

The Munich Procedure

Calculating and applying coefficient corrections following the Munich Procedure

Michaela Schauer

2024-04-02

Contents

1	Introduction	4
2	About this code	4
3	Packages	5
4	Set working directory	5
5	Preparing the spreadsheet	6
5.1	Coefcors for p-XRF and laboratory methods	6
5.2	Device-internal coefcors	7
6	Example of developing a coefcor: Aluminium in coefcor I	8
6.1	First iteration	8
6.1.1	First impressions	8
6.1.1.1	<i>SCATTER PLOT</i>	8
6.1.1.2	<i>CORRELATION</i>	9
6.1.1.3	<i>MORE KEY VALUES</i>	9
6.1.2	Ordinary Linear Regression (OLR)	10
6.1.2.1	<i>CALCULATING THE OLR</i>	10
6.1.2.2	<i>CALCULATING THE ROOT MEAN SQUARED ERROR (RMS)</i>	11
6.1.2.3	<i>CALCULATING THE RELATIVE SEE (RSEE)</i>	11
6.1.2.4	<i>CHECKING RESIDUALS</i>	12
6.1.2.5	<i>CHECKING FOR SIGNIFICANT OUTLIERS</i>	12
6.1.2.6	<i>TESTING FOR ROBUSTNESS - CONFIDENCE INTERVALS</i>	13
6.1.3	Regression Trough Origin (RTO)	14
6.1.3.1	<i>CALCULATING THE RTO</i>	14
6.1.3.2	<i>CALCULATING THE ROOT MEAN SQUARED ERROR (RMS)</i>	15
6.1.3.3	<i>CALCULATING THE RELATIVE SEE (RSEE)</i>	15
6.1.3.4	<i>CALCULATING THE COEFFICIENT OF DETERMINATION (R^2)</i>	16
6.1.3.5	<i>CHECKING RESIDUALS</i>	16
6.1.3.6	<i>CHECKING FOR SIGNIFICANT OUTLIERS</i>	17
6.1.3.7	<i>TESTING FOR ROBUSTNESS - CONFIDENCE INTERVALS</i>	18
6.1.4	Result	19

6.2	Second iteration	19
6.2.1	First impressions	20
6.2.2	Ordinary Linear Regression (OLR)	20
6.2.3	Regression Trough Origin (RTO)	23
6.2.4	Result	25
6.3	Third iteration	26
6.3.1	First impressions	26
6.3.2	Ordinary Linear Regression (OLR)	27
6.3.3	Regression Trough Origin (RTO)	29
6.3.4	Result	31
6.4	Fourth iteration	32
6.4.1	First impressions	32
6.4.2	Ordinary Linear Regression (OLR)	33
6.4.3	Regression Trough Origin (RTO)	35
6.4.4	Result	37
6.5	Fifth iteration	38
6.5.1	First impressions	38
6.5.2	Ordinary Linear Regression (OLR)	39
6.5.3	Regression Trough Origin (RTO)	41
6.5.4	Result	43
6.6	Sixth iteration	44
6.6.1	First impressions	44
6.6.2	Ordinary Linear Regression (OLR)	45
6.6.3	Regression Trough Origin (RTO)	47
6.6.4	Result	49
6.7	Seventh iteration	50
6.7.1	First impressions	50
6.7.2	Ordinary Linear Regression (OLR)	51
6.7.3	Regression Trough Origin (RTO)	54
6.7.4	Result	56
6.8	Eighth iteration	57
6.8.1	First impressions	57
6.8.2	Ordinary Linear Regression (OLR)	58
6.8.3	Regression Trough Origin (RTO)	60
6.8.4	Result	62
6.9	Ninth iteration	63
6.9.1	First impressions	63
6.9.2	Ordinary Linear Regression (OLR)	64
6.9.3	Regression Trough Origin (RTO)	66
6.9.4	Result	69
6.10	Tenth iteration	69
6.10.1	First impressions	69
6.10.2	Ordinary Linear Regression (OLR)	70
6.10.3	Regression Trough Origin (RTO)	73
6.10.4	Result	75
6.11	Eleventh iteration	76
6.11.1	First impressions	76
6.11.2	Ordinary Linear Regression (OLR)	77
6.11.3	Regression Trough Origin (RTO)	79
6.11.4	Result - RTO	82

7	Example of extracting criteria and factors	82
7.1	Extracting from OLR	83
7.2	Extracting from RTO	84
7.3	Exporting coefcor criteria (_criteria.csv)	86
7.4	Exporting coefcor factors - slope and intercept (_factors.csv)	86
8	Example of graphics export	87
8.1	Exporting all relevant descriptive graphics for OLR and RTO	87
8.2	Exporting the traditionally presented scatter plots	92
9	Example of applying coefcors obtained from p-XRF and laboratory values to a dataset	94
9.1	Loading and formating the data	94
9.2	Performing the calculations	94
9.3	Compiling and exporting the processed data (_cor_data.csv)	97
10	Example of applying device-internal coefcors to a dataset	98
10.1	Loading and formating the data	98
10.2	Performing the calculations	99
10.3	Compiling and exporting the processed data (_cor_data.csv)	101

1 Introduction

This script **presents the Munich Procedure** which is **based on the Frankfurt Procedure** developed by Dr Markus Helfert and **takes this approach a step further** (Schauer 2023 & Schauer 2024): The Frankfurt Procedure compares values measured by p-XRF (x-values or independent variable) with values of the same set of samples obtained by a laboratory method - in this case WD-RFA (y-values or dependent variable). A coefficient correction (coefcor) is applied by using the calculated slope and, if appropriate, intercept to correct the p-XRF values to match those of the laboratory method. In addition, the same approach can be used to compare data obtained from the same instrument before and after a change of set-up by the manufacturer. The Munich Procedure additionally provides the calculations as R-scripts and adds quality criteria such as rSEE, RMS and CI, as well as the associated graphical output.

This code example shows the development of a coefcor for the element **aluminium** by **comparing p-XRF and WD-RFA values** from measurements of the **Frankfurt pottery sample set**. The p-XRF data used are from the first data set produced for the Niton XL3t No 97390 - namely **coefcor I**.

As the following code is an example for the application on aluminium - for all other elements the term 'Al' can simply be replaced by the desired element (Ti, Fe etc.).

2 About this code

The code was written using RGUI version 4.3.1, RCMDR version 2.8-0 and RStudio version 2023.06.1 with Quarto version 1.3.433. It is executable, i.e. it is fully functional when handled according to the instructions:

- In order for the code to run without errors, the 'Chunk Output line' must be checked under 'Edit the R-Markdown format options for the current file' (small cogwheel to the right of the 'Render' button).
- To set the working directory, the checkbox under 'Workspace Panes' (the four-square icon above the Render button) must be set to 'Show all panes'. Then there is a panel at the bottom right of the screen, whose display in this panel's toolbar must be set to 'Files'. There you will find a blue cogwheel (More File Commands), select this and a drop down menu will appear. Select 'set as working directory' - this will set the working directory to the folder where the project is located. This only works if the entire project folder has been saved unchanged!

The sources used to create the code for each chapter can be found under 'source' and are only presented when the code is first introduced.

Schauer, Michaela 2024: Coefficient corrections for portable X-ray fluorescence data of the Niton XL3t No. 97390 (coefcor I-IV) developed according to the Munich procedure. Data in Brief 53. 109914. <https://doi.org/10.1016/j.dib.2023.109914>

Schauer, Michaela 2023: R-scripts and data of coefficient corrections developed since 2017 for the Niton XL3t No. 97390 following the Munich Procedure. 12. September 2023. Open Data LMU. <https://doi.org/10.5282/ubm/data.405>

3 Packages

source

Instructions following <https://www.projectpro.io/recipes/load-package-r> and <https://www.dataquest.io/blog/install-package-r/>

install packages

To use this script you have to install the necessary packages by using this code

```
install.packages(c('car','cowplot','dplyr','ggpmisc','ggplot2','ggpubr','grid',  
                  'tibble','Rcmdr','RcmdrMisc'))
```

The below named packages must be loaded actively at every restart (regardless of whether R-Gui or R-Studio is used to run the code).

```
library(car)  
library(cowplot)  
library(dplyr)  
library(ggpmisc)  
library(ggplot2)  
library(ggpubr)  
library(grid)  
library(tibble)  
library(Rcmdr)  
library(RcmdrMisc)
```

4 Set working directory

source

Instructions following <https://martinctc.github.io/blog/rstudio-projects-and-working-directories-a-beginner%27s-guide/>, https://www.grainge.org/pages/authoring/relative_paths/relative_paths.htm and <https://r4ds.had.co.nz/workflow-projects.html>

Next, the file path - i.e. the location where the file that R is supposed to work with can be found - must be defined. In R-terminology this is call working directory. We are working with a relative path which is defined in and by the location of the R-project-file (**MunichProcedure.Rproj**).

To adjust the working directory to the position of the files on your computer, click on ‘Files’ in the task bar of the lower right pane and select ‘Munich Procedure’ from the terms in the header written in blue. You should see the folder structure of this project in the field below. Now under ‘More’ (blue cogwheel) select ‘Set As Working Directory’.

```
knitr::opts_knit$set(root.dir = "./")
```

5 Preparing the spreadsheet

5.1 Coefcors for p-XRF and laboratory methods

source

Instructions following <https://www.digitalocean.com/community/tutorials/r-read-csv-file-into-data-frame>, <https://sparkbyexamples.com/r-programming/r-select-function-from-dplyr/>, <https://dplyr.tidyverse.org/reference/mutate.html>, <https://www.digitalocean.com/community/tutorials/replace-in-r>, <https://search.r-project.org/R/refmans/base/html/Round.html>, <https://www.statology.org/transpose-data-frame-in-r/>, <https://stackoverflow.com/questions/40947288/how-to-calculate-mean-of-all-columns-by-group>, <https://www.rdocumentation.org/packages/base/versions/3.6.2/topics/merge> and <https://datatofish.com/export-dataframe-to-csv-in-r/>

The source file must be prepared in the same way as the `**_data.csv**` files (see e.g. `coefcorI_example_data.csv`). The most important requirements are the headers of the columns - p-XRF data have headers that only give the elements (Si, Ti etc.), while the header for laboratory analyses also contains information about the analysis type - for example WD-RFA (Si.WDXRF, Ti.WDRFA etc.).

The file format is created from the analytical data in files named `**_data.csv**` (in this example `coefcorI_adata.csv`). Loading is only possible if the required file is stored in a folder that can be found in the defined working directory.

```
dataset<- read.csv("../data_analytical//coefcorI_adata.csv")
```

Next we filter for the information we need.

```
dataset_pXRF<-select(dataset,"Sample","Si","Ti","Al","Fe","Mn","Mg","Ca","K","P",  
"S","Cl","Sc","V","Cr","Co","Ni","Cu","Zn","As","Se","Rb","Sr","Y","Zr","Nb","Mo",  
"Pd","Ag","Cd","Sn","Sb","Te","Cs","Ba","La","Ce","Hf","Ta","W","Re","Au","Hg",  
"Pb","Bi","Th","U")
```

Then we replace the information that a certain value was below the limit of detection (<LOD) with 0 and define all measurement values as numeric characters.

```
dataset_pXRF <- mutate_all(dataset_pXRF, ~gsub("<LOD", "0", .))  
  
dataset_pXRF[, 2:47] <- lapply(dataset_pXRF[, 2:47], as.numeric)
```

Following that, we calculate the mean for each sample and element, then we round those values to whole numbers.

```
dataset_pXRF_mean<-summarise(dataset_pXRF, across(everything(), mean), .by = c(Sample))  
  
dataset_pXRF_mean[, 2:47] <- round(dataset_pXRF_mean[, 2:47])
```

Next, we load the file containing the WDXRF measurements of Dr. Markus Helfert (`MHelfert_FrankfurtProcedur`

```
dataset_WDXRF<- read.csv("../data_analytical//MHelfert_FrankfurtProcedure_WDXRFdata.csv")
```

And combine both files by sample number.

```
data<- merge(dataset_pXRF_mean, dataset_WDXRF, by = "Sample", all = FALSE)
```

Then, we export the file as `**_data.csv**` (in our case `coefcorI_data.csv`)

```
write.csv(data,"../data_analytical/coefcorI_data.csv",row.names=TRUE)
```

Now we are ready to go and can start to develop coefcor I.

5.2 Device-internal coefcors

source

Instructions following <https://www.digitalocean.com/community/tutorials/r-read-csv-file-into-data-frame>, <https://sparkbyexamples.com/r-programming/r-select-function-from-dplyr/>, <https://www.statology.org/r-add-suffix-to-column-names/> and <https://datatofish.com/export-dataframe-to-csv-in-r/>

If the `_data.csv` file is to be prepared for internal coefcors such as for example coefcor ItoII of the Niton XL3t No 97390 (Schauer 2023 & Schauer 2024), the files prepared in accordance with section 5.1 form the basis.

The first step is to load the file whose data are to be fitted - i.e. they correspond to the independent variable (x-axis), ergo the p-XRF data from before. Then we select the required variables. In our example this dataset corresponds to `coefcorI_data.csv`.

```
dataset_pXRF_I<- read.csv("../data_analytical//coefcorI_data.csv")
```

```
dataset_pXRF_I<-select(dataset_pXRF_I,"Sample","Si","Ti","Al","Fe","Mn","Mg","Ca","K","P",
" S","Cl","Sc","V","Cr","Co","Ni","Cu","Zn","As","Se","Rb","Sr","Y","Zr","Nb","Mo",
" Pd","Ag","Cd","Sn","Sb","Te","Cs","Ba","La","Ce","Hf","Ta","W","Re","Au","Hg","Pb",
" Bi","Th","U")
```

In the second step, the file containing the values to be fitted is loaded - i.e. they correspond to the dependent variable (y-axis), ergo the WDXRF values from before. Again, the desired variables are selected. In our example this dataset corresponds to `coefcorII_data.csv`.

```
dataset_pXRF_II<- read.csv("../data_analytical//coefcorII_data.csv")
```

```
dataset_pXRF_II<-select(dataset_pXRF_II,"Sample","Si","Ti","Al","Fe","Mn","Mg","Ca","K","P",
" S","Cl","Sc","V","Cr","Co","Ni","Cu","Zn","As","Se","Rb","Sr","Y","Zr","Nb","Mo",
" Pd","Ag","Cd","Sn","Sb","Te","Cs","Ba","La","Ce","Hf","Ta","W","Re","Au","Hg","Pb",
" Bi","Th","U")
```

As before, an abbreviation is added to the column headings of the independent variables, namely the Roman numeral of the coefcor - in our example 'II'.

```
colnames(dataset_pXRF_II)[-1] <- paste(colnames(dataset_pXRF_II)[-1], ".II", sep = "")
```

And combine both files by sample number.

```
data<- merge(dataset_pXRF_II, dataset_pXRF_II, by = "Sample", all = FALSE)
```

Then, we export the file as `**_data.csv**` (in our case `coefcorItoII_data.csv`)

```
write.csv(data,"../data_analytical//coefcorItoII_data.csv",row.names=TRUE)
```

Now we are ready to go and could start to develop coefcor ItoII - yet the following example uses the data for coefcor I.

6 Example of developing a coefcor: Aluminium in coefcor I

The following code is used to compute coefcors from a `__data.csv`-file for the element aluminium based on the yet to load data set (see section 6.1) of coefcor I (`coefcorI_data.csv`).

Device-internal coefcors

If an device-internal coefcor is to be developed, only ‘WDXRF’ needs to be replaced by the respective Roman number of the independent variables (i.e. their suffix). Thus, in the case of coefcor ItoII, ‘WDXRF’ would be replaced by ‘II’.

6.1 First iteration

For the first iteration, the entire data set containing only the values of p-XRF and WDXRF-measurements is loaded (`coefcorI_data.csv`).

```
dataset<- read.csv("../data_analytical//coefcorI_data.csv")
```

This data is now used to calculate linear regressions which can give us the coefcors we are looking for.

6.1.1 First impressions

The following information gives a first impression of the data with regard to the already existing agreement between p-XRF values and those of laboratory analysis.

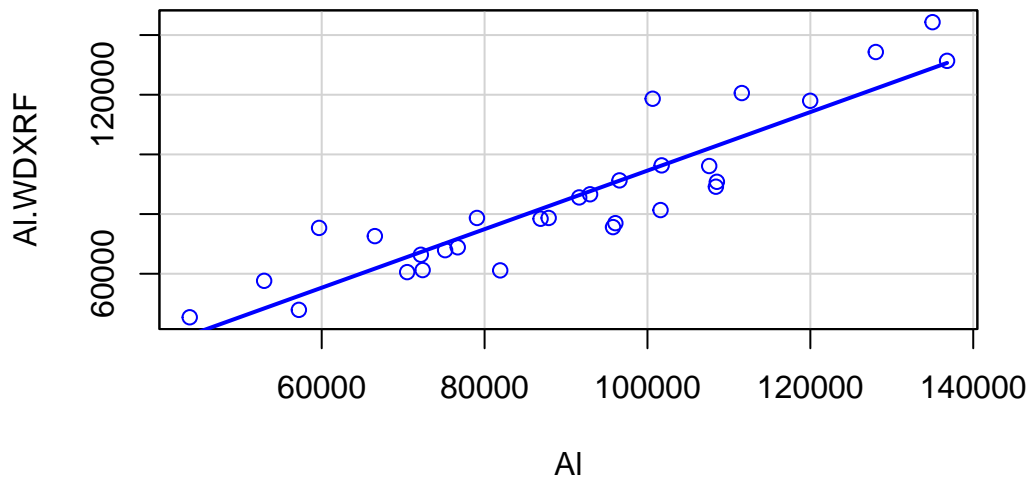
6.1.1.1 Scatter plot

source

Instructions following <http://www.sthda.com/english/wiki/scatter-plots-r-base-graphs>

The scatter plot allows an intuitive interpretation of the agreement of p-XRF- and WD-RFA-values based on the graph.


```
scatterplot(AI.WDXRF~AI, regLine=TRUE, smooth=FALSE, boxplots=FALSE, data=dataset)
```



In this case, the agreement does look quite okay already.

6.1.1.2 Correlation

source

Instructions following <https://www.statology.org/r-cor-function/>

Complementing the graph, the value of the pearson correlation coefficient gives an impression of how good the agreement of the values is. The correlation coefficient can have values between 0 and 1, whereby results close to 1 are desirable.

```
cor(dataset$AI,dataset$AI.WDXRF)
```

```
[1] 0.9105113
```

Also this figure is good - still it is likely that some iterations will be necessary to get a linear regression which complies with all the criteria defined in [Schauer 2024](#) and shortly described below.

6.1.1.3 More key values

source

Instructions following <https://www.rdocumentation.org/packages/Rcmdr/versions/2.0-4/topics/numSummary>

It is also worth taking a look at other key values such as mean value, standard deviation and the values of the quartiles. From this table, lower (0%) and upper limit (100%) are taken after completion of the analysis. They define the range of values for which the selected linear regression and thus the associated coefficients are valid.

```
numSummary(dataset[,c("Al", "Al.WDXRF"), drop=FALSE],
            statistics=c("mean", "sd", "quantiles"), quantiles=c(0,.25,.5,.75,1))
```

	mean	sd	0%	25%	50%	75%	100%	n
Al	90520.87	23619.45	43795	73085	92284.5	106119.8	136788	30
Al.WDXRF	85260.67	25454.23	45409	68140	78698.0	94906.0	144324	30

So in this case, p-XRF values of the element aluminium which range between 43795 ppm (0% - lower limit) and 136788 ppm (100% - upper limit) could be fitted with the coefficients generated in this iteration.

6.1.2 Ordinary Linear Regression (OLR)

First, an ordinary linear regression - i.e. a calculation with slope and intercept - is performed. The characteristic values generated here are then compared with those of a Regression Trough Origin (RTO) calculated subsequently.

6.1.2.1 Calculating the OLR

source

Instructions following <https://www.statology.org/OLR-regression-in-r/> and <https://search.r-project.org/R/refmans/stats/html/lm.html>

These lines of code give the most important information of the linear regression. Of particular interest are the spread of the residuals, the estimated values of intercept and slope (Al), the residual standard error (SEE) and the coefficient of determination (r^2). With regard to the latter, the 'multiple R-squared' value is relevant.

```
OLRA1<-lm(Al.WDXRF~Al, data=dataset)
summary(OLRA1)
```

Call:

```
lm(formula = Al.WDXRF ~ Al, data = dataset)
```

Residuals:

Min	1Q	Median	3Q	Max
-15694.1	-5683.2	-948.6	5667.8	23468.0

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-3.562e+03	7.870e+03	-0.453	0.654
Al	9.812e-01	8.421e-02	11.652	2.97e-12 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 10710 on 28 degrees of freedom

Multiple R-squared: 0.829, Adjusted R-squared: 0.8229

F-statistic: 135.8 on 1 and 28 DF, p-value: 2.972e-12

The residuals show a very large scatter, the necessary adjustment of the intercept is very high, but the slope correction is acceptable (the closer to 1 the better). The SEE is high, while r^2 is to low.

6.1.2.2 Calculating the Root Mean Squared error (RMS)

source

Instructions following <https://statisticsglobe.com/extract-fitted-values-from-regression-model-r> and <https://sparkbyexamples.com/r-programming/add-column-to-dataframe-in-r/>

With the root mean squared error (RMS), another relevant value of the linear regression is calculated. It allows an estimation of the deviation of the p-XRF-values from their expected position (i.e. the value of the WD-XRF measurement). It is not used as a discriminating criterion here but is shown in the factors table (`_factors.csv`).

To calculate the RMS, first the fitted values are extracted from the linear regression model and added as new columns to the dataset (now `datasetb`), then the RMS is computed.

```
fitted.OLRA1 <- fitted(OLRA1)
datasetb<-cbind(dataset,fitted.OLRA1)
RMS<-sqrt(mean((datasetb$A1-datasetb$fitted.OLRA1)^2)/nrow(datasetb))
RMS
```

```
[1] 963.6646
```

6.1.2.3 Calculating the relative SEE (rSEE)

source

Instructions following <https://www.rdocumentation.org/packages/stats/versions/3.6.2/topics/sigma> and <https://www.statology.org/r-mean-of-column/>

Of much higher importance to the Munich Procedure is the rSEE. To calculate this factor, first the SEE must be obtained from the previously generated data; the mean of the p-XRF-aluminium-values computed.

```
SEE<-sigma(OLRA1)
SEE
```

```
[1] 10711.22
```

Then the formula to calculate the relative residual standard error is defined and the rSEE for aluminium (Al) calculated.

```
mean<-mean(dataset$Al)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 11.83288
```

The value of the rSEE in this case is above 10% and thus higher than the threshold.

6.1.2.4 *Checking residuals*

source

Instructions following <https://www.rdocumentation.org/packages/stats/versions/3.6.2/topics/plot.lm> and <https://stat.ethz.ch/R-manual/R-devel/library/stats/html/plot.lm.html>

Residuals provide information on how much the measured values differ from the values estimated by linear regression. In practice, it turns out that the Q-Q plot, in which the residuals should lie on a straight line as much as possible, gives a good first impression of whether significant outliers might be present.

```
oldparOLRA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(OLRA1)
```

In the Q-Q-plot, the pair of measurements in row 25 of the table (the header counts as row '0') could be an outlier. Therefore, removing this sample might significantly lower the rSEE.

6.1.2.5 *Checking for significant outliers*

source

Instructions following <https://rdrr.io/cran/car/src/R/outlierTest.R>

To check, if significant outliers (Bonferroni $p < 0.05$) are present, a Bonferroni test for outlier (with studentized residuals) is performed.

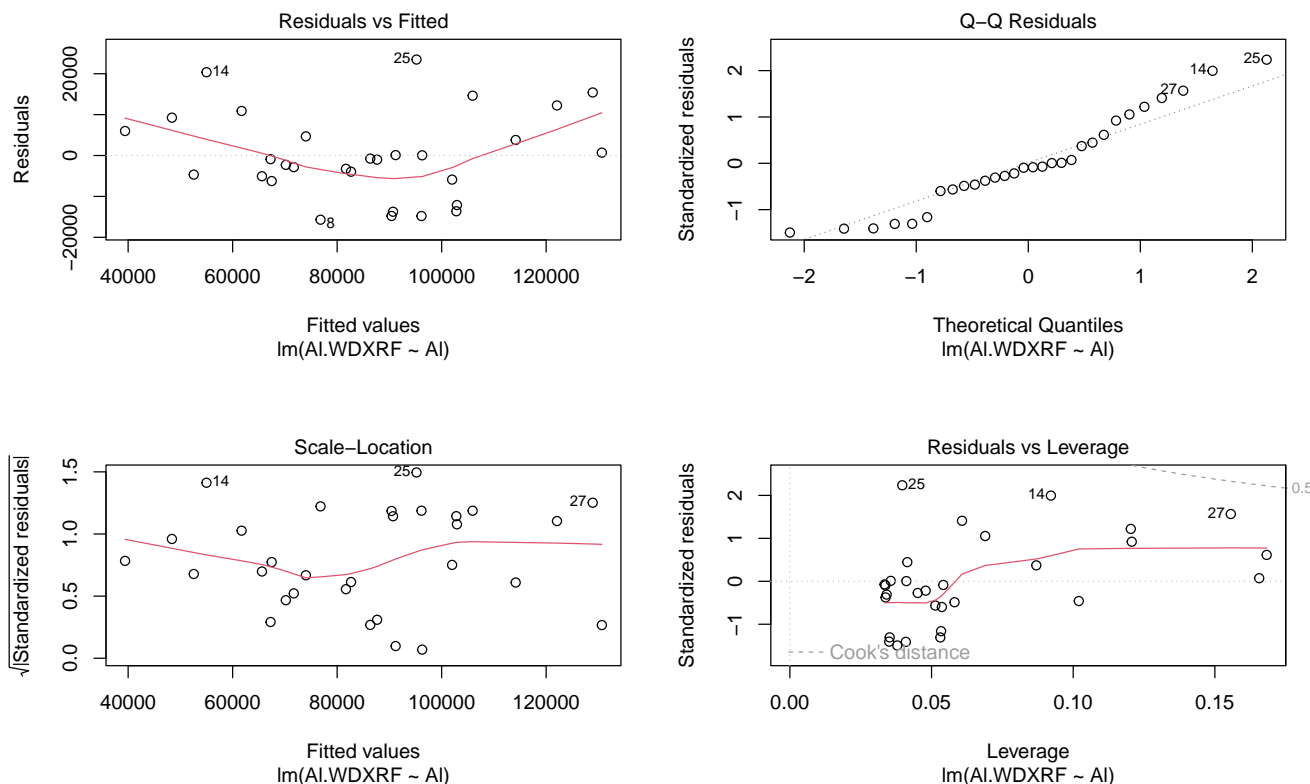
```
outlierTest(OLRA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest |rstudent|:

	rstudent	unadjusted p-value	Bonferroni p
25	2.422307	0.022409	0.67227

This analysis confirms - as already suspected - that the sample in line 25 is a outlier. But, the p-value of the Bonferroni p shows this outlier is not significant. Due to the criteria defined in [Schauer 2024](#), the sample does not have to be excluded, yet it might become necessary if this sample is also an outlier for the RTO.



NULL

6.1.2.6 Testing for robustness - confidence intervals

source

Instructions following <https://www.rdocumentation.org/packages/stats/versions/3.6.2/topics/confint>,
<https://www.rdocumentation.org/packages/boot/versions/1.3-28.1/topics/boot>,
<https://www.rdocumentation.org/packages/boot/versions/1.3-28.1/topics/plot.boot> and
https://stopsack.github.io/risks/reference/confint.margstd_boot.html

To test the robustness of the linear regression, the values of the 0.5 and 0.95 confidence levels of the slope (AI) estimated from all p-XRF aluminium values are compared with those calculated by a bootstrap algorithm. The results of both calculations should always be as close as possible to each other. First we calculate regular CI-values.

```
Confint(OLRA1, level=0.90)
```

	Estimate	5 %	95 %
(Intercept)	-3562.0894024	-1.694952e+04	9825.344971
AI	0.9812407	8.379863e-01	1.124495

Then we apply the bootstrap algorithm

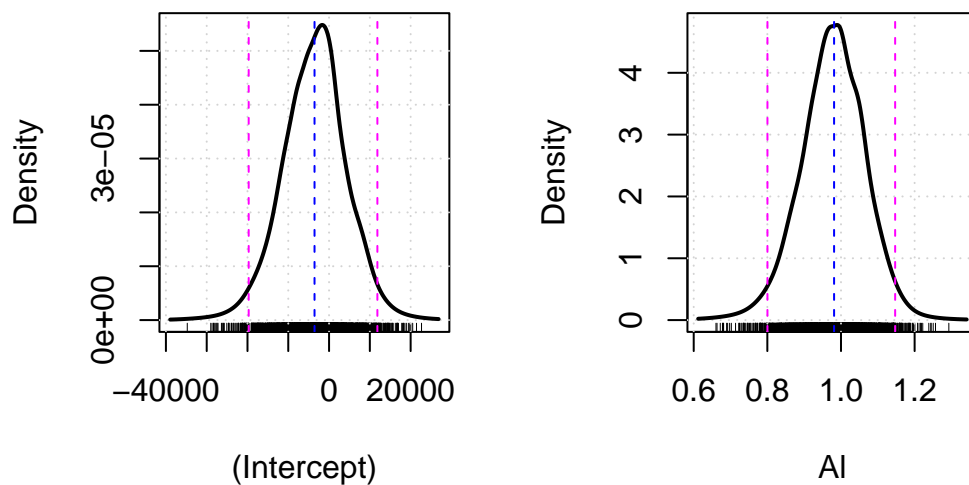
```
.bs.samplesOLRA1<- Boot(OLRA1, R=2500, method="case")
confint(.bs.samplesOLRA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
(Intercept)	-1.709252e+04	9095.108698
Al	8.345235e-01	1.123719

```
plotBoot(.bs.samplesOLRA1)
```

Bootstrap Distributions



According to the criteria defined in [Schauer 2024](#), the difference in the 0.5 confidence level of the slope (Al) values are close enough to call the linear regression robust. Also the density distribution of the slope looks not too bad.

Now we are done with all the necessary calculation for the OLR - let's start with the RTO.

6.1.3 Regression Trough Origin (RTO)

The procedure for calculating RTO and the associated parameters follows almost exactly the procedure for computing OLR. The main difference is that r^2 has to be calculated manually. Explanation for this can be found in [Schauer 2024](#).

6.1.3.1 Calculating the RTO

source

Instructions following <https://search.r-project.org/R/refmans/stats/html/lm.html>

These lines of code give the most important information of the linear regression. Of particular interest are the spread of the residuals, the estimated values of intercept and slope (Al), the residual standard

error (SEE) and the coefficient of determination (r^2). With regard to the latter, the ‘multiple R-squared’ value is relevant.

```
RTOA1<-lm(A1.WDXRF~A1-1, data=dataset)
summary(RTOA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1 - 1, data = dataset)
```

Residuals:

Min	1Q	Median	3Q	Max
-16232	-6094	-1462	4515	23622

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
A1	0.94432	0.02064	45.76	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 10560 on 29 degrees of freedom

Multiple R-squared: 0.9863, Adjusted R-squared: 0.9859

F-statistic: 2094 on 1 and 29 DF, p-value: < 2.2e-16

The residuals - as before for the OLR - show a large scatter. Again, the slope correction is acceptable, the SEE is quite high, while r^2 as calculated by the formula can't be used right away (see below).

6.1.3.2 Calculating the Root Mean Squared error (RMS)

source

Instructions following <https://statisticsglobe.com/extract-fitted-values-from-regression-model-r> and <https://sparkbyexamples.com/r-programming/add-column-to-dataframe-in-r/>

Again, in the sense of completeness, the RMS is calculated.

```
fitted.RTOA1 <- fitted(RTOA1)
datasetc<-cbind(dataset,fitted.RTOA1)
RMS<-sqrt(mean((datasetc$A1-datasetc$fitted.RTOA1)^2)/nrow(datasetc))
RMS
```

```
[1] 950.017
```

6.1.3.3 Calculating the relative SEE (rSEE)

source

Instructions following <https://www.rdocumentation.org/packages/stats/versions/3.6.2/topics/sigma> and <https://www.statology.org/r-mean-of-column/>

Same for the rSEE - with the difference that there is no need to calculate the mean as it was already defined when computing the rSEE or the OLR (see above)

```
SEE<-sigma(RTOAl)
SEE
```

```
[1] 10563.36
```

Then the formula to calculate the relative residual standard error is defined and the rSEE for aluminium (Al) calculated.

```
mean<-mean(dataset$Al)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 11.66953
```

Also in this case, the value of the rSEE is higher than 10% and thus above the threshold defined in [Schauer 2024](#).

6.1.3.4 Calculating the coefficient of determination (r^2)

source

Instructions following <https://pubs.cif-ific.org/doi/pdf/10.5558/tfc71326-3> and <https://onlinelibrary.wiley.com/doi/10.1111/1467-9639.00136>

To calculate the corrected r^2 (cr2) that correctly describes the RTO, other variables have to be determined first. We already have the first one - the SEE (see above). The next are the residual sum of squares - SSR - followed by the corrected total sum of squares - cSST. Using these three variables, cr2 can then be calculated. For the export of the graphics ([section -#sec-exgraphic]), every cr2 has to be named after the element it belongs to - e.g. Alcr2.

```
SSR<-(nrow(dataset)-1)*(SEE^2)
cSST <-sum(((dataset$Al.WDXRF)-mean(dataset$Al.WDXRF))^2)
Alcr2<-1-(SSR/cSST)
Alcr2
```

```
[1] 0.8277798
```

As for the OLR, r^2 is too low to be accepted.

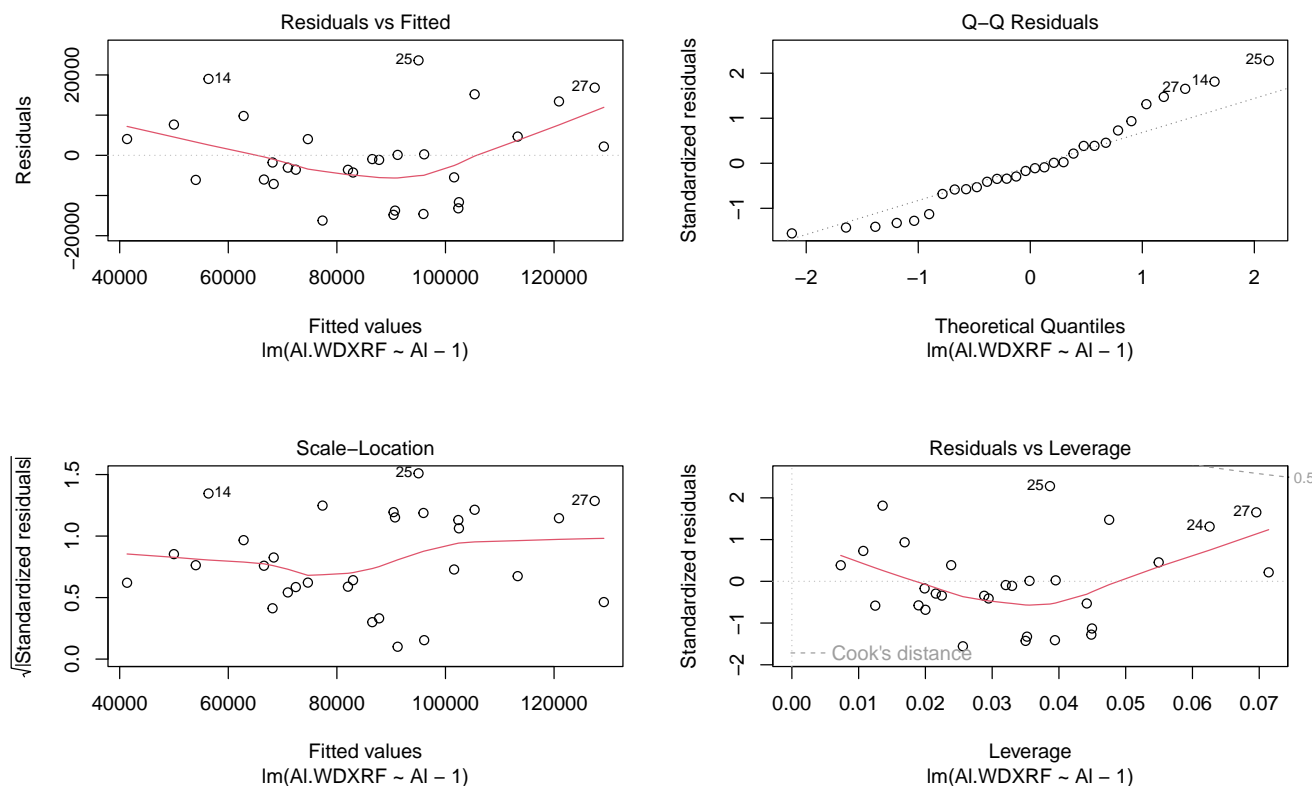
6.1.3.5 Checking residuals

source

Instructions following <https://www.rdocumentation.org/packages/stats/versions/3.6.2/topics/plot.lm>

Again we also check the residuals.

```
oldparRTOA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(RTOA1)
```



NULL

As before, in the Q-Q-plot, the sample in row 25 might be an outlier. So we have to check if it is by performing a Bonferroni test.

6.1.3.6 Checking for significant outliers

source

Instructions following <https://rdrr.io/cran/car/src/R/outlierTest.R>

Still, we perform the Bonferroni test.

```
outlierTest(RTOA1)
```

No Studentized residuals with Bonferroni $p < 0.05$
Largest `|rstudent|`:

	rstudent	unadjusted	p-value	Bonferroni	p
25	2.473844		0.019696		0.59089

As before, the sample in row 25 is a (non significant) outlier. According to the criteria defined in [Schauer 2024](#) this means we will have to perform a second iteration.

6.1.3.7 Testing for robustness - confidence intervals

source

Instructions following <https://www.rdocumentation.org/packages/stats/versions/3.6.2/topics/confint>,
<https://www.rdocumentation.org/packages/boot/versions/1.3-28.1/topics/boot>,
<https://www.rdocumentation.org/packages/boot/versions/1.3-28.1/topics/plot.boot> and
https://stopsack.github.io/risks/reference/confint.margstd_boot.html

Even, as we are clear how to proceed, checking the robustness of the RTO is always a good idea.

```
Confint(RTOA1, level=0.90)
```

	Estimate	5 %	95 %
A1	0.9443196	0.9092542	0.9793849

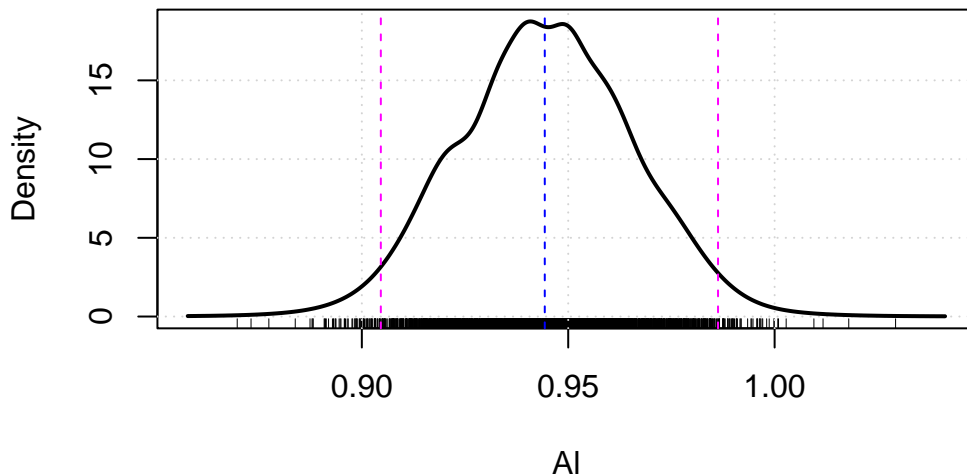
```
.bs.samplesRTOA1<- Boot(RTOA1, R=2500, method="case")
confint(.bs.samplesRTOA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
A1	0.9101073	0.9803524

```
plotBoot(.bs.samplesRTOA1)
```

Bootstrap Distributions



According to the criteria defined in [Schauer 2024](#), the difference in the slope (AI) values are small enough to call the linear regression robust. Also, the density distribution of the slope looks not too bad.

6.1.4 Result

The first iteration showed that the same outlier (sample in row 25) is visible in OLR as well as in RTO. Moreover, rSEE and r^2 for the two linear regressions are not within acceptable limits. Yet, robustness is given for both. Since three of the relevant criteria are not fulfilled, we carry out a second iteration of the analysis.

6.2 Second iteration

source

Instructions following <https://www.digitalocean.com/community/tutorials/r-read-csv-file-into-data-frame> and <https://sparkbyexamples.com/r-programming/drop-dataframe-rows-in-r/>

Since in our case we can improve the linear regression by removing the outlier - i.e. line 25 - we will do this for our second iteration.

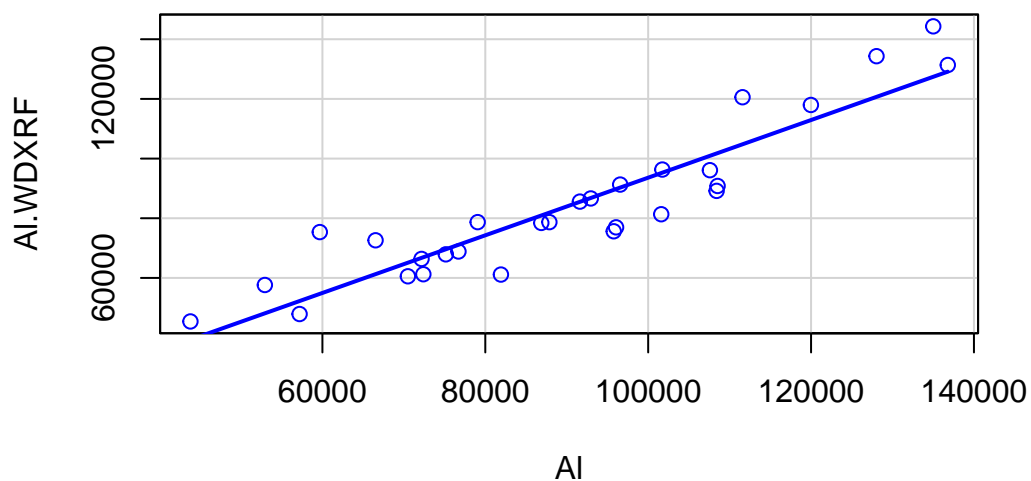
```
dataset1<- read.csv("../data_analytical//coefcorI_data.csv")
dataset<- dataset1[-c(25), ]
```

Thus, the data set we are now working with contains all samples except for the one in line 25. I.e. now all calculations for OLR and RTO are carried out again and checked for compliance with the criteria mentioned in [Schauer 2024](#). Based on this result, it is then decided whether a further iteration is necessary or if one of the two linear regressions can be selected and the coefcors accepted.

calculations

6.2.1 First impressions

```
scatterplot(A1.WDXRF~A1, regLine=TRUE, smooth=FALSE, boxplots=FALSE, data=dataset)
```



```
cor(dataset$A1,dataset$A1.WDXRF)
```

```
[1] 0.9221522
```

```
numSummary(dataset[,c("A1", "A1.WDXRF"), drop=FALSE],  
            statistics=c("mean", "sd", "quantiles"),quantiles=c(0,.25,.5,.75,1))
```

	mean	sd	0%	25%	50%	75%	100%	n
A1	90172.0	23958.74	43795	72395	91616	107579	136788	29
A1.WDXRF	84109.1	25096.90	45409	67902	78698	91294	144324	29

6.2.2 Ordinary Linear Regression (OLR)

```
OLRA1<-lm(A1.WDXRF~A1, data=dataset)  
summary(OLRA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1, data = dataset)
```

Residuals:

Min	1Q	Median	3Q	Max
-15011.0	-4813.6	-352.9	5317.8	20710.9

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-2.993e+03	7.267e+03	-0.412	0.684
A1	9.660e-01	7.798e-02	12.387	1.2e-12 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 9886 on 27 degrees of freedom

Multiple R-squared: 0.8504, Adjusted R-squared: 0.8448

F-statistic: 153.4 on 1 and 27 DF, p-value: 1.196e-12

```
fitted.OLRA1 <- fitted(OLRA1)
datasetb<-cbind(dataset,fitted.OLRA1)
RMS<-sqrt(mean((datasetb$A1-datasetb$fitted.OLRA1)^2)/nrow(datasetb))
RMS
```

[1] 1135.644

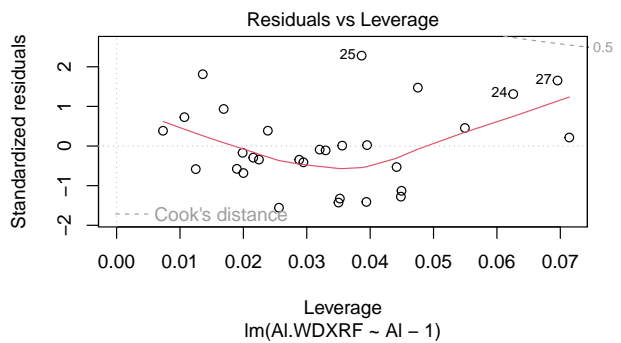
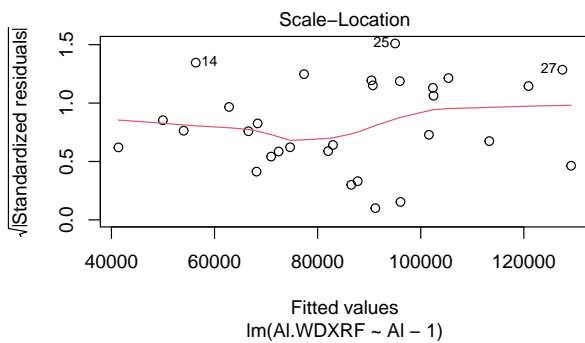
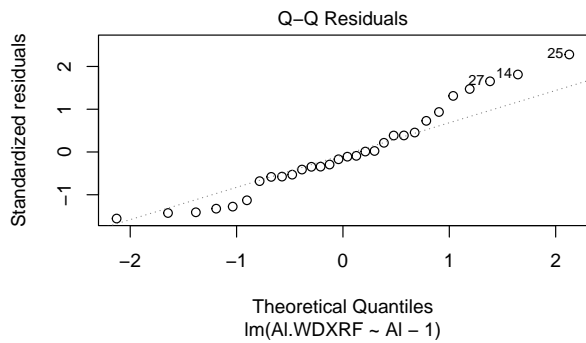
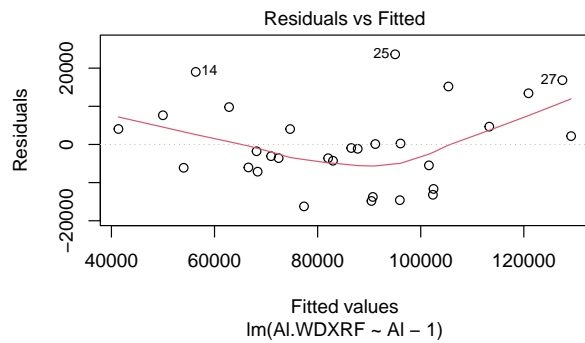
```
SEE<-sigma(OLRA1)
SEE
```

[1] 9886.31

```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

[1] 10.96384

```
oldparOLRA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(OLRA1)
```



NULL

```
outlierTest(OLRA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest `|rstudent|`:

	<code>rstudent</code>	unadjusted p-value	Bonferroni p
14	2.381525	0.024849	0.72061

```
Confint(OLRA1, level=0.90)
```

	Estimate	5 %	95 %
(Intercept)	-2993.351979	-1.537190e+04	9385.198412
AI	0.965959	8.331341e-01	1.098784

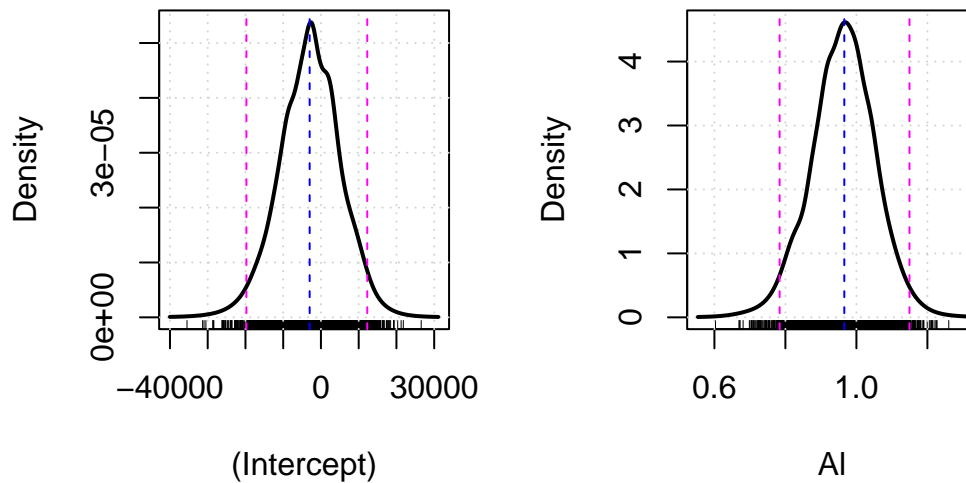
```
.bs.samplesOLRA1<- Boot(OLRA1, R=2500, method="case")
confint(.bs.samplesOLRA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
(Intercept)	-1.690438e+04	10106.961297
AI	8.135558e-01	1.116781

```
plotBoot(.bs.samplesOLRA1)
```

Bootstrap Distributions



6.2.3 Regression Trough Origin (RTO)

```
RTOA1<-lm(A1.WDXRF~A1-1, data=dataset)
summary(RTOA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1 - 1, data = dataset)
```

Residuals:

Min	1Q	Median	3Q	Max
-15458	-5576	-1103	4782	19572

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
A1	0.9349	0.0194	48.18	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 9739 on 28 degrees of freedom

Multiple R-squared: 0.9881, Adjusted R-squared: 0.9877

F-statistic: 2321 on 1 and 28 DF, p-value: < 2.2e-16

```
fitted.RTOA1 <- fitted(RTOA1)
datasetc<-cbind(dataset,fitted.RTOA1)
RMS<-sqrt(mean((datasetc$A1-datasetc$fitted.RTOA1)^2)/nrow(datasetc))
RMS
```

```
[1] 1126.93
```

```
SEE<-sigma(RTOA1)
SEE
```

```
[1] 9738.615
```

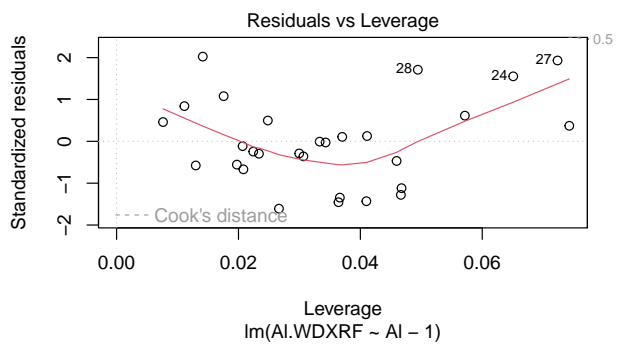
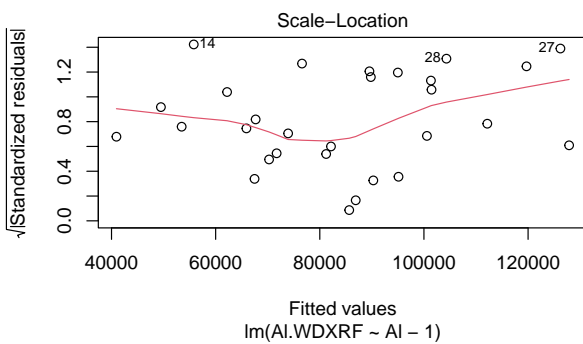
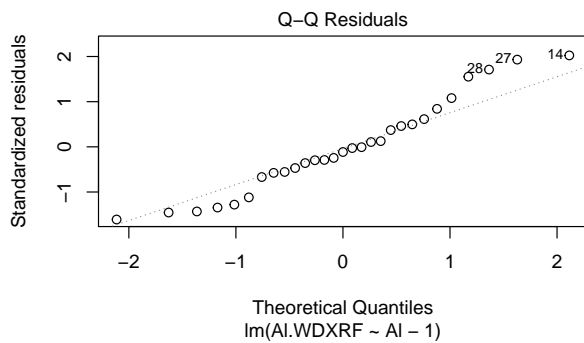
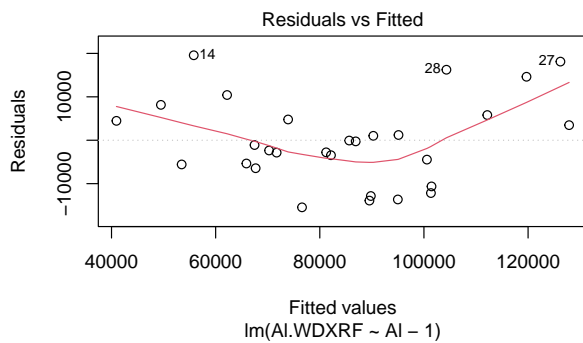
```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 10.80004
```

```
SSR<-(nrow(dataset)-1)*(SEE^2)
cSST <-sum(((dataset$A1.WDXRF)-mean(dataset$A1.WDXRF))^2)
Alcr2<-1-(SSR/cSST)
Alcr2
```

```
[1] 0.8494245
```

```
oldparRTOA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(RTOA1)
```



NULL

```
outlierTest(RTOA1)
```


No Studentized residuals with Bonferroni $p < 0.05$

Largest |rstudent|:

	rstudent	unadjusted p-value	Bonferroni p
14	2.151244	0.040564	NA

```
Confint(RTOA1, level=0.90)
```

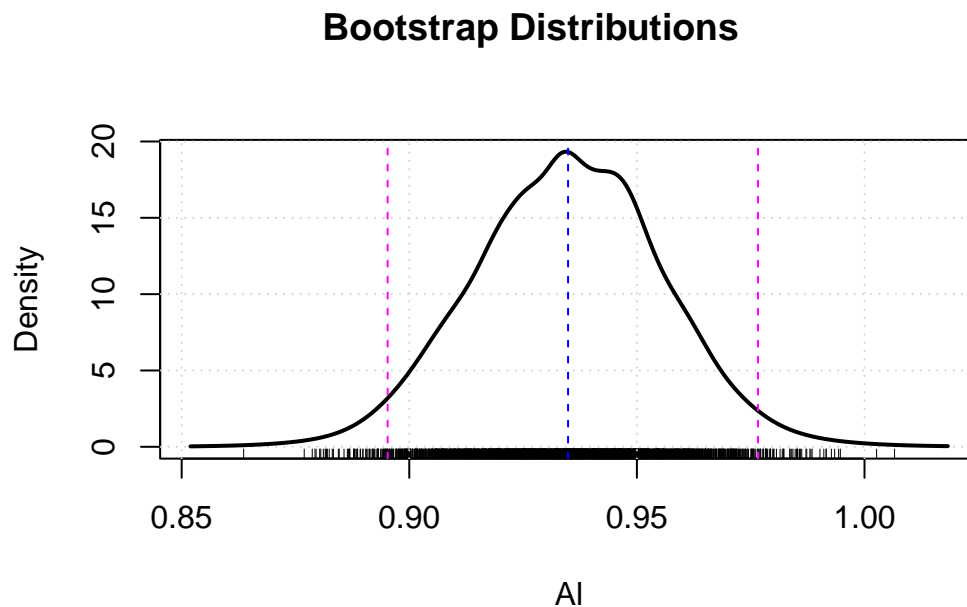
	Estimate	5 %	95 %
A1	0.9348813	0.9018713	0.9678913

```
.bs.samplesRTOA1<- Boot(RTOA1, R=2500, method="case")  
confint(.bs.samplesRTOA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
A1	0.9007452	0.9698718

```
plotBoot(.bs.samplesRTOA1)
```



6.2.4 Result

Both linear regressions give the sample in row 14 as a (non-significant) outlier. In addition, r^2 is too low, rSEE still too high for OLR and RTO, while robustness is fine. So, having seen that, we know that we need to do a third iteration.

6.3 Third iteration

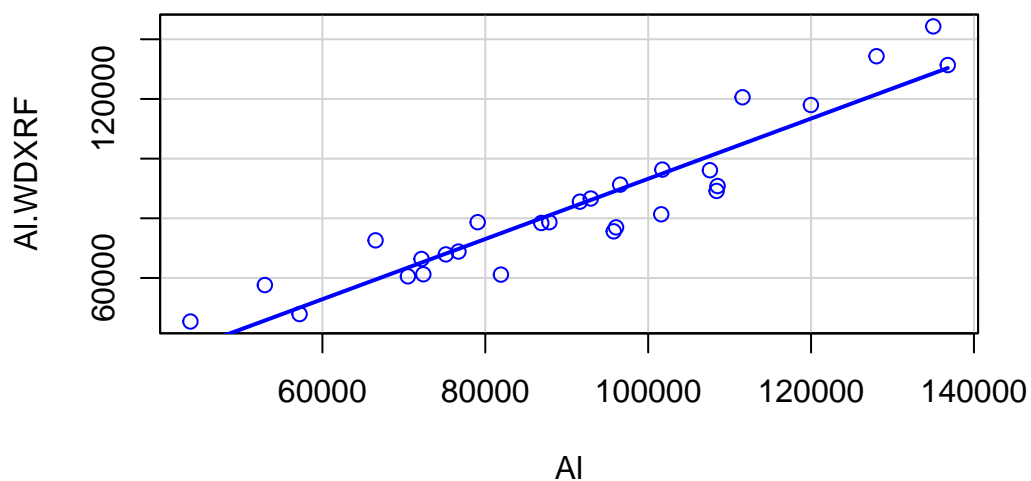
With this knowledge, for the third iteration, we now remove the sample in row 14 in addition to the one in row 25.

```
dataset1<- read.csv("../data_analytical//coefcorI_data.csv")
dataset<- dataset1[-c(25,14), ]
```

calculations

6.3.1 First impressions

```
scatterplot(AI.WDXRF~AI, regLine=TRUE, smooth=FALSE, boxplots=FALSE, data=dataset)
```



```
cor(dataset$AI,dataset$AI.WDXRF)
```

```
[1] 0.9362728
```

```
numSummary(dataset[,c("AI", "AI.WDXRF"), drop=FALSE],
            statistics=c("mean", "sd", "quantiles"),quantiles=c(0,.25,.5,.75,1))
```

	mean	sd	0%	25%	50%	75%	100%	n
AI	91261.07	23656.10	43795	74465.00	92284.5	107785	136788	28
AI.WDXRF	84421.43	25499.97	45409	67518.25	78698.0	92498	144324	28

6.3.2 Ordinary Linear Regression (OLR)

```
OLRA1<-lm(A1.WDXRF~A1, data=dataset)
summary(OLRA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1, data = dataset)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-13866.9	-4306.0	121.8	5068.6	15757.0

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-7.684e+03	6.993e+03	-1.099	0.282
A1	1.009e+00	7.426e-02	13.591	2.53e-13 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 9128 on 26 degrees of freedom

Multiple R-squared: 0.8766, Adjusted R-squared: 0.8719

F-statistic: 184.7 on 1 and 26 DF, p-value: 2.532e-13

```
fitted.OLRA1 <- fitted(OLRA1)
datasetb<-cbind(dataset,fitted.OLRA1)
RMS<-sqrt(mean((datasetb$A1-datasetb$fitted.OLRA1)^2)/nrow(datasetb))
RMS
```

```
[1] 1293.209
```

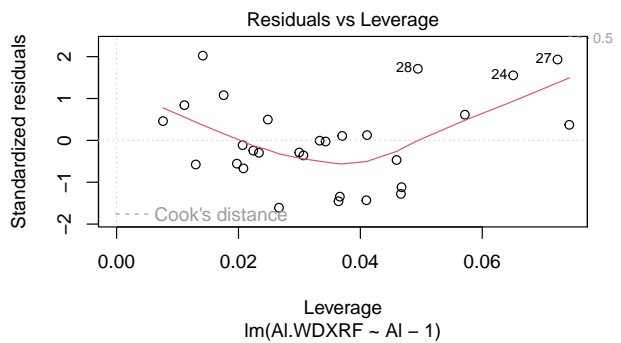
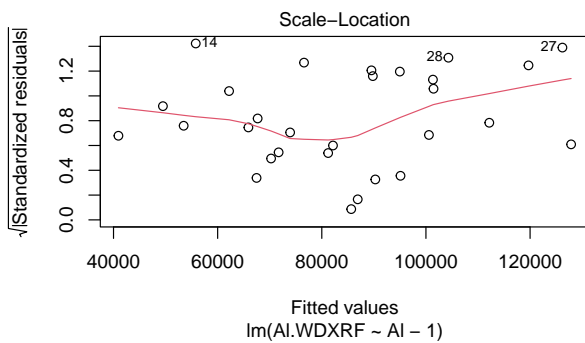
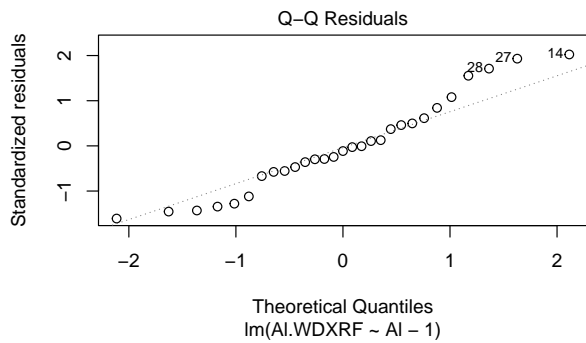
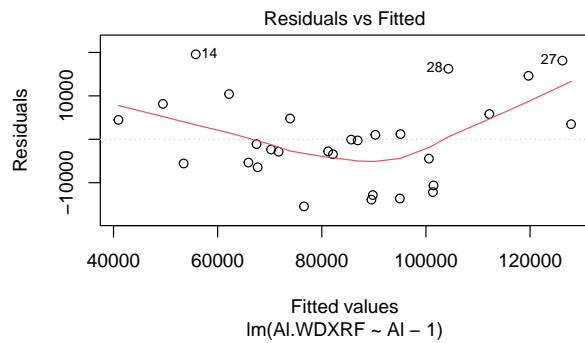
```
SEE<-sigma(OLRA1)
SEE
```

```
[1] 9128.106
```

```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 10.00219
```

```
oldparOLRA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(OLRA1)
```



NULL

```
outlierTest(OLRA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest `|rstudent|`:

	<code>rstudent</code>	unadjusted p-value	Bonferroni p
27	1.990638	0.057559	NA

```
Confint(OLRA1, level=0.90)
```

	Estimate	5 %	95 %
(Intercept)	-7683.85069	-1.961151e+04	4243.81165
AI	1.00925	8.825911e-01	1.13591

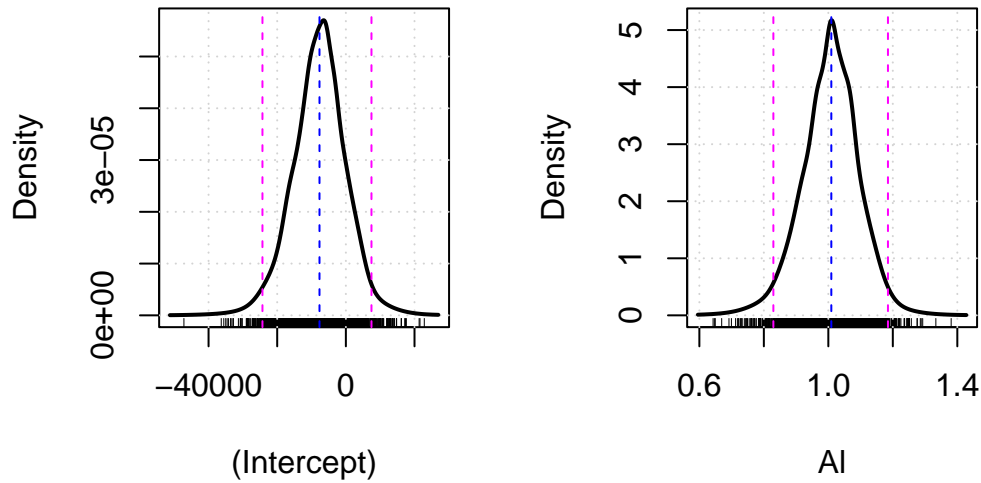
```
.bs.samplesOLRA1<- Boot(OLRA1, R=2500, method="case")
confint(.bs.samplesOLRA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
(Intercept)	-2.143322e+04	4490.556135
AI	8.596707e-01	1.156139

```
plotBoot(.bs.samplesOLRA1)
```

Bootstrap Distributions



6.3.3 Regression Trough Origin (RTO)

```
RTOA1<-lm(A1.WDXRF~A1-1, data=dataset)
summary(RTOA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1 - 1, data = dataset)
```

Residuals:

Min	1Q	Median	3Q	Max
-15073	-5507	-1385	4792	18748

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
A1	0.93018	0.01839	50.59	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 9163 on 27 degrees of freedom

Multiple R-squared: 0.9896, Adjusted R-squared: 0.9892

F-statistic: 2559 on 1 and 27 DF, p-value: < 2.2e-16

```
fitted.RTOA1 <- fitted(RTOA1)
datasetc<-cbind(dataset,fitted.RTOA1)
RMS<-sqrt(mean((datasetc$A1-datasetc$fitted.RTOA1)^2)/nrow(datasetc))
RMS
```

```
[1] 1242.61
```

```
SEE<-sigma(RTOA1)
SEE
```

```
[1] 9163.078
```

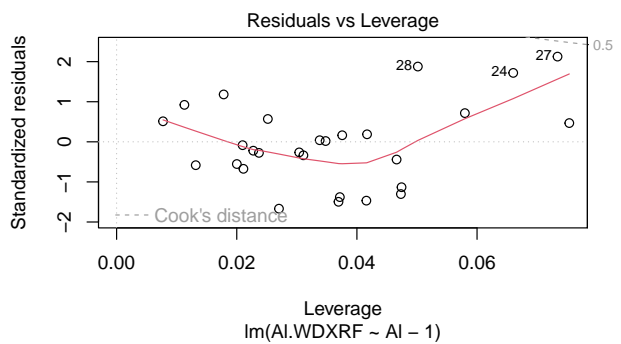
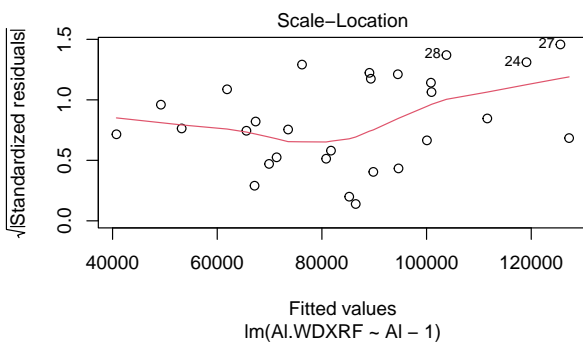
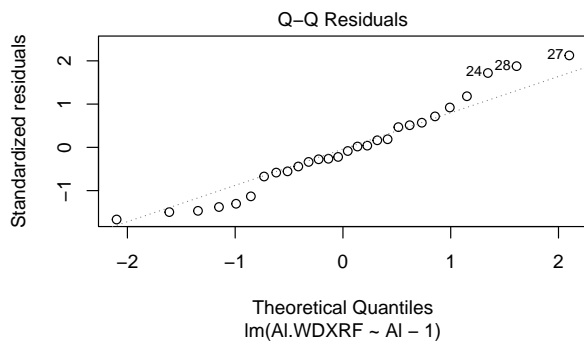
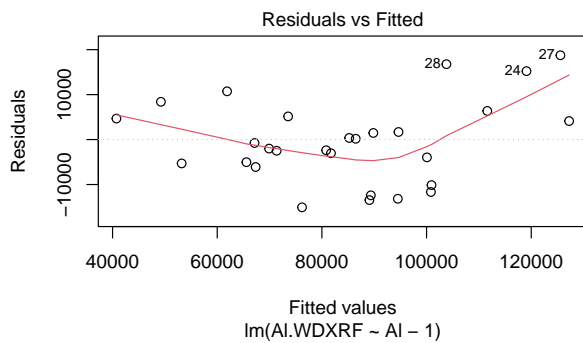
```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 10.04051
```

```
SSR<-(nrow(dataset)-1)*(SEE^2)
cSST <-sum(((dataset$A1.WDXRF)-mean(dataset$A1.WDXRF))^2)
Alcr2<-1-(SSR/cSST)
Alcr2
```

```
[1] 0.8708771
```

```
oldparRTOA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(RTOA1)
```



NULL

```
outlierTest(RTOA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest |rstudent|:

	rstudent	unadjusted p-value	Bonferroni p
27	2.285811	0.030661	0.8585

```
Confint(RTOA1, level=0.90)
```

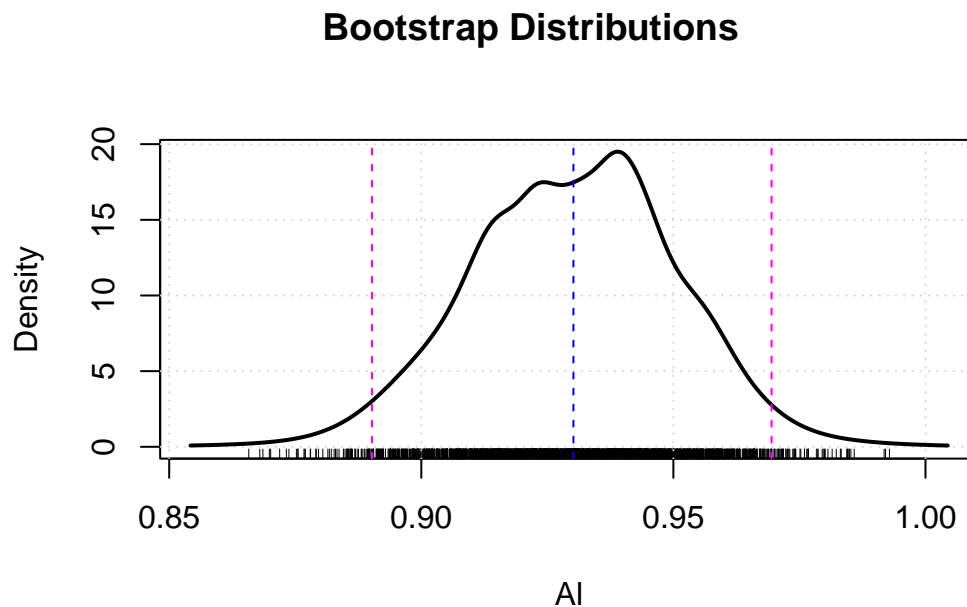
	Estimate	5 %	95 %
AI	0.9301774	0.8988566	0.9614982

```
.bs.samplesRTOA1<- Boot(RTOA1, R=2500, method="case")  
confint(.bs.samplesRTOA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
AI	0.8954861	0.9631428

```
plotBoot(.bs.samplesRTOA1)
```



6.3.4 Result

Both linear regressions give the sample in row 27 as a (non-significant) outlier. Again, r^2 is too low, rSEE still too high. Robustness is fine for OLR and RTO. So on to iteration four.

6.4 Fourth iteration

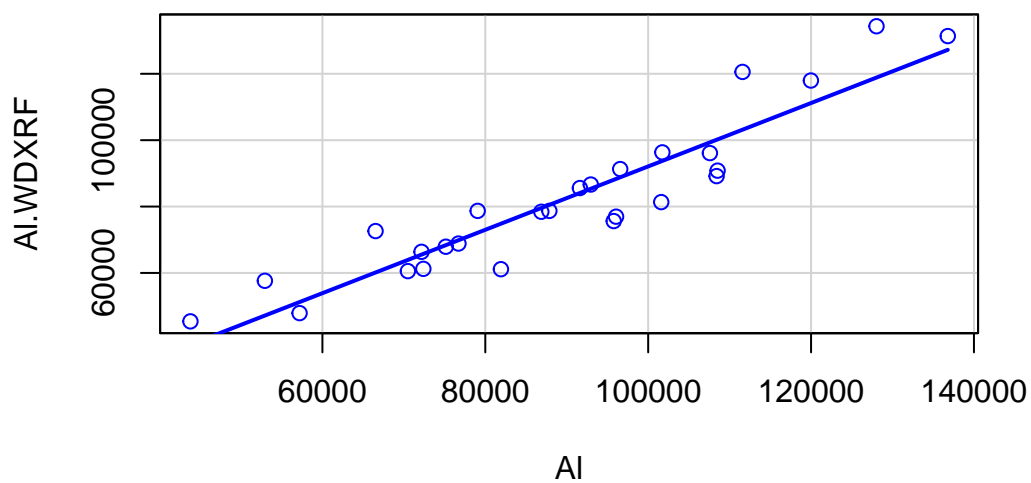
We additionally remove row 27 from the data set.

```
dataset1<- read.csv("../data_analytical//coefcorI_data.csv")
dataset<- dataset1[-c(25,14,27), ]
```

calculations

6.4.1 First impressions

```
scatterplot(A1.WDXRF~A1, regLine=TRUE, smooth=FALSE, boxplots=FALSE, data=dataset)
```



```
cor(dataset$A1,dataset$A1.WDXRF)
```

```
[1] 0.9299688
```

```
numSummary(dataset[,c("A1", "A1.WDXRF"), drop=FALSE],
            statistics=c("mean", "sd", "quantiles"),quantiles=c(0,.25,.5,.75,1))
```

	mean	sd	0%	25%	50%	75%	100%	n
A1	89641.04	22468.24	43795	73775.0	91616	104660.5	136788	27
A1.WDXRF	82202.81	23068.04	45409	67134.5	78698	91056.0	134321	27

6.4.2 Ordinary Linear Regression (OLR)

```
OLRA1<-lm(A1.WDXRF~A1, data=dataset)
summary(OLRA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1, data = dataset)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-13703.8	-3940.8	-469.6	5366.3	17405.3

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-3.386e+03	6.969e+03	-0.486	0.631
A1	9.548e-01	7.549e-02	12.648	2.31e-12 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 8649 on 25 degrees of freedom

Multiple R-squared: 0.8648, Adjusted R-squared: 0.8594

F-statistic: 160 on 1 and 25 DF, p-value: 2.308e-12

```
fitted.OLRA1 <- fitted(OLRA1)
datasetb<-cbind(dataset,fitted.OLRA1)
RMS<-sqrt(mean((datasetb$A1-datasetb$fitted.OLRA1)^2)/nrow(datasetb))
RMS
```

```
[1] 1444.281
```

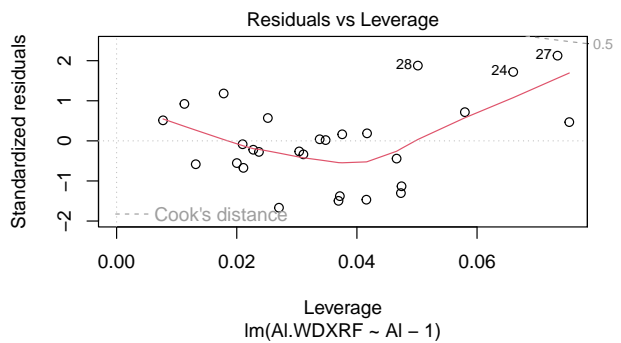
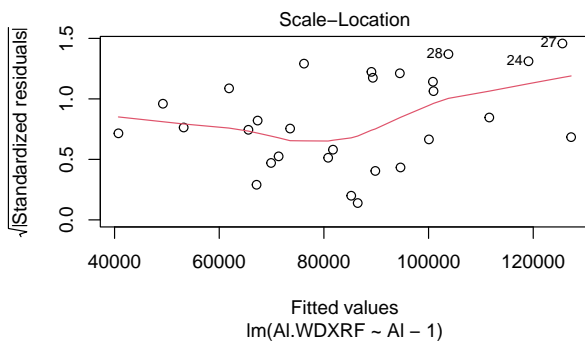
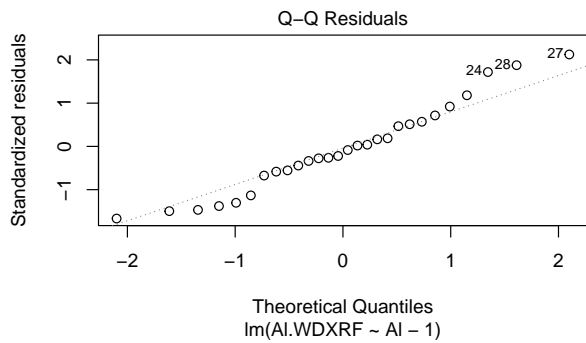
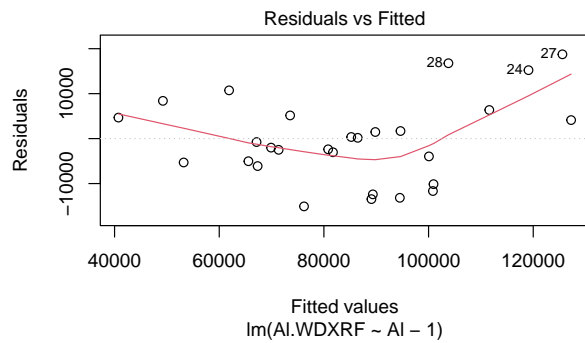
```
SEE<-sigma(OLRA1)
SEE
```

```
[1] 8648.649
```

```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 9.648091
```

```
oldparOLRA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(OLRA1)
```



NULL

```
outlierTest(OLRA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest `|rstudent|`:

	<code>rstudent</code>	unadjusted p-value	Bonferroni p
28	2.25551	0.033493	0.90431

```
Confint(OLRA1, level=0.90)
```

	Estimate	5 %	95 %
(Intercept)	-3385.9613614	-1.528955e+04	8517.624376
AI	0.9547946	8.258461e-01	1.083743

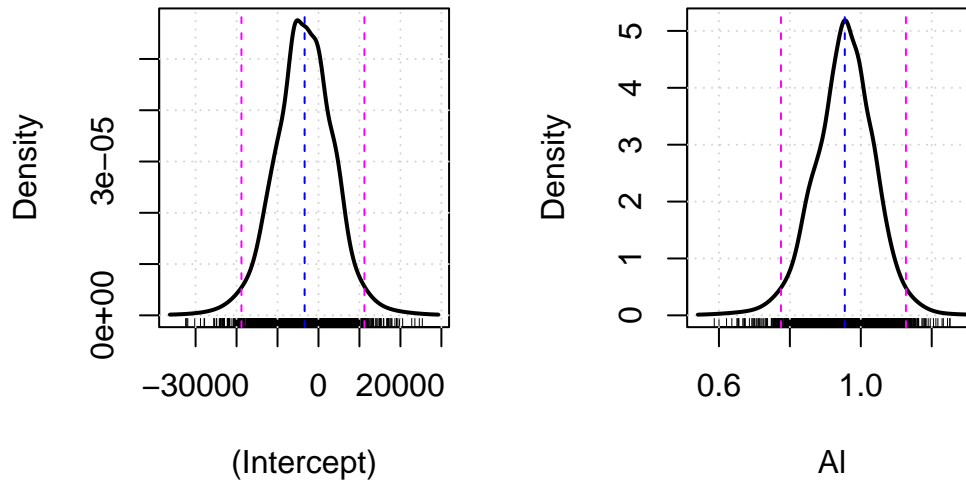
```
.bs.samplesOLRA1<- Boot(OLRA1, R=2500, method="case")
confint(.bs.samplesOLRA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
(Intercept)	-1.546008e+04	8537.623131
AI	8.116031e-01	1.091759

```
plotBoot(.bs.samplesOLRA1)
```

Bootstrap Distributions



6.4.3 Regression Trough Origin (RTO)

```
RTOA1<-lm(A1.WDXRF~A1-1, data=dataset)
summary(RTOA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1 - 1, data = dataset)
```

Residuals:

Min	1Q	Median	3Q	Max
-14172	-4994	-1179	5390	17994

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
A1	0.91918	0.01776	51.74	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 8521 on 26 degrees of freedom

Multiple R-squared: 0.9904, Adjusted R-squared: 0.99

F-statistic: 2678 on 1 and 26 DF, p-value: < 2.2e-16

```
fitted.RTOA1 <- fitted(RTOA1)
datasetc<-cbind(dataset,fitted.RTOA1)
RMS<-sqrt(mean((datasetc$A1-datasetc$fitted.RTOA1)^2)/nrow(datasetc))
RMS
```

```
[1] 1435.87
```

```
SEE<-sigma(RTOA1)
SEE
```

```
[1] 8520.646
```

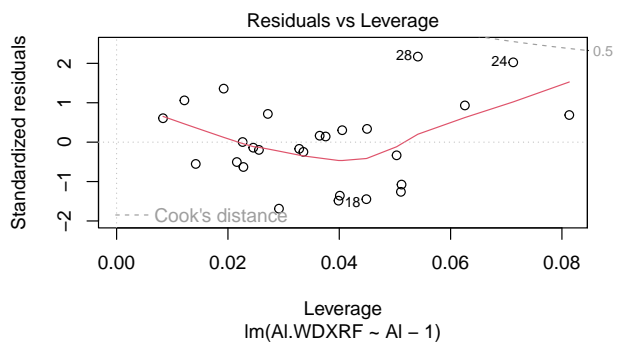
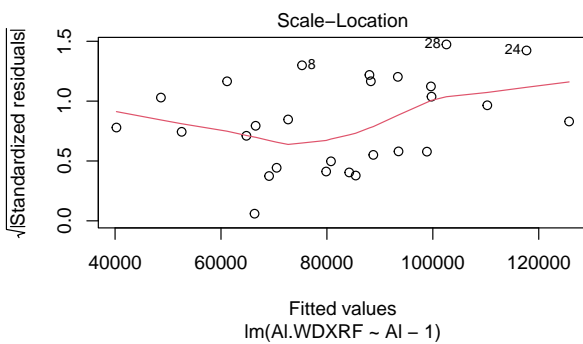
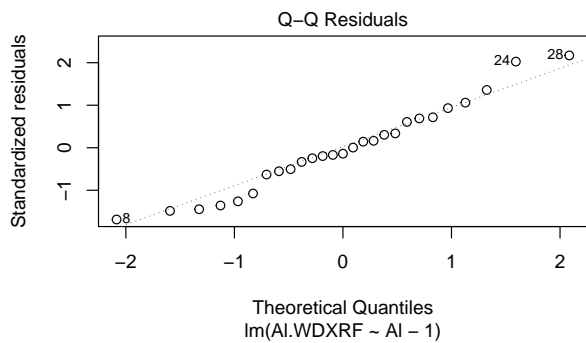
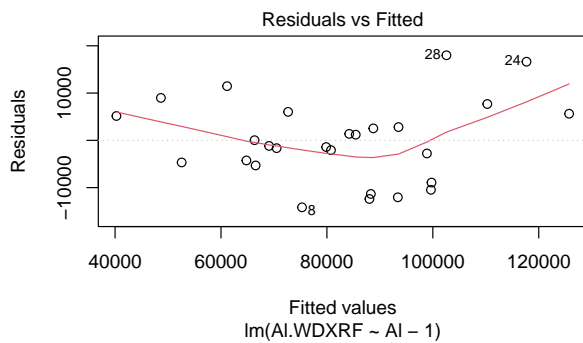
```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 9.505296
```

```
SSR<-(nrow(dataset)-1)*(SEE^2)
cSST <-sum(((dataset$A1.WDXRF)-mean(dataset$A1.WDXRF))^2)
Alcr2<-1-(SSR/cSST)
Alcr2
```

```
[1] 0.8635656
```

```
oldparRTOA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(RTOA1)
```



NULL

```
outlierTest(RTOA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest |rstudent|:

	rstudent	unadjusted p-value	Bonferroni p
28	2.35321	0.026784	0.72315

```
Confint(RTOA1, level=0.90)
```

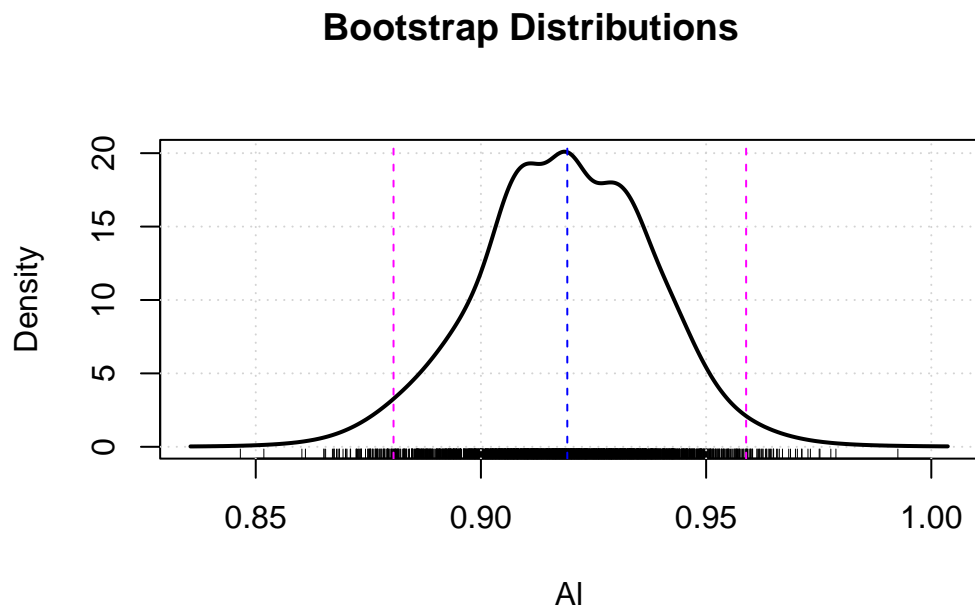
	Estimate	5 %	95 %
A1	0.9191769	0.8888791	0.9494747

```
.bs.samplesRTOA1<- Boot(RTOA1, R=2500, method="case")  
confint(.bs.samplesRTOA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
A1	0.8866721	0.951634

```
plotBoot(.bs.samplesRTOA1)
```



6.4.4 Result

Both linear regressions give the sample in row 28 as a (non-significant) outlier. Again, r^2 is too low, but rSEE and robustness are fine for OLR and RTO. So on to iteration five.

6.5 Fifth iteration

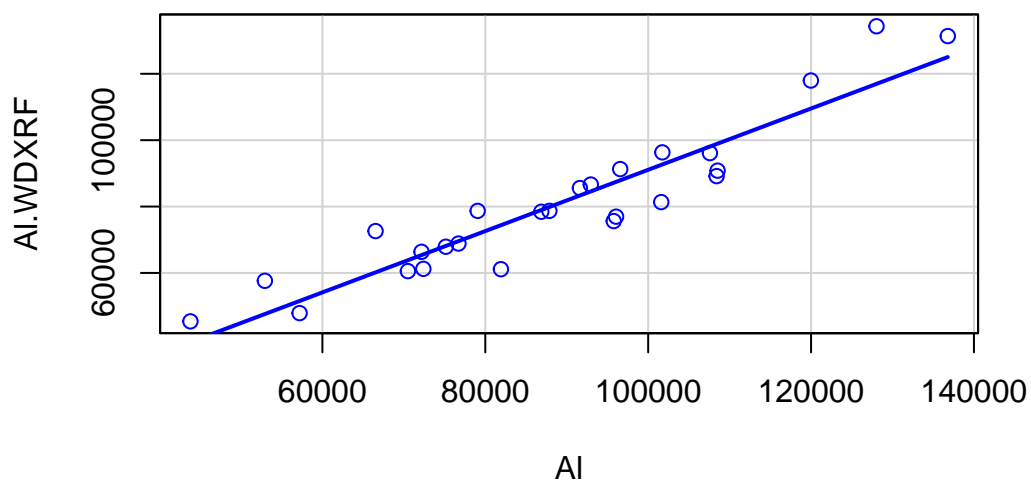
For the fifth iteration we additionally exclude row 28.

```
dataset1<- read.csv("../data_analytical//coefcorI_data.csv")
dataset<- dataset1[-c(25,14,27,28), ]
```

calculations

6.5.1 First impressions

```
scatterplot(A1.WDXRF~A1, regLine=TRUE, smooth=FALSE, boxplots=FALSE, data=dataset)
```



```
cor(dataset$A1,dataset$A1.WDXRF)
```

```
[1] 0.9352198
```

```
numSummary(dataset[,c("A1", "A1.WDXRF"), drop=FALSE],
  statistics=c("mean", "sd", "quantiles"),quantiles=c(0,.25,.5,.75,1))
```

	mean	sd	0%	25%	50%	75%	100%	n
A1	88797.0	22472.44	43795	73085.00	89743.0	101707.50	136788	26
A1.WDXRF	80727.5	22187.88	45409	66750.75	78565.5	90407.75	134321	26

6.5.2 Ordinary Linear Regression (OLR)

```
OLRA1<-lm(A1.WDXRF~A1, data=dataset)
summary(OLRA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1, data = dataset)
```

Residuals:

Min	1Q	Median	3Q	Max
-13250.4	-4175.1	-375.2	5586.8	17360.2

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-1.266e+03	6.529e+03	-0.194	0.848
A1	9.234e-01	7.136e-02	12.940	2.58e-12 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 8018 on 24 degrees of freedom

Multiple R-squared: 0.8746, Adjusted R-squared: 0.8694

F-statistic: 167.4 on 1 and 24 DF, p-value: 2.583e-12

```
fitted.OLRA1 <- fitted(OLRA1)
datasetb<-cbind(dataset,fitted.OLRA1)
RMS<-sqrt(mean((datasetb$A1-datasetb$fitted.OLRA1)^2)/nrow(datasetb))
RMS
```

```
[1] 1616.831
```

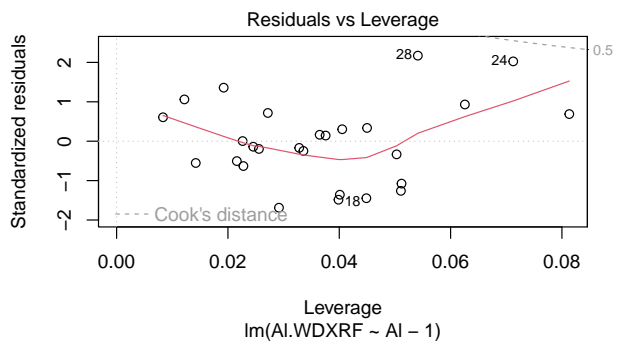
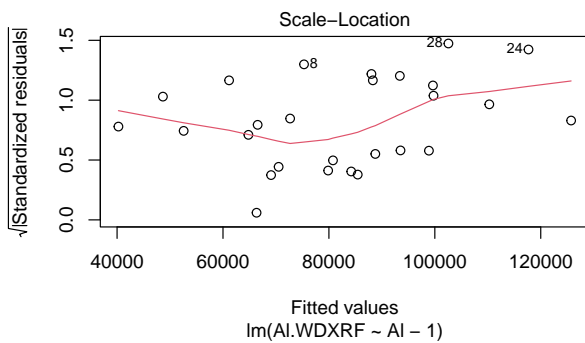
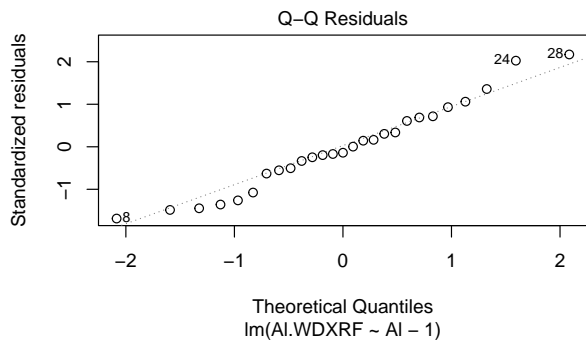
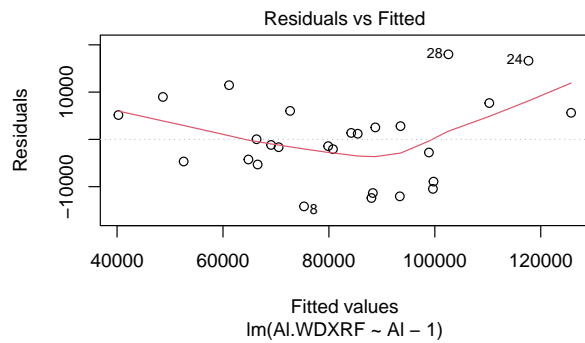
```
SEE<-sigma(OLRA1)
SEE
```

```
[1] 8018.006
```

```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 9.029591
```

```
oldparOLRA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(OLRA1)
```



NULL

```
outlierTest(OLRA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest `|rstudent|`:

	<code>rstudent</code>	unadjusted p-value	Bonferroni p
24	2.640695	0.014612	0.37991

```
Confint(OLRA1, level=0.90)
```

	Estimate	5 %	95 %
(Intercept)	-1265.6557394	-1.243536e+04	9904.052039
AI	0.9233775	8.012914e-01	1.045464

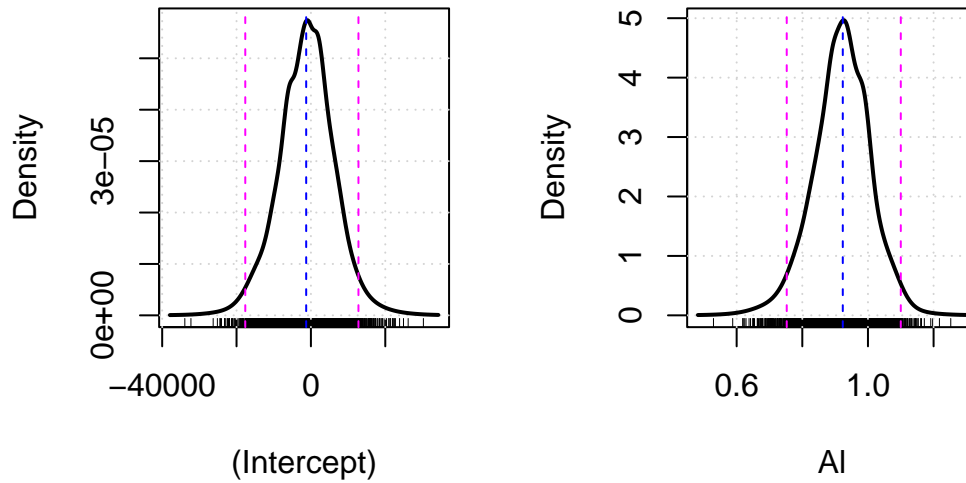
```
.bs.samplesOLRA1<- Boot(OLRA1, R=2500, method="case")
confint(.bs.samplesOLRA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
(Intercept)	-1.531969e+04	10203.353624
AI	7.814904e-01	1.076988

```
plotBoot(.bs.samplesOLRA1)
```


Bootstrap Distributions



6.5.3 Regression Trough Origin (RTO)

```
RTOA1<-lm(A1.WDXRF~A1-1, data=dataset)
summary(RTOA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1 - 1, data = dataset)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-13416.2	-4519.8	-553.1	5103.7	17813.6

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
A1	0.90995	0.01685	53.99	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 7862 on 25 degrees of freedom

Multiple R-squared: 0.9915, Adjusted R-squared: 0.9912

F-statistic: 2915 on 1 and 25 DF, p-value: < 2.2e-16

```
fitted.RTOA1 <- fitted(RTOA1)
datasetc<-cbind(dataset,fitted.RTOA1)
RMS<-sqrt(mean((datasetc$A1-datasetc$fitted.RTOA1)^2)/nrow(datasetc))
RMS
```

```
[1] 1615.725
```

```
SEE<-sigma(RTOA1)
SEE
```

```
[1] 7862.158
```

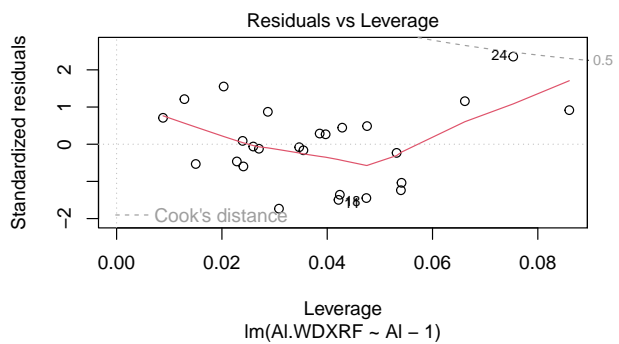
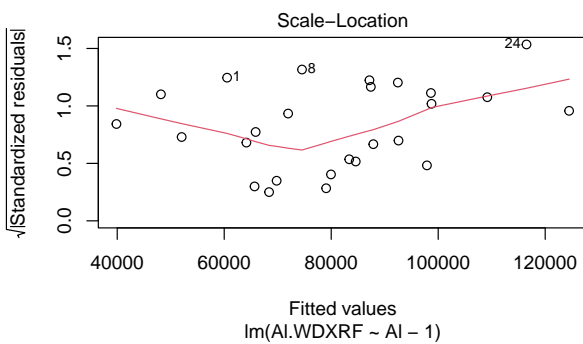
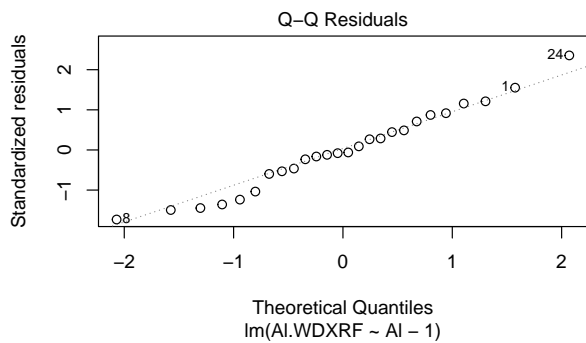
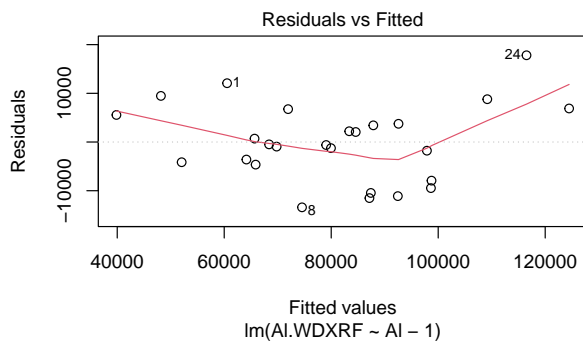
```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 8.854081
```

```
SSR<-(nrow(dataset)-1)*(SEE^2)
cSST <-sum(((dataset$A1.WDXRF)-mean(dataset$A1.WDXRF))^2)
Alcr2<-1-(SSR/cSST)
Alcr2
```

```
[1] 0.8744398
```

```
oldparRTOA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(RTOA1)
```



NULL

```
outlierTest(RTOA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest |rstudent|:

	rstudent	unadjusted p-value	Bonferroni p
24	2.617467	0.015096	0.39249

```
Confint(RTOA1, level=0.90)
```

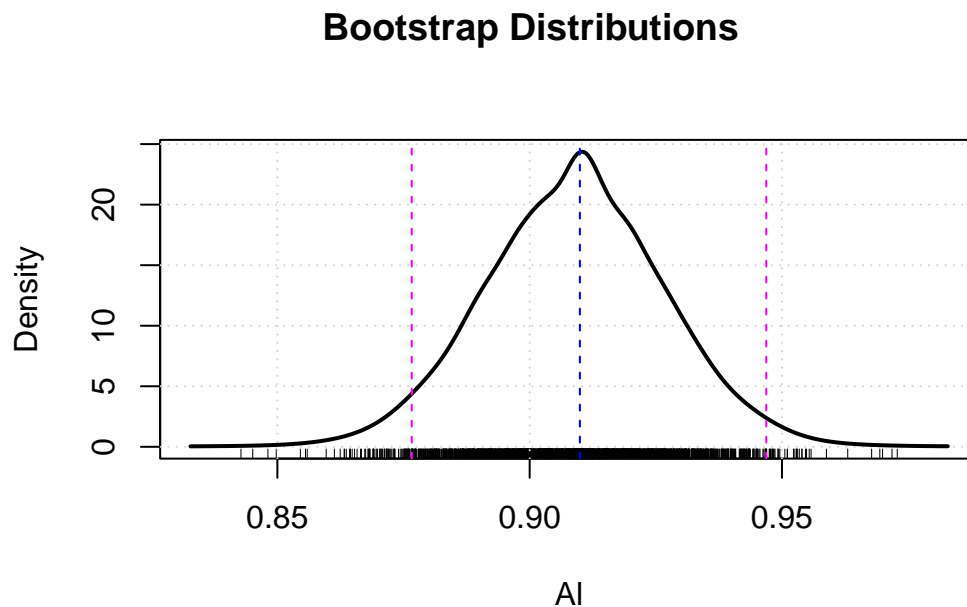
	Estimate	5 %	95 %
AI	0.909951	0.8811636	0.9387385

```
.bs.samplesRTOA1<- Boot(RTOA1, R=2500, method="case")  
confint(.bs.samplesRTOA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
AI	0.8815718	0.9421537

```
plotBoot(.bs.samplesRTOA1)
```



6.5.4 Result

In the fifth iteration, r^2 is still too low for OLR and RTO, but $rSEE$ is within the required range and robustness is also given. Both linear regressions show the sample in row 24 as an outlier, which is why it is removed for the sixth iteration.

6.6 Sixth iteration

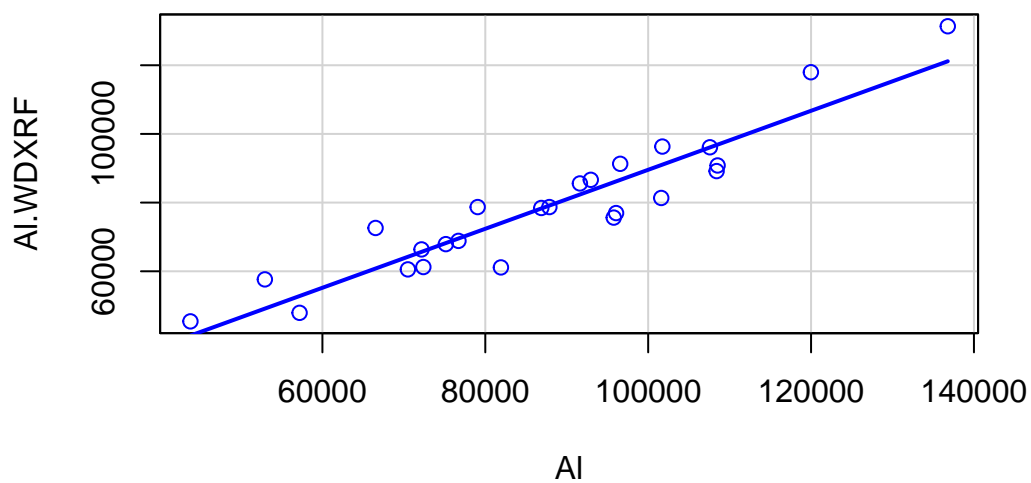
Now we exclude row 24.

```
dataset1<- read.csv("../data_analytical//coefcorI_data.csv")
dataset<- dataset1[-c(25,14,27,28,24), ]
```

calculations

6.6.1 First impressions

```
scatterplot(A1.WDXRF~A1, regLine=TRUE, smooth=FALSE, boxplots=FALSE, data=dataset)
```



```
cor(dataset$A1,dataset$A1.WDXRF)
```

```
[1] 0.9343291
```

```
numSummary(dataset[,c("A1", "A1.WDXRF"), drop=FALSE],
            statistics=c("mean", "sd", "quantiles"),quantiles=c(0,.25,.5,.75,1))
```

	mean	sd	0%	25%	50%	75%	100%	n
A1	87227.40	21431.96	43795	72395	87870	101604	136788	25
A1.WDXRF	78583.76	19706.59	45409	66367	78433	89177	131358	25

6.6.2 Ordinary Linear Regression (OLR)

```
OLRA1<-lm(A1.WDXRF~A1, data=dataset)
summary(OLRA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1, data = dataset)
```

Residuals:

Min	1Q	Median	3Q	Max
-12897.1	-4888.3	41.9	4687.3	11813.1

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	3.646e+03	6.131e+03	0.595	0.558
A1	8.591e-01	6.833e-02	12.572	8.67e-12 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 7175 on 23 degrees of freedom

Multiple R-squared: 0.873, Adjusted R-squared: 0.8674

F-statistic: 158.1 on 1 and 23 DF, p-value: 8.675e-12

```
fitted.OLRA1 <- fitted(OLRA1)
datasetb<-cbind(dataset,fitted.OLRA1)
RMS<-sqrt(mean((datasetb$A1-datasetb$fitted.OLRA1)^2)/nrow(datasetb))
RMS
```

```
[1] 1827.187
```

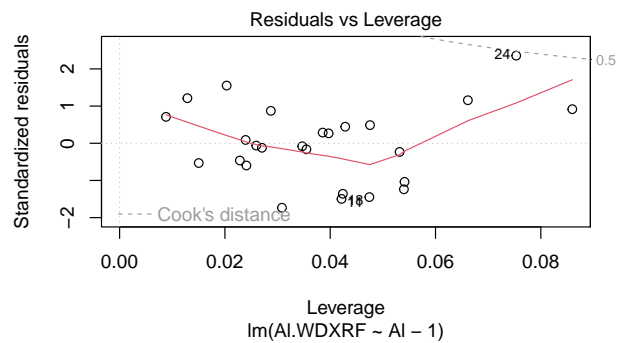
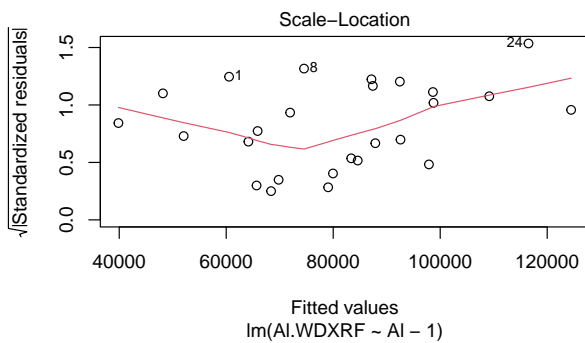
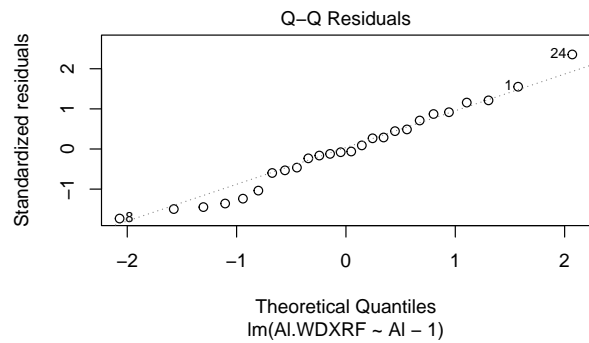
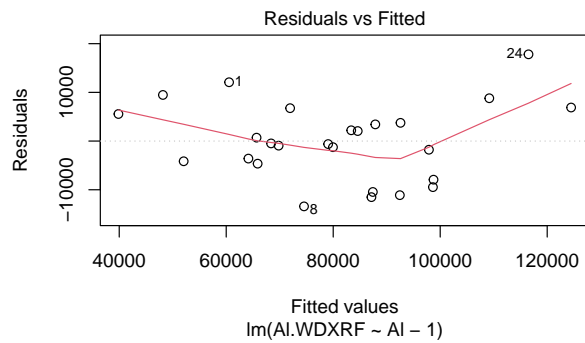
```
SEE<-sigma(OLRA1)
SEE
```

```
[1] 7174.72
```

```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 8.225305
```

```
oldparOLRA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(OLRA1)
```



NULL

```
outlierTest(OLRA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest `|rstudent|`:

	<code>rstudent</code>	unadjusted p-value	Bonferroni p
8	-1.945074	0.064657	NA

```
Confint(OLRA1, level=0.90)
```

	Estimate	5 %	95 %
(Intercept)	3645.6923078	-6861.8747902	1.415326e+04
AI	0.8591116	0.7419957	9.762274e-01

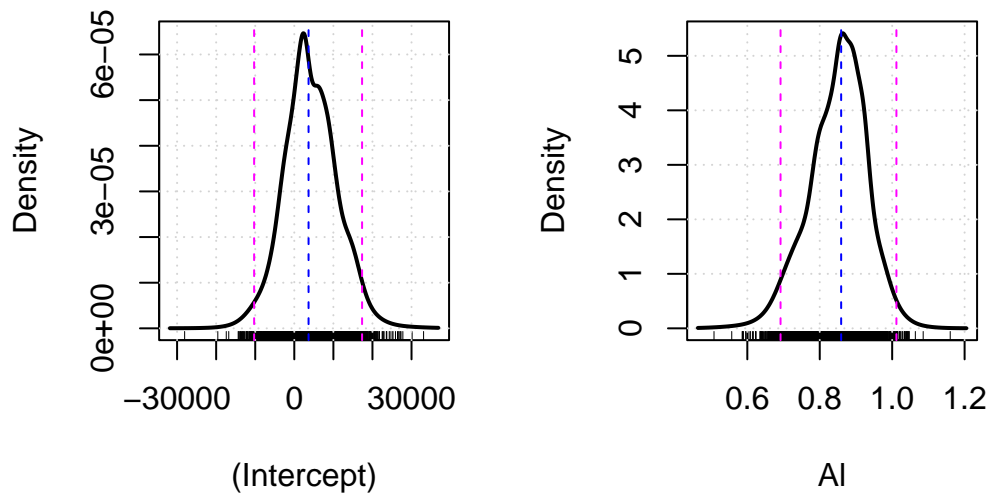
```
.bs.samplesOLRA1<- Boot(OLRA1, R=2500, method="case")
confint(.bs.samplesOLRA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
(Intercept)	-7726.5410289	1.537624e+04
AI	0.7166626	9.857368e-01

```
plotBoot(.bs.samplesOLRA1)
```

Bootstrap Distributions



6.6.3 Regression Trough Origin (RTO)

```
RTOA1<-lm(A1.WDXRF~A1-1, data=dataset)
summary(RTOA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1 - 1, data = dataset)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-12487.7	-3822.4	363.8	4894.9	12830.6

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
A1	0.89862	0.01578	56.96	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 7077 on 24 degrees of freedom

Multiple R-squared: 0.9927, Adjusted R-squared: 0.9924

F-statistic: 3244 on 1 and 24 DF, p-value: < 2.2e-16

```
fitted.RTOA1 <- fitted(RTOA1)
datasetc<-cbind(dataset,fitted.RTOA1)
RMS<-sqrt(mean((datasetc$A1-datasetc$fitted.RTOA1)^2)/nrow(datasetc))
RMS
```

```
[1] 1819.2
```

```
SEE<-sigma(RTOA1)
SEE
```

```
[1] 7077.441
```

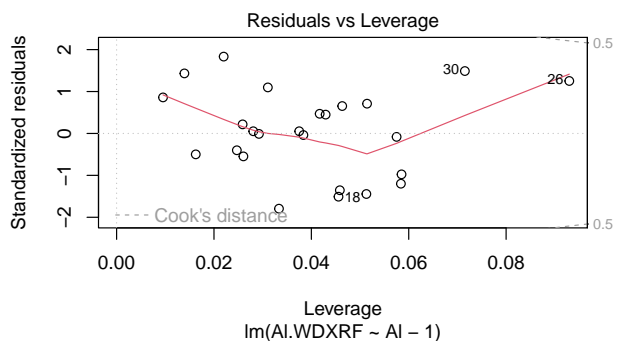
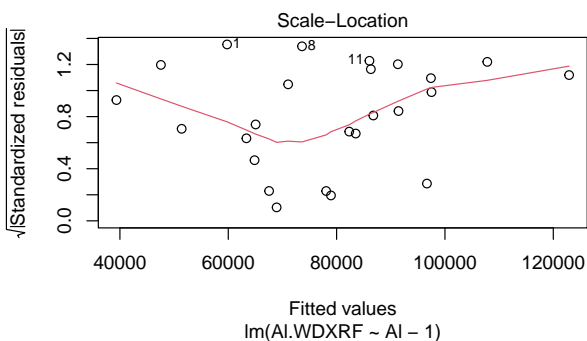
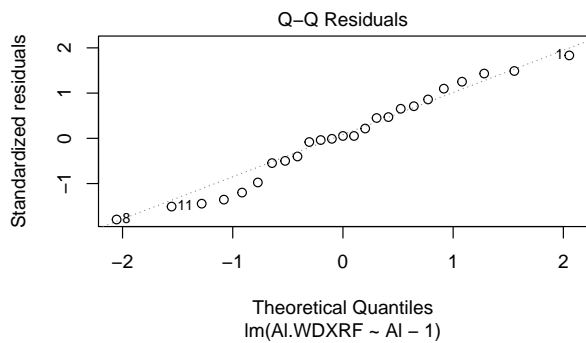
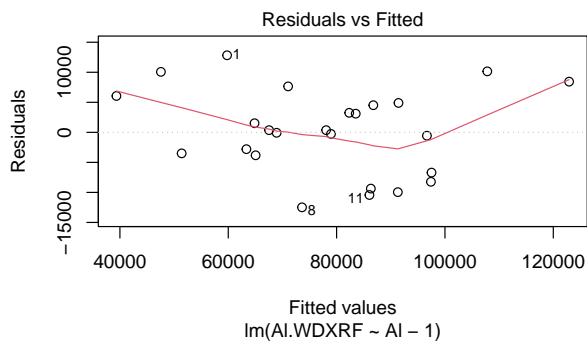
```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 8.113782
```

```
SSR<-(nrow(dataset)-1)*(SEE^2)
cSST <-sum(((dataset$A1.WDXRF)-mean(dataset$A1.WDXRF))^2)
Alcr2<-1-(SSR/cSST)
Alcr2
```

```
[1] 0.8710179
```

```
oldparRTOA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(RTOA1)
```



NULL

```
outlierTest(RTOA1)
```


No Studentized residuals with Bonferroni $p < 0.05$

Largest |rstudent|:

	rstudent	unadjusted p-value	Bonferroni p
1	1.935143	0.065363	NA

```
Confint(RTOA1, level=0.90)
```

	Estimate	5 %	95 %
A1	0.8986173	0.871625	0.9256096

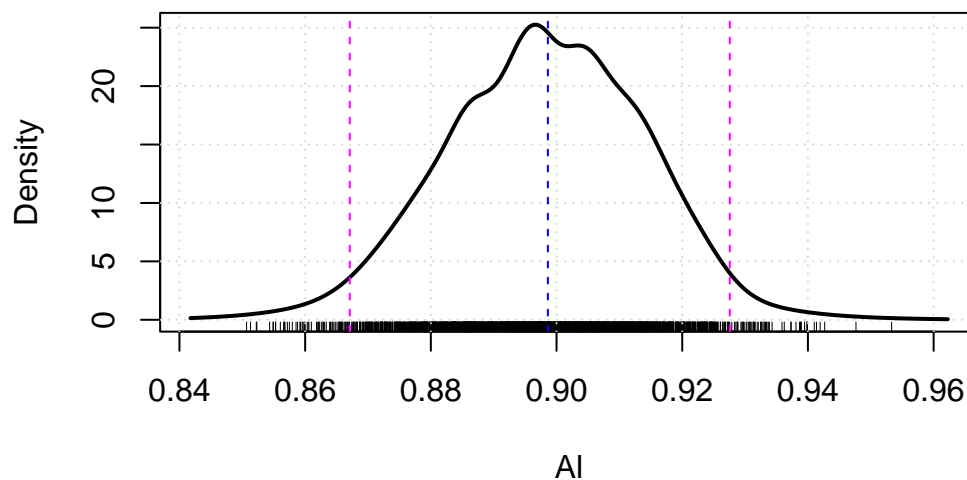
```
.bs.samplesRTOA1<- Boot(RTOA1, R=2500, method="case")  
confint(.bs.samplesRTOA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
A1	0.8721035	0.9235699

```
plotBoot(.bs.samplesRTOA1)
```

Bootstrap Distributions



6.6.4 Result

In this run we see different results: r^2 is past the required limits, robustness and rSEE are fine for both linear regressions. Yet, the sample in row 8 qualifies as a (non-significant) outlier for the OLR, sample 1 for RTO. A direct comparison of the characteristic values reveals the following:

criteria	OLR	RTO
r ²	0.873	0.8710
rSEE	8.23	8.11
CI 0.05	0.7420 - 0.7164	0.8716 - 0.8707
CI 0.95	0.9762 - 0.9920	0.9256 - 0.9231

The direct comparison shows very clearly that the RTO has the better values or proportions of values in all criteria. For this reason - and because r² is still too low - the sample in row 1 is excluded for the next iteration.

Different values by the bootstrap?

If you do the calculations yourself, the values of CIBS5 and CIBS95 will most likely differ slightly from those in the table. Don't worry - this is completely normal! As long as the differences between your values and those in the table are not too great (i.e. differences of more than 0.05) everything is fine.

6.7 Seventh iteration

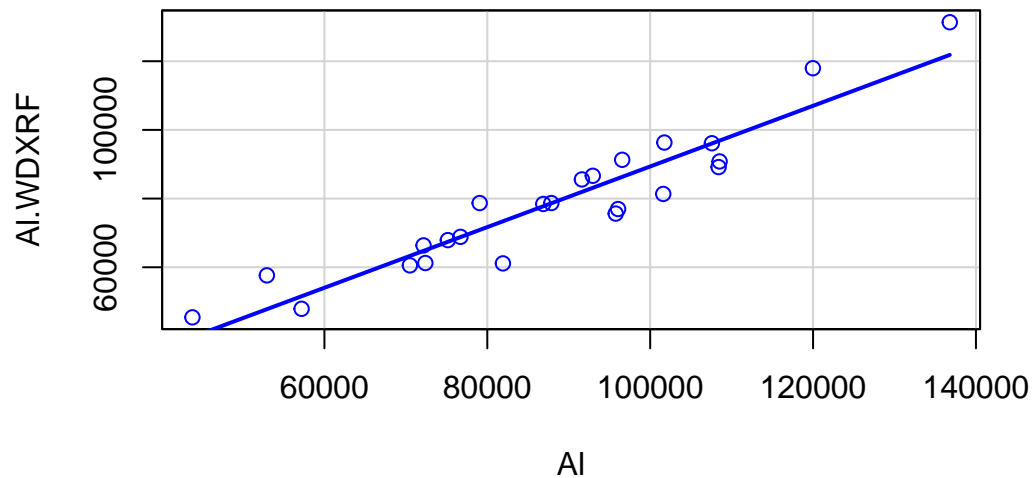
And out with row 1.

```
dataset1<- read.csv("../data_analytical//coefcorI_data.csv")
dataset<- dataset1[-c(25,14,27,28,24,1), ]
```

calculations

6.7.1 First impressions

```
scatterplot(A1.WDXRF~A1, regLine=TRUE, smooth=FALSE, boxplots=FALSE, data=dataset)
```



```
cor(dataset$AI,dataset$AI.WDXRF)
```

```
[1] 0.9427524
```

```
numSummary(dataset[,c("AI", "AI.WDXRF"), drop=FALSE],
  statistics=c("mean", "sd", "quantiles"),quantiles=c(0,.25,.5,.75,1))
```

	mean	sd	0%	25%	50%	75%	100%	n
AI	88089.96	21445.07	43795	74465.0	89743.0	101638.50	136788	24
AI.WDXRF	78832.58	20090.28	45409	65083.5	78565.5	89587.25	131358	24

6.7.2 Ordinary Linear Regression (OLR)

```
OLRA1<-lm(AI.WDXRF~AI, data=dataset)
summary(OLRA1)
```

Call:

```
lm(formula = AI.WDXRF ~ AI, data = dataset)
```

Residuals:

Min	1Q	Median	3Q	Max
-12256	-4319	285	5090	10971

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	1.032e+03	6.032e+03	0.171	0.866
AI	8.832e-01	6.661e-02	13.259	5.72e-12 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6851 on 22 degrees of freedom

Multiple R-squared: 0.8888, Adjusted R-squared: 0.8837

F-statistic: 175.8 on 1 and 22 DF, p-value: 5.716e-12

```
fitted.OLRA1 <- fitted(OLRA1)
datasetb<-cbind(dataset,fitted.OLRA1)
RMS<-sqrt(mean((datasetb$A1-datasetb$fitted.OLRA1)^2)/nrow(datasetb))
RMS
```

[1] 1954.824

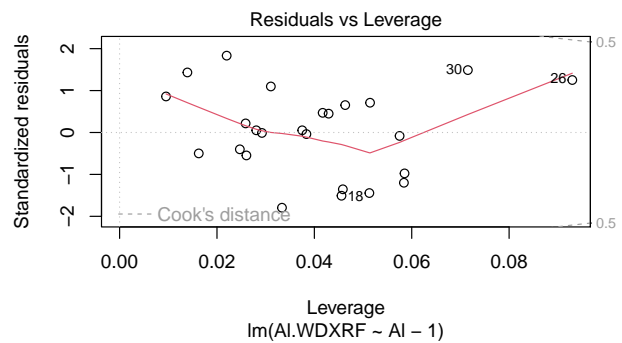
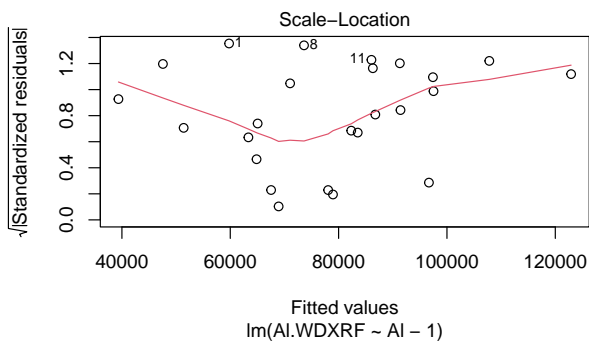
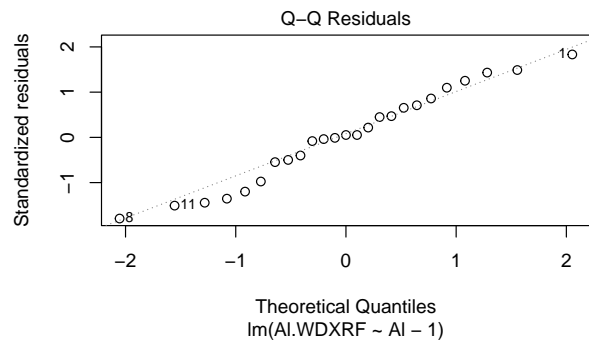
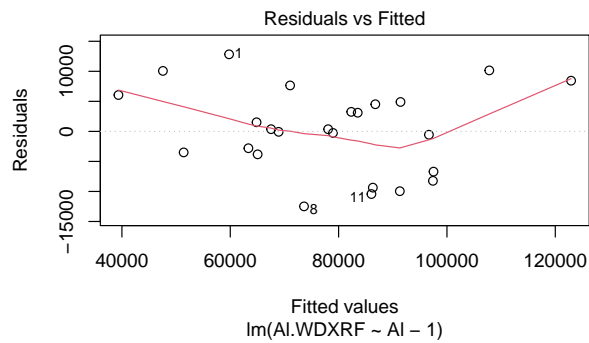
```
SEE<-sigma(OLRA1)
SEE
```

[1] 6850.562

```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

[1] 7.77678

```
oldparOLRA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(OLRA1)
```



NULL

```
outlierTest(OLRA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest `|rstudent|`:

	<code>rstudent</code>	unadjusted p-value	Bonferroni p
8	-1.943091	0.065531	NA

```
Confint(OLRA1, level=0.90)
```

	Estimate	5 %	95 %
(Intercept)	1032.0240723	-9325.6851601	1.138973e+04
AI	0.8831944	0.7688166	9.975722e-01

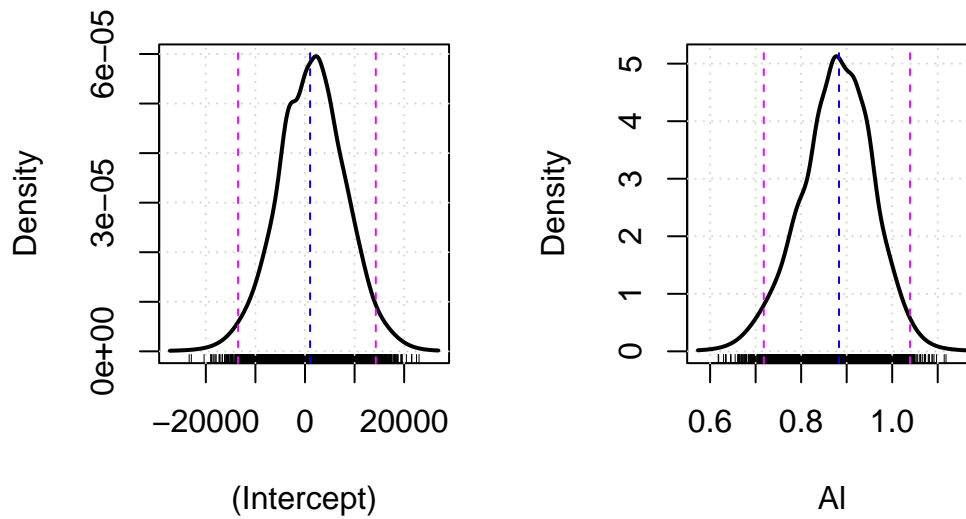
```
.bs.samplesOLRA1<- Boot(OLRA1, R=2500, method="case")
confint(.bs.samplesOLRA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
(Intercept)	-1.126613e+04	11856.108768
AI	7.472384e-01	1.017284

```
plotBoot(.bs.samplesOLRA1)
```

Bootstrap Distributions



6.7.3 Regression Trough Origin (RTO)

```
RTOA1<-lm(A1.WDXRF~A1-1, data=dataset)
summary(RTOA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1 - 1, data = dataset)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-12132.4	-4189.3	475.3	5036.7	10673.1

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
A1	0.89428	0.01511	59.17	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6704 on 23 degrees of freedom

Multiple R-squared: 0.9935, Adjusted R-squared: 0.9932

F-statistic: 3502 on 1 and 23 DF, p-value: < 2.2e-16

```
fitted.RTOA1 <- fitted(RTOA1)
datasetc<-cbind(dataset,fitted.RTOA1)
RMS<-sqrt(mean((datasetc$A1-datasetc$fitted.RTOA1)^2)/nrow(datasetc))
RMS
```

```
[1] 1954.214
```

```
SEE<-sigma(RTOA1)
SEE
```

```
[1] 6704.438
```

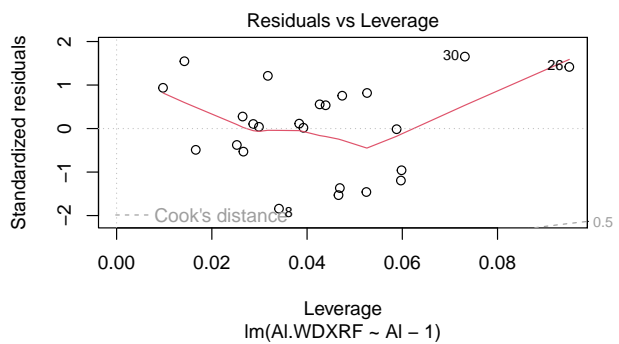
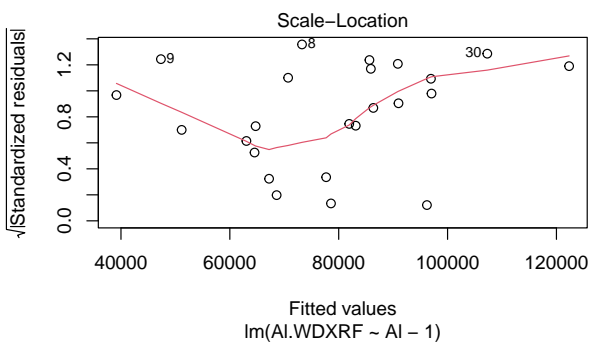
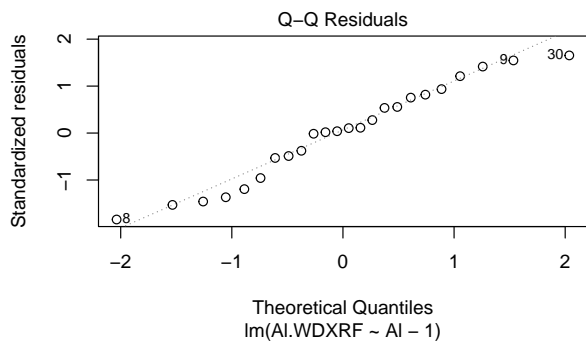
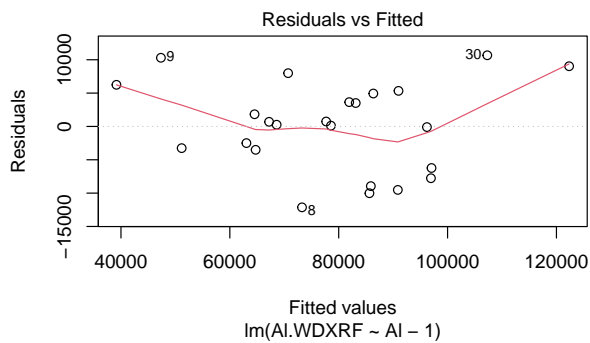
```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 7.610899
```

```
SSR<-(nrow(dataset)-1)*(SEE^2)
cSST <-sum(((dataset$A1.WDXRF)-mean(dataset$A1.WDXRF))^2)
Alcr2<-1-(SSR/cSST)
Alcr2
```

```
[1] 0.888634
```

```
oldparRTOA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(RTOA1)
```



NULL

```
outlierTest(RTOA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest |rstudent|:

	rstudent	unadjusted p-value	Bonferroni p
8	-1.950273	0.063999	NA

```
Confint(RTOA1, level=0.90)
```

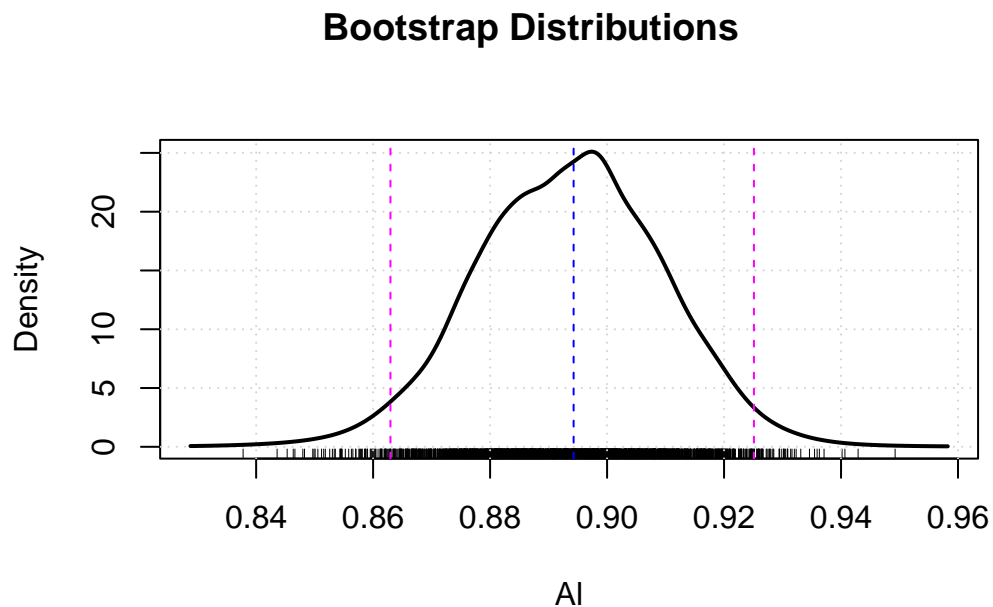
	Estimate	5 %	95 %
AI	0.8942803	0.8683796	0.9201811

```
.bs.samplesRTOA1<- Boot(RTOA1, R=2500, method="case")  
confint(.bs.samplesRTOA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
AI	0.8677046	0.9202462

```
plotBoot(.bs.samplesRTOA1)
```



6.7.4 Result

Both linear regressions give the sample in row 8 as a (non-significant) outlier. Again, r^2 is too low, but rSEE and robustness are fine for OLR and RTO.

6.8 Eighth iteration

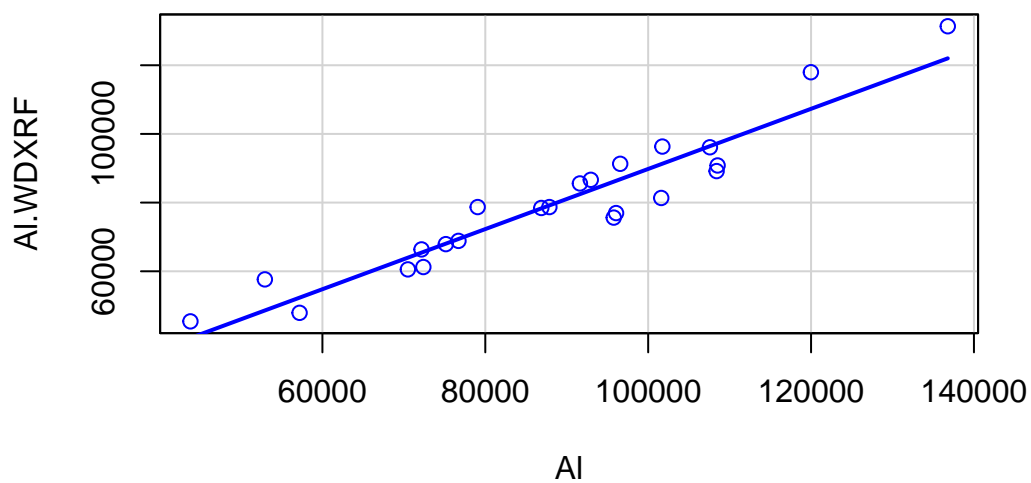
Lets exclude row 8, too.

```
dataset1<- read.csv("../data_analytical//coefcorI_data.csv")
dataset<- dataset1[-c(25,14,27,28,24,1,8), ]
```

calculations

6.8.1 First impressions

```
scatterplot(A1.WDXRF~A1, regLine=TRUE, smooth=FALSE, boxplots=FALSE, data=dataset)
```



```
cor(dataset$A1,dataset$A1.WDXRF)
```

```
[1] 0.9498882
```

```
numSummary(dataset[,c("A1", "A1.WDXRF"), drop=FALSE],
            statistics=c("mean", "sd", "quantiles"),quantiles=c(0,.25,.5,.75,1))
```

	mean	sd	0%	25%	50%	75%	100%	n
A1	88358.22	21885.83	43795	73775.0	91616	101673.0	136788	23
A1.WDXRF	79602.39	20176.64	45409	67134.5	78698	89997.5	131358	23

6.8.2 Ordinary Linear Regression (OLR)

```
OLRA1<-lm(A1.WDXRF~A1, data=dataset)
summary(OLRA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1, data = dataset)
```

Residuals:

Min	1Q	Median	3Q	Max
-10464.9	-4404.3	-138.3	4667.4	10675.1

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	2.227e+03	5.717e+03	0.389	0.701
A1	8.757e-01	6.289e-02	13.925	4.47e-12 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6455 on 21 degrees of freedom

Multiple R-squared: 0.9023, Adjusted R-squared: 0.8976

F-statistic: 193.9 on 1 and 21 DF, p-value: 4.471e-12

```
fitted.OLRA1 <- fitted(OLRA1)
datasetb<-cbind(dataset,fitted.OLRA1)
RMS<-sqrt(mean((datasetb$A1-datasetb$fitted.OLRA1)^2)/nrow(datasetb))
RMS
```

```
[1] 1908.136
```

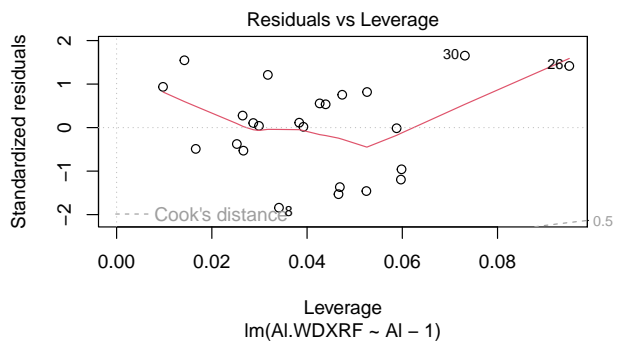
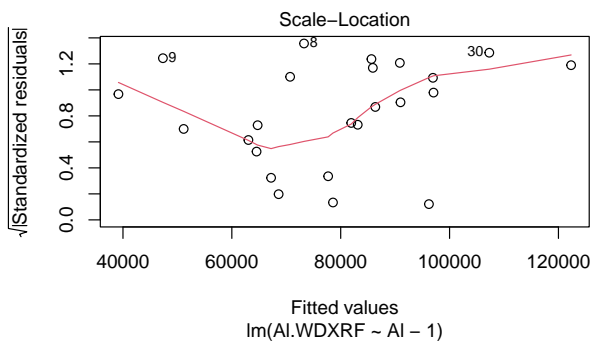
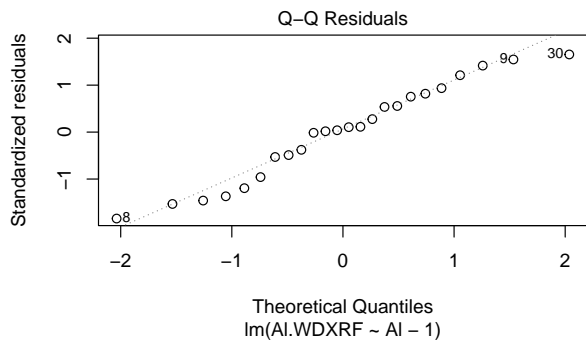
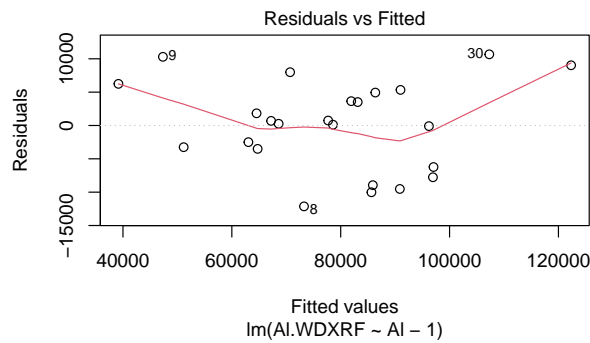
```
SEE<-sigma(OLRA1)
SEE
```

```
[1] 6455.434
```

```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 7.30598
```

```
oldparOLRA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(OLRA1)
```



NULL

```
outlierTest(OLRA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest `|rstudent|`:

	rstudent	unadjusted p-value	Bonferroni p
30	1.886986	0.073764	NA

```
Confint(OLRA1, level=0.90)
```

	Estimate	5 %	95 %
(Intercept)	2226.5502167	-7611.2380481	1.206434e+04
AI	0.8757062	0.7674963	9.839162e-01

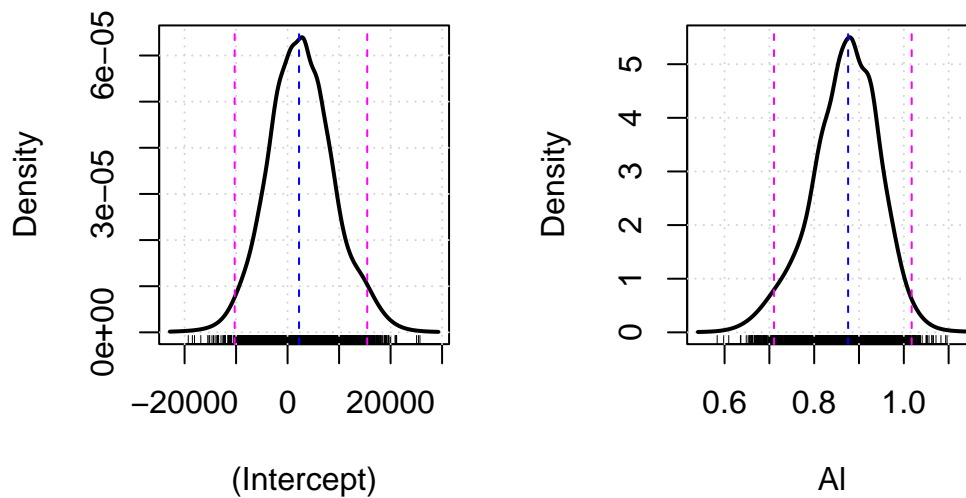
```
.bs.samplesOLRA1<- Boot(OLRA1, R=2500, method="case")
confint(.bs.samplesOLRA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
(Intercept)	-8691.0294715	1.338129e+04
AI	0.7391363	9.949074e-01

```
plotBoot(.bs.samplesOLRA1)
```

Bootstrap Distributions



6.8.3 Regression Trough Origin (RTO)

```
RTOA1<-lm(A1.WDXRF~A1-1, data=dataset)
summary(RTOA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1 - 1, data = dataset)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-10517.9	-3720.1	286.4	4618.1	10045.9

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
A1	0.89951	0.01452	61.96	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6330 on 22 degrees of freedom

Multiple R-squared: 0.9943, Adjusted R-squared: 0.994

F-statistic: 3839 on 1 and 22 DF, p-value: < 2.2e-16

```
fitted.RTOA1 <- fitted(RTOA1)
datasetc<-cbind(dataset,fitted.RTOA1)
RMS<-sqrt(mean((datasetc$A1-datasetc$fitted.RTOA1)^2)/nrow(datasetc))
RMS
```

```
[1] 1905.003
```

```
SEE<-sigma(RTOA1)
SEE
```

```
[1] 6329.748
```

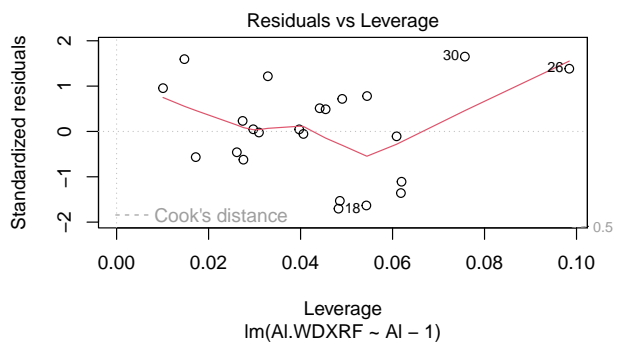
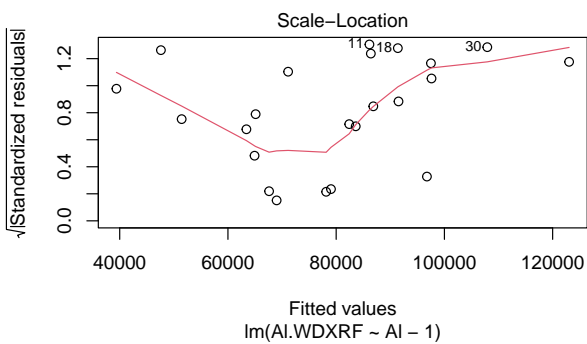
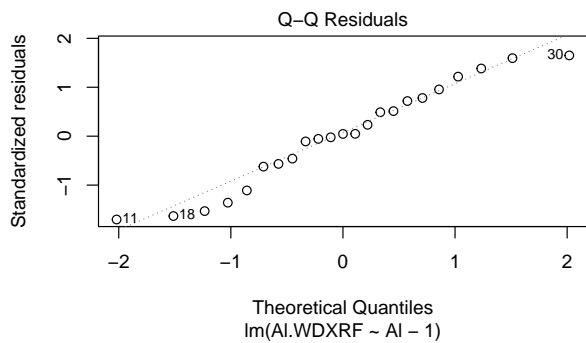
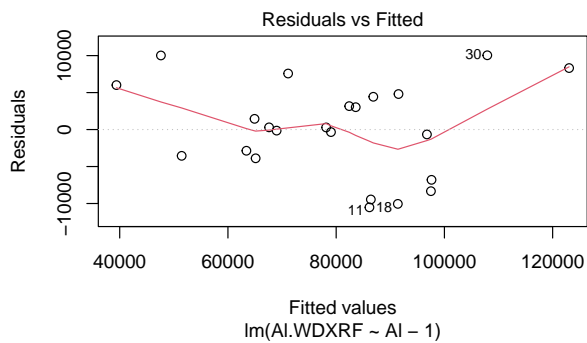
```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 7.163735
```

```
SSR<-(nrow(dataset)-1)*(SEE^2)
cSST <-sum(((dataset$A1.WDXRF)-mean(dataset$A1.WDXRF))^2)
Alcr2<-1-(SSR/cSST)
Alcr2
```

```
[1] 0.9015819
```

```
oldparRTOA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(RTOA1)
```



NULL

```
outlierTest(RTOA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest |rstudent|:

	rstudent	unadjusted p-value	Bonferroni p
11	-1.786024	0.088544	NA

```
Confint(RTOA1, level=0.90)
```

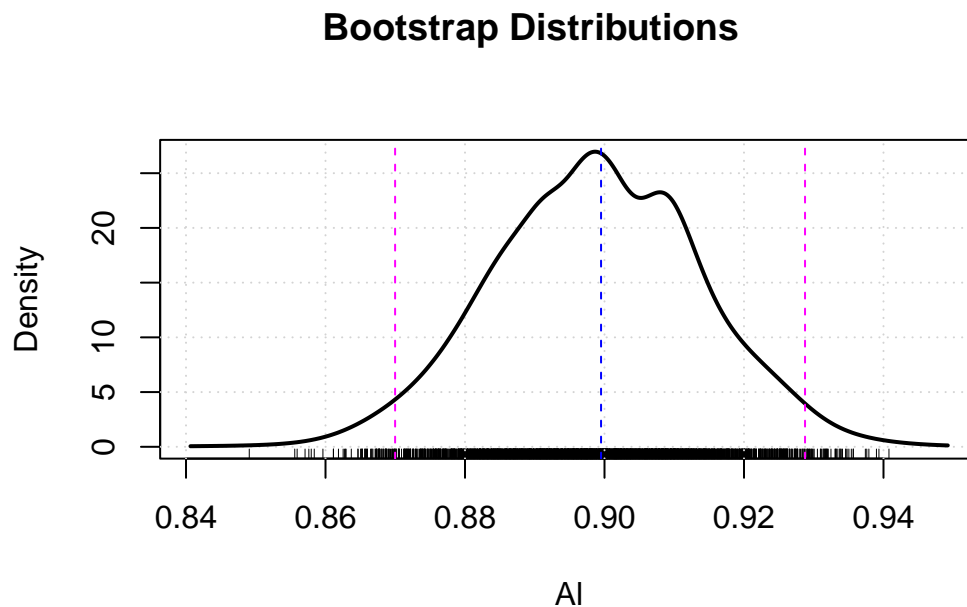
	Estimate	5 %	95 %
AI	0.8995085	0.8745799	0.9244372

```
.bs.samplesRTOA1<- Boot(RTOA1, R=2500, method="case")  
confint(.bs.samplesRTOA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
AI	0.8750248	0.925148

```
plotBoot(.bs.samplesRTOA1)
```



6.8.4 Result

Now we are getting somewhere: r^2 , robustness and rSEE are fine for both linear regressions. The sample in row 30 qualifies as a (non-significant) outlier for the OLR, sample 11 for RTO. A direct comparison of the characteristic values reveals the following:

criteria	OLR	RTO
r^2	0.9023	0.9016
rSEE	7.31	7.16
CI 0.05	0.7675 - 0.7450	0.8746 - 0.8723
CI 0.95	0.9839 - 0.9982	0.9244 - 0.9238

The direct comparison shows again that the RTO has the better values or proportions of values in all criteria except for r^2 which is very slightly better for the OLR. Still, the RTO is deemed better. As r^2 is very close to the benchmark the sample in row 11 is excluded for the next iteration to see if the values are getting better.

6.9 Ninth iteration

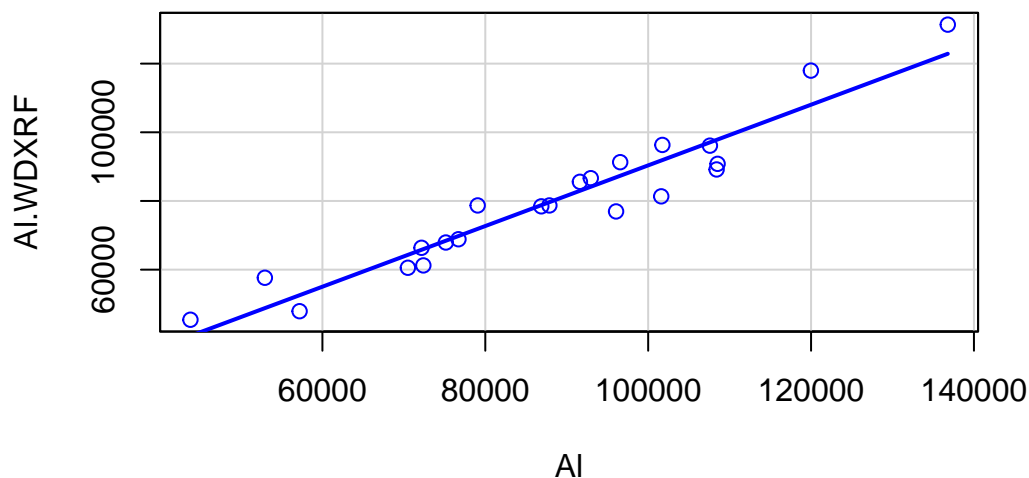
Proceeding with excluding row 11.

```
dataset1<- read.csv("../data_analytical//coefcorI_data.csv")
dataset<- dataset1[-c(25,14,27,28,24,1,8,11), ]
```

calculations

6.9.1 First impressions

```
scatterplot(AI.WDXRF~AI, regLine=TRUE, smooth=FALSE, boxplots=FALSE, data=dataset)
```



```
cor(dataset$AI,dataset$AI.WDXRF)
```

```
[1] 0.9565483
```

```
numSummary(dataset[,c("A1", "A1.WDXRF"), drop=FALSE],
            statistics=c("mean", "sd", "quantiles"), quantiles=c(0,.25,.5,.75,1))
```

	mean	sd	0%	25%	50%	75%	100%	n
A1	88021.32	22339.73	43795	73085.00	89743	101707.50	136788	22
A1.WDXRF	79783.05	20632.40	45409	66750.75	78698	90407.75	131358	22

6.9.2 Ordinary Linear Regression (OLR)

```
OLRA1<-lm(A1.WDXRF~A1, data=dataset)
summary(OLRA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1, data = dataset)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-10438.6	-4430.5	-426.7	4303.7	9952.1

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	2.021e+03	5.461e+03	0.37	0.715
A1	8.834e-01	6.022e-02	14.67	3.62e-12 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6164 on 20 degrees of freedom

Multiple R-squared: 0.915, Adjusted R-squared: 0.9107

F-statistic: 215.3 on 1 and 20 DF, p-value: 3.618e-12

```
fitted.OLRA1 <- fitted(OLRA1)
datasetb<-cbind(dataset,fitted.OLRA1)
RMS<-sqrt(mean((datasetb$A1-datasetb$fitted.OLRA1)^2)/nrow(datasetb))
RMS
```

```
[1] 1838.242
```

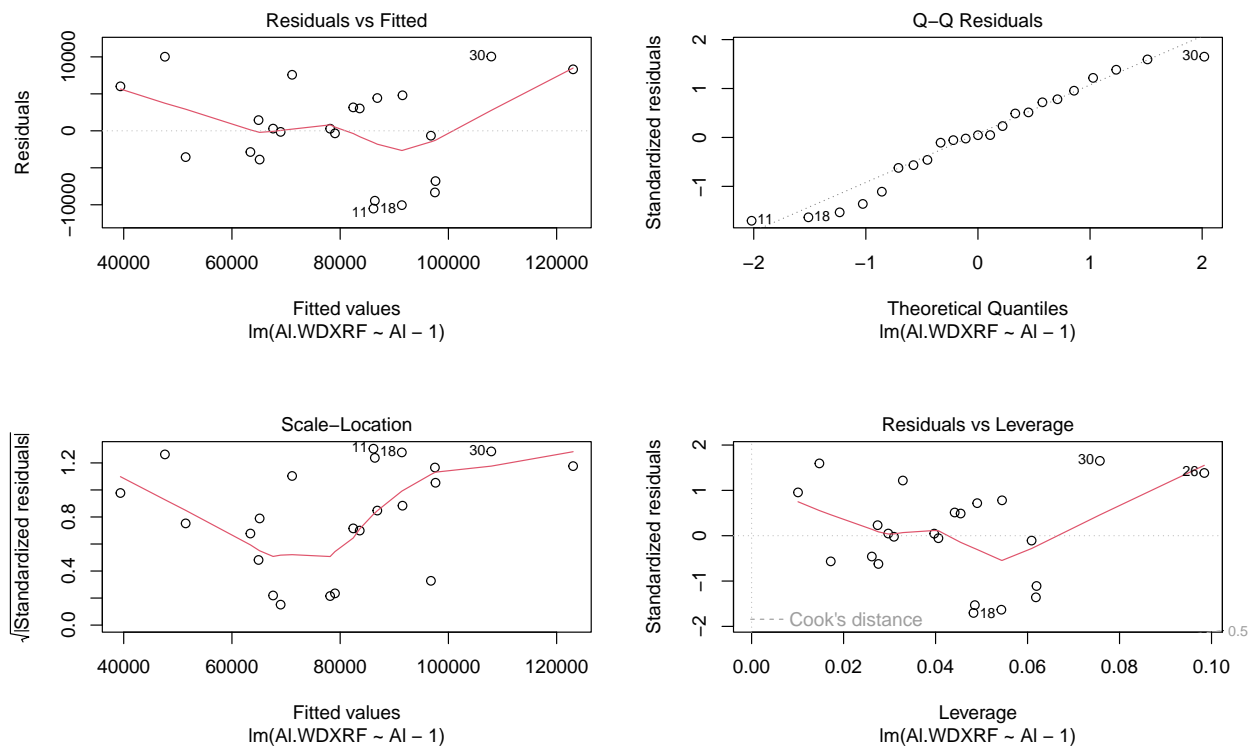
```
SEE<-sigma(OLRA1)
SEE
```

```
[1] 6164.434
```

```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```


[1] 7.003342

```
oldparOLRA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(OLRA1)
```



NULL

```
outlierTest(OLRA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest $|rstudent|$:

	$rstudent$	unadjusted p-value	Bonferroni p
18	-1.852753	0.079517	NA

```
Confint(OLRA1, level=0.90)
```

	Estimate	5 %	95 %
(Intercept)	2021.1820924	-7397.0375539	1.143940e+04
AI	0.8834435	0.7795894	9.872977e-01

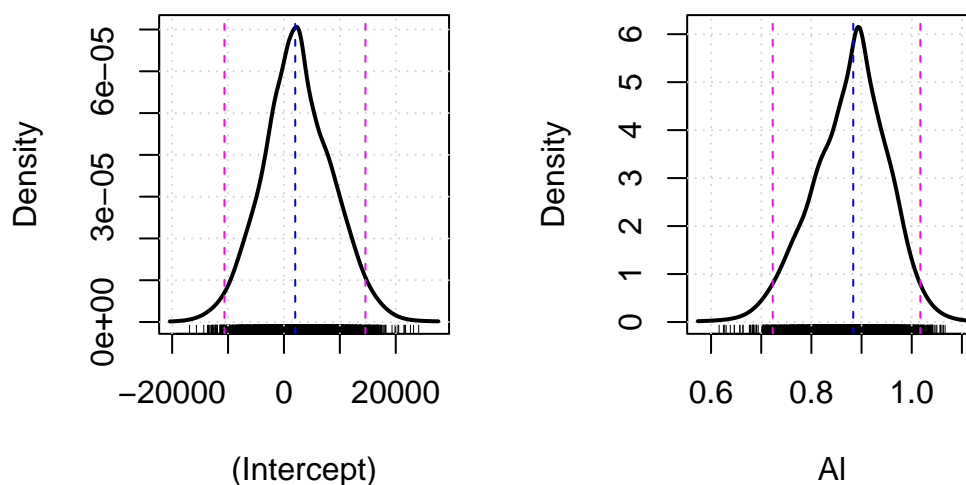
```
.bs.samplesOLRA1<- Boot(OLRA1, R=2500, method="case")
confint(.bs.samplesOLRA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
(Intercept)	-8522.0471526	1.241959e+04
Al	0.7466253	9.959083e-01

```
plotBoot(.bs.samplesOLRA1)
```

Bootstrap Distributions



6.9.3 Regression Trough Origin (RTO)

```
RTOA1<-lm(Al.WDXRF~Al-1, data=dataset)
summary(RTOA1)
```

Call:

```
lm(formula = Al.WDXRF ~ Al - 1, data = dataset)
```

Residuals:

Min	1Q	Median	3Q	Max
-10615.3	-3718.5	-158.1	4151.9	9725.6

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
Al	0.90508	0.01419	63.78	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6036 on 21 degrees of freedom

Multiple R-squared: 0.9949, Adjusted R-squared: 0.9946

F-statistic: 4067 on 1 and 21 DF, p-value: < 2.2e-16

```
fitted.RTOA1 <- fitted(RTOA1)
datasetc<-cbind(dataset,fitted.RTOA1)
RMS<-sqrt(mean((datasetc$A1-datasetc$fitted.RTOA1)^2)/nrow(datasetc))
RMS
```

```
[1] 1835.314
```

```
SEE<-sigma(RTOA1)
SEE
```

```
[1] 6036.44
```

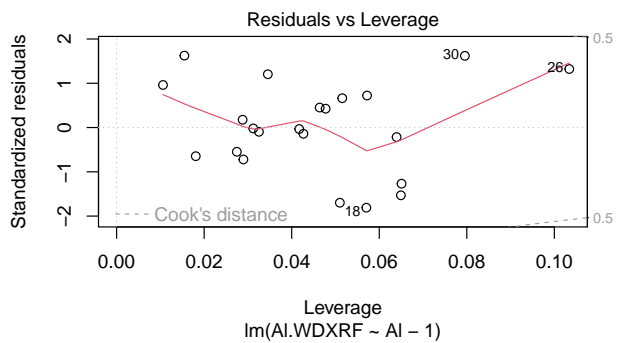
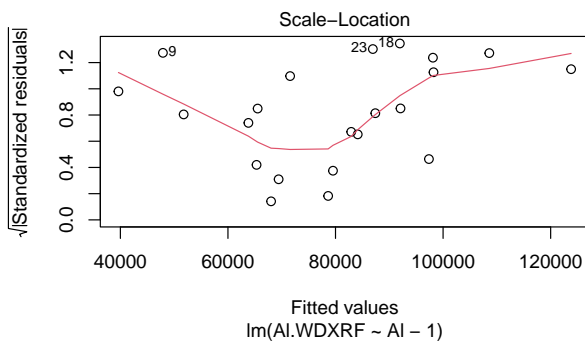
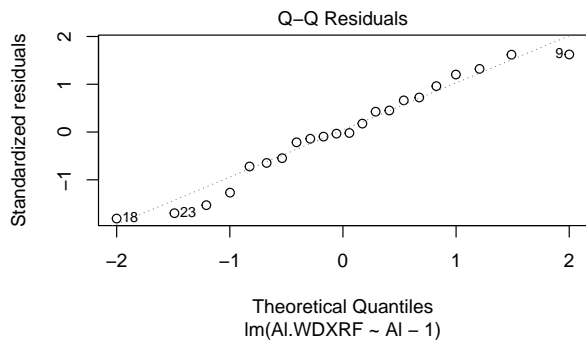
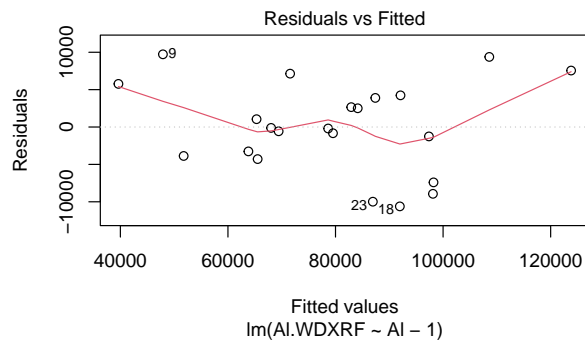
```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 6.85793
```

```
SSR<-(nrow(dataset)-1)*(SEE^2)
cSST <-sum(((dataset$A1.WDXRF)-mean(dataset$A1.WDXRF))^2)
Alcr2<-1-(SSR/cSST)
Alcr2
```

```
[1] 0.9144023
```

```
oldparRTOA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(RTOA1)
```



NULL

```
outlierTest(RTOA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest `|rstudent|`:

	<code>rstudent</code>	unadjusted p-value	Bonferroni p
18	-1.923924	0.068713	NA

```
Confint(RTOA1, level=0.90)
```

	Estimate	5 %	95 %
AI	0.9050759	0.8806561	0.9294956

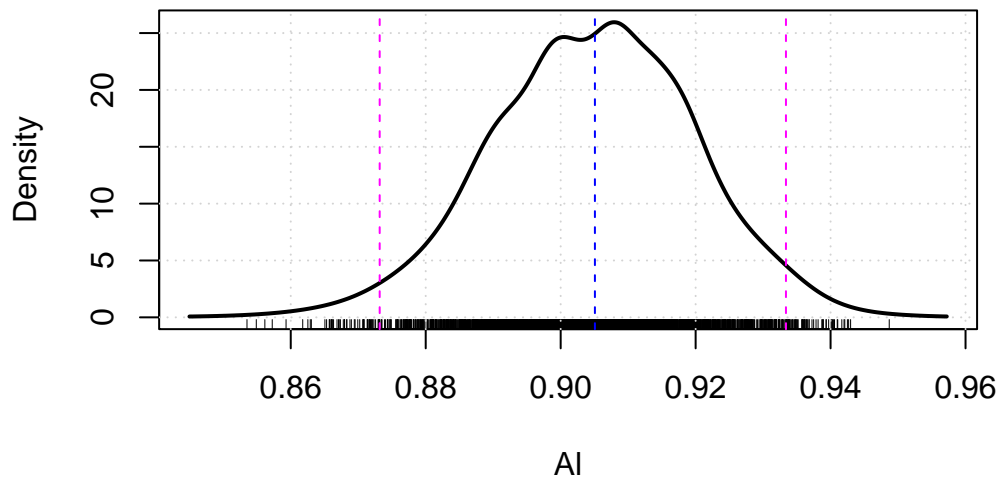
```
.bs.samplesRTOA1<- Boot(RTOA1, R=2500, method="case")
confint(.bs.samplesRTOA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
AI	0.8783731	0.9296875

```
plotBoot(.bs.samplesRTOA1)
```

Bootstrap Distributions



6.9.4 Result

In direct comparison, the values of the criteria of the RTO remain better. Also, they increased by excluding sample 11. As both linear regressions give the sample in row 18 as outlier, we are going to exclude this one, too.

6.10 Tenth iteration

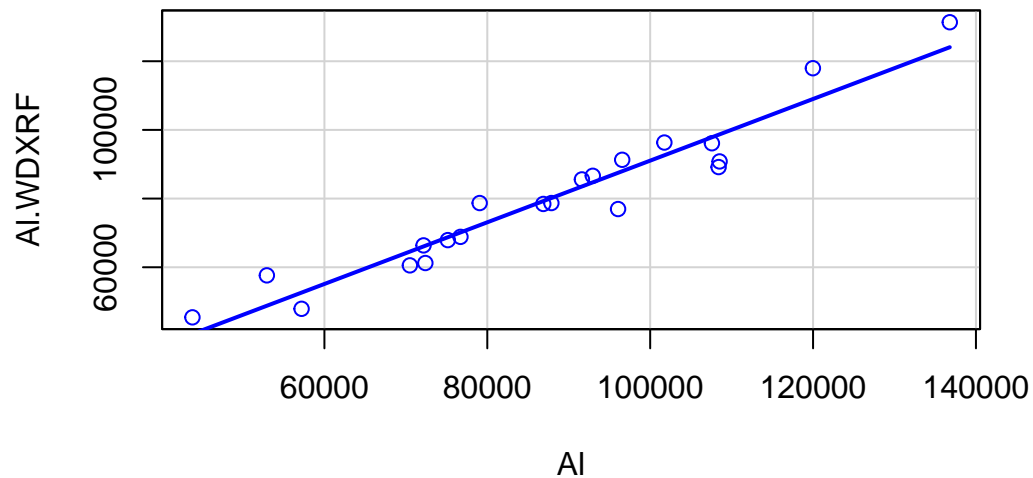
Now excluding row 18.

```
dataset1<- read.csv("../data_analytical//coefcorI_data.csv")
dataset<- dataset1[-c(25,14,27,28,24,1,8,11,18), ]
```

calculations

6.10.1 First impressions

```
scatterplot(AI.WDXRF~AI, regLine=TRUE, smooth=FALSE, boxplots=FALSE, data=dataset)
```



```
cor(dataset$AI,dataset$AI.WDXRF)
```

```
[1] 0.9633137
```

```
numSummary(dataset[,c("AI", "AI.WDXRF"), drop=FALSE],
  statistics=c("mean", "sd", "quantiles"),quantiles=c(0,.25,.5,.75,1))
```

	mean	sd	0%	25%	50%	75%	100%	n
AI	87374.52	22679.35	43795	72395	87870	101742	136788	21
AI.WDXRF	79708.71	21138.90	45409	66367	78698	90818	131358	21

6.10.2 Ordinary Linear Regression (OLR)

```
OLRA1<-lm(AI.WDXRF~AI, data=dataset)
summary(OLRA1)
```

Call:

```
lm(formula = AI.WDXRF ~ AI, data = dataset)
```

Residuals:

Min	1Q	Median	3Q	Max
-10554	-4009	-829	3713	8984

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	1.257e+03	5.173e+03	0.243	0.811
AI	8.979e-01	5.739e-02	15.646	2.62e-12 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 5821 on 19 degrees of freedom

Multiple R-squared: 0.928, Adjusted R-squared: 0.9242

F-statistic: 244.8 on 1 and 19 DF, p-value: 2.616e-12

```
fitted.OLRA1 <- fitted(OLRA1)
datasetb<-cbind(dataset,fitted.OLRA1)
RMS<-sqrt(mean((datasetb$A1-datasetb$fitted.OLRA1)^2)/nrow(datasetb))
RMS
```

[1] 1744.009

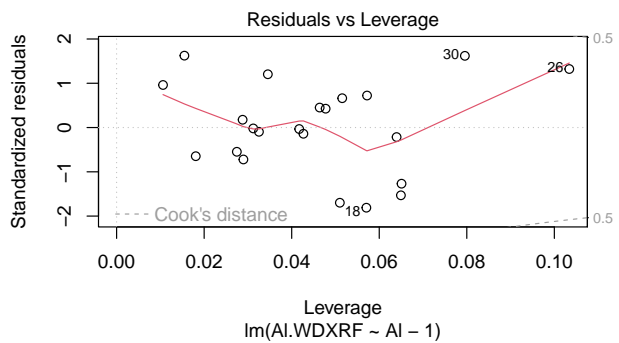
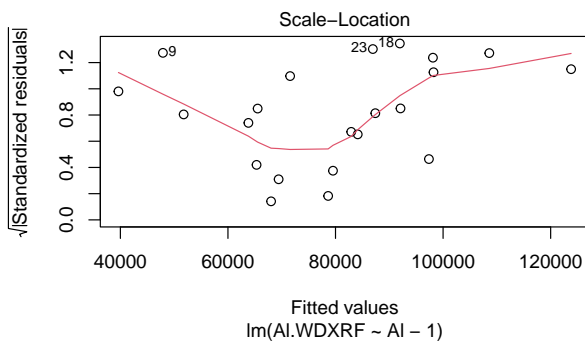
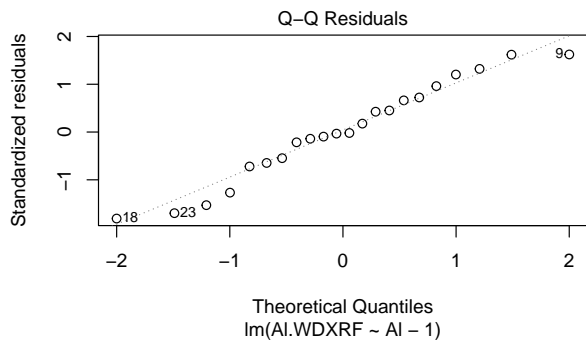
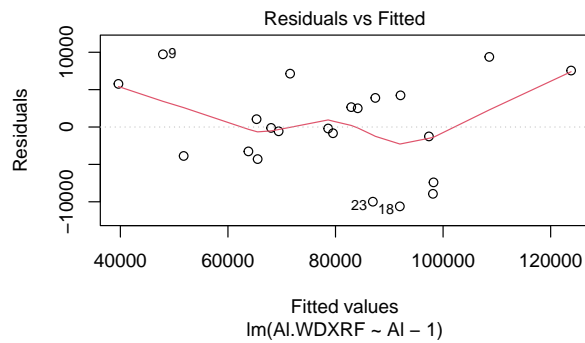
```
SEE<-sigma(OLRA1)
SEE
```

[1] 5820.597

```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

[1] 6.661664

```
oldparOLRA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(OLRA1)
```



NULL

```
outlierTest(OLRA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest `|rstudent|`:

	<code>rstudent</code>	unadjusted	p-value	Bonferroni	p
23	-2.00849	0.059836	NA		

```
Confint(OLRA1, level=0.90)
```

	Estimate	5 %	95 %
(Intercept)	1256.6511447	-7687.5108511	1.020081e+04
AI	0.8978826	0.7986509	9.971142e-01

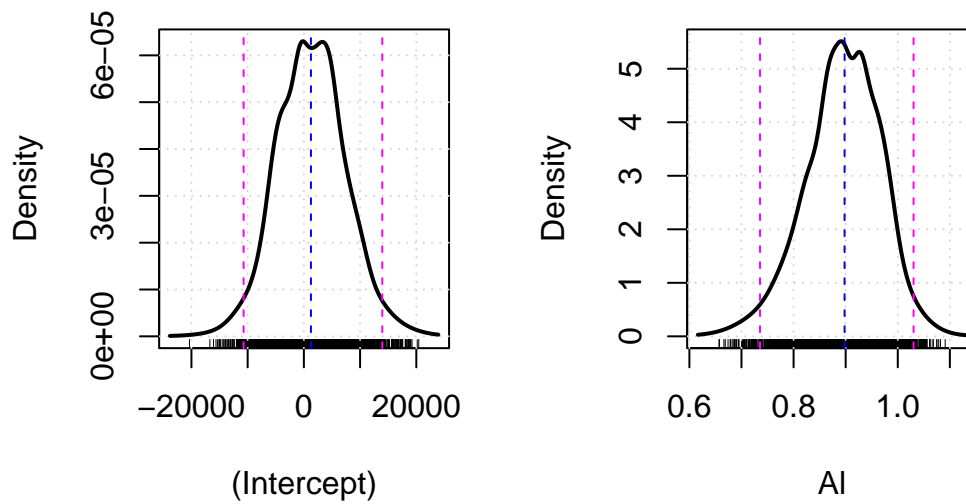
```
.bs.samplesOLRA1<- Boot(OLRA1, R=2500, method="case")
confint(.bs.samplesOLRA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
(Intercept)	-8682.3453843	11549.284151
AI	0.7672977	1.007969

```
plotBoot(.bs.samplesOLRA1)
```


Bootstrap Distributions



6.10.3 Regression Trough Origin (RTO)

```
RTOA1<-lm(A1.WDXRF~A1-1, data=dataset)
summary(RTOA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1 - 1, data = dataset)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-10595.0	-3704.9	-594.1	3594.6	9391.0

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
A1	0.91140	0.01376	66.25	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 5682 on 20 degrees of freedom

Multiple R-squared: 0.9955, Adjusted R-squared: 0.9952

F-statistic: 4389 on 1 and 20 DF, p-value: < 2.2e-16

```
fitted.RTOA1 <- fitted(RTOA1)
datasetc<-cbind(dataset,fitted.RTOA1)
RMS<-sqrt(mean((datasetc$A1-datasetc$fitted.RTOA1)^2)/nrow(datasetc))
RMS
```

```
[1] 1742.708
```

```
SEE<-sigma(RTOA1)
SEE
```

```
[1] 5682.021
```

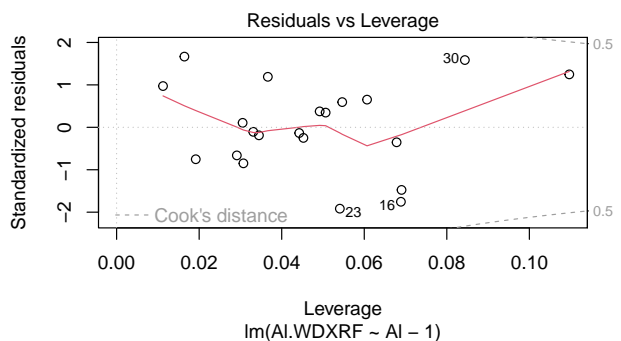
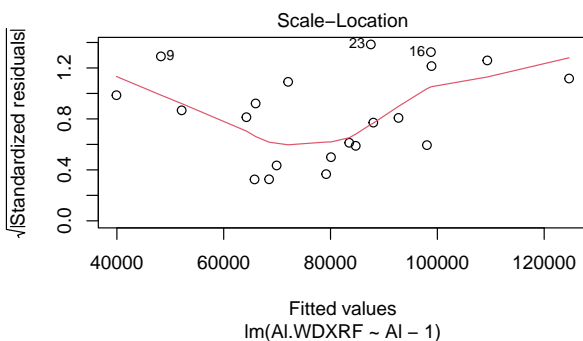
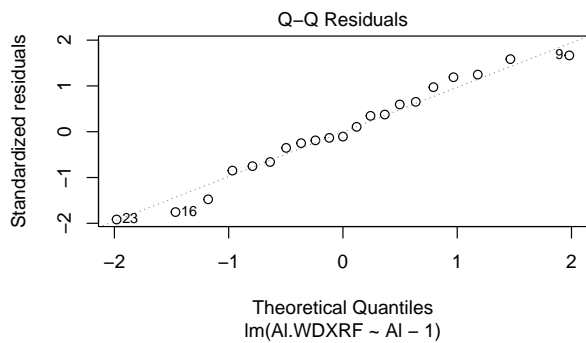
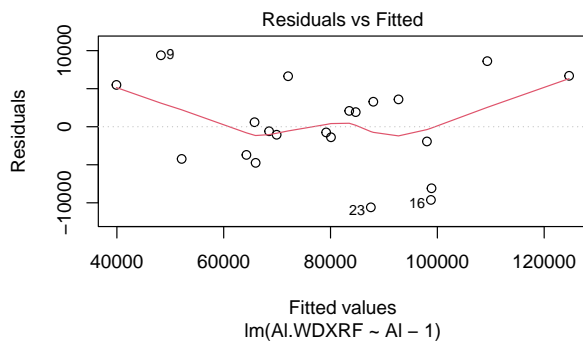
```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 6.503063
```

```
SSR<-(nrow(dataset)-1)*(SEE^2)
cSST <-sum(((dataset$A1.WDXRF)-mean(dataset$A1.WDXRF))^2)
Alcr2<-1-(SSR/cSST)
Alcr2
```

```
[1] 0.9277495
```

```
oldparRTOA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(RTOA1)
```



NULL

```
outlierTest(RTOA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest |rstudent|:

	rstudent	unadjusted p-value	Bonferroni p
23	-2.068397	0.05249	NA

```
Confint(RTOA1, level=0.90)
```

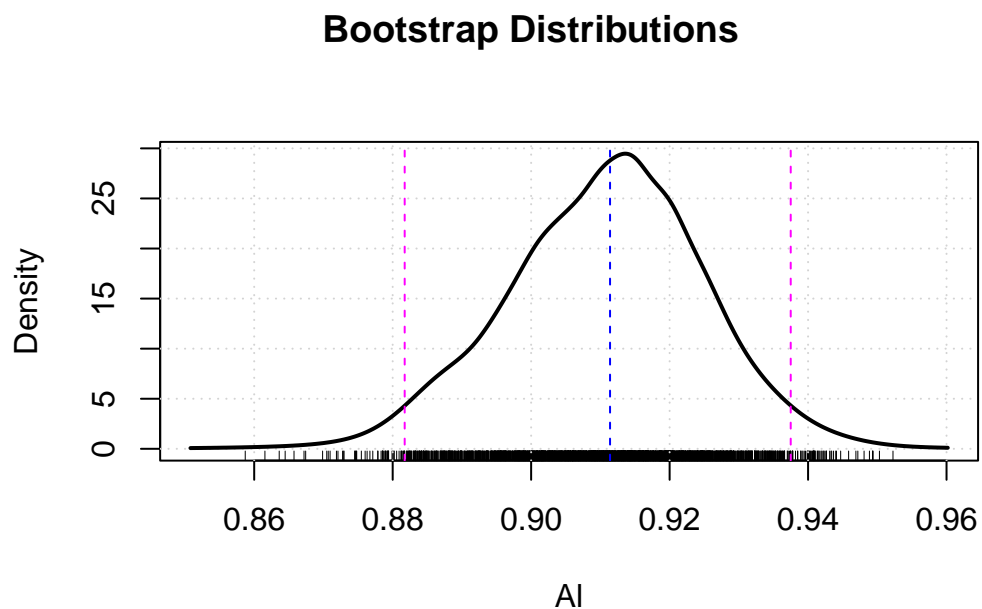
	Estimate	5 %	95 %
A1	0.9113977	0.8876719	0.9351236

```
.bs.samplesRTOA1<- Boot(RTOA1, R=2500, method="case")
confint(.bs.samplesRTOA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
A1	0.8854027	0.9335376

```
plotBoot(.bs.samplesRTOA1)
```



6.10.4 Result

By excluding row 18, something interesting is happening:

criteria	OLR	RTO
r^2	0.915 0.928	0.9144 0.9277

criteria	OLR	RTO
rSEE	7.00 6.66	6.86 6.50
CI 0.05	0.7796 - 0.7521 0.7597 - 0.7987	0.8807 - 0.8723 0.8877 - 0.8855
CI 0.95	0.9873 - 0.9982 1.0037 - 0.9971	0.9244 - 0.9273 0.9351 - 0.9348

While the criteria of the RTO are getting better, the agreement of CILR5 and CIBS5 of the OLR are getting worse. Also, the range of values is increasing. This means, the robustness less good for this OLR. As RTO is still better, this does not affect the next step but it is worth noticing. As both linear regressions are giving the sample in line 23 as outlier, we are excluding it.

6.11 Eleventh iteration

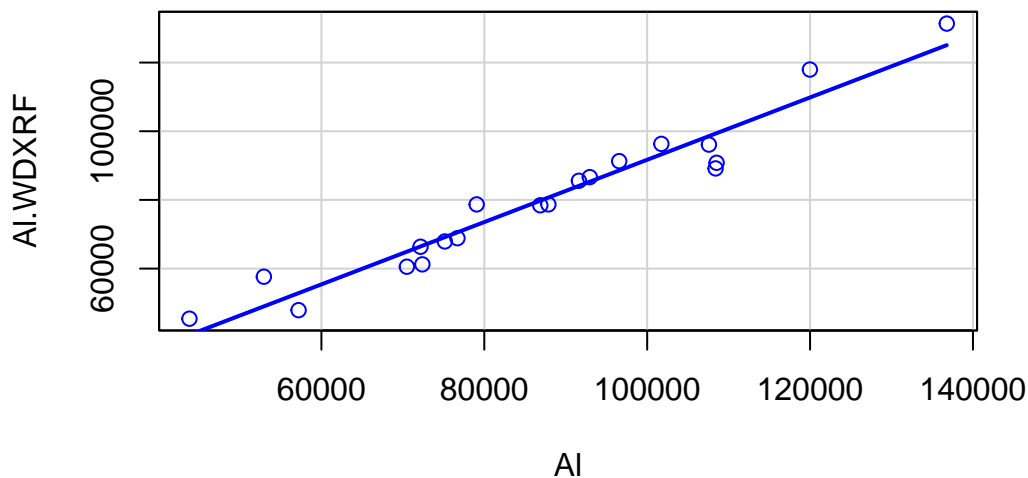
Now excluding row 23.

```
dataset1<- read.csv("../data_analytical//coefcorI_data.csv")
dataset<- dataset1[-c(25,14,27,28,24,1,8,11,18,23), ]
```

calculations

6.11.1 First impressions

```
scatterplot(AI.WDXRF~AI, regLine=TRUE, smooth=FALSE, boxplots=FALSE, data=dataset)
```



```
cor(dataset$AI,dataset$AI.WDXRF)
```

```
[1] 0.9701069
```

```
numSummary(dataset[,c("Al", "Al.WDXRF"), drop=FALSE],
            statistics=c("mean", "sd", "quantiles"), quantiles=c(0,.25,.5,.75,1))
```

	mean	sd	0%	25%	50%	75%	100%	n
Al	86940.35	23178.81	43795	72338.75	87373.5	103201.2	136788	20
Al.WDXRF	79846.55	21678.37	45409	65083.50	78698.0	90937.0	131358	20

6.11.2 Ordinary Linear Regression (OLR)

```
OLRA1<-lm(Al.WDXRF~Al, data=dataset)
summary(OLRA1)
```

Call:

```
lm(formula = Al.WDXRF ~ Al, data = dataset)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-10142.8	-2942.0	-664.9	3461.5	8642.6

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	964.7972	4805.5128	0.201	0.843
Al	0.9073	0.0535	16.960	1.62e-12 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 5405 on 18 degrees of freedom

Multiple R-squared: 0.9411, Adjusted R-squared: 0.9378

F-statistic: 287.6 on 1 and 18 DF, p-value: 1.624e-12

```
fitted.OLRA1 <- fitted(OLRA1)
datasetb<-cbind(dataset,fitted.OLRA1)
RMS<-sqrt(mean((datasetb$Al-datasetb$fitted.OLRA1)^2)/nrow(datasetb))
RMS
```

```
[1] 1653.891
```

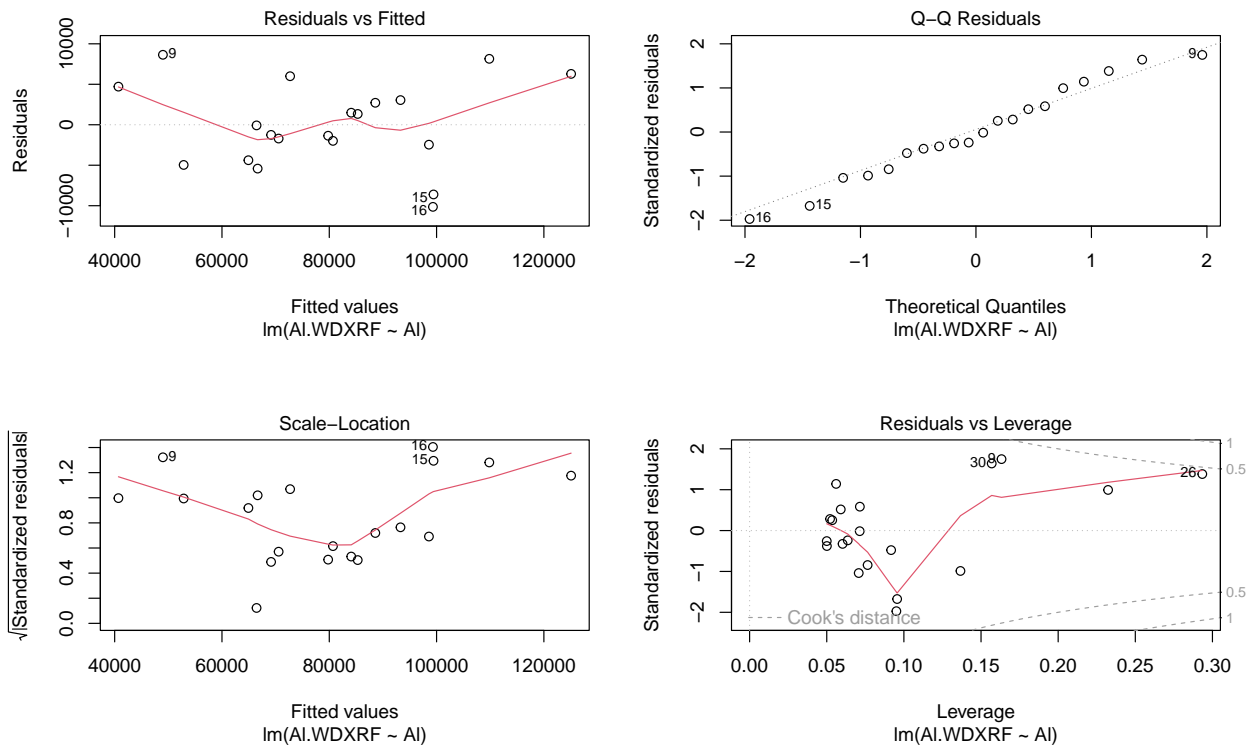
```
SEE<-sigma(OLRA1)
SEE
```

```
[1] 5405.02
```

```
mean<-mean(dataset$Al)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 6.216929
```

```
oldparOLRA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(OLRA1)
```



NULL

```
outlierTest(OLRA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest `|rstudent|`:

	rstudent	unadjusted p-value	Bonferroni p
16	-2.165474	0.044863	0.89726

```
Confint(OLRA1, level=0.90)
```

	Estimate	5 %	95 %
(Intercept)	964.7971979	-7368.2676888	9297.862085
Al	0.9073089	0.8145417	1.000076

```
.bs.samplesOLRA1<- Boot(OLRA1, R=2500, method="case")
confint(.bs.samplesOLRA1, level=0.9, type="bca")
```

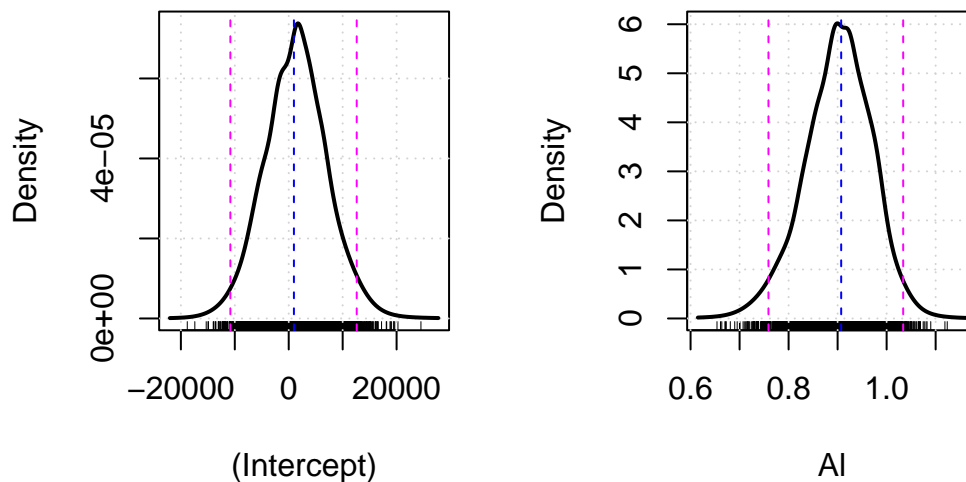
Bootstrap bca confidence intervals

	5 %	95 %

```
(Intercept) -9079.6990773 10404.183608
Al          0.7874974      1.017895
```

```
plotBoot(.bs.samplesOLRA1)
```

Bootstrap Distributions



6.11.3 Regression Trough Origin (RTO)

```
RTOA1<-lm(A1.WDXRF~A1-1, data=dataset)
summary(RTOA1)
```

Call:

```
lm(formula = A1.WDXRF ~ A1 - 1, data = dataset)
```

Residuals:

Min	1Q	Median	3Q	Max
-10304.9	-2999.1	-465.9	3519.2	9057.2

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
A1	0.91770	0.01311	70	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 5267 on 19 degrees of freedom

Multiple R-squared: 0.9961, Adjusted R-squared: 0.9959

F-statistic: 4900 on 1 and 19 DF, p-value: < 2.2e-16

```
fitted.RTOA1 <- fitted(RTOA1)
datasetc<-cbind(dataset,fitted.RTOA1)
RMS<-sqrt(mean((datasetc$A1-datasetc$fitted.RTOA1)^2)/nrow(datasetc))
RMS
```

```
[1] 1653.001
```

```
SEE<-sigma(RTOA1)
SEE
```

```
[1] 5266.747
```

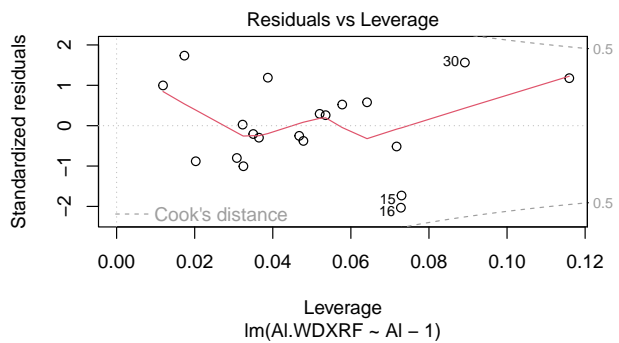
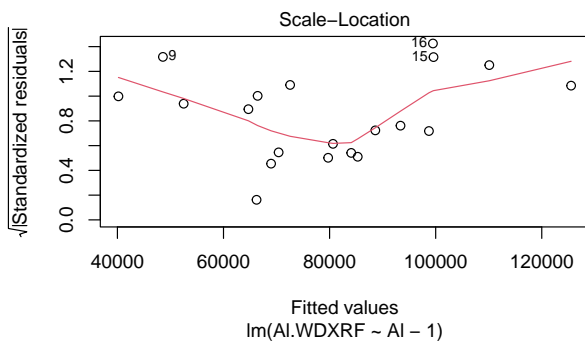
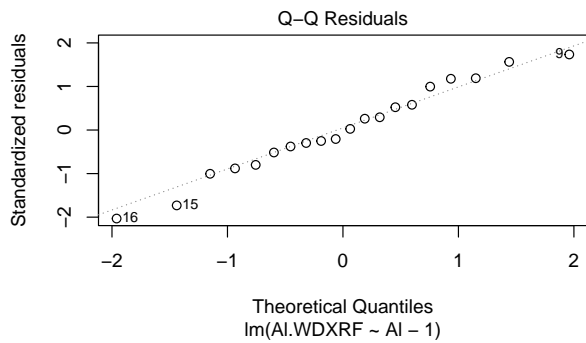
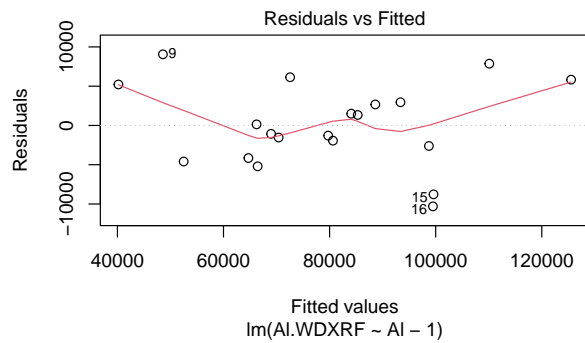
```
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
rSEE
```

```
[1] 6.057886
```

```
SSR<-(nrow(dataset)-1)*(SEE^2)
cSST <-sum(((dataset$A1.WDXRF)-mean(dataset$A1.WDXRF))^2)
Alcr2<-1-(SSR/cSST)
Alcr2
```

```
[1] 0.9409756
```

```
oldparRTOA1 <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(RTOA1)
```

NULL

```
outlierTest(RTOA1)
```

No Studentized residuals with Bonferroni $p < 0.05$

Largest `|rstudent|`:

	<code>rstudent</code>	unadjusted p-value	Bonferroni p
16	-2.235542	0.03829	0.7658

```
Confint(RTOA1, level=0.90)
```

	Estimate	5 %	95 %
AI	0.9177042	0.8950345	0.9403739

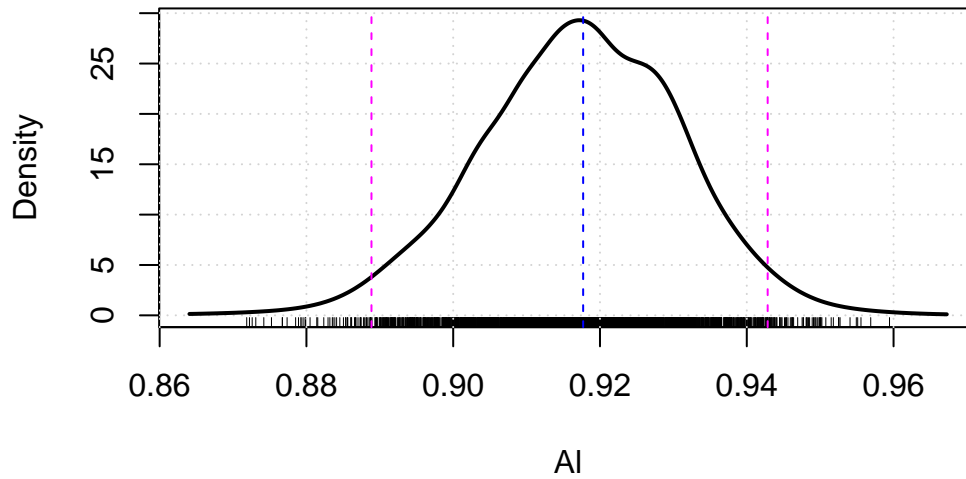
```
.bs.samplesRTOA1<- Boot(RTOA1, R=2500, method="case")
confint(.bs.samplesRTOA1, level=0.9, type="bca")
```

Bootstrap bca confidence intervals

	5 %	95 %
AI	0.8932434	0.9392519

```
plotBoot(.bs.samplesRTOA1)
```

Bootstrap Distributions



6.11.4 Result - RTO

With this step, also the robustness of the RTO starts to decrease:

criteria	OLR	RTO
r^2	0.928 0.9411	0.9277 0.9410
rSEE	6.66 6.22	6.50 6.06
CI 0.05	0.7597 - 0.7987 0.8145 - 0.7791	0.8877 - 0.8855 0.8950 - 0.8932
CI 0.95	1.0037 - 0.9971 1.0000 - 1.0113	0.9351 - 0.9348 0.9404 - 0.9390

Therefore, excluding row 23 is not purposeful and the RTO of iteration ten is determined as the final linear regression to extract the coefcords from.

7 Example of extracting criteria and factors

source

Instructions following <https://stackoverflow.com/questions/66771929/intercept-and-slope-functions-in-r>, <https://www.digialocean.com/community/tutorials/rbind-function-r>, <https://www.statology.org/number-of-rows-in-r/>, <https://statisticsglobe.com/r-max-min-function/>, <https://www.statology.org/extract-r-squared-from-lm-in-r/>, <https://sparkbyexamples.com/r-programming/select-columns-by-index-position-in-r-2/>, <https://search.r-project.org/R/refmans/base/html/Round.html>, <https://www.statology.org/transpose-data-frame-in-r/> and <https://datatofish.com/export-dataframe-to-csv-in-r/>

In order to compile all information and factors of aluminium for coefcor I in a table, the relevant criteria must first be extracted from the previous calculations. The procedure differs somewhat for OLR and RTO. For this reason, both are described below by way of example, even though only the factors of the RTO are decisive for aluminium. To get the right coefcor, we first have to re-run again some parts the calculation of iteration ten. In the normal workflow, this step is omitted because the last iteration calculated is also the one from which the coefcors are extracted.

7.1 Extracting from OLR

First, the important factors slope (a) and intercept (b) are extracted.

```
cf <- Confint(OLRA1, level=0.90)
a<-cf[2]
a<-round(a,4)
b<-cf[1]
b<-round(b,0)
```

Next, r^2 , rSEE, SEE and RMS are selected. Since the last three terms are used for OLR and RTO, they must be specified here.

```
r2<-summary(OLRA1)$r.squared
r2<-round(r2,4)

SEE<-sigma(OLRA1)
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100
SEE<-round(SEE,0)
rSEE<-round(rSEE,2)

fitted.OLRA1 <- fitted(OLRA1)
datasetb<-cbind(dataset,fitted.OLRA1)
RMS<-sqrt(mean((datasetb$A1-datasetb$fitted.OLRA1)^2)/nrow(datasetb))
RMS<-round(RMS,0)
```

Then the range of values for which the coefcor applies is defined. For this, the lowest (start) and highest measured value (end) are extracted.

```
start<-min(dataset$A1)
start<-round(start,0)
end<-max(dataset$A1)
end<-round(end,0)
```

Additionally, the values of the confidence levels (0.05 and 0.95) of the slope calculated from the filtered, whole data set (CILR) and by the bootstrap (CIBS) are to be provided.

```
CILR5<-cf[4]
CILR5<-round(CILR5,4)
CILR95<-cf[6]
```

```

CILR95<-round(CILR95,4)

cfb<-confint(.bs.samplesOLRA1, level=0.9, type="bca")
CIBS5<-cfb[2]
CIBS5<-round(CIBS5,4)
CIBS95<-cfb[4]
CIBS95<-round(CIBS95,4)

```

Last but not least, we extract percentage (outpct) and specification of the rows of the outliers (out). The number of samples in the data set (here 30) can be taken manually from the entire data set (first iteration - first impressions table).

```

outpct<-(100/30)*(30-nrow(dataset))
outpct<-round(outpct,2)
out<- '25-14-27-28-24-1-8-11-18'

```

Now all variables with their respective values are merged into one table, the column header is labelled with the element name and this column is explicitly selected and defined as its own short table. In this way, all elements can be summarised together in an overview table termed **__criteria.csv** (e.g. **coefcorI__criteria.csv**) at the end of the analysis when all coefcors for all elements are defined (see below).

```

A1OLR_tab<-rbind(a,b,r2,rSEE,SEE,RMS,start,end,CILR5,CIBS5,CILR95,CIBS95,out,outpct)
colnames(A1OLR_tab)[1] ="A1OLR"
A1OLR_criteria<-t(A1OLR_tab)
A1OLR_criteria

```

	a	b	r2	rSEE	SEE	RMS	start	end	CILR5
A1OLR	"0.8979"	"1257"	"0.928"	"6.66"	"5821"	"1744"	"43795"	"136788"	"0.7987"
	CIBS5	CILR95	CIBS95	out					outpct
A1OLR	"0.7657"	"0.9971"	"1.0093"	"25-14-27-28-24-1-8-11-18"					"30"

7.2 Extracting from RTO

The compilation of the important criteria of the RTO works according to the same principle - only that now the RTO is selected to serve as input. And r^2 and b have to be extracted in a different way. In the case of the RTO, b is "0".

```

cf <- Confint(RTOA1, level=0.90)
a<-cf[1]
a<-round(a,4)
b='0'

r2<-Alcr2
r2<-round(r2,4)

SEE<-sigma(RTOA1)
mean<-mean(dataset$A1)
rSEE<-(SEE/mean)*100

```


7.3 Exporting coefcor criteria (__criteria.csv)

source

Instructions following <https://datatofish.com/export-dataframe-to-csv-in-r/>

And finally, we can combine and export the criteria of the coefcors for AI and AIOLR in a joint table (__criteria.csv) which for this example is called **coefcorI_example_criteria.csv**.

```
criteria<- rbind(AI_criteria,AIOLR_criteria)
criteria
```

	a	b	r2	rSEE	SEE	RMS	start	end	CILR5
AI	"0.9114"	"0"	"0.9277"	"6.5"	"5682"	"1743"	"43795"	"136788"	"0.8877"
AIOLR	"0.8979"	"1257"	"0.928"	"6.66"	"5821"	"1744"	"43795"	"136788"	"0.7987"
	CIBS5	CILR95	CIBS95	out				outpct	
AI	"0.8856"	"0.9351"	"0.9346"	"25-14-27-28-24-1-8-11-18"	"30"				
AIOLR	"0.7657"	"0.9971"	"1.0093"	"25-14-27-28-24-1-8-11-18"	"30"				

```
write.csv(criteria,"../data_processed//coefcorI_example_criteria.csv",row.names=TRUE)
```

And - tada! - we are done. Well, almost...

7.4 Exporting coefcor factors - slope and intercept (__factors.csv)

source

Instructions following <https://campus.datacamp.com/courses/model-a-quantitative-trading-strategy-in-r/chapter-1-introduction-to-r-for-trading?ex=4> and <https://rdr.io/r/base/cbind.html>

In order to apply the coefcor factors of slope (a) and intercept (b) to a data set using the R-script developed for the Munich Procedure, they have to be compiled in a certain way (__factors.csv). For this purpose we need to load the __criteria.csv-file, in our case **coefcorI_example_criteria.csv**.

```
dataset2<- read.csv("../data_processed//coefcorI_example_criteria.csv")
```

Then we can extract the necessary factors for slope and intercept from AI and AI_OLR.

```
AI_a<-dataset2[1,2]
AI_b<-dataset2[1,3]
AIOLR_a<-dataset2[2,2]
AIOLR_b<-dataset2[2,3]
```

We bind them together and export a table containing only the relevant coefcor factors in a spreadsheet called __factors.csv - in our example **coefcorI_example_factors.csv**. An example of factors of a complete coefcor would be **coefcorI_factors.csv**

```
s_i<- cbind(A1_a,A1_b,A1OLR_a,A1OLR_b)
s_i
```

```
      A1_a A1_b A1OLR_a A1OLR_b
[1,] 0.9114    0  0.8979   1257
```

```
write.csv(s_i,"../data_processed//coefcorI_example_factors.csv",row.names=TRUE)
```

8 Example of graphics export

source

Instructions following <http://www.sthda.com/english/wiki/ggplot2-scatter-plots-quick-start-guide-r-software-and-data-visualization#basic-scatter-plots>, <http://www.sthda.com/english/wiki/ggplot2-title-main-axis-and-legend-titles>, https://ggplot2.tidyverse.org/reference/geom_smooth.html, <https://github.com/kassambara/ggpubr/issues/312>, <https://quarto.org/docs/computations/execution-options.html>, <https://stat.ethz.ch/R-manual/R-devel/library/graphics/html/hist.html>, <https://cran.r-project.org/web/packages/ggplotify/vignettes/ggplotify.html>, https://www.rdocumentation.org/packages/cowplot/versions/1.1.1/topics/plot_grid, <https://www.statology.org/plot-title-in-r/> and <https://ggplot2.tidyverse.org/reference/ggsave.html>

8.1 Exporting all relevant descriptive graphics for OLR and RTO

If you wish, you can compile and export all relevant graphics of the final run. For this purpose, the graphics must first be defined as variables and then combined in a joint export of OLR (`__OLR__allfig.eps`) and RTO (`__RTO__allfig.eps`) of the given element - in our case A1 of coefcor I (`coefcorI_example_A1_OLR__allfig.eps/coefcorI_example_A1_RTO__allfig.eps`). The figure will include:

- scatter plots of p-XRF values (A1) to WDXRF values (A1.WDXRF), showing respectively the linear regression of OLR (A1_scatter_OLR) or RTO (A1_scatter_RTO). The algorithm used here is rounding the intercept values to the nearest 100, R^2 to two decimal digits. The otherwise usual formula $y=ax+b$ is given in the form $y=b+ax$.

Attention - we have to take r^2 of the RTO explicitly from the previously performed calculations because it is calculated incorrectly by the algorithm section 6.1.3.4.

```
A1_scatter_OLR<-ggplot(dataset, aes(x=A1, y=A1.WDXRF))+geom_point()+
  geom_smooth(method=lm,se=FALSE,formula = y ~ x)+ ggtitle("OLR")+
  theme(plot.title = element_text(color="black", size=9,face="bold"))+
  stat_poly_eq(formula = y ~ x,aes(label = paste(..eq.label.., ..rr.label..,
    sep = "~~~")),parse = TRUE, coef.digits = 4, f.digits = 4,rr.digits=4)

Alcr2<-round(Alcr2,4)
```

```
Al_scatter_RTO<-ggplot(dataset, aes(x=Al, y=Al.WDXRF))+geom_point()+
  geom_smooth(method=lm,se=FALSE,formula = y ~ x-1)+ ggtitle("RTO")+
  theme(plot.title = element_text(color="black", size=9,face="bold"))+
  stat_poly_eq(formula = y ~ x - 1,aes(label = paste(..eq.label..,
    sep = "~~", "R2==", Alcr2)),parse = TRUE, coef.digits = 4)
```

- scatter plots of WDXRF values (Al.WDXRF) to fitted values. The latter term refers to the recalculated p-XRF values based on the specified values for slope (RTO - scatter_Al_RTO_WDXRF_fitted) and intercept (ORL - scatter_Al_ORL_WDXRF_fitted). The values shown in the graph for the linear regression model chosen above (in our case the RTO - see section 7.2) should for slope and r^2 be as near to 1 as possible, intercept close to 0. The better these criteria are met, the higher the quality of the recalculation of p-XRF measurements made possible by coefcors.

Here we can use the r^2 (Alcr2) calculated above in section 7.2 and used for Al_scatter_RTO as only the scaling and origin of the data have been changed. However, the relationship between the data points and therefore r^2 remains the same as the linear function is still the best fit to the data.

```
Al_scatter_OLR_WDXRF_fitted<-ggplot(datasetb, aes(x=fitted.OLRA1, y=Al.WDXRF))+
  geom_point()+geom_smooth(method=lm,se=FALSE,formula = y ~ x)+ ggtitle("OLR")+
  theme(plot.title = element_text(color="black", size=9,face="bold"))+
  stat_poly_eq(formula = y ~ x,aes(label = paste(..eq.label.., ..rr.label..,
    sep = "~~~"))),parse = TRUE, coef.digits = 4, f.digits = 4,rr.digits=4)
```

```
Al_scatter_RTO_WDXRF_fitted<-ggplot(datasetc, aes(x=fitted.RTOAl, y=Al.WDXRF))+
  geom_point()+geom_smooth(method=lm,se=FALSE,formula = y ~ x-1)+ ggtitle("RTO")+
  theme(plot.title = element_text(color="black", size=9,face="bold"))+
  stat_poly_eq(formula = y ~ x - 1,aes(label = paste(..eq.label..,
    sep = "~~", "R2==", Alcr2)),parse = TRUE, coef.digits = 4)
```

- the graphs of the diagnostics of the linear regressions (now compiled as OLRA1_1_2/3_5 and RTOAl_1_2/3_5) we already know from our steps in section 6.1.2.4 and section 6.1.3.5.

```
par(mar = c(0.1, 0.1, 0.1,0.1))

OLRA11<-as_grob(~plot(OLRA1,1))
OLRA12<-as_grob(~plot(OLRA1,2))
OLRA13<-as_grob(~plot(OLRA1,3))
OLRA15<-as_grob(~plot(OLRA1,5))

OLRA1_1_2<-plot_grid(OLRA11,OLRA12)
OLRA1_3_5<-plot_grid(OLRA13,OLRA15)

RTOA11<-as_grob(~plot(RTOAl,1))
RTOA12<-as_grob(~plot(RTOAl,2))
RTOA13<-as_grob(~plot(RTOAl,3))
RTOA15<-as_grob(~plot(RTOAl,5))
```



```
RTOA1_1_2<-plot_grid(RTOA11,RTOA12)
RTOA1_3_5<-plot_grid(RTOA13,RTOA15)
```

- histograms of the distribution of the residuals for OLR (Al_hist_OLR) and RTO (Al_hist_RTO). To do this, the values of the residuals must first be extracted from the linear regression calculations. This is done in the same way as in section 6.1.2.2 and section 6.1.3.2.

```
residuals.OLRA1 <- residuals(OLRA1)
datasetb<-cbind(datasetb,residuals.OLRA1)

Al_hist_OLR<-as_grob(~Hist(datasetb$residuals.OLRA1, scale="frequency",
                           breaks="Sturges", col="darkgray",xlab="OLR residuals"))

residuals.RTOA1 <- residuals(RTOA1)
datasetc<-cbind(datasetc,residuals.RTOA1)

Al_hist_RTO<-as_grob(~Hist(datasetc$residuals.RTOA1, scale="frequency",
                           breaks="Sturges", col="darkgray",xlab="RTO residuals"))
```

- finally, we assign the density distributions of the bootstrap known from steps section 6.1.2.6 and section 6.1.3.7 for each linear regression (Al_boot_OLR/Al_boot_RTO) to separate variables.

```
Al_boot_OLR<-as_grob(~plotBoot(.bs.samplesOLRA1)+title('OLR'))
Al_boot_RTO<-as_grob(~plotBoot(.bs.samplesRTOA1)+title('RTO'))
```

Now that all the graphics are defined we can put them together as **coefcorI_example_Al_OLR_allfig.eps**.

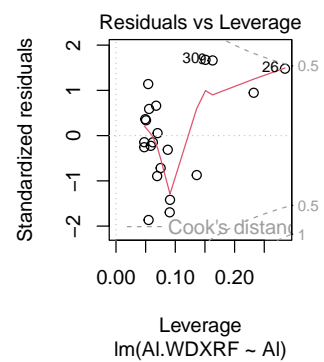
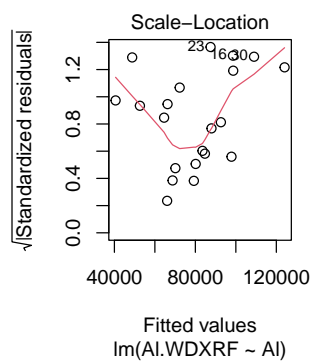
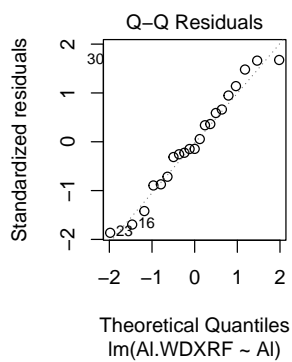
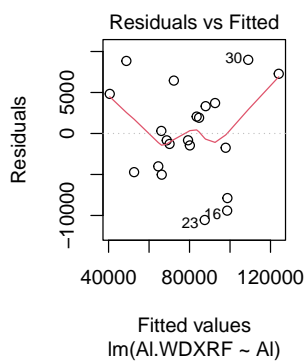
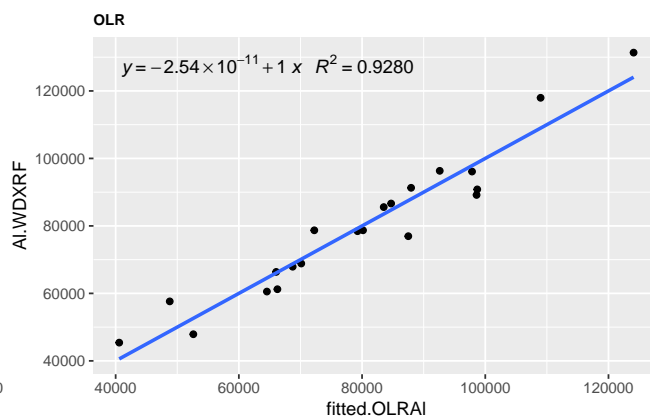
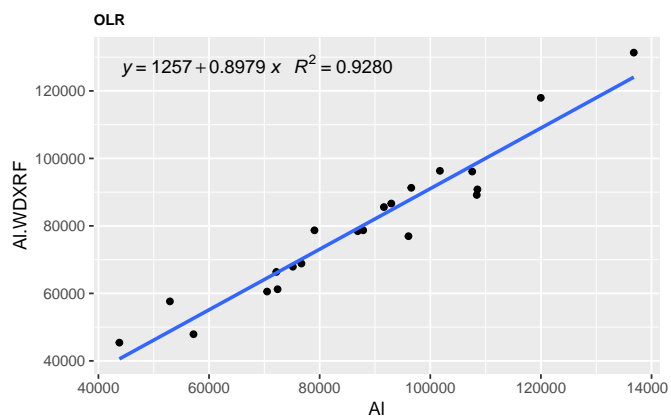
```
coefcorI_example_Al_OLR_allfig<-plot_grid(Al_scatter_OLR,Al_scatter_OLR_WDXRF_fitted,
                                           OLRA1_1_2,OLRA1_3_5,Al_boot_OLR,Al_hist_OLR,
                                           ncol=2,nrow=3,rel_heights=c(1,1.5,1))

ggsave("coefcorI_example_Al_OLR_allfig.eps",path="..//graphics"),plot=last_plot(),
        device="eps",height=25,width=30,unit=c("cm"))
```

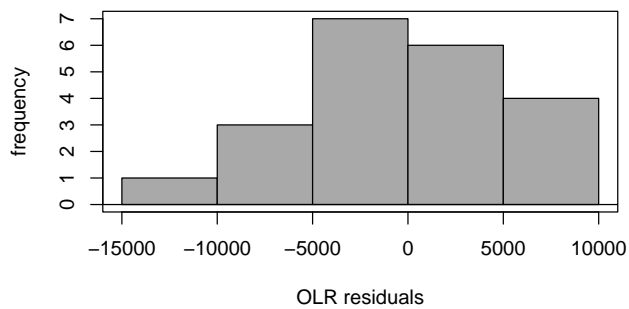
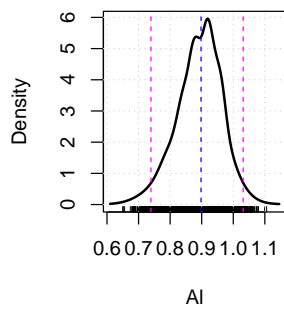
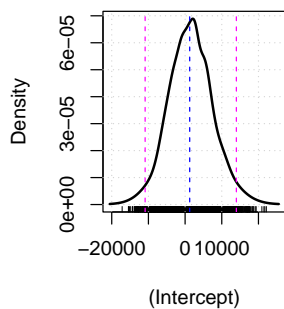
Same procedure for the RTO graphics named **coefcorI_example_Al_RTO_allfig.eps**.

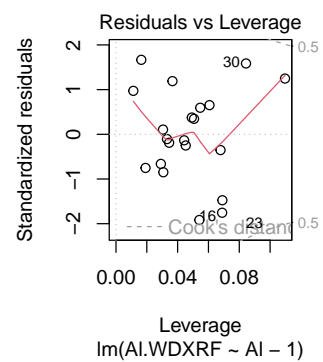
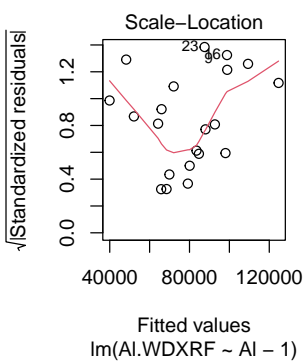
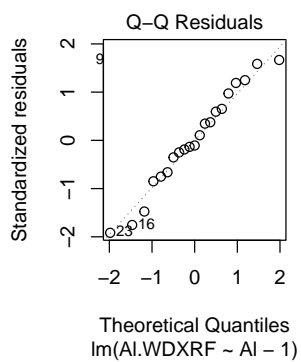
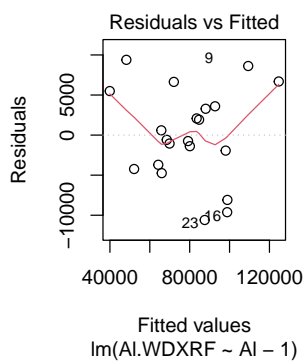
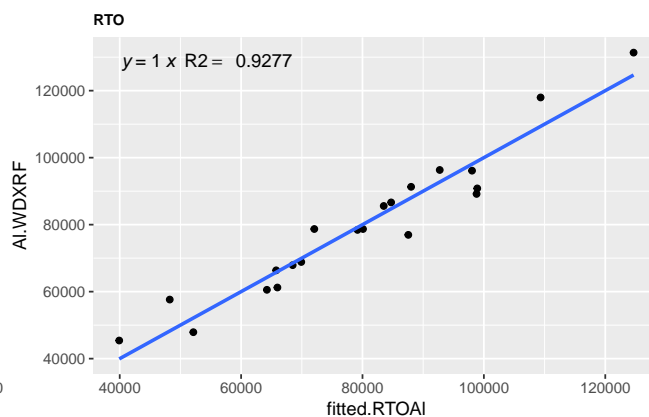
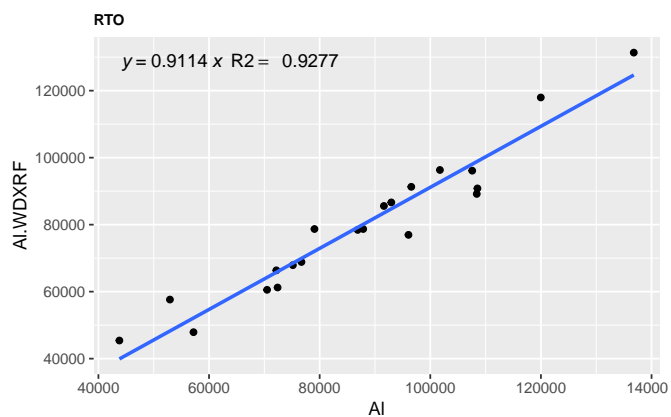
```
coefcorI_example_Al_RTO_allfig<-plot_grid(Al_scatter_RTO,Al_scatter_RTO_WDXRF_fitted,
                                           RTOA1_1_2,RTOA1_3_5,Al_boot_RTO,Al_hist_RTO,
                                           ncol=2,nrow=3,rel_heights=c(1,1.5,1))

ggsave("coefcorI_example_Al_RTO_allfig.eps",path="..//graphics"),plot=last_plot(),
        device="eps",height=25,width=30,unit=c("cm"))
```

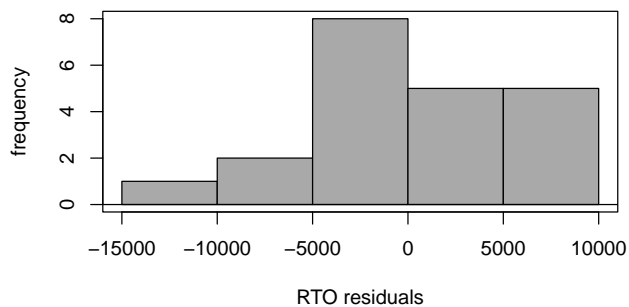
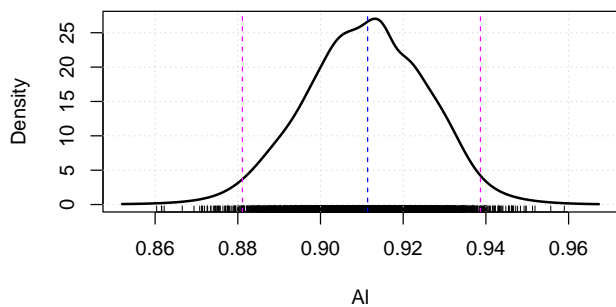


Bootstrap Distributions OLR





Bootstrap Distributions RTO



8.2 Exporting the traditionally presented scatter plots

source

Instructions following <https://www.digitalocean.com/community/tutorials/r-read-csv-file-into-data-frame>, <https://ggplot2.tidyverse.org/reference/ggsave.html>, <http://www.sthda.com/english/wiki/ggplot2-scatter-plots-quick-start-guide-r-software-and-data-visualization#basic-scatter-plots>, https://www.rdocumentation.org/packages/cowplot/versions/1.1.1/topics/plot_grid, <https://www.statology.org/plot-title-in-r/>, <https://datatofish.com/export-dataframe-to-csv-in-r/>

In addition to the named graphics which show the criteria of the selected linear regression of a particular element, you can of course create other graphs. Here we create the combined display of scatter plots, functions and r^2 values of different chemical elements that are traditionally used to present coefcors or calibrations.

In order to create such an example, a diagram for p-XRF values and WDXRF values of silicium (without any calculations to find the best linear regression but displaying a RTO of this data) is created using the data set **coefcorI_data.csv** and displayed together with the final scatter plot we defined above for aluminium. This graphics export is called **__traditional_scatterplots.eps** (in our example therefore **coefcorI_example_traditional_scatterplots.eps**)

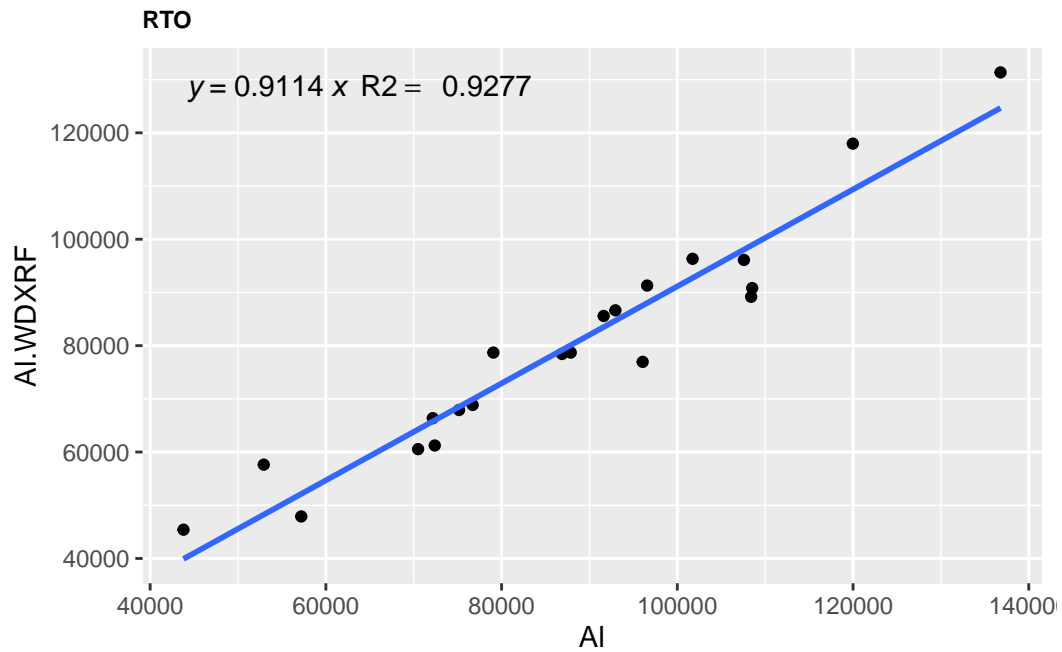
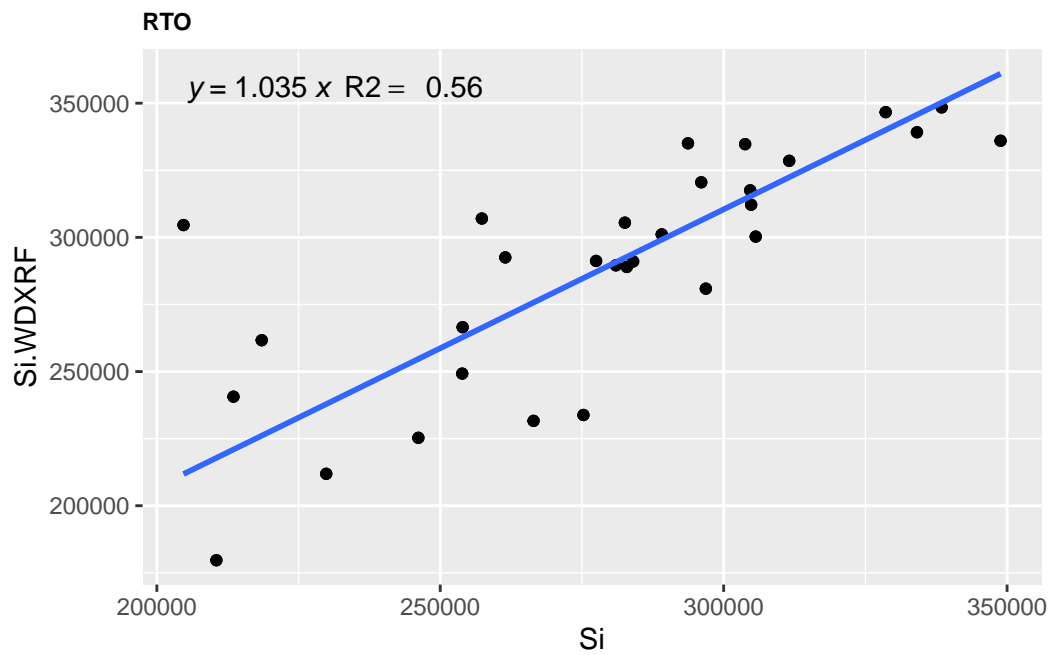
```
dataset<- read.csv("../data_analytical//coefcorI_data.csv")

RTOSi<-lm(Si.WDXRF~Si-1, data=dataset)
SEE_Si<-sigma(RTOSi)
SSR_Si<-(nrow(dataset)-1)*(SEE_Si^2)
cSST_Si <-sum(((dataset$Si.WDXRF)-mean(dataset$Si.WDXRF))^2)
Sicr2<-1-(SSR_Si/cSST_Si)
Sicr2<-round(Sicr2,2)

Si_scatter_RTO<-ggplot(dataset, aes(x=Si, y=Si.WDXRF))+geom_point()+
  geom_point()+geom_smooth(method=lm,se=FALSE,formula = y ~ x-1)+ ggtitle("RTO")+
  theme(plot.title = element_text(color="black", size=9,face="bold"))+
  stat_poly_eq(formula = y ~ x - 1,aes(label = paste(..eq.label..,
    sep = "~~", "R2==", Sicr2)),parse = TRUE, coef.digits = 4)

coefcorI_example_traditional_scatterplots<-plot_grid(Si_scatter_RTO,
  Al_scatter_RTO,ncol=2,nrow=1)

ggsave("coefcorI_example_traditional_scatterplots.eps",path=("../graphics"),
  plot=last_plot(),device="eps",height=8,width=15.3,unit=c("cm"))
```



9 Example of applying coefcors obtained from p-XRF and laboratory values to a dataset

The `__factor.csv` of a coefcor is only valid in the respective time period and is therefore used to apply the coefcors to a p-XRF data set produced during that time. In this example we use `coefcorI_factors.csv` which is valid for data of the Niton XL3t No 97390 collected between July 2017 and August 2018 (Schauer 2024).

9.1 Loading and formating the data

source

Instructions following <https://www.digitalocean.com/community/tutorials/r-read-csv-file-into-data-frame> and <https://sparkbyexamples.com/r-programming/r-select-function-from-dplyr/>

In the following, the application of the correction factors (`__factors.csv`) - in this case from coefcor I (`coefcorI_factors.csv`) - to a data set generated by p-XRF (e.g. measurements on ancient pottery) is demonstrated. `coefcorI_data.csv` is used as an example data set. First we load both files:

```
coefcor<- read.csv("../data_processed//coefcorI_factors.csv")
dataset<- read.csv("../data_analytical//coefcorI_data.csv")
```

Then we create a new dataset containing only the columns with p-XRF values and the sample names because we only want to recalculate those. We also define 'sample' as a variable to be able to add it to the recalculated data.

```
dataset<-select(dataset,"Sample","Si","Ti","Al","Fe","Mn","Mg","Ca","K","P","S",
  "Cl","Sc","V","Cr","Co","Ni","Cu","Zn","As","Se","Rb","Sr","Y",
  "Zr","Nb","Mo","Pd","Ag","Cd","Sn","Sb","Te","Cs","Ba","La","Ce",
  "Hf","Ta","W","Re","Au","Hg","Pb","Bi","Th","U")

Sample<-dataset$Sample
```

9.2 Performing the calculations

source

Instructions following <https://cran.r-project.org/web/packages/dplyr/vignettes/rowwise.html> and <https://dplyr.tidyverse.org/reference/mutate.html>

From this data, the necessary variables are extracted: new values to be calculated for the respective element (e.g. Si) as well as the slope `_a` (e.g. `Si_a`) and intercept `_b` (e.g. `Si_b`) to be applied.

```
Si<-dataset$Si
Si_a<-coefcor$Si_a
Si_b<-coefcor$Si_b
```

The recalculation is performed using these variables. It is then necessary to decide whether the values should be given in ppm (e.g. Si_ppm), weight percent (e.g. Sipct) or, in the case of the major elements Si, Ti, Al, Fe, Mn, Mg, Ca, K and P, in oxidation percentage (e.g. SiO2).

```
Sippm<-(Si_a*Si+Si_b
Sippm
```

```
[1] 282702.9 258373.1 276851.7 303960.0 316729.8 315289.7 245772.0 341986.0
[9] 249612.5 334382.8 286544.2 295039.9 309324.5 238982.8 309950.6 349988.8
[17] 315978.1 316104.4 338631.4 276907.9 270892.6 299213.1 293321.0 321274.1
[25] 279523.2 300073.3 297725.2 298952.8 243451.7 307523.2
```

```
Sipct<-(Si_a*Si+Si_b)*0.0001
Sipct
```

```
[1] 28.27029 25.83731 27.68517 30.39600 31.67298 31.52897 24.57720 34.19860
[9] 24.96125 33.43828 28.65442 29.50399 30.93245 23.89828 30.99506 34.99888
[17] 31.59781 31.61044 33.86314 27.69079 27.08926 29.92131 29.33210 32.12741
[25] 27.95232 30.00733 29.77252 29.89528 24.34517 30.75232
```

```
SiO2<-(Si_a*Si+Si_b)*0.00021393
SiO2
```

```
[1] 60.47863 55.27376 59.22688 65.02616 67.75800 67.44992 52.57800 73.16106
[9] 53.39960 71.53451 61.30039 63.11789 66.17379 51.12559 66.30773 74.87311
[17] 67.59721 67.62422 72.44342 59.23891 57.95206 64.01065 62.75017 68.73017
[25] 59.79839 64.19468 63.69236 63.95497 52.08162 65.78844
```

According to the Munich Procedure, the major elements are given in oxidation percentage.

```
Ti<-dataset$Ti
Ti_a<-coefcor$Ti_a
Ti_b<-coefcor$Ti_b
TiO2<-(Ti_a*Ti+Ti_b)*0.0001668

Al<-dataset$Al
Al_a<-coefcor$Al_a
Al_b<-coefcor$Al_b
Al2O3<-(Al_a*Al+Al_b)*0.00018895

Fe<-dataset$Fe
Fe_a<-coefcor$Fe_a
Fe_b<-coefcor$Fe_b
Fe2O3<-(Fe_a*Fe+Fe_b)*0.000143

Mn<-dataset$Mn
```

```

Mn_a<-coefcor$Mn_a
Mn_b<-coefcor$Mn_b
MnO<-(Mn_a*Mn+Mn_b)*0.00012912

Mg<-dataset$Mg
Mg_a<-coefcor$Mg_a
Mg_b<-coefcor$Mg_b
MgO<-(Mg_a*Mg+Mg_b)*0.00016583

Ca<-dataset$Ca
Ca_a<-coefcor$Ca_a
Ca_b<-coefcor$Ca_b
CaO<-(Ca_a*Ca+Ca_b)*0.00013992

K<-dataset$K
K_a<-coefcor$K_a
K_b<-coefcor$K_b
K2O<-(K_a*K+K_b)*0.00012046

P<-dataset$P
P_a<-coefcor$P_a
P_b<-coefcor$P_b
P2O5<-(P_a*P+P_b)*0.00022914

```

For sample-based normalisation, the created variables are combined into a new data set (data_norm) and supplemented with the sum of the main elements (sum) calculated line by line (data_norm_withsum). This new column is used to compute the percentage deviation from 100% of the major elements per row i.e. sample (sumpct).

```

data_norm<-data.frame(SiO2,TiO2,Al2O3,Fe2O3,MnO,MgO,CaO,K2O,P2O5)

data_norm_withsum<-data_norm %>% rowwise() %>% mutate(sum = sum(c(SiO2,TiO2,Al2O3,
                             Fe2O3,MnO,MgO,CaO,K2O,P2O5)))

sumpct<-100/data_norm_withsum$sum

```

The factor sumpct will now be used for the recalculation of the main elements.

```

SiO2<-SiO2*sumpct
TiO2<-TiO2*sumpct
Al2O3<-Al2O3*sumpct
Fe2O3<-Fe2O3*sumpct
MnO<-MnO*sumpct
MgO<-MgO*sumpct
CaO<-CaO*sumpct
K2O<-K2O*sumpct
P2O5<-P2O5*sumpct

```

Once the main elements have been recalculated, the cofcor factors (here coefcor I) are applied to the

trace elements. These are given in ppm. As examples, the calculations for rubidium and strontium are shown.

```
Rb<-dataset$Rb
Rb_a<-coefcor$Rb_a
Rb_b<-coefcor$Rb_b
Rb<-Rb_a*Rb+Rb_b

Sr<-dataset$Sr
Sr_a<-coefcor$Sr_a
Sr_b<-coefcor$Sr_b
Sr<-Sr_a*Sr+Sr_b
```

9.3 Compiling and exporting the processed data (_cor_data.csv)

source

Instructions following <https://search.r-project.org/R/refmans/base/html/Round.html>, <https://www.statology.org/transpose-data-frame-in-r/> and <https://datatofish.com/export-dataframe-to-csv-in-r/>

Finally, all the newly calculated variables are combined with the sample number in a new table called **_cor_data.csv** (in our case **coefcorI_exsample_cor_data.csv**). Major element values are rounded to one decimal place, trace element values to whole digits. This table is then exported. It contains quantitative, processed data and forms the basis for further analysis.

```
coefcorI_exsample_cor_data<-data.frame(Sample,SiO2,TiO2,Al2O3,Fe2O3,MnO,MgO,CaO,K2O,
                                         P2O5,Rb,Sr)

coefcorI_exsample_cor_data[, c("SiO2","TiO2","Al2O3","Fe2O3","MnO","MgO","CaO",
                                "K2O","P2O5")] <- round(coefcorI_exsample_cor_data[, c("SiO2","TiO2","Al2O3",
                                            "Fe2O3","MnO","MgO","CaO","K2O","P2O5")], 1)

coefcorI_exsample_cor_data[, c("Rb","Sr")] <- round(coefcorI_exsample_cor_data[, c("Rb","Sr")], 1)

coefcorI_exsample_cor_data
```

	Sample	SiO2	TiO2	Al2O3	Fe2O3	MnO	MgO	CaO	K2O	P2O5	Rb	Sr
1	H006	70.1	0.7	13.3	5.9	0.1	1.9	3.8	3.4	0.8	117	163
2	H045	56.6	0.5	10.1	3.8	0.1	2.8	24.3	1.4	0.4	141	295
3	H047	60.7	0.6	16.2	4.9	0.1	2.5	12.0	2.8	0.2	138	658
4	H052	69.0	0.7	15.9	6.1	0.2	2.3	2.4	3.1	0.4	138	135
5	H053	70.6	0.7	15.8	5.5	0.1	2.0	2.1	3.0	0.1	141	134
6	H056	70.8	1.3	16.8	7.3	0.1	0.7	1.0	1.9	0.1	118	78
7	H057	60.1	0.5	8.6	3.4	0.1	1.4	23.4	2.1	0.5	77	264
8	H058	72.1	0.6	13.9	4.4	0.0	2.5	3.9	2.4	0.1	107	109
9	H060	66.0	0.6	11.3	4.1	0.1	1.3	13.6	2.9	0.2	89	225
10	H066	75.5	0.7	13.2	5.0	0.0	1.3	2.1	2.2	0.1	100	97

11	H076	56.6	0.5	15.2	5.0	0.1	2.7	16.8	2.8	0.3	125	438
12	H077	70.3	0.7	15.2	6.7	0.2	1.5	2.5	2.2	0.8	135	189
13	H078	74.6	0.6	14.0	4.3	0.1	1.4	2.2	2.4	0.4	117	143
14	H079	71.5	0.8	14.4	7.5	0.1	1.1	0.9	3.2	0.5	104	71
15	H087	63.2	0.7	17.8	7.2	0.1	3.0	3.7	3.8	0.4	156	142
16	H088	71.4	0.7	17.8	3.5	0.0	1.8	1.0	3.6	0.1	280	99
17	H089	68.3	0.6	18.7	4.1	0.0	1.5	2.7	3.6	0.4	225	140
18	H092	66.3	0.6	17.2	3.9	0.1	3.9	4.9	2.9	0.2	144	170
19	H095	74.3	1.7	17.1	3.9	0.0	0.6	0.9	1.3	0.2	82	105
20	H100	63.9	0.7	14.2	5.1	0.1	1.9	11.1	2.6	0.4	112	480
21	H102	59.5	0.5	13.3	4.8	0.1	2.8	16.4	2.5	0.1	108	339
22	H111	63.8	0.8	17.5	6.9	0.1	3.2	4.1	3.5	0.2	172	138
23	H112	59.5	0.7	15.7	5.9	0.1	3.1	13.2	1.6	0.3	162	429
24	H114	70.2	1.1	22.5	1.5	0.0	0.5	2.4	1.7	0.2	115	117
25	H124	70.2	2.1	20.4	3.2	0.0	0.7	1.2	1.9	0.4	72	95
26	H125	66.6	1.2	24.4	3.1	0.0	1.1	1.0	2.4	0.2	155	147
27	H126	66.6	1.3	24.3	3.1	0.0	1.5	0.8	2.2	0.2	175	141
28	H127	68.7	2.2	20.6	4.7	0.1	0.6	1.0	1.9	0.2	97	294
29	H138	54.1	0.5	12.6	4.3	0.1	3.8	21.8	2.5	0.2	83	317
30	H139	71.1	1.1	22.3	3.5	0.0	0.3	0.5	1.1	0.1	63	77

```
write.csv(coefcorI_exsample_cor_data,
          "../data_processed//coefcorI_example_cor_data.csv",row.names=TRUE)
```

10 Example of applying device-internal coefcors to a dataset

The `__factor.csv` needed to correct values from the time period in which they were taken (for example from coefcor I) to fit the data of different time period (for example coefcor II; see Schauer 2023 & Schauer 2024) has to be chosen. In this example we use `coefcorItoII_factors.csv`.

10.1 Loading and formating the data

source

Instructions following <https://www.digitalocean.com/community/tutorials/r-read-csv-file-into-data-frame> and <https://sparkbyexamples.com/r-programming/r-select-function-from-dplyr/>

In the following, the application of the correction factors (`__factors.csv`) - in this case from coefcor ItoII (`coefcorItoII_factors.csv`) - to a data set generated by p-XRF (e.g. measurements on ancient pottery) is demonstrated. As before, we load both files and create the new dataset and variable.

```
coefcor<- read.csv("../data_processed//coefcorItoII_factors.csv")
dataset<- read.csv("../data_analytical//coefcorI_data.csv")

dataset<-select(dataset,"Sample","Si","Ti","Al","Fe","Mn","Mg","Ca","K","P","S",
                 "Cl","Sc","V","Cr","Co","Ni","Cu","Zn","As","Se","Rb","Sr","Y",
```

```
"Zr","Nb","Mo","Pd","Ag","Cd","Sn","Sb","Te","Cs","Ba","La","Ce",
"Hf","Ta","W","Re","Au","Hg","Pb","Bi","Th","U")
```

```
Sample<-dataset$Sample
```

10.2 Performing the calculations

As this calculation is performed to fit the data before applying the coefcor used in the period to which the data are correct (in our case this would be coefcor II), no calculation of oxidation percentage or normalisation needs to be performed. For all elements the corrections are applied and expressed in ppm.

```
Si<-dataset$Si
Si_a<-coefcor$Si_a
Si_b<-coefcor$Si_b
Si<-Si_a*Si+Si_b
```

```
Ti<-dataset$Ti
Ti_a<-coefcor$Ti_a
Ti_b<-coefcor$Ti_b
Ti<-Ti_a*Ti+Ti_b
```

```
Al<-dataset$Al
Al_a<-coefcor$Al_a
Al_b<-coefcor$Al_b
Al<-Al_a*Al+Al_b
```

```
Fe<-dataset$Fe
Fe_a<-coefcor$Fe_a
Fe_b<-coefcor$Fe_b
Fe<-Fe_a*Fe+Fe_b
```

```
Mn<-dataset$Mn
Mn_a<-coefcor$Mn_a
Mn_b<-coefcor$Mn_b
Mn<-Mn_a*Mn+Mn_b
```

```
Mg<-dataset$Mg
Mg_a<-coefcor$Mg_a
Mg_b<-coefcor$Mg_b
M<-Mg_a*Mg+Mg_b
```

```
Ca<-dataset$Ca
Ca_a<-coefcor$Ca_a
Ca_b<-coefcor$Ca_b
C<-Ca_a*Ca+Ca_b
```

```

K<-dataset$K
K_a<-coefcor$K_a
K_b<-coefcor$K_b
K<-K_a*K+K_b

P<-dataset$P
P_a<-coefcor$P_a
P_b<-coefcor$P_b
P<-P_a*P+P_b

Cl<-dataset$Cl
Cl_a<-coefcor$Cl_a
Cl_b<-coefcor$Cl_b
Cl<-Cl_a*Cl+Cl_b

V<-dataset$V
V_a<-coefcor$V_a
V_b<-coefcor$V_b
V<-V_a*V+V_b

Cr<-dataset$Cr
Cr_a<-coefcor$Cr_a
Cr_b<-coefcor$Cr_b
Cr<-Cr_a*Cr+Cr_b

Ni<-dataset$Ni
Ni_a<-coefcor$Ni_a
Ni_b<-coefcor$Ni_b
Ni<-Ni_a*Ni+Ni_b

Zn<-dataset$Zn
Zn_a<-coefcor$Zn_a
Zn_b<-coefcor$Zn_b
Zn<-Zn_a*Zn+Zn_b

As<-dataset$As
As_a<-coefcor$As_a
As_b<-coefcor$As_b
As<-As_a*As+As_b

Rb<-dataset$Rb
Rb_a<-coefcor$Rb_a
Rb_b<-coefcor$Rb_b
Rb<-Rb_a*Rb+Rb_b

Sr<-dataset$Sr
Sr_a<-coefcor$Sr_a
Sr_b<-coefcor$Sr_b

```

```

Sr<-Sr_a*Sr+Sr_b

Y<-dataset$Y
Y_a<-coefcor$Y_a
Y_b<-coefcor$Y_b
Y<-Y_a*Y+Y_b

Zr<-dataset$Zr
Zr_a<-coefcor$Zr_a
Zr_b<-coefcor$Zr_b
Zr<-Zr_a*Zr+Zr_b

Nb<-dataset$Nb
Nb_a<-coefcor$Nb_a
Nb_b<-coefcor$Nb_b
Nb<-Nb_a*Nb+Nb_b

Ba<-dataset$Ba
Ba_a<-coefcor$Ba_a
Ba_b<-coefcor$Ba_b
Ba<-Ba_a*Ba+Ba_b

Pb<-dataset$Pb
Pb_a<-coefcor$Pb_a
Pb_b<-coefcor$Pb_b
Pb<-Pb_a*Pb+Pb_b

```

10.3 Compiling and exporting the processed data (__cor_data.csv)

source

Instructions following <https://search.r-project.org/R/refmans/base/html/Round.html>, <https://www.statology.org/transpose-data-frame-in-r/> and <https://datatofish.com/export-dataframe-to-csv-in-r/>

As before, all newly calculated variables are combined with the sample number in a new table called **__cor_data.csv** (in our case **coefcorItoII_example_cor_data.csv**). All element values are rounded to whole digits. This table is then exported. It contains quantitative, processed data and forms the basis for applying the coefcor of the given time period (here coefcor II).

```

coefcorItoII_cor_data<-data.frame(Sample,Si,Ti,Al,Fe,Mn,Mg,Ca,
                                   K,P,Cl,V,Cr,Ni,Zn,As,Rb,Sr,Y,Zr,Nb,Ba,Pb)

coefcorItoII_cor_data[, c("Si","Ti","Al","Fe","Mn","Mg","Ca","K","P","Cl","V",
                           "Cr","Ni","Zn","As","Rb","Sr","Y","Zr","Nb",
                           "Ba","Pb")] <- round(coefcorItoII_cor_data[, c("Si","Ti","Al","Fe",
                                   "Mn","Mg","Ca","K","P","Cl","V","Cr","Ni","Zn","As","Rb","Sr",
                                   "Y","Zr","Nb","Ba","Pb")], 0)

```

```
write.csv(coefcorItoII_cor_data,  
          "../data_processed//coefcorItoII_example_cor_data.csv",row.names=TRUE)
```