Graphic Lasso: Scaled membership (Simple Case)

Jiaxin Hu

March 25, 2021

1 Thoughts

1. In Note 0323, I decomposed the original difference of the likelihood into 5 terms $H_1, ..., H_5$, and I tried to use the following inequality to show the MLE estimate is near to the true parameters.

$$0 \ge G(\hat{u}, \hat{\Theta}) \ge H_1 + H_5 - H_2 - |H_3| + H_4.$$

However, from G to $H_1, ..., H_5$, there are a lot of inequalities. I think this may be the reason why I can not show $\hat{u} \to u$ and $\hat{\Theta} \to \Theta$.

Therefore, in the following new proof, I would like to use the original G and show that $\hat{u}\hat{\Theta} \to u\Theta$ and further $\hat{u} \to u, \hat{\Theta} \to \Theta$.

In the discrete case, we have $\sum_{al} D_{al} \|\Delta_{al}\|_F \to 0$, where D_{al} is the entries of confusion matrix and $\Delta_{al} = \hat{\Theta}^l - \Theta^a$. Then, we know that

$$D_{al} \|\Delta_{al}\| + D_{a'l} \|\Delta_{a'l}\| \ge \min\{D_{al}, D_{a'l}\} \|\Theta^a - \Theta^{a'}\| \ge \min\{D_{al}, D_{a'l}\}\delta,$$

where δ is the minimal gap between Θ^l . Thus, for each a, there is only one l such that D_{al} does not tend to 0, i.e., with proper permutation, all the off-diagonal elements in the confusion matrix tends to 0.

In our case, $\sum_{k=1}^{K} \|\hat{u}_k \hat{\Theta} - u_k \Theta\|$ is an analogy of $\sum_{al} D_{al} \|\Delta_{al}\|_F$ in the continuous case. Since we do not have minimal gap here and $\hat{\Theta}, \Theta$ are positive definite, I think similar techniques can be applied to our case from the angle of u_k . See Step I for details.

2. The constraint $||u||_F^2 = K$ is crucial since we need $u_k \ge a > 0$ and the norm of u grows along with K.

2 Simple case

Consider the model in which K categories share the same precision matrix structure with different magnitude. The optimization problem is stated below:

$$\min_{\{u,\Theta\}} \quad \mathcal{L}(u,\Theta) = \sum_{k=1}^{K} \langle S^k, \Omega^k \rangle - \log \det(\Omega^k),$$

$$s.t. \quad \Omega^k = u_k \Theta, \quad k = 1, ..., K,$$

$$u_k \ge a, \|u\|_F^2 = K, \quad a > 0,$$

$$\Theta \text{ is positive definite with, and } \tau_1 < \varphi_{\min}(\Theta) \le \varphi_{\max}(\Theta) < \tau_2, \tau_1, \tau_2 > 0$$

Lemma 1 (Precision matrix Accuracy). Let $\{u, \Theta\}$ denote the true parameters. Consider a estimation $\{\hat{u}, \hat{\Theta}\}$ such that $\mathcal{L}(\hat{u}, \hat{\Theta}) \geq \mathcal{L}(u, \Theta)$. With probability tends to 1 as $n \to \infty$, we have the accuracy

$$\sum_{k=1}^{K} \left\| \hat{\Omega}^k - \Omega^k \right\|_F = \sum_{k=1}^{K} \left\| \hat{u}_k \hat{\Theta} - u_k \Theta \right\|_F \le 16\tau_2^2 K^{3/2} C \sqrt{\frac{p^2 \log p}{n}}$$

Remark 1. In the accuracy rate, the order of K is $\mathcal{O}(K^{3/2})$. We can consider the factor $\mathcal{O}(\sqrt{K})$ is from the estimation of Θ and K is from the estimation of u_K . I think this conclusion follows the intuition. By previous result Note 0113, the accuracy rate to estimation common precision matrix is of order $\mathcal{O}(\sqrt{K})$. For continuous values u_k , the accuracy rate to estimate K different continuous variables is about $\mathcal{O}(K)$. In our result, $\sum_{k=1}^K \left\| \hat{u}_k \hat{\Theta} - u_k \Theta \right\|_F$ can be considered as a multiplication $(\hat{u}\hat{\Theta})$ of the error from estimating Θ and u_k .

Proof. We prove the accuracy rate by two steps.

Step I: Show that $\hat{u} \to u$ and $\hat{\Theta} \to \Theta$.

First, we define

$$G(\hat{u}, \hat{\Theta}) = \mathcal{L}(\hat{u}, \hat{\Theta}) - \mathcal{L}(u, \Theta)$$

$$= \sum_{k=1}^{K} \langle S^k, \hat{u}_k \hat{\Theta} \rangle - \langle S^k, u_k \Theta \rangle - \log \det(\hat{u}_k \hat{\Theta}) + \log \det(u_k \Theta).$$

Let $\Delta_k = \hat{u}_k \Theta - u_k \Theta$. By Taylor expansion, we have

$$-\log \det(\hat{u}_k \hat{\Theta}) + \log \det(u_k \Theta) \ge -\langle (u_k \Theta)^{-1}, \Delta_k \rangle + \frac{1}{2u_k^2 \tau_2^2 + \|\Delta_k\|_F^2} \|\Delta_k\|_F^2,$$

$$\ge -\langle u_k^{-1} \Sigma^{-1}, \Delta_k \rangle + \frac{1}{2u_k^2 \tau_2^2 + \|\Delta_k\|_F^2} \|\Delta_k\|_F^2. \tag{1}$$

Plugging the inequality (1) into G, we have

$$G(\hat{u}, \hat{\Theta}) \ge \sum_{k=1}^{K} \langle S^k - u_k^{-1} \Sigma, \Delta_k \rangle + \frac{1}{2K\tau_2^2 + (\sum_{k=1}^{K} \|\Delta_k\|_F)^2} \sum_{k=1}^{K} \|\Delta_k\|_F^2.$$
 (2)

Let $X_1^k, ..., X_n^k \sim_{i.i.d.} \mathcal{N}(0, \Sigma/u_k)$. We know that

$$S_{jl}^{k} = \frac{1}{n} \sum_{i=1}^{n} \left[X_{ij}^{k} X_{jl}^{k} - X_{.j}^{k} X_{.l}^{k} \right].$$

Since $X_{.j}^k, X_{.l}^k \to 0$ almost sure when $n \to \infty$, we have

$$|S_{jl}^{k} - \Sigma_{jl}/u_{k}| = |\frac{1}{n}X_{ij}^{k}X_{jl}^{k} - \Sigma_{jl}/u_{k}| \le C\sqrt{\frac{\log p}{n}},$$
(3)

with high probability. Therefore, by the assumption $\mathcal{L}(\hat{u}, \hat{\Theta}) \geq \mathcal{L}(u, \Theta)$, we have

$$0 \ge G(\hat{u}, \hat{\Theta}) \ge \frac{1}{2K\tau_2^2 + (\sum_{k=1}^K \|\Delta_k\|_F)^2} \sum_{k=1}^K \|\Delta_k\|_F^2 - C\sqrt{\frac{\log p}{n}} \sum_{k=1}^K \|\Delta_k\|, \tag{4}$$

which implies that

$$C\sqrt{\frac{\log p}{n}}K\left[2K\tau_2^2 + (\sum_{k=1}^K \|\Delta_k\|_F)^2\right] - \sum_{k=1}^K \|\Delta_k\|_F \ge 0.$$

Note that $\sqrt{\frac{\log p}{n}} \to 0$ as $n \to \infty$. We need

$$\sum_{k=1}^{K} \|\Delta_k\|_F = \sum_{k=1}^{K} \|\hat{u}_k \hat{\Theta} - u_k \Theta\|_F \to 0, \quad n \to \infty.$$

Since $\|\Delta_k\|_F \geq 0$, we also have

$$\|\Delta_k\|_F = \|\hat{u}_k \hat{\Theta} - u_k \Theta\|_F \to 0, \quad n \to \infty, \quad \text{for all} \quad k \in [K]$$

and thus

$$\|\hat{u}_k\hat{\Theta} - u_k\Theta\|_F/u_k \to 0$$
, for all $k \in [K]$, and $\sum_{k=1}^K \|\hat{u}_k\hat{\Theta} - u_k\Theta\|_F/u_k \to 0$.

For arbitrary $k, k' \in [K]$, note that

$$\left\|\hat{u}_k\hat{\Theta} - u_k\Theta\right\|_F / u_k + \left\|\hat{u}_{k'}\hat{\Theta} - u_{k'}\Theta\right\|_F / u_{k'} \ge \left\|(\hat{u}_k / u_k - \hat{u}_{k'} / u_{k'})\hat{\Theta}\right\|_F \to 0,$$

which implies for any pair (k, k'), we need

$$\frac{\hat{u}_k}{u_k} - \frac{\hat{u}_{k'}}{u_{k'}} \to 0$$
, and thus $\hat{u} \to cu$,

for some constant c. By the assumption that $\|\hat{u}\|_F = \|u\|_F = K$, the constant c = 1 and therefore we obtain that $\hat{u} \to u$ as $n \to \infty$. On the other hand, given $\hat{u} \to u$, we also have

$$\|\Delta_k\|_F = \|u_k(\hat{\Theta} - \Theta) + (\hat{u}_k - u_k)\hat{\Theta}\|_F \to 0, \text{ for all } k \in [K],$$

which implies that $\|\hat{\Theta} - \Theta\|_{F} \to 0$.

Sanity Check: Let $S^k = u_k^{-1} \Sigma$.

The inequality (2) becomes,

$$0 \ge G(\hat{u}, \hat{\Theta}) \ge \frac{1}{2K\tau_2^2 + (\sum_{k=1}^K \|\Delta_k\|_F)^2} \sum_{k=1}^K \|\Delta_k\|_F^2,$$

which requires $\sum_{k=1}^{K} \|\Delta_k\|_F^2 \to 0$, otherwise, the right hand side tends to a positive constant as $n \to \infty$. Then, following the above steps from $\sum_{k=1}^{K} \|\Delta_k\|_F^2 \to 0$ to $\hat{u}_k \to u_k$ and $\hat{\Theta} \to \Theta$, we obtain the conclusion that MLE is near the true parameters.

Step II: Sharpen the accuracy rate.

Note that accuracy rate bound from inequality (4) is sub-optimal since it does not use the common structure of the precision matrix. Therefore, back to the inequality (2) of G.

$$G(\hat{u}, \hat{\Theta}) \ge \sum_{k=1}^{K} \langle S^k - u_k^{-1} \Sigma, \Delta_k \rangle + \sum_{k=1}^{K} \frac{1}{2u_k^2 \tau_2^2 + (\sum_{k=1}^{K} \|\Delta_k\|_F)^2} \|\Delta_k\|_F^2,$$

$$\ge \sum_{k=1}^{K} \langle \left[u_k S^k - \Sigma \right], \Delta_k / u_k \rangle + \frac{1}{4\tau_2^2} \sum_{k=1}^{K} \|\Delta_k / u_k\|_F^2,$$

$$= I_1 + I_2.$$

where the second inequality follows by the conclusion in Step I, and I_1, I_2 denote the two terms respectively. Let $\Delta = \hat{\Theta} - \Theta$. Note that

$$\Delta_k/u_k = \hat{u}_k/u_k\hat{\Theta} - \Theta = \Delta + (\hat{u}_k/u_k - 1)\hat{\Theta}.$$
 (5)

For I_1 , by the decomposition (5), we have

$$I_{1} = \sum_{k=1}^{K} \langle \left[u_{k} S^{k} - \Sigma \right], \Delta \rangle + \sum_{k=1}^{K} \left(\hat{u}_{k} / u_{k} - 1 \right) \langle \left[u_{k} S^{k} - \Sigma \right], \hat{\Theta} \rangle$$

$$\leq \sum_{k=1}^{K} \langle \left[u_{k} S^{k} - \Sigma \right], \Delta \rangle + \max_{k \in [K]} \left| \left(\hat{u}_{k} / u_{k} - 1 \right) \right| \sum_{k=1}^{K} \left| \left\langle \left[u_{k} S^{k} - \Sigma \right], \hat{\Theta} \right\rangle \right|,$$

By similar process to obtain the inequality (3), we have

$$\max_{(i,j)} |\sum_{k=1}^{K} \left[u_k S_{jl}^k - \Sigma_{jl} \right] | \le \sqrt{K} C \sqrt{\frac{\log p}{n}},$$

with high probability. Therefore, we have

$$|I_1| \le \sqrt{K}C\sqrt{\frac{p^2\log p}{n}} \left[\|\Delta\|_F + \max_{k \in [K]} |(\hat{u}_k/u_k - 1)| \|\hat{\Theta}\|_F \right].$$
 (6)

For I_2 , note that for n large enough,

$$\begin{split} \left\| \Delta_k / u_k \right\|_F &= \left\| \Delta \right\|_F + \max_{k \in [K]} \left| \left(\hat{u}_k / u_k - 1 \right) \right| \left\| \hat{\Theta} \right\|_F \\ &+ \left\| \Delta + \left| \left(\hat{u}_k / u_k - 1 \right) \right| \hat{\Theta} \right\|_F - \left(\left\| \Delta \right\|_F + \max_{k \in [K]} \left| \left(\hat{u}_k / u_k - 1 \right) \right| \left\| \hat{\Theta} \right\|_F \right) \\ &\geq \frac{1}{2} \left[\left\| \Delta \right\|_F + \max_{k \in [K]} \left| \left(\hat{u}_k / u_k - 1 \right) \right| \left\| \hat{\Theta} \right\|_F \right], \end{split}$$

where the inequality follows the fact that both $\|\Delta\|_F$, $\max_{k \in [K]} |(\hat{u}_k/u_k - 1)| \to 0$ as $n \to \infty$. This inequality makes sense since $\|A + B\|_F^2$ are near to $\|A\|_F^2 + \|B\|_F^2$ when all the entries in A, B are close to 0. Therefore, we have

$$I_{2} \geq \frac{1}{16\tau_{2}^{2}} \sum_{k=1}^{K} \left[\|\Delta\|_{F} + \max_{k \in [K]} |\left(\hat{u}_{k}/u_{k} - 1\right)| \|\hat{\Theta}\|_{F} \right]^{2}$$

$$= \frac{1}{16\tau_{2}^{2}} K \left[\|\Delta\|_{F} + \max_{k \in [K]} |\left(\hat{u}_{k}/u_{k} - 1\right)| \|\hat{\Theta}\|_{F} \right]^{2}.$$

$$(7)$$

Combining the inequality (6), (7) with the assumption that $G(\hat{u}, \hat{\Theta}) \leq 0$, we have

$$\begin{split} 0 &\geq I_2 - |I_1| \\ &\geq \frac{1}{16\tau_2^2} K \left[\|\Delta\|_F + \max_{k \in [K]} |\left(\hat{u}_k/u_k - 1\right)| \left\| \hat{\Theta} \right\|_F \right]^2 \\ &- \sqrt{K} C \sqrt{\frac{p^2 \log p}{n}} \left[\|\Delta\|_F + \max_{k \in [K]} |\left(\hat{u}_k/u_k - 1\right)| \left\| \hat{\Theta} \right\|_F \right], \end{split}$$

which implies that

$$K\left[\|\Delta\|_F + \max_{k \in [K]} |\left(\hat{u}_k/u_k - 1\right)| \left\|\hat{\Theta}\right\|_F\right] \le 16\tau_2^2 \sqrt{K}C\sqrt{\frac{p^2 \log p}{n}}.$$

Last, note that

$$\sum_{k=1}^{K} \|\Delta_k / u_k\|_F \leq K \|\Delta\|_F + \sum_{k=1}^{K} \left(\hat{u}_k / u_k - 1\right) | \left\| \hat{\Theta} \right\|_F \leq K \left[\|\Delta\|_F + \max_{k \in [K]} |\left(\hat{u}_k / u_k - 1\right)| \left\| \hat{\Theta} \right\|_F \right],$$

and

$$\sum_{k=1}^{K} \|\Delta_k\|_F \le \sum_{k=1}^{K} u_k \sum_{k=1}^{K} \|\Delta_k / u_k\|_F \le K \sum_{k=1}^{K} \|\Delta_k / u_k\|_F,$$

where the first inequality follows the fact that $\sum_{k=1}^K a_k b_k \leq \sum_{k=1}^K |a_k| \sum_{k=1}^K |b_k|$, and the second inequality follows by the Cauchy-Schwartz inequality that $\left(\sum_{k=1}^K u_k\right)^2 \leq K \sum_{k=1}^K u_k^2 = K^2$.

Finally, we have the accuracy rate

$$\sum_{k=1}^{K} \|\Delta_k\|_F \le 16\tau_2^2 K^{3/2} C \sqrt{\frac{p^2 \log p}{n}}.$$