

# Graphic Lasso: Accuracy with intercept

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Consider the model

$$\Omega_k = \Theta_0 + \sum_{l=1}^r u_{kl} \Theta_l, \quad k \in [K].$$

Let  $U = \llbracket u_{kl} \rrbracket \in \mathbb{R}^{K \times r}$  be the membership matrix and  $u_l$  denote the  $l$ -th column of  $U$ . The optimization problem is stated as

$$\begin{aligned} \min_{\{U, \Theta\}} \quad & \mathcal{L}(U, \Theta) = \sum_{k=1}^K \langle S^k, \Omega^k \rangle - \log \det(\Omega^k), \\ \text{s.t.} \quad & \Omega^k = \Theta_0 + u_k \Theta_1, \quad k = 1, \dots, K, \\ & \|U\|_F = 1, \sum_{k=1}^K u_k = 0, \end{aligned}$$

where  $\Theta_0, \Theta_1$  are positive definite and  $\tau_1 < \min\{\varphi_{\min}(\Theta_0), \varphi_{\min}(\Theta_1)\} \leq \max\{\varphi_{\max}(\Theta_0), \varphi_{\max}(\Theta_1)\} < \tau_2, \tau_1, \tau_2 > 0$ .

**Lemma 1.** Let  $Z_i \sim_{i.i.d.} \mathcal{N}(0, \Sigma)$  and  $\phi_{\max}(\Sigma) \leq \tau < \infty$ . Let  $\Sigma = \llbracket \Sigma_{ij} \rrbracket$ , then

$$P \left( \left| \sum_{i=1}^n Z_{ij} Z_{ik} - n \Sigma_{jk} \right| \geq n \nu \right) \leq c_1 e^{-c_2 n \nu^2}, \quad \text{for } |\nu| \leq \delta,$$

where  $c_1, c_2, \delta$  depends on  $\tau$  only.

## 1 Conjectures

**Lemma 2 (Conjecture).** Consider the random variables  $Z_i^k \sim_{i.i.d.} \mathcal{N}(0, \Sigma_1), i = [n], k \in [K]$ , where  $\Sigma_1$  are positive definite with bounded singular values. Then for a non-negative sequence  $c_k \geq 0, k \in [K]$ , we have

$$P \left( \left| \sum_{k=1}^K \sum_{i=1}^n \left[ c_k Z_{ij}^k Z_{il}^k - \overset{\text{sum of squares of } c\_k}{c_k \Sigma_{1,jl}} \right] \right| \geq \overset{\text{sum of squares of } c\_k}{n K \nu} \right) \leq C_1 \exp(-C_2 n K \nu^2),$$

K  $\rightarrow$  sum of squares of c\_k

for some  $\nu$  small enough.

**Theorem 1.1** (Accuracy for intercept case (Conjecture)). Let  $\{\Theta_0, \Theta_1, U\}$  denote the true parameters. There exists a local minimizer  $\{\hat{\Theta}_0, \hat{\Theta}_1, \hat{U}\}$  satisfies

$$\max \left\{ \left\| \hat{\Theta} - \Theta \right\|_F, \left\| \hat{\Theta}_1 - \Theta_1 \right\|_F, \max_{k \in [K]} |\hat{u}_k - u_k| \right\} = \mathcal{O} \left( \sqrt{\frac{p^2 \log p}{nK}} \right),$$

*Proof.* Define the function

$$G(\hat{U}, \hat{\Theta}_l) = \sum_{k=1}^K \langle S^k, \hat{\Theta}_0 + \hat{u}_k \hat{\Theta}_1 - \Theta_0 - u_k \Theta_1 \rangle - \log \det(\hat{\Theta}_0 + \hat{u}_k \hat{\Theta}_1) + \log \det(\Theta_0 + u_k \Theta_1).$$

Let  $\Delta_k = \hat{\Theta}_0 + \hat{u}_k \hat{\Theta}_1 - \Theta_0 - u_k \Theta_1$ . By Taylor Expansion, we have

$$-\log \det(\hat{\Theta}_0 + \hat{u}_k \hat{\Theta}_1) + \log \det(\Theta_0 + u_k \Theta_1) \geq -\langle (\Theta_0 + u_k \Theta_1)^{-1}, \Delta_k \rangle + \frac{1}{4\tau_2^2} \|\Delta_k\|_F^2.$$

Let  $\Sigma^k = (\Theta_0 + u_k \Theta_1)^{-1}$  denote the true precision matrix. Then, we have

$$G(\hat{U}, \hat{\Theta}_l) \geq \sum_{k=1}^K \langle S^k - \Sigma^k, \Delta_k \rangle + \frac{1}{4\tau_2^2} \|\Delta_k\|_F^2 = I_1 + I_2.$$

Consider the set  $\mathcal{A} = \left\{ (\hat{U}, \hat{\Theta}_1, \hat{\Theta}_0) : \|\Delta\|_F \leq M_1 \sqrt{\frac{p^2 \log p}{nK}}, \|\Delta_1\|_F \leq \gamma_1, \max_{k \in [K]} |\hat{u}_k - u_k| \leq \gamma_2 \right\}$ ,

where  $\gamma_1, \gamma_2 = o\left(\sqrt{\frac{p^2 \log p}{nK}}\right)$ . Let  $\partial\mathcal{A}$  denote the boundary of  $\mathcal{A}$ . Therefore, we only need to prove

$G(\hat{u}, \hat{\Theta}) > 0$  for the estimates  $\{\hat{u}, \hat{\Theta}\} \in \partial\mathcal{A}$ .

**For**  $I_1$ , let  $\Delta = \hat{\Theta}_0 - \Theta_0$ ,  $\Delta_1 = \hat{\Theta}_1 - \Theta_1$ . Then, we have

$$\Delta_k = \Delta + u_k \Delta_1 + (\hat{u}_k - u_k) \hat{\Theta}_1.$$

Then, we have

$$\begin{aligned} |I_1| &= \left| \sum_{k=1}^K \langle S^k - \Sigma^k, \Delta_k \rangle \right| \\ &\leq \left| \sum_{k=1}^K \langle S^k - \Sigma^k, \Delta \rangle \right| + \left| \sum_{k=1}^K \langle S^k - \Sigma^k, u_k \Delta_1 \rangle \right| + \left| \sum_{k=1}^K \langle S^k - \Sigma^k, (\hat{u}_k - u_k) \hat{\Theta}_1 \rangle \right| \\ &\leq \left| \left\langle \sum_{k=1}^K S^k - \Sigma^k, \Delta \right\rangle \right| + \left| \left\langle \sum_{k=1}^K S^k - \Sigma^k, \Delta_1 \right\rangle \right| + \max_{k \in [K]} |\hat{u}_k - u_k| \left| \left\langle \sum_{k=1}^K S^k - \Sigma^k, \hat{\Theta}_1 \right\rangle \right|. \end{aligned}$$

Note that

$$\Sigma^k = (\Theta_0 + u_k \Theta_1)^{-1} = \Theta_0^{-1} + \frac{u_k}{1 + u_k \langle \Theta_0^{-1}, \Theta_1 \rangle} \Theta_0^{-1} \Theta_1 \Theta_0^{-1}.$$

Let  $\Sigma_k = \Sigma_0 + c_k \Sigma_1$ , where

$$\Sigma_0 = \Theta_0^{-1} + \min_{k \in [K]} \frac{u_k}{1 + u_k \langle \Theta_0^{-1}, \Theta_1 \rangle} \Theta_0^{-1} \Theta_1 \Theta_0^{-1}, \quad \Sigma_1 = \Theta_0^{-1} \Theta_1 \Theta_0^{-1}$$

, and

$$c_k = \frac{u_k}{1 + u_k \langle \Theta_0^{-1}, \Theta_1 \rangle} - \min_{k \in [K]} \frac{u_k}{1 + u_k \langle \Theta_0^{-1}, \Theta_1 \rangle}.$$

Now, consider random variable  $Y_i^k = X_i^k + \sqrt{c_k} Z_i^k \sim_{i.i.d.} \mathcal{N}(0, \Sigma_0 + c_k \Sigma_1)$ , where  $X_i^k \sim_{i.i.d.} \mathcal{N}(0, \Sigma_0)$ ,  $Z_i^k \sim_{i.i.d.} \mathcal{N}(0, \Sigma_1)$  and  $X_i^k$  is independent with  $Z_i^k$ . Then, we have

$$\frac{1}{K} \sum_{k=1}^K S_{ab}^k - \Sigma_{0,ab} - c_k \Sigma_{1,ab} = \frac{1}{nK} \sum_{k=1}^K \sum_{i=1}^n Y_{ia}^k Y_{ib}^k - Y_{.a}^k Y_{.b}^k - \Sigma_{0,ab} - c_k \Sigma_{1,ab}.$$

Note that  $Y_{.a}^k \rightarrow_{a.s.} 0$ , and

$$\begin{aligned} Y_{ia}^k Y_{ib}^k &= [X_{ia}^k + \sqrt{c_k} Z_{ia}^k][X_{ib}^k + \sqrt{c_k} Z_{ib}^k] \\ &= X_{ia}^k X_{ib}^k + c_k Z_{ia}^k Z_{ib}^k + \sqrt{c_k} X_{ia}^k Z_{ib}^k + \sqrt{c_k} Z_{ia}^k X_{ib}^k, \end{aligned}$$

where

$$\frac{1}{n} \sum_{i=1}^n X_{ia}^k Z_{ib}^k \rightarrow_{a.s.} 0, \quad \frac{1}{n} \sum_{i=1}^n Z_{ia}^k X_{ib}^k \rightarrow_{a.s.} 0.$$

Hence, with high probability, we have

$$\begin{aligned} \left| \frac{1}{K} \sum_{k=1}^K S_{ab}^k - \Sigma_{0,ab} - c_k \Sigma_{1,ab} \right| &= \left| \frac{1}{K} \sum_{k=1}^K X_{ia}^k X_{ib}^k + c_k Z_{ia}^k Z_{ib}^k - \Sigma_{0,ab} - c_k \Sigma_{1,ab} \right| \\ &\leq \left| \frac{1}{K} \sum_{k=1}^K \sum_{i=1}^n X_{ia}^k X_{ib}^k - \Sigma_{0,ab} \right| + \left| \frac{1}{K} \sum_{k=1}^K \sum_{i=1}^n c_k Z_{ia}^k Z_{ib}^k - c_k \Sigma_{1,ab} \right|, \\ &\leq C \sqrt{\frac{\log p}{nK}}, \end{aligned}$$

where  $C$  is a constant and the last inequality follows by the lemmas 1 and 2.

Therefore, we have

$$|I_1| \leq C \sqrt{K} \sqrt{\frac{p^2 \log p}{n}} \left[ \|\Delta\|_F + \|\Delta_1\|_F + \max_{k \in [K]} |(\hat{u}_k - u_k)| \right].$$

**For  $I_2$ ,** consider the estimate  $\{\hat{\Theta}_0, \hat{\Theta}_1, \hat{U}\} \in \partial \mathcal{A}$ . By triangle inequality, we have

$$\|\Delta\|_F - \|\Delta_1\|_F - \max_{k \in [K]} |(\hat{u}_k - u_k)| \tau_2 \leq \|\Delta_k\|_F \leq \|\Delta\|_F + \|\Delta_1\|_F + \max_{k \in [K]} |(\hat{u}_k - u_k)| \tau_2,$$

and thus  $\|\Delta_k\|_F \asymp \|\Delta\|_F$ .

Therefore, for the estimate  $\{\hat{\Theta}_0, \hat{\Theta}_1, \hat{U}\} \in \partial \mathcal{A}$ , we have

$$\begin{aligned} G(\hat{\Theta}_0, \hat{\Theta}_1, \hat{U}) &\geq I_2 - |I_1| \\ &\geq \frac{C''}{4\tau_2^2} K \|\Delta\|_F - C' \sqrt{K} \sqrt{\frac{p^2 \log p}{n}} [\|\Delta\|_F] \\ &\geq \frac{C'' M^2}{4\tau_2^2} \frac{p^2 \log p}{n} - C' M \frac{p^2 \log p}{n} \\ &> 0, \end{aligned}$$

for  $M$  large enough.

□