# Gaussian Tensor Matching

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### 1 Problem Formulation and Model

Consider two random tensors  $\mathcal{A}, \mathcal{B}' \in \mathbb{R}^{d^{\otimes m}}$ , where  $\mathcal{A}(\omega)$  and  $\mathcal{B}'(\omega)$  denote the tensor entry indexed by  $\omega = (i_1, \ldots, i_m) \in [d]^m$ . Suppose  $\mathcal{A}$  and  $\mathcal{B}'$  are super-symmetric; i.e.,  $\mathcal{A}(\omega) = \mathcal{A}(f(\omega)), \mathcal{B}(\omega) = \mathcal{B}'(f(\omega))$  for any function f permutes the indices in  $\omega$  for all  $\omega \in [d]^m$ . Consider the bivariate generative model that for the entries  $\{\omega : 1 \leq i_1 \leq \cdots \leq i_m \leq d\}$ 

$$(\mathcal{A}(\omega), \mathcal{B}'(\omega)) \sim \mathcal{N}\left(0, \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix}\right), \text{ and } (\mathcal{A}(\omega), \mathcal{B}'(\omega)) \perp (\mathcal{A}(\omega'), \mathcal{B}'(\omega')), \text{ for all } \omega \neq \omega',$$

where the correlation  $\rho \in (0,1)$  and  $\perp$  denote the statistical independence. We call  $\mathcal{A}$  and  $\mathcal{B}'$  as two correlated Wigner tensors.

Suppose we observe the tensor pair  $\mathcal{A}$  and  $\mathcal{B} \stackrel{\text{def}}{=} \mathcal{B}' \circ \pi$ , where  $\pi : [d] \mapsto [d]$  denotes a permutation on [d], and by definition  $\mathcal{B}(i_1, \ldots, i_m) = \mathcal{B}'(\pi(i_1), \ldots, \pi(i_m))$  for all  $(i_1, \ldots, i_m) \in [d]^m$ .

This work aims to recover the true matching  $\pi$  given the noisy observations  $\mathcal{A}, \mathcal{B}$ .

# 2 Gaussian Tensor Matching

In this section, we develop the matching strategies for two correlated Wigner tensors.

### 2.1 Matching via Empirical Distributions

The main idea for correlated Wigner tensor matching is to use the empirical distribution of each slices and construct a distance statistics between the distributions. Specifically, for node  $i \in [d]$  and tensor A, we define the empirical distribution

$$\mu_i = \frac{1}{d^{m-1}} \sum_{(i_2, \dots, i_m) \in [d]^{m-1}} \delta_{\mathcal{A}_{i, i_2, \dots, i_m}},$$

where  $\delta_x$  is the point mass at x. Similarly, we define

$$\nu_k = \frac{1}{d^{m-1}} \sum_{(i_2, \dots, i_m) \in [d]^{m-1}} \delta_{\mathcal{B}_{k, i_2, \dots, i_m}}$$

with tensor  $\mathcal{B}$ . Intuitively, the distance between the empirical distributions  $\mu_i$  and  $\nu_k$  are small if i and k forms a true pair, and the distance is large, otherwise. Also, we define the empirical CDFs as

$$F_d^i(t) = \frac{1}{d^{m-1}} \sum_{(i_2, \dots, i_m) \in [d]^{m-1}} \mathbb{1} \{ \mathcal{A}_{i, i_2, \dots, i_m} \le t \}, \text{ and } G_d^k(t) = \frac{1}{d^{m-1}} \sum_{(i_2, \dots, i_m) \in [d]^{m-1}} \mathbb{1} \{ \mathcal{B}_{k, i_2, \dots, i_m} \le t \}.$$

We construct the distance statistic to measure the similarity between  $\mu_i$  and  $\nu_k$  as

$$d_p(\mu_i, \nu_k) = \left( \int_{\mathbb{R}} dt |F_d^i(t) - G_d^k(t)|^p \right)^{1/p},$$

for  $p \in [1, \infty)$ . Particularly, when p = 1,  $d_p(\mu_i, \nu_k)$  is equivalent to the 1-Wasserstein distance, where

$$d_1(\mu_i, \nu_k) = \sum_{j=1}^{d^{m-1}} |\operatorname{vec}(\mathcal{A}^i)_{(j)} - \operatorname{vec}(\mathcal{B}^k)_{(j)}|,$$
(1)

where  $\mathcal{A}^i$  denotes the *i*-th slice of  $\mathcal{A}$ ,  $\text{vec}(\mathcal{A}^i)_{(j)}$  denotes the *j*-th largest entry in the *i*-th slice of  $\mathcal{A}$ , and  $\mathcal{B}^k$ ,  $\text{vec}(\mathcal{B}^k)_{(j)}$  have similar definitions. Hence, we develop a Gaussian tensor matching algorithm with the distance statistics (1). See Algorithm 1.

Algorithm 1 Gaussian tensor matching via empirical distribution with 1-Wasserstein distance

Input: Gaussian tensors  $\mathcal{A}, \mathcal{B} \in \mathbb{R}^{d^{\otimes m}}$ .

- 1: Calculate the distance statistics  $d_1(\mu_i, \nu_k)$  in (1) for each pair of  $(i, k) \in [d] \times [d]$ .
- 2: Sort  $\{d_1(\mu_i, \nu_k) : (i, k) \in [d] \times [d]\}$  and let S be the set of indices of the smallest d elements.
- 3: if there exists a permutation  $\hat{\pi}$  such that  $S = \{(i, \hat{\pi}(i)) : i \in [d]\}$  then
- 4: Output  $\hat{\pi}_1$  and  $\hat{\pi}_2$
- 5: else
- 6: Output error.
- 7: end if

**Output:** Estimated permutations  $\hat{\pi}$  or error.

The theoretical guarantee for the success of Algorithm 1 is below.

**Theorem 2.1** (Guarantee of Algorithm 1). Let  $\rho = \sqrt{1 - \sigma^2}$ . Suppose  $\sigma \leq \frac{c}{\log d}$  for sufficiently small constant c. Algorithm 1 recover the true permutation  $\pi$  with probability tends to 1.

#### 2.2 Improvement with Seeded Matching

Previous strategy in Section 2.1 aims to find the matching in one shot. In this section, we improve the Algorithm 1 by seeded matching. The seeded matching includes two steps: (1) use Algorithm 1 to find the seeds with enough true pairs; (2) apply a seeded bipartite matching with the seeds.

For the first step, we consider the seeds consisting of high-degree pairs. We define the slice sum

$$a_i = \frac{1}{\sqrt{d^{m-1}}} \sum_{\omega \in [d]^{\otimes m-1}} \mathcal{A}_{i,\omega}, \quad b_k = \frac{1}{\sqrt{d^{m-1}}} \sum_{\omega \in [d]^{\otimes m-1}} \mathcal{B}_{k,\omega},$$

where  $a_i$  and  $b_k$  are considered as the counterparts of "degrees" for Gaussian tensors. By Ding et al. (2021), we have

$$\mathbb{P}(a_i \geq \xi, b_k \geq \xi) = \begin{cases} Q(\xi)^2 & \text{if } (i, k) \text{ is a fake pair} \\ Q(\xi) \exp(-C\sigma^2 \xi^2) & \text{if } (i, k) \text{ is a true pair}, \end{cases}$$

where C is a positive constant,  $Q(\cdot)$  is the complementary CDF of standard normal distribution, and  $\xi$  serves as the threshold for high-degree. Consider the high-degree set

$$S = \{(i, k) \in [d]^2 : a_i, b_k \ge \xi, d_1(\mu_i, \nu_k) \le \zeta\}.$$
(2)

with given thresholds  $\xi$  and  $\zeta$ . Suppose we need s seeds for bipartite matching success. We need

(1) S has enough true pairs, i.e.,

$$dQ(\xi) \exp(-C\sigma^2 \xi^2) \ge s.$$

(2) no fake pairs involved in S, i.e,

$$d^{2}Q(\xi)^{2} \exp(-C\sigma^{-1}) = o(1),$$

where C is a positive constant and the term  $\exp(-C\sigma^{-1})$  follows from the property of distance statistics  $d_p(\mu_i, \nu_k)$  for the fake pair (i, k) in (3), and  $d^2$  is the order of total pairs.

Choose the threshold  $\xi = \mathcal{O}(\sqrt{s})$ . We have  $Q(\xi) = \Omega(s/d)\mathcal{O}(\exp \sigma^2 s)$  by (1), and with (2) we obtain

$$\sigma \le \frac{c}{s^{1/3}},$$

for some constant c. Hence, if the number of true pairs in the seed s is smaller than  $\log^3 d$ , we release the condition from  $\sigma \leq c/\log d$  to  $\sigma \leq c/s^{1/3}$ .

For the second step, the main idea for seeded matching is to use the prior information in seed to describe distance between the unseeded pairs. Let  $\pi_0: S \mapsto T$  denotes the seeds, where  $S, T \subset [n]$  and  $\pi_0(j) = \pi(j)$  for all  $j \in S$ . Define the sets

$$\mathcal{N} = \{(i_2, \dots, i_m) : i_l \in S, \text{ for all } l = 2, \dots, m\}$$

with  $|\mathcal{N}| = |S|^{m-1}$ , and define  $\pi_0(\mathcal{N})$  by replacing  $i_l$  to  $\pi_0(i_l)$  in the definition of  $\mathcal{N}$  for all  $l = 2, \ldots, m$ . Then, we define the distance for the unseeded pairs (i, k) as

$$H_{p,ik} = \left( \int_{\mathbb{R}} dt \left| \frac{1}{|\mathcal{N}|} \sum_{\omega \in \mathcal{N}} \mathbb{1} \{ \mathcal{A}_{i,\omega} \le t \} - \frac{1}{|\pi_0(\mathcal{N})|} \sum_{\omega \in \pi_0(\mathcal{N})} \mathbb{1} \{ \mathcal{B}_{k,\omega} \le t \} \right|^p \right)^{1/p},$$

for some  $p \geq 1$ . Intuitively, the term  $\frac{1}{|\mathcal{N}|} \sum_{\omega \in \mathcal{N}} \delta_{\mathcal{A}_{i,\omega}}$  describes the empirical distribution of edges in  $\mathcal{A}$  related to the unseeded node i and the seeded nodes. The term  $H_{i,k}$  indicates the difference between the empirical distributions related to the seeds with unseeded nodes i and k in  $\mathcal{A}$  and  $\mathcal{B}$ , respectively.

See the improved matching strategy in Algorithm 2 with seeded matching as subroutine in Algorithm 3.

The theoretical guarantee for Algorithm 2 is below.

## Algorithm 2 Gaussian tensor matching with seed improvement

Input: Gaussian tensors  $\mathcal{A}, \mathcal{B} \in \mathbb{R}^{d^{\otimes m}}$ , threshold  $\xi, \zeta$ .

- 1: Calculate the distance statistics  $d_1(\mu_i, \nu_k)$  in (1) for each pair of  $(i, k) \in [d] \times [d]$ .
- 2: Obtain the high-degree set S in (2).
- 3: if there exists a permutation  $\pi_0$  such that  $S = \{(i, \pi_0(i)) : i \in [d]\}$  then
- 4: Run bipartite Algorithm with seed  $\pi_0$  and output  $\hat{\pi}$
- 5: **else**
- 6: Output error.
- 7: end if

**Output:** Estimated permutations  $\hat{\pi}$  or error.

## Algorithm 3 Seeded Gaussian tensor matching

Input: Gaussian tensors  $\mathcal{A}, \mathcal{B} \in \mathbb{R}^{d^{\otimes m}}$ , seed  $\pi_0 : S \mapsto T$ .

- 1: For  $i \in S^c$  and  $k \in T^c$ , obtain the distance  $H_{1,ik}$ .
- 2: Find the optimal bipartite permutation  $\pi'_1$  such that

$$\pi_1' = \operatorname*{arg\,min}_{\pi} \sum_{i \in S^c} H_{i\pi(i)}.$$

Let  $\hat{\pi}$  denote the matching on [d] such that  $\hat{\pi}|_S = \pi_0$  and  $\hat{\pi}|_{S^c} = \pi'_1$ .

- 3: if  $\hat{\pi}$  is a perfect matching on [d] such that  $\hat{\pi}$  is an one-to-one function from [d] to [d]. then
- 4: Output  $\hat{\pi}$ .
- 5: **else**
- 6: Output error.
- 7: end if

**Output:** Estimated permutations  $\hat{\pi}$  or error.

**Theorem 2.2** (Conjecture: guarantee for Algorithm 2). Let  $\rho = \sqrt{1 - \sigma^2}$ . Suppose  $\sigma \leq \frac{c}{\log^{1/3(m-1)}d}$  for sufficiently small constant c. Algorithm 1 recover the true permutation  $\pi$  with probability tends to 1.

Remark 1 (From matrix matching to tensor matching). The improvement of tensor matching with increasing order m is mainly indicated in the seeded algorithm. Intuitively, in tensor cases, we need less seeds to obtain the description of the unseeded pairs with the same accuracy, which results in a looser upper bound of  $\sigma$ . Note that a larger  $\sigma$  indicates a smaller correlation between two tensors and thereof a weaker "signal" in the matching problem. Therefore, we allow a weaker signal assumption  $\sigma = \mathcal{O}(\frac{1}{\log^{1/3(m-1)}d})$  as m increases.

## 3 Proof Sketches

Proof Sketch of Theorem 2.1. Without loss of generality, we assume the true permutation  $\pi$  is the identity mapping; i.e.,  $\pi(i) = i$  for all  $i \in [d]$ . For simplicity, let  $d_{ik}$  denote the distance statistics  $d_1(\mu_i, \nu_j)$  in (1). To guarantee the Algorithm 1 outputs the true permutation with probability, it suffices to show

$$\min_{i \neq k \in [d]} d_{ik} > \max_{i \in [d]} d_{ii}$$

with probability tends to 1.

According to Ding et al. (2021), for all  $i \in [d]$  we have

$$\mathbb{P}\left(d_{ii} \geq \sqrt{\frac{\sigma}{d^{m-1}}}\right) \leq \exp\left(-\frac{C_1}{\sigma}\right), \text{(needs to be verified)}$$

and for all  $i \neq k \in [d]$ 

$$\mathbb{P}\left(d_{ik} \le \sqrt{\frac{\sigma}{d^{m-1}}}\right) \le \exp\left(-\frac{C_2}{\sigma}\right). \tag{3}$$

Hence, we have

$$\mathbb{P}\left(\max_{i\in[d]}d_{ii}<\sqrt{\frac{\sigma}{d^{m-1}}}\right)\geq [1-\exp(-C_1/\sigma)]^d,$$

and

$$\mathbb{P}\left(\min_{i\neq k\in[d]}d_{ik} > \sqrt{\frac{\sigma}{d^{m-1}}}\right) \ge [1 - d\exp(-C_2/\sigma)]^d,$$

where  $C_1, C_2$  are two positive constants. Take  $\sigma \leq \frac{c}{\log d}$  for sufficiently small c. Then, we have

$$\min_{i\neq k\in[d]}d_{ik}>\sqrt{\frac{\sigma}{d^{m-1}}}>\max_{i\in[d]}d_{ii},$$

with probability  $\mathcal{O}([1-1/d^2]^d)$  that tends to 1 as  $d\to\infty$ .

4 Numerical Experiments

## References

Ding, J., Ma, Z., Wu, Y., and Xu, J. (2021). Efficient random graph matching via degree profiles. *Probability Theory and Related Fields*, 179(1):29–115.