

# Graphic Lasso: Clustering accuracy for precision matrix model

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## 1 Weakest Condition under which the clustering accuracy holds?

The following are the weakest assumptions to ensure the clustering accuracy for tensor block model.

1. Irreducibility. This implies the minimal gap between blocks  $\delta = \min_k \delta^{(k)} > 0$ , where

$$\delta^{(k)} = \min_{r_k \neq r'_k} \max_{r_1, \dots, r_{k-1}, r_{k+1}, \dots, r_K} (f(c_{r_1, \dots, r_k, \dots, r_K}) - f(c_{r_1, \dots, r'_k, \dots, r_K}))^2.$$

Otherwise, two blocks with the same mean are not able to be distinguished.

2. Convex loss function. The loss function should be convex, otherwise the estimation of  $\mathcal{C}$  would not be unique with given membership  $\{\mathbf{M}_k\}$ , and thus the stochastic error for estimation will be hard to measure.
3. Bounded variance. This implies there exist two positive constants  $a_1, a_2$  such that  $0 < a_1 < \text{Var}(\mathcal{Y}_{i_1, \dots, i_K} | \Theta_{i_1, \dots, i_K}) < a_2 < \infty$ .

### For precision matrix:

1. Irreducibility. The minimal gap between the precision matrices  $\delta > 0$ , where

$$\delta = \min_{k \neq k' \in [K], i, j \in [p]} |\Theta_{i,j}^k - \Theta_{i,j}^{k'}|^2 \geq 0.$$

2. Convex loss function. Note that for symmetric matrices  $A, B$ ,  $\text{tr}(AB) = \langle A, B \rangle$  is a convex function. Since  $\Theta^l$  are positive-definite, the  $\log \det(\cdot)$  is a concave function. Thus, the objective function for the precision matrix model is convex.
3. Bounded variance. This requires the variance of the sample covariance matrices are bounded. Since the variance  $\text{Var}(S^k) = f(\Sigma^k)$  is a function of the true covariance matrix, the assumption holds when the elements in the true covariances matrices are bounded.

Therefore, the precision matrix model satisfies the weakest assumption for the clustering accuracy.

## 2 Accuracy

Consider the optimization problem.

$$\begin{aligned} \min_{\mathbf{U}, \{\Theta^l\}} \quad & Q(\mathbf{U}, \{\Theta^l\}) = \sum_{k=1}^K \text{tr}(S^k \Omega^k) - \log |\Omega^k|, \\ \text{s.t.} \quad & \Omega^k = \sum_{l=1}^r u_{kl} \Theta^l, \\ & \mathbf{U} \in \{0, 1\}^{K \times r} \text{ is a membership matrix,} \\ & \{\Theta^l\} \text{ are irreducible and invertible.} \end{aligned}$$

*Proof.* Recall the objective function

$$Q(\mathbf{U}, \{\Theta^l\}) = \sum_{k=1}^K \text{tr}(S^k \Omega^k) - \log |\Omega^k| = \sum_{k=1}^K \langle S^k, \Omega^k \rangle - \log |\Omega^k|.$$

The deviation between the true parameters  $\{\mathbf{U}, \{\Theta^l\}\}$  and estimations  $\{\hat{\mathbf{U}}, \{\hat{\Theta}^l\}\}$  comes from two aspects: the estimation of  $\{\Theta^l\}$  and the misclassification (estimation of  $\mathbf{U}$ ). We tease apart these two parts.

1. First, we suppose the membership  $\mathbf{U}$  is given. We now assess the stochastic error due to the estimation of  $\{\Theta^l\}$ , conditional on  $\mathbf{U}$ . Define the index sets  $I_l = \{k : u_{kl} = 1\}$  for  $l = 1, \dots, r$ . Note that the objective function is convex. The optimal  $\{\Theta^l\}$  satisfies the first order condition.

$$\frac{\partial Q}{\partial \Theta_{ij}^l} = \sum_{k \in I_l} S_{ij}^k - |I_l| \frac{C(\Theta^l)_{ij}}{|\Theta^l|} = 0,$$

where  $C(A)_{ij}$  is the cofactor of matrix  $A$  corresponding to the element  $A_{ij}$ . Note that  $\Theta^l$  should be a symmetric matrix. The cofactor matrix  $C(\Theta^l) = C^T(\Theta^l)$ . Then, the derivation of the matrix  $\Theta^l$  is equal to

$$\frac{\partial Q}{\partial \Theta^l} = \sum_{k \in I_l} S^k - |I_l| \frac{C^T(\Theta^l)}{|\Theta^l|} = 0,$$

which implies that

$$\hat{\Theta}^l = \left( \frac{\sum_{k \in I_l} S^k}{|I_l|} \right)^{-1}, \quad \text{for } l \in [r].$$

Therefore, the estimation of  $\{\Theta^l\}$  is a function of  $\mathbf{U}$ . Consider the function  $F(\mathbf{U}) = Q(\mathbf{U}, \{\hat{\Theta}^l\})$ . By a straightforward calculation, we have

$$\begin{aligned} F(\mathbf{U}) &= \sum_{l=1}^r \sum_{k \in I_l} \langle S^k, \left( \frac{\sum_{k \in I_l} S^k}{|I_l|} \right)^{-1} \rangle - |I_l| \log \left| \left( \frac{\sum_{k \in I_l} S^k}{|I_l|} \right)^{-1} \right| \\ &= \sum_{l=1}^r \left\langle \sum_{k \in I_l} S^k, \left( \frac{\sum_{k \in I_l} S^k}{|I_l|} \right)^{-1} \right\rangle - |I_l| \log \left| \left( \frac{\sum_{k \in I_l} S^k}{|I_l|} \right)^{-1} \right| \\ &= \sum_{l=1}^r |I_l| p - |I_l| \log \left| \left( \frac{\sum_{k \in I_l} S^k}{|I_l|} \right)^{-1} \right|. \end{aligned}$$

Note that  $\sum_{l=1}^r |I_l|p = Kp$  is independent with the membership. We only need to consider the second term. For simplicity, we define

$$F(\mathbf{U}) = - \sum_{l=1}^r |I_l| \log \left| \left( \frac{\sum_{k \in I_l} S^k}{|I_l|} \right)^{-1} \right|.$$

Correspondingly, we define the population version of  $F(\mathbf{U})$  as following.

$$G(\mathbf{U}) = - \sum_{l=1}^r |I_l| \log \left| \left( \frac{\sum_{k \in I_l} \Sigma^k}{|I_l|} \right)^{-1} \right|,$$

where  $\Sigma^k = \mathbb{E}[S^k]$  is the true covariance matrices. Therefore, the deviation  $F(\mathbf{U}) - G(\mathbf{U})$  quantifies the stochastic error due to the estimation of  $\{\Theta^l\}$ .

2. Next, we free  $\mathbf{U}$  and quantify the total deviation. Considering the maximizer,

$$\hat{\mathbf{U}} = \arg \min_{\mathbf{U}} F(\mathbf{U}).$$

The corresponding  $G(\hat{\mathbf{U}})$  is

$$G(\hat{\mathbf{U}}) = - \sum_{l=1}^r |\hat{I}_l| \log \left| \left( \frac{\sum_{k \in \hat{I}_l} \Sigma^k}{|\hat{I}_l|} \right)^{-1} \right|,$$

and the function  $G(\mathbf{U})$  with true membership is

$$G(\mathbf{U}) = - \sum_{l=1}^r |I_l| \log \left| \left( \frac{\sum_{k \in I_l} \Sigma^k}{|I_l|} \right)^{-1} \right| = \sum_{l=1}^r |I_l| \log |\Theta^l|.$$

Then, the deviation  $G(\hat{\mathbf{U}}) - G(\mathbf{U})$  measures the stochastic error of the misclassification.

□

**Lemma 1.** Suppose the misclassification rate  $MCR(\hat{\mathbf{U}}, \mathbf{U}) = \max_{l, a \neq a' \in [r]} \min\{D_{al}, D(a'l)\} \geq \epsilon$ .

$$G(\hat{\mathbf{U}}) - G(\mathbf{U}) \leq$$

*Proof.* Recall the definition of confusion matrix  $D_{ll'} = \sum_{k=1}^K \mathbf{I}\{u_{kl} = 1, \hat{u}_{kl'} = 1\}$ . Since  $MCR(\hat{\mathbf{U}}, \mathbf{U}) \geq \epsilon$ , without the loss of generality, let  $D_{11}$  be the largest element in the first column of  $D$ ,  $D_{12}$  be the second largest element in the column, and  $D_{21} \geq \epsilon$ . □