

# STA2001 Probability and Statistics (I)

## Lecture 8

Tianshi Chen

The Chinese University of Hong Kong, Shenzhen

# Review

We are interested in the number of successes in  $n$  Bernoulli trials.

## Definition[Binomial distribution]

A RV  $X$  is said to have a binomial distribution with  $n$  Bernoulli trials and the probability of success  $p$ , if the range space  $\bar{S} = \{0, 1, \dots, n\}$  and the pmf  $f(x)$  is in the form of

$$f(x) = \binom{n}{x} p^x (1-p)^{n-x}, \quad x = 0, 1, \dots, n.$$

We can simply denote it by  $X \sim b(n, p)$ .

## Section 2.5 Negative Binomial Distribution

# Negative Binomial Distribution

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Define a RV  $X$  to denote the trial number at which the  $r$ th success is observed. Then  $X$  has the range  $\bar{S} = \{r, r + 1, \dots\}$ .

Let  $f(x)$  denote the pmf of  $X$ . Then recall  $f(x) = P(X = x)$

# Negative Binomial Distribution

$$\begin{aligned} f(x) &= P(\{\text{at the } x\text{th trial, the } r\text{th success is observed}\}) \\ &= P(\underbrace{\{\text{for the first } x - 1 \text{ trials, } r - 1 \text{ success have been observed}\}}_A \\ &\quad \cap \underbrace{\{\text{at the } x\text{th trial, the outcome is a success}\}}_B) \\ &= P(A \cap B) = P(A)P(B) \text{ (because } A \text{ and } B \text{ are independent)} \end{aligned}$$

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$$P(A) = \binom{x-1}{r-1} p^{r-1} (1-p)^{x-r}, \quad P(B) = p$$

Therefore

$$f(x) = \binom{x-1}{r-1} p^r (1-p)^{x-r}, \quad x = r, r+1, \dots$$

# Negative Binomial Distribution

## Definition[Negative Binomial Distribution]

A RV  $X$  is said to have a negative binomial distribution with the probability of success  $p$  and the number of successes  $r$  we are interested in, if the range  $\bar{S} = \{r, r+1, \dots\}$  and the pmf  $f(x)$  is in the form of

$$f(x) = \binom{x-1}{r-1} p^r (1-p)^{x-r}, \quad x = r, r+1, \dots$$

This distribution get its name due to the negative binomial series

$$(1-w)^{-r} = \sum_{x=r}^{\infty} \binom{x-1}{r-1} w^{x-r}$$



# Geometric Distribution

## Definition[Geometric Distribution]

A RV  $X$  is said to have a geometric distribution with the probability of success  $p$ , if the range  $\bar{S} = \{1, 2, \dots\}$  and the pmf  $f(x)$  is in the form of

$$f(x) = p(1 - p)^{x-1}, \quad x = 1, 2, \dots.$$

For a positive integer  $k$ ,

$$P(X > k) = \sum_{x=k+1}^{\infty} p(1 - p)^{x-1} = \frac{(1 - p)^k p}{1 - (1 - p)} = (1 - p)^k$$

$$P(X \leq k) = \sum_{x=1}^k p(1 - p)^{x-1} = 1 - P(X > k) = 1 - (1 - p)^k$$

## Example 1, page 83

Biology students are checking eye color of fruit flies. For each fly,

$$P(\text{white}) = \frac{1}{4}, \quad P(\text{red}) = \frac{3}{4}.$$

Assume the observations are independent Bernoulli trials.

To observe 1 white fly, what's the probability one has to check

at least 4 flies?

at most 4 flies?

4 flies?

## Example 1, page 83

We define  $X$  to be the number of fruit flies one has to check until the first white-eye fly is observed.

Then  $X$  has the geometric distribution with probability of success  $1/4$ . So the probability one has to check

$$\text{at least 4 flies?} \longrightarrow P(X \geq 4) = P(X > 3) = \left(1 - \frac{1}{4}\right)^3 = \left(\frac{3}{4}\right)^3$$

$$\text{at most 4 flies?} \longrightarrow P(X \leq 4) = 1 - \left(1 - \frac{1}{4}\right)^4$$

$$4 \text{ flies?} \longrightarrow P(X = 4) = \frac{1}{4} \cdot \left(\frac{3}{4}\right)^3$$

# Mathematical Expectations of Negative Binomial Distribution

Mean and Variance

$$\text{Mean : } E[X] = \frac{r}{p}$$

$$\text{Variance : } \text{Var}[X] = E[X^2] - (E[X])^2 = \frac{r(1-p)}{p^2}$$

can be calculated by using the mgf

$$\text{Mgf : } M(t) = E[e^{tX}] = \frac{(pe^t)^r}{[1 - (1-p)e^t]^r}, \text{ for } (1-p)e^t < 1$$

which can be obtained by using the negative binomial series

$$(1-w)^{-r} = \sum_{x=r}^{\infty} \binom{x-1}{r-1} w^{x-r}$$

## Section 2.6 Poisson Distribution

# Motivation

Description: There are experiments that result in counting the number of times that particular events occur within a given period or for a given physical object:

- ▶ the number of flaws in a 100 feet long wire.
- ▶ the number of customers that arrive at a ticket window between 7:00-8:00 pm.

Counting such events can be seen as observations of a RV associated with an approximate Poisson process (APP).

# Approximate Poisson Process (APP)

## Definition[Approximate Poisson Process (APP)]

Let the number of occurrences of some event in a given continuous interval be counted. Then we have an APP with parameter  $\lambda > 0$  if

- (a) The number of occurrences in non-overlapping subintervals are independent.
- (b) The probability of exactly one occurrence in a sufficiently short subinterval of length  $h$  is approximately  $\lambda h$ .
- (c) The probability of two or more occurrences in a sufficiently short subinterval is essentially 0.

# Poisson Distribution

Consider a random experiment described by APP. Let  $X$  denote the number of occurrences in **an interval with length 1**. We aim to find an approximation for  $f(x) = P(X = x)$  with  $x = 0, 1, 2, \dots$ .

To this goal,

1. Partition the unit interval into  $n$  equally spaced subintervals.



2. If  $n$  is sufficiently large ( $n \gg x$ ),  $P(X = x)$  can be approximated by the probability that exactly  $x$  of these  $n$  subintervals each has one occurrence.



# Poisson Distribution

- 2.1 By condition (c), the probability of two or more occurrences in any sufficiently short subinterval is 0. [ *$n$  Bernoulli experiments.*]
- 2.2 By condition (b), the probability of one occurrence in any subinterval (with length  $\frac{1}{n}$ ) is approximately  $\lambda \frac{1}{n}$ . [*Same probability of success  $\lambda \frac{1}{n}$ .*]
- 2.3 By condition (a), the  $n$  Bernoulli experiments are independent. [ *$n$  Bernoulli trials with probability of success  $\lambda \frac{1}{n}$ .*]

Therefore occurrence and nonoccurrence in the  $n$  subintervals are  $n$  Bernoulli trials with probability of success  $\frac{\lambda}{n}$

# Poisson Distribution

3. Therefore,  $P(X = x)$  can be approximated by the probability of  $x$  successes for  $b(n, p = \frac{\lambda}{n})$

$$\frac{n!}{x!(n-x)!} \left(\frac{\lambda}{n}\right)^x \left(1 - \frac{\lambda}{n}\right)^{n-x}$$

4. Let  $n \rightarrow \infty$ . Then

$$\begin{aligned} \lim_{n \rightarrow \infty} \frac{n!}{x!(n-x)!} \left(\frac{\lambda}{n}\right)^x \left(1 - \frac{\lambda}{n}\right)^{n-x} \\ = \lim_{n \rightarrow \infty} \frac{n!}{(n-x)!n^x} \cdot \frac{\lambda^x}{x!} \left(1 - \frac{\lambda}{n}\right)^n \left(1 - \frac{\lambda}{n}\right)^{-x} \end{aligned}$$

# Poisson Distribution

Noting

$$\lim_{n \rightarrow \infty} \frac{n!}{(n-x)!n^x} = 1$$

$$\lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^n = e^{-\lambda},$$

$$\lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^{-x} = 1$$

We have

$$P(X = x) = \lim_{n \rightarrow \infty} \frac{n!}{(n-x)!n^x} \cdot \frac{\lambda^x}{x!} \left(1 - \frac{\lambda}{n}\right)^n \left(1 - \frac{\lambda}{n}\right)^{-x} = \frac{\lambda^x e^{-\lambda}}{x!}$$

# Poisson Distribution

It can be verified

$$f(x) = \frac{\lambda^x e^{-\lambda}}{x!}, \quad x = 0, 1, \dots$$

is a well-defined pmf.

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## Question

What's the implication of  $\lambda$ ?

# Mean and Variance

The mgf of a Poisson distributed RV  $X$  is

$$\begin{aligned}M(t) &= E[e^{tX}] = \sum_{x=0}^{\infty} e^{tx} \frac{\lambda^x e^{-\lambda}}{x!} \\&= e^{-\lambda} \sum_{x=0}^{\infty} \frac{(\lambda e^t)^x}{x!} = e^{-\lambda} \cdot e^{\lambda e^t} = e^{\lambda(e^t-1)}\end{aligned}$$

$$M'(t) = \lambda e^t e^{\lambda(e^t-1)} \Rightarrow M'(0) = \lambda$$

$$M''(t) = \lambda e^t e^{\lambda(e^t-1)} + \lambda^2 e^{2t} e^{\lambda(e^t-1)} \Rightarrow M''(0) = \lambda + \lambda^2 = E[X^2]$$

$$E[X] = M'(0) = \lambda$$

$$\text{Var}[X] = E[X^2] - (E[X])^2 = \lambda + \lambda^2 - \lambda^2 = \lambda$$

$\lambda$  is the mean and variance of  $X \sim \text{Poisson}(\lambda)$ : the average number of occurrences in **the unit interval!**

## Example 1, page 91

### Question

In SZ, telephone calls to 110 come on the average of 2 calls every 3 minutes. If one models with APP, what's the probability of 5 or more calls arrive in a 9-minute period?

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We need to determine  $\lambda$ .

$$E[X] = 6 = \lambda \implies f(x) = \frac{6^x e^{-6}}{x!}$$

Therefore,

$$P(X \geq 5) = 1 - P(X \leq 4) = 1 - \sum_{x=0}^4 \frac{6^x e^{-6}}{x!}$$