#### 第一章 概率论

条件概率:  $P(B|A) = \frac{P(AB)}{P(B|A)}$ 

 $P(A) = \sum_{i=1}^{m} P(B_i) P(A \mid B_i)$ 全概率公式:

$$P(B_{j} | A) = \frac{P(B_{i})P(A|B_{i})}{\sum_{i=1}^{n} P(B_{i})P(A|B_{i})}$$

贝叶斯:

分布函数定义:  $F(x) = P\{X < x\}$ 离散型分布律(概率密度):  $P\{X=x_k\}=$   $p_k$   $P\{X=x_k\}=$   $p_k$ 离散型分布函数: 连续型分布函数:  $F(x) = \int_{-\infty}^{x} f(u)du$ ) 连续型概率密度: f(x) 0-1 分布:  $\varphi_X = q + pe^{iu}$  (特征 $\varphi_X(u)$ ) P(A)=p,P(A')=1-p E(x)=p,D(x)=pq二项分布 B(n,p): E=np,D=npq  $p_k = P\{X = k\} = C_n^k p^k q^{n-k}$  $\varphi_X = (q + pe^{iu})^n$ 

泊松分布:  $E=D=\lambda_o$   $\phi_X=e^{\lambda(e^{iu}-1)}$   $p_k=P\{X=k\}=\frac{\lambda^k}{k!}e^{-\lambda},\lambda>0$ 均匀分布 U(a,b):

$$f(x) = \begin{cases} \frac{1}{b-a} & a < x < b \\ 0 & a < x < b \end{cases}$$

$$f(x) = \begin{cases} b - a \\ 0 \text{ or } \end{cases}$$

$$F(x) = \begin{cases} \frac{x - a}{b - a} & a < x < b \\ 1 & x \ge b \end{cases}$$

$$E(x) = \frac{a+b}{2} D(x) = \frac{(b-a)^2}{12}$$
$$\varphi_X = \frac{e^{ibu} - e^{iau}}{i(b-a)u}$$
$$\mathbb{Z} a = -b \quad \varphi_X = \frac{\sin bu}{bu}$$

$$f(x) = \begin{cases} \lambda e^{-\lambda} & x > 0 \\ 0 & x \le 0 \end{cases}$$

指数分布:  

$$f(x) = \begin{cases} \lambda e^{-\lambda} & x > 0 \\ 0 & x \le 0 \end{cases}$$

$$F(x) = \begin{cases} 1 - e^{-\lambda} & x > 0 \\ 0 & x \le 0 \end{cases}$$

$$E(x) = \frac{1}{\lambda} D(x) = \frac{1}{\lambda^2}$$

$$\varphi_X = (1 - i\frac{u}{\lambda})^{-1}$$

正态分布: X~N(μ,σ²)

$$f(x) = \frac{1}{\sqrt{2\pi\sigma}}e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$
$$E(x) = \mu D(x) = \sigma^2$$

 $E(x) = \mu D(x) = \sigma^2$  $\varphi_X = e^{i\mu u - \frac{1}{2}\sigma^2 u^2}$ 

离散联合分布律:  $pij=P\{X=xi,Y=yj\}$  离散联合分布函数:  $F(x,y)=\sum_{x\in Y}\sum_{y\in Y}p_{ij}$ 离散 X 边缘分布律:  $p_i = P(X = x_i) = \sum_{j=1}^{x_i < x_j < y_j} p_{ij}$  离散 X 条件分布律:  $p_{ij} = \frac{p_{ij}}{p_{ij}}$ 

连续联合分布函数:  $F(x,y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f(u,v) du dv$ 连续 X 边缘分布函数:  $F_X(x) = F(x, +\infty)$ 连续 X 边缘概率密度:  $f_X(x) = \int_{-\infty}^{\infty} f(x,y) dy$ Y 条件概率密度:  $f_Y|X(y|x)=f(x,y)/f_X(x)$ X 和 Y 独立: F(x,y)=F<sub>x</sub>(x)F<sub>Y</sub>(y) 离散数学期望:  $E(X) = \sum_{k=1}^{\infty} x_k p_k$  连续数学期望:  $E(X) = \int_{-\infty}^{+\infty} x f(x) dx$ 注: x 为 g(x)或 g(x,y),则取代 x 的位

方差:  $D(X) = E[X-E(X)]^2 = E(X^2)-E^2(X)=c_{ii}$  标准差:  $\sigma_X = \sqrt{D(X)}$ 

协方差: cov(X<sub>i</sub>,X<sub>j</sub>)=E(X<sub>i</sub>X<sub>j</sub>)-E(X<sub>i</sub>)E(X<sub>j</sub>)=c<sub>ij</sub> 相关系数:  $\rho_{XY} = \frac{V_{COV}(X,Y)^N}{\sqrt{D(X)}\sqrt{D(Y)}}$  E(aX+b)=aE(X)+b, $D(aX+b)=a^2D(X)$ 

 $\mathbf{a}_{k}$ 为任意常数:  $E(\sum a_{k}X_{k}) = \sum^{n}a_{k}E(X_{k})$ 

 $D(\sum_{i=1}^{n} a_i X_i) = \sum_{i=1}^{n} \sum_{j=1}^{n} a_i a_j \operatorname{cov}(X_i, X_j)$ cov(aX+bY,cX+dY)=acD(X)+bdD(Y)+(ad+b)

离散 X 的条件期望.  $E(X|Y=y_j) = \sum_{i=1}^{\infty} x_i p_{ij}$  连续,  $E(X|Y=y) = \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx$ 

# 第二章 随机过程基本概念

随机过程{X(t),t∈T}的一维分布函数:  $F(t,x) = P\{X(t) < x\}, t \in T, x \in R = (-\infty, +\infty)$ 连续一维分布函数:  $F(t,x) = \int_{-\infty}^{\infty} f(t,y)dy$ 

二维:  $F(s,t;x,y) = P\{X(s) < x, X(t) < y\}$ 连续:  $F(s,t;x,y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f(s,t;x,y) dxdy$ 连续二维概率密度:  $f(s,t;x,y) = \frac{\partial^2 F(s,t;x,y)}{\partial s}$ 均值函数: m(t)=E[X(t)]

$$m(t) = \sum_{k=0}^{\infty} x_k p_k(t) = \int_{-\infty}^{+\infty} x f(t, x) dx$$

 $E[X+Y]^2=E[X^2]+E[Y^2]+2E[XY]$ XY 相互独立: E[XY]=E[X]E[Y] 方差: D(t)=E[X(t)-m(t)]2=E[X2(t)]-

协方差: C(s,t)=E[X(s)X(t)]-m(s)m(t) 相关函数: R(s,t)=E[X(s)X(t)]

$$= \frac{\alpha^2}{2\pi} \frac{1}{2} \int_0^{2\pi} [\cos \beta(t-s) + \cos(\beta(t+s) + 2\theta)] d\theta$$
$$= \frac{\alpha^2}{2} \cos \beta(t-s)$$

## 第三章 重要的随机过程

{X(t),t∈T}为平稳独立增量过程: X(t+h)-X(s+h)与 X(t)-X(s)有相同且独立的概率 分布。

增量 X(t+τ)-X(t) 仅依赖于时间区间长 度而与起始时间无关, 具有平稳性。 如果{X(t),t>=0}是平稳独立增量过程, X(0)=0, f(t)=mt,  $D(t)=\sigma^2t$ ,  $C(s,t)=\sigma^2=min(s,t)$ (m 和  $\sigma$  均为常数) 二项计数过程:

$$Y(n) = \sum_{k=0}^{\infty} X(k) \sim B(n,p)$$

 $E[Y(n)] = np \ D[Y(n)] = npq$   $C_Y(m,n) = pqmin(m,n)$ 

泊松讨程定义:

N(0)=0;具有平稳独立增量;N(t)-N(s)服 从参数为 λ(t-s)的泊松分布(或者  $P{N(h)=1}=h+o(h); P{N(h)=2}=o(h))$ ,

$$P(N(t) - N(s) = k) = \frac{\left[\lambda(t - s)\right]^k}{k!} e^{-\lambda(t - s)}$$

$$P(N(t) = k) = \frac{(\lambda t)^k}{k!} e^{-\lambda t}$$

$$m(t) = E[N(t)] = \lambda t = D(t)$$

2222222 $\phi(u) = e^{\lambda t(e^{iu} - 1)}$ 

 $C(s,t) = \lambda \min(s,t)$ 

 $R(s,t) = \lambda \min(s,t) + \lambda^2 st$ 二维概率分布:

$$P\{N(s) = j, N(t) = k\} = \frac{\lambda^k s^j (t-s)^{k-j}}{j! (k-j)!} e^{-\lambda t}, \quad 0 < s < t$$

泊松过程的充要条件:

设{N(t),t<sup>3</sup>0}是参数为 λ 的泊松过程, {T<sub>n</sub>,n=1,2,...}为点间间距序列,表示事 件第 n-1 次出现到第 n 次出现之间的 时间,则 T<sub>n</sub>,n=1,2,...是相互独立同分布 的随机变量,且都服从参数为λ的(负) 指数分布1~₽\*, t≥0

$$F_T(t) = \begin{cases} 1 - \epsilon & , & t \ge 0 \\ 0, & t < 0 \end{cases}$$

$$f_{T_1}(t) = \begin{cases} \lambda e^{-\lambda t}, & t \ge 0\\ 0, & t < 0 \end{cases}$$

 $\{\tau_n, n=1,2,...\}$ 为事件第 n 次出现的等待 时间序列,则 τ  $_{n}$  $\sim$  $\Gamma$ (n, $\lambda$ ),概率密度

$$f(t) = \begin{cases} \frac{\lambda^n}{(n-1)!} t^{n-1} e^{-\lambda t}, t \ge 0\\ 0, t < 0 \end{cases}$$

$$F(t) = \sum_{k=0}^{\infty} \frac{\lambda t^k}{k!} e^{-\lambda t}, t \ge 0$$

非齐次泊松过程:

$$P\{N(t_0 + t) - N(t_0) = k\}$$

$$=\frac{[m(t_0+t)-m(t_0)]^k}{k!}e^{-[m(t_0+t)-m(t_0)]}$$

其中, $m(t) = \int_0^t \lambda(s) ds$  复合泊松过程:

复合泪松过程: 
$$X(t) = \sum_{n=1}^{N(t)} Y_n$$
特征函数:  $\Phi X(t,u) = e^{\lambda t[\Phi_Y(u)-1]}$ 

均值函数: mX(t) = λ tE[Y] 方差函数: E[N(t)]= λt  $D[X(t)]=E[X^{2}(t)]-E^{2}[X(t)]=\lambda tE[Y^{2}]$ 更新计数过程:  $F_{N(t)}(k)=1-F_{\tau k}(t)$  $m(t) = \sum_{\tau_0} F_{\tau_0}(t)$ 

### 第四章<sup>k=</sup>马尔可夫过程

马氏过程定义(仅依赖于 X(t<sub>n-1</sub>)已知值) :  $P\{X(t_n) < x_n \mid X(t_1) = x_1, X(t_2) = x_2, ..., X(t_{n-1}) = x_1, X(t_1) = x_1, X(t_2) = x_2, ..., X(t_n) = x_1, X(t_n) =$  $_{1})=x_{n-1}\}=P\{X(t_{n})< x_{n} \mid X(t_{n-1})=x_{n-1}\}$ 转移概率: p(s,t;x,y)=  $P{X(t)< y | X(s)=x}$ .

状态空间: E ={x: X(t)=x,t∈T} 离散参数马氏链(马氏链):  $P\{X(m+k)=i_{m+k} | X(n_1)=i_{n1}, X(n_2)=i_{n2},...,X(n_j)=i_{n2},...,X(n_j)=i_{n2},...,X(n_j)=i_{n3},...,X(n_j)=i_{n4$  $i_{ni},X(m)=i_m$  =  $P\{X(m+k)=i_{m+k} | X(m)=i_m\}$ 马氏链{X(n),n=0,1,...}在 m 时刻的 k 步 转移概率: p<sub>ii</sub>(m,k)=P{X(m+k)=j|X(m)=i} 一步转移概率, 简称转移概率。 k 步转移矩阵,一步转移矩阵 P(m,1)简

称转移矩阵:  $p_{ij}^{P}(m,k) \ge 0 (p_{ij}(m \ge k)) (m \ge k) = 1$ 

齐次马氏链(与 m 无关):  $p_{ii}(m,k) = P\{X(m+k)=j \mid X(m)=i\} = p_{ij}(k);$  $p_{ii}(m,1) = p_{ij}(1) = p_{ij};$ K 步转移矩阵: P(m,k)=P(k)=(p;i(k)),i,i e E 转移矩阵:  $P(m,1) = P(1) = P = (p_{ij})_{i,j \in E}$  C-K 方程:  $p_{ij}(k+s) = \sum_{r \in E} p_{ir}(k) p_{rj}(s)$ n 步转移矩阵: P(n)=Pn 初始分布:  $p_i = P\{X(0)=i\} i \in E$ ,即 X(0)概率分布。记  $\widetilde{P}_0 = (p_i, i \in E)$ 。 绝对分布:  $p_i(\mathbf{n}) = P\{\mathbf{X}(\mathbf{n}) = \mathbf{i}\},$  记  $\widetilde{P}_n$   $p_j(\mathbf{n}) = \sum_{i=E} p_i p_{ij}(\mathbf{n}) \ (j \in E)$ 有限维劳布:

#### $P\{X(n1)\!\!=\!\!i1,\!X(n2)\!\!=\!\!i1,\!...,\!X(nk)\!\!=\!\!ik\}$

 $= \sum p_i \cdot p_{ii_1}(n_1) \cdot p_{i_1i_2}(n_2 - n_1) \cdot \cdots \cdot p_{i_{k-1}i_k}(n_k - n_{k-1})$ 

极限分布(最终分布):  $\Pi = (\pi_i, j \in E)$ 设齐次马氏链{X(n),n=0,1,2,...}的状态空 间有限 E={1,2,...,s}, 若存在正整数 n<sub>0</sub>,对任意 i,j∈E, n<sub>0</sub>步转移概率 p<sub>ii</sub>(n<sub>0</sub>)>0,则此链是**遍历的**,且极限分

有等 
$$\pi_j * \sum_{i \in E} \pi_i p_{ij}^{-1} : 或记 \Pi = \Pi P$$
 
$$\pi_j > 0, \sum_{j \in E} \pi_j = 1$$

互通:如果存在某一个 n≥1,使得 p<sub>ii</sub>(n)> 0, i,j∈E,则称从状态i可到达状态j, 记为 i→j;互通记 i↔j,状态性质相同。

 $f_{ii}(n)$ :表示从状态 i 出发经过 n 步首次 返回 i 的概率;

fii: 表示从状态 i 出发迟早返回 i 的概  $\mathcal{F}_{ij} = \sum f_{ij}(n)$ 

如果  $f_{ii}$ =1,则称状态 i 是常返状态; 正常返状态:则  $\mu_i = \sum_{\alpha^{n=1}} n f_{ii}(n) < +\infty$ 零常返状态: 则  $\mu_i = \sum_{n=1}^{\infty} n f_{ii}(n) = +\infty$  如果  $f_{ii} < 1$ ,则称状态 i 是非常返状态。

状态空间分解:

 $p_{ii}=1$  的状态 i 称为吸收状态。任何一 个吸收状态构成最小的单点闭集。

 $E = N + C_1 + C_2 + ... + C_k + ...$ 

其中 N 为非常返状态集合,  $C_1, C_2, ..., C_k, ...$ 是互不相交的常返闭集。 马氏链为不可约马氏链的充分必要条 件是任何两个状态都相通。

p<sub>ii</sub>(n)>0,且 n>1,则状态是周期的,反之。

生灭过程(针对连续参数齐次马氏链)

状态转移速度矩阵。Q= 

平稳分布:  $\Pi Q = 0 (\pi_i$ 竖乘之和) E={0,1,2,...,N}的平稳分布 Π={ $\pi_i$ ,i∈E}为:

$$\begin{cases} \pi_0 = \left(1 + \sum_{j=1}^{N} \frac{\lambda_0 \lambda_1 \cdots \lambda_{j-1}}{\mu_1 \mu_2 \cdots \mu_j}\right)^{-1} \\ \\ \pi_k = \frac{\lambda_0 \lambda_1 \cdots \lambda_{k-1}}{\mu_1 \mu_2 \cdots \mu_k} \pi_0 = \frac{\lambda_{k-1}}{\mu_k} \pi_{k-1}, k = 1, 2, \cdots, N \end{cases}$$

生灭过程在排队论中的应用:

M/M/1 损失制 (λ<sup>□</sup>μ)

一个服务员,发现服务台被占用,离

M/M/n 损失制 (λ図nμ)

$$\begin{cases} \pi_0 = \left[ \sum_{k=0}^n \frac{1}{k!} (\frac{\lambda}{\mu}) \right]^{-1} \\ \pi_k = \frac{1}{k!} (\frac{\lambda}{\mu})^k \pi_0 \end{cases}$$

损失概率 $\pi_n = \frac{1}{n!} \left(\frac{\lambda}{\mu}\right)^n \pi_0$   $\pi_0$  图图图图图 M/M/1 等待制 (λ□μ)

$$\begin{cases} \pi_0 = 1 - \rho \\ \pi_k = \rho^k (1 - \rho) \end{cases}$$

逗留平均数 $L_s = \frac{\lambda}{\mu - \lambda}$ 

等待平均数 $L_q = \frac{\lambda^2}{\mu(\mu - \lambda)}$ 

平均逗留时间 $W_s = \frac{1}{\lambda}L_s = \frac{1}{\mu - \lambda}$ 

 $_{\text{平均等待时间}}W_q = \frac{1}{\lambda}L_q = \frac{\lambda}{\mu(\mu - \lambda)}$ M/M/n 等待制 (λ<sup>2</sup>nμ)

$$M/M/n$$
 等待制( $\lambda = n$   
 $\pi_k = \frac{1}{k!} (\frac{1}{\mu})^k \pi_0 \ 0 < k \le n$   
 $\pi_k = \frac{1}{n^{k-1} n!} (\frac{1}{\mu})^k \pi_0$   
 $\pi_0 = [\sum_{k=0}^{n-1} \frac{1}{k!} (\frac{1}{\mu})^k + \frac{(\frac{1}{\mu})^n}{n!(1 - \frac{1}{n\mu})}]^{-1} (\lambda < n\mu)$ 

$$L_{q} = \frac{\rho^{n+1}}{n * n! (1 - \frac{\rho}{n})^{2}} \pi_{0}$$

 $L_s = L_q + \rho \ W_s = \frac{L_s}{\lambda} W_q = \frac{L_q}{\lambda}$ M/M/1 有限资源等待制(m  $\lambda 2 \mu 2 2 2 m$ 

$$\begin{cases} \pi_0 = [\sum_{k=0}^m \frac{m!}{(m-k)!} (\frac{\lambda}{\mu})^k]^{-1} \\ \pi_k = \frac{m!}{(m-k)!} (\frac{\lambda}{\mu})^k \pi_0 \quad 0 < k \leq m \end{cases}$$
 
$$L_s = m - \frac{\mu}{\lambda} (1 - \pi_0) L_q = L_s - (1 - \pi_0)$$
 
$$\text{PLIME} \hat{x} \hat{x}_e = \mu (1 - \pi_0)$$
 
$$W_s = \frac{L_s}{\lambda_e} W_q = \frac{L_q}{\lambda_e}$$

设备正常运行台数:  $K=m^{\lambda_{\varrho}}$  设备利用率:  $\tau = \frac{m-Ls}{m}$ 

M/M/n 有限资源等待制(m  $\lambda ? n \mu ? ? ? ? ? m, ? ? ? ? ? n$ 

$$\pi_k = \begin{cases} \frac{m!}{k!(m-k)!} (\frac{\lambda}{\mu})^k \pi_0, 0 \le k \le n \\ \frac{1}{n!n^{k-n}(m-k)!} (\frac{\lambda}{\mu})^k \pi_0, n < k < m \\ \frac{1}{n!n^{m-n}(m-k)!} (\frac{\lambda}{\mu})^m \pi_0, k = m \end{cases}$$

$$\pi_0 = [\sum_{k=0}^{n-1} C_m^k (\frac{\lambda}{\mu})^k + \sum_{k=n}^m \frac{1}{n!n^{k-n}(m-k)!} (\frac{\lambda}{\mu})^k ]$$

$$L_{q} = \sum_{k=0}^{m} (k-n)\pi_{k}, L_{s} = \sum_{k=0}^{m} k\pi_{k}$$

$$\lambda_e = \lambda (m - L_s).W_s = \frac{L_s}{\lambda_a}$$

M/M/1 有限等待空间 (顾客排队数最

大 N,否则离去。和 
$$\mu$$
 大 N,否则离去。和  $\mu$  )   

$$\begin{cases}
\pi_0 = \frac{1-\rho}{1-\rho^{N+2}} \\
\pi_k = (\frac{\lambda}{\mu})^k \pi_0 & 1 \le k \le N+1 \\
\pi_{\mathbb{B}} = \pi_{N+1} = \rho^{N+1} \pi_0 \\
L_s = \frac{\rho}{1-\rho} - \frac{(N+2)\rho^{N+2}}{1-\rho^{N+2}} \\
L_q = L_s - (1-\pi_0) \\
L_s = \frac{\rho}{1-\rho} + \frac{(N+2)\rho^{N+2}}{1-\rho^{N+2}} \\
L_q = \frac{(N+2)\rho^{N+2}}$$

$$L_q = L_s - (1 - \pi_0)$$

$$W_s = \frac{L_s}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0} \mathbb{D} W_q = W_s - \frac{1}{\mu} = \frac{1}{\lambda(1 - \rho^{N+1})\pi_0}$$

M/M/n 有限等待空间 (顾客排队数最 大 N, 否则离去。 $\lambda$ <sup>0</sup>nμ)

$$\begin{cases} \pi_k = \frac{1}{k!} \rho^k \pi_{0}.k = 1,2...n \\ \pi_{n+k} = \frac{1}{n!n^k} \rho^{n+k} \pi_{0}.k = 1,2...N \end{cases}$$

$$\pi_0 = \begin{cases} 1 + \rho + \frac{\rho^2}{2!}... + \frac{\rho^n}{n!} + \frac{\rho^n [\frac{\rho}{n} - (\frac{\rho}{n})^{N+1}]}{n!(1 - \frac{\rho}{n})} \end{cases}$$

$$L_q = \frac{\rho^{N+1} [1 - (N+1)(\frac{\rho}{n})^N + N(\frac{\rho}{n})^{N+1}] \pi_0}{n * n!(1 - \frac{\rho}{n})^2}$$

$$L_s = L_q + \rho (1 - \frac{\rho^{N+n}}{n^N * n!} \pi_0)$$

$$W_s = \frac{L_s}{\lambda} W_q = \frac{L_q}{\lambda}$$

### 排队论 第二章 无限源简单排队系统

M/M/1/∞(一个服务台,不空闲则等 待,均服从负指数分布)  $\rho = \lambda/\mu$ 

$$2000N = \frac{\rho}{1-\rho}$$
 
$$20000N_q = \sum_{j=0}^{\infty} jP_{j+1} = \frac{\rho^2}{1-\rho}$$

222222222222  $\mu b = \frac{1}{1-\rho}$ 

#### 最佳服务率: P53

可变输入率 M/M/1/∞

 $W(t) = 1 - \frac{e^{-\mu t}}{e^{\rho} - 1} \sum_{j=0}^{\infty} \frac{\rho^{j+1}}{(j+1)!} \sum_{k=0}^{j} \frac{(\mu t)^k}{k!}$  $\bar{W} = \frac{\rho e^{\rho}}{\mu (e^{\rho} - 1)}$ 单位时间内进入系统的平均顾客数:  $\lambda_e = \mu (1 - e^{-\rho})$  $N = \lambda_e W \ \ \mathbb{Z} N_q = \lambda_e W_q$ 可变服务率 M/M/1/∞ (队长小于 m,服务率 μ1,反之  $\rho_1 = \frac{\lambda}{\mu_1}, \quad \rho_2 = \frac{\lambda}{\mu_2}$ 

(看到队长为 k,进入系统的概率 ak)