

PROBABILITY THEORY II

Matthew Joseph, notes by Ramdas Singh

Second Semester

List of Symbols

Ω , a sample space.

ω , an element of a sample space.

EX , the expectation of the random variable X .

$\text{Var}X$, the variance of the random variable X .

$N(\mu, \sigma^2)$, a normal distribution with expectation μ and variance σ^2 .

Contents

| | |
|--|---|
| 1 | 1 |
| 1.1 The Law of Large Numbers | 2 |
| Appendices | 5 |
| A Appendix | 7 |
| Index | 9 |

Chapter 1

January 3rd.

Let Ω be a countable state space, and let each $\omega \in \Omega$ have a probability $P(\omega)$ associated with it.

Lemma 1.1. *For random variables X, Y such that $X(\omega) \leq Y(\omega)$ for all $\omega \in \Omega$. Then, $EX \leq EY$.*

Proof. This can easily be seen by summing over all terms via the alternate definition of the expectation,

$$EX = \sum_{\omega \in \Omega} X(\omega)P(\omega) \leq \sum_{\omega \in \Omega} Y(\omega)P(\omega) = EY. \quad (1.1)$$

■

We now state Markov's inequality.

Theorem 1.2 (*Markov's inequality*). *If X is a non-negative random variable, then for $a > 0$, we have*

$$P(X > a) \leq \frac{EX}{a}. \quad (1.2)$$

Proof. Define an indicator function $I_a(\omega)$ as 1 if $X(\omega) \geq a$, and 0 if otherwise. We then have

$$I_a(\omega) \leq \frac{X(\omega)}{a} \implies P(X \geq a) = EI_a \leq \frac{1}{a}EX. \quad (1.3)$$

■

Remark 1.3. A better upper bound here may be found by starting with $I_a(\omega)X(\omega)$ instead of just $X(\omega)$.

If we have $X \sim N(0, 1)$, then we can find an upper bound for its probability density function.

$$P(X > a) = \int_a^\infty \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \leq \int_a^\infty \frac{1}{\sqrt{2\pi}} \frac{x}{a} e^{-\frac{x^2}{2}} dx = \frac{e^{-\frac{a^2}{2}}}{\sqrt{2\pi}a}. \quad (1.4)$$

Note that X here is a random variable over a continuous state space; the previous lemma and Markov's inequality also work here. We are to show them for the continuous case instead of the discrete one.

Proof. Here, we have $0 \leq X(\omega) \leq Y(\omega)$ for all ω in our continuous state space Ω . We see that $\{X > x\} \subseteq \{Y > x\} \implies P(X > x) \leq P(Y > x)$. Integrating both sides gives us $EX \leq EY$. ■

Theorem 1.4 (*Chebyshev's inequality*). *Let X be a random variable with finite mean $\mu = EX$ and finite variance $\sigma^2 = \text{Var}(X)$. Then for $a > 0$,*

$$P(|X - \mu| > a) \leq \frac{\text{Var}(X)}{a^2}. \quad (1.5)$$

Proof. Start with the proof of Markov's inequality, replacing the indicator function with one that's unity when $|X - \mu| \geq a$. ■

Example 1.5. Suppose X_1, X_2, \dots, X_n are n independent and identically distributed random variables, with $EX_i = \mu$ and $\text{Var}X_i = \sigma^2$. If $S_n = \sum X_i$, we then have

$$P(|S_n - n\mu| > a) \leq \frac{\text{Var}S_n}{a^2} = \frac{n\sigma^2}{a^2}. \quad (1.6)$$

If we replace a with $n^{\frac{1}{2}+\varepsilon}$, we then have

$$P(|S_n - n\mu| > n^{\frac{1}{2}+\varepsilon}) \leq \frac{\sigma^2}{n^{2\varepsilon}} \rightarrow 0 \text{ as } n \rightarrow \infty. \quad (1.7)$$

Proposition 1.6. If $\text{Var}(X) = 0$, then $P(X = EX) = 1$.

Proof. For all $\varepsilon > 0$, we have

$$P(|X - EX| > \varepsilon) \leq \frac{\text{Var}X}{\varepsilon^2} = 0. \quad (1.8)$$

Define A_n as $\{|X - EX| > \frac{1}{n}\}$. Taking $P(\bigcup A_n) = \lim_{n \rightarrow \infty} P(A_n)$, the proof follows. \blacksquare

1.1 The Law of Large Numbers

We start by stating the weak law of large numbers.

Theorem 1.7 (*Weak law of large numbers*). Let $\{X_k\}_{k \geq 1}$ be a sequence of independent and identically distributed random variables with $E|X_i| < \infty$. Let $\mu = EX_i$. Then for any $a > 0$,

$$\lim_{n \rightarrow \infty} P\left(\left|\frac{X_1 + X_2 + \dots + X_n}{n} - \mu\right| > a\right) = 0. \quad (1.9)$$

Proof. For now, let us assume that Ω is countable. We begin with the case where the variance of X_i , σ^2 , is finite. Fix $a > 0$, and let $S_n = X_1 + X_2 + \dots + X_n$. Then,

$$P\left(\left|\frac{S_n}{n} - \mu\right| > a\right) = P(|S_n - n\mu| > na) \leq \frac{\text{Var}S_n}{n^2a^2} = \frac{n\sigma^2}{n^2a^2} \rightarrow 0 \text{ as } n \rightarrow \infty. \quad (1.10)$$

We now focus the case when the variance, σ^2 , is infinite. Assume that the expected value, μ , is 0; if it were non-zero, we would then instead work with $X_i - \mu$. Let $\delta > 0$; we shall choose a particular δ later. For each n , define n pairs of random variables, $U_1, V_1, \dots, U_n, V_n$, as $U_k = X_k, V_k = 0$ if $|X_k| \leq \delta n$, and $U_k = 0, V_k = X_k$ if $|X_k| > \delta n$. X_k can be rewritten as $U_k + V_k$. We then have

$$\{|X_1 + \dots + X_n| \geq na\} \subseteq \{|U_1 + \dots + U_n| \geq \frac{na}{2}\} \cup \{|V_1 + \dots + V_n| \geq \frac{na}{2}\} \quad (1.11)$$

$$\implies P(|X_1 + \dots + X_n| \geq na) \leq P\left(|U_1 + \dots + U_n| \geq \frac{na}{2}\right) + P\left(|V_1 + \dots + V_n| \geq \frac{na}{2}\right). \quad (1.12)$$

We focus on the first term on the right hand side. The U_i 's are independently and identically distributed, so

$$P\left(|U_1 + \dots + U_n| \geq \frac{na}{2}\right) \leq \frac{4E[|U_1 + \dots + U_n|^2]}{a^2n^2} = \frac{4}{a^2n^2} (\text{Var}(U_1 + \dots + U_n) + (nEU_i)^2). \quad (1.13)$$

For the variance, we have

$$\text{Var}(U_1 + \dots + U_n) = n\text{Var}U_i \leq nEU_i^2 \leq nE[|U_i||U_i|] \leq \delta n^2 E[|U_i|] \quad (1.14)$$

which transforms the previous equation as

$$P\left(|U_1 + \dots + U_n| \geq \frac{na}{2}\right) \leq \frac{4}{a^2n^2} (\delta n^2 E[|U_i|] + (nEU_i)^2). \quad (1.15)$$

A lemma (to be proven later) states that $E[|U_i|] = E[|X_i|]$ as $n \rightarrow \infty$, and $EU_i = EX_i = 0$ too. So,

$$P\left(|U_1 + \dots + U_n| \geq \frac{na}{2}\right) \leq \frac{4}{a^2n^2} (\delta n^2 E[|U_i|] + (nEU_i)^2) \leq \frac{4\delta E[|U_i|]}{a^2} + \frac{4}{a^2} (EU_i)^2. \quad (1.16)$$

For the second term on the right hand side, begin with

$$\begin{aligned} P(V_1 + \dots + V_n \neq 0) &\leq P(\{V_1 \neq 0\} \cup \dots \cup \{V_n \neq 0\}) \leq nP(V_i \neq 0) = n \sum_{|x| > \delta n} P(X_i = x) \\ &\leq n \sum_{|x| > \delta n} \frac{|x|}{\delta n} P(X_i = x) = \frac{1}{\delta} E[|V_i|]. \end{aligned} \tag{1.17}$$

The rightmost term here tends to 0 as $n \rightarrow \infty$. Now choose δ to be $\frac{\varepsilon a^2}{6E|X_i|}$, and then choose N to be large enough such that for all $n > N$, both the terms are smaller than $\frac{\varepsilon}{2}$. ■

Appendices

Chapter A

Appendix

Extra content goes here.

Index

Chebyshev's inequality, 1

Markov's inequality, 1

Weak law of large numbers, 2