# Adaptation to unknown number of clusters

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## 1 Estimation for the number of clusters

When the generating model is from  $\alpha$ -smooth probability tensor, we do not have the true number of group k. In this case, we can easily pick  $k = \lfloor n^{\frac{m}{m+2\alpha}} \rfloor$  which guarantees the convergence rate  $\mathcal{O}(n^{\frac{-2m\alpha}{m+2\alpha}} + \log n/n)$ . If we take nonparametric histogram perspective and consider k as bandwidth, we do not need to estimate k but set optimal  $k = \lfloor n^{\frac{m}{m+2\alpha}} \rfloor$ .

However, when we believe that the probability tensor has block structure, there is true k so we need to set k for the estimation. Under the stochastic block model assumption, we take a variation of a 2-folded cross validation approach. To be specific, we split the observed entries into two half with probability 1/2 and use one for the training data set and the other for the test dataset. Let  $\Omega_1$  be the training set and  $\Omega_2$  be the test set from Bernoulli(1/2) sampling. Define the training tensor  $\mathcal{A}^{(1)}$  and  $\mathcal{A}^{(2)}$  such that,

$$\mathcal{A}_{\omega}^{(1)} = \begin{cases} \mathcal{A}_{\omega} & \text{if } \omega \in \Omega_{1}, \\ 0 & \text{if } \omega \in \Omega_{2}. \end{cases} \text{ and } \mathcal{A}_{\omega}^{(2)} = \begin{cases} 0 & \text{if } \omega \in \Omega_{1}, \\ \mathcal{A}_{\omega} & \text{if } \omega \in \Omega_{2}. \end{cases}$$

For many different  $k \in [n]$ , we calculate

$$\hat{\Theta}_k^{(i)} = \underset{\Theta \in \text{cut}(\mathcal{P}_k)}{\arg \min} \|\mathcal{A}^{(i)} - \Theta\|_F^2,$$

for i = 1, 2. We select the parameter which minimizes the MSE error on the test dataset

$$k_i = \underset{k \in [n]}{\operatorname{arg \, min}} \sum_{\omega \in \Omega_i^c} |\mathcal{A}_{\omega} - (\hat{\Theta}_k^{(i)})_{\omega}|^2, \text{ for } i = 1, 2.$$

$$\tag{1}$$

The final estimation is given by

$$\hat{\Theta}_{\hat{k}} = \begin{cases} (\hat{\Theta}_{k_2}^{(2)})_{\omega} & \text{if } \omega \in \Omega_1, \\ (\hat{\Theta}_{k_1}^{(1)})_{\omega} & \text{if } \omega \in \Omega_2. \end{cases}$$
 (2)

Remark 1. The above adaptation is based on Gao et al. [2016] and different from regular cross validation approach. My previous thought was to use

$$\hat{\Theta}_{\hat{k}} = \underset{\Theta \in \mathrm{cut}(\mathcal{P}_{k})}{\mathrm{arg} \min} \|\mathcal{A} - \Theta\|_{F}^{2}, \qquad \text{replaced by sum}_{i=1,2}\sum_{j=1,2}\sum_{k=1,2}...$$

which is regular hyperparameter setting based on cross validation.

We can show that the convergence rate of the estimator (2) is the same as  $\hat{\Theta}_k$  where k is the true number of group.

**Theorem 1.1** (Stochastic block model with adaptation of the number of group k). Let  $\hat{\Theta}_{\hat{k}}$  be the estimator from (2). Suppose true probability tensor  $\Theta \in \text{cut}(\mathcal{P}_k)$  for fixed block size k Then, there exists two constants  $C_1, C_2, C_3 > 0$ , such that

$$\frac{1}{n^m} \|\hat{\Theta}_{\hat{k}} - \Theta^{\text{true}}\|_F^2 \le \frac{C_1}{\rho} \left( \left(\frac{k}{n}\right)^m + \frac{\log k}{n^{m-1}} + \left(\frac{\log n}{\rho}\right)^2 \right),$$

with probability at least  $1 - \exp(-C_2(n \log k + k^m)) - (n^m)^{-C_3}$ .

*Proof.* From theorem with known k case, we have

$$\frac{1}{n^m} \|\hat{\Theta}_k - \Theta^{\text{true}}\|_F^2 \le \frac{C_1}{\rho} \left( \left( \frac{k}{n} \right)^m + \frac{\log k}{n^{m-1}} \right), \tag{4}$$

with probability at least  $1 - \exp\left(-C_2(n\log k + k^m)\right)$ . By triangular inequality,

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$$\|\hat{\Theta}_{\hat{k}} - \Theta^{\text{true}}\|_F^2 \le 2 \underbrace{\|\hat{\Theta}_k - \hat{\Theta}_k\|_F^2}_{\text{(i)}} + 2 \underbrace{\|\hat{\Theta}_k - \Theta^{\text{true}}\|_F^2}_{\text{(i)}}. \tag{5}$$
 estimation error within Omega 1: reproduce the

Since we have the error bound (ii) as in (4), we find the upper bound of the error (i). Same within Omega 2

 $\Theta_{\hat{k}}$ , we have the following inequality,

$$\begin{split} \|\hat{\Theta}_{\hat{k}} - \frac{\hat{\Theta}_{k}}{\|\hat{\Omega}_{2}} &\leq 2 \left\langle \hat{\Theta}_{k_{1}}^{(1)} - \frac{\hat{\Theta}_{k}}{\rho}, \frac{\mathcal{A} - \rho \hat{\Theta}_{k}}{\rho} \right\rangle_{\Omega_{2}} & \text{the red inequality does not hold. Holds if yellow term is replaced by ground truth (no randomness)} \\ &\text{replace all yellow terms by Theta_true} \\ &= 2 \left( \left\langle \hat{\Theta}_{k_{1}}^{(1)} - \frac{\hat{\Theta}_{k}}{\rho}, \frac{\mathcal{A} - \rho \Theta^{\text{true}}}{\rho} \right\rangle_{\Omega_{2}} + \langle \hat{\Theta}_{k_{1}}^{(1)} - \hat{\Theta}_{k}, \Theta^{\text{true}} - \hat{\Theta}_{k} \rangle_{\Omega_{2}} \right) \end{split}$$

$$=2\left(\left\langle\hat{\Theta}_{k_1}^{(1)}-\hat{\underline{\Theta}}_k,\frac{\mathcal{A}-\rho\Theta^{\mathrm{true}}}{\rho}\right\rangle_{\Omega_2}+\langle\hat{\Theta}_{k_1}^{(1)}-\hat{\underline{\Theta}}_k,\Theta^{\mathrm{true}}-\hat{\underline{\Theta}}_k\rangle_{\Omega_2}\right)$$
 
$$\leq 2\|\hat{\Theta}_{k_1}^{(1)}-\hat{\underline{\Theta}}_k\|_{\Omega_2}\left(\left\langle\frac{\hat{\Theta}_{k_1}^{(1)}-\hat{\underline{\Theta}}_k}{\|\hat{\Theta}_{k_1}^{(1)}-\hat{\underline{\Theta}}_k\|_{\Omega_2}},\frac{\mathcal{A}-\rho\Theta^{\mathrm{true}}}{\rho}\right\rangle_{\Omega_2}+\|\Theta^{\mathrm{true}}-\hat{\underline{\Theta}}_k\|_{\Omega_2}\right).$$
 Left term in the inner product: randomness comes from Omega 1 Right term: randomness comes from Omega 2

It suffices to bound the inner product term because of (4). I haven't figure out how to derive this inner product part.

$$\max_{k_1 \in [n]} \left\langle \frac{\hat{\Theta}_{k_1}^{(1)} - \hat{\Theta}_k}{\|\hat{\Theta}_{k_1}^{(1)} - \hat{\Theta}_k\|_{\Omega_2}}, \frac{\mathcal{A} - \rho \Theta^{\text{true}}}{\rho} \right\rangle_{\Omega_2} \leq C \frac{\log n}{\rho},$$

with probability at least  $1 - (n^m)^{-C'}$  for some universal constants C, C' > 0. Assuming we proved this bound, we have

$$\|\hat{\Theta}_{\hat{k}} - \hat{\Theta}_k\|_{\Omega_2}^2 \le C_1 \left( \|\Theta^{\text{true}} - \hat{\Theta}_k\|_{\Omega_2}^2 + \left( \frac{\log n}{\rho} \right)^2 \right),$$

for some constant  $C_1 > 0$ . A symmetric argument leads to,

$$\|\hat{\Theta}_{\hat{k}} - \hat{\Theta}_k\|_{\Omega_1}^2 \le C_2 \left( \|\Theta^{\text{true}} - \hat{\Theta}_k\|_{\Omega_1}^2 + \left( \frac{\log n}{\rho} \right)^2 \right),\,$$

for some constant  $C_2 > 0$ .

Summing up the above two inequalities, we have

(i) 
$$\leq C \left( \|\Theta^{\text{true}} - \hat{\Theta}_k\|_{\text{F}}^2 + \left( \frac{\log n}{\rho} \right)^2 \right).$$

Plugging the above inequality in (5) completes the proof.

#### $\mathbf{2}$ Simulation results

Gao et al. [2016] does not estimate the number of clusters in unknown k but suggested new estimation

method. Instead, I estimated k in the simulation by following procedure. First, I calculate  $\mathcal{A}^{\text{test}}$  as

$$\mathcal{A}_{\omega}^{\text{test}} = \begin{cases} \mathcal{A}_{\omega} & \text{if } \omega \in \Omega_1, \\ 0 & \text{if } \omega \in \Omega_2. \end{cases}$$

For many different  $k \in [n]$ , we calculate

missing superscript "test"?

$$\frac{\hat{\Theta}_k}{\Theta \in \text{cut}(\mathcal{P}_k)} = \underset{\Theta \in \text{cut}(\mathcal{P}_k)}{\operatorname{arg\,min}} \|\mathcal{A}^{\text{test}} - \Theta\|_F^2,$$

Based on a series of  $\hat{\Theta}_k$ , I estimate  $\hat{k} = k_1$  such that

$$\hat{k} = \underset{k \in [n]}{\operatorname{arg \, min}} \sum_{\omega \in \Omega_2} |\mathcal{A}_{\omega} - (\hat{\Theta}_k^{\operatorname{test}})_{\omega}|^2.$$

Gao et al has a simulation in table 1, page 13.

**Remark 2.** This procedure is to calculate  $k_1$  in (1). I have not calculated the adaptive estimation of Gao et al. [2016]'s paper. I will compare (2) and (3) later and update the note.

Ground truth of the model is smooth symmetric  $\Theta$  with  $k \in \{5, 10, 15, 20\}$  and  $n \in \{50, 100, 150, 200\}$ . Table 1 summarizes the estimated number of clusters for different true k and the number of nodes. It seems that the estimated number of cluster quite close to true one when n > 50. Figure 1 shows the MSE according to the number of clusters as an input of the algorithm across different ground truth settings.

True # of clusters	5	10	15	20
node 50	4	6	4	6
node 100	5	10	14	9
node 150	5	9	12	16
node 200	5	9	9	15

Table 1: Estimation for the number of clusters according to the number of nodes and true clusters.

### References

Chao Gao, Yu Lu, Zongming Ma, and Harrison H Zhou. Optimal estimation and completion of matrices with biclustering structures. *The Journal of Machine Learning Research*, 17(1):5602–5630, 2016.

### Is the MSE in y-axis testing error or estimation error?

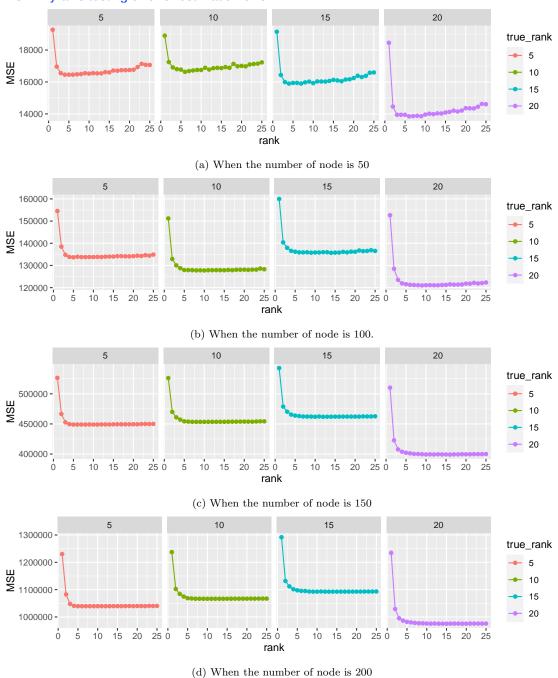


Figure 1: MSE errors on the test set with different input for the number of clusters given the number of node and true clusters.