wide variety of years; the Dominant Prey Species with the smallest number of observations tend to show either predominantly large (Atlantic Cod and Capelin) or very small (Pollock and White Hake) Pup Wean sizes and these tend to occur in a variety of years.

## Pup Wean Mass + Pup Sex

```
ggplot(seal_data) + geom_point(aes(Pup.sex, Pup.Weans.Mass, color = Year )) +
  labs(title = "Pup Wean Mass and Pup Sex", x = "Pup Sex", y = "Pup Wean Mass (Kg)") +
  geom_smooth(aes(Pup.sex, Pup.Weans.Mass), method="lm", se=T)
```

`geom\_smooth()` using formula = 'y ~ x'



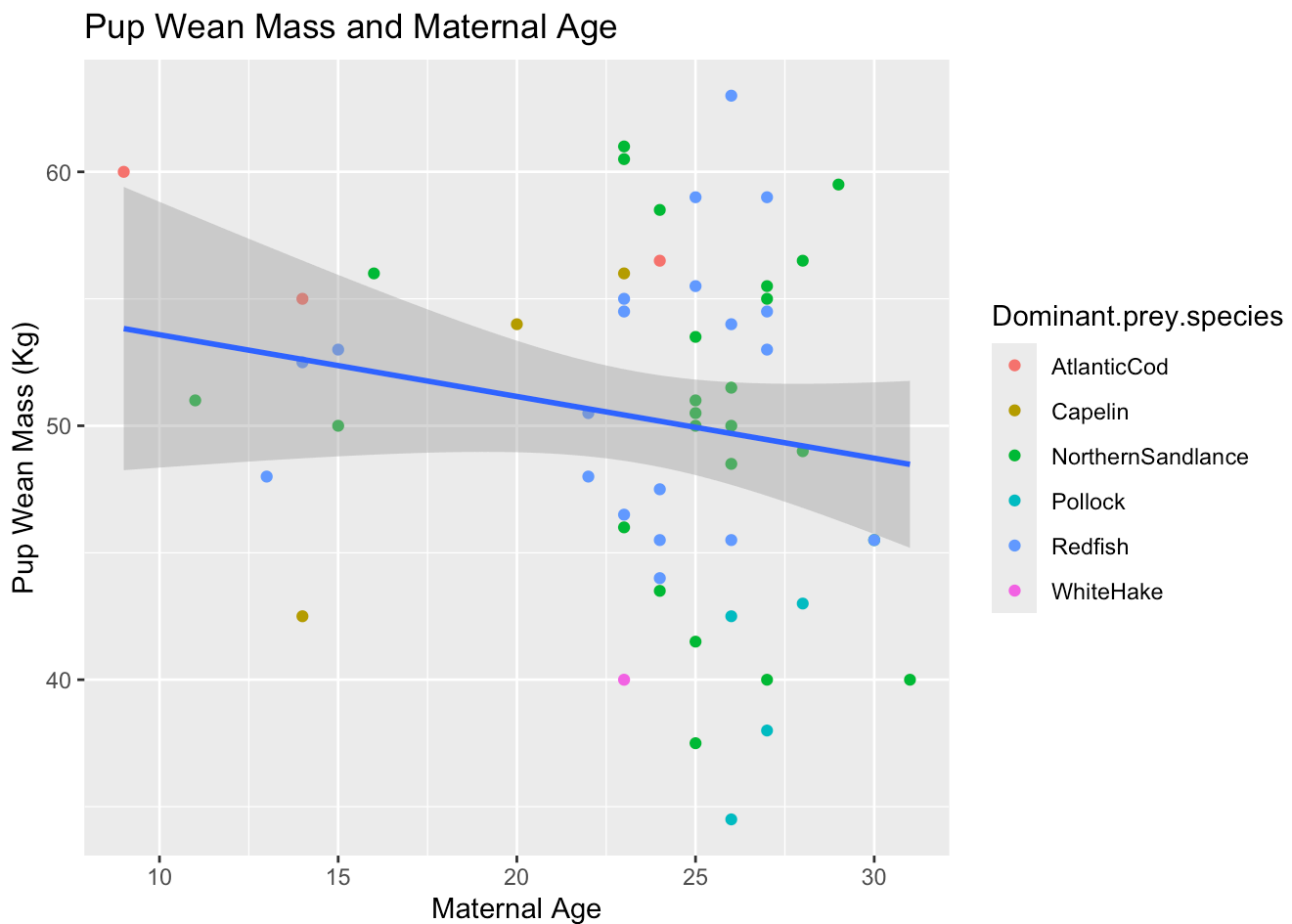
**Figure 7 - Pup Wean Mass and Pup Sex**

This plot shows that the Pup Sex seems to have little effect on the average Pup Mean Mass. Both Pup Sex seem to create a fairly even distribution of Pup Mean Mass over a variety of years. It is expected that this variable will have a small effect on the fit model.

## Pup Wean Mass + Maternal Age

```
ggplot(seal_data) + geom_point(aes(Mom.Age, Pup.Wean.Mass, colour = Dominant.prey.species)) +
  labs(title = "Pup Wean Mass and Maternal Age", x = "Maternal Age", y = "Pup Wean Mass (Kg)")
```

`geom\_smooth()` using formula = 'y ~ x'



**Figure 7 - Pup Wean Mass and Maternal Age**

A considerable amount of insight can be gained from this plot. It seems younger maternal seals have larger pups but it is also noteworthy that the majority of older mothers prey on the most observed Dominant Prey Species including Redfish and Northern Sandlance. Alternatively it seems less frequent observed Dominant Prey Species were preyed on by young maternal mothers. It seems as though there might be a social aspects to why younger mothers prey on less frequently observed Dominant Prey Species.

## Fit Linear Model

We now explore fitting general linear models to further explore the variables, the residuals and their relationships.

First we convert the catagorical data to factors and determine the number of levels within each catagorical group.

```
# convert catagorical data to factors

seal_data$MomID = as.factor(seal_data$MomID)
seal_data$Year = as.factor(seal_data$Year)
```

```
seal_data$Dominant.prey.species = as.factor(seal_data$Dominant.prey.species)
seal_data$Pup.sex = as.factor(seal_data$Pup.sex)
seal_data$Pup.Mom.Age = as.factor(seal_data$Mom.Age)

nlevels(seal_data$MomID)
```

[1] 51

```
# 51
nlevels(seal_data$Year)
```

[1] 13

```
# 13
nlevels(seal_data$Dominant.prey.species)
```

[1] 6

```
# 5
nlevels(seal_data$Pup.sex)
```

[1] 2

```
# 2
nlevels(seal_data$Pup.Mom.Age)
```

[1] 17

```
# 17
```

## Fit Linear Model

---

### All Predictor Variables

```
Q1 = lm(Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.density
        + Year + Mom.Age + Pup.sex + MomID, data = seal_data)
summary(Q1)
```

Call:

```
lm(formula = Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity +
    Dietary.energy.density + Year + Mom.Age + Pup.sex + MomID,
    data = seal_data)
```

Residuals:

ALL 56 residuals are 0: no residual degrees of freedom!

Coefficients: (16 not defined because of singularities)

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-127.5808	NaN	NaN	NaN
Dominant.prey.speciesCapelin	14.0800	NaN	NaN	NaN
Dominant.prey.speciesNorthernSandlance	2.5039	NaN	NaN	NaN
Dominant.prey.speciesPollock	-19.7677	NaN	NaN	NaN
Dominant.prey.speciesRedfish	-16.6515	NaN	NaN	NaN
Dominant.prey.speciesWhiteHake	4.6941	NaN	NaN	NaN
Diet.diversity	8.3424	NaN	NaN	NaN
Dietary.energy.density	35.8485	NaN	NaN	NaN
Year1997	19.2395	NaN	NaN	NaN
Year1998	-1.3181	NaN	NaN	NaN
Year2000	14.1916	NaN	NaN	NaN
Year2001	-0.8124	NaN	NaN	NaN
Year2002	-3.0383	NaN	NaN	NaN
Year2009	29.8803	NaN	NaN	NaN
Year2010	-2.9017	NaN	NaN	NaN
Year2011	37.2376	NaN	NaN	NaN
Year2012	23.6996	NaN	NaN	NaN
Year2013	20.8844	NaN	NaN	NaN
Year2014	-55.5366	NaN	NaN	NaN
Year2015	16.4409	NaN	NaN	NaN
Mom.Age	-1.8940	NaN	NaN	NaN
Pup.sex2	14.4153	NaN	NaN	NaN
MomID23	4.4626	NaN	NaN	NaN
MomID24	27.5950	NaN	NaN	NaN
MomID30	-7.9254	NaN	NaN	NaN
MomID45	88.7703	NaN	NaN	NaN
MomID69	26.2285	NaN	NaN	NaN
MomID93	-2.2734	NaN	NaN	NaN
MomID109	-4.5789	NaN	NaN	NaN
MomID125	NA	NA	NA	NA
MomID137	3.6242	NaN	NaN	NaN
MomID142	-4.6121	NaN	NaN	NaN
MomID146	1.8643	NaN	NaN	NaN
MomID574	NA	NA	NA	NA
MomID829	-3.1820	NaN	NaN	NaN
MomID2668	-19.0778	NaN	NaN	NaN
MomID2690	31.0252	NaN	NaN	NaN
MomID2718	16.5836	NaN	NaN	NaN
MomID2734	-15.3957	NaN	NaN	NaN
MomID2999	-31.0853	NaN	NaN	NaN
MomID3271	-22.2778	NaN	NaN	NaN
MomID3736	0.2347	NaN	NaN	NaN
MomID4266	-12.6464	NaN	NaN	NaN
MomID4269	-5.2029	NaN	NaN	NaN
MomID5108	NA	NA	NA	NA
MomID5485	NA	NA	NA	NA
MomID5681	-9.8719	NaN	NaN	NaN

MomID5846	-31.6400	NaN	NaN	NaN
MomID6116	3.4943	NaN	NaN	NaN
MomID6118	-9.0816	NaN	NaN	NaN
MomID6122	NA	NA	NA	NA
MomID6177	8.7215	NaN	NaN	NaN
MomID6630	10.9350	NaN	NaN	NaN
MomID9018	-34.0206	NaN	NaN	NaN
MomID9019	-18.3770	NaN	NaN	NaN
MomID9020	-21.8694	NaN	NaN	NaN
MomID9021	NA	NA	NA	NA
MomID9410	-2.7034	NaN	NaN	NaN
MomID9417	NA	NA	NA	NA
MomID9418	NA	NA	NA	NA
MomID9420	17.6372	NaN	NaN	NaN
MomID9928	-34.8701	NaN	NaN	NaN
MomID9930	-27.9118	NaN	NaN	NaN
MomID9931	-12.0164	NaN	NaN	NaN
MomID9932	NA	NA	NA	NA
MomID9933	NA	NA	NA	NA
MomID9934	NA	NA	NA	NA
MomID10333	NA	NA	NA	NA
MomID10687	NA	NA	NA	NA
MomID10688	NA	NA	NA	NA
MomID10689	NA	NA	NA	NA
MomID10690	NA	NA	NA	NA

Residual standard error: NaN on 0 degrees of freedom

Multiple R-squared: 1, Adjusted R-squared: NaN

F-statistic: NaN on 55 and 0 DF, p-value: NA

When this model is run we received an error suggesting we have more columns than rows. This suggest we have more variables than observations. As a result we must remove some of the categorical variables. The maternal mother ID seems to be adding a lot of categories (n -51) and did not seem very important in Figure 6 above. As a result Mom ID was removed and the model was fit again.

## All Predictor Variables Except Mom ID

```
Q2 = lm(Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.density +
summary
```

standardGeneric for "summary" defined from package "base"

```
function (object, ...)
standardGeneric("summary")
<environment: 0x1249eff90>
Methods may be defined for arguments: object
Use showMethods(summary) for currently available ones.
```

```
# Adjusted R-squared: 0.209
```



```
# p-value: 0.08393
```

Although this model was not significant, it did provide a model that works as a starting point for the Step Function.

## Step Function (Forward)

We now used the Step Function to determine the best fit model using AIC.

```
fwd.model = step (Q2, direction='forward')
```

Start: AIC=216.76

Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.density +  
Year + Mom.Age + Pup.sex

The forward step retained all variables with AIC = 216.76

## Step Function (Backward)

```
backward.model = step(Q2, direction='backward')
```

Start: AIC=216.76

Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.density +  
Year + Mom.Age + Pup.sex

	Df	Sum of Sq	RSS	AIC
- Year	12	347.70	1572.2	206.75
- Mom.Age	1	1.61	1226.1	214.83
- Pup.sex	1	17.14	1241.6	215.53
<none>			1224.5	216.76
- Dietary.energy.density	1	47.89	1272.4	216.91
- Diet.diversity	1	69.12	1293.6	217.83
- Dominant.prey.species	5	329.04	1553.5	220.08

Step: AIC=206.75

Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.density +  
Mom.Age + Pup.sex

	Df	Sum of Sq	RSS	AIC
- Mom.Age	1	1.51	1573.7	204.81
- Pup.sex	1	3.13	1575.3	204.87
- Dietary.energy.density	1	50.68	1622.9	206.53
<none>			1572.2	206.75
- Dominant.prey.species	5	364.88	1937.1	208.44
- Diet.diversity	1	162.65	1734.8	210.27

Step: AIC=204.81

Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.density +  
Pup.sex

	Df	Sum of Sq	RSS	AIC
- Pup.sex	1	2.57	1576.3	202.90
- Dietary.energy.density	1	49.35	1623.0	204.54
<none>			1573.7	204.81
- Dominant.prey.species	5	415.52	1989.2	207.93
- Diet.diversity	1	164.37	1738.1	208.37

Step: AIC=202.9

Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.density

	Df	Sum of Sq	RSS	AIC
- Dietary.energy.density	1	49.16	1625.4	202.62
<none>			1576.3	202.90
- Diet.diversity	1	166.48	1742.8	206.52
- Dominant.prey.species	5	456.31	2032.6	207.14

Step: AIC=202.62

Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity

	Df	Sum of Sq	RSS	AIC
<none>			1625.4	202.62
- Diet.diversity	1	124.06	1749.5	204.74
- Dominant.prey.species	5	493.24	2118.7	207.46

The backward step retained only the Dominant Prey Species and Diet Diversity variables with AIC 202.62.

## Explore The Resulting Linear Model From Backward Step Function

```
Q1_Reduced = lm(Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity,
                 data = seal_data)
summary(Q1_Reduced)
```

Call:

```
lm(formula = Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity,
    data = seal_data)
```

Residuals:

Min	1Q	Median	3Q	Max
-15.0508	-4.1950	0.0968	3.2694	13.3653

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	64.422	5.013	12.850	< 2e-16 ***
Dominant.prey.speciesCapelin	-5.133	4.743	-1.082	0.284509
Dominant.prey.speciesNorthernSandlance	-6.630	3.519	-1.884	0.065539 .
Dominant.prey.speciesPollock	-15.831	4.500	-3.518	0.000949 ***
Dominant.prey.speciesRedfish	-5.797	3.570	-1.624	0.110838
Dominant.prey.speciesWhiteHake	-14.093	6.838	-2.061	0.044627 *

```
Diet.diversity                -20.491      10.596   -1.934  0.058916 .
```

```
----
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Residual standard error: 5.76 on 49 degrees of freedom
```

```
Multiple R-squared:  0.3509,    Adjusted R-squared:  0.2715
```

```
F-statistic: 4.416 on 6 and 49 DF,  p-value: 0.00121
```

```
# Adjusted R-squared: 0.2715
```

```
# p-value: 0.00121
```

This found that the reduced Linear Model is significant with a p-value = 0.00121 and an Adjusted R-Squared = 0.2715 meaning the model accounts for about 27% of the variation of the Pup Wean Mass. It should also be noted that the most significant variables are Pollock with a p-value of 0.000949 and White Hake with a p-value = 0.044627.

We will now look at the residuals for the this model.

## Explore Residuals for Reduced Linear Model

```
library(ggfortify)
autoplot(Q1_Reduced)
```

```
Warning: Removed 5 rows containing missing values or values outside the scale range
(`geom_line()`).
```

```
Warning: Removed 1 row containing missing values or values outside the scale range
(`geom_point()`).
```

```
Warning: Removed 1 row containing missing values or values outside the scale range
(`geom_line()`).
```

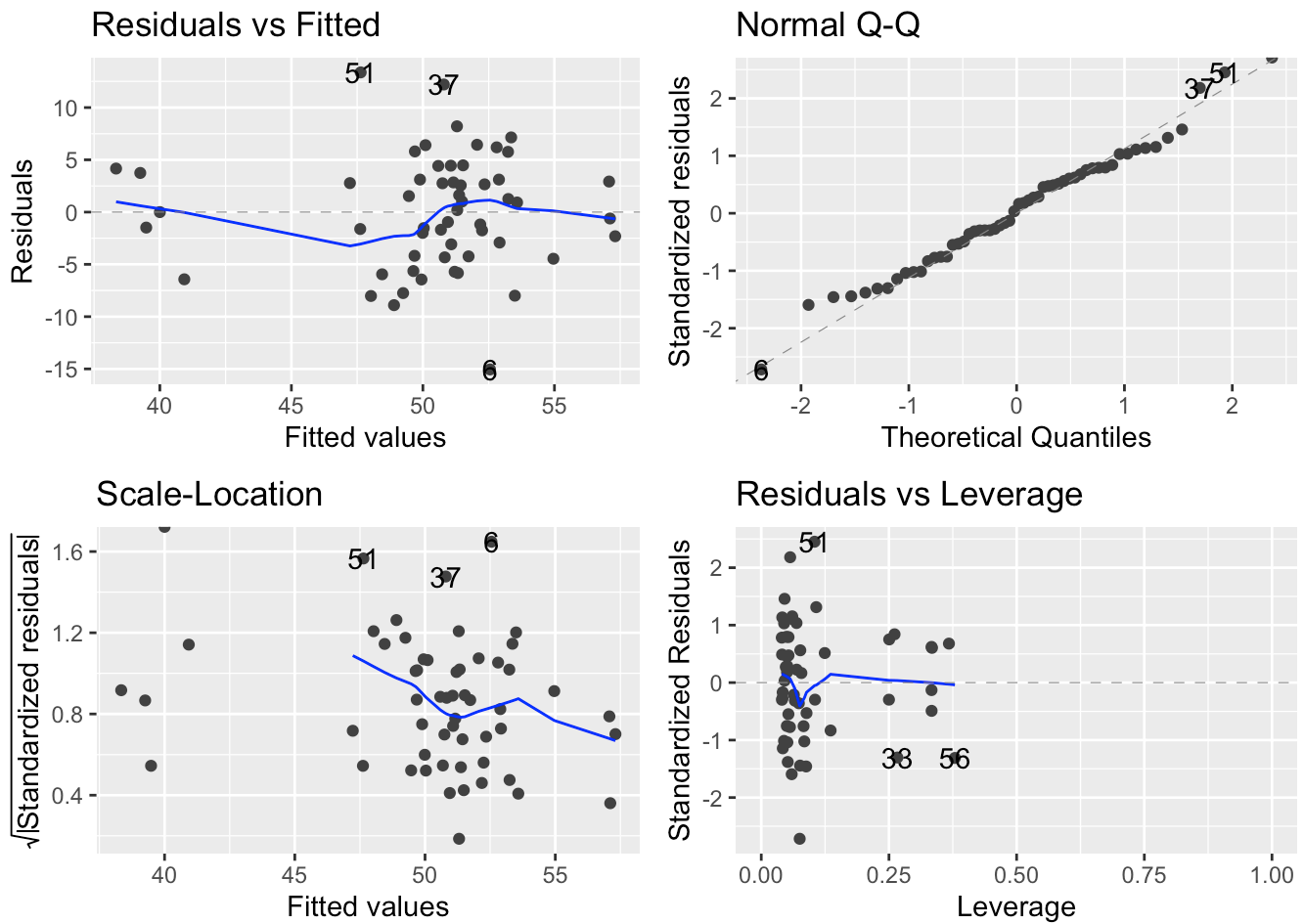


Figure 8 - Residual for Reduced Variables Linear Model

```
#residuals
res = resid(Q1_Reduced)
plot(fitted(Q1_Reduced), res)
abline(0,0)
```

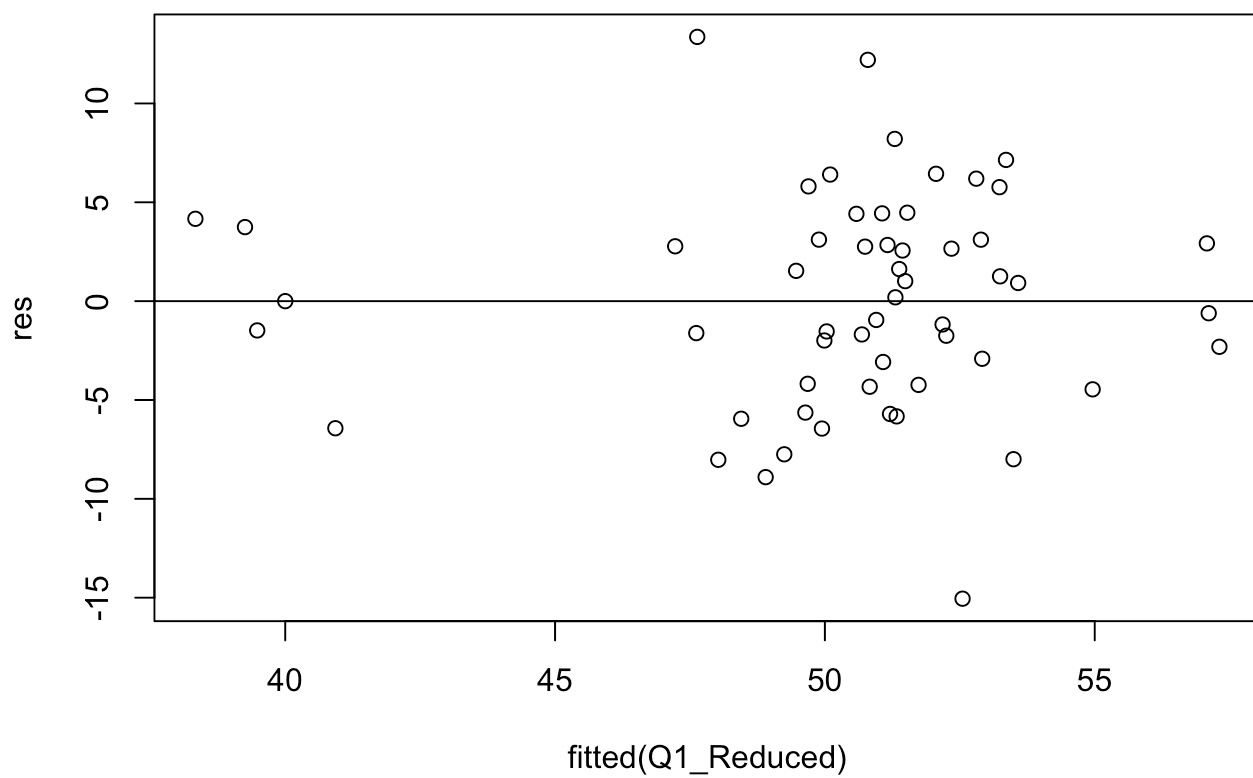
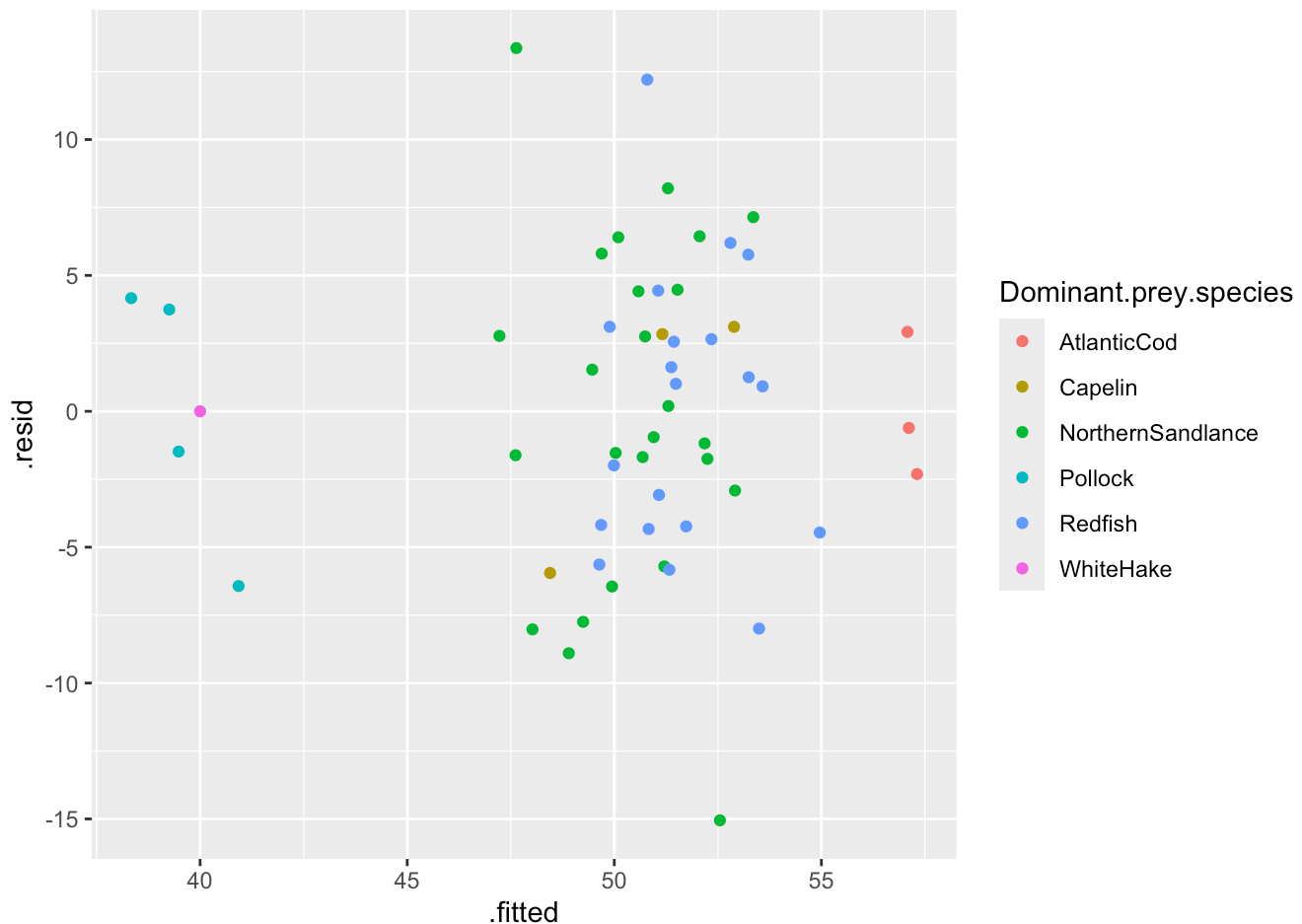


Figure 9 - Residual for Reduced Variables Linear Model

```
ggplot(Q1_Reduced, aes(x = .fitted, y = .resid, colour = Dominant.prey.species )) + geom_
```



**Figure 10 - Residual for Reduced Variables Linear Model**

From these residual visualization it seems clear that a number observations seem to be outliers and they should be removed or we should try robust regression strategies. We can see from Figure 10 that it is the observation with White Hake and Pollock as the Dominant Prey Species are clearly outliers.

We will now try to fit a robust regression first.

```
require(robustbase)
```

Loading required package: robustbase

```
Q1_Reduced_Robust = lmrob(Pup.Weann.Mass ~ Dominant.prey.species + Diet.diversity,
                           data = seal_data)
summary(Q1_Reduced_Robust)
```

Call:

```
lmrob(formula = Pup.Weann.Mass ~ Dominant.prey.species + Diet.diversity, data = seal_data)
  \--> method = "MM"
```

Residuals:

Min	1Q	Median	3Q	Max
-----	----	--------	----	-----

-15.74108 -3.76149 -0.08415 3.69887 13.97206

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	66.326	4.925	13.467	< 2e-16 ***
Dominant.prey.speciesCapelin	-4.698	2.646	-1.776	0.081995 .
Dominant.prey.speciesNorthernSandlance	-6.460	1.825	-3.540	0.000889 ***
Dominant.prey.speciesPollock	-15.208	3.194	-4.761	1.74e-05 ***
Dominant.prey.speciesRedfish	-6.031	1.719	-3.508	0.000977 ***
Dominant.prey.speciesWhiteHake	-13.271	2.405	-5.518	1.28e-06 ***
Diet.diversity	-25.898	13.446	-1.926	0.059903 .

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Robust residual standard error: 5.928

Multiple R-squared: 0.3699, Adjusted R-squared: 0.2928

Convergence in 12 IRWLS iterations

Robustness weights:

4 weights are ~1. The remaining 52 ones are summarized as

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
0.4607	0.9074	0.9590	0.9238	0.9847	0.9974

Algorithmic parameters:

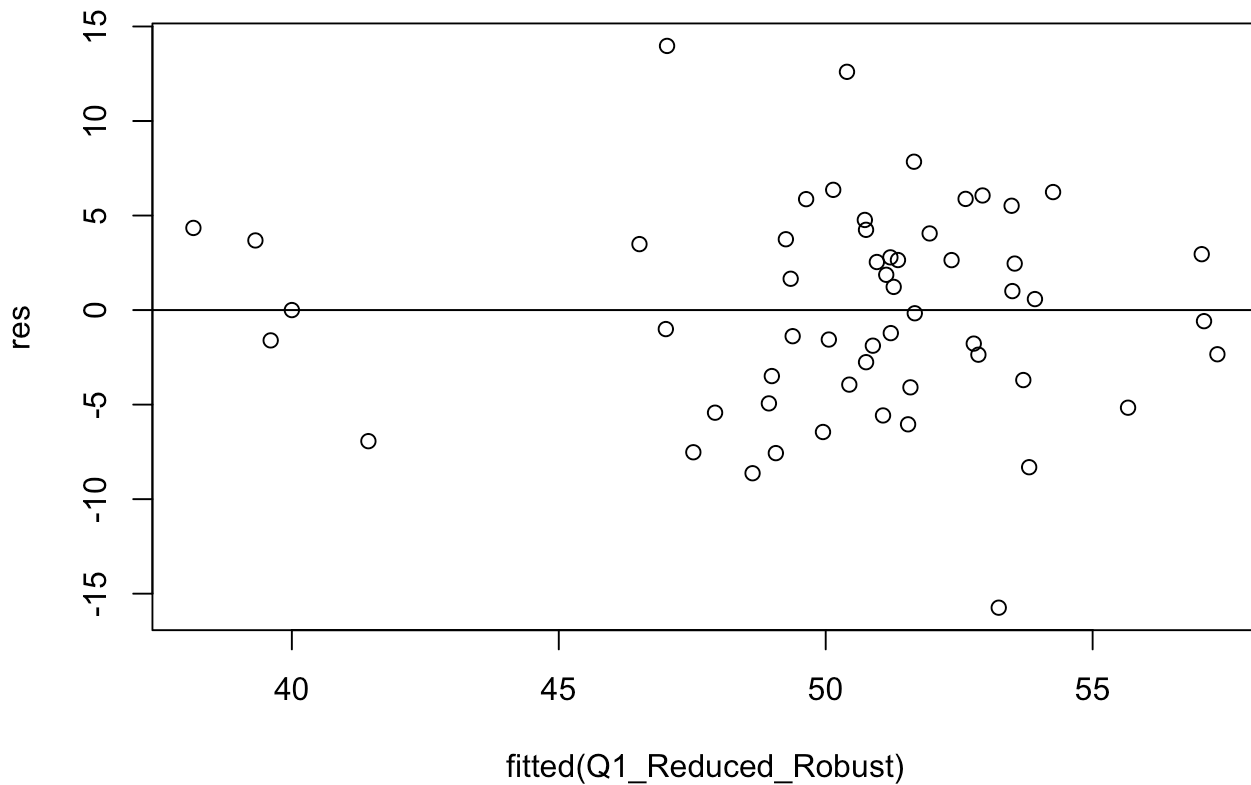
tuning.chi	bb	tuning.psi	refine.tol	
1.548e+00	5.000e-01	4.685e+00	1.000e-07	
rel.tol	scale.tol	solve.tol	zero.tol	
1.000e-07	1.000e-10	1.000e-07	1.000e-10	
eps.outlier	eps.x	warn.limit.reject	warn.limit.meanrw	
1.786e-03	1.819e-12	5.000e-01	5.000e-01	
nResample	max.it	best.r.s	k.fast.s	k.max
500	50	2	1	200
maxit.scale	trace.lev	mts	compute.rd	fast.s.large.n
200	0	1000	0	2000
psi	subsampling	cov		
"bisquare"	"nonsingular"	".vcov.avar1"		
compute.outlier.stats				
"SM"				

seed : int(0)

The robust regression returned a better Adjusted R-Squared = 2928 (approximately 29% of variation) than the regular regression with that showed approximately 27%. Interestingly all the Dominant Prey Species are now significant coefficients except Capelin. We will now look at the residuals for the robust regression model.

```
#residuals
res = resid(Q1_Reduced_Robust)
plot(fitted(Q1_Reduced_Robust), res)
abline(0,0)
```





**Figure 11 - Residual for Reduced Variables with Robust Linear Model**

The residuals look a bit better but not a huge improvement over the linear model above. We will try removing the outlier observation. A new dataset was uploaded below with Pollock and White Hake Dominant Prey Species Removed.

```
seal_data_2 <- read.csv("/Users/peterbraithwaite/Desktop/IDPhD Classes/Stat 5620_Updated/summary(seal_data_2)
```

MomID	Year	Dietary.energy.density	Diet.diversity
Min. : 19.0	Min. :1996	Min. :5.207	Min. :0.1788
1st Qu.: 701.5	1st Qu.:2002	1st Qu.:5.686	1st Qu.:0.3061
Median : 4269.0	Median :2011	Median :5.805	Median :0.3518
Mean : 4897.8	Mean :2008	Mean :5.782	Mean :0.3531
3rd Qu.: 9020.5	3rd Qu.:2013	3rd Qu.:5.904	3rd Qu.:0.3959
Max. :10690.0	Max. :2015	Max. :6.517	Max. :0.5290

Dominant.prey.species	Mom.Age	Pup.sex	Pup.Weann.Mass
Length:51	Min. : 9.00	Min. :1.000	Min. :37.50
Class :character	1st Qu.:23.00	1st Qu.:1.000	1st Qu.:47.00
Mode :character	Median :25.00	Median :1.000	Median :51.50
	Mean :23.22	Mean :1.431	Mean :51.36
	3rd Qu.:26.00	3rd Qu.:2.000	3rd Qu.:55.50
	Max. :31.00	Max. :2.000	Max. :63.00

## Re-fit Linear Model with Reduced Data-set

Removed Pollock and White Hake from Data-set

```
Q3_Full= lm(Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.densi  
summary(Q3_Full)
```

Call:

```
lm(formula = Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity +  
    Dietary.energy.density + Year + Mom.Age + Pup.sex, data = seal_data_2)
```

Residuals:

Min	1Q	Median	3Q	Max
-14.9432	-3.2854	0.9013	3.1288	14.2253

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-258.8154	396.4474	-0.653	0.5174
Dominant.prey.speciesCapelin	-0.8433	6.9966	-0.121	0.9046
Dominant.prey.speciesNorthernSandlance	-3.2757	5.2773	-0.621	0.5381
Dominant.prey.speciesRedfish	-2.9541	5.0552	-0.584	0.5621
Diet.diversity	-26.8081	11.8105	-2.270	0.0284 *
Dietary.energy.density	-7.2831	6.0669	-1.200	0.2367
Year	0.1835	0.2045	0.897	0.3747
Mom.Age	-0.1580	0.2413	-0.655	0.5160
Pup.sex	-0.1694	1.7915	-0.095	0.9251

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 5.95 on 42 degrees of freedom

Multiple R-squared: 0.1856, Adjusted R-squared: 0.03048

F-statistic: 1.197 on 8 and 42 DF, p-value: 0.3243

```
# p-value: 0.3243  
# Adjusted R-squared: 0.03048
```

## Step Function (Forward)

We now used the Step Function to determine the best fit model using AIC.

```
fwd.model = step (Q3_Full, direction='forward')
```

Start: AIC=190

```
Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.density +  
    Year + Mom.Age + Pup.sex
```

The forward step retained all variables with AIC = 190. It should be noted that this AIC is better than the best AIC 202.62 for the simple linear regression above.

## Step Function (Backward)

```
backward.model = step(Q3_Full, direction='backward')
```

Start: AIC=190

Pup.Weann.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.density +  
Year + Mom.Age + Pup.sex

	Df	Sum of Sq	RSS	AIC
- Dominant.prey.species	3	27.752	1514.5	184.94
- Pup.sex	1	0.316	1487.0	188.01
- Mom.Age	1	15.188	1501.9	188.52
- Year	1	28.503	1515.2	188.97
- Dietary.energy.density	1	51.011	1537.7	189.72
<none>			1486.7	190.00
- Diet.diversity	1	182.378	1669.1	193.90

Step: AIC=184.94

Pup.Weann.Mass ~ Diet.diversity + Dietary.energy.density + Year +  
Mom.Age + Pup.sex

	Df	Sum of Sq	RSS	AIC
- Pup.sex	1	1.573	1516.0	182.99
- Year	1	36.520	1551.0	184.16
- Mom.Age	1	52.015	1566.5	184.66
<none>			1514.5	184.94
- Dietary.energy.density	1	134.905	1649.4	187.29
- Diet.diversity	1	195.457	1709.9	189.13

Step: AIC=182.99

Pup.Weann.Mass ~ Diet.diversity + Dietary.energy.density + Year +  
Mom.Age

	Df	Sum of Sq	RSS	AIC
- Year	1	41.970	1558.0	182.39
- Mom.Age	1	52.962	1569.0	182.75
<none>			1516.0	182.99
- Dietary.energy.density	1	134.114	1650.2	185.32
- Diet.diversity	1	195.334	1711.4	187.17

Step: AIC=182.39

Pup.Weann.Mass ~ Diet.diversity + Dietary.energy.density + Mom.Age

	Df	Sum of Sq	RSS	AIC
- Mom.Age	1	17.150	1575.2	180.94
<none>			1558.0	182.39

```

- Dietary.energy.density 1    93.118 1651.1 183.35
- Diet.diversity          1    187.904 1745.9 186.19

```

Step: AIC=180.94

Pup.Weans.Mass ~ Diet.diversity + Dietary.energy.density

	Df	Sum of Sq	RSS	AIC
<none>			1575.2	180.94
- Dietary.energy.density	1	93.233	1668.4	181.88
- Diet.diversity	1	197.610	1772.8	184.97

The backward step retained the Diet Diversity and Dietary.energy.density variables with AIC 180.94. It should be noted that this AIC is better than the best AIC 202.62 for the simple linear regression above.

## Explore Backward Step Linear Model

```

Q3_Reduced = lm(Pup.Weans.Mass ~ Dietary.energy.density + Diet.diversity,
                 data = seal_data_2)
summary(Q3_Reduced)

```

Call:

```

lm(formula = Pup.Weans.Mass ~ Dietary.energy.density + Diet.diversity,
    data = seal_data_2)

```

Residuals:

Min	1Q	Median	3Q	Max
-15.3764	-4.3276	0.7659	3.9409	13.9590

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	97.382	22.844	4.263	9.39e-05 ***
Dietary.energy.density	-6.374	3.781	-1.686	0.0984 .
Diet.diversity	-25.954	10.576	-2.454	0.0178 *

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 5.728 on 48 degrees of freedom

Multiple R-squared: 0.1372, Adjusted R-squared: 0.1012

F-statistic: 3.815 on 2 and 48 DF, p-value: 0.029

```

# p-value: 0.029
# Adjusted R-squared: 0.1012

```

## Explore Residuals For Reduced DataSet Removing White Hake and Pollock

```
library(ggfortify)
autoplot(Q3_Reduced)
```

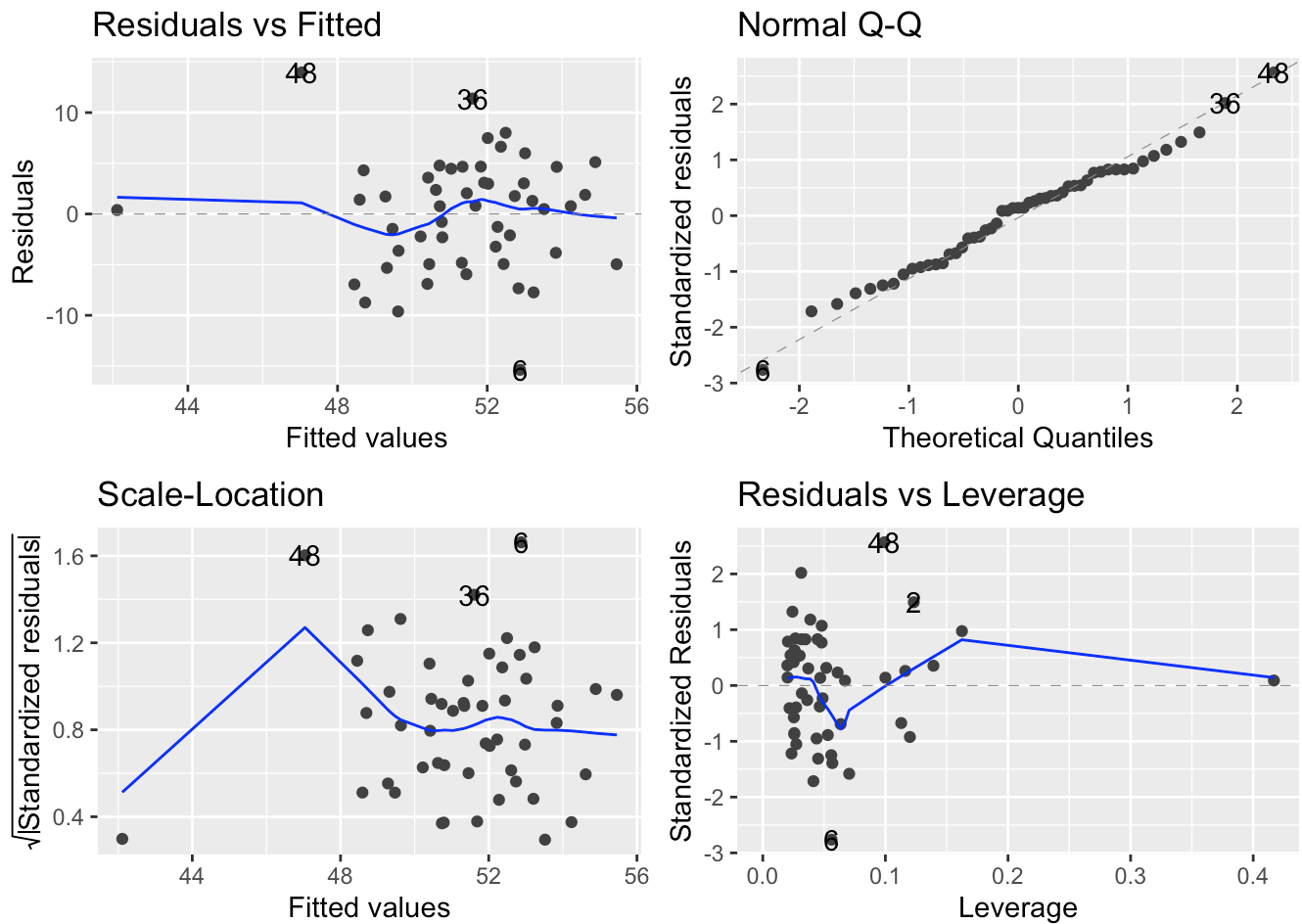
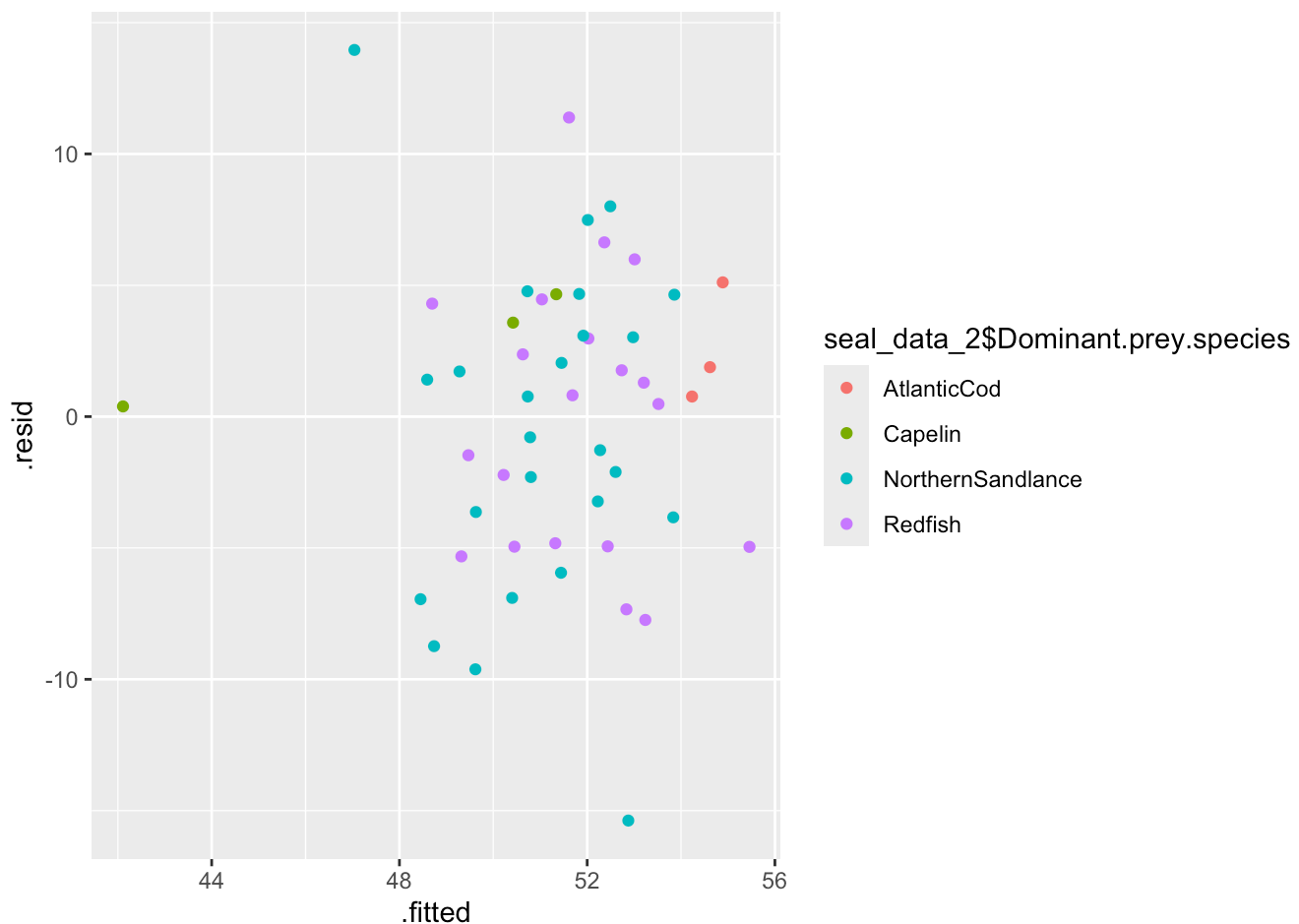


Figure 12 - Residual for Reduced Dataset (White Hake and Pollock removed).

```
ggplot(Q3_Reduced, aes(x = .fitted, y = .resid, color = seal_data_2$Dominant.prey.species
```



**Figure 13 - Residual for Reduced Dataset (White Hake and Pollock removed).**

The lowest value of Capelin is now an outlier and has very high leverage. It is most likely best to remove the lowest value of Capelin or remove the Capelin coefficient all together. Below we load a new dataset with the lowest value of Capelin removed.

```
seal_data_3 <- read.csv("/Users/peterbraithwaite/Desktop/IDPhD Classes/Stat 5620_Updated/summary(seal_data_2)
```

MomID	Year	Dietary.energy.density	Diet.diversity
Min. : 19.0	Min. :1996	Min. :5.207	Min. :0.1788
1st Qu.: 701.5	1st Qu.:2002	1st Qu.:5.686	1st Qu.:0.3061
Median : 4269.0	Median :2011	Median :5.805	Median :0.3518
Mean : 4897.8	Mean :2008	Mean :5.782	Mean :0.3531
3rd Qu.: 9020.5	3rd Qu.:2013	3rd Qu.:5.904	3rd Qu.:0.3959
Max. :10690.0	Max. :2015	Max. :6.517	Max. :0.5290

Dominant.prey.species	Mom.Age	Pup.sex	Pup.Weann.Mass
Length:51	Min. : 9.00	Min. :1.000	Min. :37.50
Class :character	1st Qu.:23.00	1st Qu.:1.000	1st Qu.:47.00
Mode :character	Median :25.00	Median :1.000	Median :51.50
	Mean :23.22	Mean :1.431	Mean :51.36
	3rd Qu.:26.00	3rd Qu.:2.000	3rd Qu.:55.50
	Max. :31.00	Max. :2.000	Max. :63.00

# Fit Full Model with Reduced Data Set

Removed Lowest Value of Capelin species from data.

```
Q4_Full= lm(Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.densi  
summary(Q4_Full)
```

Call:

```
lm(formula = Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity +  
    Dietary.energy.density + Year + Mom.Age + Pup.sex, data = seal_data_3)
```

Residuals:

Min	1Q	Median	3Q	Max
-14.7728	-3.0865	0.7264	3.0387	12.9970

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-311.0480	401.3625	-0.775	0.443
Dominant.prey.speciesCapelin	-1.3014	7.0288	-0.185	0.854
Dominant.prey.speciesNorthernSandlance	-5.1433	5.6712	-0.907	0.370
Dominant.prey.speciesRedfish	-4.6264	5.3875	-0.859	0.395
Diet.diversity	-20.3147	13.8136	-1.471	0.149
Dietary.energy.density	-3.6124	7.2924	-0.495	0.623
Year	0.1990	0.2056	0.968	0.339
Mom.Age	-0.1911	0.2445	-0.782	0.439
Pup.sex	-0.1885	1.7953	-0.105	0.917

Residual standard error: 5.962 on 41 degrees of freedom

Multiple R-squared: 0.1651, Adjusted R-squared: 0.002236

F-statistic: 1.014 on 8 and 41 DF, p-value: 0.441

```
# p-value: 0.002236  
# Adjusted R-squared:0.004332
```

## Step Function (Forward)

We now used the Step Function to determine the best fit model using AIC.

```
fwd.model = step (Q4_Full, direction='forward')
```

Start: AIC=186.61

```
Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.density +  
    Year + Mom.Age + Pup.sex
```

The forward step retained all variables with AIC = 186.61.

## Step Function (Backward)

```
backward.model = step(Q4_Full, direction='backward')
```

Start: AIC=186.61

Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.density +  
Year + Mom.Age + Pup.sex

	Df	Sum of Sq	RSS	AIC
- Dominant.prey.species	3	52.481	1509.7	182.38
- Pup.sex	1	0.392	1457.6	184.62
- Dietary.energy.density	1	8.721	1465.9	184.91
- Mom.Age	1	21.729	1478.9	185.35
- Year	1	33.316	1490.5	185.74
<none>			1457.2	186.61
- Diet.diversity	1	76.867	1534.1	187.18

Step: AIC=182.38

Pup.Weans.Mass ~ Diet.diversity + Dietary.energy.density + Year +  
Mom.Age + Pup.sex

	Df	Sum of Sq	RSS	AIC
- Pup.sex	1	1.992	1511.7	180.45
- Year	1	39.426	1549.1	181.67
- Mom.Age	1	56.342	1566.0	182.21
<none>			1509.7	182.38
- Dietary.energy.density	1	77.795	1587.5	182.89
- Diet.diversity	1	116.741	1626.4	184.10

Step: AIC=180.45

Pup.Weans.Mass ~ Diet.diversity + Dietary.energy.density + Year +  
Mom.Age

	Df	Sum of Sq	RSS	AIC
- Year	1	45.154	1556.8	179.92
- Mom.Age	1	56.672	1568.3	180.29
<none>			1511.7	180.45
- Dietary.energy.density	1	78.067	1589.7	180.97
- Diet.diversity	1	117.958	1629.6	182.20

Step: AIC=179.92

Pup.Weans.Mass ~ Diet.diversity + Dietary.energy.density + Mom.Age

	Df	Sum of Sq	RSS	AIC
- Mom.Age	1	18.078	1574.9	178.50
- Dietary.energy.density	1	48.992	1605.8	179.47
<none>			1556.8	179.92
- Diet.diversity	1	124.319	1681.1	181.76

Step: AIC=178.5



```
Pup.Weans.Mass ~ Diet.diversity + Dietary.energy.density
```

	Df	Sum of Sq	RSS	AIC
<none>			1574.9	178.50
- Dietary.energy.density	1	66.647	1641.5	178.57
- Diet.diversity	1	159.201	1734.1	181.31

The backward step retained the Diet Diversity and Dietary.energy.density variables with AIC=178.5. It should be noted that this the best AIC value so far.

## Explore Best Backward Step Linear Model

```
Q4_Reduced = lm(Pup.Weans.Mass ~ Dietary.energy.density + Diet.diversity,  
                data = seal_data_3)  
summary(Q4_Reduced)
```

Call:

```
lm(formula = Pup.Weans.Mass ~ Dietary.energy.density + Diet.diversity,  
    data = seal_data_3)
```

Residuals:

Min	1Q	Median	3Q	Max
-15.3747	-4.5794	0.8175	4.1879	14.0677

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	98.933	29.029	3.408	0.00135 **
Dietary.energy.density	-6.614	4.689	-1.410	0.16503
Diet.diversity	-26.461	12.140	-2.180	0.03432 *

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 5.789 on 47 degrees of freedom

Multiple R-squared: 0.0977, Adjusted R-squared: 0.0593

F-statistic: 2.545 on 2 and 47 DF, p-value: 0.08928

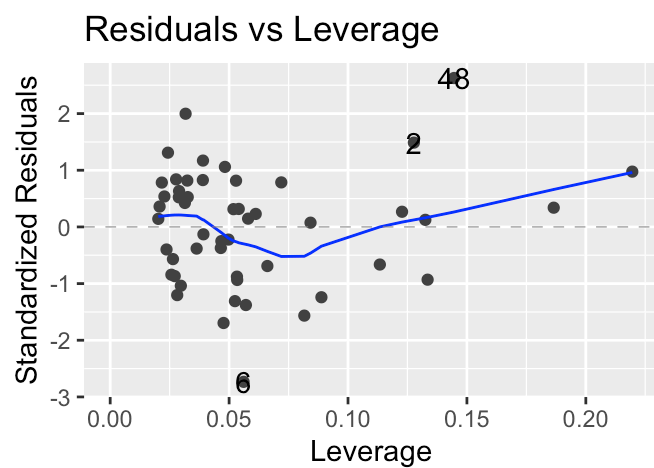
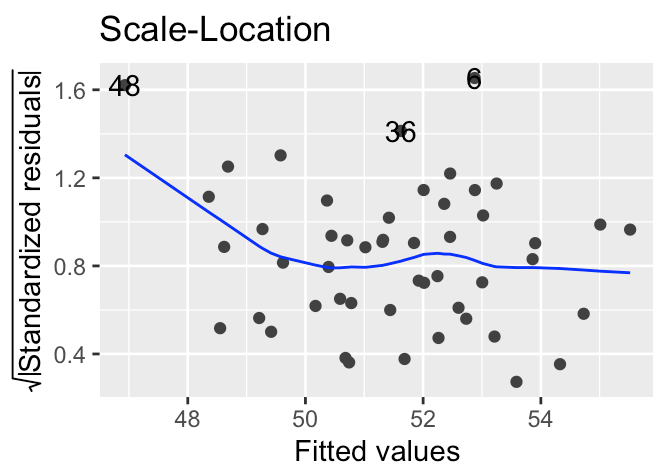
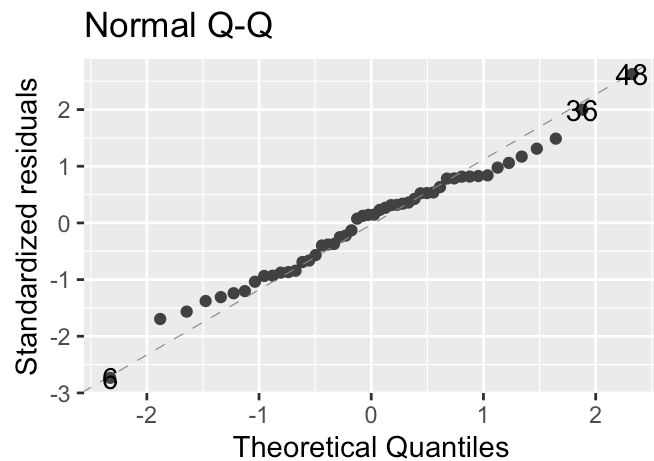
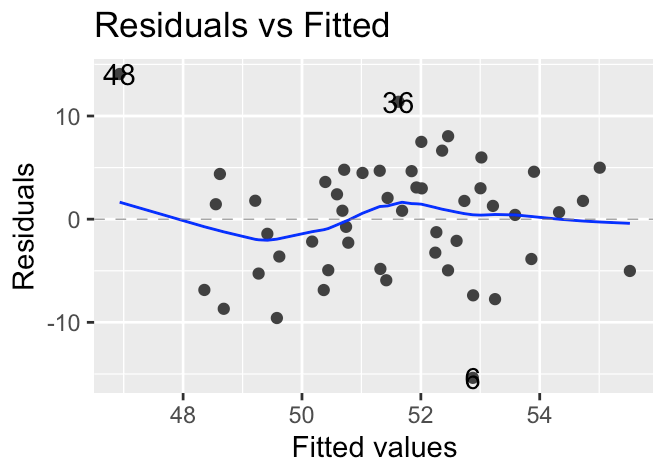
```
extractAIC(Q4_Reduced)
```

```
[1] 3.000 178.496
```

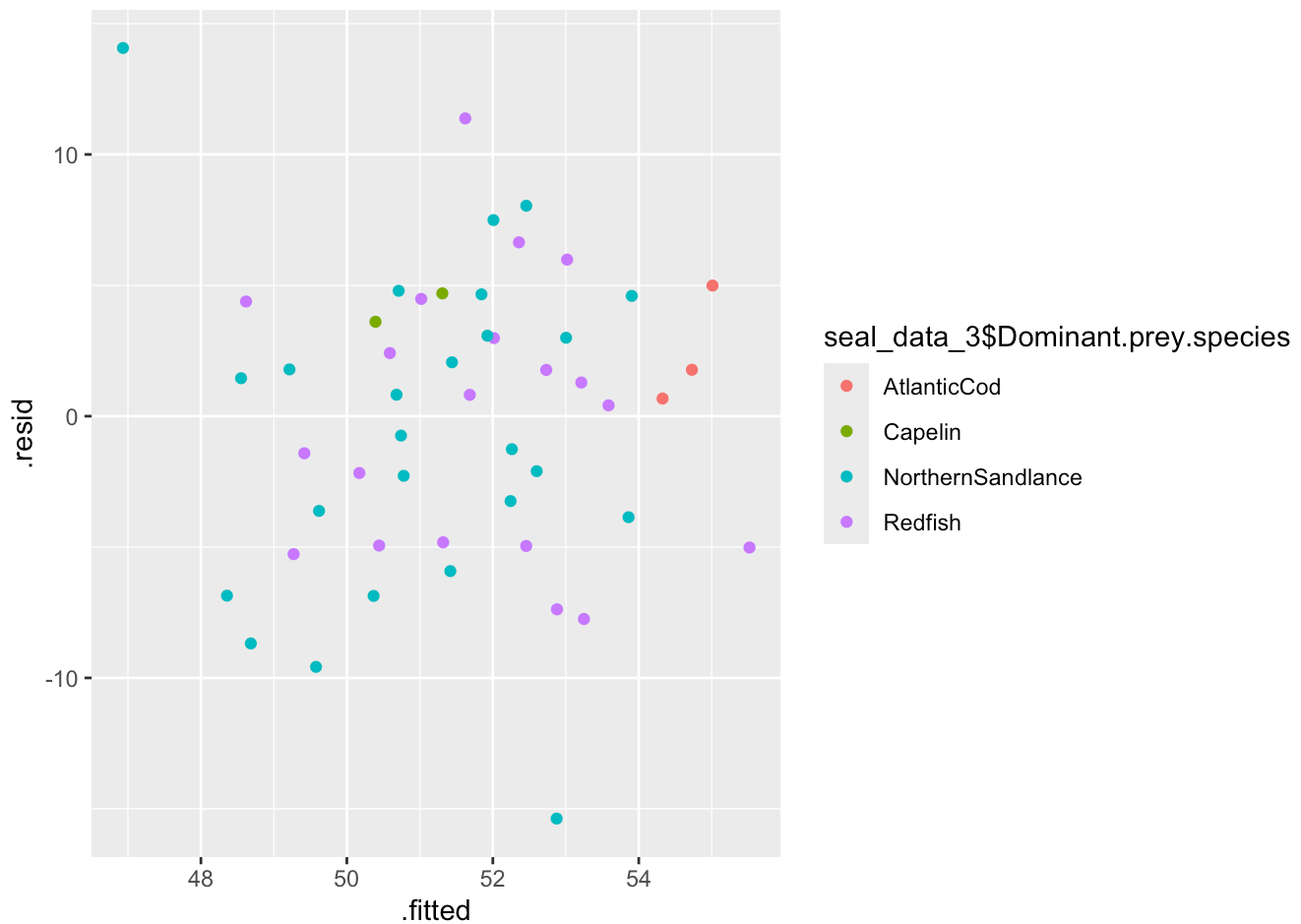
```
# p-value: 0.08928  
# Adjusted R-squared: 0.0593  
# AIC=178.496
```

## Explore Residuals

```
library(ggfortify)
autoplot(Q4_Reduced)
```



```
ggplot(Q4_Reduced, aes(x = .fitted, y = .resid, color = seal_data_3$Dominant.prey.species))
```



**Figure 16 - Residual for Reduced Dataset (White Hake and Pollock removed completely and lowest value of Capeline).**

From this figure it seems we have a fairly balanced residuals now. Although Northern Sandlance is show some extreme values (both large and small) for risk of overfitting the model we will stop removing variables and observations.

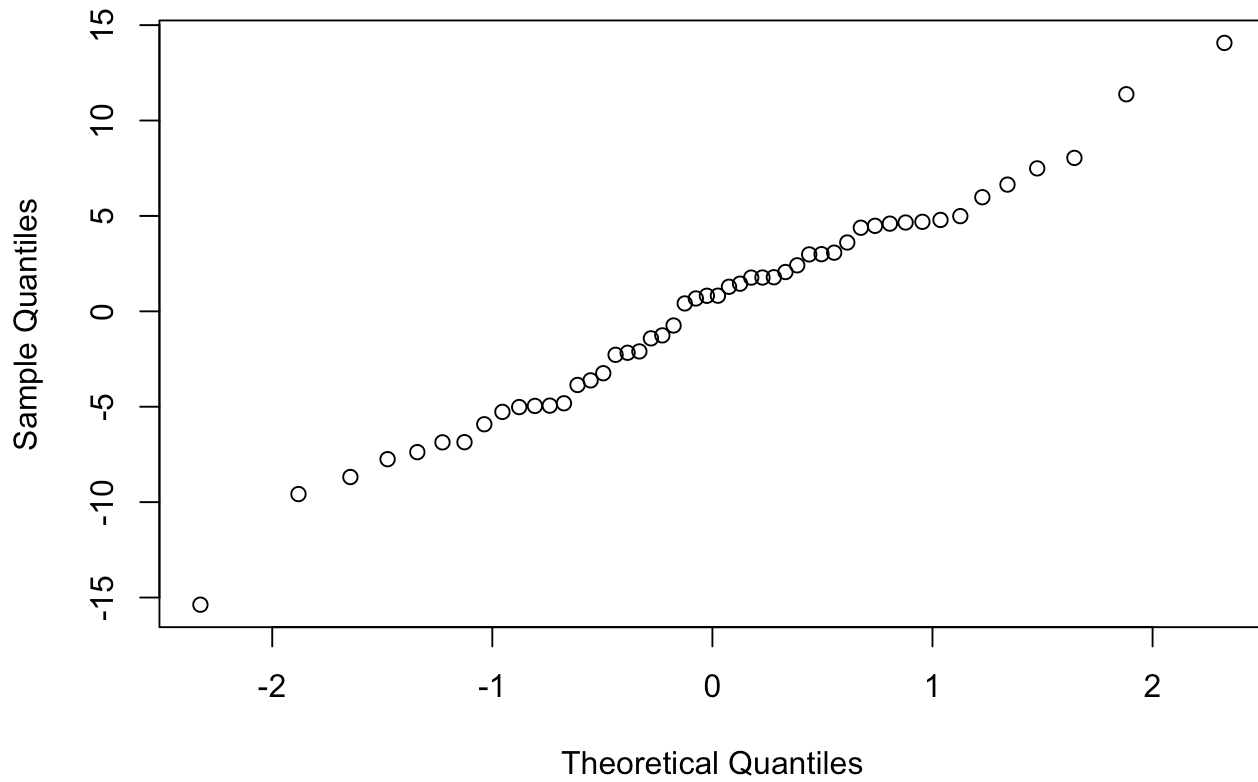
Now, since this is a General Linear Model and not a Generalized Linear Model we will need to check for the assumptions including linearity, independence of errors, homoscedasticity, and normality.

## Testing Assumptions

### Checking for Normality

```
qqnorm(Q4_Reduced$residuals)
```

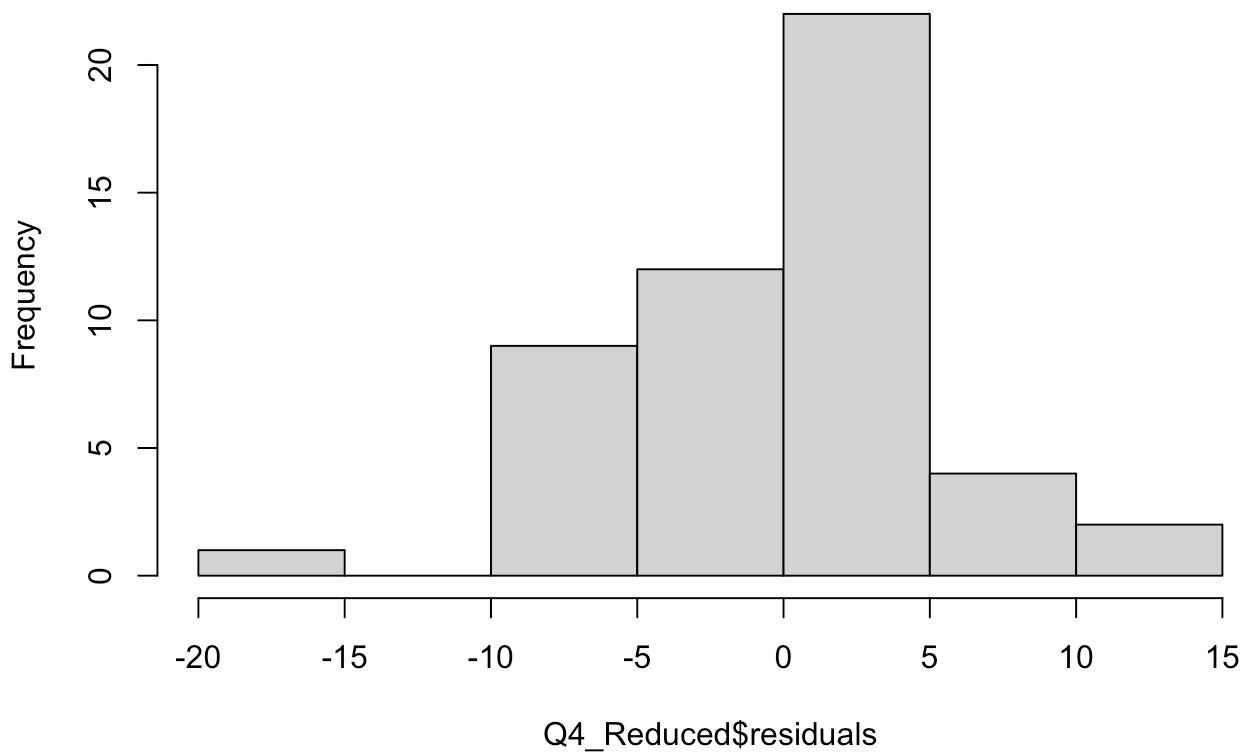
## Normal Q-Q Plot



Q-Q plot appears to show normality

```
hist(Q4_Reduced$residuals)
```

## Histogram of Q4\_Reduced\$residuals



Histogram appears to show normality

## Shapiro-Wilk's Test

```
shapiro.test(Q4_Reduced$residuals)
```

Shapiro-Wilk normality test

data: Q4\_Reduced\$residuals  
W = 0.98681, p-value = 0.8456

From this  $W = 0.98681$  and  $p\text{-value} = 0.8456$ . Given the p-value is not significant there is not a variation of the assumption of normality.

From these 3 explorations we are now confident in normality of the data.

## Checking for Homogeneity of Variance

Using Breusch\_Pagan Test

```
require(lmtest)
```

Loading required package: lmtest

Loading required package: zoo

Attaching package: 'zoo'

The following objects are masked from 'package:base':

as.Date, as.Date.numeric

```
bptest(Q4_Reduced)
```

studentized Breusch-Pagan test

data: Q4\_Reduced

BP = 2.0797, df = 2, p-value = 0.3535

```
# BP = 2.0797  
# df = 2  
# p-value = 0.3535
```

Thus the p-value is greater than 0.05 so we accept the null hypothesis and assume we have homoscedasticity.

## Checking for Independence of Predictor Variables

Using the Dublin-Watson

```
dwtest(Q4_Reduced)
```

Durbin-Watson test

data: Q4\_Reduced

DW = 2.2815, p-value = 0.792

alternative hypothesis: true autocorrelation is greater than 0

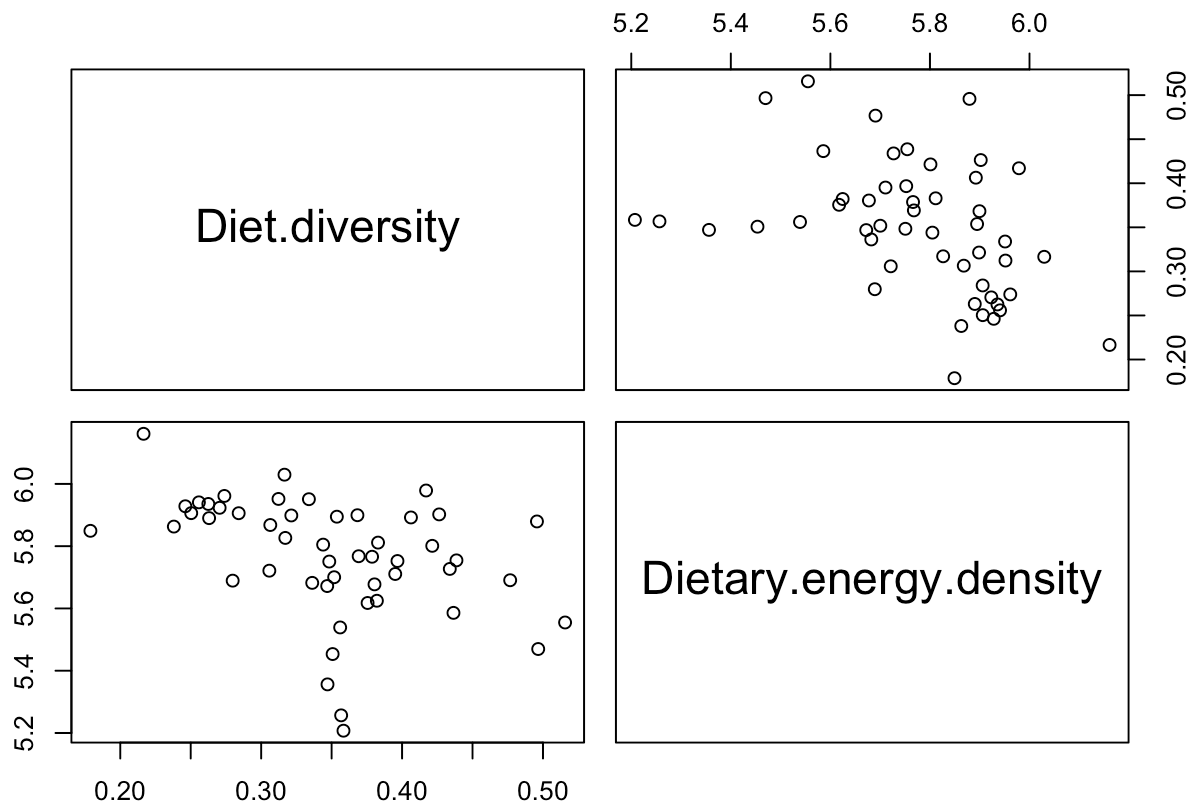
```
# DW = 2.2815  
# p-value = 0.792
```

This p-value suggests very little to no autocorrelation.

## Checking for Correlation of the Variables

Graphically Explore for Correlation

```
pairs( ~Diet.diversity + Dietary.energy.density, data = seal_data_3)
```



Data points look random and not correlated.

## Numerically Explore for Correlation

```
cor.test( ~Diet.diversity + Dietary.energy.density, data = seal_data_3)
```

Pearson's product-moment correlation

data: Diet.diversity and Dietary.energy.density

t = -3.0499, df = 48, p-value = 0.00372

alternative hypothesis: true correlation is not equal to 0

95 percent confidence interval:

-0.6125589 -0.1402952

sample estimates:

cor

-0.402909

```
# t = -0.35989
```

```
# df = 74
```

```
# p-value = 0.72
```

From this we can conclude that the variables are not correlated.

Thus, all our assumptions are met for the Linear Model and our findings can be accepted.

## Generalized Linear Model

Given we will be fitting a Gaussian (normal) distribution Generalized Linear Model it seems safe to assume that the model will demonstrate the same data issues as the Linear Model with regard to residual outliers. For this reason we start with the reduced dataset that eliminated the the Dominant Prey Species White Hake, Pollock, and the lowest value of Capelin.

### Fitting Full Generalized Linear Model (Gaussian Family)

#### All Predictor Variables

```
GLM_1 = glm (Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.dens  
summary (GLM_1)
```

Call:

```
glm(formula = Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity +  
    Dietary.energy.density + Year + Mom.Age + Pup.sex + MomID,  
    family = gaussian, data = seal_data_3)
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-1.076e+02	4.092e+02	-0.263	0.7940
Dominant.prey.speciesCapelin	-4.464e+00	7.103e+00	-0.629	0.5332
Dominant.prey.speciesNorthernSandlance	-6.486e+00	5.592e+00	-1.160	0.2530
Dominant.prey.speciesRedfish	-6.424e+00	5.363e+00	-1.198	0.2380
Diet.diversity	-2.408e+01	1.366e+01	-1.762	0.0856 .
Dietary.energy.density	-3.922e+00	7.123e+00	-0.551	0.5850
Year	9.835e-02	2.090e-01	0.471	0.6405
Mom.Age	-1.175e-01	2.425e-01	-0.484	0.6307
Pup.sex	-4.899e-01	1.762e+00	-0.278	0.7824
MomID	4.208e-04	2.432e-04	1.730	0.0913 .

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for gaussian family taken to be 33.89238)

Null deviance: 1745.4 on 49 degrees of freedom  
Residual deviance: 1355.7 on 40 degrees of freedom  
AIC: 328.9



Number of Fisher Scoring iterations: 2

```
# AIC = 328.9
```

From here we will use the GLM\_1 model for the Step Function to reduce the number of predictor variables and increase the accuracy and efficiency of the model.

## Step Function (Backward)

```
backward.model = step(GLM_1, direction='backward')
```

Start: AIC=328.9

Pup.Weans.Mass ~ Dominant.prey.species + Diet.diversity + Dietary.energy.density +  
Year + Mom.Age + Pup.sex + MomID

	Df	Deviance	AIC
- Dominant.prey.species	3	1412.7	324.96
- Pup.sex	1	1358.3	326.99
- Year	1	1363.2	327.17
- Mom.Age	1	1363.7	327.19
- Dietary.energy.density	1	1366.0	327.27
<none>		1355.7	328.90
- MomID	1	1457.2	330.51
- Diet.diversity	1	1461.0	330.64

Step: AIC=324.96

Pup.Weans.Mass ~ Diet.diversity + Dietary.energy.density + Year +  
Mom.Age + Pup.sex + MomID

	Df	Deviance	AIC
- Pup.sex	1	1416.7	323.10
- Year	1	1422.6	323.30
- Mom.Age	1	1444.5	324.07
<none>		1412.7	324.96
- MomID	1	1509.7	326.28
- Dietary.energy.density	1	1521.1	326.65
- Diet.diversity	1	1571.7	328.29

Step: AIC=323.1

Pup.Weans.Mass ~ Diet.diversity + Dietary.energy.density + Year +  
Mom.Age + MomID

	Df	Deviance	AIC
- Year	1	1430.0	321.57
- Mom.Age	1	1449.0	322.23
<none>		1416.7	323.10
- MomID	1	1511.7	324.34
- Dietary.energy.density	1	1525.0	324.78
- Diet.diversity	1	1576.8	326.45

Step: AIC=321.57

Pup.Weans.Mass ~ Diet.diversity + Dietary.energy.density + Mom.Age +  
MomID

	Df	Deviance	AIC
- Mom.Age	1	1449.0	320.23
<none>		1430.0	321.57
- Dietary.energy.density	1	1525.9	322.81
- MomID	1	1556.8	323.81
- Diet.diversity	1	1601.1	325.21

Step: AIC=320.23

Pup.Weans.Mass ~ Diet.diversity + Dietary.energy.density + MomID

	Df	Deviance	AIC
<none>		1449.0	320.23
- Dietary.energy.density	1	1569.6	322.22
- MomID	1	1574.9	322.39
- Diet.diversity	1	1661.7	325.07

The backward step retained the Diet Diversity and Dietary.energy.density variables with AIC AIC = 320.23. It should be noted this is the same predictor variables that were selected in the Step Function as the Linear Model but the AIC is much worse then the reduced Linear Model with AIC = 178.5. This is because the Linear Model is a less complex model and thus has a lower AIC.

## Explore Backward Step Linear Model

```
GLM_2 = glm(Pup.Weans.Mass ~ Dietary.energy.density + Diet.diversity,  
            data = seal_data_3, family = gaussian)  
summary(GLM_2)
```

Call:

```
glm(formula = Pup.Weans.Mass ~ Dietary.energy.density + Diet.diversity,  
     family = gaussian, data = seal_data_3)
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	98.933	29.029	3.408	0.00135 **
Dietary.energy.density	-6.614	4.689	-1.410	0.16503
Diet.diversity	-26.461	12.140	-2.180	0.03432 *

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for gaussian family taken to be 33.50838)

Null deviance: 1745.4 on 49 degrees of freedom  
Residual deviance: 1574.9 on 47 degrees of freedom  
AIC: 322.39

Number of Fisher Scoring iterations: 2

# AIC = 322.39

## Explore Residuals

```
library(ggfortify)
autoplot(GLM_2)
```

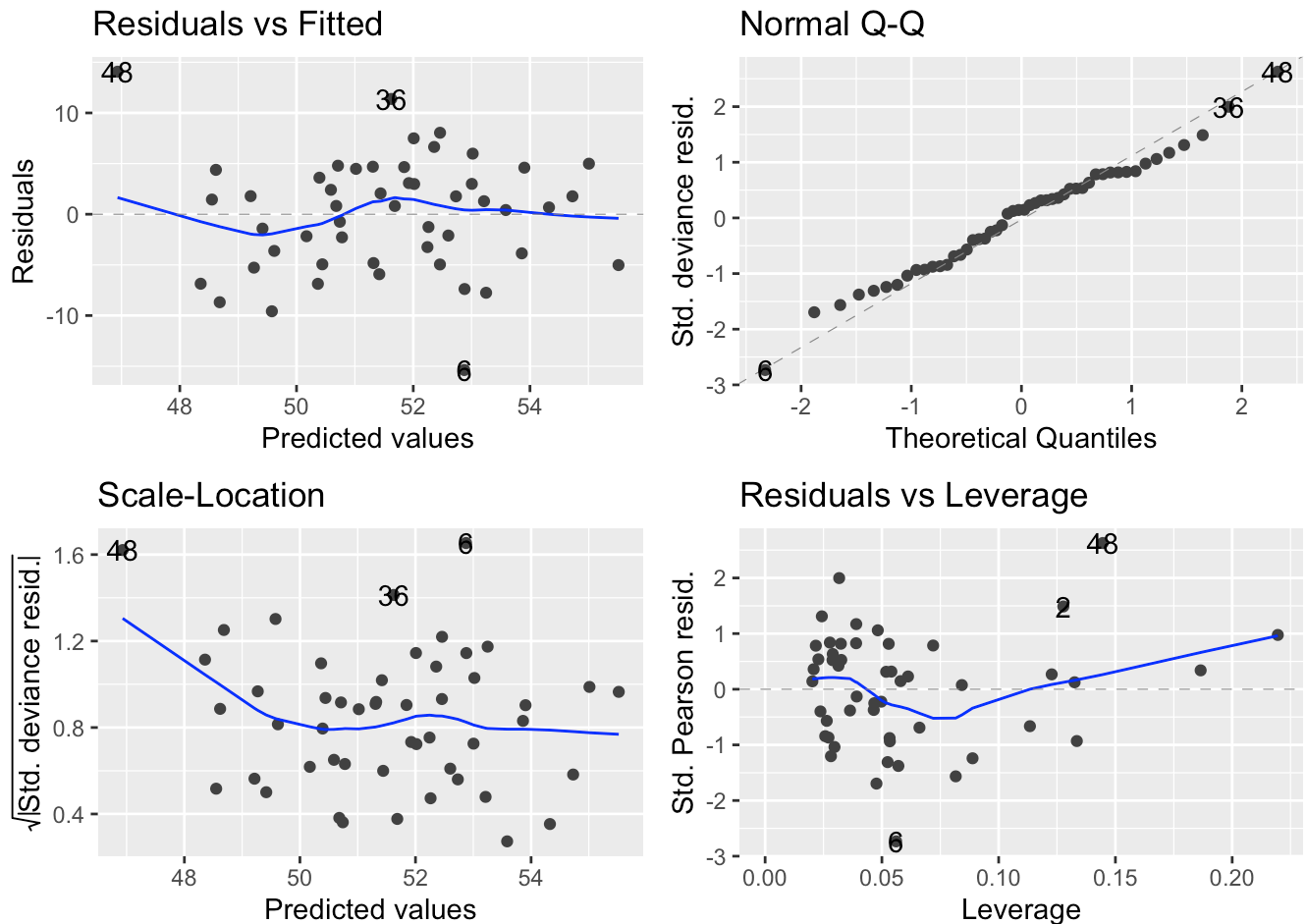
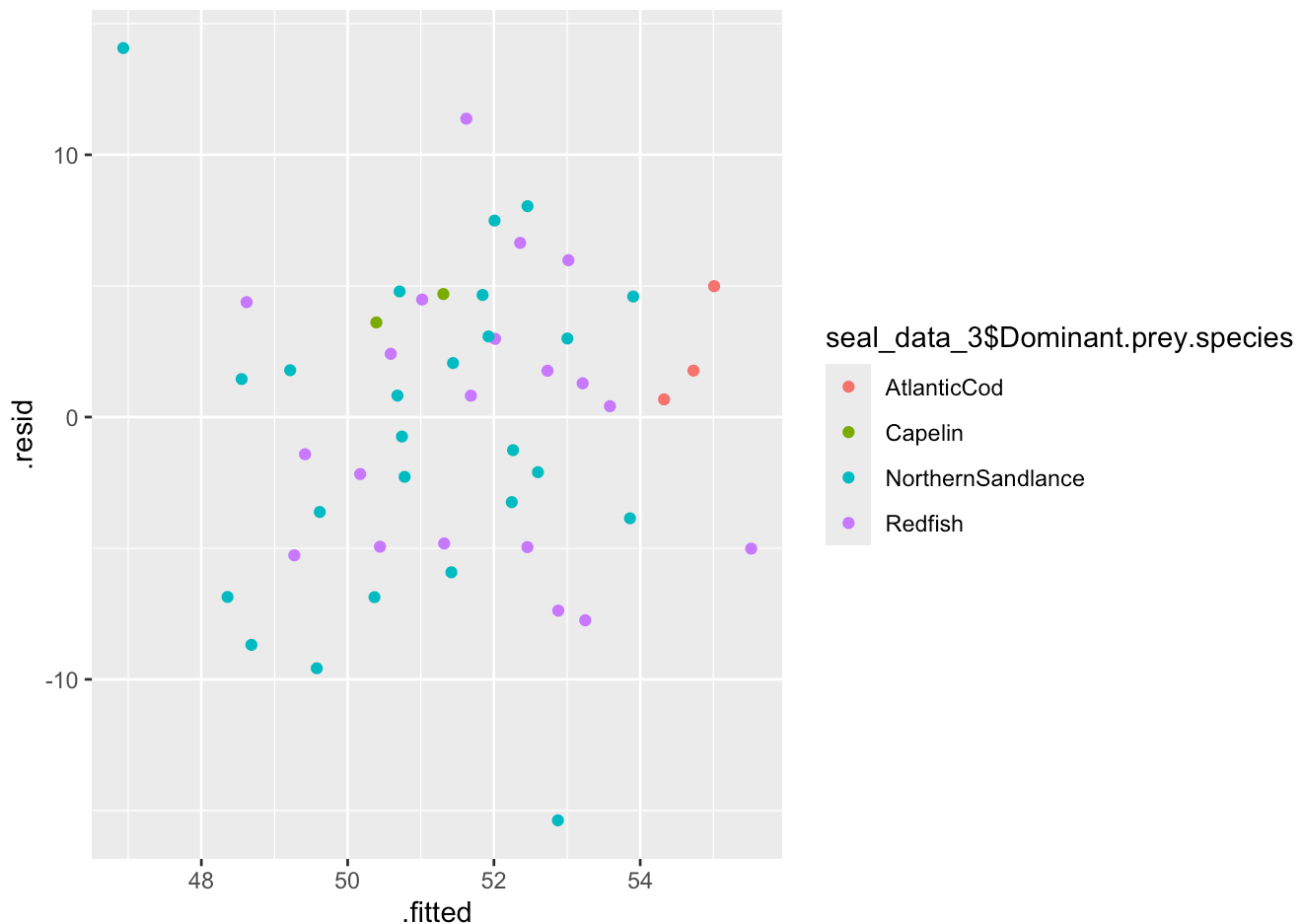


Figure 17 - Residual for Generalized Linear Model With Reduced Dataset

As we might expect the residuals for this Gaussian Generalized Linear Model are very similar to the residuals for the Linear Model above.

```
ggplot(GLM_2, aes(x = .fitted, y = .resid, color = seal_data_3$Dominant.prey.species )) +
```



**Figure 18 - Residual for Generalized Linear Model With Reduced Dataset**

Again, as we might expect the residuals for this Gaussian Generalized Linear Model are very similar to the residuals for the Linear Model above. Given the Generalized Linear Model is creating similar residuals with the a same predictor variables as the Linear Model but with a less desirable AIC (GLM = 322.39 vs. LM = 178.5), it seems the more simple Linear Model should be used.

## Fit Linear Mixed Model

The Linear Model may benefit from Random Effects on the Dominant Prey Species variable given the different prey species would have an effect on the other two remaining variables, Dietary Energy Density and Diet Diversity.

### Linear Mixed Model With Both Diet Diversity and Dietary Energy Loss Random Effects Based on Cluster Variable Dominant Prey Species

```
require(lme4)
```

```
GLMM_1 = lmer(Pup.Weans.Mass ~ Diet.diversity + Dietary.energy.density + (Dietary.energy
```

boundary (singular) fit: see help('isSingular')

```
summary(GLMM_1)
```

Linear mixed model fit by REML ['lmerMod']

Formula:

Pup.Weans.Mass ~ Diet.diversity + Dietary.energy.density + (Dietary.energy.density | Dominant.prey.species) + (Diet.diversity | Dominant.prey.species)

Data: seal\_data\_3

REML criterion at convergence: 301.5

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.6560	-0.7911	0.1412	0.7235	2.4302

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
Dominant.prey.species	(Intercept)	1.228e-02	0.11081	
	Dietary.energy.density	3.838e-04	0.01959	-1.00
Dominant.prey.species.1	(Intercept)	0.000e+00	0.00000	
	Diet.diversity	2.394e-04	0.01547	NaN
Residual		3.351e+01	5.78864	

Number of obs: 50, groups: Dominant.prey.species, 4

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	98.933	29.029	3.408
Diet.diversity	-26.461	12.140	-2.180
Dietary.energy.density	-6.613	4.689	-1.410

Correlation of Fixed Effects:

(Intr) Dt.dvr

Diet.dvrsty -0.522

Dtry.nrgy.d -0.991 0.403

optimizer (nloptwrap) convergence code: 0 (OK)

boundary (singular) fit: see help('isSingular')

```
extractAIC(GLMM_1)
```

```
[1] 10.0000 334.3899
```

```
# AIC = 337.7947
```

## Linear Mixed Model With Dietary Energy Loss Random Effects Based on Cluster Variable Dominant Prey Species

```
GLMM_2 = lmer(Pup.Wean.Mass ~ Diet.diversity + Dietary.energy.density + (Dietary.energy
```

boundary (singular) fit: see help('isSingular')

```
summary(GLMM_2)
```

Linear mixed model fit by REML ['lmerMod']

Formula:

```
Pup.Wean.Mass ~ Diet.diversity + Dietary.energy.density + (Dietary.energy.density |  
  Dominant.prey.species)  
Data: seal_data_3
```

REML criterion at convergence: 301.5

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.6560	-0.7911	0.1412	0.7235	2.4302

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
Dominant.prey.species	(Intercept)	1.420e-03	0.037686	
	Dietary.energy.density	4.228e-05	0.006502	-1.00
Residual		3.351e+01	5.788642	

Number of obs: 50, groups: Dominant.prey.species, 4

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	98.933	29.029	3.408
Diet.diversity	-26.461	12.140	-2.180
Dietary.energy.density	-6.614	4.689	-1.410

Correlation of Fixed Effects:

(Intr) Dt.dvr

Diet.dvrsty -0.522

Dtry.nrgy.d -0.991 0.403

optimizer (nloptwrap) convergence code: 0 (OK)

boundary (singular) fit: see help('isSingular')

```
extractAIC(GLMM_2)
```

```
[1] 7.0000 328.3899
```

```
# AIC = 328.3899
```

## Linear Mixed Model With Diet Diversity Random Effects Based on Cluster Variable Dominant Prey Species

```
GLMM_3 = lmer(Pup.Wean.Mass ~ Diet.diversity + Dietary.energy.density + (Diet.diversity
```

boundary (singular) fit: see help('isSingular')

```
summary(GLMM_3)
```

Linear mixed model fit by REML ['lmerMod']

Formula:

```
Pup.Wean.Mass ~ Diet.diversity + Dietary.energy.density + (Diet.diversity |  
  Dominant.prey.species)  
Data: seal_data_3
```

REML criterion at convergence: 301.5

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.6560	-0.7911	0.1412	0.7235	2.4302

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
Dominant.prey.species	(Intercept)	0.000e+00	0.000e+00	
	Diet.diversity	2.911e-17	5.395e-09	NaN
Residual		3.351e+01	5.789e+00	

Number of obs: 50, groups: Dominant.prey.species, 4

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	98.933	29.029	3.408
Diet.diversity	-26.461	12.140	-2.180
Dietary.energy.density	-6.614	4.689	-1.410

Correlation of Fixed Effects:

(Intr) Dt.dvr

Diet.dvrsty -0.522

Dtry.nrgy.d -0.991 0.403

optimizer (nloptwrap) convergence code: 0 (OK)

boundary (singular) fit: see help('isSingular')

```
extractAIC(GLMM_3)
```

```
[1] 7.0000 328.3899
```

```
# AIC = 328.3899
```

## Linear Mixed Model With Intercept Random Effects Based on Cluster Variable Dominant Prey Species

```
GLMM_4 = lmer(Pup.Wean.Mass ~ Diet.diversity + Dietary.energy.density + (1 | Dominant.pr
```

boundary (singular) fit: see help('isSingular')

```
summary(GLMM_4)
```

Linear mixed model fit by REML ['lmerMod']

Formula: Pup.Wean.Mass ~ Diet.diversity + Dietary.energy.density + (1 | Dominant.prey.species)

Data: seal\_data\_3

REML criterion at convergence: 301.5

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.6560	-0.7911	0.1412	0.7235	2.4302

Random effects:

Groups	Name	Variance	Std.Dev.
Dominant.prey.species	(Intercept)	0.00	0.000
Residual		33.51	5.789

Number of obs: 50, groups: Dominant.prey.species, 4

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	98.933	29.029	3.408
Diet.diversity	-26.461	12.140	-2.180
Dietary.energy.density	-6.614	4.689	-1.410

Correlation of Fixed Effects:

	(Intr) Dt.dvr
Diet.dvrsty	-0.522
Dtry.nrgy.d	-0.991 0.403

optimizer (nloptwrap) convergence code: 0 (OK)

boundary (singular) fit: see help('isSingular')

```
extractAIC(GLMM_4)
```

Warning in optwrap(optimizer, devfun, x@theta, lower = x@lower, calc.derivs = TRUE, : convergence code 3 from bobyqa: bobyqa -- a trust region step failed to reduce q

```
[1] 5.0000 324.3899
```

```
# AIC = 324.3899
```

So, in summary, none of the Linear Mixed Models with Random Effect on the cluster variavble Dominant Prey Species resulted in a more desirable AIC. The best Mixed Model model had an AIC = 324.3899 but



the best Linear Model had an AIC=178.496. Thus the less complex Linear Model is the most efficient and most accurate model.

## Cross Validation

It has been determined that a simple Linear Model using a maternal mother seal's Dietary Energy Density and Diet Diversity are the best predictor variables for Pup Wean Mass. The model `lm(Pup.Wean.Mass ~ Dietary.energy.density + Diet.diversity, data = seal_data_3)` has shown an AIC = 178.496.

We will now use Cross Validation to assess how well we expect this model to perform when predicting future Pup Wean Mass changes.

```
require(modelr)
```

Loading required package: modelr

```
require(caret)
```

Loading required package: caret

Loading required package: lattice

```
require(readr)

# Data
attach(seal_data_3)

# Set Random Seed
set.seed(1980)

# Create Index Matrix (80% train data and 20% test data )
index = createDataPartition(seal_data_3$Pup.Wean.Mass, p = .8, list = FALSE, times = 1)

summary(seal_data_3$Pup.Wean.Mass)
```

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
37.50	47.62	52.00	51.54	55.50	63.00

```
print(seal_data_3$Pup.Wean.Mass)
```

```
[1] 50.5 60.5 50.0 54.5 45.5 37.5 54.5 59.0 50.5 51.0 58.5 59.0 56.0 55.0 56.0
[16] 51.5 59.5 45.5 50.0 47.5 53.5 49.0 55.0 52.5 54.0 55.0 53.0 45.5 56.5 60.0
[31] 48.0 55.5 56.5 48.5 46.5 63.0 43.5 55.5 54.0 51.0 41.5 48.0 53.0 40.0 45.5
[46] 44.0 40.0 61.0 46.0 50.0
```

```
# Convert data frame
data_frame = as.data.frame(seal_data_3)

# Create a data frame for the train data that is 80%
train_data_frame = data_frame[index,]

# Create a data frame for the test data that is 20% ("-" before index mean everything but
test_data_frame = data_frame[-index,]

# Convert response variable in both train and test data frame to a factor
train_data_frame$Pup.Wean.Mass = as.factor (train_data_frame$Pup.Wean.Mass)
test_data_frame$Pup.Wean.Mass = as.factor (test_data_frame$Pup.Wean.Mass)

# Ensure the response variables classes are factors
class (train_data_frame$Pup.Wean.Mass)
```

```
[1] "factor"
```

```
# "factor"
class (test_data_frame$Pup.Wean.Mass)
```

```
[1] "factor"
```

```
# "factor"

# Specify type of training method used and the number of folds
control_specs = trainControl(method = "cv", number = 11 , savePredictions = "all")

# Set Random Seed
set.seed(1980)

require(randomForest)
```

Loading required package: randomForest

randomForest 4.7-1.2

Type rfNews() to see new features/changes/bug fixes.

Attaching package: 'randomForest'

The following object is masked from 'package:ggplot2':

margin

The following object is masked from 'package:dplyr':

combine

```
# now we train the model with the best fit GLM above
# GLM_Final = glm (seal_data$Pup.Wean.Mass ~ seal_data$Dominant.prey.species + seal_data$

model_cross = train(Pup.Wean.Mass ~ Dominant.prey.species + Diet.diversity,
                    method = "lm", data = seal_data_3, trControl = control_specs)

print (model_cross)
```

Linear Regression

50 samples  
2 predictor

No pre-processing  
Resampling: Cross-Validated (11 fold)  
Summary of sample sizes: 45, 46, 46, 44, 46, 46, ...  
Resampling results:

RMSE	Rsquared	MAE
5.434521	0.2021304	4.563434

Tuning parameter 'intercept' was held constant at a value of TRUE

```
# 50 samples
# 2 predictor

# RMSE      Rsquared    MAE
# 5.434521  0.20213048  4.563434

# We now apply the model to the test_data_frame we created from 20%
# of the data that the new model created from the 11 folds has not yet seen

# Predict outcome using model from train_data_frame applied to test_data_frame
Predict_Data = predict (model_cross, newdata = test_data_frame)

# predictions results

#. 52.74225 51.24721 50.67234 57.30204 51.49974 49.31590 50.08707 48.16223

test_data_frame$Pup.Wean.Mass
```

```
[1] 59  59.5 49  55  52.5 41.5 48  40
Levels: 40 41.5 48 49 52.5 55 59 59.5
```

```
# 59    59.5 49    55    52.5 41.5 48    40

#Creates vectors having data points
predicted_value <- (c(52.74225, 51.24721, 50.67234, 57.30204, 51.49974, 49.31590, 50.0870
expected_value <- (c(59, 59.5, 49, 55, 52.5, 41.5, 48, 40))

Cross_Validation_Corrolation = cor (predicted_value, expected_value)

### 0.6596914

R2 = (Cross_Validation_Corrolation^2 )

### 0.4351927 (43.52% of variance)
```

From this cross validation procedure we can see that the Linear Model has a validation score of about 0.66 which mean the Linear Model would do a fair job at predicting Pup Mean Mass based on the maternal mother seal's Dietary Energy Density and Diet Diversity. The estimated R-squared value suggests the model would account for about 43.5% of the variance of Pup Mean Mass.

## Results

Although it was proposed that many predictor variables (2 continuous + 3 categorical for question one and 2 continuous + 5 categorical for question two) would allow greater computational potential to establish a well fit model, the added complexity of these variables was observed to significantly increase the AIC and potentially overfit the model. It was found, particularly in question two, that the simplicity of a Linear Model with a Least Square approach resulted in a simpler linear model with a significantly lower AIC.

In question 2 it was found that best fit Linear Model for Pup Wean Mass included the predictor variables Dietary Energy Density and Diet Diversity with an AIC = 178.496. When a Generalized Linear Model was explored with a Gaussian Family and Identity Link Function it was also found that the best fit model included Dietary Energy Density and Diet Diversity predictor variable but the AIC was much less preferable at 322.39. The reasons for this decreased predicted fit is likely based on the more complex Maximum Likelihood Estimation used in Generalized Linear Models rather than the Linear Least Square method used in the Linear Models. Generalized Linear Mixed Models (GLMM) were also explored using the Dominant Prey Species as the Cluster Variable because it was assumed that this variable would likely have a effect on the prey energy density as well as the diet diversity based on feeding location ecosystems. The best fit GLMM was found to include Diet Diversity and Dietary Energy Density as fixed effects predictor variables and Dominant Prey Species as an intercept random effect variable which had a AIC of 324.3899. As a result it was determined that the best fit model to determine Pup Wean Mass was the Linear Model and this model was tested using Cross Validation. The correlation of predicted values to actual values showed a value of 0.6596914 meaning the Linear Model could be expected to predict Pup Wean Mass with about a 66 percent accuracy.

## Conclusions

## References

Data source : Sara Iverson laboratory and DFO Grey Seal Program.