## Review of the last lecture

#### Key concepts and/or techniques:

Mathematical Expectation:

$$E(g(X,Y)) = \sum_{(x,y)\in\overline{S}} g(x,y)f(x,y)$$

Covariance and correlation coefficient:

To study the relation between 2 RVs

$$Cov(X, Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y]$$
$$\rho(X, Y) = \frac{Cov(X, Y)}{\sqrt{Var(X)}\sqrt{Var(Y)}}, \ \ Var(X) > 0, \ Var(Y) > 0.$$

 Interpretation and properties of covariance and correlation coefficient

#### Review of the last lecture

- Cov(X, Y) = E(XY) − E(X)E(Y)
   When Cov(X, Y) = 0, X and Y are uncorrelated.
   When Cov(X, Y) > 0, X and Y are positively correlated.
   When Cov(X, Y) < 0, X and Y are negatively correlated.</li>
- Interpretation: Roughly speaking, a positive or negative covariance indicate that the values of X E(X) and Y E(Y) obtained in a single experiment "tend" to have the same or the opposite sign respectively.
- ▶ Independence of X and Y  $\Rightarrow$  uncorrelation of X and Y, but the converse is in general not true.

#### Review of the last lecture

#### Correlation coefficient

$$\rho(X,Y) = \frac{Cov(X,Y)}{\sqrt{Var(X)}\sqrt{Var(Y)}}$$

▶ It is a normalized version of Cov(X,Y) and in fact  $-1 \le \rho(X,Y) \le 1$  and the size of  $|\rho|$  provides a normalized measure of the extent to which this is true.

•  $\rho=1$  or  $(\rho=-1)$  if and only if there exists a positive. (or negative,respectively) constant c such that

$$Y - E(Y) = c(X - E(X))$$

# STA2001 Probability and Statistics (I)

Lecture 16

Tianshi Chen

The Chinese University of Hong Kong, Shenzhen

#### **Section 4.3 Conditional distribution**

## **Conditional Distribution**

Motivation: the conditional probability distribution is a probability distribution that describes the distribution of probability of events of a RV given the occurrence of a particular event

Assume that X and Y have a joint pmf  $f(x,y): \overline{S} \to (0,1]$ . The marginal pmf of X and Y are

$$f_X(x): \overline{S_X} \to (0,1]$$
  $f_Y(y): \overline{S_Y} \to (0,1]$ 

$$\overline{S_X} = \{\text{all possible values of } X \text{ in } \overline{S}\}$$

$$\overline{S_X}(y) = \{x | (x, y) \in \overline{S}\} \text{ for } y \in \overline{S_Y}$$

$$\overline{S_Y} = \{ \text{all possible values of } Y \text{ in } \overline{S} \}$$

$$\overline{S_Y}(x) = \{ y | (x, y) \in \overline{S} \} \text{ for } x \in \overline{S_X}$$

## **Conditional Distribution**

By definition,

$$f(x,y) = P(X = x, Y = y)$$

$$\stackrel{\triangle}{=} P(\{X = x, Y = y\}), (x,y) \in \overline{S}$$

$$f_X(x) = P_X(X = x)$$

$$\stackrel{\triangle}{=} P(\{X = x, Y \in \overline{S_Y}(x)\}) = \sum_{y \in \overline{S_Y}(x)} f(x,y)$$

$$f_Y(y) = P_Y(Y = y)$$

$$\stackrel{\triangle}{=} P(\{X \in \overline{S_X}(y), Y = y\}) = \sum_{x \in \overline{S_Y}(y)} f(x,y)$$

## **Conditional Distribution**

Let

$$A = \{X = x, Y \in \overline{S_Y}(x)\}$$
  
$$B = \{X \in \overline{S_X}(y), Y = y\}$$

Then for  $(x, y) \in \overline{S}$ ,

$$A \cap B = \{X = x, Y = y\}$$

and recall the conditional probability of event A given event B is

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{f(x,y)}{f_Y(y)}$$

under the assumption P(B) > 0, i.e.,  $f_Y(y) > 0$ .

# **Conditional pmf**

#### Definition

Conditional pmf of X given Y = y is defined by

$$g(x|y) = \frac{f(x,y)}{f_Y(y)}, \quad x \in \overline{S_X}(y)$$

provided that  $f_Y(y) > 0$ .

Similarly, the conditional pmf of Y given that X=x is defined by

$$h(y|x) = \frac{f(x,y)}{f_X(x)}, \quad y \in \overline{S_Y}(x)$$

provided that  $f_X(x) > 0$ .

- ▶ What is the interpretation of the conditional pmf g(x|y)?
- ▶ What if X and Y are independent?



#### **Some Remarks**

Conditional pmf is a well-defined pmf:

- 1. h(y|x) > 0
- $2. \sum_{y \in \overline{S_Y}(x)} h(y|x) = 1$

$$\sum_{y \in \overline{S_Y}(x)} h(y|x) = \sum_{y \in \overline{S_Y}(x)} \frac{f(x,y)}{f_X(x)} = \frac{\sum_{y \in \overline{S_Y}(x)} f(x,y)}{f_X(x)} = \frac{f_X(x)}{f_X(x)} = 1$$

3. for  $A \subseteq \overline{S_Y}(x)$ 

$$P(Y \in A|X = x) = \frac{P(X = x, Y \in A)}{P(X = x)}$$
$$= \frac{\sum_{y \in A} f(x, y)}{f_X(x)} = \sum_{y \in A} h(y|x)$$

Therefore, h(y|x) (resp. g(x|y)) determines the distribution of probability of events of Y (resp. X) given X = x (resp. Y = y).

## **Some Remarks**

If X and Y are independent, then  $f(x, y) = f_X(x)f_Y(y)$  and thus

$$g(x|y) = f_X(x)$$
, and  $h(y|x) = f_Y(y)$ ,

#### which implies

- ▶ the occurrence of the event Y = y does not change the probability of the occurrence of events of X
- ▶ the occurrence of the event X = x does not change the probability of the occurrence of events of Y

Now, the implication of independent RVs becomes clear.

## Example 1

#### Question

Let X and Y have the joint pmf

$$f(x,y) = \frac{x+y}{21}$$
,  $x = 1,2,3$ ;  $y = 1,2$ .

We have showed

$$f_X(x) = \frac{2x+3}{21}, \quad x = 1, 2, 3$$
  
 $f_Y(y) = \frac{y+2}{7}, \quad y = 1, 2.$ 

- Q1: What is the conditional pmf of X given Y = y?
- Q2: What is the conditional pmf of Y given X = x?
- Q3: What is  $P(1 \le X \le 2 | Y = 1)$ ?

## Example 1

Q1: The conditional pmf of X given Y = y is

$$g(x|y) = \frac{f(x,y)}{f_Y(y)} = \frac{x+y}{21} / (\frac{y+2}{7}) = \frac{x+y}{3(y+2)},$$
  
  $x = 1, 2, 3;$   $y = 1, 2.$ 

Q2: The conditional pmf of Y given X = x is

$$h(y|x) = \frac{f(x,y)}{f_X(x)} = \frac{x+y}{21} / (\frac{2x+3}{21}) = \frac{x+y}{2x+3}$$
  
  $x = 1, 2, 3;$   $y = 1, 2.$ 

Q3:

$$P(1 \le X \le 2|Y=1) = \sum_{x=1}^{2} g(x|1) = \sum_{x=1}^{2} \frac{x+1}{3(1+2)} = \frac{5}{9}$$

# **Conditional Mathematical Expectation**

ightharpoonup Let g(Y) be a function of Y.

Then the conditional expectation of g(Y) given X = x

$$E(g(Y)|X = x) = \sum_{y \in \overline{S_Y}(x)} g(y)h(y|x)$$

 $\blacktriangleright \text{ When } g(Y) = Y,$ 

$$E(Y|X=x) = \sum_{y \in \overline{S_Y}(x)} yh(y|x) \rightarrow \text{conditional mean}$$

# **Conditional Mathematical Expectation**

When 
$$g(Y) = [Y - E(Y|X = x)]^2$$

$$Var(Y|X = x) \stackrel{\triangle}{=} E\{[Y - E(Y|X = x)]^2 | X = x\}$$

$$= \sum_{y \in \overline{S_Y}(x)} [y - E(Y|X = x)]^2 h(y|x)$$

$$= E(Y^2|X = x) - [E(Y|X = x)]^2$$

$$\rightarrow \text{conditional variance}$$

## **Example 1, continued**

#### Question

Let X and Y have the joint pmf

$$f(x,y) = \frac{x+y}{21}, \quad x = 1,2,3; \quad y = 1,2.$$

We have showed

$$f_X(x) = \frac{2x+3}{21}, \quad x = 1, 2, 3$$
  
 $f_Y(y) = \frac{y+2}{7}, \quad y = 1, 2.$ 

- Q1: What is the expectation of Y given X = 3?
- Q2: What is the variance of Y given X = 3?

# Example 1, continued

$$Q1: E(Y|X=3) = \sum_{y \in \overline{S_Y}(3)} yh(y|3) = \sum_{y=1}^{2} y(\frac{3+y}{9}) = \frac{14}{9}$$

$$Q2: Var(Y|X=3) = \sum_{y \in \overline{S_Y}(3)} [y - E(Y|X=3)]^2 h(y|3)$$

$$= \sum_{y=1}^{2} (y - \frac{14}{9})^2 \frac{3+y}{9} = \frac{20}{81}$$

## **Example 2: Trinomial Distribution**

Assume that  $(X, Y) \sim \text{Trinomial}(n, p_X, p_Y)$ . Then what is the conditional pmf of Y given X = x, i.e., h(y|x)?

## **Example 2: Trinomial Distribution**

Assume that  $(X, Y) \sim \text{Trinomial}(n, p_X, p_Y)$ . Then what is the conditional pmf of Y given X = x, i.e., h(y|x)?

First, we recall that

$$f(x,y) = \frac{n!}{x!y!(n-x-y)!} p_X^x p_Y^y (1 - p_X - p_Y)^{n-x-y}, (x,y) \in \overline{S},$$

$$\overline{S} = \{(x,y)|x+y \le n, x = 0, 1, \dots, n, y = 0, 1, \dots, n\}$$
and  $f_X(x) = \frac{n!}{x!(n-x)!} p_X^x (1 - p_X)^{n-x}, x \in \overline{S_X} = \{0, 1, \dots, n\}.$ 

## Example 2: Trinomial Distribution

Assume that  $(X, Y) \sim \text{Trinomial}(n, p_X, p_Y)$ . Then what is the conditional pmf of Y given X = x, i.e., h(y|x)?

First, we recall that

$$f(x,y) = \frac{n!}{x!y!(n-x-y)!} p_X^x p_Y^y (1-p_X-p_Y)^{n-x-y}, (x,y) \in \overline{S},$$

$$\overline{S} = \{(x,y)|x+y \le n, x = 0, 1, \dots, n, y = 0, 1, \dots, n\}$$
and  $f_X(x) = \frac{n!}{x!(n-x)!} p_X^x (1-p_X)^{n-x}, x \in \overline{S_X} = \{0, 1, \dots, n\}.$ 

Then we have  $\overline{S_Y}(x) = \{0, \dots, n-x\}$ , and for any  $y \in \overline{S_Y}(x)$ ,

$$h(y|x) = \frac{f(x,y)}{f_X(x)} = \frac{\frac{n!}{x!y!(n-x-y)!} p_X^x p_Y^y (1 - p_X - p_Y)^{n-x-y}}{\frac{n!}{x!(n-x)!} p_X^x (1 - p_X)^{n-x}}$$

$$= \frac{(n-x)!}{y!(n-x-y)!} \frac{p_Y^y (1 - p_X - p_Y)^{n-x-y}}{(1 - p_X)^y (1 - p_X)^{n-x-y}}$$

$$= \frac{(n-x)!}{y!(n-x-y)!} \left(\frac{p_Y}{1 - p_X}\right)^y \left(1 - \frac{p_Y}{1 - p_X}\right)^{n-x-y}$$

## Example 2: Covariance for Trinomial Distribution

What is the conditional distribution of Y given X = x?

$$Y|X = x \sim b(n-x, \frac{p_Y}{1-p_X}).$$

## Example 2: Covariance for Trinomial Distribution

What is the conditional distribution of Y given X = x?

$$Y|X = x \sim b(n-x, \frac{p_Y}{1-p_X}).$$

Then what is Cov(X, Y)?

## Example 2: Covariance for Trinomial Distribution

What is the conditional distribution of Y given X = x?

$$Y|X = x \sim b(n-x, \frac{p_Y}{1-p_X}).$$

Then what is Cov(X, Y)?

$$Cov(X,Y) = \sum_{(x,y)\in\overline{S}} (x - E[X])(y - E[Y])f(x,y)$$

$$= \sum_{x\in\overline{S_X}} (x - E[X]) \sum_{y\in\overline{S_Y}(x)} (y - E[Y])f(x,y)$$

$$= \sum_{x\in\overline{S_X}} (x - E[X])f_X(x) \sum_{y\in\overline{S_Y}(x)} (y - E[Y])h(y|x)$$
expectation of function of X expectation of function of Y|X = x

The calculation of Cov(X, Y) is converted to that of mathematical expectation of functions of two univariate RV, which could be easier if the distributions of two univariate RV are known.

Section 4.4 Bivariate Distribution of Continuous Type

## **Bivariate Continuous RV**

#### Definition

Let X and Y be two continuous random variables and (X, Y) be a pair of RVs with their range denoted by  $\overline{S} \subseteq R^2$ . Then (X, Y) or X and Y is said to be a bivariate continuous RV.

Moreover, let  $\overline{S_X} \subseteq R$  and  $\overline{S_Y} \subseteq R$  denote the range of X and Y, respectively.

$$\overline{S} = \{ \text{all possible values of } (X, Y) \}$$

$$\overline{S_X} = \{ \text{all possible values of } X \} = \{ x | (x, y) \in \overline{S} \}$$

$$\overline{S_Y} = \{ \text{all possible values of } Y \} = \{ y | (x, y) \in \overline{S} \}$$

Then, it holds that

$$\overline{S} \subseteq \overline{S_X} \times \overline{S_Y} = \{(x, y) | x \in \overline{S_X}, y \in \overline{S_Y}\}$$

# Roadmap for bivariate continuous random distributions

To study the bivariate continuous random variable

 $\mathsf{discrete}\ \mathsf{RV} \longrightarrow \mathsf{continuous}\ \mathsf{RV} \qquad \mathsf{Mathematical}\ \mathsf{expectations}$ 

 $pmf \longrightarrow pdf$  mean

joint pmf  $\longrightarrow$  joint pdf variance

marginal pmf — marginal pdf covariance

conditional pmf  $\longrightarrow$  conditional pdf correlation coefficient

# Joint pdf

#### Definition

The joint pdf of two continuous RVs X and Y is a function  $f(x,y): \overline{S} \to (0,\infty)$  with the following properties:

1. 
$$f(x,y) > 0, (x,y) \in \overline{S}$$

2. 
$$\iint_{\overline{S}} f(x,y) dx dy = 1$$

3.

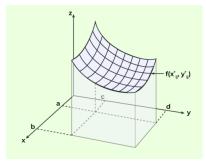
$$P((X, Y) \in A) \stackrel{\Delta}{=} P(\{(X, Y) \in A\})$$
$$= \iint_{A} f(x, y) dx dy, A \subseteq \overline{S}$$

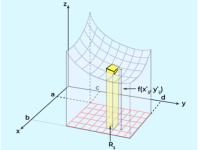
#### Remarks

Recall the geometric interpretation of double integral:

$$P((X,Y) \in A) = \iint_A f(x,y) dxdy$$

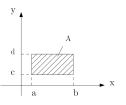
calculates the volume of the solid under the surface z = f(x, y) over the region A in the the xy-plane.





#### Remarks

- Very often, we extend the definition domain of f(x,y) from  $\overline{S}$  to  $R \times R$  by letting f(x,y) = 0, for  $(x,y) \notin \overline{S}$  and  $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x,y) dx dy = 1$
- Bottom line:



If the set A is rectangular with its line segments parallel to the coordinate axes, i.e.,

$$A = \{(x, y) | a \le x \le b, c \le y \le d\},$$

then the double integral becomes

$$P((X,Y) \in A) = \int_a^b \int_c^d f(x,y) dy dx$$