

Rural Windfall or a New Resource Curse? Coca, Income, and Civil Conflict in Colombia

Based on Angrist & Kugler (2008)

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```
knitr::opts_chunk$set(echo = TRUE, warning = FALSE, message = FALSE,
                      fig.width = 10, fig.height = 6)

# Setup
if (!require(haven)) install.packages("haven")
library(haven)
if (!require(dplyr)) install.packages("dplyr")
library(dplyr)
if (!require(foreign)) install.packages("foreign")
library(foreign)
if (!require(plm)) install.packages("plm")
library(plm)
if (!require(stargazer)) install.packages("stargazer")
library(stargazer)
if (!require(ggplot2)) install.packages("ggplot2")
library(ggplot2)
```

```

if (!require(sandwich)) install.packages("sandwich")
library (sandwich)
if (!require(lmtest)) install.packages("lmtest")
library (lmtest)
if (!require(tidyverse)) install.packages("tidyverse")
library (tidyverse)
if (!require(BART)) install.packages("BART")
library (BART)
if (!require(grf)) install.packages("grf")
library (grf)
if (!require(car)) install.packages("car")
library (car)

# Load Data and take a look at the dataset
dta <- read_delim("data00_AngristKugler.tab", delim = "\t")

```

1 Q1. Setup and Data Construction

Tasks:

1. Create grow variable (1 if $\text{dep_ocu} \in \{13, 18, 19, 50, 52, 86, 95, 97, 99\}$, 0 otherwise)

```

# Creating a new variable grow
department_list <- c(13, 18, 19, 50, 52, 86, 95, 97, 99)

dta <- dta %>%
  mutate(grow = ifelse(dep_ocu %in% department_list, 1, 0))

```

2. Subset data to years 1991, 1992, 1993 and 1996, 1997, 1998

```

# Subsetting the dataset to years 1991 - 1993 and 1996 - 1998
dta_subset <- dta %>%
  filter(year %in% c(1991, 1992, 1993, 1996, 1997, 1998))

```

3. Create after variable (1 if $\text{year} \in \{1996, 1997, 1998\}$, 0 otherwise)

```

# We create a variable called after with 1 for years 1996 - 1998
dta_subset <- dta_subset %>%
  mutate(after = ifelse(year %in% c(1996, 1997, 1998), 1, 0))

```

4. Create growafter variable ($\text{grow} \times \text{after}$)

```

# Creating growafter variable (grow * after)
dta_subset <- dta_subset %>%
  mutate(growafter = grow * after)

dta_subset %>%
  count(grow, after, growafter)

```

```

## # A tibble: 4 x 4
##   grow after growafter     n
##   <dbl> <dbl>     <dbl> <int>
## 1     0     0         0  2491
## 2     0     1         0  2566

```

```
## 3      1      0      0    852
## 4      1      1      1    907
```

5. Create outcome variable: $\log\left(\frac{\text{populati}+1}{\text{violent}+1}\right)$

```
# Create outcome variable
```

```
dta_subset <- dta_subset %>%
  mutate(outcome = log((violent + 1) / (populati + 1)))
```

2 Q2. Visualizing Violence Before and After

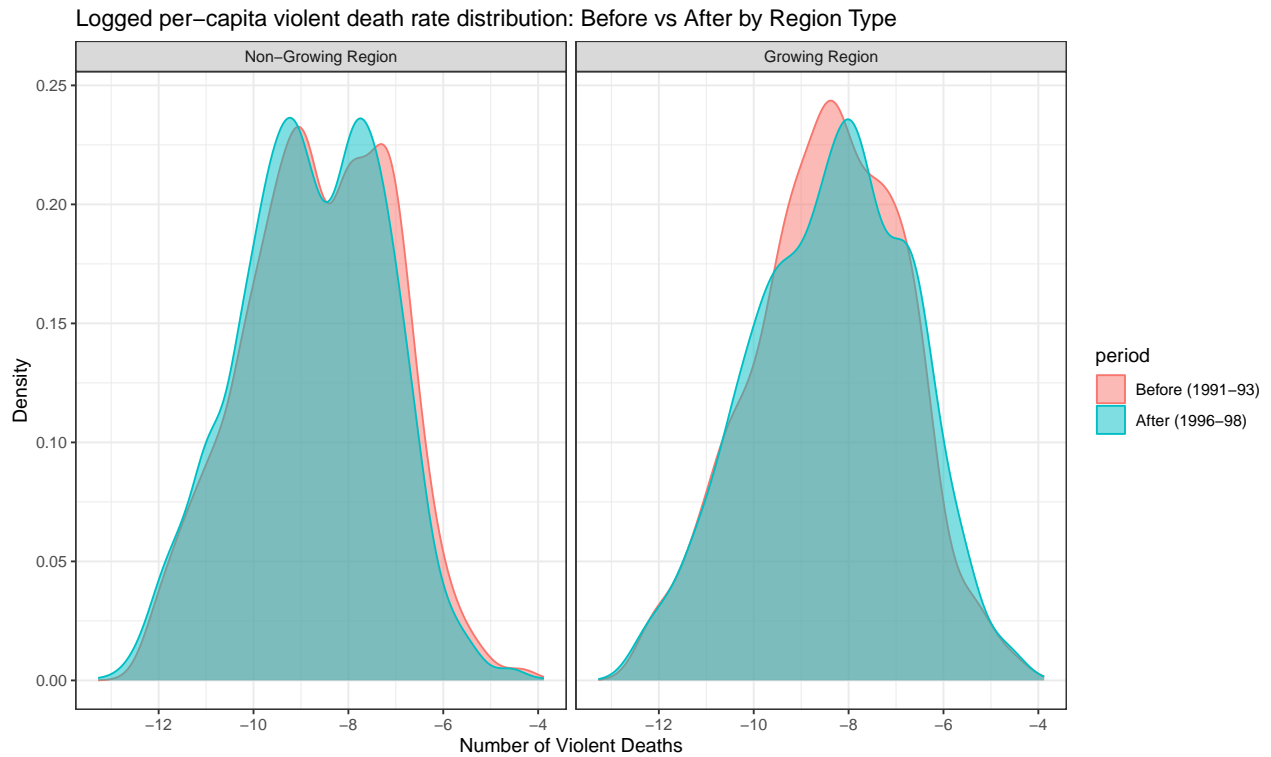
Tasks:

1. Create density plots for non-growing vs. growing regions, before vs. after
2. Extend to 2×2 grid by gender (men: `sex=1`, women: `sex=2`)
3. Interpret: Evidence of shifts in violence? Different by gender?

```
# Density plots: Non-growing vs. Growing regions
```

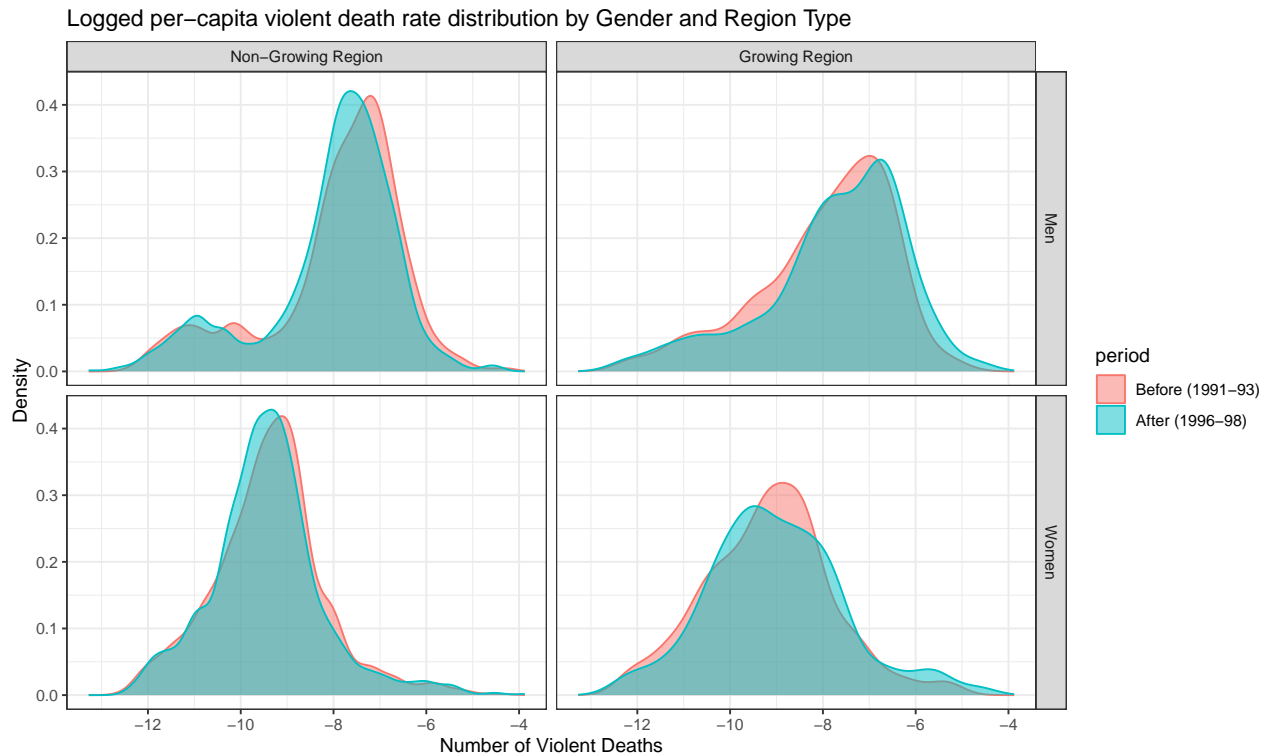
```
density_plot <- dta_subset %>%
  mutate(
    period = factor(after, levels = c(0, 1), labels = c("Before (1991-93)", "After (1996-98)")),
    region_type = factor(grow, levels = c(0, 1), labels = c("Non-Growing Region", "Growing Region")),
    gender = factor(sex, levels = c(1, 2), labels = c("Men", "Women"))
  )
```

```
density_plot %>%
  ggplot(aes(x = outcome, fill = period, color = period)) +
  geom_density(alpha = 0.5) +
  facet_wrap(~ region_type) +
  labs(title = "Logged per-capita violent death rate distribution: Before vs After by Region Type",
       x = "Number of Violent Deaths",
       y = "Density") +
  theme_bw()
```



2x2 grid: Top=men, Bottom=women; Left=non-growing, Right=growing

```
density_plot %>%
  filter(!is.na(gender)) %>%
  ggplot(aes(x = outcome, fill = period, color = period)) +
  geom_density(alpha = 0.5) +
  facet_grid(gender ~ region_type) +
  labs(title = "Logged per-capita violent death rate distribution by Gender and Region Type",
        x = "Number of Violent Deaths",
        y = "Density") +
  theme_bw()
```



Interpretation:

The evidence indicates a significant shift in violence following the air-bridge disruption, with the effect being highly specific to both region and gender. For the treatment group (**coca-growing regions**), there was a slight increase in the per-capita violent death rate among men. In contrast, the control group (**non-growing regions**) showed no meaningful change for either gender, which suggests that the increase in violence was not due to a nationwide trend. Furthermore, the pattern seems to be strongly gendered; the effect on women in the treatment group was minimal compared to a much larger effect on men. This suggests that the impact of the coca boom on violence was almost exclusively concentrated among the male population, who seem to have been the primary participants in the conflict.

3 Q3. Age-Specific Effects

Task: For coca-growing regions only, plot the change in outcome (after - before) by age group.

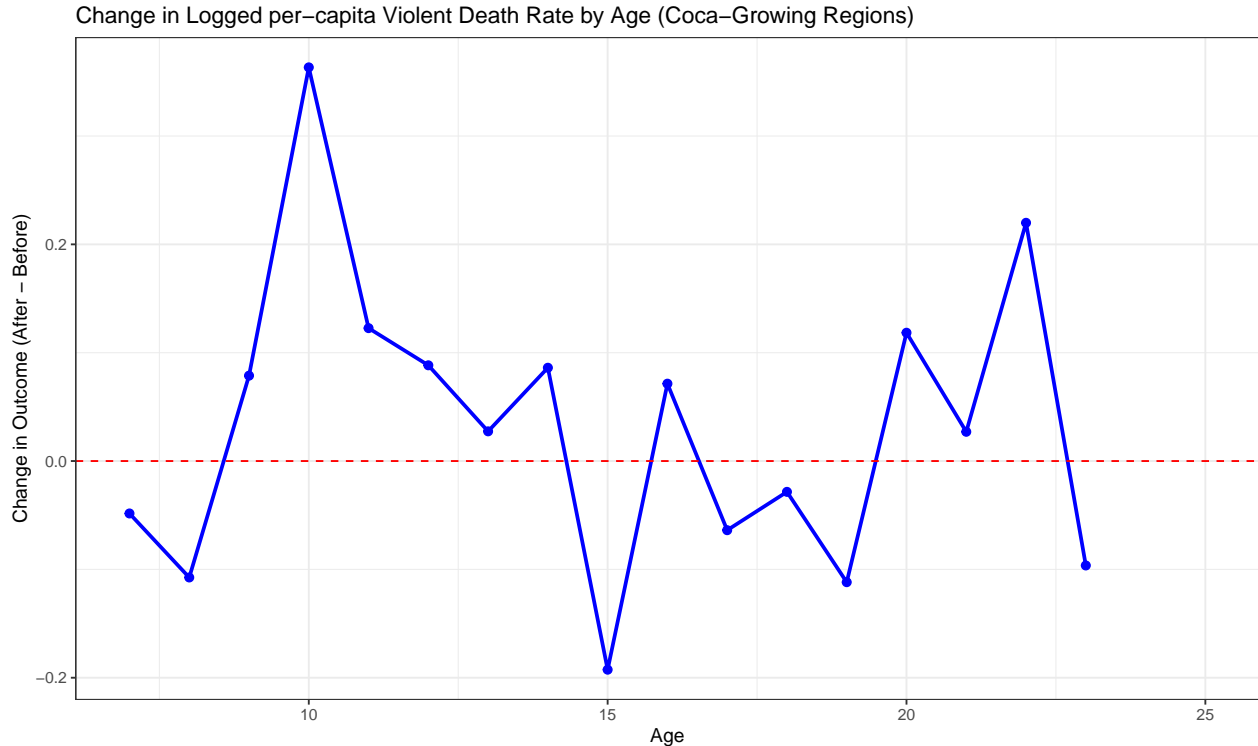
Calculate mean difference by age for growing regions only

```
age_effects <- dta_subset %>%
  filter(grow == 1) %>%
  group_by(age, after) %>%
  summarise(mean_outcome = mean(outcome, na.rm = TRUE), .groups = "drop") %>%
  pivot_wider(names_from = after, values_from = mean_outcome, names_prefix = "period_") %>%
  mutate(change = period_1 - period_0)
```

Plot age-specific effects

```
age_effects %>%
  ggplot(aes(x = age, y = change)) +
  geom_line(color = "blue", size = 1) +
  geom_point(color = "blue", size = 2) +
```

```
geom_hline(yintercept = 0, linetype = "dashed", color = "red") +
labs(title = "Change in Logged per-capita Violent Death Rate by Age (Coca-Growing Regions)",
     x = "Age",
     y = "Change in Outcome (After - Before)") +
theme_bw()
```



Interpretation:

This plot illustrates how the average violence rate changes for each age group by calculating the difference between the average violence rate for an age group after the air-bridge disruption and the average violence rate for that same group before the disruption. Clearly, the plot shows that the effect varies dramatically across different age groups, with a concentration in certain age brackets. The increase in conflict and violence seems to have disproportionately affected individuals of fighting age, as can be seen from the peaks in the graph between the ages of 17 and 22. Additionally, younger adolescents around the age of 10 seem to have been particularly involved in violent activities, resulting in a much higher death rate after the air disruption. However, it is important to note that, for a few age groups — especially those around 15 years old — the rate of violence decreased after the air disruption.

This seemingly volatile distribution across age groups reveals a key insight. If the increase in violence was solely due to increased criminal activity, we would most likely see it concentrated among cohorts that met the age of fighting and above. However, since we clearly see that also much younger age groups show an increase in violent deaths, we can conclude that this is mostly caused by conflict-related violence (e.g. civil war) after the ‘air bridge’ intervention, which corresponds to the paper’s findings.

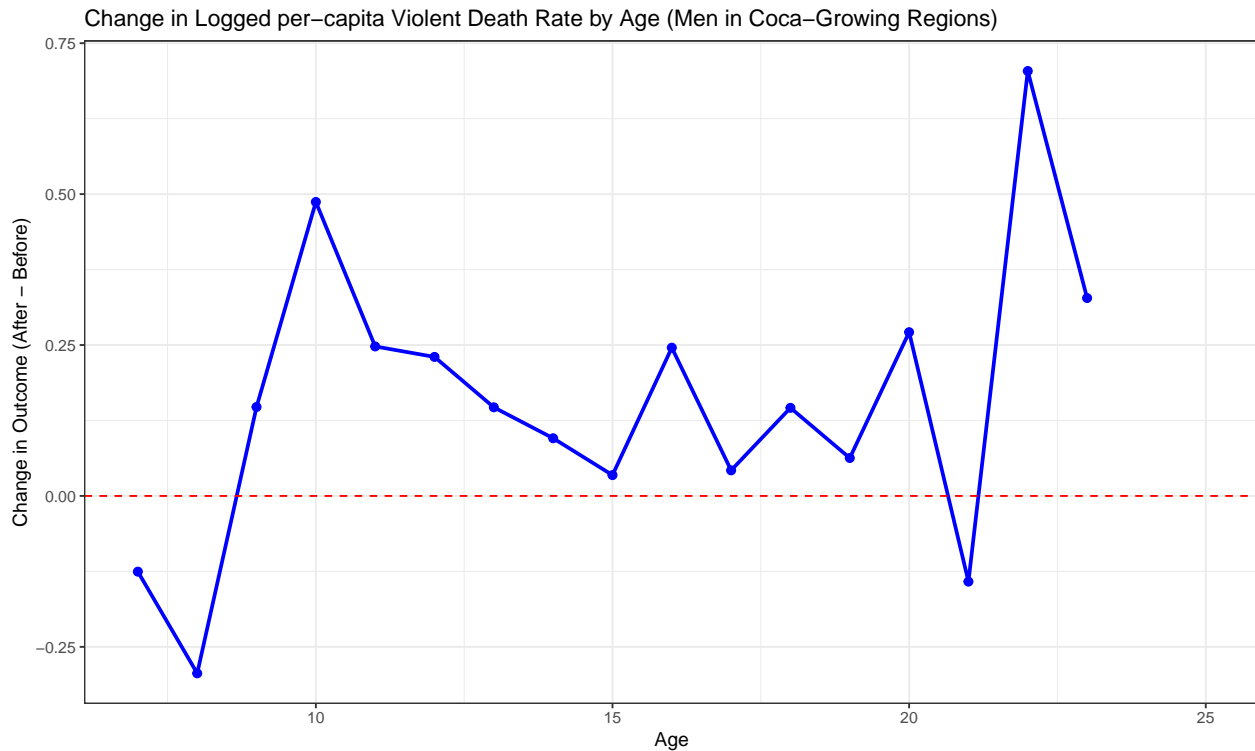
To explore this interesting trend, we created a graph for the per-capita violent death rate in coca-growing regions before and after for men only. What we can see here is that the increase in violence was not confined by the typical “fighting age” bracket but was high across the entire young male distribution, including children around 10. This pattern strongly suggests widespread, conflict-related violence. The coca boom fueled this by creating two distinct sets of victims: older cohorts were recruited as soldiers, while younger boys, drawn in as laborers, became collateral damage in the armed groups’ fight to control the coca fields and labor force

```

age_effects_men <- dta_subset %>%
  filter(grow == 1, sex == 1) %>% # <- HIER IST DIE ÄNDERUNG
  group_by(age, after) %>%
  summarise(
    mean_outcome = mean(outcome, na.rm = TRUE),
    .groups = "drop"
  ) %>%
  pivot_wider(
    names_from = after,
    values_from = mean_outcome,
    names_prefix = "period_"
  ) %>%
  mutate(change = period_1 - period_0)

# Plot der altersspezifischen Effekte für Männer
age_effects_men %>%
  ggplot(aes(x = age, y = change)) +
  geom_line(color = "blue", size = 1) +
  geom_point(color = "blue", size = 2) +
  geom_hline(
    yintercept = 0,
    linetype = "dashed",
    color = "red"
  ) +
  labs(
    title = "Change in Logged per-capita Violent Death Rate by Age (Men in Coca-Growing Regions)", # <-
    x = "Age",
    y = "Change in Outcome (After - Before)"
  ) +
  theme_bw()

```



4 Q4. Testing the Parallel Trends Assumption (Pre-Treatment)

Tasks:

1. Use pre-treatment years (1990-1993)
2. Estimate: $\text{outcome} = \alpha + \beta \cdot \text{year} + \gamma \cdot \text{grow} + \delta \cdot (\text{grow} \times \text{year}) + u$
3. Test if $\text{grow} \times \text{year}$ interactions are jointly zero (year as linear and categorical)
4. Create graph of average outcome by year and group

```
# Subset to pre-treatment years (1990-1993)
dta_pretreatment <- dta %>%
  filter(year %in% c(1990, 1991, 1992, 1993)) %>%
  mutate(grow = ifelse(dep_ocu %in% department_list, 1, 0),
         outcome = log((violent + 1) / (populati + 1)))

# Model with year as linear
year_linear <- lm(outcome ~ year + grow + year:grow, data = dta_pretreatment)

summary(year_linear)
```

```
##
## Call:
## lm(formula = outcome ~ year + grow + year:grow, data = dta_pretreatment)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -3.972 -1.057  0.029  1.214  4.494
##
## Coefficients:
```



```
##               Estimate Std. Error t value Pr(>|t|)
## (Intercept) -1.469e+01  4.921e+01  -0.299   0.765
## year        3.043e-03  2.471e-02   0.123   0.902
## grow        1.469e+00  9.748e+01   0.015   0.988
## year:grow   -6.358e-04  4.895e-02  -0.013   0.990
##
## Residual standard error: 1.552 on 4221 degrees of freedom
## (238 observations deleted due to missingness)
## Multiple R-squared:  0.003242, Adjusted R-squared:  0.002534
## F-statistic: 4.577 on 3 and 4221 DF, p-value: 0.003328

# Model with year as categorical (factor)
year_factor <- lm(outcome ~ factor(year) + grow + factor(year):grow, data = dta_pretreatment)

summary(year_factor)

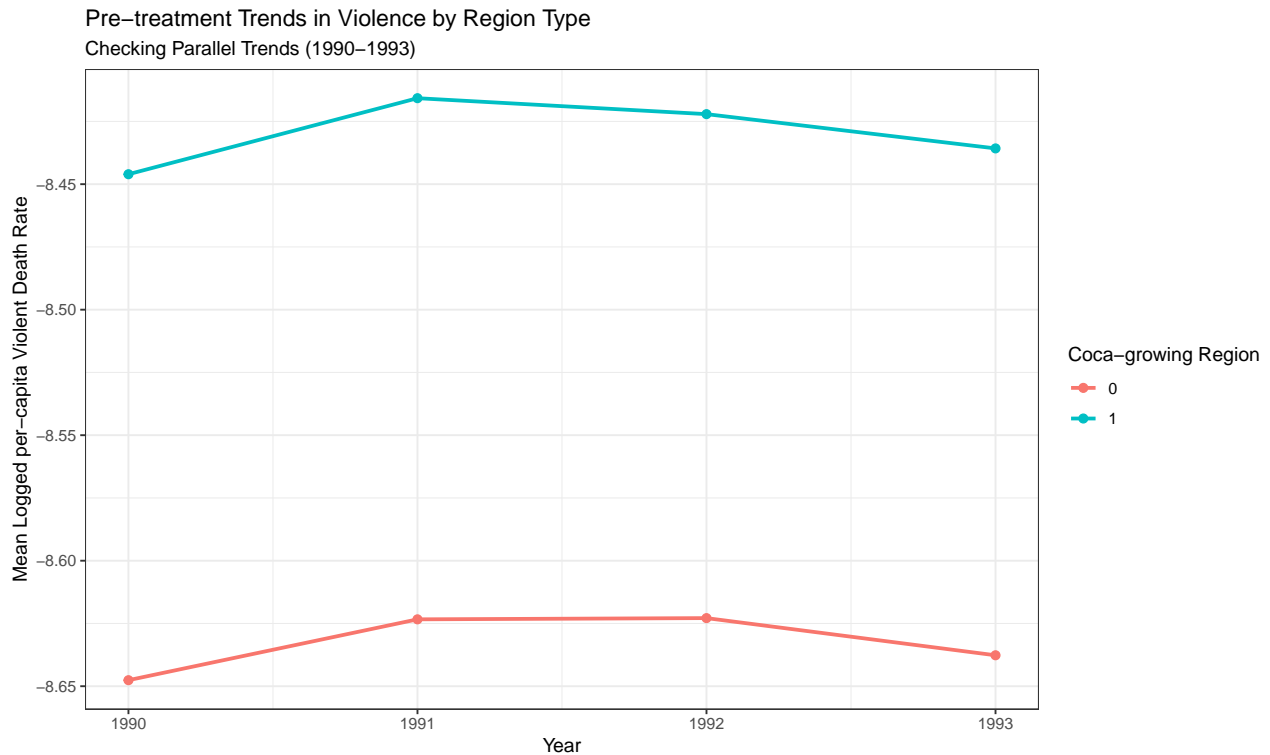
##
## Call:
## lm(formula = outcome ~ factor(year) + grow + factor(year):grow,
##     data = dta_pretreatment)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -3.9625 -1.0596  0.0296  1.2174  4.5045
##
## Coefficients:
##               Estimate Std. Error t value Pr(>|t|)
## (Intercept)    -8.6476041  0.0551588 -156.776 <2e-16 ***
## factor(year)1991  0.0242450  0.0780805   0.311   0.756
## factor(year)1992  0.0247314  0.0782302   0.316   0.752
## factor(year)1993  0.0099279  0.0781551   0.127   0.899
## grow           0.2015622  0.1100062   1.832   0.067 .
## factor(year)1991:grow 0.0060478  0.1560531   0.039   0.969
## factor(year)1992:grow -0.0007938  0.1549368  -0.005   0.996
## factor(year)1993:grow 0.0003595  0.1547948   0.002   0.998
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.552 on 4217 degrees of freedom
## (238 observations deleted due to missingness)
## Multiple R-squared:  0.003286, Adjusted R-squared:  0.001631
## F-statistic: 1.986 on 7 and 4217 DF, p-value: 0.05322

# Test if grow*year interactions are jointly zero
linearHypothesis(year_factor, matchCoefs(year_factor, ":grow"))

##
## Linear hypothesis test:
## factor(year)1991:grow = 0
## factor(year)1992:grow = 0
## factor(year)1993:grow = 0
##
## Model 1: restricted model
## Model 2: outcome ~ factor(year) + grow + factor(year):grow
##
```

```
##   Res.Df   RSS Df Sum of Sq    F Pr(>F)
## 1    4220 10162
## 2    4217 10162   3 0.0058328 8e-04    1

# Graph: Average outcome by year and group
dta_pretreatment %>%
  group_by(year, grow) %>%
  summarise(mean_outcome = mean(outcome, na.rm = TRUE), .groups = "drop") %>%
  ggplot(aes(x = year, y = mean_outcome, color = factor(grow))) +
  geom_line(size = 1) +
  geom_point(size = 2) +
  labs(
    title = "Pre-treatment Trends in Violence by Region Type",
    subtitle = "Checking Parallel Trends (1990-1993)",
    x = "Year",
    y = "Mean Logged per-capita Violent Death Rate",
    color = "Coca-growing Region"
  ) +
  theme_bw()
```



Interpretation:

To assess the validity of the parallel trend assumption, we focused on the pre-treatment period from 1990 to 1993. Both regression models formally test this parallel trends assumption by examining if the pre-treatment trends in violence are statistically different between the coca-growing regions and non-growing regions. The first model simplifies our analysis by assuming that time follows a linear trend. The key coefficient here is the interaction term **year:grow**, which measures the difference in the slopes of trend lines for the two groups. The p-value for this term is 0.99, which indicates that there is no statistically significant difference between the slopes. The second model offers a more flexible, year-by-year analysis by checking if the gap in violence between the two groups changed in any year from 1991 to 1993 compared to the baseline year of 1990, instead of assuming a straight line. The additional, joint F-test on these interactions yields a p-value

of 1 which is why we fail to reject the null hypothesis H_0 that the pre-treatment trends are perfectly parallel. This result is crucial because it validates the use of non-growing regions as a credible counterfactual. It gives us confidence that any divergence between the groups after the disruption is due to the coca boom itself, rather than pre-existing differences in trends.

5 Q5. Placebo DiD Test

Tasks:

1. Create `placebo_after` (1 if year = 1992 or 1993, 0 if year = 1990 or 1991)
2. Estimate placebo DiD model
3. Interpret `placebo_after` \times `grow` coefficient

```
# Subset and create placebo variables
dta_placebo <- dta %>%
  filter(year %in% c(1990, 1991, 1992, 1993)) %>%
  mutate(
    grow = ifelse(dep_ocu %in% department_list, 1, 0),
    outcome = log((violent + 1) / (populati + 1)),
    placebo_after = ifelse(year %in% c(1992, 1993), 1, 0)
  )

# Estimate placebo DiD model
placebo_did <- lm(outcome ~ placebo_after + grow + placebo_after:grow, data = dta_placebo)
summary(placebo_did)
```

```
##
## Call:
## lm(formula = outcome ~ placebo_after + grow + placebo_after:grow,
##     data = dta_placebo)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -3.9699 -1.0574  0.0288  1.2145  4.4924
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   -8.635505   0.039023 -221.295 < 2e-16 ***
## placebo_after    0.005216   0.055292   0.094  0.92485
## grow           0.204494   0.077990   2.622  0.00877 **
## placebo_after:grow -0.003148  0.109627  -0.029  0.97710
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.552 on 4221 degrees of freedom
## (238 observations deleted due to missingness)
## Multiple R-squared:  0.00324,    Adjusted R-squared:  0.002532
## F-statistic: 4.574 on 3 and 4221 DF,  p-value: 0.003342
```

Interpretation:

[Should the placebo effect be significant? What would significance suggest?]

6 Q6. Covariate Balance at Time 0

Task: Compare treatment and control regions on age, sex, and population using pre-treatment data.

```
# Create balance table
dta_balance <- dta_subset %>%
  filter(after == 0)

balance_table <- dta_balance %>%
  group_by(grow) %>%
  summarise(
    mean_age = mean(age, na.rm = TRUE),
    mean_sex = mean(sex, na.rm = TRUE),
    mean_pop = mean(population, na.rm = TRUE),
    .groups = "drop"
  )

balance_table
```

```
## # A tibble: 2 x 4
##   grow mean_age mean_sex mean_pop
##   <dbl>   <dbl>   <dbl>   <dbl>
## 1     0    15.6     1.51   78301.
## 2     1    15.5     1.59   40541.
```

```
# Optional: Standardized difference plot
```

Discussion:

[Why is covariate balance critical? What would imbalance imply?]

7 Q7. Why Covariate Balance Matters

Discussion Questions:

1. If covariates are balanced at time 0, what does this imply about confounding?
2. What role do these variables play after assignment?
3. If violence trends already differ before treatment, how might this bias DiD?

[Your answers here]

8 Q8. Covariate Timing and Post-Treatment Bias

Discussion Questions:

1. Should we include covariates from time 0, time 1, or both?
2. What happens if you include a covariate measured after treatment?
3. When might adjusting for post-treatment variables be appropriate?

[Your answers here]

9 Q9. Computing the DiD Estimate

Task: Compute manual DiD estimate.

```
# Mean difference (after - before) for grow=1 and grow=0
did_table <- dta_subset %>%
  group_by(grow, after) %>%
  summarise(mean_outcome = mean(outcome, na.rm = TRUE), .groups = "drop") %>%
  pivot_wider(names_from = after, values_from = mean_outcome, names_prefix = "time") %>%
  mutate(diff = time1 - time0)

# DiD estimate: subtract the two
did_estimate <- diff(did_table$diff)
did_estimate
```

```
## [1] 0.1970235
```

Interpretation:

[Why is DiD preferable to simple before-after comparison?]

10 Q10. Regression Form of DiD

Tasks:

1. Estimate: $\text{outcome} = \beta_0 + \beta_1 \cdot \text{after} + \beta_2 \cdot \text{grow} + \beta_3 \cdot (\text{after} \times \text{grow}) + u$
2. Report β_3 and p-value
3. Show analytically that β_3 equals the manual DiD estimate

```
# DiD regression model
did <- lm(outcome ~ after + grow + after:grow, data = dta_subset)
summary(did)

##
## Call:
## lm(formula = outcome ~ after + grow + after:grow, data = dta_subset)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -4.5075 -1.0675  0.0344  1.2000  4.8920
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -8.62797    0.03193 -270.247 < 2e-16 ***
## after       -0.14132    0.04485  -3.151  0.00163 **
## grow         0.20330    0.06317   3.218  0.00130 **
## after:grow   0.19702    0.08820   2.234  0.02554 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.55 on 6445 degrees of freedom
## (367 observations deleted due to missingness)
## Multiple R-squared:  0.008903, Adjusted R-squared:  0.008441
## F-statistic: 19.3 on 3 and 6445 DF, p-value: 1.877e-12
```

Interpretation:

[What does β_3 tell us about the causal effect?]

Analytical proof:

[Show that $\beta_3 = (\bar{Y}_{1,1} - \bar{Y}_{1,0}) - (\bar{Y}_{0,1} - \bar{Y}_{0,0})$]

11 Q11. Adding Covariates

Task: Estimate three models and compare.

```
# Model 1: outcome ~ grow + after + growafter
cov_did1 <- lm(outcome ~ after + grow + after:grow, data = dta_subset)

# Model 2: Add age and sex
cov_did2 <- lm(outcome ~ after + grow + after:grow + age + sex, data = dta_subset)

# Model 3: Add age, sex, and populati
cov_did3 <- lm(outcome ~ after + grow + after:grow + age + sex + populati, data = dta_subset)

# Compare models
stargazer(cov_did1, cov_did2, cov_did3,      # List your models
           type = "text",                    # Output type: "text", "html", or "latex"
           title = "Regression Results: The Effect of Treatment on Outcome",
           align = TRUE,                     # Aligns numbers on decimal points
           dep.var.labels = "Outcome",       # A clean name for the dependent variable
           column.labels = c("Base Model", "Adds Demographics", "Full Model"),
           covariate.labels = c("After Treatment", "Treatment Group (Grow)", "Age",
                                "Sex", "Population", "Interaction: After x Grow"),
           notes = "Standard errors are in parentheses.",
           notes.align = "l")
```

```
##
## Regression Results: The Effect of Treatment on Outcome
## =====
##                                     Dependent variable:
## -----
##                                     Outcome
##                                     Adds Demographics
##                                     Full Model
##                                     (1)                (2)                (3)
## -----
## After Treatment                -0.141***          -0.123***          -0.107***
##                                (0.045)             (0.036)             (0.036)
##
## Treatment Group (Grow)         0.203***          0.304***          0.231***
##                                (0.063)             (0.051)             (0.051)
##
## Age                            0.142***          0.123***
##                                (0.003)             (0.004)
##
## Sex                           -1.057***          -1.057***
##                                (0.028)             (0.028)
##
## Population                     -0.00000***
##                                (0.00000)
```

```

## Interaction: After x Grow      0.197**          0.131*          0.123*
##                               (0.088)          (0.071)          (0.071)
##
## Constant                     -8.628***         -9.162***         -8.731***
##                               (0.032)          (0.070)          (0.079)
##
## -----
## Observations                  6,449            6,449            6,449
## R2                           0.009            0.350            0.363
## Adjusted R2                  0.008            0.350            0.362
## Residual Std. Error          1.550 (df = 6445)    1.256 (df = 6443)    1.244 (df = 6442)
## F Statistic                  19.298*** (df = 3; 6445) 693.978*** (df = 5; 6443) 610.532*** (df = 6; 6442)
## =====
## Note:                        *p<0.1; **p<0.05; ***p<0.01
##                               Standard errors are in parentheses.

```

Discussion:

[Does β_3 change? Do covariates matter? Which specification is most credible?]

12 Q12. Interpretation and Reflection

Summary:

1. Did violence increase or decrease after the air-bridge disruption?
2. Does the evidence support a “resource-curse” interpretation?
3. What are the remaining identification threats?

[Your final interpretation here]