

Assignment 4

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2022-11-12

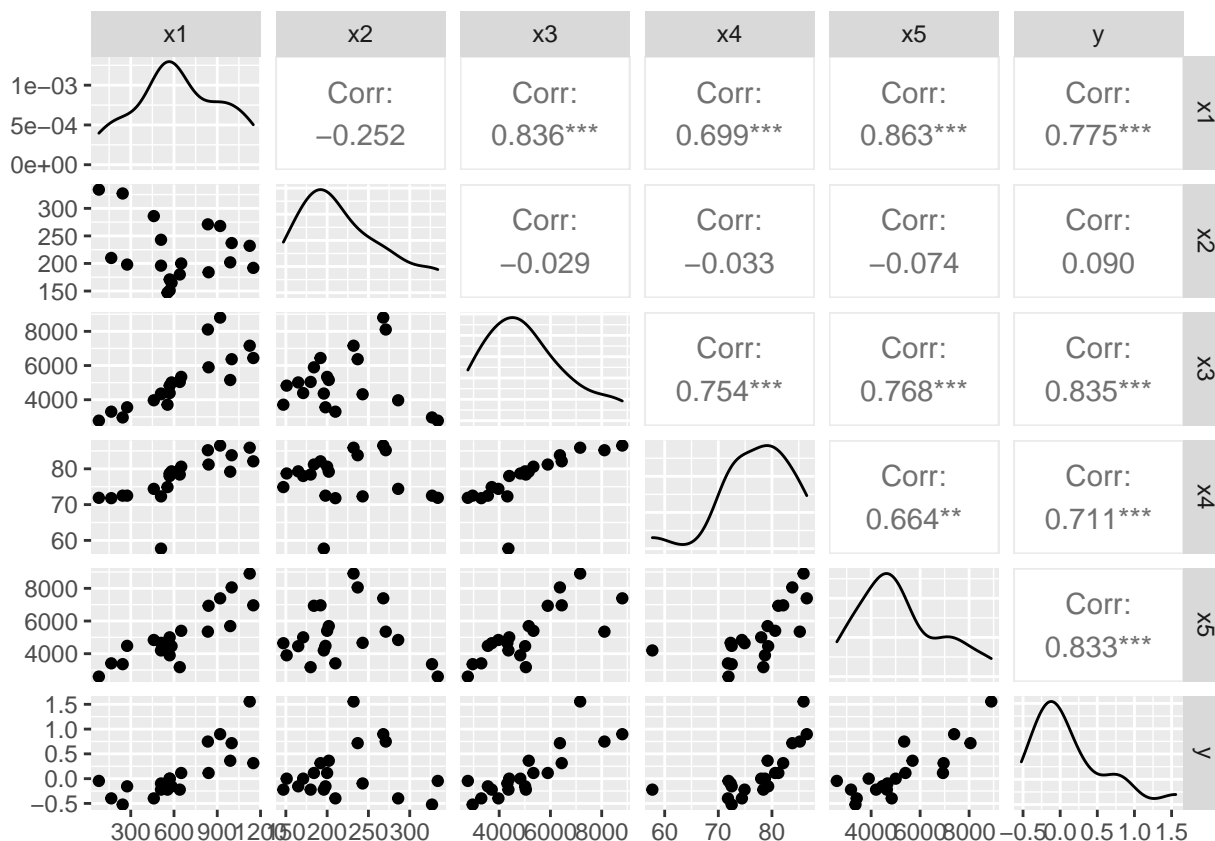
Problem 1

This is an experimental data of five laboratory measurements conducted to interpret total oxygen demand in dairy waste. Data were collected on samples kept in suspension in water in a laboratory for 220 days, and we assume that all observations are independent.

Table 1: Total oxygen demand in dairy wastes

Item	Variable	Description	Unit
1	y	<i>log(oxygen demand)</i>	mg oxygen per unit
2	x1	<i>biological oxygen demand</i>	mg/liter
3	x2	<i>Total Kjeldahl nitrogen</i>	mg/liter
4	x3	<i>Total solids</i>	mg/liter
5	x4	<i>Total volatile solids</i>	mg/liter
6	x5	<i>Chemical oxygen demand</i>	mg/liter

##	Day	x1	x2	x3
##	Min. : 0.0	Min. : 79.0	Min. :147.0	Min. :2777
##	1st Qu.: 35.0	1st Qu.: 497.5	1st Qu.:183.0	1st Qu.:3904
##	Median : 76.0	Median : 576.5	Median :201.0	Median :4918
##	Mean : 80.4	Mean : 633.5	Mean :219.7	Mean :5075
##	3rd Qu.:110.8	3rd Qu.: 860.0	3rd Qu.:249.2	3rd Qu.:6014
##	Max. :220.0	Max. :1150.0	Max. :334.0	Max. :8804
##	x4	x5	y	
##	Min. :57.70	Min. :2599	Min. : -0.5229	
##	1st Qu.:72.50	1st Qu.:4125	1st Qu.: -0.2218	
##	Median :78.55	Median :4752	Median : -0.0229	
##	Mean :77.34	Mean :5171	Mean : 0.1192	
##	3rd Qu.:81.42	3rd Qu.:6000	3rd Qu.: 0.3252	
##	Max. :86.50	Max. :8905	Max. : 1.5563	



- x_2, x_3, x_5 and y are right skewed, while x_4 is left skewed.
- From the correlation coefficient table we found that the estimator of x_4 , β_4 has a very high negative correlation associate with the intercept (β_0), which might result from the range of x_4 , which compared to other variables, is relatively small.
- As we can observe from the plot, the values of correlation coefficient between x_1 and x_3 , x_1 and x_4 , x_1 and x_5 , x_3 and x_5 , x_4 and x_5 are very high, in addition, the response variable y seemed to have be positively correlated with x_1, x_3, x_4 and x_5 .

Fit a multiple regression model $y = \beta_0 + \beta_1 * x_1 + \beta_2 * x_2 + \beta_3 * x_3 + \beta_4 * x_4 + \beta_5 * x_5 + \epsilon$, using y as the dependent variable and all x'_j s as the independent variables.

Below is the result of fitting a multiple regression model.

```
##
## Call:
## lm(formula = y ~ x1 + x2 + x3 + x4 + x5)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.39447 -0.11847  0.00053  0.08313  0.56232
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
```

```
## (Intercept) -2.156e+00  9.135e-01 -2.360  0.0333 *
## x1          -9.012e-06  5.184e-04 -0.017  0.9864
## x2           1.316e-03  1.263e-03  1.041  0.3153
## x3           1.278e-04  7.690e-05  1.662  0.1188
## x4           7.899e-03  1.400e-02  0.564  0.5815
## x5           1.417e-04  7.375e-05  1.921  0.0754 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.2618 on 14 degrees of freedom
## Multiple R-squared:  0.8107, Adjusted R-squared:  0.743
## F-statistic: 11.99 on 5 and 14 DF,  p-value: 0.0001184
##
## Correlation of Coefficients:
##      (Intercept) x1      x2      x3      x4
## x1  0.09
## x2 -0.20      0.47
## x3  0.31      -0.51 -0.27
## x4 -0.93      -0.09 -0.04 -0.39
## x5  0.01      -0.64 -0.25 -0.03 -0.09
```

a. Form a 95% confidence interval for β_3 and again for β_5 .

- 95% Critical value, $df_{\Omega} = n - p = 20 - 6 = 14$

```
## [1] 2.144787
```

- 95% confidence interval for β_3 and β_5 .

*confidence interval = estimate \pm critical value * s.e of estimate*

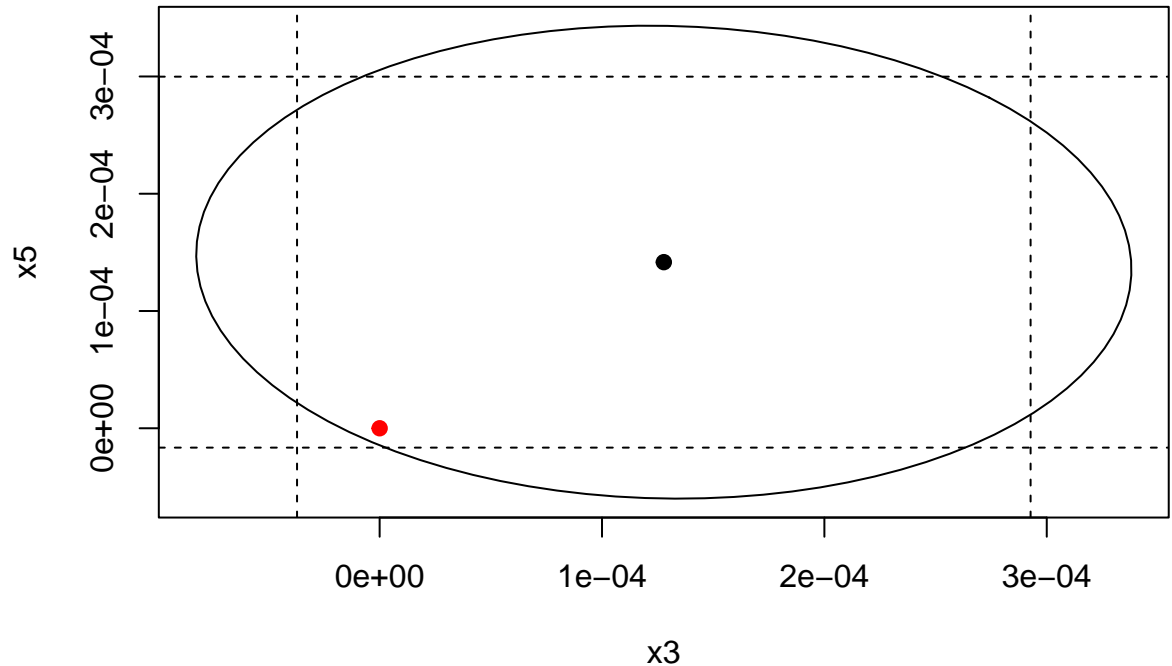
	2.5 %	97.5 %
x3	-3.71e-05	0.0002927
x5	-1.65e-05	0.0002998

- Both intervals contain 0, this indicates that the null hypotheses $H_0 : \beta_3 = 0$ and $H_0 : \beta_5 = 0$ would not be rejected at the 5% significance level. We can see from the summary where the p-value for β_3 and β_5 are 11.88% and 7.54% respectively, confirming the point.

b. Form a confidence interval for $\beta_3 + 2 * \beta_5$.

	x
2.5 %	-0.0000702
97.5 %	0.0008924

c. Show a 95% C.R. graph for β_3 and β_5 . Plot the origin. State the hypothesis test and its out-



come.

```
## [1] -0.03156963
```

- The seemingly parallelism of the major axis of the ellipse to the y-axis and that of the semi-minor axis to the x-axis suggests that the two estimators, β_3, β_5 are uncorrelated.
- We can deduced from the plot that the joint hypothesis $H_0 : \beta_3 = \beta_5 = 0$ is not reject because the origin lies right inside the ellipse
- Both of the hypotheses $H_0 : \beta_3 = 0$ and $H_0 : \beta_5 = 0$ are not rejected either because 0 does lie within the vertical and horizontal dashed lines which represents as the C.I. of x_3 and x_5 respectively.
- We can further conduct some experiments testing on the significance of β_3 where x_5 is included in the model and the significance of β_5 where x_3 is included in the model to ensure that both variables x_3, x_5 are not significant when the true model is assumed to be $y = \beta_0 + \beta_1 * x_1 + \beta_2 * x_2 + \beta_3 * x_3 + \beta_4 * x_4 + \beta_5 * x_5 + \epsilon$.

```
## Analysis of Variance Table
```

```
##
```

```
## Model 1: y ~ x1 + x2 + x4 + x5
```

```
## Model 2: y ~ x1 + x2 + x4
```

```
##   Res.Df    RSS Df Sum of Sq      F Pr(>F)
```

```
## 1      15 1.1488
```

```
## 2      16 1.4159 -1  -0.26713 3.4878 0.08149 .
```

```
## ---
```

```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
## Analysis of Variance Table
##
## Model 1: y ~ x1 + x2 + x3 + x4
## Model 2: y ~ x1 + x2 + x4
##   Res.Df    RSS Df Sum of Sq      F Pr(>F)
## 1      15 1.2124
## 2      16 1.4159 -1   -0.20355 2.5184 0.1334
```

- In conclusion, as we can observe from the above tests and visualization, all of our hypotheses are not rejected, which indicate that neither of the estimators, β_3, β_5 are significant enough to be indispensable for response variable y under the 95% confidence level.

d. If a 95% joint confidence region was computed for $\beta_1, \beta_2, \beta_3, \beta_4, \beta_5$, would the origin lie inside or outside the region? Explain.

	2.5 %	97.5 %
(Intercept)	-4.1153842	-0.1969077
x1	-0.0011208	0.0011027
x2	-0.0013940	0.0040258
x3	-0.0000371	0.0002927
x4	-0.0221278	0.0379255
x5	-0.0000165	0.0002998

	x
(Intercept)	FALSE
x1	TRUE
x2	TRUE
x3	TRUE
x4	TRUE
x5	TRUE

- As we can observe from the table, if measured individually, none of the predictors are significant; however, combining univariate tests to test on joint relationship may come up with a biased result, thus we need to conduct a test to compute the boundaries of the joint confidence region.

Joint Effect

```
##           [,1]
## [1,] 3.215305
```

Critical Value

```
## [1] 0.02956926
```

- After testing the joint effect of $\beta_1, \beta_2, \beta_3, \beta_4, \beta_5$, we found the calculated result is significantly larger than the corresponding critical value $2 * \hat{\sigma}^2 * F_{5,14} = 0.02957$, which implies that the origin $(0, 0, 0, 0, 0)$ is lying outside the region.
- Thus, we can reject the null hypothesis that $H_0 : \beta_1 = \beta_2 = \beta_3 = \beta_4 = \beta_5 = 0$.

e. Suppose non-volatile solids have no linear effect on the response. State a hypothesis that reflects this suspicion, and test it using a C.I. in your answer to one of the above questions. Explain why the chosen confidence interval can be used to do this work.

- Arrange our previous model: $y = \beta_0 + \beta_1 * x1 + \beta_2 * x2 + \beta_3 * x3 + \beta_4 * x4 + \beta_5 * x5 + \epsilon$
 - Total solids - volatile solids = non-volatile solids: $y = \beta_0 + \beta_1 * x1 + \beta_2 * x2 + \beta_3 * (x3 - x4) + \beta_4 * x4 + \beta_5 * x5 + \epsilon$
 - $y = \beta_0 + \beta_1 * x1 + \beta_2 * x2 + \beta_3 * x3 + (\beta_4 - \beta_3) * x4 + \beta_5 * x5 + \epsilon$
 - $y = \beta_0 + \beta_1 * x1 + \beta_2 * x2 + \beta_3 * x3 + \beta'_4 * x4 + \beta_5 * x5 + \epsilon$, where $\beta'_4 = \beta_4 - \beta_3$
 - As we can observe from the model equation, originally, $x4 \subset x3$, where β_3 is an estimator for both volatile and non-volatile solids while β_4 is an estimator for volatile solids, and therefore β_4 explains both regular parts from $x3$ and $x4$; yet with the arrangement, now β_3 becomes the estimator for non-volatile solids which is uncorrelated to predictor $x4$.
 - Hence, we hypothesized that $H_0 : \beta_3 = 0$, Since the arrangement is only a linear transformation performed on variable $x3$ which does not change the value of β_3 , I would apply the C.I. value calculated from question a to test this hypothesis.
 - Set $\beta_3 = 0$
 - Introduce C.I. for β_3 from model1

```
##          2.5 %          97.5 %
## -3.713929e-05  2.927368e-04
```

 - Whether we accept the null hypothesis
- ```
[1] TRUE
```
- Since 0 lies in the C.I. of  $\beta_3$ , we do not reject the null hypothesis, that is, the non-volatile solids have no linear effect on the response.

## Problem 2

This data are a random sample of home sales from Spring 1993 in Albuquerque.

Table 6: Home sales data

| Item | Variable | Description               | Unit                                                                                                                             |
|------|----------|---------------------------|----------------------------------------------------------------------------------------------------------------------------------|
| 1    | Price    | Selling price             | \$100                                                                                                                            |
| 2    | SQFT     | living space              | <i>feet</i> <sup>2</sup>                                                                                                         |
| 3    | Age      | Age of home               | <i>year</i>                                                                                                                      |
| 4    | Features | Number out of 11 features | (dish washer, refrigerator, microwave, disposer, washer, intercom, skylight(s), compactor, dryer, handicap fit, cable TV access) |

| Item | Variable | Description                         | Unit   |
|------|----------|-------------------------------------|--------|
| 5    | NE       | Located in northeast sector of city | 1 or 0 |
| 6    | Corner   | Corner location                     | 1 or 0 |
| 7    | Taxes    | Annual taxes                        | \$     |

a. There are a large number of missing values in the *Age* variable. We could either exclude *Age* from our models for the selling price or we could keep *Age* and exclude the cases that have missing values for *Age*. Which choice is better for this data? Explain your reasoning.

- Check how many missing values are in the data.

```
Price SQFT Age Features NE
Min. : 540 Min. : 837 Min. : 1.00 Min. :0.00 Min. :0.0000
1st Qu.: 780 1st Qu.:1280 1st Qu.: 5.75 1st Qu.:3.00 1st Qu.:0.0000
Median : 960 Median :1549 Median :13.00 Median :4.00 Median :1.0000
Mean :1063 Mean :1654 Mean :14.97 Mean :3.53 Mean :0.6667
3rd Qu.:1200 3rd Qu.:1894 3rd Qu.:19.25 3rd Qu.:4.00 3rd Qu.:1.0000
Max. :2150 Max. :3750 Max. :53.00 Max. :8.00 Max. :1.0000
##
NA's :49
Corner Tax
Min. :0.000 Min. : 223.0
1st Qu.:0.000 1st Qu.: 600.0
Median :0.000 Median : 731.0
Mean :0.188 Mean : 793.5
3rd Qu.:0.000 3rd Qu.: 919.0
Max. :1.000 Max. :1765.0
##
NA's :10
```

- As we can observe from the result of the summary, there are 49 na values in variable *Age* and 10 in variable *Tax*.
- Furthermore, although *Features* variable description indicated there are 11 different features in the data, yet as we can see from the table, there are only 8 features at most including in our sample data.

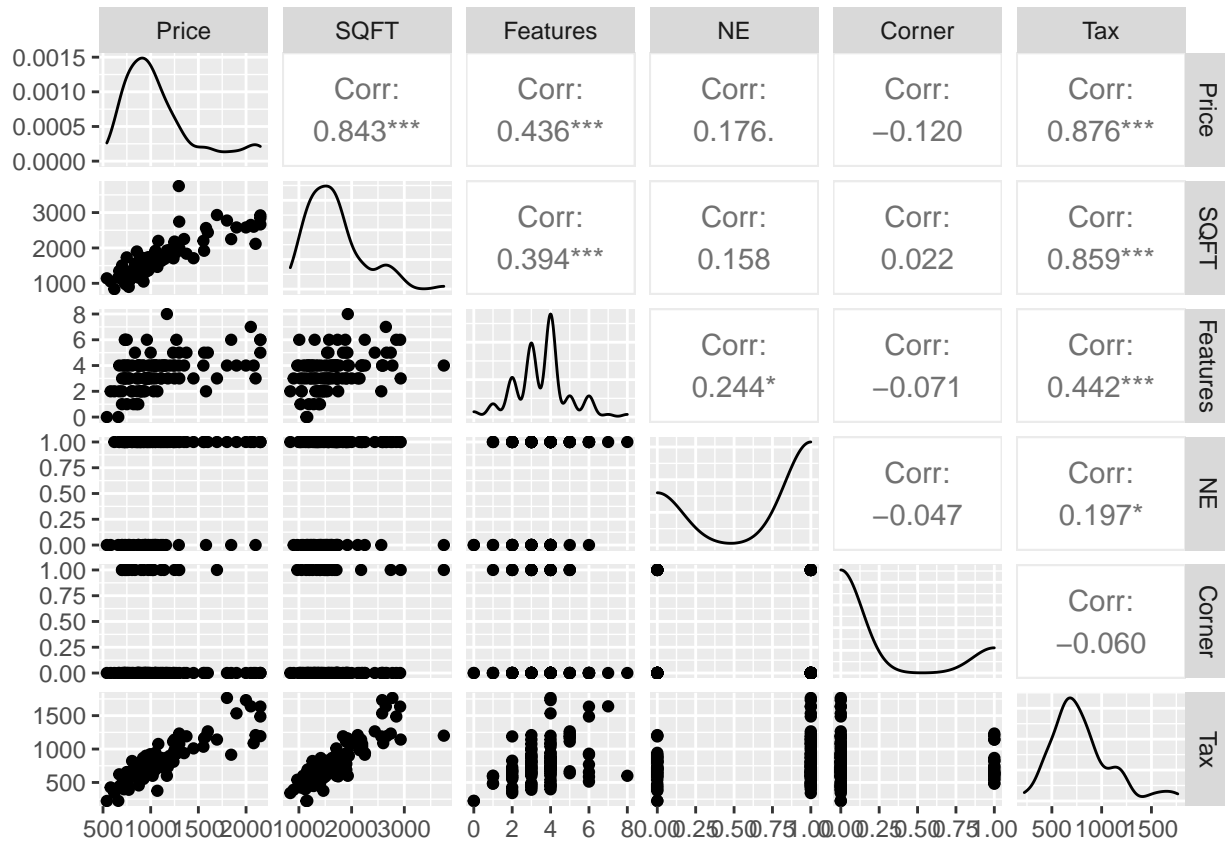
```
[1] 117
```

- However, there are only 117 samples in this data, removing all the missing values in *Age* would indicate dropping nearly half the amount of the observations, which is irrational to give up so much observations; therefore, choosing to exclude variable *Age* from the dataset is a more economical and reasonable option.

| Price | SQFT | Features | NE | Corner | Tax  |
|-------|------|----------|----|--------|------|
| 2050  | 2650 | 7        | 1  | 0      | 1639 |
| 2080  | 2600 | 4        | 1  | 0      | 1088 |
| 2150  | 2664 | 5        | 1  | 0      | 1193 |
| 2150  | 2921 | 6        | 1  | 0      | 1635 |
| 1999  | 2580 | 4        | 1  | 0      | 1732 |
| 1900  | 2580 | 4        | 1  | 0      | 1534 |

| Price        | SQFT         | Features      | NE             | Corner         | Tax            |
|--------------|--------------|---------------|----------------|----------------|----------------|
| Min. : 540   | Min. : 837   | Min. :0.000   | Min. :0.0000   | Min. :0.0000   | Min. : 223.0   |
| 1st Qu.: 815 | 1st Qu.:1290 | 1st Qu.:3.000 | 1st Qu.:0.0000 | 1st Qu.:0.0000 | 1st Qu.: 600.0 |
| Median : 975 | Median :1565 | Median :4.000 | Median :1.0000 | Median :0.0000 | Median : 731.0 |
| Mean :1077   | Mean :1667   | Mean :3.533   | Mean :0.6636   | Mean :0.1963   | Mean : 793.5   |
| 3rd Qu.:1190 | 3rd Qu.:1897 | 3rd Qu.:4.000 | 3rd Qu.:1.0000 | 3rd Qu.:0.0000 | 3rd Qu.: 919.0 |
| Max. :2150   | Max. :3750   | Max. :8.000   | Max. :1.0000   | Max. :1.0000   | Max. :1765.0   |

- New data after excluding *Age* from the dataset and removing na values from variable *Tax*.



- House *Price* seems to be positively correlated with *SQFT*, *Tax* which maps to our intuition. While the relationship between *Price* and other variables are unclear.

**b. Fit a model with *Price* as the response and *SQFT*, *Features*, *NE*, *Corner*, *Tax* as predictors. Form 95% and 99% C.I. for their coefficients. Explain how the p-value for the parameter for *Corner* relates to whether zero falls in the two corresponding C.I..**  $Price = \beta_0 + \beta_1 * SQFT + \beta_2 * Features + \beta_3 * NE + \beta_4 * Corner + \beta_5 * Tax + \epsilon$

```
##
Call:
lm(formula = Price ~ SQFT + factor(Features) + factor(NE) + factor(Corner) +
Tax)
##
```



```

Residuals:
Min 1Q Median 3Q Max
-507.41 -72.61 -14.43 64.13 607.27
##
Coefficients:
Estimate Std. Error t value Pr(>|t|)
(Intercept) 154.54826 136.56039 1.132 0.260631
SQFT 0.25275 0.06704 3.770 0.000285 ***
factor(Features)1 -34.35628 153.76729 -0.223 0.823685
factor(Features)2 -44.18014 140.34877 -0.315 0.753620
factor(Features)3 -31.40899 138.77495 -0.226 0.821436
factor(Features)4 -44.75297 139.38047 -0.321 0.748859
factor(Features)5 8.07140 156.79580 0.051 0.959055
factor(Features)6 35.90463 151.67262 0.237 0.813386
factor(Features)7 98.54855 242.25824 0.407 0.685086
factor(Features)8 118.43407 220.81098 0.536 0.592977
factor(NE)1 -4.56309 37.89202 -0.120 0.904405
factor(Corner)1 -83.75216 44.94169 -1.864 0.065503 .
Tax 0.69047 0.12462 5.540 2.74e-07 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
Residual standard error: 176.4 on 94 degrees of freedom
Multiple R-squared: 0.8128, Adjusted R-squared: 0.7889
F-statistic: 34.01 on 12 and 94 DF, p-value: < 2.2e-16
##
Correlation of Coefficients:
(Intercept) SQFT factor(Features)1 factor(Features)2
SQFT -0.39
factor(Features)1 -0.80 0.19
factor(Features)2 -0.87 0.19 0.81
factor(Features)3 -0.87 0.20 0.83 0.92
factor(Features)4 -0.87 0.20 0.83 0.92
factor(Features)5 -0.76 0.18 0.76 0.85
factor(Features)6 -0.78 0.14 0.77 0.85
factor(Features)7 -0.51 0.20 0.53 0.59
factor(Features)8 -0.48 -0.06 0.49 0.54
factor(NE)1 0.01 0.01 -0.14 -0.15
factor(Corner)1 0.06 -0.18 -0.11 -0.14
Tax 0.27 -0.85 -0.23 -0.26
factor(Features)3 factor(Features)4 factor(Features)5
SQFT
factor(Features)1
factor(Features)2
factor(Features)3
factor(Features)4 0.95
factor(Features)5 0.88 0.89
factor(Features)6 0.88 0.89 0.83
factor(Features)7 0.62 0.64 0.61
factor(Features)8 0.55 0.55 0.51
factor(NE)1 -0.14 -0.16 -0.18
factor(Corner)1 -0.13 -0.12 -0.17
Tax -0.30 -0.32 -0.34
factor(Features)6 factor(Features)7 factor(Features)8

```

```

SQFT
factor(Features)1
factor(Features)2
factor(Features)3
factor(Features)4
factor(Features)5
factor(Features)6
factor(Features)7 0.60
factor(Features)8 0.52 0.34
factor(NE)1 -0.19 -0.12 -0.16
factor(Corner)1 -0.06 -0.07 0.00
Tax -0.28 -0.37 0.00
factor(NE)1 factor(Corner)1
SQFT
factor(Features)1
factor(Features)2
factor(Features)3
factor(Features)4
factor(Features)5
factor(Features)6
factor(Features)7
factor(Features)8
factor(NE)1
factor(Corner)1 0.04
Tax -0.05 0.19

```

- **Note.** Features, Corner and NE are qualitative data, the numeric value are categorical.

|                   | 2.5 %        | 97.5 %      |
|-------------------|--------------|-------------|
| (Intercept)       | -116.5955928 | 425.6921070 |
| SQFT              | 0.1196449    | 0.3858495   |
| factor(Features)1 | -339.6648385 | 270.9522849 |
| factor(Features)2 | -322.8459148 | 234.4856306 |
| factor(Features)3 | -306.9499127 | 244.1319319 |
| factor(Features)4 | -321.4961537 | 231.9902206 |
| factor(Features)5 | -303.2503320 | 319.3931363 |
| factor(Features)6 | -265.2449074 | 337.0541723 |
| factor(Features)7 | -382.4608585 | 579.5579628 |
| factor(Features)8 | -319.9913118 | 556.8594541 |
| factor(NE)1       | -79.7985862  | 70.6724125  |
| factor(Corner)1   | -172.9849346 | 5.4806216   |
| Tax               | 0.4430308    | 0.9379164   |

|                   | 0.5 %        | 99.5 %      |
|-------------------|--------------|-------------|
| (Intercept)       | -204.4891622 | 513.5856764 |
| SQFT              | 0.0764987    | 0.4289958   |
| factor(Features)1 | -438.6331879 | 369.9206342 |
| factor(Features)2 | -413.1777790 | 324.8174948 |
| factor(Features)3 | -396.2688303 | 333.4508495 |
| factor(Features)4 | -411.2047956 | 321.6988625 |

|                   | 0.5 %        | 99.5 %      |
|-------------------|--------------|-------------|
| factor(Features)5 | -404.1679021 | 420.3107063 |
| factor(Features)6 | -362.8650747 | 434.6743396 |
| factor(Features)7 | -538.3841223 | 735.4812266 |
| factor(Features)8 | -462.1106032 | 698.9787456 |
| factor(NE)1       | -104.1868089 | 95.0606352  |
| factor(Corner)1   | -201.9104934 | 34.4061804  |
| Tax               | 0.3628201    | 1.0181270   |

- p-value for *Corner* is not significant whether the significance level is 95% or 99%. Since zero falls in both  $\alpha = 5\%$ ,  $\alpha = 1\%$  confidence intervals.

**c. Predict the *Price* of a specific house with  $SQFT = 2500$ ,  $Features = 5$ ,  $NE = 1$ ,  $Corner = 1$ ,  $Tax = 120$ . Give an appropriate 95% C.I..**

$$Price = 154.548 + 0.253 * SQFT - 34.356 * F_1 - 44.180 * F_2 - 31.409 * F_3 - 44.753 * F_4 + 8.071 * F_5 + 35.905 * F_6 + 98.549 * F_7 + 118.434 * F_8$$

- **Note.** If  $Featur_s = x$ ,  $F_x = 1$ ,  $x \in [1, 8]$

```
fit lwr upr
1 1534.741 1396.064 1673.417
```

- This is an interpolation prediction computing for the C.I. of mean response.

**d. Suppose you are only told that  $SQFT = 2500$ . Predict the *Price* and 95% C.I..**

```
Warning in predict.lm(model2, interval = "prediction"): predictions on current data refer to _future_
```

```
fit lwr upr
1 1418.1319766 889.451658 1946.8123
2 750.9011812 369.767088 1132.0353
3 916.2190550 517.773971 1314.6641
4 1450.6336437 1042.185047 1859.0822
5 1451.5846011 1052.637448 1850.5318
6 1236.1570900 844.829949 1627.4842
7 1487.4891863 1087.097484 1887.8809
8 881.4024876 483.109554 1279.6954
9 680.6166847 297.140836 1064.0925
10 914.0430196 515.608935 1312.4771
11 683.5811421 286.941138 1080.2211
12 652.3282236 269.134669 1035.5218
13 828.0896187 429.931014 1226.2482
14 538.1319766 9.451658 1066.8123
15 379.2357727 -3.027859 761.4994
16 431.0239678 51.045826 811.0021
17 472.2370781 82.793975 861.6802
18 330.3581168 -57.213176 717.9294
19 585.9591419 203.291300 968.6270
20 375.5350634 -4.700886 755.7710
21 399.9081096 17.688237 782.1280
```

|       |              |             |           |
|-------|--------------|-------------|-----------|
| ## 22 | 324.3982300  | -56.195060  | 704.9915  |
| ## 23 | 317.1780540  | -76.839875  | 711.1960  |
| ## 24 | 507.3222694  | 109.928864  | 904.7157  |
| ## 25 | 285.1057416  | -103.012937 | 673.2244  |
| ## 26 | 771.4419546  | 378.307078  | 1164.5768 |
| ## 27 | 693.2362414  | 312.642598  | 1073.8299 |
| ## 28 | 752.4258551  | 368.077808  | 1136.7739 |
| ## 29 | 210.0248045  | -179.330925 | 599.3805  |
| ## 30 | 167.1554568  | -221.510233 | 555.8211  |
| ## 31 | 293.9337335  | -86.926907  | 674.7944  |
| ## 32 | 123.5108976  | -270.297535 | 517.3193  |
| ## 33 | 190.4481989  | -197.922031 | 578.8184  |
| ## 34 | 228.2363509  | -182.888202 | 639.3609  |
| ## 35 | 217.6486422  | -170.609774 | 605.9071  |
| ## 36 | 227.1483332  | -183.975238 | 638.2719  |
| ## 37 | 100.7032331  | -283.966817 | 485.3733  |
| ## 38 | 0.1689456    | -385.333355 | 385.6712  |
| ## 39 | 385.4642833  | -13.212435  | 784.1410  |
| ## 40 | 308.2150243  | -91.595158  | 708.0252  |
| ## 41 | 62.0543929   | -328.554534 | 452.6633  |
| ## 42 | 45.8656904   | -338.671634 | 430.4030  |
| ## 43 | 1289.6070192 | 885.625374  | 1693.5887 |
| ## 44 | 994.5563318  | 595.583153  | 1393.5295 |
| ## 45 | 588.7865390  | 208.800685  | 968.7724  |
| ## 46 | 905.8586758  | 503.909939  | 1307.8074 |
| ## 47 | 755.7730503  | 365.660036  | 1145.8861 |
| ## 48 | 448.4322515  | 68.507073   | 828.3574  |
| ## 49 | 545.2658297  | 165.391616  | 925.1400  |
| ## 50 | 437.5520742  | 57.595357   | 817.5088  |
| ## 51 | 508.2732268  | 128.428310  | 888.1181  |
| ## 52 | 737.8449684  | 356.845967  | 1118.8440 |
| ## 53 | 519.1534041  | 139.306129  | 899.0007  |
| ## 54 | 332.4510102  | -49.980702  | 714.8827  |
| ## 55 | 286.7542654  | -95.933973  | 669.4425  |
| ## 56 | 473.8933155  | 91.676730   | 856.1099  |
| ## 57 | 63.2739741   | -320.919032 | 447.4670  |
| ## 58 | 353.7747088  | -26.599278  | 734.1487  |
| ## 59 | 430.3726062  | 48.183117   | 812.5621  |
| ## 60 | 259.5761426  | -145.256293 | 664.4086  |
| ## 61 | 306.8660963  | -81.249358  | 694.9816  |
| ## 62 | 407.6073758  | 20.153778   | 795.0610  |
| ## 63 | 304.6900608  | -83.424797  | 692.8049  |
| ## 64 | 235.0492124  | -154.187266 | 624.2857  |
| ## 65 | 68.7140628   | -315.373989 | 452.8021  |
| ## 66 | 104.2023294  | -307.120622 | 515.5253  |
| ## 67 | 46.3023465   | -339.229296 | 431.8340  |
| ## 68 | 229.8777475  | -171.332893 | 631.0884  |
| ## 69 | 115.4988253  | -267.738787 | 498.7364  |
| ## 70 | 251.3771919  | -136.786836 | 639.5412  |
| ## 71 | -53.2677733  | -444.054920 | 337.5194  |
| ## 72 | 826.0599038  | 431.958547  | 1220.1613 |
| ## 73 | 296.6319241  | -93.407992  | 686.6718  |
| ## 74 | 343.4244001  | -46.766065  | 733.6149  |
| ## 75 | 218.0098923  | -194.908986 | 630.9288  |

|        |             |             |           |
|--------|-------------|-------------|-----------|
| ## 76  | 240.6232216 | -144.055931 | 625.3024  |
| ## 77  | 13.2275154  | -374.057613 | 400.5126  |
| ## 78  | 170.9493787 | -223.692661 | 565.5914  |
| ## 79  | -32.9560411 | -490.806649 | 424.8946  |
| ## 80  | -10.1813630 | -401.675582 | 381.3129  |
| ## 81  | 903.2260208 | 513.371954  | 1293.0801 |
| ## 82  | 74.1565085  | -312.212471 | 460.5255  |
| ## 83  | -30.7800057 | -488.630614 | 427.0706  |
| ## 84  | 44.2195236  | -346.726134 | 435.1652  |
| ## 85  | 673.3546834 | 270.703061  | 1076.0063 |
| ## 86  | 874.3770537 | 476.157259  | 1272.5968 |
| ## 87  | 52.0629773  | -366.540520 | 470.6665  |
| ## 88  | 516.5430696 | 133.661554  | 899.4246  |
| ## 89  | 475.6350519 | 91.294346   | 859.9758  |
| ## 90  | 362.0445515 | -20.990918  | 745.0800  |
| ## 91  | 443.6458815 | 60.821729   | 826.4700  |
| ## 92  | 389.2449949 | 6.312287    | 772.1777  |
| ## 93  | 600.3204350 | 217.086277  | 983.5546  |
| ## 94  | 455.6140766 | 72.796446   | 838.4317  |
| ## 95  | 298.3104819 | -110.484701 | 707.1057  |
| ## 96  | 426.6335459 | 30.361733   | 822.9054  |
| ## 97  | 545.2681868 | 160.566084  | 929.9703  |
| ## 98  | 567.0285414 | 182.170325  | 951.8868  |
| ## 99  | 291.7600550 | -92.640833  | 676.1609  |
| ## 100 | 515.4550519 | 132.576100  | 898.3340  |
| ## 101 | 469.1069455 | 84.789312   | 853.4246  |
| ## 102 | 342.4602324 | -40.669223  | 725.5897  |
| ## 103 | 372.2733672 | -11.919943  | 756.4667  |
| ## 104 | 267.3870088 | -116.257551 | 651.0316  |
| ## 105 | 328.3160018 | -54.891736  | 711.5237  |
| ## 106 | 210.2418671 | -179.636866 | 600.1206  |
| ## 107 | 109.0562179 | -281.681236 | 499.7937  |