# STATS310B – Lecture 4

### Sourav Chatterjee Scribed by Michael Howes

01/13/22

### Contents

L	Stopped $\sigma$ -algebras	1
2	Gambler's ruin	3
3	Sub-martingales and super-martingales	5

## 1 Stopped $\sigma$ -algebras

Recall the following definition from the previous lecture.

**Definition 1.** Given a filtration  $\{\mathcal{F}_n\}_{n\geq 0}$  and a stopping time T. The stopped  $\sigma$ -algebra is defined to be

$$\mathcal{F}_T = \{ A \in \mathcal{F} : A \cap \{ T = n \} \in \mathcal{F}_n \text{ for all } n \}.$$

We saw that  $\mathcal{F}_T$  is indeed a  $\sigma$ -algebra. Informally, a typical event in  $\mathcal{F}_T$  is one that depends on  $\mathcal{F}_n$  for  $n \leq T$ .

**Example 1.** Let  $S_n$  be a simple symmetric random walk on  $\mathbb{Z}$  with  $S_0 = 0$ . Let  $\mathcal{F}_n = \sigma(S_0, S_1, \ldots, S_n)$ . We saw previously that for  $a, b \in \mathbb{Z}$  with a < 0 < b, the random variable  $T = \inf\{n : S_n = a \text{ or } S_n = b\}$  is a stopping time with respect to  $\{F_n\}_{n\geq 0}$ . We also claimed that the event  $A = \{S_k \geq 0, \text{ for } k \leq T\}$  was in  $\mathcal{F}_T$ . To see why this is true, take any n. Then,

$$A \cap \{T = n\} = \left(\bigcap_{k=0}^{n} \{S_k \ge 0\}\right) \cap \{T = n\}.$$

Since  $S_k$  is  $\mathcal{F}_n$  measurable for  $n \leq k$  and since T is a stopping time, both of the above events are in  $\mathcal{F}_n$ . Thus, the above intersection is in  $\mathcal{F}_n$  and so  $A \in \mathcal{F}_T$ .

**Proposition 1.** If S and T are stopping times with respect to  $\{F_n\}_{n\geq 0}$  and  $S\leq T$  always, then  $\mathcal{F}_S\subseteq \mathcal{F}_T$ .

*Proof.* Let  $A \in \mathcal{F}_S$  and take any  $n \in \mathbb{N}$ . Then

$$\begin{split} A \cap \{T=n\} &= A \cap \{S \leq n\} \cap \{T=n\} \\ &= \left(\bigcup_{k=0}^n A \cap \{S=k\}\right) \cap \{T=n\}. \end{split}$$

For each k we have  $A \cap \{S = k\} \in \mathcal{F}_k \subseteq \mathcal{F}_n$  and so  $\bigcup_{k=0}^n A \cap \{S = k\} \in \mathcal{F}_n$ . Also,  $\{T = n\} \in \mathcal{F}_n$  and thus  $A \cap \{T = n\} \in \mathcal{F}_n$  and so  $A \in \mathcal{F}_T$ .

Note that for the above argument, we need  $S \leq T$  always. The result may not be true if  $S \leq T$  almost surely.

**Example 2.** Suppose that  $N \in \{0, 1, ...\}$ . Then the constant random variable T = N is a stopping time with respect to any filtration  $\{\mathcal{F}_n\}_{n>0}$ . The stopped  $\sigma$ -algebra is  $\mathcal{F}_N$ .

**Definition 2.** A stopping time T is bounded if there is an integer N such that T < N always.

Note that if  $T \leq N$  always, then by proposition 1 and example 2, we have  $\mathcal{F}_T \subseteq \mathcal{F}_N$ . We will also define stopped random variables.

**Definition 3.** Let  $\{X_n\}_{n\geq 0}$  be a sequence of random variables adapted to  $\{\mathcal{F}_n\}_{n\geq 0}$ . Let T be a stopping time with respect to  $\{\mathcal{F}_n\}_{n\geq 0}$  such that  $\mathbb{P}(T<\infty)=1$ . The random variables  $X_T$  is defined on the set where  $T<\infty$  by

$$X_T(\omega) = X_{T(\omega)}(\omega).$$

Note that we can define  $X_T$  on all of  $\Omega$  by taking  $X_T(\omega)$  equal to any value on the set  $\{T = \infty\}$  which has probability zero. Note that the random variable  $X_T$  may be very different to each of the random variables  $X_n$ . For example if  $S_n$  is the SSRW on  $\mathbb{Z}$  with  $S_0 = 0$ , then the support of  $S_n$  is the set of integers in  $\{-n, -n+1, \ldots, n-1, n\}$  with the same parity as n. However, if  $T = \inf\{n : S_n = a \text{ or } S_n = b\}$ , then  $S_T \in \{a, b\}$  almost surely (we will see shortly that  $\mathbb{P}(T < \infty)$ ) = 1.

We will now state a useful theorem.

**Theorem 1** (Optional stopping theorem). Let  $\{X_n\}_{n\geq 0}$  be a martingale adapted to a filtration  $\{\mathcal{F}_n\}_{n\geq 0}$ . Let S and T be bounded stopping times with respect to  $\{F_n\}$  such that  $S\leq T$  always. Then  $X_S$  and  $X_T$  are integrable and  $\mathbb{E}(X_T|\mathcal{F}_S)=X_S$  almost surely. In particular  $\mathbb{E}[X_T]=\mathbb{E}[X_0]$ .

Note that when T and S are both constant this reduces to the result that for all n, k such that  $k \leq n$ ,

$$\mathbb{E}(X_n|\mathcal{F}_k) = X_k$$
 and  $\mathbb{E}[X_n] = \mathbb{E}[X_0]$ .

The above results can be proved directly by using induction and the tower property. We will now prove the optional stopping theorem.

*Proof.* Let N be an integer such that  $S \leq T \leq N$  always. Note that

$$|X_S|, |X_T| \le |X_0| + |X_1| + \ldots + |X_N|.$$

Thus,  $X_S$  and  $X_T$  are both integrable. We will next show that  $X_S$  is  $\mathcal{F}_S$  measurable. Fix any  $c \in \mathbb{R}$ , we wish to show that  $\{X_S \leq c\} \in \mathcal{F}_S$  which requires  $\{X_S \leq c\} \cap \{S = n\} \in \mathcal{F}_n$  for every n.

$$\{X_S < c\} \cap \{S = n\} = \{X_n < c\} \cap \{S = n\}.$$

Both  $\{S=n\}$  and  $\{X_n \leq c\}$  are in  $\mathcal{F}_n$  and so  $\{X_S \leq c\} \in \mathcal{F}_S$ . Now let  $A \in \mathcal{F}_S$ . We wish to show

$$\mathbb{E}[X_T \mathbf{1}_A] = \mathbb{E}[X_S \mathbf{1}_A].$$

Note that

$$\mathbb{E}[X_N \mathbf{1}_A] = \mathbb{E}\left[X_N \sum_{n=0}^N \mathbf{1}_{\{T=n\}} \mathbf{1}_A\right]$$
$$= \sum_{n=0}^N \mathbb{E}[X_N \mathbf{1}_{\{T=n\}} \mathbf{1}_A]$$
$$= \sum_{n=0}^N \mathbb{E}[X_N \mathbf{1}_{\{T=n\} \cap A}].$$

We know that  $\{T = n\} \cap A \in \mathcal{F}_n \text{ since } A \in \mathcal{F}_S \subseteq \mathcal{F}_T$ . Thus

$$\begin{split} \sum_{n=0}^{N} \mathbb{E}[X_N \mathbf{1}_{\{T=n\} \cap A}] &= \sum_{n=0}^{N} \mathbb{E}[\mathbb{E}(X_N | \mathcal{F}_n) \mathbf{1}_{\{T=n\} \cap A}] \\ &= \sum_{n=0}^{N} \mathbb{E}[X_n \mathbf{1}_{\{T=n\} \cap A}] \\ &= \mathbb{E}\left[\sum_{n=0}^{N} X_n \mathbf{1}_{\{T=n\}} \mathbf{1}_A\right] \\ &= \mathbb{E}[X_T \mathbf{1}_A]. \end{split}$$

By the same argument we have  $\mathbb{E}[X_N \mathbf{1}_A] = \mathbb{E}[X_S \mathbf{1}_A]$ . Thus,  $\mathbb{E}[X_T \mathbf{1}_A] = \mathbb{E}[X_S \mathbf{1}_A]$  and thus

$$\mathbb{E}(X_T|\mathcal{F}_S) = X_S.$$

If we take S = 0, then we get

$$\mathbb{E}[X_T] = \mathbb{E}[\mathbb{E}(X_T | \mathcal{F}_0)] = \mathbb{E}[X_0].$$

The requirement that S and T are bounded stopping times may seem restrictive, but the result is not true if we do not have this assumption. Fortunately we can often approximate an unbounded stopping time by a sequence of stopping times.

**Proposition 2.** Let T be a stopping time with respect to a filtration  $\{F_k\}_{k\geq 0}$ . Then for every  $n\in\mathbb{N}$ , the random variable

$$T \wedge n = \min\{T, n\},\$$

is a stopping time bounded by n. Furthermore,  $T \wedge n \to T$  on the set where  $T < \infty$ .

**Proposition 3.** To see that  $T \wedge n$  is a stopping time, note that

$$\{T \wedge n = k\} = \begin{cases} \{T = k\} & \text{if } k < n, \\ \{T \ge k\} & \text{if } k = n, \\ \emptyset & \text{if } k > n. \end{cases}$$

Thus, we immediately see that  $\{T \land n = k\} \in \mathcal{F}_k$  when  $k \neq n$ . When k = n, note that

$${T \ge k} = {T < k}^C = \left(\bigcup_{j=1}^{k-1} {T = j}\right)^C.$$

We know that  $\{T = j\} \in \mathcal{F}_j \subseteq \mathcal{F}_k$  for all j < k and so  $\{T \ge k\} \in \mathcal{F}_k$ .

### 2 Gambler's ruin

We will now study an example which shows the usefulness of the optional stopping theorem. Let  $\{S_n\}_{n\geq 0}$  be a SSRW with  $S_0=0$  and as before let  $T=\inf\{n: S_n=a \text{ or } S_n=b\}$ . We are interested in

$$\mathbb{P}(T<\infty,S_T=b).$$

We will soon see that  $\mathbb{P}(T<\infty)=1$ . Thus, the probability we are interested in is

$$\mathbb{P}(S_T = b).$$

Remark 1. This question is related to the gambler's ruin. Consider a gambler with  $x \in \{1, 2, ...\}$  dollars. The gambler repeatedly makes bets where they can either win or lose \$1. They continue playing until they make y > x dollars, or they go broke. The probability that they don't go broke is equal to the probability that they end with y dollars. This is equal to  $\mathbb{P}(S_T = b)$  when a = -x and b = y - x.

Note that  $S_T$  only takes the values a and b. Thus,

$$\mathbb{E}[S_T] = a\mathbb{P}(S_T = a) + b\mathbb{P}(S_T = b) = a(1 - \mathbb{P}(S_T = b)) + b\mathbb{P}(S_T = b) = a + (b - a)\mathbb{P}(S_T = b).$$

Thus, if we can calculate  $\mathbb{E}[S_T]$ , then we will know  $\mathbb{P}(S_T = b)$  since

$$\mathbb{P}(S_T = b) = \frac{\mathbb{E}[S_T] - a}{b - a}.$$

We wish to use the optional stopping theorem to conclude that  $\mathbb{E}[S_T] = \mathbb{E}[S_0] = 0$ . The stopping time  $T \wedge n$  is bounded for each n, and so we know that  $\mathbb{E}[S_{T \wedge n}] = 0$ . We also have  $|S_{T \wedge n}| \leq \max\{-a, b\}$ . Thus, it suffices to show that  $T \wedge n \to T$  almost surely. To do this, we need to show that  $\mathbb{P}(T < \infty) = 1$ .

**Proposition 4.** With  $S_n$  and T as above,  $\mathbb{P}(T < \infty) = 1$ .

*Proof.* We can divide the set  $\{1,2,3,\ldots\}$  into infinitely many blocks  $B_j$  of consecutive integers such that  $B_j$  contains |a|+b integers. Let  $A_j$  be the event that  $S_k-S_{k-1}=1$  for all  $k\in B_j$ . Each  $A_j$  has probability  $\mathbb{P}(A_j)=2^{-|a|-b}$  and the events  $A_j$  are independent since the sets  $B_j$  are disjoint. Thus, by the second Borel–Cantelli lemma,

$$\mathbb{P}(A_i, \text{ infinitely often}) = 1.$$

For each  $A_j$  we have  $A_j \subseteq \{T < \infty\}$ . To see this let  $k_1$  be the minimum of  $B_j$ . Suppose that  $\omega \in A_j$ . If  $T(\omega) > k_1$ , then  $S_{k_1}(\omega) \ge a$ . Since  $\omega \in A_j$ ,

$$S_{k_1+|a|+b}(\omega) = S_{k_1}(\omega) + |a| + b = b.$$

Thus,  $T(\omega) \leq k_1 + |a| + b < \infty$ . We thus have  $A_i \subseteq \{T < \infty\}$ , and so

$$\mathbb{P}(T < \infty) \geq \mathbb{P}(A_i, \text{ infinitely often}) = 1.$$

Thus,  $T \wedge n \to T$  almost surely and furthermore more  $S_{T \wedge n} \to S_T$ . Since  $S_{T \wedge n}$  is uniformly bounded we have

$$\mathbb{E}[S_T] = \lim_{n \to \infty} \mathbb{E}[S_{T \wedge n}] = \lim_{n \to \infty} \mathbb{E}[S_0] = 0.$$

We thus have

$$\mathbb{P}(S_T = b) = \frac{-a}{-a+b}.$$

In terms of the gambler's ruin this equal  $\frac{x}{y}$ . So if the gambler starts with x=900 dollars, they have a 90% of making y=1000 dollars before going broke. But as we see, the gambler will have to be very patient. We can also use the optional stopping theorem to calculate the expected value of T as well. Recall that

$$M_n = S_n^2 - n,$$

is a martingale since  $Var(X_n) = 1$ . And thus

$$\mathbb{E}[M_{T \wedge n}] = \mathbb{E}[S_{T \wedge n}^2] - \mathbb{E}[T \wedge n].$$

Page 4

The random variables  $S_{T\wedge n}^2$  are uniformly bounded by  $\max\{a^2,b^2\}$  and converge almost surely to  $S_T^2$ . Furthermore,  $T\wedge n$  are all non-negative, and they converge almost surely to T. Therefore, by the dominated and monotone convergence theorem, we have

$$\mathbb{E}[S_T^2] = \lim_{n \to \infty} \mathbb{E}[S_{T \wedge n}^2] = \lim_{n \to \infty} \mathbb{E}[T \wedge n] = \mathbb{E}[T].$$

We also know that

$$\mathbb{E}[S_T^2] = a^2 \mathbb{P}(S_T = a) + b^2 \mathbb{P}(S_T = b)$$

$$= \frac{a^2 b}{b - a} + \frac{-b^2 a}{b - a}$$

$$= \frac{ab(a - b)}{b - a}$$

$$= -ab$$

In the gambler's ruin, this equals x(y-x). So to get from \$900 to \$1000 or go broke, the gambler will have to wait on average 90000 turns.

# 3 Sub-martingales and super-martingales

We will now state two definitions that generalize martingales.

**Definition 4.** Let  $\{F_n\}_{n\geq 0}$  be a filtration and let  $\{X_n\}_{n\geq 0}$  be a sequence of adapted, integrable random variables. Then,

1.  $\{X_n\}_{n>0}$  is a *sub-martingale* if for every n,

$$\mathbb{E}(X_{n+1}|\mathcal{F}_n) \geq X_n$$
. a.s

2.  $\{X_n\}_{n>0}$  is a super-martingale if for every n,

$$\mathbb{E}(X_{n+1}|\mathcal{F}_n) \leq X_n$$
. a.s

**Remark 2.** One can use Jensen's inequality to make sub-martingales out of martingales. Let  $\{M_n\}_{n\geq 0}$  be a martingale and let  $\phi$  be a convex function such that  $X_n = \phi(M_n)$  is integrable for every  $n \geq 0$ . Then  $\{X_n\}_{n\geq 0}$  is a sub-martingale since, almost surely

$$\mathbb{E}(X_{n+1}|\mathcal{F}_n) = \mathbb{E}(\phi(M_{n+1})|\mathcal{F}_n) > \phi(\mathbb{E}(M_{n+1}|\mathcal{F}_n)) = \phi(M_n) = X_n.$$