

Wiki: https://en.wikipedia.org/wiki/Least_squares

- When the observations come from an exponential family and mild conditions are satisfied, least-squares estimates and maximum-likelihood estimates are identical. ^[1] The method of least squares can also be derived as a method of moments estimator.

1 Least Square

$$S = \sum_{i=0}^{N-1} r_i^2 = \sum_{i=0}^{N-1} [y_i - f(x_i, \beta)]^2 \quad (1-1)$$

The minimum of the sum of squares is found by setting the gradient to zero. Since the model contains m parameters, there are m gradient equations:

$$\frac{\partial S}{\partial \beta_j} = -2 \sum_{i=0}^{N-1} \left([y_i - f(x_i, \beta)] \cdot \frac{\partial f(x_i, \beta)}{\partial \beta_j} \right) = 0 \quad (1-2)$$

1.1 Linear Least Square

A regression model is a linear one when the model comprises a linear combination of the parameters, i.e.

$$f(x, \beta) = \sum_{j=0}^{m-1} \beta_j \cdot \phi_j(x) \quad (1-3)$$

Letting

$$X_{i,j} = \frac{\partial f(x_i, \beta)}{\partial \beta_j} = \phi_j(x_i) \quad (1-4)$$

Then equation (1-2) becomes to

$$\begin{aligned} \sum_{i=0}^{N-1} ([y_i - f(x_i, \beta)] \cdot X_{i,j}) &= 0 \\ \Leftrightarrow \sum_{i=0}^{N-1} y_i \cdot X_{i,j} &= \sum_{i=0}^{N-1} f(x_i, \beta) \cdot X_{i,j} \\ \Leftrightarrow \sum_{i=0}^{N-1} y_i \cdot X_{i,j} &= \sum_{i=0}^{N-1} \left[\left(\sum_{j=0}^{m-1} \beta_j \cdot X_{i,j} \right) \cdot X_{i,j} \right] \\ \Leftrightarrow \mathbf{X}_{:,j}^T \cdot \vec{y} &= \sum_{i=0}^{N-1} [X_{i,j} \cdot (\mathbf{X}_{i,:} \cdot \hat{\beta})] \\ \Leftrightarrow \mathbf{X}_{:,j}^T \cdot \vec{y} &= \mathbf{X}_{:,j}^T \cdot \mathbf{X}_{i,:} \cdot \hat{\beta} \\ \Leftrightarrow \mathbf{X}^T \cdot \vec{y} &= \mathbf{X}^T \cdot \mathbf{X} \cdot \hat{\beta} \end{aligned} \quad (1-5)$$

Then the parameters $\hat{\beta}$ can be estimated by

$$\hat{\beta} = (\mathbf{X}^T \cdot \mathbf{X})^{-1} \cdot \mathbf{X}^T \cdot \vec{y} \quad (1-6)$$

1.2 Linear Least Square Estimation/Equalization

1.2.1 Signal model

$$\mathbf{y} = \mathbf{f}(\mathbf{s}) \quad (1-7)$$

Where \mathbf{y} is the received signal, regression model to estimate \mathbf{s} using LLS

$$\hat{\mathbf{s}} = \boldsymbol{\beta}^H \mathbf{y} \quad (1-8)$$

LLS to minimize

$$J = (\mathbf{s} - \hat{\mathbf{s}})(\mathbf{s} - \hat{\mathbf{s}}) \quad (1-9)$$

Base on (1-6) and (1-9)

1.3 Non-linear Least Square

The formula (1-6) is also valid for non-linear case but with the $X_{i,j} = \frac{\partial f(x_i, \beta)}{\partial \beta_j}$.

There is no closed-form solution to a non-linear least squares problem. Instead, numerical algorithms are used to find the value of the parameters $\hat{\boldsymbol{\beta}}$ that minimizes the objective. Most algorithms involve choosing initial values for the parameters. Then, the parameters are refined iteratively, that is, the values are obtained by successive approximation:

$$\beta_j^{k+1} = \beta_j^k + \Delta\beta_j \quad (1-10)$$

Or

$$\boldsymbol{\beta}^{k+1} = \boldsymbol{\beta}^k + \Delta\boldsymbol{\beta} \quad (1-11)$$

where k is an iteration number, and the vector of increments $\Delta\beta_j$ is called the shift vector. In some commonly used algorithms, at each iteration the model may be linearized by approximation to a first-order Taylor series expansion about $\boldsymbol{\beta}^k$.

$$\begin{aligned} f(x_i, \boldsymbol{\beta}) &= f^k(x_i, \boldsymbol{\beta}) + \sum_j \frac{\partial f(x_i, \boldsymbol{\beta})}{\partial \beta_j} (\beta_j - \beta_j^k) \\ &= f^k(x_i, \boldsymbol{\beta}) + \sum_j J_{ij} \Delta\beta_j \end{aligned} \quad (1-12)$$

The Jacobian J is a function of constants, the independent variable and the parameters, so it changes from one iteration to the next. The residuals are given by

$$\begin{aligned} r_i &= y_i - f^k(x_i, \boldsymbol{\beta}) - \sum_{k=0}^{m-1} J_{ik} \Delta\beta_k \\ &= \Delta y_i - \sum_{j=0}^{m-1} J_{ij} \Delta\beta_j \end{aligned} \quad (1-13)$$

To minimize the sum of square of r_i , the gradient equation is set to zero and solved for $\Delta\beta_j$:

$$-2 \sum_{i=0}^{N-1} J_{ij} \left(\Delta y_i - \sum_{k=0}^{m-1} J_{ik} \Delta\beta_k \right) = 0 \quad (1-14)$$

Which, on rearrangement, become m simultaneous linear equations, the normal equations:

$$\sum_{i=0}^{N-1} \sum_{k=0}^{m-1} J_{ij} J_{ik} \Delta \beta_k = \sum_{i=0}^{N-1} J_{ij} \Delta y_i, j = 0, \dots, m-1 \quad (1-15)$$

The normal equations are written in matrix notation as

$$(J^T J) \Delta \beta = J^T \Delta y \quad (1-16)$$

These are the defining equations of the [Gauss-Newton algorithm](#).

By rearranging (1-16)

$$\Delta \beta = (J^T J)^{-1} J^T \Delta y \quad (1-17)$$

One example of channel estimation model, it uses a simplified LS method:

$$\Delta \beta \cong u J^T \Delta y \quad (1-18)$$

```
/* update channel estimate */
/* h[i] = h[i] + mu sum_{k=L}^{tsclen-1} tsc[k-i]*e[k+L], i=0:L */
for (i = 0; i <= L; i++)
{
    temp.r = temp.i = 0;
    for (k = L; k < tsclen; k++)
    {
        temp.r += tsc[k-i]*e[k].r;
        temp.i += tsc[k-i]*e[k].i;
    }
    h[i].r += temp.r*chanest_mul6/32768;
    h[i].i += temp.i*chanest_mul6/32768;
}
```

1.3.1 Iteration solving of the non-linear least square problem(not sure whether correct)

With formula (1-16), one can solve the non-linear least square problem by

$$\begin{aligned} \beta^k &= \beta^{k-1} + \Delta \beta \\ &= \beta^{k-1} + (J^T J)^{-1} J^T \Delta y \end{aligned} \quad (1-19)$$

Where J can be calculated by

$$J_{ij} = \frac{\partial f(x_i, \beta)}{\partial \beta_j} \quad (1-20)$$