

第5周作业

1、

(1)

$$\begin{aligned} \text{取出红球的概率} P_1 &= \frac{3}{3+4+5} = \frac{3}{12} \\ \text{取出白球的概率} P_2 &= \frac{4}{3+4+5} = \frac{4}{12} \\ \text{取出黑球的概率} P_3 &= \frac{5}{3+4+5} = \frac{5}{12} \end{aligned}$$

$$P(\text{2红3白1黑}) = \binom{6}{2\ 3\ 1} P_1^2 P_2^3 P_3 = \frac{25}{432}$$

(2)

$X + Y = 3$, 分别有 $(X, Y) = (0, 3), (1, 2), (2, 1), (3, 0)$

$$\begin{aligned} P(X = 0, Y = 3) &= \frac{C_4^3}{C_{12}^3} = \frac{1}{55} \\ P(X = 1, Y = 2) &= \frac{C_3^1 C_4^2}{C_{12}^3} = \frac{9}{110} \\ &\text{(把所有无序3元组看成基本事件)} \\ P(X = 2, Y = 1) &= \frac{C_3^2 C_4^1}{C_{12}^3} = \frac{3}{55} \\ P(X = 3, Y = 0) &= \frac{C_3^3}{C_{12}^3} = \frac{1}{220} \end{aligned}$$

事件	$X = 0, Y = 3$	$X = 1, Y = 2$	$X = 2, Y = 1$	$X = 3, Y = 0$
P	$\frac{1}{55}$	$\frac{9}{110}$	$\frac{3}{55}$	$\frac{1}{220}$

(3)

$$P(X = 1) = \frac{C_3^1 C_9^2}{C_{12}^3} = \frac{27}{55}$$

2、

$$F(x, y) = \int_{-\infty}^x \int_{-\infty}^y f(s, t) dt ds$$

$$\begin{aligned} LHS &= F(b, d) - F(a, d) - F(b, c) + F(a, c) \\ &= (F(b, d) - F(b, c)) - (F(a, d) - F(a, c)) \\ &= \left(\int_{-\infty}^b \int_{-\infty}^d f(s, t) dt ds - \int_{-\infty}^b \int_{-\infty}^c f(s, t) dt ds \right) - \left(\int_{-\infty}^a \int_{-\infty}^d f(s, t) dt ds - \int_{-\infty}^a \int_{-\infty}^c f(s, t) dt ds \right) \\ &= \int_{-\infty}^b \int_c^d f(s, t) dt ds - \int_{-\infty}^a \int_c^d f(s, t) dt ds \\ &= \int_a^b \int_c^d f(s, t) dt ds \\ &= P(a < X \leq b, c < Y \leq d) \end{aligned}$$

3、

(1)

$$\begin{aligned} S &= \pi \\ \iint_S p(x, y) &= 1 \\ p &= \frac{1}{\pi} \end{aligned}$$

$$f(X, Y) = \begin{cases} \frac{1}{\pi}, x^2 + y^2 \leq 1 \\ 0, x^2 + y^2 > 1 \end{cases}$$

(2)

$$\begin{aligned} X \text{的pdf} : f_x(X) &= \int_{-\infty}^{\infty} f(X, Y) dY = \frac{2}{\pi} \sqrt{4 - X^2} \\ \text{由对称性, } Y \text{的pdf} : f_y(Y) &= \frac{2}{\pi} \sqrt{4 - Y^2} \end{aligned}$$

(3)

记以 r 为半径的圆为 s

$$P(R \leq r) = \iint_S f(x, y) dx dy = \frac{\pi r^2}{\pi} = r^2$$

(4)

R 的cdf为 $F(r) = r^2$, pdf为 $f(r) = \frac{d}{dr} F(r) = 2r$

$$\begin{aligned} E(R) &= \int_0^1 r f(r) dr \\ &= \int_0^1 r \times 2r dr \\ &= \frac{2}{3} \end{aligned}$$

4、

$$\begin{aligned} f(x, y) &= \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)} \left[\frac{(x-\mu_x)^2}{\sigma_x^2} + \frac{(y-\mu_y)^2}{\sigma_y^2} - \frac{2\rho(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y} \right] \right) \\ f_x(x) &= \int_{-\infty}^{+\infty} \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)} \left[\frac{(x-\mu_x)^2}{\sigma_x^2} + \frac{(y-\mu_y)^2}{\sigma_y^2} - \frac{2\rho(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y} \right] \right) dy \\ &= \int_{-\infty}^{+\infty} \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)} \left[\left(\frac{y-\mu_y}{\sigma_y} - \frac{x-\mu_x}{\sigma_x}\right)^2 + (1-\rho^2) \frac{(x-\mu_x)^2}{\sigma_x^2} \right] \right) dy \\ &= \int_{-\infty}^{+\infty} \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)} \left[\left(\frac{y-\mu_y}{\sigma_y} - \frac{x-\mu_x}{\sigma_x}\right)^2 \right] - \frac{1}{2} \frac{(x-\mu_x)^2}{\sigma_x^2} \right) dy \\ &= \int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}\sigma_x} \exp\left(-\frac{1}{2(1-\rho^2)} \left[\left(\frac{y-\mu_y}{\sigma_y} - \frac{x-\mu_x}{\sigma_x}\right)^2 \right] - \frac{1}{2} \frac{(x-\mu_x)^2}{\sigma_x^2} \right) d\left(\frac{y-\mu_y}{\sqrt{2(1-\rho^2)}\sigma_y}\right) \\ &= \frac{1}{\sqrt{2\pi}\sigma_x} \sqrt{\pi} \exp\left(-\frac{1}{2} \frac{(x-\mu_x)^2}{\sigma_x^2}\right) \\ &= \frac{1}{\sqrt{2\pi}\sigma_x} e^{-\frac{(x-\mu_x)^2}{2\sigma_x^2}} \end{aligned}$$

5、

$$\begin{aligned}
 f(x|y) &= \frac{f(x, y)}{f_Y(y)} \\
 &= \frac{\frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)} \left[\frac{(x-\mu_x)^2}{\sigma_x^2} + \frac{(y-\mu_y)^2}{\sigma_y^2} - \frac{2\rho(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y}\right]\right)}{\frac{1}{\sqrt{2\pi}\sigma_y} e^{-\frac{(y-\mu_y)^2}{2\sigma_y^2}}} \\
 &= \frac{1}{\sqrt{2\pi}\sigma_x\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)} \left[-(1-\rho^2)\frac{(y-\mu_y)^2}{\sigma_y^2} + \frac{(x-\mu_x)^2}{\sigma_x^2} + \frac{(y-\mu_y)^2}{\sigma_y^2} - \frac{2\rho(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y}\right]\right) \\
 &= \frac{1}{\sqrt{2\pi}\sigma_x\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)} \left[\rho^2\frac{(y-\mu_y)^2}{\sigma_y^2} + \frac{(x-\mu_x)^2}{\sigma_x^2} - \frac{2\rho(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y}\right]\right) \\
 &= \frac{1}{\sqrt{2\pi}\sigma_x\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)} \left[\left(\frac{x-\mu_x}{\sigma_x} - \rho\frac{y-\mu_y}{\sigma_y}\right)^2\right]\right) \\
 &= \frac{1}{\sqrt{2\pi}\sigma_x\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2\sigma_x^2)} (x-\mu_x - \rho\frac{\sigma_x}{\sigma_y}(y-\mu_y))^2\right)
 \end{aligned}$$

6、

(1)

$$S = \frac{1}{2}$$

$$P(X, Y) = \begin{cases} 2, & X \geq 0 \text{ 且 } Y \geq 0 \text{ 且 } X + Y \leq 1 \\ 0, & \text{其他} \end{cases}$$

(2)

当 $Y < 0$ 或 $Y > 1$, $P(X, Y) = 0, f_Y(y) = 0$

当 $0 \leq Y \leq 1, P(X, Y) = 2, X \in [0, 1 - y], f_Y(y) = \int_0^{1-y} 2dx = 2(1 - y)$

(3)

当 $Y < 0$ 或 $Y > 1$, $f_Y(y) = 0$, 无条件密度函数

当 $0 \leq Y \leq 1, f_Y(y) = 2(1 - y)$

$$f(x|y) = \frac{f(x, y)}{f_Y(y)} = \begin{cases} \frac{1}{1-y}, & y \leq 1 \\ 0, & y > 1 \end{cases}$$

7、

(1)

$$\begin{aligned}
P(X_1 = k | X_1 + X_2 = n) &= \frac{P(X_1 = k, X_1 + X_2 = n)}{P(X_1 + X_2 = n)} \\
&= \frac{P(X_1 = k, X_2 = n - k)}{P(X_1 + X_2 = n)} \\
\text{由独立性} &= \frac{P(X_1 = k)P(X_2 = k)}{P(X_1 + X_2 = n)} \\
&= \frac{\frac{\lambda_1^k}{k!} e^{-\lambda_1} \frac{\lambda_2^{n-k}}{(n-k)!} e^{-\lambda_2}}{\frac{(\lambda_1 + \lambda_2)^n}{n!} e^{-(\lambda_1 + \lambda_2)}} \\
&= C_n^k \left(\frac{\lambda_1}{\lambda_1 + \lambda_2} \right)^k \left(\frac{\lambda_2}{\lambda_1 + \lambda_2} \right)^{n-k}
\end{aligned}$$

(2)

泊松分布可近似理解为二项分布，分别做 k 次实验一和 $n - k$ 次实验二，得到的就是类似的二项分布

8、

(1)

考虑0 - 60分钟, $P(X = k) = P(Y = k) = \frac{1}{60}, k \in [0, 60]$

因为 X, Y 独立, $P(X = k, Y = m) = \left(\frac{1}{60}\right)^2 = \frac{1}{3600}$

(2)

如果甲先到, $P = \int_0^{50} \int_{x+10}^{60} \frac{1}{3600} dy dx = \frac{25}{72}$

对称可得乙先到等甲超过十分钟的概率为 $\frac{25}{72}$

$$P_{tot} = \frac{25}{36}$$

9、

(1)

$$\begin{aligned}
H(x, y) &= \int_{-\infty}^x \int_{-\infty}^y f(s, t) dt ds \\
\frac{\partial}{\partial x} H(x, y) &= \int_{-\infty}^y f(x, t) dt \\
\lim_{y \rightarrow +\infty} \frac{\partial}{\partial x} H(x, y) &= \int_{-\infty}^{+\infty} f(x, t) dt = f_x(x)
\end{aligned}$$

$$\lim_{y \rightarrow +\infty} \frac{\partial}{\partial x} H(x, y) = G(y)(F'(x)[1 + \alpha(1 - F(x))(1 - G(y))] + F(x)[- \alpha F'(x)(1 - G(y))])$$

代入 $G(y) = 1, LHS = F'(x)$

由对称性, Y 的边际分布函数 $f_y(y) = G'(y)$

(2)

令 $F(x) = x, x \in [0, 1]$, 则 X 的边际分布函数 $F'(x) = x' = 1$, 类似定义 $G(y) = y$, 取 $\alpha = -1$ 和1

10、

令 $U = F(x), V = G(y)$, 则有

$$\begin{aligned}
C(u, v) &= P(U \leq u, V \leq v) \\
&= P(F_x(X) \leq F_x(x), F_y(Y) \leq F_y(y))
\end{aligned}$$

由累计分布函数单调性 $= P(X \leq x, Y \leq y)$

11、

(1)X连续, Y连续

$$\text{全概率公式: } \int_{-\infty}^{+\infty} f_{X|Y}(x|y)f_Y(y)dy$$

$$\text{Bayes'公式: } f_{Y|X}(y|x) = \frac{f(x,y)}{f_X(x)} = \frac{f_{X|Y}(x|y)f_Y(y)}{\int_{-\infty}^{+\infty} f_{X|Y}(x|y)f_Y(y)dy}$$

(2)X连续, Y离散

$$\text{全概率公式: } \sum_i f_{X|Y}(x|y)P_{Y_i}(y)$$

$$\text{Bayes'公式: } P_{Y|X}(y|x) = \frac{P(x,y)}{f_X(x)} = \frac{f_{X|Y}(x|y)P_Y(y)}{\sum_i f_{X|Y}(x|y)P_{Y_i}(y)}$$

(3)X离散, Y连续

$$\text{全概率公式: } \int_{-\infty}^{+\infty} P(x|y)f_Y(y)dy$$

$$\text{Bayes'公式: } f_{Y|X}(y|x) = \frac{P(x,y)}{f_X(x)} = \frac{P_{X|Y}(x|y)f_Y(y)}{\int_{-\infty}^{+\infty} P_{X|Y}(x|y)f_Y(y)}$$

(4)X离散, Y离散

$$\text{全概率公式: } \sum_i P(x|y_i)P(y_i)$$

$$\text{Bayes'公式: } P(y|x) = \frac{P(x|y)P(y)}{\sum_i P(x|y_i)P(y_i)}$$

12、

(1)

$$f(x,y) = \frac{c}{1+x^2+y^2}, x^2+y^2 \leq 1$$

$$\begin{aligned} \iint_{x^2+y^2 \leq 1} \frac{c}{1+x^2+y^2} dx dy &= \iint_{0 \leq r \leq 1, 0 \leq \theta \leq 2\pi} \frac{c}{1+r^2} r dr d\theta \\ &= \int_{0 \leq r \leq 1} \frac{2\pi r c}{1+r^2} dr \\ &= \int_{0 \leq r \leq 1} \frac{\pi c}{1+r^2} dr^2 \\ &= \pi c \ln(1+r^2) \Big|_{r=0}^1 \\ &= \pi c (\ln(2) - \ln(1)) \\ &= \pi c \ln 2 \\ \pi c \ln 2 &= 1 \\ c &= \frac{1}{\pi \ln 2} \end{aligned}$$

(2)

$$\begin{aligned}
f_X(x) &= \int_{-\infty}^{+\infty} f(x, y) dy \\
&= \int_{-\infty}^{+\infty} \frac{c}{1+x^2+y^2} dy \\
\text{令 } \alpha^2 = 1+x^2, &= \int_{-\infty}^{+\infty} \frac{c}{\alpha^2+y^2} dy \\
&= \int_{-\infty}^{+\infty} \frac{c}{\alpha(1+(\frac{y}{\alpha})^2)} d\frac{y}{\alpha} \\
&= \arctan(\frac{y}{\alpha}) \Big|_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{c}{\alpha}
\end{aligned}$$

类似可定义 $\beta^2 = 1+y^2$, $f_Y(y) = \arctan(\frac{x}{\beta}) \Big|_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} \frac{c}{\beta}$
显然 $f(x) \neq f_X(x)f_Y(y)$

13、
(1)

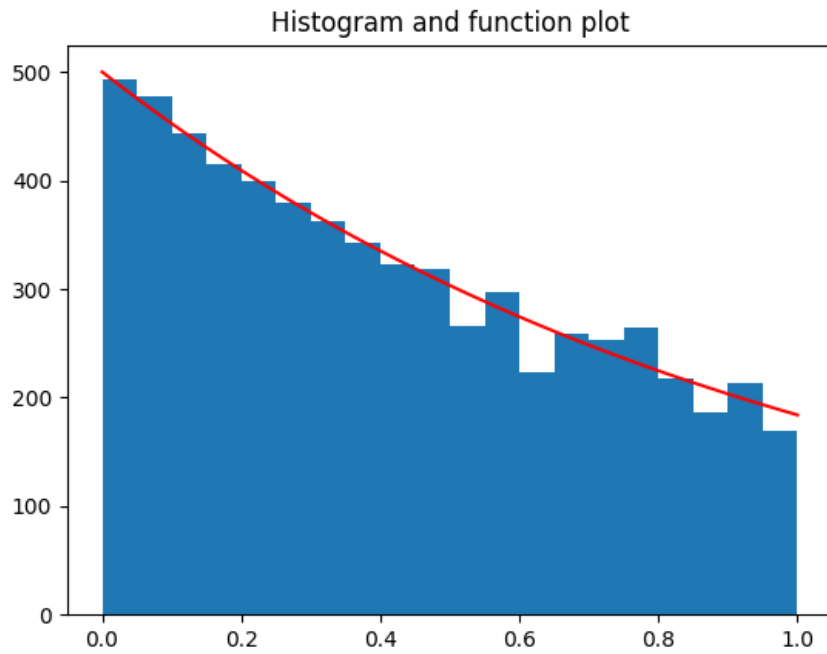
$$\begin{aligned}
f(x, y) &= \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)} \left[\frac{(x-\mu_x)^2}{\sigma_x^2} + \frac{(y-\mu_y)^2}{\sigma_y^2} - \frac{2\rho(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y} \right]\right) \\
&= \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left(-\frac{x^2+y^2-2\rho xy}{2(1-\rho^2)}\right)
\end{aligned}$$

由 X, Y 独立, $\rho = 0 = \frac{1}{2\pi} \exp(-\frac{x^2+y^2}{2})$
 $\geq \frac{1}{2\pi} e^{-1} \approx 0.059$

在 $x^2+y^2 \leq 1$ 时, $f(x, y) + \frac{xy}{100} \geq 0$, 且 $\iint_{x^2+y^2 \leq 1} \frac{xy}{100} dx dy = 0$ 满足 $\iint_{\Omega} f(x, y) = 1$, 所以 $g(x, y)$ 是二维密度函数
(2)

$$\begin{aligned}
f(x, y) &= \frac{1}{2\pi} \exp(-\frac{x^2+y^2}{2}) \\
g(x, y) &= \frac{1}{2\pi} \exp(-\frac{x^2+y^2}{2}) + \frac{xy}{100} \\
g_X(x) &= \int g(x, y) dy = \frac{1}{2\pi} e^{-\frac{x^2}{2}} \\
\text{类似得 } g_Y(y) &= \frac{1}{2\pi} e^{-\frac{y^2}{2}} \\
\text{且 } (U, V) &\text{显然不符合二元正态分布}
\end{aligned}$$

14、



```
import random
import math
import matplotlib.pyplot as plt
import numpy as np

# exp = np.random.exponential(scale=1, size=10000)
y = []
x = []
for i in range(10000):
    y.append(random.uniform(0, 1))
    x.append(-math.log(1 - y[i]))
plt.hist(x, bins=20, range=(0,1))

x_func = np.linspace(0, 1, 1000)
y_func = np.exp(-x_func) * 500
plt.title('Histogram and function plot')
plt.plot(x_func, y_func, 'r', label='500 * pdf function')
plt.savefig('2.png')
plt.show()
```