Graphic Lasso: Accuracy with intercept

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Consider the model

$$\Omega_k = \Theta_0 + \sum_{l=1}^r u_{kl} \Theta_l, \quad k \in [K].$$

Let $U = [u_{kl}] \in \mathbb{R}^{K \times r}$ be the membership matrix and u_l denote the l-th column of U. The optimization problem is stated as

$$\min_{\{U,\Theta\}} \quad \mathcal{L}(U,\Theta) = \sum_{k=1}^{K} \langle S^k, \Omega^k \rangle - \log \det(\Omega^k),$$

$$s.t. \quad \Omega^k = \Theta_0 + u_k \Theta_1, \quad k = 1, ..., K,$$

$$\|U\|_F = 1, \sum_{k=1}^{K} u_k = 0,$$

where Θ_0 , Θ_1 are positive definite and $\tau_1 < \min\{\varphi_{\min}(\Theta_0), \varphi_{\min}(\Theta_1)\} \le \max\{\varphi_{\max}(\Theta_0), \varphi_{\max}(\Theta_1)\} < \tau_2, \tau_1, \tau_2 > 0$.

Lemma 1. Let $Z_i \sim_{i.i.d.} \mathcal{N}(0,\Sigma)$ and $\phi_{max}(\Sigma) \leq \tau < \infty$. Let $\Sigma = [\![\Sigma_{ij}]\!]$, then

$$P\left(\left|\sum_{i=1}^{n} Z_{ij} Z_{ik} - n \Sigma_{jk}\right| \ge n\nu\right) \le c_1 e^{-c_2 n\nu^2}, \quad \text{for} \quad |\nu| \le \delta,$$

where c_1, c_2, δ depends on τ only.

1 Conjectures

Lemma 2 (Conjecture). Consider the random variables $Z_i^k \sim_{i.i.d.} \mathcal{N}(0, \Sigma_1), i = [n], k \in [K],$ where Σ_1 are positive definite with bounded singular values. Then for a non-negative sequence $c_k \geq 0, k \in [K]$, we have

$$P\left(\left|\sum_{k=1}^{K}\sum_{i=1}^{n}\left[c_{k}Z_{ij}^{k}Z_{il}^{k}-c_{k}\Sigma_{1,jl}\right]\right|\geq nK\nu\right)\leq C_{1}\exp\left(-C_{2}nKv^{2}\right),$$

for some v small enough.

Theorem 1.1 (Accuracy for intercept case (Conjecture)). Let $\{\Theta_0, \Theta_1, U\}$ denote the true parameters. There exists a local minimizer $\{\hat{\Theta}_0, \hat{\Theta}_1, \hat{U}\}$ satisfies

$$\max \left\{ \left\| \hat{\Theta} - \Theta \right\|_F, \left\| \hat{\Theta}_1 - \Theta_1 \right\|_F, \max_{k \in [K]} |\hat{u}_k - u_k| \right\} = \mathcal{O}\left(\sqrt{\frac{p^2 \log p}{nK}}\right),$$

Proof. Define the function

$$G(\hat{U}, \hat{\Theta}_l) = \sum_{k=1}^K \langle S^k, \hat{\Theta}_0 + \hat{u}_k \hat{\Theta}_1 - \Theta_0 - u_k \Theta_1 \rangle - \log \det(\hat{\Theta}_0 + \hat{u}_k \hat{\Theta}_1) + \log \det(\Theta_0 + u_k \Theta_1).$$

Let $\Delta_k = \hat{\Theta}_0 + \hat{u}_k \hat{\Theta}_1 - \Theta_0 - u_k \Theta_1$. By Taylor Expansion, we have

$$-\log \det(\hat{\Theta}_0 + \hat{u}_k \hat{\Theta}_1) + \log \det(\Theta_0 + u_k \Theta_1) \ge -\langle (\Theta_0 + u_k \Theta_1)^{-1}, \Delta_k \rangle + \frac{1}{4\tau_2^2} \|\Delta_k\|_F^2.$$

Let $\Sigma^k = (\Theta_0 + u_k \Theta_1)^{-1}$ denote the true precision matrix. Then, we have

$$G(\hat{U}, \hat{\Theta}_l) \ge \sum_{k=1}^{K} \langle S^k - \Sigma^k, \Delta_k \rangle + \frac{1}{4\tau_2^2} \|\Delta_k\|_F^2 = I_1 + I_2.$$

Consider the set $\mathcal{A} = \left\{ (\hat{U}, \hat{\Theta}_1, \hat{\Theta}_0) : \|\Delta\|_F \leq M_1 \sqrt{\frac{p^2 \log p}{nK}}, \|\Delta_1\|_F \leq \gamma_1, \max_{k \in [K]} |\hat{u}_k - u_k| \leq \gamma_2 \right\},$ where $\gamma_1, \gamma_2 = o\left(\sqrt{\frac{p^2 \log p}{nK}}\right)$. Let $\partial \mathcal{A}$ denote the boundary of \mathcal{A} . Therefore, we only need to prove $G(\hat{u}, \hat{\Theta}) > 0$ for the estimates $\{\hat{u}, \hat{\Theta}\} \in \partial \mathcal{A}$.

For I_1 , let $\Delta = \hat{\Theta}_0 - \Theta_0$, $\Delta_1 = \hat{\Theta}_1 - \Theta_1$. Then, we have

$$\Delta_k = \Delta + u_k \Delta_1 + (\hat{u}_k - u_k)\hat{\Theta}_1.$$

Then, we have

$$\begin{split} |I_1| &= |\sum_{k=1}^K \langle S^k - \Sigma^k, \Delta_k \rangle| \\ &\leq |\sum_{k=1}^K \langle S^k - \Sigma^k, \Delta \rangle| + |\sum_{k=1}^K \langle S^k - \Sigma^k, u_k \Delta_1 \rangle| + |\sum_{k=1}^K \langle S^k - \Sigma^k, (\hat{u}_k - u_k) \hat{\Theta}_1 \rangle| \\ &\leq |\langle \sum_{k=1}^K S^k - \Sigma^k, \Delta \rangle| + |\langle \sum_{k=1}^K S^k - \Sigma^k, \Delta_1 \rangle| + \max_{k \in [K]} |(\hat{u}_k - u_k)| |\langle \sum_{k=1}^K S^k - \Sigma^k, \hat{\Theta}_1 \rangle|. \end{split}$$

Note that

$$\Sigma^{k} = (\Theta_{0} + u_{k}\Theta_{1})^{-1} = \Theta_{0}^{-1} + \frac{u_{k}}{1 + u_{k}\langle\Theta_{0}^{-1}, \Theta_{1}\rangle}\Theta_{0}^{-1}\Theta_{1}\Theta_{0}^{-1}.$$

Let $\Sigma_k = \Sigma_0 + c_k \Sigma_1$, where

$$\Sigma_0 = \Theta_0^{-1} + \min_{k \in [K]} \frac{u_k}{1 + u_k \langle \Theta_0^{-1}, \Theta_1 \rangle} \Theta_0^{-1} \Theta_1 \Theta_0^{-1}, \quad \Sigma_1 = \Theta_0^{-1} \Theta_1 \Theta_0^{-1}$$

, and

$$c_k = \frac{u_k}{1 + u_k \langle \Theta_0^{-1}, \Theta_1 \rangle} - \min_{k \in [K]} \frac{u_k}{1 + u_k \langle \Theta_0^{-1}, \Theta_1 \rangle}$$

Now, consider random variable $Y_i^k = X_i^k + \sqrt{c_k} Z_i^k \sim_{i.i.d.} \mathcal{N}(0, \Sigma_0 + c_k \Sigma_1)$, where $X_i^k \sim_{i.i.d.} \mathcal{N}(0, \Sigma_0)$, $Z_i^k \sim_{i.i.d.} \mathcal{N}(0, \Sigma_1)$ and X_i^k is independent with Z_i^k . Then, we have

$$\frac{1}{K} \sum_{k=1}^{K} S_{ab}^{k} - \Sigma_{0,ab} - c_{k} \Sigma_{1,ab} = \frac{1}{nK} \sum_{k=1}^{K} \sum_{i=1}^{n} Y_{ia}^{k} Y_{ib}^{k} - Y_{.a}^{k} Y_{.b}^{k} - \Sigma_{0,ab} - c_{k} \Sigma_{1,ab}.$$

Note that $Y_{.a}^k \to_{a.s.} 0$, and

$$\begin{split} Y_{ia}^k Y_{ib}^k &= [X_{ia}^k + \sqrt{c_k} Z_{ia}^k] [X_{ib}^k + \sqrt{c_k} Z_{ib}^k] \\ &= X_{ia}^k X_{ib}^k + c_k Z_{ia}^k Z_{ib}^k + \sqrt{c_k} X_{ia}^k Z_{ib}^k + \sqrt{c_k} Z_{ia}^k X_{ib}^k, \end{split}$$

where

$$\frac{1}{n} \sum_{i=1}^{n} X_{ia}^{k} Z_{ib}^{k} \to_{a.s.} 0, \quad \frac{1}{n} \sum_{i=1}^{n} Z_{ia}^{k} X_{ib}^{k} \to_{a.s.} 0.$$

Hence, with high probability, we have

$$\begin{split} |\frac{1}{K} \sum_{k=1}^K S_{ab}^k - \Sigma_{0,ab} - c_k \Sigma_{1,ab}| &= |\frac{1}{K} \sum_{k=1}^K X_{ia}^k X_{ib}^k + c_k Z_{ia}^k Z_{ib}^k - \Sigma_{0,ab} - c_k \Sigma_{1,ab}| \\ &\leq |\frac{1}{K} \sum_{k=1}^K \sum_{i=1}^n X_{ia}^k X_{ib}^k - \Sigma_{0,ab}| + |\frac{1}{K} \sum_{k=1}^K \sum_{i=1}^n c_k Z_{ia}^k Z_{ib}^k - c_k \Sigma_{1,ab}|, \\ &\leq C \sqrt{\frac{\log p}{nK}}, \end{split}$$

where C is a constant and the last inequality follows by the lemmas 1 and 2. Therefore, we have

$$|I_1| \le C\sqrt{K}\sqrt{\frac{p^2\log p}{n}} \left[\|\Delta\|_F + \|\Delta_1\|_F + \max_{k \in [K]} |(\hat{u}_k - u_k)| \right].$$

For I_2 , consider the estimate $\{\hat{\Theta}_0, \hat{\Theta}_1, \hat{U}\} \in \partial \mathcal{A}$. By triangle inequality, we have

$$\|\Delta\|_F - \|\Delta_1\|_F - \max_{k \in [K]} |(\hat{u}_k - u_k)|_{\tau_2} \le \|\Delta_k\|_F \le \|\Delta\|_F + \|\Delta_1\|_F + \max_{k \in [K]} |(\hat{u}_k - u_k)|_{\tau_2},$$

and thus $\|\Delta_k\|_F \simeq \|\Delta\|_F$.

Therefore, for the estimate $\{\hat{\Theta}_0, \hat{\Theta}_1, \hat{U}\} \in \partial \mathcal{A}$, we have

$$\begin{split} G(\hat{\Theta}_{0}, \hat{\Theta}_{1}, \hat{U}) &\geq I_{2} - |I_{1}| \\ &\geq \frac{C''}{4\tau_{2}^{2}} K \|\Delta\|_{F} - C' \sqrt{K} \sqrt{\frac{p^{2} \log p}{n}} \left[\|\Delta\|_{F} \right] \\ &\geq \frac{C'' M^{2}}{4\tau_{2}^{2}} \frac{p^{2} \log p}{n} - C' M \frac{p^{2} \log p}{n} \\ &> 0, \end{split}$$

for M large enough.