Tutorial Week 8 and 9: Matching and Entropy Balancing

Problem Set - Propensity Scores, Matching, and Synthetic Control $\,$

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```
library(sandwich)
library(knitr)
library(kableExtra)
library(broom)
library(Synth)
```

2 Part I: Propensity Scores, Matching, and Robust Post-Matching Inference

2.1 Background

Research Question: Do United Nations interventions help shorten the duration of civil wars?

Gilligan and Sergenti (2008) use matching methods to re-evaluate earlier findings suggesting that UN interventions prolong conflict. They argue that this conclusion stems from **selection bias** — the UN tends to intervene in the worst conflicts.

Dataset: war_pre_snapshots.dta

Each row represents a **conflict episode** observed before a potential UN intervention. Our goal is to estimate the causal effect of UN involvement (UN) on the length of the conflict (t1 - t0), while balancing on key pre-treatment covariates.

2.2 Load Data

```
# Read the UN intervention dataset
data_un <- read_dta("war_pre_snapshots.dta")

# Display structure and summary
#glimpse(data_un)
#summary(data_un)</pre>
```

2.3 Define Treatment and Covariates

```
# Treatment variable: UN intervention (1 = Yes, 0 = No)
treat <- data_un$UN

# Outcome variable: conflict duration (t1 - t0)
outcome <- data_un$t1 - data_un$t0

# Covariates for propensity score model
covar <- c(
    "inter", "deaths", "couprev", "sos", "drugs", "t0",
    "ethfrac", "pop", "lmtnest", "milper",
    "eeurop", "lamerica", "asia", "ssafrica"
)

# Create covariate matrix
X <- data_un[, covar]</pre>
```

2.4 Q0. First Check of the Data

2.4.1 Tasks:

- 1. Why might UN interventions **not** be randomly assigned across conflicts?
- 2. Which of the listed variables are most likely to confound the relationship between UN and conflict duration? Run a quick logistic regression and check.

Logistic regression: UN intervention as function of covariates

Discussion:

[Why might UN interventions not be randomly assigned? Which variables show strong associations with UN intervention?]

2.5 Q1. Estimating Propensity Scores

2.5.1 Theoretical Background

Let $T_i \in \{0,1\}$ be the treatment indicator and $X_i = (X_{i1}, X_{i2}, \dots, X_{ip})$ the vector of pre-treatment covariates. The **propensity score** is defined as:

$$e(X_i) = P(T_i = 1 \mid X_i)$$

2.5.2 Tasks:

- 1. Define the propensity score
- 2. Estimate $\hat{e}(X_i)$ in two ways:
 - (a) Logistic regression: $logit(e(X_i)) = X_i'\beta$
 - (b) Random forest classifier
- 3. Report mean, SD, and range of $\hat{e}(X_i)$ for treated and control
- 4. Create histogram/density plot by treatment status
- # (a) Logistic regression propensity score
- # Extract predicted probabilities
- # (b) Random forest propensity score
- # Extract predicted probabilities
- # Summary statistics of propensity scores
- # Density plot of propensity scores by treatment status

Interpretation:

[Discuss the distribution of propensity scores. Are there regions of poor overlap?]

2.6 Q2. Implement 1:1 Nearest-Neighbor Matching

2.6.1 Matching Setup

For each estimated propensity score $\hat{e}(X_i)$, match each treated unit to the nearest control on the **logit of** the propensity score:

$$\ell_i = \log \left(\frac{\hat{e}(X_i)}{1 - \hat{e}(X_i)} \right)$$

Use **replacement** and a **caliper** of $0.2 \times SD(\ell_i)$ to restrict poor matches.

2.6.2 Tasks:

- 1. Implement matching using both logit and RF propensity scores
- 2. Report how many treated units fail to find a match
- 3. How does this change the estimand?
- # Matching using logit propensity scores
- # Number of matched treated units
- # Matching using RF propensity scores
- # Number of matched treated units

Discussion:

[How many treated units were dropped? What does this mean for the target estimand (ATT)?]

2.7 Q3. Standardized Mean Differences (SMDs)

2.7.1 SMD Formulas

For each covariate X^k :

Before matching (ATT version):

$${\rm SMD_{raw}}(k) = \frac{\bar{X}_{T=1}^k - \bar{X}_{T=0}^k}{\sqrt{s_{T=1}^{2,k}}}$$

After matching:

$$\mathrm{SMD}_{\mathrm{match}}(k) = \frac{\bar{X}_{\mathrm{match}}^{k,\mathrm{treated}} - \bar{X}_{\mathrm{match}}^{k,\mathrm{control}}}{\sqrt{s_{T=1}^{2,k}}}$$

2.7.2 Tasks:

- 1. Compute SMDs before and after matching for all covariates
- 2. Create a Love plot showing balance before and after matching (both methods)
- 3. Add vertical line at 0.1 (acceptable threshold)
- 4. Comment on which design achieves better covariate balance

5. Create two additional Love plots including interactions and squared terms

```
# Calculate SMDs before matching
# Calculate SMDs after matching (logit)

# Calculate SMDs after matching (RF)

# Love plot: Base covariates
# Love plot: Including interactions
```

Interpretation:

[Which method achieves better balance? Are all SMDs below 0.1?]

2.8 Q4. Overlap

2.8.1 Tasks:

1. For each method, report:

Love plot: Including squared terms

- Min and max of $\hat{e}(X_i)$ for treated and controls
- Proportion of treated units whose $\hat{e}(X_i)$ lies inside the support of controls (and vice versa)
- 2. Plot distributions of $\hat{e}(X_i)$ for treated and controls
- 3. Identify regions of poor overlap or extreme propensities
- 4. (Optional) Trim observations outside common support and re-compute ATT
- 5. Examine matched subsets do matches seem like fair counterfactuals?

```
# Min/max propensity scores by treatment group

# Proportion in common support

# Plot propensity score distributions

# Optional: Trim and re-estimate
```

Discussion:

[Is there good overlap? Which observations are on the edge of common support?]

2.9 Q5. Matched-Pair ATT

2.9.1 ATT Estimator

Let each matched pair be denoted by (i, j(i)) where i is treated and j(i) is its matched control.

The average treatment effect on the treated is:

$$\hat{\tau}_{\mathrm{ATT}} = \frac{1}{N_T^*} \sum_{i \in \mathcal{T}^*} \left(Y_i - Y_{j(i)} \right)$$

where \mathcal{T}^* is the set of treated units with a valid match.

2.9.2 Task:

Compute the ATT for both matching methods.

ATT using logit matching

ATT using RF matching

Interpretation:

[What is the estimated effect of UN intervention on conflict duration?]

2.10 Q5.5. Bias-Variance Tradeoff in Matching Ratios

2.10.1 (a) Conceptual Question

For 1-to-m nearest-neighbor matching without replacement, the ATT estimator is:

$$\hat{\tau}_{\mathrm{ATT}}^{(m)} = \frac{1}{N_T^*} \sum_{i \in \mathcal{T}^*} \left(Y_i - \frac{1}{m} \sum_{j \in \mathcal{J}(i)} Y_j \right)$$

where $\mathcal{J}(i)$ is the set of the m closest control matches for treated unit i.

Tasks:

- 1. Explain why increasing m tends to:
 - Decrease variance
 - Increase bias
- 2. Discuss how this relates to distance in covariate space
- 3. If overlap is weak, which risk dominates as m grows?

Discussion:

[Your explanation of the bias-variance tradeoff here]

2.10.2 (b) Practical Exercise

Tasks:

- 1. Re-run matching for 1:1, 2:1, and 3:1 ratios (with replacement and same caliper)
- 2. Record: number matched, mean distance, ATT estimate
- 3. Compute cluster-robust standard errors for each design
- 4. Create results table
- 5. Plot ATT vs. m with ± 1.96 SE error bars

1:1 matching

2:1 matching

3:1 matching

Create comparison table

Plot ATT by matching ratio with error bars

Interpretation:

[Do results display expected bias-variance pattern?]

2.10.3 (c) Discussion

Tasks:

- Which design (1:1, 2:1, or 3:1) is most appropriate?
- How does observed pattern relate to Abadie & Imbens (2006)?
- What would happen with infinite data and perfect overlap?

Discussion:

[Your analysis here]

2.11 Q6. Robust Post-Matching Inference (Abadie & Spiess, 2021)

2.11.1 Regression with Cluster-Robust Standard Errors

After matching, fit the regression:

$$Y_i = \alpha + \tau T_i + \varepsilon_i$$

using only matched data.

Let s(i) denote the subclass (pair id) of observation i.

Compute cluster-robust standard errors for $\hat{\tau}$ by clustering on s(i):

$$\widehat{V}_{\mathrm{CR}}(\widehat{\tau}) = (X'X)^{-1} \left(\sum_s X_s' \widehat{\varepsilon}_s \widehat{\varepsilon}_s' X_s \right) (X'X)^{-1}$$

2.11.2 Tasks:

- 1. Report $\hat{\tau}$ and its cluster-robust standard error
- 2. Compare results for logit-matched and RF-matched samples
- ${\it \# Regression \ on \ logit-matched \ data \ with \ cluster-robust \ SE}$
- # Regression on RF-matched data with cluster-robust SE
- # Compare results

Interpretation:

[Compare point estimates and standard errors across methods]

2.12 Q7. (Optional) Bootstrap Check

2.12.1 Matched-Pair Bootstrap

Warning: Bootstraps are not theoretically valid for matching estimators, but this serves as a check.

Tasks:

- 1. Resample matched pairs (subclasses) with replacement
- 2. Recompute $\hat{\tau}^{(b)}$ for each bootstrap sample $b=1,\ldots,B$
- 3. Report bootstrap mean, SD, and percentile 95% CI
- 4. Compare to cluster-robust results

Bootstrap procedure

Discussion:

[Do bootstrap and cluster-robust results tell a similar story?]

2.13 Q8. Reflection

2.13.1 Tasks:

- 1. Why does the propensity score $e(X_i)$ act as a balancing score?
- 2. How does random-forest estimation of $e(X_i)$ change matching results compared to logistic regression?
- 3. Why is overlap $(0 < e(X_i) < 1)$ necessary for identifying the ATT?

Discussion:

[Your reflection here]

3 Part II: Synthetic Control - German Reunification Study

3.1 Background

In 1990, West Germany underwent reunification with East Germany. The question: What was the economic cost (or benefit) of this event on West Germany's GDP per capita?

Using the synthetic control method, we construct a counterfactual "synthetic West Germany" from a weighted combination of other OECD countries.

Paper: Abadie, Diamond & Hainmueller (2015), Comparative Politics and the Synthetic Control Method, AJPS.

Dataset: Available via Harvard Dataverse (doi:10.7910/DVN/24714)

3.2 Load Data

```
# Read German reunification dataset
# Load the replication dataset
load("repgermany.RData")
repgermany <- x

# Display structure
#str(repgermany)
#head(repgermany)</pre>
```

3.3 (a) Conceptual Questions

3.3.1 Tasks:

- 1. Explain the intuition behind the synthetic control method. What kind of assignment problem does it address?
- 2. Why is it particularly suitable for the West Germany case?
- 3. What is the key identification assumption?

Discussion:

[Your conceptual explanation here]

3.4 (b) Mathematical/Optimization Questions

3.4.1 The Optimization Problem

The synthetic control method solves:

$$\min_{w} \sum_{t \leq T_0} \left(Y_{1t} - \sum_{j=2}^{J+1} w_j Y_{jt} \right)^2$$

subject to:

$$w_j \ge 0, \quad \sum_j w_j = 1$$

3.4.2 Tasks:

- 1. Write and explain each term in the optimization problem
- 2. What role do v-weights play in predictor balancing?
- 3. Why is the convex-combination constraint important? What if weights could be negative or sum $\neq 1$?

Mathematical Discussion:

[Your explanation of the optimization problem and constraints]

3.5 (c) Estimation, Balance Before & After

3.5.1 Tasks:

- 1. Estimate synthetic control for West Germany over pre-treatment period
- 2. Compute balance table of key predictors (GDP, trade openness, inflation, schooling, investment) showing treated vs. synthetic mean **before treatment**
- 3. Report non-zero weights w_i
- 4. Interpret: which donor countries dominate and why?
- 5. Assess whether pre-treatment fit is acceptable for credible inference

```
# Prepare data for Synth package
```

Run synthetic control estimation

Create balance table for pre-treatment predictors

Report unit weights

Interpretation:

[Which countries contribute most to synthetic West Germany? Is pre-treatment balance good?]

3.6 (d) Effect Size & Permutation Test

3.6.1 Tasks:

- 1. Plot actual vs. synthetic GDP per capita trajectory (pre- and post-treatment)
- 2. Calculate estimated effect (gap) in first few post-treatment years and average post-treatment gap
- 3. Perform **permutation (placebo) test** by reassigning treatment to each control country
- 4. Report where treated unit's gap falls in the distribution (approximate p-value)
- 5. Interpret: What does this suggest about the economic impact of reunification?

```
# Plot actual vs synthetic West Germany
```

Calculate treatment effect (gap)

Permutation test: assign treatment to each control

Calculate p-value

${\bf Interpretation:}$

What is the estimated effect? Is it statistically significant based on permutation test?

3.7 (e) Placebo Test on Earlier Years

3.7.1 Tasks:

- 1. Conduct placebo treatment year **before** actual 1990 treatment (e.g., 1975)
- 2. Re-estimate synthetic control and plot the gap
- 3. What does pre-treatment gap behavior tell you about parallel-trajectory assumption?
- 4. Comment on how convincing you find the main causal estimate

```
# Placebo test with fake treatment year
# Plot placebo gap
```

Interpretation:

[Does the placebo test support the validity of the main estimate?]

4 Conclusion

[Optional: Summarize key findings from both parts]

5 References

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- Rosenbaum, P., & Rubin, D. (1983). The Central Role of the Propensity Score in Observational Studies for Causal Effects. *Biometrika*.
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- Abadie, A. (2021). Using Synthetic Controls: Feasibility, Data Requirements, and Methodological Aspects. *Journal of Economic Literature*, 59(2), 391–425.

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