

Joint estimation of $\hat{\Theta}$ and $\hat{\mathbf{b}}$ in the tensor ordinal model

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Remark 1 (Notation). For notational convenience, we consider the special case $d_1 = \dots = d_K = d$. We focus on the asymptotic region $d \rightarrow \infty$ and treat all other quantities as constants, i.e., $\mathbf{r}, L, U_{\alpha, \beta, \Delta}, L_{\alpha, \beta, \Delta}, \dots = \mathcal{O}(1)$. In particular,

1. $\prod_k d_k \asymp d^K, \sum_k d_k \asymp d$.
2. Denote $n_{\max} = \max_{\ell} n_{\ell}$ and $n_{\min} = \min_{\ell} (n_{\ell} + n_{\ell+1})$. Then, $n_{\max} \asymp d^K$ (important!) but $\mathcal{O}(1) \leq n_{\min} \leq \mathcal{O}(d^K)$.

Note: we use $C >$ to denote a positive constant whose value may change from line to line.

1 Results

1.1 Current bound

Total MSE:

$$\begin{aligned} \text{MSE}((\hat{\Theta}, \hat{\mathbf{b}}), (\Theta^{\text{true}}, \mathbf{b}^{\text{true}})) &\stackrel{\text{def}}{=} \frac{\prod_k d_k \text{MSE}(\hat{\Theta}, \Theta^{\text{true}}) + (L-1) \text{MSE}(\hat{\Theta}, \Theta^{\text{true}})}{\prod_k d_k + L - 1} \\ &\leq \frac{c_1 \sum_k d_k + c_2 \frac{n_{\max}^2}{n_{\min}^2}}{\prod_k d_k + L - 1} \\ &= \mathcal{O}\left(\frac{\sum_k d_k}{\prod_k d_k}\right) + \mathcal{O}\left(\frac{n_{\max}^2}{n_{\min}^2} \frac{1}{\prod_k d_k}\right) \end{aligned} \quad (1)$$

Remark 2. Unfortunately, this total MSE bound does not converge to zero. In the worse case, the bound can be $\asymp d^K$; for example, the bound diverges when $n_{\min} \asymp 1$ (and $n_{\max} \asymp d^K$).

Remark 3. The total MSE bound converges to zero only when

$$\frac{n_{\max}}{n_{\min}} \ll d^{K/2}, \quad \text{or equivalently, } n_{\min} \gg d^{K/2}. \quad (2)$$

In other words, the current bound (1) tolerates only certain imbalanced classes for which $\mathcal{O}(d^{K/2}) \leq n_{\min} \leq \mathcal{O}(d^K)$.

1.2 Sharper bound

We will prove a sharper bound on the linear term

$$\begin{aligned} (\mathbf{b}^{\text{true}} - \hat{\mathbf{b}})^T \nabla_{\mathbf{b}} \mathcal{L}_{\mathcal{Y}}(\mathbf{b}^{\text{true}}) &\leq \|\mathbf{b}^{\text{true}} - \hat{\mathbf{b}}\|_F \|\nabla_{\mathbf{b}} \mathcal{L}_{\mathcal{Y}}(\mathbf{b}^{\text{true}})\|_F \\ &\leq C \|\mathbf{b}^{\text{true}} - \hat{\mathbf{b}}\|_F \sqrt{d^{K+2}}, \end{aligned}$$

where the last inequality is followed from a sharper bound on the gradient:

$$\left| \frac{\partial \mathcal{L}_{\mathcal{Y}}}{\partial b_{\ell}} \right|_{(\hat{\Theta}, \mathbf{b}^{\text{true}})} \leq C \sqrt{d^{K+2}}, \quad \text{for all } \ell \in [L-1]. \quad (3)$$

Suppose (3) holds. Following the same line in the current proof (i.e., Taylor expansion, quadratic bound, etc.), we have

$$\|\mathbf{b}^{\text{true}} - \hat{\mathbf{b}}\|_F \leq C \frac{d^{(K+2)/2}}{n_{\min}}.$$

(Final results.) Therefore, the total MSE:

$$\text{Total MSE} \leq \mathcal{O}\left(\frac{1}{d^{K-1}}\right) + \mathcal{O}\left(\frac{d^{K+2}}{n_{\min}^2 d^K}\right) = \mathcal{O}\left(\frac{d^2}{\min\{d^{K+1}, n_{\min}^2\}}\right) \quad (4)$$

which convergences to zero whenever

$$n_{\min} \gg d. \quad (5)$$

In other words, the new bound (4) tolerates highly imbalanced classes for which $\mathcal{O}(\sqrt{d}) \leq n_{\min} \leq \mathcal{O}(d^K)$.

Remark 4. The consistency condition (5) is more relaxed than (2) for all $K \geq 3$. Both bounds agree in the matrix case ($K = 2$).

Remark 5. When $n_{\min} \gg d^{(K+1)/2}$, the error in $\hat{\Theta}$ dominates the total MSE. Then $n_{\min} \ll d^{(K+1)/2}$, the error in $\hat{\mathbf{b}}$ dominates the total MSE

2 Proofs

Now we prove the inequality (3).

Lemma 1 (Sharper Bound on Gradients). *Consider the same set-up as in Theorem A.1. Then, with very high probability,*

$$\left| \frac{\partial \mathcal{L}_y}{\partial b_\ell} \Big|_{(\hat{\Theta}, \mathbf{b}^{\text{true}})} \right| \leq C_1 d^{K/2} + C_2 d^{K/2} \|\Theta^{\text{true}} - \hat{\Theta}\|_F, \quad \text{for all } \ell \in [L-1].$$

where $C_1, C_2 > 0$ are two constants. In particular, there exists a constant $d_0 \in \mathbb{N}_+$, such that for all $d \geq d_0$,

$$\left| \frac{\partial \mathcal{L}_y}{\partial b_\ell} \Big|_{(\hat{\Theta}, \mathbf{b}^{\text{true}})} \right| \leq C d^{(K+2)/2}, \quad \text{for all } \ell \in [L-1].$$

Corollary 1 (MSE for $\hat{\mathbf{b}}$). *Under the same set-up as in Theorem A.1, we have*

$$\|\hat{\mathbf{b}} - \mathbf{b}^{\text{true}}\|_F \leq \frac{C_1 d^{K/2} + C_2 d^{K/2} \|\hat{\Theta} - \Theta^{\text{true}}\|_F}{n_{\min}} \leq \frac{C d^{(K+2)/2}}{n_{\min}}.$$

Remark 6. The bound (3) is sharper than the trivial bound $|b_\ell^{\text{true}} - \hat{b}_\ell| \leq 2\beta$. In particular, $\|\hat{\mathbf{b}} - \mathbf{b}^{\text{true}}\|_F \rightarrow 0$ as $n_{\min} \asymp (d^{(K+2)/2}) \rightarrow \infty$.

Proof of Lemma 1. We only prove the case for $\ell = 1$. Other cases can be proved similarly.

Note that

$$\frac{\partial \mathcal{L}_y}{\partial b_1} \Big|_{(\hat{\Theta}, \mathbf{b}^{\text{true}})} = \underbrace{\frac{\partial \mathcal{L}_y}{\partial b_1} \Big|_{(\hat{\Theta}, \mathbf{b}^{\text{true}})} - \mathbb{E}_y \left[\frac{\partial \mathcal{L}_y}{\partial b_1} \Big|_{(\hat{\Theta}, \mathbf{b}^{\text{true}})} \right]}_{:=A} + \underbrace{\mathbb{E}_y \left[\frac{\partial \mathcal{L}_y}{\partial b_1} \Big|_{(\hat{\Theta}, \mathbf{b}^{\text{true}})} \right] - \mathbb{E}_y \left[\frac{\partial \mathcal{L}_y}{\partial b_1} \Big|_{(\Theta^{\text{true}}, \mathbf{b}^{\text{true}})} \right]}_{:=B} \quad (6)$$

where we have used the fact that the score function has mean zero, $\mathbb{E}_{\mathcal{Y}} \left[\frac{\partial \mathcal{L}_{\mathcal{Y}}}{\partial b_1} \middle|_{(\Theta^{\text{true}}, \mathbf{b}^{\text{true}})} \right] = 0$. Here all expectations are taken with respect to $\mathcal{Y} \sim \mathbb{P}(\Theta^{\text{true}}, \mathbf{b}^{\text{true}})$.

We now bound the two deviation terms in (6) separately. The term A in (6) is the stochastic deviation of log-likelihood to its expectation:

$$A = \sum_{\omega \in \Omega} \left\{ \underbrace{[\mathbb{1}_{\{y_{\omega}=1\}} - g_1(\theta_{\omega}^{\text{true}})] \frac{\dot{f}(b_1 - \hat{\theta}_{\omega})}{g_1(\hat{\theta}_{\omega})} - [\mathbb{1}_{\{y_{\omega}=2\}} - g_2(\theta_{\omega}^{\text{true}})] \frac{\dot{f}(b_1 - \hat{\theta}_{\omega})}{g_2(\hat{\theta}_{\omega})}}_{:=W_{\omega}} \right\}.$$

Note that $\{W_{\omega}\}$ are i.i.d. random variables, and each W_{ω} has zero mean and bounded variance:

$$\text{Var}(W_{\omega}) \leq C (g_1(\theta_{\omega}^{\text{true}})(1 - g_1(\theta_{\omega}^{\text{true}})) + g_2(\theta_{\omega}^{\text{true}})(1 - g_2(\theta_{\omega}^{\text{true}}))) \leq C/2.$$

By central limit theorem (verify..)

$$\sum_{\omega \in \Omega} W_{\omega} \xrightarrow{\mathcal{D}} N(0, Cd^K), \quad \text{as } d^K \rightarrow \infty.$$

Hence, with very high probability,

$$|A| = \left| \sum_{\omega \in \Omega} W_{\omega} \right| \leq Cd^{K/2}. \quad (7)$$

The second term B in (6) is the bias induced by the inaccuracy of $\hat{\Theta}$:

$$\begin{aligned} B &= \sum_{\omega \in \Omega} g_1(\theta_{\omega}^{\text{true}}) \left(\frac{\dot{f}(b_1 - \hat{\theta}_{\omega})}{g_1(\hat{\theta}_{\omega})} - \frac{\dot{f}(b_1 - \theta_{\omega}^{\text{true}})}{g_1(\theta_{\omega}^{\text{true}})} \right) - \sum_{\omega \in \Omega} g_2(\theta_{\omega}^{\text{true}}) \left(\frac{\dot{f}(b_2 - \hat{\theta}_{\omega})}{g_2(\hat{\theta}_{\omega})} - \frac{\dot{f}(b_2 - \theta_{\omega}^{\text{true}})}{g_2(\theta_{\omega}^{\text{true}})} \right) \\ &\leq \sum_{\omega \in \Omega} g_1(\theta_{\omega}^{\text{true}})(\theta_{\omega}^{\text{true}} - \hat{\theta}_{\omega}) \left\{ \frac{\partial}{\partial \theta} \left(\frac{\dot{f}(b_1 - \theta)}{g_1(\theta)} \right) \middle|_{\rho \hat{\theta}_{\omega} + (1-\rho)\theta_{\omega}^{\text{true}}} \right\} \\ &\quad - \sum_{\omega \in \Omega} g_2(\theta_{\omega}^{\text{true}})(\theta_{\omega}^{\text{true}} - \hat{\theta}_{\omega}) \left\{ \frac{\partial}{\partial \theta} \left(\frac{\dot{f}(b_2 - \theta)}{g_2(\theta)} \right) \middle|_{\rho' \hat{\theta}_{\omega} + (1-\rho')\theta_{\omega}^{\text{true}}} \right\} \\ &\leq C \sum_{\omega \in \Omega} [g_1(\theta_{\omega}^{\text{true}}) - g_2(\theta_{\omega}^{\text{true}})] (\theta_{\omega}^{\text{true}} - \hat{\theta}_{\omega}). \end{aligned}$$

By Cauchy-Schwartz inequality,

$$|B| \leq Cd^{K/2} \|\Theta^{\text{true}} - \hat{\Theta}\|_F. \quad (8)$$

Plugging (7) and (8) back to (6) yields that

$$\left| \frac{\partial \mathcal{L}_{\mathcal{Y}}}{\partial b_{\ell}} \middle|_{(\hat{\Theta}, \mathbf{b}^{\text{true}})} \right| \leq C_1 d^{K/2} + C_2 d^{K/2} \|\Theta^{\text{true}} - \hat{\Theta}\|_F$$

holds with very high probability. The second inequality in the conclusion comes from the fact that $\|\Theta^{\text{true}} - \hat{\Theta}\|_F \leq \mathcal{O}(d)$ as $d \rightarrow \infty$. \square