Some Evidence Theory and Checking Algorithm

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4/18/2019

1 Evidence theory

Consider the semi-supervised binary tensor factorization, under the scenario where treat covariate as predictors. The general model is:

$$U = logit(\mathbb{E}Y) = B \times_1 X$$

$$B = C \times_1 N_1 \times_2 N_2 \times_3 N_3$$

where Y is the given binary tensor, X is the given covariate. N_1, N_2, N_3 are the factor matrices, C is the core tensor. B is the intermediate tensor.

1.1 Conclusion 1: Each update has unique solution, and each update increases log-likelihood

To write in a general form, we have:

$$U = logit(\mathbb{E}Y) = C \times_1 X N_1 \times_2 N_2 \times_3 N_3$$

1.1.1 Multi-process of GLM

First, we prove that the update of factor matrices or core tensor is just the multi-process of regular GLM.

For updating N_3 when fix C, N_1, N_2 (update N_2 is the same) Since we have:

$$U_{(3)} = N_3 C_{(3)} [N_2 \otimes (XN_1)]^T$$

Where $U_{(3)}$, $C_{(3)}$ are the unfold matrices of tensor U,C through mode-3. Recalling the matrix form GLM, which is identical to this scenario: Consider

$$logit[E(Y))_{n*p}] = U_{n*p} = X_{n*R} \times \beta_{R*p}$$

Where $U_{n*p} = (u_1, ..., u_p), \ \beta_{R*p} = (\beta_1, ..., \beta_p)$

We have:

$$\begin{aligned} u_1^{N*1} &= X^{n*R} \times \beta_1^{R*1} \\ u_2^{N*1} &= X^{n*R} \times \beta_2^{R*1} \\ &\vdots \\ u_p^{N*1} &= X^{n*R} \times \beta_P^{R*1} \end{aligned}$$

Then we implement regular GLM on each equation (i.e. each column of matrix β). These are multiprocessing and mutually independent GLM processes.

For updating N_1 when fix C, N_2, N_3

Since we have:

$$U_{(1)} = X N_1 C_{(1)} [N_3 \otimes N_2]^T$$

Where $U_{(1)}$, $C_{(1)}$ are the unfold matrices of tensor U,C through mode-1.

Recalling the form on note **Semi-Supervised Binary Tensor Factorization on dnations data**:

$$U^{d_1 \times d_2 d_3} = Y^{d_1 p} W^{pr_1} G^{r_1 \times d_2 d_3}$$

we have already proved that through some vectorization of matrix, it can be written as regular GLM form :

$$Y = X\beta$$

where β is the vectorization of N_1 in this scenario.

Thus it's still GLM.

For updating C when fix N_1, N_2, N_3

Recalling the note **Unsupervised Binary Tensor Factorization**, we have already proved that through some vectorization of matrix, it can be written as regular GLM form:

$$Y = X\beta$$

where β is the vectorization of C in this scenario. Thus it's still GLM.

1.1.2 Conclusion 1 on GLM

Then we just need to show the conclusion is satisfied on the GLM, in this case logistic regression.

Consider the form:

$$logit(\mathbb{E}\tilde{Y}) = U = X\beta$$

Recalling the negative log-likehood (i.e. cross-entropy), we have:

$$Q = -\sum_{i=1}^{n} \left\{ \tilde{y}_i \log \left(\pi_i \right) + (1 - \tilde{y}_i) \log(1 - \pi_i) \right\}$$
$$\pi_i = \frac{\exp \left(\mathbf{x}_i' \boldsymbol{\beta} \right)}{1 + \exp \left(\mathbf{x}_i' \boldsymbol{\beta} \right)}$$

Where x_i are one observation. If our dimension is p, then $x_i, \beta \in \mathbb{R}^p$ Consider $\tilde{y}_i \in \{0, 1\}$, let

$$y_i = 2\tilde{y_i} - 1 \in \{-1, 1\}$$

Then:

$$Q = -\sum_{i=1}^{n} log \left\{ \frac{1}{1 + \exp(-y_i x_i' \beta)} \right\}$$
$$= \sum_{i=1}^{n} log \left\{ 1 + \exp(-y_i x_i' \beta) \right\}$$

Then we consider the optimization problem:

$$\beta^* = \operatorname*{arg\,min}_{\beta} Q \tag{1}$$

$$= \underset{\beta}{\operatorname{arg\,min}} \sum_{i=1}^{n} \log \left\{ 1 + \exp\left(-y_i x_i' \beta\right) \right\} \tag{2}$$

Since we have:

$$\frac{\partial Q}{\partial \beta} = -\sum_{i=1}^{n} \frac{y_i}{1 + \exp(y_i x_i' \beta)} x_i$$
$$\frac{\partial^2 Q}{\partial \beta^2} = \sum_{i=1}^{n} \left[\frac{y_i}{1 + \exp(y_i x_i' \beta)} \right]^2 x_i x_i^T$$

Since for $\forall z \in \mathbb{R}^p$, we have:

$$z^{T} \frac{\partial^{2} Q}{\partial \beta^{2}} z = \sum_{i=1}^{n} \left[\frac{y_{i}}{1 + \exp(y_{i} x_{i}' \beta)} \right]^{2} (z^{T} x_{i})^{2} \ge 0$$

The equal sign holds if and only if z=0. Then we conclude that the Hessian matrix is positive definite.

According to:

$$\nabla^2 Q(\beta) > 0$$
 for all $\beta \in \text{dom}(Q) = \mathbb{R}^p$

We can conclude Q is strictly convex with respect to β . Then the optimization problem (2) has unique solution.

Consider the update at t step:

$$\beta^{(t)} = \underset{\beta}{\operatorname{arg \, min}} \, Q(\beta^{(t-1)})$$

$$= \underset{\beta}{\operatorname{arg \, min}} \, \sum_{i=1}^{n} \log \left\{ 1 + \exp\left(-y_i x_i' \beta^{(t-1)}\right) \right\}$$

According to what we derived, each update has unique solution. And each update won't decrease the log-likelihood. If $\beta^{(t-1)}$ reach the global minimum of opbejctive function, the update will remain the same: $\beta^{(t)} = \beta^{(t-1)}$.

1.2 Conclusion 2: The log-likelihood is invariant to orthogonalization

Suppose we have:

$$U = logit(\mathbb{E}Y) = B \times_1 X$$

$$B = C \times_1 N_1 \times_2 N_2 \times_3 N_3$$

where Y is the given binary tensor, X is the given covariate. N_1, N_2, N_3 are the factor matrices, C is the core tensor. B is the intermediate tensor.

When we conduct the orthogonalization and normalization on N_1 (N_2 , N_3 are the same), such as SVD:

$$N_1 = U\Sigma V^T$$

Take $\tilde{N}_1 = U$ as orthonormal version of N_1 , let $P = \Sigma V^T$, then:

$$N_1 = \tilde{N}_1 P$$

Let $C = \tilde{C}P^{-1}$ Then we have:

$$B = C \times_1 N_1 \times_2 N_2 \times_3 N_3$$

= $\tilde{C}P^{-1} \times_1 \tilde{N}_1 P \times_2 N_2 \times_3 N_3$
= $\tilde{C} \times_1 \tilde{N}_1 \times_2 N_2 \times_3 N_3$

Then the log-likelihood is the same.