

习题 1-1

1. 指出下列命题哪些正确, 哪些不正确?

(1) $A \cup B = A \bar{B} \cup B$.

解: $A \cup B = A \bar{B} \cup A \bar{B} \cup B \cup A \bar{B} = A \bar{B} \cup [(A \cup \bar{A}) B] = A \bar{B} \cup B$. 所以正确.

(2) $A = A \bar{B} \cup A B$.

解: $A = A \Omega = A(B \cup \bar{B}) = A \bar{B} \cup A B$. 所以正确.

(3) $\bar{A} B = A \cup B$.

解: 由(1)的推导可知 $A \cup B = A \cup \bar{A} B$, 所以 $\bar{A} B \neq A \cup B$. 反命题不正确.

(4) $(\overline{A \cup B}) C = \bar{A} \bar{B} \bar{C}$.

解: $(\overline{A \cup B}) C = \bar{A} \bar{B} C \neq \bar{A} \bar{B} \bar{C}$. 反命题不正确.

(5) $(A B)(A \bar{B}) = \emptyset$.

解: $(A B)(A \bar{B}) = A(B \bar{B}) = A \emptyset = \emptyset$. 所以命题正确.

(6) 若 $A \subset B$, 则 $A = A B$.

解: 若 $A \subset B$, 则 $A \bar{B} = \emptyset$. $A = A \Omega = A(B \cup \bar{B}) = A B \cup A \bar{B} = A B \cup \emptyset = A B$. 所以命题正确.

(7) 若 $A \subset B$, 则 $A \cup B = A$.

解: $\because A \cup B = A \cup \bar{A} B$. 要 $A \cup B = A$, 则 $\bar{A} B = \emptyset$. 而若 $A \subset B$, 则 $\bar{A} B \neq \emptyset$. 所以反命题不正确.

(8) 若 $A \subset B$, 则 $\bar{B} \subset \bar{A}$.

解: 若 $A \subset B$, 则 $A \bar{B} = \emptyset$. 由 $\bar{B} = \bar{B} \Omega = \bar{B}(A \cup \bar{A}) = \bar{B} A \cup \bar{B} \bar{A} = \emptyset \cup \bar{B} \bar{A} = \bar{B} \bar{A} \subset \bar{A}$. 所以反命题正确.

(9) 若 $A B = \emptyset$, 则 $\bar{A} \bar{B} \neq \emptyset$.

解: 当 $A B = \emptyset$ 时, 若 $A \cup B = \Omega$, 则 $\bar{A} \bar{B} = \emptyset$. 所以命题不正确.

(10) 若 $A B = \emptyset$, 则 $\bar{A} \bar{B} = \emptyset$.

解: 当 $A B = \emptyset$ 时, 若 $A \cup B \neq \Omega$, 则 $\bar{A} \bar{B} \neq \emptyset$. 所以命题不正确.

2. 在分别标有号码 1~8 的八张卡片中任抽一张. 设事件 A 为“抽得一张标号不大于 4 的卡片”, 事件 B 为“抽得一张标号为偶数的卡片”, 事件 C 为“抽得一张标号为能被 3 整除的卡片”.

(1) 试写出试验的样本点和样本空间.

解: 记 $\omega_i =$ “标号为 i 的卡片”, $i = 1, 2, 3, 4, 5, 6, 7, 8$. 为试验的样本点, 样本空间 $\Omega = \{\omega_1, \omega_2, \omega_3, \omega_4, \omega_5, \omega_6, \omega_7, \omega_8\}$.

(2) 试将下列事件表示为样本点的集合, 并说明分别表示什么事件?

(a) $A B$, (b) $A \cup B$, (c) \bar{B} , (d) $A - B$, (e) $\bar{B} \bar{C}$, (f) $\overline{B \cup C}$.

解: $A = \{\omega_1, \omega_2, \omega_3, \omega_4\}$, $B = \{\omega_2, \omega_4, \omega_6, \omega_8\}$, $C = \{\omega_3, \omega_6\}$.

(a) $A B = \{\omega_2, \omega_4\} =$ “抽得一张标号不大于 4, 且为偶数的卡片”.

(b) $A \cup B = \{\omega_1, \omega_2, \omega_3, \omega_4, \omega_6, \omega_8\} =$ “抽得一张标号不大于 4, 或者为偶数”.

偶数的卡片”。

(c) $\bar{B} = \{\omega_1, \omega_3, \omega_5, \omega_7\}$ = “抽得一张标号不是偶数(标号为奇数)的卡片”。

(d) $A - B = \{\omega_1, \omega_3\}$ = “抽得一张标号不大于4,但不是偶数的卡片”。

(e) $\overline{BC} = \{\omega_1, \omega_2, \omega_3, \omega_4, \omega_5, \omega_7, \omega_8\}$ = “抽得的一张标号不能为被3整除的偶数(即除标号为6的卡片外)的卡片”。

(f) $\overline{BUC} = \{\omega_1, \omega_5, \omega_7\}$ = “抽得一张标号既不是偶数,又不是能被3整除的卡片”(用 $\bar{B}\bar{C}$ 作解释)。

3. 将下列事件用 A, B, C 表示出来:

(1) 只有 A 发生。

解: $A\bar{B}\bar{C}$ 。

(2) A 不发生,但 B, C 至少有一个发生。

解: $\bar{A} \cdot (B \cup C)$ 。

(3) 三个事件恰有一个发生。

解: $A\bar{B}\bar{C} \cup \bar{A}B\bar{C} \cup \bar{A}\bar{B}C$ 。

(4) 三个事件至少有两个发生。

解: $AB \cup AC \cup BC$ (注: AB 即 $ABC \cup AB\bar{C}$, 其余同理)

或者为 $AB\bar{C} \cup A\bar{B}C \cup \bar{A}BC \cup ABC$ 。

(5) 三个事件都不发生,

解: $\bar{A}\bar{B}\bar{C}$ 或 $\overline{A \cup B \cup C}$ 。

(6) 三个事件最多有一个发生。

解: 即至少有两个事件不会发生: $\bar{A}\bar{B} \cup \bar{A}\bar{C} \cup \bar{B}\bar{C}$ 。

(7) 三个事件不都发生。

解: 不都发生意味着至少有一个不发生, 故为 $\bar{A} \cup \bar{B} \cup \bar{C}$ 。

(8) 三个事件至少有一个发生。

解: $A \cup B \cup C$ 。

4. 设 $\Omega = \{1, 2, \dots, 10\}$, $A = \{2, 3, 5\}$, $B = \{3, 5, 7\}$, $C = \{1, 3, 4, 7\}$

求下列事件: (1) $\overline{A \cdot B}$; (2) $A(\overline{BC})$ 。

解: (1) $\overline{A \cdot B} = \overline{A \cap B} = A \cup B = \{2, 3, 5, 7\}$ 。

(2) $A(\overline{BC}) = \overline{A} \cup (\overline{BC}) = \bar{A} \cup (BC)$

$\bar{A} = \{1, 4, 6, 7, 8, 9, 10\}$, $BC = \{3, 7\}$ 。

$\therefore \overline{A \cdot B} = \overline{A \cap B} = \{1, 3, 4, 6, 7, 8, 9, 10\}$ 。

习题 1-2

1. 已知 $P(A \cup B) = 0.8$, $P(A) = 0.5$, $P(B) = 0.6$. 求 $P(AB)$, $P(\bar{A}\bar{B})$, $P(\bar{A} \cup \bar{B})$.

解: $P(AB) = P(A) + P(B) - P(A \cup B) = 0.5 + 0.6 - 0.8 = 0.3$;

$$P(\bar{A}\bar{B}) = P(\overline{A \cup B}) = 1 - P(A \cup B) = 1 - 0.8 = 0.2$$
;

$$P(\bar{A} \cup \bar{B}) = P(\overline{AB}) = 1 - P(AB) = 1 - 0.3 = 0.7.$$

2. 已知 $P(A) = 0.6$, $P(B) = 0.7$. 求 $P(AB)$ 的最大值和最小值.

解: $P(AB) = P(A) + P(B) - P(A \cup B)$.

$\because P(A) \neq 0, P(B) \neq 0$, 故 $P(A \cup B) \neq 0$. 注意到 $P(A) < P(B)$, 故当 $A \subset B$ 时 $P(A \cup B) = P(B)$, 则 $P(AB) = P(A) = 0.6$ 取到最大值; 而当 $P(A \cup B) = 1$ 即 $A \cup B = \Omega$ 时, $P(AB) = P(A) + P(B) - 1 = 0.3$ 取到最小值.

3. 已知 $P(A) = x$, $P(B) = 2x$, $P(C) = 3x$, $P(AB) = P(BC)$, 求 x 的最大值.

解: $P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(AB) - P(AC) - P(BC) + P(ABC)$. (*)

当 $A \cup B \cup C = \Omega$, 且 $A \subset B, A \subset C$ 时, 有 $AB = BC = AC = ABC$.

则 (*) 式为 $1 = x + 2x + 3x - x - x - x + x$, 即 $4x = 1$.

$\therefore x = \frac{1}{4}$ 时取到最大值.

4. 设 $P(A) > 0, P(B) > 0$. 将下列四个数: $P(A), P(AB), P(A \cup B), P(A) + P(B)$ 按由小到大的顺序排列. 用符号 \leq 联系它们, 并指出在什么情况下可能有等式成立?

解: 顺序为 $P(AB) \leq P(A) \leq P(A \cup B) \leq P(A) + P(B)$.

当 $A \subset B$ 时, $P(AB) = P(A)$; 当 $B \subset A$ 时, $P(A) = P(A \cup B)$.

当 $AB = \emptyset$ 时, $P(A \cup B) = P(A) + P(B)$.

习题 1-3

1. 电话号码由六个数字组成, 每个数字可以是 $0, 1, 2, \dots, 9$ 中的任一数字, (但第一数字不能为 0). 求电话号码是由完全不同的数字组成的概率.

解: 记 $A =$ “电话号码由完全不同的数字组成”

$$P(A) = \frac{9 \cdot A_9^5}{9 \cdot 10^5} = \frac{9 \cdot 8 \cdot 7 \cdot 6 \cdot 5}{10^5} = 0.1512.$$

(注) 记 $A_i =$ “电话号码的第 i 个数字”, $i = 1, 2, \dots, 6$
利用 §1.4 乘法公式:

$$P(A) = P(A_1 A_2 \dots A_6) = P(A_1) P(A_2/A_1) P(A_3/A_1 A_2) \dots P(A_6/A_1 A_2 A_3 A_4 A_5) \\ = \frac{9}{9} \cdot \frac{9}{10} \cdot \frac{8}{10} \cdot \frac{7}{10} \cdot \frac{6}{10} \cdot \frac{5}{10} = 0.1512.$$

2. 在 11 张卡片分别写上 Probability 这 11 个字母, 从中任意连抽 7 张, 求其排列结果是 ability 的概率.

解: 记 $A =$ “排列结果是 ability”.

$$P(A) = \frac{A_2^1 A_3^2}{A_{11}^7} = \frac{1}{415800}$$

(注) 在 A 事件的样本点中, a, l, t, y 都仅一张卡片, b, i 各有两张, 因此共有 $A_1^1 A_2^2 A_3^2 A_4^1 A_5^1 A_6^1$ 即 $A_2^2 A_3^2$ 个点. 本题也可用乘法公式解.

3. 随机地将 15 名新生平均分配到三个班级中去, 这 15 名新生中有 3 名运动员, 问 (1) 每个班级各分配到一名运动员的概率是多少? (2) 3 名运动员被分配到同班级的概率是多少?

解: (1) 记 $A =$ “各班分配到一名运动员”.

$$P(A) = \frac{C_3^1 C_{12}^4 \cdot C_2^1 C_8^4}{C_{15}^5 C_{10}^5} \quad (\text{或} = \frac{A_{12}^{12} \cdot A_3^3 (C_5^3)^3}{A_{15}^{15}}), \text{得 } P(A) = \frac{25}{91};$$

(2) 记 $B =$ “3 名运动员分配到同班级”.

$$P(B) = \frac{C_3^1 \cdot C_{12}^2 C_3^2 \cdot C_{10}^5}{C_{15}^5 C_{10}^5} \quad (\text{或} = \frac{A_{12}^{12} \cdot A_3^3 \cdot C_3^1 C_5^3}{A_{15}^{15}}), \text{得 } P(B) = \frac{6}{91}.$$

4. 某工厂生产的一批产品共 100 个, 其中有 5 个次品, 从这批产品中任取一半来检查, 求取到的次品不多于一个的概率.

解: 记 $A =$ “取到的次品不多于一个”.

$$P(A) = \frac{C_{95}^{50} C_5^0 + C_{95}^{49} C_5^1}{C_{100}^{50}} = \frac{\frac{95!}{50! 45!} + \frac{95! \cdot 5}{49! 46!}}{\frac{100!}{50! 50!}} = 0.181.$$

5. 袋内放有 2 个伍分的钱币, 3 个贰分的钱币, 5 个壹分的钱币, 任取其中 5 个, 求总数超过一角的概率

解: 记 $A = \text{"5个钱币总数超过一角"}$.

$$P(A) = \frac{1}{C_{10}^5} [C_2^1 C_3^2 C_5^2 + C_2^1 C_3^3 C_5^1 + C_2^2 C_3^0 C_5^3 + C_2^2 C_3^1 C_5^2 + C_2^2 C_3^2 C_5^1 + C_2^2 C_3^3 C_5^0] \\ = 0.5.$$

6. 一学生宿舍有6名学生, 问: (1) 6人生日都在星期天的概率是多少? (2) 6人的生日都不在星期天的概率是多少? (3) 6人的生日不都在星期天的概率是多少?

解: (1) 记 $A = \text{"6人生日都在星期天"}$.

$$P(A) = \frac{1}{7^6}. \quad (\text{注: 可利用相互独立事件的积事件的概率公式:})$$

记 $A_i = \text{"第 } i \text{ 人生日在星期天"}$, $i=1, 2, 3, 4, 5, 6$ 相互独立, 且

$$P(A_i) = \frac{1}{7}, \quad i=1, \dots, 6. \quad \text{则 } P(A) = P(A_1 A_2 \dots A_6) = P(A_1) P(A_2) \dots P(A_6) = \left(\frac{1}{7}\right)^6.$$

(2) 记 $B = \text{"6人生日都不在星期天"}$.

$$P(B) = \left(\frac{6}{7}\right)^6. \quad (\text{注: 记 } B_i = \text{"第 } i \text{ 人生日不在星期天"}, P(B_i) = \frac{6}{7}, i=1, \dots, 6 \\ \text{相互独立, 则 } P(B) = P(B_1 B_2 \dots B_6) = P(B_1) P(B_2) \dots P(B_6) = \left(\frac{6}{7}\right)^6.)$$

(3) 记 $C = \text{"6人的生日不都在星期天"}$.

$$P(C) = 1 - P(\bar{C}) = 1 - P(A) = 1 - \left(\frac{1}{7}\right)^6.$$

7. 在 $1 \sim 100$ 共一百个数中任取一个数, 求这个数能被3或5整除的概率.

解: 记 $A = \text{"这个数能被3或5整除"}$.

在 $1 \sim 100$ 中能被3整除的数共有 $\left[\frac{100}{3}\right] = 33$ 个; 能被5整除的共 $\left[\frac{100}{5}\right] = 20$ 个; 能被 3×5 整除的共有 $\left[\frac{100}{15}\right] = 6$ 个. 故

$$P(A) = \frac{1}{100} [33 + 20 - 6] = \frac{47}{100} = 0.47.$$

8. 设袋中有5个白球与4个黑球, 每次从袋中任取一个球, 取出的球不放回去, 求 (1) 第2次才取得白球的概率; (2) 第2次取得白球的概率.

解: (1) 记 $A = \text{"第2次才取得白球"}$.

这表明第一次取得的是黑球, 第2次取得的是白球. 故:

$$P(A) = \frac{A_4^1 A_5^1}{A_9^2} = \frac{5}{18}.$$

(注: 可用 §1.4 的乘法公式: 记 $A_i = \text{"第 } i \text{ 次取得白球"}$, $i=1, 2$.

$$P(A) = P(\bar{A}_1 A_2) = P(\bar{A}_1) P(A_2 | \bar{A}_1) = \frac{C_4^1}{C_9^1} \cdot \frac{C_5^1}{C_8^1} = \frac{4}{9} \cdot \frac{5}{8} = \frac{5}{18}.)$$

(2) 记 $B = \text{"第2次取得白球"}$.

这表明第一次取得的可能白球, 也可能是黑球. 因此可直

接利用例6的结论, 这里取 $a=5$, $b=4$. 故 $P(B) = \frac{5}{5+4} = \frac{5}{9}$.
 直接计算则有 $P(B) = \frac{A_5^1 \cdot A_4^1}{A_9^2} = \frac{5}{9}$.

(注: 记 $A_i =$ "第 i 次取出白球", $i=1, 2$.)

$$\text{则 } P(A) = P(A_1 A_2 \cup \bar{A}_1 A_2) = P[(A_1 \cup \bar{A}_1) \cdot A_2] = P(A_2) = \frac{5}{9}.$$

$$\text{或 } P(A) = P(A_1 A_2 \cup \bar{A}_1 A_2) = P(A_1 A_2) + P(\bar{A}_1 A_2) = P(A_1)P(A_2/A_1) + P(\bar{A}_1)P(A_2/\bar{A}_1) \\ = \frac{5}{9} \cdot \frac{4}{8} + \frac{4}{9} \cdot \frac{5}{8} = \frac{5}{9}. \text{ 此式为全概率公式.}$$

9. 猎人在距离100米处射击一动物, 击中的概率为0.6. 如果第一次未击中, 则进行第二次射击, 但由于动物逃跑而使距离变为150米; 如果第二次又未击中, 则进行第三次射击, 这时距离变为200米. 假如击中的概率与距离成反比, 求猎人击中动物的概率.

解: 记 $A =$ "猎人击中动物", 又记 $A_i =$ "猎人第 i 次击中动物", $i=1, 2, 3$.

$$\text{由题意可得: } P(A_1) = 0.6; P(A_2/\bar{A}_1) = \frac{100}{150} \times 0.6 = 0.4;$$

$$P(A_3/\bar{A}_1 \bar{A}_2) = \frac{150}{200} \times 0.4 = 0.3.$$

$$\begin{aligned} \therefore P(A) &= P\{A_1 \cup \bar{A}_1 A_2 \cup \bar{A}_1 \bar{A}_2 A_3\} \\ &= P\{A_1\} + P(\bar{A}_1 A_2) + P(\bar{A}_1 \bar{A}_2 A_3) \\ &= P(A_1) + P(\bar{A}_1)P(A_2/\bar{A}_1) + P(\bar{A}_1)P(\bar{A}_2/\bar{A}_1)P(A_3/\bar{A}_1 \bar{A}_2) \\ &= 0.6 + 0.4 \times 0.4 + 0.4 \times 0.6 \times 0.3 = 0.832. \end{aligned}$$

(注: $A_1, \bar{A}_1 A_2, \bar{A}_1 \bar{A}_2 A_3$ 是互斥事件, 显然可见

$$A_1(\bar{A}_1 A_2) = (A_1 \bar{A}_1) A_2 = \emptyset,$$

$$A_1(\bar{A}_1 \bar{A}_2 A_3) = (A_1 \bar{A}_1) \bar{A}_2 A_3 = \emptyset.$$

$$(\bar{A}_1 A_2)(\bar{A}_1 \bar{A}_2 A_3) = (A_2 \bar{A}_2) \bar{A}_1 A_3 = \emptyset.$$

习题 1-4

1. 玻璃杯成箱出售，每箱20只，假设各箱含0, 1, 2只残次品的概率相应为0.8, 0.1和0.1。一顾客欲购一箱玻璃杯，在购买时售货员随意取一箱，而顾客随机地察看4只，若无残次品，则买下该箱玻璃杯，否则退回。试求：(1) 顾客买下该箱的概率 α ；(2) 在顾客买下的

一箱中，确实没有残次品的概率 β 。

解：设 $B_i =$ “一箱玻璃杯中有 i 只残次品”， $i = 0, 1, 2$ ，

$$P(B_0) = 0.8, P(B_1) = 0.1, P(B_2) = 0.1.$$

又设 $A =$ “该箱通过察看4只无残次品”。

$$P(A/B_0) = 1, P(A/B_1) = \frac{C_{19}^4}{C_{20}^4} = \frac{4}{5}, P(A/B_2) = \frac{C_{18}^4}{C_{20}^4} = \frac{12}{19}$$

$$(1) \alpha = P(A) = P(B_0)P(A/B_0) + P(B_1)P(A/B_1) + P(B_2)P(A/B_2) \\ = \frac{4}{5} \times 1 + \frac{1}{10} \times \frac{4}{5} + \frac{1}{10} \times \frac{12}{19} = \frac{448}{475} = 0.9432$$

$$(2) \beta = P(B_0/A) = \frac{P(A/B_0)P(B_0)}{P(A)} = \frac{P(B_0)P(A/B_0)}{P(A)} = \frac{\frac{4}{5} \times 1}{448/475} = \frac{95}{112} = 0.8482.$$

2. 两台车床加工同样的零件，第一台出现废品的概率是0.03，第二台出现废品的概率是0.02，加工的零件放在一起，并且已知第一台加工的零件比第二台加工的零件多一倍，求：(1) 任意取出的零件是合格的概率；(2) 如果任意取出的零件是废品，求它是第二台车床加工的概率。

解：设 $B_i =$ “第 i 台车床加工的零件”， $i = 1, 2$ 。

$$P(B_1) = \frac{2}{3}, P(B_2) = \frac{1}{3}.$$

又设 $A =$ “任取出来的零件是合格的”。

$$P(A/B_1) = 1 - P(\bar{A}/B_1) = 1 - 0.03 = 0.97;$$

$$P(A/B_2) = 1 - P(\bar{A}/B_2) = 1 - 0.02 = 0.98.$$

$$(1) P(A) = P(B_1)P(A/B_1) + P(B_2)P(A/B_2) = \frac{2}{3} \times 0.97 + \frac{1}{3} \times 0.98 = 0.973.$$

$$(2) P(B_2/\bar{A}) = \frac{P(\bar{A}/B_2)P(B_2)}{P(\bar{A})} = \frac{P(B_2)P(\bar{A}/B_2)}{1 - P(A)} = \frac{P(B_2)(1 - P(A/B_2))}{1 - P(A)} \\ = \frac{\frac{1}{3} \times (1 - 0.98)}{1 - 0.973} = \frac{0.2/3}{0.027} = 0.25.$$

3. 为防止意外，某矿井内同时设有两种报警系统A与B，每种系统单独使用时，其有效运行的概率，系统A为0.92，系统B为0.93，而在A失灵的条件下B有效的概率为0.85，求：

(1) 发生意外时, 这两个报警系统至少有一个有效的概率:

(2) B失灵条件下, A有效的概率.

解: 已知 $P(A) = 0.92$, $P(B) = 0.93$, $P(B/\bar{A}) = 0.85$.

$$\begin{aligned}(1) P\{\text{两系统至少有一个有效}\} &= P(A \cup B) = 1 - P(\overline{A \cup B}) \\&= 1 - P(\bar{A} \bar{B}) = 1 - P(\bar{A})P(\bar{B}/\bar{A}) = 1 - [1 - P(A)][1 - P(B/\bar{A})] \\&= 1 - (1 - 0.92)(1 - 0.85) = 0.988.\end{aligned}$$

$$\begin{aligned}(2) P\{\text{B失灵条件下A有效}\} &= P(A/\bar{B}) = 1 - P(\bar{A}/\bar{B}) = 1 - \frac{P(\bar{A} \bar{B})}{P(\bar{B})} \\&= 1 - \frac{(1 - P(A))(1 - P(B/\bar{A}))}{1 - P(B)} = 1 - \frac{(1 - 0.92)(1 - 0.85)}{1 - 0.93} = 0.829.\end{aligned}$$

4. 一批电子元件中, 甲类的占80%, 乙类的占12%, 丙类的占8%. 三类元件的使用寿命能达到指定要求的概率依次为0.9, 0.8和0.7. 今任取一个元件, 求其使用寿命达到指定要求的概率.

解: 设 $B_i =$ “一批 i 类的电子元件”, $i = \text{甲, 乙, 丙}$.

$$P(B_{\text{甲}}) = 0.8, P(B_{\text{乙}}) = 0.12, P(B_{\text{丙}}) = 0.08$$

又 $A =$ “元件使用寿命达到指定要求”.

$$P(A/B_{\text{甲}}) = 0.9, P(A/B_{\text{乙}}) = 0.8, P(A/B_{\text{丙}}) = 0.7.$$

$$\begin{aligned}\text{则 } P(A) &= P(B_{\text{甲}})P(A/B_{\text{甲}}) + P(B_{\text{乙}})P(A/B_{\text{乙}}) + P(B_{\text{丙}})P(A/B_{\text{丙}}) \\&= 0.8 \times 0.9 + 0.12 \times 0.8 + 0.08 \times 0.7 = 0.872.\end{aligned}$$

5. 甲袋中有3只白球4只红球, 乙袋中有5只白球, 2只红球, 从甲袋中任取2球投入乙袋, 再从乙袋中任取2球, 求最后取出的两只球全是白球的概率.

解: 设 $B_{ij} =$ “从甲袋中取出 i 只白球, j 只红球共2球”.

$$i = 0, 1, 2, j = 0, 1, 2, i + j = 2.$$

$$P(B_{20}) = \frac{C_3^2 C_4^0}{C_7^2} = \frac{3}{7}, P(B_{11}) = \frac{C_3^1 C_4^1}{C_7^2} = \frac{4}{7}, P(B_{02}) = \frac{C_3^0 C_4^2}{C_7^2} = \frac{2}{7}.$$

又 $A =$ “从乙袋中取出2只白球”.

$$P(A/B_{20}) = \frac{C_7^2 C_2^0}{C_9^2} = \frac{7}{12}, P(A/B_{11}) = \frac{C_6^2 C_3^0}{C_9^2} = \frac{5}{12}, P(A/B_{02}) = \frac{C_5^2 C_4^0}{C_9^2} = \frac{5}{18}.$$

$$\begin{aligned}\therefore P(A) &= P(B_{20})P(A/B_{20}) + P(B_{11})P(A/B_{11}) + P(B_{02})P(A/B_{02}) \\&= \frac{3}{7} \cdot \frac{7}{12} + \frac{4}{7} \cdot \frac{5}{12} + \frac{2}{7} \cdot \frac{5}{18} = 0.4008.\end{aligned}$$

习题 1-5

1. (1) 设 A, C 独立, B, C 独立, A, B 互斥. 证明: $A \cup B$ 与 C 独立.

证: $\because P(AC) = P(A)P(C)$, $P(BC) = P(B)P(C)$, $P(AB) = 0$. 有 $P(A \cup B) = P(A) + P(B)$
 由 $P[(A \cup B)C] = P(AC \cup BC) = P(AC) + P(BC) = P(A)P(C) + P(B)P(C)$
 $= (P(A) + P(B))P(C) = P(A \cup B) \cdot P(C)$. $\therefore A \cup B$ 与 C 独立.

(2) 设 A, B, C 独立, 证明: $A \cup B$ 与 \bar{C} 独立.

证: $P[(A \cup B)\bar{C}] = P(A\bar{C} \cup B\bar{C}) = P(A\bar{C}) + P(B\bar{C}) - P(AB\bar{C})$

由 $P(A\bar{C}) = P(A) - P(AC)$, $P(B\bar{C}) = P(B) - P(BC)$, $P(AB\bar{C}) = P(AB) - P(ABC)$
 且由 A, B, C 独立. 则 $P(AB) = P(A)P(B)$, $P(AC) = P(A)P(C)$, $P(ABC) = P(A)P(B)P(C)$.
 $\therefore P[(A \cup B)\bar{C}] = (P(A) - P(AC)) + (P(B) - P(BC)) - (P(AB) - P(ABC))$

$$= P(A)(1 - P(C)) + P(B)(1 - P(C)) - P(AB)(1 - P(C))$$

$$= (P(A) + P(B) - P(AB)) \cdot (1 - P(C)) = P(A \cup B) \cdot P(\bar{C}). \therefore A \cup B \text{ 与 } \bar{C} \text{ 独立.}$$

2. 甲、乙、丙三车间生产同种产品, 次品率分别为 0.05, 0.08 和 0.1. 从三个车间各取一件产品检查, 求下列事件的概率:

(1) 恰有 2 件次品; (2) 至少有 1 件次品.

解: 设 $A_i =$ “第 i 车间的一件产品是次品”. $i = 1, 2, 3$. 相互独立.

$$P(A_1) = 0.05, P(\bar{A}_1) = 0.95; P(A_2) = 0.08, P(\bar{A}_2) = 0.92; P(A_3) = 0.1, P(\bar{A}_3) = 0.9.$$

$$(1) P(\text{恰有 2 件次品}) = P(A_1 A_2 \bar{A}_3 \cup A_1 \bar{A}_2 A_3 \cup \bar{A}_1 A_2 A_3)$$

$$= P(A_1 A_2 \bar{A}_3) + P(A_1 \bar{A}_2 A_3) + P(\bar{A}_1 A_2 A_3) = P(A_1)P(A_2)P(\bar{A}_3) + P(A_1)P(\bar{A}_2)P(A_3)$$

$$+ P(\bar{A}_1)P(A_2)P(A_3) = 0.05 \times 0.08 \times 0.9 + 0.05 \times 0.92 \times 0.1 + 0.95 \times 0.08 \times 0.1$$

$$= 0.0158.$$

$$(2) P(\text{至少有 1 件次品}) = P(A_1 \cup A_2 \cup A_3) = 1 - P(\bar{A}_1 \bar{A}_2 \bar{A}_3) = 1 - P(\bar{A}_1)P(\bar{A}_2)P(\bar{A}_3)$$

$$= 1 - 0.95 \times 0.92 \times 0.9 = 0.2134.$$

3. 一个工人看管三台车床, 在一小时内车床不需要工人照管的概率: 第一台等于 0.9, 第二台等于 0.8, 第三台等于 0.7. 求一小时内三台车床中最多有一台需要工人照管的概率.

解: 设 $A_i =$ “第 i 台车床需要工人照管”. $i = 1, 2, 3$ 且相互独立.

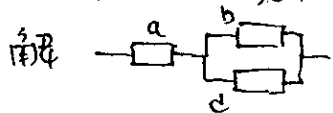
$$P(\text{最多有一台需要工人照管}) = P\{\bar{A}_1 \bar{A}_2 \bar{A}_3 \cup \bar{A}_1 \bar{A}_2 A_3 \cup \bar{A}_1 A_2 \bar{A}_3 \cup A_1 \bar{A}_2 \bar{A}_3\}$$

$$= P(\bar{A}_1 \bar{A}_2 \bar{A}_3) + P(\bar{A}_1 \bar{A}_2 A_3) + P(\bar{A}_1 A_2 \bar{A}_3) + P(A_1 \bar{A}_2 \bar{A}_3) = P(\bar{A}_1)P(\bar{A}_2)P(\bar{A}_3) + P(\bar{A}_1)P(\bar{A}_2)P(A_3)$$

$$+ P(\bar{A}_1)P(A_2)P(\bar{A}_3) + P(A_1)P(\bar{A}_2)P(\bar{A}_3) = 0.9 \times 0.8 \times 0.7 + 0.9 \times 0.8 \times (1 - 0.7)$$

$$+ 0.9 \times (1 - 0.8) \times 0.7 + (1 - 0.9) \times 0.8 \times 0.7 = 0.902$$

4. 电路由电池a与两个并联的电池b及c串联而成. 设电池a, b, c损坏的概率分别为0.3, 0.2和0.2. 求电路发生中断的概率.



解 设电池a, b, c正常工作的事件依次为A, B, C且相互独立. $P(A) = 0.7, P(\bar{A}) = 0.3; P(B) = 0.8$

$$P(\bar{B}) = 0.2; P(C) = 0.8, P(\bar{C}) = 0.2.$$

$$P\{\text{电路发生中断}\} = P(\bar{A} \cup \bar{B}\bar{C}) = P(\bar{A}) + P(\bar{B}\bar{C}) - P(\bar{A}\bar{B}\bar{C})$$

$$= P(\bar{A}) + P(\bar{B})P(\bar{C}) - P(\bar{A})P(\bar{B})P(\bar{C}) = 0.3 + 0.2 \times 0.2 - 0.3 \times 0.2 \times 0.2 = 0.328.$$

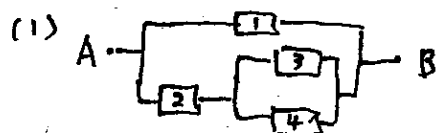
(注): $P\{\text{电路发生中断}\} = 1 - P\{\text{电路正常}\} = 1 - P\{A(B \cup C)\}$

$$= 1 - P(AB \cup AC) = 1 - [P(AB) + P(AC) - P(ABC)]$$

$$= 1 - [P(A)P(B) + P(A)P(C) - P(A)P(B)P(C)] = 1 - [0.7 \times 0.8 + 0.7 \times 0.8 - 0.7 \times 0.8 \times 0.8]$$

$$= 0.328.$$

5. 设下列系统每个部件的可靠性都是 γ . 且各部件能否正常工作是相互独立的. 已知A至B只要有一条通路正常工作, 系统便能正常运行. 求各系统的可靠性.



设 $E_i =$ "第*i*个部件正常工作", $i = 1, 2, 3, 4$.

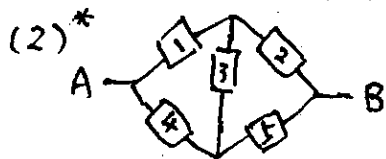
$P(E_i) = \gamma$. 且各 $E_i, i = 1, 2, 3, 4$ 相互独立.

$$P\{\text{系统正常}\} = P\{E_1 \cup [E_2(E_3 \cup E_4)]\} = P\{E_1 \cup (E_2 E_3 \cup E_2 E_4)\}$$

$$= P(E_1) + P(E_1 E_2) + P(E_2 E_4) - P(E_1 E_2 E_3) - P(E_1 E_2 E_4) - P(E_2 E_3 E_4) + P(E_1 E_2 E_3 E_4)$$

$$= P(E_1) + P(E_1)P(E_2) + P(E_2)P(E_4) - P(E_1)P(E_2)P(E_3) - P(E_1)P(E_2)P(E_4) - P(E_2)P(E_3)P(E_4) + P(E_1)P(E_2)P(E_3)P(E_4)$$

$$= \gamma + \gamma^2 + \gamma^2 - \gamma^3 - \gamma^3 - \gamma^3 + \gamma^4 = \gamma + 2\gamma^2 - 3\gamma^3 + \gamma^4.$$



设 $E_i =$ "第*i*个部件正常工作", $i = 1, 2, 3, 4, 5$

且相互独立. $P(E_1) = P(E_2) = P(E_3) = P(E_4) = P(E_5) = \gamma$.

$$P\{\text{系统正常}\} = P\{E_1 E_2 \cup E_1 E_3 E_4 \cup E_4 E_5 \cup E_2 E_3 E_5\} = P(E_1 E_2) + P(E_1 E_3 E_4)$$

$$+ P(E_4 E_5) + P(E_2 E_3 E_5) - P(E_1 E_2 E_3 E_4) - P(E_1 E_2 E_4 E_5) - P(E_1 E_2 E_3 E_5)$$

$$- P(E_1 E_3 E_4 E_5) - P(E_2 E_3 E_4 E_5) - P(E_1 E_2 E_3 E_4 E_5) + 4P(E_1 E_2 E_3 E_4 E_5)$$

$$= P(E_1)P(E_2) + P(E_1)P(E_3)P(E_4) + P(E_4)P(E_5) + P(E_2)P(E_3)P(E_5)$$

$$- P(E_1)P(E_2)P(E_3)P(E_4) - P(E_1)P(E_2)P(E_4)P(E_5) - P(E_1)P(E_2)P(E_3)P(E_5)$$

$$- P(E_1)P(E_3)P(E_4)P(E_5) - P(E_1)P(E_3)P(E_5)P(E_4) + 2P(E_1)P(E_2)P(E_3)P(E_4)P(E_5)$$

$$= \gamma^2 + \gamma^3 + \gamma^2 + \gamma^3 - \gamma^4 - \gamma^4 - \gamma^4 - \gamma^4 - \gamma^4 + 2\gamma^5$$

$$= 2\gamma^2 + 2\gamma^3 - 5\gamma^4 + 2\gamma^5$$

6. 甲、乙、丙三人向同一飞机射击，设击中的概率分别是0.4, 0.5和0.7. 如果只有一人击中，则飞机被击落的概率是0.2；如果有两人击中，则飞机被击落的概率是0.6；如果三人都击中，则飞机一定被击落. 求飞机被击落的概率.

解：设 $B_i = \text{"有 } i \text{ 人击中飞机"} \quad i = 0, 1, 2, 3$.

又记甲、乙、丙各自击中飞机的事件依次为 C_1, C_2, C_3 ，且相互独立
 则 $P(B_0) = P(\bar{C}_1 \bar{C}_2 \bar{C}_3) = P(\bar{C}_1)P(\bar{C}_2)P(\bar{C}_3) = 0.6 \times 0.5 \times 0.3 = 0.09$;

$$P(B_1) = P(C_1 \bar{C}_2 \bar{C}_3 \cup \bar{C}_1 C_2 \bar{C}_3 \cup \bar{C}_1 \bar{C}_2 C_3) = P(C_1 \bar{C}_2 \bar{C}_3) + P(\bar{C}_1 C_2 \bar{C}_3) + P(\bar{C}_1 \bar{C}_2 C_3) \\ = P(C_1)P(\bar{C}_2)P(\bar{C}_3) + P(\bar{C}_1)P(C_2)P(\bar{C}_3) + P(\bar{C}_1)P(\bar{C}_2)P(C_3) \\ = 0.4 \times 0.5 \times 0.3 + 0.6 \times 0.5 \times 0.3 + 0.6 \times 0.5 \times 0.7 = 0.36$$

$$P(B_2) = P(C_1 C_2 \bar{C}_3 \cup C_1 \bar{C}_2 C_3 \cup \bar{C}_1 C_2 C_3) = P(C_1 C_2 \bar{C}_3) + P(C_1 \bar{C}_2 C_3) + P(\bar{C}_1 C_2 C_3) \\ = P(C_1)P(C_2)P(\bar{C}_3) + P(C_1)P(\bar{C}_2)P(C_3) + P(\bar{C}_1)P(C_2)P(C_3) \\ = 0.4 \times 0.5 \times 0.3 + 0.4 \times 0.5 \times 0.7 + 0.6 \times 0.5 \times 0.7 = 0.41$$

$$P(B_3) = P(C_1 C_2 C_3) = P(C_1)P(C_2)P(C_3) = 0.4 \times 0.5 \times 0.7 = 0.14$$

且设 $A = \text{"飞机被击落"}$. 则有 $P(A/B_0) = 0$; $P(A/B_1) = 0.2$; $P(A/B_2) = 0.6$;
 $P(A/B_3) = 1$.

$$\text{则 } P(A) = P(B_0)P(A/B_0) + P(B_1)P(A/B_1) + P(B_2)P(A/B_2) + P(B_3)P(A/B_3) \\ = 0.09 \times 0 + 0.36 \times 0.2 + 0.41 \times 0.6 + 0.14 \times 1 = 0.458$$

7. 甲、乙、丙三人同时破译一份密码，已知三人能译出的概率分别是 $\frac{1}{3}$, $\frac{1}{4}$ 和 $\frac{1}{5}$. 求密码能译出的概率.

解：设甲、乙、丙各人译出密码的事件分别为 A, B, C ，相互独立. 且

$$P(A) = \frac{1}{3}, P(B) = \frac{1}{4}, P(C) = \frac{1}{5}; P(\bar{A}) = \frac{2}{3}, P(\bar{B}) = \frac{3}{4}, P(\bar{C}) = \frac{4}{5}$$

$$\text{则 } P\{\text{密码能译出}\} = 1 - P\{\text{密码不能译出}\} = 1 - P(\bar{A} \bar{B} \bar{C}) \\ = 1 - P(\bar{A})P(\bar{B})P(\bar{C}) = 1 - \frac{2}{3} \cdot \frac{3}{4} \cdot \frac{4}{5} = \frac{3}{5} = 0.6$$

$$\text{(注): } P\{\text{密码能译出}\} = P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(AB) - P(AC) \\ - P(BC) + P(ABC) = P(A) + P(B) + P(C) - P(A)P(B) - P(A)P(C) - P(B)P(C) \\ + P(A)P(B)P(C) = \frac{1}{3} + \frac{1}{4} + \frac{1}{5} - \frac{1}{3 \times 4} - \frac{1}{3 \times 5} - \frac{1}{4 \times 5} + \frac{1}{3 \times 4 \times 5} \\ = \frac{1}{60} (20 + 15 + 12 - 5 - 4 - 3 + 1) = \frac{36}{60} = 0.6$$

复习题 1 (三. 计算题与证明题)

1. 设 x 为自然数 1 至 100 中随机地选取的一个数. 求 $x + \frac{100}{x} > 50$ 的概率.

解: 由 $x + \frac{100}{x} > 50$ 得 $x^2 - 50x + 100 > 0$, 即 $(x - 2.09)(x - 47.91) > 0$.
 则有 $x < 2.09$ 或 $x > 47.91$. 在 1~100 的自然数中有 $x = 1, 2, 48, 49, 50, \dots, 100$, 共有 55 个数. 则

$$P\{x + \frac{100}{x} > 50\} = \frac{55}{100} = 0.55.$$

2. 掷两颗骰子, 求下列各事件的概率: $A =$ "两颗骰子出现点数不同",
 $B =$ "两颗骰子出现点数之和等于 7", $C =$ "两颗骰子出现点数之和大于 5 且小于 9".

解: 记两颗骰子出现的点数 $\omega = (i, j)$, $i, j = 1, 2, 3, 4, 5, 6$. 则

$$A = \{(1, 2), (1, 3), \dots, (1, 6); (2, 1), (2, 3), \dots, (2, 6); (3, 1), (3, 2), (3, 4), \dots, (3, 6);$$

$$(4, 1), \dots, (4, 3), (4, 5), (4, 6); (5, 1), \dots, (5, 4), (5, 6); (6, 1), \dots, (6, 5)\}. \text{共 } 30 \text{ 种.}$$

$$B = \{(1, 6), (2, 5), (3, 4), (4, 3), (5, 2), (6, 1)\}. \text{共 } 6 \text{ 种.}$$

$$C = \{(1, 5), (1, 6), (2, 4), (2, 5), (2, 6); (3, 3), (3, 4), (3, 5); (4, 2), (4, 3), (4, 4);$$

$$(5, 1), (5, 2), (5, 3), (6, 1), (6, 2)\}. \text{共 } 16 \text{ 种.}$$

$$\text{故 } P(A) = \frac{30}{36} = \frac{5}{6}; P(B) = \frac{6}{36} = \frac{1}{6}; P(C) = \frac{16}{36} = \frac{4}{9}.$$

3. 从 1, 2, \dots , 10 这 10 个数字中任取 1 个, 然后放回, 先后取出 6 个数字. 求下列各事件的概率: $A =$ "6 个数字全不相同"; $B =$ "6 个数字不含 10 和 1".

$$\text{解: } P(A) = \frac{A_6^{10}}{10^6} = 0.060, P(B) = \frac{8^6}{10^6} = (0.8)^6 = 0.262.$$

4. 证明: 若 $P(A/B) = P(A/\bar{B})$, 则事件 A 与 B 独立.

$$\text{证: } P(A/B) = P(A/\bar{B}) \Rightarrow P(AB)/P(B) = P(A\bar{B})/P(\bar{B})$$

$$\Rightarrow P(AB)P(\bar{B}) = P(A\bar{B})P(B) \Rightarrow P(AB)[1 - P(B)] = [P(A) - P(AB)]P(B).$$

$$\Rightarrow P(AB) - P(AB)P(B) = P(A)P(B) - P(AB)P(B) \Rightarrow P(AB) = P(A)P(B).$$

$\therefore A, B$ 独立.

5. 一个工厂有 \dots 3 个车间生产同一产品, 每个车间的产量占总产量的 45%, 35%, 20%. 如果每个车间成品中的次品率分别为 5%, 4%, 2%.

(1) 从全厂产品中任意抽取一个产品, 求取出是次品的概率;

(2) 如果从全厂产品中抽出的一个恰好是次品, 求这个产品是由一车间生产的概率.

解: 设 $B_i =$ "第 i 个车间生产的产品", $i = 1, 2, 3$

$$\text{则 } P(B_1) = 0.45, P(B_2) = 0.35, P(B_3) = 0.20.$$

又 $A =$ "抽取出的一个产品是次品".

$$\text{则 } P(A/B_1) = 0.05, P(A/B_2) = 0.04, P(A/B_3) = 0.02$$

$$(1) P(A) = P(B_1)P(A/B_1) + P(B_2)P(A/B_2) + P(B_3)P(A/B_3) \\ = 0.45 \times 0.05 + 0.35 \times 0.04 + 0.2 \times 0.02 = 0.0405.$$

$$(2) P(B_1/A) = P(AB_1)/P(A) = P(B_1)P(A/B_1)/P(A) = 0.45 \times 0.05 / 0.0405 \approx 0.56.$$

6. 寝室中有四个人, 求: (1) 至少有 2 人的生日同在 12 月的概率; (2) 至少有 2 人的生日在同一个月的概率; (3) 至少有 2 人的生日同在星期一的概率.

解: (1) $P\{\text{至少有 2 人的生日同在 12 月}\} = (C_4^2 \cdot 11^2 + C_4^3 \cdot 11 + 1) / 12^4 = 0.0372.$

$$(2) P\{\text{至少有 2 人的生日在同一个月}\} = 1 - P\{\text{四人生日不在同一个月}\} \\ = 1 - A_{12}^4 / 12^4 = 0.4271.$$

$$(3) P\{\text{至少有 2 人的生日同在星期一}\} = (C_4^2 \cdot 6^2 + C_4^3 \cdot 6 + 1) / 7^4 = 0.1004.$$

(注): (1), (3) 可用第 5 章的方法处理: 如 (1) 题:

设随机变量 $X_i = \begin{cases} 1, & \text{第 } i \text{ 人生日在 12 月;} \\ 0, & \text{第 } i \text{ 人生日不在 12 月.} \end{cases} \quad i=1, 2, 3, 4 \text{ 相互独立.}$

$$\text{且 } P\{X_i=1\} = \frac{1}{12}, \quad P\{X_i=0\} = \frac{11}{12}, \quad i=1, 2, 3, 4.$$

则随机变量 $X = \text{"4 人中生日在 12 月的人数"} = \sum_{i=1}^4 X_i \sim B(4, \frac{1}{12}).$

$$\therefore P\{\text{至少有 2 人的生日在 12 月}\} = P\{X \geq 2\} = \sum_{i=2}^4 P\{X=i\} = \sum_{i=2}^4 C_4^i \left(\frac{1}{12}\right)^i \left(\frac{11}{12}\right)^{4-i} \\ = C_4^2 \cdot \frac{11^2}{12^4} + C_4^3 \cdot \frac{11}{12^4} + C_4^4 \cdot \frac{1}{12^4} = 0.0372.$$

第 (3) 题可类似处理.

7. 从 5 双不同尺码的鞋子中任取 4 只, 求至少有两只凑成一双的概率.

解: 记 $A = \text{"所取 4 只鞋子中至少有两只凑成一双"}$.

$$P(A) = 1 - P(\bar{A}) = 1 - \frac{C_5^4 (C_2^1)^4}{C_{10}^4} = 1 - \frac{8}{21} = \frac{13}{21}.$$

$$(注): P(A) = \frac{C_5^1 \cdot C_2^2 + C_4^2 (C_2^1)^2}{C_{10}^4} + \frac{C_5^2 C_2^2 C_2^1}{C_{10}^4} = \frac{13}{21}.$$

本题易错解为 $P(A) = \frac{C_5^1 C_2^2 \cdot C_2^2}{C_{10}^4}$, 则分子中样本点有重复计数之错误.

9. 已知 $P(A) = P(B) = P(C) = \frac{1}{4}$, $P(AB) = 0$, $P(AC) = P(BC) = \frac{1}{16}$. 求下列事件的概率: (1) A, B, C 全不发生; (2) A, B, C 恰好发生一个.

解: $\because P(A) = P(B) \neq 0, P(AB) = 0, \therefore A, B$ 互斥, $AB = \emptyset$.

$$P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(AB) - P(AC) - P(BC) + P(ABC) \\ = P(A) + P(B) + P(C) - P(AC) - P(BC) = \frac{3}{4} - \frac{2}{16} = \frac{5}{8}.$$

$$(1) P\{A, B, C \text{ 全不发生}\} = P(\bar{A} \bar{B} \bar{C}) = 1 - P(A \cup B \cup C) = 1 - \frac{5}{8} = \frac{3}{8} = 0.375.$$

$$(2) P\{A, B, C \text{ 恰好发生一个}\} = P(A \bar{B} \bar{C} \cup \bar{A} B \bar{C} \cup \bar{A} \bar{B} C)$$

$$\begin{aligned}
 &= P(AB\bar{C}) + P(\bar{A}B\bar{C}) + P(\bar{A}\bar{B}C) = (P(A) - P[A(B \cup C)]) + (P(B) - P[B(A \cup C)]) \\
 &+ (P(C) - P[C(A \cup B)]) = P(A) + P(B) + P(C) - P(AB \cup AC) - P(AB \cup BC) - P(AC \cup BC) \\
 &= P(A) + P(B) + P(C) - [P(AB) + P(AC) - P(ABC)] - [P(AB) + P(BC) - P(ABC)] \\
 &\quad - [P(AC) + P(BC) - P(ABC)] = P(A) + P(B) + P(C) - P(AC) - P(BC) - P(AC) - P(BC) \\
 &= \frac{1}{4} + \frac{1}{4} + \frac{1}{4} - \frac{1}{16} - \frac{1}{16} - \frac{1}{16} - \frac{1}{16} = \frac{8}{16} = \frac{1}{2}.
 \end{aligned}$$

(注): (2) 中使用事件运算关系: 例如 $P(A\bar{B}\bar{C}) = P(A) - P(A\overline{B\bar{C}})$

$= P(A) - P(A(B \cup C))$. 是由 $A = A\Omega = A(\bar{B}\bar{C} \cup \overline{B\bar{C}}) = A\bar{B}\bar{C} \cup A\overline{B\bar{C}}$

则 $P(A) = P(A\bar{B}\bar{C}) + P(A\overline{B\bar{C}})$, 则 $P(A\bar{B}\bar{C}) = P(A) - P(A\overline{B\bar{C}})$. 其余同理

10. 假设有两箱同种零件, 第一箱内装 50 件, 其中 10 件一等品, 第二箱内装 30 件, 其中 18 件一等品, 现从两箱中先后随机取出 2 个零件 (取出的零件不放回). 试求: (1) 先取出的零件是一等品的概率; (2) 在先取出的零件是一等品的条件下, 第二次取出的零件仍是一等品的条件概率.

解: 设 B_i = "第 i 箱的零件", 且 $P(B_i) = \frac{1}{2}$, $i = 1, 2$.

又 A_i = "第 i 次取出一个零件是一等品", $i = 1, 2$.

$$(1) P(A_1/B_1) = C_{10}^1 / C_{50}^1 = \frac{1}{5}; P(A_1/B_2) = C_{18}^1 / C_{30}^1 = \frac{3}{5}.$$

$$\therefore P\{\text{先取出的零件是一等品}\} = P(A_1) = P(B_1)P(A_1/B_1) + P(B_2)P(A_1/B_2) = \frac{1}{2} \times \frac{1}{5} + \frac{1}{2} \times \frac{3}{5}$$

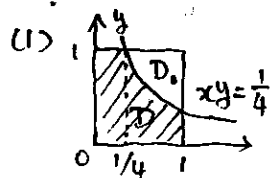
$$(2) P(A_2/A_1) = \frac{P(A_1 A_2)}{P(A_1)} = \frac{1}{P(A_1)} [P(B_1)P(A_1 A_2/B_1) + P(B_2)P(A_1 A_2/B_2)]$$

$$\text{由 } P(A_1 A_2/B_1) = C_{10}^2 / C_{50}^2 = 0.03673; P(A_1 A_2/B_2) = C_{18}^2 / C_{30}^2 = 0.35172.$$

$$\therefore P(A_2/A_1) = \frac{1}{0.4} \left[\frac{1}{2} \times 0.03673 + \frac{1}{2} \times 0.35172 \right] = 0.4856.$$

8. 在区间 $(0, 1)$ 中随机地取 2 个数, 求: (1) 两数之积小于 $\frac{1}{4}$ 的事件的概率, (2) 两数之和大于 1.2 的事件的概率.

解: 设 x, y 为所取的 2 个数, $0 \leq x \leq 1, 0 \leq y \leq 1$.

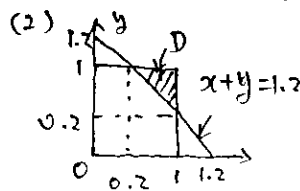


$$P(xy < \frac{1}{4}) = S_{D_0} / S_{\text{正方形}} = 1 - S_{D_1} / S_{\text{正方形}}$$

$$S_{D_0} = \iint_{D_0} dx dy = \int_{1/4}^1 dx \int_{1/4x}^1 dy = \int_{1/4}^1 (1 - \frac{1}{4x}) dx = (x - \frac{\ln x}{4}) \Big|_{1/4}^1 = \frac{3}{4} - \frac{\ln 2}{2}.$$

$$S_{\text{正方形}} = 1.$$

$$\therefore P(xy < \frac{1}{4}) = 1 - (\frac{3}{4} - \frac{\ln 2}{2}) = \frac{1}{4} + \frac{1}{2} \ln 2.$$



$$P(x + y > 1.2) = S_D / S_{\text{正方形}}$$

$$S_D = 0.8 \times 0.8 / 2 = 0.32. S_{\text{正方形}} = 1.$$

$$\therefore P\{x + y > 1.2\} = 0.32.$$

习题 2-1

1. 1014 产品中有 8 件合格品和 2 件不合格品, 从中任取 3 次, 每次取一件, 分别依照 (1) 放回, (2) 不放回方式, 求取得不合格品数 X 的分布律.

解: $X =$ "从 1014 产品中任取 3 次, 每次一件, 取得不合格品的个数".

(1) 放回方式:

$$P(X=K) = \frac{C_3^K 2^K 8^{3-K}}{10^3} = C_3^K \left(\frac{2}{10}\right)^K \left(\frac{8}{10}\right)^{3-K}, \quad K=0, 1, 2, 3.$$

(2) 不放回方式:

$$P(X=K) = \frac{C_2^K C_8^{3-K}}{C_{10}^3}, \quad K=0, 1, 2.$$

2. 掷 2 颗骰子, 记点数之和为 X , (1) 写出 X 的分布; (2) 计算 $P(X \geq 6 | X \geq 3)$.

解: $X =$ "2 颗骰子点数之和".

(1) X 的分布律:

X	2	3	4	5	6	7	8	9	10	11	12
P_K	$1/36$	$2/36$	$3/36$	$4/36$	$5/36$	$6/36$	$5/36$	$4/36$	$3/36$	$2/36$	$1/36$

$$(2) P(X \geq 6 | X \geq 3) = P((X \geq 6) \cap (X \geq 3)) / P(X \geq 3) = P(X \geq 6) / P(X \geq 3) \\ = \frac{\sum_{k=6}^{12} P(X=k)}{\sum_{k=3}^{12} P(X=k)} = \frac{26/36}{35/36} = \frac{26}{35}.$$

3. 袋中有 5 只球, 编号为 1, 2, 3, 4, 5, 从中同时取 3 只, 设 X 为取出的 3 只球中的最大号码, 写出 X 的分布律.

解: $P(X=K) = \frac{C_{K-1}^2}{C_5^3}, \quad K=3, 4, 5$ 或

X	3	4	5
P_K	$1/10$	$3/10$	$6/10$

4. 设随机变量 X 具有分布律:

X	0	1	2	3
P_K	$1/9$	$2\theta(1-\theta)$	$1/9$	$1-2\theta$

试确定常数 θ .

解: 由 $1 = \sum_{k=0}^3 P(X=k) = \frac{1}{9} + 2\theta(1-\theta) + \frac{1}{9} + (1-2\theta) = \frac{11}{9} - 2\theta^2$

$\therefore \theta^2 = \frac{1}{9}, \quad \therefore \theta > 0, \quad \text{故 } \theta = \frac{1}{3}.$

习题 2-2.

1. 一条自动生产线上产品的一级品率为 0.6, 随机检查 10 件, 求至少有两件一级品的概率.

解: 设 $X =$ “随机检查 10 件中的一级品个数”. 则 $X \sim B(10, 0.6)$

$$\therefore P\{\text{至少有 2 件一级品}\} = P\{X \geq 2\} = 1 - P\{X < 2\} = 1 - P(X=0) - P(X=1) \\ = 1 - C_{10}^0 (0.6)^0 (0.4)^{10} - C_{10}^1 0.6 \cdot (0.4)^9 = 1 - 0.000105 - 0.001573 = 0.9983.$$

2. 设从学校乘汽车到火车站的途中有 5 个十字路口, 每个十字路口遇到红灯的事件是相互独立的, 并且概率都等于 0.6. 以 X 表示途中遇到红灯的次数, 求 X 的分布律.

解: 设 $X_i = \begin{cases} 1, & \text{第 } i \text{ 个十字路口遇到红灯;} \\ 0, & \text{第 } i \text{ 个十字路口未遇到红灯.} \end{cases}$ $P(X_i=1) = 0.6, i=1, 2, \dots, 5$
而 X_1, X_2, X_3, X_4, X_5 相互独立.

则 $X =$ “途中遇到红灯的次数” $= \sum_{i=1}^5 X_i \sim B(5, 0.6)$.

有分布律 $P(X=k) = C_5^k (0.6)^k (0.4)^{5-k}, k=0, 1, 2, 3, 4, 5$.

3. 某种灯泡使用时数在 1500 小时以上的概率为 0.7, 求 5 个灯泡中至少有 3 个能使用 1500 小时以上的概率.

解: 设 $X_i = \begin{cases} 1, & \text{第 } i \text{ 个灯泡能使用 1500 小时以上;} \\ 0, & \text{第 } i \text{ 个灯泡不能使用 1500 小时以上.} \end{cases}$ $P(X_i=1) = 0.7$
且 X_1, X_2, X_3, X_4, X_5 相互独立.

又设 $X =$ “5 个灯泡中能使用 1500 小时以上的个数”

则 $X = \sum_{i=1}^5 X_i \sim B(5, 0.7)$. 有 $P\{X=k\} = C_5^k (0.7)^k (0.3)^{5-k}, k=0, 1, 2, \dots, 5$.

$$\text{故 } P\{X \geq 3\} = \sum_{k=3}^5 P\{X=k\} = \sum_{k=3}^5 C_5^k (0.7)^k (0.3)^{5-k} = C_5^3 (0.7)^3 (0.3)^2 + C_5^4 (0.7)^4 (0.3) \\ + C_5^5 (0.7)^5 = 0.3087 + 0.36015 + 0.16807 = 0.837.$$

4. 一堆种子发芽率为 0.98, 任取其中 5 粒, 求以下概率:

(1) 恰有 3 粒种子能发芽; (2) 至少有 4 粒种子能发芽.

解: 设 $X_i = \begin{cases} 1, & \text{第 } i \text{ 粒种子能发芽;} \\ 0, & \text{第 } i \text{ 粒种子不能发芽.} \end{cases}$ $P\{X_i=1\} = 0.98, i=1, 2, \dots, 5$
 X_1, X_2, X_3, X_4, X_5 相互独立.

又由 $X =$ “5 粒种子中能发芽的粒数” $= \sum_{i=1}^5 X_i \sim B(5, 0.98)$. 故

$$(1) P(\text{恰有 3 粒种子能发芽}) = P(X=3) = C_5^3 (0.98)^3 (0.02)^2 = 0.003765.$$

$$(2) P(\text{至少有 4 粒种子能发芽}) = P(X \geq 4) = P(X=4) + P(X=5)$$

$$= C_5^4 (0.98)^4 \cdot 0.02 + C_5^5 (0.98)^5 = 0.09224 + 0.90392 = 0.9962.$$

5. 一射手对同一目标独立地进行 4 次射击, 若至少命中一次的概率为 $\frac{80}{81}$, 求该射手的命中率.

解: 设 $X =$ “4次射击中击中目标的次数”, 易知 $X \sim B(4, p)$.

其中 p 为该射手的命中率. $0 < p < 1$. 由 $P(X \geq 1) = 80/81$.

$$\therefore P(X \geq 1) = 1 - P(X < 1) = 1 - P(X = 0) = 1 - C_4^0 p^0 (1-p)^4 = 1 - (1-p)^4$$

$$\therefore 1 - (1-p)^4 = \frac{80}{81} \quad \text{即} \quad (1-p)^4 = \frac{1}{81}, \quad \therefore p = \frac{2}{3}.$$

6. 一条流水线上产品合格率为 0.9, 合格品中有 80% 为一级品, 从该厂产品中任取 10 件, 求 (1) 取到 7 件合格品, 3 件不合格品的概率; (2) 至少取到 8 件一级品的概率; (3) 已知其中有一件不是一级品, 求非一级品数不超过 2 件的概率.

解: (1) 设 $X =$ “任取 10 件中的合格品个数”, 则 $X \sim B(10, 0.9)$

$$\therefore P(X=7) = C_{10}^7 (0.9)^7 (0.1)^3 = 0.0574.$$

(2) 设 $Y =$ “任取 10 件中的一级品个数”, 则 $Y \sim B(10, 0.9 \times 0.8)$

即 $Y \sim B(10, 0.72)$.

$$\begin{aligned} \therefore P(X \geq 8) &= P(X=8) + P(X=9) + P(X=10) = C_{10}^8 (0.72)^8 (0.28)^2 + C_{10}^9 (0.72)^9 (0.28) \\ &\quad + C_{10}^{10} (0.72)^{10} = 0.25479 + 0.14560 + 0.03744 = 0.4378. \end{aligned}$$

(3) 由 (2) $Y \sim B(10, 0.72)$.

$\therefore P(\text{已知其中有一件不是一级品的条件下, 非一级品数不超过 2 件})$

$$= P(X \geq 8 / Y \leq 9) = \frac{P(8 \leq Y \leq 9)}{P(Y \leq 9)} = \frac{P(Y=8) + P(Y=9)}{1 - P\{Y=10\}}$$

$$= \frac{C_{10}^8 (0.72)^8 (0.28)^2 + C_{10}^9 (0.72)^9 (0.28)}{1 - C_{10}^{10} (0.72)^{10}} = \frac{0.25479 + 0.14560}{1 - 0.03744} = 0.4160$$

习题 2-3

1. 设某本书中每页印刷错误的个数 X 服从泊松分布 $\pi(0.2)$. 求一页上至多有一个印刷错误的概率.

解: $X = \text{"每页印刷错误的个数"} , X \sim \pi(0.2) . P(X=k) = \frac{0.2^k}{k!} e^{-0.2} , k=0,1,\dots$
 则 $P(X \leq 1) = P(X=0) + P(X=1) = \frac{0.2^0}{0!} e^{-0.2} + \frac{0.2^1}{1!} e^{-0.2} = e^{-0.2} + 0.2e^{-0.2} = 1.2e^{-0.2} = 0.9825 .$

2. 设某电话总机 5 分钟内接到电话呼叫的次数 X 服从泊松分布 $\pi(2)$.

(1) 计算该总机 5 分钟内共接到 K 个电话 ($K=0,1,\dots,6$) 的概率; (2) 求 5 分钟内至多接到 3 个电话的概率.

解: $X = \text{"5 分钟内接到电话呼叫次数"} , X \sim \pi(2)$

$$P(X=k) = \frac{2^k}{k!} e^{-2} = \frac{2^k}{k!} \times 0.13535 , k=0,1,2,\dots$$

(1)	X	0	1	2	3	4	5	6
	P_k	0.13535	0.2707	0.2707	0.18047	0.09023	0.03609	0.01203

$$(2) P(X \leq 3) = P(X=0) + P(X=1) + P(X=2) + P(X=3)$$

$$= 0.13535 + 0.2707 + 0.2707 + 0.18047 = 0.85722 .$$

3. 某商店某种商品的月销售量服从参数为 5 的泊松分布. 问在月初应库存多少该种商品, 才能保证当月不脱销的概率达到 0.999?

解: $X = \text{"某商品的月销售量"} , X \sim \pi(5) , P(X=k) = \frac{5^k}{k!} e^{-5} , k=0,1,2,\dots$
 设月初应库存 K 个该商品. 使 $P(X \leq K) = 0.999$.

$$\text{即 } P(X \leq K) = \sum_{k=0}^K \frac{5^k}{k!} e^{-5} = 0.999 . \text{ 经试算得 } K=12 .$$

4. 某医院在长度为 t 的时间间隔内收治的急诊病人数 X 服从参数为 $\frac{5}{2}$ 的泊松分布, 而 t 时间间隔的起点无关 (时间以小时计).

(1) 求某一天中午 12 时至下午 3 时没有急诊病人的概率;

(2) 求某一天中午 12 时至下午 5 时至少有 2 个急诊病人的概率.

解: (1) 设 $X = \text{"中午 12 时至下午 3 时的急诊病人数"} .$ 则 $X \sim \pi(\frac{3}{2})$.

$$\therefore P\{\text{没有急诊病人}\} = P(X=0) = \frac{(\frac{3}{2})^0}{0!} e^{-\frac{3}{2}} = e^{-\frac{3}{2}} = 0.2231 .$$

(2) 设 $Y = \text{"中午 12 时至下午 5 时的急诊病人数"} .$ 则 $Y \sim \pi(\frac{5}{2})$.

$$\therefore P\{\text{至少有 2 个急诊病人}\} = P\{Y \geq 2\} = 1 - P\{Y < 2\}$$

$$= 1 - P(Y=0) - P(Y=1) = 1 - e^{-\frac{5}{2}} - \frac{5}{2} e^{-\frac{5}{2}} = 1 - \frac{7}{2} e^{-\frac{5}{2}}$$

$$= 1 - 0.287297 = 0.7127 .$$

习题 2-4

1. 设从学校乘汽车到火车站的途中有5个十字路口, 每个十字路口遇到红灯的事件是相互独立的, 并且概率都等于0.6. 以 X 表示途中遇到的红灯次数, 求 X 的分布律和分布函数.

解: 参见习题2-2第2题: $X \sim B(5, 0.6)$. 即 $P(X=K) = C_5^K (0.6)^K (0.4)^{5-K}$, $K=0, 1, 2, \dots, 5$

计算得:

X	0	1	2	3	4	5
P	0.0102	0.0768	0.2304	0.3456	0.2592	0.0778

$$F(x) = P(X \leq x) = \sum_{K \leq x} P(X=K) = \begin{cases} 0, & x < 0; \\ 0.0102, & 0 \leq x < 1; \\ 0.0870, & 1 \leq x < 2; \\ 0.3174, & 2 \leq x < 3; \\ 0.6630, & 3 \leq x < 4; \\ 0.9222, & 4 \leq x < 5; \\ 1, & 5 \leq x. \end{cases}$$

2. 设某电话总机5分钟内接到电话呼叫的次数 X 服从泊松分布 $\pi(2)$, 对 $x \leq 6$, 计算 X 的分布函数 $F(x)$.

解: $X = \text{"5分钟内的电话呼叫次数"}$, $X \sim \pi(2)$

有分布律: $P(X=K) = \frac{2^K}{K!} e^{-2} = \frac{2^K}{K!} \times 0.135335$, $K=0, 1, 2, \dots$

X	0	1	2	3	4	5	6	...
P	0.135335	0.27067	0.27067	0.180447	0.090223	0.036089	0.012030	...

$$F(x) = P(X \leq x) = \sum_{K \leq x} P(X=K) = \begin{cases} 0, & x < 0; \\ 0.135335, & 0 \leq x < 1; \\ 0.406005, & 1 \leq x < 2; \\ 0.676675, & 2 \leq x < 3; \\ 0.857122, & 3 \leq x < 4; \\ 0.947345, & 4 \leq x < 5; \\ 0.983434, & 5 \leq x < 6; \\ 0.995464, & 6 \leq x < 7; \\ \dots, & \dots \end{cases}$$

(注): 一般表示分布函数采用下列形式,

$$F(x) = P(X \leq x) = \sum_{K \leq x} P(X=K) = \begin{cases} 0, & x < 0; \\ e^{-2}, & 0 \leq x < 1; \\ 3e^{-2}, & 1 \leq x < 2; \\ 5e^{-2}, & 2 \leq x < 3; \\ \frac{19}{3}e^{-2}, & 3 \leq x < 4; \\ 7e^{-2}, & 4 \leq x < 5; \end{cases}$$

3. 设随机变量 X 具有分布律:

X	0	1	2
P	$\frac{1}{3}$	$\frac{1}{6}$	$\frac{1}{2}$

(1) 求 X 的分布函数 $F(x)$;

(2) 计算 $P(X \leq \frac{3}{2})$, $P(1 < X \leq 4)$ 和 $P(1 \leq X \leq 4)$

解: (1) X 的分布函数:

$$F(x) = P(X \leq x) = \sum_{k \leq x} P(X=k) = \begin{cases} 0, & x < 0; \\ \frac{1}{3}, & 0 \leq x < 1; \\ \frac{1}{2}, & 1 \leq x < 2; \\ 1, & 2 \leq x \end{cases}$$

$$(2) P(X \leq \frac{3}{2}) = P\{(X=0) \cup (X=1)\} = P(X=0) + P(X=1) = \frac{1}{3} + \frac{1}{6} = \frac{1}{2}$$

$$\text{或 } P(X \leq \frac{3}{2}) = F(\frac{3}{2}) = \frac{1}{2}.$$

$$P(1 < X \leq 4) = P\{X=2\} = \frac{1}{2}, \text{ 或 } P(1 < X \leq 4) = F(4) - F(1) = 1 - \frac{1}{2} = \frac{1}{2}.$$

$$P(1 \leq X \leq 4) = P\{(X=1) \cup (X=2)\} = P(X=1) + P(X=2) = \frac{1}{6} + \frac{1}{2} = \frac{2}{3}.$$

$$\text{或 } P(1 \leq X \leq 4) = P(X \leq 4) - P(X < 1) = P(X \leq 4) - P(X \leq 1) + P(X=1) \\ = F(4) - F(1) + (F(1+0) - F(1-0)) = 1 - \frac{1}{2} + (\frac{1}{2} - \frac{1}{3}) = \frac{2}{3}.$$

$$(\text{注}) F(a+0) = F(a), \text{ 而 } P(X=a) = F(a+0) - F(a-0) = F(a) - F(a-0)$$

4. 设随机变量 X 的分布函数为:

$$F(x) = \begin{cases} 0, & x < -1; \\ 0.2, & -1 \leq x < 0; \\ 0.6, & 0 \leq x < 2; \\ 0.9, & 2 \leq x < 4; \\ 1, & 4 \leq x, \end{cases} \quad \text{求 } X \text{ 的分布律}$$

解: 根据 $P\{X=a\} = F(a+0) - F(a-0) = F(a) - F(a-0)$.

$$\text{有 } P(X=-1) = F(-1) - F(-1-0) = 0.2 - 0 = 0.2,$$

$$P(X=0) = F(0) - F(0-0) = 0.6 - 0.2 = 0.4,$$

$$P(X=2) = F(2) - F(2-0) = 0.9 - 0.6 = 0.3,$$

$$P(X=4) = F(4) - F(4-0) = 1 - 0.9 = 0.1.$$

$$\therefore X \text{ 的分布律为 } \begin{array}{c|cccc} X & -1 & 0 & 2 & 4 \\ \hline P_k & 0.2 & 0.4 & 0.3 & 0.1 \end{array}$$

习题 2-5

1. 设连续型随机变量 X 的密度函数为 $f(x) = \begin{cases} Kx^2, & -1 < x < 2 \\ 0, & \text{其他} \end{cases}$

(1) 求常数 K 的值; (2) 求 X 的分布函数; (3) 用两种方法计算 $P(0 < X \leq 1)$.

解: (1) 由 $1 = \int_{-\infty}^{+\infty} f(x) dx = \int_{-1}^2 Kx^2 dx = K \left(\frac{x^3}{3} \right) \Big|_{-1}^2 = 3K$, $\therefore K = \frac{1}{3}$.

(2) X 的分布函数

$$F(x) = \int_{-\infty}^x f(x) dx = \begin{cases} 0, & x \leq -1, \\ \int_{-1}^x \frac{1}{3} x^2 dx, & -1 < x < 2, \\ 1, & 2 \leq x, \end{cases} = \begin{cases} 0, & x \leq -1, \\ \frac{x^3+1}{9}, & -1 < x < 2, \\ 1, & 2 \leq x. \end{cases}$$

$$(3) P(0 < x \leq 1) = F(1) - F(0) = \frac{1^3+1}{9} - \frac{0^3+1}{9} = \frac{1}{9};$$

$$\text{或 } P(0 < x \leq 1) = \int_0^1 \frac{1}{3} x^2 dx = \frac{x^3}{9} \Big|_0^1 = \frac{1}{9}.$$

2. 设随机变量 X 的密度函数为

$$f(x) = \begin{cases} x/2, & 0 < x \leq 1; \\ 1/2, & 1 < x \leq 2; \\ (3-x)/2, & 2 < x < 3; \\ 0, & \text{其他}. \end{cases}$$

(1) 求分布函数 $F(x)$,

(2) 画出密度函数和分布函数的图形. (略).

解:

$$F(x) = \int_{-\infty}^x f(x) dx = \begin{cases} 0, & x \leq 0; \\ \int_0^x \frac{x}{2} dx, & 0 < x \leq 1; \\ \int_0^1 \frac{x}{2} dx + \int_1^x \frac{1}{2} dx, & 1 < x \leq 2; \\ \int_0^1 \frac{x}{2} dx + \int_1^2 \frac{1}{2} dx + \int_2^x \frac{3-x}{2} dx, & 2 < x < 3; \\ 1, & 3 \leq x; \end{cases} = \begin{cases} 0, & x \leq 0; \\ \frac{x^2}{4}, & 0 < x \leq 1; \\ \frac{x-2}{4}, & 1 < x \leq 2; \\ \frac{6x-x^2-5}{4}, & 2 < x < 3; \\ 1, & 3 \leq x. \end{cases}$$

3. 设连续型随机变量 X 的分布函数为,

$$F(x) = \begin{cases} A, & x < 0; \\ Bx^2, & 0 \leq x < 1; \\ Cx - \frac{1}{2}x^2 - 1; & 1 \leq x < 2; \\ 1, & 2 \leq x. \end{cases}$$

(1) 求常数 A, B, C ,

(2) 求 X 的密度函数 $f(x)$,

(3) 用两种方法计算 $P(X > \frac{1}{2})$.

解: (1) 由连续型随机变量 X 的分布函数 $F(x)$ 为连续函数, 故有:

$$F(0-0) = F(0+0) \therefore A = B \times 0^2;$$

$$F(1-0) = F(1+0) \therefore B = C - \frac{3}{2}; \Rightarrow \begin{cases} A = 0; \\ B = \frac{1}{2}; \\ C = 2. \end{cases}$$

$$F(2-0) = F(2+0) \therefore 2C - 3 = 1.$$

$$(2) f(x) = F'(x) = \begin{cases} 0; & x < 0; \\ x; & 0 \leq x < 1; \\ 2-x; & 1 \leq x < 2; \\ 0. & 2 \leq x. \end{cases} = \begin{cases} x; & 0 \leq x < 1; \\ 2-x; & 1 \leq x < 2; \\ 0, & \text{其他}. \end{cases}$$

$$(3) P(X > \frac{1}{2}) = 1 - F(\frac{1}{2}) = 1 - (\frac{1}{2})^2 / 2 = \frac{7}{8}$$

$$\text{或 } P(X > \frac{1}{2}) = \int_{\frac{1}{2}}^{+\infty} f(x) dx = \int_{\frac{1}{2}}^1 x dx + \int_1^2 (2-x) dx = \left(\frac{x^2}{2}\right) \Big|_{\frac{1}{2}}^1 + \left(2x - \frac{x^2}{2}\right) \Big|_1^2 = \frac{7}{8}$$

4. 设采用某种保鲜技术包装的食品, 保鲜时间 X 为一随机变量 (以小时计).

具有概率密度 $f(x) = \begin{cases} \frac{20000}{(x+100)^3} & x > 0 \\ 0 & x \leq 0 \end{cases}$; 任取一包食品求: (1) 保鲜 200

小时以上的概率; (2) 保鲜时间在 80 到 120 小时之间的概率.

$$\text{解: (1) } P(X > 200) = \int_{200}^{+\infty} \frac{20000}{(x+100)^3} dx = -\frac{10000}{(x+100)^2} \Big|_{200}^{+\infty} = \frac{10000}{300^2} = \frac{1}{9}$$

$$(2) P(80 < X < 120) = \int_{80}^{120} \frac{10000}{(x+100)^3} dx = -\frac{10000}{(x+100)^2} \Big|_{80}^{120} = 0.1020$$

5. 设随机变量 X 的概率密度为 $f(x) = \begin{cases} 2x & 0 < x < 1 \\ 0 & \text{其他} \end{cases}$; 以 Y 表示对

X 的三次独立重复观察中事件 $(X \leq \frac{1}{2})$ 出现的次数, 求 Y 的分布律.

解: 设 $X_i = \begin{cases} 1, & \text{第 } i \text{ 次观察事件 } (X \leq \frac{1}{2}) \text{ 出现;} \\ 0, & \text{第 } i \text{ 次观察事件 } (X \leq \frac{1}{2}) \text{ 未出现.} \end{cases} \quad i=1, 2, 3. \text{ 且相互独立.}$

$$P(X_i=1) = P(X \leq \frac{1}{2}) = p = \int_0^{\frac{1}{2}} 2x dx = \frac{1}{4}. \text{ 则 } Y = \sum_{i=1}^3 X_i \sim B(3, \frac{1}{4}).$$

$$\text{即 } P(Y=k) = C_3^k \left(\frac{1}{4}\right)^k \left(\frac{3}{4}\right)^{3-k}, \quad k=0, 1, 2, 3.$$

6. 设某河流每年的最高洪水水位 (m) 具有概率密度: $f(x) = \begin{cases} 2/x^3, & x \geq 1 \\ 0, & \text{其他} \end{cases}$. 今要修造能防御百年一遇洪水 (即遇到的概率不超过 0.01) 的河堤. 问河堤至少要修多高?

解: 设河堤至少要修到百年一遇时的洪水水位 a (m) 高.

又设 $X = \text{"每年的最高洪水水位 (m)"} \cdot X \sim f(x) = \begin{cases} 2/x^3, & x \geq 1 \\ 0 & \text{其他} \end{cases}$

$$\text{则 } P(X \geq a) = 0.01.$$

$$\text{由 } P(X \geq a) = \int_a^{+\infty} \frac{2}{x^3} dx = -\frac{1}{x^2} \Big|_a^{+\infty} = \frac{1}{a^2}, \text{ 得 } \frac{1}{a^2} = 0.01$$

$$\therefore a = 10 \text{ (m)}.$$

习题 2-6

1. 设随机变量 $X \sim U(2, 5)$. 现对 X 进行 3 次独立观测, 求至少有两次观测值大于 3 的概率.

解: 设 $X_i = \begin{cases} 1, & \text{第 } i \text{ 次观测 } X \text{ 值大于 } 3; \\ 0, & \text{第 } i \text{ 次观测 } X \text{ 值未大于 } 3. \end{cases} \quad i=1, 2, 3. \text{ 相互独立.}$

$$\because X \sim U(2, 5). \therefore f_X(x) = \begin{cases} 1/3, & 2 < x < 5; \\ 0, & \text{其他.} \end{cases}$$

$$P(X_i=1) = P(X > 3) = \int_3^5 f_X(x) dx = \int_3^5 \frac{1}{3} dx = \frac{1}{3} x \Big|_3^5 = \frac{2}{3}, \quad i=1, 2, 3.$$

又设 $Y =$ “对 X 的 3 次观测中, 观测值大于 3 的次数”

$$\text{则 } Y = \sum_{i=1}^3 X_i \sim B(3, \frac{2}{3}), \text{ 即 } P(Y=k) = C_3^k (\frac{2}{3})^k (\frac{1}{3})^{3-k}, \quad k=0, 1, 2, 3.$$

$$\begin{aligned} \therefore P(\text{至少有两次观测值大于 } 3) &= P(Y \geq 2) = P(Y=2) + P(Y=3) \\ &= C_3^2 (\frac{2}{3})^2 (\frac{1}{3}) + C_3^3 (\frac{2}{3})^3 = \frac{4}{9} + \frac{8}{27} = \frac{20}{27} = 0.7407. \end{aligned}$$

2. 设随机变量 $K \sim U(0, 5)$. 求方程 $4x^2 + 4Kx + K + 2 = 0$ 有实根的概率.

解: 方程 $4x^2 + 4Kx + K + 2 = 0$ 有实根的充要条件是它的判别式 $\Delta \geq 0$.

$$\text{由 } \Delta = (4K)^2 - 4 \cdot 4 \cdot (K+2) \geq 0, \text{ 即 } K^2 - K - 2 \geq 0 \text{ 得 } (K \geq 2) \cup (K \leq -1)$$

$$\therefore P(\text{方程有实根}) = P\{(K \geq 2) \cup (K \leq -1)\} = P(K \geq 2) + P(K \leq -1) = P(K \geq 2)$$

$$= \int_2^5 f(x) dx. \text{ 而 } X \sim f(x) = \begin{cases} \frac{1}{5}, & 0 < x < 5; \\ 0, & \text{其他.} \end{cases}$$

$$\therefore P(\text{方程有实根}) = \int_2^5 \frac{1}{5} dx = (\frac{x}{5}) \Big|_2^5 = \frac{3}{5} = 0.6.$$

3. 设某书店收银台顾客排队等待服务的时间 X (以分计) 服从指数分布, 密度函数为 $f(x) = \begin{cases} \frac{1}{5} e^{-\frac{x}{5}}, & x > 0; \\ 0, & x \leq 0. \end{cases}$ 分别利用 X 的密度函数和分布函数计算 $P(X > 10)$.

解: X 服从参数为 $\lambda = \frac{1}{5}$ 的指数分布, 分布函数 $F(x) = \begin{cases} 1 - e^{-\frac{x}{5}}, & x > 0, \\ 0, & x \leq 0. \end{cases}$

$$\therefore P(X > 10) = 1 - F(10) = 1 - (1 - e^{-\frac{10}{5}}) = e^{-2};$$

$$\text{或 } P(X > 10) = \int_{10}^{+\infty} \frac{1}{5} e^{-\frac{x}{5}} dx = (-e^{-\frac{x}{5}}) \Big|_{10}^{+\infty} = e^{-\frac{10}{5}} = e^{-2}.$$

4. 设随机变量 X 的密度函数为

$$f(x) = \begin{cases} K e^{-3(x-1)}, & x > 1; \\ 0, & x \leq 1. \end{cases} \quad (1) \text{ 确定常数 } K; \quad (2) \text{ 计算 } P(1.5 \leq X \leq 2).$$

$$\text{解: (1) 由 } 1 = \int_{-\infty}^{+\infty} f(x) dx = \int_1^{+\infty} K e^{-3(x-1)} dx = \frac{1}{3} K (-e^{-3(x-1)}) \Big|_1^{+\infty} = \frac{K}{3}.$$

$$\therefore K = 3. \therefore f(x) = \begin{cases} 3 e^{-3(x-1)}, & x > 1; \\ 0, & x \leq 1; \end{cases}$$

$$(2) \quad X \text{ 的分布函数 } F(x) = \begin{cases} 1 - e^{-3(x-1)} & x > 1 \\ 0 & x \leq 1 \end{cases}$$

$$\therefore P(1.5 \leq X \leq 2) = F(2) - F(1.5) = (1 - e^{-3}) - (1 - e^{-1.5}) = e^{-1.5} - e^{-3}$$

$$\text{或 } P(1.5 \leq X \leq 2) = \int_{1.5}^2 3e^{-3(x-1)} dx = (-e^{-3(x-1)}) \Big|_{1.5}^2 = e^{-1.5} - e^{-3}$$

5. 设某种仪器装了3只独立工作的同型号元件, 其寿命 X (小时) 服从密度函数为 $f(x) = \begin{cases} \frac{1}{600} e^{-\frac{x}{600}} & x > 0 \\ 0 & x \leq 0 \end{cases}$ 的指数分布, 求仪器在最初200小时内至少有1只元件出故障的概率.

解: 设 $X_i = \begin{cases} 1 & \text{第 } i \text{ 只元件寿命不到 } 200 \text{ 小时,} \\ 0 & \text{第 } i \text{ 只元件寿命不少于 } 200 \text{ 小时.} \end{cases} \quad i=1, 2, 3. \text{ 相互独立.}$

$$P(X_i=1) = P(X < 200) = \int_0^{200} \frac{1}{600} e^{-\frac{x}{600}} dx = (-e^{-\frac{x}{600}}) \Big|_0^{200} = 1 - e^{-\frac{1}{3}}$$

又 $Y =$ “仪器的3只元件中, 寿命不到200小时的只数”

$$\therefore Y = \sum_{i=1}^3 X_i \sim B(3, 1 - e^{-\frac{1}{3}}), \quad P(Y=k) = C_3^k (1 - e^{-\frac{1}{3}})^k (e^{-\frac{1}{3}})^{3-k}, \quad k=0, \dots, 3.$$

$P(\text{仪器在最初200小时内至少有1只元件出故障})$

$$= P(Y \geq 1) = 1 - P(Y < 1) = 1 - P(Y=0) = 1 - C_3^0 (e^{-\frac{1}{3}})^3 = 1 - e^{-1}$$

习题 2-7

1. 设 $X \sim N(0, 1)$. 求 (1) $P(0.02 < X < 2.33)$; (2) $P(-1.85 < X < 0.04)$.

解: (1) $P(0.02 < X < 2.33) = \Phi(2.33) - \Phi(0.02) = 0.9901 - 0.5080 = 0.4821$.

$$(2) P(-1.85 < X < 0.04) = \Phi(0.04) - \Phi(-1.85) = \Phi(0.04) - (1 - \Phi(1.85)) \\ = \Phi(0.04) + \Phi(1.85) - 1 = 0.5160 + 0.9678 - 1 = 0.4798.$$

2. 设 $X \sim N(10, 3^2)$. (1) 求 $P(7 < X < 16)$; (2) 求常数 α , 使 $P(X < \alpha) = 0.9$; (3) 求常数 α , 使 $P(|X - \alpha| > \alpha) = 0.01$.

解: (1) $P(7 < X < 16) = \Phi(\frac{16-10}{3}) - \Phi(\frac{7-10}{3}) = \Phi(2) - \Phi(-1) = \Phi(2) - (1 - \Phi(1)) \\ = \Phi(2) + \Phi(1) - 1 = 0.9772 + 0.8413 - 1 = 0.8185.$

$$(2) P(X < \alpha) = \Phi(\frac{\alpha-10}{3}) = 0.9, \therefore \frac{\alpha-10}{3} = 1.285, \therefore \alpha = 13.855.$$

$$(3) P(|X - \alpha| > \alpha) = P[(X - \alpha) > \alpha] \cup (X - \alpha) < -\alpha] = P[(X > 2\alpha) \cup (X < 0)] \\ = P(X > 2\alpha) + P(X < 0) = 1 - \Phi(\frac{2\alpha-10}{3}) + \Phi(-\frac{10}{3}) \\ = 1 - \Phi(\frac{2\alpha-10}{3}) + 1 - \Phi(\frac{10}{3}) = 1 - \Phi(\frac{2\alpha-10}{3}) + 0.0004 \\ \therefore \frac{2\alpha-10}{3} = 2.34, \therefore \alpha = 8.51 \quad (\text{注: } \Phi(\frac{10}{3}) = \Phi(3.33) = 0.9996)$$

3. 某机器生产的螺栓长度 (cm) 服从参数 $\mu = 10.05$, $\sigma = 0.06$ 的正态分布. 规定长度在范围 10.05 ± 0.12 内为合格品. 求该机器生产的螺栓的合格率.

解: $P(10.05 - 0.12 < X < 10.05 + 0.12) = P(9.93 < X < 10.17) = \Phi(\frac{10.17 - 10.05}{0.06}) \\ - \Phi(\frac{9.93 - 10.05}{0.06}) = \Phi(2) - \Phi(-2) = 2\Phi(2) - 1 = 2 \times 0.9772 - 1 = 0.9544.$

4. 设一台软饮料包装机所装每罐饮料净含量 X 为一随机变量, 服从 $\mu = 200$, $\sigma = 15$ (毫升) 的正态分布. 求该包装机生产的饮料中 (1) 净含量超过 224 毫升的比例; (2) 净含量在 191 到 209 毫升之间的概率; (3) 求使 $P\{X \leq \alpha\} \leq 0.25$ 成立的最大数 α .

解: (1) $P(X > 224) = 1 - \Phi(\frac{224-200}{15}) = 1 - \Phi(1.6) = 1 - 0.9452 = 0.0548$;

$$(2) P(191 < X < 209) = \Phi(\frac{209-200}{15}) - \Phi(\frac{191-200}{15}) = \Phi(0.6) - \Phi(-0.6) \\ = 2\Phi(0.6) - 1 = 2 \times 0.7257 - 1 = 0.4514.$$

$$(3) P(X \leq \alpha) = \Phi(\frac{\alpha-200}{15}) = 1 - \Phi(\frac{200-\alpha}{15}) \leq 0.25.$$

$$\therefore \Phi(\frac{200-\alpha}{15}) \geq 0.75, \text{ 当取等号时 } \alpha \text{ 取最大值.}$$

$$\text{由 } \frac{200-\alpha}{15} = 0.675, \therefore \alpha = 189.875$$

习题 2-8

1. 设离散型随机变量 X 具有分布律

X	-2	-1	0	1	2	3
P_k	1/16	2/16	4/16	5/16	3/16	1/16

(1) 求 $Y = 6 - X^2$ 的分布律; (2) 求 $Z = \max(X+2, X^2)$ 的分布律.

解: (1)

$6 - X^2$	$6 - (-2)^2$	$6 - (-1)^2$	$6 - 0^2$	$6 - 1^2$	$6 - 2^2$	$6 - 3^2$
P_k	1/16	2/16	4/16	5/16	3/16	1/16

\Rightarrow

$6 - X^2$	2	5	6	5	2	-3
P_k	1/16	2/16	4/16	5/16	3/16	1/16

\Rightarrow

Y	-3	2	5	6
P_k	1/16	4/16	7/16	4/16

(2)

$\max(X+2, X^2)$	$\max((-2)+2, (-2)^2)$	$\max((-1)+2, (-1)^2)$	$\max(0+2, 0^2)$	$\max(1+2, 1^2)$	$\max(2+2, 2^2)$	$\max(3+2, 3^2)$
P_k	1/16	2/16	4/16	5/16	3/16	1/16

\Rightarrow

$\max(X+2, X^2)$	4	1	2	3	4	9
P_k	1/16	2/16	4/16	5/16	3/16	1/16

\Rightarrow

Z	1	2	3	4	9
P_k	2/16	4/16	5/16	4/16	1/16

2. 袋中有 K 号球 K 个, $K=1, 2, \dots, n$. 随机从中取一个球. 设 X 为所取到的球的号码. 求 $Y = X/(1+X)$ 的分布律.

解: \because 袋中共有球 $\frac{n(n+1)}{2}$ 个. $\therefore P(X=K) = \frac{K}{n(n+1)/2} = \frac{2K}{n(n+1)}, K=1, 2, \dots, n$.

\therefore 当取出的球为 K 号时, $Y = K/(1+K), K=1, 2, \dots, n$.

故 $Y = \frac{X}{1+X}$ 的分布律为

$$P(Y = \frac{K}{1+K}) = P(\frac{X}{1+X} = \frac{K}{1+K}) = P(X=K) = \frac{2K}{n(n+1)}, K=1, 2, \dots, n$$

3. 设随机变量 $X \sim U(0, 1)$. (1) 求随机变量 $Y = aX+b$ 的密度函数; (2) 求随机变量 $Z = \frac{X}{1+X}$ 的密度函数.

解: $X \sim U(0, 1), f_X(x) = \begin{cases} 1, & 0 < x < 1; \\ 0, & \text{其他} \end{cases}$

(1) 当 $a > 0$ 时.

$$F_Y(y) = P(Y \leq y) = P(aX+b \leq y) = P(X \leq \frac{1}{a}(y-b)) = F_X(\frac{1}{a}(y-b))$$

$$\therefore f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} F_X(\frac{1}{a}(y-b)) = f_X(\frac{1}{a}(y-b)) \cdot \frac{1}{a}$$

$$= \begin{cases} \frac{1}{a}, & b < y < a+b; \\ 0, & \text{其他} \end{cases}$$

当 $a < 0$ 时.

$$F_Y(y) = P(Y \leq y) = P(aX+b \leq y) = P(X \geq \frac{1}{a}(y-b)) = 1 - F_X(\frac{1}{a}(y-b))$$

$$\therefore f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} (1 - F_X(\frac{1}{a}(y-b))) = -\frac{1}{a} f_X(\frac{1}{a}(y-b))$$

$$= \begin{cases} -\frac{1}{a}, & a+b < y < b; \\ 0, & \text{其他.} \end{cases}$$

$$(2) F_Z(u) = P(Z \leq u) = P\left(\frac{X}{1+X} \leq u\right) = P\left(X \leq \frac{u}{1-u}\right) = F_X\left(\frac{u}{1-u}\right)$$

$$f_Z(u) = \frac{d}{du} F_Z(u) = \frac{d}{du} F_X\left(\frac{u}{1-u}\right) = f_X\left(\frac{u}{1-u}\right) \cdot \frac{1}{(1-u)^2} = \begin{cases} \frac{1}{(1-u)^2}, & 0 < u < \frac{1}{2}; \\ 0, & \text{其他.} \end{cases}$$

4. 设随机变量 $X \sim N(\mu, \sigma^2)$, 求 $Y = \frac{X-\mu}{\sigma}$ 的密度函数.

$$\text{解: } X \sim f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, \quad -\infty < x < +\infty.$$

$$F_Y(y) = P(Y \leq y) = P\left(\frac{X-\mu}{\sigma} \leq y\right) = P(X \leq \sigma y + \mu) = F_X(\sigma y + \mu).$$

$$\therefore f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} F_X(\sigma y + \mu) = f_X(\sigma y + \mu) \cdot \sigma = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{[(\sigma y + \mu) - \mu]^2}{2\sigma^2}} \cdot \sigma$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}, \quad -\infty < y < +\infty.$$

$$\text{即 } Y = \frac{X-\mu}{\sigma} \sim N(0, 1).$$

5. 设随机变量 X 具有密度函数 $f_X(x) = \begin{cases} \frac{3}{2}x^2, & -1 < x < 1; \\ 0, & \text{其他.} \end{cases}$

(1) 求随机变量 $Y = |X|$ 的密度函数; (2) 求随机变量 $Y = X^2$ 的密度函数.

解: (1) 当 $y \leq 0$ 时, $F_Y(y) = P(Y \leq y) = P(|X| \leq y) = 0, \therefore f_Y(y) = 0;$

当 $y \geq 0$ 时, $F_Y(y) = P(Y \leq y) = P(|X| \leq y) = P(-y \leq X \leq y) = F_X(y) - F_X(-y)$

$$\therefore f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} [F_X(y) - F_X(-y)] = f_X(y) + f_X(-y)$$

$$= \begin{cases} \frac{3}{2}y^2 + \frac{3}{2}(-y)^2, & 0 \leq y < 1; \\ 0, & y \geq 1. \end{cases} = \begin{cases} 3y^2, & 0 \leq y < 1; \\ 0, & y \geq 1. \end{cases}$$

$$\text{故有 } f_Y(y) = \begin{cases} 3y^2, & 0 \leq y < 1; \\ 0, & \text{其他.} \end{cases}$$

(2) 当 $u < 0$ 时, $F_Z(u) = P(Z \leq u) = P(X^2 \leq u) = 0, \therefore f_Z(u) = 0;$

当 $u \geq 0$ 时, $F_Z(u) = P(Z \leq u) = P(X^2 \leq u) = P(-\sqrt{u} \leq X \leq \sqrt{u}) = F_X(\sqrt{u}) - F_X(-\sqrt{u})$

$$\therefore f_Z(u) = \frac{d}{du} F_Z(u) = \frac{d}{du} [F_X(\sqrt{u}) - F_X(-\sqrt{u})] = f_X(\sqrt{u}) \cdot \frac{1}{2\sqrt{u}} + f_X(-\sqrt{u}) \cdot \frac{1}{2\sqrt{u}}$$

$$= \begin{cases} \frac{1}{2\sqrt{u}} \left[\frac{3}{2}(\sqrt{u})^2 + \frac{3}{2}(-\sqrt{u})^2 \right], & 0 \leq u < 1; \\ 0, & u \geq 1. \end{cases} = \begin{cases} \frac{3}{2}\sqrt{u}, & 0 \leq u < 1; \\ 0, & u \geq 1. \end{cases}$$

$$\text{故有 } f_Z(u) = \begin{cases} \frac{3}{2}\sqrt{u}, & 0 \leq u < 1; \\ 0, & \text{其他.} \end{cases}$$

6. 设随机变量 $X \sim U(0, 2)$, 求随机变量 $Y = 2 - (X-1)^2$ 的密度函数.

解: $X \sim U(0, 2), \therefore f_X(x) = \begin{cases} 1/2, & 0 < x < 2; \\ 0, & \text{其他.} \end{cases}$

$$F_Y(y) = P(Y \leq y) = P(2 - (X-1)^2 \leq y) = P((X-1)^2 \geq 2-y)$$

$$\text{当 } y \geq 2 \text{ 时, } P((X-1)^2 \geq 2-y) = P((X-1)^2 \geq 0) = P(\Omega) = 1.$$

$$\therefore f_Y(y) = 0.$$

$$\begin{aligned} \text{当 } y < 2 \text{ 时, } F_Y(y) &= P((X-1)^2 \geq 2-y) = P[(X-1) \geq \sqrt{2-y} \cup (X-1) \leq -\sqrt{2-y}] \\ &= P(X-1 \geq \sqrt{2-y}) + P(X-1 \leq -\sqrt{2-y}) = P(X \geq 1+\sqrt{2-y}) + P(X \leq 1-\sqrt{2-y}) \\ &= 1 - F_X(1+\sqrt{2-y}) + F_X(1-\sqrt{2-y}) \end{aligned}$$

$$\therefore f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} [1 - F_X(1+\sqrt{2-y}) + F_X(1-\sqrt{2-y})]$$

$$= -f_X(1+\sqrt{2-y}) \frac{d}{dy} (1+\sqrt{2-y}) + f_X(1-\sqrt{2-y}) \frac{d}{dy} (1-\sqrt{2-y})$$

$$= \frac{1}{2\sqrt{2-y}} [f_X(1+\sqrt{2-y}) + f_X(1-\sqrt{2-y})]$$

$$= \begin{cases} \frac{1}{2\sqrt{2-y}} \left[\frac{1}{2} + \frac{1}{2} \right], & 1 < y < 2 \\ 0, & y \leq 1 \end{cases} = \begin{cases} \frac{1}{2\sqrt{2-y}}, & 1 < y < 2 \\ 0, & y \leq 1. \end{cases}$$

$$\text{故 } f_Y(y) = \begin{cases} \frac{1}{2\sqrt{2-y}}, & 1 < y < 2, \\ 0, & \text{其他.} \end{cases}$$

复习题 = (三, 解答题).

1. 3个不同的球, 随机投入编号为1, 2, 3, 4的盒中, X 表示有球盒的最小编号, 求 X 的分布律.

解: 设有球盒的最小编号为 K , 此时3个球可投入的盒子有 $(4-K)+1$ 只, 但为保证第 K 号必有球投入, 因此3个球不能同时投入 K 号盒后的 $(4-K)$ 只盒中, 所以事件“最小编号为 K ”的样本点数为 $[(4-K)+1]^3 - (4-K)^3$. 故

$$P(X=K) = \frac{[(4-K)+1]^3 - (4-K)^3}{4^3}, \quad K=1, 2, 3, 4.$$

2. 将一颗骰子抛掷两次, 以 X 表示两次中得到的小的点数, 求 X 的分布律.

解: 设两次中得到的小的点数为 K , 此时比 K 大的点数有 $(6-K)$ 个, 第一次出现 K 点共有 $(6-K)+1$ 个样本点; 第二次出现 K 点, 除第一次出现 K 点情形有 $(6-K)$ 个样本点, 所以“两次投掷最小编号为 K ”的事件共有样本点为 $[(6-K)+1] + (6-K) = 2(6-K)+1$ 个. 故

$$P(X=K) = \frac{2(6-K)+1}{6^2}, \quad K=1, 2, 3, 4, 5, 6.$$

3. 自动生产线在调整以后出现的废品率为 p , 生产过程中出现废品时, 立即重新进行调整, 求两次调整之间生产的合格品数的分布律.

解: 设 X = “两次调整之间生产的合格品数”. 注意到调整是在出现废品时立即进行, 所以 X 可能取的值应为 $0, 1, 2, \dots$, $(X=0)$ 事件是出现废品立即调整的意思, 而不是取值为 $1, 2, \dots$. 故 X 的分布律为:

$$P(X=K) = p(1-p)^K, \quad K=0, 1, 2, \dots$$

4. 5只电池, 其中2只是次品, 每次取一只测试, 直到找出2只次品或3只正品为止, 写出所需测试次数的分布律.

解: 设 X = “直到找出2只次品或3只正品为止的测试次数”. 显然 X 可能取 $2, 3, 4$. 注意最多测试4次, 这4次中如果第2只次品在第4次找出或第3只正品在第4次找出时, 测试必结束. 故 X 的分布律为

$$\begin{array}{c|ccc} X & 2 & 3 & 4 \\ \hline P & A_2^2/A_5^2 & (C_2^1 A_3^1 A_2^2 + A_3^3)/A_5^3 & (C_3^1 A_3^2 A_2^2 + C_3^2 A_3^3 A_2^1)/A_5^4 \end{array}$$

$$\text{即 } \begin{array}{c|ccc} X & 2 & 3 & 4 \\ \hline P & 1/10 & 3/10 & 6/10 \end{array}$$

(注) 可具体写出测试的所有可能结果来求: 记 $X_i = \begin{cases} 1, & \text{第} i \text{次测试得正品;} \\ 0, & \text{第} i \text{次测试得次品;} \end{cases}$
 $i=1, 2, 3, 4, 5$. 所有结果为:

$(0, 0, 1, 1, 1), (0, 1, 0, 1, 1), (1, 0, 0, 1, 1), (1, 1, 1, 0, 0), (0, 1, 1, 0, 1)$
 $(1, 0, 1, 0, 1), (1, 1, 0, 0, 1), (1, 1, 0, 1, 0), (1, 0, 1, 1, 0), (0, 1, 1, 1, 0)$

共10个. 由此易得 X 的分布律, 比用排列组合计算便当.

5. 进行重复独立试验, 设每次试验成功的概率为 p , 失败的概率为 $q=1-p$.

($0 < p < 1$), (1) 将试验进行到出现一次成功为止, 以 X 表示所需的试验次数, 求 X 的分布律 (此时称 X 服从参数为 p 的几何分布); (2) 将试验进行到出现 r 次成功为止, 以 Y 表示所需的试验次数, 求 Y 的分布律 (此时称 Y 服从参数为 p 的巴斯卡分布); (3) 一篮球运动员的投篮命中率为 45%, 以 X 表示他首次投入时累计已投篮的次数, 写出 X 的分布律, 并计算 X 取偶数的概率.

解: (1) 设试验进行到第 K 次才出现一次成功, 则前 $K-1$ 次必都是不成功, 故

$$P(X=K) = q^{K-1} \cdot p \quad (\text{或 } (1-p)^{K-1} \cdot p), \quad K=1, 2, 3, \dots$$

(2) 设试验进行到第 K 次时出现第 r 次成功, 则前 $K-1$ 次试验中应出现 $r-1$ 次成功, 故

$$P(Y=K) = C_{K-1}^{r-1} p^{r-1} (1-p)^{(K-1)-(r-1)} \cdot p = C_{K-1}^{r-1} p^r (1-p)^{K-r}, \quad K=r, r+1, \dots$$

(3) X = “直到首次投篮命中时, 累计的投篮次数”, 服从参数 $p=0.45$ 的几何分布: $P(X=K) = (0.55)^{K-1} \cdot 0.45, \quad K=1, 2, 3, \dots$

$$\begin{aligned} \therefore P(X \text{ 取偶数}) &= \sum_{K=1}^{\infty} P(X=2K) = \sum_{K=1}^{\infty} (0.55)^{2K-1} \cdot 0.45 = 0.45 \cdot \sum_{K=1}^{\infty} (0.55)^{2K-1} \\ &= \frac{0.45}{100} \cdot \frac{\frac{55}{100}}{1 - (\frac{55}{100})^2} = \frac{11}{31}. \quad (\text{注: 利用几何数列求和公式}) \end{aligned}$$

6. 试卷中共有 10 道选择题, 其中前四题每题 3 分, 后六题每题 5 分. 每道选择题都有 4 个答案, 其中只有一个答案是正确的. 如果每题都是随机选一个答案, 问至少得 10 分的概率有多大?

解: 记 $X_i = \begin{cases} 1, & \text{第 } i \text{ 题答案正确;} \\ 0, & \text{第 } i \text{ 题答案不正确.} \end{cases}$ 且 $\begin{array}{c|cc} X_i & 0 & 1 \\ \hline P & 3/4 & 1/4 \end{array} \quad i=1, 2, \dots, 10, \text{ 相互独立.}$

设 X = “前四题中答案正确的题数” $= \sum_{i=1}^4 X_i \sim B(4, \frac{1}{4})$;

Y = “后六题中答案正确的题数” $= \sum_{i=5}^{10} X_i \sim B(6, \frac{1}{4})$.

(得 0 分)

10 个题目总得分不满 10 分的情况为: ① 10 题一题都没答正确; ② 仅答对前四题中的一题 (得 3 分); ③ 仅答对后六题中的一题 (得 5 分); ④ 仅答对前四题中的两题 (得 6 分); ⑤ 答对前四题中一题和后六题中一题 (得 8 分); ⑥ 答对前四题中的三题 (得 9 分); 对应的概率为:

$$P(\text{得 0 分}) = P(X=0)P(Y=0) = C_4^0 C_6^0 (\frac{3}{4})^{10} = 0.0563;$$

$$P(\text{得 3 分}) = P(X=1)P(Y=0) = C_4^1 C_6^0 (\frac{1}{4})(\frac{3}{4})^9 = 0.0751;$$

$$P(\text{得 5 分}) = P(X=0)P(Y=1) = C_4^0 C_6^1 (\frac{1}{4})(\frac{3}{4})^9 = 0.1126;$$

$$P(\text{得 6 分}) = P(X=2)P(Y=0) = C_4^2 C_6^0 (\frac{1}{4})^2 (\frac{3}{4})^8 = 0.0375;$$

$$P(\text{得 8 分}) = P(X=1)P(Y=1) = C_4^1 C_6^1 (\frac{1}{4})^2 (\frac{3}{4})^8 = 0.1502;$$

$$P(\text{得 9 分}) = P(X=3)P(Y=0) = C_4^3 C_6^0 (\frac{1}{4})^3 (\frac{3}{4})^7 = 0.0083.$$

$$\therefore P(\text{不满 10 分}) = 0.44$$

7. 设一厂家生产的每台仪器以概率 0.7 可以直接出厂, 以概率 0.3 需进一步调试, 经过调试后以概率 0.8 可以出厂, 以概率 0.2 定为不合格品不能出厂. 现该厂生产了 n 台仪器 ($n \geq 2$, 生产过程独立), 求 (1) 全部能出厂的概率, (2) 至少有 2 件不能出厂的概率.

解: 设 $X_i = \begin{cases} 1; & \text{第 } i \text{ 台仪器不能出厂;} \\ 0. & \text{第 } i \text{ 台仪器可以出厂;} \end{cases}$ $\frac{X_i}{P} \begin{matrix} 0 & 1 \\ 0.94 & 0.06 \end{matrix} \quad i=1, \dots, n$ 相互独立.

又 $Y = \text{"}n\text{ 台仪器中不能出厂的台数"} = \sum_{i=1}^n X_i \sim B(n, 0.06)$

$$(1) P(\text{全部能出厂}) = P(Y=0) = C_n^0 (0.06)^0 (0.94)^n = (0.94)^n$$

$$(2) P(\text{至少有 2 台不能出厂}) = P(Y \geq 2) = 1 - P(Y < 2) = 1 - P(Y=0) - P(Y=1) \\ = 1 - C_n^0 (0.94)^n - C_n^1 0.06 \cdot (0.94)^{n-1} = 1 - (0.94)^n - n \cdot 0.06 (0.94)^{n-1}$$

(注) 可以出厂的产品的概率为 $0.7 + 0.3 \times 0.8 = 0.94$.

8. 已知每天到某炼油厂的油船数 $X \sim \pi(2)$, 而港口的设备一天只能为三艘油船服务. 如果一天中到达的油船数超过三艘, 超出的油船必须转向另一港口. 求: (1) 这一天中必须有油船转走的概率, (2) 设备增加到多少才能使每天到达港口的油船有 90% 可以得到服务? (3) 每天到达港口的油船最可能有几艘?

解: $X \sim \pi(2)$. $P(X=k) = \frac{2^k}{k!} e^{-2} (= \frac{2^k}{k!} \times 0.13534)$, $k=0, 1, 2, \dots$

$$(1) P(\text{一天中必须有油船转走}) = P(X > 3) = 1 - P(X \leq 3) = 1 - P(X=0) - P(X=1) \\ - P(X=2) - P(X=3) = 1 - e^{-2} - 2e^{-2} - 2e^{-2} - \frac{4}{3}e^{-2} = 1 - \frac{19}{3}e^{-2} = 0.14285$$

$$(2) \text{设增加到一天能接 } k \text{ 艘服务, 使 } P(X \leq k) \geq 0.9.$$

$$\text{由 } P(X=0) + P(X=1) + \dots + P(X=4) = 7e^{-2} = 0.94738, \text{ 故取 } k=4.$$

$$(3) \text{因 } P(X=1) = P(X=2) = 2e^{-2}, \text{ 都大于 } P(X=0) = e^{-2}, P(X=3) = \frac{4}{3}e^{-2} \dots$$

所以在每天到达港口的油船数中, 最可能为一艘或两艘.

9. 假设某地在任何长为 t (周) 的时间内发生地震的次数 $N(t)$ 服从参数为 λt 的泊松分布. (1) 求相邻两周内至少发生 3 次地震的概率; (2) 求在连续 8 周无地震的情形下, 在未来 8 周中仍无地震的概率.

解: $N(t) \sim \pi(\lambda t)$. $P\{N(t)=k\} = (\lambda t)^k e^{-\lambda t} / k!$, $k=0, 1, 2, \dots$

$$(1) P(N(t) \geq 3) = 1 - P(N(t) < 3) = 1 - P(N(t)=0) - P(N(t)=1) - P(N(t)=2) \\ = 1 - [e^{-2\lambda} + 2\lambda e^{-2\lambda} + 2\lambda^2 e^{-2\lambda}] = 1 - (1 + 2\lambda + 2\lambda^2) e^{-2\lambda}$$

(2) 设 $X = \text{"连续 8 周的时间内发生地震的次数"};$

$Y = \text{"未来 8 周的时间内发生地震的次数"}.$

$$P(\text{连续 8 周无地震的情形下, 在未来 8 周中仍无地震}) = P(Y=0 | X=0)$$

$$= \frac{P(X=0 \cap Y=0)}{P(X=0)} = \frac{e^{-16\lambda}}{e^{-8\lambda}}$$

10. 设一大型设备在任何长度 t 的时间内发生故障的次数 $N(t)$ 服从参数为 λt 的泊松分布. (1) 求相继两次故障时间间隔 T 的分布函数; (2) 已知设备已无故障工作了 6 小时, 求再无故障工作 6 小时的概率.

解: $N(t) \sim \pi(\lambda t)$. $P(X=K) = (\lambda t)^K e^{-\lambda t} / K!$, $K=0, 1, 2, \dots$

(1) 设两次故障之间相隔的时间为 t . 则

$$P(\text{相继两次故障时间间隔 } T \text{ 不大于 } t) = P(T \leq t)$$

$$= P(\text{在时间长度 } t \text{ 内至少发生一次故障}) = P(N(t) \geq 1)$$

$$\text{即 } P(T \leq t) = P(N(t) \geq 1) = 1 - P(N(t) < 1) = 1 - P(N(t) = 0) = 1 - e^{-\lambda t}, (t > 0)$$

$$\text{而 } t < 0 \text{ 时, } P(T \leq t) = P(\emptyset) = 0.$$

$$\therefore T \text{ 的分布函数为: } F_T(t) = \begin{cases} 1 - e^{-\lambda t}, & t > 0; \\ 0, & t \leq 0. \end{cases} \text{ 即参数为 } \lambda \text{ 的指数分布.}$$

(2) 设 X = “设备在工作 6 小时内发生的故障数”;

Y = “设备在继续工作 6 小时内发生的故障数”.

$$P(Y=0/X=0) = P[(X=0) \cap (Y=0)] / P(X=0) = \frac{P(N(12)=0)}{P(N(6)=0)} = \frac{e^{-12\lambda}}{e^{-6\lambda}} = e^{-6\lambda}$$

11. 设某种昆虫产卵数 X 服从参数为 λ 的泊松分布 ($\lambda > 0$). 每个卵能够孵化成幼虫的概率为 p ($0 < p < 1$), 且彼此独立. 求该种昆虫有 r 个后代的概率.

解: $X \sim \pi(\lambda)$. 若该种昆虫产 i 个卵, 又设该种昆虫的后代数 Y , 则在产 i 个卵的条件下, 孵化出 r 个后代的概率为条件概率:

$$P(Y=r/X=i) = C_i^r p^r (1-p)^{i-r}, \quad i=0, 1, 2, \dots, r \leq i.$$

$$\begin{aligned} \therefore P(\text{昆虫有 } r \text{ 个后代}) &= P(Y=r) = \sum_{i=r}^{\infty} P(X=i) P(Y=r/X=i) \\ &= \sum_{i=r}^{\infty} \frac{\lambda^i}{i!} e^{-\lambda} \cdot C_i^r p^r (1-p)^{i-r} \\ &= \frac{(\lambda p)^r}{r!} \sum_{k=0}^{\infty} \frac{1}{k!} [\lambda(1-p)]^k \cdot e^{-\lambda} = \frac{(\lambda p)^r}{r!} \cdot e^{\lambda(1-p)} \cdot e^{-\lambda} \\ &= \frac{(\lambda p)^r}{r!} e^{-\lambda p}, \quad r=0, 1, 2, \dots \end{aligned}$$

$$\text{即 } Y \sim \pi(\lambda p).$$

12. 设随机变量 X 的密度函数为 $f(x) = \begin{cases} ax+b; & 0 < x < 1, \\ 0, & \text{其他.} \end{cases}$ 又已知:

$$P(X < \frac{1}{3}) = P(X > \frac{1}{3}), \text{ 试求常数 } a \text{ 和 } b.$$

$$\text{解: 由 } 1 = \int_{-\infty}^{+\infty} f(x) dx = \int_{\frac{1}{3}}^1 (ax+b) dx = \left(\frac{ax^2}{2} + bx \right) \Big|_{\frac{1}{3}}^1 = \frac{a}{2} + b;$$

$$\text{又 } P(X < \frac{1}{3}) = \int_{-\infty}^{\frac{1}{3}} f(x) dx = \int_0^{\frac{1}{3}} (ax+b) dx = \left(\frac{ax^2}{2} + bx \right) \Big|_0^{\frac{1}{3}} = \frac{a}{18} + \frac{b}{3};$$

$$P(X > \frac{1}{3}) = \int_{\frac{1}{3}}^{+\infty} f(x) dx = \int_{\frac{1}{3}}^1 (ax+b) dx = \left(\frac{ax^2}{2} + bx \right) \Big|_{\frac{1}{3}}^1 = \frac{8a}{18} + \frac{2b}{3}.$$

$$\therefore \text{有 } \begin{cases} \frac{a}{2} + b = 1 \\ \frac{a}{18} + \frac{b}{3} = \frac{8a}{18} + \frac{2b}{3} \end{cases}; \text{ 即 } \begin{cases} \frac{a}{2} + b = 1 \\ \frac{7a}{18} + \frac{b}{3} = 0 \end{cases}; \text{ 解得 } \begin{cases} a = -\frac{3}{2} = -1.5; \\ b = \frac{7}{3} \end{cases}$$

13. 设数据 X 和 Y 的取值是相互独立的. 它们都服从区间 $[1, 3]$ 上的均匀分布. 设事件 $A = "X < a"$, 事件 $B = "Y > a"$. 如果已知 $P(A \cup B) = \frac{7}{9}$. 求常数 a .

解. $X \sim U[1, 3]$, $Y \sim U[1, 3]$ 相互独立.

$$f_X(x) = \begin{cases} \frac{1}{2}, & 1 < x < 3; \\ 0, & \text{其他} \end{cases}, \quad f_Y(y) = \begin{cases} \frac{1}{2}, & 1 < y < 3; \\ 0, & \text{其他} \end{cases}$$

$$P(A) = P\{X < a\} = \int_{-\infty}^a f_X(x) dx = \int_1^a \frac{1}{2} dx = \left(\frac{x}{2}\right) \Big|_1^a = \frac{a-1}{2};$$

$$P(B) = P\{Y > a\} = \int_a^{+\infty} f_Y(y) dy = \int_a^3 \frac{1}{2} dy = \left(\frac{y}{2}\right) \Big|_a^3 = \frac{3-a}{2};$$

$$\text{由 } P(A \cup B) = P(A) + P(B) - P(A)P(B) = \frac{a-1}{2} + \frac{3-a}{2} - \frac{a-1}{2} \cdot \frac{3-a}{2} = \frac{a^2 - 4a + 7}{4}.$$

$$\therefore \frac{a^2 - 4a + 7}{4} = \frac{7}{9}, \text{ 即 } 9a^2 - 36a + 35 = 0. \text{ 亦即 } (3a-5)(3a-7) = 0$$

$$\therefore a = \frac{5}{3} \text{ 或 } a = \frac{7}{3}.$$

14. 设 $X \sim U(a, b)$, $0 < a < b$, 已知 $P(X > 4) = \frac{1}{2}$, $P(3 < X < 4) = \frac{1}{4}$.

(1) 求 X 的密度函数 $f(x)$; (2) 计算 $P(0 < X < 3)$.

解. $X \sim U(a, b)$, $f(x) = \begin{cases} \frac{1}{b-a}, & a < x < b; \\ 0, & \text{其他} \end{cases}$

$$(1) P(X > 4) = \int_4^{+\infty} f(x) dx = \int_4^b \frac{1}{b-a} dx = \left(\frac{x}{b-a}\right) \Big|_4^b = \frac{b-4}{b-a};$$

$$P(3 < X < 4) = \int_3^4 \frac{1}{b-a} dx = \left(\frac{x}{b-a}\right) \Big|_3^4 = \frac{1}{b-a};$$

$$\therefore \begin{cases} \frac{b-4}{b-a} = \frac{1}{2} \\ \frac{1}{b-a} = \frac{1}{4} \end{cases} \quad \therefore \begin{cases} a+b=8; \\ a-b=-4. \end{cases} \quad \therefore \begin{cases} a=2; \\ b=6. \end{cases} \quad \therefore f(x) = \begin{cases} \frac{1}{4}; & 2 < x < 6 \\ 0. & \text{其他} \end{cases}$$

$$(2) P(0 < X < 3) = \int_0^3 f(x) dx = \int_2^3 \frac{1}{4} dx = \left(\frac{x}{4}\right) \Big|_2^3 = \frac{1}{4}.$$

15. 设随机变量 $X \sim N(18, 2.5^2)$, 求: (1) $P(17 < X < 21)$; (2) 求常数 K , 使 $P(X < K) = 0.2236$; (3) 求最大的 K 使 $P(X > K) \geq 0.1814$.

$$\text{解: (1) } P(17 < X < 21) = \Phi\left(\frac{21-18}{2.5}\right) - \Phi\left(\frac{17-18}{2.5}\right) = \Phi(1.2) - \Phi(-0.4)$$

$$= \Phi(1.2) - (1 - \Phi(0.4)) = \Phi(1.2) + \Phi(0.4) - 1 = 0.8849 + 0.6554 - 1 = 0.5403;$$

$$(2) P(X < K) = \Phi\left(\frac{K-18}{2.5}\right) = 0.2236; \therefore \Phi\left(\frac{18-K}{2.5}\right) = 0.7764.$$

$$\therefore \frac{18-K}{2.5} = 0.76, \quad \therefore K = 16.1.$$

$$(3) P(X > K) = 1 - P(X < K) = 1 - \Phi\left(\frac{K-18}{2.5}\right) \geq 0.1814.$$

$$\therefore \Phi\left(\frac{K-18}{2.5}\right) \leq 0.8186, \quad \therefore \frac{K-18}{2.5} \leq 0.91, \quad \therefore K \leq 20.275.$$

则最大的 $K = 20.275$

16. 设随机变量 X 的分布律为 $\begin{array}{c|ccc} X & 1 & 2 & 3 \\ \hline P & 1/6 & 1/3 & 1/2 \end{array}$. 随机变量 $Y \sim U(0, X)$.

(1) 求 $P(Y \leq 0.5)$; (2) 计算 $P(X=1/Y \leq 0.5)$.

解: (1) 设 X 取值为 K , ($K=1, 2, 3$), 在此条件下, $(Y \leq 0.5)$ 的概率为:

$$P(Y \leq 0.5/X=K) = \int_0^{0.5} \frac{1}{K} dy = \left(\frac{y}{K}\right)\bigg|_0^{0.5} = \frac{1}{2K}, \quad K=1, 2, 3.$$

$$\therefore P(Y \leq 0.5) = \sum_{K=1}^3 P(X=K) P(Y \leq 0.5/X=K) = \frac{1}{6} \cdot \frac{1}{2} + \frac{1}{3} \cdot \frac{1}{4} + \frac{1}{2} \cdot \frac{1}{6} = \frac{1}{4}.$$

$$(2) P(X=1/Y \leq 0.5) = P[(X=1) \cap (Y \leq 0.5)] / P(Y \leq 0.5)$$

$$= P(X=1) P(Y \leq 0.5/X=1) / P(Y \leq 0.5) = \frac{1}{6} \cdot \frac{1}{2} / \frac{1}{4} = \frac{1}{3}.$$

17. 设某种电子管寿命 X (以小时计) 服从正态分布 $N(160, \sigma^2)$. 若要求

$$P(120 \leq X \leq 200) \geq 0.80, \text{ 问允许 } \sigma \text{ 最大为多少?}$$

$$\text{解: } P(120 \leq X \leq 200) = \Phi\left(\frac{200-160}{\sigma}\right) - \Phi\left(\frac{120-160}{\sigma}\right) = \Phi\left(\frac{40}{\sigma}\right) - \Phi\left(-\frac{40}{\sigma}\right)$$

$$= 2\Phi\left(\frac{40}{\sigma}\right) - 1 \geq 0.80, \therefore \Phi\left(\frac{40}{\sigma}\right) \geq 0.90.$$

$$\text{则 } \frac{40}{\sigma} \geq 1.285, \therefore \sigma \leq 31.13. \text{ 最大的 } \sigma = 31.13.$$

18. 设某人造卫星偏离预定轨道的距离 $X \sim N(0, 4^2)$ (单位: m), 观测者把偏离值超过 10 m 时, 称作“失败”. (1) 求 5 次独立观测中至少有 2 次“失败”的概率; (2) 已知 5 次独立观测中有一次“失败”, 求 5 次观测中至少有 2 次“失败”的概率.

$$\text{解: } P(\text{“失败”}) = P(|X| > 10) = P[(X > 10) \cup (X < -10)] = P(X > 10) + P(X < -10)$$

$$= 1 - P(X < 10) + P(X < -10) = 1 - \Phi\left(\frac{10}{4}\right) + \Phi\left(-\frac{10}{4}\right) = 2(1 - \Phi(2.5))$$

$$= 2(1 - 0.9938) = 0.0124.$$

$$\text{记 } X_i = \begin{cases} 1, & \text{第 } i \text{ 次“失败”;} \\ 0, & \text{第 } i \text{ 次未“失败”.} \end{cases} \quad P(X_i = 1) = P(|X| > 10) = 0.0124, \quad i=1, 2, \dots, 5.$$

$$\text{又 } Y = \text{“5 次独立观测中“失败”的次数”} = \sum_{i=1}^5 X_i \sim B(5, 0.0124).$$

$$(1) P(Y \geq 2) = 1 - P(Y < 2) = 1 - P(Y=0) - P(Y=1)$$

$$= 1 - C_5^0 (0.9876)^5 - C_5^1 (0.0124)(0.9876)^4 = 0.0015.$$

$$(2) P(Y \geq 2/Y \geq 1) = P[(Y \geq 1) \cap (Y \geq 2)] / P(Y \geq 1) = \frac{P(Y \geq 2)}{P(Y \geq 1)}.$$

$$\therefore P(Y \geq 1) = 1 - P(Y < 1) = 1 - P(Y=0) = 1 - C_5^0 (0.9876)^5 = 0.06048$$

$$\therefore P(Y \geq 2/Y \geq 1) = 0.0015 / 0.06048 = 0.0248.$$

19. 设成年男子身高 $X \sim N(170, 36)$. (1) 问应如何选择公共汽车门的高度 h ,

才能使男乘客与车门碰头的机会小于 0.01? (2) 若车门高 182 cm,

求 100 个男子中与车门碰头的人数不多于 2 个的概率.

$$\text{解: (1) } P(\text{男乘客与车门碰头}) = P(X \geq h) = 1 - \Phi\left(\frac{h-170}{6}\right) < 0.01$$

$$\therefore \Phi\left(\frac{h-170}{6}\right) > 0.99, \therefore \frac{h-170}{6} > 2.325, \therefore h > 183.95.$$

$$(2) P(X \geq 182) = 1 - \Phi\left(\frac{182-170}{6}\right) = 1 - \Phi(2) = 1 - 0.9772 = 0.0228.$$

$$\text{设 } X_i = \begin{cases} 1, & \text{第 } i \text{ 人与车门碰头;} \\ 0, & \text{第 } i \text{ 人不与车门碰头.} \end{cases} P(X_i=1) = P(X \geq 182) = 0.0228$$

$i=1, 2, \dots, 100$, 相互独立.

$$\text{又 } Y = \text{"100个男子中与车门碰头的人数"} = \sum_{i=1}^{100} X_i \sim B(100, 0.0228).$$

$$\therefore P(Y \leq 2) = P(Y=0) + P(Y=1) + P(Y=2) = C_{100}^0 (0.9772)^{100} + C_{100}^1 0.0228 (0.9772)^{99} + C_{100}^2 (0.0228)^2 (0.9772)^{98} = 0.09962 + 0.23243 + 0.26844 \approx 0.6.$$

(注) 可采用泊松分布 $\pi(np)$ 近似 $B(n, p)$. (当 $n \geq 20, p \leq 0.05$ 时).

$$\text{则 } P(Y \leq 2) = e^{-2.28} + 2.28 e^{-2.28} + \frac{(2.28)^2}{2!} e^{-2.28}$$

$$= 0.10228 + 0.23321 + 0.26586 \approx 0.6$$

20. 在电源电压不超过 200V, 200V ~ 240V 和超过 240V 三种情况下, 某种电子元件损坏的概率分别为 0.1, 0.001 和 0.2. 假设电源电压 X 服从正态分布 $N(220, 25^2)$. 试求: (1) 该电子元件损坏的概率, (2) 该电子元件损坏时, 电源电压在 200V ~ 240V 之间的概率.

$$\text{解: } P(X \leq 200) = \Phi\left(\frac{200-220}{25}\right) = \Phi(-0.8) = 1 - \Phi(0.8) = 1 - 0.7881 = 0.2119.$$

$$P(200 < X < 240) = \Phi\left(\frac{240-220}{25}\right) - \Phi\left(\frac{200-220}{25}\right) = \Phi(0.8) - \Phi(-0.8) = 2\Phi(0.8) - 1 = 2 \times 0.7881 - 1 = 0.5762,$$

$$P(X \geq 240) = 1 - \Phi\left(\frac{240-220}{25}\right) = 1 - \Phi(0.8) = 1 - 0.7881 = 0.2119.$$

(1) 由题意, 设 $A = \text{"电子元件损坏"}$, 则

$$P(A/X \leq 200) = 0.1; P(A/200 < X < 240) = 0.001; P(A/X \geq 240) = 0.2.$$

$$\text{所以 } P(A) = P(X \leq 200) \cdot P(A/X \leq 200) + P(200 < X < 240) \cdot P(A/200 < X < 240) + P(X \geq 240) \cdot P(A/X \geq 240) = 0.1 \times 0.2119 + 0.001 \times 0.5762 + 0.2 \times 0.2119 = 0.0641.$$

$$(2) P(200 < X < 240/A) = \frac{1}{P(A)} \cdot P[(200 < X < 240) \cap A] = \frac{1}{P(A)} \cdot P(200 < X < 240) \cdot P(A/200 < X < 240) = \frac{0.001 \times 0.5762}{0.0641} = 0.009.$$

21. 某学校计划招生 800 人, 按考试成绩从高分到低分依次录取, 设参加考试的 3000 人的考试成绩服从正态分布, 且 600 分以上的有 200 人, 500 分以下的有 2075 人. 求录取分数线应定为多少?

$$\therefore P(X > 600) = 1 - \Phi\left(\frac{600 - \mu}{\sigma}\right) = \frac{200}{3000} = 0.0667;$$

$$\therefore \Phi\left(\frac{600 - \mu}{\sigma}\right) = 0.9333, \therefore \frac{600 - \mu}{\sigma} = 1.505; \text{ 即 } 600 - \mu = 1.505\sigma;$$

$$\text{又 } P(X < 500) = \Phi\left(\frac{500 - \mu}{\sigma}\right) = \frac{2075}{3000} = 0.6917;$$

$$\therefore \Phi\left(\frac{500 - \mu}{\sigma}\right) = 0.6917, \therefore \frac{500 - \mu}{\sigma} = 0.505. \text{ 即 } 500 - \mu = 0.505\sigma.$$

$$\text{由 } \begin{cases} 600 - \mu = 1.505\sigma, \\ 500 - \mu = 0.505\sigma. \end{cases} \therefore \begin{cases} \mu = 449.5, \\ \sigma = 100. \end{cases} \text{ 则 } X \sim N(449.5, 100^2).$$

$$\text{设分数线为 } x \text{ 分. } P(X > x) = \frac{800}{3000} = 0.2667$$

$$\text{即 } P(X \leq x) = \Phi\left(\frac{x - 449.5}{100}\right) = 1 - P(X > x) = 1 - 0.2667 = 0.7333.$$

$$\therefore \frac{x - 449.5}{100} = 0.625. \therefore x = 512 \text{ (分)}.$$

22. 设随机变量 X 具有密度函数 $f(x) = \frac{1}{\pi(1+x^2)}, -\infty < x < \infty$.

求 (1) $Y = \arctan X$ 的密度函数; (2) $Z = 1 - \sqrt[3]{X}$ 的密度函数.

解: (1) $F_Y(y) = P(Y \leq y) = P(\arctan X \leq y) = P(X \leq \tan y) = F_X(\tan y)$

$$\therefore f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} F_X(\tan y) = f_X(\tan y) \cdot \sec^2 y.$$

$$= \begin{cases} \frac{1}{\pi(1+\tan^2 y)} \cdot \sec^2 y, & -\frac{\pi}{2} < y < \frac{\pi}{2}, \\ 0, & \text{其他.} \end{cases} = \begin{cases} \frac{1}{\pi}, & -\frac{\pi}{2} < y < \frac{\pi}{2}, \\ 0, & \text{其他.} \end{cases}$$

$$(2) F_Z(u) = P(Z \leq u) = P(1 - \sqrt[3]{X} \leq u) = P(X > (1-u)^3) = 1 - F_X((1-u)^3)$$

$$\therefore f_Z(u) = \frac{d}{du} F_Z(u) = \frac{d}{du} (1 - F_X((1-u)^3)) = -f_X((1-u)^3) \cdot (-3(1-u)^2)$$

$$= \frac{1}{\pi[1-(1-u)^3]} \cdot 3(1-u)^2 = \frac{3(1-u)^2}{\pi[1-(1-u)^3]}, \quad -\infty < u < +\infty.$$

23. 设 $X \sim U(-\frac{\theta}{2}, \theta)$, 其中 $\theta > 0$ 为常数, 求随机变量 $Y = |X|$ 的密度函数.

解: $X \sim U(-\frac{\theta}{2}, \theta), \therefore f_X(x) = \begin{cases} \frac{2}{3\theta}, & -\frac{\theta}{2} < x < \theta, \\ 0, & \text{其他.} \end{cases}$

$$F_Y(y) = P(Y \leq y) = P(|X| \leq y).$$

$$\text{当 } y \leq 0 \text{ 时, } F_Y(y) = P(\emptyset) = 0. \therefore f_Y(y) = 0;$$

$$\text{当 } y \geq 0 \text{ 时, } F_Y(y) = P(|X| \leq y) = P(-y < X < y) = F_X(y) - F_X(-y).$$

$$\therefore f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} [F_X(y) - F_X(-y)] = f_X(y) + f_X(-y)$$

$$= \begin{cases} 4/3\theta, & 0 < y < \theta/2, \\ 2/3\theta, & \theta/2 \leq y < \theta, \\ 0, & y \geq \theta. \end{cases} \therefore f_Y(y) = \begin{cases} 4/3\theta, & 0 < y < \theta/2, \\ 2/3\theta, & \theta/2 \leq y < \theta, \\ 0, & y \geq \theta. \end{cases}$$

24. 设随机变量 X 的密度函数为 $f_X(x) = \begin{cases} \frac{2x}{\pi^2}, & 0 < x < \pi; \\ 0, & \text{其他.} \end{cases}$

求 $Y = \sin X$ 的密度函数.

解: 在 $0 < y < 1$ 时.

$$\begin{aligned} F_Y(y) &= P(Y \leq y) = P(\sin X \leq y) = P[(X \leq \arcsin y) \cup (\pi - X \leq \arcsin y)] \\ &= P(X \leq \arcsin y) + P(X \geq \pi - \arcsin y) = F_X(\arcsin y) + 1 - F_X(\pi - \arcsin y) \end{aligned}$$

$$\begin{aligned} \therefore f_Y(y) &= \frac{d}{dy} F_Y(y) = \frac{d}{dy} [F_X(\arcsin y) + 1 - F_X(\pi - \arcsin y)] \\ &= f_X(\arcsin y) \cdot \frac{1}{\sqrt{1-y^2}} - f_X(\pi - \arcsin y) \cdot \left(\frac{-1}{\sqrt{1-y^2}} \right) \\ &= \frac{1}{\sqrt{1-y^2}} [f_X(\arcsin y) + f_X(\pi - \arcsin y)] \\ &= \frac{1}{\sqrt{1-y^2}} \left[\frac{2}{\pi^2} \arcsin y + \frac{2}{\pi^2} (\pi - \arcsin y) \right] = \frac{1}{\pi \sqrt{1-y^2}}. \end{aligned}$$

$$\therefore f_Y(y) = \begin{cases} \frac{1}{\pi \sqrt{1-y^2}}, & 0 < y < 1; \\ 0, & \text{其他.} \end{cases}$$

25. 设随机变量 X 的密度函数为 $f(x) = \frac{1}{2} e^{-|x|}, -\infty < x < +\infty$.

(1) 求 X 的分布函数; (2) 设 $Y = \begin{cases} 1, & X > 0; \\ -1, & X \leq 0. \end{cases}$ 求 Y 的分布函数.

解: (1) $F_X(x) = P(X \leq x) = \int_{-\infty}^x f(x) dx = \int_{-\infty}^x \frac{1}{2} e^{-|x|} dx$.

当 $x \leq 0$ 时, $F_X(x) = \int_{-\infty}^x \frac{1}{2} e^x dx = \left(\frac{1}{2} e^x \right) \Big|_{-\infty}^x = \frac{1}{2} e^x$;

当 $x > 0$ 时, $F_X(x) = \int_{-\infty}^x \frac{1}{2} e^{-|x|} dx = \int_{-\infty}^0 \frac{1}{2} e^x dx + \int_0^x \frac{1}{2} e^{-x} dx = \frac{e^x}{2} \Big|_{-\infty}^0 + \left(-\frac{e^{-x}}{2} \right) \Big|_0^x$

$$= \frac{1}{2} + \left(\frac{1}{2} - \frac{e^{-x}}{2} \right) = 1 - \frac{e^{-x}}{2},$$

$$\therefore F_X(x) = \begin{cases} \frac{e^x}{2}, & x \leq 0; \\ 1 - \frac{1}{2} e^{-x}, & x > 0. \end{cases}$$

(2) $P(Y=1) = P(X > 0) = 1 - F_X(0) = 1 - \frac{1}{2} = \frac{1}{2}$;

$P(Y=-1) = P(X \leq 0) = F_X(0) = \frac{1}{2}$.

即

Y	-1	1
P	$\frac{1}{2}$	$\frac{1}{2}$

$$\therefore F_Y(y) = \begin{cases} 0, & y < -1; \\ \frac{1}{2}, & -1 \leq y \leq 1; \end{cases}$$

习题 3-1

1. 盒中有 4 个红球 1 个白球，从盒中任取两次，每次取一球，令

$$X = \begin{cases} 1, & \text{第一次取到红球;} \\ 0, & \text{第一次取到白球;} \end{cases} \quad Y = \begin{cases} 1, & \text{第二次取到红球;} \\ 0, & \text{第二次取到白球;} \end{cases}$$

求：(1) 在有放回抽样情形下， (X, Y) 的联合分布律；

(2) 在不放回抽样情形下， (X, Y) 的联合分布律。

解：(1) 有放回的情形：

$$P(X=0, Y=0) = \frac{1}{5^2}, \quad P(X=0, Y=1) = \frac{1 \cdot 4}{5^2} = \frac{4}{5^2}$$

$$P(X=1, Y=0) = \frac{4}{5^2}, \quad P(X=1, Y=1) = \frac{4^2}{5^2}$$

$X \backslash Y$	0	1
0	$\frac{1}{25}$	$\frac{4}{25}$
1	$\frac{4}{25}$	$\frac{16}{25}$

(2) 不放回情形：

$$P(X=0, Y=0) = \frac{1}{5} \cdot \frac{0}{4} = 0, \quad P(X=0, Y=1) = \frac{1}{5} \cdot \frac{4}{4} = \frac{1}{5}$$

$$P(X=1, Y=0) = \frac{4}{5} \cdot \frac{1}{4} = \frac{1}{5}, \quad P(X=1, Y=1) = \frac{4}{5} \cdot \frac{3}{4} = \frac{3}{5}$$

$X \backslash Y$	0	1
0	0	$\frac{1}{5}$
1	$\frac{1}{5}$	$\frac{3}{5}$

2. 盒中有 4 个红球，1 个白球，3 个黑球，从盒中不放回的任取 4 球，试求取得红球数-5 白球数的联合分布。

解：设： $X = \text{"4 球中的红球数"}; Y = \text{"4 球中的白球数"}.$

$$P(X=0, Y=0) = 0, \quad P(X=0, Y=1) = \frac{C_4^0 C_1^1 C_3^3}{C_8^4}$$

$$P(X=1, Y=0) = \frac{C_4^1 C_1^0 C_3^3}{C_8^4}; \quad P(X=1, Y=1) = \frac{C_4^1 C_1^1 C_3^2}{C_8^4}$$

$$P(X=2, Y=0) = \frac{C_4^2 C_1^0 C_3^2}{C_8^4}; \quad P(X=2, Y=1) = \frac{C_4^2 C_1^1 C_3^1}{C_8^4}$$

$$P(X=3, Y=0) = \frac{C_4^3 C_1^0 C_3^1}{C_8^4}; \quad P(X=3, Y=1) = \frac{C_4^3 C_1^1 C_3^0}{C_8^4}$$

$$P(X=4, Y=0) = \frac{C_4^4 C_1^0 C_3^0}{C_8^4}, \quad P(X=4, Y=1) = 0$$

$X \backslash Y$	0	1
0	0	$\frac{1}{70}$
1	$\frac{4}{70}$	$\frac{12}{70}$
2	$\frac{18}{70}$	$\frac{18}{70}$
3	$\frac{12}{70}$	$\frac{4}{70}$
4	$\frac{1}{70}$	0

3. 设 (X, Y) 的联合分布律。

$X \backslash Y$	-1	0	1
0	0.07	0.18	0.15
1	0.08	a	0.20

求：(1) a ;

$$(2) P(X \leq 0, Y \leq 0);$$

$$(3) P(X \leq 0, Y < 0).$$

解：(1) $1 = \sum_i \sum_j P(X=x_i, Y=y_j) = 0.07 + 0.18 + 0.15 + 0.08 + a + 0.20$
 $= 0.68 + a, \quad \therefore a = 0.32$

(2) $P(X \leq 0, Y \leq 0) = P(X=0, Y=-1) + P(X=0, Y=0) = 0.07 + 0.18$
 $= 0.25$

(3) $P(X \leq 0, Y < 0) = P(X=0, Y=-1) = 0.07$

4. 甲、乙两人独立地各进行两次射击，假设甲的命中率为0.2，乙的命中率为0.5，以 X 和 Y 分别表示甲和乙的命中次数，试求 X 和 Y 的联合分布律。

解：设 X = “甲两次射击命中数”， Y = “乙两次射击命中数”

$$\therefore X \sim B(2, 0.2), P(X=k) = C_2^k (0.2)^k (0.8)^{2-k}, \quad k=0, 1, 2.$$

$$Y \sim B(2, 0.5), P(Y=k) = C_2^k (0.5)^k (0.5)^{2-k}, \quad k=0, 1, 2.$$

$$\begin{aligned} P(X=i, Y=j) &= P(X=i) \cdot P(Y=j) = C_2^i (0.2)^i (0.8)^{2-i} \cdot C_2^j (0.5)^j (0.5)^{2-j} \\ &= C_2^i C_2^j (0.2)^i (0.5)^2 (0.8)^{2-i}, \quad i, j = 0, 1, 2. \end{aligned}$$

$$P(X=0, Y=0) = C_2^0 C_2^0 (0.2)^0 (0.5)^2 (0.8)^2 = 0.16;$$

$$P(X=0, Y=1) = C_2^0 C_2^1 (0.2)^0 (0.5)^2 (0.8)^2 = 0.32;$$

$$P(X=0, Y=2) = C_2^0 C_2^2 (0.2)^0 (0.5)^2 (0.8)^2 = 0.16;$$

$$P(X=1, Y=0) = C_2^1 C_2^0 (0.2) (0.5)^2 (0.8) = 0.08;$$

$$P(X=1, Y=1) = C_2^1 C_2^1 (0.2) (0.5)^2 (0.8) = 0.16;$$

$$P(X=1, Y=2) = C_2^1 C_2^2 (0.2) (0.5)^2 (0.8) = 0.08;$$

$$P(X=2, Y=0) = C_2^2 C_2^0 (0.2)^2 (0.5)^2 (0.8)^0 = 0.01;$$

$$P(X=2, Y=1) = C_2^2 C_2^1 (0.2)^2 (0.5)^2 (0.8)^0 = 0.02;$$

$$P(X=2, Y=2) = C_2^2 C_2^2 (0.2)^2 (0.5)^2 (0.8)^0 = 0.01.$$

$X \backslash Y$	0	1	2
0	0.16	0.32	0.16
1	0.08	0.16	0.08
2	0.01	0.02	0.01

1. 设 (X, Y) 的联合分布函数

$$F(x, y) = A(B + \arctan x)(C + \arctan y), \quad -\infty < x, y < +\infty.$$

(1) 求常数 A, B, C 的值; (2) 求 (X, Y) 的联合分布密度函数.

解: (1) 由 $F(x, -\infty) = A(B + \arctan x)(C - \frac{\pi}{2}) = 0$; ①

$$F(-\infty, y) = A(B - \frac{\pi}{2})(C + \arctan y) = 0; \quad ②$$

$$F(+\infty, +\infty) = A(B + \frac{\pi}{2})(C + \frac{\pi}{2}) = 1. \quad ③$$

由 ① 得 $C = \frac{\pi}{2}$; 由 ② 得 $B = \frac{\pi}{2}$. 代入 ③ 得 $A = \frac{1}{\pi^2}$.

$$\therefore F(x, y) = \frac{1}{\pi^2}(\frac{\pi}{2} + \arctan x)(\frac{\pi}{2} + \arctan y), \quad -\infty < x, y < +\infty.$$

$$(2) f(x, y) = \frac{\partial^2 F(x, y)}{\partial x \partial y} = \frac{1}{\pi^2(1+x^2)(1+y^2)}, \quad -\infty < x, y < +\infty.$$

2. 设 (X, Y) 的联合密度函数

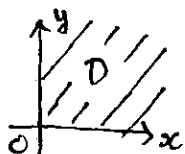
$$f(x, y) = \begin{cases} Ke^{-(3x+2y)} & x > 0, y > 0; \\ 0 & \text{其他.} \end{cases}$$

求: (1) 常数 K ;

(2) (X, Y) 的联合分布函数;

(3) $P(X \leq Y)$.

解: (1)



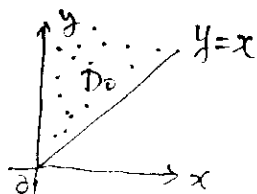
$$\begin{aligned} \text{由 } 1 &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x, y) dx dy = \iint_D Ke^{-(3x+2y)} dx dy \\ &= \int_0^{+\infty} dx \int_0^{+\infty} Ke^{-(3x+2y)} dy = \int_0^{+\infty} Ke^{-3x} \left(\frac{-e^{-2y}}{2} \right) \Big|_0^{+\infty} dx \\ &= \int_0^{+\infty} \frac{K}{2} e^{-3x} dx = \frac{K}{6} (-e^{-3x}) \Big|_0^{+\infty} = \frac{K}{6}, \quad \therefore K = 6. \end{aligned}$$

(2) 当 $x > 0, y > 0$ 时,

$$\begin{aligned} F(x, y) &= \int_{-\infty}^x \int_{-\infty}^y f(x, y) dx dy = \int_0^x \int_0^y 6e^{-(3x+2y)} dx dy = \int_0^y 3e^{-3x} dx \cdot \int_0^y 2e^{-2y} dy \\ &= (-e^{-3x}) \Big|_0^x \cdot (-e^{-2y}) \Big|_0^y = (1 - e^{-3x})(1 - e^{-2y}) \end{aligned}$$

$$\therefore F(x, y) = \begin{cases} (1 - e^{-3x})(1 - e^{-2y}), & x > 0, y > 0; \\ 0 & \text{其他.} \end{cases}$$

(3)

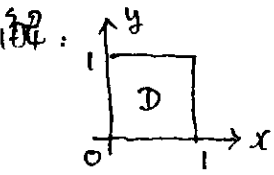


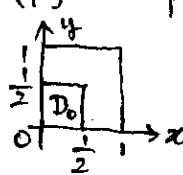
$$\begin{aligned} P(X \leq Y) &= \iint_{x \leq y} f(x, y) dx dy = \iint_{D_0} 6e^{-(3x+2y)} dx dy \\ &= \int_0^{+\infty} \int_0^y 6e^{-3x-2y} dx dy = \int_0^{+\infty} 3e^{-2y} (-e^{-3x}) \Big|_0^y dy \\ &= \int_0^{+\infty} 3e^{-2y} (-e^{-3y}) dy = \int_0^{+\infty} -3e^{-5y} dy = \int_0^{+\infty} 3e^{-5y} dy = \frac{3}{5} \end{aligned}$$

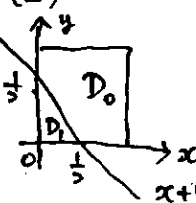
3. 设 (X, Y) 的联合密度函数 $f(x, y) = \begin{cases} Kxy, & 0 \leq x, y < 1. \\ 0, & \text{其他.} \end{cases}$

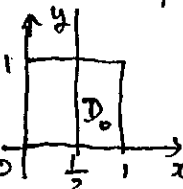
求常数 K 及下列随机事件的概率: (1) $P(X \leq \frac{1}{2}, Y \leq \frac{1}{2})$,

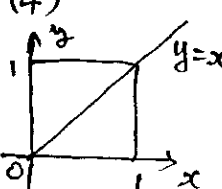
(2) $P(X+Y > \frac{1}{2})$; (3) $P(X > \frac{1}{2})$; (4) $P(X=Y)$.

解:  由 $1 = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x, y) dx dy = \iint_D Kxy dx dy = \int_0^1 Kx dx \int_0^1 y dy$
 $= K \left(\frac{x}{2} \right) \Big|_0^1 \left(\frac{y}{2} \right) \Big|_0^1 = \frac{K}{4}, \therefore K=4.$

(1)  $P(X \leq \frac{1}{2}, Y \leq \frac{1}{2}) = \iint_{D_0} f(x, y) dx dy = \iint_{D_0} 4xy dx dy = \int_0^{\frac{1}{2}} 2x dx \int_0^{\frac{1}{2}} 2y dy$
 $= (x^2) \Big|_0^{\frac{1}{2}} (y^2) \Big|_0^{\frac{1}{2}} = \frac{1}{4} \cdot \frac{1}{4} = \frac{1}{16}.$

(2)  $P(X+Y > \frac{1}{2}) = \iint_{D_0} f(x, y) dx dy = \iint_{D_0} 4xy dx dy = 1 - \iint_{D_1} 4xy dx dy$
 $= 1 - \int_0^{\frac{1}{2}} dx \int_0^{\frac{1}{2}-x} 4xy dy = 1 - \int_0^{\frac{1}{2}} 2x(y^2) \Big|_0^{\frac{1}{2}-x} dx = 1 - \int_0^{\frac{1}{2}} 2x(\frac{1}{2}-x)^2 dx$
 $= 1 - \int_0^{\frac{1}{2}} (\frac{x}{2} - 2x^2 + 2x^3) dx = 1 - (\frac{x^2}{4} - \frac{2x^3}{3} + \frac{2x^4}{4}) \Big|_0^{\frac{1}{2}} = 1 - \frac{1}{96} = \frac{95}{96}.$

(3)  $P(X > \frac{1}{2}) = \iint_{D_0} f(x, y) dx dy = \iint_{D_0} 4xy dx dy = \int_{\frac{1}{2}}^1 dx \int_0^1 4xy dy$
 $= \int_{\frac{1}{2}}^1 2x dx \int_0^1 2y dy = (x^2) \Big|_{\frac{1}{2}}^1 (y^2) \Big|_0^1 = 1 - \frac{1}{4} = \frac{3}{4}.$

(4)  $P(X=Y) = \iint_{D_0} f(x, y) dx dy = \iint_{D_0} 4xy dx dy = \int_0^1 dx \int_x^x 4xy dy$
 $= 0$
 $\begin{matrix} x=y \\ 0 \leq x \leq 1 \\ 0 \leq y \leq 1 \end{matrix}$

习题 3-3

1. 设 (X, Y) 的联合分布函数为

$$F(x, y) = \frac{1}{\pi^2} \left(\frac{\pi}{2} + \arctan \frac{x}{2} \right) \left(\frac{\pi}{2} + \arctan \frac{y}{3} \right), \quad -\infty < x, y < +\infty$$

求关于 X, Y 的边缘分布函数.

解: $F_X(x) = F(x, +\infty) = \lim_{y \rightarrow +\infty} F(x, y) = \lim_{y \rightarrow +\infty} \frac{1}{\pi^2} \left(\frac{\pi}{2} + \arctan \frac{x}{2} \right) \left(\frac{\pi}{2} + \arctan \frac{y}{3} \right)$

$$= \frac{1}{\pi^2} \left(\frac{\pi}{2} + \arctan \frac{x}{2} \right) \left(\frac{\pi}{2} + \lim_{y \rightarrow +\infty} \arctan \frac{y}{3} \right) = \frac{1}{\pi} \left(\frac{\pi}{2} + \arctan \frac{x}{2} \right), \quad -\infty < x < +\infty$$

$$F_Y(y) = F(+\infty, y) = \lim_{x \rightarrow +\infty} F(x, y) = \lim_{x \rightarrow +\infty} \frac{1}{\pi^2} \left(\frac{\pi}{2} + \arctan \frac{x}{2} \right) \left(\frac{\pi}{2} + \arctan \frac{y}{3} \right)$$

$$= \frac{1}{\pi^2} \left(\frac{\pi}{2} + \arctan \frac{y}{3} \right) \left(\frac{\pi}{2} + \lim_{x \rightarrow +\infty} \arctan \frac{x}{2} \right) = \frac{1}{\pi} \left(\frac{\pi}{2} + \arctan \frac{y}{3} \right), \quad -\infty < y < +\infty.$$

2. 设 (X, Y) 的可能取值为 $(0, 0), (-1, 1), (-1, \frac{1}{3}), (2, 0), (2, \frac{1}{3})$. 相应的概率为 $\frac{1}{6}, \frac{1}{3}, \frac{1}{12}, \frac{1}{4}, \frac{1}{6}$. (1) 列表表示其联合分布律; (2) 求关于 X, Y 的边缘分布律.

解: (1)

$X \backslash Y$	0	$\frac{1}{3}$	1
-1	0	$\frac{1}{12}$	$\frac{1}{3}$
0	$\frac{1}{6}$	0	0
2	$\frac{1}{4}$	$\frac{1}{6}$	0

(2)

X	-1	0	2
P	$\frac{5}{12}$	$\frac{2}{12}$	$\frac{5}{12}$

Y	0	$\frac{1}{3}$	1
P	$\frac{5}{12}$	$\frac{3}{12}$	$\frac{4}{12}$

3. 设 (X, Y) 的联合密度函数为 $f(x, y)$, 分别求关于 X, Y 的边缘密度函数.

(1) $f(x, y) = \begin{cases} 4xye^{-(x^2+y^2)} & x > 0, y > 0; \\ 0 & \text{其他.} \end{cases}$ (2) $f(x, y) = \begin{cases} 6xy(2-x-y), & 0 < x, y < 1, \\ 0 & \text{其他.} \end{cases}$

(3) $f(x, y) = \begin{cases} 4.8y(2-x), & 0 \leq x \leq 1, 0 \leq y \leq x; \\ 0 & \text{其他.} \end{cases}$

(4) $f(x, y) = \frac{1}{\pi^2(1+x^2)(1+y^2)}, \quad -\infty < x, y < +\infty.$

解: (1)

$$f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \int_0^{+\infty} 4xye^{-(x^2+y^2)} dy = 2xe^{-x^2} \int_0^{+\infty} 2ye^{-y^2} dy$$

$$= 2xe^{-x^2} \cdot (-e^{-y^2}) \Big|_0^{+\infty} = 2xe^{-x^2}, \quad (x > 0).$$

$\therefore f_X(x) = \begin{cases} 2xe^{-x^2}, & x > 0; \\ 0 & \text{其他.} \end{cases}$

(2) $f_x(x) = \int_{-\infty}^{+\infty} f(x,y) dy = \int_0^1 6xy(2-x-y) dy = \int_0^1 (12xy - 6x^2y - 6xy^2) dy$
 $= 6x(2-x) \left(\frac{y^2}{2} \right) \Big|_0^1 - 2x(y^3) \Big|_0^1 = 3x(2-x) - 2x = 4x - 3x^2, 0 < x < 1$

$\therefore f_x(x) = \begin{cases} 4x - 3x^2, & 0 < x < 1; \\ 0, & \text{其他} \end{cases}$

$f_y(y) = \int_{-\infty}^{+\infty} f(x,y) dx = \int_0^1 6xy(2-x-y) dx = \int_0^1 (12xy - 6x^2y - 6xy^2) dx$
 $= 6y(x^2) \Big|_0^1 - 2y(x^3) \Big|_0^1 - 3y^2(x^2) \Big|_0^1 = 6y - 2y - 3y^2 = 4y - 3y^2, 0 < y < 1$

$\therefore f_y(y) = \begin{cases} 4y - 3y^2, & 0 < y < 1; \\ 0, & \text{其他} \end{cases}$

(3) $f_x(x) = \int_{-\infty}^{+\infty} f(x,y) dy = \int_0^x 4.8y(2-x) dy = 4.8(2-x) \left(\frac{y^2}{2} \right) \Big|_0^x$
 $= 2.4x(2-x), 0 < x < 1.$

$\therefore f_x(x) = \begin{cases} 2.4x(2-x), & 0 < x < 1; \\ 0, & \text{其他} \end{cases}$

$f_y(y) = \int_{-\infty}^{+\infty} f(x,y) dx = \int_y^1 4.8y(2-x) dx = 9.6y(x) \Big|_y^1 - 4.8y \left(\frac{x^2}{2} \right) \Big|_y^1$
 $= 9.6y(1-y) - 2.4y(1-y^2) = 2.4y(3-4y+y^2), 0 < y < 1.$

$\therefore f_y(y) = \begin{cases} 2.4y(3-4y+y^2), & 0 < y < 1; \\ 0, & \text{其他} \end{cases}$

(9) $f_x(x) = \int_{-\infty}^{+\infty} f(x,y) dy = \int_{-\infty}^{+\infty} \frac{dy}{\pi^2(1+x^2)(1+y^2)} = \frac{1}{\pi^2(1+x^2)} (\arctan y) \Big|_{-\infty}^{+\infty}$
 $= \frac{1}{\pi(1+x^2)}, -\infty < x < +\infty;$

$f_y(y) = \int_{-\infty}^{+\infty} f(x,y) dx = \int_{-\infty}^{+\infty} \frac{dx}{\pi^2(1+x^2)(1+y^2)} = \frac{1}{\pi^2(1+y^2)} (\arctan x) \Big|_{-\infty}^{+\infty}$
 $= \frac{1}{\pi(1+y^2)}, -\infty < y < +\infty.$

4. 设 (X,Y) 服从二维正态分布. (1) 当参数 $\mu_1=1, \mu_2=-1, \sigma_1^2=25, \sigma_2^2=9, \rho=\frac{3}{5}$ 时, 写出 (X,Y) 的联合密度函数和边缘密度函数; (2) 若 (X,Y) 的联合密度函数 $f(x,y) = \frac{\sqrt{3}}{2\pi} \exp\left\{-\frac{1}{6}[4(x-4)^2 - 6(x-4)(y+1) + 9(y+1)^2]\right\}$, $-\infty < x, y < +\infty$. 求参数 $\mu_1, \mu_2, \sigma_1^2, \sigma_2^2$ 和 ρ . 并写出 X, Y 的边缘密度函数.

解: $(X,Y) \sim N(\mu_1, \mu_2; \sigma_1^2, \sigma_2^2; \rho)$ 的联合密度函数:

$f(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left\{-\frac{1}{2(1-\rho^2)}\left[\frac{(x-\mu_1)^2}{\sigma_1^2} - 2\rho\frac{(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + \frac{(y-\mu_2)^2}{\sigma_2^2}\right]\right\}, -\infty < x, y < +\infty$

(1) 显然 $(X, Y) \sim N(1, -1; 25, 9; \frac{9}{5})$. 联合密度函数:

$$f(x, y) = \frac{1}{24\pi} \exp\left\{-\frac{1}{288}[9(x-1)^2 - 18(x-1)(y+1) + 25(y+1)^2]\right\} \quad -\infty < x, y < +\infty.$$

$$f_X(x) = \int_{-\infty}^{+\infty} \frac{1}{24\pi} \exp\left\{-\frac{1}{288}[9(x-1)^2 - 18(x-1)(y+1) + 25(y+1)^2]\right\} dy$$

$$= \int_{-\infty}^{+\infty} \frac{1}{24\pi} \exp\left\{-\frac{25}{32}\left[\left(\frac{x-1}{5}\right)^2 - \frac{6}{5}\left(\frac{x-1}{5}\right)\left(\frac{y+1}{3}\right) + \left(\frac{y+1}{3}\right)^2\right]\right\} dy$$

$$\underline{\underline{\hat{u} = \frac{x-1}{5}, \quad v = \frac{y+1}{3}}}} \quad \int_{-\infty}^{+\infty} \frac{1}{8\pi} \exp\left\{-\frac{25}{32}[u^2 - \frac{6}{5}u \cdot v + v^2]\right\} dv$$

$$= \int_{-\infty}^{+\infty} \frac{1}{8\pi} \exp\left\{-\frac{25}{32}\left[\frac{16}{25}u^2 + \left(\frac{9}{25}u^2 - \frac{6}{5}u \cdot v + v^2\right)\right]\right\} dv$$

$$= \frac{1}{8\pi} e^{-\frac{u^2}{2}} \cdot \int_{-\infty}^{+\infty} e^{-\frac{25}{32}(v - \frac{3}{5}u)^2} dv. \quad \underline{\underline{\hat{t} = \frac{5(v - \frac{3}{5}u)}{4}}}, \quad \text{则}$$

$$f_X(x) = \frac{1}{8\pi} \cdot \frac{4}{5} e^{-\frac{u^2}{2}} \cdot \int_{-\infty}^{+\infty} e^{-\frac{t^2}{2}} dt = \frac{1}{\sqrt{2\pi} \cdot 5} e^{-\frac{u^2}{2}} \cdot \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} e^{-\frac{t^2}{2}} dt$$

$$= \frac{1}{\sqrt{2\pi} \cdot 5} e^{-\frac{u^2}{2}} \quad \underline{\underline{\frac{x-1}{5} = u}} \quad \frac{1}{\sqrt{2\pi} \cdot 5} e^{-\frac{(x-1)^2}{2 \cdot 5^2}}, \quad -\infty < x < +\infty.$$

即 $X \sim N(1, 5^2)$

同理可得: $f_Y(y) = \frac{1}{\sqrt{2\pi} \cdot 3} e^{-\frac{(y+1)^2}{2 \cdot 3^2}}, \quad -\infty < y < +\infty.$

即 $Y \sim N(-1, 3^2)$.

(注): 一般当 $(X, Y) \sim N(\mu_1, \mu_2; \sigma_1^2, \sigma_2^2; \rho)$ 时.

则可得 $X \sim N(\mu_1, \sigma_1^2)$ 和 $Y \sim N(\mu_2, \sigma_2^2)$. 反之不一定.

(2) $f(x, y) = \frac{\sqrt{3}}{2\pi} \exp\left\{-\frac{1}{6}[4(x-4)^2 - 6(x-4)(y+1) + 9(y+1)^2]\right\}$

$$= \frac{\sqrt{3}}{2\pi} \exp\left\{-\frac{4}{6}\left[\left(\frac{x-4}{1}\right)^2 - \left(\frac{x-4}{1}\right)\left(\frac{y+1}{2/3}\right) + \left(\frac{y+1}{2/3}\right)^2\right]\right\}$$

$$\therefore \mu_1 = 4, \mu_2 = -1; \sigma_1^2 = 1, \sigma_2^2 = \frac{4}{9}, \rho = \frac{1}{2}.$$

则 $f_X(x) = \frac{1}{\sqrt{2\pi} \cdot 1} e^{-\frac{(x-4)^2}{2 \cdot 1^2}}, \quad -\infty < x < +\infty.$ 即 $X \sim N(4, 1)$

$$f_Y(y) = \frac{3}{\sqrt{2\pi} \cdot 3} e^{-\frac{(y+1)^2}{2 \cdot (\frac{2}{3})^2}}, \quad -\infty < y < +\infty. \quad \text{即 } Y \sim N(-1, \frac{4}{9})$$

习题 3-4 *

1. 盒子中有6个球, 分别标有号码 1, 1, 2, 2, 2, 3. 不放回地随机取2个球, 用 X 和 Y 分别表示取得的第一个球与第二个球的号码. (1) 求 (X, Y) 的联合分布律, (2) 求关于 X, Y 的边缘分布律, (3) 分别求 X 在 $Y=1, Y=2$ 和 $Y=3$ 下的条件分布律.

解. 记 $X = \text{"第一个球的号码"}; Y = \text{"第二个球的号码"}$

(1) 注意取球有序. 故用乘法公式 $P(X=x, Y=y) = P(X=x) \cdot P(Y=y/X=x)$.

$X \backslash Y$	1	2	3
1	$\frac{2}{6} \cdot \frac{1}{5} = \frac{2}{30}$	$\frac{2}{6} \cdot \frac{3}{5} = \frac{6}{30}$	$\frac{2}{6} \cdot \frac{1}{5} = \frac{2}{30}$
2	$\frac{3}{6} \cdot \frac{2}{5} = \frac{6}{30}$	$\frac{3}{6} \cdot \frac{2}{5} = \frac{6}{30}$	$\frac{3}{6} \cdot \frac{1}{5} = \frac{3}{30}$
3	$\frac{1}{6} \cdot \frac{2}{5} = \frac{2}{30}$	$\frac{1}{6} \cdot \frac{3}{5} = \frac{3}{30}$	$\frac{1}{6} \cdot \frac{0}{5} = 0$

\Rightarrow

$X \backslash Y$	1	2	3
1	$\frac{2}{30}$	$\frac{6}{30}$	$\frac{2}{30}$
2	$\frac{6}{30}$	$\frac{6}{30}$	$\frac{3}{30}$
3	$\frac{2}{30}$	$\frac{3}{30}$	0

(2)

X	1	2	3
P	$\frac{2}{6}$	$\frac{3}{6}$	$\frac{1}{6}$

Y	1	2	3
P	$\frac{2}{6}$	$\frac{3}{6}$	$\frac{1}{6}$

(3) $P(X=1/Y=1) = P(X=1, Y=1) / P(Y=1) = \frac{2/30}{2/6} = \frac{1}{5}$;
 $P(X=2/Y=1) = P(X=2, Y=1) / P(Y=1) = \frac{6/30}{2/6} = \frac{3}{5}$;
 $P(X=3/Y=1) = P(X=3, Y=1) / P(Y=1) = \frac{2/30}{2/6} = \frac{1}{5}$.

\therefore

X	1	2	3
$P(X=x/Y=1)$	$\frac{1}{5}$	$\frac{3}{5}$	$\frac{1}{5}$

$P(X=1/Y=2) = P(X=1, Y=2) / P(Y=2) = \frac{6/30}{3/6} = \frac{2}{5}$;

$P(X=2/Y=2) = P(X=2, Y=2) / P(Y=2) = \frac{6/30}{3/6} = \frac{2}{5}$;

$P(X=3/Y=2) = P(X=3, Y=2) / P(Y=2) = \frac{3/30}{3/6} = \frac{1}{5}$.

\therefore

X	1	2	3
$P(X=x/Y=2)$	$\frac{2}{5}$	$\frac{2}{5}$	$\frac{1}{5}$

$P(X=1/Y=3) = P(X=1, Y=3) / P(Y=3) = \frac{2/30}{1/6} = \frac{2}{5}$;

$P(X=2/Y=3) = P(X=2, Y=3) / P(Y=3) = \frac{3/30}{1/6} = \frac{3}{5}$;

$P(X=3/Y=3) = P(X=3, Y=3) / P(Y=3) = 0 / \frac{1}{6} = 0$.

\therefore

X	1	2	3
$P(X=x/Y=3)$	$\frac{2}{5}$	$\frac{3}{5}$	0

2. 令 X 表示某一天新生婴儿的个数, Y 表示其中男孩的个数. 设 (X, Y) 的联合分布律:

$$P(X=i, Y=j) = \frac{(7.14)^j (6.86)^{i-j}}{j! (i-j)!} e^{-14} \quad \begin{matrix} i = 0, 1, 2, 3, \dots \\ j = 0, 1, 2, \dots, i \end{matrix}$$

(1) 求关于 X, Y 的边缘分布律; (2) 求条件分布律 $P(X=i/Y=j)$ 和 $P(Y=j/X=i)$

解: $P(X=i) = \sum_{j=0}^i \frac{(7.14)^j (6.86)^{i-j}}{j! (i-j)!} e^{-14} = \frac{e^{-14}}{i!} \sum_{j=0}^i \frac{i!}{j! (i-j)!} (7.14)^j (6.86)^{i-j}$
 $= \frac{e^{-14}}{i!} (7.14 + 6.86)^i = \frac{14^i}{i!} e^{-14} \quad i = 0, 1, 2, \dots$

$$P(Y=j) = \sum_{i=j}^{\infty} \frac{(7.14)^i (6.86)^{i-j}}{j! (i-j)!} e^{-14} = \frac{(7.14)^j}{j!} e^{-14} \sum_{i=j}^{\infty} \frac{(6.86)^{i-j}}{(i-j)!}$$

$$\text{令 } k=i-j \quad \frac{(7.14)^j}{j!} e^{-14} \sum_{k=0}^{\infty} \frac{(6.86)^k}{k!} = \frac{(7.14)^j}{j!} e^{-14} e^{6.86} = \frac{(7.14)^j}{j!} e^{-7.14} \quad j=0,1,2,\dots$$

$\therefore X \sim \pi(14), Y \sim \pi(7.14)$.

$$(2)^* P(X=i/Y=j) = \frac{P(X=i, Y=j)}{P(Y=j)} = \frac{(7.14)^j (6.86)^{i-j} e^{-14}}{\frac{(7.14)^j}{j!} e^{-7.14}}$$

$$= \frac{(6.86)^{i-j}}{(i-j)!} e^{-6.86} \quad i=j, j+1, j+2, \dots$$

$$P(Y=j/X=i) = \frac{P(X=i, Y=j)}{P(X=i)} = \frac{(7.14)^j (6.86)^{i-j} e^{-14}}{\frac{(7.14)^i}{i!} e^{-14}}$$

$$= \frac{i!}{j! (i-j)!} (7.14)^j (6.86)^{i-j} = C_i^j (7.14)^j (6.86)^{i-j} \quad j=0,1,2,\dots,i$$

3. (X, Y) 的联合密度函数为 $f(x, y)$. 分别求条件密度函数 $f_{X|Y}(x|y)$ 和 $f_{Y|X}(y|x)$

$$(1) f(x, y) = \begin{cases} 4xy e^{-(x^2+y^2)} & x>0, y>0 \\ 0 & \text{其他} \end{cases} \quad (2) f(x, y) = \begin{cases} 6xy(2-x-y) & 0<x, y<1 \\ 0 & \text{其他} \end{cases}$$

解: (1) 由习题 3-3 第 3 题 (1) 得

$$f_X(x) = \begin{cases} 2x e^{-x^2} & x>0 \\ 0 & \text{其他} \end{cases} \quad f_Y(y) = \begin{cases} 2y e^{-y^2} & y>0 \\ 0 & \text{其他} \end{cases}$$

当 $y>0$ 时:

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \begin{cases} \frac{4xy e^{-(x^2+y^2)}}{2y e^{-y^2}} & x>0 \\ 0 & \text{其他} \end{cases} = \begin{cases} 2x e^{-x^2} & x>0 \\ 0 & \text{其他} \end{cases}$$

当 $x>0$ 时:

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} = \begin{cases} \frac{4xy e^{-(x^2+y^2)}}{2x e^{-x^2}} & y>0 \\ 0 & \text{其他} \end{cases} = \begin{cases} 2y e^{-y^2} & y>0 \\ 0 & \text{其他} \end{cases}$$

(2) 由习题 3-3 第 3 题 (2) 得:

$$f_X(x) = \begin{cases} 4x-3x^2 & 0<x<1 \\ 0 & \text{其他} \end{cases} \quad f_Y(y) = \begin{cases} 4y-3y^2 & 0<y<1 \\ 0 & \text{其他} \end{cases}$$

当 $0<y<1$ 时:

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \begin{cases} \frac{6xy(2-x-y)}{4y-3y^2} & 0<x<1 \\ 0 & \text{其他} \end{cases} = \begin{cases} \frac{6x(2-x-y)}{4-3y} & 0<x<1 \\ 0 & \text{其他} \end{cases}$$

当 $0<x<1$ 时:

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} = \begin{cases} \frac{6xy(2-x-y)}{4x-3x^2} & 0<y<1 \\ 0 & \text{其他} \end{cases} = \begin{cases} \frac{6y(2-x-y)}{4-3x} & 0<y<1 \\ 0 & \text{其他} \end{cases}$$

习题 3-5

1. 判別习题 3-4 第 1 题中的 X 与 Y 是否独立? 说明理由.

解. (X, Y) 的联合分布律与边缘分布律为:

$X \backslash Y$	1	2	3	X
1	$2/30$	$6/30$	$2/30$	$10/30$
2	$6/30$	$6/30$	$3/30$	$15/30$
3	$2/30$	$3/30$	0	$5/30$
Y	$10/30$	$15/30$	$5/30$	

$$\begin{aligned} \therefore P(X=1, Y=1) &= 2/30 = \frac{1}{15} \\ \text{而 } P(X=1)P(Y=1) &= \frac{10}{30} \cdot \frac{10}{30} = \frac{1}{9} \\ \therefore P(X=1, Y=1) &\neq P(X=1)P(Y=1) \\ \therefore X, Y &\text{ 不独立. } \end{aligned}$$

2. 判別习题 3-3 第 3 题中的 X 与 Y 是否独立? 说明理由.

解: (1) $(X, Y) \sim f(x, y) = \begin{cases} 4xye^{-(x^2+y^2)} & x > 0, y > 0; \\ 0 & \text{其他.} \end{cases}$

$$X \sim f_X(x) = \begin{cases} 2xe^{-x^2} & x > 0; \\ 0 & \text{其他.} \end{cases} \quad Y \sim f_Y(y) = \begin{cases} 2ye^{-y^2} & y > 0; \\ 0 & \text{其他.} \end{cases}$$

$$\text{显然有 } f(x, y) = \begin{cases} 4xye^{-(x^2+y^2)} & x > 0, y > 0; \\ 0 & \text{其他.} \end{cases} = \begin{cases} 2xe^{-x^2} \cdot 2ye^{-y^2} & x > 0, y > 0; \\ 0 & \text{其他.} \end{cases} = f_X(x)f_Y(y)$$

$\therefore X, Y$ 相互独立.

$$(2) (X, Y) \sim f(x, y) = \begin{cases} 6xy(2-x-y) & 0 < x, y < 1; \\ 0 & \text{其他.} \end{cases}$$

$$X \sim f_X(x) = \begin{cases} 4x-3x^2 & 0 < x < 1; \\ 0 & \text{其他.} \end{cases} \quad Y \sim f_Y(y) = \begin{cases} 4y-3y^2 & 0 < y < 1; \\ 0 & \text{其他.} \end{cases}$$

\therefore 在 $0 < x, y < 1$ 时, 有

$$f(x, y) = 6xy(2-x-y) \neq (4x-3x^2)(4y-3y^2) = f_X(x)f_Y(y)$$

$\therefore X, Y$ 不相互独立.

$$(3) (X, Y) \sim f(x, y) = \begin{cases} 4.8y(2-x) & 0 \leq x \leq 1, 0 \leq y \leq x; \\ 0 & \text{其他.} \end{cases}$$

$$X \sim f_X(x) = \begin{cases} 2.4x^2(2-x) & 0 \leq x \leq 1; \\ 0 & \text{其他.} \end{cases} \quad Y \sim f_Y(y) = \begin{cases} 2.4y(3-4y+y^2) & 0 \leq y \leq 1; \\ 0 & \text{其他.} \end{cases}$$

\therefore 在 $0 \leq x \leq 1, 0 \leq y \leq x$ 时, 有

$$f(x, y) = 4.8y(2-x) \neq 2.4x^2(2-x) \cdot 2.4y(3-4y+y^2) = f_X(x) \cdot f_Y(y)$$

$\therefore X, Y$ 不相互独立.

$$(4) (X, Y) \sim f(x, y) = \frac{1}{\pi^2(1+x^2)(1+y^2)}, \quad -\infty < x, y < +\infty$$

$$X \sim f_X(x) = \frac{1}{\pi(1+x^2)}, \quad -\infty < x < +\infty; \quad Y \sim f_Y(y) = \frac{1}{\pi(1+y^2)}, \quad -\infty < y < +\infty$$

显然有 $f(x, y) = f_X(x)f_Y(y)$

3. 设 (X, Y) 的联合分布律如下表所示, 问表中 x, y 取何值时, X 与 Y 相互独立?

解: 则关于 X, Y 的边缘分布律为:

$X \backslash Y$	1	2	3
1	$1/6$	$1/9$	$1/18$
2	$1/3$	x	y

X	1	2
P	$\frac{1}{3}$	$\frac{1}{3} + x + y$

Y	1	2	3
P	$\frac{1}{2}$	$\frac{1}{9} + x$	$\frac{1}{18} + y$

当 X, Y 相互独立时, 满足:

$$P(X=1, Y=2) = P(X=1)P(Y=2), \text{ 则 } \frac{1}{9} = \frac{1}{3}(\frac{1}{9} + x) \therefore x = \frac{2}{9};$$

$$P(X=1, Y=3) = P(X=1)P(Y=3), \text{ 则 } \frac{1}{18} = \frac{1}{3}(\frac{1}{18} + y), \therefore y = \frac{1}{9}.$$

而当 $x = \frac{2}{9}, y = \frac{1}{9}$ 时, 可以验证得其余各关系也成立, 即

$$P(X=1, Y=1) = \frac{1}{6} = \frac{1}{3} \cdot \frac{1}{2} = P(X=1)P(Y=1);$$

$$P(X=2, Y=1) = \frac{1}{3} = (\frac{1}{3} + \frac{2}{9} + \frac{1}{9}) \cdot \frac{1}{2} = \frac{2}{3} \cdot \frac{1}{2} = P(X=2)P(Y=1);$$

$$P(X=2, Y=2) = \frac{2}{9} = (\frac{1}{3} + \frac{2}{9} + \frac{1}{9})(\frac{1}{9} + \frac{2}{9}) = \frac{2}{3} \cdot \frac{1}{3} = P(X=2)P(Y=2);$$

$$P(X=2, Y=3) = \frac{1}{9} = (\frac{1}{3} + \frac{2}{9} + \frac{1}{9})(\frac{1}{18} + \frac{1}{9}) = \frac{2}{3} \cdot \frac{1}{6} = P(X=2)P(Y=3);$$

故 $x = \frac{2}{9}, y = \frac{1}{9}$ 时, X, Y 相互独立.

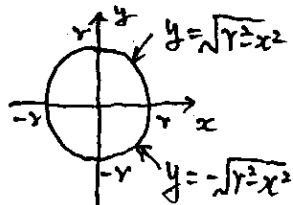
4. 设 (X, Y) 服从区域 D 上均匀分布:

(1) 若 $D: x^2 + y^2 \leq r^2$, 问 X 与 Y 是否独立?

(2) 若 $D: a < x < b, c < y < d$, 问 X 与 Y 是否相互独立?

解 (1) (X, Y) 服从区域 $D: x^2 + y^2 \leq r^2$ 上的均匀分布, $S_D = \pi r^2$.

$$\therefore (X, Y) \sim f(x, y) = \begin{cases} \frac{1}{\pi r^2}, & x^2 + y^2 \leq r^2 \\ 0, & \text{其他} \end{cases}$$



$$f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \int_{-\sqrt{r^2-x^2}}^{\sqrt{r^2-x^2}} \frac{1}{\pi r^2} dy = \frac{(y)}{\pi r^2} \Big|_{-\sqrt{r^2-x^2}}^{\sqrt{r^2-x^2}} = \frac{2}{\pi r^2} \sqrt{r^2-x^2} \quad (-r \leq x \leq r);$$

$$\therefore f_X(x) = \begin{cases} \frac{2}{\pi r^2} \sqrt{r^2-x^2}, & -r \leq x \leq r; \\ 0, & \text{其他} \end{cases}$$

同理可得

$$f_Y(y) = \begin{cases} \frac{2}{\pi r^2} \sqrt{r^2-y^2}, & -r \leq y \leq r; \\ 0, & \text{其他} \end{cases}$$

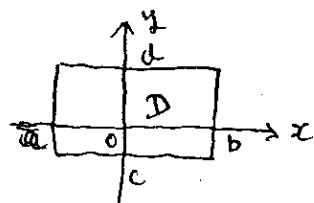
\therefore 在 $x^2 + y^2 \leq r^2$ 时, 有

$$f(x, y) = \frac{1}{\pi r^2} \neq \frac{2}{\pi r^2} \sqrt{r^2-x^2} \cdot \frac{2}{\pi r^2} \sqrt{r^2-y^2} = f_X(x) \cdot f_Y(y).$$

$\therefore X, Y$ 不独立.

(2) (X, Y) 服从区域 $D: a < x < b, c < y < d$ 上的均匀分布.

$$(X, Y) \sim f(x, y) = \begin{cases} \frac{1}{(b-a)(d-c)} & a \leq x \leq b, c \leq y \leq d, \\ 0 & \text{其他} \end{cases}$$



$$f_x(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \int_c^d \frac{dy}{(b-a)(d-c)} = \frac{1}{b-a}, \quad a \leq x \leq b,$$

$$\therefore f_x(x) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b, \\ 0 & \text{其他} \end{cases}$$

$$f_y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \int_a^b \frac{dx}{(b-a)(d-c)} = \frac{1}{d-c}, \quad c \leq y \leq d,$$

$$\therefore f_y(y) = \begin{cases} \frac{1}{d-c} & c \leq y \leq d, \\ 0 & \text{其他} \end{cases}$$

$$\text{由 } f(x, y) = \begin{cases} \frac{1}{(b-a)(d-c)} & a \leq x \leq b, c \leq y \leq d \\ 0 & \text{其他} \end{cases} = f_x(x) \cdot f_y(y),$$

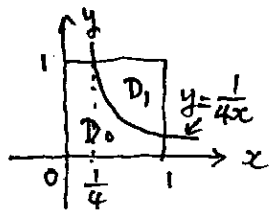
$\therefore X, Y$ 相互独立.

5. 设 $X \sim U(0, 1)$, $Y \sim U(0, 1)$, 且 X 与 Y 相互独立, 求关于 z 的一次方程 $Xz^2 + z + Y = 0$ 有实根的概率.

解: 二次方程 $Xz^2 + z + Y = 0$ 有实根的充分必要条件为它的判别式 $\Delta = 1 - 4XY \geq 0$. 故

$$P(\text{二次方程 } Xz^2 + z + Y = 0 \text{ 有实根}) = P(\Delta = 1 - 4XY \geq 0) \\ = P(XY \leq \frac{1}{4}) = 1 - P(XY > \frac{1}{4}) = 1 - \iint_{xy > \frac{1}{4}} f(x, y) dx dy$$

$$= 1 - \iint_{D_1} f_x(x) f_y(y) dx dy.$$



$$f_x(x) = \begin{cases} 1, & 0 \leq x \leq 1; \\ 0, & \text{其他} \end{cases}, \quad f_y(y) = \begin{cases} 1, & 0 \leq y \leq 1; \\ 0, & \text{其他} \end{cases}$$

$$\therefore P(XY \leq \frac{1}{4}) = 1 - \iint_{D_1} dx dy = 1 - \int_{\frac{1}{4}}^1 dx \int_{\frac{1}{4x}}^1 dy = 1 - \int_{\frac{1}{4}}^1 (1 - \frac{1}{4x}) dx \\ = 1 - (x - \frac{\ln x}{4}) \Big|_{\frac{1}{4}}^1 = 1 - [(1 - \frac{\ln 1}{4}) - (\frac{1}{4} - \frac{\ln \frac{1}{4}}{4})] \\ = 1 - (\frac{3}{4} - \frac{\ln 2}{2}) = \frac{1}{4} + \frac{1}{2} \ln 2.$$

习题 3-6

1. 设 (X, Y) 的联合分布律表为

$X \backslash Y$	-1	0	1	2
-1	$4/20$	$3/20$	$2/20$	$6/20$
1	$2/20$	0	$2/20$	$1/20$

求: (1) $Z_1 = X + Y$;

(2) $Z_2 = XY$;

(3) $Z_3 = \max(X, Y)$;

(4) $Z_4 = \min(X, Y)$ 的分布律.

解: (1) $X+Y$

P	$4/20$	$3/20$	$2/20$	$6/20$	$2/20$	0	$2/20$	$1/20$
-----	--------	--------	--------	--------	--------	---	--------	--------

$\Rightarrow X+Y$

P	$4/20$	$3/20$	$2/20$	$6/20$	$2/20$	0	$2/20$	$1/20$
-----	--------	--------	--------	--------	--------	---	--------	--------

$\Rightarrow Z_1$

P	$4/20$	$3/20$	$4/20$	$6/20$	$2/20$	$1/20$
-----	--------	--------	--------	--------	--------	--------

(2) XY

P	$4/20$	$3/20$	$2/20$	$6/20$	$2/20$	0	$2/20$	$1/20$
-----	--------	--------	--------	--------	--------	---	--------	--------

$\Rightarrow XY$

P	$4/20$	$3/20$	$2/20$	$6/20$	$2/20$	0	$2/20$	$1/20$
-----	--------	--------	--------	--------	--------	---	--------	--------

$\Rightarrow Z_2$

P	$6/20$	$4/20$	$3/20$	$6/20$	$1/20$
-----	--------	--------	--------	--------	--------

(3) $\max(X, Y)$

P	$4/20$	$3/20$	$2/20$	$6/20$
-----	--------	--------	--------	--------

$2/20$	0	$2/20$	$1/20$
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$\Rightarrow \max(X, Y)$

P	$4/20$	$3/20$	$2/20$	$6/20$	$2/20$	0	$2/20$	$1/20$
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$\Rightarrow Z_3$

P	$4/20$	$3/20$	$6/20$	$7/20$
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(4) $\min(X, Y)$

P	$4/20$	$3/20$	$2/20$	$6/20$
-----	--------	--------	--------	--------

$2/20$	0	$2/20$	$1/20$
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$\Rightarrow \min(X, Y)$

P	$4/20$	$3/20$	$2/20$	$6/20$	$2/20$	0	$2/20$	$1/20$
-----	--------	--------	--------	--------	--------	---	--------	--------

$\Rightarrow Z_4$

P	$4/20$	$3/20$	$2/20$	$6/20$	$2/20$	0	$2/20$	$1/20$
-----	--------	--------	--------	--------	--------	---	--------	--------

2. 设 $X \sim B(n_1, p)$, $Y \sim B(n_2, p)$ 且 X 与 Y 相互独立, 证明: $Z = X + Y \sim B(n_1 + n_2, p)$

证: 方法一: $\therefore P(X=i) = C_{n_1}^i p^i (1-p)^{n_1-i}$, $i = 0, 1, 2, \dots, n_1$;
 $P(Y=j) = C_{n_2}^j p^j (1-p)^{n_2-j}$, $j = 0, 1, 2, \dots, n_2$.

$$\begin{aligned} \therefore P(X+Y=k) &= \sum_{i+j=k} P(X=i, Y=j) = \sum_{i+j=k} P(X=i) P(Y=j) \\ &= \sum_{i+j=k} C_{n_1}^i p^i (1-p)^{n_1-i} \cdot C_{n_2}^j p^j (1-p)^{n_2-j} = \sum_{i+j=k} C_{n_1}^i C_{n_2}^j p^k (1-p)^{n_1+n_2-k} \\ &= C_{n_1+n_2}^k p^k (1-p)^{n_1+n_2-k}. \text{ 其中 } C_{n_1+n_2}^k = \sum_{i=0}^k C_{n_1}^i C_{n_2}^{k-i} = \sum_{i+j=k} C_{n_1}^i C_{n_2}^j \end{aligned}$$

$\therefore X+Y \sim B(n_1+n_2, p)$.

方法二: 用构造法. 设 $\frac{X_i}{p} \mid \begin{matrix} 0 & 1 \\ 1-p & p \end{matrix}$, $i = 1, 2, \dots, n_1, n_1+1, \dots, n_1+n_2$,

且相互独立. 记 $X = \sum_{i=1}^{n_1} X_i$, $Y = \sum_{j=n_1+1}^{n_1+n_2} X_j$, 则 $X \sim B(n_1, p)$, $Y \sim B(n_2, p)$ 相互独立.
 则 $X+Y = \sum_{i=1}^{n_1} X_i + \sum_{j=n_1+1}^{n_1+n_2} X_j = \sum_{k=1}^{n_1+n_2} X_k \sim B(n_1+n_2, p)$.

3. 设 $X \sim \pi(\lambda_1)$, $Y \sim \pi(\lambda_2)$, 且 X 与 Y 相互独立, 证明: $Z = X+Y \sim \pi(\lambda_1+\lambda_2)$

证: $P(X=i) = \frac{\lambda_1^i}{i!} e^{-\lambda_1}$; $P(Y=j) = \frac{\lambda_2^j}{j!} e^{-\lambda_2}$, $i, j = 0, 1, 2, \dots$

$$\begin{aligned} \therefore P(X+Y=k) &= \sum_{i+j=k} P(X=i, Y=j) = \sum_{i=0}^k P(X=i, Y=k-i) \\ &= \sum_{i=0}^k P(X=i) P(Y=k-i) = \sum_{i=0}^k \frac{\lambda_1^i}{i!} e^{-\lambda_1} \cdot \frac{\lambda_2^{k-i}}{(k-i)!} e^{-\lambda_2} \\ &= \sum_{i=0}^k \frac{k!}{i!(k-i)!} \cdot \lambda_1^i \lambda_2^{k-i} \cdot \frac{1}{k!} e^{-(\lambda_1+\lambda_2)} = \frac{1}{k!} e^{-(\lambda_1+\lambda_2)} \sum_{i=0}^k C_k^i \lambda_1^i \lambda_2^{k-i} \\ &= \frac{(\lambda_1+\lambda_2)^k}{k!} e^{-(\lambda_1+\lambda_2)}, \quad k = 0, 1, 2, \dots \end{aligned}$$

$\therefore X+Y \sim \pi(\lambda_1+\lambda_2)$.

4. 设 X, Y 的密度函数分别为

$$f_X(x) = \begin{cases} 3e^{-3x}, & x > 0; \\ 0, & \text{其他} \end{cases}; \quad f_Y(y) = \begin{cases} 2e^{-2y}, & y > 0; \\ 0, & \text{其他} \end{cases}$$

且 X, Y 相互独立. 求 $Z = X+Y$ 的分布.

解: $f_{X+Y}(u) = \int_{-\infty}^{+\infty} f_X(x) f_Y(u-x) dx, \quad -\infty < u < +\infty$

由 $\begin{cases} x > 0 \\ u-x > 0 \end{cases} \Rightarrow \begin{cases} x > 0 \\ u > x \end{cases}$ 即 $u > x > 0$ 时 $f_X(x) f_Y(u-x) \neq 0$

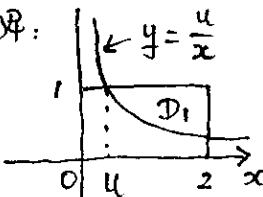
\therefore 在 $u > 0$ 时,

$$\begin{aligned} f_{X+Y}(u) &= \int_0^u 3e^{-3x} \cdot 2e^{-2(u-x)} dx = \int_0^u 6e^{-2u} \cdot e^{-x} dx = 6e^{-2u} (-e^{-x}) \Big|_0^u \\ &= 6e^{-2u} (1 - e^{-u}). \end{aligned}$$

故 $Z = X+Y$ 的分布为密度函数 $f(u) = \begin{cases} 6e^{-2u}(1-e^{-u}), & u > 0; \\ 0, & \text{其他} \end{cases}$

5. 设 (X, Y) 服从矩形 $G = \{(x, y) | 0 \leq x \leq 2, 0 \leq y \leq 1\}$ 上的均匀分布.

求边长为 X 和 Y 的矩形面积 S 的分布.

解:  $(X, Y) \sim f(x, y) = \begin{cases} \frac{1}{2}, & 0 \leq x \leq 2, 0 \leq y \leq 1; \\ 0, & \text{其他.} \end{cases}$

设 $S = XY$, 其分布函数为 $F_S(u)$.

当 $u \leq 0$ 时, $F_S(u) = P\{S \leq u\} = P(\emptyset) = 0$.

$$\begin{aligned} \text{当 } 0 < u < 2 \text{ 时, } F_S(u) &= P\{S \leq u\} = P(XY \leq u) = 1 - P(XY > u) \\ &= 1 - \iint_{xy > u} f(x, y) dx dy = 1 - \iint_{D_1} \frac{1}{2} dx dy = 1 - \int_u^2 dx \int_{u/x}^1 \frac{1}{2} dy = 1 - \int_u^2 \frac{1}{2} (1 - \frac{u}{x}) dx \\ &= \frac{1}{2} (x - u \ln x) \Big|_u^2 = 1 - \frac{1}{2} [(2 - u \ln 2) - (u - u \ln u)] \\ &= \frac{u}{2} (1 + \ln 2 - \ln u). \end{aligned}$$

$$\text{当 } u \geq 2 \text{ 时, } F_S(u) = P(XY \leq u) = \iint_D \frac{1}{2} dx dy = \int_0^2 dx \int_0^1 \frac{1}{2} dy = 1.$$

$$\therefore F_S(u) = \begin{cases} 0, & u \leq 0; \\ \frac{u}{2} (1 + \ln 2 - \ln u), & 0 < u < 2; \\ 1, & u \geq 2. \end{cases}$$

$$\therefore f_S(u) = \begin{cases} \frac{1}{2} (\ln 2 - \ln u), & 0 < u < 2; \\ 0, & \text{其他.} \end{cases}$$

6. 已知 $X \sim N(-3, 1)$, $Y \sim N(2, 1)$, 且 X 与 Y 相互独立. 求 $Z = X - 2Y$ 的分布.

解: 由定理 3.3 知: 相互独立的正态变量的线性组合服从正态分布.

即 $X_k \sim N(\mu_k, \sigma_k^2)$, $k=1, \dots, n$ 相互独立. 则对不全为零的常数 $a_k, k=1, \dots, n$ 有 $\sum_{k=1}^n a_k X_k \sim N(\sum_{k=1}^n a_k \mu_k, \sum_{k=1}^n a_k^2 \sigma_k^2)$.

故 $Z = X - 2Y \sim N[(-3) - 2 \times 2, 1 + 4 \times 1]$ 即 $Z \sim N(-7, 5)$.

(注) 根据第四章数字特征知: $X_k \sim N(\mu_k, \sigma_k^2)$, $k=1, \dots, n$ 相互独立.

则对不全为零的常数 $a_k, k=1, 2, \dots, n$. 有 $\sum_{k=1}^n a_k X_k \sim N(\mu, \sigma^2)$.

其中 $\mu = E(\sum_{k=1}^n a_k X_k)$, $\sigma^2 = D(\sum_{k=1}^n a_k X_k)$.

利用数学期望与方差的运算性质: 由 $a_k X_k, k=1, \dots, n$ 相互独立.

$$\text{则 } \mu = E(\sum_{k=1}^n a_k X_k) = \sum_{k=1}^n a_k E(X_k) = \sum_{k=1}^n a_k \mu_k;$$

$$\sigma^2 = D(\sum_{k=1}^n a_k X_k) = \sum_{k=1}^n a_k^2 D(X_k) = \sum_{k=1}^n a_k^2 \sigma_k^2.$$

复习题 3

1. 设 A, B 是两个随机事件, 且 $P(A) = \frac{1}{4}, P(B) = \frac{1}{6}, P(AB) = \frac{1}{12}$. 令

$X = \begin{cases} 1, & A \text{ 发生}; \\ 0, & A \text{ 不发生}. \end{cases} Y = \begin{cases} 1, & B \text{ 发生}; \\ 0, & B \text{ 不发生}. \end{cases}$ 求: 二维随机变量 (X, Y) 的联合分布律.

解: $P(X=1) = P(A) = \frac{1}{4}, P(X=0) = P(\bar{A}) = \frac{3}{4};$

$P(Y=1) = P(B) = \frac{1}{6}, P(Y=0) = P(\bar{B}) = \frac{5}{6};$

$P(X=0, Y=0) = P(\bar{A} \cdot \bar{B}) = 1 - P(A \cup B) = 1 - (P(A) + P(B) - P(AB))$
 $= 1 - (\frac{1}{4} + \frac{1}{6} - \frac{1}{12}) = \frac{2}{3};$

$P(X=0, Y=1) = P(\bar{A} \cdot B) = P(B) - P(AB) = \frac{1}{6} - \frac{1}{12} = \frac{1}{12};$

$P(X=1, Y=0) = P(A \cdot \bar{B}) = P(A) - P(AB) = \frac{1}{4} - \frac{1}{12} = \frac{1}{6};$

$P(X=1, Y=1) = P(A \cdot B) = \frac{1}{12}.$

$$\therefore \begin{array}{c|cc} X \backslash Y & 0 & 1 \\ \hline 0 & 2/3 & 1/12 \\ 1 & 1/6 & 1/12 \end{array}$$

2. 设随机变量 Y 服从参数 $\lambda=1$ 的指数分布, 随机变量 $X_k = \begin{cases} 0, & Y \leq k; \\ 1, & Y > k. \end{cases}$
 $(k=1, 2)$. 求 X_1, X_2 的联合分布律.

解: Y 的分布函数 $F_Y(y) = \begin{cases} 1 - e^{-y} & y > 0; \\ 0 & y \leq 0. \end{cases}$

$P(X_1=0, X_2=0) = P(Y \leq 1, Y \leq 2) = P(Y \leq 1) = F_Y(1) = 1 - e^{-1};$

$P(X_1=0, X_2=1) = P(Y \leq 1, Y > 2) = P(\emptyset) = 0;$

$P(X_1=1, X_2=0) = P(Y > 1, Y \leq 2) = P(1 < Y \leq 2) = F_Y(2) - F_Y(1)$
 $= (1 - e^{-2}) - (1 - e^{-1}) = e^{-1} - e^{-2};$

$P(X_1=1, X_2=1) = P(Y > 1, Y > 2) = P(Y > 2) = 1 - F_Y(2) = 1 - (1 - e^{-2}) = e^{-2}.$

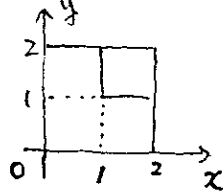
$$\therefore \begin{array}{c|cc} X_1 \backslash X_2 & 0 & 1 \\ \hline 0 & 1 - e^{-1} & 0 \\ 1 & e^{-1} - e^{-2} & e^{-2} \end{array}$$

3. 设 (X, Y) 的联合分布律 $\begin{array}{c|cc} X \backslash Y & 0 & 1 \\ \hline 0 & 1/25 & 4/25 \\ 1 & 4/25 & 16/25 \end{array}$, 求 $F(2, \frac{1}{2})$.

解: $F(2, \frac{1}{2}) = P(X \leq 2, Y \leq \frac{1}{2}) = P(X=0, Y=0) + P(X=1, Y=0)$
 $= \frac{1}{25} + \frac{4}{25} = \frac{1}{5}.$

4. 设 (X, Y) 服从区域 $G = \{(x, y) | 0 < x, y < 2\}$ 上的均匀分布, 求 X, Y 至少有一个小于 1 的概率.

解: $(X, Y) \sim f(x, y) = \begin{cases} \frac{1}{4}, & 0 < x < 2, 0 < y < 2. \end{cases}$



$$P[(X < 1) \cup (Y < 1)] = P(X < 1) + P(Y < 1) - P(X < 1, Y < 1).$$

$$\text{由 } P(X < 1) = \iint_{x < 1} f(x, y) dx dy = \int_0^1 dx \int_0^2 \frac{1}{4} dy = \int_0^1 \frac{1}{2} dx = \frac{1}{2};$$

$$P(Y < 1) = \iint_{y < 1} f(x, y) dx dy = \int_0^1 dy \int_0^2 \frac{1}{4} dx = \int_0^1 \frac{1}{4} dy = \frac{1}{4};$$

$$P(X < 1, Y < 1) = \iint_{\substack{x < 1 \\ y < 1}} f(x, y) dx dy = \int_0^1 dx \int_0^1 \frac{1}{4} dy = \int_0^1 \frac{1}{4} dx = \frac{1}{4}.$$

$$\therefore P[(X < 1) \cup (Y < 1)] = \frac{1}{2} + \frac{1}{4} - \frac{1}{4} = \frac{3}{4}.$$

5. 设 \$(X, Y)\$ 的联合密度函数 \$f(x, y) = \begin{cases} 4xy, & 0 \leq x, y \leq 1; \\ 0, & \text{其他.} \end{cases}\$ 求 \$(X, Y)\$ 的联合分布函数.

解: 在 \$x < 0, y < 0\$ 时, \$F(x, y) = \int_{-\infty}^x \int_{-\infty}^y 0 dx dy = 0\$.

在 \$0 \leq x, y \leq 1\$ 时, \$F(x, y) = \int_{-\infty}^x \int_{-\infty}^y f(x, y) dx dy = \int_0^x \int_0^y 4xy dx dy = \int_0^y 2x^2 dy = \frac{2}{3} y^3 = x^2 y^2\$.

在 \$0 \leq x \leq 1, y > 1\$ 时, \$F(x, y) = \int_{-\infty}^x \int_{-\infty}^y f(x, y) dx dy = \int_0^x \int_1^y 4xy dx dy = \int_1^y 2x^2 dy = 2x^2 \int_1^y dy = 2x^2(y - 1) = x^2 y^2 - x^2\$.

在 \$x > 1, 0 \leq y \leq 1\$ 时, \$F(x, y) = \int_{-\infty}^x \int_{-\infty}^y f(x, y) dx dy = \int_0^y \int_1^x 4xy dx dy = \int_0^y 2x^2 dy = 2x^2 \int_0^y dy = 2x^2 y = x^2 y^2 - x^2\$.

在 \$x > 1, y > 1\$ 时, \$F(x, y) = \int_{-\infty}^x \int_{-\infty}^y f(x, y) dx dy = \int_0^1 \int_0^1 4xy dx dy + \int_1^x \int_1^y 4xy dx dy = \frac{2}{3} + \int_1^x 2y^2 dx = \frac{2}{3} + 2y^2(x - 1) = x^2 y^2 - x^2 - y^2 + 1\$.

$$\therefore F(x, y) = \begin{cases} 0; & x < 0, y < 0; \\ x^2 y^2; & 0 \leq x, y \leq 1; \\ x^2; & 0 \leq x \leq 1, y > 1; \\ y^2; & x > 1, 0 \leq y \leq 1; \\ 1; & x > 1, y > 1. \end{cases}$$

6. 设二维随机变量 \$(X, Y)\$ 在区域 \$D: 0 < x < 1, |y| < x\$ 内服从均匀分布. 求 (1) 关于 \$X, Y\$ 的边缘密度函数; (2) 条件密度函数 \$f_{Y|X}(y|x)\$ 和 \$f_{X|Y}(x|y)\$; (3) \$P\{-1 < X < \frac{3}{4} / Y = \frac{1}{2}\}, P\{-1 < Y < \frac{1}{2} / X = \frac{1}{2}\}\$

解: \$S_D = \iint_D dx dy = \int_0^1 dx \int_{-x}^x dy = \int_0^1 2x dx = (x^2)|_0^1 = 1\$.

$$\therefore (X, Y) \sim f(x, y) = \begin{cases} 1; & 0 < x < 1, -1 < y < x; \\ 0. & \text{其他} \end{cases}$$

$$(1) f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_{-x}^x 1 dy, & 0 < x < 1; \\ 0. & \text{其他} \end{cases} = \begin{cases} 2x; & 0 < x < 1; \\ 0. & \text{其他} \end{cases}$$

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \int_{-y}^1 1 dx; & -1 < y \leq 0; \\ \int_y^1 1 dx; & 0 < y < 1 \\ 0. & \text{其他} \end{cases} = \begin{cases} 1+y; & -1 < y \leq 0; \\ 1-y; & 0 < y < 1; \\ 0. & \text{其他} \end{cases}$$

(注: $\because |y| < x$, \therefore 当 $-x < y$, 即 $x > -y$ 时, 由 $0 < x < 1$, 故有 $-y < x < 1$, 而 $-1 < y \leq 0$; 当 $y < x$ 时, 有 $y < x < 1$ 而 $0 < y < 1$.)

(2) 当 $|y| < 1$ 时.

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \begin{cases} \frac{1}{1+y}; & 0 < x < 1, -x < y \leq 0; \\ \frac{1}{1-y}; & 0 < x < 1, 0 < y < x \\ 0. & \text{其他} \end{cases} = \begin{cases} \frac{1}{1-|y|}; & |y| < x < 1; \\ 0. & \text{其他} \end{cases}$$

当 $0 < x < 1$ 时:

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} = \begin{cases} \frac{1}{2x}, & |y| < x; \\ 0. & \text{其他} \end{cases}$$

(3) 注意到条件密度函数是在某条件下的随机变量的密度函数, 故有公式:

$$P(a < X \leq b / Y = c) = \int_a^b f_{X|Y}(x|c) dx.$$

$$P(c < Y \leq d / X = e) = \int_c^d f_{Y|X}(y|e) dy$$

$$\text{则 } P(-1 < X < \frac{3}{4} / Y = \frac{1}{2}) = \int_{-1}^{\frac{3}{4}} f_{X|Y}(x|\frac{1}{2}) dx = \int_{-\frac{1}{2}}^{\frac{3}{4}} \frac{1}{1-\frac{1}{2}} dx = \int_{-\frac{1}{2}}^{\frac{3}{4}} 2 dx = \frac{1}{2}$$

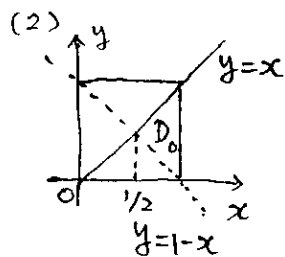
$$P(-1 < Y < \frac{1}{2} / X = \frac{1}{2}) = \int_{-1}^{\frac{1}{2}} f_{Y|X}(y|\frac{1}{2}) dy = \int_{-\frac{1}{2}}^{\frac{1}{2}} \frac{1}{2 \times \frac{1}{2}} dy = \int_{-\frac{1}{2}}^{\frac{1}{2}} dy = 1.$$

$$\text{其中 } f_{X|Y}(x|\frac{1}{2}) = \begin{cases} 2, & \frac{1}{2} < x < 1; \\ 0, & \text{其他} \end{cases}; \quad f_{Y|X}(y|\frac{1}{2}) = \begin{cases} 1, & |y| < \frac{1}{2}; \\ 0, & \text{其他} \end{cases}$$

7. 设随机变量 X 服从区间 $(0, 1)$ 上的均匀分布, 在 $X = x$ ($0 < x < 1$) 的条件下, 随机变量 Y 服从区间 $(0, x)$ 上的均匀分布. 求 (1) (X, Y) 的联合密度函数; (2) $P(X+Y > 1)$.

$$\text{解: } f_X(x) = \begin{cases} 1; & 0 < x < 1; \\ 0. & \text{其他} \end{cases}; \quad f_{Y|X}(y|x) = \begin{cases} \frac{1}{x}, & 0 < y < x; \\ 0. & \text{其他} \end{cases}$$

$$(1) f(x, y) = f_x(x) \cdot f_{y|x}(y|x) = \begin{cases} \frac{1}{x}, & 0 < y < x < 1; \\ 0, & \text{其他} \end{cases}$$



$$\begin{aligned} P(X+Y > 1) &= \iint_{x+y>1} f(x, y) dx dy = \iint_{D_0} \frac{1}{x} dx dy = \int_{1/2}^1 dx \int_{1-x}^x \frac{1}{x} dy \\ &= \int_{1/2}^1 \left(\frac{y}{x} \right) \Big|_{1-x}^x dx = \int_{1/2}^1 \frac{2x-1}{x} dx = \int_{1/2}^1 \left(2 - \frac{1}{x} \right) dx \\ &= (2x - \ln x) \Big|_{1/2}^1 = (2 - \ln 1) - (2 \times \frac{1}{2} - \ln \frac{1}{2}) = 1 + \ln \frac{1}{2} \\ &= 1 - \ln 2. \end{aligned}$$

8. 设 X 与 Y 相互独立, 有相同的分布律 $\begin{matrix} X & -1 & 1 \\ P & 1/2 & 1/2 \end{matrix}$, 则下列正确的是 ().

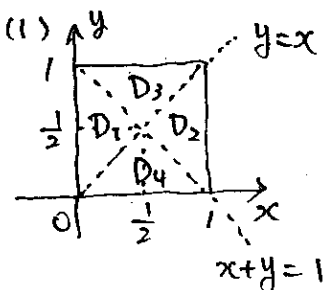
(A) $X=Y$, (B) $P(X=Y)=1$, (C) $P(X=Y)=\frac{1}{2}$, (D) $P(X=Y)=\frac{1}{4}$.

解: $P(X=Y) = P[(X=-1, Y=-1) \cup (X=1, Y=1)] = P(X=-1, Y=-1) + P(X=1, Y=1)$
 $= P(X=-1)P(Y=-1) + P(X=1)P(Y=1)$
 $= \frac{1}{2} \cdot \frac{1}{2} + \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{2}$. 故是 (C).

9. 设二维随机变量 (X, Y) 在区域 $D: 0 < x < 1, 0 < y < 1$ 上服从均匀分布, 令 $U = \begin{cases} 1, & X > Y; \\ 0, & X \leq Y. \end{cases}$ $V = \begin{cases} 1, & X+Y > 1; \\ 0, & X+Y \leq 1. \end{cases}$ 求 (1) (U, V) 的联合分布律;

(2) 关于 U 和 V 的边缘分布律; (3) 问 U, V 是否相互独立, 为什么?

解: $(X, Y) \sim f(x, y) = \begin{cases} 1, & 0 < x < 1, 0 < y < 1; \\ 0, & \text{其他} \end{cases}$



$$\begin{aligned} P(U=0, V=0) &= P[(X \leq Y) \cap (X+Y \leq 1)] \\ &= \iint_{D_1} 1 dx dy = \int_0^{1/2} dx \int_x^{1-x} dy = \int_0^{1/2} (1-2x) dx = \frac{1}{4}; \end{aligned}$$

$$\begin{aligned} P(U=0, V=1) &= P[(X \leq Y) \cap (X+Y > 1)] \\ &= \iint_{D_2} 1 dx dy = \int_{1/2}^1 dx \int_{1-x}^x dy = \int_{1/2}^1 (2x-1) dx = \frac{1}{4}; \end{aligned}$$

$$\begin{aligned} P(U=1, V=0) &= P[(X > Y) \cap (X+Y \leq 1)] = \iint_{D_3} 1 dx dy = \int_0^{1/2} dy \int_0^y dx \\ &= \int_0^{1/2} (1-2y) dy = \frac{1}{4}; \end{aligned}$$

$$\begin{aligned} P(U=1, V=1) &= P[(X > Y) \cap (X+Y > 1)] = \iint_{D_4} 1 dx dy = \int_{1/2}^1 dy \int_{1-y}^y dx \\ &= \int_{1/2}^1 (2y-1) dy = \frac{1}{4}. \end{aligned}$$

$U \backslash V$	0	1
0	$1/4$	$1/4$
1	$1/4$	$1/4$

故 (U, V) 的联合分布律为

(2) 关于 U 和 V 的边缘分布律依次为:

U	0	1
P	$\frac{1}{2}$	$\frac{1}{2}$

和

V	0	1
P	$\frac{1}{2}$	$\frac{1}{2}$

(3) $\therefore P(U=0, V=0) = P(U=0)P(V=0) = \frac{1}{4}$;

$P(U=0, V=1) = P(U=0)P(V=1) = \frac{1}{4}$;

$P(U=1, V=0) = P(U=1)P(V=0) = \frac{1}{4}$ $\therefore U$ 和 V 相互独立.

$P(U=1, V=1) = P(U=1)P(V=1) = \frac{1}{4}$

10. 设 ξ, η 是相互独立且服从同一分布的两个随机变量, 已知 ξ 的分布律为 $P(\xi=i) = \frac{1}{3}, i=1, 2, 3$, 又设 $X = \max(\xi, \eta)$, 求 X 的分布.

解: $P(\xi=i) = \frac{1}{3}, i=1, 2, 3; P(\eta=j) = \frac{1}{3}, j=1, 2, 3$ 相互独立.

$P(X=1) = P(\max(\xi, \eta)=1) = P(\xi=1, \eta=1) = P(\xi=1)P(\eta=1) = \frac{1}{3} \cdot \frac{1}{3} = \frac{1}{9}$;

$P(X=2) = P(\max(\xi, \eta)=2) = P[(\xi=1, \eta=2) \cup (\xi=2, \eta=1) \cup (\xi=2, \eta=2)]$
 $= P(\xi=1, \eta=2) + P(\xi=2, \eta=1) + P(\xi=2, \eta=2)$
 $= P(\xi=1)P(\eta=2) + P(\xi=2)P(\eta=1) + P(\xi=2)P(\eta=2) = \frac{1}{9} + \frac{1}{9} + \frac{1}{9} = \frac{1}{3}$;

$P(X=3) = P(\max(\xi, \eta)=3) = P[(\xi=1, \eta=3) \cup (\xi=2, \eta=3) \cup (\xi=3, \eta=3) \cup (\xi=3, \eta=1) \cup (\xi=3, \eta=2)]$
 $= P(\xi=1, \eta=3) + P(\xi=2, \eta=3) + P(\xi=3, \eta=3) + P(\xi=3, \eta=1) + P(\xi=3, \eta=2)$
 $= P(\xi=1)P(\eta=3) + P(\xi=2)P(\eta=3) + P(\xi=3)P(\eta=3) + P(\xi=3)P(\eta=1) + P(\xi=3)P(\eta=2)$
 $= \frac{1}{9} + \frac{1}{9} + \frac{1}{9} + \frac{1}{9} + \frac{1}{9} = \frac{5}{9}$.

$\therefore X$ 的分布律为

X	1	2	3
P	$\frac{1}{9}$	$\frac{1}{3}$	$\frac{5}{9}$

11. 设 X, Y 的密度函数分别为:

$f_X(x) = \begin{cases} 1 & 0 \leq x \leq 1 \\ 0 & \text{其他} \end{cases}, f_Y(y) = \begin{cases} e^{-y} & y > 0 \\ 0 & \text{其他} \end{cases}$; 且 X, Y 相互独立.

求 $Z = 2X + Y$ 的分布.

解: $f_Z(u) = \int_{-\infty}^{+\infty} f_X(x) f_Y(u-2x) dx, -\infty < u < +\infty$, (见注)

由 $\begin{cases} 0 \leq x \leq 1 \\ u-2x > 0 \end{cases}$ 得 $\begin{cases} 0 \leq x \leq 1 \\ \frac{1}{2}u > x \end{cases}$ 即 $\begin{cases} 0 < x < \frac{u}{2} \\ 0 < u \leq 2 \end{cases}$ 及 $\begin{cases} 0 < x < 1 \\ u > 2 \end{cases}$.

故当 $u \leq 0$ 时, $f_Z(u) = \int_{-\infty}^{+\infty} f_X(x) f_Y(u-2x) dx = 0$.

当 $0 < u \leq 2$ 时, $f_Z(u) = \int_{\frac{1}{2}u}^{+\infty} f_X(x) f_Y(u-2x) dx = \int_{\frac{1}{2}u}^{u/2} 1 \cdot e^{-(u-2x)} dx = \int_{\frac{1}{2}u}^{u/2} e^{-u+2x} dx = \int_{\frac{1}{2}u}^{u/2} e^{-u} e^{2x} dx = e^{-u} \int_{\frac{1}{2}u}^{u/2} e^{2x} dx = e^{-u} \left[\frac{1}{2} e^{2x} \right]_{\frac{1}{2}u}^{u/2} = e^{-u} \left(\frac{1}{2} e^u - \frac{1}{2} e^{\frac{1}{2}u} \right) = \frac{1}{2} (1 - e^{-\frac{1}{2}u})$.

$$= \left(\frac{1}{2} e^{2x-u} \right) \Big|_0^{+\infty} = \frac{1}{2} (1 - e^{-u}).$$

$$\begin{aligned} \text{当 } u > 2 \text{ 时, } f_2(u) &= \int_{-\infty}^{+\infty} f_x(x) \cdot f_Y(u-2x) dx = \int_0^1 1 \cdot e^{-(u-2x)} dx \\ &= \left(\frac{1}{2} e^{2x-u} \right) \Big|_0^1 = \frac{1}{2} (e^{2-u} - e^{-u}) = \frac{1}{2} e^{-u} (e^2 - 1). \end{aligned}$$

$$\therefore f_2(u) = \begin{cases} 0, & u \leq 0; \\ \frac{1}{2} (1 - e^{-u}), & 0 < u \leq 2; \\ \frac{1}{2} e^{-u} (e^2 - 1), & u > 2. \end{cases}$$

(注) ① $F_2(u) = P\{2X+Y \leq u\} = \iint f(x,y) dx dy = \iint f_x(x) \cdot f_Y(y) dx dy$
 $= \int_{-\infty}^{+\infty} dx \int_{-\infty}^{u-2x} f_x(x) f_Y(y) dy = \int_{-\infty}^{+\infty} f_x(x) \cdot F_Y(u-2x) dx$ (其中 $\int_{-\infty}^{u-2x} f_Y(y) dy = F_Y(u-2x)$)
 $\therefore f_2(u) = \frac{dF_2(u)}{du} = \int_{-\infty}^{+\infty} f_x(x) \cdot f_Y(u-2x) dx \quad (-\infty < u < +\infty)$

② 更一般情形: $Z = aX + bY$ 的分布函数与密度函数 (X, Y 相互独立)
 为 $F_Z(u) = \int_{-\infty}^{+\infty} f_x(x) F_Y(\frac{1}{b}(u-ax)) dx$, $f_Z(u) = \frac{1}{b} \int_{-\infty}^{+\infty} f_x(x) f_Y(\frac{1}{b}(u-ax)) dx$.

③ 亦可以先求出 $W = 2X$ 的密度函数 $f_W(w) = \begin{cases} \frac{1}{2}, & 0 \leq w \leq 2 \\ 0, & \text{其它} \end{cases}$

显然 W 与 Y 相互独立. 由 $Z = W + Y$, 故有

$$f_2(u) = \int_{-\infty}^{+\infty} f_W(w) f_Y(u-w) dw \quad \text{由 } \begin{cases} 0 \leq w \leq 2 \\ u-w > 0 \end{cases} \text{ 得 } \begin{cases} 0 < w < u \\ 0 < u < 2 \end{cases} \text{ 及 } \begin{cases} 0 < w < 2 \\ u \geq 2 \end{cases}.$$

可作出 $f_2(u)$, 结果相同.

12. 设随机变量 X 与 Y 相互独立, 其中 X 的分布律 $\begin{array}{c|cc} X & 1 & 2 \\ \hline P & 0.3 & 0.7 \end{array}$, Y 的密度函数为 $f_Y(y)$. 求随机变量 $Z = X+Y$ 的密度函数 $g(z)$.

$$\begin{aligned} \text{解: } F_2(u) &= P(Z \leq u) = P(X+Y \leq u) = P[(X=1, Y \leq u-1) \cup (X=2, Y \leq u-2)] \\ &= P(X=1, Y \leq u-1) + P(X=2, Y \leq u-2) \\ &= P(X=1) \cdot P(Y \leq u-1) + P(X=2) P(Y \leq u-2) \\ &= 0.3 F_Y(u-1) + 0.7 F_Y(u-2) \end{aligned}$$

$$\therefore f_2(u) = \frac{dF_2(u)}{du} = \frac{d}{du} [0.3 F_Y(u-1) + 0.7 F_Y(u-2)] = 0.3 f_Y(u-1) + 0.7 f_Y(u-2)$$

$$\text{即 } g(z) = f_2(z) = 0.3 f_Y(z-1) + 0.7 f_Y(z-2).$$

1. 略.

2. 某批产品共 24 件, 其中有次品 4 件, 其余均为合格品, 求从这批产品中任意取出的 5 件里所含次品件数的数学期望.

解: 设 $X =$ “从 24 件中任取 5 件里所含次品数”.

$$\text{则 } P\{X=K\} = \frac{C_4^K C_{20}^{5-K}}{C_{24}^5}, \quad K=0, 1, 2, 3, 4.$$

$$\therefore E(X) = \sum_{K=0}^4 K \cdot P(X=K) = \sum_{K=0}^4 K \cdot \frac{C_4^K C_{20}^{5-K}}{C_{24}^5} = 0.8334$$

注: ①

X	0	1	2	3	4
P	0.36477	0.45596	0.16093	0.01788	0.00047

② 本题的分布为超几何分布: 在 N 件产品中有次品 m 件, 从中任取 n 件, 则 n 件中所含次品数 X 的分布律为: $P(X=K) = C_m^K C_{N-m}^{n-K} / C_N^n$, $K=0, 1, 2, \dots, \min(n, m)$

由组合数 $C_N^n = \sum_{K=0}^n C_m^K C_{N-m}^{n-K}$. 因此 (设 $n \leq m$)

$$E(X) = \sum_{K=0}^n K \cdot C_m^K C_{N-m}^{n-K} / C_N^n = \sum_{K=0}^n K \cdot \frac{\frac{m}{K} \cdot C_{m-1}^{K-1} \cdot C_{N-m}^{n-K}}{C_N^n} = \frac{nm}{N} \cdot \sum_{i=0}^{n-1} \frac{C_{m-1}^i C_{N-m}^{n-1-i}}{C_{N-1}^{n-1}}$$

$$= \frac{nm}{N}. \quad \text{其中 } \sum_{i=0}^{n-1} C_{m-1}^i C_{N-m}^{n-1-i} / C_{N-1}^{n-1} = 1.$$

所以, 本题的确切值 $E(X) = \frac{5 \times 4}{24} = \frac{5}{6} = 0.8333$. 解中的 0.8334 实际上由 ① 的计算值计算所得, 有一定的计算误差.

3. 设随机变量 X 取非负整数 n 的概率为 $P(X=n) = \frac{ab^n}{n!}$, 已知 $E(X) = \lambda$. 试确定 a, b 之值.

解: $\because \sum_{n=0}^{\infty} P(X=n) = 1, \therefore \sum_{n=0}^{\infty} \frac{ab^n}{n!} = a \sum_{n=0}^{\infty} \frac{b^n}{n!} = a e^b = 1$

又 $\because \sum_{n=0}^{\infty} n P(X=n) = E(X) = \lambda, \therefore \sum_{n=0}^{\infty} n \frac{ab^n}{n!} = ab \sum_{n=1}^{\infty} \frac{b^{n-1}}{(n-1)!} = ab e^b = \lambda$

$\therefore b = \lambda, a = e^{-\lambda}$. 有 $P(X=n) = \frac{\lambda^n}{n!} e^{-\lambda}$, 即 $X \sim \pi(\lambda)$.

4. 设随机变量 X 的密度函数为

$$f(x) = \begin{cases} \frac{6}{5} x(x+1), & 0 < x < 1, \\ 0, & \text{其他.} \end{cases} \quad \text{求 } E(X).$$

解: $E(X) = \int_{-\infty}^{+\infty} x f(x) dx = \int_0^1 x \cdot \frac{6}{5} x(x+1) dx = \int_0^1 \frac{6}{5} (x^3 + x^2) dx$

$$= \frac{6}{5} \left(\frac{x^4}{4} + \frac{x^3}{3} \right) \Big|_0^1 = \frac{6}{5} \left(\frac{1}{4} + \frac{1}{3} \right) = \frac{7}{10}.$$

习题 4-2

1. 设随机变量 X 的分布律为

X	-2	-1	0	1
P	0.1	0.4	0.3	0.2

求 $E(X)$, $E(X^2)$, $E(X-1)^2$

解: $E(X) = \sum_{k=1}^{\infty} x_k P(X=x_k) = (-2) \times 0.1 + (-1) \times 0.4 + 0 \times 0.3 + 1 \times 0.2 = -0.4$;
 $E(X^2) = \sum_{k=1}^{\infty} x_k^2 P(X=x_k) = (-2)^2 \times 0.1 + (-1)^2 \times 0.4 + 0^2 \times 0.3 + 1^2 \times 0.2 = 1$;
 $E(X-1)^2 = \sum_{k=1}^{\infty} (x_k-1)^2 P(X=x_k) = [(-2)-1]^2 \times 0.1 + [(-1)-1]^2 \times 0.4 + [0-1]^2 \times 0.3 + [1-1]^2 \times 0.2$
 $= 9 \times 0.1 + 4 \times 0.4 + 1 \times 0.3 + 0 \times 0.2 = 2.8$.

(注) $E(X-1)^2 = E(X^2 - 2X + 1) = E(X^2) - 2E(X) + 1 = 1 - 2 \times (-0.4) + 1 = 2.8$.

2. 设随机变量 X 的密度函数为 $f(x) = \begin{cases} e^{-x} & x \geq 0 \\ 0 & x < 0 \end{cases}$, 求 $E(3X)$, $E(e^{-3X})$.

解: $E(3X) = \int_{-\infty}^{+\infty} 3x f(x) dx = \int_0^{+\infty} 3x e^{-x} dx = 3x(-e^{-x}) \Big|_0^{+\infty} + 3 \int_0^{+\infty} e^{-x} dx = 3(-e^{-x}) \Big|_0^{+\infty} = 3$;
 $E(e^{-3X}) = \int_{-\infty}^{+\infty} e^{-3x} f(x) dx = \int_0^{+\infty} e^{-3x} e^{-x} dx = \int_0^{+\infty} e^{-4x} dx = \frac{1}{4} (-e^{-4x}) \Big|_0^{+\infty} = \frac{1}{4}$.

3. 设随机变量 X 与 Y 相互独立, 它们的密度函数分别为

$f_X(x) = \begin{cases} 1, & 0 \leq x \leq 1 \\ 0, & \text{其他} \end{cases}$, $f_Y(y) = \begin{cases} e^{-y}, & y \geq 0 \\ 0, & y < 0 \end{cases}$, 求 $E(X+Y)$.

解: X, Y 相互独立, 故 $f(x, y) = f_X(x) \cdot f_Y(y) = \begin{cases} e^{-y}, & 0 \leq x \leq 1, y \geq 0 \\ 0, & \text{其他} \end{cases}$.
 $E(X+Y) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} (x+y) \cdot f(x, y) dx dy = \iint_D (x+y) e^{-y} dx dy = \int_0^{+\infty} dy \int_0^1 (x+y) e^{-y} dx$
 $= \int_0^{+\infty} [(x+y)e^{-y}] \Big|_0^1 + \int_0^{+\infty} e^{-y} dy dx = \int_0^{+\infty} (x+1) dx = (\frac{x^2}{2} + x) \Big|_0^{+\infty} = \frac{3}{2}$.

(注). $X \sim U(0, 1)$, 有 $E(X) = \frac{1}{2}$. $Y \sim$ 参数为 1 的指数分布, 有 $E(Y) = 1$

$\therefore E(X+Y) = E(X) + E(Y) = \frac{1}{2} + 1 = \frac{3}{2}$. 由此可见, X, Y 是否相互独立,

在利用期望性质时可以不管, 但用定义计算时要涉及 $f(x, y)$ 的计算, 故需要.

4. 设 (X, Y) 的联合分布律为

$X \backslash Y$	0	1	2
0	0.1	0.2	a
1	0.1	b	0.2

已知 $E(X^2 + Y^2) = 2.4$, 求 a, b 之值.

解: 由 $1 = \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} P(X=x_i, Y=y_j) = 0.1 + 0.2 + a + 0.1 + b + 0.2 = a + b + 0.6$

$E(X^2 + Y^2) = (0^2 + 0^2) \times 0.1 + (0^2 + 1^2) \times 0.2 + (0^2 + 2^2) \times a + (1^2 + 0^2) \times 0.1 + (1^2 + 1^2) \times b$
 $+ (1^2 + 2^2) \times 0.2 = 4a + 2b + 1.3 = 2.4$.

$\therefore \begin{cases} a + b = 0.4 \\ 4a + 2b = 1.1 \end{cases} \therefore \begin{cases} a = 1.5 \\ b = 2.5 \end{cases}$

(注). 也可利用关于 X, Y 的边缘分布分别计算 $E(X^2)$ 和 $E(Y^2)$ 再获得

$E(X^2), E(Y^2)$

习题 4-3

1. 设随机变量 X 的数学期望与方差均存在且 $D(X) > 0$, 称 $X^* = \frac{X - E(X)}{\sqrt{D(X)}}$

为 X 的标准化的随机变量, 证明: $E(X^*) = 0, D(X^*) = 1$.

$$\text{证: } E(X^*) = E\left(\frac{X - E(X)}{\sqrt{D(X)}}\right) = \frac{1}{\sqrt{D(X)}} E(X - E(X)) = \frac{1}{\sqrt{D(X)}} (E(X) - E(X)) = 0;$$

$$D(X^*) = D\left(\frac{X - E(X)}{\sqrt{D(X)}}\right) = \frac{1}{D(X)} D(X - E(X)) = \frac{1}{D(X)} (D(X) - 0) = 1.$$

2. 设随机变量 X 的密度函数为

$$f(x) = \begin{cases} \frac{A}{1+x^2}, & -1 \leq x \leq 1; \\ 0, & \text{其他.} \end{cases} \quad \text{求: (1) 常数 } A; (2) E(X); (3) D(X).$$

$$\text{解: (1) 由 } 1 = \int_{-\infty}^{+\infty} f(x) dx = \int_{-1}^1 \frac{A}{1+x^2} dx = A(\arctan x) \Big|_{-1}^1 = A\left(\frac{\pi}{4} - \left(-\frac{\pi}{4}\right)\right) = \frac{\pi A}{2}$$

$$\therefore A = \frac{2}{\pi}.$$

$$(2) E(X) = \int_{-\infty}^{+\infty} x f(x) dx = \int_{-1}^1 \frac{\frac{2}{\pi} x}{1+x^2} dx = \frac{1}{\pi} \int_{-1}^1 \frac{1}{1+x^2} d(1+x^2) = \frac{1}{\pi} \ln(1+x^2) \Big|_{-1}^1 = 0.$$

$$(3) E(X^2) = \int_{-\infty}^{+\infty} x^2 f(x) dx = \int_{-1}^1 \frac{\frac{2}{\pi} x^2}{1+x^2} dx = \frac{2}{\pi} \int_{-1}^1 \frac{x^2+1-1}{x^2+1} dx = \frac{2}{\pi} \int_{-1}^1 \left(1 - \frac{1}{1+x^2}\right) dx$$

$$= \frac{4}{\pi} \int_0^1 \left(1 - \frac{1}{1+x^2}\right) dx = \frac{4}{\pi} (x - \arctan x) \Big|_0^1 = \frac{4}{\pi} \left(1 - \frac{\pi}{4}\right) = \frac{4}{\pi} - 1.$$

$$D(X) = E(X^2) - (E(X))^2 = \left(\frac{4}{\pi} - 1\right) - 0^2 = \frac{4}{\pi} - 1.$$

3. 设随机变量 X 的密度函数为

$$f(x) = \begin{cases} 1+x, & -1 \leq x < 0; \\ 1-x, & 0 \leq x < 1; \\ 0, & \text{其他.} \end{cases} \quad \text{求: (1) } E(X),$$

(2) $D(X)$;

(3) $P(|X - E(X)| \leq 2\sqrt{D(X)})$.

$$\text{解: (1) } E(X) = \int_{-\infty}^{+\infty} x f(x) dx = \int_{-1}^0 x(1+x) dx + \int_0^1 x(1-x) dx = \left(\frac{x^2}{2} + \frac{x^3}{3}\right) \Big|_{-1}^0 + \left(\frac{x^2}{2} - \frac{x^3}{3}\right) \Big|_0^1 = \left(-\frac{1}{2} + \frac{1}{3}\right) + \left(\frac{1}{2} - \frac{1}{3}\right) = 0.$$

$$(2) D(X) = E(X^2) - (E(X))^2 = E(X^2) - 0^2 = E(X^2) = \int_{-\infty}^{+\infty} x^2 f(x) dx$$

$$= \int_{-1}^0 x^2(1+x) dx + \int_0^1 x^2(1-x) dx = \left(\frac{x^3}{3} + \frac{x^4}{4}\right) \Big|_{-1}^0 + \left(\frac{x^3}{3} - \frac{x^4}{4}\right) \Big|_0^1 = \frac{2}{3} - \frac{1}{2} = \frac{1}{6}$$

$$(3) P(|X - E(X)| \leq 2\sqrt{D(X)}) = P(|X| \leq \frac{1}{3}) = P\left(-\frac{1}{3} \leq X \leq \frac{1}{3}\right) = \int_{-1/3}^{1/3} f(x) dx$$

$$= \int_{-1/3}^0 (1+x) dx + \int_0^{1/3} (1-x) dx = \left(x + \frac{x^2}{2}\right) \Big|_{-1/3}^0 + \left(x - \frac{x^2}{2}\right) \Big|_0^{1/3} = \frac{5}{18} + \frac{5}{18} = \frac{5}{9}.$$

4. 设随机变量 X 与 Y 相互独立, 且 $X \sim N(0, 1), Y \sim U[0, 2]$. 求:

(1) $E(X - 2Y)$; (2) $D(X - 2Y)$; (3) $E[(X + Y)^2]$.

$$\text{解: } X \sim N(0, 1). \therefore E(X) = 0, D(X) = 1, E(X^2) = D(X) + (E(X))^2 = 1.$$

$$Y \sim U[0, 2]. \therefore E(Y) = 1, D(Y) = \frac{2^2}{12} = \frac{1}{3}, E(Y^2) = D(Y) + (E(Y))^2 = \frac{4}{3}.$$

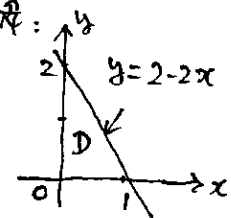
$$(1) E(X - 2Y) = E(X) - 2E(Y) = 0 - 2 \times 1 = -2.$$

$$(2) D(X-2Y) = D(X) + 4D(Y) = 1 + 4 \times \frac{1}{3} = \frac{7}{3}; \quad (X, Y \text{ 相互独立的})$$

$$(3) E[(X+Y)^2] = E(X^2 + Y^2 + 2XY) = E(X^2) + E(Y^2) + 2E(XY), \quad (X, Y \text{ 独立})$$

$$= E(X^2) + E(Y^2) + 2E(X)E(Y)$$

5. 设二维随机变量 (X, Y) 的密度函数为 $f(x, y) = \begin{cases} 6xy, & 0 < x < 1, 0 < y < 2(1-x) \\ 0, & \text{其他} \end{cases}$
求 $E(X)$, $E(Y)$, $D(Y)$ 及 $E(XY)$.

解: 

$$E(X) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} x f(x, y) dx dy = \iint_D x \cdot 6xy dx dy = \int_0^1 dx \int_0^{2-2x} 6x^2 y dy$$

$$= \int_0^1 [3x^2(y^2)]_0^{2(1-x)} dx = \int_0^1 12x^2(1-x)^2 dx = \int_0^1 12(x^2 - 2x^3 + x^4) dx$$

$$= (4x^3 - 6x^4 + \frac{12}{5}x^5) \Big|_0^1 = \frac{2}{5}$$

$$E(Y) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} y f(x, y) dx dy = \iint_D y \cdot 6xy dx dy = \int_0^1 dx \int_0^{2-2x} 6xy^2 dy$$

$$= \int_0^1 [2x(y^3)]_0^{2(1-x)} dx = \int_0^1 16x(1-x)^3 dx = \int_0^1 16(x - 3x^2 + 3x^3 - x^4) dx$$

$$= (8x^2 - 16x^3 + 12x^4 - \frac{16}{5}x^5) \Big|_0^1 = \frac{4}{5}$$

$$E(Y^2) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} y^2 f(x, y) dx dy = \iint_D y^2 \cdot 6xy dx dy = \int_0^1 dx \int_0^{2-2x} 6xy^3 dy$$

$$= \int_0^1 [\frac{3}{2}x(y^4)]_0^{2(1-x)} dx = \int_0^1 24x(1-x)^4 dx = \int_0^1 24(x - 4x^2 + 6x^3 - 4x^4 + x^5) dx$$

$$= (12x^2 - 32x^3 + 36x^4 - \frac{96}{5}x^5 + 4x^6) \Big|_0^1 = \frac{4}{5}$$

$$\therefore D(Y) = E(Y^2) - (E(Y))^2 = \frac{4}{5} - (\frac{4}{5})^2 = \frac{4}{25}$$

$$E(XY) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f(x, y) dx dy = \iint_D xy \cdot 6xy dx dy = \int_0^1 dx \int_0^{2-2x} 6x^2 y^2 dy$$

$$= \int_0^1 [2x^2(y^3)]_0^{2(1-x)} dx = \int_0^1 16x^2(1-x)^3 dx = \int_0^1 16(x^2 - 3x^3 + 3x^4 - x^5) dx$$

$$= (\frac{16}{3}x^3 - 12x^4 + \frac{48}{5}x^5 - \frac{8}{3}x^6) \Big|_0^1 = \frac{4}{15}$$

(注): $E(Y) = \iint_D 6xy^2 dx dy = \int_0^2 dy \int_0^{1-\frac{1}{2}y} 6xy^2 dx = \int_0^2 [3y^2(x^2)]_0^{1-\frac{1}{2}y} dy$

$$= \int_0^2 3y^2(1-\frac{y}{2})^2 dy = \int_0^2 3(y^2 - y^3 + \frac{y^4}{4}) dy = (y^3 - \frac{3}{4}y^4 + \frac{3}{20}y^5) \Big|_0^2 = \frac{4}{5}$$

$$E(Y^2) = \iint_D 6xy^3 dx dy = \int_0^2 dy \int_0^{1-\frac{1}{2}y} 6xy^3 dx = \int_0^2 [3y^3(x^2)]_0^{1-\frac{1}{2}y} dy$$

$$= \int_0^2 3y^3(1-\frac{y}{2})^2 dy = \int_0^2 3(y^3 - y^4 + \frac{y^5}{4}) dy = (\frac{3}{4}y^4 - \frac{3}{5}y^5 + \frac{1}{40}y^6) \Big|_0^2 = \frac{4}{5}$$

$$P\{Y_k=1\} = P\{\sum_{i=1}^{100} X_i > 3\} = 0.0017.$$

$$P\{Y_k=0\} = P\{\sum_{i=1}^{100} X_i \leq 3\} = 0.9983.$$

$$\sum_{k=1}^{10000} Y_k = \text{"10000箱中不能通过验收的箱数"} \text{近似服从 } E(\sum_{k=1}^{10000} Y_k) \\ = 10000 \times 0.0017 = 17; D(\sum_{k=1}^{10000} Y_k) = 10000 \times 0.0017 \times 0.9983 = 16.9983 \\ \text{的正态分布 } N(17, 16.9983).$$

$$\therefore P\{\sum_{k=1}^{10000} Y_k > 25\} = 1 - \Phi\left(\frac{25-17}{\sqrt{16.9983}}\right) = 1 - \Phi(1.94) \\ = 1 - 0.9738 = 0.0262.$$

7. 用棣莫弗-拉普拉斯中心极限定理证明伯努利大数定理.

解: 伯努利大数定理的条件与棣莫弗-拉普拉斯中心极限定理的条件一致, 即 $X_i (i=1, 2, \dots)$ 服从参数为 $p (0 < p < 1)$ 的 0-1 分布, 且相互独立. 则 $\sum_{i=1}^n X_i \sim B(n, p)$. 由中心极限定理得 $\sum_{i=1}^n X_i$ 近似服从正态分布 $N(np, np(1-p))$. 即

$$P\left\{\frac{\sum_{i=1}^n X_i - np}{\sqrt{np(1-p)}} \leq x\right\} = \Phi(x).$$

$$\text{由 } P\left\{\left|\frac{1}{n}\sum_{i=1}^n X_i - p\right| < \varepsilon\right\} = P\left\{-\varepsilon < \frac{1}{n}\sum_{i=1}^n X_i - p < \varepsilon\right\} \\ = P\left\{-\frac{\sqrt{n}\varepsilon}{\sqrt{p(1-p)}} < \frac{\sum_{i=1}^n X_i - np}{\sqrt{np(1-p)}} < \frac{\sqrt{n}\varepsilon}{\sqrt{p(1-p)}}\right\} = \Phi\left(\frac{\sqrt{n}\varepsilon}{\sqrt{p(1-p)}}\right) - \Phi\left(\frac{-\sqrt{n}\varepsilon}{\sqrt{p(1-p)}}\right) \\ = 2\Phi\left(\frac{\sqrt{n}\varepsilon}{\sqrt{p(1-p)}}\right) - 1.$$

$$\therefore P\left\{\left|\frac{1}{n}\sum_{i=1}^n X_i - p\right| < \varepsilon\right\} = 2\Phi\left(\frac{\sqrt{n}\varepsilon}{\sqrt{p(1-p)}}\right) - 1.$$

$$\therefore \lim_{n \rightarrow +\infty} P\left\{\left|\frac{1}{n}\sum_{i=1}^n X_i - p\right| < \varepsilon\right\} = 2 \lim_{n \rightarrow +\infty} \Phi\left(\frac{\sqrt{n}\varepsilon}{\sqrt{p(1-p)}}\right) - 1 \\ = 2\Phi(+\infty) - 1 = 1.$$

即为伯努利大数定理的结论成立.

$$E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} \sum_{i=1}^n E(X_i) = 0.8; D\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n^2} \sum_{i=1}^n D(X_i) = \frac{0.16}{n}$$

$$(1) \text{ 根据切比雪夫不等式, } P\left(\left|\frac{1}{n} \sum_{i=1}^n X_i - E\left(\frac{1}{n} \sum_{i=1}^n X_i\right)\right| < \varepsilon\right) > 1 - \frac{D\left(\frac{1}{n} \sum_{i=1}^n X_i\right)}{\varepsilon^2}$$

$$\text{即 } P\{0.8 - \varepsilon < \frac{1}{n} \sum_{i=1}^n X_i < 0.8 + \varepsilon\} > 1 - \frac{0.16}{\varepsilon^2 \cdot n}$$

$$\text{现求 } P\{0.76 < \frac{1}{n} \sum_{i=1}^n X_i < 0.84\} > 0.9 \text{ 中的 } n. \text{ 对上式}$$

$$\text{得 } \varepsilon = 0.04, \text{ 且 } 1 - \frac{0.16}{(0.04)^2 \cdot n} = 0.9.$$

$$\text{得 } n = 0.16 / (0.1 \times 0.0016) = 1000.$$

$$(2) \text{ 根据中心极限定理: } \frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{\text{分布}} N(0.8, \frac{0.16}{n}).$$

$$\therefore P\{0.76 < \frac{1}{n} \sum_{i=1}^n X_i < 0.84\} = \Phi\left(\frac{0.84 - 0.8}{\sqrt{0.16/n}}\right) - \Phi\left(\frac{0.76 - 0.8}{\sqrt{0.16/n}}\right)$$

$$= \Phi(0.1\sqrt{n}) - \Phi(-0.1\sqrt{n}) = 2\Phi(0.1\sqrt{n}) - 1 > 0.9$$

$$\therefore \Phi(0.1\sqrt{n}) > 0.95, \therefore 0.1\sqrt{n} = 1.645, \therefore n = 270.6 \approx 271.$$

由结果可见(2)的结果比(1)更接近于实际.

6. 某工厂生产的一种产品其次品率为0.005, 每只产品是否为次品相互独立. 产品按每100只包装成为一箱, 一箱中若次品数超过3只就不能通过验收. 各箱是否能通过验收相互独立. 今有10000箱产品. 求多于25箱不能通过验收的概率.

解: $X_i = \begin{cases} 1, & \text{第 } i \text{ 个产品为次品;} \\ 0, & \text{第 } i \text{ 个产品为正品.} \end{cases} i = 1, \dots, 100 \text{ 相互独立.}$

$$P\{X_i = 1\} = 0.005, P\{X_i = 0\} = 0.995.$$

$$\sum_{i=1}^{100} X_i = \text{"一箱100只产品中的次品数"} \sim B(100, 100 \times 0.005)$$

$$P\left\{\sum_{i=1}^{100} X_i > 3\right\} = 1 - \left[P\left\{\sum_{i=1}^{100} X_i = 0\right\} + P\left\{\sum_{i=1}^{100} X_i = 1\right\} + P\left\{\sum_{i=1}^{100} X_i = 2\right\} + P\left\{\sum_{i=1}^{100} X_i = 3\right\}\right]$$

$$= 1 - \left[C_{100}^0 (0.005)^0 (0.995)^{100} + C_{100}^1 0.005 \cdot (0.995)^{99} + C_{100}^2 (0.005)^2 (0.995)^{98} + C_{100}^3 (0.005)^3 (0.995)^{97}\right]$$

$$= 1 - 0.9983 = 0.0017. \text{ (亦可用泊松分布作近似计算)}$$

又设 $Y_k = \begin{cases} 1, & \text{第 } k \text{ 箱不通过验收;} \\ 0, & \text{第 } k \text{ 箱通过验收.} \end{cases} k = 1, \dots, 10000, \text{ 相互独立.}$

习题 4-4

4. 设二维随机变量 (X, Y) 分别具有下列密度函数, 求 $\text{Cov}(X, Y)$, ρ_{XY} .

(1) $f(x, y) = \begin{cases} \frac{1}{8}(x+y), & 0 \leq x \leq 2, 0 \leq y \leq 2; \\ 0, & \text{其他}; \end{cases}$ (2) $f(x, y) = \begin{cases} 2, & x > 0, y > 0, x+y < 1; \\ 0, & \text{其他}. \end{cases}$

解: (1) $E(X) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} x f(x, y) dx dy = \iint_D x \cdot \frac{1}{8}(x+y) dx dy = \int_0^2 dx \int_0^2 \frac{1}{8}(x^2 + xy) dy$
 $= \int_0^2 [\frac{1}{8}x^2(y) + \frac{1}{16}x(y^2)] dx = \int_0^2 \frac{1}{4}(x^2 + x) dx = (\frac{x^3}{12} + \frac{x^2}{8}) \Big|_0^2 = \frac{7}{6}$

$E(Y) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} y f(x, y) dx dy = \iint_D y \cdot \frac{1}{8}(x+y) dx dy$
 $= \int_0^2 dy \int_0^2 \frac{1}{8}(xy + y^2) dx = \int_0^2 [\frac{y}{16}(x^2) + \frac{y^2}{8}(x)] dy = \int_0^2 (\frac{y}{4} + \frac{y^2}{4}) dy = (\frac{y^2}{8} + \frac{y^3}{12}) \Big|_0^2 = \frac{7}{6}$

$E(XY) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f(x, y) dx dy = \iint_D xy \cdot \frac{1}{8}(x+y) dx dy = \int_0^2 dx \int_0^2 \frac{1}{8}(x^2y + xy^2) dy$
 $= \int_0^2 [\frac{x^2}{16}(y^2) + \frac{x}{24}(y^3)] dy = \int_0^2 (\frac{x^2}{4} + \frac{x}{3}) dx = (\frac{x^3}{12} + \frac{x^2}{6}) \Big|_0^2 = \frac{4}{3}$

$\therefore \text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \frac{4}{3} - \frac{7}{6} \cdot \frac{7}{6} = -\frac{1}{36}$

$E(X^2) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} x^2 f(x, y) dx dy = \iint_D x^2 \cdot \frac{1}{8}(x+y) dx dy = \int_0^2 dx \int_0^2 \frac{1}{8}(x^3 + x^2y) dy$
 $= \int_0^2 [\frac{x^3}{8}(y) + \frac{x^2}{16}(y^2)] dy = \int_0^2 \frac{1}{4}(x^3 + x^2) dx = (\frac{x^4}{16} + \frac{x^3}{12}) \Big|_0^2 = \frac{5}{3}$

$E(Y^2) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} y^2 f(x, y) dx dy = \iint_D y^2 \cdot \frac{1}{8}(x+y) dx dy = \int_0^2 dy \int_0^2 \frac{1}{8}(y^2x + y^3) dx$
 $= \int_0^2 [\frac{y^2}{16}(x^2) + \frac{y^3}{8}(x)] dy = \int_0^2 \frac{1}{4}(y^2 + y^3) dy = (\frac{y^3}{12} + \frac{y^4}{16}) \Big|_0^2 = \frac{5}{3}$

$\therefore D(X) = E(X^2) - (E(X))^2 = \frac{5}{3} - (\frac{7}{6})^2 = \frac{11}{36}$; $D(Y) = E(Y^2) - (E(Y))^2 = \frac{5}{3} - (\frac{7}{6})^2 = \frac{11}{36}$

$\therefore \rho_{XY} = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)} \cdot \sqrt{D(Y)}} = \frac{-\frac{1}{36}}{\sqrt{\frac{11}{36}} \cdot \sqrt{\frac{11}{36}}} = -\frac{1}{11}$

(2) $E(X) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} x \cdot f(x, y) dx dy = \iint_D x \cdot 2 dx dy = \int_0^1 dx \int_0^{1-x} 2x dy$
 $= \int_0^1 2x(1-x) dx = (x^2 - \frac{2}{3}x^3) \Big|_0^1 = \frac{1}{3}$

$E(Y) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} y f(x, y) dx dy = \iint_D y \cdot 2 dx dy = \int_0^1 dy \int_0^{1-y} 2y dx$
 $= \int_0^1 2y(1-y) dy = (y^2 - \frac{2}{3}y^3) \Big|_0^1 = \frac{1}{3}$

$E(XY) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f(x, y) dx dy = \iint_D xy \cdot 2 dx dy = \int_0^1 dx \int_0^{1-x} 2xy dy$

$$= \int_0^1 x(1-x)^2 dx = \int_0^1 (x^2 - 2x^3 + x^4) dx = \left(\frac{x^3}{3} - \frac{2}{4}x^4 + \frac{x^5}{5} \right) \Big|_0^1 = \frac{1}{12}$$

$$\therefore \text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \frac{1}{12} - \frac{1}{3} \cdot \frac{1}{3} = -\frac{1}{36};$$

$$E(X^2) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} x^2 f(x, y) dx dy = \iint_D x^2 \cdot 2 dx dy = \int_0^1 dx \int_0^{1-x} 2x^2 dy = \int_0^1 2x^2(1-x) dx$$

$$= \left(\frac{2}{3}x^3 - \frac{1}{2}x^4 \right) \Big|_0^1 = \frac{1}{6};$$

$$E(Y^2) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} y^2 f(x, y) dx dy = \iint_D y^2 \cdot 2 dx dy = \int_0^1 dy \int_0^{1-y} 2y^2 dx = \int_0^1 2y^2(1-y) dy$$

$$= \left(\frac{2}{3}y^3 - \frac{1}{2}y^4 \right) \Big|_0^1 = \frac{1}{6};$$

$$\therefore D(X) = E(X^2) - (E(X))^2 = \frac{1}{6} - \frac{1}{3} \cdot \frac{1}{3} = \frac{3}{54}; D(Y) = E(Y^2) - (E(Y))^2 = \frac{1}{6} - \frac{1}{3} \cdot \frac{1}{3} = \frac{3}{54};$$

$$\therefore \rho_{XY} = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}} = \frac{-\frac{1}{36}}{\sqrt{\frac{3}{54}}\sqrt{\frac{3}{54}}} = -\frac{1}{2}.$$

(注) 在计算两重积分时, 可视被积函数情况选择逐次积分的先后顺序来减少计算量.

2. 设 X 与 Y 为随机变量, $D(X)=25$, $D(Y)=36$, $\rho_{XY}=0.4$. 求 $D(X+Y)$, $D(X-Y)$.

解: 由 $\text{Cov}(X, Y) = \rho_{XY} \sqrt{D(X) \cdot D(Y)} = 0.4 \sqrt{25 \times 36} = 0.4 \times 30 = 12$.

$$\therefore D(X+Y) = D(X) + D(Y) + 2\text{Cov}(X, Y) = 25 + 36 + 2 \times 12 = 85;$$

$$D(X-Y) = D(X) + D(Y) - 2\text{Cov}(X, Y) = 25 + 36 - 2 \times 12 = 37.$$

3. 设 X 与 Y 为随机变量, 若 $X_1 = aX + b$, $Y_1 = cY + d$, 其中 a, b, c, d 均为常数, 且 $ac > 0$. 试证: $\rho_{X_1, Y_1} = \rho_{XY}$.

证: $E(X_1) = E(aX + b) = aE(X) + b$; $D(X_1) = D(aX + b) = a^2 D(X)$;

$$E(Y_1) = E(cY + d) = cE(Y) + d; D(Y_1) = D(cY + d) = c^2 D(Y).$$

$$E(X_1 Y_1) = E[(aX + b)(cY + d)] = E(acXY + adX + bcY + bd)$$

$$= acE(XY) + adE(X) + bcE(Y) + bd.$$

$$\therefore \text{Cov}(X_1, Y_1) = E(X_1 Y_1) - E(X_1)E(Y_1) = a^2 D(X) + c^2 D(Y)$$

$$= acE(XY) + adE(X) + bcE(Y) + bd - (aE(X) + b)(cE(Y) + d)$$

$$= ac[E(XY) - E(X)E(Y)] = ac \text{Cov}(X, Y).$$

$$\therefore \rho_{X_1, Y_1} = \frac{\text{Cov}(X_1, Y_1)}{\sqrt{D(X_1)}\sqrt{D(Y_1)}} = \frac{ac \text{Cov}(X, Y)}{\sqrt{a^2 D(X)}\sqrt{c^2 D(Y)}} = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}} = \rho_{XY}.$$

习题 4-5

1. 设二维随机变量 (X, Y) 分别具有下列联合密度函数. 问 X 与 Y 是否相互独立, X 与 Y 是否相关? 为什么?

$$(1) f(x, y) = \begin{cases} 4xy, & 0 \leq x \leq 1, 0 \leq y \leq 1; \\ 0, & \text{其他} \end{cases}; (2) f(x, y) = \begin{cases} \frac{1}{\pi}, & x^2 + y^2 \leq 1; \\ 0, & \text{其他} \end{cases}$$

解: (1) $f_x(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \int_0^1 4xy dy = 2x(y^2)|_0^1 = 2x, 0 \leq x \leq 1;$

$$f_y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \int_0^1 4xy dx = 2y(x^2)|_0^1 = 2y, 0 \leq y \leq 1;$$

$$\therefore f_x(x) = \begin{cases} 2x, & 0 \leq x \leq 1; \\ 0, & \text{其他} \end{cases}; f_y(y) = \begin{cases} 2y, & 0 \leq y \leq 1; \\ 0, & \text{其他} \end{cases}$$

$$\therefore f(x, y) = \begin{cases} 4xy, & 0 \leq x \leq 1, 0 \leq y \leq 1; \\ 0, & \text{其他} \end{cases} = f_x(x) f_y(y), \therefore X, Y \text{ 相互独立.}$$

则有 $E(XY) = E(X)E(Y) \therefore \text{Cov}(X, Y) = 0, \therefore \rho_{XY} = 0, \therefore X, Y \text{ 不相关.}$

$$(2) f_x(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{1}{\pi} dy = \frac{2}{\pi} \sqrt{1-x^2}, -1 \leq x \leq 1;$$

$$f_y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} \frac{1}{\pi} dx = \frac{2}{\pi} \sqrt{1-y^2}, -1 \leq y \leq 1.$$

$$\therefore f_x(x) = \begin{cases} \frac{2}{\pi} \sqrt{1-x^2}, & -1 \leq x \leq 1; \\ 0, & \text{其他} \end{cases}; f_y(y) = \begin{cases} \frac{2}{\pi} \sqrt{1-y^2}, & -1 \leq y \leq 1; \\ 0, & \text{其他} \end{cases}$$

因在 $x^2 + y^2 \leq 1$ 的圆域内

$$f(x, y) = \frac{1}{\pi} \neq \frac{2}{\pi} \sqrt{1-x^2} \cdot \frac{2}{\pi} \sqrt{1-y^2} = f_x(x) \cdot f_y(y), \therefore X, Y \text{ 不相互独立,}$$

而 $E(X) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} x f(x, y) dx dy = \iint_D \frac{x}{\pi} dx dy = \int_{-1}^1 dx \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{x}{\pi} dy = \int_{-1}^1 \frac{2x}{\pi} \sqrt{1-x^2} dx \stackrel{(*)}{=} 0;$

$$E(Y) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} y f(x, y) dx dy = \iint_D \frac{y}{\pi} dx dy = \int_{-1}^1 dy \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} \frac{y}{\pi} dx = \int_{-1}^1 \frac{2y}{\pi} \sqrt{1-y^2} dy \stackrel{(*)}{=} 0;$$

$$E(XY) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f(x, y) dx dy = \iint_D \frac{xy}{\pi} dx dy = \int_{-1}^1 dy \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} \frac{xy}{\pi} dx = \int_{-1}^1 \left[\frac{y}{2\pi} (x^2) \right]_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} dy = 0;$$

$$\therefore E(XY) = E(X)E(Y), \therefore \text{Cov}(X, Y) = 0, \therefore \rho_{XY} = 0, \therefore X, Y \text{ 不相关.}$$

(注) * 处用奇函数在对称区间上的定积分为 0 的性质.

联系 (1) (2) 的结论可见: X, Y 相互独立仅仅是 X, Y 不相关的充分条件; 反之, X, Y 不相关仅仅是 X, Y 相互独立的必要条件.

2. 设随机变量 X_1, X_2, X_3, X_4 相互独立, 且具有相同的分布. 数学期望为 0, 方差为 σ^2 . 令 $X = X_1 + X_2 + X_3, Y = X_2 + X_3 + X_4$. 求 ρ_{XY} .

解: $E(X) = E(X_1 + X_2 + X_3) = E(X_1) + E(X_2) + E(X_3) = 0;$

$$E(Y) = E(X_2 + X_3 + X_4) = E(X_2) + E(X_3) + E(X_4) = 0;$$

$$E(XY) = E[(X_1 + X_2 + X_3)(X_2 + X_3 + X_4)] = E[X_1X_2 + X_1X_3 + X_1X_4 + 2X_2X_3 + X_2X_4 + X_3X_4 + X_2^2 + X_3^2] = E(X_1X_2) + E(X_1X_3) + E(X_1X_4) + 2E(X_2X_3) + E(X_2X_4) + E(X_3X_4) + E(X_2^2) + E(X_3^2) = 0 + 0 + 0 + 2\sigma^2 + 0 + 0 + \sigma^2 + \sigma^2 = 4\sigma^2;$$

$$+ E(X_3 X_4) + E(X_2^2) + E(X_3^2) = E(X_1)E(X_2) + E(X_1)E(X_3) + E(X_1)E(X_4) + 2E(X_2)E(X_3) \\ + E(X_2)E(X_4) + E(X_3)E(X_4) + D(X_2) + D(X_3) = 2\sigma^2;$$

$$E(X^2) = E[(X_1 + X_2 + X_3)^2] = E(X_1^2 + X_2^2 + X_3^2 + 2X_1X_2 + 2X_1X_3 + 2X_2X_3) = E(X_1^2) + E(X_2^2) + E(X_3^2) \\ + 2E(X_1)E(X_2) + 2E(X_1)E(X_3) + 2E(X_2)E(X_3) = 3\sigma^2;$$

$$E(Y^2) = E[(X_2 + X_3 + X_4)^2] = E(X_2^2 + X_3^2 + X_4^2 + 2X_2X_3 + 2X_2X_4 + 2X_3X_4) = E(X_2^2) + E(X_3^2) + E(X_4^2) \\ + 2E(X_2)E(X_3) + 2E(X_2)E(X_4) + 2E(X_3)E(X_4) = 3\sigma^2.$$

$$\therefore \text{Cov}(X, Y) = E(XY) - E(X)E(Y) = 2\sigma^2; \quad D(X) = E(X^2) - (E(X))^2 = 3\sigma^2; \quad D(Y) = E(Y^2) - (E(Y))^2 = 3\sigma^2$$

$$\therefore \rho_{XY} = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}} = \frac{2\sigma^2}{\sqrt{3\sigma^2} \cdot \sqrt{3\sigma^2}} = \frac{2}{3}.$$

3. 设二维随机变量 (X, Y) 的联合分布律为

$X \backslash Y$	0	1
0	0.1	0.3
1	0.2	0.4

求 (X, Y) 的协方差矩阵.

解: 关于 X, Y 的边缘分布律为

X	0	1
P	0.3	0.7

Y	0	1
P	0.4	0.6

$$E(X) = 0 \times 0.3 + 1 \times 0.7 = 0.7; \quad E(Y) = 0 \times 0.4 + 1 \times 0.6 = 0.6;$$

$$E(X^2) = 0^2 \times 0.3 + 1^2 \times 0.7 = 0.7; \quad E(Y^2) = 0^2 \times 0.4 + 1^2 \times 0.6 = 0.6;$$

$$E(XY) = 0 \times 0 \times 0.1 + 0 \times 1 \times 0.2 + 1 \times 0 \times 0.3 + 1 \times 1 \times 0.4 = 0.4.$$

$$\therefore \text{Cov}(X, X) = D(X) = E(X^2) - (E(X))^2 = 0.7 - (0.7)^2 = 0.21;$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = 0.4 - 0.7 \times 0.6 = 0.02;$$

$$\text{Cov}(Y, X) = E(XY) - E(X)E(Y) = 0.4 - 0.7 \times 0.6 = 0.02;$$

$$\text{Cov}(Y, Y) = D(Y) = E(Y^2) - (E(Y))^2 = 0.6 - (0.6)^2 = 0.24.$$

$$\therefore (X, Y) \text{ 的协方差矩阵为 } \begin{pmatrix} \text{Cov}(X, X) & \text{Cov}(X, Y) \\ \text{Cov}(Y, X) & \text{Cov}(Y, Y) \end{pmatrix} = \begin{pmatrix} 0.21 & 0.02 \\ 0.02 & 0.24 \end{pmatrix}.$$

复习题 4 (仅对三. 计算题作解答)

1. 设随机变量 X 的密度函数为

$$f(x) = \begin{cases} ax, & 0 < x < 2; \\ bx+c, & 2 \leq x \leq 4; \\ 0, & \text{其他.} \end{cases} \quad \text{已知 } E(X)=2, P(1 < X < 3) = \frac{3}{4}.$$

解: (1) 由 $1 = \int_{-\infty}^{+\infty} f(x) dx = \int_0^2 ax dx + \int_2^4 (bx+c) dx = 2a+6b+2c;$

$$E(X) = \int_{-\infty}^{+\infty} xf(x) dx = \int_0^2 x \cdot ax dx + \int_2^4 x(bx+c) dx = \frac{8}{3}a + \frac{56}{3}b + 6c = 2;$$

$$P(1 < X < 3) = \int_1^3 f(x) dx = \int_1^2 ax dx + \int_2^3 (bx+c) dx = \frac{3}{2}a + \frac{5}{2}b + c = \frac{3}{4}.$$

得方程组 $\begin{cases} a+3b+c = \frac{1}{2}; \\ 4a+28b+6c = 2; \\ 3a+5b+2c = \frac{3}{2}. \end{cases}$ 解得 $\begin{cases} a = \frac{1}{4}; \\ b = -\frac{1}{4}; \\ c = 1. \end{cases}$

$$\begin{aligned} (2) E(e^X) &= \int_{-\infty}^{+\infty} e^x f(x) dx = \int_0^2 e^x \cdot \frac{x}{4} dx + \int_2^4 e^x (1 - \frac{x}{4}) dx \\ &= \left[\frac{x}{4}(-e^{-x}) \right]_0^2 + \frac{1}{4} \int_0^2 e^{-x} dx + \left[(1 - \frac{x}{4})(-e^{-x}) \right]_2^4 - \frac{1}{4} \int_2^4 e^{-x} dx \\ &= -\frac{1}{4}e^{-x} \Big|_0^2 + \frac{1}{4}e^{-x} \Big|_2^4 = 1 - \frac{1}{2}e^{-2} + \frac{1}{4}e^{-4} = \frac{1}{4}(2 - e^{-2})^2. \end{aligned}$$

2. 设 X 与 Y 是随机变量, 为使 $E\{[Y - (aX + bY)]^2\}$ 达到最小值, 求常数 a, b 之值.

解: 记 $g(a, b) = E\{[Y - (aX + bY)]^2\} = (1-b)^2 E(Y^2) + a^2 E(X^2) - 2a(1-b)E(XY)$

利用二元函数求极值的方法:

$$\begin{cases} \frac{\partial g(a, b)}{\partial a} = 2aE(X^2) - 2(1-b)E(XY) = 0 \\ \frac{\partial g(a, b)}{\partial b} = -2(1-b)E(Y^2) + 2aE(XY) = 0 \end{cases} \quad \text{得} \begin{cases} a = 0; \\ b = 1. \end{cases}$$

(注) 本题书上答案错的, 且当 $a=0, b=1$ 时, 该期望值为 0, 此结果用直接观察即可获得. 故本题作者可能出得不妥, 但解题方法可借鉴.

3. 设随机变量 X 的密度函数为

$$f(x) = \begin{cases} ax^2 + bx + c, & 0 < x < 1; \\ 0, & \text{其他.} \end{cases} \quad \text{已知 } E(X) = \frac{1}{2}, D(X) = \frac{3}{20}.$$

求 a, b, c 之值.

解: 由 $\int_{-\infty}^{+\infty} f(x) dx = \int_0^1 (ax^2 + bx + c) dx = \frac{a}{3} + \frac{b}{2} + c = 1$;

$$E(X) = \int_{-\infty}^{+\infty} x f(x) dx = \int_0^1 x(ax^2 + bx + c) dx = \frac{a}{4} + \frac{b}{3} + \frac{c}{2} = \frac{1}{2};$$

$$E(X^2) = \int_{-\infty}^{+\infty} x^2 f(x) dx = \int_0^1 x^2(ax^2 + bx + c) dx = \frac{a}{5} + \frac{b}{4} + \frac{c}{3};$$

$$D(X) = E(X^2) - (E(X))^2 = \frac{a}{5} + \frac{b}{4} + \frac{c}{3} - \frac{1}{4} = \frac{3}{20}.$$

得 $\begin{cases} 2a + 3b + 6c = 6; \\ 3a + 4b + 6c = 6; \\ 12a + 15b + 20c = 24. \end{cases}$ 解得 $\begin{cases} a = 12; \\ b = -12; \\ c = 3. \end{cases}$

4. 设随机变量 X 与 Y 相互独立, 且具有有限的方差, 试证明:

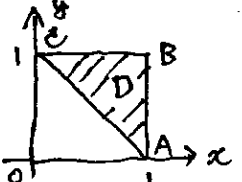
$$D(XY) = D(X)D(Y) + [E(X)]^2 D(Y) + [E(Y)]^2 D(X).$$

由此说明 $D(XY) \geq D(X)D(Y)$

解: $D(XY) = E[(XY)^2] - [E(XY)]^2 = E(X^2Y^2) - [E(X)E(Y)]^2$
 $= E(X^2)E(Y^2) - [E(X)]^2[E(Y)]^2$
 $= [(E(X^2) - (E(X))^2) + (E(X))^2][(E(Y^2) - (E(Y))^2) + (E(Y))^2] - [E(X)]^2[E(Y)]^2$
 $= [D(X) + (E(X))^2][D(Y) + (E(Y))^2] - [E(X)]^2[E(Y)]^2$
 $= D(X)D(Y) + [E(X)]^2 D(Y) + [E(Y)]^2 D(X)$

$$\because [E(X)]^2 D(Y) \geq 0, [E(Y)]^2 D(X) \geq 0, \therefore D(XY) \geq D(X)D(Y).$$

5 设随机变量 X 与 Y 的联合分布在以 $(0, 1)$, $(1, 0)$, $(1, 1)$ 为顶点的三角形区域上服从均匀分布, 试求随机变量 $U = X + Y$ 的方差.

解  $f(x, y) = \begin{cases} 2, & 0 \leq 1-x < y < 1; \\ 0, & \text{其他.} \end{cases} \quad (S_{\triangle ABC} = \frac{1}{2})$

$$D(X+Y) = D(X) + D(Y) + 2[E(XY) - E(X)E(Y)]$$

$$E(X) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} x f(x, y) dx dy = \iint_D x \cdot 2 dx dy = \int_0^1 dx \int_{1-x}^1 2x dy = \frac{2}{3};$$

$$E(Y) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} y f(x, y) dx dy = \iint_D y \cdot 2 dx dy = \int_0^1 dx \int_{1-x}^1 2y dy = \frac{2}{3};$$

$$E(X^2) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} x^2 f(x, y) dx dy = \iint_D x^2 \cdot 2 dx dy = \int_0^1 dx \int_{1-x}^1 2x^2 dy = \frac{1}{2};$$

$$E(Y^2) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} y^2 f(x, y) dx dy = \iint_D y^2 \cdot 2 dx dy = \int_0^1 dx \int_{1-x}^1 2y^2 dy = \frac{1}{2};$$

$$E(XY) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f(x, y) dx dy = \iint_D xy \cdot 2 dx dy = \int_0^1 dx \int_{1-x}^1 2xy dy = \frac{5}{12}.$$

$$\therefore D(X) = E(X^2) - (E(X))^2 = \frac{1}{2} - \left(\frac{2}{3}\right)^2 = \frac{1}{18};$$

$$D(Y) = E(Y^2) - (E(Y))^2 = \frac{1}{2} - \left(\frac{2}{3}\right)^2 = \frac{1}{18};$$

$$\begin{aligned}\therefore D(X+Y) &= D(X) + D(Y) + 2[E(XY) - E(X)E(Y)] \\ &= \frac{1}{18} + \frac{1}{18} + 2\left[\frac{5}{12} - \frac{2}{3} \cdot \frac{2}{3}\right] = \frac{2}{18} - \frac{1}{18} = \frac{1}{18}.\end{aligned}$$

6. 设 X, Y 为随机变量, $E(X)=1, D(X)=1; E(Y)=2, D(Y)=4; \rho_{XY}=\frac{1}{2}$.

记 $Z = \frac{X}{2} + \frac{Y}{3}$, 求 $E(Z), D(Z), \text{CoV}(X, Z)$.

$$\text{解: } E(Z) = E\left(\frac{X}{2} + \frac{Y}{3}\right) = \frac{1}{2}E(X) + \frac{1}{3}E(Y) = \frac{1}{2} \times 1 + \frac{1}{3} \times 2 = \frac{7}{6};$$

$$\begin{aligned}D(Z) &= D\left(\frac{X}{2} + \frac{Y}{3}\right) = \frac{1}{4}D(X) + \frac{1}{9}D(Y) + 2\left[\frac{1}{6}E(XY) - \frac{1}{6}E(X)E(Y)\right] \\ &= \frac{1}{4} \times 1 + \frac{1}{9} \times 4 + 2 \times \frac{1}{6}(E(XY) - E(X)E(Y)) = \frac{37}{36}\end{aligned}$$

$$\text{其中 } E(XY) - E(X)E(Y) = \text{CoV}(X, Y) = \rho_{XY} \sqrt{D(X)} \sqrt{D(Y)} = \frac{1}{2} \times 1 \times 2 = 1.$$

$$\begin{aligned}\text{CoV}(X, Z) &= E(XZ) - E(X)E(Z) = E\left(X\left(\frac{X}{2} + \frac{Y}{3}\right)\right) - E(X)E(Z) \\ &= \frac{1}{2}E(X^2) + \frac{1}{3}E(XY) - E(X)E(Z) = \frac{1}{2} \times 2 + \frac{1}{3} \times 3 - 1 \times \frac{7}{6} = \frac{5}{6};\end{aligned}$$

$$\text{其中 } E(XY) = \text{CoV}(X, Y) + E(X)E(Y) = 1 + 1 \times 2 = 3.$$

7. 设随机变量 X 与 Y 相互独立, $D(X)=4D(Y), U=2X+3Y, V=2X-3Y$ 求 ρ_{UV} .

$$\text{解: } E(U) = E(2X+3Y) = 2E(X) + 3E(Y);$$

$$E(V) = E(2X-3Y) = 2E(X) - 3E(Y);$$

$$E(U^2) = E((2X+3Y)^2) = E(4X^2 + 9Y^2 + 12XY) = 4E(X^2) + 9E(Y^2) + 12E(XY);$$

$$E(V^2) = E((2X-3Y)^2) = E(4X^2 + 9Y^2 - 12XY) = 4E(X^2) + 9E(Y^2) - 12E(XY);$$

$$E(UV) = E[(2X+3Y)(2X-3Y)] = E(4X^2 - 9Y^2) = 4E(X^2) - 9E(Y^2);$$

$$\begin{aligned}\therefore \text{CoV}(U, V) &= E(UV) - E(U)E(V) = (4E(X^2) - 9E(Y^2)) - (2E(X))^2 - (3E(Y))^2 \\ &= 4(E(X^2) - (E(X))^2) - 9(E(Y^2) - (E(Y))^2) = 4D(X) - 9D(Y);\end{aligned}$$

$$\begin{aligned}D(U) &= E(U^2) - (E(U))^2 = (4E(X^2) + 9E(Y^2) + 12E(XY)) - (2E(X) + 3E(Y))^2 \\ &= 4(E(X^2) - (E(X))^2) + 9(E(Y^2) - (E(Y))^2) = 4D(X) + 9D(Y);\end{aligned}$$

$$\begin{aligned}D(V) &= E(V^2) - (E(V))^2 = (4E(X^2) + 9E(Y^2) - 12E(XY)) - (2E(X) - 3E(Y))^2 \\ &= 4(E(X^2) - (E(X))^2) + 9(E(Y^2) - (E(Y))^2) = 4D(X) + 9D(Y);\end{aligned}$$

$$\begin{aligned}\therefore \rho_{UV} &= \frac{\text{CoV}(U, V)}{\sqrt{D(U)} \cdot \sqrt{D(V)}} = \frac{4D(X) - 9D(Y)}{\sqrt{(4D(X) + 9D(Y))^2}} = \frac{4D(X) - 9D(Y)}{4D(X) + 9D(Y)} = \frac{16D(Y) - 9D(Y)}{16D(Y) + 9D(Y)} \\ &= \frac{7}{25}.\end{aligned}$$

8. 设 A, B 为两个随机事件. 随机变量

$$X = \begin{cases} 1, & \text{若 } A \text{ 发生;} \\ -1, & \text{若 } A \text{ 不发生.} \end{cases} \quad Y = \begin{cases} 1, & \text{若 } B \text{ 发生;} \\ -1, & \text{若 } B \text{ 不发生.} \end{cases}$$

试证明随机变量 X 和 Y 不相关的充分必要条件是随机事件 A 和 B 相互独立.

证: $P(X=1) = P(A); P(X=-1) = 1 - P(A);$

$$P(Y=1) = P(B); P(Y=-1) = 1 - P(B).$$

$$E(X) = 1 \cdot P(X=1) + (-1) \cdot P(X=-1) = P(A) - (1 - P(A)) = 2P(A) - 1;$$

$$E(Y) = 1 \cdot P(Y=1) + (-1) \cdot P(Y=-1) = P(B) - (1 - P(B)) = 2P(B) - 1.$$

$$\begin{aligned} E(XY) &= (-1)(-1)P(X=-1, Y=-1) + (-1)(1)P(X=-1, Y=1) \\ &\quad + (1)(-1)P(X=1, Y=-1) + (1)(1)P(X=1, Y=1) \\ &= P(X=-1, Y=-1) - P(X=-1, Y=1) - P(X=1, Y=-1) + P(X=1, Y=1) \\ &= P(\bar{A} \cdot \bar{B}) - P(\bar{A} \cdot B) - P(A \cdot \bar{B}) + P(A \cdot B) \\ &= P(\overline{A \cup B}) - [P(B) - P(AB)] - [P(A) - P(AB)] + P(AB) \\ &= 1 - (P(A) + P(B) - P(AB)) - (P(B) - P(AB)) - (P(A) - P(AB)) + P(AB) \\ &= 1 - 2P(A) - 2P(B) + 4P(AB) \end{aligned}$$

$$\begin{aligned} \text{则 } \text{Cov}(X, Y) &= E(XY) - E(X)E(Y) = (1 - 2P(A) - 2P(B) + 4P(AB)) \\ &\quad - (2P(A) - 1)(2P(B) - 1) = 4(P(AB) - P(A)P(B)). \end{aligned}$$

故当 A, B 相互独立时, 有 $P(AB) = P(A)P(B)$, 则 $\text{Cov}(X, Y) = 0, \therefore \rho_{XY} = 0$
即 X, Y 不相关;

而当 X, Y 不相关时, 有 $\rho_{XY} = 0, \therefore \text{Cov}(X, Y) = 0$,

即有 $P(AB) - P(A)P(B) = 0$, 亦即 $P(AB) = P(A)P(B)$.

$\therefore A, B$ 相互独立.

习题 5-1

1. 设随机变量 X 的数学期望为 $E(X)$, 已知方差 $D(X) = 0.009$, 若用切比雪夫不等式可估出 $P\{|X - E(X)| < \varepsilon\} \geq 0.9$, 试问 ε 的最小值是多少?

解: 切比雪夫不等式 $P\{|X - E(X)| < \varepsilon\} \geq 1 - \frac{D(X)}{\varepsilon^2}$ 取等号时为最小.
 则由 $1 - \frac{D(X)}{\varepsilon^2} = 0.9$, 即 $\varepsilon^2 = \frac{D(X)}{0.1} = \frac{9}{100} = 0.09$
 $\therefore \varepsilon = 0.3$.

2. 设随机变量 X 和 Y 的数学期望分别为 -2 和 2 , 方差分别为 1 和 4 , 相关系数为 -0.5 . 试根据切比雪夫不等式求 $P(|X + Y| \geq 6)$ 的近似值.

解: 切比雪夫不等式 $P\{|(X+Y) - E(X+Y)| \geq \varepsilon\} \leq \frac{D(X+Y)}{\varepsilon^2}$
 $E(X+Y) = E(X) + E(Y) = (-2) + 2 = 0$
 $D(X+Y) = D(X) + D(Y) + 2\text{Cov}(X, Y) = D(X) + D(Y) + 2\rho_{XY}\sqrt{D(X)D(Y)}$
 $= 1 + 4 + 2(-\frac{1}{2})\sqrt{1 \times 4} = 3$
 $\therefore P\{|X+Y| \geq \varepsilon\} \leq \frac{3}{\varepsilon^2}$, 故 $P\{|X+Y| \geq 6\} \leq \frac{3}{6^2} = \frac{1}{12}$.

3. 设随机变量 X 的概率密度函数为 $f(x) = \begin{cases} e^{-x}, & x > 0; \\ 0, & x \leq 0. \end{cases}$

(1) 求 $P(|X - E(X)| \geq \frac{3}{2})$;

(2) 利用切比雪夫不等式求 $P(|X - E(X)| \geq \frac{3}{2})$ 的上界;

(3) 试比较 (1)、(2) 的结果.

解: (1) $E(X) = \int_{-\infty}^{+\infty} x f(x) dx = \int_0^{+\infty} x e^{-x} dx = x(-e^{-x}) \Big|_0^{+\infty} + \int_0^{+\infty} e^{-x} dx = \int_0^{+\infty} e^{-x} dx$
 $= (-e^{-x}) \Big|_0^{+\infty} = 1$.

$E(X^2) = \int_{-\infty}^{+\infty} x^2 f(x) dx = \int_0^{+\infty} x^2 e^{-x} dx = x^2(-e^{-x}) \Big|_0^{+\infty} + 2 \int_0^{+\infty} x e^{-x} dx = 2 \int_0^{+\infty} x e^{-x} dx$
 $= 2$.

$D(X) = E(X^2) - (E(X))^2 = 2 - 1 = 1$.

$\therefore P(|X - E(X)| \geq \frac{3}{2}) = P(|X - 1| \geq \frac{3}{2}) = P((X \geq \frac{5}{2}) \cup (X \leq -\frac{1}{2}))$
 $= P(X \geq \frac{5}{2}) + P(X \leq -\frac{1}{2}) = \int_{\frac{5}{2}}^{+\infty} e^{-x} dx + \int_{-\infty}^{-\frac{1}{2}} 0 dx = \int_{\frac{5}{2}}^{+\infty} e^{-x} dx$
 $= (-e^{-x}) \Big|_{\frac{5}{2}}^{+\infty} = e^{-\frac{5}{2}} \approx 0.082$

$$(2) P\{|X - E(X)| \geq \frac{3}{2}\} \leq \frac{D(X)}{(\frac{3}{2})^2} = \frac{1}{(\frac{3}{2})^2} = \frac{4}{9} = 0.44.$$

(3). 0.082 是 $P(|X - E(X)| \geq \frac{3}{2})$ 的精确值, 而 0.44 为 $P(|X - E(X)| \geq \frac{3}{2})$ 的估计值, 且为最大估计值.

4. 若随机变量 X 服从 $[-1, b]$ 上的均匀分布, 且由切比雪夫不等式得 $P\{|X - 1| < \varepsilon\} \geq \frac{2}{3}$, 求数 b 和 ε .

解: $f(x) = \begin{cases} \frac{1}{1+b}, & -1 \leq x \leq b \\ 0, & \text{其他} \end{cases}$; $\therefore P(|X - 1| < \varepsilon) \geq \frac{2}{3}$, $\therefore E(X) = 1$

$$\text{由 } E(X) = \int_{-\infty}^{+\infty} x f(x) dx = \int_{-1}^b \frac{x}{1+b} dx = \frac{(x^2)^b}{1+b} - 1 = \frac{b^2 - 1}{2(1+b)} = \frac{b-1}{2} = 1$$

得 $b = 3$.

$$E(X^2) = \int_{-\infty}^{+\infty} x^2 f(x) dx = \int_{-1}^3 \frac{x^2}{4} dx = \frac{(x^3)^3}{12} - 1 = \frac{7}{3}$$

$$D(X) = E(X^2) - (E(X))^2 = \frac{7}{3} - 1 = \frac{4}{3}$$

$$\text{由 } \frac{2}{3} = 1 - \frac{D(X)}{\varepsilon^2}, \quad \varepsilon^2 = 3D(X), \quad \therefore \varepsilon^2 = 4, \quad \varepsilon = 2.$$

5. 设随机变量 $X_1, X_2, \dots, X_n, \dots$ 相互独立, 且有分布律

X_n	$-na$	0	na
P	$\frac{1}{2n^2}$	$1 - \frac{1}{n^2}$	$\frac{1}{2n^2}$

证明: $\lim_{n \rightarrow +\infty} P(|\frac{1}{n} \sum_{i=1}^n X_i| > \varepsilon) = 0.$

$$\text{解: } E(X_n) = (-na) \cdot \frac{1}{2n^2} + 0 \cdot (1 - \frac{1}{n^2}) + (na) \cdot \frac{1}{2n^2} = 0; \quad n = 1, 2, 3, \dots$$

$$E(X_n^2) = (-na)^2 \cdot \frac{1}{2n^2} + 0^2 \cdot (1 - \frac{1}{n^2}) + (na)^2 \cdot \frac{1}{2n^2} = a^2, \quad n = 1, 2, 3, \dots$$

$$D(X_n) = E(X_n^2) - (E(X_n))^2 = a^2.$$

$n = 1, 2, 3, \dots$

$$E(\frac{1}{n} \sum_{i=1}^n X_i) = \frac{1}{n} \sum_{i=1}^n E(X_i) = 0; \quad D(\frac{1}{n} \sum_{i=1}^n X_i) = \frac{1}{n^2} \sum_{i=1}^n D(X_i) = \frac{a^2}{n}$$

由切比雪夫不等式:

$$P\{|\frac{1}{n} \sum_{i=1}^n X_i| > \varepsilon\} = P\{|\frac{1}{n} \sum_{i=1}^n X_i - E(\frac{1}{n} \sum_{i=1}^n X_i)| > \varepsilon\} < \frac{1}{\varepsilon^2} D(\frac{1}{n} \sum_{i=1}^n X_i)$$

$$\text{即 } P\{|\frac{1}{n} \sum_{i=1}^n X_i| > \varepsilon\} \leq \frac{1}{\varepsilon^2} \cdot \frac{a^2}{n}$$

$$\text{则 } \lim_{n \rightarrow +\infty} P\{|\frac{1}{n} \sum_{i=1}^n X_i| > \varepsilon\} \leq \lim_{n \rightarrow +\infty} \frac{1}{\varepsilon^2} \cdot \frac{a^2}{n} = 0.$$

$$\text{由 } P\{|\frac{1}{n} \sum_{i=1}^n X_i| > \varepsilon\} \geq 0,$$

$$\therefore \lim_{n \rightarrow +\infty} P\{|\frac{1}{n} \sum_{i=1}^n X_i| > \varepsilon\} = 0.$$

习题 5-2

1. 设 $X_i (i=1, 2, \dots, 50)$ 是相互独立的随机变量, 且它们都服从参数为 $\lambda = 0.03$ 的泊松分布, 记 $Z = X_1 + X_2 + \dots + X_{50}$. 试利用中心极限定理求 $P(Z \geq 3)$ 的近似值.

解: $P\{X_i = k\} = \frac{\lambda^k}{k!} e^{-\lambda}$; 即 $P\{X_i = k\} = \frac{0.03^k}{k!} e^{-0.03}, k=0, 1, 2, \dots, i=1, 2, \dots, 50$.
有 $E(X_i) = \lambda, D(X_i) = \lambda$. 即 $E(X_i) = 0.03, D(X_i) = 0.03$.

$$\therefore E(Z) = E\left(\sum_{i=1}^{50} X_i\right) = \sum_{i=1}^{50} E(X_i) = 50 \times 0.03 = 1.5,$$

$$D(Z) = D\left(\sum_{i=1}^{50} X_i\right) = \sum_{i=1}^{50} D(X_i) = 50 \times 0.03 = 1.5.$$

由 $X_i (i=1, 2, \dots, 50)$ 相互独立服从同一分布, 且 $E(X_i) = 0.03, D(X_i) = 0.03$ 满足中心极限定理. 故知 Z 近似服从 $N(1.5, 1.5)$ 分布.

$$\therefore P(Z \geq 3) = 1 - F_Z(3) = 1 - \Phi\left(\frac{3-1.5}{\sqrt{1.5}}\right) = 1 - \Phi(1.2247) \\ = 1 - 0.8888 = 0.1112.$$

2. 设随机变量 X_1, X_2, \dots, X_{100} 相互独立, 且都服从相同的指数分布, 概率密度函数为

$$f(x) = \begin{cases} \frac{1}{2} e^{-\frac{1}{2}x}, & x > 0; \\ 0, & x \leq 0. \end{cases} \text{ 试求概率 } P\left(\sum_{i=1}^{100} X_i < 240\right) \text{ 的近似值.}$$

解: X_1, \dots, X_{100} 满足中心极限定理条件, 由 $E(X_i) = 2, D(X_i) = 4$.

$$\therefore E\left(\sum_{i=1}^{100} X_i\right) = 100 \times 2 = 200; D\left(\sum_{i=1}^{100} X_i\right) = 100 \times 4 = 400.$$

由中心极限定理知: $\sum_{i=1}^{100} X_i$ 近似服从 $N(200, 400)$ 分布.

$$\therefore P\left(\sum_{i=1}^{100} X_i < 240\right) = F(240) = \Phi\left(\frac{240-200}{\sqrt{400}}\right) = \Phi(2) = 0.9772.$$

3. 设发芽率为 0.95 的一批种子中随机取出 400 粒. 试求其不发芽的种子不多于 25 粒的概率.

解: 设 $X_i = \begin{cases} 1, & \text{第 } i \text{ 粒种子不发芽;} \\ 0, & \text{第 } i \text{ 粒种子发芽;} \end{cases} i=1, 2, \dots, 400. \text{ 相互独立.}$

$$P\{X_i = 1\} = 1 - 0.95 = 0.05; P\{X_i = 0\} = 0.95;$$

$$\therefore E(X_i) = 1 \times 0.05 + 0 \times 0.95 = 0.05;$$

$$E(X_i^2) = 1^2 \times 0.05 + 0^2 \times 0.95 = 0.05;$$

$$\therefore D(X_i) = E(X_i^2) - (E(X_i))^2 = 0.05 - (0.05)^2 = 0.0475.$$

$$E\left(\sum_{i=1}^{400} X_i\right) = 400 \times 0.05 = 20; D\left(\sum_{i=1}^{400} X_i\right) = 400 \times 0.0475 = 19$$

由 X_1, \dots, X_{400} 服从中心极限定理, 所以 $\sum_{i=1}^{400} X_i$ 近似服从正态分布 $N(20, 19)$ 则 $\sum_{i=1}^{400} X_i =$ "400粒中不发芽的种子粒数".

$$P\left(\sum_{i=1}^{400} X_i \leq 25\right) = F(25) = \Phi\left(\frac{25-20}{\sqrt{19}}\right) = \Phi(1.15) = 0.8749.$$

4. 某学校有20000名住校生, 每人以80%的概率去该校食堂就餐, 每个学生去就餐相互独立. 问食堂至少设多少个座位, 才能以99%的概率保证去就餐的同学都有座位?

解: $X_i = \begin{cases} 1, & \text{第 } i \text{ 名学生去就餐;} \\ 0, & \text{第 } i \text{ 名学生没有去就餐.} \end{cases} \quad i=1, 2, \dots, 20000, \text{ 相互独立,}$

$$P(X_i=1) = 0.8; P(X_i=0) = 0.2, \quad i=1, \dots, 20000. \quad E(X_i) = 0.8, D(X_i) = 0.8 \times 0.2$$

$X_1, X_2, \dots, X_{20000}$ 服从中心极限定理: $\sum_{i=1}^{20000} X_i =$ "就餐学生人数"

$\sum_{i=1}^{20000} X_i$ 近似服从 $N(E(\sum_{i=1}^{20000} X_i), D(\sum_{i=1}^{20000} X_i))$ 分布.

$$E\left(\sum_{i=1}^{20000} X_i\right) = \sum_{i=1}^{20000} E(X_i) = 20000 \times E(X_i) = 20000 \times 0.8 = 16000;$$

$$D\left(\sum_{i=1}^{20000} X_i\right) = \sum_{i=1}^{20000} D(X_i) = 20000 \times D(X_i) = 20000 \times 0.8 \times 0.2 = 3200.$$

即 $\sum_{i=1}^{20000} X_i \sim N(16000, 3200)$. 设至少设 x 个座位.

$$\text{则 } P\left\{\sum_{i=1}^{20000} X_i \leq x\right\} = F(x) = \Phi\left(\frac{x-16000}{\sqrt{3200}}\right) = 0.99.$$

$$\therefore \frac{x-16000}{\sqrt{3200}} = 2.325. \quad x = 16000 + 2.325 \sqrt{3200} = 16131.5$$

即至少设16132个座位.

5. 设一条自动生产线的产品合格率是0.8, 要使一批产品的合格率在76%与84%之间的概率不小于90%, 试用

(1) 切比雪夫不等式; (2) 中心极限定理两种方法求这批产品至少要生产多少件? 试比较两种方法.

解: $X_i = \begin{cases} 1, & \text{第 } i \text{ 个产品合格;} \\ 0, & \text{第 } i \text{ 个产品不合格;} \end{cases} \quad i=1, 2, \dots, n \text{ 相互独立.}$

$$P(X_i=1) = 0.8, \quad P(X_i=0) = 0.2.$$

$$E(X_i) = 0.8, \quad D(X_i) = 0.8 \times 0.2 = 0.16.$$

$\sum_{i=1}^n X_i =$ "n个产品中的合格品数". 则合格率为 $\frac{\sum_{i=1}^n X_i}{n}$.

$$E\left(\frac{1}{n}\sum_{i=1}^n X_i\right) = \frac{1}{n}\sum_{i=1}^n E(X_i) = 0.8; D\left(\frac{1}{n}\sum_{i=1}^n X_i\right) = \frac{1}{n^2}\sum_{i=1}^n D(X_i) = \frac{0.16}{n}$$

(1) 根据切比雪夫不等式, $P\left(\left|\frac{1}{n}\sum_{i=1}^n X_i - E\left(\frac{1}{n}\sum_{i=1}^n X_i\right)\right| < \varepsilon\right) > 1 - \frac{D\left(\frac{1}{n}\sum_{i=1}^n X_i\right)}{\varepsilon^2}$

$$\text{即 } P\{0.8 - \varepsilon < \frac{1}{n}\sum_{i=1}^n X_i < 0.8 + \varepsilon\} > 1 - \frac{0.16}{\varepsilon^2 n}$$

现求 $P\{0.76 < \frac{1}{n}\sum_{i=1}^n X_i < 0.84\} > 0.9$ 中的 n . 对照上式得 $\varepsilon = 0.04$. 且 $1 - \frac{0.16}{(0.04)^2 n} = 0.9$.

$$\text{得 } n = 0.16 / (0.1 \times 0.0016) = 1000.$$

(2) 根据中心极限定理: $\frac{1}{n}\sum_{i=1}^n X_i \xrightarrow{D} N(0.8, \frac{0.16}{n})$.

$$\therefore P\{0.76 < \frac{1}{n}\sum_{i=1}^n X_i < 0.84\} = \Phi\left(\frac{0.84 - 0.8}{\sqrt{0.16/n}}\right) - \Phi\left(\frac{0.76 - 0.8}{\sqrt{0.16/n}}\right)$$

$$= \Phi(0.1\sqrt{n}) - \Phi(-0.1\sqrt{n}) = 2\Phi(0.1\sqrt{n}) - 1 > 0.9$$

$$\therefore \Phi(0.1\sqrt{n}) > 0.95, \therefore 0.1\sqrt{n} = 1.645, \therefore n = 270.6 \approx 271.$$

由结果可见(2)的结果比(1)更接近于实际.

6. * 某工厂生产的一种产品其次品率为 0.005, 每只产品是否为次品相互独立. 产品按每 100 只包装成一箱, 一箱中若次品数超过 3 只就不能通过验收. 各箱是否能通过验收相互独立. 今有 10000 箱产品. 求多于 25 箱不能通过验收的概率.

解: $X_i = \begin{cases} 1, & \text{第 } i \text{ 个产品为次品;} \\ 0, & \text{第 } i \text{ 个产品为正品.} \end{cases} i = 1, \dots, 100 \text{ 相互独立.}$

$$P\{X_i = 1\} = 0.005, P\{X_i = 0\} = 0.995.$$

$$\sum_{i=1}^{100} X_i = \text{"一箱 100 只产品中的次品数"} \sim B(100, 100 \times 0.005)$$

$$P\left\{\sum_{i=1}^{100} X_i > 3\right\} = 1 - \left[P\left\{\sum_{i=1}^{100} X_i = 0\right\} + P\left\{\sum_{i=1}^{100} X_i = 1\right\} + P\left\{\sum_{i=1}^{100} X_i = 2\right\} + P\left\{\sum_{i=1}^{100} X_i = 3\right\}\right]$$

$$= 1 - \left[C_{100}^0 (0.005)^0 (0.995)^{100} + C_{100}^1 (0.005)^1 (0.995)^{99} + C_{100}^2 (0.005)^2 (0.995)^{98} + C_{100}^3 (0.005)^3 (0.995)^{97}\right]$$

$$= 1 - 0.9983 = 0.0017. \quad (\text{亦可用泊松分布作近似计算})$$

又设 $Y_k = \begin{cases} 1, & \text{第 } k \text{ 箱不通过验收;} \\ 0, & \text{第 } k \text{ 箱通过验收.} \end{cases} k = 1, \dots, 10000, \text{相互独立.}$