Stat 4201 Homework 4

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Problem 1

- a. The scatter plot is shown in Fig 1.
- b. The output of the simple linear regression model from R is shown below:

Call:

lm(formula = cesium\$Mushrooms ~ cesium\$Soils)

Residuals:

```
Min 1Q Median 3Q Max -47.279 -30.319 -9.483 31.000 54.271
```

Coefficients:

```
Estimate Std. Error t value Pr(>|t|)
(Intercept) 16.72569 12.41954 1.347 0.19807
cesium$Soils 0.09590 0.02993 3.205 0.00591 **
```

Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1

Residual standard error: 36.56 on 15 degrees of freedom Multiple R-squared: 0.4064, Adjusted R-squared: 0.3668 F-statistic: 10.27 on 1 and 15 DF, p-value: 0.005909

c. The output of the simple linear regression model excluding sample number 17 from R is shown below:

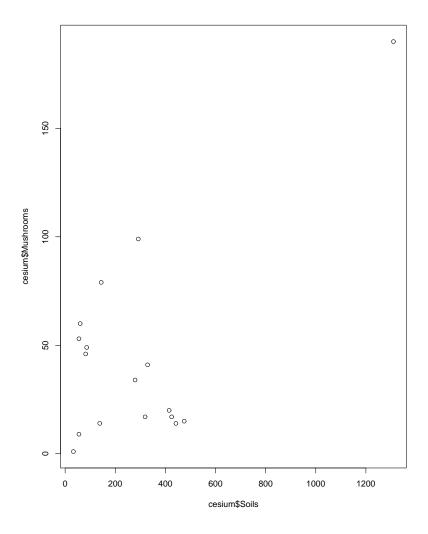


Figure 1: Problem 1-a: scatterplot

Call: lm(formula = cesium.trim\$Mushrooms ~ cesium.trim\$Soils)

Residuals:

Min 1Q Median 3Q Max -41.658 -13.938 -4.061 9.744 65.908

Coefficients:

```
Estimate Std. Error t value Pr(>|t|)
(Intercept) 43.87726 12.25571 3.580 0.00301 **
cesium.trim$Soils -0.03693 0.04454 -0.829 0.42085
---
Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1
```

Residual standard error: 27.76 on 14 degrees of freedom Multiple R-squared: 0.04682, Adjusted R-squared: -0.02126 F-statistic: 0.6877 on 1 and 14 DF, p-value: 0.4208

d. Based on the scatterplot from part a, we can see that example number 17 is very likely to be an outlier since it is far from every other samples. After excluding example number 17, we get the linear model described in part c. Note that this model's R-squared is 0.0468, which is very close to 0. This indicates that the simple linear regression model is poorly fitted. From the coefficients' p-value, we can see that it is very likely that the model is a constant model.

In conclusion, there is no linear relationship between Cesium concentration in soil and Cesium concentration in mushrooms after Chernobyl accident.

e.

About linearity, based on the R-squared from part b, we can see that the simple linear regression model is poorly fitted.

About normality, I use Shapiro-Wilk normality test on the residuals, here is the result from R:

Shapiro-Wilk normality test

```
data: res.p1 W = 0.9113, p-value = 0.1053
```

The p-value is 0.1053. This indicates the data is normal.

About homoscedasticity, I use the Residual vs Fitted plot to do the analysis. The plot is shown in Fig 2. Based on the plot, we can see that the data is not homoscedastic.

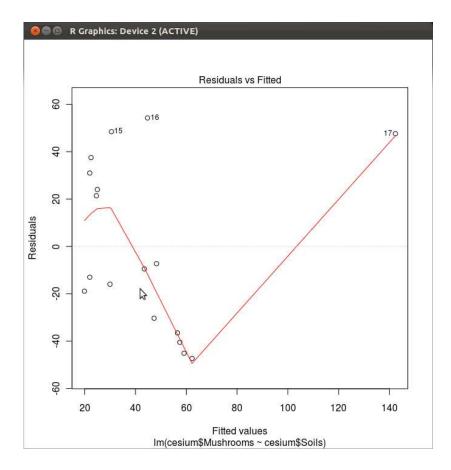


Figure 2: Problem 1-e: Residual vs Fitted

About uncorrelated error, I use Perason's product-moment test to do the analysis. And here is the result from R:

Pearson's product-moment correlation

data: cesium\$Mushrooms and cesium\$Soils
t = 3.2045, df = 15, p-value = 0.005909
alternative hypothesis: true correlation is not equal to 0
95 percent confidence interval:

```
0.2261250 0.8558834 sample estimates: cor 0.6374843
```

As we can see, Cesium concentrations in mushrooms and soils are correlated.

Problem 2

a. The output of the simple linear regression model from R is shown below:

Call:

b.

Residuals:

```
Min 1Q Median 3Q Max -7.2377 -1.7117 -0.4551 2.3614 5.6978
```

Coefficients:

```
Estimate Std. Error t value Pr(>|t|)
(Intercept) -39.9197 11.8960 -3.356 0.00375 **
stack.x[, 1] 0.7156 0.1349 5.307 5.8e-05 ***
stack.x[, 2] 1.2953 0.3680 3.520 0.00263 **
stack.x[, 3] -0.1521 0.1563 -0.973 0.34405
---
```

Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1

Residual standard error: 3.243 on 17 degrees of freedom Multiple R-squared: 0.9136, Adjusted R-squared: 0.8983 F-statistic: 59.9 on 3 and 17 DF, p-value: 3.016e-09

About linearity, the R-squared for this model is 0,9136. This indicates that the simple linear regression model is fitted well.

About normality, I use Shapiro-Wilk normality test on the residuals, here is the result from R:

Shapiro-Wilk normality test

data: e W = 0.974, p-value = 0.8186

The p-value is 0.8186. This indicates the data is normal.

About homoscedasticity, I use the Residual vs Fitted plot to do the analysis. The plot is shown in Fig 3. Based on the plot, we can see that the data is nearly homoscedastic.

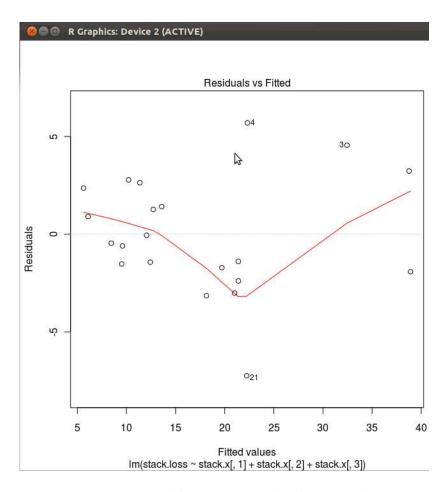


Figure 3: Problem 2-a: Residual vs Fitted

About uncorrelated error, I use Perason's product-moment test to do the analysis. And here is the result from R:

```
Pearson's product-moment correlation
```

```
data: stack.loss and stack.x[, 1]
t = 10.2079, df = 19, p-value = 3.774e-09
alternative hypothesis: true correlation is not equal to 0
95 percent confidence interval:
    0.8092570 0.9673185
sample estimates:
        cor
0.9196635
```

Pearson's product-moment correlation

```
data: stack.loss and stack.x[, 2]
t = 7.8977, df = 19, p-value = 2.028e-07
alternative hypothesis: true correlation is not equal to 0
95 percent confidence interval:
    0.7134686    0.9486536
sample estimates:
        cor
0.8755044
```

Pearson's product-moment correlation

```
data: stack.loss and stack.x[, 3]
t = 1.9014, df = 19, p-value = 0.07252
alternative hypothesis: true correlation is not equal to 0
95 percent confidence interval:
   -0.03850282   0.70912123
sample estimates:
        cor
0.3998296
```

As we can see, all three explanatory variables – "Air Flow", "Water Temp" and "Acid Conc" are correlated with "stack.loss".

About outliers, here is the studentized deleted residuals from R:

1 2 3 4 5 6

As we can see, only $|T_{21}| = 3.33 > t_{.975.17} = 2.110$. So only example 21 is an outlier.

About influential points, here is the DFFITS from R:

As we can see, only $|DFFITS_{21}| > 1$. So only example 21 is influential.

Appendices

The code is listed below:

```
cesium.reg.summary <- summary(cesium.reg)</pre>
print(cesium.reg.summary)
#(c)
cesium.trim <- cesium[c(1:16),]</pre>
cesium.trim.reg <- lm(cesium.trim$Mushrooms~cesium.trim$Soils)</pre>
cesium.trim.reg.summary <- summary(cesium.trim.reg)</pre>
print(cesium.trim.reg.summary)
# additional
lmi.p1 <- lm.influence(cesium.reg)</pre>
res.p1 <- resid(cesium.reg)</pre>
nt.p1 <- shapiro.test(res.p1)</pre>
print(nt.p1)
cor.p1 <- cor.test(cesium$Mushrooms, cesium$Soils)</pre>
print(cor.p1)
# problem2
attach(stackloss)
#a)
fit1 <- lm(stack.loss~stack.x[,1] + stack.x[,2] + stack.x[,3])</pre>
lmi <- lm.influence(fit1)</pre>
lms <- summary(fit1)</pre>
print(lms)
#b)
e <- resid(fit1)
nt.p2 <- shapiro.test(e)</pre>
print(nt.p2)
cor1.p2 <- cor.test(stack.loss, stack.x[,1])</pre>
print(cor1.p2)
cor2.p2 <- cor.test(stack.loss, stack.x[,2])</pre>
print(cor2.p2)
cor3.p2 <- cor.test(stack.loss, stack.x[,3])</pre>
print(cor3.p2)
```

```
s <- lms$sigma
si <- lmi$sigma
xxi <- diag(lms$cov.unscaled)
h <- lmi$hat

bi <- coef(fit1) - t(coef(lmi))
stand.resid <- e/(s*(1-h)^0.5)
student.resid <- e/(si*(1-h)^0.5)
DFFITS <- h^0.5*e/(si*(1-h))</pre>
```