## New distance and its tail bound

Jiaxin Hu

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### 1 New distance and its tail bound

#### 1.1 Definitions

Suppose that we have i.i.d. samples  $(X_1, Y_1), \ldots, (X_n, Y_n)$  following the multivariate zero-mean Gaussian distribution with variance 1 and correlation  $\rho \in [0, 1)$ ; i.e,

$$(X_i, Y_i) \sim \mathcal{N}\left(\mathbf{0}, \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix}\right), \text{ and } (X_i, Y_i) \perp (X_j, Y_j), \text{ for all } i \neq j.$$
 (1)

Define the L-distance for the empirical distributions as

$$d_L = \sum_{l \in [L]} |F_n(I_l) - G_n(I_l)|,$$

where L is a positive integer,  $I_l = \left[ -\frac{1}{2} + \frac{l-1}{L}, -\frac{1}{2} + \frac{l}{L} \right]$  for all  $l \in [L]$  are the uniform partition of [-1/2, 1/2], and

$$F_n(I_l) = \frac{1}{n} \sum_{i \in [n]} \mathbb{1}\{X_i \le I_l\}, \quad G_n(I_l) = \frac{1}{n} \sum_{i \in [n]} \mathbb{1}\{Y_i \le I_l\}$$

are empirical distributions for X and Y, respectively.

**Remark 1.** Note that the distance  $d_L$  is the direct analogy of Ding's distance  $Z_{ik}$  (equation (27) in Ding et al. (2021)) for the Bernoulli case.

#### 1.2 Tail bounds

**Lemma 1** (Large deviation of *L*-distance with true pairs). Suppose we have i.i.d. samples  $(X_1, Y_1), \ldots, (X_n, Y_n)$  from the model (1). Let  $\sigma = \sqrt{1 - \rho^2}$ . We have, for all t > 0

$$\mathbb{P}\left(d_L \ge 2L\sqrt{\frac{\sigma}{n}} + c_1\sqrt{\frac{t}{n}}\right) \le e^{-t},$$

where  $c_1$  is an absolute constant.

**Lemma 2** (Small deviation of *L*-distance with fake pairs). Suppose we have i.i.d. samples  $(X_1, Y_1), \ldots, (X_n, Y_n)$  from the model (1) with  $\rho = 0$ . Let  $\sigma = \sqrt{1 - \rho^2}$ . Assume  $L \ge L_0$  for sufficiently large constant  $L_0$  and  $L = \mathcal{O}(n)$ . We have, for all t > 0

$$\mathbb{P}\left(d_L \le c_2 \sqrt{\frac{L}{n}} - c_3 \sqrt{\frac{t}{n}}\right) \le e^{-t},$$

where  $c_2, c_3$  are absolute constants.

**Remark 2** (Guarantee for Algorithm 1). Let  $d_{ik,L}$  denote the L-distance for the pair (i,k), and  $\pi^*$  be the identity mapping. Take  $\sigma \leq \sigma_0/\log n, L = L_0\log n$  such that  $\sqrt{\sigma_0 L_0} \leq c_2/4$ . Let  $\xi_{\rm true} = 2L\sqrt{\frac{\sigma}{n}} + c_1\sqrt{\frac{t}{n}}$  and  $\xi_{\rm fake} = c_2\sqrt{\frac{L}{n}} - c_3\sqrt{\frac{t}{n}}$ . Take

$$t = \left(\frac{c_2 - \sqrt{\sigma_0 L_0}}{(c_1 + c_3)}\right)^2 L.$$

Then, we will have  $\xi_{\text{fake}} \geq \xi_{\text{true}}$ , and

$$\mathbb{P}(d_{ii,L} \ge \xi_{\text{true}}) \le C_1 \exp(-\log n)$$
, and  $\mathbb{P}(d_{ik,L} \le \xi_{\text{fake}}) \le C_2 \exp(-\log n)$ ,

which will lead to our desired guarantee for Algorithm 1.

Proof of Lemma 1. Recall that  $\sigma = \sqrt{1-\rho^2}$  and  $I_l = \left[-\frac{1}{2} + \frac{l-1}{L}, -\frac{1}{2} + \frac{l}{L}\right]$  for all  $l \in [L]$ . Notice that for arbitrary  $l \in [L]$ 

$$\mathbb{P}(X_{1} \in I_{l}, Y_{1} \notin I_{l}) \leq \mathbb{P}(X_{1} \in I_{l}, Y_{1} > -\frac{1}{2} + \frac{l}{L}) + \mathbb{P}(X_{1} \in I_{l}, Y_{1} < -\frac{1}{2} + \frac{l-1}{L}) \\
\leq \mathbb{P}(X_{1} \leq -\frac{1}{2} + \frac{l}{L}, Y_{1} > -\frac{1}{2} + \frac{l}{L}) + \mathbb{P}(X_{1} \geq -\frac{1}{2} + \frac{l-1}{L}, Y_{1} < -\frac{1}{2} + \frac{l-1}{L}) \\
\leq 2 \sup_{t \in \mathbb{R}} \mathbb{P}(X_{1} \leq t, Y_{1} > t) \\
\leq 2\sigma.$$

where the third inequality follows from the fact that  $X_1, Y_1$  are identical distributed, and the last inequality follows from Proposition 1. By symmetry, we have

$$\mathbb{P}(X_1 \in I_l, Y_1 \notin I_l) + \mathbb{P}(X_1 \notin I_l, Y_1 \in I_l) < 4\sigma.$$

Take  $\nu, \nu'$  as standard Gaussian distributions. By Lemma 3, we have

$$\mathbb{P}\left(d_L \ge 2L\sqrt{\frac{\sigma}{n}} + c_1\sqrt{\frac{t}{n}}\right) \le e^{-t},$$

for all t > 0.

Proof of Lemma 2. Take  $\nu, \nu'$  as standard Gaussian distributions, and recall that  $I_l = \left[ -\frac{1}{2} + \frac{l-1}{L}, -\frac{1}{2} + \frac{l}{L} \right]$  for all  $l \in [L]$ . Notice that for arbitrary  $l \in [L]$ , we have  $|I_l| = 1/L$  and

$$\frac{1}{L\sqrt{2\pi}}e^{-1/8} \le \nu(I_l) = \frac{1}{\sqrt{2\pi}} \int_{I_l} \exp(-x^2/2) dx \le \frac{1}{L\sqrt{2\pi}}.$$

With the assumption that  $L \geq L_0$  and  $L = \mathcal{O}(n)$ , by Lemma 4, we have

$$\mathbb{P}\left(d_L \le c_2 \sqrt{\frac{L}{n}} - c_3 \sqrt{\frac{t}{n}}\right) \le e^{-t},$$

for all t > 0.

#### 1.3 Useful Lemmas for the proofs of Lemma 1 and 2.

**Lemma 3** (Lemma 7 in Ding et al. (2021)). Let  $(X_1, Y_1), \ldots, (X_n, Y_n)$  be i.i.d. so that  $X_i \sim \nu$  and  $Y_i \sim \nu'$ . Let  $\pi = \frac{1}{n} \sum_{i \in [n]} \delta_{X_i} - \nu$  and  $\pi' = \frac{1}{n} \sum_{i \in [n]} \delta_{Y_i} - \nu'$ . Assume that for all  $l \in [L]$ ,

$$\mathbb{P}(X_1 \in I_l, Y_1 \notin I_l) + \mathbb{P}(X_1 \notin I_l, Y_1 \in I_l) \leq \beta.$$

Then, for any  $\Delta > 0$ ,

$$d_L(\pi, \pi') := \sum_{l \in [L]} |\pi(I_l) - \pi'(I_l)| \le L\sqrt{\frac{\beta}{n}} + c_1\sqrt{\frac{\Delta}{n}},$$

with probability at least  $1 - e^{-\Delta}$ , where  $c_1$  is an absolute constant.

**Remark 3.** The  $\beta$  in original Lemma 7 of Ding et al. (2021) has a very complex definition (in equation (67)). But after checking the proof of Lemma 7, the lemma holds for any positive  $\beta$ .

**Lemma 4** (Lemma 6 in Ding et al. (2021)). Let  $X_1, \ldots, X_n$  and  $Y_1, \ldots, Y_n$  be two independent sequence of real-valued random variables, where  $X_i \sim \nu$  independently and  $Y_i \sim \nu'$  independently. Suppose the partition  $I_1, \ldots, I_L$  is chosen so that for all  $l \in [L]$ 

$$\frac{C_1}{L} \le \nu(I_l) \le \frac{C_2}{L},$$

for some absolute constants  $C_1, C_2 \in (0,1]$ . Given any  $\nu, \nu'$  in the real line, let  $\pi = \frac{1}{n} \sum_{i \in [n]} \delta_{X_i} - \nu$  and  $\pi' = \frac{1}{n} \sum_{i \in [n]} \delta_{Y_i} - \nu'$ . Assume that  $n \geq CL$  and  $L \geq L_0$  for some sufficiently large constants  $C, L_0$ . Then for any  $\Delta > 0$ ,

$$d_L(\pi, \pi') := \sum_{l \in [L]} |\pi(I_l) - \pi'(I_l)| \le c_2 \sqrt{\frac{L}{n}} - c_3 \sqrt{\frac{\Delta}{n}},$$

with probability at least  $1 - e^{-\Delta}$  and  $c_2, c_3$  are two absolute constants.

**Remark 4.** Original Lemma 6 in Ding et al. (2021) discusses a more general situation than above. The special case that  $X_i$  and  $Y_i$  are i.i.d. distributed is enough in our case.

**Proposition 1.** Suppose that we have samples  $(X_1, Y_1), \ldots, (X_n, Y_n)$  from (1); i.e.,  $(X_i, Y_i)$  i.i.d. follow the multivariate zero-mean Gaussian distribution with variance 1 and correlation  $\rho \in (0, 1)$ . Then, for all  $t \in \mathbb{R}$ , we have

$$p(t) := \mathbb{P}(X_1 \le t, Y_1 > t) \le \sqrt{1 - \rho^2}.$$

Proof of Proposition 1. See note 0403.

# References

Ding, J., Ma, Z., Wu, Y., and Xu, J. (2021). Efficient random graph matching via degree profiles. *Probability Theory and Related Fields*, 179(1):29–115.