Stat 4201 Homework 8

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Question 1

```
a) Here is the logistic regression of carrier on CK and H:
```

```
> summary(fit.p1)
```

Call:

glm(formula = Y ~ CK + H, family = binomial)

Deviance Residuals:

```
Min 1Q Median 3Q Max -2.44845 -0.38658 -0.19898 0.00193 2.44680
```

Coefficients:

Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 149.840 on 119 degrees of freedom Residual deviance: 62.224 on 117 degrees of freedom

AIC: 68.224

Number of Fisher Scoring iterations: 8

Here is the confidence intervals:

I notice that using the H and CK will lead to the following warning:

```
fit.p1 <- glm(Y~CK+H, family=binomial)
Warning message:
glm.fit: fitted probabilities numerically 0 or 1 occurred</pre>
```

I find out that this indicates that some of the CK values are too large to fit the model and CK needs a transformation. Since problem 2 explores about which predictor is appropriate, I just leave the appropriate logistic regression model to problem 2.

b) β_{CK} : If CK increases by 1 unit, the odds Y = 1 will change by a multiplicative factor of $exp(\beta_{CK})$, other variables being the same.

 β_H : If H increases by 1 unit, the odds Y = 1 will change by a multiplicative factor of $exp(\beta_H)$, other variables being the same.

Question 2

- a) The scatterplot of H versus log(CK) is shown in Fig-1. As we can see, most of the squares in the plot are in the district where log(CK) is low. As to H, there isn't such concentration on the plot. So log(CK) might be useful predictors of whether a woman is a carrier.
- b)
 The logistic regression of carrier on CK and CK-squared:

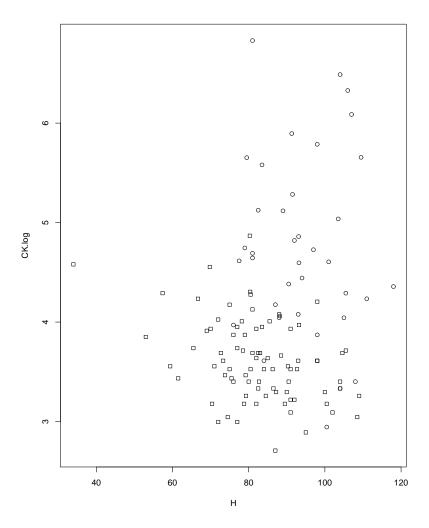


Figure 1: Scatterplot: H versus log(CK)

```
> summary(fit.b1)
```

Call:
glm(formula = Y ~ CK + CK.square, family = binomial)

Deviance Residuals:

Min 1Q Median 3Q Max -2.27518 -0.51824 -0.37943 0.03892 2.50614

Coefficients:

Estimate Std. Error z value Pr(>|z|)

(Intercept) -4.181e+00 7.272e-01 -5.749 8.96e-09 ***

CK 5.805e-02 1.301e-02 4.460 8.18e-06 ***

CK.square -5.060e-05 3.286e-05 -1.540 0.124

Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 149.840 on 119 degrees of freedom Residual deviance: 85.435 on 117 degrees of freedom

AIC: 91.435

Number of Fisher Scoring iterations: 9

As we can see, CK-squared term does not significantly differ from 0.

The logistic regression of carrier on log(CK) and $[log(CK)]^2$:

Call:

glm(formula = Y ~ CK.log + CK.log.square, family = binomial)

Deviance Residuals:

Min 1Q Median 3Q Max -2.28852 -0.50190 -0.38037 0.03075 2.39251

Coefficients:

Estimate Std. Error z value Pr(>|z|)

(Intercept) 9.830 16.309 0.603 0.547 CK.log -8.568 8.366 -1.024 0.306 CK.log.square 1.453 1.064 1.365 0.172

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 149.84 on 119 degrees of freedom Residual deviance: 84.98 on 117 degrees of freedom

AIC: 90.98

Number of Fisher Scoring iterations: 7

As we can see, the square term does not significantly differ from 0.

I think the log(CK) is more appropriate because the second regression model has a smaller AIC 90.98 compared with the first model (91.435). This means that use log(CK) fits the model better.

c) Here is the logistic regression of carrier on log(CK) and H:

> summary(fit.c)

Call:

glm(formula = Y ~ CK.log + H, family = binomial)

Deviance Residuals:

Median Min 10 3Q Max -1.89707 -0.38782 -0.16697 0.09903 2.60372

Coefficients:

Estimate Std. Error z value Pr(>|z|)5.80030 -4.985 6.20e-07 *** (Intercept) -28.91300 4.849 1.24e-06 *** 4.02041 0.82909 CK.log Η 0.13652 0.03654 3.736 0.000187 ***

Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 149.840 on 119 degrees of freedom Residual deviance: 61.992 degrees of freedom on 117

AIC: 67.992

Number of Fisher Scoring iterations: 7

d) Here is the drop-in-deviance test for the hypothesis that neither log(CK) nor H are useful predictors of whether a woman is a carrier:

Single term deletions

As we can see, both log(CK) and H are useful predictors of whether a woman is a carrier.

e) From part c, we get the logit regression:

$$logit(\pi) = -28.91 + 4.02 \cdot \log CK + 0.14 \cdot H \tag{1}$$

Applying this equation, we get

$$odds_{suspected} = \exp(-28.91 + 4.02 \cdot \log 300 + 0.14 \cdot 100) = 37.07$$

$$odds_{typical} = \exp(-28.91 + 4.02 \cdot \log 80 + 0.14 \cdot 85) = 1.85$$

$$\frac{odds_{suspected}}{odds_{typical}} = 20.02$$

The value that odds that the suspected carrier is a carrier relative to the odds that a woman with typical values is a carrier is 20.02.

Appendices

The R code is listed below:

Problem 1

```
data.p1 <- read.csv("ex2012.csv", header=TRUE)</pre>
attach(data.p1)
fit.p1 <- glm(Y~CK+H, family=binomial)</pre>
confint(fit.p1)
# Problem 2
CK.log <- log(CK)</pre>
postscript(file="~/Documents/LaTeX/stat4201-hmwk8/scatter.eps",
            onefile=FALSE, horizontal=FALSE)
plot(H, CK.log, pch=c(21, 22)[GROUP])
dev.off()
Y <- 1*(GROUP=="Case")
CK.square <- CK^2
fit.b1 <- glm(Y~CK+CK.square, family=binomial)</pre>
CK.log.square <- CK.log^2</pre>
fit.b2 <- glm(Y~CK.log+CK.log.square, family=binomial)</pre>
fit.c <- glm(Y~CK.log+H, family=binomial)</pre>
drop1(fit.c, test="Chisq")
odd1 <- exp(-28.9 + 4.02 * log(100) + 0.14 * 100)
odd2 <- exp(-28.9 + 4.02 * log(80) + 0.14 * 85)
ratio <- odd1 / odd2
```