

1. The MLE,  $\hat{p}$ , for  $p$  is given by:

$$\hat{p} = \frac{12}{70} = \frac{6}{35}$$

2. The likelihood for  $\theta$ , where  $X_1, \dots, X_n \sim \text{Bern}(\theta)$  and  $0 < \theta < 1$ , is given by:

$$L(\theta) = \prod_{i=1}^n p^{x_i} (1-p)^{1-x_i} = \prod_{i=1}^n \theta^{x_i} (1-\theta)^{1-x_i}$$

$$L(\theta) = (1-\theta)^n \quad \text{if } x_i = 0$$

and

$$= \theta^n \quad \text{if } x_i = 1$$

Taking the log likelihood:

$$l(\theta) = n \ln(1-\theta) \quad (\text{if every observed value is } 0)$$

Taking the first derivative to get the MLE of  $\theta$ :

$$\frac{d}{d\theta} [l(\theta)] = \frac{-n}{1-\theta} = 0$$

from the above equation, we cannot solve for  $\hat{\theta}$ .

Similarly, if every observation is 1,

$$\frac{d}{d\theta} [l(\theta)] = \frac{n}{\theta} = 0$$

$\therefore$  when every observed value is either 0 or 1, the MLE of  $\theta$  does not exist.

3.  $X_1, \dots, X_n \sim \text{Poisson}(\lambda)$ ,  $\lambda > 0$

$$L(\lambda) \rightarrow \text{Likelihood} \rightarrow f(x_1, \dots, x_n | \lambda) = \prod_{i=1}^n \frac{\lambda^{x_i} e^{-\lambda}}{x_i!}$$

$$\begin{aligned} \log \text{ likelihood: } l(\lambda) &= \sum_{i=1}^n (x_i \ln(\lambda) - n\lambda - \ln(x_i!)) \\ &= \ln(\lambda) \sum_{i=1}^n x_i - n\lambda - \sum_{i=1}^n \ln(x_i!) \end{aligned}$$

taking the first derivative:

$$\frac{d}{d\lambda} [l(\lambda)] = \frac{1}{\lambda} \sum_{i=1}^n x_i - n = 0$$

$$\therefore \hat{\lambda} = \frac{\sum_{i=1}^n x_i}{n} = \bar{x}$$

If every observed value is 0, then we cannot find MLE  $\hat{\lambda}$  of  $\lambda$  from  $\frac{1}{\lambda} \sum_{i=1}^n x_i - n = 0$  as  $\sum_{i=1}^n x_i - n < 0$  for all  $x_i = 0$ ; and we cannot find <sup>any</sup> one value of  $\lambda$  to satisfy the inequality.

4.  $X_1, \dots, X_n \sim N(\mu, \sigma^2)$ ,  $\sigma^2$  is unknown and  $\mu$  is known.

$$\begin{aligned} \log \text{ likelihood: } l(\mu, \sigma^2, x_1, \dots, x_n) &= -\frac{n}{2} \ln(2\pi) - \frac{n}{2} \ln(\sigma^2) \\ &\quad - \frac{1}{2\sigma^2} \sum_{j=1}^n (x_j - \mu)^2 \end{aligned}$$

taking the 1<sup>st</sup> derivative:

$$\begin{aligned} \frac{d}{d\sigma^2} [l(\mu, \sigma^2, x_1, \dots, x_n)] &= \frac{-n(2)}{2\sigma^2} - \frac{(-2)}{2\sigma^3} \sum_{j=1}^n (x_j - \mu)^2 \\ &= -\frac{n}{\sigma} + \frac{1}{\sigma^3} \left( \sum_{j=1}^n (x_j - \mu)^2 \right) = 0 \end{aligned}$$

$$\Rightarrow \frac{\sum_{j=1}^n (x_j - \mu)^2}{\sigma^3} = \frac{n}{\sigma}$$

$$\Rightarrow \hat{\sigma}^2 = \frac{\sum_{j=1}^n (x_j - \mu)^2}{n}$$