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## Honesty Pleage:

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Mayork Dahiya

Q-1 Griven points = 
$$(1,2)$$
,  $(3,4)$ ,  $(4,5)$ ,  $(5,5)$ ,  $(5,6)$ ,  $(6,10)$ 

Mean: =  $\overline{X} = \frac{1}{n} \le x_i = \frac{1}{6} (1+3+4+5+5+10) = 4.67$ 
 $\overline{Y} = \frac{1}{n} \le y_i^2 = \frac{1}{6} (2+4+5+5+6+10) = 5.33$ 

=) Mean Vector =  $(4.67, 5.33)$ 

## CoVarience Hatrix:

Variance 
$$(x)$$
:
$$\begin{bmatrix}
Var(x) & CoV(x,y) \\
Var(y)
\end{bmatrix}$$
Variance  $(x)$ :
$$\frac{2}{N} (xe - \overline{x})^2 = \frac{1}{6} [(-3.67)^2 + (1.67)^2 + (0.67)^2 + (0.33)^2 + (5.33)^2]$$

$$= 7.55$$

$$Var(y) = \frac{1}{N} \left[ (y_i - \bar{y})^2 = \frac{1}{6} \left[ (-3.33)^2 + (-1.33)^2 + (0.33)^2 + (-0.33)^2 + (-0.33)^2 + (-0.67)^2 + (-0.67)^2 \right]$$

$$= 5.88$$

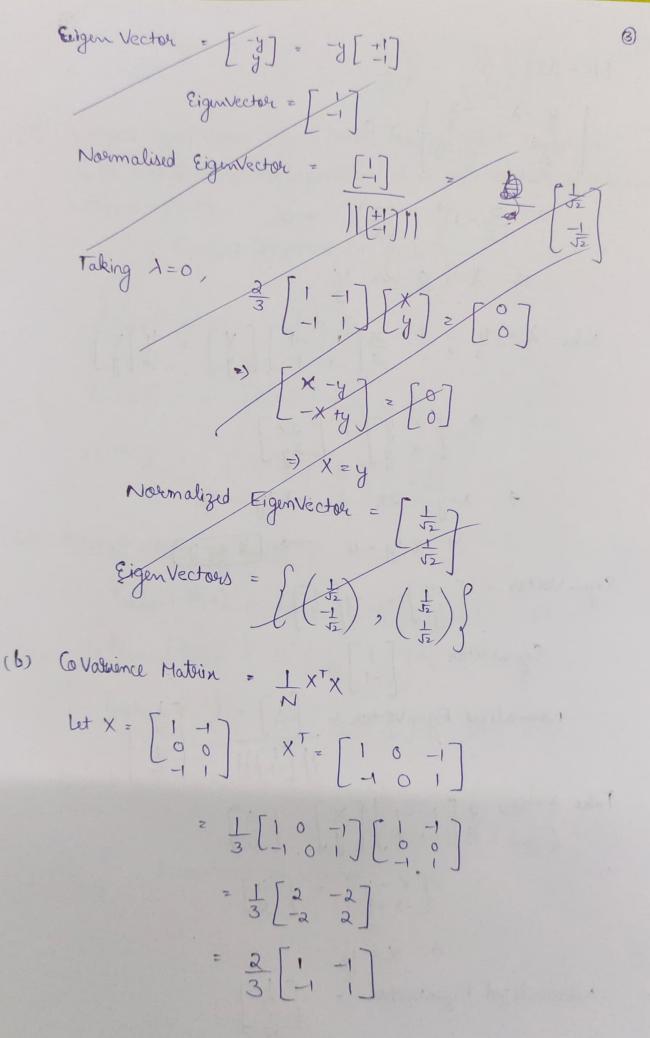
$$CoV(x,y) = \frac{1}{N} \sum_{l=1}^{N} (x_{l} - \bar{x})(y_{l} - \bar{y})$$

$$= \frac{1}{N} \left[ (-3.67)(-3.33) + (-1.67)(-1.33) + (-0.67)(-0.33) + (0.33)(-0.33) + (0.33)(0.67) + (5.33)(4.67) \right]$$

$$= 6.61$$

(a) Zero Mean Data points are calculated as follows:  $(x',y') = (x_i,y_i) - (\bar{x},\bar{y})$   $\bar{x} = \frac{1}{3}(-1-2-3) = -2$   $\bar{y} = \frac{1}{3}(1+2+3) = +2$   $(\bar{x},\bar{y}) = (-2,2)$ 

$$(x^2, y^2) = (1, -1)$$
,  $(0, 0)$ ,  $(-1, 1)$   $\rightarrow$  after performing above subtraction



$$\begin{vmatrix} \frac{2}{3} - \lambda & -\frac{2}{3} \\ -\frac{2}{3} & \frac{2}{3} - \lambda \end{vmatrix} = 0$$

$$= \frac{(2 - 1)^2}{(3 - 1)^2} = \frac{(2 - 1)^2}{(3 - 1)^2} = 0$$

Take 
$$\lambda = \frac{4}{3}$$
,  $\frac{2}{3} \begin{bmatrix} 1 & -1 \end{bmatrix} \begin{bmatrix} 1 & 1 \\ 2 & 3 \end{bmatrix} \begin{bmatrix} 1 & 1$ 

$$\Rightarrow \begin{bmatrix} x - y \\ -x \end{bmatrix} = \begin{bmatrix} 2x \\ 2y \end{bmatrix}$$

$$x-y=2x , -x+y=2y$$

$$\Rightarrow X + y = 0 \rightarrow X = -y$$

Take 
$$\lambda = 0$$
,  $\frac{3}{3} \begin{bmatrix} 1 & -1 \end{bmatrix} \begin{bmatrix} x \\ y \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}$ 

Normalized EigenVector . [ 1]

Eigen Vectors = 
$$\left\{ \begin{pmatrix} \frac{1}{\sqrt{2}} \\ \frac{1}{\sqrt{2}} \end{pmatrix}, \begin{pmatrix} \frac{1}{\sqrt{2}} \\ \frac{1}{\sqrt{2}} \end{pmatrix} \right\}$$

(d) Largest eigen Value = 4 , hence the principal component will correspond to this eigenValue as it will capture maximum variance of data.

Principal Component = 
$$\begin{pmatrix} \frac{1}{12} \\ -\frac{1}{12} \end{pmatrix}$$

P(1,+1) =  $\begin{bmatrix} 1 & -1 \end{bmatrix} \begin{pmatrix} \frac{1}{12} \\ -\frac{1}{12} \end{pmatrix} = \begin{bmatrix} \frac{2}{12} \\ -\frac{1}{12} \end{pmatrix} = \begin{bmatrix} \frac{2}{12} \\ -\frac{1}{12} \end{pmatrix} = \begin{bmatrix} \frac{1}{12} \\ -\frac{1}{12} \end{pmatrix} = \begin{bmatrix} -\frac{1}{12} \\ -\frac{1}{12} \end{pmatrix} = \begin{bmatrix} -\frac{1}{12} \\ -\frac{1}{12} \end{pmatrix} = \begin{bmatrix} -\frac{1}{12} \\ -\frac{1}{12} \\ -\frac{1}{12} \end{pmatrix} = \begin{bmatrix} -\frac{1}{12} \\ -\frac{1}{12} \\ -\frac{1}{12} \end{pmatrix} = \begin{bmatrix} -\frac{1}{12} \\ -\frac{1}{12} \\ -\frac{1}{12} \\ -\frac{1}{12} \\ -\frac{1}{12} \end{pmatrix} = \begin{bmatrix} -\frac{1}{12} \\ -\frac{1}{1$ 

(e) Reconstruction of Points:

Pricons 
$$(+1,-1) = \sqrt{2} \left[ \frac{1}{12}, -\frac{1}{52} \right] = \left[ \frac{1}{12}, -\frac{1}{12} \right]$$
Pricons  $(0,0) = 0 \left[ \frac{1}{12}, -\frac{1}{52} \right] = \left[ 0,0 \right]$ 
Pricons  $(-1,1) = -\sqrt{2} \left[ \frac{1}{52}, -\frac{1}{52} \right] = \left[ -1,1 \right]$ 

Reconstruction evor = 11 Original P - Reconstructed P11 = 11011 = 0 (any)

Reconstruction evor = 0%

0-3 Given:

X -> zero mean data

V > eight vectors of the covarience motion of X.

So, X - nxp dimension

V → PXP dimension { P≤n}

Projected data (x1) = X.V = [x, -- xn] [V, v2 -- Vp]

\* X, V, X, V2 ---)

So, mean of any column =  $S = X_1 V_1 + X_2 + \cdots \times_n V_j$ 

(Exi) Vj

Mean of the given projected data = 1.5

 $= \left(\sum_{i=1}^{n} x_i\right) V_i = 0 \quad \text{as } \hat{\xi} x_i = 0$ 

Hence Proved, the projected data is also zero mean data.

Q-4

 $J_1 = \int_{n_1 n_2} \underbrace{\xi}_{n_1 n_2} \underbrace{\xi}_{j_1 \in Y_1} \underbrace{(y_i - y_i)^2}_{j_1 \in Y_2}$ 

m, , m2 -> means of y, , y2

S,2, S22 -> sample variance of Y, Y2

(a) Prove: 
$$J_1 = (m_1 - m_2)^2 + \frac{S_1^2 + S_2^2}{n_1}$$

$$T_1 = \underbrace{1}_{n_1 n_2} \underbrace{\xi}_{j_1 \in Y_1} \underbrace{\xi}_{j_2 \in Y_2} \underbrace{(y_1 - y_1)^2}_{n_1 n_2} = \underbrace{1}_{j_1 \in Y_1} \underbrace{\xi}_{j_2 \in Y_2} \underbrace{(y_1^2 - 2y_1 y_1 + y_1^2)}_{j_1 \in Y_2}$$

According to over data initiation step:

We can see that

$$\begin{cases}
\xi \ \forall i = n_1 m_1 \\
\forall i \in Y_1
\end{cases}$$

$$\begin{cases}
\xi \ \forall i = n_2 m_2 \\
\forall i \in Y_2
\end{cases}$$

$$\begin{cases}
\xi \ \forall i^2 = S_1^2 + n_1 m_1^2 \\
\forall i \in Y_1
\end{cases}$$

$$\begin{cases}
\xi \ \forall i = n_2 m_2 \\
\forall i \in Y_2
\end{cases}$$

$$\begin{cases}
\xi \ \forall i = n_2 m_2 \\
\forall i \in Y_2
\end{cases}$$

$$\begin{cases}
\xi \ \forall i = n_2 m_2 \\
\forall i \in Y_2
\end{cases}$$

$$\begin{cases}
\xi \ \forall i = n_2 m_2
\end{cases}$$

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\end{cases}$$

$$\begin{cases}
\xi \ \forall i = n_2 m_2
\end{cases}$$

$$\begin{cases}
\xi \ \forall i = n_2 m_2
\end{cases}$$

$$\begin{cases}
\xi \ \forall i \in Y_2
\end{cases}$$

$$\begin{cases}
\xi \ \exists = n_2
\end{cases}$$

$$\begin{cases}
\xi \ \exists = n_2
\end{cases}$$

$$\begin{cases}
\xi \ \exists = n_2
\end{cases}$$

Using above :

$$J_{1} = \frac{1}{n_{1}n_{2}} \left( (S_{1}^{2} + n_{1}m_{1}^{2}) n_{2} - 2(n_{1}m_{1})(n_{2}m_{2}) + (S_{2}^{2} + n_{2}m_{2}^{2}) \right)$$

$$= \left( \frac{S_{1}^{2}}{n_{1}} + \frac{S_{2}^{2}}{n_{2}} + m_{1}^{2} - 2m_{1}m_{2} + m_{2}^{2} \right)$$

$$\Rightarrow J_1 = \left[ \left( m_1 - m_2 \right)^2 + \frac{S_1^2 + S_2^2}{n_1} \right]$$

(b) For total scatter materia

$$J_2 = \frac{1}{n_1} \frac{\xi (y_1 - m_1)^2 + 1}{\eta_2 y_3 \in Y_2} \frac{\xi (y_3 - m_2)^2}{\eta_2 y_3 \in Y_2}$$

$$S_i^2 = \frac{\xi}{y_i \in Y_i} (y_i - m_i)^2$$
,  $S_2^2 = \frac{\xi}{y_i \in Y_2} (y_i - m_2)^2$ 

$$\int_{1} = \frac{S_{1}^{2} + S_{2}^{2}}{n_{1}}$$

hagrangian 
$$\rightarrow h(w, x) = J_1 + J_2$$
  
Computing quadient of x wort w.

$$\nabla \omega h = \nabla \omega J_1 + \lambda \nabla \omega J_2$$

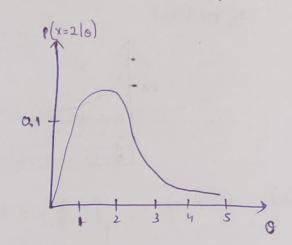
$$\nabla \omega h = -(y - \omega^T x) n + \lambda(\omega - \omega_0)$$

$$= -(y - \omega x) n + \lambda (\omega - \omega_0) = 0$$

Hure, 
$$m_1$$
 and  $m_2$  are constants and  $S_1$  and  $S_2$  are defined as 
$$S_1 = \sum_{x \in D} \frac{-1}{n_2 S_2(m_1 - m_2)}$$

$$S_2 = \sum_{x \in D} \frac{1}{n_1 (m_1 - m) (m_2 - m)^T}$$

Substitute 1 into enpression for a to get, find expression



(b) log likelihood 
$$\Rightarrow$$
  $l(o) = \sum_{k=1}^{n} ln p(x_{k} | o)$ 

$$\sum_{k=1}^{2} \left[ \ln o - o_{xk} \right] = n \ln o - o \sum_{k=1}^{2} x_{k}$$

N20

otherwise

$$\nabla_0 l(0) = \frac{\partial}{\partial 0} (n ln 0 - 0 \sum_{k=1}^n x_k)$$

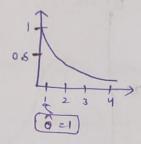
$$= \frac{n}{0} - \sum_{k=1}^{n} n_k = 0$$

$$\hat{O} = \frac{n}{\hat{Z} \times_{k}}$$
 f moximum likelihood function  $S$ 

Using Integration , 
$$\int x p(x) dx$$

$$= \int x e^{-x} dx = 1$$

Using these we can see that  $\hat{o}=1$ , so platting it on the graph of part (a)



(a) Let us assume I(.) indicator function which outputs 1.0 if condition is satisfied else 0=0.

$$P(N|6) = \prod_{k=1}^{n} \frac{1}{p}(X_{k}|0)$$

$$= \prod_{k=1}^{n} \frac{1}{p}I(0 \le M_{k} \le 0)$$

$$= \frac{1}{p}I(0 \ge \max_{k} M_{k})I(0 \le \min_{k} X_{k})$$

P(0|0) decreases as value of 0 increases and also  $I(0 \ge \max_{K} X_{K})$  will become zero if 0 goes below max value of  $X_{K}$ .

The likelihood function is maximized at \$ = max(Xx)

Q-6 Discrete sandom variable with support - f-2, -1, 0, 1, 25

X -> follows uniform distribution

$$P(X=x) = \begin{cases} \frac{1}{n} & n \in \{-2,-1,0,1,2\} \\ 0 & \text{else} \end{cases}$$

$$\begin{cases} \frac{1}{5} & \kappa \in \{-2, -1, 0, 1, 2\} \\ 0 & \text{else} \end{cases}$$

$$\frac{2}{5}\left(-2\right)\frac{1}{5}+\left(-1\right)\frac{1}{5}+6\left(\frac{1}{5}\right)+1\left(\frac{1}{5}\right)+\left(\frac{2}{5}\right)=0$$

Var 
$$(x) = E[x^2] - (E[x])^2 = [(\frac{2}{5})^2 + (\frac{1}{5})^2 + (0)^2 + (\frac{1}{5})^2 + (\frac{3}{5})^2] - (0)^2$$

$$Var(x) = 2$$

$$Q-7$$
  $g=w_1^{(3)}a_1^{(2)}+w_2^{(3)}a_2^{(2)}; \quad \partial [\sigma(z)] = [\sigma(z)][1-\sigma(z)]$ 

(i) 
$$\frac{\partial f}{\partial z_i} = \frac{\partial f}{\partial a_i} \cdot \frac{\partial a_i^{(2)}}{\partial z_i^{(2)}}$$
 (chain suite)

(ii) 
$$\frac{\partial f}{\partial z_{1}^{(2)}} = \frac{\partial f}{\partial q_{1}^{(2)}} \times \frac{\partial q_{2}^{(2)}}{\partial z_{2}^{(2)}}$$

$$\Rightarrow \left[\frac{\partial f}{\partial z_{2}^{(2)}} = W_{2}^{(3)}, \left[\sigma(z_{2}^{(2)})\right] \left[1 - \sigma(z_{2}^{(2)})\right]\right]$$

$$\frac{\partial \omega_{ij}}{\partial \omega_{ij}} = \left(\frac{\partial \xi_{ij}}{\partial \xi_{ij}} \cdot \frac{\partial \xi_{ij}}{\partial \xi_{ij}} \cdot \frac{$$

$$\frac{\partial \mathcal{C}_{13}}{\partial \mathcal{C}_{13}} = \left(\frac{\partial \mathcal{C}_{13}}{\partial z_{13}}\right) \cdot \left[\frac{\partial \mathcal{C}_{13}}{\partial \mathcal{C}_{13}} \cdot \frac{\partial \mathcal{C}_{13}}{\partial z_{13}} + \frac{\partial \mathcal{C}_{13}}{\partial \mathcal{C}_{13}} + \frac{\partial \mathcal{C}_{23}}{\partial z_{13}} \cdot \frac{\partial \mathcal{C}_{23}}{\partial z_{13}}\right]$$

Answer ->

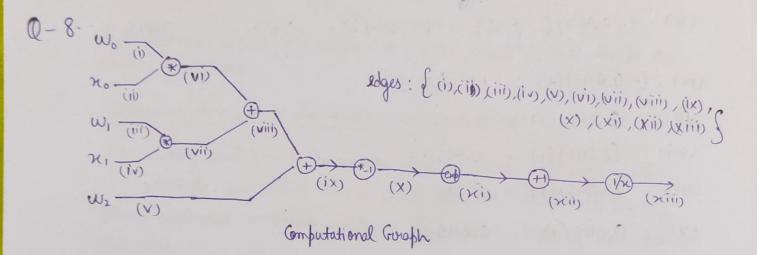
$$\frac{\partial \mathcal{L}_{1}}{\partial \mathcal{L}_{1}} = \left[ \chi_{1} \cdot \left[ \sigma(z_{1}^{(1)}) \right] \cdot \left[ 1 - \sigma(z_{1}^{(1)}) \right] \cdot \left[ (\omega_{1}^{(2)}) \cdot \sigma(z_{1}^{(2)}) \right] \cdot \left[ (\omega_{1}^{(2)}) \cdot \sigma(z_{1}^{(2)}) \right] \cdot \left[ (\omega_{1}^{(2)}) \cdot (\sigma(z_{1}^{(2)})) \cdot (\omega_{1}^{(2)}) \cdot (\omega_{1}^{(2)}) \right] \cdot \left[ (\omega_{1}^{(2)}) \cdot (\sigma(z_{1}^{(2)})) \cdot (\omega_{1}^{(2)}) \cdot (\omega_{1}^{(2)}) \cdot (\omega_{1}^{(2)}) \right] \cdot \left[ (\omega_{1}^{(2)}) \cdot (\sigma(z_{1}^{(2)})) \cdot (\omega_{1}^{(2)}) \cdot (\omega_{1}$$

$$(iv) \frac{\partial f}{\partial \omega_{22}^{(1)}} = \left(\frac{\partial z_{1}^{(1)}}{\partial \omega_{22}^{(1)}}\right) \left[\frac{\partial f}{\partial z_{1}^{(2)}} \cdot \frac{\partial z_{1}^{(2)}}{\partial z_{2}^{(1)}} + \frac{\partial f}{\partial z_{1}^{(2)}} \cdot \frac{\partial z_{2}^{(2)}}{\partial z_{2}^{(2)}}\right]$$

Answer: 
$$\frac{\partial f}{\partial \omega_{22}^{(1)}} = \left[ \mathcal{H}_{2} \left[ \sigma(z_{2}^{(2)}) \right] \left[ 1 - \sigma(z_{2}^{(2)}) \right] \right] \left[ \left( \omega_{1}^{(3)} \right) \cdot \left( \omega_{12}^{(2)} \right) \left( \sigma(z_{2}^{(2)}) \right) \left( 1 - \sigma(z_{2}^{(2)}) \right) \right] \left[ \left( \omega_{12}^{(3)} \right) \cdot \left( \omega_{12}^{(2)} \right) \left( \sigma(z_{2}^{(2)}) \right) \left( 1 - \sigma(z_{2}^{(2)}) \right) \right]$$

$$\frac{\partial w_{13}}{\partial \theta} = \left(\frac{\int w_{13}}{\int g_{13}}\right) \cdot \left(\frac{\int g_{13}}{\int g_{13}} \cdot \frac{\int g_{13}}{\int g_{13}} + \frac{\int g_{13}}{\int g_{13}} + \frac{\int g_{13}}{\int g_{13}}\right)$$

Answer: 
$$\frac{\partial f}{\partial \omega_{12}^{(1)}} = \left[ \chi_2 \left[ \sigma(z_1^{(1)}) \right] \left[ 1 - \sigma(z_1^{(2)}) \right] \right] \cdot \left[ \left( \omega_1^{(3)} \right) \left( \omega_{11}^{(2)} \right) \left( \sigma(z_1^{(2)}) \right) \left( 1 - \sigma(z_1^{(2)}) \right) \right] + \left( \left( \omega_2^{(3)} \right) \left( \omega_{21}^{(3)} \right) \left( \sigma(z_1^{(2)}) \right) \left( 1 - \sigma(z_2^{(2)}) \right) \right]$$



[a) Forward past on the computation graph 
$$w = [w_0, w_1, w_2] = [1,1,1]$$
 adopt  $w_0 \times w_0 \times w_0 = 1$ 

(vi)  $w_0 \times w_0 \times w_1 \times w_1 = 1$ 

(i)  $w_0 = 1$ 

(iii)  $w_0 = 1$ 

(iv)  $w_0 \times w_0 + w_1 \times w_1 = 2$ 

(iv)  $w_0 \times w_0 + w_1 \times w_1 + w_2 = 3$ 

(iv)  $w_1 = 1$ 

(x)  $-(w_0 \times w_0 + w_1 \times w_1 + w_2) = -3$ 

(xi)  $e^{-(w_0 \times w_0 + w_1 \times w_1 + w_2)} = 0.0497$ 

(xii)  $1 + e^{-(w_0 \times w_0 + w_1 \times w_1 + w_2)} = 0.952$ 

(b) local gradient times upstream gradient

eldges Value = (upstruam gradient) x (lovel geodient)

(Xiii) : 1

(Xii): 
$$I\left[\frac{d}{dn}(\frac{1}{n})\right] = \frac{1}{\lambda^2} = -0.907$$

(x) 
$$(-0.907)(\frac{de^n}{dn}) = (-0.907)e^n = -0.907e^{-3} = -0.045$$

- Q-9. Steps of k means clastering algorithm.
  - Degine the number of clusters (K)

    There's no single best value of K and hence it often involves some exploration and experimentation to find the optimal number of clusters that best suppresent the structure of data.
- Randomly chase K datapoints from your dataset to be the initial centroids.
- For every datapoint, measure the enclided distance with all centraids and label them with the one nearest to them. Other distance metrics such as Manhattan distance can distance be used.
- Recompute the centroids coolesponding to each cluster

  Once all data points have been labelled according to their nearest

  centroid, secompute new centroids for each cluster averaging

  the co-ordinates of all such datapoints belonging to the

  respective clusters.
- Expeat steps 3 and 4 until convergence
  Repeat steps 3 and 4 until the centraids no longer change or
  the convergence condition is achieved.
- > Limitations of k means clustering algorithm
- Devedefined number of clusters (x)

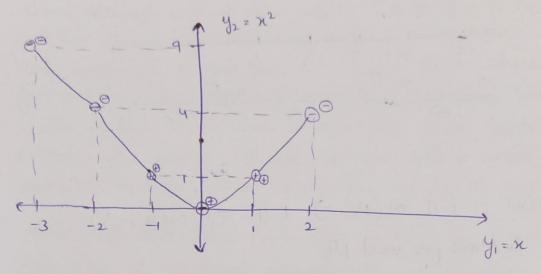
  The optimal value of K is many times not clearly unsible from data but the K means clustering algorithm demands you to specify a value of K initially which is very challenging.

- 0 Sensitive to initial centraids and outliers Different initializations might lead to slightly different cluster assignments and also, the algorithm is highly effected by outliers as it trues to provide a definite label to every datapoint.
- 3 Not Ideal fare High Dimensional Data K Heans becomes less effective as the no. of dimensions in our data increases.
- Spherical Cluster Assumption K means works best for data forming roughly spherical clusters and is not very ideal for data with warying densities or eny other shape.
- Distance Based Clustering K Hears Clustering is solely based on distance as a similarity metric between two points neglecting any other attributes.
- 6) No guarantee of optimal solution K Means finds the local minima which not be the best solution globally. Therefore, you cannot be sure if the solution you have reached is optimal.

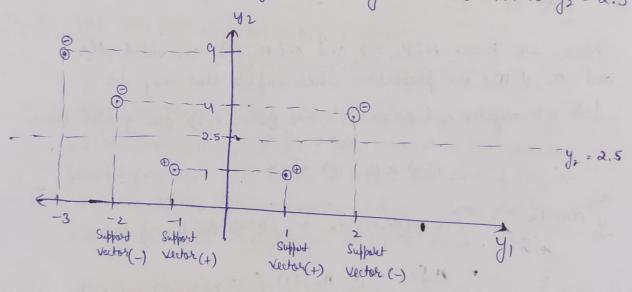
Q-10 e e o

(a) Elature map to make the data linearly seperable is as Rollows: ->

Clotting in 2D space & y, y = §



(b) w2 y2 + w, y, +w. by hard margin linear SVM is y2 = 2.5



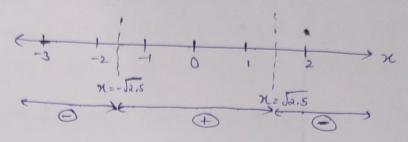
Positive support vectors  $\rightarrow \xi(-1,1), (1,1) \xi$ Negative support vectors  $\rightarrow \xi(-2,4), (2,4) \xi$ 

(c) Equivalent Decision Boundary in the original input space is  $\pi_2 \pm \sqrt{2.5} \Rightarrow \pi_2^2 = 2.5$ 

ne [-Jais, Jais]: prediction is the

x & [-12.5, Ja.5]: prediction is -ve

Input space is ID and hence the decision boundary is an interval.



## Ques 118)

(a) 
$$P(n) = (\pi_0 N(\mu_0, \sigma_0) + (1 - \pi_0)(N(\mu_1, \sigma_1))$$
  
Conditions for valid pdf

Now, we know  $N(H_0, \sigma_0)$  and  $N(H_1, \sigma_1)$  are valid pdfs and  $T_0$ ,  $(1-T_0)$  are proportions which implies that they are both non-negative. ( $N(H_0, \sigma_0)$ ) and  $(N(H_1, \sigma_1))$  are greater than zero for all n

 $P(n) \ge 0$  for all n

$$\int_{-\infty}^{\infty} P(n) dn = \int_{-\infty}^{\infty} \pi_0 N(\mu_0, \sigma_0) dn + \int_{-\infty}^{\infty} (1 - \pi_0) N(\mu_1, \sigma_1) dn$$

 $\int P(n) dn = \pi_0 + (1-\pi_0) = 1$   $\int P(n) dn = \pi_0 + (1-\pi_0) = 1$ 

P(n) is valid probability distribution function.

- (b) The clustering results from Kreans and Gitt might not be the same. Although both of them will be able to find the clusters well but the cluster centres might differ because Kreans uses hard assignment of each point to a specific cluster whereas Gitt goes with a probabilistic soft assignment that causes Gitt's cluster centres to be closer to each other as the points in Kreans other cluster will also contribute with a small weight associated to them while calculating Gitti's cluster center.
  - (c) Circle shifts to the deft while stars shifts to the right in the next stage of expectation maximization

Q-12 (a) For SVM1,  $W^TX+b\geq 1$  ( then  $n\rightarrow +$ )  $W^TX+b\leq -1$  ( than  $n\rightarrow -$ )

There is some finite moch zero value of masyin given by 2/11w11 but in SVM-2, masyin is zero, vary point above the decision boundary is (+) and below it is (-) and even if they are very close to the decision boundary (d  $\rightarrow$ 0), then also they are classified. Hence, the masyin in SVM2 can be zero.

Hence, Margin of SVM-1 > Margin of SVM 2 (0) [ given data is linearly septrable

(b) Now, if we take a SVM3, with w= w (SVM-1 trained wight)

b= b (SVM-1 trained wight)

we know that all points abready satisfy,

WTX + 6 21 ( BOOL +)

So, if we kept w = w/2 and 6 = 6/2

Decision function equation becomes (wx +6) >1 , B(n) >1 >0 (box +)

f(n) c o ( for -)

We assign points class  $\Theta$  if wixtb>0 and  $\Theta$  otherwise, as changing the weight from  $W \to W/2$  and  $bP \to b/2$  did not after the decision of labels of points as the sign of surnams same,

New, wtx+b > 1 and wtx +b < 1 are the new margins is.
magnitude of decision boundary function get scaled but sign is
because wed. So, SVM 3 correctly identifies all training points because
decision boundary remains same.

(C) As w + w/2 in svm-3 and margin = 2/11w11So, if margin of svm + is 2/11w11Margin of  $svm 3 = \frac{2}{11w11} = \frac{4}{11w11} = 2(2/11w11)$ 

=) Margin of SVM3 is twice that of SVM-1