## Graphic Lasso: Miscellaneous

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## 1 Corrected proof for sufficient condition

Consider the

$$\mathbb{E}[\mathcal{Y}] = f(\Theta), \text{ where } \Theta = \mathcal{C} \times_1 \mathbf{M}_1 \times_2 \cdots \times_K \mathbf{M}_K,$$

and the optimization problem

$$\max_{\Theta = \mathcal{C} \times_1 \mathbf{M}_1 \times_2 \dots \times_K \mathbf{M}_K} \mathcal{L}_{\mathcal{Y}}(\Theta) = \langle \mathcal{Y}, \Theta \rangle - \sum_{(i_1, \dots, i_K)} g(\Theta_{i_1, \dots, i_K}). \tag{1}$$

**Theorem 1.1** (Sufficient condition). Let  $\{C, M_k\}$  denote the true parameters and  $\{\hat{C}, \hat{M}_k\}$  denote the maximizer of the objective function (1). The minimal sufficient conditions to obtain the clustering accuracy in form of

$$\mathbb{P}(MCR(\hat{\boldsymbol{M}}_k, \boldsymbol{M}_k) \geq \epsilon) \leq p(\epsilon, \delta), \quad where \quad p(\epsilon, \delta) \to 0, \quad as \quad \epsilon \to 1$$

include

- 1. The function g is convex,  $\sup_{x=f(c_{r_1,\ldots,r_K})} (g')^{-1}(x) \leq p(\mathcal{C})$ , where  $p(\cdot)$  is a function of the true parameter  $\mathcal{C}$ , and  $\sup_x g''(x) \leq a$ , where a is a positive constant.
- 2. The minimal gap between blocks is strictly larger than 0, i.e.,  $\delta = \min_k \delta^{(k)} > 0$  , where

$$\delta^{(k)} = \min_{r_k \neq r'_k} \max_{r_1, \dots, r_{k-1}, r_{k+1}, \dots, r_K} (f(c_{r_1, \dots, r_k, \dots, r_K}) - f(c_{r_1, \dots, r'_k, \dots, r_K}))^2.$$

3. The observation satisfies the assumptions for Hoeffding's inequality, i.e., each entry of  $\mathcal{Y}$  is bounded in [a,b] or sub-Gaussian with parameter  $\sigma$ .

*Proof.* We proof the sufficiency in following steps:

1. With given membership matrix  $\hat{M}_k$ , the estimate to  $\hat{C}$  is

$$\hat{c}_{r_1,...,r_K}(\hat{M}_k) = (g')^{-1} \left( \frac{1}{\prod_k d_k \prod_k \hat{p}_{r_k}^{(k)}} [\mathcal{Y} \times_1 \hat{M}_1^T \times_2 \cdots \times_K M_K^T]_{r_1,...,r_K} \right).$$

The estimate is unique since g is convex.

2. We define following useful functions. First, let  $F(\hat{M}_k) = \mathcal{L}_{\mathcal{Y}}(\hat{\mathcal{C}}, \hat{M}_k)$ , where  $\hat{\mathcal{C}} = \hat{\mathcal{C}}(\hat{M}_k)$  is the estimate depends on  $\hat{M}_k$ . We have

$$F(\hat{M}_{k}) = \langle \mathcal{Y} \times_{1} \hat{M}_{1}^{T} \times_{2} \cdots \times_{K} M_{K}^{T}, \hat{\mathcal{C}} \rangle - \sum_{i_{1}, \dots, i_{K}} g([\hat{\mathcal{C}} \times_{1} \hat{M}_{1} \times_{2} \cdots \times_{K} M_{K}]_{i_{1}, \dots, i_{K}})$$

$$= \sum_{r_{1}, \dots, r_{K}} \prod_{k} d_{k} \prod_{k} \hat{p}_{r_{k}}^{(k)} [g'(\hat{c}_{r_{1}, \dots, r_{K}}) \hat{c}_{r_{1}, \dots, r_{K}} - g(\hat{c}_{r_{1}, \dots, r_{K}})]$$

$$= \sum_{r_{1}, \dots, r_{K}} \prod_{k} d_{k} \prod_{k} \hat{p}_{r_{k}}^{(k)} h(g'(\hat{c}_{r_{1}, \dots, r_{K}})),$$

where  $h(x) = x(g')^{-1}(x) - g((g')^{-1}(x))$ . Correspondingly, we define

$$G(\hat{M}_k) = \sum_{r_1, \dots, r_K} \prod_k d_k \prod_k \hat{p}_{r_k}^{(k)} h(\mathbb{E}[g'(\hat{c}_{r_1, \dots, r_K})]),$$

where 
$$\mathbb{E}[g'(\hat{c}_{r_1,\dots,r_K})] = \frac{1}{\prod_k \hat{p}_{r_k}^{(k)}} [f(\mathcal{C}) \times_1 D_1^T \times_2 \dots \times_K D_K^T].$$

Then, for the true parameters  $\{C, M_k\}$ , we have

$$F(\boldsymbol{M}_k) = \mathcal{L}_{\mathcal{Y}}(\hat{\mathcal{C}}(\boldsymbol{M}_k), \boldsymbol{M}_k) = \sum_{r_1, \dots, r_K} \prod_k d_k \prod_k p_{r_k}^{(k)} h(g'(\hat{c}_{r_1, \dots, r_K})),$$

and

$$G(\mathbf{M}_k) = \sum_{r_1, \dots, r_K} \prod_k d_k \prod_k p_{r_k}^{(k)} h(\mathbb{E}[g'(\hat{c}_{r_1, \dots, r_K})]) = \sum_{r_1, \dots, r_K} \prod_k d_k \prod_k p_{r_k}^{(k)} h(f(c_{r_1, \dots, r_K})).$$

3. Consider the difference between  $F(M'_k)$  and  $G(M'_k)$ . Since g is convex and  $h''(x) = \frac{1}{g''((g')^{-1}(x))} > 0$ , then the function h is convex and thus h is local Lipschitz. Note that  $h'(x) = (g')^{-1}(x)$ . Therefore, we have

$$|F(M'_k) - G(M'_k)| \le p(\mathcal{C}) \|g'(\hat{c}_{r_1,\dots,r_K}) - \mathbb{E}[g'(\hat{c}_{r_1,\dots,r_K})]\|_{\max}.$$
 (2)

4. Consider the misclassification error. With assumption 1,2, we satisfy the condition for Lemma 1. Therefore, we have

$$G(\hat{\mathbf{M}}_k) - G(\mathbf{M}_k) \le -\frac{\epsilon}{4a} \tau^{K-1} \delta. \tag{3}$$

5. Combining step (2) with step (3), we obtain the accuracy

$$\mathbb{P}(MCR(\hat{\boldsymbol{M}}_{k}, \boldsymbol{M}_{k}) \geq \epsilon) \leq \mathbb{P}\left(\sup_{\{\boldsymbol{M}_{k}\}} \left\| g'(\hat{c}_{r_{1},\dots,r_{K}}) - \mathbb{E}[g'(\hat{c}_{r_{1},\dots,r_{K}})] \right\|_{\max} \geq \frac{\epsilon}{8ap(\mathcal{C})} \tau^{K-1} \delta\right) \\
\leq \mathbb{P}\left(\sup_{I_{r_{1},\dots,r_{K}}} \frac{\sum_{(i_{1},\dots,i_{K}) \in I_{r_{1},\dots,r_{K}}} \mathcal{Y}_{i_{1},\dots,i_{K}} - \mathbb{E}[\mathcal{Y}_{i_{1},\dots,i_{K}}]}{|I_{r_{1},\dots,r_{K}}|} \geq \frac{\epsilon}{8ap(\mathcal{C})} \tau^{K-1} \delta\right) \\
\leq 2^{1+\sum d_{k}} \exp\left(-\frac{\epsilon^{2} \tau^{2K-2} \delta^{2} L}{C\sigma^{2} ap(\mathcal{C})^{2}}\right).$$

**Remark 1.** Note that the proof does not utilize the self-consistency property. For misclassification error, we have

$$G(\hat{M}_k) = \sum_{r_1, \dots, r_K} \prod_k d_k \prod_k \hat{p}_{r_k}^{(k)} h\left(\frac{1}{\prod_k \hat{p}_{r_k}^{(k)}} [f(\mathcal{C}) \times_1 D_1^T \times_2 \dots \times_K D_K^T]_{r_1, \dots, r_K}\right),$$

and

$$G(\mathbf{M}_k) = \sum_{r_1, \dots, r_K} \prod_k d_k \prod_k p_{r_k}^{(k)} h(f(c_{r_1, \dots, r_K}))$$

$$= \sum_{r_1, \dots, r_K} \prod_k d_k \prod_k \hat{p}_{r_k}^{(k)} \frac{1}{\prod \hat{p}_{r_k}^{(k)}} [h(f(\mathcal{C})) \times_1 D_1^T \times_2 \dots \times_K D_K^T]_{r_1, \dots, r_K}.$$

The true parameter  $\{M_k\}$  is the maximizer of  $G(M_k)$  because of the convexity of h. The linearity of  $g'(\hat{c}_{r_1,\dots,r_K})$  is also crucial to take the advantage of Jensen's inequality.

**Lemma 1.** Suppose minimal gap between blocks is strictly larger than 0, i.e.,  $\delta = \min_k \delta^{(k)} > 0$ , and  $h''(x) \geq \frac{1}{a}$ . For an fixed  $\epsilon > 0$ , suppose  $MCR(\hat{M}_k, M_k) \geq \epsilon$  for some  $k \in [K]$ . We have

$$G(\hat{\mathbf{M}}_k) - G(\mathbf{M}_k) \le -\frac{\epsilon}{4a} \tau^{K-1} \delta.$$

Proof. We provide the proof for k = 1. The proof for other  $k \in [K]$  is similar. Since  $MCR(\hat{M}_1, M_1) \ge \epsilon$ , there exist some  $r_1 \in [R_1]$  and  $a_1 \ne a'_1$  such that  $\min\{D_{a_1,r_1}^{(1)}, D_{a'_1,r_1}^{(1)}\} \ge \epsilon$ . Let  $\mathcal{N} = [h(g'(c_{r_1,\ldots,r_K}))]$  and  $W = \prod_k \hat{p}_{r_k}^{(k)}$ . Then, there exists  $c^*$  such that

$$\begin{split} & [\mathcal{N} \times_{1} \boldsymbol{D}^{(1),T} \times_{2} \cdots \times_{K} \boldsymbol{D}^{(K),T}]_{r_{1},\dots,r_{K}} \\ & = D_{a_{1},r_{1}}^{(1)} D_{a_{2},r_{2}}^{(2)} \cdots D_{a_{K},r_{K}}^{(K)} h(g'(c_{a_{1},\dots,a_{K}})) + D_{a_{1}',r_{1}}^{(1)} D_{a_{2},r_{2}}^{(2)} \cdots D_{a_{K},r_{K}}^{(K)} h(g'(c_{a_{1}',\dots,a_{K}})) \\ & + (W - D_{a_{1},r_{1}}^{(1)} D_{a_{2},r_{2}}^{(2)} \cdots D_{a_{K},r_{K}}^{(K)} - D_{a_{1}',r_{1}}^{(1)} D_{a_{2},r_{2}}^{(2)} \cdots D_{a_{K},r_{K}}^{(K)}) c^{*}. \end{split}$$

Define  $\mu_{r_1,\dots,r_K} = \frac{1}{\prod_k \hat{p}_{r_k}^{(k)}} [f(\mathcal{C}) \times_1 D_1^T \times_2 \dots \times_K D_K^T]$ . Then, by Taylor Expansion of function  $h(\cdot)$  at the point  $\mu_{r_1,\dots,r_K}$ , we have

$$\frac{1}{W} [\mathcal{N} \times_{1} \mathbf{D}^{(1),T} \times_{2} \cdots \times_{K} \mathbf{D}^{(K),T}]_{r_{1},\dots,r_{K}} - h(\mu_{r_{1},\dots,r_{K}})$$

$$\geq \frac{1}{2W} D_{a_{1},r_{1}}^{(1)} D_{a_{2},r_{2}}^{(2)} \cdots D_{a_{K},r_{K}}^{(K)} h''(\mu_{r_{1},\dots,r_{K}}) (g'(c_{a_{1},\dots,a_{K}}) - \mu_{r_{1},\dots,r_{K}})^{2}$$

$$+ \frac{1}{2W} D_{a_{1},r_{1}}^{(1)} D_{a_{2},r_{2}}^{(2)} \cdots D_{a_{K},r_{K}}^{(K)} h''(\mu_{r_{1},\dots,r_{K}}) (g'(c_{a'_{1},\dots,a_{K}}) - \mu_{r_{1},\dots,r_{K}})^{2}$$

$$+ \frac{1}{2W} (W - D_{a_{1},r_{1}}^{(1)} D_{a_{2},r_{2}}^{(2)} \cdots D_{a_{K},r_{K}}^{(K)} - D_{a'_{1},r_{1}}^{(1)} D_{a_{2},r_{2}}^{(2)} \cdots D_{a_{K},r_{K}}^{(K)}) h''(\mu_{r_{1},\dots,r_{K}}) (c^{*} - \mu_{r_{1},\dots,r_{K}})^{2},$$

where  $h''(x) = \frac{1}{a''(a',-1(x))} \frac{1}{a}$ . By the inequality  $a^2 + b^2 \ge \frac{(a+b)^2}{2}$ , we obtain that

$$\frac{1}{W} [\mathcal{N} \times_{1} \mathbf{D}^{(1),T} \times_{2} \cdots \times_{K} \mathbf{D}^{(K),T}]_{r_{1},\dots,r_{K}} - h(\mu_{r_{1},\dots,r_{K}})$$

$$\geq \frac{1}{a4W} \min\{D_{a_{1},r_{1}}^{(1)}, D_{a'_{1},r_{1}}^{(1)}\} D_{a_{2},r_{2}}^{(2)} \cdots D_{a_{K},r_{K}}^{(K)} (g'(c_{a_{1},\dots,a_{K}}) - g'(c_{a'_{1},\dots,a_{K}}))^{2}. \tag{4}$$

Noted  $h(\cdot)$  is a convex function, for other  $r'_1 \in [R_1]/\{r_1\}$ , by Jensen's inequality, we have

$$\frac{1}{W} [\mathcal{N} \times_1 \mathbf{D}^{(1),T} \times_2 \dots \times_K \mathbf{D}^{(K),T}]_{r'_1,\dots,r_K} - h(\mu_{r'_1,\dots,r_K}) \ge 0.$$
 (5)

Combing the inequality (4) and (5), we obtain that

$$G(\hat{\mathbf{M}}_k) - G(\mathbf{M}_k) \le -\frac{\epsilon}{4a} \tau^{K-1} \delta,$$

where the inequality follows by the fact that  $\sum_{r_k} D_{a_k r_k}^{(k)} = p_{a_k}^{(k)} \ge \tau$ .

## 2 General Loss Function

Consider the model

$$\mathbb{E}[\mathcal{Y}] = f(\Theta), \text{ where } \Theta = \mathcal{C} \times_1 \mathbf{M}_1 \times_2 \cdots \times_K \mathbf{M}_K.$$

**Theorem 2.1** (General property for loss function to guarantee the clustering accuracy). Let  $\{C, M_k\}$  denote the true parameters, and  $\mathcal{L}_{\mathcal{Y}}(C', M_k')$  denote the sample-based loss function to estimate  $\{C, M_k\}$ . Define the population-based loss function as

$$l(\mathcal{C}', \mathbf{M}'_k) = \mathbb{E}_{\mathcal{V}}[\mathcal{L}_{\mathcal{V}}(\mathcal{C}', \mathbf{M}'_k)].$$

For all  $\{C', M'_k\}$  in the parameter space, suppose the sample-based and population based satisfies the following properties

1. (Self-consistency) Suppose  $MCR(\mathbf{M}'_k, \mathbf{M}_k) \geq \epsilon$  for  $\epsilon > 0$ . We have

$$l(C', M'_k) - l(C, M_k) \le -C(\epsilon),$$
 (6)

where  $C(\cdot)$  is the function of  $\epsilon$  which takes positive value.

2. (Bounded difference between sample- and population-based loss) The difference between sample-based and population-based loss function is bounded in probability, i.e.,

$$p(t) = \mathbb{P}(|\mathcal{L}_{\mathcal{Y}}(\mathcal{C}', \mathbf{M}'_k) - l(\mathcal{C}', \mathbf{M}'_k)| \ge t) \to 0, \quad as \quad t \to \infty.$$
 (7)

Let  $\{\hat{\mathcal{C}}, \hat{M}_k\}$  denote the maximizer of the  $\mathcal{L}_{\mathcal{Y}}$ . Then, we obtain the clustering accuracy, for any  $\epsilon > 0$ ,

$$\mathbb{P}(MCR(\hat{M}_k, M_k) \ge \epsilon) \le p\left(\frac{C(\epsilon)}{2}\right).$$

*Proof.* Since  $\{\hat{\mathcal{C}}, \hat{M}_k\}$  is the maximizer of the population-based objective function  $\mathcal{L}_{\mathcal{Y}}$ , we have

$$0 \leq \mathcal{L}_{\mathcal{Y}}(\hat{\mathcal{C}}, \hat{\mathbf{M}}_k) - \mathcal{L}_{\mathcal{Y}}(\mathcal{C}, \mathbf{M}_k)$$
  
=  $\mathcal{L}_{\mathcal{Y}}(\hat{\mathcal{C}}, \hat{\mathbf{M}}_k) - l(\hat{\mathcal{C}}, \hat{\mathbf{M}}_k) + l(\hat{\mathcal{C}}, \hat{\mathbf{M}}_k) - l(\mathcal{C}, \mathbf{M}_k) + l(\mathcal{C}, \mathbf{M}_k) - \mathcal{L}_{\mathcal{Y}}(\mathcal{C}, \mathbf{M}_k).$ 

Suppose  $MCR(\hat{M}_k, M_k) \ge \epsilon$ . By the property (6), we have

$$0 \le 2r - C(\epsilon),$$

where  $r = \sup_{\mathcal{C}', \mathbf{M}_k'} |\mathcal{L}_{\mathcal{Y}}(\mathcal{C}', \mathbf{M}_k') - l(\mathcal{C}', \mathbf{M}_k')|$ . Therefore, we have

$$\begin{split} \mathbb{P}(MCR(\hat{\boldsymbol{M}}_k, \boldsymbol{M}_k) \geq \epsilon) &= \mathbb{P}(l(\hat{\mathcal{C}}, \hat{\boldsymbol{M}}_k) - l(\mathcal{C}, \boldsymbol{M}_k) \leq -C(\epsilon)) \\ &\leq \mathbb{P}(C(\epsilon) \leq 2r) \\ &= p\left(\frac{C(\epsilon)}{2}\right), \end{split}$$

where the last equation follows the second property (7).