Rademacher complexity and tuning parameter

1 Rademacher complexity

Theorem 1.1. Let $K : \mathbb{R}^{d_1 \times d_2} \times \mathbb{R}^{d_1 \times d_2} \to \mathbb{R}^{d_1 \times d_1}$ be bounded with $\sqrt{tr(K(\boldsymbol{X}, \boldsymbol{X}))} \leq G$ and let $h : \mathbb{R}^{d_1 \times d_2} \to \mathbb{R}^{d_1 \times d_2}$ be a corresponding feature mapping such that $K(\boldsymbol{X}, \boldsymbol{X}') = h(\boldsymbol{X})h(\boldsymbol{X}')^T$. Define

$$\mathcal{F}_r = \{ f : \mathbb{R}^{d_1 \times d_2} \to \mathbb{R} : f(\mathbf{X}) = \langle \mathbf{B}, h(\mathbf{X}) \rangle \text{ with } \mathbf{B} \in \mathcal{B} \},$$

where $\mathcal{B} = \{ \mathbf{B} \in \mathbb{R}^{d_1 \times d_2'} : rank(\mathbf{B}) \leq r, \lambda_1(\mathbf{B}) \leq M \}$. Then

$$\mathcal{R}_n(\mathcal{F}_r) \leq \frac{2MG\sqrt{r}}{\sqrt{n}}.$$

Proof. Since the hinge loss is an 1-Lipschitz function,

$$\mathcal{R}_n(\mathcal{F}_r) = 2\mathbb{E} \sup_{\boldsymbol{B} \in \mathcal{B}} \frac{1}{n} \sum_{i=1}^n \sigma_i (1 - y_i \langle \boldsymbol{B}, h(\boldsymbol{X}_i) \rangle)_+ \le 2\mathbb{E} \sup_{\boldsymbol{B} \in \mathcal{B}} \frac{1}{n} \sum_{i=1}^n \sigma_i \langle \boldsymbol{B}, h(\boldsymbol{X}_i) \rangle,$$

where $\{\sigma_i\}_{i=1}^n$ is independent Rademacher random variables with $\mathbb{P}(\sigma_i = \pm 1) = 1/2$. The result follows by observing

$$2\mathbb{E}\sup_{\boldsymbol{B}\in\mathcal{B}}\frac{1}{n}\sum_{i=1}^{n}\sigma_{i}\langle\boldsymbol{B},h(\boldsymbol{X}_{i})\rangle = \frac{2}{n}\mathbb{E}\sup_{\boldsymbol{B}\in\mathcal{B}}\langle\boldsymbol{B},\sum_{i=1}^{n}\sigma_{i}h(\boldsymbol{X}_{i})\rangle$$

$$\leq \frac{2}{n}\mathbb{E}\sup_{\boldsymbol{B}\in\mathcal{B}}\|\boldsymbol{B}\|\left\|\sum_{i=1}^{n}\sigma_{i}h(\boldsymbol{X}_{i})\right\|, \text{ by Cauchy-Schwartz inequality}$$

$$\leq \frac{2}{n}\mathbb{E}\sup_{\boldsymbol{B}\in\mathcal{B}}\lambda_{1}\sqrt{r}\left\|\sum_{i=1}^{n}\sigma_{i}h(\boldsymbol{X}_{i})\right\|$$

$$\leq \frac{2M\sqrt{r}}{n}\mathbb{E}\left\|\sum_{i=1}^{n}\sigma_{i}h(\boldsymbol{X}_{i})\right\|$$

$$\leq \frac{2M\sqrt{r}}{n}\sqrt{\mathbb{E}\left\langle\sum_{i=1}^{n}\sigma_{i}h(\boldsymbol{X}_{i}),\sum_{i=1}^{n}\sigma_{i}h(\boldsymbol{X}_{i})\right\rangle} \text{ by Jensen's inequality}$$

$$\leq \frac{2M\sqrt{r}}{n}\sqrt{\sum_{i=1}^{n}\|h(\boldsymbol{X}_{i})\|^{2}}$$

$$= \frac{2M\sqrt{r}}{n}\sqrt{\sum_{i=1}^{n}\text{tr}\left(K(\boldsymbol{X}_{i},\boldsymbol{X}_{i})\right)} \leq \frac{2MG\sqrt{r}}{\sqrt{n}}.$$

Remark 1. If we choose K as a linear kernel, Theorem 1.1 reduces to the linear SMM Rademacher complexity case.

Corollary 1.1. Assume the same condition in Theorem 1.1. Then, with probability at least $1 - \delta$, the generalization error of low-rank SMM is

$$\mathbb{P}\{Y^{new} \neq sign(\hat{f}(\boldsymbol{X}^{new}))\} \leq training \ error + \frac{MG\sqrt{r}}{\sqrt{n}} + \sqrt{\frac{\ln(\frac{1}{\delta})}{2n}}.$$

2 Tuning parameter

In the probability estimation, we assume that we selected the optimal tuning parameter λ and the rank r. In practice, the tuning parameter selection can be done using an independent validation set or cross-validation. From the available dataset N, I propose to use one half for training and the other half for tuning, i.e. $n_{\text{train}} = n_{\text{tune}} = \frac{N}{2}$.

Detail procedure for parameter tuning is as follows

- 1. We obtain the tuning grid $\{(\lambda_i, r_j) : \lambda_1 < \cdots < \lambda_M, r_j = j \in \{1, \dots, \min(d_1, d_2)\}\}$
- 2. For a given (λ_i, r_i) we obtain probability estimates

$$\hat{p}^{(\lambda_i,r_j)}(\boldsymbol{X}_k) = \hat{\mathbb{P}}^{(\lambda_i,r_j)}(y=1|\boldsymbol{X}_k), \quad k=1,\ldots,n_{\text{tune}}.$$

3. We evaluate the log-likelihood

$$L(\lambda_i, r_j) = \sum_{k=1}^{n_{\text{tune}}} \log(\hat{p}^{(\lambda_i, r_j)}(\boldsymbol{X}_k)).$$

+ sum_{y=-1} log(1-\hat p)

4. We choose the optimal tuning parameter $(\lambda_{\hat{i}}, r_{\hat{j}})$ that minimizes BIC value based on the log-likelihood

$$(\hat{i}, \hat{j}) = \underset{i,j}{\arg\min} -2L(\lambda_i, r_j) + r_j(d_1 + d_2 - r_j) \log(n_{\text{tune}}).$$

This grid search might requires too many calculations because we have to perform $M \times \min(d_1, d_2)$ times probability estimation. One way to avoid this grid search is to find the best tuning parameters using profile method. First, fix rank r first and find the best λ . Second, find the best rank r fixing the obtained λ . In this way, we can reduce the number of trials to $M + \min(d_1, d_2)$.

3 Consistency of Probability estimation

Our estimation method is based on the following optimization problem.

$$\min_{f \in \mathcal{F}} n^{-1} \left[(1 - \pi) \sum_{y_i = 1} (1 - y_i f(\mathbf{X}_i))_+ + \pi \sum_{y_i = -1} (1 - y_i f(\mathbf{X}_i))_+ \right] + \lambda J(f).$$
 (1)

In (1), when $n \to \infty$, the first component approaches

$$\mathbb{E}\left[S(Y)(1-Yf(\boldsymbol{X}))_{+}\right] \quad \text{where } S(Y) = 1-\pi \text{ if } Y = 1, \text{ and } \pi \text{ otherwise.}$$

We prove that minimizing (2) with respect to f yields the Bayes rule $\bar{f}_{\pi}(\mathbf{X}) = \text{sign}(p(\mathbf{X}) - \pi)$ where $p(\mathbf{X}) = \mathbb{P}(Y = 1|\mathbf{X})$. The following theorem is referred from [1]. Every argument works through

on our setting because this theorem is specified in terms of complexity of considered function space.

Define $e_V(f, \bar{f}_{\pi}) = \mathbb{E}\{V(f, \mathbf{X}, y) - V(\bar{f}_{\pi}, \mathbf{X}, y)\}$ where $V(f, \mathbf{X}, y) = S(y)(1 - yf(\mathbf{X}))_+$. There are three assumptions to be made for the theorem.

Assumption 1. For some positive sequence such that $s_n \to 0$ as $n \to \infty$, there exists $f_{\pi}^* \in \mathcal{F}$ such that $e_V(f_{\pi}^*, \bar{f}_{\pi}) \leq s_n$.

Assumption 1 ensures that the Bayes rule \bar{f}_{π} is well approximated by \mathcal{F} .

Define a truncated V by $V^T(f, \mathbf{X}, y) = V(f, \mathbf{X}, y) \mathbb{1}\{V(f, \mathbf{X}, y) \leq T\} + T\mathbb{1}\{V(f, \mathbf{X}, y) > T\}$ for some truncation constant T such that $\max\{V(\bar{f}_{\pi}, \mathbf{X}, y), V(f_{\pi}^*, \mathbf{X}, y)\} \leq T$ almost surely, and $e_{V^T}(f, \bar{f}_{\pi}) = \mathbb{E}\{V^T(f, \mathbf{X}, y) - V(\bar{f}_{\pi}, \mathbf{X}, y)\}.$

Assumption 2. There exist constants $0 \le \alpha < \infty, 0 \le \beta \le 1, a_1 > 0$ and $a_2 > 0$ such that, for any sufficiently small $\delta > 0$,

$$\sup_{\{f \in \mathcal{F}: e_{V^T}(f, \bar{f}_{\pi}) \leq \delta\}} \|sign(f) - sign(\bar{f}_{\pi})\|_1 \leq a_1 \delta^{\alpha},$$

$$\sup_{\{f \in \mathcal{F}: e_{V^T}(f, \bar{f}_{\pi}) \leq \delta\}} var\{V^T(f, \boldsymbol{X}, y) - V(\bar{f}_{\pi}, \boldsymbol{X}, y)\} \leq a_2 \delta^{\beta}.$$

Assumption 2 describe local smoothness within a neighborhood of \bar{f}_{π} .

We define the L_2 metric entropy with bracketing that measures the cardinality of \mathcal{F} . Given any $\epsilon > 0$, define $\{(f_m^\ell, f_m^u)\}_{m=1}^M$ to be an ϵ -bracketing function set of \mathcal{F} if for any $f \in \mathcal{F}$, there exists an m such that $f_m^\ell \leq f \leq f_m^u$ and $\|f_m^\ell - f_m^u\|_2 \leq \epsilon$ for $m = 1, \ldots, M$. Then L_2 -metric entropy with bracketing $H_2(\epsilon, \mathcal{F})$ is defined as the logarithm of the cardinality of the smallest ϵ -bracketing function set of \mathcal{F} . Let $\mathcal{F}^V(k) = \{V^T(f, \mathbf{X}, y) - V(f_\pi^*, \mathbf{X}, y) : f \in \mathcal{F}(k)\}$ where $\mathcal{F}(k) = \{f \in \mathcal{F} : \frac{1}{2} \|f\|_k^2 \leq k\}$ and $J_\pi^* = \max\{J(f_\pi^*), 1\}$.

Assumption 3. For some constant $a_3, a_4, a_5 > 0$, and $\epsilon_n > 0$,

$$\sup_{k\geq 2} \int_{a_4L}^{\sqrt{a_3L^{\beta}}} \sqrt{H_2(\omega,\mathcal{F}^V(k))} d\omega/L \leq a_5\sqrt{n}, \text{ where } L = L(\epsilon,\lambda,k) = \min\{\epsilon^2 + \lambda(k/2-1)J_{\pi}^*,1\}.$$

Theorem 3.1. Under Assumptions 1-3, for the estimator \hat{p} obtained from our method, there exists a constant $a_6 > 0$ such that

$$\mathbb{P}\left\{\|\hat{p} - p\|_1 \ge \frac{1}{2m} + \frac{1}{2}a_1(m+1)\delta_n^{2\alpha}\right\} \le 15\exp\{-a_6n(\lambda J_\pi^*)^{2-\beta}\},\,$$

provided that $\lambda^{-1} \ge 4\delta_n^{-2}J_\pi^*$, where $\delta_n^2 = \min\{\max(\epsilon_n^2, s_n), 1\}$

4 Covering number bounds of linear function classes

From theorems in [2], we can calculate the entropy of linear function class in our setting. This following lemma might be helpful for checking assumptions in Section 3.

Lemma 1. Define $\mathcal{F} = \{f : \mathbb{R}^{d_1 \times d_2} \to \mathbb{R} : f(\boldsymbol{X}) = \langle \boldsymbol{B}, \boldsymbol{X} \rangle \text{ with } \boldsymbol{B} \in \mathcal{B} \}$ under the condition that $\|\boldsymbol{X}\| \leq G$, there exists constraints c, c' > 0 such that for all $n \in \mathbb{N}$ and all $\epsilon > 0$,

$$\log_2 H_2(\epsilon, \mathcal{F}) \le \left\lfloor \frac{M^2 G^2 r}{\epsilon^2} \right\rfloor \log_2(2d_1d_2 + 1).$$

References

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Convergence rate for non-parametric regression with high-dimensional matrix predictors

1 Linear case

Define

$$\mathcal{F}_r = \{ f : \mathbb{R}^{d_1 \times d_2} \to \mathbb{R} : f(\boldsymbol{X}) = \langle \boldsymbol{B}, \boldsymbol{X} \rangle \text{ with } \boldsymbol{B} \in \mathcal{B} \},$$

where $\mathcal{B} = \{ \boldsymbol{B} \in \mathbb{R}^{d_1 \times d_2} : \operatorname{rank}(\boldsymbol{B}) \leq r, \lambda_1(\boldsymbol{B}) \leq M \}$. Let \bar{f}_{π} be a Bayes rule. In addition, let $e_V(f, \bar{f}_{\pi}) = \mathbb{E}\{V(f, \boldsymbol{X}, y) - V(\bar{f}_{\pi}, \boldsymbol{X}, y)\}$ with $V(f, \boldsymbol{X}, y) = S(y)L\{yf(\boldsymbol{X})\}$.

Based on function class \mathcal{F}_r , we have the following theorem.

Theorem 1.1 (linear case). Assume that

- 1. For some positive sequence such that $s_n \to 0$ as $n \to \infty$, there exists $f_{\pi}^* \in \mathcal{F}_r$ such that $e_V(f_{\pi}^*, \bar{f}_{\pi}) \leq s_n$.
- 2. There exists $0 \le \alpha < \infty$ and $a_1 > 0$ such that, for any sufficiently small $\delta > 0$,

$$\sup_{\{f \in \mathcal{F}: e_V(f, \bar{f}_\pi) \le \delta\}} \|sign(f) - sign(\bar{f}_\pi)\|_1 \le a_1 \delta^{\alpha},$$

3. Considered feature space is uniformly bounded such that there exists $0 < G < \infty$ satisfying $\|X\| \leq G$

Then, for the estimator \hat{p} obtained from our algorithm, there exists a constant a_2 such that

$$\mathbb{P}\left\{\|\hat{p} - p\|_1 \ge \frac{1}{2m} + \frac{1}{2}a_1(m+1)\delta_n^{2\alpha}\right\} \le 15\exp\{-a_2n(\lambda J_{\pi}^*)\},\,$$

provided that
$$\lambda^{-1} \ge \frac{rGJ_{\pi}^*}{2\delta_n^2}$$
 where $J_{\pi}^* = \max(J(f_{\pi}^*), 1)$ and $\delta_n = \max\left(\mathcal{O}\left(rG\exp\left(-\frac{\sqrt{n}}{rG}\right)^{2/3}\right), s_n\right)$.

Proof. We apply Theorem 3 in [2] to our case. First, notice that trucation on the loss function V is not needed by third assumption:

$$||uf(X)|| = ||\langle B, X \rangle|| < ||B|| ||X|| < rGM,$$

which implies uniformly boundness of V. Let V be bounded by T. For the second equation of Assumption 2 in [2],

$$\operatorname{var}\{V(f, \boldsymbol{X}, y) - V(\bar{f}_{\pi}, \boldsymbol{X}, y)\} \leq \mathbb{E}|V(f, \boldsymbol{X}, y) - V(\bar{f}_{\pi}, \boldsymbol{X}, y)|^{2}$$

$$\leq T\mathbb{E}|V(f, \boldsymbol{X}, y) - V(\bar{f}_{\pi}, \boldsymbol{X}, y)|$$

$$= Te_{V}(f, \bar{f}_{\pi}).$$

Therefore, β in [2] can be replaced by 1 from the following inequality.

$$\sup_{\{f \in \mathcal{F}: e_V(f, \bar{f}_\pi) \leq \delta\}} \operatorname{var}\{V(f, \boldsymbol{X}, y) - V(\bar{f}_\pi, \boldsymbol{X}, y)\} \leq \sup_{\{f \in \mathcal{F}: e_V(f, \bar{f}_\pi) \leq \delta\}} Te_V(f, \bar{f}_\pi) \leq T\delta$$

Now we check Assumption 3 in [2]. Notice that

$$H_2(\epsilon, \mathcal{F}^V(k)) \le H_2(\epsilon, \mathcal{F}(k)) \le \log N_\infty(\epsilon, \mathcal{F}(k)),$$
 (1)

because for functions f_{ℓ} and f_{u} , $||V(f_{\ell},\cdot) - V(f_{u},\cdot)||_{2} \le ||f_{\ell} - f_{u}||_{2}$. The last inequality in (1) is from Lemma 9.22 in [1]. From [3], we have

$$\log N_{\infty}(\epsilon, \mathcal{F}(k)) \le \mathcal{O}\left(\left(\frac{k}{\epsilon}\right)^2 \log\left(\frac{k}{\epsilon}\right)\right).$$

Therefore,

$$\phi(\epsilon, k) \approx \int_{\mathcal{O}(L)}^{\mathcal{O}(\sqrt{L})} \left(\frac{k}{\omega}\right) \sqrt{\log\left(\frac{k}{\omega}\right)} d\omega \approx \mathcal{O}\left(k\left(\log\left(\frac{k}{L}\right)\right)^{3/2}\right),$$

where $L = \min\{\epsilon^2 + \lambda(k/2 - 1)H_\pi^*, 1\}$. Solving Assumption 3 in [2] gives us $\epsilon_n^2 = \mathcal{O}\left(rG\exp\left(-\frac{\sqrt{n}}{rG}\right)^{2/3}\right)$ when $\epsilon_n^2 \geq \lambda rGJ_\pi^*$. Plugging each variable into Theorem 3 proves the theorem. Notice that condition of λ is replaced because $\{\epsilon_n^2 \geq \lambda rGJ_\pi^*\} \subset \{\epsilon_n^2 \geq 2\lambda J_\pi^*\}$ when $rG \geq 2$.

Remark 1. In the proof, we use discrete version of the covering number based on ℓ observations defined by

$$N_{\infty}(\epsilon, \mathcal{F}, \ell) = \sup_{\{\boldsymbol{X}\}_{i=1}^{\ell} \subset \mathbb{R}^{d_1 \times d_2}} N_{\infty}(\epsilon, \mathcal{F}, \{\boldsymbol{X}\}_{i=1}^{\ell}).$$

We can check the following relationship in [4].

$$N_{\infty}(\mathcal{F}, \epsilon, \ell) \le N_{\infty}(\mathcal{F}, \epsilon) \le N_{\infty}\left(\mathcal{F}, \epsilon - \frac{cMG\sqrt{r}}{\ell^{1/d_1d_2} - 1}, \ell\right), \text{ where } c > 0 \text{ is a constant}$$
 (2)

(The above upper bound makes sense only for $\ell > (cMG\sqrt{r}/\epsilon + 1)^{d_1d_2}$ i.e. for sufficiently large ℓ (2) holds). Therefore, the covering number $N_{\infty}(\epsilon, \mathcal{F})$ is almost equivalent to the covering number $N_{\infty}(\epsilon, \mathcal{F}, \ell)$.

2 Nonlinear case

Define

$$\mathcal{F}_r = \{ f : \mathbb{R}^{d_1 \times d_2} \to \mathbb{R} : f(\boldsymbol{X}) = \langle \boldsymbol{B}, h(\boldsymbol{X}) \rangle \text{ with } \boldsymbol{B} \in \mathcal{B} \},$$

where $h: \mathbb{R}^{d_1 \times d_2} \to \mathbb{R}^{d_1 \times d_2'}$ and $\mathcal{B} = \{\sum_{i=1}^r \lambda_i u_i v_i^T : u_i \in \mathbb{R}^{d_1}, v_i \in \mathbb{R}^{d_2'}, \|u_i\| = \|v_i\| = 1$, and $0 < \lambda_r \le \cdots \le \lambda_1 \le M\}$. The reason of changing the set \mathcal{B} from previous lecture note is that rank concept becomes ambiguous when $d_2' = \infty$. In this case, I am not sure that there is an unified way to calculate the covering number. Above all, we cannot use the covering number of linear class because Equation (2) becomes meaningless when extended feature space has infinity dimension. Therefore, the covering number of kernel case should be calculated individually.

References

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