# Polynomial-time estimation of permutation-equivalence tensors

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### 1 Permutation-equivalence tensor model

Given a symmetric tensor  $\Theta \in \mathbb{R}^{d \times \cdots \times d}$  and a permutation  $\sigma \colon [d] \to [d]$ , we use  $\Theta \circ \sigma$  to denote the permuted tensor

$$(\Theta \circ \sigma)(i_1, \dots, i_m) = \Theta(\sigma(i_1), \dots, \sigma(i_m)), \text{ for all } (i_1, \dots, i_m) \in [d]^m.$$

**Definition 1** (Lipschitz smooth tensor).

$$\mathcal{P}(L) = \left\{ \Theta \colon |\Theta(\omega) - \Theta(\omega')| \le \frac{L|\omega - \omega'|_1}{d}, \text{ for all } \omega, \omega' \in [d]^m \right\}. \tag{1}$$

For simplicity, we consider L=1 thoughout this note.

**Model 1** (Permutated tensor model). Let  $\mathcal{Y} \in \mathbb{R}^{d \times \cdots \times d}$  be a data tensor generated from the model

$$\mathcal{Y} = \Theta \circ \sigma + \mathcal{E} \tag{2}$$

where  $\Theta \in \mathcal{P}(L)$  is an unknown structured tensor,  $\sigma \colon [d] \to [d]$  is an unknown permutation, and  $\mathcal{E} \in \mathbb{R}^{d \times \cdots \times d}$  is a noise tensor consisting of zero-mean standard normal entries.

**Remark 1** (Random design vs. fixed design). Our results below assume no randomness in the signal tensor  $\Theta$ . This is the major distinction between our model and the hypergraphon model. In the graphon model, the data tensor  $\mathcal{Y}$  has two randomnesses: one from the noise tensor  $\mathcal{E}$ , and the other from signal tensor  $\Theta$ ,

$$\Theta(i_1, \dots, i_m) = f(U_{i_1}, \dots, U_{i_m}), \quad \text{with } (U_{i_\ell})_{\ell \in [m]} \sim_{\text{i.i.d}} \text{Unif}[0, 1].$$
(3)

We refer to (3) as the random design, and refer to the grid samples

$$\Theta(i_1, \dots, i_m) = f\left(\frac{i_1}{d}, \dots, \frac{i_m}{d}\right), \quad \text{for all } (i_1, \dots, i_m) \in [d]^m,$$
(4)

as the fixed design. Our permutation-equivalence tensor model specified by (1) and (2) is equivalent to classical hypergraphon model with fixed design (4).

**Assumption 1** ( $\beta$ -monotonicity of degree). Define the degree function  $deg(\cdot)$  associated with tensor  $\Theta$ ,

deg: 
$$[d] \to \mathbb{R}$$
 
$$i \mapsto \frac{1}{d^{m-1}} \sum_{\ell=2}^{m} \sum_{i=1}^{d} \Theta(i, i_2, \dots, i_m)$$

Assume the degree function is strictly increasing and satisfies  $\beta$ -monotonicity with  $\beta \in [0, 1]$ ,

$$\left|\frac{i-j}{d}\right|^{1/\beta} \le \deg(i) - \deg(j), \quad \text{for all } i > j \in [d].$$

**Remark 2.** A lower value of  $\beta$  implies less steepness of the degree function. We make the convention that a constant degree function is 0-monotonic.

Polynomial-time algorithm for estimating  $\Theta$ : Input:  $\mathcal{Y}$ , k; Output:  $\hat{\sigma}$  and  $\hat{\Theta}_{LS}$ .

- 1. Sorting: Sort the nodes based on the empirical degree of  $\mathcal{Y}$ . The sorting returns the node permutation  $\hat{\sigma}^{-1}$ :  $[d] \to [d]$  for which the degree function associated with  $\mathcal{Y} \circ \hat{\sigma}^{-1}$  is non-decreasing in  $i \in [d]$ .
- 2. Blocking: Estimate  $\Theta$  based on block tensor approximation

$$\hat{\Theta}_{LS} = Block_k(\mathcal{Y} \circ \hat{\sigma}^{-1}),$$

where the operator  $\operatorname{Block}_k(\cdot)$  converts a tensor to a block tensor with k equal-sized blocks; i.e,

$$\hat{\Theta}_{\mathrm{LS}}(\omega') := \mathrm{Block}_k(\mathcal{Y} \circ \hat{\sigma}^{-1})(\omega') = \mathrm{Average}\left\{\Theta(\omega) \colon \lfloor \omega k/d \rfloor = \lfloor \omega' k/d \rfloor\right\}, \quad \text{for all } \omega' \in [d]^m.$$

We quantify the estimation error using risk

$$\mathcal{R}(\hat{\Theta}, \Theta) = \frac{1}{d^m} \mathbb{E}_{\mathcal{Y}} \| \hat{\Theta} - \Theta \|_F^2.$$

**Theorem 1.1** (Sorting-and-blocking under  $\beta$ -monotonicity of degree). Consider model 2 under Assumption 1. Set  $k = d^{\frac{m}{2+m}}$  in Algorithm 1. Then, with probability at least  $1 - d^{-1}$ ,

$$\mathcal{R}(\hat{\Theta}_{\mathrm{LS}},\Theta) \leq \underbrace{d^{-\frac{2m}{2+m}}}_{\text{statistical error}} + \underbrace{d^{-\beta(m-1)}}_{\text{algorithmic error}}.$$

**Remark 3.** When  $\beta \ge \frac{2m}{(m-1)(m+2)}$ , the statistical error dominates the algorithmic error. In this regime, we have

$$\mathcal{R}(\hat{\Theta}_{LS}, \Theta) \leq d^{-\frac{2m}{2+m}}.$$

The rate agrees with the best possible rate known for this problem [1]. However, the estimate proposed in [1] is based on a combinatoric search with exponentially computational complexity. In contrast, our estimate is polynomial-time solveable. We show that, under the degree monotonicity assumption, our estimate achieves both statical accuracy and computational efficiency.

Furthermore, the required  $\beta$ -monotonicity becomes weaker as the tensor order m increases. Recall that a lower value of  $\beta$  implies less constrained degree function. We find that the required lower bound threshold  $\beta \geq \frac{2m}{(m-1)(m+2)}$  vanishes to zero as  $m \to \infty$ .

*Proof of theorem 1.1.* We decompose the estimation error into three terms,

$$\|\hat{\Theta} - \Theta\|_F^2 \leq \underbrace{\|\mathrm{Block}_k(\mathcal{Y} \circ \hat{\sigma}^{-1}) - \mathrm{Block}_k(\mathcal{Y} \circ \sigma^{-1})\|_F^2}_{\text{Permutation error; Lemmas 1-2}} + \underbrace{\|\mathrm{Block}_k(\mathcal{Y} \circ \sigma^{-1}) - \mathrm{Block}_k(\Theta)\|_F^2}_{\text{Nonparametric error; Lemma 3}} + \underbrace{\|\mathrm{Block}(\Theta) - \Theta\|_F^2}_{\text{Approximation error; Lemma 4}}$$

$$\leq d^{m} \operatorname{Loss}^{2}(\sigma, \hat{\sigma}) + k^{m} + \frac{d^{m}}{k^{2}}$$

$$\leq d^{-\beta(m-1)+m} + k^{m} + \frac{d^{m}}{k^{2}}$$

$$\leq d^{-\beta(m-1)+m} + d^{\frac{m^{2}}{m+2}}$$

Lemma 1 (Permutation error). Step 1 in the algorithm yields the permutation error

$$\operatorname{Loss}(\sigma, \hat{\sigma}) := \frac{1}{d} \max_{i} |\sigma(i) - \hat{\sigma}(i)| \le d^{-(m-1)\beta/2},$$

with probability at least  $1 - \exp(-d)$ .

**Lemma 2** (Estimation error due to permutation; Lemma 3 in [2]). With probability at least  $1 - \exp(-d)$ ,

$$\|\operatorname{Block}_k(\mathcal{Y}\circ\hat{\sigma}^{-1})-\operatorname{Block}_k(\mathcal{Y}\circ\sigma^{-1})\|_F^2\leq d^m\operatorname{Loss}^2(\sigma,\hat{\sigma}).$$

**Remark 4.** Lemma 2 quantifies the estimation error due to permutation error. **Lemma 3** (Denoising error; Lemma 4 in [2]). With probability at least  $1 - \exp(-d)$ ,

$$\|\operatorname{Block}_k(\mathcal{Y} \circ \sigma^{-1}) - \operatorname{Block}_k(\Theta)\|_F^2 \le k^m.$$

**Lemma 4** (Approximation error from Lee's 0225 note; corrected Lemma 1 in [2]). Suppose the true parameter  $\Theta$  is from (1). For every fixed integer  $k \leq d$ , we have

$$\|\operatorname{Block}_k(\Theta) - \Theta\|_F^2 \le \frac{d^m}{k^2}.$$

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#### 2 Proofs

*Proof of Lemma 1.* By definition, deg(i) is the sample average of roughly  $d^{(m-1)}$  i.i.d. terms except for at most a few diagonal terms. With high probability, the stochastic deviation satisfies

$$\deg(i) - \widehat{\deg}(i) \lesssim d^{-(m-1)/2}.$$

Without loss of generality, assume that  $\sigma$  is the identity permutation. By definition,

$$\deg(1) \le \deg(2) \le \dots \le \deg(d-1) \le \deg(d). \tag{5}$$

The estimated permutation  $\hat{\sigma}$  is defined for which

$$\widehat{\operatorname{deg}}(\widehat{\sigma}(1)) \le \widehat{\operatorname{deg}}(\widehat{\sigma}(2)) \le \dots \le \widehat{\operatorname{deg}}(\widehat{\sigma}(d-1)) \le \widehat{\operatorname{deg}}(\widehat{\sigma}(d)). \tag{6}$$

Now, for any given index i, we examine the error  $|i - \hat{\sigma}(i)|$ . By (5) and (6), we have

$$i = |\underbrace{\{j \colon \deg(j) \le \deg(i)\}}_{=:\mathbf{I}}|, \quad \text{and} \quad \widehat{\sigma}(i) = |\underbrace{\{j \colon \widehat{\deg}(\widehat{\sigma}(j)) \le \widehat{\deg}(\widehat{\sigma}(i))\}}_{=:\mathbf{II}}|,$$

where  $|\cdot|$  denotes the cardinality of the set. We claim that the sets I and II differ only in at most  $d^{(m-1)\beta/2}$  elements. To prove this, we partition the nodes in [d] in two cases.

1. long-distance nodes in  $\{j: |j-i| \gg d^{1-(m-1)\beta/2}\}$ . In this case, the ordering of (i,j) remains the same in (6) and (5), i.e,

$$\deg(i) < \deg(j) \iff \widehat{\deg}(\widehat{\sigma}(i)) < \widehat{\deg}(\widehat{\sigma}(j)). \tag{7}$$

The  $\Longrightarrow$  in (7) is because

$$\widehat{\operatorname{deg}}(\widehat{\sigma}(j)) - \widehat{\operatorname{deg}}(\widehat{\sigma}(i)) \geq \underbrace{\left\{\widehat{\operatorname{deg}}(\widehat{\sigma}(j)) - \operatorname{deg}(j)\right\}}_{\leq d^{-(m-1)/2}} - \underbrace{\left\{\widehat{\operatorname{deg}}(\widehat{\sigma}(i)) - \operatorname{deg}(i)\right\}}_{\leq d^{-(m-1)/2}} + \underbrace{\left\{\operatorname{deg}(j) - \operatorname{deg}(i)\right\}}_{\gg d^{-(m-1)/2}} > 0,$$

where the third term in the inequality is due to  $\beta$ -smoothness of deg(·) and the assumption  $|j-i| \gg d^{1-\beta(m-1)/2}$ . The other direction in (7) can be similarly proved. Therefore, we conclude that none of long-distance nodes belong to I  $\Delta$  II.

2. short-distance nodes in  $\{j: |j-i| \le d^{1-\beta(m-1)/2}\}$ . In this case, (6) and (5) may yield different ordering of (i,j).

Combining the above two cases gives that

$${j: |j-i| \le d^{1-\beta(m-1)}} \supset I\Delta II.$$

Therefore,

$$\operatorname{Loss}(\sigma, \hat{\sigma}) := \frac{1}{d} \max_{i} |i - \hat{\sigma}(i)| \le \frac{1}{d} |\operatorname{I}\Delta\operatorname{II}| \le d^{-\beta(m-1)/2}.$$

# 3 Further thoughts

Step 1 is equivalent to

$$\hat{\sigma} = \underset{\sigma : [d] \to [d]}{\operatorname{arg \, min}} \sum_{i \in [d-1]} \operatorname{dist}(\sigma(i), \sigma(i+1)), \quad \text{where} \quad \operatorname{dist}(x, y) := |\widehat{\operatorname{deg}}(x) - \widehat{\operatorname{deg}}(y)|. \tag{8}$$

The optimization (8) has closed form solution under the degree-based distance function. Specifically, the optimizer of (8) is uniquely determined by the sorting

$$\widehat{\operatorname{deg}}(\widehat{\sigma}(1)) \le \cdots \le \widehat{\operatorname{deg}}(\widehat{\sigma}(d)).$$

Can the above framework incorporate the neighborhood estimator in [3]? List the corresponding Lemmas 1-4 for the estimator in [3]. Which steps make the estimate [3] less optimal?

## References

- [1] Krishnakumar Balasubramanian, Nonparametric modeling of higher-order interactions via hypergraphons, arXiv preprint arXiv:2105.08678 (2021).
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- [3] Yuan Zhang, Elizaveta Levina, and Ji Zhu, Estimating network edge probabilities by neighborhood smoothing, Biometrika 104 (2015), no. 4, 771–783.