## STA2001 Tutorial 13

- 1. 5.7-13. Let  $X_1, X_2, \dots, X_{36}$  be a random sample of size 36 from the geometric distribution with pmf  $f(x) = (1/4)^{x-1}(3/4), x = 1, 2, 3, \dots$ . Approximate
  - (a)  $P(46 \le \sum_{i=1}^{36} X_i \le 49)$ .
  - (b)  $P(1.25 < \bar{X} < 1.50)$ .

Hint: Observe that the distribution of the sum is of the discrete type.

## **Solution:**

Let 
$$Y = \sum_{i=1}^{36} X_i$$
.

This question is similar as the previous one, so we have two ways to derive the expectation and variance of Y. The first way is to prove that Y follows a negative binomial distribution with parameters p=3/4 and r=36 (see Exercise 5.4-6), so that  $E(Y)=\frac{36}{3/4}=48$ ,  $Var(Y)=\frac{36\cdot(1-3/4)}{(3/4)^2}=16$ .

The second way, which is more general, starts from the expectation and variance of  $X_i$  and makes use of theorem 5.3-2

$$E(X_i) = \frac{1}{3/4} = \frac{4}{3}$$

$$Var(X_i) = \frac{1 - 3/4}{(3/4)^2} = \frac{4}{9}$$

$$E(Y) = E\left(\sum_{i=1}^{36} X_i\right) = \frac{4}{3} \cdot 36 = 48$$

$$Var(Y) = Var\left(\sum_{i=1}^{36} X_i\right) = \frac{4}{9} \cdot 36 = 16$$

Next, we use N(0,1) to approximate the distribution of Y according to the CLT.

(a) Since Y takes value in  $36, 37, 38, \dots$ , it has a discrete distribution.

$$P(46 \le Y \le 49) \approx P\left(\frac{46 - 48 - 0.5}{4} \le Z \le \frac{49 - 48 + 0.5}{4}\right)$$
$$= \Phi(0.375) - \Phi(-0.625)$$
$$= 0.3802$$

where the half-unit correction is used above.

(b) Since  $\bar{X} = \frac{Y}{36}$  and Y is a discrete distribution,  $\bar{X}$  is a discrete distribution as well.

We transform the probability related to  $\bar{X}$  into the probability related to Y, and then we can use the same method in question (a).

$$\begin{split} P(1.25 \leq \bar{X} \leq 1.50) &= P(1.25 \cdot 36 \leq 36 \cdot \bar{X} \leq 1.5 \cdot 36) \\ &= P(45 \leq Y \leq 54) \\ &\approx P\left(\frac{45 - 48 - 0.5}{4} \leq Z \leq \frac{54 - 48 + 0.5}{4}\right) \\ &= \Phi(1.625) - \Phi(-0.875) \\ &= 0.7571 \end{split}$$

2. 5.9-3. Let  $S^2$  be the sample variance of a random sample of size n from  $N(\mu, \sigma^2)$ . Show that the limit, as  $n \to \infty$ , of the mgf of  $S^2$  is  $e^{\sigma^2 t}$ .

## **Solution:**

Since  $S^2$  is the sample variance of a random sample of size n from  $N(\mu, \sigma^2)$ , we know that  $W = \frac{(n-1)S^2}{\sigma^2} \sim \chi^2(n-1)$  according to Theorem 5.5-2. a

Thus, the mgf of W is given by

$$M_W(t) = E(e^{tW}) = (1 - 2t)^{-(n-1)/2}, t < \frac{1}{2}$$

Note that  $S^2 = \frac{\sigma^2}{n-1}W$ , by a simple transformation the mgf of  $S^2$  can be written as

$$M_{S^2}(t) = M_W \left( \frac{\sigma^2}{n-1} \cdot t \right) = \left( 1 - \frac{\sigma^2 t}{(n-1)/2} \right)^{-(n-1)/2}$$

Take the limit for both sides and by the definition of  $e^x$  (actually, one of the definitions of exponential function), we have

$$\lim_{n \to \infty} M_{S^2}(t) = \lim_{n \to \infty} \left(1 - \frac{\sigma^2 t}{(n-1)/2}\right)^{-(n-1)/2} = \lim_{n \to \infty} \left(1 + \frac{\sigma^2 t}{-(n-1)/2}\right)^{-(n-1)/2} = e^{\sigma^2 t}$$

3. Let  $X_n \stackrel{d}{\to} X$  where  $X \equiv x$  is a constant random variable. Prove that  $X_n \stackrel{p}{\to} X$ . Note that  $\stackrel{d}{\to}$  is the convergence in distribution and  $\stackrel{p}{\to}$  is the convergence in probability.

**Solution:** Fix  $\varepsilon > 0$ . We have

$$\Pr(|X_n - X| \le \varepsilon) = \Pr(|X_n - x| \le \varepsilon)$$

$$= \Pr(x - \varepsilon \le X \le x + \varepsilon)$$

$$\geq \Pr(x - \varepsilon < X \le x + \varepsilon)$$

$$= F_{X_n}(x + \varepsilon) - F_{X_n}(x - \varepsilon)$$

$$\xrightarrow{n \to \infty} F_X(x + \varepsilon) - F_X(x - \varepsilon)$$

$$= \Pr(x - \varepsilon < X \le x + \varepsilon) = 1$$

Hence we have

$$\lim_{n \to \infty} \Pr\left(|X_n - X| \le \varepsilon\right) = 1$$

for any  $\varepsilon > 0$ , proving the result.