

# Unit 5: Sampling Distributions

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# Unit 5 Outline: Sampling Distributions

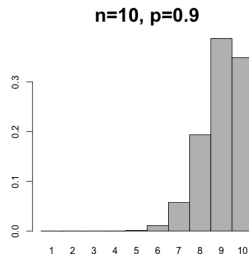
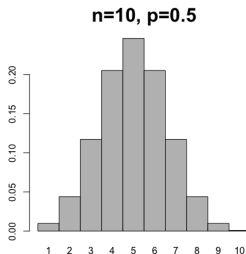
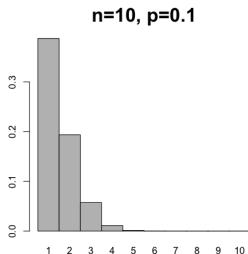
- 1 Distributions that converge to normal
- 2 Sampling distributions
- 3 Point estimates and bias
- 4 Sampling distribution for  $\hat{p}$
- 5 Sampling distribution for a difference of proportions ( $\hat{p}_1 - \hat{p}_2$ )
- 6 Sampling distribution for  $\mu$
- 7 Sampling distribution for a difference of means ( $\mu_1 - \mu_2$ )

# Normal Approximation to the Binomial Distribution

- A discrete binomial variable can be approximated by a continuous normal variable.
- This is useful when the binomial formula becomes computationally intensive for large  $n$ .
- This concept will be very important later in statistical inference.

# Binomial Shape at Small $n = 10$

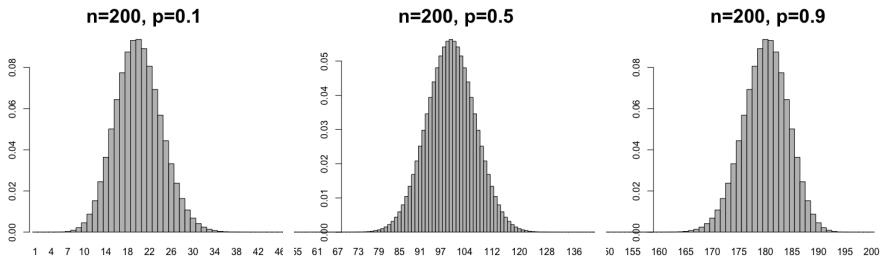
Let's consider several values of  $p$  with  $n = 10$ .



- $p = 0.5$ : symmetric, bell-shaped
- $p = 0.1$ : right-skewed
- $p = 0.9$ : left-skewed

# Effect of Increasing $n$ (to 200)

Let's keep the same values of  $p$ , but increase  $n$  to 200:



As  $n$  increases, the binomial distribution looks more normal - even for skewed  $p$  values.

# When is the Normal Approximation Valid?

## Rule of Thumb

The normal approximation is appropriate if:

$$np \geq 10 \quad \text{and} \quad n(1 - p) \geq 10$$

- Interpreted as having at least 10 expected successes and 10 expected failures.
- Use:

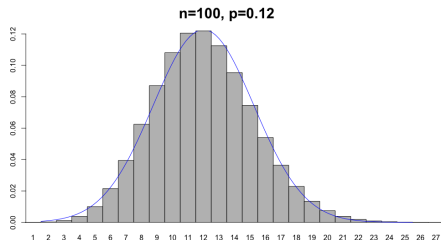
$$\mu = np, \quad \sigma = \sqrt{np(1 - p)}$$

## Example: No Continuity Correction

A basketball player has a 12% chance of making a free throw. Estimate the probability they make 18 or more in 100 shots.

Check approximation validity:

$$np = 12, \quad n(1 - p) = 88 \Rightarrow \text{valid}$$



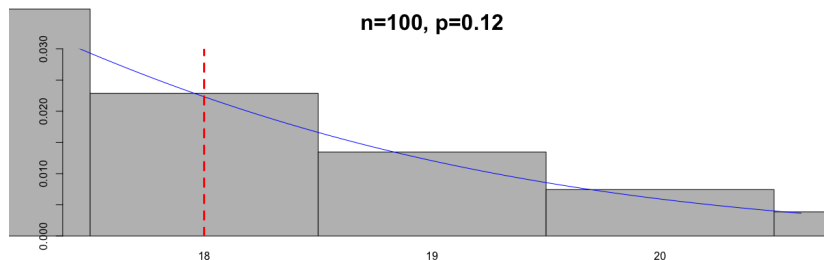
$$\mu = 12, \quad \sigma = \sqrt{100(0.12)(0.88)} \approx 3.25$$

$$z = \frac{18 - 12}{3.25} \approx 1.85 \Rightarrow P(Z > 1.85) \approx 0.0324$$

About 3.2% chance of making at least 18 shots.

# Why Use a Continuity Correction?

We're approximating a discrete variable with a continuous one. So we may also apply a continuity correction.



- Without correction, we ignore part of the probability mass.
- For better accuracy, use  $x = 17.5$  instead of 18.



# Continuity Correction Summary

| Binomial      | Normal Approximation             |
|---------------|----------------------------------|
| $P(X = x)$    | $P(x - 0.5 \leq X \leq x + 0.5)$ |
| $P(X \leq x)$ | $P(X \leq x + 0.5)$              |
| $P(X < x)$    | $P(X \leq x - 0.5)$              |
| $P(X > x)$    | $P(X \geq x + 0.5)$              |
| $P(X \geq x)$ | $P(X \geq x - 0.5)$              |

*Tip: Don't memorize - just sketch the histogram and think logically!*

# Distributions That Converge to Normal

- Many important probability distributions become approximately normal under the right conditions (typically as sample size or degrees of freedom increase).

## Examples:

- **Binomial:** Normal approximation valid as  $n \rightarrow \infty$ .
- **Hypergeometric:** As  $n \rightarrow \infty$  for small  $n$  relative to  $N$ .
- **Poisson:** For large  $\lambda$ , the distribution becomes approximately normal.
- **Chi-Square ( $\chi^2$ ):** Becomes more symmetric and bell-shaped as degrees of freedom increase.
- **t-distribution:** Approaches standard normal as degrees of freedom increase.

**Takeaway:** The normal distribution plays a central role in inference because many statistics follow a normal distribution in large samples.

# Point Estimation

- Statistics are used to estimate **population parameters**.
- A **point estimate** is a single value used to estimate a target parameter.

$\bar{x}$  is a point estimate for  $\mu$ ,  $\hat{p}$  is a point estimate for  $p$

## Bias of a Point Estimator

We say  $\hat{\theta}$  is an **unbiased estimator** of the parameter  $\theta$  if:

$$E(\hat{\theta}) = \theta$$

The bias of an estimator is defined as:

$$B(\hat{\theta}) = E(\hat{\theta}) - \theta$$

- We also care about the **variance** and **distribution** of estimators.

# What is a Sampling Distribution?

- A **sampling distribution** is the distribution of a statistic over all possible samples.
- Imagine repeating a random sample process infinitely many times and recording a statistic each time.
- The distribution of all these sample statistics forms the sampling distribution.

## Why It Matters

Sampling distributions are **essential** for statistical inference. They allow us to:

- Understand variability in estimates
- Construct confidence intervals (Unit 6-9)
- Perform hypothesis testing (Unit 6-9)

# Sampling Distribution for $\hat{p}$

- Take a random sample of size  $n$  from a population of size  $N$ .
- Let  $X$  be the number of sample elements with a certain characteristic.

$$\hat{p} = \frac{X}{n}$$

- The population has  $r$  total successes, so:

$$p = \frac{r}{N}$$

- $X \sim \text{Hypergeometric}(r, n, N)$

# Expected Value of $\hat{p}$ (Hypergeometric)

Prove  $\hat{p}$  is an unbiased estimator for  $p$ :

$$\begin{aligned} E(\hat{p}) &= E\left(\frac{X}{n}\right) \\ &= \frac{1}{n}E(X) \\ &= \frac{1}{n} \cdot n \left(\frac{r}{N}\right) \\ &= \frac{r}{N} = p \end{aligned}$$

- $\hat{p}$  is an **unbiased estimator** of  $p$ .
- The sampling distribution of  $\hat{p}$  is centered at the true population proportion.

# Variance of $\hat{p}$ (Hypergeometric)

Determine the variance of the statistics  $\hat{p}$ :

$$\begin{aligned}\text{Var}(\hat{p}) &= \text{Var}\left(\frac{X}{n}\right) \\ &= \frac{1}{n^2} \cdot n \cdot \frac{r}{N} \left(1 - \frac{r}{N}\right) \left(\frac{N-n}{N-1}\right) \\ &= \frac{p(1-p)}{n} \cdot \left(\frac{N-n}{N-1}\right)\end{aligned}$$

$$SD(\hat{p}) = \sqrt{\frac{p(1-p)}{n} \cdot \left(\frac{N-n}{N-1}\right)}$$

- This is most appropriate for small, finite populations.
- We **DON'T** use this for AP statistics (why not)?

# When Can We Use the Binomial Approximation?

- We approximate  $X \sim \text{Binomial}(n, p)$ , which is valid when:

## Independence Condition

Sample size  $n$  is less than 10% of the population:  $n < 0.1N$

Assuming  $X \sim \text{Binomial}(n, p)$ , we get:

$$E(\hat{p}) = p, \quad \text{Var}(\hat{p}) = \frac{p(1-p)}{n}, \quad SD(\hat{p}) = \sqrt{\frac{p(1-p)}{n}}$$

- 1 Why is  $X \sim \text{Binomial}(n, p)$  when  $n < 10\%N$ ?
- 2 Prove the expected value for  $\hat{p}$  and variance for  $\hat{p}$  using the binomial approximation.



# Normal Approximation to the Sampling Distribution of $\hat{p}$

- Even with the binomial model, exact computations can be complex.
- So we use a normal approximation for  $\hat{p}$ , if the following condition is met:

## Normality Condition

$$np > 10 \quad \text{and} \quad n(1 - p) > 10$$

(At least 10 expected successes and failures)

$$\frac{\hat{p} - p}{\sqrt{\frac{p(1 - p)}{n}}} \sim N(0, 1)$$

# Conditions for Using the Normal Sampling Distribution of $\hat{p}$

To use the normal model for  $\hat{p}$ , the following must be true:

- **Random Sampling:** Sample is collected randomly.
- **Independence:** Population is at least 10 times larger than the sample ( $n < 0.1N$ ).
- **Normality:**  $np > 10$  and  $n(1 - p) > 10$

These are assumptions - they are not always verifiable but are necessary to use this model.

# Example: Sampling Distribution for $\hat{p}$

It is known that across North America, 65% of university students take longer than four years to complete their undergraduate degree. You survey 100 University of Calgary graduates.

**(a) Distribution for  $X$ :**

Since  $n = 100 < 0.1N$ , we approximate using a binomial model:

$$X \sim \text{Binomial}(n = 100, p = 0.65)$$

**(b) Sampling distribution for  $\hat{p}$ :**

Conditions:

- Independence:  $n = 100 < 0.1N$  ✓
- Normality:  $np = 65 > 10$ ,  $n(1 - p) = 35 > 10$  ✓

$$\hat{p} \sim \text{Normal}\left(0.65, \frac{0.65(0.35)}{100}\right)$$

**(c) Probability that  $\hat{p} > 0.70$ :**

$$z = \frac{0.70 - 0.65}{\sqrt{\frac{0.65 \cdot 0.35}{100}}} = 1.048$$

$$P(\hat{p} > 0.70) = P(Z > 1.048) \approx 0.147$$

There is about a 15% chance that more than 70% of your sample took over four years to graduate.

# The Sampling Distribution of $\bar{x}$

Suppose  $X_1, X_2, \dots, X_n$  are independent and identically distributed random variables.

$$\bar{x} = \frac{X_1 + X_2 + \dots + X_n}{n}$$

Assume each  $X_i \sim \text{Normal}(\mu, \sigma^2)$ . Then:

$$E(\bar{x}) = \mu \quad (\text{Unbiased})$$

$$\text{Var}(\bar{x}) = \frac{\sigma^2}{n}, \quad \text{SD}(\bar{x}) = \frac{\sigma}{\sqrt{n}}$$

So:

$$\bar{x} \sim \text{Normal}\left(\mu, \frac{\sigma^2}{n}\right)$$

**Independence condition:**  $n < 10\%$  of the population

# Standardizing the Sampling Distribution

We often standardize  $\bar{x}$  using:

$$Z = \frac{\bar{x} - \mu}{\frac{\sigma}{\sqrt{n}}}$$

Where:

- $Z \sim \text{Normal}(0, 1)$
- Requires known  $\sigma$
- Assumes random sampling and independence

If  $\sigma$  is unknown, we will require a new distribution - this will be **covered later**.

## Example: Pale-Throated Sloths (Setup)

The weights of pale-throated sloths are normally distributed:

$$\mu = 4.5 \text{ kg}, \quad \sigma = 1.1 \text{ kg}$$

You randomly sample 20 sloths.

- a) **Describe the sampling distribution of  $\bar{x}$ :**

Since  $n = 20 < 0.1N$ , and the parent distribution is normal:

$$\bar{x} \sim \text{Normal} \left( 4.5, \frac{(1.1)^2}{20} \right)$$

- b) What is the probability the sample mean is between 2.3 kg and 4.3 kg?

$$z_{\text{low}} = \frac{2.3 - 4.5}{\frac{1.1}{\sqrt{20}}} = -8.94, \quad z_{\text{high}} = \frac{4.3 - 4.5}{\frac{1.1}{\sqrt{20}}} = -0.81$$

$$P(2.3 \leq \bar{x} \leq 4.3) = P(-8.9 \leq Z \leq -0.8) \approx 0.2119$$

There is approximately a 21.2% chance the sample mean falls in this range.

# The Central Limit Theorem (CLT)

**Question:** What happens when the parent distribution is not normal?

## *The Central Limit Theorem*

Let  $X_1, X_2, \dots, X_n$  be i.i.d. random variables with

$$E(X_i) = \mu, \quad \text{Var}(X_i) = \sigma^2$$

Then:

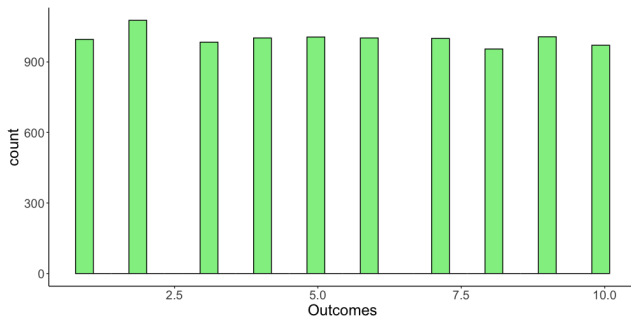
$$\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n} \xrightarrow{n \rightarrow \infty} \text{Normal} \left( \mu, \frac{\sigma^2}{n} \right)$$

That is, the sampling distribution of  $\bar{X}$  becomes normal as  $n$  increases - regardless of the parent distribution.

$$\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim Z \quad \text{for } n \geq 30$$

## Example: 10-Sided Die

Let  $X$  represent the outcome of a 10-sided die roll. The parent distribution is uniform.

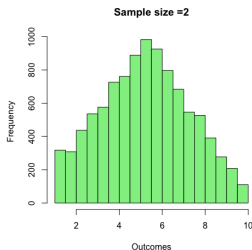


Even though this parent distribution is not normal, the CLT applies as  $n$  increases.

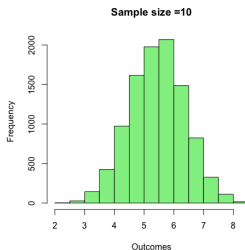


# Sampling Distributions of $\bar{X}$

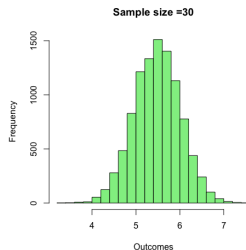
Below are sampling distributions from 10,000 samples for sample sizes of 2, 10, and 30:



Sample size = 2



Sample size = 10



Sample size = 30

As the sample size increases, the sampling distribution of  $\bar{X}$  becomes more normal **regardless of the parent population.**

## Example: Carnival Game - Profit Distribution

A carnival game has the following profit distribution:

|             |      |      |      |      |
|-------------|------|------|------|------|
| Profit (\$) | -1   | 1    | 5    | 20   |
| Probability | 0.95 | 0.03 | 0.02 | 0.01 |

Let  $X$  be your profit from a single play.

**a) Determine expected value for  $X$ :**

$$E(X) = -1(0.95) + 1(0.03) + 5(0.02) + 20(0.01) = -0.62$$

**b) Determine variance for  $X$ :**

$$E(X^2) = 1(0.95) + 1(0.03) + 25(0.02) + 400(0.01) = 5.48$$

$$\text{Var}(X) = E(X^2) - (E(X))^2 = 5.48 - (-0.62)^2 = 5.4556$$

# Carnival Game: CLT Approximation

- Suppose you play the game 30 times ( $n = 30$ ).
- CLT applies: large sample size.
- Then:

$$\mu_{\bar{x}} = -0.62, \quad \sigma_{\bar{x}} = \sqrt{\frac{5.4556}{30}} = 0.4264$$

$$\bar{x} \sim \text{Normal}(-0.62, 0.4264)$$

- 1 What is the probability that your profit is positive after playing the 30 games?

$$\begin{aligned} P(\bar{x} > 0) &= P\left(\frac{\bar{x} - \mu}{\frac{\sigma}{\sqrt{n}}} > \frac{0 - (-0.62)}{\frac{0.4264426}{\sqrt{30}}}\right) \\ &= P(Z > 7.963275) \\ &= 1 - P(Z \leq 7.963275) \\ &\approx 0 \end{aligned}$$

# Assumptions for Using a Normal Model

To use the normal model for  $\bar{x}$ , we must assume:

- **Normality:** Either the parent population is normal or  $n \geq 30$
- **Independence:** Sample size  $n < 10\%$  of population size  $N$
- **Random Sampling:** Sample is collected using a random method

$$\bar{x} \sim \text{Normal} \left( \mu, \frac{\sigma^2}{n} \right)$$

Similar to assumptions for the sampling distribution of  $\hat{p}$

# A Sampling Distribution Involving $s^2$

Suppose  $X_1, X_2, \dots, X_n$  is a random sample from a normal population with mean  $\mu$  and variance  $\sigma^2$ .

Then the following distribution holds:

$$\frac{(n-1)s^2}{\sigma^2} \sim \chi_{n-1}^2$$

## Example: Pale-Throated Sloths

The weights of sloths are normally distributed with  $\mu = 4.5$  kg,  $\sigma = 1.1$  kg. A random sample of  $n = 20$  sloths is taken. What is the probability that the sample standard deviation is at least 0.9?

$$\begin{aligned} P(s^2 > 0.9^2) &= P\left(\chi_{19}^2 > \frac{(0.9)^2 \cdot 19}{(1.1)^2}\right) \\ &= P(\chi_{19}^2 > 12.72) = 0.8526 \end{aligned}$$

**Conclusion:** There's an 85% chance of observing a sample standard deviation of 0.9 or greater.

# Standard Deviation vs. Standard Error

**Problem:** Many sampling distributions involve unknown population parameters.

- For the sampling distribution of the sample mean:

$$\sigma_{\bar{x}} = \frac{\sigma}{\sqrt{n}}$$

- But the population standard deviation  $\sigma$  is usually unknown.
- We estimate it using the sample standard deviation  $s$ .

## Standard Error

The **standard error** is the estimated standard deviation of a statistic:

$$SE_{\bar{x}} = \frac{s}{\sqrt{n}}$$

**What happens to the distribution?**

$$\frac{\bar{x} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim Z \quad (\text{when } \sigma \text{ is known})$$

$$\frac{\bar{x} - \mu}{\frac{s}{\sqrt{n}}} \sim ??? \quad (\text{new distribution})$$

# Using the $t$ -Distribution

Consider a standard normal random variable  $Z$ , and a chi-square random variable with  $k$  degrees of freedom. The  $t$ -distribution is defined as:

$$t = \frac{Z}{\sqrt{\frac{\chi_k^2}{k}}}$$

Recall the following known distributions (when assumptions are met):

$$\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim Z, \quad \frac{(n-1)s^2}{\sigma^2} \sim \chi_{n-1}^2$$

Using these, we construct the  $t$ -statistic:

$$\frac{\left( \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \right)}{\sqrt{\frac{\left( \frac{(n-1)s^2}{\sigma^2} \right)}{n-1}}} = \frac{\bar{X} - \mu}{\frac{s}{\sqrt{n}}} \sim t_{n-1}$$

# Assumptions for Using the $t$ -Distribution

To use the  $t$ -distribution, the following assumptions must hold:

- **Simple Random Sampling**
- **Independence:**  $n < 0.1N$
- **Normality:**

Ideally, the population is normal with mean  $\mu$ , variance  $\sigma^2$ . Then:

$$\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim Z, \quad \frac{(n-1)s^2}{\sigma^2} \sim \chi_{n-1}^2$$

If the parent population is unknown:

- If  $n \geq 30$ , the CLT allows:

$$\frac{\bar{X} - \mu}{s/\sqrt{n}} \sim t_{n-1}$$

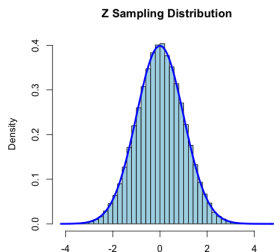
- If  $n < 30$ , we require the population to be approximately normal (unimodal, symmetric, no outliers).

**Caution:** Small, skewed, or heavy-tailed samples may make the  $t$ -distribution inappropriate.

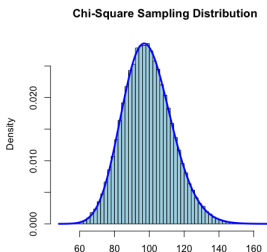


# Case I: Normal Parent Distribution, Large Sample ( $n = 100$ )

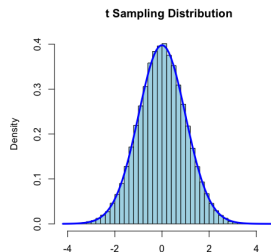
Assume  $X_1, X_2, \dots, X_{100} \sim \text{Normal}(10, 2)$  Histograms below show the sampling distributions (100,000 simulations), with theoretical curves superimposed.



$$\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim Z$$



$$\frac{(n-1)s^2}{\sigma^2} \sim \chi_{99}^2$$

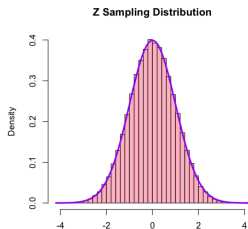


$$\frac{\bar{X} - \mu}{s/\sqrt{n}} \sim t_{99}$$

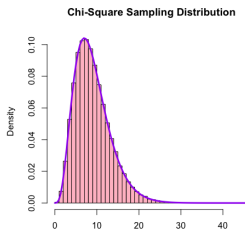
**Conclusion:** With a large sample size and normal parent population, the theoretical distributions are a very good fit.

## Case II: Normal Parent Distribution, Small Sample ( $n = 10$ )

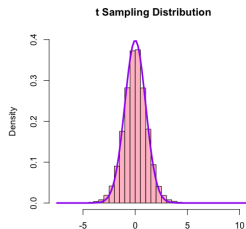
Assume  $X_1, X_2, \dots, X_{10} \sim \text{Normal}(10, 2)$ . Again, histograms show empirical sampling distributions with theoretical curves.



$$\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim Z$$



$$\frac{(n-1)s^2}{\sigma^2} \sim \chi_9^2$$



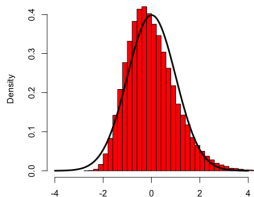
$$\frac{\bar{X} - \mu}{s/\sqrt{n}} \sim t_9$$

**Conclusion:** Even with a small sample, normality in the parent distribution ensures that the  $t$ -distribution is appropriate.

# Case III: Skewed Parent Distribution, Small Sample ( $n = 10$ )

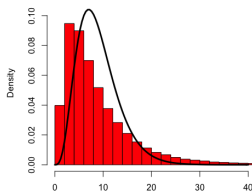
Assume  $X_1, X_2, \dots, X_{10} \sim \text{Exponential}(3)$ , a **highly right skew distribution**.

Z Sampling Distribution



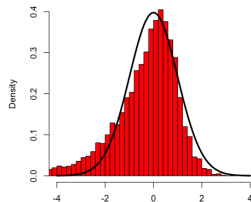
$$\frac{\bar{x} - \mu}{\sigma/\sqrt{n}} \sim ?$$

Chi-Square Sampling Distribution



$$\frac{(n-1)s^2}{\sigma^2} \sim ?$$

t Sampling Distribution



$$\frac{\bar{x} - \mu}{s/\sqrt{n}} \sim ?$$

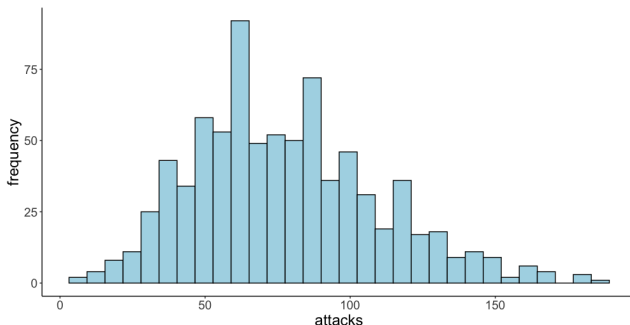
**Conclusion:** With a skewed parent distribution and small  $n$ , the theoretical distributions do not fit. Use caution when applying the  $t$ -distribution in this scenario.

# Example: Pokémon Attack Scores

A random sample of  $n = 801$  Pokémon has:

$$\bar{x} = 78, \quad s = 32$$

Suppose the true population mean is  $\mu = 70$ . The sample distribution is shown below:



- a) What is the probability that a future sample has a mean attack score less than 70?

## Example: Pokémon Attack Scores Solution

**Solution:** Large  $n$  and approximately normal data  $\rightarrow$  use the  $t$ -distribution.

$$\begin{aligned}P(\bar{x} < 70) &= P\left(\frac{\bar{x} - \mu}{s/\sqrt{n}} < \frac{70 - 78}{32/\sqrt{801}}\right) \\&= P(t_{800} < -2.6533) = 0.0041\end{aligned}$$

**Conclusion:** There's about a 0.41% chance that a random sample of 801 Pokémon would have a mean attack below 70.

# Summary of Sampling Distributions

Let's summarize the sampling distributions we've developed so far:

| Distribution  | Assumptions   |
|---|---|
| $\hat{p} \sim \text{Normal} \left( p, \frac{p(1-p)}{n} \right)$     | Random sampling, independence, and normality condition: $np > 10$ , $n(1-p) > 10$                   |
| $\frac{\hat{p} - p}{\sqrt{\frac{p(1-p)}{n}}} \sim Z$                | Same as above (standardized version)  |
| $\bar{x} \sim \text{Normal} \left( \mu, \frac{\sigma^2}{n} \right)$ | Random sampling, independence ( $n < 0.1N$ ), and normal population or large $n \geq 30$            |
| $\frac{\bar{x} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim Z$              | When $\sigma$ is known, with same assumptions as above  |
| $\frac{\bar{x} - \mu}{\frac{s}{\sqrt{n}}} \sim t_{n-1}$             | When $\sigma$ is unknown. Requires normality or large sample, plus random sampling and independence |

# Sampling Distribution of a Difference in Proportions

We often compare two sample proportions:

- $\hat{p}_1$ : the sample proportion from a group of size  $n_1$
- $\hat{p}_2$ : the sample proportion from a second group of size  $n_2$

**Our goal:** Understand the behavior of the statistic  $\hat{p}_1 - \hat{p}_2$ . **Assumptions:**

- **Random Sampling:** Each sample is drawn using a random method.
- **Independence:** Observations are independent within and between samples. Assume this if:

$$n_1 < 0.1N_1 \quad \text{and} \quad n_2 < 0.1N_2$$

- **Normality:** Each sample must have at least 10 successes and 10 failures:

$$n_1 p_1 > 10, \quad n_1(1 - p_1) > 10, \quad n_2 p_2 > 10, \quad n_2(1 - p_2) > 10$$

- What is  $E(\hat{p}_1 - \hat{p}_2)$ ?
- What is  $\text{Var}(\hat{p}_1 - \hat{p}_2)$ ?
- What distribution does  $\hat{p}_1 - \hat{p}_2$  follow?

## Example: Difference in Proportions - Two Towns

In one town, 51% of voters are conservative; in another, 44% are conservative. A random sample of 100 voters is taken from each town.

a) **Is a normal model appropriate for  $\hat{p}_1 - \hat{p}_2$ ?**

- **Simple Random Sample:** Assumed for both towns.
- **Independence:**  $n_1 = n_2 = 100 < 0.1N$  so we assume independence.
- **Normality:**

$$n_1 p_1 = 51, \quad n_1(1 - p_1) = 49$$

$$n_2 p_2 = 44, \quad n_2(1 - p_2) = 56$$

b) **What is the probability that  $\hat{p}_1 < \hat{p}_2$ ?**

$$\begin{aligned} P(\hat{p}_1 - \hat{p}_2 < 0) &= P\left(Z < \frac{0 - (0.51 - 0.44)}{\sqrt{\frac{0.51(0.49)}{100} + \frac{0.44(0.56)}{100}}}\right) \\ &= P(Z < -0.994) = 0.1602 \end{aligned}$$

**Conclusion:** There is about a 16% chance the first sample yields a lower proportion than the second.



# Sampling Distribution for a Difference in Sample Means

Suppose we take two independent random samples:

- $\bar{x}_1$  is the mean of a sample of size  $n_1$ , from a population with mean  $\mu_1$  and standard deviation  $\sigma_1$
- $\bar{x}_2$  is the mean of a sample of size  $n_2$ , from a population with mean  $\mu_2$  and standard deviation  $\sigma_2$

We are interested in the statistic  $\bar{x}_1 - \bar{x}_2$

## Assumptions:

- **Random Sampling:** Each sample is randomly drawn
- **Independence:** Each sample satisfies  $n_1 < 0.1N_1$ ,  $n_2 < 0.1N_2$
- **Normality:** Either:
  - Both populations are approximately normal
  - OR sample sizes are large:  $n_1 \geq 30$  and  $n_2 \geq 30$
- What is  $E(\bar{x}_1 - \bar{x}_2)$ ?
- What is  $\text{Var}(\bar{x}_1 - \bar{x}_2)$ ?
- What is the sampling distribution of  $\bar{x}_1 - \bar{x}_2$ ?

# Difference in Sample Means (Unknown Variances)

When population standard deviations  $\sigma_1$  and  $\sigma_2$  are unknown, we use the sample standard deviations  $s_1$  and  $s_2$  to estimate them.

**Sampling Distribution:**

$$\frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} \sim t_{df}$$

**Degrees of Freedom (df):**

$$\min(n_1 - 1, n_2 - 1) \leq df \leq n_1 + n_2 - 2$$

Which degree of freedom would be the most conservative?

**Conditions:**

- **Random Sampling:** Both samples are independently and randomly drawn.
- **Independence:**  $n_1 < 10\%$  of  $N_1$ ,  $n_2 < 10\%$  of  $N_2$
- **Normality:** Each sample is from a normal population or both  $n_1, n_2 \geq 30$

# Welch-Satterthwaite Approximation

When population variances are unknown and unequal, we estimate the degrees of freedom using the Welch-Satterthwaite formula:

$$df = \frac{\left( \frac{s_1^2}{n_1} + \frac{s_2^2}{n_2} \right)^2}{\frac{\left( \frac{s_1^2}{n_1} \right)^2}{n_1 - 1} + \frac{\left( \frac{s_2^2}{n_2} \right)^2}{n_2 - 1}}$$

**Use in:**

$$\frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} \sim t_{df}$$

**Note:** This formula often gives a non-integer  $df$ ; statistical software typically handles this automatically.

# Sampling Distribution with Pooled Variance (Enrichment)

Suppose we take two independent random samples from two populations, and we assume that the population variances are equal:

$$\sigma_1^2 = \sigma_2^2 = \sigma^2$$

We estimate the common variance using the **pooled sample variance**:

$$s_p^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2}$$

If all assumptions are satisfied, then the sampling distribution of the difference in sample means is:

$$\frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim t_{n_1 + n_2 - 2}$$

## Assumptions:

- Random sampling
- Independence:  $n_1 < 10\%N_1$ ,  $n_2 < 10\%N_2$
- Normal populations or large sample sizes
- **Equal population variances**