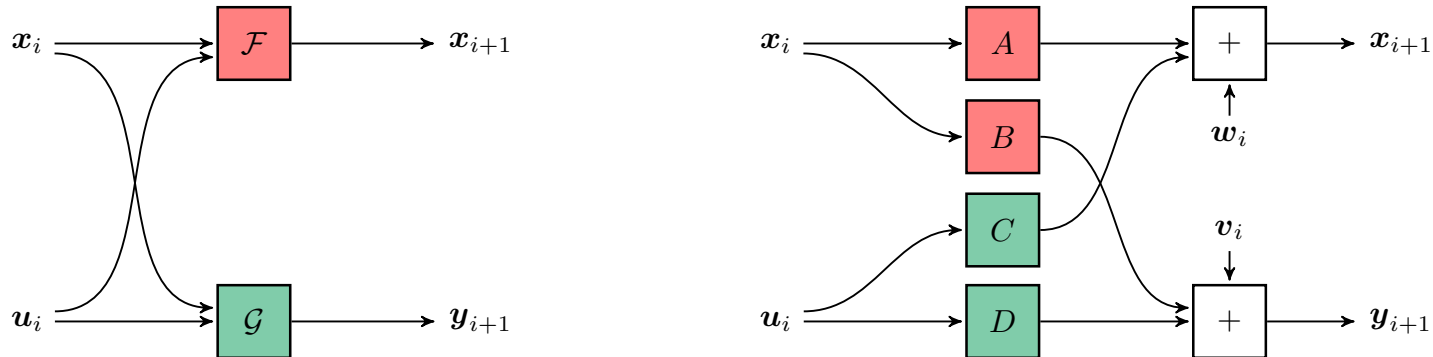


MTAT.03.227 MACHINE LEARNING

**Expectation-Maximisation algorithm
for sequential models**

Sven Laur
University of Tartu

Discrete time systems

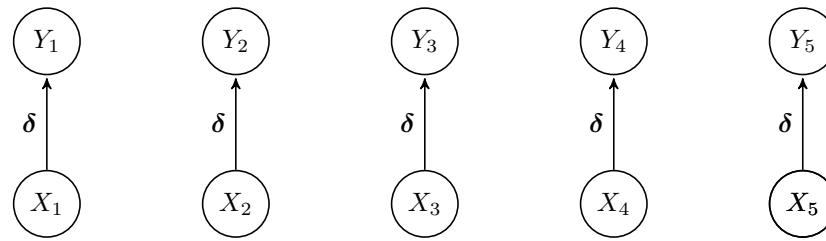


Sequential models describe evolution of discrete time systems.

- ▷ System has an hidden state x_i evolving over state space \mathcal{X} .
- ▷ We can make observations of the system y_i by measuring it.
- ▷ We can influence the system by changing the control signal u_i .
- ▷ For linear system, uncontrollable noise w_i perturbs the state x_{i+1} .
- ▷ For linear system, uncontrollable noise v_i perturbs the observation y_i .

Enforcing temporal consistency

Multinomial mixture model



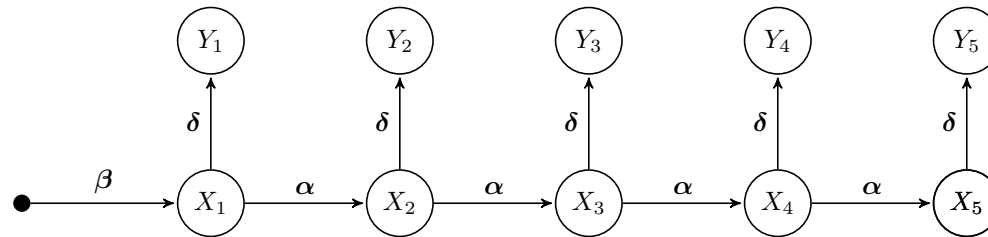
Multinomial mixture model is a discrete time-system.

- ▷ The state space \mathcal{X} is finite.
- ▷ All states x_1, \dots, x_n are independently and identically distributed.
- ▷ Mixture proportions $(\lambda_x)_{x \in \mathcal{X}}$ quantify the corresponding probabilities.
- ▷ Emission matrix $(\delta_{xy})_{x \in \mathcal{X}, y \in \mathcal{Y}}$ the conditional probability of outcomes.

$$\lambda_x = \Pr[x_i = x]$$

$$\delta_{xy} = \Pr[y_i = y | x_i = x]$$

Discrete Hidden Markov Model



Discrete HMM is a refinement of multinomial mixture model.

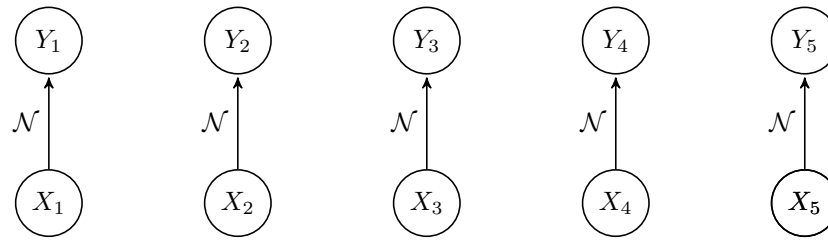
- ▷ State transition probabilities $(\alpha_{x,x'})_{x,x' \in \mathcal{X}}$ are non-trivial.
- ▷ Initial state probabilities $(\beta_x)_{x \in \mathcal{X}}$ become important now.
- ▷ Marginal state probabilities $(\lambda_{xi})_{x \in \mathcal{X}, i \in \mathbb{N}}$ change over time.

$$\beta_x = \Pr[x_1 = x]$$

$$\alpha_{xx'} = \Pr[x_{i+1} = x' | x_i = x]$$

$$\delta_{xy} = \Pr[y_i = y | x_i = x]$$

Gaussian mixture model



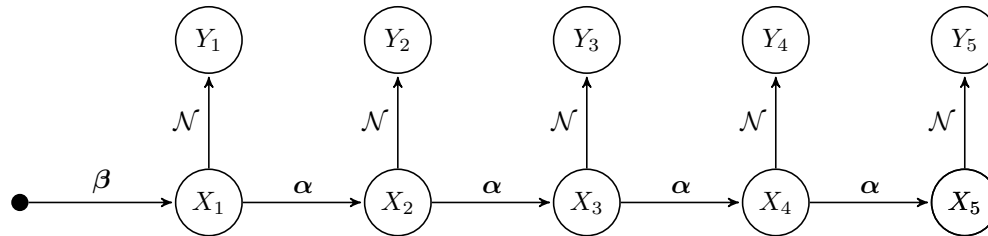
Gaussian mixture model is a discrete time-system.

- ▷ The state space \mathcal{X} is finite.
- ▷ All states x_1, \dots, x_n are independently and identically distributed.
- ▷ Mixture proportions $(\lambda_x)_{x \in \mathcal{X}}$ quantify the corresponding probabilities.
- ▷ Observations \mathbf{y}_i are determined by multivariate normal distributions.

$$\lambda_x = \Pr [x_i = x]$$

$$\mathbf{y}_i \sim \mathcal{N}(\boldsymbol{\mu}_{x_i}, \boldsymbol{\Sigma}_{x_i})$$

Continuous Hidden Markov Model



Continuous HMM is a refinement of Gaussian mixture model.

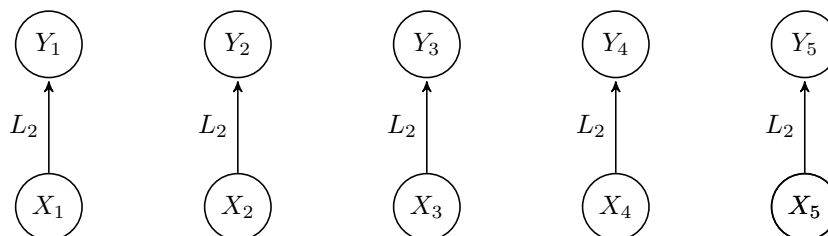
- ▷ State transition probabilities $(\alpha_{x,x'})_{x,x' \in \mathcal{X}}$ are non-trivial.
- ▷ Initial state probabilities $(\beta_x)_{x \in \mathcal{X}}$ become important now.
- ▷ Marginal state probabilities $(\lambda_{xi})_{x \in \mathcal{X}, i \in \mathbb{N}}$ change over time.

$$\beta_x = \Pr[x_1 = x]$$

$$\alpha_{xx'} = \Pr[x_{i+1} = x' | x_i = x]$$

$$y_i \sim \mathcal{N}(\mu_{x_i}, \Sigma_{x_i})$$

Multivariate linear transformations



Multivariate linear transformation is a discrete time-system.

- ▷ The merged input and state space $\mathbb{R}^d \times \mathbb{R}^k$ is infinite.
- ▷ All states $\mathbf{x}_1, \dots, \mathbf{x}_n$ are independently and identically distributed.
- ▷ States and observations disturbed by white gaussian noise.
- ▷ Observations \mathbf{y}_i are linear in inputs \mathbf{u}_i and states \mathbf{x}_i .

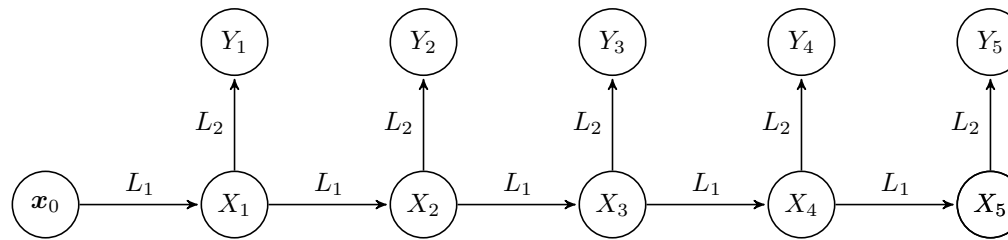
$$\mathbf{x}_{i+1} = \mathbf{0}\mathbf{x}_i + \mathbf{0}\mathbf{u}_i + \mathbf{w}_i,$$

$$\mathbf{w}_i \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$$

$$\mathbf{y}_i = \mathbf{C}\mathbf{x}_i + \mathbf{D}\mathbf{u}_i + \mathbf{v}_i,$$

$$\mathbf{v}_i \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$$

Kalman filter



Kalman filter is a refinement of multivariate linear transformation.

- ▷ Linear state evolution equation is non-trivial.
- ▷ The initial state x_0 is assumed to be fixed value.
- ▷ States and observations disturbed by colored gaussian noise.

$$\begin{aligned}x_{i+1} &= Ax_i + Bu_i + w_i, & w_i &\sim \mathcal{N}(\mathbf{0}, \Sigma_1) \\ y_i &= Cx_i + Du_i + v_i, & v_i &\sim \mathcal{N}(\mathbf{0}, \Sigma_2)\end{aligned}$$

EM-algorithm for HMM

Lower bound function

The lower bound function used in the EM algorithm is current notation

$$F(\mathbf{q}, \Theta) = - \sum_{\mathbf{x}} q(\mathbf{x}) \cdot \log q(\mathbf{x}) + \sum_{\mathbf{x}} q(\mathbf{x}) \cdot \log (p[\Theta, \mathbf{x}|\mathbf{y}])$$

If we assign non-informative prior to the model parameters then in M-step it is sufficient to maximise the function

$$F_* = \sum_{\mathbf{x}} q(\mathbf{x}) \cdot \log (p[\mathbf{y}, \mathbf{x}|\Theta])$$

Probability assignment in E-step

According to the theory the optimal probability assignment is

$$q(\mathbf{x}) = \Pr [\mathbf{x}|\mathbf{y}, \Theta_*] = \frac{p [\mathbf{y}, \mathbf{x}|\Theta_*]}{p [\mathbf{y}|\Theta_*]}$$

and thus we get

$$F_* = \frac{1}{p [\mathbf{y}|\Theta_*]} \cdot \sum_{\mathbf{x}} p [\mathbf{y}, \mathbf{x}|\Theta_*] \cdot \log (p [\mathbf{y}, \mathbf{x}|\Theta])$$

and thus it is sufficient to maximise

$$Q = \sum_{\mathbf{x}} p [\mathbf{y}, \mathbf{x}|\Theta_*] \cdot \log (p [\mathbf{y}, \mathbf{x}|\Theta])$$

Further decomposition

As the log-likelihood decomposes into three independent parameter groups

$$\log(p[\mathbf{y}, \mathbf{x} | \Theta]) = \log \beta_{x_1} + \sum_{i=2}^n \log(\alpha_{x_{i-1}x_i}) + \sum_{i=1}^n \log(p[y_i | x_i])$$

we can solve three independent maximisation tasks in the M-step:

$$Q_1 = \sum_{\mathbf{x}} p[\mathbf{y}, \mathbf{x} | \Theta_*] \cdot \log \beta_{x_1}$$

$$Q_2 = \sum_{\mathbf{x}} p[\mathbf{y}, \mathbf{x} | \Theta_*] \cdot \sum_{i=2}^n \log(\alpha_{x_{i-1}x_i})$$

$$Q_3 = \sum_{\mathbf{x}} p[\mathbf{y}, \mathbf{x} | \Theta_*] \cdot \sum_{i=1}^n \log(p[y_i | x_i])$$

Simplification of the first term

As

$$\begin{aligned} Q_1 &= \sum_{\mathbf{x}} p[y_1, x_1 | \Theta_*] p[y_2 \dots, y_n, x_2, \dots, x_n | x_1, \Theta_*] \cdot \log \beta_{x_1} \\ &= \sum_{x_1} p[y_1, x_1 | \Theta_*] \cdot p[y_2, \dots, y_n | x_1, \Theta] \cdot \log \beta_{x_1} \\ &= \sum_{x_1} p[\mathbf{y}, x_1 | \Theta_*] \cdot \log \beta_{x_1} \end{aligned}$$

we can establish

$$\beta_x = \frac{p[\mathbf{y}, x_1 = x | \Theta_*]}{p[\mathbf{y} | \Theta_*]} = p[x_1 = x | \Theta_*, \mathbf{y}]$$

Simplification of the second term

For the term

$$Q_2 = \sum_{i=2}^n \sum_{\mathbf{x}} p[y_1, x_1 | \Theta_*] \cdot \prod_{j=2}^n p[y_j, x_j | x_{j-1}, \Theta_*] \cdot \log(\alpha_{x_{i-1}x_i})$$

we can use general equality $\sum_{\mathbf{x}} \prod_{\ell=1}^n a_{\ell x_\ell} = \prod_{\ell=1}^n \sum_{j=1}^k a_{\ell j}$ for getting

$$Q_2 = \sum_{i=2}^n \sum_{x \in \mathcal{X}} \sum_{x' \in \mathcal{X}} \log \alpha_{x, x'} \cdot p[\mathbf{y}, x_{i-1} = x, x_i = x' | \Theta_*]$$

The latter allows to establish

$$\alpha_{xx'} = \frac{\sum_{i=2}^n \Pr[\mathbf{y}, x_{i-1} = x, x_i = x' | \Theta_*]}{\sum_{i=2}^n \Pr[\mathbf{y}, x_{i-1} = x | \Theta_*]}$$

Simplification of the third term

For the third term

$$Q_3 = \sum_{i=1}^n \sum_{\mathbf{x}} p[y_1, x_1 | \Theta_*] \cdot \prod_{j=2}^n p[y_j, x_j | x_{j-1}, \Theta_*] \cdot \log(p[y_i | x_i])$$

we can still use the general equality for getting

$$Q_3 = \sum_{i=1}^n \sum_{x \in \mathcal{X}} \log(p[y_i | x_i = x]) \cdot p[\mathbf{y}, x_i = x | \Theta_*]$$

This term is identical to the term we maximise in the clustering algorithm.

Full recipe for discrete HMM

E-step. Compute the following marginal probabilities

$$\gamma_x(i) = \Pr[x_i = x | \mathbf{y}, \Theta]$$

$$\xi_{xx'}(i) = \Pr[x_i = x, x_{i+1} = x' | \mathbf{y}, \Theta]$$

M-step. Compute the following parameters

$$\beta_x = \gamma_x(1)$$

$$\alpha_{xx'} = \frac{\sum_{j=1}^{n-1} \xi_{xx'}(j)}{\sum_{j=1}^{n-1} \gamma_x(j)}$$

$$\delta_{xy} = \frac{\sum_{j=1}^{n-1} \gamma_x(j) \cdot [y_j = y]}{\sum_{j=1}^n \gamma_x(j)}$$

Full recipe for continuous HMM

E-step. Compute the following marginal probabilities

$$\gamma_x(i) = \Pr [x_i = x | \mathbf{y}, \Theta]$$

$$\xi_{xx'}(i) = \Pr [x_i = x, x_{i+1} = x' | \mathbf{y}, \Theta]$$

M-step. Compute the following parameters

$$\beta_x = \gamma_x(1)$$

$$\alpha_{xx'} = \frac{\sum_{j=1}^{n-1} \xi_{xx'}(j)}{\sum_{j=1}^{n-1} \gamma_x(j)}$$

and find parameters $\boldsymbol{\mu}_j, \Sigma_j$ for the normal distribution by doing maximum likelihood fit for the datapoints with weights $w_{ix} = \Pr [x_i = x | \mathbf{y}, \Theta_*]$.

EM-algorithm for Kalman filter

Lower bound function

The lower bound function used in the EM algorithm is current notation

$$F(\mathbf{q}, \boldsymbol{\Theta}) = - \int_{\mathbf{x}} q(\mathbf{x}) \cdot \log q(\mathbf{x}) d\mathbf{x} + \int_{\mathbf{x}} q(\mathbf{x}) \cdot \log (p[\boldsymbol{\Theta}, \mathbf{x} | \mathbf{y}, \mathbf{u}]) d\mathbf{x}$$

If we assign non-informative prior to the model parameters then in M-step it is sufficient to maximise the function

$$F_* = \int_{\mathbf{x}} q(\mathbf{x}) \cdot \log (p[\mathbf{y}, \mathbf{x} | \boldsymbol{\Theta}, \mathbf{u}]) d\mathbf{x}$$

Probability assignment in E-step

According to the theory the optimal probability assignment is

$$q(\mathbf{x}) = \Pr[\mathbf{x}|\mathbf{y}, \mathbf{u}, \Theta_*] = \frac{p[\mathbf{y}, \mathbf{x}|\Theta_*, \mathbf{u}]}{p[\mathbf{y}|\Theta_*, \mathbf{u}]}$$

and thus we get

$$F_* = \frac{1}{p[\mathbf{y}|\Theta_*, \mathbf{u}]} \cdot \int_{\mathbf{x}} p[\mathbf{y}, \mathbf{x}|\Theta_*, \mathbf{u}] \cdot \log(p[\mathbf{y}, \mathbf{x}|\Theta, \mathbf{u}]) d\mathbf{x}$$

and thus it is sufficient to maximise

$$Q = \int_{\mathbf{x}} p[\mathbf{y}, \mathbf{x}|\Theta_*, \mathbf{u}] \cdot \log(p[\mathbf{y}, \mathbf{x}|\Theta, \mathbf{u}]) d\mathbf{x}$$

Further decomposition

As the log-likelihood decomposes into two independent parameter groups

$$\log(p[\mathbf{y}, \mathbf{x} | \Theta, \mathbf{u}]) = \sum_{i=1}^n \log(p[\mathbf{x}_i | \mathbf{x}_{i-1}, \mathbf{u}_{i-1}, \Theta]) + \sum_{i=1}^n \log(p[\mathbf{y}_i | \mathbf{x}_i, \mathbf{u}_i, \Theta])$$

where

$$p[\mathbf{x}_i | \mathbf{x}_{i-1}, \mathbf{u}_{i-1}, \Theta] = p_{\mathcal{N}}[\mathbf{x}_i - A\mathbf{x}_{i-1} - B\mathbf{u}_{i-1} | \Sigma_1]$$
$$p[\mathbf{y}_i | \mathbf{x}_i, \mathbf{u}_i, \Theta] = p_{\mathcal{N}}[\mathbf{y}_i - C\mathbf{x}_i - D\mathbf{u}_i | \Sigma_2]$$

we can solve two independent maximisation tasks in the M-step. Again, finding Q-function seems to be a daunting task but the minimisation task can be reduced to finding marginal distributions as for the HMM.