## Session: DSGE models and rational expectations

Aurélien Poissonnier<sup>1</sup>

<sup>1</sup>Insee-European Commission

Applied Macroeconometrics - ENSAE

## Applied Macroeconometrics Part III: DSGE models

Dynamic
Stochastic
General
Equilibrium

# Plan of part III of the course "Rational Expectation and DSGE models" - Lectures 4,5,6,7

### Lecture 4 (This lecture)

- Introduction to DSGE models (examples, linearizing)
- Solving a DSGE model (i.e. solving for unobserved expectations)

Lecture 5-6 (HLB) Estimation of DSGE models Lecture 7 (AP) Simulation and use of DSGE models with Dynare.

#### Plan

Rational expectations

DSGE models

Solving DSGE models

Wrapping-up

## Rational expectation models

#### Rational expectation models (e.g. DSGE) differ from

- Traditional structural models, "Cowles Commission"-type (taught by P.O. Beffy in Lecture 1 and 2)
- Structural VAR models (Lecture 3)

## **Common point:** dynamic time-series models **Key differences:**

- RE models make expectations explicit
- RE models are more structural than "structural VARs" (which impose only few restrictions)
- RE models are in general derived from micro foundations.

#### Proeminent examples: DSGE models

NB: expectations are (most often) non-observed variables.

→ rational expectations; help circumvent this

## Expectations in macro models

- Expectations are everywhere in macro!
- Consumption-saving decision:  $U'(c_t) = \beta(1 + r_t)E_tU'(c_{t+1})$  ( $c_t$  consumption,  $r_t$  real interest rate)
- asset yields; in the risk-neutral case:  $r_t = \frac{E_t p_{+1} p_t}{p_t} + \frac{d_t}{p_t}$  ( $r_t$  riskless rate,  $d_t$  dividend,  $p_t$  stock price)
- money demand (Cagan)  $ln(M_t/P_t) = -\alpha E_t(\Delta P_{t+1}/P_t) + \varepsilon_t$

## The rational expectation assumption

**Notation**:  $E_t Y_{t+1}$  expectation of  $Y_{t+1}$  formulated at date t

Rational Expectations:

$$E_t Y_{t+1} = E(Y_{t+1}|I_t),$$

 $E(.|I_t)$  mathematical expectation conditional on  $I_t$ , information set available at date t,

## Rational Expectation Hypothesis: a discussion

#### Justifications:

- RE expectations are "model-consistent": no reason to assume economic agents have less knowledge than the econometrician about the structure of the model
- Lucas critique (Lucas, 1976) : traditional reduced form models are non invariant to an economic policy intervention

#### • Remarks:

- RE is a strong assumption
- rational expectations  $\neq$  perfect expectations
- RE can be tested (direct tests, using expectations data like surveys or indirect tests)

## Some alternatives to rational expectations

• Adaptatives expectations

$$E_t Y_{t+1} = \lambda Y_t + (1 - \lambda) E_{t-1} Y_t,$$

- Naive expectations, an extreme particular case:  $E_t Y_{t+1} = Y_t$  (case  $\lambda = 1$ ).
- Many others: rational learning, limited rationality...

## Some alternatives to rational expectations (continued)

Expectations data that can be used for direct tests of the REH  $E_t Y_{t+1}$  Such data are either directly *observed* or built.

#### Examples:

- quantitative or qualitative answers to business survey (in France surveys of consumers or firms by Insee)
- expectations data inferred from financial market data (e.g. inflation swap)

#### Some limits:

- assumptions underlying quantification
- directly observed expectations are most often forecast of some specific agents (e.g. professional forecasters like OECD experts or Consensus Forecast)
- in practice, expectations are most often unoserved

#### These alternatives are not developed in this lecture

## Rational expectations in DSGE

$$E_t Y_{t+1} = E(Y_{t+1}|I_t),$$

We will assume for  $E(\bullet|I_t)$  that agents know the model structure and past data when forming expectations.

Expectations will be internally consistent

#### Plan

#### Rational expectations

#### DSGE models

A simple RBC model
The simplest neo-Keynesian model
(large) DSGE models in public administrations and
international institutions
Linearizing

Solving DSGE models

Wrapping-up

## This course is not about deriving DSGE models

For detailed material on the derivation of such models, see

- 2A Macroéconomie 2: fluctuations with Franck Malherbet
- 3A Monetary Economics with Olivier Loisel
- 3A Structural macroeconomics with Edouard Challe
- appendix slides for this class from last year by Benoît Campagne (Smets and Wouters, 2003)

#### This course focuses on linear models

- linear models ... or linearized around the stationary equilibrium
- Alternative to linearization:
- value function iteration,
- second-order (or higher order) approximations

These approaches will not be detailed here.

See (Canova, 2011), (DeJong and Dave, 2011), (Juillard and Ocaktan, 2008).

NB: in some cases you may lose relevant information by linearizing (Lindé and Trabandt, 2018)

#### What is this course about?

This course will show in the next session how linearized DSGE models can be put under the form of a restricted SVAR.

Before that we will describe DSGE models as macroeconomic tools.

#### Plan

#### **DSGE** models

#### A simple RBC model

The simplest neo-Keynesian model (large) DSGE models in public administrations and international institutions
Linearizing

## A simple RBC model

Keynes-Ramsey's rules: 
$$\frac{1}{c_t} = \tilde{\beta} E_t \left[ \frac{1}{c_{t+1}(1+r_{t+1})} \right]$$
Hours worked: 
$$H_t = 1 - \phi \frac{c_t}{w_t}$$
Production function: 
$$y_t = A_t k_t^{\alpha} H_t^{1-\alpha}$$
Real interest rate: 
$$r_t = \alpha \frac{y_t}{k_t} - \delta$$
Real wage: 
$$w_t = (1-\alpha) \frac{y_t}{H_t}$$
Investment: 
$$i_t = \gamma k_{t+1} - (1\delta) k_t$$
Market clearing: 
$$y_t = c_t + i_t$$
Productivity shock: 
$$log(A_t) = \eta log(A_{t-1} + (1-\eta)log(\bar{A}) + \varepsilon_t$$

cf. Franck Malherbet's class on *Fluctuations* (or King and Rebelo, 1999)

#### Plan

#### **DSGE** models

A simple RBC model

The simplest neo-Keynesian model

(large) DSGE models in public administrations and international institutions
Linearizing

## The simplest neo-Keynesian model

$$\tilde{y}_{t} = E_{t}(\tilde{y}_{t+1}) - \frac{1}{\sigma}(i_{t} - E_{t}(\pi_{t+1}) - r_{t}^{n})$$
 (1)

$$\pi_t = \beta E_t \left( \pi_{t+1} \right) + \kappa \tilde{y}_t \tag{2}$$

$$i_t = \phi_p \pi_t + \phi_y \tilde{y}_t + \varepsilon_t^i \tag{3}$$

cf. Olivier Loisel's class on *Monetary economics* (or Jordi Galí (2015). *Monetary policy, inflation, and the business cycle: an introduction to the new Keynesian framework and its applications*. Chap. 3)

## RBC-neoK controversy

There is a bit of a *dogmatic* controversy between both types of models...

#### Don't get involved

Keep a **scientific** approach: the model does or does not fit the data, channel X or Y is or is not important in explaining macro fluctuations...

You should not care whether Prof. W or Z had the right intuition decades ago!

Anyway, the *truth* will (dis)agree partially with both so be **nuanced**.

#### Plan

#### DSGE models

A simple RBC model
The simplest neo-Keynesian model

(large) DSGE models in public administrations and international institutions

Linearizing

## The Smets-Wouters/Christiano Eichenbaum and Evans core

The most popular mid-size model is (Smets and Wouters, 2003; Christiano, Eichenbaum, and Evans, 2005) (see Campagne's annex or Challe's class).

- Closed economy
- Capital and labour
- Nominal rigidities (monetary policy has a role) and real rigidities
- Exogenous fiscal policy
- 12 equations 10 shocks 7 observables

## DSGE in policy institutions

DSGE models are very popular in central banks and other economic policy institutions. The most popular mid-size model is Smets-Wouters/CEE (see Campagne's annex or Challe's class).

- IMF-GIMF (Kumhof et al., 2010)
- IMF-GEM (Bayoumi et al., 2004)
- European Commission-Quest (Ratto, Roeger, and Veld, 2009)
- ECB-NAWM (Warne, Coenen, and Christoffel, 2008)
- ECB-EAGLE (Gomes, Jacquinot, and Pisani, 2012)
- Fed-Sigma (Erceg, Guerrieri, and Gust, 2006)
- DG Trésor-Omega3 (Carton and Guyon, 2007)
- Insee-Mélèze (Campagne and Poissonnier, 2016)
- ...

There are libraries of models macromodelbase.com and private collections (J. Pfeifer)

#### Main blocks in a DSGE model

#### Firms

- --> Production function
- → Balancing factors of production
- → Set prices / marginal cost
- → Banking sector (optional)

#### Households

- --> Consume
- → Save-Invest
- --> Supply labour
- → Negotiate wages / marginal productivity
- → Heterogenous agents (optional)
- Policy (fiscal and monetary, optional)
- Laws of nature
  - → Market clearing
  - --> Capital dynamics
  - → Other countries and trade (optional)

## What would you want to include in a DSGE model?

## A general form

A generalized DSGE model (non linear)

$$E_t f(y_{t+1}, y_t, y_{t-1}, v_t) = 0 (4)$$

with ... endogenous variables, ... exogenous variables, and f containing ... equations for *Idots* endogenous variables.

In general only a subset  $y_{t+1}^+$  ( $y_{t-1}^-$ ) of the ..... variables are forward (reps. backward) looking.

The endogenous variables need to be solved for, as a function of the *predetermined variables* and the shocks:  $y_t = g(...,...)$  with g the ".....".

#### Plan

#### DSGE models

A simple RBC model
The simplest neo-Keynesian model
(large) DSGE models in public administrations and
international institutions

Linearizing

## Why linear?

The simplest RBC and neo-Keynesian model are "naturally" linear when taken in logs.

For large DSGE models this cannot be true.

We perform a first order Taylor development around a **Steady State**.

You need to think twice about trends and the definition of equilibrium in your model.

## (Standard) Notations

 $X_t$  a model variable,  $\bar{X}$  its steady state,  $\hat{X}_t = \frac{X_t - \bar{X}}{\bar{X}}$  the deviation from the steady state.

In a linearized model you assume that  $\hat{X} \ll 1$  (e.g. 5% at most) so that  $\hat{X}^2 \simeq 0$  (is negligible).

 $\hat{X}_t = \frac{X_t - \bar{X}}{\bar{X}} \simeq log(X_t) - log(\bar{X})$  (but I recommend that you do not use the log definition of  $\hat{X}$ , cf. pen& paper). You will see why in a minute.

$$Y = f(X) \Longrightarrow \hat{Y} = \frac{f'(\bar{X})\bar{X}}{\bar{Y}}\hat{X}$$

## Pen and paper

Let's take 2 simple equations:

$$Y_t = K_t^{\alpha} (A_t L_t)^{1-\alpha} \tag{5}$$

and the growth rate

$$dY_t = \frac{Y_t - Y_{t-1}}{Y_{t-1}} \approx \log\left(\frac{Y_t}{Y_{t-1}}\right) \tag{6}$$

Let's assume that  $A_t = A_0 e^{at}$  is a deterministic trend. We denote  $y_t = Y_t/A_t$  and  $k_t = K_t/A_t$  the stationary component of output and capital.  $L_t$  is stationary.

Compute a log-linearisation of the 2 equations above

- with  $\hat{y}_t = log(y_t/\bar{y})$
- with  $\hat{y}_t = \frac{y_t \bar{y}}{\bar{y}}$

## Pen and paper - Solution

The first equation gives the same result in both cases:

$$y_t = k_t^{\alpha} \mathcal{L}_t^{1-\alpha}$$
 ;  $\bar{y} = \bar{k}^{\alpha} \bar{\mathcal{L}}^{1-\alpha}$  ;  $\hat{y}_t = \alpha \hat{k}_t + (1-\alpha)\hat{\mathcal{L}}_t$  (7)

The second does not:

$$dY_t \approx log\left(\frac{Y_t}{Y_{t-1}}\right) = log(y_t) - log(y_{t-1}) + a = \hat{y}_t - \hat{y}_{t-1} + a$$
 (8)

$$dY_{t} = \frac{Y_{t} - Y_{t-1}}{Y_{t-1}} = \frac{y_{t}e^{a} - y_{t-1}}{y_{t-1}} = \frac{e^{a}(1 + \hat{y}_{t})}{(1 + \hat{y}_{t-1})} - 1$$

$$\approx e^{a}(1 + \hat{y}_{t} - \hat{y}_{t-1}) - 1 = e^{a}(\hat{y}_{t} - \hat{y}_{t-1}) + (e^{a} - 1)$$

$$(9)$$

#### Linearized outcome

The generalized DSGE model (non linear)

$$E_t f(y_{t+1}, y_t, y_{t-1}, v_t) = 0 (10)$$

can then be turned into a matrix expression (linear)

$$AY_{t} = BE_{t}Y_{t+1} + CY_{t-1} + DV_{t}$$
 (11)

see (Villemot, 2011) in the Dynare documentation to understand how it's done automatically.

Note for the future, linearized model makes an implicit certainty equivalence assumption.

#### Plan

Rational expectations

DSGE models

Solving DSGE models

Wrapping-up

## Solving rational expectation models

• A benchmark (simple) example:

$$y_t = \beta E_t y_{t+1} + \gamma x_t + \varepsilon_t$$

• "Forward "solution :

$$y_t = \beta E_t (\beta E_{t+1} y_{t+2} + \gamma x_{t+1} + \varepsilon_{t+1}) + \gamma x_t + \varepsilon_t$$
$$y_t = \gamma \sum_{i=0}^T \beta^i E_t x_{t+i} + \beta^{T+1} E_t y_{t+T+1} + \varepsilon_t$$

ullet The "fundamental" solution (in the case eta < 1) :

$$y_t = \gamma \sum_{i=0}^{\infty} \beta^i E_t x_{t+i} + \varepsilon_t$$

• An interpretation: Discounted Present Value

## Computing a closed form

Let's assume the model gives us a backward dynamic for  $x_t$  e.g. AR(1)  $x_t = \rho x_{t-1} + \nu_t$ 

$$E_t x_{t+i} = \rho^i x_t$$

Implied reduced form:

Property:

$$y_t = \gamma \frac{1}{(1 - \beta \rho)} x_t + \varepsilon_t$$

## Models with a lagged endogenous variable

#### Model specification

$$y_t = bE_t y_{t+1} + ay_{t-1} + cx_t + \varepsilon_t$$
  
$$x_t = \rho x_{t-1} + u_t$$

#### **Examples:**

- new "hybrid" Phillips curve;
- consumption Euler equation with habit formation in utility,
- models with adjustment costs,...

## Example of a resolution method:

the undetermined coefficient approach

"Guess" the form of the solution:

$$y_t = \varphi y_{t-1} + \alpha x_t + \tilde{\varepsilon}_t$$

This implies

$$E_t y_{t+1} = \varphi y_t + \alpha E_t x_{t+1} = \varphi (\varphi y_{t-1} + \alpha x_t + \tilde{\varepsilon}_t) + \alpha \rho x_t$$
  
=  $\varphi^2 y_{t-1} + \alpha (\varphi + \rho) x_t + \varphi \tilde{\varepsilon}_t$ 

Substitute this in  $y_t = bE_t y_{t+1} + ay_{t-1} + cx_t + \varepsilon_t$ 

This yields a second degree equation for  $\varphi$ :

$$\varphi = b\varphi^2 + a$$

And  $\alpha=c+b\alpha(\varphi+\rho)$  We choose the root such that  $\varphi<1$  It results that  $\varphi=\frac{1-\sqrt{1-4ba}}{2b}, \text{ and } \alpha=\frac{c}{1-b(\varphi+\rho)}$ 

$$y_t = \left(\frac{1 - \sqrt{1 - 4ba}}{2b}\right) y_{t-1} + \left(\frac{c}{1 - b(\varphi + \rho)}\right) x_t + \left(\frac{1}{1 - b\varphi}\right) \varepsilon_t$$

**Shortcoming**: this approach is not systematic

#### The multivariate case

A general formulation for a multivariate model:

$$AY_t = BE_t Y_{t+1} + CY_{t-1} + DX_t$$
  

$$X_t = \Phi X_{t-1} + V_t$$

where A,B,C,D,  $\Phi$  are matrices,  $Y_t$ , a vector of endogenous variables,  $X_t$ , a vector of exogenous variables (e.g. shocks),  $V_t$  i.i.d. inovation

• Other general formulations are possible:

$$AY_t = BE_t Y_{t+1} + DX_t$$
$$X_t = \Phi X_{t-1} + V_t$$

The two formulations are equivalent up to a redifinition of vector  $Y_t$ .

• In general, no analytical solution is available

#### Various procedures that produce numerical soultions:

- (Uhlig, 1995): undetermined coefficient approach
- Using matrices decompositions (Jordan, Schur, QZ...) to provide forward solutions: (Blanchard and Kahn, 1980), (Klein, 2000), (Sims, 2002)

# Sketch of the undetermined coefficient approach (Uhlig, 1995).

Canonical formulation of a multivariate model

$$FE_tY_{t+1} + GY_t + HY_{t-1} + MX_t = 0$$
$$X_t = NX_{t-1} + V_t$$

where F ,G ,H ,M ,N are matrices,  $X_t$ , is a vector of endogenous variables

"Guess" the solution:

$$Y_t = PY_{t-1} + QX_t$$

P and Q fulfill:

$$FP^2 + GP + H = 0 \tag{12}$$

$$(FQ + L)N + (FP + G)Q + M = 0$$
 (13)

 $\rightarrow$  Numerical solution of quadratic equation (12)

## The Blanchard and Kahn (1980) method for the Multivariate case

Formulation of the model

$$E_t Y_{t+1} = AY_t + CX_t$$
$$X_t = \Phi X_{t-1} + V_t$$

- ightarrow Distinguish in  $Y_t$  between variables  $y_{1t}$  predetermined  $(n_1)$  and  $y_{2t}$  non-predetermined  $(\equiv$  no initial condition ) .
- $\rightarrow y_{1t}$  with size  $n_1$ ,  $y_{2t}$  with size  $n_2$

## Note: unicity/stability of the solution

#### Blanchard and Kahn (1980) Conditions:

In the above model:

Let n the number of eigenvalues of matrix A that are larger than 1 If  $n = \text{number of non -predetetermined variables (ie <math>n_2$ ): then the stable soution is unique (saddle-path solution)

If  $n > n_2$  (number of non-predetermined variables): instability (unit-root, explosive dynamics for the variables)

If  $n < n_2$  number of non-predetermined variables : multiplicity of solutions

## Example: the simple model

$$y_t = \beta E_t y_{t+1} + \gamma x_t + \varepsilon_t$$
$$x_t = \rho x_{t-1} + v_t$$

with  $x_t$  predetermined variable,  $y_t$  non predetermined variable,  $Y_t = (x_t, y_t)'$ 

$$E_t Y_{t+1} = A Y_t + V_t$$

with 
$$A = \begin{bmatrix} \rho & 0 \\ -\gamma/\beta & 1/\beta \end{bmatrix}$$

Eigenvalues of A are  $\rho$ ,  $(1/\beta)$ .

Unicity and stability if  $-1 < \rho < 1, -1 < \beta < 1$ 

If  $\beta>1$ , "rational stationary bubbles" are possible (multiplicity of solutions)

Example of a bubble:  $b_t = (1/\beta)b_{t-1} + e_t$ 

## Blanchard and Kahn (1980) method

Formulation of the model

$$E_t Y_{t+1} = AY_t + CX_t$$
$$X_t = \Phi X_{t-1} + V_t$$

• We seek a solution that has the form :

$$y_{2t} = Hy_{1t} + NX_t$$
$$y_{1t} = My_{1t-1} + LX_{t-1}$$

## Blanchard and Kahn (1980) method

The system writes:

$$\begin{bmatrix} y_{1t+1} \\ E_t y_{2t+1} \end{bmatrix} = A \begin{bmatrix} y_{1t} \\ y_{2t} \end{bmatrix} + CX_t$$

Perform a Jordan decomposition of A:

$$A = \Lambda^{-1} J \Lambda$$

J is a bloc-diagonal matrix containing the eigenvalues of A

$$J = \left[ \begin{array}{cc} J_1 & 0 \\ 0 & J_2 \end{array} \right]$$

where  $J_1$  collects eigenvalues smaller than 1 and  $J_2$  collects eigenvalues larger than 1 Note J is diagonal if eigenvalues are distinct, otherwise J contains both "0"'s and "1"'s on the line above the diagonal.

#### Implementing Blanchard and Kahn (1980)

The system can be re written as:

$$\Lambda \left[ \begin{array}{c} y_{1t+1} \\ E_t y_{2t+1} \end{array} \right] = J \Lambda \left[ \begin{array}{c} y_{1t} \\ y_{2t} \end{array} \right] + \Lambda \left[ \begin{array}{c} C_1 \\ C_2 \end{array} \right] X_t$$

We define auxiliary variables

$$\left[\begin{array}{c} \widetilde{y}_{1t} \\ \widetilde{y}_{2t} \end{array}\right] = \left[\begin{array}{cc} \Lambda_{11} & \Lambda_{12} \\ \Lambda_{21} & \Lambda_{22} \end{array}\right] \left[\begin{array}{c} y_{1t} \\ y_{2t} \end{array}\right]$$

Then we can write a "decoupled" system.

$$\begin{bmatrix} \widetilde{y}_{1t+1} \\ E_t \widetilde{y}_{2t+1} \end{bmatrix} = \begin{bmatrix} J_1 & 0 \\ 0 & J_2 \end{bmatrix} \begin{bmatrix} \widetilde{y}_{1t} \\ \widetilde{y}_{2t} \end{bmatrix} + \begin{bmatrix} D_1 \\ D_2 \end{bmatrix} X_t$$

The sub-system associated with eigenvalues larger than one can be solved forward :

$$E_t \widetilde{y}_{2t+1} = J_2 \widetilde{y}_{2t} + D_2 X_t$$

hence

$$\widetilde{y}_{2t} = J_2^{-1} E_t \widetilde{y}_{2t+1} - J_2^{-1} D_2 X_t$$

Solving:

$$\widetilde{y}_{2t} = -\sum_{k=0}^{\infty} J_2^{-1-k} D_2 E_t X_{t+k}$$

Using the properties of the forcing process  $X_t$ : VAR(1)  $E_t X_{t+k} = \Phi^k X_t$ 

$$\widetilde{y}_{2t} = -\sum_{k=0}^{\infty} J_2^{-1-k} D_2 \Phi^k X_t$$

An explicit form for this sum (using properties of the Kronecker product):

$$\widetilde{y}_{2t} = -J_2^{-1}(I - \Phi' \otimes J_2^{-1})D_2X_t$$

Expression as a fonction of the predetermined variables  $y_{2t}=\Lambda_{22}^{-1}\widetilde{y}_{2t}-\Lambda_{22}^{-1}\Lambda_{21}y_{1t}$  So

$$y_{2t} = -\Lambda_{22}^{-1}J_2^{-1}(I - \Phi' \otimes J_2^{-1})D_2X_t - \Lambda_{22}^{-1}\Lambda_{21}y_{1t}$$

This expression is indeed of the form

$$y_{2t} = Hy_{1t} + NX_t$$

#### Last step (!)

Solving for predetermined variables  $y_{1t}$ 

$$y_{1t+1} = A_{11}y_{1t} + A_{12}y_{2t} + D_1X_t$$

where 
$$A = \begin{bmatrix} A_{11} & A_{12} \\ A_{21} & A_{22} \end{bmatrix}$$

Replacing  $y_{2t}$  with its solved form, the process for  $y_{1t}$  is indeed of the form:

$$y_{1t} = My_{1t-1} + LX_{t-1}$$

#### Extensions and alternatives to Blanchard-Kahn method

• (Sims, 2002), (Klein, 2000): use of Schur and QZ decomposition  $BE_t Y_{t+1} = AY_t + CX_t$ 

$$X_t = \Phi X_{t-1} + V_t$$

- $\rightarrow$  In a general model B may be non invertible
- (if B invertible, back to B-K case with  $\widetilde{A} = B^{-1}A$ )
- ightarrow Sims method does not require specification of predetermined/non-predetermined variables
- $\rightarrow$  The various procedures are available in the form of Matlab, Gauss,... routines
- $\rightarrow$  The Sims/Klein method is implemented in the **Dynare** toolbox

### Other procedures

• (Anderson and Moore, 1985) : multiple leads and lags of endogenous variable

$$Y_t = \sum_{j=1}^{J} A_j E_t Y_{t+j} + \sum_{i=1}^{I} B_i Y_{t-i} + V_t$$

AIM Algorithm (Federal Reserve) with Matlab, GAUSS

• DYNARE allows the user to write multiple leads and lags

#### Plan

Rational expectations

DSGE models

Solving DSGE models

 $Wrapping\hbox{-} up$ 

#### Bottom line of this lecture

Start with a non linear model

$$E_t f(y_{t+1}, y_t, y_{t-1}, v_t) = 0$$

Linearize it into

$$AY_t = BE_t Y_{t+1} + CY_{t-1} + DV_t$$

Obtain the reduced form

$$Y_t = MY_{t-1} + D\eta_t$$

You obtain a constrained VAR model • In practice done by computer routines (eg Dynare)

- From there you can do the same as with a SVAR
- + you can derive normative results

## What are A,B,C,D numerically?

#### **Calibration**

- Associated to RBC models.
- Main principle of calibration: set values of parameters (including shocks standard deviations) Then compare model predictions with second moments of the data
- Cf. DeJong and Dave (2011) chap.6
- This approach is not strictly speaking an econometric one.

It can be viewed as a particular case, and less formalized version of these approaches: Minimum Distance Estimation, Bayesian Approach

Next Lectures with HLB, you will discuss the estimation of such models

## Prepare for the next Lectures

- Lecture 5-6: have a look at (Christiano, Eichenbaum, and Evans, 2005)
- Lecture 7: Install Dynare + Matlab/Octave (if not already done)

try the examples provided by the dynare team

#### References I

- Anderson, Gary and George Moore (1985). "A linear algebraic procedure for solving linear perfect foresight models". In: *Economics Letters* 17.3, pp. 247–252.
- Bayoumi, Tamim et al. (Nov. 2004). *GEM: A New International Macroeconomic Model*. IMF Occasional Papers 239. International Monetary Fund.
- Blanchard, Olivier Jean and Charles M Kahn (1980). "The Solution of Linear Difference Models under Rational Expectations". In: *Econometrica* 48.5, pp. 1305–1311.
- Campagne, B. and A. Poissonnier (2016). *MELEZE: A DSGE model for France within the Euro Area*. Documents de Travail de l'Insee INSEE Working Papers g2016-05. Institut National de la Statistique et des Etudes Economiques.
- Canova, Fabio (2011). Methods for applied macroeconomic research. Princeton university press Princeton, NJ.

#### References II

- Carton, Benjamin and Thibault Guyon (2007). Divergences de productivité en union monétaire Présentation du modèle Oméga3. Tech. rep. Technical Report 2007/08, Direction Générale du Trésor et de la Politique Économique.
- Christiano, Lawrence J., Martin Eichenbaum, and Charles L. Evans (2005). "Nominal Rigidities and the Dynamic Effects of a Shock to Monetary Policy". In: *Journal of Political Economy* 113.1, pp. 1–45.
- DeJong, David N and Chetan Dave (2011). Structural macroeconometrics. Princeton University Press.
- Erceg, Christopher J., Luca Guerrieri, and Christopher Gust (2006). "SIGMA: A New Open Economy Model for Policy Analysis". In: *International Journal of Central Banking* 2.1.

#### References III

- Galí, Jordi (2015). Monetary policy, inflation, and the business cycle: an introduction to the new Keynesian framework and its applications. Princeton University Press.
- Gomes, S., P. Jacquinot, and M. Pisani (2012). "The EAGLE. A model for policy analysis of macroeconomic interdependence in the euro area". In: *Economic Modelling* 29.5, pp. 1686–1714.
- Juillard, Michel and Tarik Ocaktan (2008). "Méthodes de simulation des modèles stochastiques d'équilibre général". In: Economie & Prévision 0.2, pp. 115–126.
  - King, Robert G. and Sergio T. Rebelo (1999). "Resuscitating real business cycles". In: *Handbook of Macroeconomics*. Ed. by J. B. Taylor and M. Woodford. Vol. 1. Handbook of Macroeconomics. Elsevier. Chap. 14, pp. 927–1007.

#### References IV

- Klein, Paul (2000). "Using the generalized Schur form to solve a multivariate linear rational expectations model". In: *Journal of Economic Dynamics and Control* 24.10, pp. 1405–1423.
- Kumhof, Michael et al. (2010). "The global integrated monetary and fiscal model (GIMF)- Theoretical structure". In: *IMF Working Paper*.
- Lindé, Jesper and Mathias Trabandt (2018). "Should we use linearized models to calculate fiscal multipliers?" In: Journal of Applied Econometrics 33.7, pp. 937–965.
- Lucas, Robert Jr (1976). "Econometric policy evaluation: A critique". In: Carnegie-Rochester Conference Series on Public Policy 1.1, pp. 19–46.

#### References V

- Ratto, Marco, Werner Roeger, and Jan in 't Veld (2009). "QUEST III: An estimated open-economy DSGE model of the euro area with fiscal and monetary policy". In: *Economic Modelling* 26.1, pp. 222–233.
- Sims, Christopher A (2002). "Solving Linear Rational Expectations Models". In: Computational Economics 20.1-2, pp. 1–20.
- Smets, Frank and Raf Wouters (2003). "An Estimated Dynamic Stochastic General Equilibrium Model of the Euro Area". In:

  Journal of the European Economic Association 1.5,
  pp. 1123–1175.
- Uhlig, Harald (1995). A toolkit for analyzing nonlinear dynamic stochastic models easily. Discussion Paper / Institute for Empirical Macroeconomics 101. Federal Reserve Bank of Minneapolis.

#### References VI



Villemot, Sébastien (Apr. 2011). Solving rational expectations models at first order: what Dynare does. Dynare Working Papers 2. CEPREMAP.



Warne, Anders, Günter Coenen, and Kai Christoffel (Oct. 2008). The new area-wide model of the euro area: a micro-founded open-economy model for forecasting and policy analysis. Working Paper Series 944. European Central Bank.

## An alternative solution method : Factorisation

Rely on "Forward" operator:

$$Fy_t = E_t y_{t+1}$$

Benchmark equation writes:

$$y_t - bFy_t - aLy_t = cx_t + \varepsilon_t$$

hence

$$P(F)Ly_t = -(1/b)(cx_t + \varepsilon_t)$$

where 
$$P(F) = F^2 - (1/b)F + (a/b)$$
  
  $P(F)$  can be factored  $P(F) = (F - \varphi_1)(F - \varphi_2)$  with  $\varphi_1 < 1$ 

Dividing by polynomial  $(F - \varphi_2)$  $\rightarrow$  Solving "forward" for the root  $\varphi_2 > 1$ 

$$(1 - \varphi_1 L)y_t = \frac{1}{b\varphi_2} \frac{1}{1 - (1/\varphi_2)F} (\beta x_t + \varepsilon_t)$$

where  $\varphi_1=\frac{1-\sqrt{1-4ba}}{2b}$  (note  $\varphi_1=\varphi$ , same as the undetermined coefficient approach) and

$$y_t = \varphi_1 y_{t-1} - \frac{1}{b\varphi_2} \sum_{k=0}^{\infty} (1/\varphi_2)^k E_t(cx_{t+k} + \varepsilon_{t+k})$$
  
Using  $E_t(x_{t+k}) = \rho^h x_t$ , we find (as in the UC approach):

$$\alpha = \frac{c}{1 - b(\varphi_1 + \rho)}$$