P8160 - Project 3 Baysian modeling of hurricane

Renjie Wei, Hao Zheng, Xinran Sun Wentong Liu, Shengzhi Luo

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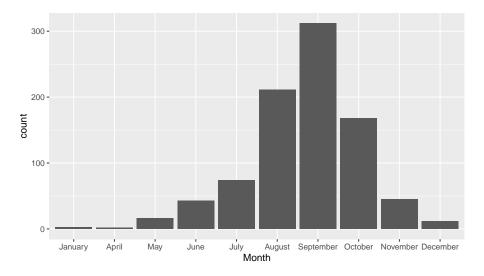
Introduction

- Hurricanes can result in death and economical damage
- There is an increasing desire to predict the speed and damage of the hurricanes
- Use Bayesian Model and Markov Chain Monte Carlo algorithm
 - Predict the wind speed of hurricanes
 - Study how hurricanes is related to death and financial loss

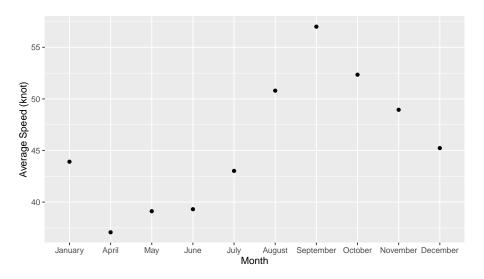
Dataset

- Hurrican703 dataset: 22038 observations × 8 variables
 - ▶ 702 hurricanes in the North Atlantic area in year 1950-2013
- Processed dataset: add 5 more variables into hurrican703
- Hurricanoutcome2 dataset: 43 observations × 14 variables

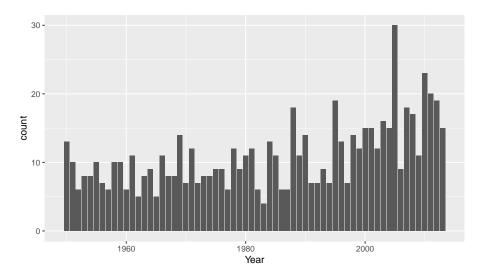
EDA - Count of Hurricanes in Each Month



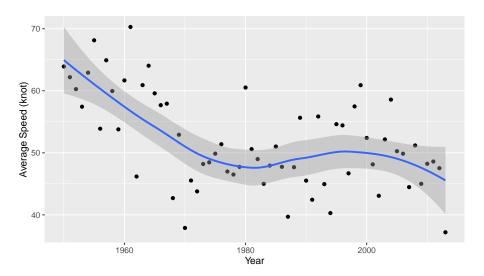
EDA - Average Speed (knot) of Hurricanes in Each Month



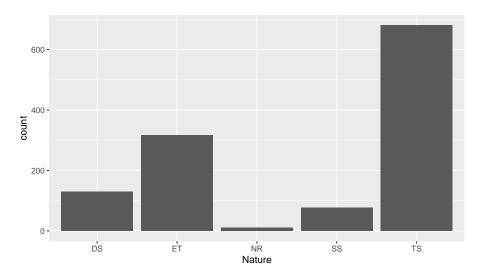
EDA - Count of Hurricanes in Each Year



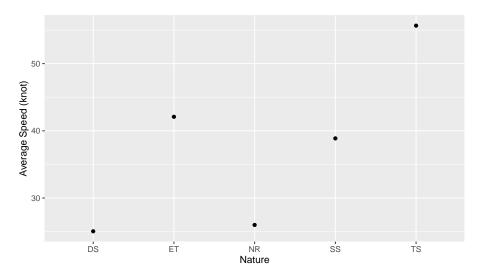
EDA - Average Speed (knot) of Hurricanes in Each Year



EDA - Count of Hurricanes in Each Nature



EDA - Average Speed (knot) of Hurricanes in Each Nature



Bayesian Model Setting

Model

The suggested Bayesian model is

$$Y_i(t+6) = \beta_{0,i} + \beta_{1,i} Y_i(t) + \beta_{2,i} \Delta_{i,1}(t) + \beta_{3,i} \Delta_{i,2}(t) + \beta_{4,i} \Delta_{i,3}(t) + \epsilon_i(t)$$

- where $Y_i(t)$ the wind speed at time t (i.e. 6 hours earlier), $\Delta_{i,1}(t)$, $\Delta_{i,2}(t)$ and $\Delta_{i,3}(t)$ are the changes of latitude, longitude and wind speed between t and t-6, and $\epsilon_{i,t}$ follows a normal distributions with mean zero and variance σ^2 , independent across t.
- $\beta_i=(\beta_{0,i},\beta_{1,i},...,\beta_{5,i})$, we assume that $\beta_i\sim N(\mu,\Sigma_{d\times d})$, where d is dimension of β_i .

Priors

$$P(\sigma^2) \propto \frac{1}{\sigma^2}; \quad P(\mu) \propto 1; \quad P(\Sigma^{-1}) \propto |\Sigma|^{-(d+1)} \exp(-\frac{1}{2}\Sigma^{-1})$$

Posterior

 $\bullet \ \, \mathsf{Derive} \,\, \pi(\Theta|Y) \mathsf{,} \,\, \mathsf{where} \,\, \Theta = (\mathbf{B}^\top, \mu^\top, \sigma^2, \Sigma), \,\, \mathbf{B} = (\boldsymbol{\beta}_1^\top, ..., \boldsymbol{\beta}_n^\top)^\top$

Joint posterior

Notations

- $\bullet \ \, X_i(t)\boldsymbol{\beta}_i^\top = \beta_{0,i} + \beta_{1,i}Y_i(t) + \beta_{2,i}\Delta_{i,1}(t) + \beta_{3,i}\Delta_{i,2}(t) + \beta_{4,i}\Delta_{i,3}(t)$
- \bullet For i^{th} hurricane, there may be m_i times of record (excluding the first and second observation), let

$$\boldsymbol{Y}_i = \begin{pmatrix} Y_i(t_0+6) \\ Y_i(t_1+6) \\ \vdots \\ Y_i(t_{m_i-1}+6) \end{pmatrix}_{m_i \times 1}$$

- $\bullet \ \ \text{Hence,} \ Y_i \mid X_i, \beta_i, \sigma^2 \sim N(X_i \beta_i^\top, \sigma^2 I)$
- Where, X_i is a $m_i \times d$ dimensional matrix

$$X_i = \begin{pmatrix} 1 & Y_i(t_0) & \Delta_{i,1}(t_0) & \Delta_{i,2}(t_0) & \Delta_{i,3}(t_0) \\ 1 & Y_i(t_1) & \Delta_{i,1}(t_1) & \Delta_{i,2}(t_1) & \Delta_{i,3}(t_1) \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ 1 & Y_i(t_{m_i-1}) & \Delta_{i,1}(t_{m_i-1}) & \Delta_{i,2}(t_{m_i-1}) & \Delta_{i,3}(t_{m_i-1}) \end{pmatrix}$$

Joint posterior

Posterior

$$\begin{split} \pi(\Theta|Y) &= \pi(\mathbf{B}^\top, \mu^\top, \sigma^2, \Sigma \mid Y) \\ &\propto \underbrace{\prod_{i=1}^n f(Y_i \mid \beta_i, \sigma^2)}_{\text{likelihood of } Y} \underbrace{\prod_{i=1}^n \pi(\beta_i \mid \mu, \Sigma)}_{\text{distribution of } \mathbf{B}} \underbrace{P(\sigma^2)P(\mu)P(\Sigma^{-1})}_{\text{priors}} \\ &\propto \prod_{i=1}^n \left\{ (2\pi\sigma^2)^{-m_i/2} \exp\big\{ -\frac{1}{2}(Y_i - X_i\beta_i^\top)^\top (\sigma^2I)^{-1}(Y_i - X_i\beta_i^\top) \right. \\ &\times \prod_{i=1}^n \left\{ \det(2\pi\Sigma)^{-\frac{1}{2}} \exp\big\{ -\frac{1}{2}(\beta_i - \mu)\Sigma^{-1}(\beta_i - \mu)^\top \big\} \right\} \\ &\times \frac{1}{\sigma^2} \times \det(\Sigma)^{-(d+1)} \exp\big\{ -\frac{1}{2}\Sigma^{-1} \big\} \end{split}$$

MCMC Algorithm

- Monte Carlo Method
 - ▶ Random sampling method to estimate quantity
- Markov Chain
 - ► Generates a sequence of random variables where the current state only depends on the nearest past
- Example: Gibbs Sampler
 - ▶ MCMC approaches with known conditional distributions
 - Samples from each random variables in turn given the value of all the others in the distribution

Conditional Posterior

- To apply MCMC using Gibbs sampling, we need to find conditional posterior distribution of each parameter, then we can implement Gibbs sampling on these conditional posterior distributions.
 - $\pi(\mathbf{B}|Y,\mu^{\top},\sigma^2,\Sigma)$
 - $\qquad \qquad \pi(\sigma^2|Y,\mathbf{B}^\top,\mu^\top,\Sigma)$
 - $\blacktriangleright \ \pi(\Sigma|Y,\mathbf{B}^\top,\mu^\top,\sigma^2)$
 - $\blacktriangleright \ \pi(\mu|Y,\mathbf{B}^\top,\sigma^2,\Sigma)$

MCMC Algorithm - Conditional Posterior

- $\bullet \ \beta_i \colon \pi(\beta_i | Y, \mu^\top, \sigma^2, \Sigma) \sim \mathcal{N}(\hat{\boldsymbol{\beta}}_i, \hat{\boldsymbol{\Sigma}}_{\beta_i})$
 - $\text{ where } \hat{\boldsymbol{\beta}}_i = (\boldsymbol{\Sigma}^{-1} + \boldsymbol{X}_i^\top (\sigma^2 \boldsymbol{I})^{-1} \boldsymbol{X}_i)^{-1} \boldsymbol{Y}_i^\top (\sigma^2 \boldsymbol{I})^{-1} \boldsymbol{X}_i + \mu \boldsymbol{\Sigma}^{-1}, \hat{\boldsymbol{\Sigma}}_{\boldsymbol{\beta}_i} = (\boldsymbol{\Sigma}^{-1} + \boldsymbol{X}_i^\top (\sigma^2 \boldsymbol{I})^{-1} \boldsymbol{X}_i)^{-1}$
- \bullet σ^2 :

$$\pi(\sigma^2|\boldsymbol{Y}, \mathbf{B}^\top, \boldsymbol{\mu}^\top, \boldsymbol{\Sigma}) \sim IG(\tfrac{1}{2} \sum_{i=1}^n m_i, \tfrac{1}{2} \sum_{i=1}^n (\boldsymbol{Y}_i - \boldsymbol{X}_i \boldsymbol{\beta}_i^\top)^\top (\boldsymbol{Y}_i - \boldsymbol{X}_i \boldsymbol{\beta}_i^\top))$$

- $\bullet \ \Sigma \colon \pi(\Sigma|Y,\mathbf{B}^\top,\mu^\top,\sigma^2) \sim IW(n+d+1,\ I+\sum_{i=1}^n(\beta_i-\mu)(\beta_i-\mu)^\top)$
- $\bullet \ \mu \colon \pi(\mu|Y,\mathbf{B}^{\top},\sigma^2,\Sigma) \sim \mathcal{N}(\tfrac{1}{n} \sum_{i=1}^n \beta_i,\tfrac{1}{n}\Sigma)$

MCMC Algorithm - Parameter Updates

The update of parameters is component wise, at $(t+1)^{\rm th}$ step, updating parameters in the following the order:

- $\textbf{ Sample B}^{(t+1)} \text{, i.e., sample each } \boldsymbol{\beta}_i^{(t+1)} \text{ from } \mathcal{N}(\boldsymbol{\hat{\beta}}_i^{(t)}, \boldsymbol{\hat{\Sigma}}_{\boldsymbol{\beta}_i}^{(t)})$
- $\textbf{2} \ \, \mathsf{Then, sample} \ \, \sigma^2 \ \, \mathsf{from}$

$$IG(\tfrac{1}{2}\sum_{i=1}^{n}m_{i},\tfrac{1}{2}\sum_{i=1}^{n}(\boldsymbol{Y}_{i}-\boldsymbol{X}_{i}\boldsymbol{\beta}_{i}^{(t+1)^{\top}})^{\top}(\boldsymbol{Y}_{i}-\boldsymbol{X}_{i}\boldsymbol{\beta}_{i}^{(t+1)^{\top}}))$$

3 Next, sample $\Sigma^{(t+1)}$ from

$$IW(n+d+1,\ I+\sum_{i=1}^{n}(\beta_{i}^{\ (t+1)}-\mu^{(t)})(\beta_{i}^{\ (t+1)}-\mu^{(t)})^{\top})$$

 $\textbf{ § Finally, sample } \mu^{(t+1)} \text{ from } \mathcal{N}(\frac{1}{n}\sum_{i=1}^n \boldsymbol{\beta_i}^{(t+1)}, \frac{1}{n}\boldsymbol{\Sigma}^{(t+1)})$

MCMC Algorithm - Train-Test split and Inital Values

Train-test split

- Drop the data of hurricane with less than 3 observations. Results in 697 hurricanes
- Within each hurricane's data, randomly 80% train, 20% test

Initial Values

① For initial value of ${\bf B}$, we run multivariate linear regressions for each hurricane and use the regression coefficients β_i^{MLR} as the initial value for β_i . Then, the initial value of ${\bf B}$ can be represented as

$$\mathbf{B}_{init} = (\beta_1^{MLR^{\top}}, \dots, \beta_n^{MLR^{\top}})^{\top}.$$

② For initial value of μ , we take the average of β_i^{MLR} , that is

$$\mu_{init} = \frac{1}{n} \sum_{i=1}^{n} \beta_n^{MLR}$$

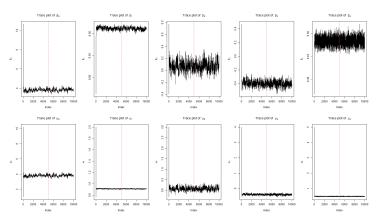
- **3** For initial value of σ^2 , we take the average of the MSE for i hurricanes.
- For initial value of Σ , we just set it to a simple diagonal matrix, i.e. $\Sigma_{i=1} = diaa(1,2,3,4,5)$

MCMC Results

Details

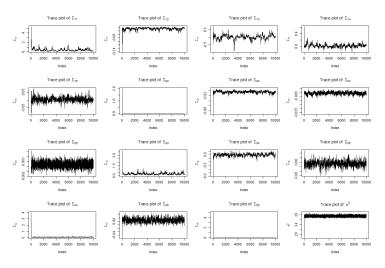
- 10000 iterations
- First 5000 iterations as burn-in period
- Estimates and inferences based on last 5000 MCMC samples

MCMC Results - Trace Plots 1



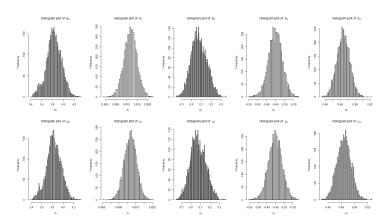
Trace plots of model parameters, based on 10000 MCMC sample

MCMC Results - Trace Plots 2



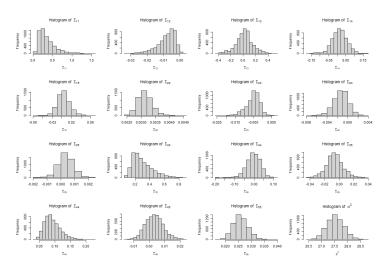
Trace plots of variance parameters, based on 10000 MCMC sample

MCMC Results - Histograms 1



Histograms of model parameters, based on last 5000 MCMC sample

MCMC Results - Histograms 2



Histograms of variance parameters, based on last 5000 MCMC sample

MCMC Results - Model Parameter Estimations and Inferences

Variables	$ar{eta}_i$	$\operatorname{Var}(\bar{\beta}_i)$	95% CI of $\bar{\beta}_i$	$\bar{\mu}$	$\mathrm{Var}(\bar{\mu})$	95% CI of $\bar{\mu}$
intercept	3.8252	0.0185	(3.5588,4.0916)	3.8166	0.0190	(3.5468,4.0865)
Wind_prev	0.9118	0.0000	(0.9059, 0.9177)	0.9121	0.0000	(0.9049, 0.9194)
Lat_change	0.0744	0.0060	(-0.0776, 0.2264)	0.0720	0.0065	(-0.0857, 0.2298)
Long_change	-0.4014	0.0015	(-0.4771,-0.3257)	-0.3968	0.0016	(-0.4759,- 0.3177)
${\bf Wind_change}$	0.4841	0.0001	(0.4674, 0.5009)	0.4847	0.0001	(0.464, 0.5053)

Bayesian posterior estimates for model parameters

MCMC Results - Variance Parameter Estimations and Inferences

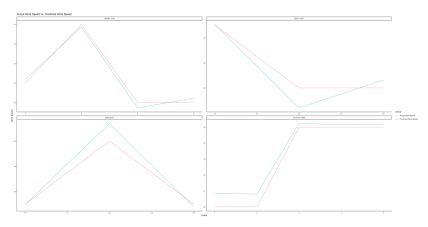
$$\Sigma = \begin{pmatrix} 0.349 & -0.008 & 0.020 & 0.013 & 0.004 \\ -0.008 & 0.003 & -0.005 & -0.001 & 0.0004 \\ 0.020 & -0.005 & 0.296 & -0.003 & -0.006 \\ 0.013 & -0.001 & -0.003 & 0.092 & 0.003 \\ 0.004 & 0.0004 & -0.006 & 0.003 & 0.026 \end{pmatrix}$$

$$\rho = \begin{pmatrix} 1 & -0.245 & 0.063 & 0.073 & 0.037 \\ -0.245 & 1 & -0.174 & -0.078 & 0.041 \\ 0.063 & -0.174 & 1 & -0.019 & -0.069 \\ 0.073 & -0.078 & -0.019 & 1 & 0.070 \\ 0.037 & 0.041 & -0.069 & 0.070 & 1 \end{pmatrix}$$

Bayesian Model Performance

	ID	r_square	rmse
1	GUSTAV.1996	0.952	0.537
2	LORENZO.2001	0.914	0.733
3	ERIN.2013	0.878	0.823
4	JOSE.2011	0.970	0.872
5	GRETA.1970	0.980	0.876
6	DELTA.1972	0.825	0.904
7	EDITH.1967	0.826	0.983
8	FABIAN.1997	0.955	1.002
9	DEBBY.2006	0.984	1.045
10	CRISTOBAL.2002	0.956	1.053

Bayesian Model Performance



Actual Wind Speed vs. Predicted Wind Speed

Seasonal Difference Exploration

	Beta	0	Beta 1		Beta	2	Beta	3	Beta 4	
	Estimate	Pr(> t)								
(Intercept)	4.4810021	0.0000000	1.3431063	0.0000000	0.0413063	0.9506172	-0.8336700	0.0185275	0.2890273	0.4482640
monthApril	0.0232609	0.8346449	0.0147943	0.6696787	0.0165579	0.9306863	0.0416468	0.6796126	0.0361823	0.7393892
monthMay	0.0259813	0.7827813	-0.0001180	0.9967888	0.0708822	0.6597505	0.0632772	0.4581672	-0.0162907	0.8594231
monthJune	0.0275693	0.7650618	0.0053935	0.8509869	-0.0070875	0.9641298	0.0556884	0.5047909	0.0237694	0.7918014
monthJuly	0.0125400	0.8914489	0.0154032	0.5901741	-0.0090910	0.9538180	0.0361214	0.6640154	0.0130817	0.8840332
monthAugust	-0.0198034	0.8284715	0.0233206	0.4124181	-0.0522548	0.7378961	0.0123691	0.8811234	0.0312427	0.7261962
monthSeptember	-0.0070528	0.9384385	0.0261005	0.3585599	-0.0361073	0.8169707	0.0212965	0.7966351	0.0444835	0.6177631
monthOctober	0.0093435	0.9185853	0.0210829	0.4587183	-0.0286163	0.8546050	0.0341549	0.6796975	0.0350505	0.6944480
monthNovember	0.0145692	0.8748155	0.0246144	0.3925264	0.0239972	0.8792681	0.0263450	0.7529105	0.0209069	0.8168323
monthDecember	0.0057977	0.9526542	0.0088244	0.7715305	-0.0543131	0.7447475	0.0422468	0.6326060	0.0114196	0.9046290
year	-0.0003419	0.0717253	-0.0002252	0.0001471	0.0000365	0.9101708	0.0002184	0.2032812	0.0000905	0.6249586
natureET	0.0008449	0.9774141	0.0037334	0.6877086	-0.0702038	0.1687975	-0.0263888	0.3286540	-0.0209217	0.4726774
natureNR	0.0008122	0.9866387	-0.0146142	0.3331114	0.0058967	0.9432660	0.0030556	0.9444979	-0.0217275	0.6462854
natureSS	0.0141564	0.4904257	-0.0033299	0.6021721	-0.0013517	0.9692484	0.0126339	0.4964264	-0.0238538	0.2339965
natureTS	0.0118370	0.4785102	-0.0059979	0.2486925	-0.0154533	0.5880814	-0.0231521	0.1258337	-0.0174987	0.2832214

Seasonal Difference Exploration

year

-0.0003543

0.0497902

-0.0002178

0.0001332

	E	Beta 0		Beta 1		Beta 2		Beta 3		Beta 4	
	Estima	te Pr(> t)	Estimate	Pr(> t)	Estimate	Pr(> t)	Estimate	Pr(> t)	Estimate	Pr(> t)	
(Intercept)	3.836550	0.0000000	0.8942250	0.0000000	0.1606506	0.0000610	-0.3500900	0.0000000	0.4422452	0.0000000	
seasonSummer	-0.03050	0.2048954	0.0152377	0.0440074	-0.0979486	0.0167511	-0.0466127	0.0338037	0.0361669	0.1203099	
seasonAutumn	-0.023534	46 0.3248438	0.0209616	0.0053662	-0.0909590	0.0253577	-0.0434764	0.0463302	0.0487052	0.0354139	
seasonWinter	-0.01865	42 0.6535827	0.0034158	0.7936540	-0.0984181	0.1637856	-0.0094850	0.8023902	0.0149135	0.7107131	
	Beta 0		Beta 1		Beta 2		Beta 3		Beta 4	1	
	Estimate	Pr(> t)	Estimate	Pr(> t)	Estimate	Pr(> t)	Estimate	Pr(> t)	Estimate	Pr(> t)	
(Intercent)	4 5142875	0.0000000	1 3448481	0.0000000	-0.1056332	0.8629385	-1 0267628	0.001781	0.3051312	0.3817170	

0.0000878

0.7757368

0.0003188

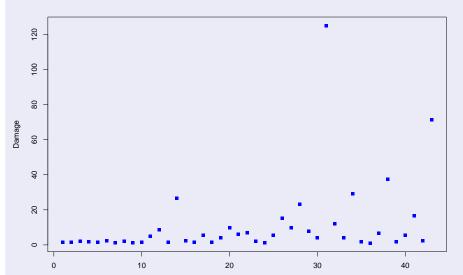
0.053474

0.0000902

0.6072986

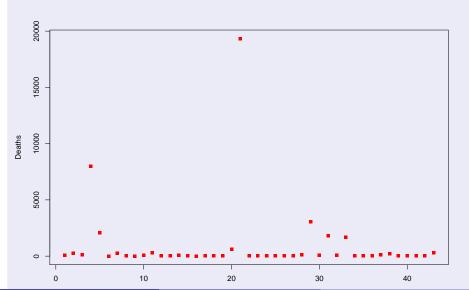
Predictions of Damage and Deaths

Basic plot of Damage and Deaths



Predictions of Damage and Deaths

Basic plot of Damage and Deaths



Generalized Linear Model - Poisson

The poisson model used in predicting deaths and damage is:

$$log(Damage*1000 or Deaths) = \beta_i X_i$$

- ullet where X_i includes $eta_0 \sim eta_4$ and the predictors in new data
- convert Damage units from billions to millions to get integer data

Coefficient Table

Table 1: Coefficient estimates table from Bayesian model

ID	β_0	β_1	β_2	β_3	β_4
agnes.1972	3.951	0.922	0.006	-0.310	0.545
alex.2010	3.799	0.937	0.070	-0.394	0.540
alicia.1983	3.897	0.904	-0.075	-0.399	0.548
allen.1980	3.687	0.966	0.131	-0.546	0.547
andrew.1992	3.676	0.938	-0.284	-0.578	0.537
betsy.1965	3.808	0.951	-0.450	-0.389	0.424
bob.1991	3.629	0.923	0.028	-0.575	0.438
camille.1969	3.994	0.936	0.073	-0.573	0.670
charley.2004	3.639	0.948	-0.180	-0.696	0.182
david.1979	3.790	0.958	-0.046	-0.382	0.685

Predict Damage

Table 2: Coefficients of damage prediction model

term	estimate	std.error	statistic	p.value
(Intercept)	-211.035	2.017	-104.623	0
β_0	5.045	0.028	182.820	0
β_1	62.835	0.444	141.656	0
β_2	-1.096	0.013	-81.665	0
β_3	3.378	0.026	130.910	0
β_4	-1.393	0.034	-41.399	0
nobs	0.049	0.000	193.646	0
Season	0.075	0.000	187.765	0
MonthJuly	0.548	0.019	29.460	0
MonthJune	-3.416	0.024	-141.750	0
MonthNovember	-1.902	0.025	-76.221	0
MonthOctober	-1.291	0.009	-136.870	0
MonthSeptember	-1.764	0.008	-229.409	0
NatureNR	-4.317	0.036	-121.180	0
NatureTS	-2.038	0.014	-142.332	0
Maxspeed	0.050	0.000	235.831	0
Meanspeed	-0.066	0.000	-134.784	0
Maxpressure	-0.007	0.001	-5.368	0
Meanpressure	0.000	0.000	-3.818	0
Total.Pop	0.000	0.000	49.870	0
Percent.Poor	-0.038	0.000	-206.165	0
Percent.USA	-0.005	0.000	-63.246	0

Predict Deaths

Table 3: Coefficients of death prediction model

term	estimate	std.error	statistic	p.value
(Intercept)	116.498	12.580	9.261	0.000
β_0	11.675	0.256	45.530	0.000
β_1	114.119	2.200	51.869	0.000
β_2	5.529	0.123	45.084	0.000
β_3	8.562	0.285	30.007	0.000
β_4	-10.492	0.306	-34.307	0.000
nobs	0.003	0.001	3.073	0.002
Season	0.006	0.002	2.914	0.004
MonthJuly	-1.184	0.145	-8.171	0.000
MonthJune	-1.292	0.090	-14.402	0.000
MonthNovember	-2.533	0.155	-16.323	0.000
MonthOctober	-1.547	0.065	-23.918	0.000
MonthSeptember	-0.275	0.046	-5.995	0.000
NatureNR	2.349	0.129	18.205	0.000
NatureTS	3.563	0.121	29.451	0.000
Meanspeed	-0.037	0.003	-11.696	0.000
Maxpressure	-0.269	0.010	-27.775	0.000
Meanpressure	0.005	0.000	26.759	0.000
Total.Pop	0.000	0.000	36.369	0.000
Percent.Poor	0.036	0.001	44.860	0.000
Percent.USA	-0.007	0.001	-12.950	0.000

Conclusions

- Based on posterior estimates of μ , an increase in current wind speed and the change in wind speed is associated with increase in the wind speed in the upcoming future.
- Our MCMC algorithm successfully estimates the high-dimensional parameters
 - ▶ All the parameters converges quickly under a good initial values setting
 - lacktriangle The overall \mathbb{R}^2 is relatively large, our model fits the data well
- For different months, there is no significant differences observed.
 Over years, the effect the wind speed 6 months ago has on the current wind speed may decrease a little.
- ullet The eta_i coefficients estimated from the Bayesian model is powerful when predicting the damage and deaths caused by hurricanes