

# Summary for Probability Theory

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## 1 Preliminary

- *DeMorgan's Laws:* Let  $\{A_i\}_{i=1}^{\infty}$  be a collection of set. Then,  $(\bigcup_{i=1}^{\infty} A_i)^c = \bigcap_{i=1}^{\infty} A_i^c$  and  $(\bigcap_{i=1}^{\infty} A_i)^c = \bigcup_{i=1}^{\infty} A_i^c$ .
- *Some set operations:* Suppose  $A, B$  are two sets. (1)  $A - B = A \cap B^c$ ; (2)  $\bigcup_{i=1}^{\infty} A_i = \{x : x \in A_i \text{ for some } i\}$ ,  $\bigcap_{i=1}^{\infty} A_i = \{x : x \in A_i \text{ for all } i\}$ .

## 2 Single variable

### 2.1 Probability and conditional probability

**Definition 1** (*Sample space*). The set  $S$  containing all possible outcomes is called the sample space.

**Definition 2** ( *$\sigma$ -field*). A collection  $\mathcal{F}$  of subsets of a sample space  $S$  is called a  $\sigma$ -field (or  $\sigma$ -algebra) if and only if (**iff**) it has the following properties:

- (1) The empty set  $\emptyset \in \mathcal{F}$ ;
- (2) If  $A \in \mathcal{F}$ , then the complement  $A^c \in \mathcal{F}$ ;
- (3) If  $A_i \in \mathcal{F}, i = 1, 2, \dots$ , then their union  $\bigcup_i A_i \in \mathcal{F}$ .

If  $A \in \mathcal{F}$ , then  $A$  is called an *event*.

**Definition 3** (*Measure and probability*). A set function  $v$  defined on a  $\sigma$ -field  $\mathcal{F}$  is called a measure **iff** it has the following properties:

- (1)  $0 \leq v(A) \leq \infty$  for any  $A \in \mathcal{F}$ ;
- (2)  $v(\emptyset) = 0$ ;
- (3) If  $A_i \in \mathcal{F}, i = 1, 2, \dots$  and  $A_i$ 's are disjoint, i.e.  $A_i \cap A_j = \emptyset, \forall i \neq j$ , then

$$v\left(\bigcup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} v(A_i).$$

If  $v(\mathcal{F}) = 1$ , then  $v$  is a probability defined on  $\mathcal{F}$  and we use notation  $P$  instead of  $v$ .

**Theorem 2.1** (Probability). Let the sample space  $S = \{s_1, s_2, \dots\}$  and  $\mathcal{F}$  be all subsets of  $S$ . Let  $p_1, p_2, \dots$  be non-negative numbers that  $\sum_i p_i = 1$ . The following defines a probability on  $\mathcal{F}$

$$P(A) = \sum_{i: s_i \in A} p_i, \quad A \in \mathcal{F}.$$

**Theorem 2.2** (Properties of probability). *Let  $P$  be a probability,  $A, B$  be events and  $\{A_i\}_{i=1}^{\infty}$  be a collection of event. Let  $\{C_i\}_{i=1}^{\infty}$  be a partition of sample space  $S$ , i.e.  $C_i \cap C_j, \forall i \neq j$  and  $\bigcup_{i=1}^{\infty} C_i = S$ . Then,*

1.  $P(A) \leq 1; P(A^c) = 1 - P(A); P(A) = P(A \cap B) + P(A \cap B^c);$

2. *If  $A \subset B$ , then  $P(A) \leq P(B);$*

3. *(General addition formula)*

$$P\left(\bigcup_{i=1}^n A_i\right) = \sum_{i=1}^n P(A_i) - \sum_{i < j} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j \cap A_k) + \cdots + (-1)^{n-1} P(A_1 \cap \cdots \cap A_n);$$

4.  $P(A) = \sum_{i=1}^{\infty} P(A \cap C_i);$

5. *(Boole's inequality)  $P\left(\bigcup_{i=1}^{\infty} A_i\right) \leq \sum_{i=1}^{\infty} P(A_i);$*

6. *(Bonferroni's inequality)  $P\left(\bigcap_{i=1}^n A_i\right) \geq \sum_{i=1}^n P(A_i) - (n - 1).$*

**Definition 4** (Conditional Probability). If  $A$  and  $B$  are events with  $P(B) > 0$ , then the conditional probability of  $A$  given  $B$  is

$$P(A|B) = \frac{P(A \cap B)}{P(B)}.$$

For convenience, we define  $P(A|B) = 0$  when  $P(B) = 0$ .

**Theorem 2.3** (Useful formulas for conditional probability). *Let  $A, \{A_i\}_{i=1}^{\infty}, B, \{B_i\}_{i=1}^{\infty}, C$  be events. Then we have:*

1.  $P(A|B) = \frac{P(A)P(B|A)}{P(B)};$

2.  $P(A^c|B) = 1 - P(A|B); P(A \cup C|B) = P(A|B) + P(C|B) - P(A \cap C|B);$

3.  $P\left(\bigcap_{i=1}^n A_i\right) = P(A_1)P(A_2|A_1) \cdots P(A_n|\bigcap_{i=1}^{n-1} A_i);$

4. *If  $\{B_i\}_{i=1}^{\infty}$  is a partition of  $S$ ,  $P(A) = \sum_{i=1}^{\infty} P(B_i)P(A|B_i).$*

**Theorem 2.4** (Bayes formula). *Let  $A$  be an event and  $\{B_i\}_{i=1}^{\infty}$  be a partition of  $S$ . Then,*

$$P(B_i|A) = \frac{P(A|B_i)P(B_i)}{\sum_{j=1}^{\infty} P(A|B_j)P(B_j)}.$$

**Definition 5** (Independence). Two events  $A, B$  are independent **iff**

$$P(A \cap B) = P(A)P(B) \quad \text{or} \quad P(A|B) = P(A) \quad \text{or} \quad P(B|A) = P(B).$$

If  $A, B$  are independent, then the following pairs are also independent:  $A$  and  $B^c$ ,  $A^c$  and  $B$ ,  $A^c$  and  $B^c$ .

**Definition 6** (*Mutual and pairwise independence*). A collection of events  $A_1, \dots, A_n$  are mutually independent **iff** for any sub-collection  $A_{i_1}, \dots, A_{i_k}$ ,

$$P(A_{i_1} \cap \dots \cap A_{i_k}) = P(A_{i_1}) \dots P(A_{i_k}).$$

The events  $A_1, \dots, A_n$  are pairwise independent **iff**  $A_i$  and  $A_j$  are independent for all  $i \neq j$ . Mutual independence is stronger than pairwise independence.

**Definition 7** (*Conditional independence*). Events  $A$  and  $B$  are conditionally independent given event  $C$  **iff**

$$P(A \cap B|C) = P(A|C)P(B|C).$$

Let  $A, B, C$  are events. Independence does not imply conditional independence:

$$P(A \cap B) = P(A)P(B) \not\Rightarrow P(A \cap B|C) = P(A|C)P(B|C).$$

Conditional independence does not imply independence:

$$P(A \cap B|C) = P(A|C)P(B|C) \not\Rightarrow P(A \cap B) = P(A)P(B).$$

Mutual independence implies conditional independence:

$$A, B, C \text{ mutually independent} \Rightarrow P(A \cap B|C) = P(A|C)P(B|C).$$

## 2.2 Random variable and distribution

**Definition 8** (*Random variable and distribution*). A random variable  $X$  is a function from  $S$  to  $\mathbb{R}$  such that, for any Borel set  $\mathcal{B} \subset \mathbb{R}$ ,

$$\{X \in \mathcal{B}\} = \{\omega \in S : X(\omega) \in \mathcal{B}\}.$$

The induced probability of  $X$  is

$$P_X(\mathcal{B}) = P(X \in \mathcal{B}) = P(\omega \in \{\omega \in S : X(\omega) \in \mathcal{B}\}).$$

The probability  $P_X$  is called the distribution of  $X$ .

**Definition 9** (*Cumulative distribution function(cdf)*). The cdf of a random variable  $X$ , denoted by  $F_X(x)$ , is defined as

$$F_X(x) = P(X \leq x), \quad x \in \mathbb{R}.$$

**Theorem 2.5** (Cdf). The function  $F(x)$  is a cdf **iff**

1.  $\lim_{x \rightarrow -\infty} F(x) = 0$  and  $\lim_{x \rightarrow \infty} F(x) = 1$ ;
2.  $F(x)$  is non-decreasing in  $x$ ;
3.  $F(x)$  is right-continuous:  $\lim_{\epsilon > 0, \epsilon \rightarrow 0} F(x + \epsilon) = F(x), \quad \forall x \in \mathbb{R}.$

**Definition 10** (*Continuity of random variable*). A random variable  $X$  is continuous if  $F_X(x)$  is continuous in  $x$ . A random variable  $x$  is discrete if  $F_X(x)$  is a step function of  $x$ .

Note that the continuity of a random variable depends on the cdf rather than pdf or pmf. There are random variable's that are mixtures of these two types.

**Definition 11** (*Probability mass function(pmf)*). The pmf of a discrete random variable  $X$  is

$$f_X(x) = P(X = x), \quad x \in \mathbb{R}.$$

The cdf of  $X$ ,  $F_X(x) = P(X \leq x) = \sum_{k \leq x} f_X(k)$ .

**Definition 12** (*Probability density function(pdf)*). The pdf of a continuous random variable  $X$  is the function  $f_X(x)$  such that

$$F_X(x) = \int_{-\infty}^x f_X(t)dt, \quad x \in \mathbb{R},$$

if  $f_X(x)$  exists. The continuous random variable  $X$  has a pdf **iff**  $F_X$  is absolutely continuous. If  $f$  is a pdf, the set  $\{x : f(x) > 0\}$  is called its support.

If  $F_X$  is differentiable, then  $f_X(x) = \frac{d}{dx}F_X(x)$ .

**Theorem 2.6** (Pdf). A function  $f(x)$  is a pdf **iff**:

1.  $f(x) \geq 0, \quad \forall x \in \mathbb{R};$
2.  $\int_{-\infty}^{\infty} f(x)dx = 1.$

How to find pdf given cdf? (1)  $f_X(x) = F'_X(x)$  for  $x$  at which  $F_X$  is differentiable; (2)  $f_X(x)$  can be any  $c \geq 0$  for  $x$  at which  $F_X$  is not differentiable.

### 2.3 Transformation

Let  $X$  be a random variable and  $Y = g(X)$ , where  $g$  is function  $\mathbb{R} \mapsto \mathcal{Y}$  and  $\mathcal{Y}$  is the domain of  $Y$ . For any  $A \in \mathcal{Y}$ ,

$$P(Y \in A) = P(g(X) \in A) = P(X \in g^{-1}(A)), \quad \text{where } g^{-1}(A) = \{x : g(x) \in A\}.$$

Given  $F_X$  or  $f_X$ , we want to obtain  $f_Y(y)$ . If  $X$  is discrete, then

$$f_Y(y) = \sum_{x \in g^{-1}(\{y\})} P(X = x)$$

If  $f_X$  is continuous and  $g$  is a continuously differentiable monotone function, then

$$f_Y(y) = \begin{cases} f_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right| & y \in \mathcal{Y} \\ 0 & \text{otherwise} \end{cases}$$

**Theorem 2.7** (Transformation for continuous random variable). Let  $X$  be a continuous random variable with pdf  $f_X$ . Suppose there are disjoint  $\{A_i\}_{i=1}^k$  such that  $P(X \in \bigcup_{i=1}^k A_i) = 1$ , and  $f_X$  is continuous on each  $A_i, i \in [k]$ . There are functions  $g_1(x), \dots, g_k(x)$  defined on  $A_i, i \in [k]$  respectively, satisfying

1.  $g(x) = g_i(x), \forall x \in A_i;$
2.  $g_i(x)$  is strictly monotone on  $A_i;$
3. The set  $\mathcal{Y} = \{y : y = g_i(x) \text{ for some } x \in A_i\}$  is the same for each  $i;$
4.  $g_i^{-1}(y)$  has a continuous derivative on  $\mathcal{Y}$  for each  $i.$

Then

$$f_Y(y) = \begin{cases} \sum_{i=1}^k f_X(g_i^{-1}(y)) \left| \frac{d}{dy} g_i^{-1}(y) \right| & y \in \mathcal{Y} \\ 0 & \text{otherwise} \end{cases}$$

**Example 1.** Suppose a random variable  $X \sim N(0, 1)$ . Obtain the distribution of  $Y = |X|$ .

*Proof.* Let  $A_1 = (-\infty, 0)$ ,  $A_2 = (0, +\infty)$  and  $\mathcal{Y} = (0, +\infty)$ . On  $A_1$ ,  $g_1(x) = x$  and  $g_1^{-1}(x) = x$ . On  $A_2$ ,  $g_2(x) = -x$  and  $g_2^{-1}(x) = -x$ .

By theorem 2.7,

$$f_Y(y) = \frac{1}{\sqrt{2\pi}}e^{-x^2/2} + \frac{1}{\sqrt{2\pi}}e^{-x^2/2} = \sqrt{\frac{2}{\pi}}e^{-x^2/2}.$$

□

## 2.4 Expectation

**Definition 13** (*Expectation*). The expected value or mean of a random variable  $g(x)$  is

$$\mathbb{E}[g(x)] = \begin{cases} \int_{-\infty}^{+\infty} g(x)f_X(x)dx \stackrel{Y=g(x)}{=} \int_{-\infty}^{+\infty} yf_Y(y)dy & \text{if } X \text{ has pdf } f_X \\ \sum_x g(x)f_X(x) \stackrel{Y=g(x)}{=} \sum_x yf_Y(y) & \text{if } X \text{ has pmf } f_X \end{cases}$$

provided that  $\mathbb{E}[|g(x)|] < \infty$ . Otherwise, the expected value of  $g(X)$  does not exist.

**Theorem 2.8** (Properties of expectation). Let  $X, Y$  be random variables whose expectations exist. Let  $a, b, c$  be constants.

1.  $\mathbb{E}(aX + bY + c) = a\mathbb{E}(X) + b\mathbb{E}(Y) + c$ ;
2. If  $X \geq Y$ , then  $\mathbb{E}(X) \geq \mathbb{E}(Y)$ .

**Theorem 2.9** (Relationship between expectation and cdf). Let  $F$  be the cdf of a random variable  $X$ . If  $X$  has a pdf or pmf, then,

$$\mathbb{E}[|X|] = \int_0^{+\infty} [1 - F(x)]dx + \int_{-\infty}^0 F(x)dx,$$

and  $\mathbb{E}[|X|] < +\infty$  **iff** both integrals are finite. In case where  $\mathbb{E}[|X|] < +\infty$ ,

$$\mathbb{E}[X] = \int_0^{+\infty} [1 - F(x)]dx - \int_{-\infty}^0 F(x)dx.$$

*Proof.* Without the loss of generality, suppose random variable  $X$  has a pdf. Let  $f$  be the pdf of  $X$ . We have

$$\begin{aligned} \mathbb{E}[|X|] &= \int_0^{+\infty} xf(x)dx - \int_{-\infty}^0 xf(x)dx \\ &= \int_0^{+\infty} \int_0^x f(x)dt dx + \int_{-\infty}^0 \int_x^0 f(x)dt dx \\ &= \int_0^{+\infty} \int_t^{+\infty} f(x)dx dt + \int_{-\infty}^0 \int_{-\infty}^t f(x)dx dt \\ &= \int_0^{+\infty} [1 - F(t)]dt + \int_{-\infty}^0 F(t)dt. \end{aligned}$$

Therefore,  $\mathbb{E}[|X|] < +\infty$  **iff** the two integrals are finite. Similarly,

$$\begin{aligned} \mathbb{E}[X] &= \int_0^{+\infty} xf(x)dx + \int_{-\infty}^0 xf(x)dx \\ &= \int_0^{+\infty} [1 - F(t)]dt - \int_{-\infty}^0 F(t)dt. \end{aligned}$$

□

**Corollary 1** (Relationship between expectation and cdf). Let  $F$  be the cdf of a random variable  $X$ . If  $X$  has a pdf or pmf, then,

$$\mathbb{E}[|X|] = \int_0^{+\infty} P(X > x) + P(-X \leq x) dx,$$

and

$$\sum_{n=1}^{\infty} P(|X| \geq n) \leq \mathbb{E}[|X|] \leq 1 + \sum_{n=1}^{\infty} P(|X| \geq n).$$

*Proof.* By the proof of theorem 2.9,

$$\begin{aligned} \mathbb{E}[|X|] &= \int_0^{+\infty} [1 - F(t)] dt + \int_{-\infty}^0 F(t) dt \\ &= \int_0^{+\infty} [1 - F(t)] dt + \int_0^{+\infty} F(-t) dt \\ &= \int_0^{+\infty} P(X > t) + P(-X \leq t) dt \end{aligned}$$

To show  $\sum_{n=1}^{\infty} P(|X| \geq n) \leq \mathbb{E}[|X|]$ , we have

$$\begin{aligned} \mathbb{E}[|X|] &= \int_0^{+\infty} P(X > t) + P(-X \leq t) dt \\ &\geq \int_0^{+\infty} P(|X| > t) dt \\ &= \sum_{n=0}^{+\infty} \int_n^{(n+1)} P(|X| > t) dt \\ &\geq \sum_{n=0}^{+\infty} \int_n^{(n+1)} P(|X| \geq (n+1)) dt \\ &= \sum_{n=1}^{+\infty} P(|X| \geq n) \end{aligned}$$

To show  $\mathbb{E}[|X|] \leq 1 + \sum_{n=1}^{\infty} P(|X| \geq n)$ , we have

$$\begin{aligned} \mathbb{E}[|X|] &= \int_0^{+\infty} P(X > t) + P(-X \leq t) dt \\ &\leq \int_0^{+\infty} P(|X| \geq t) dt \\ &= \sum_{n=0}^{+\infty} \int_n^{(n+1)} P(|X| \geq t) dt \\ &\leq \sum_{n=0}^{+\infty} \int_n^{(n+1)} P(|X| \geq n) dt \\ &\leq 1 + \sum_{n=1}^{+\infty} P(|X| \geq n). \end{aligned}$$

□