## Graphic Lasso: Accuracy with intercept

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Consider the model

$$\Omega_k = \Theta_0 + \sum_{l=1}^r u_{kl} \Theta_l, \quad k \in [K].$$

Let  $U = [\![u_{kl}]\!] \in \mathbb{R}^{K \times r}$  be the membership matrix and  $u_l$  denote the l-th column of U. The optimization problem is stated as

$$\min_{\{U,\Theta\}} \quad \mathcal{L}(U,\Theta) = \sum_{k=1}^{K} \langle S^k, \Omega^k \rangle - \log \det(\Omega^k),$$

$$s.t. \quad \Omega^k = \Theta_0 + u_k \Theta_1, \quad k = 1, ..., K,$$

$$\|U\|_F = 1, \sum_{k=1}^{K} u_k = 0,$$

where  $\Theta_0$ ,  $\Theta_1$  are positive definite and  $\tau_1 < \min\{\varphi_{\min}(\Theta_0), \varphi_{\min}(\Theta_1)\} \le \max\{\varphi_{\max}(\Theta_0), \varphi_{\max}(\Theta_1)\} < \tau_2, \tau_1, \tau_2 > 0$ .

**Lemma 1** (Accuracy with intercept). Let  $\{U, \Theta_l\}$  denote the true parameter. Suppose the estimate  $\{\hat{U}, \hat{\Theta}_l\}$  satisfies  $\mathcal{L}(\hat{U}, \hat{\Theta}_l) \leq \mathcal{L}(U, \Theta_l)$ . Then, with probability tends to 1, we have the accuracy rate

$$\sum_{k=1}^{K} \|\Delta_k\|_F \le KC\sqrt{\frac{\log p}{n}}.$$

**Remark 1.** Note that by matrix inverse lemma, we have

$$\Sigma^{k} = (\Theta_{0} + u_{k}\Theta_{1})^{-1} = \Theta_{0}^{-1} + \frac{u_{k}}{1 + u_{k}\langle\Theta_{0}^{-1}, \Theta_{1}\rangle}\Theta_{0}^{-1}\Theta_{1}\Theta_{0}^{-1}.$$

Since  $\sum_{k=1}^{K} \frac{u_k}{1+u_k\langle\Theta_0^{-1},\Theta_1\rangle}$  may not equal to 0, the true covariance matrices do not share the same pattern and we can not use all sample covariance matrices  $S^k$  to improve the estimation error from K to  $\sqrt{K}$ .

*Proof.* Define the function

$$G(\hat{U}, \hat{\Theta}_l) = \sum_{k=1}^K \langle S^k, \hat{\Theta}_0 + \hat{u}_k \hat{\Theta}_1 - \Theta_0 - u_k \Theta_1 \rangle - \log \det(\hat{\Theta}_0 + \hat{u}_k \hat{\Theta}_1) + \log \det(\Theta_0 + u_k \Theta_1).$$

Let  $\Delta_k = \hat{\Theta}_0 + \hat{u}_k \hat{\Theta}_1 - \Theta_0 - u_k \Theta_1$ . By Taylor Expansion, we have

$$-\log \det(\hat{\Theta}_0 + \hat{u}_k \hat{\Theta}_1) + \log \det(\Theta_0 + u_k \Theta_1) \ge -\langle (\Theta_0 + u_k \Theta_1)^{-1}, \Delta_k \rangle + \frac{1}{4\tau_2^2} \|\Delta_k\|_F^2.$$

Let  $\Sigma^k = (\Theta_0 + u_k \Theta_1)^{-1}$  denote the true precision matrix. Then, we have

$$G(\hat{U}, \hat{\Theta}_l) \ge \sum_{k=1}^K \langle S^k - \Sigma^k, \Delta_k \rangle + \frac{1}{4\tau_2^2} \|\Delta_k\|_F^2 = I_1 + I_2.$$

By Lemma 1 in A.J. Rothman et al, we have

$$\max |S_{jk}^k - \Sigma^{jk}| \le C\sqrt{\frac{\log p}{n}},$$

with high probability. Therefore, we have

$$0 \ge \frac{1}{4\tau_2^2} \frac{1}{K} \left( \sum_{k=1}^K \|\Delta_k\|_F \right)^2 - C\sqrt{\frac{\log p}{n}} \sum_{k=1}^K \|\Delta_k\|_F,$$

which implies that

$$\sum_{k=1}^{K} \|\Delta_k\|_F \le KC\sqrt{\frac{\log p}{n}}.$$