Class 3: Linear Regression and GLMs

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Learning outcomes

- Know how to fit and interpret a linear regression
- Know the difference between a linear regression model and a generalised linear model (GLM)
- Know what a link function is and why it is used
- Be able to interpret the output of a simple GLM

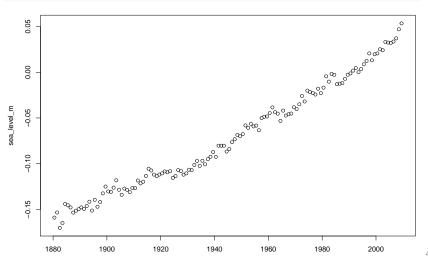
Using more of the data

- It's very rare that we are given a data set with just a single variable
- More often we're given multiple variables and asked to predict one or more of the variables from the others
- ► This is an example of *conditional* inference
- It might look more complicated, but this is still just fitting a probability distribution to some data

Linear regression example

Here is some data on sea level rise:

```
sl = read.csv('../../data/tide_gauge.csv')
with(sl, plot(year_AD, sea_level_m))
```



Linear regression models

- ▶ The simplest version of a linear regression model has:
 - ► A response variable (y) which is what we are trying to predict/understand
 - ▶ An *explanatory variable* or *covariate* (x) which is what we are trying to predict the response variable from
 - ▶ Some residual uncertainty (ϵ) which is the leftover uncertainty that is not accounted for by the explanatory variable
- ▶ Our goal is to predict the response variable from the explanatory variable, *or* to try and discover if the explanatory variable *causes* some kind of change in the response

The linear models in maths

We write the linear model as:

$$y_i = \alpha + \beta x_i + \epsilon_i$$

where α is the intercept, β the slope, and i = 1, ..., N represents each of the N observations

- ▶ Usually we make the additional assumption that $\epsilon_i \sim N(0, \sigma^2)$ where σ^2 is the residual standard deviation
- Writing this in probability distributions:

$$y_i|x_i,\alpha,\beta,\sigma\sim N(\alpha+\beta x_i,\sigma^2)$$

.

Fiting linear regression models

▶ We can create a likelihood as before by guessing some values of the parameters and then using the dnorm function to compute the likelihood value

```
alpha = 2
beta = 1.5
sigma = 0.6
y = sl$sea_level_m
x = sl$year_AD
sum(dnorm(y, mean = alpha + beta*x, sd = sigma, log = TRUE)
```

[1] -1539607458

Not a very high value of the likelihood!

Finding the best values

Luckily, R has the 1m function to find the best fitting values of the parameters

```
summarv(lm(v ~ x))
##
## Call:
## lm(formula = v \sim x)
##
## Residuals:
         Min
##
                     10
                           Median
## -0.0167787 -0.0051874 -0.0003646
          30
                    Max
## 0.0063022 0.0252621
##
## Coefficients:
##
                Estimate Std. Error
## (Intercept) -3.062e+00 3.937e-02
## x
              1.538e-03 2.024e-05
              t value Pr(>|t|)
##
## (Intercept) -77.78 <2e-16 ***
## x
               75.99 <2e-16 ***
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.01 '*' 0.05
## '.' 0.1 ' ' 1
##
## Residual standard error: 0.008659 on 128 degrees of freedom
## Multiple R-squared: 0.9783, Adjusted R-squared: 0.9781
## F-statistic: 5775 on 1 and 128 DF, p-value: < 2.2e-16
```

Checking the likelihood

[1] 433.9335

A much higher value of the likelihood!

Other notes about 1m

- Usually we would store the output from lm in another object to allow us to manipulate the output, e.g. my_model = lm(y ~ x)
- ▶ We can use the confint function to get confidence intervals on the parameters
- ▶ We can predict future values of sea level from the model by giving it new x values, e.g.

```
my_model = lm(y ~ x)
predict(my_model, newdata = data.frame(x = 2050))
```

```
## 1
## 0.09055375
```

Checking the model

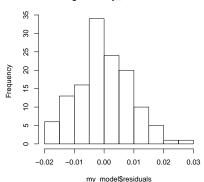
- Just finding the best values of the parameters and their uncertainty is not the whole story
- We need to check the fit of the model
- We can do this by analysing whether the assumed probability distribution is correct or not
- Can look at the probability distribution we have fitted to the data, but most people just look at the leftover bits - the residuals

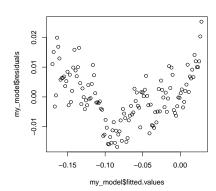
Residual plot

Common to plot a histogram of the residuals, and a scatter plot of the residuals vs the fits

```
par(mfrow=c(1, 2))
hist(my_model$residuals)
plot(my_model$fitted.values, my_model$residuals)
```

Histogram of my_model\$residuals





From LMs to GLMs

- ▶ If a normal distribution is not suitable for the residuals we need to choose another probability distribution
- ▶ Here is some data from an experiment on whitefly:

```
whitefly = read.csv('../../data/whitefly.csv')
head(whitefly, 4)
```

```
## imm week block trt n live plantid

## 1 0 5 3 5 6 6 3

## 2 0 1 2 3 10 2 4

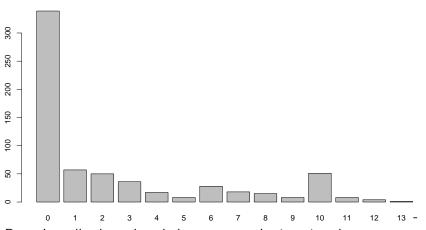
## 3 0 3 2 3 10 0 4

## 4 0 4 2 3 10 6
```

- ► The live and n columns indicate how many whitefly survived and were used in the experiment respectively
- ▶ We have a fixed total and a number of surviving whitefly out of this total. Which probability distribution might be appropriate?

Plotting the whitefly data

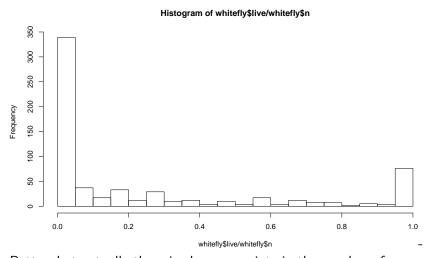
barplot(table(whitefly\$live))



Doesn't really show the whole story as we're ignoring the total

Second go

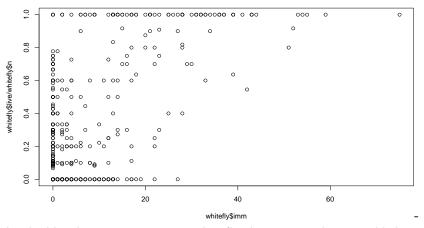
hist(whitefly\$live/whitefly\$n, breaks = 30)



Better, but actually there is also a covariate in the number of immature whitefly that were included (variable ${\tt imm}$)

Third go

plot(whitefly\$imm, whitefly\$live/whitefly\$n)



Looks like the more immature whitefly there were, the more likely they were to die

Fitting a model

- Let's call y the number of live whitefly at the end of the experiment, and n the number of whitefly used in the experiment
- Let's fit a binomial distribution:

$$y \sim Bin(n, p)$$

- We know the value of n so it is not really a parameter but a fixed part of the data
- We need to estimate p
- We could use method of moments or maximum likelihood. If we use method of moments we get $\hat{p} = 0.245$

Fitting a better model

- ▶ What if we wanted to include imm as a covariate?
- ▶ One way would be to let $p = \alpha + \beta x$ where x is the number of immature whitefly
- ▶ We could fit this using maximum likelihood to get estimates of $\hat{\alpha}$ and $\hat{\beta}$
- ► This is now a Generalised Linear Model (GLM)
- The likelihood would be:

```
y = whitefly$live
n = whitefly$n
x = whitefly$imm
alpha = 0.5
beta = 0.0001
sum(dbinom(y, n, alpha + beta * x, log = TRUE))
```

```
## [1] -3330.495
```

Fitting an even better model

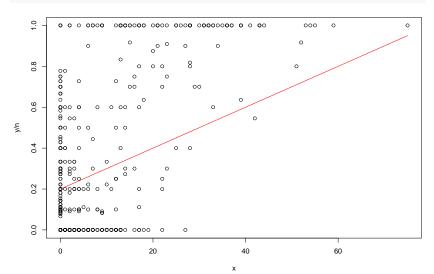
▶ Suppose after using maximum likelihood we got $\hat{\alpha}=0.2$ and $\hat{\beta}=0.01$. Then the likelihood would be

```
alpha = 0.2
beta = 0.01
sum(dbinom(y, n, alpha + beta * x, log = TRUE))
## [1] -1971.889
```

ightharpoonup Can anyone see any problems with this model? (Hint: suppose I wanted to predict what proportion would die when x=100)

Plotting the fitted model

```
plot(x, y/n)
lines(x, alpha + beta * x, col = 'red')
```



The logit function

- To stop the lines going out of the range (0, 1) people often use the *logit* transformation:

$$\log \left[\frac{p}{1-p} \right] = \alpha + \beta x \text{ or } p = \frac{e^{\alpha + \beta x}}{e^{\alpha + \beta x} + 1}$$

- The latter is known as the inverse logit function - We now maximise the likelihood:

```
p = exp(alpha + beta * x) / (exp(alpha + beta * x) + 1)
sum(dbinom(y, n, p, log = TRUE))
```

```
## [1] -3630.322
```

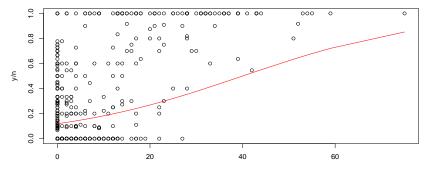
▶ These logit and inverse logit functions are in the boot package:

```
library(boot)
sum(dbinom(y, n, inv.logit(alpha + beta * x), log = TRUE))
## [1] -3630.322
```

Plotting the fit

- ▶ Suppose under this method we got maximum likelihood estimates of $\hat{\alpha} = -2$ and $\hat{\beta} = 0.05$
- ▶ A plot of the fitted values is now:

```
alpha = -2
beta = 0.05
plot(x, y/n)
lines(x, inv.logit(alpha + beta * x), col = 'red')
```



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Finding the maximum likelihood values

- R has a function called glm to find the maximum likelihood values for us
- ▶ For binomial model with a logit *link function* we would type:

```
glm(cbind(y, n) ~ x, family = binomial(link = logit))
##
## Call: glm(formula = cbind(y, n) ~ x, family = binomial
##
## Coefficients:
## (Intercept)
## -1.95517 0.05391
##
## Degrees of Freedom: 639 Total (i.e. Null); 638 Residual
## Null Deviance:
                       2389
## Residual Deviance: 1688 AIC: 2541
```

Further details about the fit

- ➤ You'll see amongst the output something called the *deviance*. This is minus twice the log-likelihood
- ▶ It's a common measure used to compare models as the deviance for a linear regression model is just the mean square error
- Next to it you'll see the Akaike Information Criterion or AIC value, which penalises the deviance by adding on twice the number of parameters (i.e. a measure of the complexity of the model)
- ► Often, you would fit multiple models with different covariates and choose the one(s) with the smallest AIC

Other glms

- Another common glm is the Poisson, useful for count data
- ► For example, suppose we treated the number of whitefly who died as a count, and temporarily ignored the *n* values
- ▶ We could fit:

```
glm(y ~ x, family = poisson(link = log))
##
## Call: glm(formula = y ~ x, family = poisson(link = log)
##
## Coefficients:
## (Intercept)
                         х
      0.38559 0.04608
##
##
## Degrees of Freedom: 639 Total (i.e. Null); 638 Residual
## Null Deviance:
                       3055
## Residual Deviance: 2154 AIC: 3124
```

A final word on glms

- ► There are lots of different types of GLMs We can do Gamma, Negative Binomial, Beta, Inverse Gaussian, . . .
- Each has a link function which transforms the main parameter into an unrestricted range through which we can include covariates
- ▶ It's also simple to include extra covariates or interactions:

```
glm(y \sim x1 + x2 + x1:x2, family = poisson(link = log))
```

Residual checks are still important, and R will create them for us We can get at them via e.g.

```
my_model = glm(cbind(y, n) ~ x, family = binomial(link = lo
plot(my_model)
```

Summary

- Linear Regression and Generalised Linear Models are two common ways to extend standard probability distributions to include covariates
- We estimate the parameters via maximum likelihood using e.g. lm or glm
- ► We sometimes need to include a link function which transforms the parameters into an unrestricted range
- There are lots of different types of GLM for every flavour of probability distribution