

Lab3

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Overview

Today we are looking at data on infant deaths (deaths in the first year of life) in the US. The dataset `infant` contains information on all deaths to the 2012 birth cohort. For today, we are interested in investigating differences in neonatal deaths (i.e. deaths in the first month of life) and cause of death.

What to hand in

As with last week, please push your Rmd and compiled document (html or pdf) to GitHub. **The questions for this week are dispersed throughout the lab.**

The dataset

Read it in and have a look to see what's in there. Variables are

- `sex`: sex of baby
- `aged`: age at death (in days)
- `race`: race of mother
- `gest`: gestation in weeks
- `ucod`: cause of death (ICD-10 code)
- `cod`: cause of death, descriptive groups
- `mom_age`: mother age in years
- `mom_age_group`: mother age group

```
library(tidyverse)
library(here)
d <- read_rds(here("data", "infant.RDS"))
head(d)

## # A tibble: 6 x 8
##   sex    aged race    gest ucod   cod      mom_age mom_age_group
##   <chr> <dbl> <chr> <dbl> <chr> <chr>    <dbl> <fct>
## 1 F      0 NHW     27 P832 peri_oth    30 30
## 2 M      0 NHW     36 Q913 cong_mal    32 30
## 3 M      8 NHW     44 P360 peri_inf    25 25
## 4 F      0 NHB     21 P072 peri_comp    29 25
## 5 M      8 NHB     26 P220 peri_resp    23 20
## 6 M     17 NHW     39 Q249 cong_mal    34 30
```

Descriptives

Let's create some new variables that will be useful:

- `neo_death`: equals 1 if the death occurred in the first 28 days

- **preterm**: equals 1 if gestational age is less than 37 weeks
- **cod_group**: reduced number of categories of cause of death

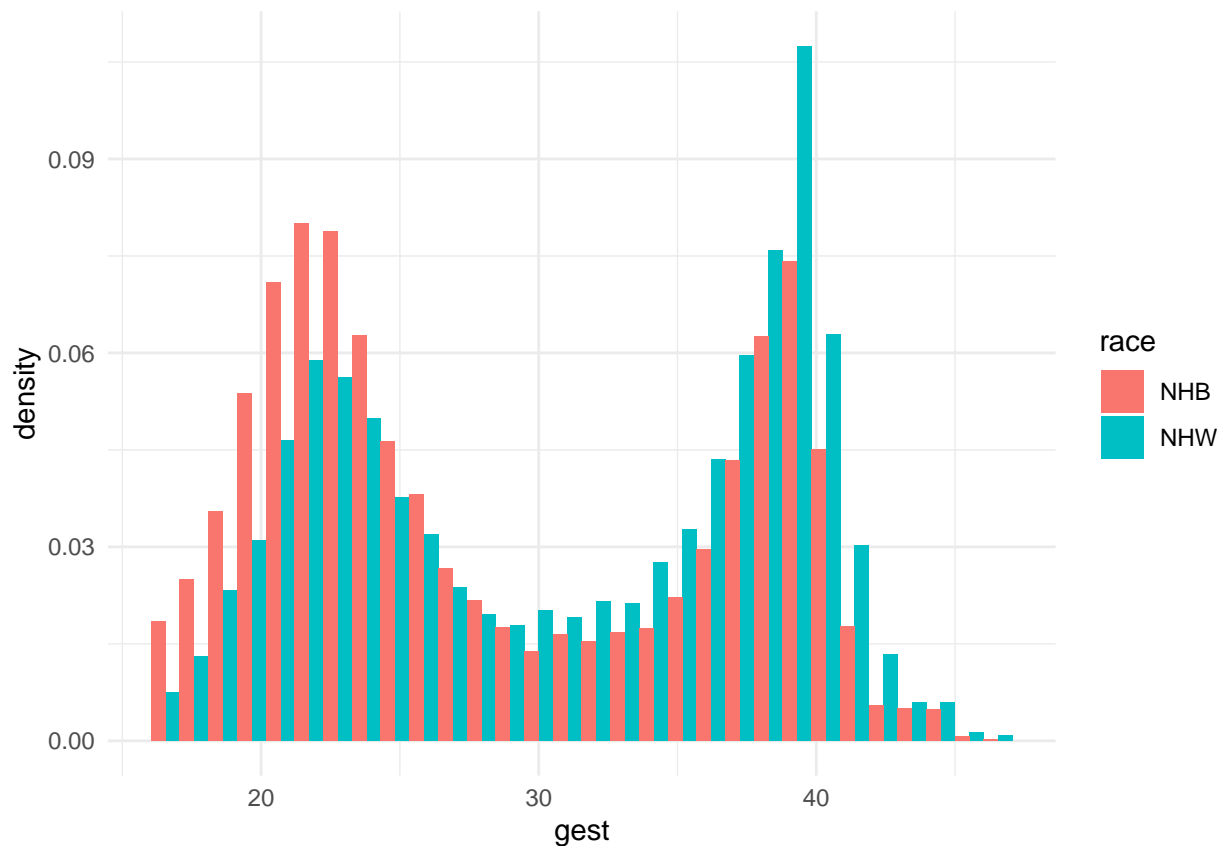
Also, removing the observations where we don't know gestational age or the mother's age.

```
d <- d %>%
  mutate(
    neo_death = ifelse(aged <= 28, 1, 0),
    cod_group = case_when(
      str_starts(cod, "peri") ~ "perinatal",
      cod %in% c("other", "unknown") ~ "oth_unk",
      cod %in% c("sids", "maltreatment", "infection") ~ "exogenous",
      cod %in% c("resp", "heart") ~ "resp_heart",
      TRUE ~ cod
    ),
    preterm = ifelse(gest < 37, 1, 0)
  ) %>%
  filter(gest < 99, !is.na(mom_age_group))
```

Distribution of gestational ages

Let's plot the distribution of gestational ages by race. It's quite bi-modal. Notice the difference in densities by race.

```
d %>% ggplot(aes(gest, fill = race)) +
  geom_histogram(position = 'dodge', aes(y = ..density..)) +
  theme_minimal()
```



Question 1

Calculate the proportion of deaths that are neonatal by race and prematurity. Which group has the highest proportion of neonatal deaths?

```
q1_df <- d %>% group_by(race, preterm) %>%  
  summarize(prop_neonatal = mean(neo_death))  
  
kableExtra::kable(q1_df)
```

race	preterm	prop_neonatal
NHB	0	0.3089953
NHB	1	0.8021413
NHW	0	0.3823013
NHW	1	0.8242140

Highest proportion of neonatal deaths are in the race “NHW” (non-hispanic white) and for the premature births. In general it seems that the premature births have an incredibly higher neonatal death rate.

Causes of death

Let's make `cod_group` a factor with congenital malformations as the reference.

```
d <- d %>%  
  mutate(cod_group = factor(cod_group, levels = c("cong_mal", "perinatal", "resp_heart", "exogenous", "oth_unk"),  
    labels = c("cong_mal", "perinatal", "resp_heart", "exogenous", "oth_unk")))
```

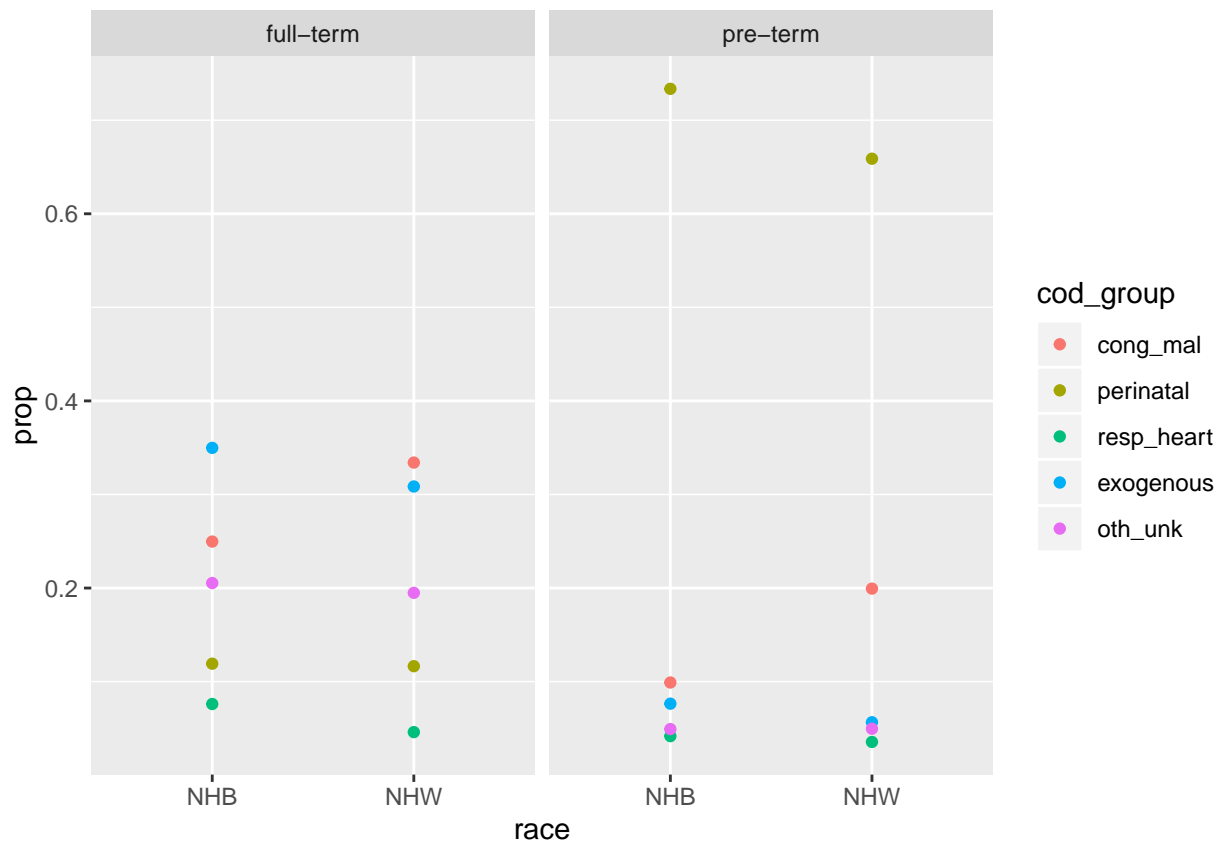
The following code calculates the proportion of deaths by cause group, race, sex and prematurity

```
prop_cause <- d %>%  
  group_by(race, preterm, sex, cod_group) %>%  
  summarise(n = n()) %>%  
  group_by(race, preterm, sex) %>%  
  mutate(prop = n / sum(n)) %>%  
  ungroup() %>%  
  mutate(preterm = ifelse(preterm == 1, "pre-term", "full-term"))
```

Question 2

Using the `prop_cause` above, filter to just look at female babies, and make a graph to help visualize differences in cause by race and prematurity.

```
prop_cause %>% filter(sex == "F") %>%  
  ggplot(aes(  
    x = race,  
    facet = preterm,  
    y = prop,  
    color = cod_group )) +  
  geom_point() +  
  facet_wrap(. ~ preterm)
```



Logistic regression

First, let's do logistic regression to explore differences in neonatal deaths. Here's a model with prematurity, sex, race, and mom's age

```
mod <- glm(neo_death ~ preterm + sex + race + race:preterm + mom_age, data = d, family = binomial)
summary(mod)
```

```
##
## Call:
## glm(formula = neo_death ~ preterm + sex + race + race:preterm +
##      mom_age, family = binomial, data = d)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.1160  -0.9126   0.6028   0.6825   1.6966
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  -1.467493   0.095651 -15.342  < 2e-16 ***
## preterm       2.191541   0.064204  34.134  < 2e-16 ***
## sexM        -0.131361   0.036994  -3.551 0.000384 ***
## raceNHW       0.290680   0.061988   4.689 2.74e-06 ***
## mom_age       0.028682   0.002997   9.569  < 2e-16 ***
## preterm:raceNHW -0.191417  0.078973  -2.424 0.015359 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##
## (Dispersion parameter for binomial family taken to be 1)
##
## Null deviance: 21723 on 16986 degrees of freedom
## Residual deviance: 18115 on 16981 degrees of freedom
## AIC: 18127
##
## Number of Fisher Scoring iterations: 4
```

Question 3

Rerun the model above with instead of `mom_age`, include a new variable `mom_age_c` which centers mother's age around its mean.

```
d_model <- d %>% mutate(mom_age_c = scale(mom_age, scale = FALSE))

mod_age_norm <- glm(neo_death ~ preterm + sex + race + race:preterm + mom_age_c, data = d_model, family =
summary(mod_age_norm)
```

```
##
## Call:
## glm(formula = neo_death ~ preterm + sex + race + race:preterm +
## mom_age_c, family = binomial, data = d_model)
##
## Deviance Residuals:
## Min 1Q Median 3Q Max
## -2.1160 -0.9126 0.6028 0.6825 1.6966
##
## Coefficients:
## Estimate Std. Error z value Pr(>|z|)
## (Intercept) -0.690052 0.056100 -12.300 < 2e-16 ***
## preterm 2.191541 0.064204 34.134 < 2e-16 ***
## sexM -0.131361 0.036994 -3.551 0.000384 ***
## raceNHW 0.290680 0.061988 4.689 2.74e-06 ***
## mom_age_c 0.028682 0.002997 9.569 < 2e-16 ***
## preterm:raceNHW -0.191417 0.078973 -2.424 0.015359 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
## Null deviance: 21723 on 16986 degrees of freedom
## Residual deviance: 18115 on 16981 degrees of freedom
## AIC: 18127
##
## Number of Fisher Scoring iterations: 4
```

Question 4

```
coefs <- mod_age_norm$coefficients
```

Interpret the `preterm`, `race` and the interaction `preterm:race` coefficients.

The `preterm` coefficient with a value of 2.1915407 which means that babies born prematurely are 8.9489903 time more likely to have a neonatal death.

The race coefficient with a value of 0.29068 which means that babies born as to non hispanic whites are 1.3373366 time more likely to have a neonatal death.

The preterm:raceNHW coefficient with a value of -0.1914166 which means that babies born preterm to non hispanic whites are a further 0.8257885 time more likely to have a neonatal death on top of the other coefficients (for race:NHW and preterm).

Multinomial regression

Now let's do multinomial regression with cause of death as the outcome. We need to get the data in a different format to run the regression:

```
d$mom_age_c <- d$mom_age - mean(d$mom_age)
d_wide <- d %>%
  group_by(sex, race, cod_group, preterm, mom_age_c) %>%
  summarise(deaths = n()) %>%
  pivot_wider(names_from = cod_group, values_from = deaths) %>%
  mutate_all(.funs = funs(ifelse(is.na(.), 0, .)))

d_wide$Y <- as.matrix(d_wide[, c("cong_mal", "perinatal", "resp_heart", "exogenous", "oth_unk")])
```

Now run the regression

```
library(nnet)
mod2 <- multinom(Y ~ sex + race + mom_age_c + preterm, data = d_wide)

## # weights: 30 (20 variable)
## initial value 27339.521819
## iter 10 value 22475.496335
## iter 20 value 19882.612578
## iter 30 value 19389.722462
## final value 19389.720141
## converged

summary(mod2)

## Call:
## multinom(formula = Y ~ sex + race + mom_age_c + preterm, data = d_wide)
##
## Coefficients:
## (Intercept) sexM raceNHW mom_age_c preterm
## perinatal -0.53315841 0.0657566 -0.6249840 -0.01906239 2.4190484
## resp_heart -1.21149941 0.1350905 -0.6303309 -0.03765107 0.1683872
## exogenous 0.40732759 0.2070469 -0.5359725 -0.07602388 -1.0125410
## oth_unk -0.09571594 0.1682889 -0.4746567 -0.04122105 -0.7189038
##
## Std. Errors:
## (Intercept) sexM raceNHW mom_age_c preterm
## perinatal 0.06335013 0.04452244 0.04846619 0.003515557 0.05452877
## resp_heart 0.09493043 0.08128540 0.08477048 0.006591926 0.08175145
## exogenous 0.06236608 0.05376561 0.05842537 0.004532928 0.05637655
## oth_unk 0.06948526 0.06015565 0.06522548 0.004905661 0.06185463
##
## Residual Deviance: 38779.44
## AIC: 38819.44
```

Question 5

Using the `predict` function, find the predicted probabilities of each cause by race, sex and prematurity for the mothers of mean age. You can use this prediction dataframe to get all the combinations you need.

```
pred_df <- tibble(
  preterm = c(rep(0, 4), rep(1, 4)),
  sex = rep(c(rep("F", 2), rep("M", 2)), 2),
  race = rep(c("NHB", "NHW"), 4),
  mom_age_c = 0)

predictions <- as_tibble(predict(mod2, type = "probs", newdata = pred_df))

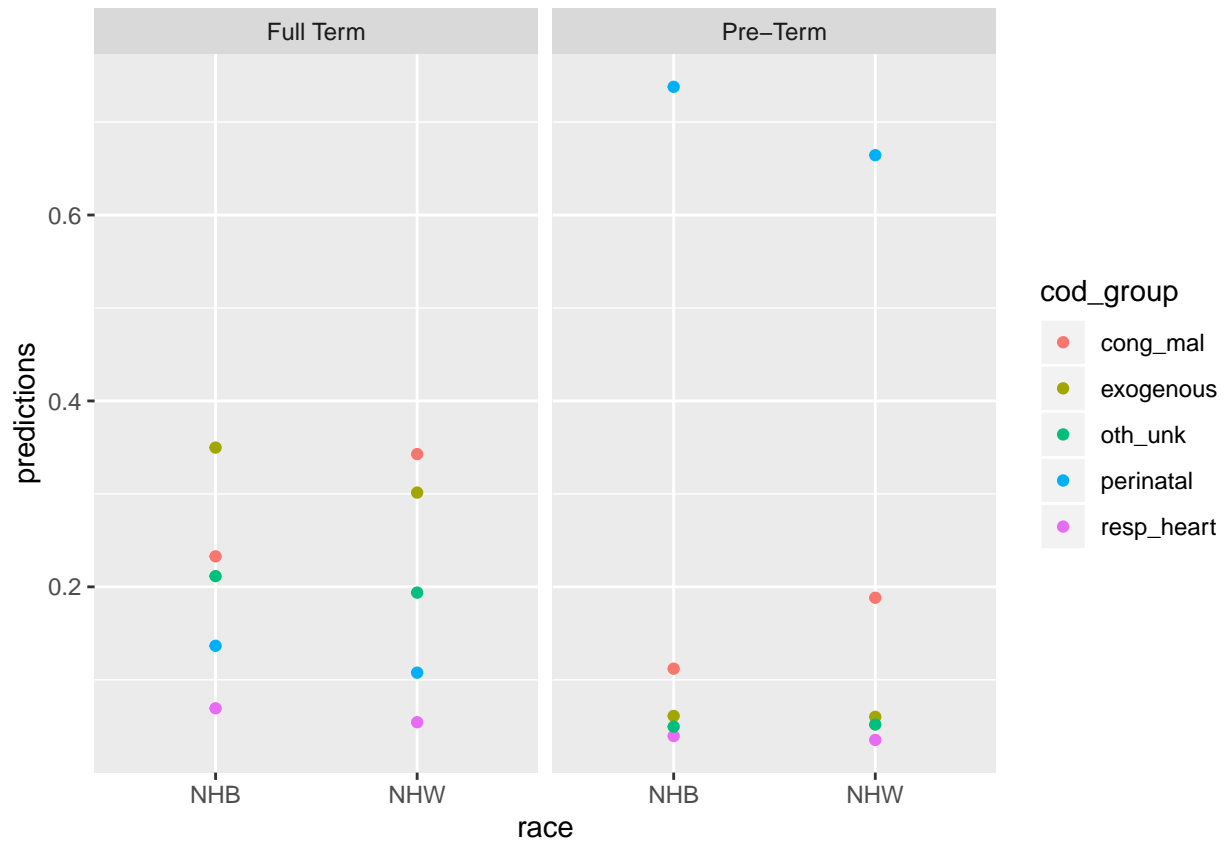
pred_df <- cbind(pred_df, predictions) %>%
  pivot_longer(cols = c(names(predictions)), names_to = "cod_group", values_to = "predictions")
```

Question 6

Plot the predicted probabilities for female babies.

```
pred_df_plot <- pred_df %>%
  mutate(preterm = as_factor(if_else(preterm == 0, "Full Term", "Pre-Term")))

pred_df_plot %>% filter(sex == "F") %>%
  ggplot(aes(
    x = race,
    y = predictions,
    facet = preterm,
    color = cod_group
  )) +
  geom_point() +
  facet_wrap(. ~ preterm)
```



Question 7

What race/prematurity/ cause group has the highest probability? How does this compare to the observed proportion in the same group?

The group with the highest probability is the Pre-term babies birth by Non-Hispanic-Black Women that are born female that have perinatal congenital malformations.

Observed proportion for this group is 0.969157 while the predicted proportion is 0.7379343 which is significantly lower. This is probably not very good.